

Disentangling type 2 diabetes and metformin treatment signatures in the human gut microbiota

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In recent years, several associations between common chronic human disorders and altered gut microbiome composition and function have been reported^{1,2}. In most of these reports, treatment regimens were not controlled for and conclusions could thus be confounded by the effects of various drugs on the microbiota, which may obscure microbial causes, protective factors or diagnostically relevant signals. Our study addresses disease and drug signatures in the human gut microbiome of type 2 diabetes mellitus (T2D). Two previous quantitative gut metagenomics studies of T2D patients that were unstratified for treatment yielded divergent conclusions regarding its associated gut microbial dysbiosis^{3,4}. Here we show, using 784 available human gut metagenomes, how antidiabetic medication confounds these results, and analyse in detail the effects of the most widely used antidiabetic drug metformin. We provide support for microbial mediation of the therapeutic effects of metformin through short-chain fatty acid production, as well as for potential microbiota-mediated mechanisms behind known intestinal adverse effects in the form of a relative increase in abundance of *Escherichia* species. Controlling for metformin treatment, we report a unified signature of gut microbiome shifts in T2D with a depletion of butyrate-producing taxa^{3,4}. These in turn cause functional microbiome shifts, in part alleviated by metformin-induced changes. Overall, the present study emphasizes the need to disentangle gut microbiota signatures of specific human diseases from those of medication.

T2D is a disorder of elevated blood glucose levels (hyperglycaemia) primarily due to insulin resistance and inadequate insulin secretion, with rising global prevalence. Genetic and environmental risk factors are known, the latter including dietary habits and a sedentary lifestyle⁵. Gut microbiota involvement is also increasingly recognized^{3,4,6,7}, although findings diverge between studies⁸; for example, Qin *et al.*³ report several *Clostridium* species enriched in T2D, whereas Karlsson *et al.*⁴ instead report enrichment of several lactobacilli species (see Supplementary Discussion). Treatment involves medication and

lifestyle intervention, which may confound reported gut dysbiosis. Many T2D patients receive metformin, an oral blood-glucose-lowering non-metabolizable compound whose primary and dominant metabolic effect is the inhibition of liver glucose production⁹. At least 30% of patients report adverse effects including diarrhoea, nausea, vomiting and bloating, with underlying mechanisms poorly understood. Studies in animals¹⁰ and humans¹¹ suggest that some beneficial effects of metformin on glucose metabolism may be microbially mediated. Here, we built a multi-country T2D metagenomic data set, starting with gut microbial samples from a nondiabetic Danish cohort of 277 individuals within the MetaHIT project¹² and additional novel Danish MetaHIT metagenomes from 75 T2D and 31 type 1 diabetes (T1D) patients, sequenced using the same protocols (samples abbreviated as MHD). Treatment information was obtained for all MHD samples, as well as for samples from a previously reported⁴ cohort of 53 female Swedish T2D patients, along with 92 nondiabetic individuals (43 with normal glucose tolerance, 49 with impaired glucose tolerance) (SWE) and a subgroup of 71 Chinese T2D patients with available information on antidiabetic treatment as well as 185 nondiabetic Chinese individuals³ (CHN). For these 784 gut metagenomes (Supplementary Table 1), taxonomic and functional profiles were determined (see Methods), verifying our meta-analysis framework to be appropriate and robust in the context of theoretical considerations and through simulations (Supplementary Discussion 1 and Extended Data Fig. 1a), as well as characterizing differences between the data sets (Extended Data Fig. 2). Initial analysis unstratified for treatment but controlling for demographic and technical variation between data sets (Supplementary Discussion 2 and Supplementary Table 2) recovered a majority of previously reported associations (Supplementary Discussion 2 and Supplementary Table 3) but with large divergence between data sets. Suspecting confounding treatments, we tested for influence of diet and antidiabetic medications (Supplementary Discussion 3, Supplementary Table 4 and Extended Data Fig. 1b), finding an effect resulting only from use of metformin. As the fraction

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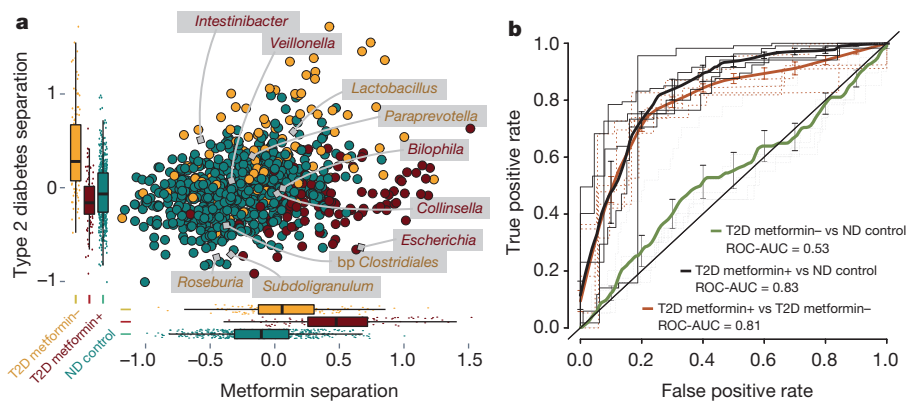


Figure 1 | Type 2 diabetes is confounded by metformin treatment.

Major treatment effects are seen in multivariate analysis and in classifier performance. **a**, Projection of genus-level gut microbiome samples from Danish, Chinese and Swedish studies constrained by diabetic state and metformin treatment. Multivariate analysis (dbRDA plot based on Canberra distances between bacterial genera) reveals a T2D dysbiosis, which overlaps only in part with taxonomic changes in metformin-treated patients. The ordination projects all T2D metformin+ ($n = 93$, dark red), T2D metformin- ($n = 106$, orange) and ND control ($n = 554$, teal) gut metagenomes, with confounding country effect adjusted for. Bacterial genera that show significant effects of metformin treatment and T2D status compared to ND control, respectively (limited to top five for each), are interpolated into the plane of maximal separation based on their abundances across all samples. Marginal box/scatter plots show the separation of the constrained projection coordinates (boxes show medians/quartiles, error bars extend to most extreme value within

1.5 interquartile ranges). The T2D separation is significant (PERMANOVA $FDR < 0.005$) in the joint data set and independently significant in CHN and MHD samples. The metformin separation is significant (PERMANOVA $FDR < 0.1$; Canberra distances) in MHD and SWE samples. bp, butyrate-producing. **b**, Classifying T2D and metformin treatment status based on gut microbiome profiles. Support vector machine classifiers were used to separate T2D metformin+ ($n = 93$), T2D metformin- ($n = 106$) and ND control ($n = 554$) gut metagenomes from each other based on genus-level gut microbiome taxonomic composition. Bold curves represent mean performance in hold-out testing of 1 out of 5 of the data each time, with separate tests shown as dashed curves. Error bars show ± 1 s.d. Metformin-treated T2D samples can be well separated from controls (using *Intestinibacter* abundance as the only feature), whereas distinguishing T2D metformin- samples from ND control samples works poorly even in the best case, requiring 63 distinct microbial features to achieve this separation. ROC-AUC, area under the receiving operating characteristic curve.

of medicated patients (denoted as T2D metformin+) varied strongly (21% CHN, 38% SWE and 77% MHD), samples were stratified on metformin treatment status. Multivariate analysis showed significant (permutational multivariate analysis of variance (PERMANOVA) false discovery rate ($FDR < 0.005$) differences in gut taxonomic composition between metformin-untreated T2D (T2D metformin-) ($n = 106$) patients and nondiabetic controls (ND control) ($n = 554$), consistent with a broad-range dysbiosis in T2D (Fig. 1a and Supplementary Table 5; see also Extended Data Table 1a and Supplementary Discussion 3 for an analysis of variances broken down by source). While metformin treatment status could be reliably recovered from microbial composition using support vector machines, metformin-untreated T2D status itself could not (Fig. 1b and Supplementary Table 6). In contrast, in all three cohorts, drug-treatment-blinded T2D samples could be separated from ND control samples with similar accuracy as previously reported^{3,4}, suggesting that the T2D metformin+ classifier robustly outperforms T2D metformin- classifiers across data sets (Supplementary Table 7).

