RESEARCH ARTICLE

Modelling the impacts of climate change on cereal crop production in East Africa: evidence from heterogeneous panel cointegration analysis

Abdikaf Hassan Abdi1,2 · Abdimalik Ali Warsame1,2 · Ibrahim Abdukadir Sheik‑Ali2,3

Received: 16 August 2022 / Accepted: 11 December 2022 / Published online: 17 December 2022 © The Author(s), under exclusive licence to Springer-Verlag GmbH Germany, part of Springer Nature 2022

Abstract

Climate change has become an issue of concern for sustainable agriculture production. East African nations are heavily reliant on the agriculture sector, which accounts for a substantial amount of their gross domestic product (GDP) and employment. Due to climatic fuctuations, the output of the sector became very unpredictable. Hence, this study investigates the efects of climate change on cereal crop production in nine East African nations between 1990 and 2018. The study implemented pooled mean group (PMG) approach to examine the long-run and short-run dynamic impacts of the varying climatic circumstances on the output of cereal crops. The results reveal that rainfall and carbon emissions have favourable and signifcant long-run efects on cereal crop output, even though their short-run impacts are negligible. Additionally, cultivated land area and rural population have a constructive role in enhancing agricultural output both in the long-run and short-run. However, average temperatures have negative repercussions on cereal crop production in the long-run and short-run, even though the magnitude of sensitivity is greater in the short-run. Dynamic ordinary least squares (DOLS) and fully modifed ordinary least squares (FMOLS) validated the robustness of the long-run fndings of the PMG technique. Besides, the Dumitrescu–Hurlin panel causality outcomes indicate that cereal crop output has a bidirectional causality with temperature, carbon emissions, and cropped area. The study further demonstrated unidirectional causation from rural population to cereal crop yield. The study recommends that East African policymakers improve the quality of farm inputs, the adoption of climate-resilient farming practices, the development of water retention facilities and the establishment of crop diversifcation initiatives.

Keywords Cereal crop production · Climate change · Pooled mean group · East Africa · Carbon emissions

Introduction

The adverse effects of climatic variations on health, environment, agriculture, and livelihoods have become a global concern in the twenty-frst century. The surge in the average temperature on the planet's surface has been discovered by many researchers (Porter and Semenov [2005;](#page-10-0) Sommer et al.

Responsible Editor: Eyup Dogan

 \boxtimes Abdikafi Hassan Abdi abdikafhasan79@gmail.com

- ¹ Faculty of Economics, SIMAD University, Mogadishu, Somalia
- ² Garaad Institute for Social Research and Development Studies, Mogadishu, Somalia
- Faculty of Economics and Management, Jamhuriya University of Science and Technology, Mogadishu, Somalia

[2013](#page-11-0); Attiaoui and Boufateh [2019\)](#page-9-0). The average global temperature is forecasted to climb by 2.6–4.8 °C in the twentyfrst century (IPCC [2013\)](#page-10-1). More frequent disasters such as excessive rainfall, extreme weather events, a declining ice cover, and modifcations to winter snow are some of the calamitous repercussions of climatic fuctuations that became recurrent threats to the environment (Mohamed et al. [2022;](#page-10-2) Warsame et al. [2022c](#page-11-1)). Agriculture production remains the most vulnerable sector to rainfall volatilities and a spike in average surface temperature (Gay et al. [2006](#page-10-3); Deschenes and Kolstad. [2011](#page-10-4); Van Passel et al. [2016](#page-11-2)). Consequently, agriculture yields plummeted due to climate change factors that undermined the global food security and sustainable livelihoods of low-income countries dependent on the agriculture sector (Warsame et al. [2022b;](#page-11-3) Zhao et al. [2017](#page-11-4)).

Climatic fuctuations are the most signifcant risk source that directly or indirectly influences crop production.

According to Pickson et al. [\(2020\)](#page-10-5), the rising temperature and changes in rainfall patterns have a straight impact on crop growth timing. Arable land becomes less appropriate for agricultural production as a result of surging temperature, which has long-run unfavourable eventuality on crop productivity. Moreover, variations in rainfall patterns are expected to result in both a short-run loss of crops and a long-run decline in crop production (Pickson et al. [2020](#page-10-5)). On the other hand, climatic fuctuations increase insects, weeds, and illnesses, which have an indirect adverse impact on agricultural yield since they make crop management difficult and costly. As a result, farmers face frequent foods, heavy rain, droughts, and fuctuations in market prices (Chandio et al. [2020\)](#page-10-6). Besides these factors, the literature was controversial about the repercussions of rainfall and temperature changes on agricultural production across diferent regions. For instance, Chandio et al. [\(2020](#page-10-6)) confrmed that precipitation and temperature in China have a favourable impact on agrarian value added in the short-run but ruin in the longrun. In contrast, Pickson et al. ([2020](#page-10-5)) noticed that average precipitation and the cropped land area had a signifcant and positive impact on cereal crop output in the short-run and long-run, whereas mean temperature and temperature variations in China have a considerable adverse efect on the long-run cereal production.

East Africa region is encountering an unprecedented increase in temperature. The major cities in this region have witnessed a temperature increase that is almost double the global average of 1.1 °C since the pre-industrial era. For instance, since 1860, Mogadishu (Somalia) has warmed by 1.9 °C, Khartoum (Sudan) by 2.09 °C, Addis Ababa (Ethiopia) by 2.2 °C, Nairobi (Kenya) by 1.9 °C, and Da res Salaam (Tanzania) by 1.9 °C (ICPAC [2021](#page-10-7)). Given the fact that climate change is strongly linked to the farming sector, rain-fed agriculture represents the primary means of livelihood in the East Africa region. Rural population livelihoods and food security in this region are becoming more exposed to the consequences of climatic variations (WWF [2006](#page-11-5)). Agriculture constitutes 40% of the gross domestic product (GDP) and offers livelihoods for 80% of the East African population. Temperature increases and rainfall variability have already afected the agriculture sector in this region (WWF [2006\)](#page-11-5). Nevertheless, a deep understanding of climate change efects, vulnerabilities, and adaptation strategies is required to mitigate global warming and its consequences. Empirical research and initiatives to strengthen the capacity to generate climate data and disseminate have become topical. In this regard, this study empirically examines climate change factors—mean rainfall, temperature, and $CO₂$ emissions—impacts on agricultural output in East African countries.

The agriculture sector is a key ingredient for sustainable economic growth and development in developing countries.

Climate change poses a signifcant threat to the agriculture sector. Comprehensive empirical studies on the climate change-agriculture nexus have been carried out over the globe using various analysis methods. Most studies have concluded that climate change has signifcantly impacted the agricultural output of diferent regions and countries, although the magnitude of the efect varies with the kind of crops. Among studies outside Africa, Chandio et al. ([2022a\)](#page-10-8) ascertained the efect of climate change on maize output in Nepal. The short-run and long-run empirical fndings of the study disclosed that rainfall stimulates maize production, whereas temperature and carbon emissions hamper it. Comparable results were found by Chandio et al. [\(2022b\)](#page-10-9) in India, Sadiq et al. [\(2018](#page-10-10)) in Pakistan, and Chandio et al. ([2022c](#page-10-11)) in Bangladesh. The growing global population, deforestation, and the utilization of fossil fuels are attributed to the elevated level of greenhouse gases (GHGs), which derives climate change and hampers agriculture production (Pickson et al. [2022\)](#page-10-12). Indeed, climatic variations' impact on agriculture varies depending on the crop production type. Sarker et al. [\(2012](#page-10-13)) found that maximum temperature has a significant adverse effect on Boro rice but has a considerable favourable impact on Aman and Aus rice. Minimum temperature undermines Aman rice but stimulates Boro rice production. However, precipitation has consistently favourable implications for rice production in Bangladesh. Similar results were found by Satari Yuzbashkandi and Khalilian ([2019](#page-10-14)) in Iran. On the contrary, Abbas and Mayo [\(2020\)](#page-9-1) uncovered that maximum temperature undermines rice production while rainfall stimulates it in Pakistan. Using the Ricardian model, Hossain et al. [\(2019\)](#page-10-15) revealed that precipitation and temperature enhance the net crop income in Bangladesh. Thus, this proves that climate change's imminent efects on agriculture are heterogeneous.

