SYSTEMATIC REVIEW



Efficacy and Safety of Mipomersen: A Systematic Review and Meta-Analysis of Randomized Clinical Trials

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Abstract

Aim Our aim was to assess the efficacy and safety of mipomersen through a systematic review of the literature and a metaanalysis of the available clinical studies.

Methods A systematic literature search in SCOPUS, PubMed Medline, ISI Web of Science and Google Scholar databases was conducted up to January 20, 2019, in order to identify clinical trials assessing the effect of mipomersen on lipoproteins, and the safety profile of mipomersen. Effect sizes for lipid changes were expressed as weighted mean differences (WMD) and 95% confidence intervals (CI). For safety analysis, odd ratios (OR) and 95% CI were calculated using the Mantel–Haenszel method. Data were pooled from 13 clinical studies comprising 49 arms, which included 1053 subjects overall, with 729 in the active-treated arm and 324 in the control arm.

Results Meta-analysis of data suggested that mipomersen significantly reduced low-density lipoprotein cholesterol (WMD – 1.52, 95% CI – 1.85 to – 1.19; p < 0.001), total cholesterol (WMD – 1.55, 95% CI – 1.97 to – 1.13; p < 0.001), non-high-density lipoprotein cholesterol (non-HDL-C) (WMD – 1.66, 95% CI – 2.06 to – 1.27; p < 0.001), lipoprotein(a) (WMD – 0.99, 95% CI – 1.37 to – 0.62; p < 0.001), apolipoprotein B (WMD – 1.66, 95% CI – 2.04 to – 1.27; p < 0.001), triglycerides (WMD – 0.61, 95% CI – 0.76 to – 0.46, p < 0.001), very-low-density lipoprotein cholesterol (WMD – 0.58, 95% CI – 0.73 to – 0.43; p < 0.001) and apolipoprotein A-I (WMD – 0.25, 95% CI – 0.51 to – 0.001; p = 0.049) without affecting HDL-C levels (WMD 0.11, 95% CI – 0.03 to 0.26; p = 0.124). However, treatment with mipomersen was positively associated with an increased risk of discontinuation of treatment (OR 3.02, 95% CI 1.96–4.65; p < 0.001), injection-site reaction (OR 11.41, 95% CI 7.88–16.52; p < 0.001), hepatic steatosis (OR 4.96, 95% CI 1.99–12.39; p = 0.001), hepatic enzymes elevation (OR 3.61, 95% CI 2.09–6.24; p < 0.001) and flu-like symptoms (OR 2.02, 95% CI 1.45–2.81; p < 0.001).

Conclusion Despite favourable effects on the lipid profile, some concerns are reinforced from the safety profile. As a matter of fact, mipomersen therapy is more likely discontinued and associated with increased risk of injection-site reactions, hepatic steatosis, hepatic enzyme elevation, and flu-like symptoms.

Federica Fogacci and Nicola Ferri contributed equally to the paper.

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Key Points

Mipomersen has recently been developed as adjunctive therapy for the treatment of homozygous familial hypercholesterolaemia.

In 2013, the European Medicines Agency (EMA) refused mipomersen marketing authorization in Europe due to safety concerns.

Despite favourable effects on the lipid profile, mipomersen is more likely discontinued and is associated with increased risk of injection-site reactions, hepatic steatosis, hepatic enzyme elevation, and flu-like symptoms.

To date, there is no apparent reason to stop taking mipomersen when it is effective and well tolerated, even though a careful monitoring of liver function is needed.

1 Background

Familial hypercholesterolaemia (FH) is highly prevalent throughout the world and represents a major public health concern [1, 2].

According to the latest research, FH is most commonly caused (>95%) by mutations in the low-density lipoprotein (LDL) receptor gene (*LDLR*, MIM 606945), while secondary genes are, in descending order of frequency, apolipoprotein B (*APOB*, MIM 107730) (2.5%), proprotein convertase subtilisin/kexin type 9 (*PCSK9*, MIM 607786) (<1%) and low-LDLR adaptor protein 1 (*LDLRAP1*, MIM 695747) (<1%) [3], which lead to impaired low-density lipoproteincholesterol (LDL-C) uptake or increased LDL receptor (LDLR) degradation [4].

Regardless of whether the screening is clinical or genetic, the prevalence of heterozygous FH (HeFH) seems to be higher than previously thought, being estimated at 1 in 200–250 individuals [1, 5], while homozygous FH (HoFH) is much rarer, estimated at 1 in 160,000–300,000 [3]. The burden of markedly elevated LDL-C levels from birth underlies the sequelae of atherosclerotic cardiovascular disease complications, which are very early for untreated HoFH, typically occurring in childhood or at most within the second decade of life [1, 3, 6].

Current treatment of HoFH with conventional drugs consists of high doses of high-intensity statins, often in combination with ezetimibe, which have recently been shown to reduce CV and all-cause mortality, with minimal adverse effects and relatively low costs [3]. Combinations of statins with other cholesterol-lowering medications, including fibrates, bile acid sequestrants, probucol and niacin have been studied in HoFH and can be considered to further lower LDL-C levels, although their use may be limited by poor tolerability [3]. However, the LDL-C goal is often not reached with this therapeutic regimen, underlying the need for additional LDL-lowering options. One of these is PCSK9 inhibition, with promising results of a robust additional LDL-C reduction and relevant reduction in major adverse cardiac events (MACE) obtained by administering evolocumab 420 mg every 4 weeks, with or without LDL apheresis [7–9].

Most recently, two additional drugs, lomitapide and mipomersen, have been developed and approved as adjunctive therapies for the treatment of HoFH [10]. These drugs reduce the production and secretion of apoB-containing lipoproteins [11, 12], rather than increasing their clearance from the circulation. In particular, mipomersen is an antisense oligonucleotide directed against apolipoprotein-B 100 (apo B-100) mRNA in the liver, ultimately resulting in decreased serum levels of apo B-100-containing lipoproteins such as LDL and lipoprotein(a) [Lp(a)] [13].

