

# **Assessing air pollution exposure misclassifcation using high‑resolution PM2.5 concentration model and human mobility data**

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### **Abstract**

Due to a paucity of human movement data, the traditional method for estimating pollution exposure is static: Exposure is based on place of residence. However, local air quality varies over both time and space. This study explores exposure measurement errors associated with ignoring human mobility and its impact on exposure-health efect estimates. Using a random forest classifcation model, this study examines the impact of a variety of factors on potential measurement errors in personal exposure to outdoor PM<sub>2.5</sub>. Mobility data at the individual level was combined with hourly PM<sub>2.5</sub> surfaces at the neighborhood level to estimate and compare residence-based and mobility-based exposures for 100,784 Los Angeles County residents. The results show that exposure measurement errors increase for individuals with high mobility levels. Signifcant sociodemographic disparities are observed across diferent exposure classifcation groups. Exposures of low-income people who have high mobility and reside in polluted neighborhoods tend to be overestimated. In contrast, exposures of high-income people living in neighborhoods with cleaner air are likely to be underestimated. The result on the exposure-health efect suggests that health risks of the socially disadvantaged after exposure to  $PM<sub>25</sub>$  is likely to be underestimated due to the exposure measurement error introduced by ignoring human mobility.

**Keywords** Exposure measurement error  $\cdot PM_{2.5} \cdot M_{\text{obility}} \cdot$  Exposure-health effect

# **Introduction**

Personal air pollution exposures are very challenging to quantify accurately. Traditional approaches to quantifying exposure to outdoor air pollution assume that concentrations at the residential address are adequate surrogates of personal exposure to air pollution of outdoor origin (Bae et al. [2007](#page-11-0); Bell and Ebisu [2012](#page-11-1); Elliott and Smiley [2019](#page-12-0); Houston et al. [2004;](#page-12-1) Rowangould [2013\)](#page-13-0). The underlying assumption is that individuals spend most of their time indoors at the residence (Klepeis et al. [2001](#page-12-2)) and that outdoor air pollution infltrates into the indoor environment where exposure occurs. However, given that people are mobile and their exposure to air pollution can occur in various locations, this static residential approach will inevitably introduce exposure measurement error and potential bias in air pollution and health

 $\boxtimes$  Yougeng Lu yougengl@stanford.edu assessments, which may lead to inefective public health policy interventions.

Measurement errors in air pollution exposure can come from several sources. Recent studies report that exposure measurement may be substantially biased low if not considering human mobility (Gurram et al. [2019](#page-12-3); Lu [2021](#page-12-4); Park and Kwan [2017](#page-13-1); Tayarani and Rowangould [2020](#page-13-2)). Park and Kwan [\(2017\)](#page-13-1) argue that the individual's time-activity pattern determines their exposure levels as personal exposure to air pollution occurs through dynamic spatiotemporal interactions between individuals and air pollutant distribution. Although people generally spend more time at home, the majority of their exposure occurs in other places (Park [2020\)](#page-12-5). For example, workers generally spend more time exposed in traffic during commuting. Overlooking human mobility will lead to inaccurate air pollution exposure measurements.

Beyond ignoring human mobility, spatiotemporal resolution of air pollution surfaces or prediction models used is another important factor that may cause exposure measurement error. Coarse resolution cannot refect important spatial gradients (Clark et al. [2022](#page-12-6); Korhonen et al. [2019](#page-12-7); Li

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et al. [2016](#page-12-8)). The spatial resolution of assigned air pollution concentrations in recent studies varied substantially, ranging from  $0.25$  to  $10 \text{ km}^2$  (Dewulf et al.  $2016$ ; Gurram et al.  $2019$ ; Park and Kwan [2017](#page-13-1); Yu et al. [2020](#page-13-3)). The temporal resolution of outdoor  $PM<sub>2.5</sub>$  surfaces is another important factor, with several studies relying on daily, monthly, or annual averaged pollution models to estimate residential exposure (M. Nyhan et al. [2016](#page-12-10); Pennington et al. [2017](#page-13-4); Setton et al. [2008](#page-13-5)). Greater temporal aggregation does not capture how concentrations vary over time. Evidence shows that personal exposure to outdoor air pollution tended to be overestimated at the residence and underestimated at daily activity spaces when daily, monthly, or annual average air pollution concentrations are used to estimate personal exposure (Dhondt et al. [2012](#page-12-11)).

Moreover, exposure measurement error will further distort the air pollution-health efect estimates (Basagaña et al. [2013](#page-11-2); S. Y. Kim et al. [2009;](#page-12-12) Samoli and Butland [2017](#page-13-6); Sellier et al. [2014\)](#page-13-7), leading to biased estimates between exposure and health outcome. Prior studies have documented the impact of potential exposure measurement errors on the air pollution-health relationship. Jerrett et al. [\(2005](#page-12-13)) show that inaccurate personal exposure measurements resulting from poorly spatially resolved air pollution prediction models may signifcantly impact the relationship between air pollution exposure and mortality. Inconsistent results in efect estimates of  $NO<sub>2</sub>$  on newborns' birthweight were obtained when diferent spatially resolved air pollution prediction models were used (Sellier et al. [2014](#page-13-7)). Another study presents that exposure measurement error might lead to bias of regression coefficients and to inflation of their variance when personal exposures assessed through air pollution prediction models with diferent spatial and temporal resolutions are used as explanatory variables in models for exposure-health estimates (Basagaña et al. [2013\)](#page-11-2).