The existing empirical studies on climate change–agriculture nexus in Africa also revealed that agriculture production decreases due to climatic fuctuations. The adverse efects of climate change are more felt in developing countries, including Africa, which rely on agriculture for a signifcant portion of their economic activities than developed nations. Notably, Sub-Saharan African (SSA) countries are more exposed to climate variations than other nations due to their lack of climate adaptability and high dependence on agriculture (Sarkodie and Strezov [2019\)](#page-10-16). Among early undertakings, Barrios et al. ([2008](#page-9-2)) conducted a comparative study analyzing the consequences of climate change on agriculture output in SSA and non-SSA countries. They concluded that temperature and average rainfall have substantial negative and positive efects on agriculture production in SSA nations, respectively, whereas climate change is incon-sequential in non-SSA countries. Later, Sultan et al. ([2013\)](#page-11-6) ascertained the infuence of climate change on millet and sorghum output in West African countries. They revealed that rainfall and temperature changes hamper agriculture yield in West African nations. Likewise, Ben Zaied and Ben Cheikh ([2015](#page-9-3)) revealed that mean temperature rise undermines cereal and date yields, except in upland areas, while rainfall enhances cereal and date productions in Tunisia. Alboghdady and El-Hendawy ([2016](#page-9-4)) found that temperature and rainfall changes undermine agricultural yield in the Middle East and North Africa (MENA). In the same vein, Ngoma et al. [\(2021\)](#page-10-17) have observed that temperature and rainfall variabilities have dreadful efects on agriculture production in Zambia. Recently, Warsame et al. ([2021](#page-11-7), [2022a\)](#page-11-8) found that rainfall has a constructive role in enhancing agriculture production in Somalia, whereas average temperature impedes it in Somalia. Using dynamic ARDL simulation, Ntiamoah et al. ([2022\)](#page-10-18) estimated the effect of $CO₂$ emissions and precipitation on maize and soybean production in Ghana. The outcomes of the study indicated that both $CO₂$ emissions and rainfall have a constructive role in enhancing maize and soybean output in Ghana.

The intensity of climate change impacts on agriculture production difers due to the data discrepancy, income level, adaptation capacity, and prevailing climatic conditions. Due to their limited ability to adapt to climatic variations, excessive reliance on agriculture, and the majority of their populations living in rural areas, East African countries are the most vulnerable to climate change (Sarkodie and Strezov [2019](#page-10-16)). Nevertherless, cross-country studies about the region are limited in the existing literature. To guide the formation and carrying out of common climate change policies at the regional level, it is necessary to conduct a panel study because the fndings of the existing individual country studies could not be generalized to large number of countries. Hence, the current study aims to assess the efect of climate change—measured for rainfall, temperature, and $CO₂$ emissions—on cereal crop production in East Africa using panel data during the period 1990 and 2018. The study utilized pooled mean group (PMG) technique to refect the dynamic long-run and short-run impacts of climatic variation on the output of cereal crops. The study employs several panel cointegration methods, such as DOLS and FMOLS, as a robustness test to validate the reliability of the panel ARDL results. Moreover, the study uses Dumitrescu and Hurlin ([2012\)](#page-10-19) panel causality test to determine the direction of the causal relationship between the explanatory variables and cereal crop production.

The subsequent sections of the study are organized as follows: the ["Methodology"](#page-2-0) section presents the data and methodology of the study, empirical analysis and discussion will be presented in the ["Empirical results and discussion"](#page-4-0) section, and fnally, the conclusion and policy implications will be reported in the ["Conclusion](#page-8-0)" section.

Methodology

Data description

The study utilized panel data covering the period 1990–2018 to examine the aftermath of climate change on the agricultural output of nine East African countries (Burundi, Djibouti, Eritrea, Ethiopia, Kenya, Rwanda, Somalia, Tanzania, and Uganda). Due to the consequences of global warming, the region has experienced horrendous repercussions of climate change on agricultural output since it relies hugely upon its livelihoods and employment. According to Nicholson [\(2017](#page-10-20)), the most common climatic scenario in East African counties is the frequent occurrence of acute droughts, mainly resulting from changes in rainfall patterns. Additionally, some of Africa's food-prone countries are located in the region (Li et al. [2016](#page-10-21)). Therefore, the study uses cereal crop production to measure agricultural output, which is infuenced by climatic factors such as average rainfall, temperature, and carbon dioxide emissions (Warsame et al. [2021](#page-11-7)). To avoid model misspecifcation and variable omissions, we incorporate cultivated land area and rural population as crucial determinants of agricultural production (Chandio et al. [2020](#page-10-6); Pickson et al. [2022\)](#page-10-12). Table [1](#page-2-1) demonstrates the variables, symbols, measurement units, and data sources.

Model specifcation

The cereal crop output function utilized in this study was specifed by following Chandio et al. ([2020\)](#page-10-6), Warsame et al. ([2021\)](#page-11-7), Kumar et al. ([2021\)](#page-10-22), and Chandio et al. [\(2021](#page-10-23)). Regarding this, we employed the following panel model to

Table 1 Variable illustration and data sources Variable Variable Symbol Unit of measurement Source of the Source Cereal crop production CP Cereal production (metric tons) WDI Average rainfall **ARF** Average annual precipitation (mm) CCKP Average temperature **AT** Average annual temperature in (°C) CCKP Carbon emissions CO₂ Metric tons per capita of CO₂ emission WDI Land under cereal cultivation LUC Land under cereal production (hectares) WDI Rural population RP people living in rural areas (millions) WDI

investigate the effects of average rainfall, average temperature, carbon dioxide emissions, cultivated land area, and agriculture labor on cereal crop production in the selected countries.

$$
CP = f(ARF, AT, CO2, LUC, RP)
$$
 (1)

where CP denotes cereal crop production, which is the explained variable. The regressors include ARF, which indicates average rainfall, AT signifies temperature, $CO₂$ represents carbon emissions, LUC denotes land under cereal output, and RP indicates the rural population. We converted all the inspected variables into a natural logarithm to ensure that the variables are interpreted as a percentage and reduce the heterogeneity issues frequently appearing in heterogeneous panel data. The reduced form is as follows:

$$
ln CP_{it} = \beta_0 + \beta_1 lnARF_{it} + \beta_2 lnAT_{it}
$$

+ $\beta_3 lnCO_{2it} + \beta_4 lnLUC_{it} + \beta_5 lnRP_{it} + \varepsilon_{it}$ (2)

Based on the panel dataset with groups of countries $i = 1$, 2, 3,...,*N* and periods $t = 1, 2, 3, \ldots, T$. The ln signifies the natural logarithm, ε signifies the white noise error term, and β_0 is the intercept. β_1 , β_4 , and β_5 are the coefficients of mean rainfall, land under cereal output, and rural population, which are expected to have a favourable infuence on cereal crop production. However, β_1 and β_3 are the coefficients of average temperature and $CO₂$ emissions, which are expected to have an unfavourable impact on agricultural output.

