CHRONIC DAILY HEADACHE (S-J WANG AND S-P CHEN, SECTION EDITORS)



Utility of Repetitive Transcranial Magnetic Stimulation for Chronic Daily Headache Prophylaxis: A Systematic Review and Meta-Analysis

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Abstract

Purpose of Review Management of chronic daily headaches (CDH) remains challenging due to the limited efficacy of standard prophylactic pharmacological measures. Several studies have reported that repetitive transcranial magnetic stimulation (rTMS) can effectively treat chronic headaches. The objective was to determine the utility of rTMS for immediate post-treatment and sustained CDH prophylaxis.

Recent Findings All procedures were conducted per PRISMA guidelines. PubMed, Scopus, Web of Science, and ProQuest databases were searched for controlled clinical trials that have tested the efficacy of rTMS on populations with CDH. DerSimonian-Laird random-effects meta-analyses were performed using the 'meta' package in R to examine the post- vs. pre-rTMS changes in standardized headache intensity and frequency compared to sham-control conditions. Thirteen trials were included with a combined study population of N=538 patients with CDH (rTMS, N=284; Sham, N=254). Patients exposed to rTMS had significantly reduced standardized CDH intensity and frequency in the immediate post-treatment period (Hedges' g=-1.16 [-1.89, -0.43], p=0.002 and Δ =-5.07 [-10.05, -0.11], p=0.045 respectively). However, these effects were sustained marginally in the follow-up period (Hedges' g=-0.43 [-0.76, -0.09], p=0.012 and Δ =-3.33 [-5.52, -1.14], p=0.003). Significant between-study heterogeneity was observed, at least partially driven by variations in rTMS protocols. **Summary** Despite the observed clinically meaningful and statistically significant benefits in the immediate post-treatment period, the prophylactic effects of rTMS on CDH do not seem to sustain with discontinuation. Thus, the cost-effectiveness of the routine use of rTMS for CDH prophylaxis remains questionable.

Registration Protocol preregistered in PROSPERO International Prospective Register of Systematic Reviews (CRD42021250100)

Keywords Chronic daily headache · Prophylaxis · Transcranial magnetic stimulation · Systematic review · Meta-analysis

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Abbreviations

CDH	Chronic daily headaches
СМ	Chronic migraine
CTTH	Chronic tension-type headaches
TCA	Tricyclic antidepressants
BTX-A	Botulinum toxin A
CGRP	Calcitonin gene-related peptide
rTMS	Repetitive transcranial magnetic stimulation
dlPFC	Dorsolateral prefrontal cortex
PRISMA	Preferred Reporting Items for Systematic
	Reviews and Meta-analyses
MCID	Minimal clinically important difference
GRADE	Grading of Recommendations, Assessment,
	Development, and Evaluation
MC	Motor cortex

Introduction

Chronic daily headaches (CDH) are a disabling condition affecting approximately 4-5% of the global population, significantly affecting quality-of-life [1–3]. Resistance of CDH to standard therapies results in substantial social and financial burdens [1, 4, 5]. For instance, chronic migraine (CM) patients incur higher healthcare costs than patients with episodic headaches due to more consultations, psychiatric interventions, hospitalizations, and medication costs [5]. Furthermore, 90% of patients with CDH have psychiatric comorbidities, such as depression, anxiety, or disordered sleep [2, 4], which worsen outcomes [6]. CDH, by definition, lasts \geq 15 days/month for at least three months, for \geq 4 h a day when untreated, which was initially defined by Silberstein and Lipton in 1996 [7]. However, later on, CDH became more of a category of headaches rather than a diagnosis, and the International Classification of Headache Disorders, 3rd edition (ICHD-3) does not define CDH [4, 8]. Chronic tension-type headaches (CTTH) and chronic migraine (CM) comprise most cases of primary CDH and frequently evolve from episodic headaches [1, 4]. Less common types include hemicrania continua and new daily persistent headaches [1]. Secondary CDH may be post-traumatic brain injury (TBI) or medication-induced, and these causes must be identified before diagnosing and treating primary CDH [9].

CDH is often managed prophylactically using antidepressants, anticonvulsants, and antihypertensives [1]. While tricyclic antidepressants (TCA) have shown the greatest potential for reducing headache symptoms, they are poorly tolerated (> 30% experience side effects) [10]. Whereas selective serotonin and norepinephrine-reuptake inhibitors may be more tolerable, evidence regarding efficacy is inconsistent [10]. Similarly, the effectiveness of anticonvulsants is not entirely clear, except for topiramate [11, 12]. Recently, botulinum toxin A (BTX-A) and monoclonal calcitonin

gene-related peptide (CGRP) antibodies have shown promise, as these agents directly target the pathological processes of CDH [1, 9, 13, 14]. BTX-A is deemed safe and effective in preventing CM; however, patient response varies, and sustained effects are not consistently observed in patients with severe, refractory CDH [15]. Furthermore, few studies have documented CGRP antibody use in CDH [13, 14], particularly for CM [16]. Moreover, abortive drugs like non-steroidal anti-inflammatory drugs can induce medication overuse headaches, complicating management [1]. Collectively, pharmacologic treatments have limited efficacyimproving headache duration and frequency in only 10% of CDH patients-and are associated with adverse effects that limit compliance by 25% [17]. Thus, a lack of reliable abortive and prophylactic treatments necessitates establishing clinically effective, evidence-based solutions for CDH.

Repetitive transcranial magnetic stimulation (rTMS) temporally alters cortical neuron excitability through noninvasive neurostimulation [18•]. rTMS has utility in treating psychiatric, movement, and chronic pain disorders and has recently gained attention as a potential treatment for headaches [19–21]. High-frequency rTMS (e.g., 10–20 Hz) results in excitatory effects, which increase cortical excitability and neuronal firing, while low-frequency rTMS (e.g., <1 Hz) is generally inhibitory, suppressing cortical excitability and neuronal activity in target brain regions [22, 23]. Furthermore, the effects of rTMS extend beyond the target area to interconnected brain networks, altering functional connectivity and neuroplasticity within and between networks [21, 24].

The mechanism underlying rTMS for reducing CDH remains unclear. rTMS can influence cortical areas involved in pain processing by modulating cortical excitability [23, 25, 26]. Additionally, rTMS may modulate pain pathways, including descending inhibitory pathways, which influence supraspinal pain tracts and social-affective regions of the brain, like the right temporal lobe [27, 28]. Particularly, rTMS demonstrated analgesic effects in migraine and tension headaches by modulating the dorsolateral prefrontal cortex (dlPFC), which regulates pain perception and pain-related emotional/cognitive processes [18•, 25, 26, 29-31]. Moreover, rTMS may induce analgesia by increasing endogenous opioids such as β endorphin in the anterior cingulate cortex, hypothalamus, and peri-aqueductal gray matter [32, 33], dopamine in the hippocampus and caudate nucleus [34, 35] and glutamate levels in the neocortex [36]. Finally, rTMS may promote neurogenesis and modulate synaptic plasticity in cortex and diencephalon [37]. These neuroplastic changes may underlie the therapeutic effects of rTMS on CDH.

Several studies have compared CDH patients receiving high-frequency rTMS in multiple sessions for several weeks with placebo controls [18•, 29, 38, 39]. Here, we aimed to

conduct a systematic review and meta-analysis of the literature to examine the effectiveness of rTMS in managing CDH, with a focus on its impact on headache intensity and frequency, its effects on CDH subtypes, and factors influencing its efficacy through subgroup and meta-regression analyses. We specifically hypothesized that exposure to regular treatment with rTMS will not only decrease the intensity and frequency of acute headache episodes in the immediate post-intervention period but will also elicit a sustained prophylactic effect among patients with CDH.

