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Environmental Filtering Drives Plant Community Assembly Processes in the Riparian Marsh of Downstream Yellow River, China

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Abstract

The driver of plant community assembly in riparian marshes is vital but still unclear. In this study, we report the results of sixty plots located along the downstream Yellow River to identify the driver of plant community assemblages. Seven functional traits were measured and compiled to calculate functional beta diversity, and phylogenetic tree was constructed to assess phylogenetic dissimilarity. The analyses show that all of traits had weak phylogenetic signals using the Blomberg's *K* statistic. The phylogenetic and functional beta diversity present clustering patterns, with the exception of leaf nitrogen concentration and leaf phosphorus concentration; and the standardized effect sizes based on multi-traits, maximum height, leaf dry matter content are well correlated with the environmental gradient. The findings imply that environment filtering was responsible for the variation of beta diversity, and drove community assembly in the riparian marsh. Of all environmental factors, soil salinity is the most crucial indicator to dominate the pattern of beta diversity.

Keywords Community assembly \cdot Environmental filtering \cdot Variance partitioning \cdot Environmental gradient \cdot Riparian marsh of the Yellow River

Introduction

Community assembly studies are imperative in explaining the formation and maintenance of biodiversity, which have been continually pursued in the field of ecology (Rosindell et al. 2011). Two major theories have been proposed to elucidate the mechanism of species assemblages, the niche theory based on deterministic processes (Diamond 1975) and the neutral theory based on stochastic processes (Hubbell 2001). As the relative importance of ecological processes varies largely with different plant communities, however, the quantification and assessment of individual assembly processes are of vital importance.

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The original explication of assembly mechanisms is often based on taxonomic identity. But it has proven to be poorly responsive to the process of species assemblage though it was a reasonable starting point (Siefert et al. 2013; Jiang et al. 2018). An alternative approach, using phylogenetic and functional diversity to test which deterministic or stochastic processes drive the assembly of species, is widely applied and offers another justification of testing ecological processes in different communities (Cornwell and Ackerly 2009; Purschke et al. 2013). For instance, the lower functional beta diversity occurred in a scale of less than 600 km and the higher functional beta diversity was shown in a scale of more than 1800 km in the eastern North American forest community (Siefert et al. 2013). In the grassland regions of China, community assembly processes are remarkably driven by environmental filtering using phylogenetic relatedness among species (Chi et al. 2014). Such observations highlight that beta diversity, reflecting the dissimilarity of phylogenetic and functional components between communities, is an effective tool for ecologists to analyze and interpret the mechanisms of community formation and maintenance (Anderson et al. 2011).

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Nonetheless, the mechanisms of community assembly are still a hotly debatable topic. Under niche-based processes, environmental filtering and the species interactions are considered as the most fundamental drivers of community assembly (Cornwell and Ackerly 2009; López-Angulo et al. 2018). Environmental filtering processes are shown to sort the potential species from a regional species pool to a suitable habitat (Diamond 1975). A common view is shared that there is a potential relationship between phylogenetic or functional diversity and environmental gradients (Spasojevic and Suding 2012; Satdichanh et al. 2015; Cadotte and Tucker 2017). In particular, environmental filtering is frequently determined as the driving force of community assemblage processes along an extensive environmental gradient (Swenson et al. 2012a; Cadotte and Tucker 2017; López-Angulo et al. 2018). Conversely, the definition and inappropriate use of environmental filtering have been questioned in recent researches (Cavender-Bares et al. 2009; Maire et al. 2012; Kraft et al. 2015). Phylogeny and trait convergence are no longer merely the result of environmental filtering processes, as biotic interactions and dispersal processes may lead to the same outcomes (Mayfield and Levine 2010; Kraft et al. 2015).

