



Does financial development reinforce ecological footprint in Singapore? Evidence from ARDL and Bayesian analysis

Bui Hoang Ngoc¹ · Ashar Awan^{2,3}

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Abstract

Singapore has been ranked in the most dynamic financial market and the highest ecological deficit country, indicating that the trade-off hypothesis may exist. The main goal of the present study is to probe the impact of financial development, economic growth, and human capital on ecological footprint in Singapore from 1980 to 2016. The outcomes obtained from the Autoregressive Distributed Lag (ARDL) method have failed to provide a clear impact of financial sector development on ecological footprint. However, the Bayesian analysis reveals that both financial development and economic growth have a harmful influence on EF, while the impact of human capital is beneficial. A theoretical conclusion derived is that monetary expansion policies should be associated with improving human capital to achieve the United Nations SDGs in the context of Singapore. The findings of the study are of particular interest to policymakers for developing sound policy decisions for sustainable economic progress which is not at the cost of environment.

Keywords Financial development · Ecological footprint · ARDL · Bayesian analysis · Singapore

JEL Classification B26 · Q01 · Q28

Introduction

Climatic change is one of the biggest issues of the twenty-first century and a topic of overwhelming interest among the research community, policymakers, and professionals to achieve Sustainable Development Goals (Bayar and Maxim 2020; Murshed et al. 2021b). Therefore, environmental protection is the top priority of countries, particularly signatories of the Paris Agreement (Saud et al. 2020; Sohail et al. 2021; Usman et al. 2020a). The biggest threat to the environment is the emission of Green House Gases (GHG) which

causes an increase in the temperature of the earth and consequently disturbing ecological balance (Ahmed et al. 2019; Baloch et al. 2019). Although CO₂ emission is the major contributor of GHG and cause of climate change (Bilgili et al. 2021), anthropogenic actions particularly, fossil fuel-based energy consumption, water waste management, and production of fertilizers are causing damage to the ecosystem too. Earlier literature used the ecological footprints as a proxy of pressure on nature by human activities (Al-Mulali and Ozturk 2015). The ecological footprint is a comprehensive measure of pressure on the ecosystem due to various human activities (Ahmed et al. 2019; Sharif et al. 2020).

While on the other side, keeping an equilibrium between economic growth and environmental damage is a key issue that policymakers are facing. Therefore, it is emphasized to keep the environmental consequences of growth policies (Adedoyin et al. 2021; Destek et al. 2018; Ozturk et al. 2016). In addition, factors that are essential for economic activities may or may not damage the ecosystem. While human capital is regarded as an environment-friendly determinant of economic growth (Ahmed and Le 2021; Danish et al. 2019), financial sector development may (Saud et al. 2020) or may not (Shahbaz et al. 2013; Uddin et al. 2017) be

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✉ Ashar Awan
asharawan786@hotmail.com

Bui Hoang Ngoc
ngoc.bh@ou.edu.vn

¹ Ho Chi Minh City Open University, Ho Chi Minh City, Vietnam

² Graduate School, NisanTasi University, Istanbul, Turkey

³ University of Azad Jammu and Kashmir, AJK, Muzaffarabad, Pakistan

harmful; there is a negative relationship between financial sector development and pollution (Hashmi and Alam 2019; Meirun et al. 2021).

An efficient financial system contributes to economic growth, and it enables people to buy houses, home appliances, and automobiles; however, all this puts pressure on nature by increasing energy demand (Baloch et al. 2019). Furthermore, a financial development boosts investment in new plants and factories and consequently causes more water waste and pollution (Danish et al. 2018). However, financial development is also credited to reducing pollution by boosting investment in the research and development of green technologies and energy-efficient machines (Shahbaz et al. 2016). Financial development serves as a strong policy tool for the government to control pressure on the environment. Government can use their influence on financial institutions on their credit provision to less pollute production activities. The financial sector can contribute to environmental protection by discouraging loans to those investment initiatives that produce massive pollution. In addition, investment in a green environmental project, loans to socially responsible firms, and credit to eco-friendly projects curb environmental degradation (Saud et al. 2020).

Indisputably, financial sector development — the crucial factor for economic development — is associated with ecological quality through technique, scale, and composition effects (Saud et al. 2020). However, findings of previous literature about the influence of financial sector development on the natural environment are mixed. On the one hand, literature reported the negative effect of financial development on EF; for instance, in a panel of 27 countries (Uddin et al. 2017), for China (Destek and Sarkodie 2019), for Malaysia (Furuoka 2015), and for Nigeria (Omoke et al. 2020). On the other hand, other studies reported a positive influence of financial development on EF, for instance (Usman et al. 2020a) for a panel of 20 highest economies, (Mrabet and Alsamara 2017) for Qatar, and (Godil et al. 2020) for Turkey.

Interestingly, Singapore is a country with rapid economic growth, industrialization, and structural change experience (Katircioğlu 2014; Tan et al. 2014). Though the country is small in size 721 km², it faces serious challenges of environmental pollution due to its dense population (Han 2017). While the country has achieved enormous economic growth targets, the government of Singapore is much concerned about negative externalities such as risk to the environment (Ridzuan et al. 2017). Though the literature has been indicating the factors that put pressure on the environment in Singapore, and the government has been introducing regulations, the voices are raised to curb environmental pollution in a way that does not compromise on the economic growth of Singapore (Meirun et al. 2021).