Methods

All procedures followed the Preferred Reporting Items for Systematic Reviews and Meta-analyses (PRISMA) guidelines [40]. The study protocol and analysis plan were preregistered in the PROSPERO registry (CRD42021250100).

Search Strategy

PubMed, Scopus, ProQuest, and Web-of-Science databases were searched on June 6, 2021, for peer-reviewed interventional studies using pre-defined keyword combinations (Table S1). The search was not limited by language or year of publication. The records identified from the initial database search were pooled, and the duplicate records were removed using an in-house pipeline used for several prior systematic reviews and meta-analyses. Titles and abstracts of all records were screened by one reviewer (ES) based on pre-defined eligibility criteria (Fig. 1, Table S2). All records were randomly divided among three additional reviewers (AB, AH, or RH) to ensure that at least two reviewers screened each record. Agreements between the reviewers were examined, and a senior tie-breaker (CNK) resolved discrepancies in judgment. Full-text articles of the records that were deemed eligible during initial screening were thoroughly examined for eligibility by two study personnel (ES and CSD), and the discrepancies were resolved by a senior tie-breaker (CNK) (Table S3). An additional systematic search was performed through June 26, 2023, broadening the search to include Cochrane database. A manual search was conducted on the reference lists and citations of eligible articles; 11 additional records were added to the pool. Records published in other languages were translated into English using Google Translate[®]. The full-text articles that



Fig. 1 PRISMA flowchart

met eligibility criteria were included for quality assessment, review, and data extraction for meta-analyses.

Data Extraction

Data were extracted from the eligible manuscripts into predefined data fields in a spreadsheet. Immediate post-intervention (i.e., soon after discontinuation of rTMS or control intervention) versus pre-intervention changes in headache intensity and frequency were extracted as primary outcomes from the intervention and control groups in the included studies. When available, post-follow-up (i.e., re-examining headache at least > 2 weeks after the discontinuation of the rTMS or control intervention) versus pre-intervention changes in headache intensity and frequency of intervention and control groups were also extracted. Quality of life measures were extracted as secondary outcomes. Migraine Disability Assessment (MIDAS) and Headache Impact Test-6 (HIT-6) were extracted as measures of quality of life [41]. Year of publication, mean age, percentage of females in the intervention group, types of CDH observed, mean duration of headache, duration of intervention, frequency of rTMS sessions per week, length of follow-up, and total number of sessions (i.e., exposure to rTMS) in a given study were extracted. Additional data regarding the rTMS protocol (i.e., rTMS pulse frequency, total rTMS pulses delivered per session, the intensity of magnetic impulses, the anatomical site used for rTMS, type of coil used for rTMS pulse delivery, and the equipment used) were also extracted to include in meta-regression analyses.

Data Analysis

Four separate DerSimonian-Laird random-effects metaanalyses were performed using the 'meta' package (version 4.11-0) in R software (version 4.0.3) to examine the immediate post- versus pre-intervention changes and postfollow-up versus pre-intervention changes in standardized headache intensity and frequency of CDH following rTMS. When interpreting the findings, a conservative estimate of the standardized between-group difference of 0.5 was considered the threshold for minimal clinically important difference (MCID) [42, 43]. In addition, a decrease in headache frequency by at least 1 day/month was considered the MCID for interpretation of outcomes regarding headache frequency [44, 45]. When sufficient data were available, subgroup meta-analyses were performed to explore the effects of rTMS on each primary outcome variable within the patient populations of subtypes of CDH (i.e., CM, TBI, or CTTH). Leave-one-out sensitivity analyses confirmed the consistency of findings. Additionally, a series of subgroup analyses were conducted to explore the effects of frequency subtypes and anatomical site of rTMS application on primary outcomes. The likelihood of publication bias was explored using funnel plots, and symmetry was assessed using Egger's tests [46]. Effect-sizes of missing (i.e., unpublished/unreported) studies were imputed via the trim-and-fill method. Heterogeneity of effect-sizes was quantified with the Higgins' I² statistic [47, 48]. Exploratory univariate random-effects meta-regression analyses were performed to explain heterogeneity using potential moderator variables described above [49].

Quality Check and Grading Quality of Evidence

The risk of bias was assessed within the individual studies using the Cochrane Collaboration's Tool RoB 2: A revised Cochrane risk-of-bias tool for randomized trials [50]. The quality of the evidence of the short-term and long-term outcomes was assessed according to the Grading of Recommendations, Assessment, Development, and Evaluation (GRADE) methodology for risk of bias, inconsistency, indirectness, imprecision, and publication bias [51]. Each quality criterion was rated very low, low, moderate, or high. Summary tables were constructed using the GRADE Profiler (GRADEpro, version 3.6) [52].

Results

Study Characteristics

The PRISMA flow diagram depicting the outcomes of the database search and screening of the literature based on the eligibility criteria is shown in Fig. 1. Fourteen randomized controlled trials (15 study arms) examining the effects of rTMS on the intensity and frequency of headache episodes among a total of 538 patients with CDH (284 exposed to rTMS; 254 exposed to sham control interventions) were included in the systematic review (Table 1) [18•, 29, 30, 38, 39, 53–60, 61••]. The sample sizes of rTMS intervention groups of the included studies ranged from 6 to 52, and the sample sizes of the control groups ranged from 6 to 46. The percentage of female participants in the rTMS groups of the included studies ranged from 17–87%, and the mean age of the participants of the rTMS groups was 40.19 years (range 32.93–53.28 years).

rTMS Protocols

When the protocols of rTMS were considered, only three studies used low-frequency rTMS, whereas the rest used highfrequency rTMS. The frequency of rTMS sessions per week ranged from 3–5 sessions per week, with a total duration of treatment ranging from 1–8 weeks. Therefore, the mean total number of sessions was 12.4 (range 3–24 rTMS sessions),

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Table 1 Sum	umary of artic	cles included in the s	systematic review	', meta-analyses	and meta-1	regre	sion	analyses						
Author	Age (years), Female (%)	Diagnosis, Type, Headache duration, Headache frequency	rTMS Parameters	Control	Measure	z	C J	Colerability and ide effects	Presence of Psychiatric comorbidities	Analysis	Headache Intensity	Headache Frequency	Headache Intensity follow up	Headache Frequency follow up
AbdElkader et al. (2001) [18•]	34.19, 81.3	IHS, CTTH & CM, 22, Chronic persistent TTH or migraine for > 3 months	 5 Hz, 100 pulses, 3 sessions/week for 4 weeks, 90%RMT, using a figure of eight, MagStim Rapid magnetic stimulator to L/ dlPFC 	Sham (same coil placed perpendicularly to vertex, same intensity and protocol)	4-point Likert scale	16		No side effects	Not assessed	IIA	×	×		
Brighina et al. (2004) [29]	47, 63.6	IHS, CM, NA,> 15 days per month for > 3 months	20 Hz, 400 pulses, 3 sessions/week for 4 weeks, 90%RMT using a figure of eight Cadwell high-frequency magnetic stimulator to L/ dlPFC	Sham (same coil placed perpendicularly to vertex, same intensity and protocol)	4-point Likert scale	Q	5	Vell tolerated No side effects	Not assessed	Ч		×		×
Conforto et al. (2014) [38]	41.4, NA	ICHD-2, CM, 16.7, > 15 days per month for > 3 months	10 Hz, 1600 pulses, 3 sessions/week for 8 weeks, 110% RMT using a figure of eight. Alpine Biomed MagPro X100 to L/dlPFC	Sham (same coil placed perpendicularly to vertex)	Diary		A 1	Vorsening or onset of headache or local pain (rTMS group-78%: and sham group- 33%) Sleepiness in both groups (78%)	BDI and STAI scores did not show any improvement	IF	×	×		
Granato et al. (2019) [39]	46.1, NA	NA, CM, 17.6, > 15 days per month for > 3 months	20 Hz, 400 pulses, 5 sessions/week for 2 weeks, 100%RMT using a figure of eight coil to L/dIPFC	Sham (a figure- of-eight coil and a sham stimulator, able to induce the same skin vibratory sensation)	Ч N	1	N L	digraine with complex aura (14%) Mild disconfort (rTMS group- 29%)	Not assessed	Completers				×

