# Missing Output Identification Model Based Recursive Least Squares Algorithm for a Distributed Parameter System

Jing Chen, Bin Jiang\*, and Juan Li

**Abstract:** This paper proposes a recursive least squares algorithm for a distributed parameter system with missing observations. By using the finite difference method, the distributed parameter system can be turned into a lumped parameter system. Then a missing output identification model based recursive least squares algorithm is derived to estimate the unknown parameters of the lumped parameter system. Furthermore, the parameters of the distributed parameter system can be computed by the estimated parameters of the lumped parameter system. The simulation results indicate that the proposed method is effective.

**Keywords:** Distributed parameter system, finite difference method, missing output identification model, parameter estimation, recursive least squares.

# 1. INTRODUCTION

Systems can be roughly divided into two classes: the lumped parameter systems (LPSs) and the distributed parameter systems (DPSs). The LPSs are described by difference equations and have finite dimension. The DPSs are described by partial differential equations and are of infinite dimension. Recently, LPSs identification has received much attention, and there exist a lot of identification methods for LPSs [1–3], including the least squares (LS) algorithms [4–6], the stochastic gradient algorithms [7–9] and the iterative algorithms [10–12].

The DPSs are widely existed in engineering practice, e.g., in semiconductor manufacturing, nanotechnology, biotechnology and chemical engineering [13, 14]. System identification is the first step for many applications such as prediction, control and fault tolerant control [15–19]. Unfortunately, the identification methods for LPSs are not suitable for DPSs, because the LPSs identification methods ignore the important features of the spatial response exhibited by the DPSs.

Over the past few decades, several methods have been developed for DPSs identification [20–23], and these methods use proper basis functions to turn DPSs into LPSs. For example, Zill and Cullen proposed a weighted residual method for DPSs, the spatio-temporal variable of the DPSs can be expanded as a set of spatial basis func-

tions [24]. Li and Qi used a Karhunen-Loeve method to turn a DPS into an LPS, then applied an LS algorithm to estimate the unknown parameters of the LPS [25]. However, when use these methods to truncate the infinite dimension to a finite dimension, the selection of dimension is difficult, a small dimension may lead a large identification error, but a large dimension may reduce the computational efficiency. Recently, Chen and Jiang developed a RLS algorithm and an SG algorithm for a twodimensional system with the assumption that the system is a single-rate system [26].

Dual-rate systems which have different sampled rates of input and output, are widely existed in many engineering applications [27-29]. For example, in polymer reactors, the manipulated variables can be adjusted at relatively fast rate, whereas the composition, density or molecular weight distribution measurements are typically obtained after several minutes of analysis [30]. The lifting technique and the polynomial transformation technique are two methods which are generally used for dual-rate systems identification [31-34]. However, these two methods can only estimate the parameters of the transformed system and can increase the number of the unknown parameters. In order to overcome these difficulties, Chen presented a missing output identification (MOI) method for dual-rate one-dimensional systems, the method can estimate the parameters of the dual-rate systems directly and

Manuscript received September 27, 2016; revised December 26, 2016 and March 2, 2017; accepted May 24, 2017. Recommended by Associate Editor Yongping Pan under the direction of Editor Duk-Sun Shim. This work is supported by the National Natural Science Foundation of China (Nos. 61403165,61374126), the Natural Science Foundation of Jiangsu Province (No. BK20131109), the Natural Science Foundation of Shandong Province (ZR2013FM021), the Post Doctoral Foundation of Jiangsu Province (No. 1501015A) and the Natural Science Fund for Colleges and Universities in Jiangsu Province (No. 16KJB120006).

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keep the number of the unknown parameters unchanged [2]. In this paper, we will extend the MOI method to a dual-rate two-dimensional system.

The contributions of this paper are as follows.

- Propose a missing output identification model based recursive least squares (MOI-RLS) algorithm for dualrate two-dimensional systems.
- The proposed method can estimate the unknown parameters and the missing outputs simultaneously; Furthermore, cannot increase the number of the unknown parameters.
- 3) Compared with the work in [35], the MOI-RLS method has less computation efforts.

The rest of this paper is organized as follows. Section 2 introduces the DPS and the finite difference method. Section 3 presents an MOI-RLS algorithm for a general dual-rate two-dimensional system. Section 4 provides an illustrative example. Finally, concluding remarks are given in Section 5.

