Family History Is Not Useful in Screening Children for Dyslexia

Emilio Ferrer¹ • Bennett A. Shaywitz² • John M. Holahan² • Sally E. Shaywitz²

Received: 25 March 2021 / Revised: 12 June 2021 / Accepted: 19 June 2021 / Published online: 21 October 2021 (© The Author(s) 2021

Abstract

Accurate and efficient early screening is important for providing effective early intervention for dyslexic readers. While family history is often considered a contributing risk factor for dyslexia, some have suggested that it could serve as a proxy for identification of dyslexia. We examined the classification accuracy of family history as a screening measure for dyslexia using an epidemiologic sample of 398 children followed from age 5 through adulthood. Sensitivity of family history for predicting dyslexia was unacceptably low for all family member groups. Moreover, results from receiver operating characteristic curves indicate that predicting dyslexia using family history does not improve the value of using an evidence-based early screening measure alone. Together, these analyses indicate that family history is inadequate as a screening measure for dyslexia; and thus, the use of positive family history as a proxy for dyslexia is unwarranted.

Keywords Dyslexia \cdot Screening \cdot Family history \cdot Reading \cdot Longitudinal studies

Accurate early screening is particularly important for providing effective early intervention for dyslexic readers, and some have suggested that family history could serve as a proxy for identification of dyslexia. Family history has been considered as a contributing risk factor for dyslexia. Scottish ophthalmologist James Hinshelwood's descriptions of dyslexia in the early twentieth century included reports that, in some cases, dyslexia occurs in families (Hinshelwood, 1900, 1911). Recent studies provide conflicting evidence regarding the potential usefulness of family history as a predictor in screening for dyslexia (Carroll et al., 2014; Dilnot et al., 2017; Thompson et al., 2015). In a high-risk sample, while family



Statement of Relevance

Accurate early screening is important for providing effective early intervention for dyslexic readers. It is in these early years of school that the slope for reading acquisition is greatest, only to plateau at a much slower rate as the child goes on in school. Some have suggested using family history as a proxy for identification of dyslexia. However, the classification accuracy of family history as a screening measure for dyslexia is unknown. Here, we show that using family history as a proxy for dyslexia is unwarranted. Using an epidemiologic sample of 398 children followed from age 5 through adulthood, we found that sensitivity of family history for predicting dyslexia was unacceptably low for all family member groups. Receiver operating characteristics (ROC) curves for family history alone, an evidence-based early screening measure alone, and the combination of screener and family history indicate that the evidence-based screener yields the best results for predicting at-risk for dyslexia while the addition of family history does not improve the value of using the screener alone. These findings indicate that family history is not effective as a screening measure for dyslexia and that family history does not improve the classification accuracy provided by an evidence-based early screening measure.

Sally E. Shaywitz sally.shaywitz@yale.edu

¹ Department of Psychology, University of California, Davis, Davis, CA, USA

² Yale Center for Dyslexia & Creativity, Department of Pediatrics, Yale University School of Medicine, 129 York Street, Suite 1P, New Haven, CT 06511, USA

history was a significant predictor of dyslexia during the preschool years, by school entry, family history no longer was predictive, and other measures, including letter knowledge and phonological awareness, were better predictors of dyslexia (Thompson et al., 2015).

Among 7- to 9-year-old children, family history was a significant predictor of reading accuracy after controlling for early speech and language (Carroll et al., 2014). In contrast, family risk of dyslexia did not predict reading readiness (a composite of word reading, letter-sound knowledge, phoneme deletion, and rapid automatized naming) once other risks were controlled (Dilnot et al., 2017). A meta-analysis of the effects of family history on reading found an average prevalence rate of dyslexia of 45% (varying between 26 and 66%) in children with a first-degree relative with dyslexia, compared to just under 12% risk of dyslexia in samples of children without such a family history, though samples of dyslexic readers without a family history were not examined (Snowling & Melby-Lervåg, 2016). None of these previous studies reported sensitivity, specificity, and area under the curve (AUC) analyses for family history as a risk factor for dyslexia.