Even though mipomersen has been approved by the US Food and Drug Administration (FDA) as adjunct therapy for HoFH patients aged ≥ 12 years in the United States, in 2013 the European Medicines Agency (EMA) refused its marketing authorization in Europe because of safety concerns [14, 15]. Consequently, we aimed to perform a meta-analysis on clinical evidence available to date to better define its efficacy and tolerability profile.

2 Methods

The study was designed according to guidelines of the 2009 preferred reporting items for systematic reviews and meta-analysis (PRISMA) statement [16], and was registered in the PROSPERO database (Registration number CRD42019121505). Because of the study design (meta-analysis), neither Institutional Review Board (IRB) approval, nor patient informed consent were required.

2.1 Search Strategy

PubMed, SCOPUS, Google Scholar and ISI Web of Science by Clarivate databases were searched, with no language restriction, using the following search terms: 'Mipomersen' AND ('Clinical trial' OR 'Clinical study'). The wild-card term (*) was used to increase the sensitivity of the search strategy, which was limited to studies in humans. The reference lists of identified papers were manually checked for additional relevant articles. In particular, additional searches for potential trials included the references of review articles on that issue and the abstracts from selected congresses on the subject of the meta-analysis. Literature was searched from inception to January 20, 2019. All paper abstracts were screened by two reviewers (FF and AFGC) in an initial process to remove ineligible articles. The remaining articles were obtained in full-text and assessed again by the same two researchers, who evaluated each article independently and carried out data extraction and quality assessment. Disagreements were resolved by discussion with a third party (NF).

2.2 Study Selection Criteria

Original studies were included if they met the following criteria: (1) being a clinical trial with either multicentre or single-centre design, (2) having an appropriate controlled design for mipomersen treatment, (3) investigating the effect of mipomersen on plasma lipids, (4) testing the safety of mipomersen short and middle-term administration, (5) reporting all the adverse events that occurred during the treatment.

Exclusion criteria were (1) lack of a control group for mipomersen administration, and (2) lack of sufficient information about the prevalence and nature of the adverse events. Studies were also excluded if they contained overlapping subjects with other studies.

2.3 Data Extraction

Data abstracted from the eligible studies were (1) first author's name; (2) year of publication; (3) study design; (4) main inclusion criteria and underlying disease; (5) treatment duration; (6) study groups; (7) number of participants in the active and control group; (8) age and sex of study participants; and (9) discontinuation of treatment and adverse events occurring during the trials. Missing or unpublished data were sought by trying to contact authors or sponsors via e-mail and repeated messages were sent in case of no response. All data extraction and database typing were reviewed by the principal investigator (AFGC) before the final analysis, and doubts were resolved by mutual agreement among the authors.

2.4 Quality Assessment

A systematic assessment of risk of bias in the included studies was performed using the Cochrane criteria [17]. The following items were used: adequacy of sequence generation, allocation concealment, blinding addressing of dropouts (incomplete outcome data), selective outcome reporting, and other probable sources of bias [18]. Risk-of-bias assessment was performed independently by two reviewers (FF and AFGC); disagreements were resolved by a consensusbased discussion.

2.5 Data Synthesis

Meta-analysis was entirely conducted using Comprehensive Meta-Analysis (CMA) V3 software (Biostat, NJ, USA) [19].

Net changes in the investigated parameters (change scores) were calculated by subtracting the value at baseline from the value after intervention, in both the active-treated and the control group. All values were collated as percent change from baseline. Standard deviations (SDs) of the mean difference were obtained as per the following, reported by Follmann and colleagues: $SD = square root [(SD_{pre-treatment})^2 +$ $(SD_{post-treatment})^2 - (2R \times SD_{pre-treatment} \times SD_{post-treatment})],$ assuming a correlation coefficient (R) = 0.5 [20]. If the outcome measures were reported in median and range (or 95% confidence interval [CI]), mean and SD values were estimated using the method described by Wan et al. [21]. Where standard error of the mean (SEM) was only reported as a dispersion measure, SD was estimated using the following formula: $SD = SEM \times square root(n)$, with n being the number of subjects. To avoid a double-counting problem, in trials comparing multiple treatment arms versus a single control group, the number of subjects in the control group was divided by the required comparisons. Studies' findings were combined using a fixed-effect model or a randomeffect model (using the DerSimonian-Laird method) and the generic inverse variance method, based on the level of interstudy heterogeneity, which was quantitatively assessed using the Higgins index (I^2) [22]. Effect sizes for lipid changes were expressed as weighted mean differences (WMD) and 95% CI. For safety analysis, odd ratios (OR) and 95% CI intervals were calculated using the Mantel-Haenszel method [23]. Safety analysis was performed by excluding studies with zero events in both arms. If one or more outcomes could not be extracted from a study, the study was removed only from the analysis involving those outcomes. Adverse events were considered for the analysis only if occurring in at least two of the included clinical trials. In order to evaluate the influence of each study on the overall effect size, sensitivity analysis was conducted using the leave-one-out method (i.e. removing one study at a time and repeating the analysis) [24]. Two-sided p values ≤ 0.05 were considered as statistically significant for all tests.

2.6 Publication Biases

Potential publication biases were explored using visual inspection of Begg's funnel plot asymmetry, Begg's rank correlation test, and Egger's weighted regression test [25]. The Duval and Tweedie 'trim and fill' method was used to adjust the analysis for the effects of publication bias [26]. Two-sided p values ≤ 0.05 were considered statistically significant.

