#### **REVIEW ARTICLE**





# Endoscopic Dilation of Bariatric RNY Anastomotic Strictures: a Systematic Review and Meta-analysis

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#### Abstract

Gastrojejunostomy anastomotic strictures are a complication of Roux-en-Y gastric bypass surgery without an established treatment guideline. A systematic review and meta-analysis were performed to determine the safety and efficacy of endoscopic dilation in their management. PubMed, Web of Science, and Cochrane Central (1994–2017) were searched. Data was analyzed with random effects meta-analysis and mixed effects meta-regression. Twenty-one observational studies (896 patients) were included. The stricture rate for laparoscopic patients was 6% (95% CI, 5–9%). Only 38% (95% CI, 30–47%) required greater than one dilation. Symptom improvement occurred in 97% (95% CI, 94–98%). The complication rate was 4% (95% CI, 3–6%). Endoscopic dilation of GJA strictures is safe, effective, and sustaining. This study can guide endoscopists in the treatment of a common bariatric surgical complication.

**Keywords** Anastomosis, surgical/adverse effects · Anastomosis, Roux-en-Y · Endoscopy, gastrointestinal · Constriction, Pathologic/etiology/therapy · Adults · Humans

# Introduction

Roux-en-Y (RNY) gastric bypass is a common bariatric procedure during which a surgeon aims to create an anastomosis that is a tight seal, but free of tension, with an adequate blood supply, and a luminal restriction that supports weight loss. Despite this intent, gastrojejunostomy anastomotic (GJA) strictures are a

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common complication in the early postoperative period. RNY gastric bypass patients suffer from symptoms and incur direct and indirect costs for medical care related to GJA strictures. There is no reported safe and effective management strategy identified. To date, there is only one meta-analysis describing endoscopic dilation of RNY GJA strictures [1]. Systematic review and meta-analysis were performed to understand the safety and effectiveness of endoscopic dilation in treating bariatric RNY gastric bypass patients with GJA strictures.

## Methods

The present study was conducted according to the recommendations of the Cochrane Handbook for Systematic Reviews of Interventions [2] and the Meta-analysis Of Observational Studies in Epidemiology (MOOSE) reporting guidelines for observational studies [3].

# Selection Criteria

## Participants

Studies with patients who met institutional qualification for bariatric surgery, underwent RNY gastric bypass, were symptomatic (i.e., nausea, vomiting, oral intolerance, abdominal pain, and dysphagia) and diagnosed with GJA strictures by upper endoscopy were included. Laparoscopic RNY surgeries were targeted by selecting studies between 1994 and 2017; however, studies that included open technique were not excluded.

#### **Endoscopic Intervention**

Studies with detailed descriptions of endoscopic dilation of GJA strictures, including diameter of dilation, number of dilation sessions, timing of diagnosis or intervention, and the type of dilator used were included. Both Savary-Gilliard and balloon dilation were accepted; however, combinations of interventions were excluded.

## Outcomes

Studies that addressed patient symptomatic improvement after dilation as well as those that needed surgical revision for either complications of dilation or for failure of improvement were included. Studies were not excluded based on follow-up period.

# **Study Design**

Randomized controlled trials, cohort studies, and case-control studies were included, but meeting abstracts, case reports, editorials, and reviews were excluded.

# **Database Search**

Database searches targeting articles that addressed the clinical question were performed using PubMed, Web of Science, and Cochrane Central during July and August 2017. The PubMed search phrase used was "Gastric Bypass" [Mesh] OR "Anastomosis, Roux-en-Y" [Mesh] OR "gastric bypass" [Text Word] OR Roux-en-Y [Text Word] OR Anastomosis[Text Word]) AND ("Constriction, Pathologic" [Mesh] OR stenosis [Text Word] OR stricture[Text Word]) AND ("Dilatation" [Mesh] OR "Endoscopy, Gastrointestinal" [Mesh] OR dilation [Text Word] OR dilatation[Text Word] OR endoscopic[Text Word] OR endoscopy[Text Word]) with limits English language, 1994-present, and Adult 19+. The Web of Science search phrase used was ("gastric bypass" OR Roux-en-Y OR Anastomosis) and (stenosis OR stricture) and (dilation OR dilatation OR endoscopic OR endoscopy) with limits 1994-2017. The Cochrane Central search used Mesh and text words listed above and publication year 1994-2017 (Supplementary Material A). Citations were saved in a citation management software (EndNote) and duplicates were removed.