It is well recognized that environmental filtering as a deterministic pattern of community assemblages, inevitably leads to a decreasing range of traits value and increasing niche overlapping (Weiher and Keddy 1995; Cornwell and Ackerly 2009). From the perspective of phylogeny and traits, environmental filtering leads to a tighter correlation of coexisting species or more similarity of specific traits, i.e., phylogenetic clustering or trait convergence. Determining a clustered pattern of community dispersion, therefore, is the first step in detecting environmental filtering processes. But such determination alone is not comprehensive, since certain species with higher intrinsic growth rate having a fitness advantage over those with lower intrinsic growth rate shall bring about a competitive exclusion which is similar with the consequence drove by environmental filtering. Moreover, the modern coexistence theory argues that the differences between population growth rate and niche overlap are vital indicators inferring species coexistence and persistence pattern (Chesson 2000; Cadotte and Tucker 2017). It is understood that species may have different growth rates in the optimal and suboptimal habitats as there is variation of fitness along environmental gradient. So it is observable a pattern that the population growth rate is continuously varying along the environmental gradient (Cadotte and Tucker 2017). It is thus crucial to ensure that the observed relationships between the growth rates and phylogenetic and functional similarity are revealed by the environment and is not the result of other processes, such as plastic traits. In consequence, confirming the existence of such nonindependence in traits requires to quantify the phylogenetic signals to acquire a powerful piece of evidence in the trait data to combine phylogeny, traits, and environment.

In this work, we conduct studies concerning the high environmental pressures in the riparian marsh along the downstream of Yellow River (high salinity and low nutrients in soil). We postulate that environmental filtering is the driver of community assemblages in the study area, and adopt the guidelines proposed by Cadotte and Tucker (2017) to confirm the leading role of environmental filtering with community phylogenetic and functional beta diversity. Based on the hypothesis, a variation partitioning approach is applied to distinguish edaphic, climatic, and spatial variables to explicate the variation of beta diversity. From previous studies, the variation partitioning approach of trait-based response variables proves to advance our understanding of deterministic ecological processes (Legendre et al. 2009; Siefert et al. 2013; Jiang et al. 2018). Furthermore, we endeavor to find out key environmental factors that drive the variation of phylogenetic and functional beta diversity separately in the riparian marsh of Yellow River. We focus on three issues: (1) Are there any significant phylogenetic signals in functional trait data? (2) Does environmental filtering drive the community assembly processes along the environmental gradient? (3) What factor is the most crucial in phylogenetic and functional beta diversity in the riparian marsh of Yellow River?

Materials and Methods

Study Site

The study area is located in the central and eastern China $(34^{\circ}56'-37^{\circ}45'N; 113^{\circ}29'-119^{\circ}09'E)$. The mean annual temperature is 16.3 °C throughout the study area, with January and July being the coldest (-1.6 °C) and warmest (28.2 °C) months, respectively. The mean annual precipitation is about 650 mm, with a peak during the summer and a dip between December and February. A large amount of sediment is deposited and formed sandbars resulting from the decrease of river's velocity, providing habitats for wetland species within the channel downstream Yellow River.

Field Sampling

60 plots (10 m × 10 m) were established along 786 km of the Yellow River channel (Fig. 1), from the Peach Valley to the estuary, from August to September 2016. Within each plot, we set up five 1 m² quadrats to investigate plant communities using random sampling methods. We recorded the species name and coverage of the species occurring in each quadrat. The plant communities were mainly composed of herbs and immature shrubs. We sorted out the species covered at least 80% of the plant community per plot and included a sufficient range of traits (Pakeman and Quested 2007; Raevel et al. 2012). Finally, 48 species were selected for measurement of



Fig. 1 Distribution of the sampling sites in the riparian marsh of downstream Yellow River, China (2016)

functional traits representing a regional species pool. Seven functional traits, quantitatively describing certain independent axes of a species niche, were measured and compiled: maximum height (H_{max} , cm), leaf thickness (LT, mm), leaf dry matter content (LDMC, g/g), specific leaf area (SLA, cm²/g), seed mass (SM, g), leaf nitrogen concentration (LNC, %), and leaf phosphorus concentration (LPC, %) (Table 1). We measured plant traits on 10–20 individuals of each species per plot following a standard protocol (Cornelissen et al. 2003) with the exception of LT for which another method (Seelig et al. 2012) was followed (See supplementary material S1). Subsequently, we created a species \times traits matrix with mean trait value and generated trait dendrograms by the R software performing a hierarchical clustering to implement the same structure as phylogeny (Swenson et al. 2012b; Yang et al. 2014) with multi-traits and univariate trait, respectively. The trait data were log1p-transformed prior to the construction of the trait dendrograms.