3 Results

3.1 Flow and Characteristics of the Included Studies

After database searches performed strictly according to inclusion and exclusion criteria, 247 published articles were identified, and the abstracts reviewed. Of these, 192 were excluded because they were non-original articles. Another 37 were eliminated because they did not meet the inclusion criteria. Thus, 18 articles were carefully assessed and reviewed. An additional five studies were excluded because of incomplete data (n=2) or substantial sample overlap (n=3) (Appendix 1, see electronic supplementary material [ESM]). Finally, 13 studies were eligible and included in the meta-analysis [27–40]. The study selection process is shown in Fig. 1. Data were pooled from 13 clinical trials comprising 49 treatment arms, which included 1053 subjects, with 729 in the active-treated arm and 324 in the control arm.

Eligible studies were published between 2006 and 2019. Follow-up periods ranged between 3 and 60 weeks and different treatment schedules were tested. All selected trials were designed with parallel groups and were multicentre [27, 31–33, 36–38] or single-centre [28, 30, 35, 39, 40] clinical studies. Enrolled subjects were adults and young adults with a diagnosis of HeFH [27, 28, 33, 37, 39] or HoFH [38], patients with coronary heart disease, [27, 31–33] or with a good status of health [30, 35, 36, 40]. The baseline characteristics of the evaluated studies are summarized in Table 1.

3.2 Risk of Bias Assessment

Almost all of the included studies were characterized by sufficient information regarding sequence generation, allocation concealment, personal and outcome assessments. All studies showed low risk of bias as regards incomplete outcome data and selective outcome reporting. Details of the quality of bias assessment are reported in Table 2.

3.3 Effect of Mipomersen on Lipids

Meta-analysis of data suggested that mipomersen significantly reduced LDL-C (WMD – 1.52, 95% CI – 1.85 to – 1.19; p < 0.001; $I^2 = 73.6\%$) (Fig. 2), total cholesterol (TC) (WMD – 1.55, 95% CI – 1.97 to – 1.13; p < 0.001; $I^2 = 80.3\%$), non-high density lipoprotein cholesterol (non-HDL-C) (WMD – 1.66, 95% CI – 2.06 to – 1.27; p < 0.001; $I^2 = 77\%$), Lp(a) (WMD – 0.99, 95% CI – 1.37 to – 0.62; p < 0.001; $I^2 = 77.3\%$), Apo B (WMD – 1.66, 95% CI – 2.04 to – 1.27; p < 0.001; $I^2 = 80.1\%$) (Fig. 3), HDL-C (WMD 0.11, 95% CI – 0.03 to 0.26; p = 0.124; $I^2 = 18.4\%$), triglycerides (TG) (WMD – 0.61, 95% CI – 0.76 to – 0.46;



Fig. 1 Flow chart of the number of studies identified and included in the meta-analysis

Study	Study design	Main inclusion criteria	Treatment duration	Study group	Patients (n)	Age, years (mean±SD)	Male, <i>n</i> (%)
Reeskamp et al.	Multicentre, randomized,	НеFH	60 weeks	Mipomersen 200 mg once weekly	67	55.2 ± 10.1	25 (37.3)
2019 [27]	double-blind, placebo-	≥18 years of age		Placebo once weekly	34	56.2 ± 10.8	13 (38.2)
	controlieu, paraliei-group clinical study	レレレーン こうりり mg/ut of レレレーン ことり0 mg/ df _ nhis documented CHD or CHD risk		Mipomersen 70 mg thrice weekly	99	51.7 ± 12.8	27 (40.9)
		equivalents Maximally tolerated lipid-lowering treatment		Placebo thrice weekly	33	56.1 ± 8.9	14 (42.4)
		HeFH	60 weeks	Mipomersen 200 mg once weekly	37	58.5 ± 9	21 (56.8)
		\geq 18 years of age		Placebo once weekly	17	54.1 ± 10	9 (52.9)
		LDL-U ≥ 100 mg/dL or ≤ 200 mg/dL Documented CHD or CHD risk equivalents		Mipomersen 70 mg thrice weekly	36	55.8 ± 9.8	21 (58.3)
		Maximally tolerated lipid-lowering treatment		Placebo thrice weekly	19	51.5 ± 11.1	11 (57.9)
Waldmann et al. 2017 [28, 29]	Single-centre, randomized, controlled, phase II clini-	HeFH ≥18 years of age	26 weeks	Mipomersen 200 mg once weekly at day 4 after apheresis	11	64.5±7	6 (54.5)
	cal study	LDL-C≥ 130 mg/dL despite maximal possible lipid-lowering treatment Lipoprotein apheresis		No treatment	3	53.7±11.6	3 (100)
Flaim et al. 2014	Single-centre, randomized,	18–75 years of age	3 weeks	Mipomersen 30 mg once daily	21	47 (28–61) ^a	18 (85.7)
[30]	double-blind, placebo-	Body weight > 50 kg		Mipomersen 70 mg twice weekly	21	50 (22–69) ^a	14 (66.7)
	controlled, phase I clinical study	Skin type 1–111 based on the Fitzpatrick scale		Mipomersen 200 mg once weekly	21	52 (19–70) ^a	16 (76.2)
	auuy			Placebo	21	48 (22–70) ^a	14 (66.7)
Thomas et al.	Multicentre, randomized,	≥ 18 years of age	26 weeks	Mipomersen 200 mg once weekly	105	59.3 ± 10	52 (49.5)
2013 [31]	double-blind, placebo- controlled, parallel-group clinical study	LDL-C≥ 100 mg/dL CHD or at high risk for CHD according to the NCEP-ATP III criteria Maximally tolerated statin dose		Placebo once weekly	52	59.3±9.5	29 (55.8)
McGowan et al.	Multicentre, randomized,	LDL-C $\ge 197.2 \text{ mg/dL}$ with known CHD or	26 weeks	Mipomersen 200 mg once weekly	39	51.8 ± 14.3	18 (46.2)
2012 [32]	double-blind, placebo- controlled clinical study	LDL-C≥301.6 mg/dL in absence of known CHD		Placebo once weekly	19	47.9±13.5	7 (36.8)
Stain at al 2012	Multioantra randomizad	Maximally tolerated lipid-lowering treatment	J6 weeks	Minomercen 200 mg once weekly	83	267407	50 (60 3)
[33]	double-blind, placebo-	HeFH by either genetic confirmation of an	500 m 07	Placebo once weekly	41	55.9 ± 9.3	28 (68.3) 28 (68.3)
	controlled, pnase III clini- cal study	LUL-K detect or according to the SBK criteria plus documented and stable CAD and untrasted 1 D1 C > 100 model					
		Maximally tolerated statin dose, with or with- out other lipid-lowering treatment					
Visser et al. 2012	Randomized, double-blind,	High CV risk according to the NCEP-ATP III	26 weeks	Mipomersen 200 mg once weekly	21	55 (46–69) ^a	11 (52)
[34]	placebo-controlled, paral- lel-group clinical study	criteria Statin intolerance T DT C > 121 S model		Placebo	12	52 (39–68) ^a	4 (33)