#### **Selection Process**

Two authors (A.B. and A.A. or D.E.) independently reviewed all collected titles/abstracts and decided yes/no/uncertain regarding relevance to the study. All investigators then met to resolve differences and determine final selection of abstracts for full-text retrieval. Two authors (A.B. and A.A. or D.E. or R.H.) then independently reviewed the selected full-text articles and decided yes/no/uncertain. Authors again met to resolve differences and identify the final set of articles for inclusion and data retrieval. Following final article selection, clinical trial registries (metaRegister of Controlled Trials, ClinicalTrials.gov, Australian Clinical Trials Registry, UK Clinical Trials, and WHO Portal) were searched, and citation lists of six key studies used and three relevant review articles were reviewed for possibly missed articles of relevance (i.e., snowballing).

## **Assessment of Study Quality**

The quality of all studies was assessed by two authors (A.B. and D.E.) using the Newcastle-Ottawa Scale (NOS) for Cohort Studies [4].

# **Data Extraction**

Four authors (A.B., R.H., M.F., and A.C.) independently extracted and input data into a standardized Excel data sheet containing the following: (1) patient characteristics: demographics (age, sex, and BMI), number of laparoscopic and open RNY patients with and without anastomotic strictures, number of patients with anastomotic ulcers, and description of stenosis; (2) intervention characteristics: timing of diagnosis/intervention, type of dilator used, use of fluoroscopy or steroids, diameters of dilator used on first session and during the total treatment period, number of dilation sessions, and time between dilation sessions; and (3) outcomes: symptom improvement, surgical revision, complications of dilation, duration of follow-up, and percentage of excess weight loss (% EWL) or change in BMI for stricture and non-stricture patients.

## **Data Missing**

Data was collected in its reported form and converted to the most common unit whenever possible. Data that could not be attained or converted was imputed for analysis. A descriptive summary or weighted estimates of the remaining available data was made. Data was retrieved only from articles; authors of studies were not contacted to confirm or obtain data. literature search results

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#### **Statistical Analysis**

All analysis was conducted with R statistical software [5], including the meta [6] and metafor packages [7]. Missing observations were imputed using the mice R package [8].

Descriptive statistics for number of: female and male patients; mean: age, BMI, and timing of diagnosis/ intervention; minimum and maximum: dilator diameter on first endoscopy, dilator diameter during total treatment period, time between dilation sessions, and duration of follow-up; and maximum: number of dilation sessions were summarized using weighted means, standard errors, and range (minimum, maximum).

Meta-analysis and meta-regression were implemented by fitting random effects and mixed effects regression models to account for both within and between study variations. The restricted maximum likelihood, inverse variance, and Clopper-Pearson confidence interval methods were used. Test statistics for estimating heterogeneity included the  $\tau^2$ ,  $I^2$ ,  $H^2$ , and  $R^2$ . Chi-square test (p value < 0.05) and a test of significance of coefficients was conducted to test for the significance of the residual heterogeneity. Both the logit transformed proportions, with continuity correction of 0.5 for studies with zero cell frequencies, and the Freeman-Tukey double arcsine transformed proportions of the outcome were compared [9]. Results were reported as odds ratios and their 95% confidence intervals.

Publication bias and sensitivity analysis were conducted using the rank correlation test and the regression test for funnel plot asymmetry [10, 11].