Econometric technique

Cross‑sectional dependence test

There is a signifcant possibility of cross-sectional dependence (CD) among the countries in a panel data setting (Appiah et al. [2018\)](#page-9-5). However, disregarding the CD test in the study might lead to invalid test statistics. Based on the test results, we can decide which econometric approach is appropriate for the panel data. Hence, this study investigates the dependency of the cross-sections in the sample using the Pesaran CD test proposed by Pesaran ([2004](#page-10-24)). The Pesaran cross-sectional dependence test is preferred to other tests, including the Breusch-Pagan Lagrange Multiplier (LM) test, because it can account for common components or spillover efects that are not observed in the sample observations (Appiah et al. [2018](#page-9-5)). Besides, the null hypothesis of crosssectional independence H_0 : $\rho_{ij} = \rho_{ji} = cor(e_{it}, e_{jt}) = 0, i \neq j$ is tested against the alternative hypothesis of cross-sectional dependence (*H*₁ ∶ $\rho_{ij} = \rho_{ji} \neq 0$, *forsomei* $\neq j$); ρ_{ji} denotes pair-wise disturbances' correlation coefficient, which is calculated by following the specifcation of Simionescu et al. [\(2021\)](#page-11-9).

$$
\rho_{ij} = \rho_{ji} = \frac{\sum_{t=1}^{T} e_{it} e_{jt}}{\sqrt{\sum_{t=1}^{T} e_{it}^2 \sqrt{\sum_{t=1}^{T} e_{jt}^2}}}
$$
(3)

The CD statistic of Pesaran could be computed as follows in the case of balanced panel data.

$$
CD = \sqrt{\frac{2}{N(N-1)}} \cdot \sum_{i=1}^{N-1} \sum_{j=1+1}^{N} \hat{\rho}_{ij}
$$
 (4)

where *T* is the entire amount of observations shared between nations *i* and *j*.

$$
\widehat{\rho}_{ij} = \widehat{\rho}_{ij} = \frac{\Sigma_t \in T_i \cap T_j(\widehat{e}_{it} - \overline{e}_i)(\widehat{e}_{jt} - \overline{e}_j)}{\sqrt{\Sigma_t \in T_i \cap T_j(\widehat{e}_{it} - \overline{e}_i)^2} \cdot \sqrt{\Sigma_t \in T_i \cap T_j(\widehat{e}_{jt} - \overline{e}_j)^2}}
$$
(5)

$$
\overline{e}_i = \frac{\Sigma_t \in T_i \cap T_j(\hat{e}_{it})}{\#(T_i \cap T_j)}
$$
(6)

Panel unit root tests

The stationarity test of the variables is a prerequisite prior to the panel cointegration test to determine the existence of a unit root and the order of integration. The frst-generation panel unit root tests that presume cross-sectional independence could produce misleading results when there is crosssectional dependence (Breitung and Das [2008\)](#page-10-25). Due to the dependency of the individual panel units, Pesaran [\(2007\)](#page-10-26) proposed the use of second-generation panel unit root tests, such as the cross-sectional ADF (CADF) and the crosssectional augmented IPS (CIPS), to assure the validity and reliability of the results. These tests take into account the potential for cross-sectional dependency due to unknown common factors. The null hypothesis of the two tests assume that the entire panels are non-stationary against the alternative hypothesis that presume some panels are stationary. Hence, the null hypothesis is refuted if the probability values are less than 1%, 5%, and 10% signifcant levels. Unlike traditional panel cointegration methods, the panel ARDL has the capability to handle variables of various orders of integration.

Panel cointegration analysis

Before the central estimation of the panel ARDL model, it is also required to examine the existence of a cointegration relationship among the variables by applying Pedroni ([1999,](#page-10-27) [2004\)](#page-10-28) and Kao [\(1999\)](#page-10-29). We have compared the alternative hypothesis, which indicates the presence of cointegration among the variables of interest, to the null hypothesis, which states that there is no cointegration between the climatic and non-climatic factors on the one hand and crop production in particular countries on the other. Therefore, the null hypothesis of no cointegration relationship would be rejected if the probability value is signifcant at 1%, 5%, and 10%, respectively.

Panel autoregressive distributed lag (ARDL) technique

The mean group (MG) and the pooled mean group (PMG) approaches developed by Pesaran [\(2004](#page-10-24)) and Pesaran et al. [\(1999](#page-10-30)) have been implemented to discover the long-run and short-run estimates of the study. Explicitly, the PMG estimator imposes equality of the long-run parameters across the entire country groups even though it permits the slope coeffcients to vary between countries in the short-run. The PMG is a consistent and efficient estimator when there is a longrun homogeneity. However, the MG approach is more reliable when the slope and constants difer for individual country groups. Moreover, the PMG technique has the beneft of estimating the long-run and short-run parameters regardless of the series integrated at *I*(0), *I*(1), or both. Based on a comparison between the PMG and MG estimators, Hausman [\(1978\)](#page-10-31) is employed to ascertain the null hypothesis of homogeneity constraint on the long-run coefficients.

The study implemented a panel ARDL approach to capture the effects of climatic variability on agricultural output in East African nations. In accordance with Pesaran et al. ([1999](#page-10-30)), the panel ARDL framework of the study (*p*, *q*,*q*,*q*,*q*,*q*) is specifed as follows:

$$
\Delta lnCP_{ii} = \alpha_0 + \gamma_1 lnCP_{ii-1} + \gamma_2 lnARF_{ii-1} + \gamma_3 lnAT_{ii-1} \n+ \gamma_4 lnCO_{2ii-1} + \gamma_5 lnLUC_{ii-1} + \gamma_6 lnRP_{ii-1} \n+ \sum_{i=1}^{p} \delta_1 \Delta lnCP_{ii-k} + \sum_{i=1}^{q} \delta_2 \Delta lnARF_{ii-k} + \sum_{i=1}^{q} \delta_3 \Delta lnAT_{ii-k} \n+ \sum_{i=1}^{q} \delta_4 \Delta lnCO_{2ii-k} + \sum_{i=1}^{q} \delta_5 \Delta lnLUC_{ii-k} + \sum_{i=1}^{q} \delta_6 \Delta lnRP_{ii-k} + \mu_i + \varepsilon_t
$$
\n(7)

where α_0 is the intercept, γ is the coefficient of long-run variables, δ signifies the coefficient of short-run variables, *p* and *q* represent the number of lags, Δ is the first difference operator, ε_t is the error term, and μ_i captures country-specific effects.