Table 1	Measured functional
traits and	d their importance and/or
function	ing for ecological
strategy	

Plant functional traits	Importance(s) and/or function(s)
H _{max}	Tolerance to environmental stress and completive vigor
LT	Water retention
LDMC	Potential relative growth rate and light-saturated photosynthetic rate
SLA	Potential relative growth rate and light-saturated photosynthetic rate
SM	Reproduction niche and dispersal approach
LNC	Nutrient acquisition and maximum photosynthetic rate
LPC	Nutrient acquisition and maximum photosynthetic rate

Abbreviations: H_{max}, maximum height; LT, leaf thickness; LDMC, leaf dry matter content; SLA, specific leaf area; SM, seed mass; LNC, leaf nitrogen concentration; LPC, leaf phosphorus concentration

Environmental Data

To characterize the spatial distribution of the plots in the study area, we obtained the geo-location information in the form of latitude and longitude by GPS. It was not found that there is significant spatial autocorrelation in species distributional data analyzed by the SAM software (Dormann et al. 2007; Qin et al. 2017). One soil sample taken at a depth of 0-20 cm was collected in each quadrat using a 5 cm diameter soil auger, stored and sealed in a plastic bag, and taken back to the laboratory to determine its physical and chemical properties. Five soil samples in total from a plot were composited to measure nine different soil properties: soil moisture content (SMC, g/g), pH, soil salinity (SS, mg/L), total nitrogen (TN, %), total phosphorus (TP, %), total carbon (TC, %), available nitrogen (AN, %), extractable phosphorus (EP, %), and effective cation exchange capacity (CEC, cmol/kg) with three replicates. The parameters were obtained in the following way, SMC as the oven-dry mass of a soil sample divided by its fresh mass, pH measured using a pH meter (FE28, Mettler-Toledo Company, Switzerland), and SS calculated by the conductivity in soil measured by a conductivity meter (FE38, Mettler-Toledo Company, Switzerland). While the Kiadhal digestion method and the ammonium acetate method were used to measure TN and CEC in soil via automatic azotometer (9860, Hanon Company, China), respectively, TP was assessed by the Mo-Sb colorimetric method using a spectrophotometer (UV-9100, Unico Company, China) and TC by a carbon sulfur analyzer (CS2800, NCS Testing Technology Company, China). AN and EP were determined, separately, by a diffuser and 'Avio 200 ICP-OES' (PerkinElmer Avio Company, USA). In addition, the climate data of study area with 1 km spatial resolution was obtained from the WorldClim2 database (Fick and Hijmans 2017). We extracted climatic variables to characterize the climate conditions for each plot from the WorldClim2 database: annual mean temperature (AMT, °C), mean temperature of coldest quarter (MTC, °C), mean temperature of warmest quarter (MTW, °C), annual precipitation (AP, mm), precipitation of warmest quarter (PW, mm), precipitation of coldest quarter (PC, mm), temperature seasonality (TS), and precipitation seasonality (PS). Besides, we calculated annual biotemperature (AB, °C), warmth index (WI, °C), and coldness index (CI, °C) using monthly mean temperature data at the same spatial resolution (Fang et al. 2011) (refer to supplementary material S2 for detailed information).

Phylogenetic Tree

In order to infer the pattern of community assemblages in the riparian marsh of the Yellow River, we explicated the causes for species coexistence through species evolution. The classification information of all 48 species was retrieved and input into the *Phylomatic* software (http://www. phylodiversity.net/phylomatic), generating the megatrees according to Zanne et al. (2014) to identify the distribution of species on an existing phylogenetic framework. Followed was the construction of a species-level phylogenetic tree based on the APG III system.

