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C677T and A1298C methylenetetrahydrofolate reductase polymorphisms and breast cancer susceptibility among Latinos: a meta-analysis

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Abstract

Background Previous meta-analyses have shown an ethnic dependency of the C677T and the A1298C methylenetetrahydrofolate reductase (MTHFR) polymorphisms, with no focus on the Latino population. For Latinos, many studies have examined these polymorphisms and breast cancer susceptibility, yielding no concise result. Therefore, we undertook this meta-analysis to determine the effect these polymorphisms have on breast cancer risk for Latinos.

Methods PubMed, EBSCO, LILACS, Scopus, and Latin American-specific databases were searched for studies exploring the association between the MTHFR polymorphisms and breast cancer susceptibility in Latinos until January 2019. Genotype distributions were extracted and, depending on the level heterogeneity determined by the ψ^2 -based Q test and the I^2 test, fixed-effects or random-effects models were used to calculate pooled odds ratios (ORs) with 95% confidence intervals (95% CIs) for the heterozygous, homozygous, dominant, recessive, and allelic genetic models. No publication bias was detected by the Begg–Mazumdar's test and Egger's test.

Results Of the 280 retrieved publications, 9 studies were included: 9 for the C677T polymorphism and 5 for the A1298C polymorphism. For the C677T polymorphism, there was an elevated risk for the homozygous (OR 1.42, 95% CI 1.05–1.92), the dominant (OR 1.16, 95% CI 1.02–1.31), the recessive (OR 1.33, 95% CI 1.01–1.75), and the allelic model (OR 1.17, 95% CI 1.03–1.33, p < 0.01). No association between the A1298C polymorphism and the risk to develop breast cancer was determined.

Conclusion The results indicated that, for Latinos, the C677T polymorphism is associated with a significant risk for developing breast cancer, whereas the A1289C polymorphism does not.

Keywords MTHFR · Mutation · Carcinogenesis · Latin America

Perla Meneses-Sanchez and Samantha C. Garcia-Hernandez contributed equally for the first authorship.

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Introduction

Cancer is one of the leading causes of death worldwide. In Latin America, breast cancer is the most frequently diagnosed cancer, with over 150,000 new cases annually

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[1]. Numerous reports have demonstrated the association between developing breast cancer and a patient's exposure to ultraviolet and ionizing radiation, chemical and biological carcinogens, as well as diet [2]. One key enzyme associated with the dietary influences on breast cancer development is methylenetetrahydrofolate reductase (MTHFR) [3]. MTHFR, which is encoded by the *mthfr* gene on Chromosome 1, plays an important role for the reduction of 5,10-methylenetetrahydrofolate to 5-methyltetrahydrofolate, a key step in the one-carbon metabolism pathway. 5-methyltetrahydrofolate is utilized by homocysteine to make methionine. Afterwards, methionine transfers a methyl group to SAM, which is eventually used for DNA methylation, promoting the expression of certain oncogenes and suppression of tumor suppressor genes [4, 5]. Thus, MTHFR provides a key link between folate consumption and DNA methylation, which is associated with cancer development. However, the exact mechanism remains unclear. Some have posited that, during cell division, low folate availability results in increased production and incorporation of uracil into the DNA, due to the limited availability of thymidine [5]. In support of this, Gao et al. and Shrubsole et al. both demonstrated that low folate consumption was associated with an increased risk of developing breast cancer [6, 7]. For a comprehensive review about MTHFR and cancer, please refer Crider et al. 2012 [5].