Dumitrescu–Hurlin panel causality test

The Dumitrescu–Hurlin causality test developed by Dumitrescu and Hurlin ([2012](#page-10-19)) is a non-causality test for heterogeneous models of panel data with constant coeffcients. The study implements this test to determine the causal association between average precipitation, mean temperature, $CO₂$ emissions, cultivated land area, agricultural labor, and cereal crop yield. It is outstanding to note that Dumitrescu–Hurlin panel causality is applicable for heterogenous panels whether *N*>*T* or *N*<*T*. The causality test of Dumitrescu and Hurlin ([2012](#page-10-19)) is specifed as follows:

$$
y_{it} = \omega_i + \sum_{n=1}^{K} \varphi_i^{(n)} y_{i,t-n} + \sum_{n=1}^{K} \psi_i^{(n)} x_{i,t-n} + \varepsilon_{i,t}
$$
(8)

where ω_i indicates the intercept term, $\varphi_i^{(n)}$ represents the autoregressive parameter, $\psi_i^{(n)}$ signifies the slope coefficient that varies among cross-sections, *x* and *y* denote two stationary variables observed for *N* countries on *T* periods, and *K* denotes the lag length. Additionally, the null and alternative hypothesis for evaluating the Dumitrescu–Hurlin panel causality is expressed as follows:

$$
H_0: \delta_i = 0 \forall i = 1, ..., N H_1: \delta_i = 0 \forall i = 1, ..., N H_1: \delta_i \neq 0 \forall i = N + 1, N + 2, ..., N
$$
 (9)

The null hypothesis of the test claims that there is no homogenous causality among the whole cross-sections, while the alternative hypothesis confrms that there is evidence of at least one causal link in the panel data.

Empirical results and discussion

Descriptive statistics and correlation analysis

The descriptive statistics demonstrate the essential characteristics of the data from the selected East African countries. The mean, standard deviation, maximum, minimum, and the number of observations of the data are illustrated in Panel A of Table [2](#page-5-0). The average cereal crop production exhibits the highest mean value of 5.516, whereas $CO₂$ emissions had the lowest mean value of−0.942. The average values of land under cereal cultivation and rural population were 5.375 and 0.887, respectively. Moreover, the mean values of climate change factors such as rainfall, and temperature, were 2.802 and 1.378, respectively. The land under cereal output shows the highest standard deviation of 1.807 compared to other variables, while temperatures demonstrated the lowest dispersion value at 0.053. Additionally, cereal production had the highest maximum value of all variables, although land under cereal cultivation had the lowest minimum value. On the other hand, the pair-wise correlation fndings assess the degree to which two variables move together or avoid each other. As shown in Panel B of Table [2](#page-5-0), all variables were positively correlated with cereal production except the average temperature and carbon emissions. Average rainfall had a negative association with temperature and carbon emissions. However, the pair-wise correlation matrix in Table [2](#page-5-0) reveals that carbon emissions are positively related to average temperatures.

Table 2 Descriptive statistics and correlations analysis

Variable	lnCP	lnARF	lnAT	lnCO ₂	lnLCC	lnRP
	Panel A: Descriptive statistics summary					
Mean	5.516	2.802	1.378	-0.942	5.375	0.887
Std. dev	1.704	0.290	0.053	0.389	1.807	0.740
Min	0.954	2.112	1.280	-1.663	0.000	-0.848
Max	7.415	3.174	1.464	0.000	7.022	1.937
Obs	258	261	261	261	261	261
	Panel B: Pairwise correlations					
lnCRLP	1					
InARF	0.616	1				
lnAT	-0.477	-0.873	1			
lnCO ₂	-0.525	-0.576	0.650	1		
lnLCC	0.992	0.568	-0.441	-0.559	1	
lnRP	0.943	0.688	-0.499	-0.506	0.888	1

Table 3 Test of cross-sectional dependency and homogeneity outcomes

Test of cross‑sectional dependence and homogeneity

Examining the CD and slope heterogeneity of the variables impacting cereal production in East Africa is the initial step of the analysis. The CD test results detailed in Table [3](#page-5-1) demonstrate that the null hypothesis of cross-sectional independence is rejected at the 1% signifcance level. This suggests that all variables under investigation have cross-sectional dependence. Moreover, the homogeneity of the slope coefficients was tested using Pesaran and Yamagata [\(2008\)](#page-10-32). As shown in Table [3,](#page-5-1) the results confrm the rejection of the null hypothesis of homogeneity based on the statistical values of delta-tilde and delta-tilde adjusted regarding their probability values. The findings indicate that heterogeneity of the slope coefficients exists across diferent cross-sections. This proposes the appropriateness of heterogeneous panel estimators in our analysis.

Table 4 Panel unit root test fndings

	CIPS		CADF	
	Level	1st difference	Level	1st difference
lnCP	$-3.262***$	$-4.662***$	$-2.782***$	$-5.132***$
lnARF	$-5.054***$	$-6.135***$	$-3.175***$	$-7.483***$
lnAT	$-3.176***$	$-5.714***$	0.177	$-2.102**$
lnCO ₂	-1.482	$-4.662***$	1.130	$-3.065***$
lnCC	$-2.711***$	$-5.662***$	$-5.628***$	$-7.927***$
lnRP	-1.123	$-2.258*$	2.179	$-1.408*$

***, **, and * signify significance levels at 1%, 5%, and 10%, respectively

Stationarity analysis

After determining the cross-sectional dependence and homogeneity, the subsequent step is to examine the stationarity of the variables and their corresponding orders of integration. Table [4](#page-5-2) exhibits the fndings of the panel unit root analysis. The results of the CIPS and CADF unit root tests indicate that the series are stationarity under diferent integration arrangements. The estimated fndings reveal mixed stationarity of the variables, i.e. some are integrated at *I*(0) while others are stationary at *I*(1). Traditional cointegration approaches might lead to inaccurate inferences in the case of varying orders of integration. Hence, this indicates the appropriateness of the PMG technique for the current study.

Panel cointegration analysis

Subsequently, we tested the cointegration relationship between the series by conducting Pedroni and Kao cointegration tests. The examined output of the Pedroni test demonstrated in Table [5](#page-6-0) indicates that the null hypothesis of no cointegration is rejected for the test statistics of panel

Table 5 Panel cointegration tests

Phillips-Perron (PP) and ADF at the 1% signifcance level. To validate the cointegration results of the Pedroni test, the study implemented the Kao cointegration test which considers heterogeneity and cross-sectional dependence under the null hypothesis of no cointegration between the series. Thus, the outcome demonstrates that the null hypothesis of no cointegration is rejected at the 1% signifcance level for the modifed Dickey-Fuller (DF), DF, and ADF tests, as shown in Table [5.](#page-6-0) Consequently, the fndings of both tests confrm the long-run interconnections between the explained variable and the regressors.

Long‑run estimates

We approximated the long-run and short-run impacts of average rainfall, temperature, carbon emissions, cropped land area, and rural population on cereal output in East African economies. The fndings displayed in Table [6](#page-6-1) reveal that the chi² statistic of the Hausman test is 4.42 with a probability value of 0.3525. This recommends that the PMG estimator is robust and more consistent to be interpreted for this analysis.

