

Optimal multi-dimensional stochastic harvesting with density-dependent prices

Luis H. R. Alvarez1 · Edward Lungu² · Bernt Øksendal3,**⁴**

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Abstract We prove a verification theorem for a class of singular control problems which model optimal harvesting with density-dependent prices or optimal dividend policy with capital-dependent utilities. The result is applied to solve explicitly some examples of such optimal harvesting/optimal dividend problems. In particular, we show that if the unit price *decreases* with population density, then the optimal harvesting policy may not exist in the ordinary sense, but can be expressed as a "chattering policy", i.e. the limit as Δx and Δt go to 0 of taking out a sequence of small quantities of size Δx within small time periods of size Δt .

Keywords Optimal harvesting · Interacting populations · Itô diffusions · Singular stochastic control · Verification theorem · Density-dependent prices · Chattering policies

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Edward Lungu lungu@mopipi.ub.bw

¹ Department of Accounting and Finance, Turku School of Economics, University of Turku, 20014 Turku, Finland

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 \boxtimes Bernt Øksendal oksendal@math.uio.no

Luis H. R. Alvarez luis.alvarez@tse.fi

² Department of Mathematics, University of Botswana, B.P. 0022, Gaborone, Botswana

³ Department of Mathematics, University of Oslo, Box 1053, Blindern, 0316 Oslo, Norway

⁴ Norwegian School of Economics, Helleveien 30, 5045 Bergen, Norway

1 Introduction

The determination of an optimal harvesting policy of a stochastically fluctuating renewable resource is typically subject to at least three key factors affecting either the intertemporal evolution of the resource stock or the incentives of a rational risk neutral harvester. First, the exact size of the harvested stock evolves stochastically due to environmental or demographical randomness. Second, the interaction between different populations has obviously a direct effect on the density of the harvested stocks. Third, most harvesting decisions are subject to density dependent costs and prices. The price of the harvested resource is typically decreasing as a function of the prevailing stock due to the decreasing marginal utility of consumption. The more abundant a resource gets, the less consumers are prepared to pay from an extra unit of that particular resource and vice versa. In a completely analogous fashion the costs associated with harvesting depend typically on the abundance of the harvested resource. The scarcer a resource becomes, the higher are the costs associated with harvesting due to costly search or other similar factors. Our objective in this study is to investigate the optimal harvesting policy of a risk neutral decision maker facing all the three key factors mentioned above.

The problem of determining an optimal harvesting policy of a risk neutral decision maker can be viewed as a singular stochastic control problem. In an unstructured one-dimensional setting where the marginal profitability of a marginal unit of the harvested stock is a constant, the existing literature usually delineates circumstances under which the optimal harvesting policy is to deplete the entire resource stock immediately or to maintain it at all times below a critical threshold at which the expected present value of the cumulative yield is maximized [\[1](#page-14-0)[,3](#page-14-1)[,4,](#page-14-2)[6](#page-14-3)[–8](#page-14-4)]. As intuitively is clear, the optimal policy is altered as soon as the marginal profitability becomes state-dependent (cf. [\[2\]](#page-14-5)) or population interaction (cf. [\[9\]](#page-15-0)) is incorporated into the analysis. In [\[2](#page-14-5)] it is shown within a one-dimensional setting that the state dependence of the instantaneous yield from harvesting results into the emergence of circumstances under which the policy resulting into the maximal value constitutes a chattering policy which does not belong into the original class of admissible càdlàg-harvesting policies. On the other hand, in [\[9](#page-15-0)] it is shown that the presence of interaction between the harvested resource stocks leads to a harvesting strategy where the decision maker generically harvests only a single resource at a time.

In this paper we combine the approaches developed in $[2,9]$ $[2,9]$ and consider the problem of determining the optimal harvesting policy from a collection of interacting populations, described by a coupled system of stochastic differential equations, when the price per unit for each population is allowed to depend on the densities of the populations. In Sect. [2](#page-1-0) we give a general verification theorem for such optimal harvesting problems (Theorem [2.1\)](#page-4-0), and in Sect. [3](#page-7-0) we study in detail some examples where the price is a decreasing function of the density and we show, perhaps surprisingly, that in such cases the optimal harvesting strategy may not exist in the ordinary sense, but can be described as a "chattering policy". See Theorems [3.2](#page-9-0) and [3.4.](#page-12-0)

2 The main result

We now describe our model in detail. This presentation follows [\[9\]](#page-15-0) closely. Consider *n* populations whose sizes or densities $X_1(t), \ldots, X_n(t)$ at time *t* are described by a system of *n* stochastic differential equations of the form

$$
dX_i(t) = b_i(t, X(t))dt + \sum_{j=1}^{m} \sigma_{ij}(t, X(t))dB_j(t); \quad 0 \le s \le t \le T
$$
 (2.1)

$$
X_i(s) = x_i \in \mathbb{R}; \quad 1 \le i \le n,
$$
\n^(2.2)

where $B(t) = (B_1(t), \ldots, B_m(t)); t > 0, \omega \in \Omega$ is *m*-dimensional Brownian motion on a filtered probability space $(\Omega, \mathcal{F}, \mathbb{F} := {\mathcal{F}_t}_{t>0}, P)$ and the differentials (i.e. the corresponding integrals) are interpreted in the Itô sense. We assume that $b = (b_1, \ldots, b_n) : \mathbb{R}^{1+n} \to \mathbb{R}^n$ and $\sigma = (\sigma_{ij})_{\substack{1 \le i \le n \\ 1 \le j \le m}}$ and $\sigma = (\sigma_{ij})_{1 \le i \le n} : \mathbb{R}^{1+n} \to \mathbb{R}^{n \times m}$ are given continuous functions. We also assume that the terminal time $T = T(\omega)$ has the form

$$
T(\omega) = \inf \{ t > s; (t, X(t)) \notin S \}
$$
\n
$$
(2.3)
$$

where $S \subset \mathbb{R}^{1+n}$ is a given set. For simplicity we will assume in this paper that

$$
S = (0, T) \times U
$$

where *U* is an open, connected set in \mathbb{R}^n . We may interprete *U* as the *survival set* and *T* is the *time of extinction* or simply the *closing/terminal* time.

We now introduce a *harvesting strategy* for this family of populations:

A *harvesting strategy* γ is a stochastic process $\gamma(t) = \gamma(t, \omega) = (\gamma_1(t, \omega), \dots, \gamma_n(t, \omega))$ $\in \mathbb{R}^n$ with the following properties:

For each $t \geq s\gamma(t, \cdot)$ is measurable with respect to the σ -algebra \mathcal{F}_t generated by ${B(s, \cdot)}$; $s < t$ }. In other words: $\gamma(\cdot)$ is F-*adapted*. (2.4) $\gamma_i(t, \omega)$ is non-decreasing with respect to *t*, for a.a. $\omega \in \Omega$ and all $i = 1, \ldots, n$ (2.5) $t \to \gamma(t, \omega)$ is right-continuous, for a.a. ω (2.6) $\gamma(s, \omega) = 0$ for a.a. ω . (2.7)

Component number *i* of $\gamma(t, \omega)$, $\gamma_i(t, \omega)$, represents *the total amount harvested from population number i up to time t*.