Data Analysis

Following the guidelines of Cadotte and Tucker (2017), an appropriate approach is to quantify the phylogenetic and functional structures of the community via the standardized effect size. We first calculated the beta diversity (Dnn') and standardized effect size (S.E.S. Dnn') on the basis of phylogeny, multitraits and all univariate traits. Then, we implemented an appropriate null model and randomized shuffled taxa names across the tips of the phylogenetic tree and trait dendrograms 999 times, while maintaining species occurrence frequency and abundance. This approach has been proven to have the best ability to discover filtering based on multivariate functional traits (Spasojevic and Suding 2012). Next, we used the Student's t test to check whether the phylogenetic clustering/ overdispersion or trait convergence/divergence is significant with the standardized effect size. The calculation was based on the equation below,

$$S.E.S. Dnn' = \frac{Dnn'_{(\text{obs})} - Dnn'_{(\overline{\text{null}})}}{sd_{(\text{null})}},$$

where $Dnn'_{(obs)}$ is the observed value of the weighted nearest neighbor distance between communities, $Dnn'_{(\overline{null})}$ is the mean of the weighted nearest neighbor distance by the null model, and $sd_{(null)}$ is the standard deviation of the weighted nearest neighbor distance by the null model.

As the influence of the environment on species coexistence is both relevant in extraction of a species from the species pool to the habitat and in the competition process by covariation with the population growth rate and environment, we focused on the variation of population growth rate along environmental gradients, which is associated to community structure (Cadotte and Tucker 2017). Simultaneously, the principal component analysis was utilized to reduce redundant variables of edaphic and climatic variables. We defined the scores of the first principal component (37.4% of variance) as the environmental gradient of the study area (supplementary material S3), and observed the correlation of *S.E.S. Dnn'* and the environmental gradient by linear regressions. Since it is necessary to explore whether the evolution of traits conform to the principle of phylogenetic conservatism, we searched the phylogenetic signals in every univariate trait using the *K* statistic (Blomberg et al. 2003). The three cases of the *K* value is distinguished, (1) K = 0 suggests that there are no phylogenetic signals in trait evolution; (2) K < 1 indicates that weaker similarities exist in closely related species than expected based on the Brownian motion model; and (3) K > 1 indicates more phylogenetic signals than expected (Blomberg et al. 2003). All of the calculations and tests were carried out with R package *vegan*, *phytools*, and *picante* (R Core Team 2017). Should the values of *S.E.S. Dnn'* be negative and correlated significantly with the environmental gradient simultaneously, significant phylogenetic signals appeared in the univariate trait data. If it is the case, we may conclude that environmental filtering drives the community assembly processes.

In order to explore the relationship of beta diversity with environmental and spatial variables in depth, we used the method of variation partitioning over edaphic, climatic, and spatial variables as explanatory variables to dissect the beta diversity as the response variables (Siefert et al. 2013; Spasojevic et al. 2016; Jiang et al. 2018). For the spatial variables, we applied the principal coordinates of neighbor matrices eigenvectors, which are recurrently implemented as a tool for quantifying stochastic processes (Borcard and Legendre 2002) in terms of centralized geographic information data. In the present study, we extracted 20 positive eigenvalue variables as a proxy for total spatial variables. For the environmental variables, we used 9 edaphic variables and 11 climatic variables to represent the environmental conditions of the study area. Prior to variation partitioning, the redundancy analysis (Borcard et al. 1992) and the forward selection (Blanchet et al. 2008) were executed to reduce redundant components and retain variables which were significantly associated with beta diversity in edaphic, climatic, and spatial variables, respectively. All the calculations and analyses were accomplished by the software R (R Core Team 2017).

Results

Dispersions of Phylogenetic and Functional Beta Diversity

The standardized effect sizes of phylogeny and multi-traits were significantly lower than zero (P < 0.05) (Table 2). For univariate traits, all traits passed a significant test except for LPC, which had no significant difference from zero. Of interest, we found that LNC had a particular pattern in contrast to other traits, which was significantly greater than zero.