Substantial effort has been invested in air pollution epidemiology research to develop statistical models to predict personal exposures at subjects' locations in situations where measurements at the desired locations are not available. However, most existing air pollution epidemiology studies focus on the impact of air pollution prediction models on the exposure-health relationship (Basagaña et al. [2013](#page-11-2); S. Y. Kim et al. [2009;](#page-12-12) Samoli and Butland [2017;](#page-13-6) Sellier et al. [2014\)](#page-13-7). Few studies have examined the effects of human mobility in exposure measurements on the association between air pollution exposure and health outcomes. Yu et al. ([2020](#page-13-3)) fnd that people with higher mobility levels tend to have larger exposure measurement errors. Shareck et al. ([2014](#page-13-8)) argue that unequally distributed features and resources across spaces may induce air pollution exposure disparities by constraining sites where people perform their everyday activities. For example, due to accessibility limitations, low-income groups usually travel shorter distances from their homes than their high-income counterparts (Morency et al. [2011](#page-12-14); Vallée et al. [2010](#page-13-9)). Compared to whites, blacks and Latinos usually have lower mobility levels (Hu et al. [2020\)](#page-12-15). Full-time employees tend to travel longer daily distances (Järv et al. [2015](#page-12-16); Morency et al. [2011](#page-12-14); Páez et al. [2010\)](#page-12-17), whereas part-time employees and unemployed people are more place-bound (Lu [2021;](#page-12-4) Vallée et al. [2010\)](#page-13-9). Little is known about how distinct travel behaviors and mobility patterns of diferent sociodemographic groups may infuence their air pollution exposure levels. The impact of human mobility on the exposure-health efect is also less studied in previous research.

This study aims to examine the impact of human mobility in exposure measurement errors and health effects associated with air pollution and disentangle the complex relationship between sociodemographic variables and exposure measurement errors. In this study, residence-based and mobilitybased  $PM<sub>2.5</sub>$  exposures for a sample of Los Angeles County residents on a typical weekday were estimated by coupling hourly  $500 \times 500$ m PM<sub>2.5</sub> surfaces at the neighborhood level and simulated daily mobility data at the individual level. The study samples were classifed into three exposure groups based on diferences between their residence- and mobilitybased  $PM<sub>2.5</sub>$  exposures: individuals with similar residence and mobility exposure, individuals whose exposures were overestimated, and individuals whose exposures were underestimated. Random forest classifcation models were used to examine the impacts of a series of mobility and sociodemographic variables on exposure classifcation results. Last, sensitivity analysis was conducted to examine the impact of human mobility on exposure-health effects across exposure classifcation groups and sociodemographic groups.

The remainder of the paper is organized as follows. The "Method" section presents the data and methods used in this study. The "Results" section summarizes the results. The "Discussion" section discusses the fndings. Conclusions are presented in the "Conclusion" section.

### **Method**

### **Study area**

Los Angeles is well recognized for its notoriously severe air pollution problem as one of the US metropolitan regions with the highest level of particulate matter pollution (American Lung Association [2020\)](#page-11-3). Los Angeles has the most developed highway system and the busiest traffic in the US  $PM<sub>2.5</sub>$  pollution, a primary air pollutant created by vehicles, which has been a serious public health problem in Los Angeles for decades. In Los Angeles,  $PM<sub>2.5</sub>$  concentrations vary spatially and temporally, with the highest pollution observed during peak hours and within core urban

areas (Lu et al. [2021](#page-12-18)). Los Angeles is, therefore, a good case study to examine variation in exposure levels across time and space taking daily travel patterns into account and test whether exposure patterns are related to sociodemographic population characteristics.

### **Data**

#### **PM2.5 pollution modeling**

In this study, ground-level  $PM<sub>2.5</sub>$  concentrations were estimated from our recently developed  $PM_{2.5}$  model (Lu et al. [2021](#page-12-18)). We have created an hourly, 0.25-km gridded  $PM_{2.5}$ model for Los Angeles County that incorporates low-cost air sensor data (i.e., PurpleAir) and machine learning techniques. Ambient air pollution has been traditionally monitored at regulatory stations at high instrumentation and maintenance costs. Sparse and uneven regulatory monitoring has a limited ability to refect pollution details, especially in unmonitored areas. Dense deployment allows low-cost PurpleAir sensors to capture spatiotemporal variations of localized PM<sub>2.5</sub> concentrations at finer resolution than regulatory air quality stations (Bi et al. [2020](#page-11-4); Lu et al. [2021](#page-12-18); Mousavi and Wu [2021](#page-12-19)). A number of recent studies have used PurpleAir sensors to develop  $PM<sub>2.5</sub>$  modeling at fine spatiotemporal resolution (Bi et al. [2020](#page-11-4); Lu et al. [2021](#page-12-18); Mousavi and Wu [2021](#page-12-19)).

