REVIEW ARTICLE

Relationship Between CASP9 and CASP10 Gene Polymorphisms and Cancer Susceptibility: Evidence from an Updated Meta-analysis

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Abstract

Caspase-9 (CASP9) and caspase-10 (CASP10) polymorphisms were associated with human cancers; however, the results remain controversial. In this meta-analysis, we aimed to estimate the relationship among CASP9 (rs1052576, rs1052571, rs4645978, rs4645981, rs4645982, rs2308950) and CASP10 (rs13006529, rs13010627, rs3900115) polymorphisms and the overall risk of cancers. Relevant studies were obtained from Web of Science, MEDLINE, PubMed, Scopus, and Google scholar databases (updated January 1, 2021). Odds ratio (OR) and 95% confidence intervals (CIs) were measured to estimate the strength of association. Our meta-analysis included 40 studies. The rs4645981 significantly enhanced the risk of cancer under TT vs. CC ($OR = 2.42$), TC vs. CC $(OR = 1.55)$, TT + TC vs. CC (OR = 1.66), TT vs. TC + CC (OR = 1.91), and T vs. C (OR $= 1.57$) inheritance models. As for the rs1052571 variant, increased risk of cancer was observed under TT vs. CC (OR = 1.22), TC vs. CC (OR = 1.17), and TT+ TC vs. CC (OR = 1.18) models. The stratified analysis showed a significant correlation between rs4645978 or rs4645981 polymorphisms and cancer risk, while in Asians rs4645978 conferred an increased risk of colorectal, lung, and prostate cancer. Both rs4645981 and rs1052576 polymorphisms were correlated with an enhanced risk of lung cancer. In conclusion, our meta-analysis suggested that CASP9 rs4645981 and rs1052571 polymorphisms are associated with overall cancer risk. More studies on larger populations are warranted to validate these associations.

Keywords $Cancer \cdot Caspase-9 \cdot Caspase-10 \cdot Meta-analysis \cdot Polymorphism$

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Introduction

Cancer is ranked as the second leading cause of death among adults in the USA and is considered a global public health concern with great importance [[1\]](#page-20-0). Only in the USA, a total of 1,806,590 new cancer cases and more than six hundred thousand cancer-related deaths are projected to occur in 2020 [[2](#page-20-0)]. Moreover, it has been approximated that the worldwide incidence of cancer will exceed 25 million by 2032 [\[3\]](#page-20-0). As a multifactorial disease, genetics and environmental factors serve pivotal roles in cancer etiology [[4](#page-20-0)].

Apoptosis is a programmed cell death mechanism that modulates tissue hemostasis in different organisms [[5\]](#page-20-0). The escape from apoptosis is known to be a hallmark of malignancy, and it regulates the development and progression of tumors [[5](#page-20-0)–[7\]](#page-20-0). Historically, major pathways of apoptosis have been characterized that lead to activation of effector caspases and cell death: the intrinsic (mitochondrial) pathway and the extrinsic (receptor-mediated) pathway. Apoptotic pathways converge at the activation of effector caspases (CASP-3, -6, and -7) [[5,](#page-20-0) [8\]](#page-20-0). CASP 8 and 10 are initiator caspases activated after ligand binding to death receptors (i.e., tumor necrosis factor receptor superfamily) [[9\]](#page-20-0). Downregulation of $CASP9$ and $CASP10$ is frequently observed in cancer patients and correlates with resistance to chemotherapy and/or poor clinical outcome [\[10,](#page-20-0) [11](#page-20-0)].

The gene encoding human CASP9 is mapped on chromosome 1p36.2, spans \sim 33 kb in length, and consists of 8 introns and 9 exons. The human CASP10 gene resides on chromosome 2q33.1 with ~46 kb length, 13 exons, and 11 introns and located approximately 20–30 kb to the 5′ untranslated region of the CASP8 gene [\[12](#page-21-0)]. Genomic mutations in CASP10 and allelic imbalance and epigenetic modifications in the CASP9 gene have been reported in multiple human tumors [[13,](#page-21-0) [14\]](#page-21-0). Singlenucleotide polymorphisms (SNPs) are the most common type of single-base variations that predicts the risk of multiple diseases, including cancer [\[15\]](#page-21-0). Among the candidate SNPs, variations mapped in the promoteric regions of genes are well-studied because they most probably influence gene expression and might affect cancer susceptibility [[5\]](#page-20-0). Figure [1](#page-2-0) and Table [1](#page-3-0) illustrate the genetic location of CASP9 and CASP10 polymorphisms.