Primarily, the estimated long-run fndings of the PMG model exhibited in Table [6](#page-6-1) reveal that average rainfall has a favourable and signifcant impact on cereal yield in East African economies. Interpretively, a percentage rise in average rainfall, keeping other factors constant, contributes a 0.451% increase in cereal output at the 1% signifcance level. The positive infuence of average rainfall on cereal production is consistent with the outcomes of Ozdemir ([2022\)](#page-10-33) and Warsame et al. ([2021](#page-11-7)). However, the result contradicts the fnding of Chandio et al. ([2020\)](#page-10-6), who found that rainfall hampers agriculture production in China. However, one striking outcome of the study shows that average temperature has negative and statistically signifcant infuence on cereal output in the long run. This implies that a 1% rise in average temperature causes cereal production to decline by 0.296%. Recent studies by Chandio et al. [\(2020\)](#page-10-6), Pickson **Table 6** Results from the PMG estimator

***, **, and * denote significance levels at 1%, 5%, and 10%, respectively

et al. ([2020](#page-10-5)), and Warsame et al. ([2021\)](#page-11-7) fndings coincided with the unfavourable consequences of rising temperature on agriculture production that was observed in the current study. These outcomes elaborate that the cereal crop production in East African countries is subject to climate change. The climbing global temperature had apparent adverse efects on these countries' agriculture production, which showed a downward trend for the past decade. Notably, our fndings counter Janjua et al. ([2014](#page-10-34)), who concluded that global climate variability does not infuence the agricultural output in Pakistan.

Moreover, carbon emissions were observed to expand cereal production in East African economies. This indicates that a 1% rise in carbon emissions signifcantly increases cereal production, on average, by 0.099%. The outcome that $CO₂$ emissions in East African countries trigger agricultural output is comparable with Ahsan et al. [\(2020](#page-9-6)). They observed that carbon emissions boost the production of cereal crops in Pakistan. Additionally, the empirical results of Ntiamoah et al. (2022) assert that $CO₂$ emissions improve cereal output in Ghana. In the long-run analysis, the study found that cropped area positively changes cereal crop production. This suggests that a percentage rise in the cultivated area improves crop production by 0.825%. Likewise, Ahsan et al. [\(2020](#page-9-6)) and Chandio et al. ([2020\)](#page-10-6) inspected that land under cereal crops favourably infuences the agricultural output of China in the long-run. In addition, the study found that the rural population positively and signifcantly boosts cereal production. This implies that a 1% rise in the agricultural labour force increases cereal crop yield by 0.26% in the long-run. This fnding is similar to Pickson et al. ([2020](#page-10-5)), who observed that increased rural population contributes crucially to cereal production.

Short‑run estimates

On the other hand, the short-run estimates presented in Table [6](#page-6-1) outline that average temperature has an adverse and signifcant infuence on cereal output. This indicates that a percentage increase in average temperature causes the cereal crops to decline by 1.830%. Pickson et al. ([2020\)](#page-10-5) discovered from China that average temperature rise negatively infuences agriculture production in the short-run. Moreover, cultivated land area favourably and signifcantly impacts cereal yield. Interpretively, a 1% rise in cereal cultivated area in the short-run increases agricultural production by 0.676%. Similarly, the rural population has a constructive infuence on cereal production. This implies that with a 1% rise in rural population, cereal production will increase by 2.403%. The outcome that land under cereal production and rural population enhance agricultural production in the short-run is in line with the recent outcomes of Ahsan et al. ([2020](#page-9-6)) and Pickson et al. ([2020\)](#page-10-5). The short-run estimates from the study indicate that average rainfall and carbon emissions do not have any statistical impact on cereal production in the short run. The speed of adjustment toward long-run equilibrium from any short-run shock in the regressors is denoted by the error correction term (ECT). The most remarkable evidence reveals that the ECT is negative and signifcant, which indicates that short-run deviations that occur in crop production will be adjusted by the interested explanatory variables by about 43.4% annually.

Robustness analysis

The study utilized various cointegration techniques as robustness tests to validate the estimated parameters of the PMG estimator. The long-run estimates of fully modifed ordinary least square (FMOLS) and dynamic ordinary least square (DOLS) are presented in Table [7.](#page-7-0) The results of the sensitivity analysis indicate that the sign and signifcance of the coefficients are in line with the PMG model estimates. Therefore, this supports that the PMG results of the study are reliable for policy making.

Dumitrescu–Hurlin panel causality test

In determining the causality direction of the scrutinized variables, the study has adopted the Dumitrescu-Hurlin panel causality technique. The empirical data presented in Table [8](#page-8-1) show that, at the 10% signifcance level, the null hypothesis

***, **, and * denote significance levels at 1%, 5%, and 10%, respectively. Values in parenthesis represent the t-statistic

that average temperature variations do not homogeneously cause cereal crop production is rejected. This suggests that patterns of cereal crop output are considerably afected by fuctuations in temperature levels. Similarly, there is evidence to refute the null hypothesis that cereal crop production does not homogeneously cause average temperature changes at the 10% threshold. This fnding demonstrates that variations in cereal crop production result in temperature changes. Likewise, Kumar et al. ([2021\)](#page-10-22) validated the presence bidirectional causality among average temperature and cereal crops in lower-middle-income countries. In contrast, the panel causality results of Pickson et al. ([2022](#page-10-12)) indicate a unidirectional causality from temperature to crop output across 30 Chinese provinces. Even though the evidence supports the null hypothesis that variations in rainfall patterns do not cause changes in cereal output, we reject the null hypothesis that changes in cereal crop yield do not homogeneously cause changes in precipitation at the 5% threshold. Therefore, these fndings imply that variations in crop production infuence changes in average rainfall. However, some studies such as Attiaoui and Boufateh [\(2019](#page-9-0)) and Chandio et al. ([2022a,](#page-10-8) [b](#page-10-9), [c\)](#page-10-11) found that changes in rainfall patterns lead to a crop production variations. According to IPCC [\(2001\)](#page-10-35), droughts and foods are caused by changes in rainfall patterns and an increase in temperature levels.

The fndings in Table [8](#page-8-1) demonstrate evidence to reject the null hypothesis that carbon dioxide emissions do not homogeneously cause a shift in cereal crop output at the 1% signifcance level. This suggests that carbon emissions across East African economies cause changes in the production of cereal crops. In addition, the analysis suggests rejecting the null hypothesis that changes in cereal crop production do not

***, **, and * denote significance levels at 1%, 5%, and 10%, respectively

homogeneously cause changes in $CO₂$ at the 1% threshold. Therefore, we can infer that changes in agricultural output granger cause the levels of emissions in East Africa. This is comparable to the fndings of Chandio et al. ([2022a,](#page-10-8) [b,](#page-10-9) [c\)](#page-10-11) who observed that there is a bidirectional causality between carbon emissions and cereal production. Moreover, the study results indicate no evidence to reject the null hypothesis that land under cereal cultivation does not homogeneously cause changes in cereal crop production at the 1% threshold. This implies that the variation in the cropped land area considerably results in changes in cereal crop production. Likewise, the study suggests rejecting the null hypothesis that changes in cereal crop production do not homogeneously cause changes in the cropped area at the 1% threshold. The analysis infers that a shift in the output of cereal crops results in changes in land under cereal cultivation. Similarly, Pickson et al. [\(2022\)](#page-10-12) reveal that there is a two-way causality between cultivated land area and crop yield. On the other hand, the study proposes rejecting the null hypothesis that shifts in the rural population do not homogeneously cause changes in the output of cereal crops at the 1% signifcance level. However, there is no evidence to reject the null hypothesis that cereal crop production does not homogeneously cause changes in rural population. This demonstrates that alterations in agricultural labour cause changes in cereal crop production in East Africa. Likewise, the evidence from Kumar et al. [\(2021\)](#page-10-22) using a panel of lowermiddle-income countries supports our fndings that shifts in rural population lead to changes in cereal crops output.