In this study, twenty-four hourly  $PM<sub>2.5</sub>$  concentration surfaces over the course of a typical weekday in 2019 were generated at a 500 m  $\times$  500 m grid level for Los Angeles County. A suite of spatiotemporal variables, including meteorological conditions, land use variables, and traffic counts, was integrated with the random forest method to estimate  $PM<sub>2.5</sub>$  concentration at the sub-daily and neighborhood level. Estimated  $PM_{2.5}$  concentrations were then validated against measured  $PM<sub>2.5</sub>$  concentrations by the 10-fold crossvalidation method. The results showed that the PurpleAirbased PM<sub>2.5</sub> prediction model could capture more than 90% of variations. A comprehensive description and validation results can be found in Lu et al. [\(2021](#page-12-18)).

#### **Activity‑based travel demand modeling**

Daily travel trajectories for 100,784 Los Angeles County residents were simulated using an activity-based travel demand model developed by the Southern California Association of Governments (SCAG) for an average weekday in 2019 (Pendyala et al. [2012;](#page-13-10) Ziemke et al. [2015](#page-13-11)). American Community Survey (ACS) 2003 and Census 2000 have been used to validate this SCAG simulated travel trajectory data. Validation results show that the SCAG activity-based travel demand model has a good performance in predicting "activity purpose-number" and mimicking corresponding population features at the individual level. According to the validation results, the majority of the synthetic population deviated less than 5% from the reference group in terms of demographic and socioeconomic characteristics (Pendyala et al. [2012](#page-13-10)).

The SCAG travel trajectory dataset contains 387,398 trip records for 100,784 Los Angeles County residents (approximately 10% of the total Los Angeles County population). Each trip record includes a personal ID, origin-destination pair of the trip, trip purpose, trip departure and arrival timestamps, trip duration, and travel mode. The personal ID is unique for each individual and is used to connect with synthetic demographic features provided by the SCAG (Bhat et al. [2013](#page-11-5); Lu [2021](#page-12-4)). The origin and destination of each trip are allocated to the geographic unit of the traffic analysis zone (TAZ), whose size is similar to the census tract. The centroid of TAZ is assumed to be each trip's origin or destination point.

However, the travel trajectory dataset lacks information on travel paths between the activity sites. This study used the OSMnx python package to estimate probable travel paths between two activity TAZs in the shortest path distance (Boeing [2017](#page-11-6); Lu [2021](#page-12-4)).

### **Individual exposure assessment**

#### **Static and dynamic exposure assessment**

The study region was subdivided into  $0.25 \text{ km}^2$  hexagon grids. As noted earlier, hourly  $PM<sub>2.5</sub>$  concentrations of each grid (twenty-four hours in total) were generated by utilizing the PM<sub>2.5</sub> model developed by Lu et al.  $(2021)$  $(2021)$  for Wednesday, September 18, 2019. Due to limitations in computing resources,  $PM_{2.5}$  concentrations are assumed to be constant during an hour within each hexagon grid. The hourly  $PM_{2.5}$ concentrations were spatially matched to each TAZ in compliance with the travel trajectory data and averaged within each TAZ if multiple  $PM<sub>2.5</sub>$  hexagon grids locating in the same TAZ. Two types of individual  $PM<sub>2.5</sub>$  exposures were then assessed: (1) static  $PM_{2.5}$  exposure at residence and (2) dynamic  $PM<sub>2.5</sub>$  exposure that considers individuals' daily mobility patterns.

The individual static and dynamic exposures are estimated as in Eqs.  $(1)$  $(1)$  and  $(2)$  $(2)$ :

<span id="page-2-0"></span>
$$
\text{Static}_{i} = \frac{\sum_{t=1}^{T} PM_{h,t}}{T} \tag{1}
$$

<span id="page-2-1"></span>
$$
\text{Dynamic}_{i} = \frac{\sum_{t=1}^{T} \sum_{n=1}^{N} PM_{n,t} \cdot P_{n}}{T}
$$
 (2)

where  $PM_{h, t}$  is  $PM_{2.5}$  concentration in hour *t* at TAZ *h*, where individual *i*'s home is located. *T* denotes 24 h of a

day.  $PM_{n, t}$  is  $PM_{2.5}$  concentration in hour *t* at TAZ *n*, where individual *i* is located within hour *t*. *N* represents the total number of TAZs (microenvironments) individual *i* has stayed during hour  $t$  ( $N \ge 1$ ).  $P_n$  denotes the percentage of time during hour *t* that individual *i* stays in TAZ *n*.

### **Exposure classifcation based on exposure measurement error**

Prior research has documented the occurrence of exposure misclassifcation when human mobility is not taken into account in exposure assessment (Guo et al. [2020](#page-12-20); Yu et al. [2020\)](#page-13-3). Exposure of individuals who have high residencebased exposures is likely to be reduced by their mobility, while exposure of individuals who have relatively low residence-based exposures is likely to be increased (J. Kim and Kwan [2021;](#page-12-21) Kwan [2018](#page-12-22)). All study subjects were subdivided into three groups according to their exposure measurement errors shown in Eq.  $(3)$  $(3)$ :  $(1)$  individuals with similar dynamic and static exposures, which is referred to as the "Accurate" group; (2) individuals with higher static exposures than their dynamic exposures, which is referred to as the "Overestimated" group; and (3) individuals with higher dynamic exposures than their static exposures, which is referred to as the "Underestimated" group.