Standardized Effect Sizes of Beta Diversity (S.E.S. Dnn') and Environmental Gradient

The functional beta diversity of multi-traits (Fig. 2b) showed a significant negative relationship with a continuous environmental gradient (P < 0.001, $R^2 = 0.209$). In comparison, the relationship between phylogeny and the environment gradient was nonlinear (Fig. 2b). For individual traits, only the maximum height (P < 0.05, $R^2 = 0.086$) (Fig. 2c) and leaf dry matter content (P < 0.001, $R^2 = 0.337$) (Fig. 2e) showed a significant correlation. Remarkably, the maximum height increased with the environmental gradient, whereas the opposite trend was apparent for the multi-traits and leaf dry matter content.

Phylogenetic Signals

The *K* values of seven functional traits were less than one (Table 3), ranging from 0.129 (LPC) to 0.295 (SM), in which a mean weaker phylogenetic signal than expected was obtained during the process of trait evolution. The *P* values of the permutation test for univariate traits were all less than 0.05, indicating that all of the selected traits had significant phylogenetic signals.

Table 2Student's t test forstandardized effect sizes of betaStandardized effect sizes of betadiversity (S.E.S. Dnn') in plantmatched sizescommunity of riparian marsh ofPhdownstream Yellow RiverM

Standardized effect sizes of beta diversity (S.E.S. Dnn')	95% confidence interval	<i>t</i> -value	Р
Phylogeny	(-0.575, -0.430)	-13.877	0.000**
Multi-traits	(-0.671, -0.466)	-11.079	0.000**
H _{max}	(-0.262, -0.001)	-2.014	0.049*
LT	(-0.350, -0.096)	-3.523	0.001**
LDMC	(-0.442, -0.064)	-2.678	0.009**
SLA	(-0.505, -0.333)	-9.733	0.000**
SM	(-0.194, -0.002)	-2.043	0.046*
LNC	(0.224, 0.542)	4.829	0.000**
LPC	(-0.044, 0.225)	1.349	0.183

For abbreviations of traits, see Table 1. Signs: * for P < 0.05; ** for P < 0.01



Fig. 2 Relationships between standardized effect size of beta diversity (SES Dnn') and environmental gradient in the riparian marsh of downstream Yellow River

Variance Partitioning of Phylogenetic and Functional Beta Diversity

We applied the forward selection and redundancy analysis preceding variance partitioning for explanatory variables (climatic, spatial and edaphic variables), which effectively

Table 3Phylogeneticsignal tests for sevenfunctional traits amongspecies in riparian marshof downstream YellowRiver

Univariate trait	Κ	Р	
H _{max}	0.206	0.002**	
LT	0.278	0.011*	
LDMC	0.144	0.013*	
SLA	0.187	0.006**	
SM	0.295	0.024*	
LNC	0.139	0.002**	
LPC	0.129	0.018*	

For abbreviations of traits, see Table 1. Signs: * for P < 0.05; ** for P < 0.01 reduces the collinearity between variables. We found that all the facets of beta diversity were associated with SS in edaphic variables (Table 4). The correlations were ubiquitous in the forward selection results, and even the most critical, shed light on most of the variance in each facet of beta diversity. The variables sets of phylogenetic beta diversity were consistent with that of multi-trait based beta diversity (Table 4). For individual traits, SS was a widespread and robust factor for all traits in the selected sets, accounting for the overall variance of the edaphic variables. In the sets of LT ($R^2 = 0.040$) and LDMC ($R^2 = 0.034$) and LPC ($R^2 = 0.028$), CEC could be responsible for a little fraction of variance (Table 4). AN (R^2 = 0.014) explained tiny fraction of LT and TN ($R^2 = 0.032$) had little contribution to LPC (Table 4). Regarding the climatic variables, we found that TS represented the climatic variables in SLA ($R^2 = 0.324$) and LNC ($R^2 = 0.352$), whereas WI is an important climatic factor in LDMC ($R^2 = 0.413$) and LPC $(R^2 = 0.534)$, CI $(R^2 = 0.354)$ is highly associated with LT. Because the spatial variables did not pass the test, all of them were utilized in variance partitioning.