The main contribution of the present study is that earlier literature on Singapore examining the influence of financial

sector development on EF has mixed findings. Some studies have found that financial development has increased EF (Destek and Sarkodie 2019), while the others show an opposite result (Naqvi et al. 2020; Saud et al. 2020). These contradictable findings have made a big question about the impact of financial development on EF in Singapore. Therefore, the purpose of this study is to further investigate the impact of financial development, economic growth, and human capital on EF in Singapore. However, our research differs from some previous studies in many ways. Firstly, previous studies have attempted to reveal the relationship between financial development and EF (Destek and Sarkodie 2019; Naqvi et al. 2020; Saud et al. 2020), but no studies included economic growth and human capital as control variables. Secondly, the impact of financial development on EF in Singapore has seemingly been ambiguous or even contradictable (Destek and Sarkodie 2019) found positive and (Khan et al. 2019b) insignificant and (Naqvi et al. 2020; Saud et al. 2020) negative. This ambiguity might be the result of the adoption of a frequentist inference, where parameters are unknown but fixed. Therefore, the present study applied two statistical inference types: frequentist inference and Bayesian inference, to provide probability interpretations of uncertainty and various effects of financial development, economic growth, and human capital on EF.

The remainder of the study is presented as follows. The next section shows discussions related to relevant literature about the nexus between economic growth, financial development, pollution, and human capital. This section is followed by the “Data and methodology” section; after which, the results and findings are discussed in the “Empirical results and discussion” section. Lastly, the conclusion and policy recommendations based on empirical findings are presented in the “Conclusion and recommendations” section of the present study.

Literature review

The extant literature on the environment merely used the novel ecological footprint proxy for environmental degradation (Ahmad et al. 2021; Murshed et al. 2021b; Tillaguango et al. 2021; Xue et al. 2021; Zeraibi et al. 2021). Most of the studies on environmental degradation use carbon emission as a measure of environmental decay (Bandyopadhyay and Rej 2021; Rej and Nag 2021). In this section, an overview of studies related to the empirical model is presented.

Financial development and EF

The previous studies showed that there were two effects (technological effect and structural effect) of financial development on the environment, which means financial

development may be beneficial or harmful to EF (Du et al. 2012; Saud et al. 2020). On the positive side, financial development boosts a country's financial structure, brings about financial channels, and attracts FDI, which in turn brings green-environment technology and fosters R&D activities. As a result, global commercials, renewable energy, and technology advancements take place (Ahmed and Le 2021; Hsueh et al. 2013). Thus, financial development can improve environmental quality and decrease EF. On the other side, financial development can cause some scale effects on the economic growth process. This point of view suggests that financial development could increase pollution which is caused by a high level of production of the economic-liberalization and higher energy consumption (Pazienza 2015; Saud et al. 2020). According to Ha et al. (2020), it is impossible to do economic and household activities without having a harmful influence on the natural ecosystem or environment.

Surprisingly, some empirical evidence supports both of these views based on the various development policies in each country and region. For instance, Uddin et al. (2017) has applied the FMOLS and DOLS methods on the panel data of 27 leading world EF contributors from 1991 to 2012 and found that financial development has improved environmental quality by decreasing EF. Similarly, Ahmed et al. (2019) has researched the connection between financial development and EF in Malaysia from 1971 to 2014. By adopting the Bayer-Hanck cointegration test and ARDL method, the outcomes have verified that financial development mitigates EF. Omoke et al. (2020) has discovered the negative relationship between financial development and EF in Nigeria from 1971 to 2014. However, some studies have revealed a contrary result regarding the relationship between financial development and EF. Khan et al. (2019a) used five Belt and Road initiative (BRI) regions as a research context for the association between financial development and EF. They used the augmented mean group (AMG) and the common correlated effect mean group (CCEMG) approaches, and then they found that EF has been fostered by financial development.

In the same line, Usman et al. (2020a, 2020b) examined the 20 highest economies from 1995 to 2017. The results showed that financial development deteriorates environmental quality by increasing EF. Godil et al. (2020) has also revealed a similar finding while testing the financial development — EF nexus in Turkey between 1986 and 2018. A comprehensive literature review is presented in Table 1.

Economic growth, human capital, and EF

Literature on the EKC hypothesis mainly discusses the relationship between environment and economic growth (Bilgili et al. 2021; Kocoglu et al. 2021). Moreover, there is

evidence that financial development has a mutual relationship with economic growth and human capital (Hsueh et al. 2013). Economic growth (Ahmed et al. 2020a; Al-Mulali and Ozturk 2015; Alola et al. 2019; Murshed et al. 2021c; Usman et al. 2020b) and human capital (Ahmed et al. 2020a; Ahmed et al. 2020b; Pata and Caglar 2021)) also have influences on EF. Accordingly, the investigation of the interaction between financial development and EF cannot provide a clear understanding without integrating human capital and economic growth.

