



Parenting a Child with Mental Health Problems: the Role of Self-Compassion

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Abstract

Objectives Parenting children with mental health problems poses multiple challenges, including coping with difficult behavior and negative child emotions. The impact on parents includes financial strain, negative social stigma, and negative feelings of guilt or blame, resulting in significant stress and lower levels of well-being. Given findings that self-compassion plays a significant role in reducing stress and improving well-being, the current study examined the role of self-compassion in the experience of parents raising a child with mental health problems. The study tested (1) whether child behavioral/emotional problem severity is associated with higher parental stress and lower parental well-being; (2) whether self-compassion is associated with lower parental stress and higher parental well-being; and (3) whether self-compassion is a stronger predictor of parental stress and well-being than child behavioral/emotional problem severity.

Methods Three hundred and six mothers and two hundred and fifty-six fathers of children attending a hospital child and adolescent psychiatric center were assessed at admission. Consenting parents completed four questionnaires: child strength and difficulty, parent version; self-compassion; parent feeling inventory; and well-being.

Results Child behavioral/emotional problem severity was associated with higher parental stress and lower parental well-being, and self-compassion was a stronger predictor of parental stress and well-being levels than child behavioral/emotional problem severity. For children with internalizing but not externalizing behavioral/emotional problems, parental self-compassion was the only predictor of parental well-being beyond the severity of child behavioral/emotional problems.

Conclusions Cultivating self-compassion is important in reducing parental stress and increasing parental well-being, particularly with internalizing presentations, and should be considered when designing therapeutic interventions for parents.

Keywords Parenting children with mental health problems · Self-compassion · Child behavioral/emotional problem severity · Parental stress · Feeling and well-being

The rate of children and adolescents diagnosed with mental health problems has been increasing substantially, signifying an issue of serious proportions (Collishaw, 2015). Global epidemiological data from the past decade have consistently reported that up to 20% of children and adolescents suffer from a disabling mental illness (Belfer, 2008; Polanczyk et al., 2015; Westerlund et al., 2020) that affects nearly all

aspects of parental life (Angold et al., 1998; Vohra et al., 2014). Parents must cope with financial strain and time constraints caused by their child's condition (Lecavalier et al., 2006). They face social stigma related to mental disorders (e.g., Hinshaw & Stier, 2008; Moses, 2010), often exacerbated by feelings of blame for their child's mental illness due to perceived deficits in their parenting (Crowe et al., 2011). Additionally, parents may internalize negative stereotyping and discrimination, which, in turn, may be reflected in feelings of guilt or blame related to their child's illness (e.g., Hinshaw, 2007; Mak & Cheung, 2012). These feelings can lead to social isolation and rejection from society (Eaton et al., 2016; Hasson-Ohayona et al., 2014).

A significant source of parents' stress resides in coping with their children's difficult behavior and aroused emotions. These negative aspects often cloud positive or

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natural characteristics and might cause parents to become more impulsive and irritable toward their children (Bögels et al., 2010). Mental health problems in adolescents challenge parenting practices and may increase a variety of negative feelings, including frustration, helplessness, and blame for being an inadequate parent or for being unable to meet their child's emotional needs (Crowe et al., 2011; Lindgren et al., 2016). Parenting a child with emotional and behavioral problems is accompanied by long-term stress and can compromise quality of life (e.g., Gerkensmeyer et al., 2011; Vaughan et al., 2013). Parents of children with mental health problems experience significantly lower levels of life satisfaction, an indicator of well-being (Athay, 2012), and marginally poorer well-being than parents raising children without disabilities (Oruche et al., 2012; Smith & Grzywacz, 2014).

One commonality between well-being and caregiver stress is their relationship to the severity of the child's symptoms. A child's symptom severity is a significant predictor of parenting stress (e.g., Barroso et al., 2018; McDonald et al., 1997) and parental well-being (e.g., Athay, 2012; Early et al., 2002). This finding is evident for various special needs populations (e.g., disabilities, autism; Ekas & Whitman, 2010; Neff & Faso, 2015). Since parental stress and parental well-being are influenced by children's mental health problems and influence parents, parent–child relationships, and child adjustment, it is important to identify potential factors that help parents reduce stress and increase well-being. However, to the best of our knowledge, few studies have focused specifically on factors that may help parents of children with mental health problems reduce stress and enhance well-being (Rodríguez-Meirinhos et al., 2018).