Several reports have studied the relationship between CASP9 and CASP10 polymorphisms and cancer [[5](#page-20-0), [15](#page-21-0)–[53](#page-22-0)]. However, the implication of these two initiator caspases and the risk of developing human cancers remains ill-defined. Hence, gaining a better understanding of the role of CASP9 and CASP10 variations in cancer incidence will expand our horizons for designing such curative strategies. Therefore, given the amount of accumulated data, we conducted a comprehensive meta-analysis by including the most relevant and recent publications (updated January 1, 2021) to identify statistical evidence.

Methods

Literature Search

We retrieved a list of the case–control studies through a comprehensive Internet-based literature search of Web of Knowledge, PubMed, Scopus, Google Scholar, and

Fig. 1 Flow diagram of selecting studies for the meta-analysis

Embase databases. The used keywords were as follows: ("cancer" OR "carcinoma" OR "tumor" OR "neoplasm" OR "neoplasia") AND ("caspase-10" OR "caspase-9" OR "CASP 10" OR "CASP 9" OR "caspase10" OR "caspase9" OR "CASP10" OR "CASP9" OR "caspase 10" OR "caspase 9" OR "CASP-10" OR "CASP-9") AND ("gene polymorphism" OR "polymorphism" OR "SNP" OR "gene mutation" OR "gene variant" OR "mutation" OR "variant"). No language, country, or ethnicity restrictions were imposed. Additional publications were retrieved using a hand search. If the results of studies on different tumors or gene polymorphisms were reported in the same literature, they were regarded as a separate study to report. The included studies met the following criteria; [[1\]](#page-20-0) original case–control study about CASP9 and CASP10 polymorphisms and cancer susceptibility; [[2\]](#page-20-0) studies with sufficient published data to enable the estimation of odds ratios (ORs) with confidence intervals (CIs); [[3\]](#page-20-0) the frequency distribution of genotypes in the control group conformed to Hardy– Weinberg equilibrium (HWE). Review articles, case reports, duplicate publications, and studies with too little information were excluded.

Document Quality Assessment

Quality assessment (QA) of the included publications was carried out by two researchers as described previously [\[54](#page-23-0)]. Each publication was scored carefully. Low-quality studies were scored equal to or less than 9, while high-quality studies were scored more than 9 (Table [1](#page-3-0)). In case of disagreement, the two researchers would settle through information exchange and ultimately reached an agreement.

Data Extraction

Two independent researchers (S.S and A.Z.A) abstracted the relevant data according to the standard protocol and the study's criteria. A third author (H.S) joined the study later to settle possible discrepancies. The following information were extracted from each included study: the first author's name, publication date, ethnicity, country, cancer type, the method for genotyping of CASP9 and CASP10 polymorphisms, allele, and genotype distribution in the studied groups, and results of the HWE test (Table [1\)](#page-3-0).

Statistical Analysis

Data analysis was carried out by both Stata15.0 statistical software and MetaGenyo web tool [[55](#page-23-0)]. Deviation from HWE was examined in controls via a χ^2 test. Pooled ORs with 95% CIs were calculated to estimate the strength of association between CASP9 and CASP10 variants and susceptibility to cancer under allelic, homozygous/heterozygous codominant, dominant, and recessive contrasted genetic models. $P < 0.05$ was considered statistically significant. Heterogeneity between-studies was evaluated via $I²$ statistics. We applied a fixed-effects model if I^2 < 50%, and if heterogeneity was present (I^2 > 50%), analyses were repeated using a random-effects model. Publication bias was determined via Egger's test and visual inspection of funnel plots. The sensitivity analysis was done by sequentially omitting each study and calculating the pooled OR to investigate the effect of each study on overall estimates.