755

Akdim et al. 2011 Single-centre, randomized, double-blind, placebo- controlled, dose-escalation clinical study Akdim et al. Multicentre, randomized, double-blind, placebo- controlled, dose-esca- lation, phase II clinical study Akdim et al. Multicentre, randomized, double-blind, placebo- controlled, dose-esca- lation, phase II clinical study				Parients (n)		Male n (%)
Akdim et al. 2011Single-centre, randomized, double-blind, placebo- controlled, dose-escalation clinical studyAkdim et al.Multicentre, randomized, double-blind, placebo- controlled, dose-esca- lation, phase II clinical studyAkdim et al.Multicentre, randomized, double-blind, placebo- controlled, dose-esca- lation, phase II clinical study		duration	ound group		mge, years (mean±SD)	Maic, n (70)
 [35] double-blind, placebo- controlled, dose-escalation clinical study Akdim et al. Multicentre, randomized, double-blind, placebo- controlled, dose-esca- lation, phase II clinical study Akdim et al. Multicentre, randomized, double-blind, placebo- 	18–65 years of age	13 weeks	Mipomersen 50 mg once weekly	8	47.4 ± 7.2	7 (87.5)
Akdim et al. Multicentre, dose-escalation cilinical study 2010a [36] double-blind, placebo- controlled, dose-esca- lation, phase II clinical study 2010b [37] double-blind, placebo- 2010b [37] double-blind, placebo-	BMI $\ge 25 \text{ kg/m}^2 \text{ and } \le 32 \text{ kg/m}^2$		Mipomersen 100 mg once weekly	8	40.5 ± 8.2	7 (87.5)
Akdim et al.Multicentre, randomized,2010a [36]double-blind, placebo- controlled, dose-esca- lation, phase II clinical studyAkdim et al.Multicentre, randomized, double-blind, placebo-	on Untreated LDL-C≥130 mg/dL No linid_lowering treatment		Mipomersen 200 mg once weekly	8	49.6 ± 8.5	7 (87.5)
 Akdim et al. Multicentre, randomized, 2010a [36] double-blind, placebo- controlled, dose-esca- lation, phase II clinical study Akdim et al. Multicentre, randomized, 2010b [37] double-blind, placebo- 			Mipomersen 300 mg once weekly	8	51.6 ± 9	8 (100)
 Akdim et al. Multicentre, randomized, 2010a [36] double-blind, placebo-controlled, dose-escalation, phase II clinical study Akdim et al. Multicentre, randomized, 2010b [37] double-blind, placebo- 			Mipomersen 400 mg once weekly	8	55.4 ± 9.8	8 (100)
Akdim et al.Multicentre, randomized,2010a [36]double-blind, placebo- controlled, dose-esca- lation, phase II clinical studyAkdim et al.Multicentre, randomized, double-blind, placebo-			Placebo	10	52.3 ± 7.4	8 (80)
 2010a [36] double-blind, placebo- controlled, dose-esca- lation, phase II clinical study Akdim et al. Multicentre, randomized, double-blind, placebo- 	18–65 years of age	5 weeks	Mipomersen 30 mg once weekly	8	58 ± 3.9	6 (75)
controlled, dose-esca- lation, phase II clinical study Akdim et al. Multicentre, randomized, 2010b [37] double-blind, placebo-	LDL-C $\geq 100 \text{ mg/dL}$ and $\leq 220 \text{ mg/dL}$ on sta-		Mipomersen 100 mg once weekly	8	57.4 ± 4.1	4 (50)
Akdim et al. Multicentre, randomized, 2010b [37] double-blind, placebo-	ble statin treatment at a dose of $\leq 40 \text{ mg}$ daily		Mipomersen 200 mg once weekly	16	58.3 ± 3.8	11 (68.8)
Akdim et al. Multicentre, randomized, 2010b [37] double-blind, placebo-			Mipomersen 300 mg once weekly	8	56.9 ± 4.3	4 (50)
Akdim et al. Multicentre, randomized, 2010b [37] double-blind, placebo-			Mipomersen 400 mg once weekly	6	61.4 ± 3.2	7 (77.8)
Akdim et al. Multicentre, randomized, 2010b [37] double-blind, placebo-			Placebo	13	NA	NA
Akdim et al. Multicentre, randomized, 2010b [37] double-blind, placebo-		13 weeks	Mipomersen 200 mg once weekly	10	59 ± 4	6 (60)
Akdim et al.Multicentre, randomized,2010b [37]double-blind, placebo-			Placebo	2	NA	NA
2010b [37] double-blind, placebo-	HeFH	6 weeks	Mipomersen 50 mg once weekly	8	49 ± 12	5 (62.5)
	18–75 years of age		Mipomersen 100 mg once weekly	8	53 ± 11	5 (62.5)
controlled, dose-esca- lation whase II clinical	LDL-C > 130 mg/dL Stable conventional linid-lowering treatment		Mipomersen 200 mg once weekly	11	56 ± 13	4 (36.4)
study	DIADIC CORVERIENTIAL INPIR-LOWCLING REALINCIN		Mipomersen 300 mg once weekly	6	47 ± 7	6 (66.7)
`			Placebo	8	54 ± 10	6 (75)
Raal et al. 2010 Multicentre, randomized,	НоFH	26 week	Mipomersen 200 mg once weekly	34	30.4 ± 11.5	15 (44)
[38] double-blind, placebo- controlled, phase III clini- cal study	 ≥12 years of age LDL-C≥131.5 mg/dL Body weight > 40 kg Maximally tolerated lipid-lowering treatment 		Placebo	17	33 ±14.1	7 (41)
Visser et al. 2010 Singe-centre, randomized,	НеҒН	13 weeks	Mipomersen 200 mg once weekly	10	49 ± 12	6 (60)
[39] double-blind, placebo- controlled clinical study	18–75 years of age LDL-C≥100.5 mg/dL		Placebo	11	46 ± 1	3 (27.3)
Kastelein et al. Randomized, double-blind,	, 18–65 years of age	55 days	Mipomersen 50 mg once weekly	8	45 ± 11	4 (50)
2006 [40] placebo-controlled, dose-	• TC < 300 mg/dL		Mipomersen 100 mg once weekly	8	48±9	6 (75)
escalation clinical study			Mipomersen 200 mg once weekly	6	58±6	3 (33.3)
			Mipomersen 400 mg once weekly	4	51 ± 9	2 (50)
			Placebo	L	57±3	3 (42.9)