Two single nucleotide polymorphisms have been shown to effect MTHFR enzymatic activity. The first is the C677T polymorphism (Ala222Val, rs1801133), located in exon 4. This polymorphism was shown to decrease enzymatic activity by 75% for the homozygote genotype [8]. The other polymorphism is A1298C (Glu429Ala, rs1801131), located in exon 7. This polymorphism was also shown to decrease enzymatic activity but to a lesser degree [8]. Carriers of these MTHFR polymorphisms were shown to be associated with breast cancer risk. When considering the C677T genotypes, compared to high folate consumption with CC genotype, Gao et al. and Shrubsole et al. both demonstrate that the risk was significantly augmented for the CC, CT and TT genotypes with low folate consumption [6, 7]. Interestingly, these polymorphisms appear to have an ethnic dependency. For example, for the Asian population, the C677T polymorphism was associated with an increased risk for developing breast cancer [9, 10]. However, the results for the A1298C polymorphism are conflicting, with one study determining a benefit for the heterozygous genetic model [11] and others demonstrating no effect at all [12, 13]. Whereas, for Caucasians, the A1298C polymorphism was associated with an increased risk of developing breast cancer [12] and not the C677T polymorphism [9, 10, 13, 14]. Some meta-analyses have tried to elucidate the effect, either determining detrimental effects or no effect at all, with no specific analysis aimed towards Latinos. Nonetheless, there is no definitive assessment among the Latin American population.

Many Latin American countries have examined the risk associated with these MTHFR polymorphisms and breast cancer; however, conflicting results were obtained [15–18]. With so many studies focusing on the C677T and the A1298C polymorphisms and breast cancer susceptibility in Latinos yielding no concise result, we undertook this meta-analysis to determine the effect these MTHFR polymorphisms have on breast cancer risk for Latinos.

Materials and methods

Publication search strategy

PubMed, EBSCO, LILACS, and Scopus databases were searched for published studies that examined the association between the MTHFR polymorphisms and breast cancer susceptibility in Latinos until January 21, 2019. The search strategy included keywords and any of their derivation for MTHFR (methylenetetrahydrofolate reductase, *MTHFR*), polymorphism (polymorphism, SNP, mutation, C677T, rs1801133, A1298C, rs1801131), cancer (cancer, cáncer, carcinoma), and Latin Americans (Latino, Hispanic, etc.) as well for Latin American countries. Furthermore, we manually reviewed the reference lists for more relevant studies. We increased our coverage by also including other Latino databases (DOAJ, Periodica, Bibliat, Latin index, Imbiomed). No language limitation was applied in the literature search.

Inclusion and Exclusion Criteria

Two authors independently examined each publication for inclusion. If the two authors disagreed about a publication, a third author assessed the publication and made a final decision. The inclusion criteria were: (1) evaluated the association between the C677T and/or the A1298C polymorphisms and breast cancer susceptibility; (2) case-control studies; (3) genotype data were available for cases and controls; (4) data based on the Latino population. The exclusion criteria were: (1) no detailed information of genotype data; (2) duplicate of a previously published study; (3) no Latino population or was mixed with another ethnicity; (4) cancer was not confirmed by histological and laboratorial methods. When duplicate data appeared in different publications, this meta-analysis only adopted the most recent study or the study with the most complete information.

Data extraction

To guarantee the accuracy of the extracted information, two authors individually reviewed each publication. The following data were collected: author's name, year of publication, country, polymorphism(s), type of breast cancer, source of the controls (hospital or population based), number of cases and controls, and genotype frequency. Two authors independently assessed the quality of the studies using the Newcastle–Ottawa Quality Assessment Scale [19]. The following aspects of each study were appraised: selection of cases and controls, comparability, and outcome or exposure. For analysis, the quality scores ranged from 0 to 9 (Electronic supplementary material 1 and Electronic supplementary material 2). Studies that scored more than six stars were considered as a high-quality study.

Statistical analysis

The relationship between the C677T and the A1289C polymorphisms and breast cancer susceptibility was analyzed using five genetic models: the dominant model (12+22 v. 11), the recessive model (22 v. 12+11), the homozygous model (22 v. 11), the heterozygous model (12 v. 11), and the allele model (2 v. 1), where 1 is the wild-type and 2 is the mutation. Hardy–Weinberg Equilibrium (HWE) was tested by the ψ^2 -test for the control groups of each study. The level of heterogeneity was determined by the Cochran O test and the Inconsistency index (I^2) . Crude odds ratios (ORs) with 95% confidence intervals (95% CIs) were used to evaluate the strength of association between the C677T and the A1298C polymorphisms and breast cancer susceptibility. Based on the level of heterogeneity, either the Fixed-effect model (Mantel-Haenszel method, p < 0.10 and $I^2 < 40\%$ [20] or the Random-effects model (Dersimonian-Laird method) was selected [21]. To examined publication bias, Egger's test, Begg-Mazumdar's test, and examining the funnel plots were performed. Sensitivity analysis was conducted by removing one study at a time and recalculating the pooled ORs to observe the stability of the crude pooled OR. All the statistical analyses were conducted using Review Manager Software (RevMan V5.3, Copenhagen, Denmark) and StatsDirect Statistical Software (StatsDirect V2.8, Cheshire, United Kingdom). Unless noted otherwise, p values < 0.05 (two-tailed) were considered statistically significant. This study was conducted in accordance with the PRISMA guidelines, see Electronic supplementary material 7.