In this report, we examine the classification accuracy of family history as a screening measure for dyslexia using a unique population, an epidemiologic sample of 398 school children representative of children entering public kindergarten, including typical readers (TR) and dyslexic readers (DR), followed longitudinally from age 5 through adulthood. Our study addresses three fundamental questions: What is the classification accuracy of family history as a screening measure for dyslexia in young children? How do dyslexia and family history compare when predicting longitudinal changes in reading skills from childhood to adolescence? What is the longitudinal predictive value of family history within typical readers (TR) and dyslexic readers (DR) considered separately?

Methods

Participants

We use data from The Connecticut Longitudinal Study, an epidemiologic sample survey of schoolchildren representative of children entering public kindergarten (Shaywitz et al., 1990). Of the 398 participants with complete data, 52.8% are females and 47.2% males. The sample contains European Americans or Whites (85.2%), African Americans or Blacks (11.8%), Asians (1.0%), Hispanics or Latinos (2.0%), and other children with unreported race or ethnicity (0.3%). The composition of this sample was similar to the racial and ethnic composition of the USA at the time of the study. All participants were primary English speakers. This cohort, assembled from a 2-stage probability sample, has been

followed longitudinally from school entry into adulthood to study the development of reading, learning, and attention (Ferrer et al., 2007, 2010, 2015; Shaywitz et al., 1990, 1992a, 1992b, 1992c, 1999). Parents or caretakers provided written consent for their children to participate in the study, and children also provided assent. The study was approved by the Institutional Review Board at Yale University and was conducted in accordance with the ethical principles that have their origin in the Declaration of Helsinki and are consistent with good clinical practices and applicable laws and regulations.

Measures

Before kindergarten entry, the participants' parents completed the Yale Children's Inventory (YCI), a comprehensive overview of the child's prenatal and perinatal history; the family's behavioral, cognitive, and medical history; the child's behavior, development, language, habits, and preschool experiences; parental education and employment; and significant life events (Shaywitz et al., 1986, 1988, 1992a, 1992b). Reading skills were measured using the WJ Reading Cluster (composite of Letter-Word Identification, Word Attack, and Passage Comprehension subtests) from the Woodcock-JohnsonPsycho-Educational Test Battery (WJ; Woodcock & Johnson, 1989), and IQ was measured using the Wechsler Intelligence Scale for Children-Revised (WISC-R; Wechsler, 1981). To compare family history and an evidence-based early screening measure for identification as at-risk for dyslexia, we used the kindergarten and first-grade teachers' ratings from the Shaywitz DyslexiaScreen (Shaywitz, 2016).

Criteria for Family History and Dyslexia

Family history criteria were established by responses to two questions from the YCI specifically related to family history of dyslexia: Did the family member have (1) trouble reading and (2) trouble spelling? Three criteria were set for trouble reading, trouble spelling, and trouble reading or spelling. The respondent provided answers (coded as 1 for yes; 0 for no) for each family member group: siblings, parents, or grandparents, for each criterion. Two variables (coded as 1 for yes; 0 for no) indicated whether any 1st-degree relatives (parent or sibling) had trouble reading, trouble spelling, and trouble reading or spelling. The last criterion indicated whether any 1st- or 2nd-degree relative (parent, sibling, or grandparent) had trouble reading or spelling. The 1st- and 2nd-degree relative definition yielded a family history positive group (FH⁺, n = 119) and a family history negative group (FH⁻, n = 279).

Dyslexia was defined using the WJ Reading Cluster scores and the WISC-R Full Scale IQ score. Dyslexic children met criteria based on low achievement (Reading Cluster Age Standard score <90) or IQ-achievement discrepancy criteria (a reading cluster >1.5 standard deviation lower than that predicted by Full Scale IQ) in grade 2 or 4 (Ferrer et al., 2015). Both definitions validly identify children as poor readers, and there is little evidence of differences between subgroups formed by one definition versus the other (Shaywitz et al., 1992a, 1992b, 1992c). This definition of dyslexia status yielded a dyslexic readers group (DR, n = 97) and a typical readers group (TR, n = 301).