Table 1 (con	ntinued)													
Author	Age (years), Female (%)	Diagnosis, Type, Headache duration, Headache frequency	rTMS Parameters	Control	Measure	z	C Toler side e	ability and effects	Presence of Psychiatric comorbidities	Analysis	Headache Intensity	Headache Frequency	Headache Intensity follow up	Headache Frequency follow up
Kalita et al. (2016) [30]	33.4, 80.8	ICDH-2R, CTTH+CM, 8.89,>15 days per month for > 3 months	 5 Hz, 100 pulses, 3 sessions/week for 4 weeks, 90%RMT using a figure of a figure of eight Mag Stim Rapid magnetic stimulator to L/ dIPFC 	l rTMS + 2 Sham (a figure of eight sham coil producing similar sound)	Diary	52	46 Well All re acc dis (10	tolerated eported ceptable comfort due noise of rTMS 00%)	Not assessed	Ē			×	×
Kumar et al. (2021) [53]	33.2, 60	ICHD-3, CM, 7.3,> 10 episodes per month for > 3 months	10 Hz, 600 pulses, 5 sessions/week for 2 weeks, 70%RMT using a figure of eight, Neurosoft Neuros	Sham (same coil placed perpendicularly to vertex at minimum stimulation strength of the stimulator)	VAS	10	IO Well No si	tolerated ide effects	Not assessed	AI	×	×	×	×
Leung et al. (2016) [54]	41.2, 16.7	ICHD-2, TBI, 14.8, Chronic persistent (24/7) for > 3 months	10 Hz, 2000 pulses, 3 sessions/week for 1 week, 80%RMT using a figure of eight, Alpine Biomed MagPro R30 to L/MC	Sham (same coil placed 180 degrees away from scalp with coil side facing scalp shielded with a molded cover containing two layers of Giron magnetic shielding film)	NPRS	12	12 Trans ten (rT gro diz gro	sient local derness iMS ups-8.3%) ups-8.3%) ups 8.3%)	Moderate levels of depression and higher interference in daily life in whole cohort No difference in HAM-D between groups	Completers	×		×	

	Headache Frequency follow up	×		
	Headache Intensity follow up	×		
	Headache Frequency	×		
	Headache Intensity	×	×	×
	Analysis	Completers	IF	AI
	Presence of Psychiatric comorbidities	Very severe degree of depression in whole cohort A significant improvement in HAM-D score	Significant improvement in SATI scores No improvement in WHOQOL scores and HAM-D	Not assessed
	Tolerability and side effects	Well tolerated	Well tolerated No side effects reported	Well tolerated No side effects reported
	C	15	15	10
	Z	14	15	10
	Measure	NPRS	NPRS	VAS
	Control	Sham (same coil placed 180 degrees away from scalp with coil side facing scalp shielded with a molded cover containing two layers of Giron magnetic shielding film)	Sham (same coil placed perpendicularly at least stimulation strength)	Sham (same coil placed perpendicularly without magnetic field penetration to brain)
	rTMS Parameters	10 Hz, 2000 pulses, 4 sessions/week for 1 week, 80%RMT using a figure of eight, Alpine Biomed MagPro R31 to L/dIPFC	1 Hz, 1200 pulses, 5 sessions/week for 4 weeks, 110%RMT using a figure of eight, Neurosoft NeuroMS/D TMS device to R/dIPFC	1 Hz, 5 sessions/ week for 2 weeks, 110%RMT using a figure of eight, Neurosoft NeuroSoft NeuroMS/D TMS device to R/dIPFC
	Diagnosis, Type, Headache duration, Headache frequency	ICHD-2, TBI, 7.9, Chronic persistent (24/7) for > 3 months	ICDH-3, CTTH, 6, > 15 days per month for > 3 months	IHS, CTTH, 6.35,>4 h with moderate intensity
ntinued)	Age (years), Female (%)	32.93, 85.7	34.7, 86.7	37.5, 60
Table 1 (con	Author	Leung et al. (2017) [55]	Mattoo et al. (2019) [56]	Rajain et al. (2023) [59]

Table 1 (con	ntinued)													
Author	Age (years), Female (%)	Diagnosis, Type, Headache duration, Headache frequency	rTMS Parameters	Control	Measure	z	C Si I	olerability and ide effects	Presence of Psychiatric comorbidities	Analysis	Headache Intensity	Headache Frequency	Headache Intensity follow up	Headache Frequency follow up
Rapinesi et al. (2016) [57]	53.28, 71.4	ICHD-3, CM, NA, > 15 days per month for > 3 months	10 HZ, 360 pulses, 3 sessions/week for 4 weeks, 100%RMT using a H1 deep TMS coil, Brainsway H1 coil deep TMS System to L/ dIPFC, mPFC, OFC	Standard treatment (medication)	VAS	7		vo side effects	Significant reduction in HAM-D scores	All	×	×	×	×
Stilling et al. (2020) [58]	40.3, 90	ICHD-3, TBI, 24,>15 days per month for > 3 months	10 Hz, 600 pulses, 5 sessions/week for 2 weeks, 70%RMT using a figure of eight Magstim Super Rapid2 to L/dIPFC	Sham (air-film coil applied to sealp using same location and stimulation protocol Participants heard sound and felt vibration	NPRS	10	0 1	Aild aggravation of headache (4.23%), scalp discomfort (0.96%), toothache (0.675%), and dizziness (0.30%)	Significant improvement in PHQ-9 and QOLIBRI	AII	×	×	×	×
Todorov et al. (2018) [61 ●]	38.7, 81.1	ICHD-3, CM, 19.3, 24.0 days/ month for the last 3 months	15 Hz, 1200 pulses, 5 sessions/week for 1 week, 70%RMT using a figure of eight coil to L/ dIPFC	Sham (same coil placed perpendicularly at same parameters)	VAS	37	28 L	dlPFC-59.5% dlPFC-59.5% and sham-32.1%) Generally well tolerated	Not assessed	AII			×	×
Todorov et al. (2018) [61••]	40.2, 81.6	ICHD-3, CM, 19.3, 25.4 days/ month for the last 3 months	15 Hz, 1200 pulses, 5 sessions/week for 1 week, 70%RMT using a figure of eight coil to L/MC	Sham (same coil placed positioning coil perpendicularly at same parameters)	VAS	38	28 I	ransient dizziness and short- lasting increase in headache intensity 39.5%)	Statistically significant improvement in HAM-A and HAM-D scores	AII			×	×