We further explored T2D gut microbiome alterations in 106 metformin-untreated T2D compared with 554 ND control samples through univariate tests of microbial taxonomic and functional differences, with significant trends shown in Fig. 2a. Metformin-untreated T2D was associated with a decrease in genera containing known butyrate producers such as *Roseburia* spp., *Subdoligranulum* spp. and a cluster of butyrate-producing Clostridiales spp. (Supplementary Table 8), consistent with previous indications^{3,4}. More fine-grained taxonomic analysis indicated some driver species (Supplementary Discussion 4 and Supplementary Table 9), and further found changes in abundance of several unclassified Firmicutes, often reduced or reversed under metformin treatment (see Supplementary Discussion 4). Although an increase in *Lactobacillus* spp. was seen in treatment-unstratified T2D samples (as previously found experimentally¹³), this trend was eliminated or reversed when controlling for metformin. Functionally, we found enrichment of catalase (conceivably a response to increased peroxide stress under inflammation) and modules for ribose, glycine and tryptophan amino acid degradation,

but a decrease in threonine and arginine degradation, and in pyruvate synthase capacity (Supplementary Table 10). While these functional differences could result from strain-level composition changes or be a compound effect of subtle enrichment/depletion of larger ecological guilds, the abundance of most of these modules correlated with abundance of the significantly altered microbial genera (Fig. 2a).

To interpret our findings on T2D gut microbiota shifts further, we compared them with 31 adult T1D patients (Supplementary Table 1; for further discussion of this sub-cohort, see also Supplementary Discussion 5 and Supplementary Tables 6 and 11). This group is dysglycaemic like T2D patients, allowing us to separate purely glycaemic phenotype effects from T2D-specific microbial features. Gene richness was significantly increased in the T1D microbiomes (Wilcoxon rank sum test $FDR < 0.1$) (Fig. 2b), but was reduced in T2D (Supplementary Table 10), as reported previously⁶. Features found to distinguish metformin-untreated T2D from ND control microbiomes did not replicate when comparing T1D to ND control. Instead, most differences between metformin-untreated T2D samples and ND controls were reversed in adult T1D patients. In contrast, some microbial functions differentially abundant between metformin-untreated T2D and controls showed similar trends in T1D samples (Fig. 2a), although not significantly, possibly owing to lower statistical power. We therefore conclude that the majority of gut microbiota shifts visible in metformin-untreated T2D are not simply effects of dysglycaemia, but rather directly or indirectly associated with the causes or progression of T2D.

Suspecting microbial mediation of some of the therapeutic effects of metformin, we next compared T2D metformin-treated ($n = 93$) and T2D metformin-untreated ($n = 106$) samples to characterize the treatment effect in more detail. Multivariate contrasts of T2D metformin-treated with T2D metformin-untreated samples appeared weaker than those between T2D metformin-untreated and ND control samples, the former only significant at the bacterial family level (PERMANOVA $FDR < 0.1$), suggesting that the effects of metformin treatment on gut microbial composition are poorly captured by multivariate analysis. Univariate tests of the effects of metformin treatment showed a significant increase of *Escherichia* spp. and a reduced

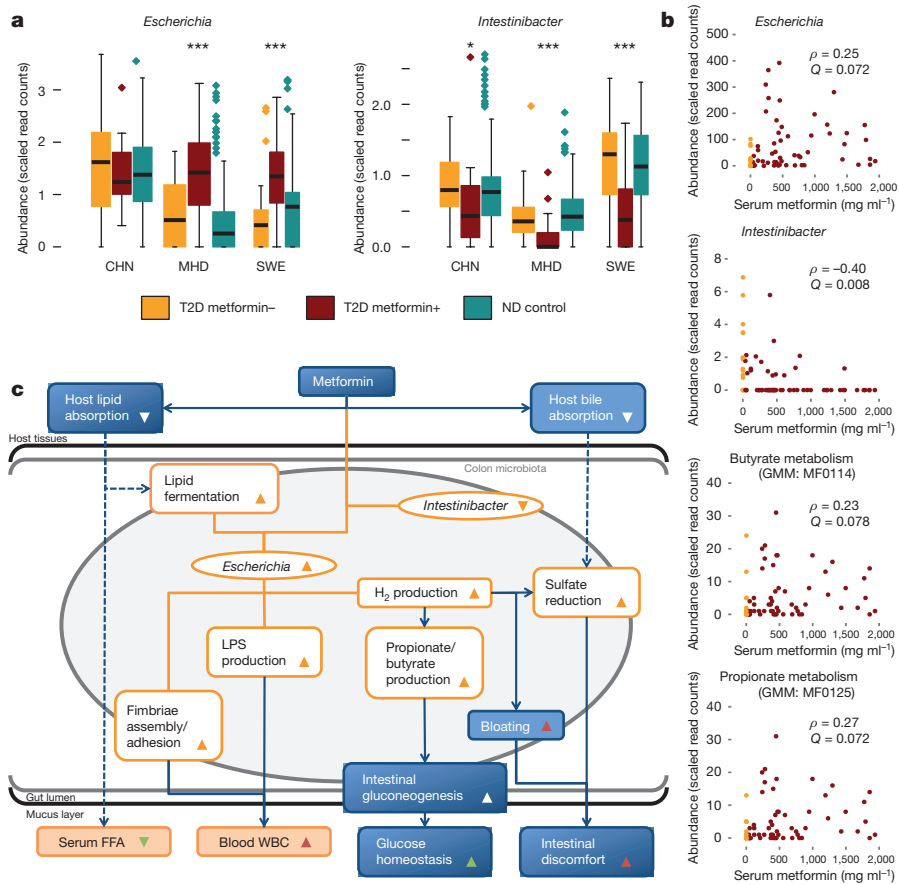


Figure 3 | Impact of metformin on the human gut microbiome.

Characterization of the microbially mediated therapeutic and adverse effects of metformin. **a**, Gut microbial shifts under metformin treatment. Metformin treatment significantly (study-source-adjusted Kruskal–Wallis test and post-hoc Mann–Whitney U -test, *FDR < 0.05, ***FDR < 0.001) increases *Escherichia* and lowers *Intestinibacter* abundance. Box plots show median/quartile abundances, whiskers extend to $1.58 \times$ interquartile range/ \sqrt{n} , for T2D metformin+ ($n_{\text{CHN}} = 15$, $n_{\text{MHD}} = 58$, $n_{\text{SWE}} = 20$), T2D metformin– ($n_{\text{CHN}} = 56$, $n_{\text{MHD}} = 17$, $n_{\text{SWE}} = 33$) and ND control ($n_{\text{CHN}} = 185$, $n_{\text{MHD}} = 277$, $n_{\text{SWE}} = 92$) gut metagenome samples. **b**, Correlations between serum levels of metformin and gut microbiota in Danish MetaHIT samples, including short-chain fatty acid production modules. Serum metformin levels of T2D patients ($n = 75$ gut metagenomes) are significantly (Spearman FDR < 0.1) positively correlated with *Escherichia* abundance, and in significant negative correlation with *Intestinibacter* abundance. Bacterial gene function modules for butyrate and propionate

production increase in abundance as serum metformin levels increase. Dot markers are shown for all MHD samples for which serum metformin concentration was measured. Metformin-untreated T2D samples (serum concentration < 10 mg ml⁻¹) are shown in orange, treated samples in dark red. Spearman coefficients (ρ ; calculated for treated samples only) and FDRs (Q) are shown. **c**, Microbial shifts under metformin treatment contribute to improved glucose control and to adverse effects. Schematic illustration of gut microbial changes and their impact on host health. Observed associations (orange lines) between microbial taxa abundances (orange ellipses), microbial functional potential (orange boxes), and blood values (filled orange boxes) and metformin treatment are linked with literature-derived metformin- or microbiota-induced host physiological effects (blue boxes and arrows; dashed arrows indicate hypothesized causality). Drug–host–microbiota interactions can contribute to previously described therapeutic (green triangles) and side (red triangles) effects of metformin treatment.

Our study of T2D illustrates the need to disentangle specific disease dysbioses from effects of treatment on the human-associated microbiota. The importance of this point was further shown by the fact that the previously reported high accuracy^{3,4} of gut microbial signatures for identifying patients with treatment-unstratified T2D decreased markedly when considering a large set of metformin-naive patients only, highlighting a general need to bear treatment regimens in mind both when developing and applying microbiome-based diagnostic and prognostic tools for common disorders or their pre-morbidity states.

Online Content Methods, along with any additional Extended Data display items and Source Data, are available in the online version of the paper; references unique to these sections appear only in the online paper.

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- Shreiner, A. B., Kao, J. Y. & Young, V. B. The gut microbiome in health and in disease. *Curr. Opin. Gastroenterol.* **31**, 69–75 (2015).
- Cho, I. & Blaser, M. J. The human microbiome: at the interface of health and disease. *Nature Rev. Genet.* **13**, 260–270 (2012).