Beginning with its first description over a century ago (Morgan, 1896), continuing through the early part of the twentieth century (Hinshelwood, 1917) and through the beginning of the twenty-first century (Lyon et al., 2003), dyslexia has always been defined as an unexpected difficulty in reading in a person who has the intelligence to be a much better reader. Ferrer and associates provided empiric evidence for dyslexia's unexpected nature (Ferrer et al., 2010) and recent federal law ("First Step Act,", 2018) has codified dyslexia as "an unexpected difficulty in reading for an individual who has the intelligence to be a much better reader." This definition fits some current revised methods to identify dyslexia based on discrepancy "...seriously low reading ability, average or better cognitive ability, and a standard score difference of 15 to 29 points [for likely] and 30 points or more [for very likely]" (Hammill & Allen, 2020).

The definition used in our study follows directly from the over a century of dyslexia research and conceptualizes dyslexia as an unexpected difficulty in reading. Investigators, including ourselves, have operationalized the definition to include unexpected for ability and unexpected for age. Specifically, dyslexic readers were identified by an observed Woodcock–Johnson Reading Cluster score 1.5 standard errors below the score predicted from their Full Scale IQ or with a Reading Cluster score below 90. Both of these definitions validly identify children as dyslexic, and there is little evidence of differences between subgroups of children formed with one criterion versus the other (Shaywitz et al., 1992c). This operational definition has been used by us in many previous peer-reviewed publications (Estrada et al., 2018; Ferrer et al., 2015; Herrera-Araujo et al., 2017) (S. E. Shaywitz et al., 2003).

Statistical Analysis

To examine the classification accuracy of the family history definitions (trouble reading, trouble spelling, and trouble reading or spelling) for each family member group as the sole predictor of reader group status (TR and DR), we performed classification analyses yielding the following statistics: sensitivity, specificity, area under the curve (AUC), 95% confidence intervals, and p value from ROC (receiver operating characteristic curve) analysis.

To characterize the normative differences from grades 1 to 9 for the WJ Reading Cluster scores, we carried out a repeated-measures ANOVA, with reader group and family history serving as between-subject effects, and grades (1, 3, 5, 7, and 9) as the repeated measure. The two between-group main effects comparing the overall reader groups (TR vs. DR) and the overall (FH⁻ and FH⁺) groups were tested using the pooled within-subjects standard deviation as the denominator for the calculated Hedge's *g* effect sizes (Clearinghouse, 2017). Two simple main effects were calculated for the differences between (FH⁻ and FH⁺) groups within the DR and the TR groups using the pooled within-subject standard deviation, the F statistic of the multivariate model, and Hedge's *g* effect sizes. To control for type I error for the four comparisons, the criterion for statistical significance was set at p = .05/4 = .0125.

Finally, to examine the predictive utility of the family history variable (positive family history (for any 1st- and 2nddegree relative)) alone, an early screening measure (Shaywitz DyslexiaScreen) alone, and the combined screen and family history variables on dyslexia status, we used ROC curve analysis. Results of the three models for kindergarten and first grade are reported in Table 3.

Results

Table 1 reports analyses for sensitivity, specificity, and area under the curve (AUC using receiver operating characteristic curves). True positives (TP, dyslexic readers classified as dyslexic by positive family history), true negatives (TN, typical readers classified as typical), false positives (FP, typical readers classified as dyslexic), and false negatives (FN, dyslexic readers classified as typical) are presented for each family member grouping (parents, siblings, 1st-degree relatives, and grandparents). These analyses are carried out considering family history definitions based on trouble reading, trouble spelling, and the combined trouble reading or spelling.