If we apply a harvesting strategy γ to our family $X(t) = (X_1(t), \ldots, X_n(t))$ of populations the harvested family $X^{(\gamma)}(t)$ will satisfy the *n*-dimensional stochastic differential equation

$$
\begin{cases} dX^{(\gamma)}(t) = b(t, X^{(\gamma)}(t))dt + \sigma(t, X^{(\gamma)}(t))dB(t) - d\gamma(t); \quad s \le t \le T\\ X^{(\gamma)}(s^-) = x = (x_1, \dots, x_n) \in \mathbb{R}^n \end{cases}
$$
(2.8)

We let Γ denote the set of all harvesting strategies γ such that the corresponding system [\(2.7\)](#page-2-0) has a unique strong solution $X^{(\gamma)}(t)$ which does not explode in the time interval [s, T] and such that $X^{(\gamma)}(t) \in U$ for all $t \in [s, T]$.

Since we do not exclude immediate harvesting at time $t = s$, it is necessary to distinguish between $X^{(\gamma)}(s)$ and $X^{(\gamma)}(s^-)$: Thus $X^{(\gamma)}(s^-)$ is the state right before harvesting starts at time $t = s$, while

$$
X^{(\gamma)}(s) = X^{(\gamma)}(s^-) - \Delta \gamma
$$

is the state immediately after, if γ consists of an immediate harvest of size $\Delta \gamma$ at $t = s$.

Suppose that *the price per unit of population number i*, when harvested at time *t* and when the current size/density of the vector $X^{(\gamma)}(t)$ of populations is $\xi = (\xi_1, \ldots, \xi_n) \in \mathbb{R}^n$, is given by

$$
\pi_i(t,\xi); \quad (t,\xi) \in S, \quad 1 \le i \le n,
$$
\n(2.9)

where the $\pi_i : S \to \mathbb{R}$; $1 \leq i \leq n$, are lower bounded continuous functions. We call such prices *density-dependent* since they depend on ξ . The total expected discounted utility harvested from time *s* to time *T* is given by

$$
J^{(\gamma)}(s, x) := E^{s, x} \left[\int\limits_{[s, T]} \pi(t, X^{(\gamma)}(t^{-})) \cdot d\gamma(t) \right]
$$
(2.10)

where $\pi = (\pi_1, \dots, \pi_n)$, $\pi \cdot d\gamma = \sum_{i=1}^n \pi_i d\gamma_i$ and $E^{s,x}$ denotes the expectation with respect to the probability law $Q^{s,x}$ of the time-state process

$$
Y^{s,x}(t) = Y^{\gamma,s,x}(t) = (t, X^{(\gamma)}(t)); \quad t \ge s \tag{2.11}
$$

assuming that $Y^{s,x}(s^-) = x$.

The *optimal harvesting problem* is to find the *value function* $\Phi(s, x)$ and an *optimal harvesting strategy* $\gamma^* \in \Gamma$ such that

$$
\Phi(s, x) := \sup_{\gamma \in \Gamma} J^{(\gamma)}(s, x) = J^{(\gamma^*)}(s, x). \tag{2.12}
$$

This problem differs from the problems considered in [\[1](#page-14-0)[,3](#page-14-1)[,4](#page-14-2)[,8](#page-14-4)[,9\]](#page-15-0) in that the prices $\pi_i(t,\xi)$ are allowed to be density-dependent. This allows for more realistic models. For example, it is usually the case that if a type of fish, say population number *i*, becomes more scarce, the price per unit of this fish increases. Conversely, if a type of fish becomes abundant then the price per unit goes down. Thus in this case the price $\pi_i(t,\xi) = \pi_i(t,\xi_1,\ldots,\xi_n)$ is a *nonincreasing* function of ξ_i . One can also have situations where $\pi_i(t, \xi)$ depends on all the other population densities ξ1,...,ξ*ⁿ* in a similar way.

It turns out that if we allow the prices to be density-dependent, a number of new – and perhaps surprising – phenomena occurs. The purpose of this paper is not to give a complete discussion of the situation, but to consider some illustrative examples.

Remark Note that we can also give the problem [\(2.12\)](#page-3-0) an economic interpretation: We can regard $X_i(t)$ as the value at time *t* of an economic quantity or asset and we can let $\gamma_i(t)$ represent the total amount paid in dividends from asset number *i* up to time *t*. Then *S* can be interpreted as the *solvency set*, *T* as the *time of bankruptcy* and $\pi_i(t, \xi)$ as the *utility rate* of dividends from asset number *i* at the state (t, ξ) . Then [\(2.12\)](#page-3-0) becomes the problem of finding the *optimal stream of dividends*. This interpretation is used in [\[5](#page-14-6)] (in the density-independent utility case). See also [\[9\]](#page-15-0).

In the following H^0 denotes the interior of a set H , \bar{H} denotes its closure.

If $G \subset \mathbb{R}^k$ is an open set we let $C^2(G)$ denote the set of real valued twice continuously differentiable functions on *G*. We let $C_0^2(G)$ denote the set of functions in $C^2(G)$ with compact support in *G*.

If we do not apply any harvesting, then the corresponding time-state population process $Y(t) = (t, X(t))$, with $X(t)$ given by [\(2.1\)](#page-2-1)–[\(2.2\)](#page-2-1), is an Itô diffusion whose generator coincides on $C_0^2(\mathbb{R}^{1+n})$ with the partial differential operator *L* given by

$$
Lg(s,x) = \frac{\partial g}{\partial s}(s,x) + \sum_{i=1}^{n} b_i(s,x) \frac{\partial g}{\partial x_i}(s,x) + \frac{1}{2} \sum_{i,j=1}^{n} \left(\sigma \sigma^T\right)_{ij}(s,x) \frac{\partial^2 g}{\partial s \partial x}
$$
(2.13)

for all functions $g \in C^2(S)$.

The following result is a generalization to the multi-dimensional case of Theorem 1 in [\[2\]](#page-14-5) and a generalization to density-dependent prices of Theorem 2.1 in [\[9\]](#page-15-0). For completeness we give the proof.

Theorem 2.1 *Assume that*

$$
\pi(t,\xi)
$$
 is non increasing with respect to ξ_1,\ldots,ξ_n , for all t. (2.14)

(a) *Suppose* $\varphi > 0$ *is a function in* $C^2(S)$ *satisfying the following conditions*

(i) $\frac{\partial \varphi}{\partial x_i}(t, x) \ge \pi_i(t, x)$ *for all* $(t, x) \in S$ (ii) $L\varphi(t, x) \leq 0$ *for all* $(t, x) \in S$.

