

Methylenetetrahydrofolate Reductase (MTHFR) C677T Polymorphism and Alzheimer Disease Risk: a Meta-Analysis

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Abstract Methylenetetrahydrofolate reductase (MTHFR) is key enzyme of folate/homocysteine pathway. Case control association studies on MTHFR C677T polymorphism and Alzheimer's disease (AD) have been repeatedly performed over the last two decades, but the results are inconclusive. The aim of the present study was to assess the risk of MTHFR C677T polymorphism for AD. Forty-one studies were identified by a search of PubMed, Google Scholar, Elsevier, and Springer Link databases, up to January 2015. Odds ratios (ORs) with corresponding 95 % confidence interval (CI) were calculated using fixed effect model or random effect model. The subgroup analyses based on ethnicity were performed. MTHFR C677T polymorphism had a significant association with susceptibility to AD in all genetic models (for T vs C OR = 1.29, 95 % CI = 1.07–1.56, p = 0.003; for TT + CT vs CC OR = 1.29, 95 % CI = 1.19–1.40, p = 0.0004; for TT vs CC OR = 1.31, 95 % CI = 1.16–1.48, p = 0.001; for CT vs CC OR = 1.24, 95 % CI = 1.13–1.35, p < 0.004; and for TT vs CT + CC OR = 1.13, 95 % CI = 1.00–1.28, p = 0.02). Results of present meta-analysis supported that the MTHFR C677T polymorphism was associated with an increased risk of AD.

Keywords Alzheimer's disease · MTHFR · C677T · Neurodegeneration · Homocysteine · Folate

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Introduction

Alzheimer's disease (AD) is a neurodegenerative disorder, contributing to about two thirds of all dementias in the elderly population [1]. It causes progressive memory loss in mid-to-late adult life. Some individuals inherit this form of dementia before the age of 65 (known as early-onset or familial AD); but most often, AD occurs late in life [2]. It is a heterogeneous disorder with both familial (about 1 % of cases) and sporadic forms, results from selective damage of specific neuronal circuits in the neocortex, hippocampus, and basal forebrain cholinergic system. Affected regions show senile plaques, comprised of neurites displayed around extracellular deposits of beta-amyloid peptides, and many neurons develop neurofibrillary tangles, which reflect the local accumulation of abnormal intracytoplasmic filaments, composed of hyperphosphorylated isoforms of the tau protein [3]. AD is a multifactorial pathology resulting of the interaction of both genetics and environmental factors. Over 100 rare, highly penetrant mutations have been described in three genes (amyloid beta precursor protein, presnilin 1, and presnilin 2) for earlyonset familial AD [4], and for late-onset AD, only the association with the apolipoprotein E (APOE) gene has been convincingly replicated. Several clinical and epidemiological studies show a relation between vascular disorders and late-onset AD [1, 5]. Vascular risk factors such as diabetes, hyperlipidemia, hypertension, heart diseases, and high serum homocysteine are also reported risk factors for AD [6, 7].

Elevated levels of homocysteine (Hcy) have been linked to AD [5, 8]. Folate is a cofactor in one-carbon metabolism, during which it promotes the remethylation of homocysteine. Numerous epidemiological and experimental studies have linked folate deficiency and resultant increased homocysteine levels with AD [5]. Methylenetetrahydrofolate reductase (MTHFR) is an important enzyme involved in the folate-dependent metabolism of homocysteine.

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MTHFR enzyme mediates the irreversible conversion of 5, 10-methylenetetrahydrofolate (5,10-MTHF) to 5-methyltetrahydrofolate (5-MTHF) [9, 10]. Several polymorphisms in the MTHFR gene have been identified, out of which, the most studied and clinically important polymorphism is C677T in exon 4, resulted in substitution of alanine amino acid by valine at position 222 in MTHFR protein [11]. MTHFR functions in dimeric form and flavin adenine dinucleotide (FAD) work as a co-factor, but variant MTHFR (222 valine) dissociates into monomers and its enzymatic activity reduces [12]. This substitution makes enzyme thermolabile with reduced enzymatic activity. The mean activity of TT variant enzyme was 40–50 % that of the CC variant enzyme at 37 °C [13].