Sensitivities (correct classification of DRs as dyslexic) range from 5% for grandparents using trouble reading as the family history criterion to a maximum of 51% for the combination of 1st- and 2nd-degree relatives, using trouble with reading or spelling as the criterion for positive family history. Specificities (correct classification of TRs as typical readers) were substantially higher, ranging from 77% for 1st- and 2nddegree relatives using trouble with reading or spelling as the criterion for positive family history to 97% for 1st-degree relatives using trouble spelling as the family history criterion. The largest observed AUC was 80% (p < .001), which was obtained for the sibling family member group using trouble spelling. The smallest AUCs were obtained for the grandparent family member group: 55%, p = .51, for trouble reading; 59%, p = .11, for trouble reading or spelling; and 62%, p = .05, for trouble spelling.

 Table 1 Family history

 classification analysis of dyslexia

	п			%				
	ТР	TN	FP	FN	Sensitivity	Specificity	AUC	р
Trouble reading								
Parents	19	285	16	78	20	95	70	<.001
Siblings	29	279	22	68	30	93	74	<.001
1st Deg. Rel.	38	264	37	59	39	88	73	<.001
Grandparents	5	290	11	92	5	96	55	.55
Trouble spelling								
Parents	24	271	30	73	25	90	67	<.001
Siblings	27	289	12	70	28	96	80	<.001
1st Deg. Rel.	41	261	40	56	42	87	74	<.001
Grandparents	10	287	14	87	10	95	62	.051
Trouble reading or spelli	ng							
Parents	27	266	35	70	28	88	67	<.001
Siblings	34	273	28	63	35	91	75	<.001
1st Deg. Rel.	47	241	60	50	48	80	72	<.001
Grandparents	11	282	19	86	11	94	59	.120
1st + 2nd Deg. Rel.	49	231	70	48	51	77	70	<.001

Note: *TP*, true positive; *TN*, true negative; *FP*, false positive; *FN*, false negative; *AUC*, area under the curve; *p*, Fisher's exact probability

Results of repeated measures ANOVA examining longitudinal differences in reading scores for the reading and family history groups from grades 1 to 9 are presented in Table 2. The overall difference in reading achievement between typical and dyslexic readers (Figure 1, panel A) was very large and statistically significant (p < .001, effect size (ES) = 1.42 standard deviations). The overall difference between the family history groups (FH⁺ = having a 1st- or 2nd-degree relative with dyslexia; FH⁻ = not) (Figure 1, panel B) was also statistically

Table 2 Repeated-measures analysis of variance

Source	df	MS	F	р
Between subjects				
Reader group (RG)	1	131,279.2	273.8	<.001
Family history group (FHG)	1	4757.8	9.92	.002
RG*FHG	1	85.5	.18	.67
Error	394	479.5		
Within subjects				
Grades (G)	4	641.2	14.01	<.001
G*RG	4	89.5	1.96	.10
Gr*FHG	4	17.7	.39	.82
G*RG*FHG	4	9.00	.20	.94
Error	1576	45.8		
Greenhouse-Geisser epsilon	.62			
Huynh-Feldt epsilon	.63			

significant (p = .002, ES = .59). The FH⁻ and FH⁺ difference within typical readers (Figure 1, panel C, upper) was small and, after adjusting the p values for multiple comparisons, not statistically significant ($p = .015^1$, ES = .29), as was the FH⁻ and FH⁺ difference within dyslexic readers (Figure 1, panel C, lower) (p = .61, ES = .22). The grade main effect was the only within-subject effect that was statistically significant (p < .001). Inspection of panels A, B, and C in Figure 1 indicates that reading scores tend to decline slightly from grades 1 to 5 and increase from grades 5 to 9 in each comparison. Although statistically significant, the differences across grades are small relative to the differences due to groups. Notably, none of the interactions is statistically significant, indicating that differences between reading groups, family history groups, and between grades are independent of each other.