Then

$$
\varphi(s, x) \ge \Phi(s, x) \quad \text{for all } (s, x) \in S. \tag{2.15}
$$

(b) *Define the non intervention region D by*

$$
D = \left\{ (t, x) \in S; \frac{\partial \varphi}{\partial x_i}(t, x) > \pi_i(t, x) \text{ for all } i = 1, \dots, n \right\}.
$$
 (2.16)

Suppose that, in addition to (i) *and* (ii) *above,*

(iii) $L\varphi(t, x) = 0$ *for all* $(t, x) \in D$

and that there exists a harvesting strategy $\hat{\gamma} \in \Gamma$ *such that the following, (iv)–(vii), hold:*

- (iv) $X^{(\hat{\gamma})}(t) \in \overline{D}$ for all $t \in [s, T]$
- (v) $\left(\frac{\partial \varphi}{\partial x_i}(t, X^{(\hat{\gamma})}(t)) \pi_i(t, X^{(\hat{\gamma})}(t))\right) \cdot d\hat{\gamma}_i^{(c)}(t) = 0; \quad 1 \le i \le n$ (i.e. $\hat{\gamma}_i^{(c)}$ increases only *when* $\frac{\partial \varphi}{\partial x_i} = \pi_i$ *)* and
- $(vi) \varphi(t_k, X^{(\hat{\gamma})}(t_k)) \varphi(t_k, X^{(\hat{\gamma})}(t_k^-)) = -\pi_i(t_k, X^{(\hat{\gamma})}(t_k^-)) \cdot \Delta \hat{\gamma}(t_k)$ *at all jumping times* $t_k \in [s, T)$ *of* $\hat{\gamma}(t)$ *, where*

$$
\Delta \hat{\gamma}(t_k) = \hat{\gamma}(t_k) - \hat{\gamma}(t_k^-)
$$

and

 (vii) $E^{s,x} \big[\varphi(T_R, X^{(\hat{\gamma})}(T_R)) \big] \to 0$ *as* $R \to \infty$ *where*

$$
T_R = T \wedge R \wedge \inf \left\{ t > s; |X^{(\hat{\gamma})}(t)| \ge R \right\}; \qquad R > 0.
$$

Then

$$
\varphi(s, x) = \Phi(s, x) \quad \text{for all } (s, x) \in S \tag{2.17}
$$

and

 $\gamma^* := \hat{\gamma}$ *is an optimal harvesting strategy.*

Proof (a) Choose $\gamma \in \Gamma$ and $(s, x) \in S$. Then by Itô's formula for semimartingales (the Doléans-Dade-Meyer formula) [\[10,](#page-15-1) Th. II.7.33] we have

$$
E^{s,x}[\varphi(T_R, X^{(\gamma)}(T_R^-))] = E^{s,x}[\varphi(s, X^{(\gamma)}(s))]
$$

+
$$
E^{s,x}[\int_s^T \frac{\partial \varphi}{\partial t}(t, X^{(\gamma)}(t))dt + \int_{(s,T_R)} \sum_{i=1}^n \frac{\partial \varphi}{\partial x_i}(t, X^{(\gamma)}(t^-))dX_i^{(\gamma)}(t)
$$

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$$
+ \sum_{i,j=1}^{n} \int_{s}^{T_R} \frac{1}{2} (\sigma \sigma^T)_{ij} (t, X^{(\gamma)}(t)) \frac{\partial^2 \varphi}{\partial x_i \partial x_j} (t, X^{(\gamma)}(t)) dt + \sum_{s < t_k < T_R} \left\{ \varphi(t_k, X^{(\gamma)}(t_k)) - \varphi(t_k, X^{(\gamma)}(t_k^-)) - \sum_{i=1}^{n} \frac{\partial \varphi}{\partial x_i} (t_k, X^{(\gamma)}(t_k^-)) \Delta X_i^{(\gamma)}(t_k) \right\} \right],
$$
\n(2.18)

where the sum is taken over all jumping times $t_k \in (s, T_R)$ of $\gamma(t)$ and

$$
\Delta X_i^{(\gamma)}(t_k) = X_i^{(\gamma)}(t_k) - X_i^{(\gamma)}(t_k^-).
$$

Let $\gamma^{(c)}(t)$ denote the continuous part of $\gamma(t)$, i.e.

$$
\gamma^{(c)}(t) = \gamma(t) - \sum_{s \le t_k \le t} \Delta \gamma(t_k).
$$

Then, since $\Delta X_i^{(\gamma)}(t_k) = -\Delta \gamma_i(t_k)$ we see that [\(2.18\)](#page-5-0) can be written

$$
E^{s,x}[\varphi(T_R, X^{(\gamma)}(T_R^-))] = \varphi(s, x)
$$

+
$$
E^{s,x}[\int_s^T \left\{ \frac{\partial \varphi}{\partial t} + \sum_{i=1}^n b_i \frac{\partial \varphi}{\partial x_i} + \frac{1}{2} \sum_{i,j=1}^n (\sigma \sigma^T)_{ij} \frac{\partial^2 \varphi}{\partial x_i \partial x_j} \right\} (t, X^{(\gamma)}(t)) dt]
$$

-
$$
E^{s,x}[\int_s^T \sum_{i=1}^n \frac{\partial \varphi}{\partial x_i} (t, X^{(\gamma)}(t)) d\gamma_i^{(c)}(t)]
$$

+
$$
E^{s,x}[\sum_{s \le t_k < T_R} \Delta \varphi(t_k, X^{(\gamma)}(t_k))]
$$
(2.19)

where

$$
\Delta \varphi(t_k, X^{(\gamma)}(t_k)) = \varphi(t_k, X^{(\gamma)}(t_k)) - \varphi(t_k, X^{(\gamma)}(t_k^{-})).
$$

Therefore

$$
\varphi(s, x) = E^{s, x} [\varphi(T_R, X^{(\gamma)}(T_R^-))] - E^{s, x} \Big[\int_s^{T_R} L \varphi(t, X^{(\gamma)}(t)) dt \Big] + E^{s, x} \Big[\int_s^{T_R} \sum_{i=1}^n \frac{\partial \varphi}{\partial x_i}(t, X^{(\gamma)}(t)) d\gamma_i^{(c)}(t) \Big] - E^{s, x} \Big[\sum_{s \le t_k < T_R} \Delta \varphi(t_k, X^{(\gamma)}(t_k)) \Big].
$$
\n(2.20)

Let $y = y(r)$; $0 \le r \le 1$ be a smooth curve in *U* from $X^{(\gamma)}(t_k)$ to $X^{(\gamma)}(t_k^-) = X^{(\gamma)}(t_k) +$ $\Delta \gamma(t_k)$. Then 1

$$
-\Delta\varphi(t_k, X^{(\gamma)}(t_k)) = \int\limits_o^1 \nabla\varphi(t_k, y(r)) dy(r).
$$
 (2.21)

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We may assume that

$$
dy_i(r) \ge 0 \qquad \text{for all } i, r.
$$

Now suppose that (i) and (ii) hold. Then by (2.20) and (2.21) we have

$$
\varphi(s,x) \ge E^{s,x} \Big[\int_s^{T_R} \sum_{i=1}^n \pi_i(t, X^{(\gamma)}(t)) d\gamma_i^{(c)}(t) \Big] + E^{s,x} \Big[\sum_{s \le t_k < T_R} \Big(\int_0^1 \sum_{i=1}^n \pi_i(t_k, y(r)) d\gamma_i(r) \Big) \Big]
$$
\n(2.22)

Since we have assumed that $\pi_i(t,\xi)$ is *nonincreasing* with respect to ξ_1,\ldots,ξ_n we have

$$
\pi_i(t_k, X^{(\gamma)}(t_k^-)) \leq \pi_i(t_k, y(r)) \leq \pi_i(t_k, X^{(\gamma)}(t_k))
$$

for all *i*, *k* and $r \in [0, 1]$. Hence

$$
\int_{0}^{1} \pi_{i}(t_{k}, y(r)) dy_{i}(r) \geq \pi_{i}(t_{k}, X^{(\gamma)}(t_{k}^{-})) \cdot \Delta \gamma_{i}(t_{k}).
$$
\n(2.23)

Combined with [\(2.22\)](#page-6-0) this gives

$$
\varphi(s, x) \ge E^{s, x} \Big[\int_{0}^{T_R} \pi(t, X^{(\gamma)}(t)) d\gamma^{(c)}(t) + \sum_{s \le t_k < T} \pi(t_k, X^{(\gamma)}(t_k^{-})) \cdot \Delta \gamma(t_k) \Big] \\ = E^{s, x} \Big[\int_{[s, T_R)} \pi(t, X^{(\gamma)}(t^{-})) d\gamma(t) \Big]. \tag{2.24}
$$

Letting $R \to \infty$ we obtain $\varphi(s, x) \geq J^{(\gamma)}(s, x)$. Since $\gamma \in \Gamma$ was arbitrary we conclude that (2.15) holds. Hence (a) is proved.