- Qin, J. *et al.* A metagenome-wide association study of gut microbiota in type 2 diabetes. *Nature* **490**, 55–60 (2012).
- Karlsson, F. H. *et al.* Gut metagenome in European women with normal, impaired and diabetic glucose control. *Nature* **498**, 99–103 (2013).
- Schellenberg, E. S., Dryden, D. M., Vandermeer, B., Ha, C. & Korownyk, C. Lifestyle interventions for patients with and at risk for type 2 diabetes: a systematic review and meta-analysis. *Ann. Intern. Med.* **159**, 543–551 (2013).
- Larsen, N. *et al.* Gut microbiota in human adults with type 2 diabetes differs from non-diabetic adults. *PLoS ONE* **5**, e9085 (2010).
- Zhang, X. *et al.* Human gut microbiota changes reveal the progression of glucose intolerance. *PLoS ONE* **8**, e71108 (2013).
- de Vos, W. M. & Nieuwdorp, M. Genomics: A gut prediction. *Nature* **498**, 48–49 (2013).
- Pernicova, I. & Korbonits, M. Metformin—mode of action and clinical implications for diabetes and cancer. *Nat. Rev. Endocrinol.* **10**, 143–156 (2014).
- Shin, N. R. *et al.* An increase in the *Akkermansia* spp. population induced by metformin treatment improves glucose homeostasis in diet-induced obese mice. *Gut* **63**, 727–735 (2014).
- Napolitano, A. *et al.* Novel gut-based pharmacology of metformin in patients with type 2 diabetes mellitus. *PLoS ONE* **9**, e100778 (2014).
- Le Chatelier, E. *et al.* Richness of human gut microbiome correlates with metabolic markers. *Nature* **500**, 541–546 (2013).
- Sato, J. *et al.* Gut dysbiosis and detection of “live gut bacteria” in blood of Japanese patients with type 2 diabetes. *Diabetes Care* **37**, 2343–2350 (2014).

14. Cabreiro, F. *et al.* Metformin retards aging in *C. elegans* by altering microbial folate and methionine metabolism. *Cell* **153**, 228–239 (2013).
15. Gerritsen, J. *et al.* Characterization of *Romboutsia ilealis* gen. nov., sp. nov., isolated from the gastro-intestinal tract of a rat, and proposal for the reclassification of five closely related members of the genus *Clostridium* into the genera *Romboutsia* gen. nov., *Intestinibacter* gen. nov., *Terrisporobacter* gen. nov. and *Asaccharospora* gen. nov. *Int. J. Syst. Evol. Microbiol.* **64**, 1600–1616 (2014).
16. Song, Y. L., Liu, C. X., McTeague, M., Summanen, P. & Finegold, S. M. *Clostridium bartlettii* sp. nov., isolated from human faeces. *Anaerobe* **10**, 179–184 (2004).
17. Messori, S., Trevisi, P., Simongiovanni, A., Priori, D. & Bosi, P. Effect of susceptibility to enterotoxigenic *Escherichia coli* F4 and of dietary tryptophan on gut microbiota diversity observed in healthy young pigs. *Vet. Microbiol.* **162**, 173–179 (2013).
18. Czyzyk, A., Tawecki, J., Sadowski, J., Ponikowska, I. & Szczepanik, Z. Effect of biguanides on intestinal absorption of glucose. *Diabetes* **17**, 492–498 (1968).
19. Winter, S. E. *et al.* Host-derived nitrate boosts growth of *E. coli* in the inflamed gut. *Science* **339**, 708–711 (2013).
20. Everard, A. *et al.* Cross-talk between *Akkermansia muciniphila* and intestinal epithelium controls diet-induced obesity. *Proc. Natl Acad. Sci. USA* **110**, 9066–9071 (2013).
21. Lee, H. & Ko, G. Effect of metformin on metabolic improvement and gut microbiota. *Appl. Environ. Microbiol.* **80**, 5935–5943 (2014).
22. De Vadder, F. *et al.* Microbiota-generated metabolites promote metabolic benefits via gut-brain neural circuits. *Cell* **156**, 84–96 (2014).
23. Croset, M. *et al.* Rat small intestine is an insulin-sensitive gluconeogenic organ. *Diabetes* **50**, 740–746 (2001).

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Author Contributions O.P., S.D.E. and P.B. devised the project, designed the study protocol and supervised all phases of the project. T.N., T.H., T.J., H.V., J.L. and O.P. carried out patient phenotyping and clinical data analyses. T.N. and F.L. performed sample collection and DNA extraction. J.D. supervised DNA extraction, J.W., K.K. supervised DNA sequencing and gene profiling, A.Y.V. and R.H. performed additional microbial DNA extraction and amplicon sequencing. J.R., H.B.N., S.B., S.D.E., P.B. and O.P. designed and supervised the data analyses. K.F., F.H., G.F., E.L.C., S.S., E.P., S.S.-V., V.G., H.K.P., M.A., P.I.C., J.R.K. and H.B.N. performed the data analyses. K.F., F.H., T.N., P.B., S.D.E. and O.P. wrote the paper. All authors contributed to data interpretation, discussions and editing of the paper. All authors are members of the MetaHIT consortium. Additional consortium members contributed to the design and execution of the study.

Author Information Raw nucleotide data can be found for all samples used in the study in the Sequence Read Archive (accession numbers: SRA045646 and SRA050230, CHN samples) and the European Nucleotide Archive (accession numbers: ERP002469, SWE samples; ERA000116, ERP003612, ERP002061 and ERP004605, MHD samples). Reprints and permissions information is available at www.nature.com/reprints. The authors declare no competing financial interests. Readers are welcome to comment on the online version of the paper. Correspondence and requests for materials should be addressed to S.D.E. (dusko.ehrlich@jouy.inra.fr) or P.B. (bork@embl.de) or O.P. (oluf@sund.ku.dk).

METHODS

No statistical methods were used to predetermine sample size.

Danish MetaHIT diabetic study. *Patient recruitment, enrolment and processing.* Patients with T2D were either recruited from the Inter99 study population²⁴ or from the out-patient clinic at Steno Diabetes Center, Gentofte, Denmark. Patients with known T2D were included if the patient had clinically defined T2D on the day of examination according to the WHO definition²⁵. Inclusion criteria were fasting serum C-peptide above 200 pmol l⁻¹ and negative testing for serum glutamic acid decarboxylase (GAD) 65 antibodies (to exclude T1D, latent autoimmune diabetes in adults), no secondary forms of diabetes like chronic pancreatitis diabetes or syndromic diabetes, no antibiotic treatment 2 months before inclusion, and no known gastro-intestinal diseases, no previous bariatric surgery or medication known to affect the immune system.

All patients with T1D were recruited from the out-patient clinic at Steno Diabetes Center, Gentofte, Denmark ($n = 31$). Inclusion criteria were dependence on insulin treatment from time of diagnosis, fasting serum C-peptide below 200 pmol l⁻¹, glycated haemoglobin (HbA1c) above 8.0% (64 mmol l⁻¹) to ensure current hyperglycaemia, T1D duration and dependence on insulin treatment > 5 years, no antibiotic treatment at least 2 months before inclusion, and no known gastrointestinal diseases. All study participants were of North European ethnicity.

The study participants were examined on 2 days that were approximately 14 days apart. On the first day, study participants were examined after an overnight fast. Height was measured without shoes to the nearest 0.5 cm, and weight was measured without shoes and wearing light clothes to the nearest 0.1 kg. Hip and waist circumference was measured using a non-expandable measuring tape to the nearest 0.5 cm. Waist circumference was measured midway between the lower rib margin and the iliac crest. Hip circumference was measured as the largest circumference between the waist and the thighs. Blood pressure was assessed while the participant was lying in an up-right position after at least 5 min of rest using a cuff of appropriate size (A&D, UA-787 plus digital or A&D, UA-779). Blood pressure was measured at least twice and the average of the measurements was calculated. On the second day of examination, all participants provided a stool sample which was immediately frozen after home collection and stored at -80°C .