The frequency of MTHFR C677T mutation varies among racial and ethnic groups of the world. T allele frequency ranges from 0.20 to 0.55 in Europeans, 0.11 to 0.35 in Americans, 0.063 to 0.094 in Africans, from 0.04 to 0.38 in Asian population, and 0.10 to 0.47 in Australians [14–18]. Several studies showed that the mutant T allele increases homocysteine levels particularly in folate deficiency state [11, 13, 19, 20]. Impact of MTHFR C677T polymorphism on development and pathogenesis of AD have been conflicting and inconclusive [21–28]. Hence, meta-analyses of all published case control studies investigating C677T polymorphism as risk factor for AD were carried out to shed some lights on conclusive role of MTHFR C677T polymorphism in AD.

Methods

Meta-analyses of published case control articles were carried out according to MOOSE guidelines [29].

Literature and Search Strategy

Electronic databases Pubmed, Springer Link, Elsevier, and Google Scholar were searched up to January 2015 for suitable articles using keywords "MTHFR and Alzheimer's disease," "folate metabolism and Alzheimers disease," and "MTHFR C677T polymorphism and Alzheimers disease." The references from the eligible articles were also reviewed to find other potential articles. If more than one study by the same author using the same case series was published, either the studies with the largest sample size or the most recently published study was included.

Inclusion and Exclusion Criteria

The following inclusion criteria were set for the meta-analysis: (i) each study should be an independent case-control study, (ii) the purpose of all the studies and statistical methods should be similar, (iii) studies should reported enough information to



Fig. 1 Flow diagram of selection of studies included in the meta-analysis

calculate the odds ratio with 95 % confidence interval (CI), and (iv) inclusion of the patients should be done according to the standard diagnosis parameter (Diagnostic and Statistical Manual of Mental Disorders (DSM-IV) and Mini Mental State Examination (MMSE)). The exclusion criteria were: (i) only case studied; (ii) review papers, editorial, and letter to editor; (iii) containing overlapping data; and (iv) study not providing enough information to estimate odds ratio (OR) with 95 % CI (incomplete raw data) and not well described.

Data Extraction

The following information was extracted from each study: the first author's family name, year of publication, country, race, sample size, outcome, characteristics of controls, case and control diagnostic criteria, genotyping method, and genotype distribution in cases and controls. Detailed information, wherever not available, was collected by contacting authors.

Statistical Analysis

The strength of association between the MTHFR C677T polymorphisms and AD risk was evaluated by OR and 95 % CI according to allele contrast (T vs C), homozygote (TT vs CC), heterozygote/co-dominant (CT vs CC), recessive (TT vs TC + CC), and dominant (TT + CT vs CC) models. A chi squarebased "Q" test defined by Cochran was used to assess the heterogeneity (between study variability) in the metaanalysis [30]. Since the Q statistics is only useful for testing the existence of heterogeneity qualitatively but not quantitatively, another index " I^2 ," calculated as the percentage of the