Conclusion

Climate change has become an issue of concern for sustainable agriculture production. The low-income countries in East Africa are highly dependent on the agriculture sector, which constitutes a large portion of their GDP and employment. The sector has recently experienced a decline due to climatic variabilities. Hence, this study aims to investigate the efects of climate change on cereal crop production in East African countries during the period 1990 to 2018. The study implemented a panel ARDL cointegration approach to investigate the long-run and short-run relationship between climate change and cereal production in East African countries. The study has discovered the dependence of cross-sections, and the null hypothesis of homogeneity of the slope coefficients was rejected. Moreover, the fndings of the CIPS and CADF unit root tests ascertained the order of integration of the variables to be a mixed order of stationarity, i.e. *I*(0) and *I*(1). Furthermore, the Pedroni and Kao cointegration tests have confrmed the long-run cointegration relationship between cereal crop production, average rainfall, temperature, carbon emissions, cultivated land area, and rural population.

The fndings of the PMG estimator demonstrate that precipitation and carbon emissions have positive and signifcant consequences on agricultural production in the long run, although their short-term impacts were negligible. Besides, land under cereal cultivation and rural population is favourably associated with agricultural output both in the long-run and short-run. However, average temperatures negatively affect agricultural production in the long-run and short-run, even though the magnitude of agricultural output sensitivity to temperature is higher in the short-run. The fndings of the PMG estimator had been validated by utilizing the panel FMOLS and panel DOLS; their long-run estimates showed similar signs and signifcance levels. Moreover, the results of the Dumitrescu-Hurlin panel causality test reveal a bidirectional causality between the output of cereal crops and temperature. The study further demonstrated a unidirectional causality running from cereal crop production to average rainfall. Additionally, a bidirectional causality was observed between carbon emissions and cereal crop output. Also, there is a bidirectional causal relationship between cropped areas and the production of cereal crops. However, there is a one-way causality running from rural population to cereal crop production.

The primary climate-related difficulties for East African countries are indeed the changing precipitation patterns and rising temperatures, which afected the livelihoods of agriculture-dependent households. Since the outcomes of the study emphasized that climatic events are critical to crop production, adaptive measures must be taken at the national and regional levels. Based on the study fndings, several policy initiatives are necessary to overcome the adverse efects of climatic variations on sustainable agricultural yield. Firstly, governments should increase the capacity to adopt climate-resilient agricultural practices that diminish the magnitude of adverse climatic events, including foods and droughts. In this regard, implementing climate-resilient farming practices can provide sustainable solutions to unpredictable climatic efects. Second, since irregular rainfall patterns cause decreased crop yield, the East African authorities should assert the implementation of water management policies. One of the most efective ways to avoid the vulnerability of cereal crop production to rainfall fuctuations is to develop water retention facilities to irrigate rainfall-reliant farms during times of rainfall failure. This strategy is benefcial during the seasons, where the expected volume of rain is minimal or completely dry seasons. Another practical approach is establishing alternative irrigation systems, such as drilling dams, to ensure crop yields can survive the strains of changing climatic conditions. Subsidizing these irrigation facilities can reduce the complete dependence on rainfall for farming activities. Third, policymakers should improve the quality of the farm inputs by ofering training and agricultural machinery to the farmers to boost their productivity. This can be achieved by enhancing the capacity of farmers to cope with climate change through training programs and developing their technical skills. Besides, the governments can also subsidize the acquisition of agricultural inputs such as tractors and fertilizers. Fourth, governments should establish crop diversifcation initiatives in cultivated land areas through agricultural investments to combat the negative consequences of climate change in East Africa. Expanding the crop production system requires research capabilities to create new cultivars and farming methods. Investments in efficient and resilient agricultural output through $R&D$, market-driven production systems, and sustainable agriculture technologies might lead to better varieties of cereal crops and create overall food security for the region. This will also diminish the efects of harsh weather on crops and could broaden the range of products that farmers may grow to sell. Fifth, there should be a consideration of farmers' income and crop yield losses due to the uncertainties related to climatic fuctuations. The increasing temperature and varying precipitation patterns are associated with more signifcant agricultural income losses. Therefore, policymakers should implement careful policy interventions, including pricing

policies that could combat the consequences of income decline for the farmers.

One major limitation of this study that future studies should consider is the efects of fertilizer consumption on cereal crop production in East African nations. The use of fertilizers is crucial for raising the sector's productivity even though its excessive use might lead to land degradation. Moreover, future studies should also consider the efects of climate change on disaggregated crops since climatic factors' impacts vary from one crop yield to another.

Author contribution Abdikaf Hassan Abdi was responsible for the study's conception, design, development, data collection, analysis, and interpretation and reviewed and edited the entire manuscript. Abdimalik Ali Warsame wrote the introduction and edited the paper. Ibrahim Abdukadir Sheik-Ali wrote some parts of the methodology.

Data availability The datasets used and/or analyzed during the current study are available from the corresponding author on reasonable request.

Declarations

Ethical approval Not applicable.

Consent to participate Not applicable.

Consent for publication Not applicable.

Competing interests The authors declare no competing interests.

References

- Abbas S, Mayo ZA (2020) Impact of temperature and rainfall on rice production in Punjab, Pakistan. Environ Dev Sustain 23(2):1706– 1728. <https://doi.org/10.1007/s10668-020-00647-8>
- Ahsan F, Chandio AA, Fang W (2020) Climate change impacts on cereal crops production in Pakistan: evidence from cointegration analysis. International Journal of Climate Change Strategies and Management 12(2):257–269
- Alboghdady M, El-Hendawy SE (2016) Economic impacts of climate change and variability on agricultural production in the Middle East and North Africa region. International Journal of Climate Change Strategies and Management 8(3):463–472. [https://doi.org/](https://doi.org/10.1108/IJCCSM-07-2015-0100) [10.1108/IJCCSM-07-2015-0100](https://doi.org/10.1108/IJCCSM-07-2015-0100)
- Appiah K, Du J, Poku J (2018) Causal relationship between agricultural production and carbon dioxide emissions in selected emerging economies. Environ Sci Pollut Res 25(25):24764–24777. [https://](https://doi.org/10.1007/s11356-018-2523-z) doi.org/10.1007/s11356-018-2523-z
- Attiaoui I, Boufateh T (2019) Impacts of climate change on cereal farming in tunisia: a panel ARDL–PMG approach. Environ Sci Pollut Res 26(13):13334–13345. [https://doi.org/10.1007/](https://doi.org/10.1007/s11356-019-04867-y) [s11356-019-04867-y](https://doi.org/10.1007/s11356-019-04867-y)
- Barrios S, Ouattara B, Strobl E (2008) The impact of climatic change on agricultural production: is it diferent for Africa? Food Policy 33(4):287–298.<https://doi.org/10.1016/j.foodpol.2008.01.003>
- Ben Zaied Y, Ben Cheikh N (2015) Long-run versus short-run analysis of climate change impacts on agricultural crops.