Table 1 (coi	utinued)													
Author	Age (years), Female (%)	Diagnosis, Type, Headache duration, Headache frequency	rTMS Parameters	Control	Measure	z	C	Tolerability and side effects	Presence of Psychiatric comorbidities	Analysis	Headache Intensity	Headache Frequency	Headache Intensity follow up	Headache Frequency follow up
Wei et al. (2023) [60]	53.95, 60.5	ICHD-3, CTTH, 10.26, 23.81 h for 19.93 attacks/ month	1 Hz, 5 sessions/ week for 4 weeks, 60%RMT using a figure of eight, MagPro R30 type to R/ dIPFC	Sham (pulse emission without magnetic field)	VAS	43	43	No side effects	Not assessed	IIA	×		×	
<i>rTMS</i> repeti cortex, <i>RM1</i> national Cla Depression I	ive transcram resting motc ssification of tating Scale,	ial magnetic stimula r threshold, <i>CTTH</i> Headache Disorder <i>WHOOOL</i> World H	ation, N sample s chronic tension- rs, VAS visual an lealth Organizatio	ize of the intervi- type headache, (alogue scale; <i>NI</i>	ention gro <i>CM</i> chroni <i>PRS</i> Nume e, <i>HAM-A</i>	up, C c mi eric F Ham	sam grain ain	ple size of the c e, <i>TBI</i> traumation Rating Scale, <i>Bi</i> Anxiety Scale,	control group wh c brain injury, <i>H</i> <i>DI</i> Beck Depres <i>PHO-9</i> Personal	nen present, <i>A</i> <i>HS</i> Internatic sion Index, <i>S</i> Health Oues	<i>AC</i> Motor cc mal Headach <i>TAI</i> State A tionnaire-9.	ortex, dlPFC he Society c anxiety Inver OOLIBRI O	Dorsolate Titeria, <i>ICH</i> Dory, <i>HAM</i>	al pre frontal <i>ID</i> The Inter- <i>ID</i> Hamilton fe after Brain

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and the mean rTMS pulses per session was 989.3 (range 100–2000). The intensity of magnetic impulses was ranging from 60–110%RMT. Nine study arms used left dlPFC, three studies used right dlPFC, and three study arms used left motor cortex (L/MC) as sites of stimulation. Out of all the included studies, ten studies used angle manipulation as a sham, whereas only three studies used a sham coil stimulator.

Quality of Studies

The quality of the included RCTs was evaluated using the Cochrane risk of bias assessment tool (Fig. S1). Six out of the 14 studies included were considered high quality, three were considered to be of moderate quality, and five were low quality. Quality ratings and the risks of bias of included studies are summarized in Fig. S1. The GRADE evidence profiles for the considered outcomes are summarized in Table S4. The GRADE Working Group grades of the level of evidence were low for all considered variables.

Effects of rTMS on the Intensity of Headaches in the Immediate Post-intervention Period

Compared to sham control interventions, exposure to rTMS significantly decreased the standardized headache intensity in the immediate post-intervention period versus pre-intervention state among patients with CDH (10 studies [18•, 38, 53–60], pooled Hedges' g=-1.16 [-1.89, -0.43], p=0.002; Fig. 2A), which exceeded the MCID threshold of 0.5. Subgroup analyses revealed a significant reduction in standardized headache intensity following rTMS in patients with TBI and CTTH (pooled Hedges' g=-0.58 [-1.05, -0.11], p=0.016 and pooled Hedges' g=-1.73 [-2.63; -0.82], p<0.001, respectively). However, there was no significant difference in post-intervention standardized headache intensity between patients with CM who received rTMS and those who received a sham control intervention (pooled Hedges' g=-0.90 [-3.37, 1.57], p=0.475).

Additional subgroup analyses were performed, separating studies conducted using high-frequency rTMS and low-frequency rTMS. There was a clinically and statistically significant reduction in headache intensity following exposure to low-frequency rTMS, while no significant difference was noted with the implementation of high-frequency rTMS (pooled Hedges' g = -1.73 [-2.63, -0.82], p < 0.001 and pooled Hedges' g = -0.87 [-1.81, 0.06], p = 0.067, respectively; Fig. S2). When the site of rTMS application was considered, right dlPFC and L/MC had statistically significant reductions in headache intensity (pooled Hedges' g = -1.73 [-2.63, -0.82], p < 0.001 and pooled Hedges' g = -1.73 [-2.63, -0.82], p < 0.001 and pooled Hedges' g = -1.73 [-2.63, -0.82], p < 0.001 and pooled Hedges' g = -1.73 [-2.63, -0.82], p < 0.001 and pooled Hedges' g = -1.73 [-2.63, -0.82], p < 0.001 and pooled Hedges' g = -1.73 [-2.63, -0.82], p < 0.001 and pooled Hedges' g = -0.73 [-2.63, -0.82], p < 0.001 and pooled Hedges' g = -0.73 [-2.63, -0.82], p < 0.001, respectively; Fig. S3), whereas left dlPFC did not show a difference (pooled Hedges' g = -0.78 [-2.18, 0.61], p = 0.272).

A. Study	Total	Mean	rTMS SD	Total	Mean	Control SD	Standardised Mean Difference	SMD	95%-CI	Weight (common)	Weight (random)
Headache = CM & CTTH AbdElkader et al. (2021)	16	-1.63	0.7867	11	-0.19	0.6288		-1.92	[-2.87; -0.98]	7.8%	10.2%
Headache = CM Conforto et al. (2014) Kumar et al. (2021) Rapinesi et al. (2016) Common effect model Random effects model Heterogeneity: $l^2 = 90\%$, r^2	7 10 7 24 = 4.31	-3.60 -3.80 -37.86	4.3655 2.1387 10.4497	7 10 7 24	-14.73 -1.10 0.00	9.2072 1.6972 13.6326		1.45 -1.34 -2.92 -0.72 -0.90	[0.23; 2.67] [-2.33; -0.35] [-4.56; -1.28] [-1.41; -0.02] [-3.37; 1.57]	4.7% 7.1% 2.6% 14.4%	9.2% 10.0% 7.6% 26.8%
Headache = TBI Leung et al. (2016) Leung et al. (2017) Stilling et al. (2020) Common effect model Random effects model Heterogeneity: $J^2 = 5\%$, $\tau^2 =$	12 14 10 36	-3.58 -1.40 -0.61	2.9025 2.2782 1.5660 0.35	12 15 10 37	-1.10 0.00 -0.52	2.1148 1.7146 1.1196	***	-0.94 -0.68 -0.06 -0.58 -0.58	[-1.79; -0.09] [-1.43; 0.07] [-0.94; 0.81] [-1.05; -0.11] [-1.05; -0.11]	9.6% 12.3% 9.1% 31.0%	10.6% 10.9% 10.5% 31.9%
Headache = CTTH Mattoo et al. (2019) Rajain et al. (2023) Wei et al. (2023) Common effect model Random effects model Heterogeneity: $I^2 = 66\%, \tau^2$	15 10 43 68 = 0.44	-4.60 -4.90 -1.79	2.3409 1.6593 1.2650	15 10 43 68	-1.20 -0.05 -0.30	2.0778 1.4333 1.1444	\$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$	-1.49 -3.00 -1.22 -1.43 -1.73	[-2.32; -0.67] [-4.35; -1.64] [-1.69; -0.76] [-1.81; -1.04] [-2.63; -0.82]	10.3% 3.8% 32.6% 46.8%	10.7% 8.7% 11.7% 31.1%
Common effect model Random effects model Heterogeneity: $l^2 = 77\%$, τ^2 Test for subgroup difference Test for subgroup difference	144 = 1.12 es (fixe es (rand	09, p < 0 d effect) dom effe	0.01 : $\chi_3^2 = 11.4$ ects): $\chi_3^2 =$	140 14, df = 9.13, d	3 (p < 0 f = 3 (p =	.01) = 0.03)	-4 -2 0 2 4	-1.10 -1.16	[-1.37; -0.84] [-1.89; -0.43]	100.0% 	 100.0%