Table 1	Details of studies	included in the	present meta-analysis
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Study	Country	Case samples	Control samples	P value of HWE	Year	Reference
Chapman et al.	Israel	49	40	0.25	1998	Stroke 29: 1401–1404
Nishiyama et al.	Japan	24	33	0.84	2001	J Epidemiol 10: 163–172
Pollak et al.	Israel	92	82	0.50	2001	J Am Geriatr Soc 48: 664–668
Brunelli et al.	Italy	231	137	0.56	2001	Neurosci Lett 315: 103-105
Postiglione et al.	Italy	74	74	0.99	2001	Gerontology 47: 324-329
Prince et al.	Sweden	204	172	0.23	2001	Eur J Hum Genet 9: 437–444
Zuliani et al.	Italy	40	54	0.63	2001	Acta Neurol Scand 103: 304-308
MacIlroy et al.	Ireland	83	71	0.90	2002	Stroke 33: 2351–2356
Wakutani et al.	Japan	241	352	0.76	2002	Ann N Y Acad Sci 977: 232–238
Religa et al.	Poland	99	100	0.90	2003	Dement Geriatr Cogn Disord 16: 64-70
Seripa et al.	USA	124	97	0.78	2003	Neurobiol Aging 24: 933–939
Seripa et al.	Italy	126	106	0.68	2003	Neurobiol Aging 24: 933–939
Anello et al.	Italy	180	181	0.98	2004	Neuroreport 15: 859-861
Bi et al.	China	42	40	$0.045^{\rm a}$	2004	Chin J Mod Med 977: 15–18
Jiang et al.	China	75	72	0.81	2004	Shanghai Arch Psyc 16: 196–208
Kida et al.	Japan	194	379	0.06	2004	Psychogeriatrics 4: 4-10
Liao et al.	China	66	143	0.19	2004	Zhong guo You Sheng Yi Yi Chuan 12: 13-15
Linnebank et al.	Germany	162	169	0.23	2004	Am J Med Genet A 131: 101–102
Wang et al.	China	127	138	0.38	2004	Chin J Geriatr 23: 460–463
Fernandez et al.	Brazil	30	29	0.65	2005	Arq Neuropsiquiatr 63: 1-6
Wang et al.	China	104	130	0.34	2005	J Mol Neurosci 27: 23–27
Zhang et al.	China	105	102	0.85	2005	Chin Med Sci J 20: 247–251
da Silva et al.	Brazil	43	50	0.24	2006	Arq Neuropsiquiatr 64: 941–945
Keikhace et al.	Iran	117	125	0.07	2006	Neurochem Res 31: 1079–1083
Wehr et al.	Poland	99	141	0.36	2006	Dement Geriatr Cogn Disord 22: 1-7
Dorszewska et al.	Poland	38	50	0.49	2007	Acta Neurobiol Exp (Wars) 67: 113-129
Kim et al.	Korea	86	625	0.08	2007	Int J Geriatr Psychiatry 23: 454–459
Yuan et al.	China	30	80	0.80	2007	Chin J Geriatr 26: 767–769
Zhang	China	68	68	0.45	2007	Chin Med Sci. J 20: 1–28
Styczyaska et al.	Poland	154	176	0.92	2008	Folia Neuropathol 46: 249–254
Zhang et al.	China	43	40	0.08	2008	Chin J Postgrad Med 31: 40–42
Bi et al.	China	386	375	0.12	2009	Neurobiol Aging 30: 1601–1607
Giedraitis et al.	Sweden	85	399	0.23	2009	Dement Geriatr Cogn Disord 27: 59-68
Li et al.	China	198	240	0.02^{a}	2009	Dement Geriatr Cogn Disord 27: 286-291
Ferlazzo et al.	Italy	69	69	0.92	2011	Neuromolecular Med 13: 167–174
Coppede et al.	Italy	378	305	0.44	2012	Antioxid Redox Signal 17: 195-204
Deng et al.	China	75	71	0.87	2012	Shi Yong Yi Xue 28: 3545–3548
Mansoori et al.	India	80	120	0.84	2012	Neurobiol Aging 33: 1003 e1001–1008
Divyakolu et al.	India	25	50	0.54	2013	Neurol Disord, 2 (1): 1000142
Elhaway et al.	Egypt	43	32	0.63	2013	Dis Markers 35: 439–446
Mansouri et al.	Tunisia	38	100	0.67	2013	Clin Neurol Neurosurg 115: 1693–1696
Chillar et al.	India	100	100	0.23	2014	Ann Indian Acad Neurol,17(): 308–312

^a Control populations were not in HWE

total variability in a set of effect sizes due to true heterogeneity, was used to quantify the degree of heterogeneity [31]. A tentative classification of l^2 values proposed by Higgins and Thompson has been used to interpret the magnitude, viz., 25, 50, and 75 %, which corresponds to low, medium, and high heterogeneities, respectively [32]. In the absence of significant heterogeneity determined by the results of Q test, the Mantel-Haenszel fixed effect model (Peto method) was used for the

 Table 2
 The distributions of MTHFR C677T genotypes and allele numbers in Alzheimer's cases and controls