Environ Model Assess 20(3):259–271. [https://doi.org/10.1007/](https://doi.org/10.1007/s10666-014-9432-4) [s10666-014-9432-4](https://doi.org/10.1007/s10666-014-9432-4)

- Breitung J, Das S (2008) Testing for unit roots in panels with a factor structure. Economet Theor 24(1):88–108
- Chandio AA, Jiang Y, Rehman A, Rauf A (2020) Short and long-run impacts of climate change on agriculture: an empirical evidence from China. International Journal of Climate Change Strategies and Management 12(2):201–221. [https://doi.org/10.1108/](https://doi.org/10.1108/IJCCSM-05-2019-0026) [IJCCSM-05-2019-0026](https://doi.org/10.1108/IJCCSM-05-2019-0026)
- Chandio AA, Gokmenoglu KK, Ahmad M, Jiang Y (2021) Towards sustainable rice production in Asia: the role of climatic factors. Earth Systems and Environment. [https://doi.org/10.1007/](https://doi.org/10.1007/s41748-021-00210-z) [s41748-021-00210-z](https://doi.org/10.1007/s41748-021-00210-z)
- Chandio AA, Akram W, Bashir U, Ahmad F, Adeel S, Jiang Y (2022a) Sustainable maize production and climatic change in Nepal: robust role of climatic and non-climatic factors in the long-run and short-run. Environ Dev Sustain.[https://doi.org/10.1007/](https://doi.org/10.1007/s10668-022-02111-1) [s10668-022-02111-1](https://doi.org/10.1007/s10668-022-02111-1)
- Chandio AA, Jiang Y, Amin A, Akram W, Ozturk I, Sinha A, Ahmad F (2022b) Modeling the impact of climatic and non-climatic factors on cereal production: evidence from Indian agricultural sector. Environ Sci Pollut Res 29(10):14634–14653. [https://doi.org/10.](https://doi.org/10.1007/s11356-021-16751-9) [1007/s11356-021-16751-9](https://doi.org/10.1007/s11356-021-16751-9)
- Chandio AA, Jiang Y, Fatima T, Ahmad F, Ahmad M, Li J (2022c) Assessing the impacts of climate change on cereal production in Bangladesh: evidence from ARDL modeling approach. International Journal of Climate Change Strategies and Management 14(2):125–147.<https://doi.org/10.1108/IJCCSM-10-2020-0111>
- Deschenes O, Kolstad C (2011) Economic impacts of climate change on California agriculture. Clim Change 109(S1):365–386. [https://](https://doi.org/10.1007/s10584-011-0322-3) doi.org/10.1007/s10584-011-0322-3
- Dumitrescu E-I, Hurlin C (2012) Testing for Granger non-causality in heterogeneous panels. Econ Model 29(4):1450-1460. [https://doi.](https://doi.org/10.1016/j.econmod.2012.02.014) [org/10.1016/j.econmod.2012.02.014](https://doi.org/10.1016/j.econmod.2012.02.014)
- Gay C, Estrada F, Conde C, Eakin H, Villers L (2006) Potential impacts of climate change on agriculture: a case of study of cofee production in Veracruz, Mexico. Climatic Change 79(3–4):259–288. <https://doi.org/10.1007/s10584-006-9066-x>
- Hausman JA (1978) Specifcation tests in econometrics. Econometrica: Journal of the Econometric Society 1251–1271
- Hossain MS, Qian L, Arshad M, Shahid S, Fahad S, Akhter J (2019) Climate change and crop farming in Bangladesh: an analysis of economic impacts. International Journal of Climate Change Strategies and Management 11(3):424–440. [https://doi.org/10.1108/](https://doi.org/10.1108/IJCCSM-04-2018-0030) [IJCCSM-04-2018-0030](https://doi.org/10.1108/IJCCSM-04-2018-0030)
- ICPAC (2021) [https://www.icpac.net.](https://www.icpac.net) Retrieved from IGAD Climate Prediction and Applications Centre: [https://www.icpac.net/news/](https://www.icpac.net/news/10-climate-change-impacts-in-east-africayou-didnt-know-about/) [10-climate-change-impacts-in-east-africayou-didnt-know-about/](https://www.icpac.net/news/10-climate-change-impacts-in-east-africayou-didnt-know-about/)
- IPCC (2001) Climate Change 2001: impacts, adaptation, and vulnerability. intergovernmental panel on climate change. Cambridge University Press, Cambridge
- IPCC (2013) Climate Change 2021—the physical science basis.In Climate Change 2013: The physical science basis. Contribution of working group I to the ffth assessment report of the intergovernmental panel on climate change.[https://doi.org/10.](https://doi.org/10.1515/ci-2021-0407) [1515/ci-2021-0407](https://doi.org/10.1515/ci-2021-0407)
- Janjua PZ, Samad G, Khan N (2014) Climate change and wheat production in Pakistan: an autoregressive distributed lag approach. NJAS-Wageningen Journal of Life Sciences 68:13–19
- Kao C (1999) Spurious regression and residual-based tests for cointegration in panel data. Journal of Econometrics 90(1):1–44. [https://doi.org/10.1016/S0304-4076\(98\)00023-2](https://doi.org/10.1016/S0304-4076(98)00023-2)
- Kumar P, Sahu NC, Kumar S, Ansari MA (2021) Impact of climate change on cereal production: evidence from lower-middleincome countries. Environ Sci Pollut Res 28(37):51597–51611
- Li C, Chai Y, Yang L, Li H (2016) Spatio-temporal distribution of food disasters and analysis of infuencing factors in Africa. Nat Hazards 82(1):721–731.<https://doi.org/10.1007/s11069-016-2181-8>
- Mohamed J, Abdi D, Abdimalik A, Warsame A, Bakri M (2022) Two phases of long - term shift in extreme precipitation in Somalia. Meteorol Atmos Phys. [https://doi.org/10.1007/](https://doi.org/10.1007/s00703-022-00896-4) [s00703-022-00896-4](https://doi.org/10.1007/s00703-022-00896-4)
- Ngoma H, Lupiya P, Kabisa M, Hartley F (2021) Correction to: impacts of climate change on agriculture and household welfare in Zambia: an economy-wide analysis (Climatic Change, (2021), 167, 3–4, (55), [https://doi.org/10.1007/s10584-021-03168](https://doi.org/10.1007/s10584-021-03168-z) [z](https://doi.org/10.1007/s10584-021-03168-z)). Climatic Change 169(3–4), 1–20. [https://doi.org/10.1007/](https://doi.org/10.1007/s10584-021-03251-5) [s10584-021-03251-5](https://doi.org/10.1007/s10584-021-03251-5)
- Nicholson SE (2017) Climate and climatic variability of rainfall over eastern Africa. Rev Geophys 55(3):590–635. [https://doi.org/10.](https://doi.org/10.1002/2016RG000544) [1002/2016RG000544](https://doi.org/10.1002/2016RG000544)
- Ntiamoah EB, Li D, Appiah-Otoo I, Twumasi MA, Yeboah EN (2022) Towards a sustainable food production: modelling the impacts of climate change on maize and soybean production in Ghana. Environ Sci Pollut Res<https://doi.org/10.1007/s11356-022-20962-z>
- Ozdemir D (2022) The impact of climate change on agricultural productivity in Asian countries: a heterogeneous panel data approach. Environ Sci Pollut Res 29(6):8205–8217
- Pedroni P (1999) Critical values for cointegration tests in heterogeneous panels with multiple regressors. Oxford Bull Econ Stat 61(S1):653–670
- Pedroni P (2004) Panel cointegration: asymptotic and fnite sample properties of pooled time series tests with an application to the PPP hypothesis. Econ Theory 20(03):597–625. [https://doi.org/10.](https://doi.org/10.1017/S0266466604203073) [1017/S0266466604203073](https://doi.org/10.1017/S0266466604203073)
- Pesaran MH (2004) General diagnostic tests for cross section dependence in panels. SSRN Electron J. [https://doi.org/10.2139/ssrn.](https://doi.org/10.2139/ssrn.572504) [572504](https://doi.org/10.2139/ssrn.572504)
- Pesaran MH (2007) A simple panel unit root test in the presence of cross-section dependence. J Appl Economet 22(2):265–312
- Pesaran MH, Yamagata T (2008) Testing slope homogeneity in large panels. Journal of Econometrics 142(1):50–93
- Pesaran MH, Shin Y, Smith RP (1999) Pooled mean group estimation of dynamic heterogeneous panels. J Am Stat Assoc 94(446):621–634
- Pickson RB, He G, Ntiamoah EB, Li C (2020) Cereal production in the presence of climate change in China. Environ Sci Pollut Res 27(36):45802–45813. [https://doi.org/10.1007/](https://doi.org/10.1007/s11356-020-10430-x) [s11356-020-10430-x](https://doi.org/10.1007/s11356-020-10430-x)
- Pickson RB, He G, Boateng E (2022) Impacts of climate change on rice production: evidence from 30 Chinese provinces. Environ Dev Sustain 24(3):3907–3925.<https://doi.org/10.1007/s10668-021-01594-8>
- Porter JR, Semenov MA (2005) Crop responses to climatic variation. Philosophical Transactions of the Royal Society b: Biological Sciences 360(1463):2021–2035. [https://doi.org/10.1098/rstb.2005.](https://doi.org/10.1098/rstb.2005.1752) [1752](https://doi.org/10.1098/rstb.2005.1752)
- Sadiq S, Saboor A, Mohsin AQ, Khalid A, Tanveer F (2018) Ricardian analysis of climate change–agriculture linkages in Pakistan. Climate Dev 11(8):679–686. [https://doi.org/10.1080/17565529.](https://doi.org/10.1080/17565529.2018.1531746) [2018.1531746](https://doi.org/10.1080/17565529.2018.1531746)
- Sarker MdAR, Alam K, Gow J (2012) Exploring the relationship between climate change and rice yield in Bangladesh: an analysis of time series data. Agric Syst 112:11–16. [https://doi.org/10.](https://doi.org/10.1016/j.agsy.2012.06.004) [1016/j.agsy.2012.06.004](https://doi.org/10.1016/j.agsy.2012.06.004)
- Sarkodie SA, Strezov V (2019) Economic, social and governance adaptation readiness for mitigation of climate change vulnerability: evidence from 192 countries. Sci Total Environ 656:150–164. <https://doi.org/10.1016/j.scitotenv.2018.11.349>
- Satari Yuzbashkandi S, Khalilian S (2019) On projecting climate change impacts on soybean yield in Iran: an econometric