(b) Next, suppose that (iii)–(vii) also hold. Then if we apply the argument above to $\gamma = \hat{\gamma}$ we get in [\(2.20\)](#page-5-1) the following:

$$
\varphi(s, x) = E^{s,x} [\varphi(T_R, X^{(\hat{\gamma})}(T_R^-))]
$$
\n
$$
+ E^{s,x} \Big[\int_0^{T_R} \pi(t, X^{(\hat{\gamma})}(t)) \cdot d\hat{\gamma}^{(c)}(t) + \sum_{s \le t_k < T_R} \pi(t_k, X^{(\hat{\gamma})}(t_k^-)) \cdot \Delta \hat{\gamma}(t_k) \Big]
$$
\n
$$
= E^{s,x} [\varphi(T_R, X^{(\hat{\gamma})}(T_R^-))] + E^{s,x} \Big[\int_{[s, T_R)} \pi(t, X^{(\hat{\gamma})}(t)) \cdot d\hat{\gamma}(t) \Big]
$$
\n
$$
\longrightarrow J^{(\hat{\gamma})}(s, x) \quad \text{as } R \to \infty.
$$

Hence $\varphi(s, x) = J^{(\hat{\gamma})}(s, x) \leq \Phi(s, x)$. Combining this with [\(2.14\)](#page-4-2) from (a) we get the conclusion [\(2.16\)](#page-4-3) of part (b). This completes the proof of Theorem [2.1.](#page-4-0)

If we specialize to the 1-dimensional case with just one population $X^{(\gamma)}(t)$ given by

$$
\begin{cases} dX^{(\gamma)}(t) = b(t, X^{(\gamma)}(t))dt + \sigma(t, X^{(\gamma)}(t))dB(t) - d\gamma(t); \quad t \ge s \\ X^{(\gamma)}(s^-) = x \in \mathbb{R} \end{cases}
$$
(2.25)

then Theorem [2.1](#page-4-0) (a) gets the form (see also [\[2,](#page-14-5) Lemma 1])

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Corollary 2.2 *Assume that*

$$
\xi \to \pi(t,\xi); \ \xi \in \mathbb{R} \ \ \text{is nonincreasing for all } t \in [0,T] \tag{2.26}
$$

$$
\varphi(t, x) \ge 0 \qquad \text{is a function in } C^2(S) \text{ such that } \tag{2.27}
$$

$$
\frac{\partial \varphi}{\partial x}(t, x) \ge \pi(t, x) \quad \text{for all } (t, x) \in S \tag{2.28}
$$

and

$$
L\varphi(t,x) \le 0 \quad \text{ for all } (t,x) \in S. \tag{2.29}
$$

Then

$$
\varphi(s, x) \ge \Phi(s, x) \quad \text{for all } (s, x) \in S. \tag{2.30}
$$

3 Examples

In this section we apply Theorem [2.1](#page-4-0) or Corollary [2.2](#page-7-1) to some special cases.

Example 3.1 Suppose $X^{(\gamma)}(t) = (X_1^{(\gamma)}(t), X_2^{(\gamma)}(t))$ is given by

$$
dX_i^{(y)}(t) = \mu_i dt + \sigma_i dB_i(t) - d\gamma_i(t); \quad t \ge s
$$

\n
$$
X_i^{(y)}(s) = x_i > 0
$$
\n(3.1)

where $\mu_i > 0$ and $\sigma_i \neq 0$ are constants; $i = 1, 2$, and $\gamma = (\gamma_1, \gamma_2)$.

We want to maximize the total discounted value of the harvest, given by

$$
J^{(\gamma)}(s,x) = E^{s,x} \left[\int\limits_{[s,T)} e^{-\rho t} \{g_1(X_1^{(\gamma)}(t^-))d\gamma_1(t) + g_2(X_2^{(\gamma)}(t^-))d\gamma_2(t) \right]
$$
(3.2)

where $g_i : \mathbb{R} \to \mathbb{R}$ are given nonincreasing functions (the density-dependent prices) and

$$
T = \inf \left\{ t > s; \min(X_1^{(\gamma)}(t), X_2^{(\gamma)}(t)) \le 0 \right\}
$$
 (3.3)

is the time of extinction, i.e. $S = \{(t, x); x_i > 0; i = 1, 2\}$. The corresponding 1-dimensional case with *g constant* was solved in [\[5](#page-14-6)]. Then it is optimal to do nothing if the population is below a certain treshold *x*∗ > 0 and then harvest according to *local time* of the downward reflected process $\bar{X}(t)$ at $\bar{X}(t) = x^*$.

Now consider the case when

$$
g_i(x) = \theta_i x^{-1/2}
$$
, i.e. $\pi_i(t, x) = e^{-\rho t} \theta_i x^{-1/2}$; $x > 0$, (3.4)

where $\theta_i > 0$ are given constants; $i = 1, 2$. Then the prices increase as the population sizes x_i go to 0, so [\(2.24\)](#page-6-1) holds. Suppose we apply the "*take the money and run*"-strategy $\hat{\gamma}$. This strategy empties the whole population immediately. It can be described by

$$
\stackrel{\circ}{\gamma}(s) = (X_1(s^-), X_2(s^-)) = (x_1, x_2). \tag{3.5}
$$

Such a strategy gives the harvest value

$$
J^{\hat{\psi}}(s,x) = e^{-\rho s} \left(\theta_1 x_1^{-1/2} x_1 + \theta_2 x_2^{-1/2} x_2 \right) = e^{-\rho s} \left(\theta_1 \sqrt{x_1} + \theta_2 \sqrt{x_2} \right); \qquad x_i > 0. \tag{3.6}
$$

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However, it is unlikely that this is the best strategy because it does not take into account that the prices increase as the population sizes go down. So for the two populations $X_i(t)$; $i = 1, 2$, we try the following "chattering policy", denoted by $\widetilde{\gamma}_i = \widetilde{\gamma}_i^{(m,\eta)}$, where *m* is a fixed natural number and $n > 0$. number and $\eta > 0$:

At the times

$$
t_k = \left(s + \frac{k}{m}\eta\right) \wedge T; \qquad k = 1, 2, \dots, m \tag{3.7}
$$

we harvest an amount $\Delta \widetilde{\gamma}_i(t_k)$ which is the fraction $\frac{1}{m}$ of the current population. This gives the expected harvest value the expected harvest value

$$
J^{(\tilde{\gamma}(m,\eta))}(s,x) = E^{s,x} \Big[\sum_{k=1}^{m} e^{-\rho t_k} \Big[\theta_1 \big(X_1^{(\tilde{\gamma})}(t_k^-) \big)^+ \Big)^{-1/2} \Delta \tilde{\gamma}_1(t_k) + \theta_2 \big(X_2^{(\tilde{\gamma})}(t_k^-) \big)^+ \Big)^{-1/2} \Delta \tilde{\gamma}_2(t_k) \Big] \Big], \tag{3.8}
$$

where we have used the notation

$$
x_i^+ = \max(x_i, 0); \qquad x_i \in \mathbb{R}.
$$

Now let $\eta \to 0$, $m \to \infty$. Then all the t_k 's converge to *s* and we get

$$
J^{(\tilde{\gamma}(m,0))}(s, x) := \lim_{\eta \to 0, m \to \infty} J^{(\tilde{\gamma}(m,\eta))}(s, x)
$$

\n
$$
= \lim_{m \to \infty} e^{-\rho s} \left[\sum_{k=1}^{m} \theta_1 \left(x_1 - \frac{k}{m} x_1 \right)^{-1/2} \frac{1}{m} x_1 + \sum_{k=1}^{m} \theta_2 \left(x_2 - \frac{k}{m} x_2 \right)^{-1/2} \frac{1}{m} x_2 \right]
$$

\n
$$
= e^{-\rho s} \left[\theta_1 x_1^{\frac{1}{2}} \int_0^1 (1 - y)^{-\frac{1}{2}} dy + \theta_2 x_2^{\frac{1}{2}} \int_0^1 (1 - y)^{-\frac{1}{2}} dy \right]
$$

\n
$$
= 2e^{-\rho s} \left[\theta_1 \sqrt{x_1} + \theta_2 \sqrt{x_2} \right].
$$
 (3.9)

We conclude that

$$
\sup_{\gamma} J^{(\gamma)}(s, x) \ge 2e^{-\rho s} \left[\theta_1 \sqrt{x_1} + \theta_2 \sqrt{x_2} \right]. \tag{3.10}
$$

We call this policy of applying $\tilde{\gamma}^{(m,\eta)}$ in the limit as $\eta \to 0$ and $m \to \infty$ the *policy of immediate chattering down to 0.* (This limit does not exist as a strategy in Γ .) From [\(3.10\)](#page-8-0) we conclude that

$$
\Phi(s, x) \ge 2e^{-\rho s} \left[\theta_1 \sqrt{x_1} + \theta_2 \sqrt{x_2} \right]. \tag{3.11}
$$

On the other hand, let us check if the function

$$
\varphi(s,x) := 2e^{-\rho s} \left[\theta_1 \sqrt{x_1} + \theta_2 \sqrt{x_2} \right]
$$
\n(3.12)

satisfies the conditions of Theorem [2.1:](#page-4-0) Condition [\(2.14\)](#page-4-2) holds trivially, and (i) of Part (a) holds, since

$$
\frac{\partial \varphi}{\partial x_i}(s, x) = e^{-\rho s} \theta_1 x_1^{-1/2} = \pi_i(s, x).
$$

Now

$$
L = \frac{\partial}{\partial s} + \mu_1 \frac{\partial}{\partial x_1} + \mu_2 \frac{\partial}{\partial x_2} + \frac{1}{2} \sigma_1^2 \frac{\partial^2}{\partial x_1^2} + \frac{1}{2} \sigma_2^2 \frac{\partial^2}{\partial x_2^2},
$$

and therefore

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$$
L\varphi(s, x) = 2e^{-\rho s} \Big[-\rho \left(\theta_1 x_1^{1/2} + \theta_2 x_2^{1/2} \right) + \mu_1 \theta_1 \frac{1}{2} x_1^{-1/2} + \mu_2 \theta_2 \frac{1}{2} x_2^{-1/2} + \frac{1}{2} \sigma_1^2 \frac{1}{2} \left(-\frac{1}{2} \right) \theta_1 x_1^{-3/2} + \frac{1}{2} \sigma_2^2 \frac{1}{2} \left(-\frac{1}{2} \right) x_2^{-3/2} \Big] = -2\rho e^{-\rho s} \Bigg[\theta_1 x_1^{-3/2} \Bigg(x_1^2 - \frac{\mu_1}{2\rho} x_1 + \frac{\sigma_1^2}{8\rho} \Bigg) + \theta_2 x_2^{-3/2} \Bigg(x_2^2 - \frac{\mu_2}{2\rho} x_2 + \frac{\sigma_2^2}{8\rho} \Bigg) \Bigg].
$$

So (ii) of Theorem [2.1](#page-4-0) (a) holds if $\mu_i^2 \le 2\rho \sigma_i^2$ for $i = 1, 2$. By Theorem 2.1 we conclude that $\varphi = \Phi$ in this case.

We have proved part (a) of the following result:

Theorem 3.2 *Let* $X^{(\gamma)}(t)$ *and* T *be given by* [\(3.1\)](#page-7-2) *and* [\(3.3\)](#page-7-3)*, respectively.*

(a) *Assume that*

$$
\mu_i^2 \le 2\rho \sigma_i^2, \quad i = 1, 2. \tag{3.13}
$$

Then

$$
\Phi(s, x) := \sup_{\gamma \in \Gamma} E^{s, x} \Big[\int_{[s, T)} e^{-\rho t} \left\{ \theta_1 X_1^{(\gamma)} (t^-)^{-1/2} d\gamma_1(t) + \theta_2 X_2^{(\gamma)} (t^-)^{-1/2} d\gamma_2(t) \right\} \Big]
$$

= $2e^{-\rho s} \Big[\theta_1 \sqrt{x_1} + \theta_2 \sqrt{x_2} \Big].$ (3.14)

This value is achieved in the limit if we apply the strategy $\widetilde{\gamma}^{(m,\eta)}$ *above with* $\eta \to 0$ *and* $m \to \infty$, *i.e.* by applying the policy of immediate chattering down to 0.

(b) *Assume that*

$$
\mu_i^2 > 2\rho \sigma_i^2; \quad i = 1, 2. \tag{3.15}
$$

Then the value function has the form

$$
\Phi(s,x) = \begin{cases}\ne^{-\rho s} \Big[C_1 \left(e^{\lambda_1^{(1)} x_1} - e^{\lambda_2^{(1)} x_1} \right) + C_2 \left(e^{\lambda_1^{(2)} x_2} - e^{\lambda_2^{(2)} x_2} \right) \Big]; & x_1 \le x_1^*, x_2 \le x_2^* \\
e^{-\rho s} \Big[2\theta_1 \sqrt{x_1} - 2\theta_1 \sqrt{x_1^*} + C_2 \left(e^{\lambda_1^{(2)} x_2} - e^{\lambda_2^{(2)} x_2} \right) + A_1 \Big]; & x_1 > x_1^*, x_2 \le x_2^* \\
e^{-\rho s} \Big[C_1 \left(e^{\lambda_1^{(1)} x_1} - e^{\lambda_2^{(1)} x_1} \right) + 2\theta_2 \sqrt{x_2} - 2\theta_2 \sqrt{x_2^*} + A_2 \Big]; & x_1 \le x_1^*, x_2 > x_2^* \\
e^{-\rho s} \Big[2\theta_1 \sqrt{x_1} - 2\theta_1 \sqrt{x_1^*} + 2\theta_2 \sqrt{x_2} - 2\theta_2 \sqrt{x_2^*} + A_1 + A_2 \Big]; & x_1 > x_1^*, x_2 > x_2^* \\
(3.16)\n\end{cases}
$$

for constants $C_i > 0$, $A_i > 0$ *and* $x_i^* > 0$; $i = 1, 2$ *satisfying the following system of 6 equations (see Remark below):*

$$
C_i \left(e^{\lambda_1^{(i)} x_i^*} - e^{\lambda_2^{(i)} x_i^*} \right) = A_i; \quad i = 1, 2
$$