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Study	Total	Mean	rTMS SD	Total	Mean	Control SD	Standardised Mean Difference	SMD	95%-CI	Weight (common)	Weight (random)
										(,	(,
Headache = CM & CTTH											
Kalita et al. (2016)	52	-1.70	1.5548	46	-2.10	1.9566		0.23	[-0.17; 0.62]	24.2%	15.9%
Usedeska – CM											
Kumar et al. (2021)	10	3 20	2 3008	10	1 00	1 6627		1.05	[1 00· 0 10]	1 3%	7 7%
Raninesi et al. (2016)	7	-25.00	10 8625	7	-5.00	11 1619 -		-1.05	[-1.99, -0.10]	2.3%	5.1%
Todorov et al. (2018) dIPEC group	38	-1 60	2 2826	28	-1 20	2 1707	1	-0.18	[-0.67, 0.31]	16.0%	14.2%
Todorov et al. (2018) M1 group	38	-2.10	2.4863	28	-1.20	2.1707		-0.38	[-0.87: 0.12]	15.8%	14.2%
Common effect model	93			73			-	-0.45	[-0.76; -0.13]	38.5%	
Random effects model								-0.61	[-1.15; -0.07]		41.2 %
Heterogeneity: $I^2 = 54\%$, $\tau^2 = 0.1638$,	p = 0.0	9									
Headache = TBI											
Leung et al. (2016)	12	-2.49	3.1403	12	-1.30	2.2428		-0.42	[-1.23; 0.39]	5.8%	9.3%
Leung et al. (2017)	14	-1.30	2.2782	15	-0.10	1.8439		-0.56	[-1.31; 0.18]	6.9%	10.1%
Stilling et al. (2020)	10	-0.21	1.7736	10	-0.41	1.1792		0.13	[-0.75; 1.00]	5.0%	8.5%
Common effects model	30			37				-0.32	[-0.79; 0.14]	17.7%	27.09/
Heterogeneity: $J^2 = 0\%$, $z^2 = 0$, $p = 0.4$	g							-0.32	[-0.79; 0.14]		21.9%
Therefore the second s	0										
Headache = CTTH											
Wei et al. (2023)	43	-3.07	1.4751	43	-1.88	1.2247		-0.87	[-1.31; -0.43]	19.5%	15.1%
Common effect model	224			199			*	-0.35	[-0.54; -0.15]	100.0%	
Random effects model								-0.43	[-0.76; -0.09]		100.0%
Heterogeneity: $I^2 = 63\%$, $\tau^2 = 0.1412$,	p < 0.0	1									
Test for subgroup differences (fixed ef	fect): χ	$\frac{5}{3} = 13.72$	$2, df = 3 (\mu)$	o < 0.01	l)		-2 -1 0 1 2				
Test for subgroup differences (random	i effects	s): χ ₃ = 1	4.22, df =	3 (p <	0.01)						

Fig. 2 Forest plots depicting results of the random-effects metaanalyses examining the effects of rTMS on A. the intensity of headaches in the immediate post-intervention period and B. the intensity of headaches after a standard-care follow-up period. Black dots and the horizontal line indicate individual trial-specific estimates and their 95% CIs; the size of the grey squares denotes the weight of the trials in the meta-analysis. The pooled estimates and the corresponding 95% CI are denoted by the center of diamonds and the width of the diamonds, respectively



Fig.3 Funnel plots depicting publication bias in the literature included in the random-effects meta-analysis examining the effects of rTMS on \mathbf{A} . the intensity of headaches in the immediate post-intervention period, \mathbf{B} . the frequency of headaches in the immediate post-intervention period, \mathbf{C} . the intensity of headaches after a standard-

Leave-one-out sensitivity analyses did not significantly change the pooled estimates (Table S5). A funnel plot of the effect-sizes indicated minimal publication bias (Fig. 3A, Table 2). Imputing one effect-size using the trim-and-fill method to adjust for publication bias and re-analyzing the data using the imputed effect-size revealed a significant decrease in immediate post-verses pre-intervention headache intensity following exposure to rTMS versus

care follow-up period, and **D**. the frequency of headaches after a standard-care follow-up period. Effect-sizes of studies included in the literature are shown in gray-colored circles. Effect-sizes imputed using the trim-and-fill method to maintain funnel plot symmetry to correct for publication bias are shown in white-colored circles

control conditions (pooled Hedges' g = -1.00 [-1.75, -0.24], p=0.010, Table 2). Significant heterogeneity was observed among the standardized mean differences pooled in the random-effects model (τ^2 = 1.121; I² = 77.5%, p < 0.001). None of the considered covariates showed significant moderator effects in a series of univariate meta-regression analyses that attempted to explain between-study heterogeneity (Table S6).

DV	Funnel Asymmetry	Egger's test t and p-value	Number of Imputes	Random-effects model
Headache intensity immediate post- intervention period	Symmetric	t = -0.28, p = 0.788	1	pooled Hedges' g=-1.00 [-1.75, -0.24], p=0.010
Headache frequency immediate post-intervention period	Asymmetric	t = -1.69, p = 0.152	2	$\Delta = -1.64$ [-7.84, 4.56], p=0.604
Headache intensity after a routine care follow-up	Symmetric	t=-1.37, p=0.212	2	pooled Hedges' g=-0.29 [-0.64, 0.06], p=0.100
Headache frequency after a routine care follow-up	Symmetric	t = -1.03, p = 0.337	2	$\Delta = -2.16$ [-4.68, 0.35], p=0.092
Quality of Life	Asymmetric	t=0.48, p=0.653	2	pooled Hedges' g=-1.04 [-1.69, -0.39], p=0.002

Table 2 Summary of assessments for publication bias and adjusted outcomes

Effects of rTMS on the Frequency of Headache Episodes in the Immediate Post-intervention Period

Exposure to rTMS decreased the frequency of headaches by approximately 5 days/month during the immediate postintervention period compared with a sham control intervention. This difference was statistically and clinically significant (8 studies—9 study arms $[30, 53-55, 57, 58, 60, 61 \bullet \bullet], \Delta = -5.07$ [-10.05, -0.11], p=0.045, MCID=1 day/month, Hedges' g=-0.85 [-1.53 -0.17], p=0.014, MCID=0.5; Fig. 4A). On subgroup analyses, this rTMS-associated reduction in the frequency of headaches was more marked among patients with CM (Δ =-6.51 [-12.93, -0.10], p=0.047) but was not significant among patients with TBI ($\Delta = -0.89$ [-2.37, 0.58], p=0.234). Leave-one-out sensitivity analyses did not significantly change the pooled estimates (Table S5). The funnel plot indicated possible publication bias (Fig. 3B, Table 2). When the data were re-analyzed, including two effect-sized imputed using the Trim-and-fill method to correct for funnel plot asymmetry, the previously observed significant post-versus preintervention decrease in headache frequency was not observed $(\Delta = -1.64 [-7.84, 4.56], p = 0.604)$. Significant betweenstudy heterogeneity remained a concern ($\tau^2 = 37.33$; $I^2 = 86\%$, p < 0.001). Attempts to explain the significant between-study heterogeneity in a series of univariate meta-regression analyses remained unfruitful (Table S6). Subgroup analyses were not conducted for rTMS frequency and site of application due to a limited number of studies.