Information on medication status was obtained by questionnaire and interview on the first day of examination. Of the 75 T2D patients, 10 patients (13%) received no hyperglycaemic medications and 58 patients (77%) received the biguanide metformin; of these 75 T2D patients, 28 patients (37%) received metformin as the only anti-hyperglycaemic medication, 10 patients (13%) received sulfonylurea alone or in combination with metformin, 14 patients (19%) received a combination of oral antidiabetic drugs and insulin treatment and 4 patients (5%) were on insulin treatment only. Eleven patients (15%) received dipeptidyl peptidase-4 (DPP4) inhibitors or glucagon-like peptide-1 (GLP1), all of them in combination with metformin. Patients were reported as receiving anti-hypertensive treatment if at least one of the following drugs was reported: spironolactone, thiazides, loop diuretics, beta blockers, calcium channel blockers, moxonidine or drugs affecting the renin-angiotensin system ($n = 55$ for T2D (73%) and $n = 23$ (74%) for T1D). Patients receiving statins, fibrates and/or ezetimibe were reported as receiving lipid-lowering medication ($n = 56$ for T2D (75%; all on statin treatment), and $n = 24$ for T1D (77%; 74% on statin treatment)). All T1D patients were on insulin treatment as their only blood glucose lowering treatment.

All biochemical analyses were performed on blood samples drawn in the morning after an over-night fast of at least 10 h. Plasma glucose was analysed by a glucose oxidase method (Granustest, Merck) with a detection limit of 0.11 mmol l⁻¹ and intra- and interassay coefficients of variation (CV) of <0.8% and <1.4%, respectively. HbA1c was measured on G7 HPLC Analyzer (Tosoh) by ion-exchange high-performance liquid chromatography. Serum C-peptide was measured using a time-resolved fluoroimmunoassay with the AutoDELFIA C-peptide kit (PerkinElmer, Wallac), with a detection limit of 5 pmol l⁻¹ and intra- and interassay CV of <4.7% and <6.4%, respectively. Serum insulin (excluding des and intact proinsulin) was measured using the AutoDELFIA insulin kit (PerkinElmer, Wallac) with a detection limit of 3 pmol l⁻¹ and with intra- and interassay CV of <3.2% and <4.5%, respectively. Plasma cholesterol, plasma high-density lipoprotein cholesterol and plasma triglycerides were all measured on Vitros 5600 using reflect-spectrophotometrics. Plasma low-density lipoprotein cholesterol was calculated using Friedewald's equation. Blood leukocytes and white blood cell differential count were measured on Sysmex XS 1000i using flow cytometrics. Plasma metformin was determined by high performance liquid chromatography followed by tandem mass spectrometry. Briefly, the proteins were precipitated

with acetonitrile containing the deuterated internal standard, metformin-d6, hydrochloride and the supernatant diluted by acetonitrile. The analysis was performed on a Waters Acquity UPLC I-class system connected to a Xevo TQ-S tandem mass spectrometer in electrospray positive ionization mode. Separation was achieved on a Waters XBridgeT BEH Amide 2.5- μm column and gradient elution with 100 mM ammonium formate (pH 3.2), and with acetonitrile. The multiple reaction monitoring transitions used for metformin and metformin-d6 were 130.2 > 71.0 and 136.2 > 60.0. Calibrators were prepared by spiking drug-free serum with metformin to a concentration of 2,000 ng ml⁻¹. B12 was measured using Vitros Immunodiagnostic Products. GAD65 was measured on serum samples by a sandwich ELISA (RSR Ltd.). Inter- and intra-assay CV were < 16.6% and < 6.7% respectively, and with a detection limit of 0.57 U ml⁻¹.

Stool samples were obtained at the homes of each participant and samples were immediately frozen by storing them in their home freezer. Frozen samples were delivered to Steno Diabetes Center using insulating polystyrene foam containers, and then they were stored at -80°C until analysis. The time span from sampling to delivery at the Steno Diabetes Center was intended to be as short as possible and no more than 48 h.

A frozen aliquot (200 mg) of each faecal sample was suspended in 250 μl of guanidine thiocyanate, 0.1 M Tris, pH 7.5, and 40 μl of 10% *N*-lauroylsarcosine. Microbial DNA extraction was then performed as previously described¹². The DNA concentration and its molecular size were estimated using nanodrop (Thermo Scientific) and agarose gel electrophoresis.

Generation and availability of metagenomic samples. Already available Danish metagenomic samples were those reported in ref. 26 and references therein (excluding 14 samples removed due to average read length below 40 nucleotides, and with 5 Chinese and 21 Swedish samples with less than the rarefaction threshold of 7 million reads in total excluded from functional profile or diversity analyses), with newly sequenced samples deposited in the European Bioinformatics Institute Sequence Read Archive under accession ERP004605.

All information on Swedish samples was retrieved from previously published data⁴. In addition to published data on Chinese individuals³, we retrieved information on metformin treatment in a subset of 71 Chinese T2D patients. One-hundred and twelve samples from ref. 3 lacked metformin treatment metadata and were therefore discarded, except for measuring differences between the country data sets disregarding treatment or diabetic status. Characteristics of all study participants included in the present protocol are given in Supplementary Table 1.

Validation cohort recruitment and sample processing. Additional Danish T2D patients were recruited at the Novo Nordisk Foundation Center for Basic Metabolic Research, University of Copenhagen throughout 2014 as a part of the ongoing MicroDiab study (<http://metabol.ku.dk/research-project-sites/microdiab/>). T2D patients were included in the study if the time of T2D diagnosis was less than 5 years ago, they were between 35 and 75 years of age, Caucasian and they had not received antibiotics within the past 4 months of inclusion. In total, 30 T2D patients (21 male and 9 female) were identified. Faecal samples were collected at the home of the patients, followed by immediate freezing of samples in home freezers, and transport of samples to the hospital stored on dry ice. The samples were stored at -80°C until DNA extraction. Information of medication was obtained from questionnaires. In total, 21 (70%) of the T2D patients received metformin.

Ethics statement. All individuals in both the Danish MetaHIT study and the Danish validation study gave written informed consent before participation in the studies. Both studies were approved by the Ethical Committees of the Capital Region of Denmark (MetaHIT study: HC-2008-017; validation study: H-3-2013-102). Both studies were conducted in accordance with the principles of the Declaration of Helsinki.

Construction of a non-redundant metagenomic reference gene catalogue. Illumina shotgun sequencing was applied to DNA extracted from 620 faecal samples originating from the MetaHIT project (Supplementary Table 1). Raw sequencing data were processed using the MOCAT (version 1.1) software package²⁷. Reads were trimmed (option `read_trim_filter`) using a quality and length cut-off of 20 and 30 bp, respectively. Trimmed reads were subsequently screened against a custom database of Illumina adapters (option `screen_fastfile`) and the human genome version 19 using a 90% identity cut-off (option `screen`). The resulting high-quality reads were assembled (option `assembly`) and assemblies revised (option `assembly_revision`). Genes were predicted on scaffolds with a minimum length of 500 bp (option `gene_prediction`).

Predicted protein-coding genes with a minimum length of 100 bp were clustered at 95% sequence identity using Cd-hit (version 4.6.1)²⁸ with parameters set to: `-c 0.95`, `-G 0`, `-a 0.9`, `-g 1`, `-r 1`. The representative genes of the resulting clusters were 'padded' (that is, extended up to 100 bp at each end of the sequence using

the sequence information available from the assembled scaffolds), resulting in the final reference gene catalogue used in this study.

The reference gene catalogue was functionally annotated using SmashCommunity²⁹ (version 1.6) after aligning the amino acid sequence of each gene to the KEGG³⁰ (version 62) and eggNOG³¹ (version 3) databases.

Profiling of metagenomic samples. Raw insert (sequenced fragments of DNA represented by single or paired-end reads) count profiles were generated using MOCAT²⁷ by mapping high-quality reads from each metagenome to the reference gene catalogue (option screen) using an alignment length and identity cut-off of 45% and 95%, respectively. For each gene, the number of inserts that matched the protein-coding region was counted. Counts of inserts that mapped with the same alignment score to multiple genes were distributed equally among them. Taxonomic abundances were computed at the level of metagenomic operational taxonomic units (mOTUs)³², normalized to the length of the concatenated marker genes for each mOTU to yield the abundances used for the study, and subsequently binned at broader taxonomic levels (genus, family, class, etc.).

Rarefaction of metagenomic data and microbial diversity measurements. For all metagenome-derived measures except the mOTU taxonomic assignments, read counts were 'rarefied' in order to avoid any artefacts of sample size on low-abundance genes. Rarefied matrices were obtained as follows. Data matrices were rarefied to 7 million reads per sample. This threshold was chosen to include most samples, but 5 Chinese and 21 Swedish samples were excluded due to having less than 7 million reads per sample. Rarefactions were performed using a C++ program developed for the Tara project³³. In total we performed 30 repetitions, and in each of these we measured the richness, evenness, chao1 and Shannon diversity metrics within a rarefaction. The median value of these was taken as the respective diversity measurement for each sample. The first of 30 rarefactions of each sample were used to create a rarefied gene abundance matrix and KEGG orthologue abundance profiles were calculated by summing the rarefied abundance of genes annotated to the respective KEGG orthologue gene.