lesterol, *LDL-R* low-density lipoprotein receptor, *NA* not available, *NCEP-ATP III* National Cholesterol Education Program Adult Treatment Panel III, *SBR* Simon Broome Register, *SD* standard deviation, *TC* total cholesterol

^aExpressed as median (minimum-maximum)

 Table 2
 Quality of bias assessment of the included studies according to Cochrane guidelines

Study	Sequence generation	Allocation concealment	Blinding of participants, personnel and outcome assess- ment	Incomplete outcome data	Selective out- come reporting	Other poten- tial threats to validity
Reeskamp et al. 2019 [27]	L	L	L	L	L	L
Waldmann et al. 2017 [28, 29]	L	L	Н	L	L	Н
Flaim et al. 2014 [30]	L	L	L	L	L	L
Thomas et al. 2013 [31]	L	L	L	L	L	L
McGowan et al. 2012 [32]	L	L	L	L	L	L
Stein et al. 2012 [33]	L	L	L	L	L	L
Visser et al. 2012 [34]	L	L	L	L	L	L
Akdim et al. 2011 [35]	L	L	L	L	L	L
Akdim et al. 2010a [36]	L	L	L	L	L	L
Akdim et al. 2010b [37]	L	L	L	L	L	L
Raal et al. 2010 [38]	L	L	L	L	L	L
Visser et al. 2010 [39]	L	L	L	L	L	L
Kastelein et al. 2006 [40]	L	L	L	L	L	L

L low risk of bias, H high risk of bias, U unclear risk of bias



Fig. 2 Forest plot displaying weighted mean difference and 95% confidence intervals (CI) for the impact of mipomersen on plasma LDL-C concentrations. *LDL-C* low-density lipoprotein cholesterol





 $p < 0.001; I^2 = 45.1\%$], very-low-density lipoprotein cholesterol (VLDL-C) (WMD – 0.58, 95% CI – 0.73 to – 0.43; $p < 0.001; I^2 = 18.1\%$) and Apo A-I (WMD – 0.25, 95% CI – 0.51 to – 0.001; $p = 0.049; I^2 = 54.8\%$), without affecting HDL-C levels (WMD 0.11, 95% CI – 0.03 to 0.26; $p = 0.124; I^2 = 18.4\%$) (Fig. 4). The effect sizes were robust in the leave-one-out sensitivity analysis and not mainly driven by a single study (Figs. S1, S2, see ESM).

Visual inspection of Begg's funnel plot revealed a slight asymmetry, suggesting potential publication bias for the effect of mipomersen on serum HDL-C concentrations (Fig. 5). This asymmetry was imputed to six potentially missing studies on the right side of the funnel plot, which increased the estimated effect size to 0.16 (95% CI 0.01–0.30), reaching statistical significance. In addition, Duval and Tweedie's 'trim and fill' method yielded three potentially missing studies on the left side of the funnel plot, increasing the effect size on TC to -1.69 (95% CI -2.10to -1.29; two potentially missing studies on the left side of the plot increasing the effect size on TG to -0.64 (95%) CI - 0.79 to - 0.49); three potentially missing studies on the left side of the plot increasing the effect size on LDL-C to -1.65 (95% CI -1.98 to -1.32); one potentially missing study on the left side of the plot increasing the effect size on non HDL-C to -1.74 (95% CI -2.14 to -1.34); four potentially missing studies on the left side of the funnel increasing the effect size on VLDL to - 0.61 (95% CI -0.75 to -0.46); two potentially missing studies on the left side of the plot increasing the effect size on Lp(a) to -1.11(95% CI - 1.48 to - 0.73), four potentially missing studies on the left side of the plot increasing the effect size on Apo B to -1.83 (95% CI -2.21 to -1.45) and seven potentially missing studies on the left side of the plot that lowered the effect size on Apo A-I to -0.45 (95% CI -0.68 to -0.21). However, Begg's rank correlation confirmed the presence of publication bias only for LDL-C (p = 0.03), non-HDL-C (p = 0.047) and Apo B (p = 0.03) and similar results were not observed even with Egger's linear regression method (p > 0.05 always).