Self-compassion plays an important role in individual stress and subjective well-being development (Neff, 2003). Self-compassion is an adaptive way of self-to-self relating (Gilbert & Procter, 2006; Neff, 2009, 2012). Neff (2003, 2009) described three dimensions of self-compassion: (1) self-kindness versus self-judgment, meaning being kind and understanding toward oneself in instances of hardship or perceived inadequacy rather than being harshly self-critical; (2) a sense of common humanity versus isolation, meaning perceiving one's experiences as part of the larger human experience rather than seeing them as isolating; and (3) mindfulness versus overidentification, meaning holding painful thoughts and feelings in mindful awareness rather than overidentifying with them. The self-compassion dimensions facilitate an individual's openhearted stance toward negative aspects of oneself and one's experience, fostering a kind, forgiving, and empathetic attitude toward others (Neff, 2012). Therefore, self-compassion has been associated with a variety of positive individual outcomes, such as higher levels of optimism, happiness, and positive affect (Neff, 2003; Neff et al., 2007), as well as greater acceptance and caring

in relationships and better relational well-being (Gouveia et al., 2016; Yarnell & Neff, 2013).

In the context of the parent–child relationship, self-compassion has the same positive relation to the feelings, behavior, and well-being of the parent (Moreira et al., 2015a, 2015b; Sirois et al., 2019). For example, Sirois et al. (2019) found that self-compassion improves parental well-being when dealing with the negative feelings of shame and blame. Gouveia et al. (2016) reported that higher levels of parental mindfulness and self-compassion are associated with lower levels of parenting stress and higher levels of an effective parenting style.

When facing an ongoing life stressor such as raising a child with mental health problems, parental self-compassion can be essential for developing an “immune system” to help parents manage their own as well as their child's negative feelings (Gilbert, 2014). Bögels et al. (2010) proposed a mindful parenting program to foster changes in interactions between parents and children with mental health problems. The program emphasizes mindfulness and self-kindness, two of three components of self-compassion. Mindfulness and self-kindness help parents manage their stress and adopt a more accepting and nonjudgmental stance toward themselves as parents and toward their children, especially when their children express negative emotions. Mindfulness and self-kindness also increase parents' positive attention to themselves and their children (Bögels et al., 2010). This overall positive effect of self-compassion might be reflected in parental well-being, although to the best of our knowledge, this specific variable has not been tested in the context of parents of children with mental health problems. For parents raising children with autism and learning disabilities, parental self-compassion has been found to contribute significantly to parental well-being. Neff and Faso (2015) found that among parents of children with autism aged 4 to 12, self-compassion predicted parental well-being over and above the effects of child symptom severity. Sumiati (2018) reported that among parents of children with learning disabilities, self-compassion was strongly and positively associated with well-being and moderated the association between emotion regulation strategy (problem or emotion focus) and well-being.

Self-compassion has positive potential for managing individual stress and subjective well-being (Neff, 2003); however, the bulk of previous research does not measure self-compassion specifically in relation to parenting a child with mental health problems. Our research focused on this particular junction between parental self-compassion and parenting children with mental health problems and examined whether parents' self-compassion is associated with their levels of stress and well-being. To capture the subjective experience of parental stress, we used data on negative and positive feelings of parenting given that negative feelings

toward the self and toward the child have been demonstrated to indicate parenting stress (Deater-Deckard, 1998). In light of prior findings that the severity of symptoms in children with mental health problems negatively impacts parental stress and well-being (Athay, 2012), our second objective was to determine whether self-compassion is a more accurate predictor of parental stress and well-being than symptom severity. We hypothesized that (1) child behavioral/emotional problem severity is associated with greater negative emotion and with overall lower levels of parental well-being; (2) self-compassion is associated with less negative and more positive parental emotion and higher levels of parental well-being; and (3) self-compassion is a stronger predictor of parental emotion and well-being than the child's behavioral/emotional problem severity.