 Table 4
 Results for forward

 selection of edaphic, climatic, and
 spatial variables with

 phylogenetic and functional beta
 diversity

Beta diversity	Selected variables	R^2	Cumulative adjusted R^2	F	Р
Phylogeny	Edaphic				
	SS	0.733	0.728	158.886	0.000**
	Climatic				
	TS	0.333	0.322	28.965	0.001**
Multi-traits	Edaphic				
	SS	0.702	0.696	136.395	0.000**
	Climatic				
	TS	0.371	0.360	34.222	0.000**
H _{max}	Edaphic				
	SS	0.288	0.275	23.433	0.003**
LT	Edaphic				
	SS	0.811	0.807	248.112	0.000**
	CEC	0.040	0.845	15.290	0.000**
	AN	0.014	0.857	5.809	0.022*
	Climatic				
	CI	0.354	0.343	31.768	0.005**
LDMC	Edaphic				
	SS	0.881	0.879	429.288	0.000**
	CEC	0.034	0.912	23.169	0.001**
	Climatic				
	WI	0.413	0.403	40.854	0.002**
SLA	Edaphic				
	SS	0.505	0.497	59.247	0.001**
	Climatic				
	TS	0.324	0.312	27.757	0.001**
SM	Edaphic				
	SS	0.291	0.279	23.798	0.005**
LNC	Edaphic				
	SS	0.683	0.678	125.041	0.000**
	Climatic				
	TS	0.352	0.341	34.551	0.000**
LPC	Edaphic				
	SS	0.661	0.655	113.004	0.000**
	CEC	0.028	0.678	5.159	0.031*
	TN	0.032	0.706	6.354	0.016*
	Climatic				
	WI	0.534	0.525	66.340	0.001*

Abbreviations: H_{max} , maximum height; LT, leaf thickness; LDMC, leaf dry matter content; SLA, specific leaf area; SM, seed mass; LNC, leaf nitrogen concentration; LPC, leaf phosphorus concentration; SS, soil salinity; TS, temperature seasonality; CEC, effective cation exchange capacity; AN, available nitrogen; CI, coldness index; WI, warmth index; TN, total nitrogen. Signs: * for P < 0.05; ** for P < 0.01

We employed the approach of variance partitioning using selected variable sets. The results of phylogenetic beta diversity were similar to that of multi-traits functional beta diversity (Fig. 3). The sole edaphic variable fraction could explain most of the variance in phylogenetic (0.575) and multi-traits functional (0.488) beta diversity. But it had no significant correlation with phylogenetic and multi-traits functional beta diversity. For single traits (Fig. 3), the sole edaphic variables indicated the best explanatory power for all facets of functional diversity. The climatic variables had minute contribution to the justification of the univariate functional diversity. Moreover, the sole spatial variables were not associated to each functional diversity.



Fig. 3 Results of variance partitioning in beta diversity in the riparian marsh of the Yellow River. The values shown in Venn diagrams are

adjusted R^2 , indicating the variance explained of each fraction. R^2 values less than zero were not displayed

Discussion

In the present study we observed three lines of evidences that could be in the support of a community assembly process driven by environmental filtering, and then expounded key factors affecting the phylogenetic and functional turnover in the riparian marsh of the Yellow River. First, we found that the general phylogenetic and functional clustering patterns, apart from LNC and LPC, pointed to the fact that deterministic assembly processes played a core role of structuring a community. Second, a pertinent mechanism between the growth rate and environmental gradient was set up when the deterministic assembly process was taken into account. Third, we observed the presence of significant phylogenetic signals in the trait data. In consequence, we could come to the conclusion that the environment filtering drove the community assembly processes in the study area. Of the environmental factors, we found that the soil salinity in the edaphic variables was the most crucial factor in environmental filtering, leading to form the pattern of phylogenetic and functional dissimilarity between communities.