Finally, results of ROC curve analyses examining the predictive utility of the family history variable (positive family history for any 1st- and 2nd-degree relative) alone, early screening measure alone, and the combined early screening measure and family history variables on dyslexia status for kindergarten and first grade are reported in Table 3. As a reference, included in this table are also classification accuracy values from the screen manual (Shaywitz, 2016, p. 10). These results indicate that the screener is superior to family history in sensitivity but inferior in specificity. Moreover,

¹ To control for type I error for the four comparisons, the criterion for statistical significance was set at p = .05/4 = .0125

Fig. 1 Panel A is a plot of means \pm 95% confidence intervals of WJ Reading Cluster scores for typical and dyslexic readers over grades 1 to 9. The grand means over grades are 109.8 and 88.8, for the TR and DR groups, respectively. The difference between the TR and DR of 20.9 is statistically significant (F₁₃₉₄ = 273.8, p <.001). Panel B is a plot of means \pm 95% confidence intervals of WJ Reading Cluster scores for FH and FH⁺ groups over grades 1 to 9. The grand means over grades are 107.3 and 98.6 for FH⁻ and FH⁺, respectively. The difference between the FH⁻ and FH⁺ of 8.6 is statistically significant ($F_{1394} =$ 9.9, p = .002). Panel C is a plot of means $\pm 95\%$ confidence intervals of WJ Reading Cluster scores for FH⁻ and FH⁺ groups over grades 1 to 9 within the TR group and within the DR group. For the TR group, the FH⁻ grand mean over grades is 110.8, and the FH⁺ grand mean over grades is 106.5. The difference between the FH⁻ and FH⁺ of 4.3 is not statistically significant ($F_{5390} =$ 2.86, p = .015). For the DR group, the FH⁻ grand mean over grades is 90.5, and the FH^+ grand mean over grades is 87.2. The difference between the FH- and FH⁺ of 3.3 is not statistically significant ($F_{5390} = .72, p = .61$)

150 130 Mean +/- 95% CI 90 Reader Group 70 Typical Dyslexic 50 WJ Reading Cluster Scores By Grade B Family History By Grade 150 130 Mean +/- 95% CI 10 90 Family History 70 FH-FH+ 50 5 WJ Reading Cluster Scores By Grade С Reader Group x Family History 150 130 \overline{O} Mean +/- 95% 10 Reader Group, Family History Typical, FH-70 Typical, FH+ Dyslexic, FH-Dyslexic, FH+ 50 WJ Reading Cluster Scores By Grade

A

Reader Group By Grade

although the screener has a higher false positivity rate than family history, the predictive indices for screener alone are very similar to those obtained by adding family history, with little value added.

Discussion

In this paper, using a longitudinal epidemiologic sample of schoolchildren, we examine the classification accuracy of family history as a screening measure for at-risk for dyslexia. In addition, we investigate the predictive value of an evidencebased early screening measure (Shaywitz DyslexiaScreen) to identify at-risk for dyslexia and determine whether family history provided added value.

Using an epidemiologic sample of 398 children followed from age 5 through adulthood, we found that sensitivity of family history for predicting dyslexia was unacceptably low for all family member groups, even when using the highest sensitivity (combining 1st- and 2nd-degree relatives). These results indicate that an evidence-based screener is superior to family history in sensitivity, the primary metric used in screening. ROC curves for family history alone, early screening measure alone, and the combination of the screener and family
 Table 3
 At-risk classification

 summary statistics for family
 history alone, screener alone, and

 their combination in kindergarten
 and grade 1

	n				%			
	ТР	TN	FP	FN	Sensitivity	Specificity	AUC	
Kindergarten (n = 398)								
Family history (FH)	49	231	70	47	51	77	70	
Early screening (ES)	66	218	83	30	69	72	78	
FH + ES	65	228	73	31	68	76	79	
First grade $(n = 371)$								
Family history (FH)	38	228	69	36	51	77	71	
Early screener (ES)	55	240	57	19	74	80	86	
FH + ES	52	250	47	22	70	70	84	
Manual reference*								
NCS kindergarten					71	71	81	
NCS first grade					70	88	89	
CTCS kindergarten					69	72	74	
CTCS first grade					74	81	84	

Note: NCS, National clinical study (n = 115); CTCS, Connecticut clinical study (n = 414)

*Shaywitz DyslexiaScreen (Shaywitz, S. E., 2016)

history indicate that predicting dyslexia using family history does not improve the value of using the screener alone and, in fact, for first-grade data, the addition of family history to the early screener appears to make the prediction worse.