#### 2. PROBLEM FORMULATION

Let "A =: X" or "X := A" stand for "A is defined as X", the superscript T denote the matrix transpose.

Consider the following DPS,

$$\frac{\partial y(x,t)}{\partial t} = a_1 \frac{\partial^2 y(x,t)}{\partial x^2} + a_2 \frac{\partial y(x,t)}{\partial x} + a_3 y(x,t) + a_4 u(x,t), \tag{1}$$

where y(x,t) is the system output, u(x,t) is the system input,  $a_1$ ,  $a_2$ ,  $a_3$  and  $a_4$  are unknown parameters to be estimated, and with the assumption that  $0 \le a_1$ ,  $0 \le a_2 \le L$ and  $K \le a_3 \le 0$ , L is a positive constant and K is a negative constant, both L and K are known in prior.

Using the finite difference method gives

$$\begin{split} &\frac{\partial^2 y(x,t)}{\partial x^2}|_{x_i,t_j} \approx \frac{y_{i+1,j+1} - 2y_{i,j+1} + y_{i-1,j+1}}{\Delta x^2}, \\ &\frac{\partial y(x,t)}{\partial t}|_{x_i,t_j} \approx \frac{y_{i,j+1} - y_{i,j}}{\Delta t}, \\ &\frac{\partial y(x,t)}{\partial x}|_{x_i,t_j} \approx \frac{y_{i+1,j} - y_{i,j}}{\Delta x}, \end{split}$$

where  $y_{i,j} = y(x_i, t_j)$  is the output at discretization node (i, j),  $\Delta x$  and  $\Delta t$  are two small intervals at space and time, respectively. By this difference algorithm, derivatives at each discretization node (i, j) are approximated by the difference over a small interval. Then the DPS in (1) can be transformed as follow,

$$\frac{y_{i,j+1} - y_{i,j}}{\Delta t} \approx a_1 \frac{y_{i+1,j+1} - 2y_{i,j+1} + y_{i-1,j+1}}{\Delta x^2} + a_2 \frac{y_{i+1,j} - y_{i,j}}{\Delta x} + a_3 y_{i,j} + a_4 u_{i,j},$$

and can be simplified as an LPS (or two-dimensional system),

$$(1+2r_1)y_{i,j+1} - r_1y_{i+1,j+1} - r_1y_{i-1,j+1} = r_2y_{i+1,j} + (r_3 - r_2 + 1)y_{i,j} + r_4u_{i,j},$$
(2)

where  $r_1 = a_1 \frac{\Delta t}{\Delta x^2}$ ,  $r_2 = a_2 \frac{\Delta t}{\Delta x}$ ,  $r_3 = a_3 \Delta t$  and  $r_4 = a_4 \Delta t$ .

The finite difference method cannot guarantee that the DPS be approximated by the LPS because of the approximation error. Therefore, in order to keep the finite difference method convergent, Von. Neumann stability analysis is introduced.

Based on Von. Neumann stability analysis, the error equation has the same structure with the LPS [36], and can be written as

$$(1+2r_1)z_{i,j+1} - r_1z_{i+1,j+1} - r_1z_{i-1,j+1}$$
  
=  $r_2z_{i+1,j} + (r_3 - r_2 + 1)z_{i,j},$  (3)

where  $z_{i,j+1}$  is the approximation error at (i, j+1), and

$$z_{i,j} = V^{j}(k)e^{ikx_{i}}, \qquad z_{i,j+1} = V^{j+1}(k)e^{ikx_{i}},$$
  

$$z_{i+1,j+1} = V^{j+1}(k)e^{ikx_{i}}e^{ik\Delta x},$$
  

$$z_{i-1,j+1} = V^{j+1}(k)e^{ikx_{i}}e^{-ik\Delta x},$$
  

$$z_{i+1,j} = V^{j}(k)e^{ikx_{i}}e^{ik\Delta x},$$
  
(4)

in which V(k) is a growth factor, k is the number of waves and e is short for exp. When  $|V(k)| \leq 1$ , the difference algorithm is convergent; when |V(k)| > 1, the difference algorithm is divergent.