Assembly Processes Drove by Environmental Filtering

As shown in the negative *S.E.S. Dnn'* values of phylogeny (Table 2) and multi-traits, the general phylogenetic and functional clustering were observed in the riparian marsh of the Yellow River, signifying that phylogenetic and multivariate functional turnover were driven by deterministic assembly processes. Still, the pattern of phylogenetic turnover was in accordance, though not perfectly aligned, with that of multi-traits functional turnover. The findings could be reasonably

interpreted as a result of evolutionary lability of functional traits due to the phylogenetic signals in the trait data, as similar results reported in other studies (Swenson 2011; Yang et al. 2014). Compared to whole-plant traits and leaf morphological traits, our results of leaf nutrient concentrations showed inconsistent patterns, while LNC and LPC had divergence and random patterns, respectively. The underlying reason could be the relative high intraspecific trait variation in leaf nutrient traits responding to available resources in habitat (Siefert et al. 2015).

Along with the guidelines, we identified the relationships between phylogenetic and functional S.E.S. Dnn' and environmental gradient to provide additional evidence of environmental filtering (Fig. 2). The result of the S.E.S. Dnn' based on multi-traits was correlated with the environmental gradient (Fig. 2b), yet it is interesting to notice that phylogeny did not show a correlation with the environmental gradient (Fig. 2a). Since significant phylogenetic signals had been discovered in the trait data, a consistent pattern should be expected in the phylogenetic turnover and functional turnover (Webb et al. 2002; Swenson 2011), but the actual results were in contrary, instead. As a matter of fact, it has been argued that although phylogeny is habitually used as a tool to deduce the key ecological processes, it is not as perfect as traits (Cadotte et al. 2017; Xu et al. 2017) and even biased to apply the phylogenetic pattern to infer the community assemblages (Gerhold et al. 2015). Moreover, sympatric species might reflect the accordance of phylogenetic similarity and functional similarity, which not only contains conserved traits but also plastic traits (Pavoine and Bonsall 2011; Xu et al. 2017). In previous works on different ecosystems, some cases in reality showed inconsistent ecological processes for phylogenetic and functional inference. In the Panama and Puerto Rico tropical rain forests at different succession stages, phylogeny displayed a stochastic pattern and traits a deterministic one (Swenson et al. 2012b). For the grassland in Algeria, similar life form to our study area, researchers described phylogeny as a poor surrogate for traits (Pavoine et al. 2013). Manifestly, the relationship between phylogeny and the assemblage mechanism is complex. Contrariwise, we recognized that the environmental filtering process was essentially the screening of species with appropriate niche from the species pool into habitats (Kraft et al. 2008; Cornwell and Ackerly 2009), and plant functional traits were the appropriate proxies for species niches (McGill et al. 2006; Violle et al. 2007). Accordingly, we were convinced that the functional turnover revealed the essence of the assembly mechanism more than the phylogenetic turnover in the riparian marsh of the Yellow River. For the univariate traits, only H_{max} (Fig. 2c) and LDMC (Fig. 2e) had significant correlations with the environmental gradient. The results proved that these two traits were the key ones in determining the community dispersion and population growth rates along the environmental gradient. H_{max} is associated with the capability of tolerating environmental stress, and LDMC is related to the potential relative growth rate (Cornelissen et al. 2003). But SLA (Fig. 2f), though showing strongly correlation with potential relative growth rate and proven in many other ecosystems (Westoby et al. 2002; Iida et al. 2014; Gibert et al. 2016), did not give the same result as LDMC. We attribute the discordance to that the difficulty in measuring the leaf area of certain species reduced the accuracy of SLA (Cornelissen et al. 2003). For example, *Suaeda glauca* and *Suaeda salsa* in *Suaeda* genus, *Typha angustata* and *Typha laxmannii* in *Typha* genus are all have high leaf thicknesses depending on the fleshy leaves, which may affect the measurement of SLA (Wilson et al. 1999). In such a case, LDMC may yield a more meaning result.

All of the phylogenetic and functional analyses were based on the evolutionary conservatism assumption as proposed. Actually, we did find a significant phylogenetic signal in the univariate trait data (Table 3). In contrast, the phylogenetic signals were weak (K < 1), which is usually interpreted as evolutionary lability (Blomberg et al. 2003; Yang et al. 2014), and might lead to an ambiguous relationship between phylogeny and functional traits (Swenson 2011).