Study ID	Genotype						Alleles			
	CC		СТ		TT		С		Т	
	Case	Control	Case	Control	Case	Control	Case	Control	Case	Control
Chapman et al. 1998	12	15	31	16	6	9	55	46	43	34
Nishiyama et al. 1999	4	13	14	15	6	5	22	41	26	25
Pollak et al. 2000	30	29	49	37	13	16	109	95	75	69
Brunelli et al. 2001	64	39	120	65	47	33	248	143	214	131
Postiglione et al. 2001	22	25	37	36	15	13	81	86	67	62
Prince et al. 2001	98	83	90	68	16	21	286	234	122	110
Zuliani et al. 2001	14	17	18	25	8	12	46	59	34	49
MacIlroy et al. 2002	59	50	20	19	4	2	138	119	28	23
Wakutani et al. 2002	93	137	112	163	36	52	298	437	184	267
Religa et al. 2003	53	55	38	38	8	7	144	148	54	52
Seripa et al. 2003	49	45	63	43	12	9	161	133	87	61
Seripa et al. 2003	30	28	67	55	29	23	127	111	125	101
Anello et al. 2004	54	53	89	90	37	38	197	196	163	166
Bi et al. 2004	23	21	13	12	6	7	59	54	25	26
Jiang et al. 2004	22	24	46	36	7	12	90	84	60	60
Kida et al. 2004	64	144	98	193	32	42	226	481	162	277
Liao et al. 2004	26	90	37	50	3	3	89	230	43	56
Linnebank et al. 2004	71	79	75	68	16	22	217	226	107	112
Wang et al. 2004	33	41	74	73	20	24	140	155	114	121
Fernandez et al. 2005	11	15	17	11	2	3	39	41	21	17
Wang et al. 2005	50	79	38	47	16	4	138	205	70	55
Zhang et al. 2005	39	34	45	49	21	19	123	117	87	87
da Silva et al. 2006	19	25	19	23	5	2	57	73	29	27
Keikhace et al. 2006	57	84	44	33	16	8	158	201	76	49
Wehr et al. 2006	51	63	38	66	10	12	140	192	58	90
Dorszewska et al. 2007	14	28	20	20	4	2	48	76	28	24
Kim et al. 2007	11	122	43	332	32	171	65	576	107	674
Yuan et al. 2007	11	27	13	38	6	15	35	92	25	68
Zhang 2007	6	11	44	29	18	28	56	51	80	85
Styczyaska et al. 2008	79	97	58	67	17	12	216	261	92	91
Zhang et al. 2008	21	97	14	67	8	12	216	261	92	91
Bi et al. 2009	82	20	179	13	125	7	56	53	30	27
Giedraitis et al. 2009	42	90	37	172	6	113	343	352	429	398
Li et al. 2009	37	206	92	168	69	25	121	580	49	218
Ferlazzo et al. 2011	17	69	32	103	20	68	166	241	230	239
Coppede et al. 2012	99	26	193	33	86	10	391	85	365	53
Deng et al. 2012	45	106	27	142	3	57	391	354	365	256
Mansoori et al. 2012	51	48	26	21	3	2	117	117	33	25
Divyakolu et al. 2013	19	89	5	29	1	2	128	207	32	33
Elhaway et al. 2013	18	20	17	10	8	2	53	50	33	14
Mansouri et al. 2013	13	52	24	39	1	9	50	143	26	57
Chillar et al. 2014	31	69	58	26	11	5	120	164	80	36

combination of data, while in the presence of significant heterogeneity, the Dersimonian-Laird random effect model (DL method) was used for combining the data [33, 34]. High-resolution plots (forest plots) were generated to estimate the

Table 3 Summary estimates for the odds ratio (OR) of MTHFR C677T in various allele/genotype contrasts, the significance level (p value) of heterogeneity test (Q test), and the l^2 metric and publication bias p value (Egger test) in total, Asian, and Caucasian studies