Regarding the influence of economic growth on EF, most previous studies have demonstrated a trade-off between economic growth and EF since the rapid economic development has generated an unprecedented rise in energy demand, especially non-renewable energy (S. Nathaniel and Khan 2020; Udemba 2020; Zafar et al. 2019). Furthermore, economic growth could facilitate urban migration and urbanization (Ahmad et al. 2019; Nathaniel et al. 2020; Ozturk et al. 2016; Wu et al. 2019), which means that it could certainly bring more pressures to urban infrastructure and ecological assets (Wu et al. 2019). However, there are also some studies suggesting that economic growth would improve EF in Africa and Europe (Usman et al. 2020b) or Pakistan (Hassan et al. 2019).

Earlier literature also verifies that human capital mitigates environmental degradation, including EF (Nathaniel et al. 2020; Nathaniel and Khan 2020; Pata and Caglar 2021). Some scholars argue that human capital plays a significant role in fostering the adaption of technology change, so it could probably make sustainable growth (Ackah and Kizys 2015; Consoli et al. 2016). Moreover, human capital generates concerns about environmental problems (Adil 2018; Asongu 2018; Reynolds et al. 2010; Ulucak and Li 2020). Nevertheless, some studies, such as Croes et al. (2021) and Ahmed et al. (2021), postulated that the beneficial outcomes of economic growth are insufficiently invested in human capital. As the result, human capital is not giving a significant effect on sustainable development (Dietz et al. 2007). In a recent study, Kassouri and Altıntaş (2020) indicated that human capital increases EF in MENA countries.

Besides, the impact of financial development on EF in Singapore is likely ambiguous or even contradictable (positive; Destek and Sarkodie (2019); insignificant; Khan et al. (2019a, 2019b); negative; Naqvi et al. (2020); Saud et al. (2020). In addition, Singapore has been ranked in the most dynamic financial markets in the world, but Singapore has also been listed in the highest ecological deficit countries. It implies that the trade-off hypothesis between financial development and ecological assets may be valid. Therefore, further investigation is necessary to provide probabilistic interpretations of model uncertainty and various influences of financial development, economic growth, and human capital on EF in Singapore.

Table 1 Summary of previous studies

No.	Authors	Duration	Countries/areas	Econometric approach	Findings
Financial development (FD) — EF nexus					
1	(Uddin et al. 2017)	1991–2012	27 countries	FMOLS, DOLS	FD → EF (-);
2	(Ahmed et al. 2019)	1971–2014	Malaysia	ARDL	
3	(Destek and Sarkodie 2019)	1977–2013	11 countries	Augmented mean group (AMG)	
4	(Omoke et al. 2020)	1971–2014	Nigeria	NARDL	
5	(Pata and Yilanci 2020)	1980–2015	G7 countries	Panel hreshold regression	
6	(Saud et al. 2020)	1990–2014	45 countries	Pooled mean group (PMG)	
7	(Charfeddine, 2017)	1970–2015	Quatar	Markov switching models	FD → EF (+)
8	(Mrabet and Alsamara 2017)	1980–2011	Quatar	ARDL	
9	(Usman et al. 2020a)	1995–2017	20 countries	DCCEMG, AMG	
10	(Naqvi et al. 2020)	1990–2017	152 countries	AMG	
11	(Godil et al. 2020)	1986–2018	Turkey	Quantile ARDL	
12	(Shujah Ur et al., 2019)	1991–2014	CCE countries	SUR regression	
13	(Baloch et al. 2019)	1990–2016	59 countries	Driscoll-Kraay regression	
14	(Khan et al. 2019a)	1990–2016	5 countries	AMG, CCEMG	
Economic growth (EG) — EF nexus					
1	Alola et al. (2019)	1997–2014	16 EU countries	Panel ARDL	EG → EF (+)
2	Danish et al. (2019)	1971–2014	Pakistan	ARDL	
3	Zafar et al. (2019)	1970–2015	United States	ARDL	
4	Nathaniel et al. (2020)	1990–2016	MENA countries	AMG	
5	Destek and Sinha (2020)	1980–2014	24 OECD countries	FMOLS, DOLS	
6	Ahmed et al. (2020a)	1970–2016	China	Bootstrap ARDL	
7	Nathaniel and Khan (2020)	1990–2016	ASEAN countries	AMG	
8	Ahmed et al. (2020b)	1971–2014	G7	CUP-FM and CUP-BC	
9	Usman et al. (2020b)	1994–2017	33 countries	FMOLS, DOLS, AMG	EG → EF (-) (Africa and Europe)
10	Baz et al. (2020)	1971–2014	Pakistan	NARDL	EG → EF (ns)
11	Hassan et al. (2019)	1970–2014	Pakistan	ARDL	EG → EF (+) Then EG → EF (-)

(+) positive impact; (-) negative impact; *ns* no-significant

Data and methodology

Research model and data sources

The role of financial development in economic growth is enormous, but it is undeniable to say that it could be harmful to natural ecosystems. Mutually, a rise in financial development brings about an increase in economic growth, which is directly proportional to EF. Besides, the study incorporated the human capital per person index (labeled, HC) as the control variable. According to Neumayer (2012), it is recognized that the concern on environmental quality in a high human capital country is better than in a low human capital country. Therefore, to assess the impact of financial development, economic growth, and human capital on EF in the case of Singapore, the study has followed the previous works of Baloch et al. (2019); Godil et al. (2020); Pata

and Yilanci (2020); Saud et al. (2020); Chen et al. (2019); Usman et al. (2020a, b); and Zhao et al. (2019) to propose an initial model, detailed as:

$$EF_t = \beta_0 + \beta_1 \ln FD_t + \beta_2 \ln GDP_t + \beta_3 HC_t + u_t \quad (1)$$

where β_1 , β_2 , and β_3 are the long-run coefficients, while t is the time (from 1980 to 2016) and u is the error term. The EF variable is the ecological footprint index (units: gha per capita) collected from the Global Footprint Network. The FD variable is the financial development index (units: point) obtained from the International Monetary Fund (IMF).