Study	RNY strictures ( <i>n</i> ) Total open/lap	Associated anastomotic ulcers (n)	Time to diagnosis/ intervention (days) Mean Min/max	One dilation session (n	> 1 dilation ) session ( <i>n</i> )	Dilator diameter first endoscopy (mm) Min/max	Dilator diameter treatment period (mm) Min/max	Duration follow-up (months) Mean (Min/max)	Symptom improvement (n)	Required surgery (n)	Complication ( <i>n</i> )
Ahmad et al. (2003) [13]	14	0	81	6	5	15/18	15/25	18	14	0	0
Alasfar et al. (2009) <b>[14</b> ]	0/14 29	1	9/4/1 52 22/200	25	4	NA/NA	10/12	NA NA	29	0	0
3arba et al. (2003) [15]	0/29 24	0	25/309 NA	16	8	12/NA	12/18	NA/NA NA	24	0	0
Caro et al. (2008) [16]	14/10 111 2/11	0	30/2/0 56 2/277	76	35	NA/NA	NA/NA	0/NA NA	111	0	ŝ
Carrodeguas et al. (2006) [17]	94 94	0	52.7 52.7	NA	NA	NA/NA	10/16.5	NA/NA NA	94	0	2
Catalano et al. (2007) [ <b>18</b> ]	NA/NA 26	1	120 120 20/540	7	19	8/10	10/15	NA/NA 26	25	1	0
Da Costa et al. (2011) [ <b>19</b> ]	0/20 105	27	90 90 20020	60	45	8/18	8/20	0/38 NA	105	0	4
Dolce et al. (2009) [20]	0/105 15 0/1 2	2	30/2/0 52	9	5	NA/NA	10/18	NA/NA NA	14	1	0
<sup>5</sup> Spinel et al. (2011) [21]	0/15 22	4	24/99 126	15	L	12/15	12/20	NA/NA 27	22	0	1
Frutos et al. (2009) [22]	0/22	0	26/768 46.8	18	5	NA/NA	10/15	33.7	23	0	0
Go et al.	0/23 38 7/21	0	15/93 53.9 21/178	NA	NA	12/15	NA/NA	0.9/8/.3 NA	36	2	1
(2004) [25] Huang et al. (2003) [24]	9	5	21/108 539	5	2	NA/NA	11/15	NA NA	6	1	0
Lee et al.	NA/NA 40	NA	35/4921 NA	17	23	NA/NA	NA/18	NA/NA NA	40	0	0
(2002) [26] (2009) [26] Mathew et al. (2009)	NA/NA 58	0	NA/NA 66.2	24	34	6/20	6/20	NA/12 NA	58	0	4
Vguyen et al. (2003) [27]	15/43 29 2000	0	12/365 46	24	5	NA/NA	NA/NA	NA/12 37	29	0	0
Peifer et al. (2007) [28]	0/29 43 26/17	З	2017/20 2017/2	34	6	NA/NA	9/20	NA/NA NA 2/12	42	1	1
Ribeiro-Parenti et al. (2015) [29	51 51 55	0	24/13/ 30 20/00	44	٢	13/15	13/15	21/12 NA 24/146	51	0	0
Rondan et al. (2012) [30]	16/0 16	0	30/90 33 12/00	12	4	NA/NA	12/12	14.4 14.4 2.5.4.2	16	0	0
Ryskina et al. (2010) [31]	0/10 58	10	49	19	39	8/15	NA/NA	19.9	58	0	0
Schwartz et al. (2004) [ <b>32</b> ]	NA/NA 32 MA/NA	×	NA/NA NA NA/NA	13	17	NA/NA	NA/NA	NA/NA NA NA MA	20	12	4
Ukleja et al. (2008) <b>[33</b> ]	61	8	09	17	44	NA/NA	6/18	NA	58	ю	3
	0/61		30/180					NA/NA			

#### Results

# **Study Characteristics**

The initial electronic search retrieved 2452 articles after duplicates were removed, which were screened by title and abstract (Fig. 1). Of those, 59 were selected for full-article review. Upon final review, 21 articles met inclusion criteria and were chosen for data extraction and analysis [12–32]. The gray literature search and snowballing revealed no additional relevant studies. No randomized controlled trials were identified. Over 50% of studies had an NOS score of  $\geq 6$  (Supplementary Material B).

# **Patient Characteristics**

Results were pooled for 896 stricture patients (635 laparoscopic, 62 open, and 201 unreported) with available data summarized in Table 1. A weighted summary of the observed data available revealed a weighted mean of mean age  $41.37 \pm 2.28$  (range 38.40-46.00; n = 566) and mean BMI  $47.30 \pm 3.20$  (range 43.79-55.20; n = 485). Studies tended to have a higher number of female stricture patients with a weighted mean number of females  $39.80 \pm 17.90$ (range 11.00-71.00 patients; n = 559) compared to a weighted mean number of males  $25.40 \pm 27.10$  (range 1.00-75.00 patients; n = 559). The remaining studies did