\n
$$
C_i \left(\lambda_1^{(i)} e^{\lambda_1^{(i)} x_i^*} - \lambda_2^{(i)} e^{\lambda_2^{(i)} x_i^*} \right) = (x_i^*)^{-1/2}; \quad i = 1, 2
$$

\n
$$
C_i \left(\left(\lambda_1^{(i)} \right)^2 e^{\lambda_1^{(i)} x_i^*} - \left(\lambda_2^{(i)} \right)^2 e^{\lambda_2^{(i)} x_i^*} \right) = -\frac{1}{2} (x_i^*)^{-3/2}; \quad i = 1, 2, \quad (3.17)
$$

where

$$
\lambda_1^{(i)} = \sigma_i^{-2} \left[-\mu_i + \sqrt{\mu_i^2 + 2\rho \sigma_i^2} \right] > 0, \quad \lambda_2^{(i)} = \sigma_i^{-2} \left[-\mu_i - \sqrt{\mu_i^2 + 2\rho \sigma_i^2} \right] < 0. \tag{3.18}
$$

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The corresponding optimal policy is the following, for $i = 1, 2$ *:*

If $x_i > x_i^*$ *it is optimal to apply immediate chattering from* x_i *down to* x_i^* *ⁱ* . (3.19)

if $0 < x_i \leq x_i^*$ *it is optimal to apply the harvesting equal to the local time of* (3.20) *the downward reflected process* $X_i(t)$ *at* x_i^* .

(c) *Assume that*

$$
\mu_1^2 > 2\rho\sigma_1^2 \text{ and } \mu_2^2 \le 2\rho\sigma_2^2. \tag{3.21}
$$

Then the value function has the form

$$
\Phi(s, x) = \begin{cases} e^{-\rho s} \Big[C_1 (e^{\lambda_1 x_1} - e^{\lambda_2 x_1}) + 2\theta_2 \sqrt{x_2} \Big]; & 0 \le x_1 < x_1^* \\ e^{-\rho s} \Big[2\sqrt{x_1} - 2\sqrt{x_1^*} + A_1 + 2\theta_2 \sqrt{x_2} \Big]; & x_1^* \le x_1 \end{cases} \tag{3.22}
$$

for constants $C_1 > 0$, $A_1 > 0$ *and* $x_1^* > 0$ *specified by the 3 equations*

$$
C_1 \left(e^{\lambda_1 x_1^*} - e^{\lambda_2 x_1^*} \right) = A_1 \tag{3.23}
$$

$$
eq3.24C_1 \left(\lambda_1 e^{\lambda_1 x_1^*} - \lambda_2 e^{\lambda_2 x_1^*}\right) = \left(x_1^*\right)^{-1/2} \tag{3.24}
$$

$$
C_1 \left(\lambda_1^2 e^{\lambda_1 x_1^*} - \lambda_2^2 e^{\lambda_2 x_1^*} \right) = -\frac{1}{2} \left(x_1^* \right)^{-3/2}, \tag{3.25}
$$

where

$$
\lambda_1 = \sigma_1^{-2} \left[-\mu_1 + \sqrt{\mu_1^2 + 2\rho \sigma_1^2} \right] > 0, \quad \lambda_2 = \sigma_1^{-2} \left[-\mu_1 - \sqrt{\mu_1^2 + 2\rho \sigma_1^2} \right] < 0. \tag{3.26}
$$

The corresponding optimal policy $\gamma^* = (\gamma_1^*, \gamma_2^*)$ *is described as follows:*

If
$$
x_1 > x_1^*
$$
 the optimal γ_1^* is to apply immediate chattering from x_1 down to x_1^* . (3.27)

if $0 < x_1 \le x_1^*$ *the optimal* γ_1^* *is to apply the harvesting equal to the local time of* (3.28) *the downward reflected process* $X_1(t)$ *at* x_1^* .

The optimal policy γ_2^* is to apply immediate chattering from x_2 down to 0.

Proof (b). First note that if we apply the policy of immediate chattering from x_i down to x_i^* , where $0 < x_i^* < x_i$, then the value of the harvested quantity is

$$
e^{-\rho s}\theta_i \int\limits_{0}^{x_i - x_i^*} (x_1 - y)^{-1/2} dy = e^{-\rho s} \theta_i \int\limits_{x_i^*}^{x_i} u^{-1/2} du = 2e^{-\rho s} \theta_i \left(\sqrt{x_i} - \sqrt{x_i^*}\right). \tag{3.29}
$$

This follows by the argument (3.7) – (3.12) above.

To verify [\(3.16\)](#page-9-1)–[\(3.18\)](#page-9-2), first note that $\lambda_1^{(i)}$, $\lambda_2^{(i)}$ are the roots of the quadratic equation

$$
-\rho + \mu_i \lambda + \frac{1}{2} \sigma_i^2 \lambda^2 = 0. \tag{3.30}
$$

Hence, with $\varphi(s, x)$ defined to be the right hand side of (3.16) we have

$$
L\varphi(s, x) = 0 \quad \text{for } x_1 < x_1^*, x_2 < x_2^* \tag{3.31}
$$
\n
$$
L\varphi(s, x) \le 0 \quad \text{for } x_1 > x_1^* \text{ or } x_2 > x_2^* \tag{3.32}
$$

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and

$$
\varphi(s,0) = 0.\tag{3.32}
$$

Note that equations [\(3.17\)](#page-9-3) imply that φ is C^2 at $x_1 = x_1^*$ and at $x_2 = x_2^*$.

We conclude that with this choice of C_i , A_i , x_i^* ; $i = 1, 2$ the function $\varphi(s, x)$ becomes a $C²$ function and the nonintervention region *D* given by [\(2.16\)](#page-4-3) is seen to be

$$
D = \left\{ (s, x) = (s, x_1, x_2); 0 < x_1 < x_1^*, 0 < x_2 < x_2^* \right\}.
$$

Thus we obtain that φ satisfies conditions (i), (ii) of Theorem [2.1](#page-4-0) and hence

$$
\varphi(s, x) \ge \Phi(s, x) \quad \text{for all } s, x. \tag{3.33}
$$

Also, by (3.31) we know that (iii) holds.