The magnitude and direction of two statistical indicators were employed to categorize exposure classifcation groups: (1) exposure measurement error and (2) mean absolute percentage error (MAPE). The exposure measurement error was calculated by subtracting an individual's static exposure from their dynamic exposure (i.e., Dynamic*i*− Static*<sup>i</sup>* ). A positive exposure measurement error indicates an individual's exposure is underestimated, while a negative measurement error indicates overestimated exposure. MAPE was adopted as an additional criterion to evaluate the degree of agreement between an individual's static and dynamic exposures: | | | Static<sub>i</sub> | Dynamic*i*−Static*<sup>i</sup>* | Static*<sup>i</sup>* |  $\times$  100%. Higher MAPE values indicate diferences between static and dynamic exposures as a result of overestimated or underestimated exposures. The thresholds for exposure measurement error and MAPE were set to  $\pm 0.5 \,\mu$ g/m<sup>3</sup> and 10%, respectively, to determine an individual's exposure classification group. A comprehensive description of the classifcation method is shown in Eq. [\(3](#page-3-0)).

$$
E_i = \begin{cases} \text{Overestimated} & \text{if Error}_i < -0.5 \text{ and } MAPE_i > 10\%\\ \text{Accurate} & \text{if } -0.5 \leq \text{Error}_i \leq 0.5 \text{ and } MAPE_i < 10\%\\ \text{Underestimated} & \text{if Error}_i > 0.5 \text{ and } MAPE_i > 10\% \end{cases}
$$

where  $E_i$  denotes the exposure classification group that individual *i* belongs to; Error*<sup>i</sup>* is the exposure measurement error for individual *i*.

<span id="page-3-0"></span>(3)

#### **Random forest classifcation model**

This study utilized the random forest classifcation model to examine associations of a variety of mobility and sociodemographic variables with exposure classifcation results. In contrast to traditional linear regression, the random forest model can capture nonlinear relationships between response variables and predictors and provide a fexible and automated process for predicting target variables (Breiman [2001\)](#page-11-7). The random forest model generates a number of decision trees and trains each decision tree independently using a random sample of the data. This randomness contributes to the model being more robust than a single decision tree and less prone to overftting the training data. Furthermore, the random forest model avoids the probable multicollinearity across sociodemographic variables, which violates the underlying premise of independence in many regression models.

In this study, 90% of samples were randomly subsampled as training set and the remaining 10% as a testing set to evaluate the model performance. Since the classes were unbalanced (79% of the study sample was classifed as the Accurate group, 9% as the Overestimated group, and 12% as the Underestimated group), a combination of the Synthetic Minority Over-sampling Technique (SMOTE) and random under-sampling methods was utilized to resample the dataset until balanced training classes were achieved (Chawla et al. [2002](#page-12-23); He and Garcia [2009\)](#page-12-24). The class-wise sensitivity and specifcity, as well as the mean classifcation accuracy, were calculated to evaluate the random forest model performance. Confusion matrices were utilized to calculate the specifcity and sensitivity of candidate models.

The optimal number of randomly sampled features at each node (*m*) and decision trees (*k*) were determined by minimizing the out-of-bag (OOB) error rate through iterative crossvalidation (Lu et al.  $2021$ ). The relative importance of each predictor variable was determined using the mean decrease in accuracy based on OOB error. Partial dependence plots were produced to depict the correlations between predictor variables and the probability of being classifed into a given class. A partial dependence plot demonstrates the marginal efect of a predictor variable on the predicted response while controlling for all other variables in the model (Friedman [2001](#page-12-25)). Figure [1](#page-4-0) shows the framework and process of assessing exposure classifcation error and factors afecting it.

## **Exposure and health efect across exposure classifcation groups**

Ordinary least squares (OLS) regression models were run as multivariate models to assess the association between personal exposures and health outcomes for static and dynamic exposures, respectively. The exposure-health efects were



<span id="page-4-0"></span>**Fig. 1** Research conceptual framework

further assessed across diferent exposure classifcation groups, racial groups, and income groups. Recent studies have shown that exposure to  $PM<sub>2.5</sub>$  is linked to acute respiratory symptoms (Bose et al. [2015\)](#page-11-8) and cardiovascular disease (Madrigano et al. [2013](#page-12-26); Neophytou et al. [2014](#page-12-27)). Thus, two health outcomes were adopted as the dependent variables for the OLS models: the rate of emergency department visits for asthma and the rate of emergency department visits for heart attacks per 10,000 persons.

The health outcome data were obtained from CalEnviroScreen at the census tract level (California Office of Environmental Health Hazard Assessment [2023\)](#page-12-28). Since the individual-level health outcome data were not available, the rate of asthma and rate of heart attack was assigned to each study subject based on their residential locations. That is, study subjects who live in the same TAZ are assumed to have the same health outcomes. The relationship between exposure and the two health indicators was estimated by OLS models adjusted several confounding variables including demographic variables (gender, age, race) and socioeconomic status (income, employment status, education). Table [2](#page-7-0) presents a summary of these variables.