Metagenomic species (MGS) construction. Clustering of the catalogue genes by co-abundance, as described in ref. 34, defined 10,754 co-abundance gene groups (CAGs) with very high correlations (Pearson correlation coefficient > 0.9). The 925 largest of these, with more than 700 genes, were termed metagenomic species (MGS). The abundance profiles of the CAGs and MGSs were determined as the medium gene abundance (downsized to 7 million reads per sample) throughout the samples. Furthermore, the CAGs and MGS were taxonomically annotated by sequence similarity to known reference genomes.

Functional annotation/binning of metagenomes. To avoid drawing false conclusions about gut microbial functions from high abundance of single genes remotely homologous to members of a functional pathway, we used an approach that required presence of multiple pathway members. Functional pathway abundance was calculated from gene catalogue KEGG orthologue annotation and MGS abundances per sample. Thus KEGG orthologues present in each MGS were used to determine for that CAG/MGS which functional modules were represented within its genetic repertoire. This required that >90% of KEGG orthologues necessary for the completion of a reaction pathway should be present, when also taking alternative enzymatic pathways into account. The module abundance within a sample was calculated from CAG abundance in each respective sample, summing over all CAGs which had the module present. Rarefied median coverages of CAG/MGS were used, so no further normalization of the module abundance matrix was required. Abundance of genetic potential falling under the same higher-order functional levels was calculated by summing up all abundances of the lower-level functional modules within each sample.

Existing functional annotation databases cover gut metabolic pathways relatively poorly. To account for this, a number of additional bacterial gene functional modules were curated and annotated, extending the KEGG system; these are referred to in result tables as GMMs (gut microbial modules) and were previously described in ref. 12.

16S amplicon processing. 16S amplicons from frozen samples were sequenced 300 bp and 200 bp paired-end reads using an Illumina miSeq machine. We used the LotuS³⁵ pipeline in short amplicon mode with default quality filtering, clustering and denoising operational taxonomic units (OTUs) with UPARSE³⁶, removing chimaeric OTUs against the RDP reference database (http://drive5.com/uchime/rdp_gold.fa) with uchime³⁷, merging reads with FLASH³⁸ and assigning a taxonomy against the SILVA 119 rRNA database³⁹, and further refined by BLAST searches against the NCBI rRNA database⁴⁰ to identify *Intestinibacter* OTUs, using the following LotuS command line options: '-p miSeq -refDB SLV -doBlast blast -amplicon_type SSU -tax_group bacteria -derepMin 2 -CL 2 -thr 14'.

Univariate tests of taxonomic or functional abundance differences. Microbial taxa where mean abundance over all samples was less than 30 reads, or that were

present in less than 3 samples, were excluded from univariate and classifier analyses. All abundances were normalized by total sample sum. For module tables, no feature filters were used except requiring the module to be present in at least 20 samples. Filtered data tables were made available online (<http://vm-lux.embl.de/~forslund/t2d/>).

Univariate testing for differential abundances of each taxonomic unit between two or more groups was tested using Mann-Whitney-*U* or Kruskal-Wallis tests, respectively, corrected for multiple testing using the Benjamini-Hochberg false discovery rate control procedure (*Q* values)⁴¹. Post-hoc statistical testing for significant differences between all combinations of two groups was conducted only for taxa with abundances significantly different at $P < 0.2$. Wilcoxon rank-sum tests were calculated for all possible group combinations and corrected for multiple testing again using the Benjamini-Hochberg false discovery rate, as implemented in R. When controlling for potential confounders such as source study, we used blocked 'independence_test' function calls with options 'ytrafo = rank, teststat = scalar' for blocked WRST and 'ytrafo = rank, teststat = quad' for blocked Kruskal-Wallis test, as implemented in the COIN software package⁴² for R. Similarly, we applied these independence tests in the framework of post-hoc testing as described above.

Analysis of correlations between taxonomic or functional features, community diversity indices and sample metadata variables were conducted using Spearman correlation tests as implemented in R, and corrected for multiple testing using the Benjamini-Hochberg false discovery rate control procedure. To control for confounders such as source study in univariate correlation analyses, blocked Spearman tests as implemented in COIN (settings 'independence_test', options 'ytrafo = rank, xtrafo = rank, distribution = asymptotic') were used.

In some analyses, taxa were corrected for the influence of a continuous confounder variable such as microbial community richness; in these cases, the residual of a linear model between normalized log-transformed taxa abundances and overall sample gene richness was used to correct for the confounding variable. Power analysis was conducted by randomly subsampling to a given sample number, repeated 5 times to achieve robust results.

Ordinations and multivariate tests. All ordinations (NMDS, dbRDA) and subsequent statistical analyses were calculated using the R package vegan⁴³ using Canberra distances on normalized taxa abundance matrices, then visualized using the ggplot2 R package⁴⁴. Community differences were calculated using a permutation test on the respective NMDS reduced feature space, as implemented in vegan.

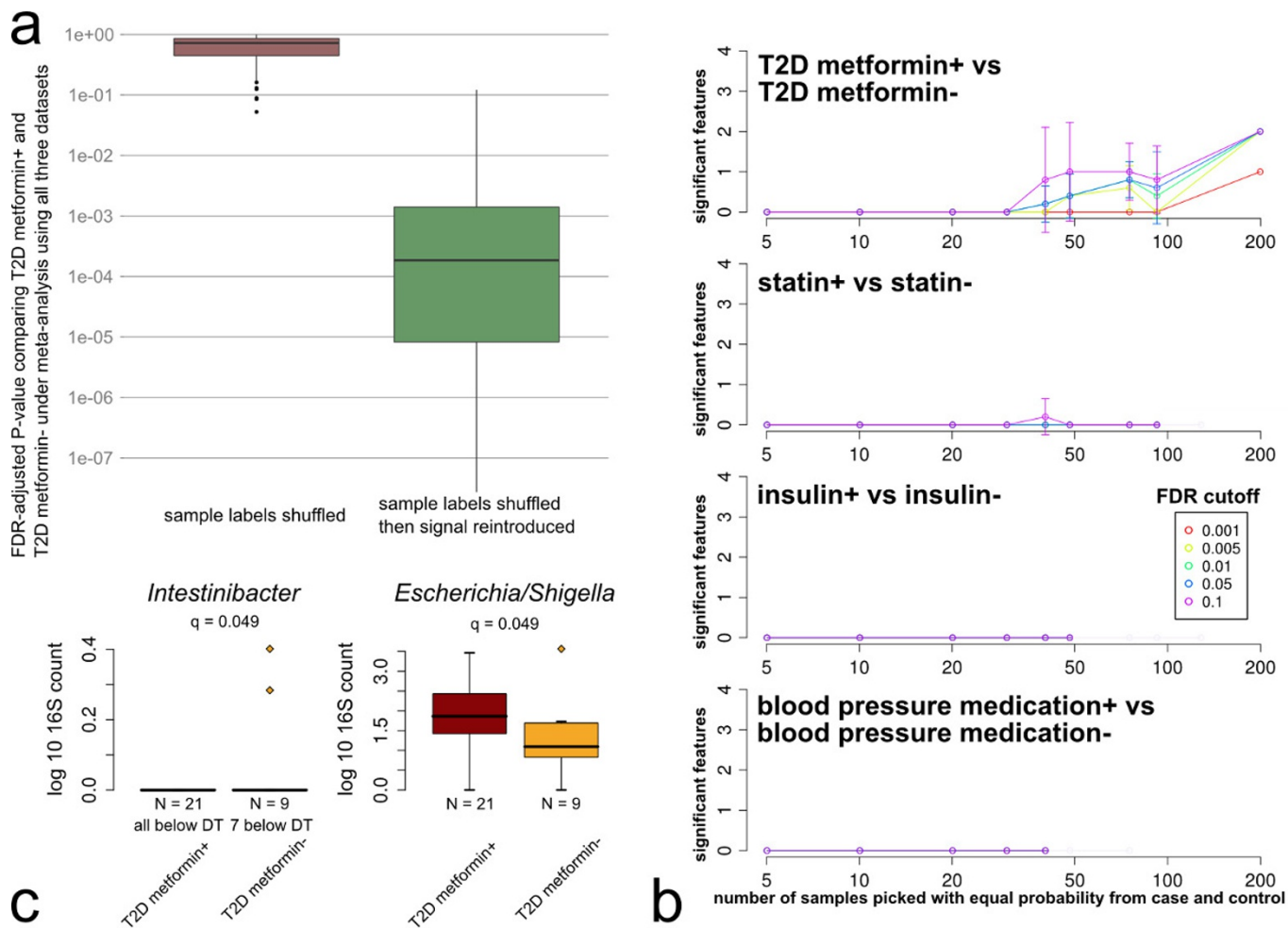
Furthermore, we calculated intergroup differences for the microbiota using PERMANOVA⁴⁵ as implemented in vegan. This test compares the intragroup distances to the intergroup distances in a permutation scheme and from this calculates a *P* value. For all PERMANOVA tests, we used 2×10^5 randomizations and a normalized genus-level mOTU abundance matrix, using Canberra intersample distances. PERMANOVA post-hoc *P* values were corrected for multiple testing using the Benjamini-Hochberg false discovery rate control procedure. Analysis of variance broken down by cohort, treatment and disease status was conducted by fitting these distances to a linear model of sample metadata distances, as further described in Supplementary Discussion 3.2.