Results

Selection of eligible studies

A total of 280 publications were retrieved from searching the multiple databases and from reviewing the publications' bibliographies (Fig. 1). 252 publications were exclude because they were not an original research articles, focused on animals or cell lines, did not focus on breast cancer, or did not examine for MTHFR or its polymorphisms. The remaining 28 publications were extensively evaluated. 16 publications did not focus solely on the Latino population, 1 publication lacked sufficient information, and 2 publications used previously published data. This resulted in 9 publications that were included in this meta-analysis: 9 studies for the C677T polymorphism (case = 2136 and controls = 2436) [15–18, 22–26] and 5 studies for the A1298C polymorphism (cases = 1226 and controls = 1739) [22–26]. Most studies determine the presence of either polymorphism using the restriction fragment length polymorphism method (RFLP; n=6), with the remaining three using other technologies (Table 1). We determined that the Calderon-Garcidueñas study [16] contained a high level of bias (Electronic supplementary material 1 and Electronic supplementary material 2). The genotype distributions of the controls were



Fig. 1 Flow chart for literature review of studies to be included in the meta-analysis

Table 1 Characteristics of inclu	ded studies										
References	Country	Type of breast cancer ^a	SNP	Group	11^{b}	12 ^b	22 ^b	HWE ^c	Source ⁶	Score ^e	Genotyping
Batschauer et al. [15]	Brazil	62 invasive ductal, 6 in situ (<i>Post-biopsy immunohistological diagnosis</i>)	C677T	Control	42 77	34	6	0.592	PB	6	RFLP
Calderón-Garcidueñas et al. [16] Mexico	Neoplasia in stage IIB y IIIB (<i>Histological diagnosis</i>)	C677T	Control Cases	37	t 9 <u>1</u>	21	0.695	HB	* 9	Microarrangment PHARMA Chip®
Cavalho Barbosa et al. [22]	Brazil	164 invasive ductal, 8 invasive lobulillar, 5 Medullary, 3 Apocirine, 2 Mucinous, 1	C677T	Control	87 76	70	19	0.388	PB	8	RFLP
		papiullary (<i>Pathological diagnosis</i>)	A1298C	Control Cases	72 88	8 8	6 1	0.014*			RFLP
Le Marchand et al. [23]	NSA	Invasive breast cáncer (Registry reports diagnosis)	C677T	Control Cases	218 70	319 117	127 49	0.593	PB	٢	Taqman PCR Core Reagent kit
			A1298C	Control Cases	423 146	212 77	29 13	0.709			
Lopez-Cortes et al. [24]	Ecuador	55 Luminal A, 26 Luminal B, 14 HER positive, 19 Triple negative (<i>Case history</i>	C677T	Control Cases	69 25	113 75	13 14	< 0.001*	PB	L	RFLP
		diagnosis)	A1289C	Control Cases	191 110	ς β		< 0.001*			RFLP
Ma et al. [25]	Brazil	458 invasive cancers (Clinic-histological diagnosis)	C677T A1298C	Control Cases Control	222 225 279	187 188 157	49 45 22	0.309 0.988	HB	×	TaqMan [®] SNP Genotyping Assays
Ramos-Silva et al. [17]	Mexico	438 invasive ductal, 53 invasive lobililar, 6	C677T	Cases Control	269 167	168 140	21 32	0.735	PB	L	RFLP
Rezende et al. [26]	Brazil	mixed (Clinical-histological diagnosis) 174 Invasive ductal, 79 mixed (Histological diagnosis)	C677T	Cases Control Cases	216 111 120	178 119	103 27 30	0.555	PB	8	RFLP
			A1298C	Control Cases	133 135	112	12 14	0.055			RFLP
Zara-Lopes et al. [18]	Brazil	Not indicated (Imaging studies, histopatho- logical analysis, biopsies)	C677T	Control Cases	66 35	65 48	13	0.596	HB	×	RFLP
<i>RFLP</i> restriction fragment lengt ^a The type of breast cancer is rep ^b 11, 12, and 22 indicate the free polymorphism, and 2 is the T-al	th polymorp sorted accor quency of the lele for the	bhism ding to the articles. The methods used to determ he wild-type, heterozygote, and homozygote mu C677T polymorphism and C-allele for the A125	uine the tyr atant, respo 8C polym	pe of breas ectively, w orphism	st canc /here 1	ers ar	e indi e C-a	cated in it llele for th	alics ie C677	[polymo	rphism and A-allele for the A1298C