Method

Participants

The sample included 446 families of children attending the child and adolescent psychiatric center of a public hospital that serves a regional multiethnic population. The center provides assessment and treatment in individual outpatient clinics for a variety of mental health problems. Families attending the psychiatric center are assessed by a clinician and referred to the relevant clinic for a deeper assessment and treatment plan. The families of the current study were referred to anxiety and post-trauma, ADHD, eating disorders, depression, and sexual abuse clinics where they were given the questionnaires. In total, 406 mothers and 256 fathers completed the questionnaires. Of the mothers, 314 (77%) were married, 62 (15%) were divorced, 11 (3%) were widows, and 19 (5%) were single mothers. Of the fathers, 226 (88%) were married, 27 (11%) were divorced, and 3 (1%) were single fathers. For 216 (48%) children, both the mother and the father answered the questionnaires; for 190 (43%) children, only the mother responded; and for 40 (9%) children, only the father responded. The children, who were in elementary and secondary school grades, ranged in age from 10 to 18 years ($M = 12.6$, $SD = 3.2$). The mothers ranged in age from 27 to 65 ($M = 43.0$, $SD = 6.5$), and the fathers ranged in age from 29 to 68 ($M = 46.5$, $SD = 6.7$).

Procedure

The questionnaires are part of a broader set of questionnaires completed by parents and children attending the center and are used by center professionals in the evaluation process, not only for research purposes. The Helsinki Committee of Ziv Medical Center, Tzfat, Israel, approved the study.

Measures

Parental Feelings Inventory

The parental feelings inventory (PFI; Bradley et al., 2013) assesses parental emotions in the parenting role. Parents are presented with emotion adjectives and are asked to indicate the degree to which they experienced that emotion during the last week in their role as parents. Thirty-one emotion adjectives are assessed using a 7-point Likert scale (1 = not at all to 7 = extremely). The emotion adjectives are separated into three groups (happy, angry, anxious/sad). For the purpose of this study, parents' feelings were separated into positive and negative groups. Correlation coefficients between PFI anger and PFI sadness were high ($r = 0.65$). The scale demonstrated good reliability (internal consistency in the current sample was $\alpha = 0.90$ for the PFI-positive scale and $\alpha = 0.93$ for the PFI-negative scale) and correlated with other measures of parent and child functioning, supporting the overall utility of the PFI as a measure of emotional experiences in the parenting role.

Personal Well-being Index (PWI-A; International Wellbeing Group, 2006)

The PWI-A for assessing the subjective sense of well-being is composed of one question inquiring about general life satisfaction and eight items measuring satisfaction in specific life domains: standard of living, personal health, achieving in life, personal relationships, personal safety, community connectedness, future security, and religion. All items are rated on a scale ranging from 0 = completely dissatisfied to 10 = completely satisfied. The scale has a Cronbach's alpha ranging from 0.70 to 0.85 (International Wellbeing Group., 2006), and internal consistency in the current sample was $\alpha = 0.87$.

Strengths and Difficulties Questionnaire Parent Version (SDQ; Goodman, 1997)

The SDQ is a widely used screening tool for assessing the symptom severity of behavioral and emotional disorders in children and adolescents (Niclasen et al., 2012), with good convergent validity with clinician-rated diagnoses (He et al., 2013). The 25-item measure includes five subscales that assess internalizing (emotional problems subscale) and externalizing (hyperactivity and conduct subscales) symptoms, a subscale measuring difficulties with peers, and a subscale assessing prosocial behavior. High scores on the four difficulty subscales represent a high degree of difficulty; a high score on the prosocial scale represents a high degree of prosocial behavior. The four difficulty subscales have been shown to be indicative of youth mental health service use

(Koskelainen et al., 2000). Internal consistency in the current sample was $\alpha=0.74$ for prosocial, $\alpha=0.78$ for externalizing, and $\alpha=0.71$ for internalizing.

Self-Compassion Scale (SCS; Neff, 2003)

Self-compassion was measured using a 26-item self-report measure with responses ranging from 1 = almost never to 5 = almost always. The SCS contains six subscales: self-kindness (e.g., I try to be loving toward myself when I'm feeling emotional pain), self-judgment (e.g., I'm disapproving and judgmental about my own flaws and inadequacies), common humanity (e.g., When things are going badly for me, I see the difficulties as part of life that everyone goes through), isolation (e.g., When I think about my inadequacies, it tends to make me feel more separate and cut off from the rest of the world), mindfulness (e.g., When I'm feeling down, I try to approach my feelings with curiosity and openness), and overidentification (e.g., When I'm feeling down, I tend to obsess and fixate on everything that's wrong). The SCS subscales may be examined separately, or a total self-compassion score can be used given that a single higher-order factor of "self-compassion" has been found to explain the intercorrelations between subscales (Neff, 2003). Note that the self-judgment, isolation, and overidentification subscales of the SCS are reverse coded so that higher scores indicate higher levels of self-compassion. The scale demonstrates convergent validity (e.g., correlates with partner ratings), discriminant validity (e.g., no correlation with social desirability), and good test–retest reliability (Neff, 2003; Neff & Beretvas, 2013). Internal consistency in the current sample for the SCS total score was $\alpha=0.90$.