Classifier construction and evaluation. To create classifiers for separating samples from different subsets, an L1 restricted LASSO using the R glmnet package⁴⁶ was carried out to test for an optimal value of lambda (number of features to be used in the final predictor) in a fivefold cross-validated and internally fourfold cross-validated LASSO run on all data. After this, the previously determined value of lambda was manually controlled for number of features used against the root mean square error of the classifier. In a fivefold cross-validation, an independent LASSO classifier was trained on 4/5 of the data using the previously determined value of lambda, and response values were predicted on 1/5 of the data. LASSO models with a Poisson response type were used in all cases.

Binary classifications between T2D and ND control samples were performed with an R reimplementation of the robust recursive feature elimination support vector machine (rRFE-SVM)⁴⁷ procedure. The SVM was performed in an outer cross-validation scheme on 4/5 of the data. Of these, 90% were randomly selected 200 times in each cross-validation for the RFE, to create a feature ranking from an average over these runs. Classifier performance was validated on the remaining 1/5 of samples using the pre-established feature ranking. In case of several cohorts, the area under the receiver operating characteristic curve (ROC-AUC) scores were measured for each cohort separately.

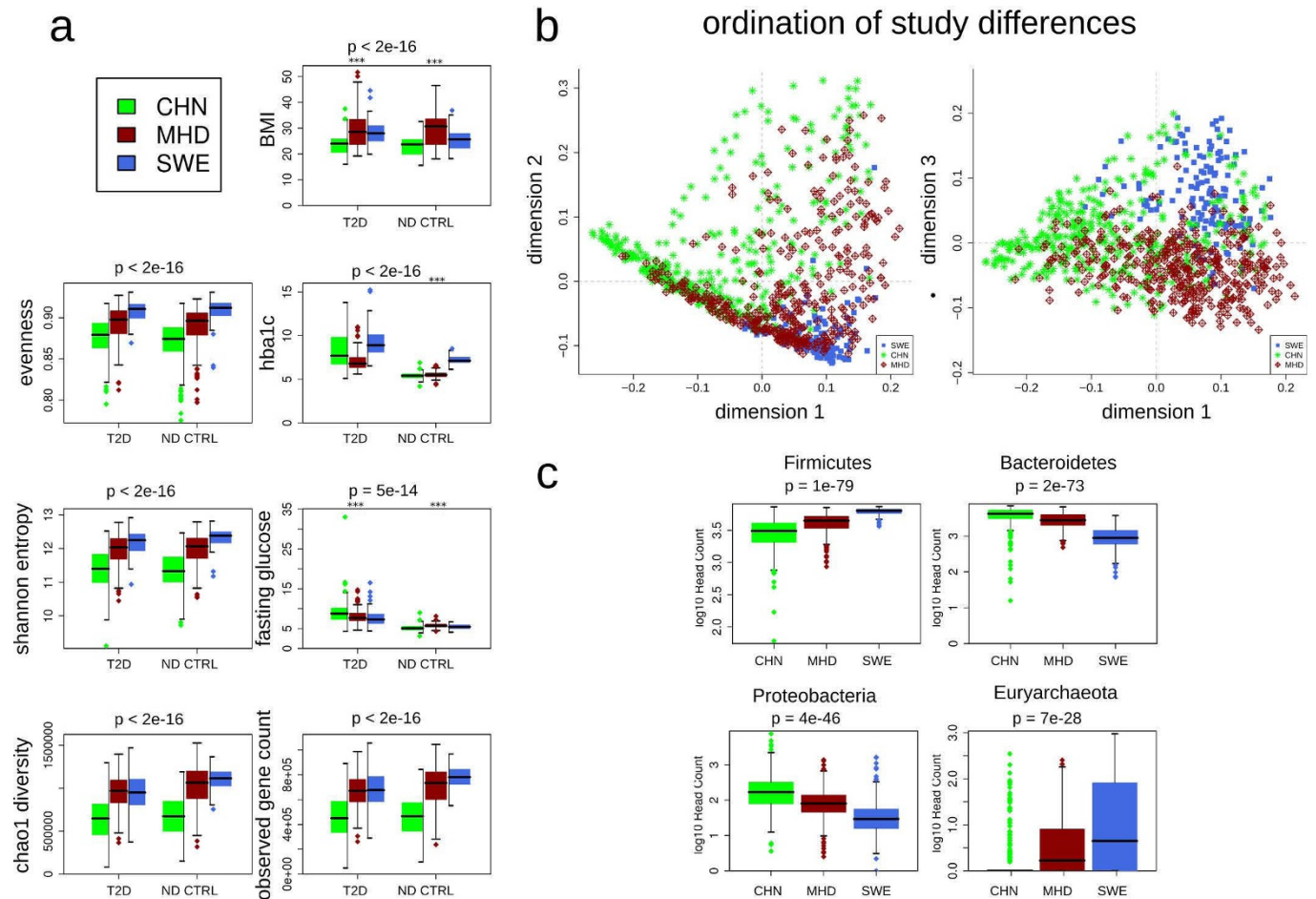
Code availability. The MGS technology has previously been described³⁴ and is available online (<http://git.dworzynski.eu/mgs-canopy-algorithm/wiki/Home>). The mOTU resource has been made publically available (<http://www.bork.embl.de/software/mOTU/>) and was analysed using MOCAT²⁷ which is also publically