**Fig. 2** Forest plots showing **a** the low laparoscopic stricture rate (proportion 0.06; 95% CI, 0.05–0.09) and **b** the low-ulcer rate in all stricture patients (proportion 0.06; 95% CI, 0.03–0.12)

a	Study	Events	Total	Proportion	95%-CI	Weight (fixed)	Weight (random)
	otady	Liento	rotai		0070-01	(incea)	(rundoni)
	Ahmad 2003	14	450	- 0.03	[0.02; 0.05]	2.5%	6.3%
	Alasfar 2009	29	126		[0.16; 0.31]	4.2%	6.7%
	Barba 2003	10	90	0.11	[0.05; 0.19]	1.7%	5.8%
	Caro 2008	111	1800	0.06	[0.05; 0.07]	19.5%	7.2%
	Da Costa 2011	105	1330	0.08	[0.07; 0.09]	18.1%	7.2%
	Dolce 2009	15	159	0.09	[0.05; 0.15]	2.5%	6.3%
	Espinel 2011	22	525	.0.04	[0.03; 0.06]	3.9%	6.6%
	Frutos 2009	23	676	- 0.03	[0.02; 0.05]	4.2%	6.7%
	Mathew 2009	43	385	0.11	[0.08; 0.15]	7.1%	6.9%
	Nguyen 2003	29	185	0.16	[0.11; 0.22]	4.6%	6.7%
	Peifer 2007	17	507	- 0.03	[0.02; 0.05]	3.1%	6.4%
	Ribeiro-Parenti 2015	51	1500	+ 0.03	[0.03; 0.04]	9.2%	7.0%
	Rondan 2012	16	338	* 0.05	[0.03; 0.08]	2.9%	6.4%
	Schwartz 2004	32	959	⊢ <u> </u> 0.03	[0.02; 0.05]	5.8%	6.8%
	Ukleja 2008	61	1012	0.06	[0.05; 0.08]	10.7%	7.1%
	Fine deffect medel		40040		TO 00. 0 071	400.00/	
	Fixed effect model		10042	0.06	[0.06; 0.07]	100.0%	400.00
	Random effects mode	2 - 0 4265		0.06	[0.05; 0.09]		100.0%
	Heterogeneity. $I = 92\%$ , 1	t - 0.4305	, <i>ρ</i> < 0.0	05 01 015 02 025 03			
				0.05 0.1 0.15 0.2 0.25 0.5			
b	Study	Evente	Total	Proportion	95%-CI	Weight (fixed)	(random)
	Study	Evenus	Total	Fioporadi	90 /0-CI	(lixed)	(ranuoni)
	Ahmad 2003	0	14 •	0.00	[0.00; 0.23]	0.8%	3.4%
	Alasfar 2009	1	29	0.03	[0.00; 0.18]	1.7%	4.7%
	Barba 2003	0	24 •	- 0.00	[0.00; 0.14]	0.8%	3.5%
	Caro 2008	0	111・	0.00	[0.00; 0.03]	0.9%	3.5%
	Carrodeguas 2006	0	94 •	0.00	[0.00; 0.04]	0.9%	3.5%