TSA and FPRP Analyses

In this study, trial sequential analysis (TSA) was used to enhance the robustness of the conclusion and decrease the random errors caused by sparse data and repetitive testing. We used the TSA software version 0.9.5.10 ([http://www.ctu.dk/tsa/\)](http://www.ctu.dk/tsa/) to calculate the required information size (RIS) (meta-analysis sample size) [[56\]](#page-23-0) under the assumption of a plausible relative risk of 10% with low-risk bias, and the significance of 5% for type I error and 20% for type II error (power 80%). The TSA monitoring boundaries were plotted based on the required information size and the risk for type I and type II errors. The robustness of the conclusion is confirmed when the cumulative Z-curve (blue line) passes the TSA monitoring boundary (dotted red lines sloping inward) before the required information size is obtained. Otherwise, the data is insufficient to get a robust conclusion, and more trials are required. False-positive report probability (FPRP) values were assessed with different prior probabilities (0.25, 0.1, 0.01, 0.001, and 0.0001) [\[57\]](#page-23-0). An FPRP value < 0.2 indicated a significant correlation.

Result

Basic Information of Research Data

A total of 40 articles published between 2004 and 2020, including 15 case–control studies on CASP9 rs4645978, 13 studies on CASP9 rs1052576, 8 studies on CASP rs4645981, 7 studies on CASP9 rs4645982 and CASP10 rs13006529, 5 studies on CASP9 s1052571, 4 studies on CASP10 rs3900115, and 3 studies on either CASP9 rs2308950 or CASP[1](#page-2-0)0 rs13010627 were included in this study (Table 1). Figure 1 illustrates the specific screening process of the retrieved studies. Figure 2 and Supplementary Table 1 show the position of analyzed SNPs within the CASP9 and CASP10 genes. The basic information of the included studies and their QA scores are represented in Supplementary Table 2.

Main Analysis Results

Table [2](#page-8-0) demonstrates the main results of the meta-analysis on the association of CASP9 and CASP10 variants with cancer susceptibility. Our pooled analysis revealed no significant association between CASP9 rs4645978, rs1052576, rs4645982, and rs2308950 polymorphisms and cancer incidence under different inheritance patterns. However, the pooled OR from 6 studies showed that CASP9 rs4645981 enhanced the risk of developing cancer under allelic [OR = 1.57; 95% CI, 1.25–1.97; $P < 0.001$, T vs. C], codominant homozygous $[OR = 2.42; 95\% \text{ CI}, 1.46-4.00; P < 0.001, TT \text{ vs. CC}, codominant$ heterozygous [OR =1.56; 95% CI, 1.21–2.00; $P < 0.001$, TC vs. CC], dominant [OR = 1.66; 95% CI, 1.26–2.17; P < 0.001, TT+ TC vs. CC], and recessive [OR = 1.92; 95% CI,

Fig. 2 Schematic representation and information of the examined variants on 1p36.2 and 2q33.1