٢)

 $^{\circ}$ Hardy–Weinberg equilibrium were calculated by ourselves. A p value <0.05 was considered not in agreement with HWE and indicated with *

^dThe source of the controls: *PB* population base, *HB* Hospital base, and *N/I* Not indicated

^eThe bias score was calculated using the Newcastle-Ottawa scale. Values <7 were considered to have a high-level of bias and indicated with *

consistent with HWE, except for the Batschahuer study [15] and the Lopez-Cortes study [24]. The publication years of the involved studies ranged from 2001 to 2017. The characteristics of the included studies are shown in Table 1. For the C677T polymorphism, four different countries are represented in this meta-analysis, with the most representative population being Brazil with 5 studies [15, 18, 22, 23, 25], followed by Mexico with 2 studies [16, 17], 1 study from USA, which focused on the Latino population [23], and 1 study from Ecuador [24]. For the A1289C polymorphism, the most representative country was Brazil with 3 studies [22, 25, 26], followed by USA [23] and Ecuador [24] with 1 study each.

MTHFR C677T polymorphism increased the risk of breast cancer for Latinos

With respect to the C677T polymorphism, the homozygous, recessive, and allelic genetic models presented with significant heterogeneity (p < 0.10, $I^2 > 30\%$). Using the Random Effects model for these genetic models, we determined that for the homozygous (OR 1.42, 95% CI 1.05–1.92, p = 0.02), the recessive (OR 1.33, 95% CI 1.01–1.85, p = 0.04), and the allelic models (OR 1.17, 95% CI 1.03–1.33, p = 0.01, Fig. 2), the polymorphism was significantly associated with an increased risk of developing breast cancer. With the remaining models, using the fixed effects model, we determined that the dominant model was associated with an increased risk of developing breast cancer (OR 1.16, 95% CI 1.02–1.31, p = 0.02), whereas the heterozygous model did not show any association (OR 1.09, 95% CI 0.96-1.24, p = 0.20). However, the results are similar independent of the effects models used.

When the sensitivity of the results was determined by removing one study and recalculating the ORs, the heterozygous and allelic models were resistant (Electronic supplementary material 3). However, for the homozygous model, the results were sensitive to the Lopez-Cortes study (OR 1.34, 95% CI 0.99-1.80), the Ramos-Silva study (OR 1.24, 95% CI 0.96-1.60), and the Zara-Lopes study (OR 1.35, 95% CI 0.99-1.85). For the dominant model, the results were sensitive to the Lopez-Cortes study (OR 1.12, 95%) CI 0.99-1.27) and the Ramos-Silva study (OR 1.13, 95% CI 0.99-1.30). Lastly, for the recessive model, the results were sensitive to the Calderon-Garcidueñas study (OR 1.33, 95% CI 0.97–1.82), the Le Marchand study (OR 1.37, 95% CI 0.99-1.90), the Lopez-Cortes study (OR 1.28, 95% CI 0.96-1.72), the Ramos-Silva study (OR 1.15, 95% CI 0.94–1.40), the Rezende study (OR 1.35, 95% CI 0.99–1.85), and the Zara-Lopes study (OR 1.28, 95% CI 0.96-1.70).