In the study, we have used the financial development index as a proxy of financial development because it is a financial inclusion index, which is calculated based on the depth, access, and efficiency of financial institutions and financial markets of a country. The GDP variable is the income per capita (at a fixed price of 2010, units: US

dollar) abstracted by the World Bank, while the HC variable is the human capital per person index (units: point), quoted by the Federal Reserve Bank of St. Louis. A scale of zero to ten is applied for the human capital per person index, where zero is the lowest educated economy and 10 is the highest educated economy. In this work, two variables (FD and GDP) are used by following the logarithm to clarify smooth data, while the EF and HC variables have original data in use. The descriptive statistics of all variables are demonstrated in Table 2.

According to Table 2, the mean of Singapore’s ecological footprint was 5.843 gha per capita, which is more than 2.8 gha per capita, the global average. Singapore has been listed in a group of countries where the ecological deficit is severe. Recently, Singapore has implemented many positive steps to reduce the ecological deficit situation and has made significant advancements in renewable energy technology. However, based on the International Energy Agency (IEA, 2018) data, Singapore has been ranked as 27th out of 142 countries in terms of emissions per capita. Likewise, the mean of the lnGDP variable was 10.287, while the maximum value of the HC variable was 3.809. These data showed that Singapore had been a developed education system and high-income country.

Methodology

The estimated coefficients of Eq. (1) have only provided the long-run effects of financial development, economic growth, and human capital on ecological footprint. To analyze the short-run impacts, the study has applied the autoregressive distributed lag (ARDL) model, introduced by. The ARDL model has some advantages, such as firstly, the estimated coefficients are unbiased and reliable in the case of small sample size; secondly, it could be applied in all three cases whether the variables are stationary at I(0), I(1), or a mixture of both; thirdly, it provides both the short and long-run estimated coefficients; and fourthly, it could be used in two cases, where the cointegration among all variables exists or not (Nkoro and Uko 2016).

Therefore, (Eq.1) is written by the ARDL(p,q) model, as follows:

$$\begin{aligned} \Delta EF_t = & \alpha_0 + \beta_1 EF_{t-1} + \beta_2 \ln FD_{t-1} + \beta_3 \ln GDP_{t-1} \\ & + \beta_4 HC_{t-1} + \sum_{k=1}^{p-1} \alpha_{1k} \Delta EF_{t-k} + \sum_{k=0}^q \alpha_{2k} \Delta \ln FD_{t-k} \\ & + \sum_{k=0}^q \alpha_{3k} \Delta \ln GDP_{t-k} + \sum_{k=0}^q \alpha_{4k} \Delta HC_{t-k} + \varepsilon_t \end{aligned} \tag{2}$$

where:

Δ is the first difference.

$\beta_1, \beta_2, \beta_3,$ and β_4 are coefficients of the long-run impacts.

$\alpha_1, \alpha_2, \alpha_3,$ and α_4 are coefficients of the short-run impacts.

ε_t is the error.

To achieve the research objectives, our approach is summarized in a four-step process. *First*, the ordinary least square (OLS) method is applied to Eq.(2) to estimate the long-run coefficients. *Second*, the cointegration test is used to verify the integration among variables, in which the null hypothesis is stated: ($H_0: \beta_1 = \beta_2 = \beta_3 = \beta_4 = 0$), while the alternative hypothesis is written ($H_1: \beta_1 \neq \beta_2 \neq \beta_3 \neq \beta_4 \neq 0$). If the *F*-statistic value lies below the *F*-critical values, the null hypothesis 4) is accepted accordingly. That means there is no cointegration among variables in the long run. On the other hand, if the *F*-statistic value exceeds the *F*-critical value (the upper Bound value, I(1)), the null hypothesis is denied, and Eq. (2) would be re-parameterized as an error correction model (ECM) as:

$$\begin{aligned} \Delta EF_t = & \alpha_0 + \lambda \cdot ECM_{t-1} + \sum_{k=1}^{p-1} \alpha_{1k} \cdot \Delta EF_{t-k} + \sum_{k=0}^q \alpha_{2k} \cdot \Delta \ln FD_{t-k} \\ & + \sum_{k=0}^q \alpha_{3k} \cdot \Delta \ln GDP_{t-k} + \sum_{k=0}^q \alpha_{4k} \cdot \Delta HC_{t-k} + \varepsilon_t \end{aligned} \tag{3}$$

where *p* and *q* are the lag order of each variable collected from the stationary test result. In Eq. (3), if the estimated coefficient of λ is negative, and belongs [0;1] and significant, it means that the EF variable can itself re-adjust to long-run equilibrium point after short-run shocks caused by financial development, economic growth, or human capital.