Alinau 2005	0	14					0.00	[0.00, 0.23]	0.070
Alasfar 2009	1	29 -+	<u>+</u>				0.03	[0.00; 0.18]	1.7%
Barba 2003	0	24					0.00	[0.00; 0.14]	0.8%
Caro 2008	0	111 ⊢	1				0.00	[0.00; 0.03]	0.9%
Carrodeguas 2006	0	94 ⊢					0.00	[0.00; 0.04]	0.9%
Catalano 2007	1	26 -++	<u> </u>				0.04	[0.00; 0.20]	1.7%
Da Costa 2011	27	105		_			0.26	[0.18: 0.35]	34.7%
Dolce 2009	2	15 -	-				0.13	[0.02: 0.40]	3.0%
Espinel 2011	4	22					0.18	[0.05: 0.40]	5.7%
Frutos 2009	0	23 ⊷	_				0.00	[0.00: 0.15]	0.8%
Go 2004	0	38					0.00	[0.00; 0.09]	0.9%
Huang 2003	5	7					0.71	[0.29: 0.96]	2.5%
Lee 2009	0	40 ⊷	-				0.00	[0.00; 0.09]	0.9%
Mathew 2009	0	58 -					0.00	[0.00; 0.06]	0.9%
Nauven 2003	0	29	- 1				0.00	[0.00: 0.12]	0.8%
Peifer 2007	3	43 -	÷				0.07	[0.01: 0.19]	4.8%
Ribeiro-Parenti 2015	0	51 H					0.00	[0.00: 0.07]	0.9%
Rondan 2012	0	16	<u> </u>				0.00	[0.00; 0.21]	0.8%
Ryskina 2010	10	58	-18-				0.17	[0.09; 0.29]	14.3%
Schwartz 2004	8	32					0.25	[0.11: 0.43]	10.4%
Ukleja 2008	8	61					0.13	[0.06; 0.24]	12.0%
Fixed effect model		896	ė.				0.16	[0.12; 0.19]	100.0%
Random effects model		4	$\geq$				0.06	[0.03; 0.12]	
Heterogeneity: $I^2 = 70\%$ , $\tau^2$ :	= 1.7226,	p < 0.01	1	1	1				
	,	0	0.2	0.4	0.6	0.8			

4.7%

7.4%

5.7%

6.4%

3.5%

3.5%

5.4%

3.5%

3.5%

3.5%

6.3%

3.5%

3.5%

7.1%

6.9%

7.0%

100.0%

Fig. 3 a Forest plot describing the rate of patients requiring greater than one endoscopic dilation session during the total treatment period (proportion 0.38; 95% CI, 0.30-0.47). b Bubble chart showing the relationship between the proportion of patients requiring greater than one endoscopic dilation and the minimum dilator diameter (mm) used on the first endoscopy. Sizes of circles are proportional to the study cohort size. A solid regression line is plotted with dotted lines indicating the 95% CI (OR, 0.86; 95% CI, 0.75-0.98; p = 0.03)



not separate demographic data for stricture and nonstricture patients. All patients were symptomatic, but five studies also used imaging in evaluation prior to endoscopy [16, 17, 23–25]. Eight studies mentioned stricture diameters (data not shown), but the majority reported that the diameter was less than 10 mm or that the scope was unable to pass (8.5–10 mm endoscope) [16–18, 24, 26, 27, 29, 32]. Not all studies reported the number of laparoscopic versus open RNY patients with anastomotic strictures, but for the data available (n = 578), there was an anastomotic stricture rate for laparoscopic patients of 6% (95% CI, 5–9%;  $l^2 = 92\%$ ; p < 0.01; Fig. 2a).

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8

10

Minimum Dilator Diameter (mm) on First Endoscopy

#### Intervention

12

The weighted mean for mean timing of diagnosis or intervention of the observed data available was  $67.06 \pm 54.82$  days after surgery (range 30.00-539.00; n = 802). The minimum and maximum timing reported in any study was 3.00 and 4921.00 days after surgery. Only 51 of 896 patients received the Savary-Gilliard dilation while the clear majority received the through-the-scope (TTS) balloon dilation [28]. Only one study (26 patients) used steroid injections to supplement dilation [17], and three studies (106 patients) dilated under fluoroscopy [17, 26, 28]. The weighted mean minimum and

16

14

maximum diameter of the dilator used on the first endoscopy was  $9.45 \pm 2.62 \text{ mm}$  (range 6.00-15.00; n = 396) and  $16.39 \pm 2.57 \text{ mm}$  (range 10.00-20.00; n = 372) respectively. The weighted mean minimum and maximum diameter of the dilator used during the total treatment period was  $9.37 \pm 2.33 \text{ mm}$  (range 6.00-15.00; n = 590) and  $17.71 \pm 2.63 \text{ mm}$ (range 12.00-25.00; n = 630) respectively. The weighted mean minimum and maximum weeks between endoscopic dilations was  $1.64 \pm 1.18 \text{ mm}$  (range 0.40-7.00; n = 245) and  $6.04 \pm 6.58 \text{ mm}$  (range 2.00-25.00; n = 245) respectively. The weighted mean maximum sessions of dilation for the observed data available was  $4.43 \pm 1.20$  sessions (range 3.00-7.00; n = 787).