Rewrite (3) as

$$(1+2r_1)V^{j+1}(k)e^{ikx_i} - r_1V^{j+1}(k)e^{ikx_i}e^{ik\Delta x} -r_1V^{j+1}(k)e^{ikx_i}e^{-ik\Delta x} = r_2V^j(k)e^{ikx_i}e^{ik\Delta x} + (r_3 - r_2 + 1)V^j(k)e^{ikx_i},$$
(5)

and the growth factor V(k) can be expressed as

$$V(k) = \frac{r_2 e^{ik\Delta x} + r_3 - r_2 + 1}{1 + 2r_1 - r_1 e^{ik\Delta x} - r_1 e^{-ik\Delta x}}.$$
(6)

Since

$$e^{ik\Delta x} + e^{-ik\Delta x} = 2\cos(k\Delta x),$$
  

$$e^{ik\Delta x} - e^{-ik\Delta x} = 2\sin(k\Delta x)i,$$
  

$$1 - \cos(k\Delta x) = 2\sin^2\frac{k\Delta x}{2},$$

we have  $e^{ik\Delta x} = \cos(k\Delta x) + \sin(k\Delta x)i$ . Simplifying Equation (6) gives

$$V(k) = \frac{r_2 e^{ik\Delta x} + r_3 - r_2 + 1}{1 + 2r_1 - 2r_1 \cos(k\Delta x)}$$
$$= \frac{r_2 \cos(k\Delta x) + r_3 - r_2 + 1 + r_2 \sin(k\Delta x)i}{1 + 4r_1 \sin^2 \frac{k\Delta x}{2}}$$

$$=\frac{r_{2}\cos(k\Delta x) + r_{3} - r_{2} + 1}{1 + 4r_{1}\sin^{2}\frac{k\Delta x}{2}} + \frac{r_{2}\sin(k\Delta x)}{1 + 4r_{1}\sin^{2}\frac{k\Delta x}{2}}i.$$
(7)

Let

$$a = \frac{r_2 \cos(k\Delta x) + r_3 - r_2 + 1}{1 + 4r_1 \sin^2 \frac{k\Delta x}{2}}$$
$$b = \frac{r_2 \sin(k\Delta x)}{1 + 4r_1 \sin^2 \frac{k\Delta x}{2}}.$$

Then we have

$$\begin{split} V(k)|^2 &= a^2 + b^2 \\ &= \frac{(r_2 \cos(k\Delta x) + r_3 - r_2 + 1)^2}{(1 + 4r_1 \sin^2 \frac{k\Delta x}{2})^2} + \frac{(r_2 \sin(k\Delta x))^2}{(1 + 4r_1 \sin^2 \frac{k\Delta x}{2})^2} \\ &= \frac{4r_2 \sin^2(\frac{k\Delta x}{2})(r_2 - r_3 - 1) + (r_3 + 1)^2}{(1 + 4r_1 \sin^2 \frac{k\Delta x}{2})^2}. \end{split}$$

Since  $r_1 = a_1 \frac{\Delta t}{\Delta x^2}$ ,  $r_2 = a_2 \frac{\Delta t}{\Delta x}$ ,  $r_3 = a_3 \Delta t$ ,  $a_1 \ge 0$ ,  $a_2 \ge 0$  and  $a_3 \le 0$ , we can choose two small positive numbers  $\Delta t$  and  $\Delta x$  to keep  $r_1 \ge 0$ ,  $0 \le r_2 \le \frac{1}{2}$  and  $-\frac{1}{2} \le r_3 \le 0$ . Then we can get

$$|V(k)|^{2} = \frac{4r_{2}^{2}\sin^{2}(\frac{k\Delta x}{2})(r_{2} - r_{3} - 1) + (r_{3} + 1)^{2}}{(1 + 4r_{1}\sin^{2}\frac{k\Delta x}{2})^{2}} \\ \leqslant \frac{(r_{3} + 1)^{2}}{(1 + 4r_{1}\sin^{2}\frac{k\Delta x}{2})^{2}} \leqslant 1.$$
(8)

Here,  $r_1 \ge 0$ ,  $0 \le r_2 \le \frac{1}{2}$  and  $-\frac{1}{2} \le r_3 \le 0$  are equivalent to  $a_1 \ge 0$ ,  $L \le \frac{\Delta x}{2\Delta t}$  and  $-\frac{1}{2\Delta t} \le K$ . Thus, based on Von. Neumann stability analysis and the assumption of  $a_1$ ,  $a_2$  and  $a_3$ , we can conclude that the DPS in (1) can be approximated by the LPS in (2).