Polymorphism		No. Genetic model	Association test		Heterogeneity			Egger's
			OR (95% CI)	\boldsymbol{P}	Model	I^2 $(\%)$	P _h	test \boldsymbol{P}
CASP9 rs4645978	14	AA vs. GG	$1.31(0.92-1.8)$	0.1320	Random	86.32	0.000	0.912
		AG vs. GG	0.83(0.59; 1.15)	0.276	Random	84.65	0.000	0.151
		AA+ AG vs. GG	1.10(0.85; 1.41)	0.445	Random	78.1	0.000	0.923
		AA vs. AG + GG	1.40(0.96; 2.03)	0.076	Random	93.77	0.000	0.437
		A vs. G	1.21(0.94; 1.55)	0.131	Random	93.48	0.000	0.649
CASP9 rs1052576	14	GG vs. AA	0.79(0.56; 1.10)	0.172	Random 67.74		0.000	0.143
		GA vs. AA	1.00(0.87; 1.16)	0.908	Fixed	34.24	0.101	0.571
		GG+ GA vs. AA	0.92(0.73; 1.14)	0.468	Random 53.78		0.008	0.239
		GG vs. $GA + AA$	0.80(0.61;1.06)	0.122	Random 69.34		0.000	0.230
		G vs. A	0.88(0.7;1.05)	0.176	Random	71.9	0.000	0.199
CASP9 rs4645981	6	TT vs. CC	2.42(1.46; 4.00)	0.001	Random 55.16		0.048	0.433
		TC vs. CC	1.56(1.21; 2.00)	< 0.001	Random	67.92	0.008	0.593
		TT+ TC vs. CC	1.66(1.26; 2.17)	0.001	Random	74.13	0.001	0.384
		TT vs. TC + CC	1.92(1.42; 2.58)	0.001	Fixed	39.65	0.141	0.561
		T vs. C	1.57(1.25; 1.97)	0.001	Random	75.25	0.001	0.375
CASP9 rs4645982	5	II vs. DD	1.04(0.84; 1.27)	0.707	Fixed	17.4	0.303	0.942
		ID vs. DD	0.90(0.76; 1.06)	0.227	Fixed	0	0.558	0.897
		$II + ID$ vs. DD	0.95(0.81; 1.11)	0.521	Fixed	$\boldsymbol{0}$	0.644	0.664
		II vs. $ID + DD$	1.13(0.94; 1.35)	0.178	Fixed	40.45	0.151	0.789
		I vs. D	1.01(0.91; 1.13)	0.722	Fixed	17.29	0.304	0.698
CASP9 rs1052571	5	TT vs. CC	1.22(1.00; 1.50)	0.046	Fixed	0	0.491	0.289
		TC vs. CC	1.17(1.00; 1.38)	0.049	Fixed	θ	0.598	0.816
		TT+ TC vs. CC	1.18(1.00; 1.38)	0.032	Fixed	0	0.542	0.776
		TT vs. $TC + CC$	1.07(0.90; 1.26)	0.403	Fixed	θ	0.414	0.107
		T vs. C	1.09(0.99; 1.21)	0.062	Fixed	$\boldsymbol{0}$	0.460	0.189
CASP9 rs2308950	3	AA vs. GG	4.77(0.49; 46.19)	0.176	Fixed	θ	0.799	0.218
		AG vs. GG	0.85(0.35; 2.04)	0.721	Fixed	33.92	0.220	0.197
		AA+ AG vs. GG	1.06(0.46; 2.47)	0.878	Fixed	54.9	0.108	0.182
		AA vs. AG + GG	$4.73(0.49)$; 45.81)	0.178	Fixed	0	0.802	0.223
		A vs. G	1.33(0.34; 5.23)	0.678	Random	64.72	0.058	0.166
CASP10 rs3900115	$\overline{4}$	GG vs. AA	1.01(0.77; 1.33)	0.914	Fixed	22.02	0.278	0.932
		AG vs. AA	1.08(0.89; 1.32)	0.387	Fixed	θ	0.393	0.127
		$GG + AG$ vs. AA	1.05(0.87; 1.26)	0.569	Fixed	0	0.580	0.264
		GG vs. $AG + AA$	0.93(0.74; 1.17)	0.555	Fixed	28.03	0.243	0.900
		G vs. A	1.00(0.88; 1.13)	0.965	Fixed	0	0.799	0.459
CASP ₁₀ rs13010627	3	AA vs. GG	0.94(0.73; 1.22)	0.682	Fixed	32.96	0.225	0.280
		AG vs. GG	0.81(0.57; 1.13)	0.229	Random	77.63	0.011	0.111
		AA+ AG vs. GG	0.79(0.54; 1.14)	0.210	Random	81.39	0.004	0.136
		AA vs. $AG + GG$	0.94(0.73; 1.22)	0.672	Fixed	27.87	0.25	0.284
		A vs. G	0.78(0.54; 1.13)	0.200	Random	83.44	0.002	0.156
CASP10 rs13006529	6	AA vs. TT	1.14(0.97; 1.34)	0.095	Fixed	0.340	11.71	0.272
		AT vs. TT	1.12(0.98; 1.30)	0.091	Fixed	0	0.629	0.547
		AA+ AT vs. TT	$1.137(0.99)$; 1.29)	0.060	Fixed	0	0.465	0.379
		AA vs. AT + TT	1.05(0.94; 1.19)	0.340	Fixed	$\mathbf{0}$	0.529	0.145
		A vs. T	1.07(0.99; 1.15)	0.086	Fixed	11.66	0.340	0.163

Table 2 The pooled ORs and 95% CIs for the association between CASP9 and CASP10 polymorphisms and overall risk of cancer

CI confidence interval, OR odds ratio, P_h P value from the heterogeneity test

1.42–2.58; $P < 0.001$, TT vs. TC + CC genetic patterns. Figure 3 demonstrates the forest plot for the relationship between CASP9 rs4645981 polymorphism and cancer risk using different genetic contrasted models. Regarding CASP9 rs1052571, an increased risk of developing cancer was observed under codominant homozygous [OR = 1.22; 95% CI, 1.00–1.50; $P = 0.046$, TT vs. CC], codominant heterozygous [OR = 1.17; 95% CI, 1.00– 1.38; $P = 0.049$, TC vs. CC], and dominant [OR = 1.18; 95% CI, 1.00–1.38; $P = 0.032$, TT+ TC vs. CC] modes of inheritance. No significant association was found between CASP10 SNPs and overall risk of cancer.