The classic fail-safe *N* test suggested that 1311 studies with negative results would be needed to bring the estimated effect size on TC to a non-significant level (p > 0.05); 147 studies with negative results would be needed to bring the estimated effect size on TG to a non-significant level (p > 0.05); 2476 studies with negative results would be needed to bring the estimated effect size on non LDL-C to a non-significant level (p > 0.05); 170 studies with negative results would be needed to bring the estimated effect size on VLDL to a non-significant level (p > 0.05); 419 studies with negative results would be needed to bring the estimated effect size on Lp(a) to a non-significant level (p > 0.05); 2789 studies with negative results would be needed to bring the estimated effect size on Apo B to a non-significant level (p > 0.05) and 27 studies with negative results would be needed to bring the estimated effect size on Apo A-I to a non-significant level (p > 0.05).

3.4 Safety Analysis

Primary outcomes were adverse events leading to discontinuation of treatment, death, MACE (i.e. acute myocardial infarction, coronary artery disease, angina pectoris, unstable angina, supraventricular extrasystoles, cardiac failure, ischaemic stroke), injection-site reaction (i.e. injection-site erythema, discolouration, pain, swelling, pruritus, haematoma, induration, discomfort, inflammation, nodule), fatigue, headache, dizziness, flu, flu-like symptoms, nasopharyngitis, cough, muscle symptoms (i.e. muscle fatigue, myalgia, muscle spasms, muscle stiffness), back pain, pain in extremity, chest pain, hepatic steatosis, nausea, constipation, diarrhoea, upper abdominal pain, lower abdominal pain, upper respiratory tract infection, urinary tract infection, creatinine elevation, hepatic enzymes elevation and C-reactive protein (CRP) elevation.

Mipomersen was positively associated with an increased risk of discontinuation of treatment, injection-site reaction, hepatic steatosis, hepatic enzymes elevation and flu-like symptoms (Fig. 6; Table S1 [see ESM]). These findings were robust in the leave-one-out sensitivity analyses (Fig. S3, see ESM). The incidence of the other adverse events did not differ between groups (Table S1, see ESM).

Visually, the funnel plot of standard error by log odds ratio was slightly asymmetric for risk of injection-site reaction and hepatic steatosis. This asymmetry was imputed to ten potentially missing studies on the left side of the funnel plot that lowered the estimated risk of injection-site reaction to 9.71 (95% CI 6.84-13.79) and two potentially missing studies on the left side of the plot that lowered the estimated risk of hepatic steatosis to 3.92 (95% CI 1.74-8.84) (Fig. S4, see ESM). In addition, Duval and Tweedie's 'trim and fill' method yielded three potentially missing studies on the right side of the funnel plot, increasing the estimated risk of hepatic enzymes elevation to 4.46 (95% CI 2.64-7.53), and two potentially missing studies always on the right side of the plot, increasing the estimated risk of flu-like symptoms to 2.07 (95% CI 1.49-2.87) (Fig. S5, see ESM). However, Egger's linear regression and Begg's rank correlation did not confirm the presence of any publication bias in the current meta-analysis (p > 0.05 always). Finally, the classic failsafe N test suggested that 55 studies with negative results would be needed to bring the estimated risk of treatment discontinuation to a non-significant level (p > 0.05); 1037 studies with negative results would be needed to bring the estimated risk for injection-site reaction to a non-significant level (p > 0.05); 12 studies with negative results would be needed to bring the estimated risk of hepatic steatosis to







Fig. 5 Funnel plot detailing publication bias in the studies reporting the impact of mipomersen on serum lipid concentrations. *HDL* high-density lipoprotein, *LDL* low-density lipoprotein, *VLDL* very-low-density lipoprotein

a non-significant level (p > 0.05); 28 studies with negative results would be needed to bring the estimated risk of hepatic enzymes elevation to a non-significant level (p > 0.05) and 34 studies with negative results would be needed to bring the estimated risk of flu-like symptoms to a non-significant level (p > 0.05).

4 Discussion

By analysing data from 13 randomized control studies including a total of 1053 patients, this meta-analysis showed that mipomersen significantly reduced TC (-21.4%); LDL-C (-26.4%); TG (-16.2%); Lp(a) (-22.7%);

VLDL-C (-19.6%); and non-HDL-C (-28.1%) without affecting HDL-C levels (+1.4%). These findings strengthen those previously reported by Panta et al., based on 462 patients [41].

Despite favourable effects on the lipid profile, including Lp(a), some concerns are reinforced from the safety profile. As a matter of fact, mipomersen therapy is more likely discontinued and associated with increased risk of injection-site reactions, hepatic steatosis, hepatic enzyme elevation, and flu-like symptoms. Elevated liver enzymes can be attenuated by reducing the mipomersen dose, reducing dosing frequency, or temporarily stopping mipomersen [27, 32, 35, 42]. However, there is limited evidence on the effect of mipomersen on hepatic steatosis or fibrosis assessed by liver