Only 38% (95% CI, 30–47%;  $I^2 = 81\%$ ; p < 0.01; Fig. 3a) of patients required greater than one dilation session during their treatment period. BMI (OR, 0.86; 95% CI, 0.77–0.97; p = 0.01), minimum dilator diameter on first endoscopy (MiDF) (OR, 0.79; 95% CI, 0.70–0.89; p < 0.001), and during the total treatment period (MiDT) (OR, 0.85; 95% CI, 0.76–0.95; p = 0.01) negatively predicted the need for repeat dilation. After adjusting for multiple variables, MiDF was still predictive of outcome. (OR, 0.86; 95% CI, 0.75–0.98; p = 0.03; Fig. 3b).

#### Outcomes

The weighted mean for mean duration of follow-up of the observed data available was  $25.29 \pm 7.31$  months (range 14.40–37.00; n = 188). The minimum and maximum followup duration reported in any study was 0.90 and 146.00 months. Meta-analysis showed that 97% (95% CI, 94–98%;  $I^2 = 67\%$ ; p < 0.01; Fig. 4a) of patients' symptoms responded safely to endoscopic dilation compared to 3% (95% CI, 2–6%;  $I^2$  = 67%; p < 0.01; Fig. 4c) that failed to respond (n = 15) or suffered a complication (n = 8) and required surgery. For symptom improvement, maximum dilator diameter on the first endoscopy (MxDF) was a strong positive predictor (OR, 1.45; 95% CI, 1.21–1.70; p < 0.001) while MiDT was a strong negative predictor (OR, 0.78; 95% CI, 0.63–0.97; *p* = 0.02). However, after adjusting for multiple variables in meta-regression, only MxDF was a strong predictor (OR, 1.36; 95% CI, 1.11-1.70; p = 0.004; Fig. 4b). The same predictors, in opposing direction, were true for requirement of surgical revision (MxDF: OR, 0.69; 95% CI, 0.57–0.83; *p* < 0.001; MiDT: OR, 1.28; 95% CI, 1.04–1.59; p = 0.02) with MxDF as a strong predictor after adjusting for multiple variables (OR, 0.74; 95% CI, 0.60-0.90; p = 0.004).

There was an overall low-complication rate (4%; 95% CI, 3-6%;  $l^2 = 7.1\%$ ; p = 0.86; Fig. 5a). There were 21 (91%) patients with perforation, one with hemorrhage, and one with a subepithelial hematoma. Eight patients underwent surgery for perforations; however, three of these patients did not require revision. The remaining complications were managed

conservatively. There was no strong predictor of complication using logit transformation; however, MiDF trended toward significance (OR, 0.86; 95% CI 0.74–1.00; p = 0.05). On the other hand, using Freeman-Tukey transformation, MiDF was a strong negative predictor of complication (OR, 0.99; 95% CI 0.98–1.00; p = 0.04; Fig. 5b). There was an anastomotic ulcer rate in all stricture patients of 6% (95% CI, 3–12%;  $f^2 = 70\%$ ; p < 0.01; Fig. 2b). Presence of an anastomotic ulcer at the time of diagnosis of anastomotic stricture did not predict complications during endoscopic dilation.

Six studies reported that the percentage of % EWL was not different between stricture and non-stricture patients within their studies [16, 19, 26–28, 30]; however, there was not enough pooled data for analysis.

#### **Sensitivity Analysis**

Sensitivity analysis detected potential bias with asymmetry (p < 0.05) in the funnel plots for ulcer rate and the outcomes of symptom improvement or surgery after dilation. Stricture rate, patient need for greater than one dilation session, and complication rate showed no bias with the absence of significant asymmetry in the funnel plots (Supplementary Material C).

# Discussion

Based on the current available literature, endoscopic dilation of GJA strictures is safe and effective with lowcomplication rates and high rates of symptom improvement after one dilation. Endoscopic dilation response is maintained long-term as GJA stricture patients are followed on average for 2 years.