Genetic models	Fixed effect OR (95 % CI), p	Random effect OR (95 % CI), p	Heterogeneity p value (Q test)		Publication bias (<i>p</i> of Egger's test)	
Total studies						
Allele contrast (T vs C)	1.2 (1.13–1.27), <0.0001	1.29 (1.07–1.56), 0.0003	< 0.0001	71.58	0.25	
Co-dominant (CT vs CC)	1.24 (1.13–1.35), <0.0001	1.24 (1.13–1.35), <0.0001	0.68	0	0.03	
Homozygote (TT vs CC)	1.3 (1.15–1.47), <0.0001	1.31 (1.16–1.48), 0.0002	0.29	10.02	0.3	
Dominant (TT + CT vs CC)	1.26 (1.16–1.37), <0.0001	1.29 (1.19–1.40), <0.0001	0.47	0	0.06	
Recessive (TT vs CT + CC)	1.14 (1.02–1.26), 0.02	1.13 (1.00–1.28), 0.04	0.21	14.56	0.61	
Asian studies						
Allele contrast (T vs C)	1.19 (1.10–1.28), <0.0001	1.21 (1.1–1.34), 0.0003	0.065	34.02	0.46	
Co-dominant (CT vs CC)	1.29 (1.14–1.46), <0.0001	1.27 (1.0–1.64), <0.0001	0.43	1.5	0.06	
Homozygote (TT vs CC)	1.36 (1.16–1.59), 0.0001	1.42 (1.16–1.76), 0.001	0.21	19.3	0.41	
Dominant (TT + CT vs CC)	1.32 (1.17–1.48), <0.0001	1.33 (1.17–1.51), <0001	0.03	11.54	0.16	
Recessive (TT vs CT + CC)	1.16 (1.02–1.32), 0.03	1.15 (0.96–1.38), 0.11	0.09	29.76	0.86	
Caucasian studies						
Allele contrast (T vs C)	1.26 (1.15–1.37), <0.0001	1.34 (1.01–1.68), 0.01	< 0.0001	83.19	0.43	
Co-dominant (CT vs CC)	1.17 (1.03–1.33), 0.01	1.17 (1.03–1.33), 0.02	0.73	0	0.48	
Homozygote (TT vs CC)	1.2 (0.99-1.45), 0.06	1.19 (0.98–1.45), 0.006	0.41	3.21	0.51	
Dominant (TT + CT vs CC)	1.18 (1.04–1.33), 0.008	1.18 (1.04–1.33), 0.009	0.57	0	0.28	
Recessive (TT vs CT + CC)	1.1 (0.92–1.29), 0.32	1.1 (0.91–1.29), 0.34	0.50	0	0.47	



Fig. 2 Forest plots for the association between MTHFR C677T polymorphism and Alzheimer's disease for allele contrast model (T vs C) with random effect model



Fig. 3 Forest plots for the association between MTHFR C677T polymorphism and Alzheimer's disease for homozygote model (TT vs CC) with random effect model

pooled odds ratio corresponding to 95 % CI and the p value. Stratified analyses were performed by ethnicity. Sensitivity analysis was performed to evaluate the stability of the results by removing the studies not in Hardy-Weinberg equilibrium (HWE), studies with very high p value, and studies with small sample size. Cumulative meta-analysis was also done to observe the effect of subsequent addition of each study. Hardy-Weinberg equilibrium and chi-squared test methods were used to test the distribution of genotypes in the control group of each study. For this purpose, data was analyzed using calculator available at http://ihg.gsf.de/cgi-bin/hw/hwa1.pl.

Publication Bias

Publication bias was investigated by using the funnel plots, viz., funnel plot of standard error by log odds ratio and funnel plot of precision by log odds ratio. Different statistical tests such as Begg and Mazumdar rank correlation [35] and Egger's regression intercept [36] were adopted to assess and quantify the publication bias and its impact on the analysis. P < 0.05 was considered to indicate statistical significance, and all the *p* values were two sided. All statistical analyses,

forest, and funnel plots were performed by the computer program MIX version 1.7 [37].

Result

The preliminary search resulted in 119 publications from Pubmed, Google Scholar, Elsevier, and Springer Link. Out of which, 51 were irrelevant for the present meta-analysis, which includes reviews, book chapters, case reports, editorials, and articles mentioning other genes. After initial exclusion, a total of 68 article publications were identified. Out of which, 16 articles were irrelevant for the present meta-analysis, and in 11 articles only, cases were discussed. Thus, a total of 41 articles were included in present meta-analysis. The search workflow was shown in Fig. 1.