Data Analyses

The outcome measures (parental well-being (PWB), PFI happy, and PFI negative) were reviewed for extreme values using boxplots, scatterplots, histograms, and normal probability plots and found to be normally distributed. Of the 662 respondents, 44 individuals did not complete the PWB questionnaire. No discernable patterns were found by visually inspecting the missing data, which suggested that they were missing at random. In a random pattern, a small percentage of data points missing from a large dataset is insignificant (Tabachnick & Fidell, 2007).

The internal consistency of each scale was calculated using Cronbach's alpha. Pearson regression coefficients were calculated to assess the association between all subscales. Comparison of the results between mothers and fathers across the variables of interest was performed using a two-sample *T*-test. To determine whether SDQ subscales (prosocial, externalizing, internalizing) predicted the outcome measures, we conducted a series of mixed linear

regressions. The predictor variables were pairwise interactions of parents (father vs. mother), each of the SDQ subscales and the corresponding main effects. To determine whether self-compassion predicted the outcome measures beyond those accounted for by SDQ subscales, we followed the first-stage mixed linear regressions described above with a second-stage series of mixed linear regressions. In the second stage, an additional pairwise interaction of parent and self-compassion and the corresponding main effect were added to the model. In both steps, child age and the mean parent age were entered as control variables. The significance of the interaction effects signifies a moderation effect of the parent in the relationship between each predictor variable and the outcome variable. Generalized linear models (using PROC GLIMMIX) were applied for the two-stage mixed linear regression models since they enable the modeling of hierarchical data (in this case, a mother and father nested in the same family and reporting for the same child). The multilevel model has the capacity to handle unbalanced data when the data are missing at random (Goldstein, 2003).

Post hoc power analysis was performed using a Gpower computer program (Buchner et al., 1998) to generate a linear multiple regression random model. The input parameters for the power calculation for each of the prediction models were the predictors' correlation coefficients with the outcome measure and the total sample size of 622 participants. Power was above 95% for all three models. A *p* value of 0.05 was considered significant. Statistical analysis was performed by SAS for Windows version 9.4.

Results

All data are available as supplementary materials to this article. Consistent with our first hypothesis, the SDQ total score and its two subscales, internalizing and externalizing severity, were positively correlated with PFI negative, $r=0.38$, $r=0.29$, and $r=0.31$, and negatively correlated with PFI positive, $r=-0.18$, $r=-0.11$, and $r=0.18$, and PWB, $r=-0.17$, $r=-0.12$, and $r=-0.14$. SDQ prosocial was not significantly correlated with PFI negative but was positively correlated with PFI positive, $r=0.19$, and PWB, $r=0.12$. Consistent with our second hypothesis, self-compassion was negatively correlated with PFI negative, $r=-0.34$, and positively correlated with PFI positive, $r=0.30$, and PWB, $r=0.56$ (Table 1). Table 2 presents how mothers and fathers differed across the variables of interest. Fathers had significantly higher PFI-negative values than mothers ($M=3.49$ vs. $M=3.08$, $p<0.001$). Given that the total SDQ score was found to be highly correlated with the SDQ externalizing and internalizing subscales and to gain more specific understanding as suggested by previous studies (Becker

Table 1 Subscales and totals of all scale descriptive statistics and between-scale Pearson correlation coefficients, $N=662$

Variable	<i>M</i>	<i>SD</i>	1	2	3	4	5	6	7	
1. PFI negative	3.33	1.36	.93)							
2. PFI positive	4.48	1.19	-.43	(.90)						
3. SDQ prosocial	7.14	2.22	-.09*	.19	(.74)					
4. SDQ externalizing	7.62	4.08	.31	-.11	-.31	(.78)				
5. SDQ internalizing	8.51	4.04	.29	-.18	-.24	.28	(.71)			
6. SDQ_T	16.13	6.48	.38	-.18	-.35	.80	.80	(.78)		
7. Self-compassion	3.58	.65	-.34	.30	.09	-.12	-.13	-.16	(.90)	
8. PWB	7.30	1.70	-.29	.44	.12	-.12	-.14	-.17	.56	(.87)