Simplifying (2) gets

$$y(i+1, j+1) = a_{0,1}y(i+1, j) + a_{1,0}y(i, j+1) + a_{2,0}y(i-1, j+1) + a_{1,1}y(i, j) + b_{1,1}u(i, j),$$
(9)

where  $y(i, j) = y_{i,j}$ ,  $a_{0,1} = -\frac{r_2}{r_1}$ ,  $a_{1,0} = \frac{1+2r_1}{r_1}$ ,  $a_{2,0} = -1$ ,  $a_{1,1} = -\frac{r_3 - r_2 + 1}{r_1}$  and  $b_{1,1} = -\frac{r_4}{r_1}$ . Rewrite (9) as

$$y(i, j) + y(i-2, j) = a_{0,1}y(i, j-1) + a_{1,0}y(i-1, j) + a_{1,1}y(i-1, j-1) + b_{1,1}u(i-1, j-1).$$
(10)

## 3. THE MOI-RLS ALGORITHM

Consider the general dual-rate two-dimensional system as

$$y(x,tq) = a_{0,1}y(x,tq-1) + a_{0,2}y(x,tq-2) + \cdots$$

$$+ a_{0,q}y(x,tq-q) + \dots + a_{0,n}y(x,tq-n) + a_{1,0}y(x-1,tq) + a_{2,0}y(x-2,tq) + \dots + a_{m,0}y(x-m,tq) + a_{1,1}y(x-1,tq-1) + a_{1,2}y(x-1,tq-2) + \dots + a_{1,q}y(x-1,tq-q) + \dots + a_{m,1}y(x-m,tq-n) + \dots + a_{m,1}y(x-m,tq-1) + a_{m,2}y(x-m,tq-2) + \dots + a_{m,q}y(x-m,tq-q) + \dots + a_{m,n}y(x-m,tq-n) + b_{1,1}u(x-1,tq-1) + v(x,tq),$$
(11)

where u(x,t) and y(x,t) are input and output, respectively. v(x,t) is a stochastic white noise with zero mean. *n* means that the output at time tq depends on all the outputs from tq - n to tq - 1, while *m* means that the output at space *x* depends on all the outputs from x - m to x - 1. For the dual-rate sampled data system, the input data u(x,t)is sampled at a quicker rate than the output y(x,t), thus all the input data  $\{u(x,t), x = 1, 2, \dots, M, t = 0, 1, 2, \dots\}$ and only the scarce output data  $\{y(x,tq), (q \ge 2)\}$  are measureable. The intersample outputs  $\{y(x,tq+h), h = 1, 2, \dots, q-1\}$  are unavailable. *M* is the total number of the points in the space at time *t*, as shown in Fig. 1.

Define the parameter vector  $\theta$  and the information vector  $\varphi(x,tq)$  as

$$\begin{aligned} \boldsymbol{\theta} &:= [a_{0,1}, a_{0,2}, \cdots, a_{0,n}, a_{1,0}, a_{2,0}, \cdots, \\ & a_{m,0}, a_{1,1}, a_{1,2}, \cdots, a_{1,n}, \cdots, \\ & a_{m,1}, a_{m,2}, \cdots, a_{m,n}, b_{1,1}]^{\mathsf{T}} \in \mathbb{R}^{p}, \\ & p = (m+1)(n+1), \end{aligned}$$
(12)  
$$\boldsymbol{\varphi}(x,tq) &:= [y(x,tq-1), y(x,tq-2), \cdots, \\ & y(x,tq-q), \cdots, y(x,tq-n), y(x-1,tq), \\ & y(x-2,tq), \cdots, y(x-m,tq), \\ & y(x-1,tq-1), y(x-1,tq-2), \cdots, \\ & y(x-1,tq-q), \cdots, \\ & y(x-1,tq-q), \cdots, \\ & y(x-m,tq-2), \cdots, y(x-m,tq-1), \\ & y(x-m,tq-2), \cdots, y(x-m,tq-q), \cdots, \\ & y(x-m,tq-n), \\ & u(x-1,tq-1)]^{\mathsf{T}} \in \mathbb{R}^{p}. \end{aligned}$$
(13)