Stratified Analysis Results

The subgroup analyses of CASP9 and CASP10 variants on risk of cancer are shown in Tables [3](#page-11-0) and [4.](#page-14-0) Regarding ethnicity of cancer patients, the CASP9 rs4645978 variant conferred an increased risk of cancer in Asians, under codominant homozygous [OR = 1.26; 95% CI, 1.06– 1.50; $P = 0.008$, AA vs. GG], recessive [OR = 1.16; 95% CI, 1.02–1.31; $P = 0.015$, AA vs. AG + GG], and allelic $[OR = 1.12; 95\% \text{ CI}, 1.03-1.23; P = 0.006, A \text{ vs. } G]$ genetic patterns. With respect to CASP9 rs4645981, the findings from five studies suggested a noteworthy

Fig. 3 The forest plot for the association between CASP9 rs4645981 polymorphism and cancer risk using different genetic contrasted models

association between this variant and increased risk of cancer in the Asian population ($P \le$ 0.001 for all the assessed genetic models) (Table [3\)](#page-11-0).

Stratifying according to cancer type indicated that CASP9 rs4645978 significantly enhanced the risk of developing colorectal cancer (under all examined genetic models), and lung and prostate cancer (under codominant heterozygous and dominant genetic patterns). We also observed an enhanced risk of lung cancer ($n = 4$ studies) regarding codominant homozygous $(OR = 2.25)$, codominant heterozygous $(OR = 1.34)$, dominant $(OR = 1.39)$, and recessive $(OR = 1.92)$ models of *CASP9* rs4645981 polymorphism. The T allele of this polymorphism increased susceptibility to lung cancer by 1.43-fold. Interestingly, the G allele of CASP9 rs1052576 was associated with a diminished risk of developing lung cancer $[OR = 0.70; 95\%]$ CI, 0.56–0.90; $P = 0.004$] (Table [4\)](#page-14-0).

Heterogeneity and Publication Bias

The heterogeneity results of all the studied polymorphisms are summarized in Tables [2](#page-8-0) and [3](#page-11-0). As shown in Fig. [4](#page-16-0), a symmetrical-shaped funnel plot was generated for the association between CASP9 rs4645981 polymorphism and cancer risk, which indicates no publication bias. Regarding CASP9 rs4645981 and rs1052571 polymorphisms, Egger's linear regression analysis detected no publication bias for the current meta-analysis under different genetic models (P values for bias > 0.05). No publication bias was also detected for significant findings of the stratified analysis, except for the relationship between three high-quality studies on CASP9 rs4645981 and cancer susceptibility under the codominant homozygous model (P value for bias $= 0.008$).

Sensitivity Analysis

Regarding TT vs. CC and TT vs. TC + CC models of CASP9 rs4645981, Lee et al.'s study had the most profound impact on pooled ORs (Fig. [5\)](#page-17-0); thus, these findings should be interpreted carefully. As for the other significant results of this meta-analysis, the relevant pooled ORs indicated no significant change in the assessed genetic models (data not shown). Except for TT vs. CC and TT vs. TC + CC models of CASP9 rs4645981, the final summary ORs are reliable and stable.

Results of TSA and FPRP Analyses

Our meta-analysis indicated a significant association between CASP9 polymorphisms (rs4645981 and rs1052571) and overall cancer risk. For rs1052571, rs4645978, and rs1052576, the cumulative Z-curve did not pass the TSA boundary lines illustrating that the cumulative evidence is insufficient. More trials are warranted to confirm the effect of these SNPs on cancer susceptibility. However, the TSA analysis of rs4645981 indicated the crossing of cumulative Z-curve (blue line) over the trial sequential monitoring boundary (dotted red line) ($P < 0.05$), suggesting reliable evidence for the rs4645981 effect on cancer risk.

Table [5](#page-18-0) represents the calculated FPRP values regarding the main significant findings in the present meta-analysis. With the assumption of a prior probability of 0.25, most FPRP values were less than 0.2, indicating the observed significant associations were notable.