* $p < .05$, for the remaining coefficients $p < .001$

Table 2 Comparison of parent subscales and demographics

	Mother $N=406$ M (SD)	Father $N=256$ M (SD)	<i>p</i> Vvalue
SDQ prosocial	7.00 (2.29)	7.23 (2.17)	.189
SDQ externalizing	7.48 (4.15)	7.71 (4.04)	.488
SDQ internalizing	8.24 (4.08)	8.68 (4.01)	.177
SDQ total	16.40 (6.38)	15.73 (14.91)	.199
PFI negative	3.08 (1.33)	3.49 (1.35)	<.001
PFI positive	4.60 (1.20)	4.41 (1.18)	.052
Self-compassion	3.61 (0.58)	3.55 (0.68)	.237
PWB	7.38 (1.73)	7.24 (1.69)	.341
Age	45.1 (6.45)	44.0 (6.40)	.035

et al., 2018), the subsequent analysis refers to these SDQ subscales.

Predicting PWB and PFI

To determine whether self-compassion predicted PWB and PFI beyond that accounted for by SDQ subscales (the third hypothesis), we conducted a series of mixed linear regressions in two steps. Pairwise interactions of parent (father vs. mother), each of the SDQ subscales (prosocial, externalizing, internalizing) and their corresponding main effects were entered in the first step, and pairwise interactions of the parent and self-compassion and the corresponding main effect were entered in the second step. In both steps, child age and mean parent age were entered as control variables (Tables 3, 4, and 5).

Table 3 presents regression results for the outcome of negative parental emotion (PFI negative). In Step 1, SDQ subscales contributed significantly to the regression model, $F(9,473) = 14.45$, $p < 0.001$, and accounted for 21.6% of the variance in PFI negative. The following predictors significantly predicted PFI negative in step 1: parent, $\beta = -0.28$, $p < 0.01$, indicating higher PFI, SDQ externalizing, $\beta = 0.26$, $p < 0.01$, SDQ internalizing, $\beta = 0.21$, $p < 0.05$, indicating higher PFI negative values for increased SDQ externalizing and internalizing values. Adding self-compassion to the regression model accounted for an additional 9.1% of the

variance in PFI negative, and this change in R^2 was significant, $F(1,593) = 43.95$, $p < 0.001$. Adding self-compassion did not change the significance of the predictors in Step 1, with the exception of child age, which became a significant predictor, $\beta = 0.09$; $p < 0.05$, indicating higher PFI negative values for older children. Self-compassion was a significant predictor, $\beta = -0.27$; $p < 0.05$, indicating lower PFI negative values for increased self-compassion values. An examination of their standardized regression coefficients revealed that parent and self-compassion were stronger predictors than any other variables.

Table 4 presents the regression results for the outcome of positive parental emotion (PFI positive). In Step 1, SDQ subscales contributed significantly to the regression model, $F(9,476) = 8.28$, $p < 0.001$, and accounted for 13.5% of the variance in PFI positive. The following predictors significantly predicted PFI positive in step 1: SDQ prosocial, $\beta = 0.11$, $p < 0.05$, indicating that SDQ prosocial values were associated with higher PFI happiness values, parent, $\beta = 0.18$, $p < 0.01$, indicating higher PFI positive values for fathers than mothers, SDQ internalizing, $\beta = -0.12$, $p < 0.05$, indicating lower PFI positive values for increased SDQ internalizing values, and the interaction SDQ externalizing *parent, $\beta = -0.16$, $p < 0.05$. Figure 1 depicts the interaction between parent and SDQ externalizing. For mothers, as SDQ externalizing scores increased, PFI positive scores decreased,

Table 3 Standardized regression coefficients for SDQ and self-compassion predicting PFI negative (controlling for child age and mean parent age)