So far we have exposed, following the guidelines for multivariate functional beta diversity and part univariate functional beta diversity, all three aspects of evidences, (1) a clustered pattern of community dispersion, (2) a covariant relationship between the population growth rate and the environmental gradient, and (3) significant phylogenetic signals in trait data. Thus, the environmental filtering was determined to be dominating in the community assembly processes in the riparian marsh of the Yellow River.

Variance Partitioning of Phylogenetic and Functional Beta Diversity

In light of the result of variance partitioning, we found that the environment, especially edaphic variables, affected phylogenetic and functional beta diversity in the riparian marsh of the Yellow River. The sole edaphic variables could always be accountable for the most variances (Fig. 3). Such observation was in consistence with reports that edaphic variables were generally crucial at a local scale (John et al. 2007; Chi et al. 2014; Yang et al. 2014; López-Angulo et al. 2018), in the support of our inference. From the results of forward selection (Table 4), we revealed that SS was the most critical indicator in edaphic variables. As a result, soil salinization steered a most direct and intense environmental pressure for coexisting species. Besides, we found that climatic variables had a weaker explanatory power on phylogenetic and functional beta diversity (Fig. 3). Although the species were generally influenced by edaphic conditions at the fine scale and by climate at the broad spatial scale (Díaz et al. 1998), a harsh edaphic environment, which had strong environmental pressure of high soil salinity and low nutrient content, might perform a more rigorous filtering of species than climate for wetland species. With respect to the spatial variables, the sole spatial variable in variance partitioning analyses might represent the dispersal limitation or the spatial structure of unmeasured environment variables (Legendre et al. 2009; Siefert et al. 2013; Spasojevic et al. 2016; Jiang et al. 2018). Different from previous findings, our study showed that the spatial variables could not explain the phylogenetic and functional turnover (Fig. 3), which might result in the neglect of dispersal limitation in the neutral theory. We still could not rule out the effect of dispersal, however, because most of species are spread seeds through the wind depending on small and light seeds.

As given in Fig. 3a, b, we observed a similar pattern to multivariate functional and phylogenetic beta diversity in variance partitioning, that is, approximately half of the variance was explainable by edaphic variables and around 30% of variance by edaphic, climatic, and spatial variables together. The results revealed that the environment, particularly edaphic conditions, controlled the phylogenetic and multivariate functional turnover between the communities. In spite of the report that herbaceous species showed to be more responsive to edaphic variables than climatic variables (Murphy et al. 2016), we realized that about 30% of the variance had no reasonable explanation still, which might be prompted by unmeasured environmental factors and spatial structure or ecological drift (Myers et al. 2013; Chi et al. 2014). Also, we ascertained the relative importance of edaphic variables in univariate functional beta diversity. Clearly, these functional traits disclosed distinct dimensions representing different ecological strategies for sympatric species (Spasojevic and Suding 2012; Gross et al. 2013; Wellstein et al. 2014). Henceforth, we deliberated that edaphic conditions played a decisive role in diverse ecological strategies of plants in the riparian marsh of the Yellow River, comprising resource acquisition, environmental tolerance, competitive exclusion and reproduction. It is amazing to find that, the whole univariate functional beta diversity was well explicated by SS in the edaphic variables, or the soil salinity was a pivotal environmental factor in influencing diverse ecological strategies.

In conclusion, our results and analyses suggest that environmental filtering plays an indispensable role in the community assembly processes in the riparian marsh of the Yellow River. The phylogenetic clustering and trait convergence alone are insufficient to infer the driver of ecological processes. As a result, we have tested the relationships between *S.E.S. Dnn'* and the environmental gradient, and phylogenetic signals in the trait data to determine the dominance of environmental filtering in the assembly processes. We have observed the soil salinity in edaphic variables as a core filter influences the community composition and species coexistence. Concurrently, some inconsistent patterns or non-significant relationships between community dispersions and environmental gradient, likely to be affected by other ecological processes, have been identified, which request further work to gauge their effects.

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