CI confidence interval, OR odds ratio, NA not applicable, QA quality assessment CI confidence interval, OR odds ratio, NA not applicable, QA quality assessment

CASP10 rs39010 rs39001-Hodgkin lymphoma 2 AA 4 CG 1.56) 1.56 0.541 Pixed 0 0.541 NA 1.12(0.579; 1.56) 0.420 0.

CI confidence interval, OR odds ratio, NA not applicable CI confidence interval, OR odds ratio, NA not applicable

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Fig. 4 Funnel plot for the association between CASP9 rs4645981 polymorphism and cancer risk using different genetic contrasted models

Discussion

As the role of cell death in the pathophysiology of cancer is gaining ground, it appears crucial to study the genetic variations of the apoptosis-associated gene in human malignancies. For example, variations in apoptosis-related genes (i.e., death receptors and TNF superfamily ligands) have been investigated in hematological malignancies as well as solid tumors [\[58](#page-23-0), [59](#page-23-0)]. Apoptosis, a genetically mediated cell suicide program and an essential physiological response, serves an indispensable role in maintaining tissue hemostasis and discarding harmful and/or unnecessary cells [[60\]](#page-23-0). Characterization of the intrinsic and extrinsic pathways of apoptosis has straightened out how apoptosis is deregulated in most human malignancies [[61](#page-23-0)]. Some studies have recommended that restoring the function of these caspases by overexpressing them could be a beneficial curative approach toward cancer [[62,](#page-23-0) [63](#page-23-0)].

It has been shown that CASP9 is necessary for p53-dependent apoptosis [\[64\]](#page-23-0). Failure of CASP9 activation also induced a higher threshold for apoptotic cell death [\[65](#page-23-0)]. Horn and

Fig. 5 Sensitivity analysis for studies on CASP9 rs4645981 using different genetic contrasted models

coworkers have shown that CASP10 negatively regulates death receptor-mediated activation of CASP8, and therefore perturbing cell death [\[66](#page-23-0)]. Based on another hypothesis, CASP10 makes cancer cells more susceptible to TRAIL-induced apoptosis following a CASP3 dependent fashion [\[67](#page-23-0)]. It is assumed that any alteration in the genetic formation of caspases could potentially impact the rate of apoptosis [[68](#page-23-0)]. Nevertheless, the mechanisms by which caspases, including CASP9 and 10, get regulated are not fully understood. Moreover, CASP10 mutations have been associated with impaired apoptotic function, suggesting that CASP10 deficiency might be the reason why cancer cells evade apoptosis [\[13\]](#page-21-0).

Few SNPs (i.e., rs1052576, rs1052571, and rs2308950) are mapped within the coding region of CASP9 gene. Ulybina et al. had suggested that genetic variations in this region might not be linked to the risk of developing breast cancer [[23\]](#page-21-0). Instead, most studied SNPs in the CASP9 gene have resided in non-coding regions, i.e., promoter site and intronic regions. The rs4661636, rs6685648, and rs2020902 polymorphisms (located within introns of CASP9) or rs4645982 (an insertion/deletion SNP situated in a splice donor site) were shown to be functional and with the potential to alter CASP9 mRNA splicing patterns [\[69,](#page-23-0) [70](#page-23-0)]. Besides, alternative splicing of CASP9 might have different impacts on apoptosis and affect cancer cells' tumorigenicity, as this gene produces two protein isoforms (CASP9a and CASP9b)

through inclusion/exclusion of four exons [\[71,](#page-23-0) [72](#page-23-0)]. Promoteric SNPs of CASP9 (e.g., rs4645981 and rs4645978) were more intensively studied. This evidence shed light on the relevance of transcriptional regulation of this gene [\[31,](#page-21-0) [33](#page-22-0)] since the $G_{rs4645978}$ C_{rs4645981} haplotype showed an elevated promoter activity compared with $G_{rs4645978}$ T_{rs4645981} and $A_{rs4645978}$ C_{rs4645981} combinations [\[5\]](#page-20-0). We then pursued the hypothesis suggesting the SNPs within the CASP10 gene might affect overall cancer susceptibility. It has been shown that the rs13006529 polymorphism impacts the very last amino acid of the protein [\[73](#page-23-0)]. However, this SNP was not correlated with breast cancer incidence [\[48\]](#page-22-0). The rs13010627 resides 5 amino acids upstream of the cleavage site of mature CASP10; hence, this variation could impact CASP10 activation and disrupt its function [[45,](#page-22-0) [74](#page-23-0)].