Variables	β	t	R^2	ΔR^2
Step 1			.216	.216
SDQ prosocial	.07	1.36		
Parent (father vs. mother)	-.28	-4.31**		
SDQ prosocial *parent (father)	-.09	-1.26		
SDQ externalizing	.26	5.22**		
SDQ externalizing *parent (father)	.05	0.66		
SDQ internalizing	.21	4.46**		
SDQ internalizing *parent (father)	-.03	-0.48		
Child age	.09	1.93		
Mean parent age	-.01	-0.22		
Step 2			.307	.091
SDQ prosocial	.09	1.94		
Parent (father vs. mother)	-.28	-4.44**		
SDQ prosocial *parent (father)	-.13	-1.84		
SDQ externalizing	.23	4.73**		
SDQ externalizing *parent (father)	.02	0.31		
SDQ internalizing	.21	4.42**		
SDQ internalizing *parent (father)	-.04	-0.65		
Self-compassion	-.27	-6.63**		
Self-compassion *parent (father)	.02	0.21		
Child age	.09	2.13*		
Mean parent age	-.01	-0.37		

* $p < .05$, ** $p < .01$

whereas for fathers, SDQ externalizing scores had a small positive effect. Additionally, younger child age, $\beta = -0.13$, $p < 0.01$, and mean parent age, $\beta = -0.10$, $p < 0.05$, were associated with higher PFI happiness values.

Adding self-compassion to the regression model accounted for an additional 7.5% of the variance in PFI positive, and this change in R^2 was significant, $F(1,590) = 27.48$, $p < 0.001$. Adding self-compassion did not change the significance of the predictors in Step 1, and self-compassion was a significant predictor, $\beta = 0.22$, $p < 0.001$. An examination of the standardized regression coefficients revealed that self-compassion was a stronger predictor than all other predictors.

Table 5 presents the regression results for the outcome of personal well-being (PWB). In Step 1, SDQ subscales contributed significantly to the regression model, $F(9,410) = 2.68$, $p < 0.01$, and accounted

Table 4 Standardized regression coefficients for SDQ and self-compassion predicting PFI positive (controlling for child age and mean parent age)

Variables	β	t	R^2	ΔR^2
Step 1			.135	.135
SDQ prosocial	.11	2.28*		
Parent (father vs. mother)	.18	2.77**		
SDQ prosocial *parent (father)	.09	1.22		
SDQ externalizing	-.03	-0.61		
SDQ externalizing *parent (father)	-.16	-2.21*		
SDQ internalizing	-.13	-2.53*		
SDQ internalizing (father)	.08	1.13		
Child age	-.13	-2.80**		
Mean parent age	-.10	-2.15*		
Step 2			.210	.075
SDQ prosocial	.10	2.02*		
Parent (father vs. mother)	.19	2.89**		
SDQ prosocial *parent (father)	.07	0.93		
SDQ externalizing	.02	0.41		
SDQ externalizing *parent (father)	-.17	-2.38*		
SDQ internalizing	-.11	-2.18*		
SDQ internalizing *parent (father)	.09	1.32		
Self-compassion	.22	5.24**		
Self-compassion *parent (father)	.08	1.03		
Child age	-.13	-2.92**		
Mean parent age	-.09	-2.00*		

* $p < .05$, ** $p < .01$

for 5.5% of the variance in PWB. SDQ internalizing significantly predicted PWB, $\beta = -0.14$, $p < 0.01$, indicating that increased SDQ internalizing values were associated with decreased PWB values. Adding self-compassion to the regression model accounted for an additional 39.7% of the variance in PWB, and this change in R^2 was significant, $F(1,593) = 154.44$, $p < 0.001$. After adding self-compassion, SDQ externalizing was no longer a significant predictor of PWB. Self-compassion was the only remaining significant predictor of PWB, $\beta = 0.50$, $p < 0.001$, indicating that increased self-compassion values were associated with increased PWB values. An examination of the standardized regression coefficients revealed that self-compassion was a significantly stronger predictor than all other predictors. This result mirrors the finding for the zero-order correlation analysis (Table 1), where the highest correlation coefficient was for self-compassion and PWB.