Study characteristics were summarized in Table 1. Fortyone articles that investigated the association between C677T polymorphism and AD were found suitable for the inclusion in the present meta-analysis [3, 21–28, 38–68]. One author [25] studied two different population; their data were included as two independent studies. A total of 42 studies were found suitable for the inclusion in present meta-analysis.



Fig. 4 Forest plots for the association between MTHFR C677T polymorphism and Alzheimer's disease for dominant model (TT + CT vs CC) with random effect model

The studies were published between 1998 and 2014. All these 42 studies were performed in different countries—Brazil [49, 51], China [27, 44, 45, 47, 48, 50, 56, 57, 59, 60, 62, 65], Egypt [68], Germany [26], India [66, 67, 69], Iran [52], Ireland [41], Israel [21, 39], Italy [22, 23, 25, 40, 43, 63, 64], Japan [38, 42, 46], Korea [55], Poland [24, 53, 54, 58], Sweden [3, 61], Tunisia [70], and USA [25].

Summary Statistics

In all 42 studies, total cases were 4888 with CC (1676), CT (2290), and TT (922), and controls were 6142 with CC (2389), CT (2719), and TT (1034) genotypes. In control genotypes, percentages of CC, CT, and TT were 38.90, 44.27, and 16.80 %, respectively. In total cases, genotype percentages of CC, CT, and TT were 34.29, 46.85, and 18.86 %, respectively. Frequency of CC genotype was highest in both cases and controls (Table 2). In cases and controls, the allele C was the most common. Except two studies [44, 62], the genotype distributions among the controls of all 39 studies followed Hardy-Weinberg equilibrium. Out of 42 studies, ORs of 12 studies [3, 22, 23, 26, 39, 43–45, 50, 53, 56, 57] were below

one; i.e., these studies did not show any association between MTHFR C677T polymorphism and AD.

Allele Contrast Meta-Analysis

The main results of meta-analysis and the heterogeneity test were shown in Table 3. Allele contrast meta-analysis showed significant association with fixed effect model (allele contrast model OR_{T vs C}=1.20, 95 % CI 1.13–1.27, p<0.0001, I^2 =71.58 %, $P_{\text{heterogeneity}}$ <0.0001, P_{Pb} =0.25) and random effect model (OR_{T vs C}=1.29, 95 % CI=1.07–1.56, p=0.0003) (Table 3 and Fig. 2). High significant heterogeneity was found, so random effect model was adopted. In cumulative meta-analysis using random effect model, the association of T allele with AD turned statistically significant with the addition of study of Liao et al. [47] and remained significant thereafter.

Genotype Contrast Meta-Analysis

Table 3 summarizes the ORs with corresponding 95 % CIs for association between MTHFR C677T polymorphism and risk of AD in dominant, recessive, homozygote, and co-dominant



Fig. 5 Forest plots for the association between MTHFR C677T polymorphism and Alzheimer's disease for homozygote model (TT vs CC) in Asian studies with random effect model

models. There was evidence of association of the MTHFR TT genotype with the risk of AD relative to the MTHFR CC genotype. Since there was no significant heterogeneity (p=0.29, $I^2=10.02$ %) between the studies, the fixed effect pooled OR was considered (homozygote model OR_{TT vs} $_{CC}=1.31$, 95 % CI=1.16–1.48, p=0.0002) (Fig. 3). Significant association was observed between C677T polymorphism and AD using co-dominant model (OR_{CT vs} $_{CC}=1.24$, 95 % CI=1.13–1.35, p<0.0001). The overall analysis of the recessive model for T allele showed insignificant association (OR_{TT vs} CT + CC=1.13, 95 % CI=1.00–1.28, p=0.04). However, the dominant model for the effect of T allele produced significant association overall (OR_{TT + CT vs} CC=1.29, 95 % CI=1.19–1.40, p<0.0001) (Fig. 4).