In summary, this study frst measures the residence-based static and mobility-based dynamic exposures for study subjects, respectively. The study subjects are then classifed to three groups according to the magnitude and direction of their exposure measurement errors. The random forest classifcation model is used to examine the impact of human mobility and sociodemographic characteristics on exposure measurement error. Last, this study explores the exposurehealth effect by using static and dynamic exposure measurements through OLS regression models among diferent cohorts (Fig. [1](#page-4-0)).

### **Results**

### **PM2.5 and population activity distribution**

Figure [2a](#page-5-0)–d presents the estimated hourly  $PM_{2.5}$  concentrations across Los Angeles County at the neighborhood level on a typical weekday in 2019. A heterogeneous spatiotemporal distribution pattern of  $PM<sub>2.5</sub>$  pollution can be observed. In general,  $PM<sub>2.5</sub>$  concentrations are higher in daytime than evening and night, especially in the morning peak hours (Fig. [2](#page-5-0)b), and most pollution concentrates in urban cores and along highways. These patterns are in line with the fndings of previous studies (Lu et al. [2021\)](#page-12-18).

Figure [2](#page-5-0)e–h presents the simulated activity patterns of Los Angeles residents at four diferent hours on a typical weekday in 2019. These fgures refect the distribution pattern of people's residences and workplaces. Overall, Los Angeles residents generally travel from their sparsely distributed places of residence (Fig. [2e](#page-5-0)) to urban cores (Downtown Los Angeles, Wilshire-Santa Monica corridor, Long Beach) in the early morning (Fig. [2f](#page-5-0)) and stay all day (Fig. [2](#page-5-0)g) until they return to their residences again in the evening (Fig. [2](#page-5-0)h).

### **Exposure classifcation error analysis**

To investigate potential exposure measurement errors resulting from ignoring human mobility, fows between diferent quartiles of study subject's static exposure and dynamic exposure were plotted in Fig. [3](#page-6-0)a. A high percentage of the population was misclassifed into other quartiles, especially for study subjects with static exposures in middle quartiles (Q2 and Q3). About one-third of populations in the middle quartiles was classifed into other quartiles when human mobility was omitted in exposure measurement. Three exposure classifcation groups were identifed by quantifying the diference between individuals' static and dynamic expo-sures based on Eq. ([3](#page-3-0)). Table [1](#page-6-1) gives summary statistics. Figure [3b](#page-6-0)–d shows the distributions of static and dynamic exposures for each group.

Table [1](#page-6-1) shows that the Accurate group is the largest. For about 80% of the observations, there is no diference between static and dynamic exposures (the diference is statistically signifcant but not meaningful). Figure [3](#page-6-0)b shows how close the two distributions are. The Overestimated group is the smallest (9% of the study sample). For individuals in the Overestimated group, mean static exposure was  $0.96 \mu$ g/m<sup>3</sup> higher than their dynamic exposure. This diference is large (about 10%) and signifcant. This group has the highest static exposure level of all groups (Fig. [3c](#page-6-0)). The Underestimated group accounts for the remaining 12% of observations. The mean diference between static and dynamic estimates is 1.15  $\mu$ g/m<sup>3</sup> or about 17%, even larger than the difference for the Overestimated group (Fig. [3d](#page-6-0)).



<span id="page-5-0"></span>Fig. 2 Estimated hourly  $PM_{2.5}$  concentrations at (a) midnight, (b) 8 AM, (c) noon, and (d) 6 PM and distribution of population activity at (e) midnight, (**f**) 8 AM, (**g**) noon, and (**h**) 6 PM on a typical weekday in Los Angeles County



<span id="page-6-0"></span>**Fig. 3** The distribution of exposure measurement error: (**a**) direction of potential PM<sub>2.5</sub> exposure misclassifications between static exposure and dynamic exposure; (**b**) distribution of static and dynamic

exposures for the Accurate group; (**c**) distribution of static and dynamic exposures for the Overestimated group; (**d**) distribution of static and dynamic exposures for the Underestimated group

		Mean	Std. dev		Median Minimum	Maximum
Accurate $(N = 80, 411)$	Dynamic exposure	8.17	1.38	8.51	2.71	12.27
	Static exposure	8.16	1.39	8.48	2.65	11.97
	Difference	$0.01***$				
Overestimated $(N = 8655)$	Dynamic exposure	8.03	1.06	8.16	3.35	11.46
	Static exposure	8.99	1.01	9.15	4.20	11.97
	<b>Difference</b>	$-0.96***$				
Underestimated $(N = 11,718)$	Dynamic exposure	7.84	1.16	8.06	3.06	13.86
	Static exposure	6.69	1.33	6.93	2.65	11.43
	Difference	$1.15***$				

### **Random forest results**

<span id="page-6-1"></span>**Table 1** Summary statistics of static and dynamic  $PM<sub>2.5</sub>$ exposures (μg/m*<sup>3</sup>* ) across exposure classifcation groups

 $(***p < 0.001)$ 

### **Model performance and variable importance**

The descriptive analysis has revealed mobility and sociodemographic diferences across exposure classifcation groups. Random forest models were further trained using the same set of mobility and sociodemographic variables to examine their correlation with exposure classifcation errors. Table [2](#page-7-0) lists all mobility and sociodemographic variables and their summary statistics used for the random forest model. As noted earlier, three exposure classifcation groups were defned, and the random forest classifcation algorithm was used to develop a predictive classifcation model based on individual's mobility patterns, residential pollution level, and sociodemographic characteristics. The hyperparameters for the random forest model were set to 1500 decision trees with a minimum sample leaf of 50.