Rewrite (11) as an identification model,

 $y(x,tq) = \boldsymbol{\varphi}^{\mathrm{T}}(x,tq)\boldsymbol{\theta} + v(x,tq).$ 

Using the following RLS algorithm proposed in [26] to estimate the parameter vector  $\theta(x, tq)$ :

$$\begin{split} \hat{\theta}(x,tq) &= \hat{\theta}(x-1,tq-1) + P(X,Tq)P^{-1}(X-1,tq) \\ &\times (\hat{\theta}(x-1,tq) - \hat{\theta}(x-1,tq-1)) \\ &+ P(X,Tq) \times P^{-1}(x,Tq-1)(\hat{\theta}(x,tq-1)) \end{split}$$



•: available observations; o: missing observations

Fig. 1. The missing observations pattern

$$-\hat{\theta}(x-1,tq-1)) + P(X,Tq)\varphi(x,tq)(y(x,tq)) - \varphi^{\mathrm{T}}(x,tq)\hat{\theta}(x-1,tq-1)), \qquad (14)$$

$$P^{-1}(X, Tq) = P^{-1}(0, 0) + \sum_{i=1}^{x} \sum_{j=1}^{tq} \varphi(i, j) \varphi^{\mathsf{T}}(i, j),$$
$$P(0, 0) = p_0 I, \quad p_0 = 10^6, \tag{15}$$

$$P^{-1}(X-1,tq) = P^{-1}(0,0) + \sum_{i=1}^{x-1} \varphi(i,tq) \varphi^{\mathsf{T}}(i,tq),$$
(16)

$$P^{-1}(x, Tq - 1) = P^{-1}(0, 0) + \sum_{j=1}^{tq-1} \varphi(x, j) \varphi^{\mathsf{T}}(x, j),$$
(17)

where  $P(\cdot)$  is the covariance matrix,  $P^{-1}(\cdot)$  is the inverse matrix of  $P(\cdot)$  and *I* is an identity matrix. Unfortunately, this RLS algorithm is impossible to implement, because the information vector  $\varphi(x,tq)$  on the right-hand side of (14) contains the missing outputs  $y(x,tq-i), i = 1, 2 \cdots, q - 1$ . If the system is a one-dimensional system, the lifting technique and the polynomial transformation technique can be used to overcome this difficulty. However, due to the complexity of the polynomial of the output, the lifting technique and the polynomial transformation technique cannot be used for this dual-rate two-dimensional system. In this paper, we use the MOI-RLS algorithm to overcome this difficulty.

The missing outputs are replaced with the outputs of an MOI model,

$$\begin{split} \hat{y}(x,tq-q+h) &= \hat{\varphi}^{\mathsf{T}}(x,tq-q+h)\hat{\theta}(x,tq-q), \\ x &= 1,\cdots, M, \ h = 1,2,\cdots,q-1, \\ \hat{\varphi}(x,tq-q+h) &= [\hat{y}(x,tq-q+h-1), \\ \hat{y}(x,tq-q+h-2),\cdots,y(x,tq-q),\cdots, \\ \hat{y}(x,tq-q+h-n), \hat{y}(x-1,tq-q+h), \\ \hat{y}(x-2,tq-q+h),\cdots, \hat{y}(x-m,tq-q+h), \end{split}$$

$$\begin{split} \hat{y}(x-1,tq-q+h-1), \hat{y}(x-1,tq-q+h-2), \cdots, \\ y(x-1,tq-q), \cdots, \hat{y}(x-1,tq-q+h-n), \cdots, \\ \hat{y}(x-m,tq-q+h-1), \hat{y}(x-m,tq-q+h-2), \cdots, \\ y(x-m,tq-q), \cdots, \hat{y}(x-m,tq-q+h-n), \\ u(x-1,tq-q+h-1)]^{\mathsf{T}}, \end{split}$$

where  $\hat{y}(x,tq-q+h)$  represents the estimate of y(x,tq-q+h) at (x,tq-q+h),  $\hat{\theta}(x,tq-q)$  represents the estimate of  $\theta$  at (x,tq-q), and  $\hat{\varphi}(x,tq-q+h)$  represents the estimate of  $\varphi$  at (x,tq-q+h).