Laparoscopic procedures are predominant in this review due to the defined selection criteria; thus, comparisons between laparoscopic and open procedures were not made. The overall reported stricture rate in laparoscopic patients is low with the absence of a publication bias. By design, all stricture patients included were symptomatic. Diagnosis of GJA strictures and endoscopic intervention occurs early, on average 2 months after surgery, when RNY patients transition from soft to solid food. Anastomotic strictures can be classified as membranous due to extended fasting; cicatricial from chronic inflammation related to foreign body material, ulceration, and anastomotic leak; and granular from ischemic tissue necrosis related to anastomotic tension, smoking, marginal ulceration, and lack of apposition [33, 34]. Membranous strictures respond well to endoscopic dilation [33, 34]. This is likely the most common stricture type with over 60% of patients responding to one dilation session.

TTS balloon dilation is far more common than wire-guided bougie dilation as it allows for direct targeting of the surgical

Fig. 4 a Forest plot describing the high rate of symptom improvement with endoscopic dilation (proportion 0.97; 95% CI, 0.94-0.98). b Bubble chart showing the relationship between the proportion of patients with symptom improvement and the maximum dilator diameter (mm) on the first endoscopy. Sizes of circles are proportional to the study cohort size. A solid regression line is plotted with dotted lines indicating the 95% CI (OR, 1.36; 95% CI, 1.11-1.70; p = 0.004). **c** Forest plot describing the low rate of surgery after endoscopic dilation (proportion 0.03; 95% CI, 0.02-0.06)



							Weight	Weight
Study	Events	Total			Proportion	95%-CI	(fixed)	(random)
Ahmad 2003	14	14			1.00	[0.77; 1.00]	2.1%	3.8%
Alasfar 2009	29	29			1.00	[0.88; 1.00]	2.1%	3.8%
Barba 2003	24	24			1.00	[0.86; 1.00]	2.1%	3.8%
Caro 2008	111	111		-	1.00	[0.97; 1.00]	2.2%	3.8%
Carrodeguas 2006	94	94			1.00	[0.96; 1.00]	2.2%	3.8%
Catalano 2007	25	26			0.96	[0.80; 1.00]	4.2%	5.5%
Da Costa 2011	105	105			1.00	[0.97; 1.00]	2.2%	3.8%
Dolce 2009	14	15			0.93	[0.68; 1.00]	4.1%	5.5%
Espinel 2011	22	22			1.00	[0.85; 1.00]	2.1%	3.8%
Frutos 2009	23	23			1.00	[0.85; 1.00]	2.1%	3.8%
Go 2004	36	38			0.95	[0.82; 0.99]	8.3%	7.2%
Huang 2003	6	7 -			0.86	[0.42; 1.00]	3.7%	5.2%
Lee 2009	40	40			1.00	[0.91; 1.00]	2.2%	3.8%
Mathew 2009	58	58			1.00	[0.94; 1.00]	2.2%	3.8%
Nguyen 2003	29	29			1.00	[0.88; 1.00]	2.1%	3.8%
Peifer 2007	42	43			0.98	[0.88; 1.00]	4.3%	5.6%
Ribeiro-Parenti 2015	51	51			1.00	[0.93; 1.00]	2.2%	3.8%
Rondan 2012	16	16			1.00	[0.79; 1.00]	2.1%	3.8%
Ryskina 2010	58	58			1.00	[0.94; 1.00]	2.2%	3.8%
Schwartz 2004	20	32		- 11	0.62	[0.44; 0.79]	32.8%	9.4%
Ukleja 2008	58	61			0.95	[0.86; 0.99]	12.5%	8.1%
Fixed effect model		896			0.93	[0.90; 0.95]	100.0%	
Random effects mode	1			$\diamond$	0.97	[0.94; 0.98]		100.0%
Heterogeneity: $I^2 = 67\%$ , a	<sup>2</sup> = 1.1680	p < 0.0	)1					
- / /			0.5 0.6 0.7	0.8 0.9 1				

Fig. 5 a Forest plot showing the low-complication rate of endoscopic dilation (proportion 0.04: 95% CI. 0.03-0.06). b Bubble chart showing the relationship between the proportion of patients with complication and the minimum dilator diameter (mm) on first endoscopy. Sizes of circles are proportional to the study cohort size. A solid regression line is plotted with dotted lines indicating the 95% CI (OR, 0.86; 95% CI 0.74–1.00; p = 0.05)



1.9%

20%

20%

10.3%

7.2%

2.0%

12.9%

1.9%

37%

2.0%

3.8%

1 9%

2.0%

12.6%

2.0%

3.8%

20%

2.0%

2.0%

12.0%

10.1%

100.0%

anastomotic stricture. Larger dilator diameter selection promotes positive outcomes. Early response to endoscopic therapy is predicted by a larger MiDF without publication bias. Likewise, a larger MxDF promotes overall symptom improvement and lessens the need for surgery. It is safe to choose a 2-6-week interval between dilations.