approach. Environ Processes 7(1):73–87. [https://doi.org/10.1007/](https://doi.org/10.1007/s40710-019-00400-y) [s40710-019-00400-y](https://doi.org/10.1007/s40710-019-00400-y)

- Simionescu M, Păuna CB, Niculescu M-DV (2021) The relationship between economic growth and pollution in some New European Union member states: a dynamic panel ARDL approach. Energies 14(9):2363.<https://doi.org/10.3390/en14092363>
- Sommer R, Glazirina M, Yuldashev T, Otarov A, Ibraeva M, Martynova L, Bekenov M, Kholov B, Ibragimov N, Kobilov R, Karaev S, Sultonov M, Khasanova F, Esanbekov M, Mavlyanov D, Isaev S, Abdurahimov S, Ikramov R, Shezdyukova L, de Pauw E (2013) Impact of climate change on wheat productivity in Central Asia. Agr Ecosyst Environ 178:78–99. [https://](https://doi.org/10.1016/j.agee.2013.06.011) doi.org/10.1016/j.agee.2013.06.011
- Sultan B, Roudier P, Quirion P, Alhassane A, Muller B, Dingkuhn M, Ciais P, Guimberteau M, Traore S, Baron C (2013) Assessing climate change impacts on sorghum and millet yields in the Sudanian and Sahelian savannas of West Africa. Environ Res Lett 8(1):1–10.<https://doi.org/10.1088/1748-9326/8/1/014040>
- Van Passel S, Massetti E, Mendelsohn R (2016) A Ricardian analysis of the impact of climate change on European agriculture. Environ Resource Econ 67(4):725–760. [https://doi.org/10.1007/](https://doi.org/10.1007/s10640-016-0001-y) [s10640-016-0001-y](https://doi.org/10.1007/s10640-016-0001-y)
- Warsame AA, Sheik-Ali IA, Mohamed J, Sarkodie SA (2022b) Renewables and institutional quality mitigate environmental degradation in Somalia. Renewable Energy 194:1184–1191. [https://doi.org/10.](https://doi.org/10.1016/j.renene.2022.05.109) [1016/j.renene.2022.05.109](https://doi.org/10.1016/j.renene.2022.05.109)
- Warsame, Ali A, Sheik-ali IA, Jama OM, Abdirahman A, Mohamed G (2022c) Assessing the efects of climate change and political instability on sorghum production: empirical evidence from Somalia. J Clean Prod 360(April):131893. [https://doi.org/10.](https://doi.org/10.1016/j.jclepro.2022.131893) [1016/j.jclepro.2022.131893](https://doi.org/10.1016/j.jclepro.2022.131893)
- Warsame AA, Sheik-Ali IA, Ali AO, Sarkodie SA (2021) Climate change and crop production nexus in Somalia: an

empirical evidence from ARDL technique. Environ Sci Pollut Res 1–13[.https://doi.org/10.1007/s11356-020-11739-3](https://doi.org/10.1007/s11356-020-11739-3)

- Warsame AA, Sheik-Ali IA, Hassan AA, Sarkodie S A (2022a) Extreme climatic efects hamper livestock production in Somalia. Environ Sci Pollut Res 1–13.<https://doi.org/10.1007/s11356-021-18114-w>
- WWF (2006) Climate change impacts on East Africa. In Climate Change Impacts on East Africa A Review of the Scientifc LiteratureClimate change
- Zhao C, Liu B, Piao S, Wang X, Lobell DB, Huang Y, Huang M, Yao Y, Bassu S, Ciais P, Durand JL, Elliott J, Ewert F, Janssens IA, Li T, Lin E, Liu Q, Martre P, Müller C, …, Asseng S (2017) Temperature increase reduces global yields of major crops in four independent estimates. Proc Natl Acad Sci USA 114(35):9326– 9331[.https://doi.org/10.1073/pnas.1701762114](https://doi.org/10.1073/pnas.1701762114)

Publisher's note Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional afliations.

We declare that this manuscript is original, has not been published before, and is not currently being considered for publication elsewhere. We confrm that the manuscript has been read and approved by all named authors and that there are no other persons who satisfed the criteria for authorship but are not listed. We further confrm that the order of authors listed in the manuscript has been approved by all of us.

Springer Nature or its licensor (e.g. a society or other partner) holds exclusive rights to this article under a publishing agreement with the author(s) or other rightsholder(s); author self-archiving of the accepted manuscript version of this article is solely governed by the terms of such publishing agreement and applicable law.