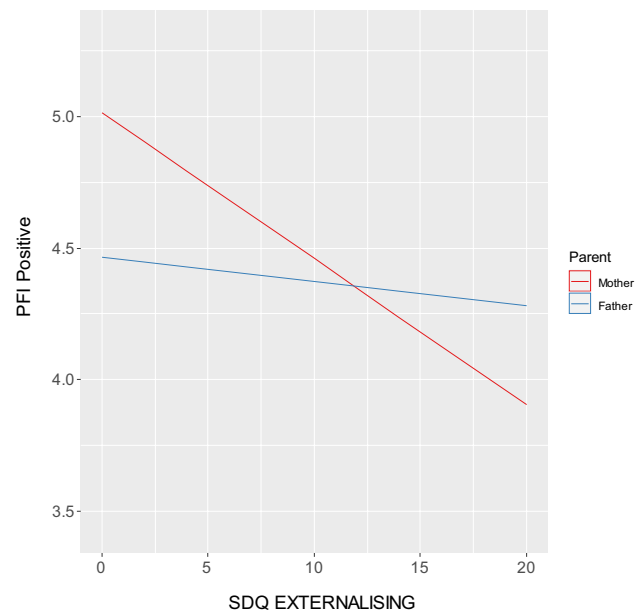
Table 5 Standardized regression coefficients for SDQ and self-compassion predicting PWB (controlling for child age and mean parent age)

Variables	β	t	R^2	ΔR^2
Step1				
			0.005	0.005
SDQ prosocial	.06	1.20		
Parent (father vs. mother)	.06	0.80		
SDQ prosocial *parent (father)	.03	0.38		
SDQ externalizing	-.10	-1.85		
SDQ externalizing *parent (father)	.03	0.43		
SDQ internalizing	-.14	-2.59**		
SDQ internalizing *parent (father)	.12	1.49		
Child age	-.04	-0.88		
Mean parent age	-.04	-0.84		
Step2				
			.402	.397
SDQ prosocial	.04	0.86		
Parent (father vs. mother)	.02	0.31		
SDQ prosocial *parent (father)	.03	0.47		
SDQ externalizing	-.05	-1.10		
SDQ externalizing *parent (father)	.03	0.43		
SDQ internalizing	-.08	-1.77		
SDQ internalizing *parent (father)	.09	1.36		
Self-compassion	.50	12.43**		
Self-compassion *parent (father)	.09	1.31		
Child age	-.02	-0.88		
Mean parent age	-.01	-0.84		

* $p < .05$, ** $p < .01$

Discussion

This study examined the potential role of parents' self-compassion in reducing their stress and increasing their well-being when raising children with mental health problems. Parent stress was examined by negative and positive feelings relating to parenthood. The first hypothesis was supported as follows: For parents of children with mental health problems, child externalizing and internalizing problems were positively correlated with negative parental feelings and negatively correlated with positive parental feelings and with parental well-being. This finding is consistent with previous studies that found that parenting stress is associated with both child externalizing and internalizing behavior problems (Dubois-Comtois et al., 2013; Rodriguez, 2011). Additionally, it is consistent with prior findings that child symptom severity is related to increased experiences of parenting stress (e.g., Angold et al., 1998; McDonald et al., 1997). Given that our sample included parents of children with a

**Fig. 1** Interaction of parent and SDQ externalizing predicting PFI positives

variety of psychiatric disorders, the results may indicate that parental stress relates to the severity of the clinical condition beyond the specific disorder. Similarly, in a meta-analysis conducted by Barroso et al. (2018), the authors documented the association between parenting stress and child behavior problems for different clinical groups.

The hypothesis that parent self-compassion is positively associated with less negative and more positive emotions and higher levels of well-being were supported by the study results. This finding highlights the important role of self-compassion in parenting children with special needs and expands it to children with mental health problems. Neff's pioneering definition of self-compassion, comprising the three primary components of self-kindness, humanity, and mindfulness, enables researchers to measure and explain why and how self-compassion contributes to reducing parental stress and increasing well-being (Neff, 2003; Neff & Faso, 2015).

When raising a child with mental health problems, parents who experience their child with kindness and perceive their lives in the context of the larger human experience can cope with the social stigma and the feelings of shame and blame for their child's mental illness in a more accepting, nonjudgmental way (e.g., Hinshaw & Stier, 2008; Moses, 2010; Van der Sanden et al., 2013). This could lead, in turn, to openness toward society, replacing the isolation and rejection that are common feelings among parents raising a child with mental health problems (Rodríguez-Meirinhos et al., 2018; Van der Sanden et al., 2013). Moreover, holding painful thoughts and feelings in mindful awareness rather than

ruminating over them may reduce parents' likelihood of being overwhelmed by the negative feelings and behaviors of their child and lead to greater awareness of the positive or natural aspects of their child's life. The same can be true for the negative parental feelings of frustration and helplessness generated when dealing with negative feelings and behaviors of children with mental health problems (Crowe et al., 2011; Lindgren et al., 2016).