Publication bias was assessed by examining the funnel plot for each of the genetic model. The funnel plots demonstrated no significant asymmetry and the shape of the funnel plots suggested no evidence of publication bias (Electronic supplementary material 4). However, a significant publication bias was determined by the Begg–Mazumdar's test and Egger's test only for the heterozygous model (Fig. 2). No correlation was determined by the Begg–Mazumdar's test or bias by Egger's test for the remaining models.

MTHFR A1289C polymorphism not increased the risk of breast cancer for Latinos

With respect to the A1298C polymorphism, none of the genetic models presented with significant heterogeneity (p < 0.10, $l^2 0\%$). Using either the fixed-effects or random-effect models, no association between the A1298C polymorphism and the risk to develop breast cancer was observed (Fig. 3). Removing one study had no effect on pooled ORs for any of the genetic models (Electronic supplementary material 5). When publication bias was assessed, no significant asymmetry was determined by examining the funnel plots (Electronic supplementary material 6); furthermore, this was confirmed by the Begg–Mazumdar's test and Egg-er's test (Fig. 3).

Discussion

Folate deficiency can influence the genetic stability of DNA, which increases the risk of cancer development [4, 27]. Moreover, dysregulation of this pathway is associated with many pathological outcomes, ranging from atherothrombosis to the risk of ischemic stroke to various forms of cancer [14, 28]. Although numerous case–control studies and meta-analyses have investigated the association between the C677T and the A1298C polymorphisms and the risk for developing breast cancer, there is no definite conclusion or a sole focus for the Latin American population. Here, we determined that the C677T polymorphism was associated with an increased risk of developing breast cancer, whereas the A1298C polymorphism had no effect.

For the C677T polymorphism, there was a significant presence of heterogeneity among the studies included in this meta-analysis. The diagnosis of breast cancer was confirmed by histology or by biochemical/molecular techniques (Table 1). Thus, it is possible that the level of heterogeneity is due to the different forms and proportions of breast cancer cases. The method used to determine the polymorphism could increase the heterogeneity of the sample; however, most studies used a similar technique—RFLP. Another cause of heterogeneity could be due to the populations themselves. For example, in Mexico the genetic composition of Mexicans—which is composed of Native American, Europeans, and Africans—varies from region to region [29], this is also seen among different regions of Brazil [30]. These

Fig. 2 Forest plots of breast cancer risk associated with the MTHFR C677T polymorphism for the heterozygous (a), homozygous (b), dominant (c), recessive (**d**), and allelic (**e**) genetic models. The squares and horizontal lines correspond to the study-specific odds ratio (OR) and 95% confidence interval (95% CI), respectively. The area of the squares reflects the study-specific weight. The diamond represents the pooled OR and 95% CI, determined using either the fixed effects or random effects, depending on the level of heterogeneity. Plots were generated using Review Manager Software (RevMan V5.3, Copenhagen, Denmark)