Moreover, if $x_i \leq x_i^*$ it is well-known that the local time $\hat{\gamma}_i$ at x_i^* of the downward reflected process $X_i(t)$ at x_i^* satisfies (iv)–(vi) (see e.g. [\[8](#page-14-4)] for more details). And (vii) follows from [\(3.16\)](#page-9-1). By Theorem thm1 (b) we conclude that if $x_i \leq x_i^*$ then $\gamma_i^* := \hat{\gamma}_i$ is optimal for $i = 1, 2$ and $\varphi(s, x) = \Phi(s, x)$. Finally, as seen above, if $x_i > x_i^*$ then immediate chattering from *x_i* down to *x*^{*}_{*i*} gives the value $2e^{-\rho s}\theta_i(\sqrt{x_i} - \sqrt{x_i^*}) + \Phi(s, x^*)$. Hence

$$
\Phi(s, x) \ge 2e^{-\rho s} \theta_i \left(\sqrt{x_i} - \sqrt{x_i^*} \right) + \Phi(s, x^*) \quad \text{for } x_i > x_i^*; i = 1, 2.
$$

Combined with [\(3.33\)](#page-11-0) this shows that

$$
\varphi(s, x) = \Phi(s, x) \quad \text{for all } s, x
$$

and the proof of (b) is complete.

The proof of the mixed case (c) is left to the reader. \Box

Remark Dividing the second equation of [\(3.17\)](#page-9-3) by the third, we get the equation

$$
\frac{\lambda_1^{(i)} e^{\lambda_1^{(i)} x_i^*} - \lambda_2^{(i)} e^{\lambda_2^{(i)} x_i^*}}{\left(\lambda_1^{(i)}\right)^2 e^{\lambda_1^{(i)} x_i^*} - \left(\lambda_2^{(i)}\right)^2 e^{\lambda_2^{(i)} x_i^*}} = -2x_i^*.
$$
\n(3.34)

Since the left hand side of [\(3.34\)](#page-11-1) goes to $(\lambda_1^{(i)} + \lambda_2^{(i)})^{-1} < 0$ as $x_i^* \to 0^+$, and goes to $(\lambda_1^{(i)})^{-1} > 0$ as $x_i^* \to \infty$, we see by the intermediate value theorem that there exist $x_i^* > 0$; *i* = 1, 2 satisfying this equation. With these values of x_i^* ; *i* = 1, 2 we see that there exists a unique solution C_i , A_i ; $i = 1, 2$ of the system [\(3.17\)](#page-9-3).

Example 3.3 The Brownian motion example is perhaps not so good as a model of a biological stock, since Brownian motion is a poor model for population growth. Instead, let us consider a standard population growth model (in the sense that it can be generated from a classic birth-death-process), like the logistic diffusion considered in [\[4\]](#page-14-2). That is, let us consider the problem

$$
V(0, x) = V(x) = \sup_{\gamma \in \Gamma} E^x \int_{[0, T)} e^{-\rho t} X^{-1/2} (t^-) d\gamma(t)
$$
 (3.35)

subject to

$$
dX(t) = \mu X(t)(1 - K^{-1}X(t))dt + \sigma X(t)dB(t) - d\gamma(t), \qquad X(0^-) = x > 0, \quad (3.36)
$$

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where $\mu > 0$, $K^{-1} > 0$, and $\sigma > 0$ are known constants, $B(t)$ denotes a Brownian motion in R, and $T = \inf\{t \ge 0 : X(t) \le 0\}$ denotes the extinction time. We define the mapping $H: \mathbb{R}_+ \mapsto \mathbb{R}_+$ as

$$
H(x) = \int_{0}^{x} y^{-1/2} dy = 2\sqrt{x}.
$$
 (3.37)

The generator *A* of $X(t)$ is given by

$$
A = \frac{1}{2}\sigma^2 x^2 \frac{d^2}{dx^2} + \mu x \left(1 - K^{-1}x\right) \frac{d}{dx}
$$

and we find that

$$
G(x) := ((A - \rho)H)(x) = \sqrt{x} \left[\mu - 2\rho - \sigma^2/4 - \mu K^{-1} x \right].
$$
 (3.38)

Thus, if $\mu < 2\rho + \sigma^2/4$ then by the same argument as in Example 3.2 we see that the optimal policy is *immediate chattering down to 0*. We then have $T = 0$, and the value reads as

$$
V(x) = 2\sqrt{x}.\tag{3.39}
$$

However, if $\mu > 2\rho + \sigma^2/4$, then we see that the mapping $G(x)$ satisfies the conditions of Theorem 2 in [\[2](#page-14-5)] and, therefore we find that there is a unique threshold *x*∗ satisfying the condition

$$
x^* \psi''(x^*) + \frac{1}{2} \psi'(x^*) = 0,
$$
\n(3.40)

where $\psi(x)$ denotes the increasing fundamental solution of the ordinary differential equation $((A - \rho)u)(x) = 0$, that is, $\psi(x) = x^{\theta}M(\theta, 2\theta + \frac{2\mu}{\sigma^2}, \frac{2\mu K^{-1}}{\sigma^2}x)$, where $\theta = \frac{1}{2} - \frac{\mu}{\sigma^2} + \frac{2\mu}{\sigma^2}x$ $\sqrt{(\frac{1}{2} - \frac{\mu}{\sigma^2})^2 + \frac{2r}{\sigma^2}}$, and *M* denotes the confluent hypergeometric function. In this case, the value reads as

$$
V(x) = \begin{cases} 2\left(\sqrt{x} - \sqrt{x^*}\right) + \sqrt{x^*} \left(\mu(1 - K^{-1}x^*) - \sigma^2/4\right) / r, & x \ge x^*\\ \frac{\psi(x)}{\sqrt{x^*}\psi'(x^*)}, & x < x^*. \end{cases} \tag{3.41}
$$

Especially, the value is a solution of the variational inequality

$$
\min\left\{((\rho - A)V)(x), V'(x) - x^{-1/2}\right\} = 0.
$$

We summarize this as follows:

Theorem 3.4 (a) *Assume that*

$$
\mu \le 2\rho + \sigma^2/4. \tag{3.42}
$$

Then the value function $V(x)$ *of problem* [\(3.29\)](#page-10-1) *is*

$$
V(x) = 2\sqrt{x}.\tag{3.43}
$$

This value is obtained by immediate chattering down to 0.

(b) *Assume that*

$$
\mu > 2\rho + \sigma^2/4. \tag{3.44}
$$

Then V(*x*) *is given by* [\(3.35\)](#page-11-2)*. The corresponding optimal policy is immediate chattering from x* down to x^* *if* $x > x^*$ *, and local time at* x^* *of the downward reflected process* $\bar{X}(t)$ *at* x^* *if* $x < x^*$ *, where* x^* *is given by* [\(3.34\)](#page-11-1)*.*

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4 Discussion on a special case