Third, the major problem with the frequentist inference (e.g., ARDL model) is that estimated coefficients are unknown but fixed. More importantly, it is impossible to assess the link between two variables if the estimated coefficient is insignificant. To solve this case, the Bayesian inference approach through the integrated Markov chain Monte Carlo sample is applied to provide probabilistic interpretations of model uncertainty and to differ effects of financial development, economic growth, and human capital on EF. And in the last process, the modified Wald test introduced by Toda and Yamamoto (1995) is used to reveal the causal relationship between a pair of variables in our proposed model. The Toda and Yamamoto procedure is based on the vector autoregressive model (VAR). The causal relationship between the EF variable and lnFD variable is illustrated as follows:

Table 2 Descriptive statistic

Variables	Obs	Mean	Std.Error	Minimum	Maximum
EF	37	5.843	1.465	3.441	8.310
lnFD	37	15.262	0.996	11.167	15.894
lnGDP	37	10.287	0.422	9.497	10.881
HC	37	2.528	0.572	1.651	3.809

$$EF_t = \alpha_0 + \sum_{i=1}^h \alpha_{1i} \cdot EF_{t-i} + \sum_{j=h+1}^p \alpha_{2i} \cdot EF_{t-j} + \sum_{i=1}^h \delta_{1i} \cdot \ln FD_{t-i} + \sum_{j=h+1}^p \delta_{2i} \cdot \ln FD_{t-j} + \mu_{1t} \quad (4)$$

$$\ln FD_t = \beta_0 + \sum_{i=1}^h \beta_{1i} \cdot \ln FD_{t-i} + \sum_{j=h+1}^p \beta_{2i} \cdot \ln FD_{t-j} + \sum_{i=1}^h \alpha_{1i} \cdot EF_{t-i} + \sum_{j=h+1}^p \alpha_{2i} \cdot EF_{t-j} + \mu_{2t} \quad (5)$$

where h is the optimal lag order of the VAR model, and $p = (h + d_{\max})$, with d_{\max} is the maximum lag order. According to Toda and Yamamoto (1995), the maximum of d_{\max} is 2, because if $d_{\max} > 2$, the tests based on F -statistic are not reliable (Davoud et al. 2013; Nkoro and Uko 2016)

In Eq. (4), if an estimated coefficient of δ_j is significant, there will be a uni-directional causality running from financial development to ecological footprint. The advantage of the Toda and Yamamoto (1995) procedure is beneficial in minimizing the risk of wrong determination of each variable's lag order and being able to be applied to the variable that is stationary at $I(0)$ or $I(1)$, or a combination of both cointegration and no-cointegration (Sankaran et al. 2019; Toda and Yamamoto 1995), or either of them.

Empirical result and discussion

Empirical results

Unit-root test Nelson and Plosser (1982) suggested that it is necessary to check the stationary of the time-series variable because most of the economic variable is non-stationary. Therefore, to avoid the empirical results being spurious, the three tests are employed, including the ADF test (Dickey and Fuller 1981), the PP test (Phillips and Perron 1988), and the GLS-ADF test proposed by Elliott et al. (1992) to confirm the stationary of each variable. Compared to the ADF test, the PP test is advantageous in accounting for the potential serial correlation and heteroskedasticity in the residuals. Similarly, the GLS-ADF test has an advantage in allowing the series to be stationary around a linear time trend; or it is to allow the series to be stationary around a possible nonzero mean with no time trend. Results of the stationary tests are presented in Table 3.

The stationary test presented in Table 3 provides evidence to reject the null hypothesis of a random walk with drift. More particularly, the $\ln FD$ variable remains stationary at $I(0)$, while three variables (EF, $\ln GDP$, and HC) remain unchanging at $I(1)$ obtained from three tests. No variable is stationary at $I(2)$. In a brief conclusion, the condition in

applying the ARDL model is satisfied accordingly (Nkoro and Uko 2016; Pesaran et al. 2001). Hence, the cointegration test could proceed for further analysis.

Cointegration test The stationary test provided an inconsistent result of lag order between variables (mixture of $I(0)$ and $I(1)$). In the next step, it is necessary to check the long-run association of all variables. Hence, a new cointegration test, Bound-testing, has been employed to verify the cointegration in financial development, economic growth, human capital, and ecological footprint in the case of Singapore. The Bound-testing technique was proposed by Pesaran et al. (2001) with the null hypothesis quoted that H_0 : No cointegration, against the alternative hypothesis H_1 : there is cointegration between examined variables. The result of the Bound-testing is given in Table 4. Accordingly, the F -statistic value (= 4.602) exceeds the F -critical value (= 4.35) at a significant level of 5%. Likewise, the t -critical value (= -3.46) is higher than the t -statistic value (= -3.671) at a significant level of 10%. These results give evidence to reject the null hypothesis, which means that a long-run association between variables in our proposed model exists. Thus, Eq. (2) must be estimated by using the ECM model.