Effects of rTMS on the Intensity of Headaches After a Routine Care Follow-up

The meta-analysis that compared the standardized postfollow-up versus pre-intervention changes in headache intensity between rTMS versus sham control groups revealed a statistically significant yet clinically not significant (less than MCID=0.5) beneficial effect (7 studies [$18\bullet$, 29, 30,

38, 55, 57, 58], pooled Hedges' g = -0.43 [-0.76, -0.09], p = 0.012; Fig. 2B). Subgroup analyses performed within patient populations with CM had significant rTMS versus control differences, whereas CDH due to TBI also did not yield significant rTMS versus control differences (pooled Hedges' g = -0.62 [-1.15, -0.07], p = 0.028 and pooled Hedges' g = -0.32 [-0.79, 0.14], p = 0.173 respectively). Leave-one-out sensitivity analyses did not change the observed pooled effectsize (Table S5). The funnel plot suggested minimal publication bias (Fig. 3C, Table 2). Statistical significance in the follow-up versus pre-intervention difference in headache intensity was no longer observed after adjusting for publication bias by imputing two effect-sizes to restore funnel plot symmetry (pooled Hedges' g=-0.29 [-0.64, 0.06], p=0.100). Significant between-study heterogeneity remained a concern ($\tau^2 = 0.14$, $I^2 = 63.1\%$, p=0.006). The duration of intervention and total number of sessions had negative moderator effects ($\beta = -0.273$, SE=0.112, p=0.015 and β =-0.049, SE=0.022, p=0.027, respectively), indicating that an increased duration of rTMS and increased number of sessions seem to result in a longlasting decrease in headache intensity after the cessation of the rTMS intervention. Furthermore, including either the duration of intervention or the total number of sessions in the model decreased the residual heterogeneity ($\tau^2 = 0.054$, $I^2 = 34\%$, p=0.152 and $\tau^2=0.053$, $I^2=33\%$, p=0.12, respectively) (Table S6).

Effects of rTMS on the Frequency of Headache Episodes After a Routine Care Follow-up

The frequency of headaches after a routine care follow-up period following an rTMS intervention decreased by approximately 3 days/month compared to exposure to a sham control condition (8 studies—9 study arms [29, 30, 39, 53, 55, 57, 58, 61••], Δ =-3.33 [-5.52, -1.14], p=0.003, MCID=1 day/month, Hedges' g=-0.54 [-0.91, -0.17], p=0.004, MCID=0.5; Fig. 4B). The post-follow-up versus

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А.											
Study	Total	Mean	rTMS SD	Total Mean	Control SD	Mean Difference	MD	95%-CI	Weight (common)	Weight (random)	
Headache = CM											
AbdElkader et al. (2021)	16	-20.31	6.5769	11 -1.64	8.5043	.	-18.67 [-24	4.64; -12.70]	3.1%	13.8%	
Brighina et al. (2004)	6	-14.20	6.5529	5 -7.50	3.4132		-6.70 [-1	2.74; -0.66]	3.0%	13.7%	
Conforto et al. (2014)	7	-2.34	2.1431	7 -3.24	3.4924	- <u>-</u>	0.90 [·	-2.14; 3.94]	11.9%	16.2%	
Kumar et al. (2021)	10	-7.20	6.0767	10 0.40	4.9108	i	-7.60 [-1	2.44; -2.76]	4.7%	14.8%	
Rapinesi et al. (2016)	7	-4.72	0.9713	7 -2.56	2.5703		-2.16 [-	4.20; -0.12]	26.6%	16.7%	
Common effect model	46			40		\diamond	-3.25 [-	4.74; -1.75]	49.3%		
Random effects model							-6.51 [-1	2.93; -0.10]		75.2%	
Heterogeneity: $I^2 = 90\%$, τ^2	² = 48.1	1456, p <	0.01								
Headache = TBI											
Loung at al. (2017)	11	0.00	1 0 4 5 2	15 0.02	2 2220		0.00 1	0.07. 0.501	E0 00/	17 00/	

Headache = TBI																
Leung et al. (2017)	14	-0.92	1.8453	15 -0	0.03	2.2230						-0.89	[-2.3	7; 0.59]	50.0%	17.0%
Stilling et al. (2020)	10	-3.70	13.3608	10 -2	2.50 1	6.3677						-1.20	[-14.30	; 11.90]	0.6%	7.8%
Common effect model	24			25					\diamond			-0.89	[-2.3]	7; 0.58]	50.7%	
Random effects model									\diamond			-0.89	[-2.3	7; 0.58]		24.8%
Heterogeneity: $I^2 = 0\%$, $\tau^2 =$	0, p =	0.96														
Common effect model	70			65					\$			-2.06	[-3.10); -1.01]	100.0%	
Random effects model								\langle	\geq			-5.08	[-10.0	5; -0.11]		100.0%
Heterogeneity: $I^2 = 86\%$, $\tau^2 =$	= 37.3	298, p <	< 0.01					1		1						
Test for subgroup difference	s (fixe	d effect)	$\chi_1^2 = 4.84$	df = 1 (p)	o = 0.03	3)	-20	-10	0	10	20					

Test for subgroup differences (fixed effect): χ_1^2 = 4.84, df = 1 (p = 0.03) -20 Test for subgroup differences (random effects): $\chi_1^2 = 2.80$, df = 1 (p = 0.09)

B. Study	Total	Mean	rTMS SD	Total	Mean	Control SD	Mean Difference	MD	95%-CI	Weight (common)	Weight (random)
Headache = CM Brighina et al. (2004) Granato et al. (2019) Kalita et al. (2016) Kumar et al. (2021)	6 7 52 10	-12.90 -3.00 -10.84 -7.60	6.5529 10.5830 8.5831 6.0466	5 7 46 10	-5.80 -7.00 -11.30 0.90	3.4132 5.2915 9.1356 4.9700		-7.10 4.00 0.46 -8.50	[-13.14; -1.06] [-4.77; 12.77] [-3.06; 3.98] [-13.35; -3.65]	2.8% 1.3% 8.2% 4.3%	8.0% 4.8% 13.3% 10.2%
Rapinesi et al. (2016) Todorov et al. (2018) dIPFC group Todorov et al. (2018) M1 group Common effect model Random effects model Heterogeneity: $l^2 = 62\%$, $\tau^2 = 7.0107$,	7 38 38 158 p = 0.0	-2.72 -6.00 -6.80	1.2785 6.6198 7.2865	28 28 131	-0.90 -0.90	2.5356 8.1934 8.1934		-2.87 -5.10 -5.90 -3.56 -3.90	[-4.97; -0.77] [-8.79; -1.41] [-9.72; -2.08] [-4.94; -2.19] [-6.47; -1.33]	23.0% 7.5% 7.0% 54.0%	17.2% 12.9% 12.6%
Headache = TBI Leung et al. (2017) Stilling et al. (2020) Common effect model Random effects model Heterogeneity: $I^2 = 0\%$, $\tau^2 = 0$, $\rho = 0.5$	14 10 24	-1.06 -5.20	1.8005 13.2816	15 10 25	0.06 -3.30	2.2975 16.3677 -		-1.12 -1.90 -1.13 -1.13	[-2.62; 0.38] [-14.96; 11.16] [-2.62; 0.36] [-2.62; 0.36]	45.4% 0.6% 46.0%	18.6% 2.5% 21.1%
Common effect model Random effects model Heterogeneity: $J^2 = 62\%$, $\tau^2 = 6.1103$, Test for subgroup differences (fixed e	182 <i>p</i> < 0.0 ffect): χ	$1^{2}_{1} = 5.56$, df = 1 (p	156 = 0.02)			-10 -5 0 5 10	-2.45 -3.33	[-3.45; -1.44] [-5.52; -1.14]	100.0% 	 100.0%

Test for subgroup differences (random effects): $\chi_1^2 = 3.33$, df = 1 (p = 0.07)