-		Expe	rimental	Co	ntrol		Odds Ratio				
а	Study	Events	Total	Events	Tota	l Weight	M-H, Fixed, 95%	CI			
	Batschauer 2011	34	61	34	76	3.1%	1.56 [0.79, 3.06]]		+	
	Calderon-Garcidueñas 2017	113	183	60	97	6.9%	1.00 [0.60, 1.65]			Begg= 0.04
	Carvalho Barbosa 2012	83	159	/0	15/	7.8%	1.36 [0.87, 2.11]]		T .	Egger= 0.04
sno	Le Marchano 2004	75	107	113	182	14.3%	1.14 [0.01, 1.01]] 1			55
, Ś	Ma 2009	188	413	187	409	23.7%	0.99 [0.75, 1.31]	1		+	
ZOZ	Ramos-Silva 2015	178	394	140	307	20.0%	0.98 [0.73, 1.33]	1		+	
etel	Rezende 2017	103	223	119	230	14.6%	0.80 [0.55, 1.16]]			
Ť	Zara-Lopes 2016 Breast	48	83	65	131	4.9%	1.39 [0.80, 2.42]]		+	
	Total (95% CI)	020	1803	1107	2126	100.0%	1.09 [0.96, 1.24]				
	Heterogeneity: Chi ² = 10.05 d	939 f = 8 (P =	0.26) · 12 ;	= 20%						T	
	Test for overall effect: Z = 1.27	7 (P = 0.2	0.20), I D)	20,0				+			
		Evnorin	ontol	Cont	rol		Odda Batia	0.02	0.1	1	10 50
b	Study or Subaroup	Experin	Total F	vents	Total N	Weight	M-H. Random, 95	5% CI			
	Batschauer 2011	7	34	9	51	5.6%	1.21 [0.40, 3.63	3]		<u> </u>	
	Calderon-Garcidueñas 2017	51	121	21	58	11.0%	1.28 [0.67, 2.45	5]		+	
	Carvalho Barbosa 2012	17	93	19	106	9.8%	1.02 [0.50, 2.11	I]			
sno	Le Marchand 2004	49	119	127	345	15.5%	1.20 [0.79, 1.84	1]			Begg= 0.35
Ŋ	Lopez-Cortes 2015	14	39	13	82	7.6%	2.97 [1.23, 7.19	9]			Egger= 0.61
ZO	Ma 2009 Damaa Cilua 2015	45	2/0	49	2/1	15.0%	0.91 [0.58, 1.41] 21			
Hom	Ramos-Silva 2015 Rezende 2017	30	150	32 27	138	12.0%	2.49 [1.59, 3.60	0] 11			
	Zara-Lopes 2016 Breast	17	52	13	79	8.3%	2.47 [1.07, 5.66	5]			
	Total (95% CI)		1197		1329	100.0%	1.42 [1.05, 1.92	i i			
	Total events	333		310						•	
	Heterogeneity: Tau ² = 0.11; Chi ²	= 17.11,	df = 8 (P =	= 0.03); l	² = 53%	0					
	Test for overall effect: $Z = 2.30$ (F	² = 0.02)						0.01	0.1	1	10 100
		Exper	imental	Cor	ntrol		Odds Ratio				
С	Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95%	CI			
	Batschauer 2011	41	68	43	85	3.2%	1.48 [0.78, 2.83]			+	
	Calderon-Garcidueñas 2017	164	234	81	118	6.8%	1.07 [0.66, 1.73]			T	Begg= 0.11
ant	Carvaino Barbosa 2012	100	1/0	89	1/0	8.2%	1.29 [0.85, 1.96]			L.	Egger= 0.10
ir	Le marchand 2004	89	114	126	195	4.3%	1.10 [0.04, 1.00]				
οŭ	Ma 2009	233	458	236	458	24.6%	0.97 [0.75, 1.26]			+	
	Ramos-Silva 2015	281	497	172	339	18.9%	1.26 [0.96, 1.67]				
	Rezende 2017	133	253	146	257	14.6%	0.84 [0.59, 1.19]				
	Zara-Lopes 2016 Breast	65	100	78	144	4.7%	1.57 [0.93, 2.66]				
	Total (95% CI)	4070	2136	4447	2436	100.0%	1.16 [1.02, 1.31]				
	Heterogeneity: Chi ² = 11 16 df	12/2 = 8 (P = () 10)· l² =	1417						ľ	
	Test for overall effect: Z = 2.30	(P = 0.02)	2070				<u> </u>	<u>+</u>		+
		Fynerim	ntal	Contro			Odds Ratio	0.01	0.1	1	10 100
d	Study or Subaroup Ev	vents 1	fotal Ev	ents To	n tal W	eiaht	M-H. Random. 95%	CI			
-	Batschauer 2011	7	68	9	85	5.3%	0.97 [0.34, 2.75]		-		
	Calderon-Garcidueñas 2017	51	234	21 1	18 1	1.6%	1.29 [0.73, 2.26]			+	D
ive	Carvalho Barbosa 2012	17	176	19 1	76	9.4%	0.88 [0.44, 1.76]			-	Begg= 0.76 Egger= 0.90
ess	Le Marchand 2004	49	236	127 6	64 1	6.2%	1.11 [0.77, 1.60]			-	Lggei - 0.50
Šec	Lopez-Cortes 2015	14	114	13 1	95	7.9%	1.96 [0.89, 4.33]				
-	Ma 2009 Ramos-Silva 2015	40 103	400	49 4	1 001	4.1%	2.51 [1.64, 3.83]]	
	Rezende 2017	30	253	27 2	257 1	4.0 <i>%</i> 1.9%	1.15 [0.66, 1.99]			-	
	Zara-Lopes 2016 Breast	17	100	13 1	44	8.2%	2.06 [0.95, 4.47]				
	Total (95% CI)	2	136	24	36 10	0.0%	1.33 [1.01, 1.75]			•	
	Lotal events	333	- 8 (D - 4	310 0.03\-12-	- 520/			+			-++
	Test for overall effect: 7 = 2 02 (P	= 0.04)	= 0 (P = 1	0.03); 1	- 52%			0.02	0.1	1	10 50
		Exnerime	ontal	Contro			Odds Ratio				
е	Study or Subgroup Ev	vents 1	fotal Ev	ents To	otal W	eight	M-H, Random, 95%	CI			
-	Batschauer 2011	48	136	52 1	70	5.5%	1.24 [0.77, 2.00]				
	Calderon-Garcidueñas 2017	215	468	102 2	236	9.9%	1.12 [0.81, 1.53]				
	Carvalho Barbosa 2012	117	352	108 3	352	9.8%	1.12 [0.82, 1.54]			-	Begg= 0.18
	Le Marchand 2004	215	472	573 13	328 1	4.8%	1.10 [0.89, 1.36]			+	Eager= 0.44
elic	Lopez-Cortes 2015	103	228	139 3	390	9.2%	1.49 [1.07, 2.08]				
All	Ma 2009 Romon Silva 2015	2/8	916	285 9	16 1	5.6% 5.0%	0.96 [0.79, 1.18]			T+	
	ramus-oliva 2015 Rezende 2017	304 163	394 506	204 0	070 1 31/ 1	0.0% 2.2%	1.40 [1.19, 1.80]			1	
	Zara-Lopes 2016 Breast	82	200	91 2	288	7.9%	1.50 [1.03 2.19]			_	
	Total (95% CI)	4	272	48	72 10	0.0%	1.17 [1.03, 1.33]				
	Total events	1605	1	1727			• • •			•	
	Heterogeneity: Tau ² = 0.02; Chi ² =	15.03, df	= 8 (P = 0	0.06); l² :	= 47%						
	rest for overall effect: Z = 2.48 (P =	- U.U1)						0.02	0.1	1	10 50