Subgroup Analysis

Subgroup analysis based on ethnicity was performed. In all eligible studies, 21 studies were conducted in Asians, 18 studies were conducted in Caucasians, and 2 studies were conducted in other population. In Asian population (number of studies = 21; 2588 cases/3630 controls), allele contrast meta-

analysis showed significant association adopting both fixed (OR_{T vs C} = 1.19, 95 % CI = 1.10–1.28, p = <0.0001, $I^2 = 34.02$ %, p value of heterogeneity (P_{hetero}) = 0.06, p value of Eggers test (P_{Pb}) = 0.46), and random (OR_{T vs C} = 1.21, 95 % CI = 1.1–1.34, p = 0.0003) effect models. Combined mutant genotypes (dominant model) also showed significant association with fixed (OR_{TT + CT vs CC} = 1.32, 95 % CI = 1.17–1.48, p < 0.0001) and random (OR_{TT + CT vs CC} = 1.33, 95 % CI = 1.17–1.51, p < 0.0001) effect models. In this subgroup, heterogeneity between studies ($I^2 = 11.54$ %, $P_{hetero} = 0.03$) was low and publication bias ($P_{Pb} = 0.16$) was absent (Table 3 and Fig. 5).

In European population (number of studies = 18; 2227/2433 cases/controls), allele contrast meta-analysis showed significant association with both fixed (OR_{T vs} $_{C}$ =1.26, 95 % CI=1.15–1.37, p<0.0001) and random (OR_{T vs} $_{C}$ =1.34, 95 % CI=1.01–1.68, p=0.01) effect models with significant heterogeneity (I^{2} =83.19 %, $P_{heterogeneity}$ =0.43). The combined mutant genotype (dominant model) showed statistically significant association with fixed effect model (OR_{TT + CT vs} CC=1.18, 95 % CI=1.04–1.33, p=0.008) with no heterogeneity and no publication bias (P_{Pb} =0.28) (Table 3 and Fig. 6).

Fig. 6 Forest plots for the association between MTHFR C677T polymorphism and Alzheimer's disease for homozygote model (TT vs CC) in Caucasian studies with random effect model



Odds Ratio 95% Confidence Interval

Sensitivity Analysis

Sensitivity analyses were conducted to determine whether modification of the inclusion criteria of the meta-analysis affected the final results. In allele contrast meta-analysis, sensitivity analysis performed by exclusion of the studies in which control population was not in Hardy Weinberg equilibrium, studies with small sample size and studies with high *p* values. Control population of only two studies [44, 62] were not in HW equilibrium, and heterogeneity did not decreased after exclusion of these two studies (P_{hetero} , $I^2 = 72.59$ %). Exclusion of ten studies involving small sample size, less than 50 (Chapman et al. [21], n = 49; Nishiyama et al. [38], n = 24; Zuliani et al. [23], n = 40; Bi et al. [44], n = 42; Fernandez et al. [49], n = 30; da Silva et al. [51], n = 43; Dorszewska et al. [54], n = 38; Yuan et al. [56], n = 30; Zhang et al. [59], n = 43; Elhaway et al. [68], n = 43, also did not decreased heterogeneity ($P_{hetero} < 0.0001$, $I^2 = 77.27$ %).

Publication Bias

Begg's funnel plot and the Egger's test were conducted to estimate the publication bias of articles. Both the results of

Begg's and Egger's test showed evidence of publication bias in allele contrast and co-dominant models, and absence of publication bias was observed in homozygote, dominant, and recessive genetic models (T vs C Begg's test p=0.02, Egger's test p=0.24; CT vs CC Begg's test p=0.02, Egger's test p=0.03; TT vs CC Begg's test p=0.18, Egger's test p=0.3; dominant model TT + CT vs CC Begg's test p=0.06, Egger's test p=0.06; and recessive model TT vs CT + CC Begg's test p=0.47, Egger's test p=0.61) (Fig. 7 and Table 3).

Discussion

Results of present meta-analysis showed strong association between MTHFR C677T polymorphism and Alzheimer's disease. Several epidemiological and experimental evidences have linked derangements of one-carbon metabolism to vascular, neurodegenerative, and neuropsychiatric diseases, including strokes [71]. Folate deficiency and thermolabile MTHFR enzyme (TT genotype) with low activity increase concentration of homocysteine. The deficiency of MTHFR