Study name		Statistic	s for eacl	n study		PO	ls ratio and	1 95% CI	Study name		Statistics	for each sti	A pr	Odds ratio and 95% CI
	Odds	Lower	Upper	7 1/6/100	And and a					Odds ratio	Lower U _I limit li	pper imit Z-Va	lue p-Value	
	ralio	Í		-value	D-value				Reeskamp 2019 - I & III	5 426	0.675 4:	3.589 1.5	591 0.112	
Reeskamp, 2019 - I & III	4,234	1,653	10,846	3,007	0,003		<u> </u>	— •	Reeskamp, 2019 - II & N	/ 4,839	0,596 31	9,292 1,4	175 0,140	
Reeskamp, 2019 - II & IV	1,934	0,771	4,850	1,406	0,160		Ŧ	T	Thomas, 2013	2,926	0,624 1,	3,718 1,5	362 0,173	
Waldmann, 2017	15.000	0.645	348.927	1.687	0.092		+		McGowan 2012	6217	0.326 11	8 532 1 3	15 0 2 2 4	
Elaim 2014	1 008	010	30,000	0.055	0 956		-		Viscar 2012	12 100	1 315 11	1 300 2 5	000 0008	
	000,1		100,00	2020	0,000		_			4 964	1 080	7 301 3 4	133 0.001	
	1,320	700,0	144,341	0,40	0,004					201	-		2000	_
I homas, 2013	3,583	1,585	8,102	3,066	0,002		<u> </u>							0,01 0,1 1 10 100
McGowan, 2012	4,645	0,537	40,217	1,395	0,163		+	•						Favours Mipomersen Favours Control
Stein, 2012	10,584	0,601	186,490	1,612	0,107		+		HEPATIC STEATO	SIS				
Visser, 2012	1,176	0,182	7,622	0,170	0,865									
Akdim. 2011 - V	5.000	0.183	136.321	0.954	0.340			Ĩ						
Akdim 2010h - III	2 059	0 077	54 803	0.431	0.666	-+								
	000,1	1000	20146	0200	0,004				Study name	St	atistics for	r each stuk	₽	Odds ratio and 95% CI
	200,0	0,027	01-07		110,0		-		0	ods Low	ver Uppe	ï		
Kaal, 2010	5,164	0,262	101,695	1,080	0,280			•		atio lin	lit limit	t z-Valu	e p-Value	
Visser, 2010	3,632	0,132	99,845	0,763	0,446		ł	ļ			0 10			-
Kastelein. 2006 - III	0.882	0.027	29.146	-0.070	0.944	1	ł	+	Waldmann, 2017	1,286 0,0	044 37,5	17 0°17	588'N G	
Kostoloin 2006 N/		0200	115 220	0 200	0 EEE				Raal, 2016	2,000 0,0	090 44,3	350 0,4:	8 0,661	
					0000				Thomas, 2013 1	7,983 1,0	054 306,7	759 1,96	6 0,046	
	3,018	868'L	4,649	5,UTT	0,000	_	-	_	McGowan, 2012 19	9,868 1,	112 354,8	357 2,00	2 0,042	
					0,0	N 0,1	-	10 100	Stein, 2012	7,602 3,3	223 17,9	331 4,65	3 0,000	•
						Favorire Mine	mercen	Favorire Control	Visser, 2012 32	2,895 1,7	722 628,4	123 2,32	1 0,020	
DISCONTINUATION OF 1	REATMEN	١T							Akdim, 2011 - 111	1.000 0.1	030 33.3	118 0.00	0 1,000	
										000 00	030 33.3	18 0.00	0 1 000	-
										, 000 000	200 000 000			
									Akalm, 2011 - V	'n /co',	284 21/,1	1, 2, 1	1 0,223	
Study name	Statis	tics for eac	th study			Odds ra	io and 95% (Akdim, 2010a - III (0,133 0,0	3,0	381 -1,25	8 0,209	
Odds	ower	llnner							Akdim, 2010a - IV	2,692 0,0	099 /3,1	195 0,56	1990 8	
ratic	limit	limit	Z-Value F	-Value					Akdim, 2010a - V	2,333 0,0	087 62,6	382 0,5(5 0,614	
					-	-	-	-	Akdim, 2010b - I	1,000 0,0	030 33,5	318 0,00	0 1,000	
Keeskamp, 2019 - I & III 4,4,	21 1,460	13,391	2,629	600'0			•	+	Akdim, 2010b - IV (0,500 0,1	023 11,0	388 -0,45	8 0,661	•
Reeskamp, 2019 - II & IV 8,3.	38 2,429	28,766	3,367	0,001					Raal, 2010	1,436 0,4	443 4,6	359 0,60	12 0,547	•
Waldmann, 2017 7,6	15 0,331	175,006	1,269	0,204						3,613 2,0	093 6,2	37 4,61	1 0,000	•
Flaim, 2014 - 1 56,2	36 2,411	1267,821	2,511	0,012										0.01 01 1 10 100
Flaim, 2014 - II 26,6	37 2,178	326,453	2,569	0,010										Equirie Minomoreon Equirie Control
Flaim, 2014 - III 33,4	44 1,458	766,953	2, 196	0,028					HEPATIC ENZYME	S ELEVA	TION			
Thomas, 2013 8,0,	22 3,794	16,961	5,450	0,000			1							
McGowan, 2012 18,9.	58 4,600	78,142	4,072	0000'0			<u> </u>							
Stein, 2012 18,1	18 6,421	51,125	5,473	0,000				•						
Visser, 2012 4,0	0,323	49,596	1,079	0,280			ļ							
Akdim, 2011 - II 17,0	0,446	648,202	1,525	0,127		1			Study name		Statistics	for each stu	Apr	Odds ratio and 95% CI
Akdim, 2011 - III 17.0	20 0,446	648,202	1.525	0,127		1		Î		Odds	lower IIr	ner		
Akdim 2011 - IV 17.00	0.446	648,202	1.525	0.127		1				ratio	limit	imit 7-Va	auley-u au	
Akrim 2011 - V 17 0	D 0.446	CUC 878	1 525	0 127		-								-
Akrim 2010a - 1 2010	need 0	404 B01	1 735	0.085					Reeskamp, 2019 - I & III	3,080	1,395	6,799 2,	785 0,005	•
	22 0 0E0	025.016	1 074	0061					Reeskamp, 2019 - II & N	/ 1,403	0,616	3,192 0,	306 0,420	+
		235,000	1 024	0.067				-	Waldmann, 2017	5,400	0,232 12	5,614 1,	050 0,294	
		022 0202	5	200				-	Thomas, 2013	1,945	0,893	4,234 1,	375 0,094	
		1213,173	2,211	0,020					McGowan, 2012	3,214	0,903 1	1,445 1,4	302 0,072	•
	6/7'L /C	100,6271	101 /z	ann'n					Stein, 2012	2,103	0,958	4,614 1,8	353 0,064	•
Akdim, 2010b - 1 85,0	00 1,319	5478,065	2,090	0,037					Visser, 2012	0,909	0.220	3.758 -0.	132 0,895	•
Akdim, 2010b - II 17,0.	0,446	648,202	1,525	0,127		•			Akdim. 2011 - IV	0.600	0.027 1.	3.582 -0.5	321 0.748	•
Akdim, 2010b - III 23,0	0,613	862,864	1,695	06000					Akdim 2011 - V	1 923	0,066,55	5 839 0	380 0.704	
Akdim, 2010b - IV 28,3.	33 0,858	935,916	1,874	0,061					Abdim 2010a - 1	1 400	0.045 4	3 780 0.5	102 0848	
Raal 2010 10.54	33 2.677	41.672	3,366	0.001						- ,+00	1 1 0 0 0 0 0	0,100 0,	132 0,040	
Visser 2010 8.6	17 0.390	191,584	1.365	0.172				Î		260,2	- een'n	0,130 0,		
Kastelein 2006 - I 7.8	57 0.284	217,106	1.217	0.223						110'0	N 10000		200 0,020	
Kaetalain 2006 - 11 5.00	n 0.183	136 321	0.054	0.340						2,092	0,033 7	3,180 U,	100,000	
		005 046	1020				•	•	Akdim, 2010a - V	1,600	0,104 2	4,703 0,	337 0,736	
		900 A 10	1,0/4	0,001					Akdim, 2010b - III	0,714	0,022 2	3,305 -0,	189 0,850	
		7005'/CON7	1,491	0, 130					Raal, 2010	1,354	0,354	5,180 0,-	143 0,658	
11,4	10 7,8/8	16,524	12,884	0,000	_		_	•	Visser, 2010	10,500	1,360 8	1,053 2,	255 0,024	
					0,01	0,1	-	10 100		2,016	1,447	2,807 4,	148 0,000	•
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	2									civio				