differences do lead to various development rates and pathologies of similar diseases.

Other meta-analyses have determined that for Asians, the C677T polymorphism was shown to augment breast cancer risk [9, 10]. Due to the high level of influence of the European and African immigration, it was not expected that the C677T polymorphism would also increase the risk of developing breast cancer in Latin Americans. Interestingly, the C677T polymorphism was associated with an increased risk of developing breast cancer, except for the heterozygous model. The proposed oncogenic mechanism of the C677T polymorphism is highly correlated with low folate consumption [5]. Many Asian countries and Latin American countries are plagued with low folate consumption [31, 32], thus providing a link between the observed results here and the results for the Asian population. This does posit that folate consumption could attenuate the effects of the C677T polymorphism.

The A1298C polymorphism was not associated with an increased risk of developing breast cancer. Due to different degrees of potency the C677T and the A1298C polymorphisms have on MTHFR enzymatic activity, this does suggests that for breast cancer, the A1298C polymorphism does not inhibit MTHFR activity sufficiently enough to switch to an oncogenic promoting condition, as seen with the C677T polymorphism. Interestingly, others meta-analyses have determined a significant association between the risk of developing breast cancer and the A1298C polymorphism. Intriguingly, this effect is typically seen with the Caucasians [11, 12]. As mention above, Asians have minimal influence on the genetic make-up of Latinos, as do Europeans. Moreover, one study found that the A1298C polymorphism was associated with a decreased risk of developing breast cancer among the Asian population [11]. Taken together, these results suggest that the Latin American population represents a unique group, whose dietary intake must be taken into consideration.