The short and long-run impacts by the error correction model Another advantage of the ARDL model is that it could auto-select the optimal lag of each variable. Based on the Akaike information criterion (AIC), Schwarz Bayesian information criterion (SBIC), and Hannan-Quinn information criterion (HQIC), the empirical result in Table 5 showed that the optimal lag of EF and $\ln GDP$ variable is one and $\ln FD$ variable is zero. Simultaneously, the obtained result also indicates that the volatility of current-period human capital is associated with two-period previous human capital.

After confirmation of cointegration, the error correction model is applied to study the impact of financial development, economic growth, and human capital on EF both in the short and long run. The optimal lag of each variable is set to select the final ARDL specification. Applying to the data of Singapore, the best ARDL model is the ARDL (1,1,0,0). The coefficients of the short and long-run impacts are shown in Table 6.

The empirical result showed in Table 6 indicates that the coefficient of $\text{CoinEq}(-1)$ is negative and significant level at 1% ($\lambda = -0.5647$, p value = 0.001). This finding supports that ecological footprint can re-adjust itself to the long-run equilibrium point after the short-run is suddenly affected by financial development, or economic growth, or human capital. Two years is the time needed for an adjustment ($=1/|\lambda|$). Table 6 also indicates that economic growth has a positive and significant impact on EF in the short and long run. More specifically, a 1% rise in economic growth leads to a 4.39 gha per capita increase in the short-run ecological footprint

Table 3 Results of the unit root test

Variables	Test statistic	1% Critical value	5% Critical value	10% Critical value
ADF test				
EF	-0.605	-4.297	-3.564	-3.218
lnFD	-2.019			
lnGDP	-1.328			
HC	-1.628			
Δ.EF	-4.335***	-4.306	-3.568	-3.221
Δ.lnFD	-4.427***			
Δ.lnGDP	-3.187*			
Δ.HC	-1.542			
PP test				
EF	-1.502	-4.279	-3.556	-3.214
lnFD	-7.197***			
lnGDP	-1.722			
HC	0.267			
Δ.EF	-6.719***	-4.288	-3.560	-3.216
Δ.lnFD	-14.431***			
Δ.lnGDP	-5.338***			
Δ.HC	-1.562			
GLS-ADF test				
EF	-1.652	-3.770	-3.283	-2.968
lnFD	-3.946***			
lnGDP	-1.645			
HC	-2.460			
Δ.EF	-5.776***	-3.770	-3.293	-2.976
Δ.lnFD	-5.262***			
Δ.lnGDP	-4.887***			
Δ.HC	-2.993*			

All tests are chosen based on the Akaike information criterion with intercept and trend, without structural breaks. ***, ** and *, respectively, denote significance levels of 1%; 5% and 10%.

Table 4 Result of cointegration test

Test statistic	Value	Significantly level	I(0)	I(1)
<i>F</i> -statistic	4.602	10%	2.72	3.77
<i>k</i>	3	5%	3.23	4.35
		1%	4.29	5.61
<i>t</i> -statistic	-3.671	10%	-2.57	-3.46
<i>k</i>	3	5%	-2.86	-3.78
		1%	-3.43	-4.37

and a 7.78 gha per capita increase in the long-run ecological footprint. Hence, our analysis reveals that the long-run impact of economic growth on EF is greater than the short-run impact. Similarly, the estimated coefficient of the HC variable is -2.10 in the short run and is -3.72 in the long run, respectively. These results imply that EF is affected by human capital in the case of Singapore. However, contrary to the influence of economic growth, an increase in the human

capital per person induces a fall in ecological footprint. We believe that these fascinating findings could provide Singapore policymakers with more insights into sustainable development strategies. More details about these findings will be discussed in the next section.

Nevertheless, the main aim of our research is to study the influence of financial development on EF in Singapore, which means that whether there is a trade-off between economic growth and environmental destruction. Interestingly, the empirical outcome reveals that financial development has a positive impact on EF, but not yet significant. In the view of frequentist inference, the obtained outcome from the ECM approach has failed to demonstrate the influence of financial development on EF in the case of Singapore. Besides, the major diagnostic tests were used to confirm the above conclusions, including the heteroskedasticity test, the autocorrelation test, the distribution of residuals, and the functional form test. The results of diagnostic tests are given at the bottom of Table 6. Accordingly, the four tests have a *p* value that is higher than 0.05, and it is evident to reject the

Table 5 The optimal lag of each variable

Lag	AIC	HQIC	SBIC
EF variable			
0	3.491	3.507	3.537
1	2.083*	2.113*	2.174*
2	2.139	2.185	2.275
3	2.142	2.203	2.323
4	2.198	2.274	2.425
lnFD variable			
0	2.975*	2.991*	3.021*
1	1.173	2.997	3.027
2	1.245	3.057	3.103
3	1.252	3.062	3.123
4	1.307	3.103	3.179
lnGDP variable			
0	0.851	0.867	0.897
1	0.002*	-3.618*	-3.587*
2	0.002	-3.562	-3.516
3	0.002	-3.587	-3.526
4	0.002	-3.540	-3.464
HC variable			
0	1.592	1.607	1.637
1	0.001	-3.668	-3.638
2	0.001*	-4.672*	-4.627*
3	0.001	-4.625	-4.564
4	0.001	-4.582	-4.506