In 2012, a meta-analysis was carried out by Yan and coworkers on 7 case–control studies, including 3962 subjects. They reported a positive association between the A allele carriers of CASP9 rs1052576 polymorphism and cancer incidence in American and Chinese populations $(OR = 0.63)$; thus, A allele and A allele carriers of this polymorphism had established protective roles against cancer [[75](#page-24-0)]. In our study, we included 14 studies and enrolled 4877 subjects to enhance statistical power. In contrast, we found no relationship between this variant and the overall risk of cancer in Asians or Caucasians. In the same year, another comprehensive meta-analysis was performed by Xu et al. on 9 studies with 5528 subjects for rs4645978, 6 studies with 2403 subjects for rs105276, and 2 studies with 2304 subjects for rs4645981. By pooling the results of included studies, they observed a protective effect for CASP9 rs1052576 under AA vs. GG (OR = 0.75) and A vs. G models (OR = 0.85) against cancer susceptibility. As for CASP9 rs4645981, they found an increased incidence of lung cancer among Asians under allelic (OR = 1.23, T vs. C) and recessive (OR = 1.22, CC vs. CT+TT) contrasted genetic models. However, they found no evidence of the association between CASP9 rs4645978 and overall cancer incidence [[76](#page-24-0)]. In our updated meta-analysis, we enrolled 5910 subjects for examining the link between CASP9 rs4645981 polymorphism and cancer risk and observed significant results under different contrasted models. Still, by including 14 studies for each SNP, we found no association between CASP9 rs1052576 and rs4645978 and the overall risk of cancer.

In 2012, Yan and colleagues reviewed and conducted a meta-analysis on the relationship between CASP10 variants and cancer incidence. They pooled the results from 8 studies with 29,936 cases of cancer and 34,041 healthy subjects. Concerning CASP10 rs13006529, they included 3751 subjects and found that the T allele was associated with a 1.17-fold increased risk of cancer. Simultaneously, the other two CASP10 polymorphisms, rs3900115 and rs13010627, were not associated with cancer risk. Moreover, by performing a stratified analysis, they observed a positive association of CASP10 rs13006529*T carriers with breast cancer incidence (OR=1.17) [\[77\]](#page-24-0). In our study, we enrolled 5648 subjects and found no noteworthy association between this variant and cancer incidence.

Dysregulated apoptosis might lead to tumorigenesis [[78](#page-24-0)]. In this respect, somatic or non-somatic mutations within caspase genes are frequent in a wide range of malignancies [[79\]](#page-24-0). In 2019, Hashemi et al. showed that CASP3 polymorphisms are associated with the overall risk of cancer [[80\]](#page-24-0). One year later, Hashemi and colleagues performed a meta-analysis and showed that common SNPs within CASP8, including rs3769818, rs3769821, rs3769825, rs3834129, and rs1045485, are also correlated with susceptibility to cancer [\[68](#page-23-0)]. In 2013, Yan and colleagues reported that variations in CASP7, another caspase that contributed to cell proliferation and cytokine maturation, are involved in the pathogenesis of cancer [[81](#page-24-0)]. The possible association between variations in other caspases and cancer risk is currently being investigated.

Heterogeneity partly determines the difficulty in drawing overall conclusions, and it might affect the results of our meta-analysis. Hence, our results should be interpreted with attention because of some limitations. This might happen because of differences in the genetic background of the subjects or allele frequencies between the studied populations. Moreover, the observed heterogeneity might be due to the lifestyle and age of cancer patients, dissimilar methods of diagnosis, and cancer types. On the other hand, most of the included studies were conducted on a limited number of populations. Cancer is a multifactorial disorder, and environmental factors play crucial roles in its susceptibility. Despite these limitations, the findings of the pooled analysis provided a conclusive estimate for the impact of CASP9 and CASP10 polymorphisms on cancer susceptibility.

Conclusion

In conclusion, our meta-analysis suggested that CASP9 rs4645981 and rs1052571 polymorphisms are associated with overall cancer risk. However, more studies on larger populations are warranted to validate these associations.

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Declarations

Conflict of Interest The authors declare no competing interests.

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