Using the following MOI-RLS algorithm to estimate the parameter vector  $\theta$  in (12):

$$\hat{\theta}(x,tq) = \hat{\theta}(x-1,tq-q) + P(X,Tq)P^{-1}(X-1,tq) \\ \times (\hat{\theta}(x-1,tq) - \hat{\theta}(x-1,tq-q)) \\ + P(X,Tq)P^{-1}(x,Tq-q)(\hat{\theta}(x,tq-q) \\ - \hat{\theta}(x-1,tq-q)) \\ + P(X,Tq)\hat{\phi}(x,tq)(y(x,tq) \\ - \hat{\phi}^{\mathrm{T}}(x,tq)\hat{\theta}(x-1,tq-q)),$$
(18)

$$\hat{\theta}(x,tq-q+h) = \hat{\theta}(x,tq-q), \ h = 1,2,\cdots,q-1,$$
(19)

$$P^{-1}(X, Tq) = P^{-1}(0, 0) + \sum_{i=1}^{x} \sum_{j=1}^{tq} \hat{\varphi}(i, j) \hat{\varphi}^{\mathsf{T}}(i, j),$$
$$P(0, 0) = p_0 I, \quad p_0 = 10^6, \tag{20}$$

$$P^{-1}(X-1,tq) = P^{-1}(0,0) + \sum_{i=1}^{x-1} \hat{\varphi}(i,tq) \hat{\varphi}^{\mathsf{T}}(i,tq),$$
(21)

$$P^{-1}(x, Tq - q) = P^{-1}(0, 0) + \sum_{j=1}^{tq-q} \hat{\varphi}(x, j) \hat{\varphi}^{\mathsf{T}}(x, j),$$
(22)

$$\hat{y}(x,tq-q+h) = \hat{\phi}^{T}(x,tq-q+h)\hat{\theta}(x,tq-q), \quad (23) 
\hat{\phi}(x,tq-q+h) 
= [\hat{y}(x,tq-q+h-1), \hat{y}(x,tq-q+h-2), \cdots,$$

$$\begin{array}{ll} y(x,tq-q), & \cdots, & \hat{y}(x,tq-q+h-n), \\ \hat{y}(x-1,tq-q+h), & \hat{y}(x-2,tq-q+h), \cdots, \\ \hat{y}(x-m,tq-q+h), & \hat{y}(x-1,tq-q+h-1), \\ \hat{y}(x-1,tq-q+h-2), & \cdots, & y(x-1,tq-q), & \cdots, \\ \hat{y}(x-1,tq-q+h-2), & \cdots, & y(x-1,tq-q), & \cdots, \\ \hat{y}(x-m,tq-q+h-1), & \hat{y}(x-m,tq-q+h-2), \\ & \cdots, & y(x-m,tq-q), & \cdots, & \hat{y}(x-m,tq-q+h-n) \\ u(x-1,tq-q+h-1)]^{\mathrm{T}}. \end{array}$$

The steps of computing the parameter estimate  $\hat{\theta}(x,tq)$  by the MOI-RLS algorithm are listed in the following.