Our verification Theorem[2.1](#page-4-0) covers a large class of state dependent singular stochastic control problems arising in the literature on the rational management of renewable resources. It is worth emphasizing that there is an interesting subclass (including the case of Example [3.1\)](#page-7-4) of problems where we can utilize our results in order to provide both a lower as well as an upper boundary for the maximal attainable expected cumulative harvesting yield. In order to shortly describe this case, assume that the underlying dynamics are time homogeneous and independent of each other and, accordingly, that the drift coefficient satisfies $b_i(t, x) = b(x_i)$ and that the volatility coefficient, in turn, satisfies $\sigma_i(t, x) = \sigma_i(x_i)$. Assume also that the price $\pi_i(t, x) = \pi_i(x_i)$ per unit of harvested stock $x_i \in \mathbb{R}_+$ is nonnegative, nonincreasing, and continuously differentiable as a function of the prevailing stock. Given these assumptions, define the nondecreasing and concave function

$$
\Pi_i(x_i) = \int_0^{x_i} \pi_i(v) dv \ge \pi_i(x_i) x_i.
$$

It is now a straightforward example in basic analysis to show by relying on a chattering policy described in our Example [3.1.](#page-7-4) that in the present case we have

$$
J^{(\tilde{\gamma}(m,0))}(0,x) = \sum_{i=1}^{n} \Pi_{i}(x_{i}).
$$

Consequently, under the assumed time homogeneity we observe that the maximal attainable expected cumulative harvesting yield satisfies the inequality

$$
\sup_{\gamma} J^{(\gamma)}(0, x) \ge \sum_{i=1}^{n} \Pi_i(x_i). \tag{4.1}
$$

On the other hand, utilizing the generalized Itô-Döblin-formula to the mapping Π_i , invoking the nonnegativity of the value Π_i , and reordering terms yields

$$
\Pi_{i}(x_{i}) \geq -E_{x} \int_{0}^{T_{N}^{*}} e^{-\rho s} \left(\mathcal{G}_{\rho}^{i} \Pi_{i} \right) (X_{i}(s)) ds + E_{x} \int_{0}^{T_{N}^{*}} e^{-\rho s} \pi_{i} (X_{i}(s)) d\gamma_{i}(s) \n- E_{x} \sum_{0 \leq s \leq T_{N}^{*}} e^{-\rho s} \left[\Pi_{i} (X_{i}(s)) - \Pi_{i} (X_{i}(s-)) - \pi'_{i} (X_{i}(s-)) \Delta X_{i}(s) \right],
$$

where T_N^* is an increasing sequence of almost surely finite stopping times converging to T and

$$
\left(\mathcal{G}_{\rho}^{i}\Pi_{i}\right)(x) = \frac{1}{2}\sigma_{i}^{2}(x)\pi_{i}'(x) + b_{i}(x)\pi_{i}(x) - \rho\Pi_{i}(x).
$$

The concavity of the mapping Π_i then implies that

$$
\Pi_i(X_i(s)) \le \Pi_i(X_i(s-)) + \pi_i(X(s-))(X_i(s) - X_i(s-))
$$

= $\Pi_i(X_i(s-)) - \pi_i(X_i(s-)) \Delta X_i(s).$

Hence, we find that for any admissible harvesting strategy γ_i we have

$$
E_x \int_0^{T_N^*} e^{-\rho s} \pi_i(X_i(s)) d\gamma_i(s) \leq \Pi_i(x_i) + E_x \int_0^{T_N^*} e^{-\rho s} (\mathcal{G}_{\rho}^i \Pi_i)(X_i(s)) ds.
$$

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Summing up the individual values then finally yields

$$
\sum_{i=1}^n E_x \int_0^{T_N^*} e^{-\rho s} \pi_i(X_i(s)) d\gamma_i(s) \leq \sum_{i=1}^n \Pi_i(x_i) + E_x \int_0^{T_N^*} e^{-\rho s} \sum_{i=1}^n (\mathcal{G}_{\rho}^i \Pi_i)(X_i(s)) ds.
$$

Letting $N \uparrow \infty$ and invoking monotone convergence then shows that in the present setting

$$
\sup_{\gamma} J^{(\gamma)}(0, x) \le \sum_{i=1}^{n} \Pi_i(x_i) + \sup_{\gamma} E_x \int_0^T e^{-\rho s} \sum_{i=1}^{n} (\mathcal{G}_{\rho}^i \Pi_i)(X_i(s)) ds.
$$
 (4.2)

Consequently, in the time homogeneous and independent setting the value which can be attained by a chattering policy can be utilized for the derivation of both a lower as well as an upper boundary for the value of the optimal harvesting policy. Moreover, in case the generators $(G_{\rho}^{i} \Pi_{i})(X_{i}(s))$ are bounded above by M_{i} we observe that

$$
\sup_{\gamma} J^{(\gamma)}(0, x) \le \sum_{i=1}^{n} \Pi_i(x_i) + \sum_{i=1}^{n} \frac{M_i}{\rho} \left(1 - E_x \left[e^{-\rho T} \right] \right). \tag{4.3}
$$

For example, if the underlying evolves as in our 2-dimensional BM example [3.1,](#page-7-4) we observe that

$$
\left(\mathcal{G}_{\rho}^{i}\Pi_{i}\right)(x) = x^{-3/2}\theta_{i}\left(\mu_{i}x - \frac{\sigma_{i}^{2}}{4} - 2\rho x^{2}\right).
$$

Hence, $(G_{\rho}^i \Pi_i)(x) \leq (G_{\rho}^i \Pi_i)(\tilde{x}_i)$, where

$$
\tilde{x}_i = -\frac{\mu_i}{4\rho} + \frac{1}{4\rho} \sqrt{\mu_i^2 + 6\sigma_i^2 \rho}.
$$

Consequently, we have that

$$
\sup_{\gamma} J^{(\gamma)}(s, x) \leq 2e^{-\rho s} \left[\theta_1 \sqrt{x_1} + \theta_2 \sqrt{x_2} \right] + e^{-\rho s} \left(\left(\mathcal{G}_{\rho}^1 \Pi_1 \right) (\tilde{x}_1) + \left(\mathcal{G}_{\rho}^2 \Pi_2 \right) (\tilde{x}_2) \right) \left(1 - E \left[e^{-\rho T} \right] \right).
$$

References

- 1. Alvarez, L.H.R.: Optimal harvesting under stochastic fluctuations and critical depensation. Math. Biosci. **152**, 63–85 (1998)
- 2. Alvarez, L.H.R.: Singular stochastic control in the presence of a state-dependent yield structure. Stoch. Process. Appl. **86**, 323–343 (2000)
- 3. Alvarez, L.H.R.: On the option interpretation of rational harvesting planning. J. Math. Biol. **40**, 383–405 (2000)
- 4. Alvarez, L.H.R., Shepp, L.A.: Optimal harvesting of stochastically fluctuating populations. J. Math. Biol. **37**, 155–177 (1998)
- 5. Jeanblanc-Picqué, M., Shiryaev, A.: Optimization of the flow of dividends. Russ. Math. Surveys **50**, 257–277 (1995)
- 6. Lande, R., Engen, S., Sæther, B.-E.: Optimal harvesting, economic discounting and extinction risk in fluctuating populations. Nature **372**, 88–90 (1994)
- 7. Lande, R., Engen, S., Sæther, B.-E.: Optimal harvesting of fluctuating populations with a risk of extinction. Am. Nat **145**, 728–745 (1995)
- 8. Lungu, E.M., Øksendal, B.: Optimal harvesting from a population in a stochastic crowded environment. Math. Biosci. **145**, 47–75 (1996)
- 9. Lungu, E.M., Øksendal, B.: Optimal harvesting from interacting populations in a stochastic environment. Bernoulli **7**, 527–539 (2001)
- 10. Protter, P.: Stochastic integration and differential equations. Springer, New York (2004)