The random forest model yielded a mean classifcation accuracy (adjusted across all classes) of 71%. Figure [4a](#page-7-1) presents the confusion matrix for evaluating the random forest model's performance. The sensitivity values for the Accurate, Overestimated, and Underestimated groups are 73%, 71%, and 70%, respectively, implying good agreement between actual and predicted classifcations.

The relative contribution value of predictor variables to the random forest classifcation results is shown in Fig. [4b](#page-7-1), sorted in order of importance. The variable importance

	Variables	Description	Mean		Std. dev. Minimum Maximum	
Travel behavior	Trip number	The number of daily trips made by an individual		2.05	$\overline{c}$	18
	Trip distance	The daily trip distance in miles an individual travels	24.48	21.44	0.19	234.28
	Hours stay out-of-home	The hours an individual spends out-of-home per day	8.07	3.49	0	22.43
	%Driving trip	Proportion of daily trips of an individual made by driving	84.3%	32.2%	$0\%$	100%
	%Public transit trip	Proportion of daily trips of an individual made by public transit	2.5%	14.1%	$0\%$	100%
	%Walk/bike trip	Proportion of daily trips of an individual made by walking or bicycling	$10.0\%$	25.8%	$0\%$	100%
	Residential pollution Residence pollution level	The standard index showing overall air pollution level at places of residence (between 1 and 10): higher value means higher pollution		1.16	2.32	9.62
Sociodemographic	Age	Age	33.99	21.43	$\overline{0}$	94
	Income	Household income (\$1000)	72.22	30.12	12.27	230.90
	Male	Dummy variable: $1 =$ male: $0 =$ female			0	1
	Non-Hispanic White	Dummy variable: $1 = \text{non-Hispanic White}; 0 =$ otherwise			0	1
	<b>Black</b>	Dummy variable: $1 = Black$ ; $0 = otherwise$			0	1
	Hispanic	Dummy variable: $1 = \text{Hispanic/Latino}$ ; $0 = \text{other}$ wise			$\Omega$	1
	Asian	Dummy variable: $1 = \text{Asian}$ ; $0 = \text{otherwise}$			0	1
	Worker	Dummy variable: $1 =$ employed: $0 =$ otherwise (unemployed, homemaker, student, or retired)			$\mathbf{0}$	1
	College	Dummy variable: $1 =$ college or higher degree; $0 =$ otherwise			$\mathbf{0}$	$\mathbf{1}$

<span id="page-7-0"></span>**Table 2** Descriptive statistics of the mobility and sociodemographic variables used in the analysis of exposure classifcation errors



<span id="page-7-1"></span>**Fig. 4** Random Forest model performance: (a) confusion matrix of predicted exposure classifcation groups; (b) variable importance rank

rank shows that daily trip distance, hours stay out of home, household income, and residential pollution level are among the most important features. By contrast, ethnicity, employment status, and education play weaker roles in afecting exposure classifcation errors. These fndings suggest that individual's exposure measurement error is mainly afected by their mobility levels, income, and pollution levels at residence.

#### **Partial dependence analysis**

The partial dependence plots illustrate the marginal effect of a single variable on the predicted classifcation outcome. According to variable importance results, the partial dependence of daily trip distance, hours stay out of home, household income, and residential pollution levels on the probability of classifcation results were examined. Figure [5](#page-8-0) plots



<span id="page-8-0"></span>**Fig. 5** Partial dependence (PD) plots for the most important variables in the random forest classifcation model for (**a**) the Accurate group, (**b**) the Overestimated group, and (**c**) the Underestimated group

the partial dependence of the abovementioned variables for all groups.

As shown in Fig. [5](#page-8-0)a, increasing probabilities of an individual belonging to the Accurate group were associated with shorter daily trip distance, fewer hours spent out of home, and higher residential pollution levels. Household income displays a nonlinear relationship with probabilities of the Accurate group. The most signifcant marginal infuence was depicted at around \$80,000. The middle column of Fig. [5a](#page-8-0) shows a two-dimensional partial dependence plot of daily trip distance and hours stay out of home to explore the effects of combining two mobility variables on probabilities of the Accurate group. The color scheme represents diferent probability levels. Yellow tones indicate a lower probability, and purple tones denote a higher probability. The two-dimensional plot shows that individuals who travel longer distances and time away from home are least likely to be categorized to the Accurate group.

A similar efect of mobility variables on classifcation probability can be observed in Fig. [5](#page-8-0)b,c. Both probabilities of the Overestimated group and the Underestimated group grow with the daily trip distance and hours stay out of home. The two-dimensional plots indicate that exposure is more likely to be overestimated or underestimated for people with high mobility levels. Although the effects of mobility variables on the magnitude of exposure classifcation error were similar for the Overestimated group and the Underestimated group, diferent household income and residential pollution levels for the two groups resulted in completely opposite directions of exposure classifcation errors. Increasing probabilities of the Overestimated group were associated with lower household income and higher residential pollution levels (Fig. [5](#page-8-0)b). By contrast, lower household income and higher residential pollution levels were associated with reduced probabilities of the Underestimated group (Fig. [5](#page-8-0)c). The opposite associations of household income and residential pollution level with the probabilities of the Overestimated and Underestimated groups suggest that as mobility levels increased, exposures were more likely to be overestimated for low-income residents living in highly polluted areas, while exposures were more likely to be underestimated for high-income residents living in areas with cleaner air.