Self-compassion was found to be a stronger predictor of reduced negative and increased positive parental feelings and parental well-being than the severity of the child's externalizing and internalizing symptoms, supporting the third hypothesis. These findings are consistent with previous studies of parents rearing children with autism (e.g., Neff & Faso, 2015) and expand the contribution of self-compassion to improving the experience of parenting children with a variety of emotional and behavioral problems. Significantly, self-compassion can serve as an acquirable, internal resource that enhances parents' ability to cope with the unexpected and uncontrollable external demands that accompany raising a child with mental health problems (Bögels et al., 2010).

Our study yielded two interesting findings regarding parental well-being. First, child internalizing but not externalizing problems were associated with a lower level of parental well-being. This finding requires further differentiation between the association of child internalizing and externalizing problems with parental well-being. Previous studies have called for similar examinations of child internalizing and externalizing problems and their different links to parental stress (e.g., Barroso et al., 2018).

Second, after adding self-compassion to the regression model, the level of the internalizing behavior problem was no longer a significant predictor of parental well-being, and self-compassion remained the only significant predictor of parental well-being. This finding suggests that with regard to a child's internalizing difficulties, self-compassion might act as a protective buffer between the child's mental health problems and parental well-being. Previous research supports this idea (Bohadana et al., 2019; Neff & Faso, 2015).

Why parental self-compassion might be especially helpful for internalizing compared with externalizing problems remains unclear and should be further investigated. Interestingly, a meta-analysis conducted by Muris (2016) revealed that self-compassion is substantially linked with internalizing versus externalizing psychopathology, which may reinforce the assumption that self-compassion has a different effect on different mental health problems both at the interpersonal level and in the parental context.

The implications for clinical practice are significant, highlighting the treatment potential of integrating a self-compassion component in parenting guidance and therapy for parents of children with mental health problems (e.g., Jefferson et al., 2020).

Limitations and Future Research

The study findings should be interpreted with caution due to several limitations. First, since the relation between parenting factors and child mental health problems is bidirectional, as reported in relation to both parent stress (Neece et al., 2012) and mindful parenting (Bögels et al., 2010), it is difficult to determine the directionality of the association between parenting factors and child mental health problems. It is also possible that there is a transactional relationship in which both parenting factors and child mental health problems impact one another over time. For example, it is possible that parents who have higher self-compassion perceive their child's emotional and behavior problems as less problematic. Longitudinal studies can measure parenting factors and child mental health problems and track the trajectory of these factors over time. Moreover, our study focuses on the role of self-compassion on parental stress and PWB. The role of parents' self-compassion should be further investigated with regard to parental practices to gain more understanding of the role of self-compassion in parenting.

Second, given the multiplicity of types of diagnoses and the small sample size for some of the diagnostic categories, we were unable to separately examine the impact of different types of mental health problems. Future studies should refer to specific diagnostic categories to examine differential effects of various mental health problems and the extent to which they affect parental feeling and well-being and the role of parent self-compassion. Another limitation relates to the fact that the data are based on self-report measures and are subject to both response bias and common method bias. In future studies, a mixed methodology, including measures from different sources (e.g., clinician assessment) and qualitative investigation (e.g., open-ended questions), may help reduce the potential for such bias (Podsakoff et al., 2003).

Furthermore, an analysis of demographic variables not considered here, such as parental education and ethnicity, can contribute important information.

An additional direction for future research would be to examine each of the three self-compassion components separately to determine which variable contributes most significantly to parental feeling and well-being. These and other future studies may have important implications for mental healthcare practice, suggesting parental self-compassion as a target for assessment and therapy.

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Author Contribution VS designed and executed the study, assisted with the data analyses, and wrote the paper. NW designed and executed the study, assisted with the data analyses, and wrote the paper. UY directed

execution of the research onsite, including data collection. All authors approved the final version of the manuscript for submission.

Data Availability The data that support the findings of this study are available from the corresponding author upon reasonable request.

Declarations

Ethics Approval This research was approved by the Ethics Committee of Ziv Medical Center in Israel. Informed consent was obtained from all individual participants included in the study. The procedures used in this study adhere to the tenets of the Declaration of Helsinki.

Conflict of Interest The authors declare no competing interests.

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