Folate's effect on DNA methylation is dependent on the level of serum folate [27], which comes from its dietary intake. Many women and children from Latin American countries suffer from low folate consumption and other members of the vitamin B family [32–34]. In Cuba, 80% of adults suffer from folate deficiency, whereas in Mexico, Venezuela, and Brazil, the prevalence of fotale deficiency is > 30%. Interestingly, using folate fortification procedures have reduced folate deficiency in selected populations to <5% [33]. Nevertheless, the availability and compliance of these programs to the majority of the population, or food supplemented with folate, remain inadequate. Therefore, it is possible that for Latinos, which typically consume lower quantities of folate [30, 32, 34] their diets do not provide sufficient amount of folate to inhibit an oncogenic state when MTHFR activity is compromised. This does posit that augmented folate consumption could alter the risk of developing breast cancer.

Hormone changes that precede menopause have been as associated with decreased folic acid absorption, affecting DNA methylation [35]. Therefore, MTHFR polymorphisms are proposed to be a more heightening factor associated with breast cancer development during post-menopause. Indeed, with post-menopausal women, the C677T polymorphism does increase the risk of developing breast cancer [36, 37]; nonetheless, many reports have demonstrated that pre-menopausal status is associated with an increased risk [38, 39]. To add to the confusion, some reports have shown no effect [40]. To address this, Kumar and colleagues performed a meta-analysis for the C677T polymorphism. They determined that there were no associations with breast cancer and the C677T polymorphism for pre-menopausal women (9 studies) and post-menopausal women (9 studies) [10]. Two years later, Naushad and colleagues examined the affect menopausal status has on the association between the C677T polymorphism and breast cancer development [36]. With their meta-analysis, they determined that for post-menopausal women (11 studies), the C677T polymorphism was associated with breast cancer development and not for premenopausal women (8 studies). However, for both studies, there was no sub-analysis by ethnicity. For Latin Americans, this effect remains elusive. Here, four studies did examine if menopausal status affects MTHFR polymorphism's effect on breast cancer development. For the Le Marchand study, the analysis was mixed with other ethnicities; thus, no clear conclusion can be made about Latin Americans [23]. For the Batschauer study, they did examine for this effect, determining that the C677T polymorphism had no effect [15]. Interestingly, only two studies, both from Brazil, provided sufficient information for an analysis-the Ma study and the Carvalho Barbosa study. For the Ma study, independent of the menopausal status, neither polymorphism affected the risk [25]. However, the Carvalho Barbosa study indicates that only for post-menopausal women, the CT genotype of C677T is associated with an increase risk [22]. For pre-menopausal women, no associations were found. With a lack of coverage of Latin America and few studies examining this affect, more studies are required.

Our study has a few limitations. First, only five countries are represented in this meta-analysis, which does suggest that parts of the Latin American community are underrepresented. Second, we calculate the crude ORs and they are unadjusted estimations. Third, we did not perform a subanalysis on the types of breast cancer. As indicated by the elevated level of heterogeneity, the different forms could be possibly affecting the associations. However, even with the crude pooled ORs, we did demonstrate the association between the C677T polymorphism and breast cancer for Latinos.



◄Fig. 3 Forest plots of breast cancer risk associated with the MTHFR A1298C polymorphism for the heterozygous (a), homozygous (b), dominant (c), recessive (d), and allelic (e) genetic models. The squares and horizontal lines correspond to the study-specific odds ratio (OR) and 95% confidence interval (95% CI), respectively. The area of the squares reflects the study-specific weight. The diamond represents the pooled OR and 95% CI, determined using either the fixed effects or random effects, depending on the level of heterogeneity. Plots were generated using Review Manager Software (RevMan V5.3, Copenhagen, Denmark)

Conclusion

Overall, the results indicated that, for Latinos, the C677T polymorphism is associated with a significant risk for developing breast cancer. On the other hand, the A1289C polymorphism does not show an association. In Latin American women with a history of breast cancer, examining for the C677T polymorphism could identify at risk subjects, who should consider folate monitoring and supplementation management.

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Compliance with ethical standards

Conflict of interest The authors declare that they have no conflict of interest.

Ethical approval This article does not contain any studies with human participants performed by any of the authors.

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