6

8

10

Minimum Dilator Diameter (mm) on First Endoscopy

12

The overall complication rate is low (4%) without publication bias, and most complications are conservatively managed with few requiring surgical revision. Although the rate is low, 91% of complications are perforations. This rate is higher than the perforation rate seen in esophageal dilations for eosinophilic esophagitis (<1%) [35, 36]. The complication rate found in this meta-analysis is more comparable to that associated with dilation of Crohn's disease strictures (4% complication rate and 3% perforation rate), in which 65% of patients had anastomotic strictures [37]. Thus, anastomotic strictures may carry higher risk of perforation than intrinsic strictures.

16

14

Complication, interestingly not predicted by an associated ulcer, is predicted by smaller MiDF. The reason is likely related to the small size of the stricture itself. As a stricture becomes smaller in diameter, the squared area decreases dramatically. Thus, when dilating a small stricture incrementally there is higher risk with each increment compared to when

dilating a larger stricture. For instance, a 5-mm diameter stricture (78.50 mm<sup>2</sup>) dilated to 7 mm (153.86 mm<sup>2</sup>) is a 96% increase in the squared area; whereas, a 9-mm diameter stricture (254.34 mm<sup>2</sup>) dilated to 11 mm (379.94 mm<sup>2</sup>) is only a 33% increase in the square area. Although larger dilator diameters on the first endoscopy promote symptom improvement and lessen the need for repeat sessions, when facing a small GJA stricture, the endoscopist should begin with the smallest dilator and use extra caution with each incremental increase as this is the time at the highest risk of perforation.

Despite that the main objective of bariatric surgery is weight loss, there was only six comparative (one matched) studies reporting % EWL, but at various follow-up intervals. In the future, researchers should pay attention to % EWL in stricture and non-stricture patients and how it is affected by endoscopic dilation.

There are several limitations to this systematic review and meta-analysis that reflect the restrictions of the included studies. First, all the studies included are observational cohort studies. Significant heterogeneity (p < 0.05) was present in meta-analysis, except for complication rate, indicating that multiple effect sizes are likely present, potentially due to variability in populations, interventions, and reporting parameters of the studies. Random effects and mixed effects regression models were fitted to account for both within and between study variations given the observational nature of the studies. Data was imputed when not available in the meta-regression analysis. Lastly, sensitivity analysis showed potential for publication bias for ulcer rate and outcomes of symptom improvement or surgery after dilation, but not for stricture rate, complication rate, and patients receiving only one endoscopic dilation. Reporting parameters for the former data varied between studies, and therefore data extraction may over or underestimate incidence or treatment effects.

The strengths of this systematic review and meta-analysis are the large number of included studies, its detailed data extraction, the rating of the quality of the cohort studies, and statistical methods used to account for heterogeneity and bias in the studies. No other systematic review and meta-analysis currently exists that provides a comprehensive guide on endoscopic treatment of GJA strictures.

# Conclusion

Endoscopic dilation of GJA strictures, typically occurring in the early post-operative period, is safe and effective with most patients responding to one dilation session. Strictures should be objectively measured before each dilation. With small strictures, the smallest dilator diameter should be selected and incrementally increased with caution. Dilations may be safely repeated every 2–6 weeks based on symptom resolution. Objective stricture measurement can also aid in evaluating the need for adjunctive therapy, such as steroid injection, in refractory strictures when following patients prospectively. If a complication, like perforation, occurs, it can generally be managed conservatively. Further studies are needed to standardize the approach to assessment, treatment, and follow-up of GJA strictures. This systematic review and meta-analysis provides a guide to endoscopists in treatment of a common bariatric surgical complication.

#### **Compliance with Ethical Standards**

**Conflict of Interest** The authors declare that they have no conflict of interest.

**Ethical Statement** This article does not contain any studies with human participants or animals performed by any of the authors.

**Informed Consent Statement** For this type of study formal consent is not required.

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