* denotes the optimal lag

Table 6 The short- and long-run estimated coefficients

Variables	Coefficient	Standard error	T-ratio [Prob]
CoinEq(-1)	-0.5647	0.1538	-3.67 [0.001]
lnFD(-1)	0.2631	0.2901	0.91 [0.372]
lnGDP(-1)	7.7833	1.6444	4.73 [0.000]
HC(-1)	-3.7203	1.2682	-2.93 [0.006]
ΔlnFD	0.0044	0.1096	0.04 [0.968]
ΔlnGDP	4.3956	1.1761	3.74 [0.001]
ΔHC	-2.1011	0.6705	-3.13 [0.004]
Intercept	-38.851	10.791	-3.60 [0.001]
R ²	0.42	R ² -adj	0.41
χ ² _{SC}	3.501 [0.174]	χ ² _{FF}	0.746 [0.395]
χ ² _{NORM}	14.56 [0.439]	χ ² _{HET}	3.695 [0.594]
CUSUM test	Stability	CUSUMSQ test	Stability

χ²_{SC}, χ²_{FF}, χ²_{NORM}, and χ²_{HET} denote the LM tests for serial correlation, functional form, normality, and heteroskedasticity. The value in brackets is the corresponding *p* value, respectively.

null hypothesis. Besides, the work has conducted the cumulative sum of recursive residuals (CUSUM) and the cumulative sum of squares of recursive residuals (CUSUMSQ)

test to check the stability of the long-run coefficients and the short-run dynamics. Both CUSUM and CUSUMSQ lie within the Bound-critical value at the level of 5% significance (see Fig. 1a, b). When all diagnostic tests are satisfied, it is possible to conclude that our proposed model is stable and the obtained coefficients by the ECM approach are reliable.

As the result, the failure in defining the effect of financial development on EF leads to difficulty makes some difficulties in suggesting efficient environmental protection policies. Therefore, the Bayesian analysis for the generalized linear model (GLM) was employed to re-examine the influence of financial development on EF in the context of Singapore.

The empirical results by Bayesian inference Contrary to frequentist inference (i.e., where observed are assumed to be random and estimation parameters are unknown but have fixed quantities), the Bayesian inference assumes that the observed data is fixed, and estimation parameters are random (Bernardo and Smith 1994). The Bayesian analysis is based on the Bayes’s rule and the posterior distribution results from using the prior information about model parameters with evidence from the observed data. An advantage of Bayesian analysis is that the Bayesian paradigm allows us to prove some probability statements, as a variable is likely or unlikely to impact on another, or the true value of a parameter falls into a certain interval with a pre-specified probability (Bernardo and Smith 1994; Thompson 2012).

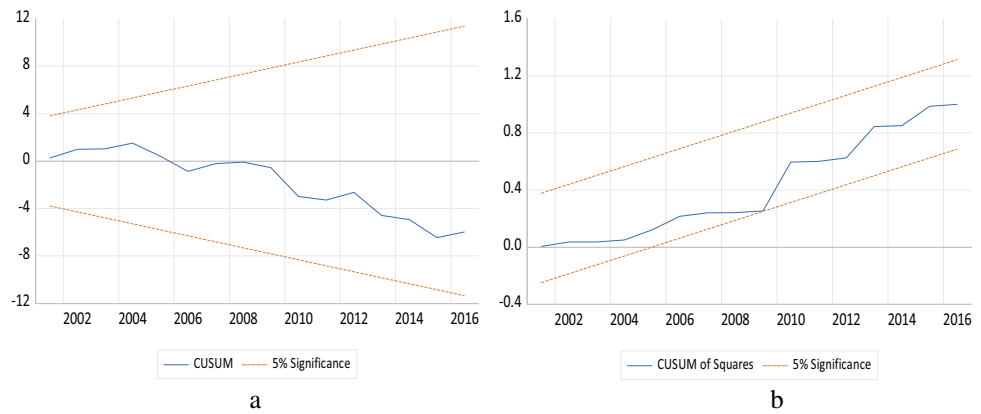
The specification of the Bayesian GLM regression is expressed as

$$y_t \sim N(\beta^T X_t, \delta^2 I) \tag{6}$$

where y_t is the ecological footprint drawn from normal Gaussian distribution and X_t is the matrix of the independent variables. β^T denotes the transposed weight matrix, while δ^2 is the variance and I is the identity matrix, to give the model a multi-dimensional formulation. Generally, the prior distribution is defined as pre-existing information about model parameters and is often derived from theoretical or expert knowledge. Fortunately, Lemoine (2019) suggested that the estimated coefficients obtained from the OLS approach could be used in the case of weakly informative priors in Bayesian analysis. By this brilliant suggestion, the study adopts the estimated coefficient of variables from the OLS estimator to set the initial information of the Bayesian GLM model with assuming a normal distribution.