Fig.4 Forest plots depicting results of the random-effects metaanalyses examining the effects of rTMS on A. the frequency of headaches in the immediate post-intervention period and B. the frequency of headaches after a standard-care follow-up period. Black dots and the horizontal line indicate individual trial-specific estimates

and their 95% CIs; the size of the grey squares denotes the weight of the trials in the meta-analysis. The pooled estimates and the corresponding 95% CI are denoted by the center of diamonds and the width of the diamonds, respectively

pre-intervention changes between rTMS versus control conditions were statistically significant within the subgroups of patients with CM but not significant for TBI ($\Delta = -3.90$ [-6.47, -1.33], p = 0.003 and $\Delta = -1.13$ [-2.62, 0.36], p = 0.136, respectively). Leave-one-out sensitivity analyses did not significantly change the pooled outcome (Table S5). The funnel plot appeared symmetrical and did not suggest publication bias (Fig. 3D, Table 2). Nevertheless, the statistical significance of follow-up versus pre-intervention difference in headache frequency was lost with the addition of two effect sizes imputed based on trim-and-fill analysis $(\Delta = -2.16 [-4.68, 0.35], p = 0.092)$. Significant betweenstudy heterogeneity was a concern for this meta-analysis as well ($\tau^2 = 6.110$; I² = 62.4%, p = 0.007). Meta-regression

analyses revealed that the anatomical site (i.e., L/MC versus L/dlPFC) was a negative moderator, indicating that the frequency of headache episodes seems to be significantly lower with an application of rTMS over the L/MC as compared to the L/dlPFC (β = -4.829, SE = 2.018, p = 0.017). The anatomical site also decreased residual heterogeneity $(\tau^2 = 1.831; I^2 = 39\%, p = 0.110)$. On the contrary, having a sham coil instead of an angle manipulation as a sham had a positive moderator effect ($\beta = 5.853$, SE = 2.80, p = 0.037), indicating that the frequency of headache episodes seems to be significantly higher in the group with sham coil compared to the group with angle manipulation as a sham. Furthermore, an increase in the duration of the follow-up period decreased the frequency of headaches ($\beta = -0.874$, SE = 0.441, p = 0.048), which may indicate a long-term effect. However, we did not explore it further due to a limited number of studies (Table S6).

Side Effects and Tolerability

Only a few studies reported side effects, which included a short-lasting increase in headache intensity, mild localized pain, dizziness, sleepiness, and toothache [38, 58, 61••, 62]. rTMS was well tolerated among the patients except for one report by Granato et al. [39] describing an episode of migraine with complex aura (visual, sensitive, and aphasic) in a patient. A summary of all the side effects is included in Table 1.

Effects of rTMS on Quality of Life

Compared to sham control interventions, exposure to rTMS significantly decreased the standardized quality of life measures in the post-intervention period versus pre-intervention state among patients with CDH (5 studies-6 study arms [38, 53, 56, 58, 61••], pooled Hedges' g = -0.73 [-1.36, -0.10], p = 0.024; Fig. S4), which exceeded the MCID threshold of 0.5. Leave-one-out sensitivity analyses did not significantly change the pooled estimates (Table S5). A funnel plot of the effect-sizes indicated minimal publication bias (Fig. S5, Table 2). Imputing two effect-size using the trim-and-fill method to adjust for publication bias and re-analyzing the data using the imputed effect-size revealed a significant decrease in post-versus pre-intervention headache intensity following exposure to rTMS versus control conditions (pooled Hedges' g = -1.04[-1.69, -0.39], p=0.002, Table 2). Significant heterogeneity was observed among the standardized mean differences pooled in the random-effects model ($\tau^2 = 0.459$; $I^2 = 69.3\%$, p = 0.006). Subgroup and meta-regression analyses were not conducted as there were only a limited number of studies.

Discussion

While meta-analyses exist on rTMS for CM [63•, 64, 65], to our knowledge, this is the first comprehensive systematic review and meta-analysis examining the effects of rTMS on immediate and long-term outcomes for CDH. Our meta-analyses showed rTMS has a statistically and clinically meaningful impact on improving CDH intensity and frequency post-treatment, though effects tend to wear off after discontinuation.

rTMS is well documented to treat neuropathic pain and headaches like migraine, cluster headaches, and trigeminal neuralgia by transiently suppressing central pain perception and increasing pain stimulus thresholds [66, 67]. Several meta-analyses have shown that rTMS is beneficial in treating migraine prophylactically and therapeutically $[63 \bullet, 64, 65]$. Lan et al. [65] found that rTMS was only beneficial for treating migraine attacks, not CM symptoms. In contrast, Zhong et al. [63•] found that headache frequencies were reduced in episodic and CM types. Furthermore, Mohamad Safiai et al. [64] found that high-frequency rTMS reduces acute medication intake and functional disability associated with migraine but not headache days or pain intensity. All these meta-analyses have focused primarily on episodic migraine or a combination of episodic and CM, not purely on CM. According to our subgroup analyses, rTMS seems to have limited short-term and long-term impact in improving CM headache frequency, and data regarding improvement of headache intensity is abstruse.

One possible reason could be that CM is more of a problem of threshold where certain predisposing factors combined with frequent headache pain lower the threshold of migraine attacks and increase central and peripheral sensitization [68, 69]. For instance, genetic, anatomical, functional, and inflammatory factors change during the progression of a migraine attack or the transformation of episodic to CM [70–73]. As a result, CM becomes independent of triggers and depends more on fronto-limbic sensitization. Therefore, the application of rTMS should reset or reduce fronto-limbic dysfunction and cortical plasticity, which requires repeated exposure and long-term follow-up [17, 29, 74]. For instance, Fumal et al. [75] showed that daily rTMS induces cortical excitability and habituation patterns in migraine patients, which may contribute to its long-term efficacy in controlling headaches. Furthermore, it is essential to note that a small number of studies drove our results of the CM subgroup, and one study showed a strong placebo effect, which affected the overall estimates. Hence, further studies are needed to conclude whether this observation is generalizable.

CDH secondary to TBI is a constellation of debilitating chronic neuropathic pain, also called post-traumatic headaches [54, 55]. These patients are prone to adverse effects from chronic use of analgesics [76]. Therefore, a non-systemic, targeted therapy is ideal for these patients. Our subgroup analysis showed that TBI patients may benefit from rTMS by immediately lowering headache intensity. However, the number of studies and the sample size were limited, restricting our ability to confirm the effect [54, 55, 58]. Therefore, more research is needed on the applicability of rTMS. Similarly, rTMS showed limited evidence for reducing headache intensity for CTTH immediately postintervention; a limited number of studies conducted on patients with CTTH was the primary limiting factor [59, 60].

Our meta-analysis showed that with rTMS, there is an immediate effect of reducing headache frequency by 5 days/ month, and after one month, it remained reduced up to 3 days/month. Despite these effect-sizes being derived from a moderately heterogeneous limited number of small studies, our results substantiated the utility of rTMS in CDH. In contrast, current literature indicates that the therapeutic effects of rTMS for CDH are limited, with most studies being on CM [77-81]. Even for CM, BTX-A and anti-CGRPs are more promising, significantly reducing headache days and intensity short and long-term, unlike rTMS [82]. For instance, Sacco et al. [83] showed that anti-CGRPs reduce CM headache frequency by 2.39 days/month [-2.69, -2.08, n = 8902]. Similarly, BTX-A reduced headache frequency by 2.0 days/month [-2.8, -1.1, n=1384] [79]. Although we observed that rTMS may decrease headache frequency by days/month, the effect-sizes regarding BTX-A and anti-CGRP effect sizes were more robust and were derived from larger samples [77, 84].

Our study found that the effects of rTMS on CDH wears off over time. Thus, while rTMS may help as an abortive strategy for refractory CDH, its utility for CDH prophylaxis seems limited. On the contrary, BTX-A and anti-CGRPs demonstrated persistent effects in CM patients [82]. For instance, Lanteri-Minet et al. [84] showed that BTX-A has a lasting effect on headache frequency even at 24 weeks and 52 weeks based on long-term follow-up studies. Shehatha et al. [85] showed that compared to rTMS, BTX-A had sustained improvement in CM symptoms even after three months. However, some studies observed a lack of sustained efficacy for CDH, indicating a need for further investigations [15].