Our report is the first to include sensitivity, specificity, and AUC analyses in an epidemiologic sample examining the classification accuracy of family history for determining at-risk status for dyslexia. Consistent with previous analyses based on growth modeling (Ferrer et al., 2015), the present analyses indicate that the persistent normative reading achievement gap between typical and dyslexic readers from first grade to adolescence is more than twice that of the persistent achievement gap between individuals with and without family history of dyslexia. The significant overall FH achievement gap and the nonsignificant FH differences when examining within TR and DR groups are consistent with the low sensitivity of family history as a predictor of dyslexia and undermine the use of family history for universal screening.

One important way to advance our understanding of the current results would be to conduct analyses of sensitivity and specificity separately by the various SES, sex, and ethnic groups. These additional results would expand the overall results and illuminate the usefulness of family history for different groups. Furthermore, they may also affect the ecological validity of the results and translational practicality of using family history during screening for educators working in diverse areas.² In our current sample, however, including SES, sex, and ethnic groups into the analyses would result in such small subgroups and reduced power that statistical comparisons would be compromised. Future research using large epidemiologic samples that oversample small groups should pursue such analyses.

We conclude that the proposed use of positive family history as a proxy for dyslexia is unwarranted. Moreover, if family history were to be used for screening to determine at-risk for dyslexia, this could have harmful consequences. For example, many children who are dyslexic may never be identified by their schools, as they and their parents may not be aware of any family history of difficulties in reading or spelling, an issue especially problematic for children from single-family households and in children from economically disadvantaged circumstances.

Open Access This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by/4.0/.

² We thank an anonymous reviewer for raising this point.

References

Carroll, J. M., Mundy, I. R., & Cunningham, A. J. (2014). The roles of family history of dyslexia, language, speech production and phonological processing in predicting literacy progress. *Developmental Science*, 17, 727–742.

- Dilnot, J., Hamilton, L., Maughan, B., & Snowling, M. J. (2017). Child and environmental risk factors predicting readiness for learning in children at high risk of dyslexia. *Development and Psychopathology*, 29, 235–244.
- Estrada, E., Ferrer, E., Shaywitz, B., Holahan, J., & Shaywitz, S. (2018). Identifying atypical change at the individual level from childhood to adolescence. *Developmental Psychology*, 54(11), 2193–2206. https://doi.org/10.1037/dev0000583.
- Ferrer, E., et al. (2007). Longitudinal models of developmental dynamics between reading and cognition from childhood to adolescence. *Developmental Psychology*,43, 1460–1473. https://doi.org/10. 1037/0012-1649.43.6.1460.
- Ferrer, E., Shaywitz, B. A., Holahan, J. M., Marchione, K., & Shaywitz, S. E. (2010). Uncoupling of reading and IQ over time: Empirical evidence for a definition of dyslexia. *Psychological Science*, 21, 93– 101.
- Ferrer, E., Shaywitz, B., Holahan, J., Marchione, K., Michaels, R., & Shaywitz, S. (2015). Achievement gap in reading is present as early as first grade and persists through adolescence. *Journal of Pediatrics*, 167(5), 1121–1125.
- First Step Act (2018). Pub. L. No. 115-391. Office of the Federal Register through the Congressional Printing Management Division, U.S. Government Printing Office.
- Hammill, D. D., & Allen, E. A. (2020). A revised discrepancy method for identifying dyslexia. *Journal of Pediatric Neuropsychology*, 6, 27– 43. https://doi.org/10.1007/s40817-020-00079-2.
- Herrera-Araujo, D., Shaywitz, B. A., Holahan, J. M., Marchione, K. E., Michaels, R., Shaywitz, S. E., & Hammitt, J. K. (2017). Evaluating willingness to pay as a measure of the impact of dyslexia in adults. *Journal of Benefit-Cost Analysis*, 8(1), 24–48. https://doi.org/10. 1017/bca.2017.3.
- Hinshelwood, J. (1900). Congenital word-blindness. The Lancet, 155, 1506–1508.
- Hinshelwood, J. (1917). Congenital word-blindness. Lewis
- Hinshelwood, J. (1911). Two cases of hereditary congenital word-blindness. *British Medical Journal*, 1, 608.
- Lyon, G. R., Shaywitz, S. E., & Shaywitz, B. A. (2003). A definition of dyslexia. Annals of Dyslexia, 53, 1–14. https://doi.org/10.1007/ s11881-003-0001-9.
- Morgan, W. P. (1896). A case of congenital word blindness. British Medical Journal, 1378.