- available (<http://vm-lux.embl.de/~kultima/MOCAT/>). The 16S pipeline LotuS³⁵ is freely available online (<http://psbweb05.psb.ugent.be/lotus>). The novel gene catalogue has been deposited online (http://vm-lux.embl.de/~kultima/share/gene_catalogs/620mht2d/), as have the raw amplicon sequences (<http://vm-lux.embl.de/~forslund/t2d/>). Statistical analysis and data visualization was conducted using freely available R libraries: vegan, COIN and ggplot2 and is described in more details elsewhere^{48,49}. Data matrices and R source code for replicating the central tests conducted on the data have been deposited online (<http://vm-lux.embl.de/~forslund/t2d/>).
- Evaluation of dietary habits.** A subset of the Danish study participants answered a validated food frequency questionnaire in order to obtain information on the habitual dietary habits. A complete data set was obtained for 66% of the nondiabetic individuals and 88% of T2D patients. When evaluating the dietary data, the consumed quantity was determined by multiplying portion size by the corresponding consumption frequency reported. Standard portion sizes for women and men, separately, were used in this calculation^{50,51}. All food items in the questionnaire were linked to food items in the Danish Food Composition Databank⁵². Estimation of daily intake of macro- and micronutrients for each participant was based on calculations in the software program FoodCalc version 1.3⁵³.
24. Jørgensen, T. *et al.* A randomized non-pharmacological intervention study for prevention of ischaemic heart disease: baseline results Inter99. *Eur. J. Cardiovasc. Prev. Rehabil.* **10**, 377–386 (2003).
 25. WHO. Definition, *Diagnosis and Classification of Diabetes Mellitus and its Complications. Part 1: Diagnosis and Classification of Diabetes Mellitus*. Report No. WHO/NCD/NCS/99.2 (World Health Organization, 1999).
 26. Li, J. *et al.* An integrated catalog of reference genes in the human gut microbiome. *Nature Biotechnol.* **32**, 834–841 (2014).
 27. Kultima, J. R. *et al.* MOCAT: a metagenomics assembly and gene prediction toolkit. *PLoS ONE* **7**, e47656 (2012).
 28. Li, W. & Godzik, A. Cd-hit: a fast program for clustering and comparing large sets of protein or nucleotide sequences. *Bioinformatics* **22**, 1658–1659 (2006).
 29. Arumugam, M., Harrington, E. D., Foerstner, K. U., Raes, J. & Bork, P. SmashCommunity: a metagenomic annotation and analysis tool. *Bioinformatics* **26**, 2977–2978 (2010).
 30. Kanehisa, M. *et al.* KEGG for linking genomes to life and the environment. *Nucleic Acids Res.* **36**, D480–D484 (2008).
 31. Powell, S. *et al.* eggNOG v3.0: orthologous groups covering 1133 organisms at 41 different taxonomic ranges. *Nucleic Acids Res.* **40**, D284–D289 (2012).
 32. Sunagawa, S. *et al.* Metagenomic species profiling using universal phylogenetic marker genes. *Nature Methods* **10**, 1196–1199 (2013).
 33. Sunagawa, S. *et al.* Structure and function of the global ocean microbiome. *Science* **348**, (2015).
 34. Nielsen, H. B. *et al.* Identification and assembly of genomes and genetic elements in complex metagenomic samples without using reference genomes. *Nature Biotechnol.* **32**, 822–828 (2014).
 35. Hildebrand, F. *et al.* LotuS: an efficient and user-friendly OTU processing pipeline. *Microbiome* **2**, 30 (2014).
 36. Edgar, R. C. UPARSE: highly accurate OTU sequences from microbial amplicon reads. *Nature Methods* **10**, 996–998 (2013).
 37. Edgar, R. C. *et al.* UCHIME improves sensitivity and speed of chimera detection. *Bioinformatics* **27**, 2194–2200 (2011).
 38. Magoč, T. & Salzberg, S. L. FLASH: fast length adjustment of short reads to improve genome assemblies. *Bioinformatics* **27**, 2957–2963 (2011).
 39. Quast, C. *et al.* The SILVA ribosomal RNA gene database project: improved data processing and web-based tools. *Nucleic Acids Res.* **41**, D590–D596 (2013).
 40. Madden, T. in *The NCBI Handbook [Internet]*. (eds, McEntyre J. & Ostell J.) Ch. 16 (National Center for Biotechnology Information, 2002) <http://www.ncbi.nlm.nih.gov/books/NBK21097/>
 41. Benjamini, Y. & Hochberg, Y. Controlling the false discovery rate: a practical and powerful approach to multiple testing. *J. R. Stat. Soc. Ser. A Stat. Soc.* **57**, 289–300 (1995).
 42. Hothorn, T., Hornik, K., van de Wiel, M. A. & Zeileis, A. A Lego system for conditional inference. *Am. Stat.* **60**, 257–263 (2006).
 43. Dixon, P. VEGAN, a package of R functions for community ecology. *J. Veg. Sci.* **14**, 927–930 (2003).
 44. Wickham H. *ggplot2: Elegant Graphics for Data Analysis*. (Springer, 2009).
 45. Anderson, M. J. A new method for non-parametric multivariate analysis of variance. *Austral. Ecol.* **26**, 32–46 (2001).
 46. Friedman, J. *et al.* Regularization paths for generalized linear models via coordinate descent. *J. Stat. Softw.* **33**, 1–22 (2010).
 47. Abeel, T., Helleputte, T., Van de Peer, Y., Dupont, P. & Saeyns, Y. Robust biomarker identification for cancer diagnosis with ensemble feature selection methods. *Bioinformatics* **26**, 392–398 (2010).
 48. Hildebrand, F. *et al.* A comparative analysis of the intestinal metagenomes present in guinea pigs (*Cavia porcellus*) and humans (*Homo sapiens*). *BMC Genomics* **13**, 514 (2012).
 49. Hildebrand, F. *et al.* Inflammation-associated enterotypes, host genotype, cage and inter-individual effects drive gut microbiota variation in common laboratory mice. *Genome Biol.* **14**, R4 (2013).
 50. Haraldsdóttir, J. *et al.* *Portionsstorleker - Nordiska standardportioner av mat och livsmedel* (Nordisk Ministerråd, 1998)
 51. Biltoft-Jensen, A. *et al.* *Danskernes kostvaner 2000–2002*. DFVF publication No. 11 (Danmarks Fødevareforskning, Afdeling for Ernæring, 2005)
 52. Møller, A. *et al.* Fødevaredatabanken version 5.0. *Fødevareinformatik, Institut for Fødevaresikkerhed og Ernæring, Fødevaredirektoratet* <http://www.foodcomp.dk> (2002).
 53. Lauritsen, J. FoodCalc. www.ibt.ku.dk/jesper/FoodCalc/ (2004).



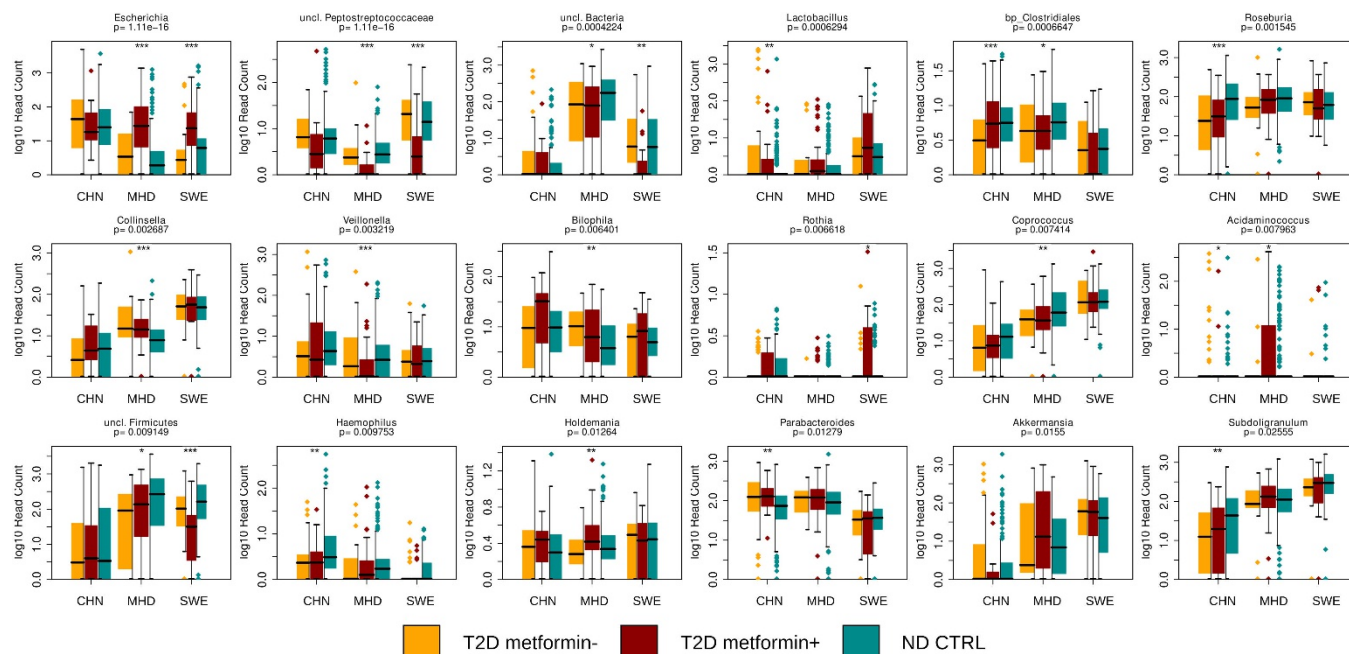
Extended Data Figure 1 | Validation of meta-analysis pipeline on simulated data. **a**, As a positive control for the meta-analysis pipeline, true signal was removed from the data by randomly reshuffling sample labels. Artificial contrast was thereafter introduced between random groups containing as many such reshuffled samples as were in the original sets of T2D metformin+ ($n_{\text{CHN}} = 15$, $n_{\text{MHD}} = 58$, $n_{\text{SWE}} = 20$) and T2D metformin- ($n_{\text{CHN}} = 56$, $n_{\text{MHD}} = 17$, $n_{\text{SWE}} = 33$) samples in each original study subset, using the genus *Akkermansia* as an example feature. Samples randomly assigned to the sets of fake ‘metformin-treated’ and ‘control’ categories had their *Akkermansia* genus abundances adjusted to match the scale of the metformin effect on *Escherichia* genus abundance reported here (metformin-treated samples were roughly 150% as likely to have non-zero abundance, with a roughly threefold higher abundance where present), while retaining their data set origin labels. The full meta-analysis pipeline (study set blocked Kruskal–Wallis test, post-hoc Wilcoxon rank-sum test) was applied to these samples. Benjamini–Hochberg-corrected P values (FDR scores/ Q values) from testing for a metformin effect on *Akkermansia* abundance are plotted in logarithmic scale on the vertical axis for 100 randomizations of the entire shuffled data set, either without (left box plot) or with (right box plot) the artificial *Akkermansia* metformin signal added after shuffling the data to remove original signal. Box plot borders show medians and quartiles, with points outside this range shown as vertical whisker lines and point markers. Whiskers extend to $1.58 \times$ interquartile range/ \sqrt{n} . Horizontal guide lines are shown for ease of visualization corresponding to different false discovery rate thresholds. For randomly reshuffled data, no significant contrast is detected as expected, whereas the artificially introduced signal is reliably detected,