Fig. 6 Forest plot comparing the risk of adverse events associated with mipomersen treatment. Cl confidence intervals

biopsies [43], and so longer-term follow-up studies with mipomersen are needed to evaluate the longer-term risk for permanent hepatic injury and adverse histological change.

Interestingly, only one of the currently available controlled clinical trials enrolled patients with HoFH [38]; consequently, the current pharmacological indication of mipomersen seems not to be supported by an adequate body of evidence.

Results from clinical trials conducted with drugs affecting serum LDL-C levels have shown that every 1 mmol/L (39 mg/dL) reduction in LDL-C levels is associated with a 19% reduction in coronary mortality as well as a 22% reduction in MACE [44]. In accordance with the latest European guidelines [45], LDL-C targets in HoFH are < 2.5 mmol/L (<100 mg/dL) [<3.5 mmol/L (<135 mg/dL) in children], or <1.8 mmol/L (<70 mg/dL) in adults with clinical atherosclerotic cardiovascular disease. However, the majority of FH patients do not reach their LDL-C guideline recommended goals, although mipomersen has been shown to increase the attainment rates. Nevertheless, the MACE rate in the mipomersen trials was 9.5 events/1000 months of treatment, which is equivalent to an 11.4% annualized event rate [46]. Conversely, the lomitapide studies reported a rate of 1.7 MACE events/1000 months (equivalent to a 2.0% annualized event rate) [47, 48], and the TAUSSIG (Trial Assessing Long-Term Use of PCSK9 Inhibition in Subjects with Genetic LDL Disorders) trial had a MACE rate of 1.8 events/1000 months (equivalent to a 2.1% annualized event rate) [49]. For this reason, lomitapide and the PCSK9 inhibitor evolocumab may represent an optimal therapeutic opportunity for homozygous FH patients as an alternative to mipomersen. As a matter of fact, both TESLA (Trial Evaluating PCSK9 Antibody in Subjects With LDL Receptor Abnormalities) and TAUSSIG clinical trials showed that the PCSK9 inhibitor evolocumab promotes a 20-30% reduction in LDL-C in HoFH patients on top of conventional lipidlowering therapies, and also lomitapide seems to be pretty promising [49-52], with a better safety profile than mipomersen. In addition, the use of evolocumab as an adjunctive therapy for HoFH subjects seems to be further supported by the markedly elevated PCSK9 levels that HoFH patients have in comparison with HeFH or non-FH subjects [53].

Our analysis also shows that mipomersen reduces Lp(a) by 22.7%, which represents an independent risk factor for the development of cardiovascular disease [54]. This effect, which has just been investigated by Nandakumar et al. [55], seems to be due to an increased fractional catabolic rate (FCR) of Lp(a) and could definitely contribute to the reduction of the CV risk, especially in patients with high Lp(a) levels at baseline. However, in the light recent statements by Burgess et al., this essay has to be critically evaluated [56].

Finally, mipomersen antisense inhibition of apoB synthesis reduces plasma concentrations of apolipoprotein C-III

(apo C-III) and apo C-III-containing lipoproteins by which TG and VLDL-C decrease [57]. Remarkably, this is of particular importance since lower concentrations of apoC-III and LDL with apoC-III have been associated with reduced risk of cardiovascular disease regardless of other traditional risk factors [57, 58].

In conclusion, the current meta-analysis demonstrates the positive effects of mipomersen on lipid profile, but also emphasizes the multiple adverse effects exerted by mipomersen with uncertainty regarding long-term effects, such as risk for hepatic injury. To date, there is no apparent reason to stop taking mipomersen when it is effective and well tolerated, even though a careful monitoring of liver function is needed. Furthermore, the PCSK9 monoclonal antibodies should be considered as a valid alternative to mipomersen for the treatment of HoFH patients. As a matter of fact, the development of the PCSK9 inhibitor evolocumab showed efficacy similar to mipomersen with respect to LDL-C levels, with a better safety profile.

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