- 1) Let  $y(-i, -j) = 0, v(-i, -j) = 0, u(-i, -j) = 0, i = 0, 1, 2, \dots, m-1, j = 0, 1, 2, \dots, n-1$  and give a small positive numbers  $\varepsilon$ .
- 2) Let t = 1, r(0) = 1, and  $P(0,0) = p_0 I$  and  $\hat{\theta}(0,0) = 1/p_0$  with 1 being a column vector whose entries are all unity and  $p_0 = 10^6$ .
- 3) Let h = 1, x = 1, collect the input-output data y(x,tq), u(x,t).
- 4) Compute  $\hat{y}(x, tq q + h)$  by (23).
- 5) Increase *x* by 1 and compare *x* with *M*, if  $x \le M$ , form  $\hat{\varphi}(x, tq q + h)$  by (24) and go to step 4; otherwise, go to step 6.
- 6) Let x = 1 again, and increase h by 1, if  $h \le q 1$ , go to step 4; otherwise, go to next step.
- 7) Form  $\hat{\varphi}(x,tq)$  by (24), compute  $P^{-1}(X,Tq)$ ,  $P^{-1}(X-1,tq)$  and  $P^{-1}(x,Tq-q)$  by (20), (21) and (22), respectively.
- 8) Update the parameter estimation vector  $\hat{\theta}(x,tq)$  by (18).
- 9) Compare  $\hat{\theta}(x,tq)$  and  $\hat{\theta}(x-1,tq-q)$ : if  $\|\hat{\theta}(x,tq) \hat{\theta}(x-1,tq-q)\| \le \varepsilon$ , then terminate the procedure and obtain  $\hat{\theta}(x,tq)$ ; otherwise, increase x by 1 and go to next step.
- 10) Compare *x* and *M*: if  $x \le M$ , go to step 7; otherwise, increase *t* by 1 and go to step 3.

Next, we will compute the original parameters of the DPS. Rewrite (10) as

$$\begin{split} y(x,tq) + y(x-2,tq) = & a_{0,1}y(x,tq-1) \\ &+ a_{1,0}y(x-1,tq) \\ &+ a_{1,1}y(x-1,tq-1) \\ &+ b_{1,1}u(x-1,tq-1). \end{split}$$

Define the parameter vector  $\theta$  and the information vector  $\varphi(x,tq)$  as

$$\boldsymbol{\theta} := [a_{0,1}, a_{1,0}, a_{1,1}, b_{1,1}]^{\mathrm{T}}, \tag{25}$$

$$\varphi(x,tq) := [y(x,tq-1), y(x-1,tq), y(x-1,tq-1), u(x-1,tq-1)]^{\mathsf{T}}.$$
(26)

By using the estimated parameters  $\{\hat{a}_{0,1}, \hat{a}_{1,0}, \hat{a}_{1,1}, \hat{b}_{1,1}\}$ , we can compute  $\hat{r}_1 = \frac{1}{\hat{a}_{1,0}-2}$ ,  $\hat{r}_2 = -\hat{r}_1\hat{a}_{0,1}$ ,  $\hat{r}_3 = -\hat{r}_1\hat{a}_{1,1} + \hat{r}_2 - 1$  and  $\hat{r}_4 = -\hat{r}_1\hat{b}_{1,1}$ . Because  $\Delta x$  and  $\Delta t$  are already known, we can compute  $\hat{a}_1 = \frac{\hat{r}_1\Delta x^2}{\Delta t}$ ,  $\hat{a}_2 = \frac{\hat{r}_2\Delta x}{\Delta t}$ ,  $\hat{a}_3 = \frac{\hat{r}_3}{\Delta t}$  and  $\hat{a}_4 = \frac{\hat{r}_4}{\Delta t}$  by the estimates  $\hat{r}_1$ ,  $\hat{r}_2$ ,  $\hat{r}_3$  and  $\hat{r}_4$ .

### 4. NUMERICAL RESULTS

Consider the following temperature distributed parameter system of large-scale vertical quench furnace in [37],

$$\frac{\partial y(x,t)}{\partial t} = a_1 \frac{\partial^2 y(x,t)}{\partial x^2} + a_2 \frac{\partial y(x,t)}{\partial x} + a_3 y(x,t) + a_4 u(x,t),$$

where y(x,t) is the temperature of the large-scale vertical quench furnace and u(x,t) is the current of the heating element,  $[a_1, a_2, a_3, a_4] = [0.1, 0.4, -5, 6]$ . Assume L = 0.4 and K = -5. Since  $L \leq \frac{\Delta x}{2\Delta t}$  and  $-\frac{1}{2\Delta t} \leq K$ , one can choose  $\Delta t = 0.1$  and  $\Delta x = 0.1$ . Then, we have  $r_1 = 1$ ,  $r_2 = 0.4$ ,  $r_3 = -0.5$  and  $r_4 = 0.6$ , and

$$\begin{split} \theta &= [a_{0,1}, a_{1,0}, a_{1,1}, b_{1,1}]^{\mathsf{T}} = [-0.4, 3, -0.1, -0.6]^{\mathsf{T}},\\ \varphi(x,t) &= [y(x,t-1), y(x-1,t), y(x-1,t-1), \\ u(x-1,t-1)]^{\mathsf{T}}. \end{split}$$