roughly matching expectations from the definition of the false discovery rate itself. **b**, To investigate statistical power for the other medications tracked, five random sub-samplings were made of pairs of medicated and non-medicated samples at each increasing number of included sample pairs and the overall analysis was replicated for each. We tested each genus for significantly differential abundance between cases and controls (Kruskal–Wallis test followed by post-hoc Wilcoxon rank-sum test) at different Benjamini–Hochberg FDR significance cut-offs, which are represented by different colours. Of the total number of samples for which medication status was known, equal numbers (n) of medicated and unmedicated samples were chosen randomly in repeated iterations. This number n was varied up to its largest possible value (smallest of either number of medicated or unmedicated samples in the overall data set) and is shown on the x axis. The y axis shows the number of significant features relative to each cut-off. Error bars show ± 1 s.d. of each set of five randomized samples. **c**, The graphs show *Intestinibacter* and *Escherichia* median and quartile abundances as box plots, whiskers extend to $1.58 \times$ interquartile range/ \sqrt{n} , with samples that are extreme relative to the interquartile range shown as point markers, and with samples below detection threshold (DT) plotted at $y = 0$, in 21 additional T2D metformin+ and 9 additional T2D metformin- samples. Differences in abundance between sample categories are significant (Wilcoxon rank-sum test, Benjamini–Hochberg FDR < 0.1). All samples in which *Intestinibacter* was detected fall among the 9 out of 30 untreated rather than the 21 out of 30 metformin-treated samples, consistent with severe depletion under treatment; whereas *Escherichia* abundances increase under treatment, likewise consistent with observations from the main data set.



Extended Data Figure 2 | Differences in physiological variables and microbiome characteristics between gut metagenome sample sets. Chinese ($n = 368$), Danish MetaHIT ($n = 383$) and Swedish ($n = 145$). **a**, Several participant metadata variables are significantly different between cohorts. A subselection is shown as box plots displaying median and quartiles, with samples outside this range shown as point markers and whiskers. Whiskers extend to $1.58 \times$ interquartile range/ \sqrt{n} . **b**, In a principal coordinates analysis ordination of Bray–Curtis distances between samples on bacterial family level, clear differences between samples from the different cohorts become apparent. These are largely

explained by taxonomic differences as summarized at the phylum level. **c**, Box plots for gut microbial taxa show medians and quartiles of log-transformed read counts for mOTUs summarized at the level of bacterial genera for the three country subsets across sample categories, with samples outside this range shown as point markers and whiskers. Whiskers extend to $1.58 \times$ interquartile range/ \sqrt{n} . For all box plots, tests for significant differences (Kruskal–Wallis test adjusted for study source) were performed, with P values shown at the head of each figure. Asterisks denote statistical significance of tests done for each country subset separately (***) ($P < 0.001$).



Extended Data Figure 3 | Microbiome taxonomic composition comparison between gut metagenomes with particular focus on possible taxonomic restoration under metformin treatment for certain taxa. T2D metformin- ($n = 106$), T2D metformin+ ($n = 93$) and ND control ($n = 554$). Box plots show medians and quartiles log-transformed read counts for mOTUs summarized at the level of bacterial genera, for the

three country subsets across sample categories, with samples outside this range shown as point markers and whiskers. Whiskers extend to $1.58 \times$ interquartile range/ \sqrt{n} . Tests for significant differences (Kruskal–Wallis test adjusted for study source) were performed, with P values shown at the head of each figure. Asterisks denote statistical significance of tests for each country subset separately (* $P < 0.05$; ** $P < 0.01$; *** $P < 0.001$).

Extended Data Table 1 | Analysis of variances

a

	Degrees of freedom	Sum of squares	Explained variation	F-statistic	Pr(>F)
Treatment	1	128	3.2%	21454.01	<2E-016 ***
Disease	1	42	1.1%	7039.86	<2E-016 ***
Country	1	376	9.4%	63206.96	<2E-016 ***
Treatment x Disease	1	1	0.0%	192.62	<2E-016 ***
Treatment x Country	1	67	1.7%	11209.33	<2E-016 ***
Disease x Country	1	1	0.0%	218.97	<2E-016 ***
Treatment x Disease x Country	1	0	0.0%	22.79	0.00000181 ***
Residuals	567001	3375	84.6%		
	Total:	3990	100.0%		

b

Database OTU identifier	MWU P-value	Enriched in	Mean abundance (%)
OTU_45	0.048968332	T2D metformin+	0.803960725
OTU_1038	0.0319637913	T2D metformin-	0.000185722

Database OTU identifier	OTU_45	OTU_1038
Domain	<i>Bacteria</i>	<i>Bacteria</i>
Phylum	<i>Proteobacteria</i>	<i>Firmicutes</i>
Class	<i>Gammaproteobacteria</i>	<i>Clostridia</i>
Order	<i>Enterobacteriales</i>	<i>Clostridiales</i>
Family	<i>Enterobacteriaceae</i>	<i>Peptostreptococcaceae</i>
Genus	<i>Escherichia-Shigella</i>	<i>Intestinibacter</i>

a, The analysis of variances table shows the results of modelling the Canberra distances between T2D metformin- ($n = 106$), T2D metformin+ ($n = 93$) and ND control ($n = 554$) samples with predictor variables encoding same/different diabetes status, same/different treatment, and same/different study source/country. Fractions of explained variance are taken as fractions of sum of square deviations from the model relative to the total deviation. **b**, Bacterial taxa found significantly different in gut abundance under metformin treatment were tested for significant differential relative abundance in a separate cohort under 16S amplicon sequencing between T2D metformin+ ($n = 21$) and T2D metformin- ($n = 9$) samples (Wilcoxon rank-sum test).

CORRECTIONS & AMENDMENTS

CORRIGENDUM

doi:10.1038/nature22318

Corrigendum: Disentangling type 2 diabetes and metformin treatment signatures in the human gut microbiota

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Nature **528**, 262–266 (2015); doi:10.1038/nature15766

In the Supplementary Information to this Letter, data from two previous studies were used in the meta-analysis. However, the unit conversions used to make the data comparable were inconsistent for two of the included phenotype measures. Although this error does not affect the data used to generate the conclusions of the Letter, it might affect follow-up studies using the glycated haemoglobin (HbA1c) and serum insulin phenotypes. A revised Supplementary Table 1 ('phenotypes' tab) and corrected cohort summary statistics in Supplementary Tables 2 ('metadata' tab) and 12 ('metadata differences' tab) are provided in Supplementary Information to this Corrigendum.

In addition, in the Supplementary Information on page 6, the sentence "For clinically relevant markers, we found glycated haemoglobin A1c (HbA1c), reflecting long-term blood glucose levels, to be lowest in MHD T2D and highest in SWE ND CTRL samples ($P < 2 \times 10^{-16}$). Fasting plasma glucose levels were highest in CHN T2D patients and lowest in CHN ND CTRL, compared to corresponding SWE & MHD samples ($P = 5.6 \times 10^{-4}$)." should have read "For clinically relevant markers, we found fasting plasma glucose levels were highest in CHN T2D patients and lowest in CHN ND CTRL, compared to corresponding SWE & MHD samples ($P = 5.6 \times 10^{-4}$)." On page 18, the sentence "To convert mU/l to pmol/l for insulin levels a factor of 6.0 was used." should have read "To convert mU/l to pmol/l for insulin levels a factor of 6.945 was used."

We would like to thank the readers who brought these issues to our attention. The original Letter has not been corrected online.

Supplementary Information is available in the online version of the Corrigendum.