Our meta-regression analyses also showed that a longer duration of treatment and a higher number of sessions have better long-term effects in terms of sustained reduction in headache intensity. Furthermore, our meta-regression analyses showed a profound and sustained effect with increased duration and number of sessions. This dosedependent effect—a notable observation—needs further exploration. Zhan et al. [86] showed that at least five sessions of rTMS treatment are required for long-lasting motor functional recovery in the injured upper limb of stroke patients. Similarly, the motor evoked potentials last longer with an increased number of pulses of rTMS [23, 87]. Furthermore, the temporal pattern of rTMS induction (simple protocols versus patterned rTMS protocols) also seems to have a substantial impact [88••]. Long-lasting effects of rTMS are thought to originate from synaptic plasticity, which produces either potentiation or depression of synaptic strength. Long-term potentiation involves an increase in synaptic strength that can last for days or even weeks and months. Conversely, long-term depression encompasses a long-lasting weakening of synaptic connections. Studies have shown that rTMS-induced changes can last for at least eight days, as evidenced by altered uptake of F-fluorodeoxyglucose, indicating changes in neuronal excitability [89]. rTMS is also capable of increasing the expression of genes involved in synaptic plasticity, such as c-Fos and zif268, which play a role in the induction of long-term potentiation [90]. Similarly, activation of the NMDA receptor leads to a post-synaptic influx of calcium ions, playing a pivotal role in these longlasting changes [91]. Furthermore, dopamine receptor activation was found to be involved in the maintenance of plasticity [92].

Of note, the treatment protocols included in the metaregressions were highly heterogeneous-for instance, the majority of studies used coil perpendicular as a sham instead of a sham stimulator. A sham stimulator mimics the sensation of active stimulation without inducing the actual neurophysiological effects [93]. However, recent evidence suggests that sham stimulation may introduce variability, confound results, and pose challenges in blinding [94]. Using angle manipulation as a sham may provide more effective control conditions, as it minimizes the direct neurophysiological effects while maintaining the sensation of stimulation [94, 95]. Nevertheless, accurate placement and orientation of the coil require precision and understanding of the orientation of the induced electric field, adding complexity to the methodology [96]. Therefore, we cannot avoid the bias endorsed by different sham methods in this metaanalysis. We have tried to explore the moderator effect by doing a meta-regression analysis, and we found that angle manipulation seems to be the better option.

There were variations in the target region in the protocol. How the mechanism of action differs according to the target region in the brain is an important question worth addressing. A strong focal activation was observed in the thalamus, insula, cingulate-orbitofrontal junction, and a periaqueductal gray area in the brainstem following rTMS to MC, suggesting that a direct top–down activation of descending pain control system mediating via motor cortex, thalamus, insula, anterior cingulate cortex, and periaqueductal gray matter, which are components of pain modulation pathways [25, 97, 98]. Accordingly, rTMS applied to the MC reduces CM headache frequency [30, 99]. Conversely, rTMS to dlPFC exerts a top-down inhibitory neural circuit along the ascending midbrain-thalamic-cingulate pathway through the descending fibers from the PFC. dlPFC is involved in cognitive components of the pain experience, such as pain inhibition, perceived control of pain, and pain anticipation [98]. dlPFC and limbic cortex have been proposed to be extremely important in the pathophysiology of many chronic neurobehavioral conditions, such as addiction, depression, bipolar disorder, and migraine [29, 100]. Stimulation of dlPFC could reset or reduce fronto-limbic dysfunction in CM, leading to pain reduction [18•]. Todorov et al. [61••] compared rTMS to dIPFC and MC in CM and showed comparative results for both regions. Interestingly, our subgroup analyses showed that applying rTMS to MC or right dlPFC has a better outcome in reducing headache intensity, and our meta-regression analysis found rTMS to MC more effective versus dIPFC in reducing headache frequency. However, this analysis had < 10 studies, so more studies are needed to verify our dIPFC versus MC findings. Since our results highlighted several factors representing treatment protocols (e.g., anatomical location of stimulation, stimulation frequency, duration of stimulation), further exploration is needed to identify potential factors/parameters.

Guidelines have reiterated the high safety and tolerability profile of rTMS [101]. The most common side effect of rTMS is scalp discomfort or pain during treatment (~40%) [102], followed by headaches after treatment (20-30%) [103] and fatigue (15–20%) [104]. rTMS has also been associated, albeit rarely, with more severe adverse events such as seizures [105]. However, seizure risk is currently estimated to be minuscule overall < 1%. A large population-based study reported 24 seizures in 300,000 rTMS sessions (standardized risk of 7/100,000 sessions). Of those, 79% (n = 19) of seizures have occurred in patients with pre-existing risk factors (medication, neurological condition, epilepsy) [106]. This study estimated that rTMS delivered within published guidelines to individuals without risk factors appears to cause fewer than one seizure per 60,000 sessions. Apart from that, hearing impairment necessitates the use of hearing protection during treatment, EEG after-effects or abnormalities without overt clinical symptoms, and syncope or fainting episodes are some of the other side effects related to rTMS [101]. In conclusion, rTMS is an acceptable treatment modality with overall safety and tolerability. The studies included in this meta-analysis reported minimal number of side effects with good tolerability.

This meta-analysis has notable limitations. First, only a few eligible studies met inclusion, with several excluded due to unavailable variance estimates or combining episodic and CM. Second, we initially intended to focus only on CTTH, but the number of studies was limited, so we combined all CDH subtypes. Therefore, the studies included were heterogeneous in various aspects. Variability in treatment response and potential heterogeneity of CDH in terms of neuroanatomical and neurophysiological differences limit our ability to draw solid conclusions. Furthermore, the lack of consensus regarding brain targets and variation in stimulation parameters caused difficulties in comparing and combining all the studies. However, we attempted to explore these variations using meta-regression analyses that partially explained the heterogeneity. Third, only a limited number of studies focused specifically on the effect of rTMS on CDH sub-types such as CTTH, limiting the ability to make inferences on subgroups of CDH. Finally, outcomes such as the impact on disability, absolute or relative risk, and number needed to treat could not be determined due to limited studies reporting these outcomes. Nevertheless, the efficacy of rTMS may vary depending on individual patient characteristics and the specific parameters of the treatment.

In conclusion, rTMS demonstrates an immediate effect on reducing CDH intensity and frequency, indicating a potential for CDH symptom control. Importantly, rTMS is non-invasive, targeted, and safe, making it favorable for patients who cannot tolerate medications. While beneficial for short-term headache control, its effect does not seem to persist; thus, the costeffectiveness of rTMS as a primary treatment is questionable. As the synthesized evidence stems from small, low-quality studies, adequately powered randomized controlled trials (especially for CTTH) are necessary to establish the effects of rTMS on CDH. Furthermore, identifying the best treatment protocols (in terms of frequency, motor threshold, anatomical site, and a minimum number of interventions) and developing a consensus statement/guideline is essential. Based on the limited available evidence, rTMS appears to be safe and has the potential utility as part of a comprehensive treatment approach that may include other interventions such as lifestyle modifications, oral medications, BTX-A, and anti-CGRP antibodies in the management of particularly refractory CDH, potentially leading to better patient outcomes.

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Author Contributions CSD and CNK have full access to all of the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis. CSD, CLS, CRL, and CNK designed the study; ES, AB, RH, AR and CSD collected the data; CLS, CRL and CNK supervised data collection; CSD and CNK extracted and analyzed the data; ES, CSD and CNK wrote the manuscript; all authors read, revised and helped finalize the manuscript. All authors accept full responsibility for all aspects of to the work described.

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Compliance with Ethical Standards

Conflicts of Interest The authors have no potential conflicts of interest to declare pertinent to the content in the manuscript.

Human and Animal Rights and Informed Consent This article does not contain any studies with human or animal subjects performed by any of the authors.

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