Then the DPS can be simplified as a two-dimensional system. Assume that the two-dimensional system is a dualrate system with two different updating periods q = 2 and q = 3, then we can get

$$\mathbf{y}(\mathbf{x},t) + \mathbf{y}(\mathbf{x}-2,t) = \boldsymbol{\varphi}^{\mathrm{T}}(\mathbf{x},t)\boldsymbol{\theta} + \mathbf{v}(\mathbf{x},t).$$

In simulation, the input  $\{u(x,t)\}$  is taken as a persistent excitation signal sequence with zero mean and unit variance, and  $\{v(x,t)\}$  is taken as a white noise sequence with zero mean and variance  $\sigma^2 = 0.10^2$ . The noise-to-signal ratio is  $\delta_{ns} = 16.667\%$ .

Define the parameter estimation error as

$$\delta := \|\hat{\theta} - \theta\| / \|\theta\|.$$

Assume M = 50, then apply the MOI-RLS algorithm to estimate the parameters. The parameter estimates and their errors are shown in Tables 1-2, and the parameter estimation errors  $\delta$  versus (x,t) are shown in Fig. 2.

From Tables 1-2 and Fig. 2, we can conclude:

- 1. The parameter estimation errors become smaller and smaller and go to zero with (x, t) increasing.
- 2. When q = 2, we can compute the original parameters  $[a_1, a_2, a_3, a_4]^{\mathsf{T}} = [0.099, 0.399, -5.012, 5.992]^{\mathsf{T}}$  by using the parameter estimate  $\hat{\theta}(50, 50)$ .
- 3. When *q* becomes larger, the parameter estimation errors also become larger because of too much outputs are missing.

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(x,t)	$a_{0,1}$	$a_{1,0}$	$a_{1,1}$	$b_{1,1}$	$\delta$ (%)
(5,5)	-0.33974	4.01138	-0.03762	-0.31474	32.55981
(10,10)	-0.40872	2.76224	-0.09443	-0.42255	9.15148
(15,15)	-0.39827	2.82413	-0.11600	-0.48578	6.48163
(20,20)	-0.39900	3.00628	-0.09972	-0.51248	2.70632
(30,30)	-0.40003	3.01096	-0.09859	-0.56634	1.09169
(40,40)	-0.40007	3.00411	-0.09937	-0.58447	0.49553
(50,50)	-0.40004	3.00037	-0.09995	-0.59977	0.40364
True Values	-0.40000	3.00000	-0.10000	-0.60000	

Table 1. The MOI-RLS algorithm estimates and errors (q = 2).

Table 2. The MOI-RLS algorithm estimates and errors (q = 3).

(x,t)	$a_{0,1}$	$a_{1,0}$	$a_{1,1}$	$b_{1,1}$	$oldsymbol{\delta}(\%)$
(5,5)	-0.83309	4.92312	-0.48469	0.28315	89.05208
(10,10)	-0.32622	3.68315	-0.07661	-0.37172	22.74399
(15,15)	-0.45332	2.32576	-0.19272	-0.51357	22.03352
(20,20)	-0.40905	2.40601	-0.15711	-0.52564	18.59932
(30,30)	-0.39676	2.93020	-0.10138	-0.56119	2.46592
(40,40)	-0.40034	2.96097	-0.09805	-0.57887	1.36904
(50,50)	-0.40006	2.94766	-0.10002	-0.59576	1.61822
True Values	-0.40000	3.00000	-0.10000	-0.60000	



Fig. 2. The parameter estimation errors  $\delta$  versus (x,t)

## 5. CONCLUSIONS

This paper proposes an identification method for a distributed parameter system with missing observations. Based on Von. Neumann stability analysis, the distributed parameter system can be simplified as a lumped parameter system. Then a missing output identification model based recursive least squares algorithm is derived to estimate the unknown parameters of the lumped parameter system. Different from the lifting technique and the polynomial transformation technique, the method in this paper can be used for dual-rate two-dimensional system identification, and can estimate the unknown parameters directly. Furthermore, this method can also be extended to fault detection, diagnosis and estimation [38–42].

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