### **Exposure‑health efect analysis**

Figure  $6$  shows the correlation coefficient (95 percent confdence interval (CI)) between exposures (static and dynamic exposure) and health outcomes (rate of asthma and heart attack) across three diferent cohorts: exposure <span id="page-9-0"></span>**Fig. 6** Sensitivity analysis of exposure-health effect



classifcation groups, racial groups, and income groups. A positive association between exposure and adverse health outcomes was observed in Fig. [6](#page-9-0). However, the coefficients of static and dynamic exposures were signifcantly diferent across diferent groups. Exposure measurement errors associated with omitting human mobility can result in bias in the correlation between exposures and health outcomes. For the Accurate group, the correlation coefficient of static exposure is similar to dynamic exposure (2.42 vs. 2.50 for asthma and 0.21 vs. 0.22 for heart attack). However, for people whose exposure is overestimated, the efect of  $PM_{2.5}$  exposure on the risk of asthma (1.85) and heart attack (0.27) is greater than estimated by static exposure (1.35 and 0.20). Conversely, for people whose exposure is underestimated, their health risks related to exposure to  $PM_{2.5}$  tended to be overstated. For the Underestimated group, a 1  $\mu$ g/m<sup>3</sup> increase in static exposure to PM<sub>2.5</sub>

would lead to a 1.21% increase in emergency department visits for asthma. At the same time, the rate decreases to 0.74% when considering human mobility in the exposure measurement (Fig. [6a](#page-9-0)).

The effect of  $PM_{2.5}$  exposure on the risk of asthma and heart attack is found to be underestimated for most racial and income groups if ignoring human mobility in exposure measurement (Fig. [6c](#page-9-0)–f). Hispanics, blacks, and the lowincome are found to be disproportionately burdened with health risks associated with air pollution, which is consistent with fndings obtained from existing literature (Bae et al. [2007](#page-11-0); Gilbert and Chakraborty [2011;](#page-12-29) Houston et al. [2004](#page-12-1)). The sensitivity analysis on the exposure-health effect suggests that health risks of the socially disadvantaged after exposure to  $PM<sub>2.5</sub>$  are likely to be underestimated due to the exposure mismeasurement introduced by ignoring human mobility.

#### **Discussion**

Overlooking human mobility may lead to incorrect exposure assessment and misleading conclusions and thus results in inefficient public health policy solutions (J. Kim and Kwan [2021](#page-12-21); Park and Kwan [2017\)](#page-13-1). A growing amount of research has highlighted the importance of human mobility in air pollution exposure assessment (Dewulf et al. [2016;](#page-12-9) Ma et al. [2020;](#page-12-30) M. M. Nyhan et al. [2019;](#page-12-31) Park [2020](#page-12-5)), but little is known about the impact of distinct mobility patterns on exposure measurement errors and how these errors infuence exposure-health efect. This study offers important insights into the literature by investigating the underlying factors contributing to exposure measurement errors. This study indicates that the individual's mobility level is the most critical factor in determining exposure measurement errors. Exposure measurement errors increase with mobility. Individuals with high mobility have the most signifcant exposure measurement errors, especially those who travel long distances and spend more time out of the home.

There is also a signifcant correlation between individuals' sociodemographic characteristics and exposure measurement errors. Household income has the greatest effect on exposure measurement errors, likely due to the key role of wealth in determining where people live, their occupations, and places people often visit (Sampson [2019](#page-13-12)). According to the results, household income is more inclined to drive the direction of exposure measurement errors. Air pollution at residence is another critical factor infuencing exposure measurement errors. On average, as mobility increases, exposure is likely overestimated for low-income residents of neighborhoods with poor air quality, while exposure is typically underestimated for high-income residents of neighborhoods with cleaner air. This fnding is consistent with the conclusion obtained from prior empirical research: Exposures of individuals who are less exposed at residence are likely amplifed by their mobility, and exposure of people with high residential exposure is usually attenuated (Dewulf et al. [2016](#page-12-9); Picornell et al. [2019;](#page-13-13) Tayarani and Rowangould [2020;](#page-13-2) Yu et al. [2020](#page-13-3)). One probable explanation is that people from neighborhoods with cleaner air are more likely to carry out their daily activities in neighborhoods with poorer air quality (Boeing et al. [2023\)](#page-11-9). In contrast, residents of neighborhoods with high air pollution tend to engage in daily activities in neighborhoods with less air pollution than their residential neighborhoods (J. Kim and Kwan [2021;](#page-12-21) Lu [2021](#page-12-4)).