- Shaywitz, S. E. (2016). Shaywitz DyslexiaScreen. NCS Pearson. Bloomingon, MN.
- Shaywitz, S. E., Schnell, C., Shaywitz, B. A., & Towle, V. R. (1986). Yale Children's Inventory (YCI): An instrument to assess children with attentional deficits and learning disabilities I. Scale development and psychometric properties. *Journal of Abnormal Child Psychology*, 14, 347–364.
- Shaywitz, S. E., Shaywitz, B. A., Schnell, C., & Towle, V. R. (1988). Concurrent and predictive validity of the Yale Children's Inventory: An instrument to assess children with attentional deficits and learning disabilities. *Pediatrics*, 81, 562–571.
- Shaywitz, S. E., Shaywitz, B. A., Fletcher, J. M., & Escobar, M. D. (1990). Prevalence of reading disability in boys and girls. Results of the Connecticut Longitudinal Study. *JAMA*, 264, 998–1002.
- Shaywitz, S. E., Escobar, M. D., Shaywitz, B. A., Fletcher, J. M., & Makuch, R. (1992a). Evidence that dyslexia may represent the lower tail of a normal distribution of reading ability. *New England Journal* of *Medicine*,326, 145–150.
- Shaywitz, S., Holahan, J., Marchione, K., Sadler, A., & Shaywitz, B. (1992b). The Yale Children's Inventory: Normative data and their implications for the diagnosis of attention deficit disorder in children. . In S. E. Shaywitz & B. A. Shaywitz (Eds.), Attention deficit disorder comes of age (pp. 29–67). Austin, TX: pro-ed
- Shaywitz, B. A., Fletcher, J. M., Holahan, J. M., & Shaywitz, S. E. (1992c). Discrepancy compared to low achievement definitions of reading disability: Results from the Connecticut Longitudinal Study. *Journal of Learning Disabilities*, 25, 639–648.
- Shaywitz, S. E., et al. (1999). Persistence of dyslexia: The Connecticut Longitudinal Study at adolescence. *Pediatrics*, 104, 1351–1359.
- Shaywitz, S. E., Shaywitz, B. A., Fulbright, R. K., Skudlarski, P., Mencl, W. E., Constable, R. T., & Gore, J. C. (2003). Neural systems for compensation and persistence: Young adult outcome of childhood reading disability. *Biological Psychiatry*, 54(1), 25–33.
- Snowling, M. J., & Melby-Lervåg, M. (2016). Oral language deficits in familial dyslexia: A meta-analysis and review. *Psychological Bulletin*,142, 498.
- Thompson, P. A., et al. (2015). Developmental dyslexia: Predicting individual risk. *Journal of Child Psychology and Psychiatry*, 56, 976– 987.
- Wechsler, D. (1981). Wechsler Adult Intelligence Scale—revised. Psychological Corporation.
- Woodcock, R. W., & Johnson, M. B. (1989). Woodcock–Johnson Psycho-Educational Battery—revised. DLM.

Publisher's Note Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Clearinghouse, W. W. (2017).