Fig. 7 Funnel plots of total studies **a** precision versus OR for T vs C model, **b** standard error versus OR for T vs C model, **c** precision versus OR for TT vs CC model, **d** standard error versus OR for TT vs CC model,

e precision versus OR for TT + CT vs CC model, and f standard error versus OR for TT + CT vs CC model

causes an accumulation of 5,10-methylenetetrahydrofolate as well as the inhibition of 5-methyltetrahydrofolate synthesis. Reduced synthesis of 5-methyltetrahydrofolate will cause decreased homocysteinemia, hypomethioninemia, and reduced *S*-adenosylmethionine occur frequently in severe MTHFR deficiency [72]. There is several evidences that MTHFR mutant TT genotype is in accordance with elevated homocysteine level, which in itself is an independent risk factor for vascular diseases and cognitive impairment [73]. Folic acid supplementation has a protective effect on homocysteine-induced oxidative stress by reducing intracellular superoxide levels and, to a lesser extent, quenching hydrogen peroxide [74].

Higher concentration of homocysteine could affect adult brain and cause degeneration in adult brain by several mechanisms like (i) it altered DNA methylation pattern [71] and affects gene expression; (ii) it impaired DNA repair [71]; (iii) it increased oxidative stress [74], generated free oxygen radicals, accelerated DNA damage, and eventually lead to neuron apoptosis [74–77]; (iv) it reduced antioxidant reserves of the cell; (v) it enhanced beta-amayloid peptide generation [78]; (vi) it sensitized neurons to amyloid toxicity [75]; (vii) it released inflammatory mediators such as nuclear factor (NF)kappa B, interleukin (IL)-1beta, IL-6, and IL-8 [66, 79]; (viii) it modulated IL-6 genes [80], which is a proinflammatory cytokine and promoted neuronal expression of neurofilaments, tau protein, and beta-amyloid precursor protein; all are involved in pathogenesis of AD [66, 81]; and (ix) Hcy impaired vascular endothelial function [82] (Stuhlinger et al. 2003), stimulated vascular smooth muscle proliferation [83], breaks balance between coagulation and bleeding pathway [84], and mediated thrombosis. Ultimately, these adverse effects reduce blood supply to the brain and accelerated the neuron apoptosis [77]. Meta-analysis is the statistical analysis of a large collection of analysis results for the purpose of integrating the findings, and it is a powerful tool for systematic review of a focused topic in the literature that provides a quantitative estimate for the effect of a gene, treatment intervention, or exposure [85]. Numerous meta-analyses were published considering MTHFR C677T as risk for several diseases and defects like neural tube defects [10], Down syndrome [86], congenital heart defects [87], schizophrenia [88], bipolar [89], anxiety [89], depression [90], and cancer [91]. One meta-analysis was identified during literature search on the same topic [77]. Zhang et al. [77] pooled 19 studies and reported significant association between AD and MTHFR 677 T polymorphism (OR = 1.15, 95 % CI = 1.08-1.39) using allele contrast genetic model. There are several published studies available but not included in the previous meta-analysis. So, author conducted a comprehensive meta-analysis with the largest number of studies to date to investigate the possible relationship between maternal MTHFR C677T polymorphism and the risk of AD.

The results of this meta-analysis should be interpreted with some caution because there were few limitations in present analysis like (i) results were based on unadjusted OR values that lack the original data from the eligible studies, which could lead to relatively weak power to estimate the real relationship; (ii) sample size in few included studies was relatively small to investigate the association between MTHFR C677T polymorphism and AD risk; (iii) significant between-study heterogeneity was detected and may be distorting the metaanalysis; and (iv) present meta-analysis was based on singlefactor estimates, which overlooked the interactions of genegene and gene-environment in the development of AD. Present meta-analysis had several strengths also. First, substantial number of cases and controls were pooled from different studies, which significantly increased the statistical power of the analysis. Second, no publication biases were detected, indicating that the whole pooled results may be unbiased.

In conclusion, present comprehensive meta-analysis indicates that there is a conclusive association between the MTHFR C677T polymorphism and the risk of AD, whereas significant heterogeneity was evident across the individual studies. The subgroup analysis also indicated that there is a significant association between MTHFR C677T gene polymorphic variation and AD patients in Asian and Caucasian population. Acknowledgments The authoress is highly grateful to Leon Bax (Chief Scientific Officer at BiostatXL, UMC Utrecht) for his valuable suggestions, which help me in statistical analysis.

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