The results show that the relative exposure measurement error is larger for wealthier people because they often reside in neighborhoods with cleaner air and their residence-based exposures start much lower than those with fnancial restrictions. However, the overall exposure and burden of health risks are much higher for the more disadvantaged populations as they are likely from more polluted neighborhoods. People tend to spend more time at home, even those with high mobility levels (Lu [2021](#page-12-4); Park [2020\)](#page-12-5). If the socially disadvantaged stay within their residential neighborhoods or vicinity most of the day and spend a lot more time in transit to move shorter distances, either or both of these mobility patterns can lead to much worse exposures but fewer exposure measurement errors. Given the signifcant contribution of residential air pollution to an individual's overall exposure, although the exposure of people who live in more polluted neighborhoods may be overestimated, they are still likely to have relatively higher exposures than those living in less polluted neighborhoods.

Moreover, exposure mismeasurement can result in bias in the correlation between air pollution exposure and health outcomes, which may further bias estimates of public health impact. The direction and magnitude of exposure measurement error can lead to incorrect estimates of the exposurehealth effect. Our results show that the exposure-health effect may be underestimated for individuals with overestimated exposure. Conversely, for those whose exposure is underestimated, their health risks after exposure to  $PM<sub>2.5</sub>$ tend to be overstated. Inefective public health and environmental interventions can be introduced by biased exposure-health efects as a result of exposure mismeasurement. Low-income and ethnic minorities have been burdened with more fnancial restrictions (Bae et al. [2007\)](#page-11-0). They are also found to be exposed to high air pollution, which makes them doubly disadvantaged (Bae et al. [2007;](#page-11-0) Elliott and Smiley [2019](#page-12-0); Gilbert and Chakraborty [2011;](#page-12-29) Houston et al. [2004;](#page-12-1) J. Kim and Kwan [2021\)](#page-12-21). Low-income people generally spend more time exposed to traffic during commuting or live in areas with poorer air quality, thus increasing exposure. Also, higher incidence of co-morbidities, nutritional deficiency, and less access to information and education due to lack of economic resources impose an increased vulnerability for socially disadvantaged groups. It is important for policymakers to account for individual's exposure at not only places of residence but also all other activity locations. Accurate exposure measurements can help policymakers develop public health policies that refect the interests of all people.

Several limitations in this study need to be addressed in future research. First, the human movement data used in this study were simulated from an activity-based travel demand model. Given this model simulated people's daily travel trajectories for a typical weekday in 2019, it was assumed that individuals have constant activity patterns throughout the year. However, people's daily mobility patterns are not consistent over time and may vary across weekdays, weekends, or seasons (Susilo and Kitamura [2005;](#page-13-14) Xianyu et al. [2017](#page-13-15)). It is debatable whether people's varied travel behaviors on a diferent day (e.g., weekend, holiday) can generate similar exposure classification patterns identified in this study. Thus, more effort should be placed into studying how diferent travel behaviors over time can afect exposure measurement error by collecting human movement data covering multiple time periods.

Second, in this study, only ambient  $PM<sub>2.5</sub>$  exposure was estimated as indoor  $PM_{2.5}$  data were not available. Recent evidence shows that people spend more time indoors (e.g., at home, workplace, and school) during the day, especially those with less mobility (Lu [2021;](#page-12-4) Park [2020\)](#page-12-5). Staying indoors may provide some protection from sources of ambient air pollution (e.g., traffic emission), leading to different results in air pollution exposure assessments. Future exposure research should consider indoor and outdoor  $PM_{2.5}$ concentrations to measure individual exposure accurately.

Third, the unique characteristics of demographic composition and land use layout are recognized for Los Angeles. As a result, the spatiotemporal mobility patterns and ground-level  $PM<sub>2.5</sub>$  concentration distribution depicted in this study only represent the study area's distinct features. The population mobility pattern, spatiotemporal variabilities of air pollution concentrations, sociodemographic mix, and land use layout are expected to vary across diferent regions. Further research is needed to examine whether fndings from this study can be applied to other areas.

# **Conclusion**

Ignoring human mobility in exposure estimates can lead to erroneous exposure assessments and inefective policy implications. Prior research has emphasized the importance of human mobility in estimating air pollution exposure, but little is known about how human mobility might lead to exposure measurement errors. To fll the literature gap, this study assesses residence-based and mobilitybased  $PM_{2.5}$  exposures for 100,784 Los Angeles County residents. It examines the impact of mobility and sociodemographic variables on potential exposure measurement errors. Detailed human mobility data was integrated with hourly  $PM_{2.5}$  surfaces at the neighborhood level. The finding suggests that the magnitude of exposure measurement error is linked to people's mobility levels, and individuals' sociodemographic variables drive the direction of exposure measurement errors. Individuals with high mobility levels are likely to have increased exposure measurement errors. High income and low residential pollution are associated with exposure underestimation, and low income and high residential pollution levels are associated with exposure overestimation. The exposure measurement error introduced by the residence-based method can further lead to erroneous conclusions on the relationship between exposure and health risks. Policymakers should take into account human mobility and sociodemographic characteristics in exposure assessment and ensure that their policies refect not only the preferences of socially advantaged populations but also the interests of disadvantaged populations.

**Author contributions** Yougeng Lu: conceptualization, formal analysis, methodology, software, data curation, writing – original draft preparation

**Data availability** The datasets of this study are available upon request.

### **Declarations**

**Ethics approval** Not applicable.

**Consent to participate** Not applicable.

**Consent for publication** Not applicable.

**Competing interests** The author declares no competing interests.

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