In the Bayesian GLM model, the posterior distribution of the estimated model parameters is generated from a probability distribution based on the prior information, observed data, and outputs. The posterior distribution can be calculated as the following equation:

Fig. 1 **a** The CUSUM test and **b** the CUSUMSQ test



$$P(\beta | y_t, X_t) = \frac{P(y_t | \beta, X_t) * P(\beta | X_t)}{P(y_t | X_t)} \tag{7}$$

where $P(y_t|\beta, X_t)$ is the likelihood of the data and $P(\beta|X_t)$ denotes the prior probability information of the model parameters, while $P(y_t|X_t)$ represents the normalization constant. In this work, the adaptive random-walk Metropolis-Hastings algorithm was used to avoid the spurious convergence and provide probabilistic interpretations of model uncertainty and varying effects of financial development, economic growth, and human capital on EF. The result of Bayesian analysis is presented in Table 7.

The empirical outcome given in Table 7 reveals a positive effect of financial development and economic growth on EF, while the impact of human capital is negative. These results are in line with the coefficients obtained by the ECM model. More importantly, the linkage between financial development and ecological footprint has been established. Additionally, the acceptance rate of 0.37 is larger than the optimal acceptance rate of 0.234 (Gelman et al. 1997), whereas standard deviation values of the parameters are small, and the Monte Carlo chain standard errors (MCSE) are close to one decimal. Besides, the CUSUM plots of the parameters are jagged, not smooth, and cross the X-axis (see Fig. 2).

These obtained outcomes provide evidence to accept the Metropolis-Hastings algorithm and confirm the high accuracy of the parameter estimates (Cowles and Carlin 1996). Thus, Bayesian inference is valid.

Contrary to frequentist inference, in Bayesian inference, 95% credible intervals indicate which range the true value of a certain parameter belongs to. For example, the mean value of the financial development variable (lnFD) lies in an interval between -0.1329 and 0.3213 with a 95% probability. As expected, given probability, we may state that financial development has a strongly positive effect on ecological footprint with a 79% probability (Block et al. 2011). Our robustness analysis using FMOLS, DOLS, and CCR method shown in Table 8, also supported positive effect of financial development on EF. The outcome is justified by the historical trend in money supply strategies and the serious ecological deficit situation in Singapore.

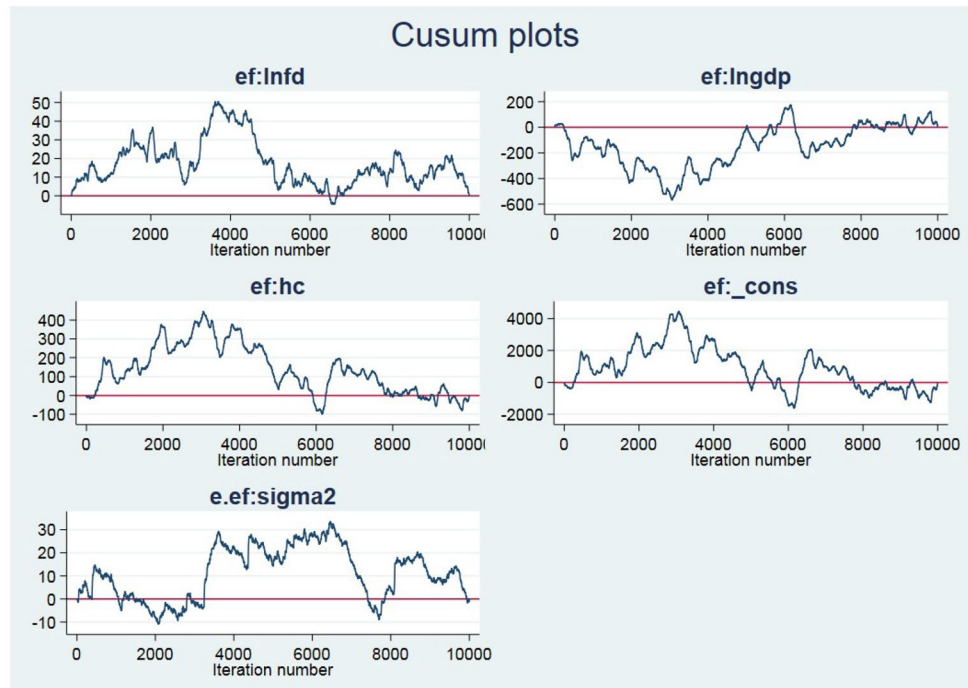
The Granger causality test The ECM or GLM approach does not guide us about the causal relationship among the variables. In the final step, the Toda and Yamamoto (1995) procedure is applied to check the Granger causality between numbers of pairs of variables. The equation is used to test has been shown in the “Data and methodology” section, and the empirical result is given in Table 9 and Fig. 3. According to Table 9, there is bi-directional causality between either

Table 7 The empirical results by Bayesian analysis

Variables	Mean	Std. Dev.	MCSE	Prob of mean > 0	[95% Cred. Interval]
Dependent variable: EF					
lnFD	0.0886	0.1145	0.0053	0.793	[-0.1329; 0.3213]
lnGDP	7.0411	0.9407	0.0687	1	[5.0955; 8.9038]
HC	-3.1321	0.7001	0.0512	1*	[-4.5258; -1.6912]
Intercept	-60.026	7.8752	0.5864	1*	[-75.158; -43.902]
e.EF	0.4305	0.1242	0.0044		[0.2601; 0.7158]
sigma2					
MCMC sample size = 10,000				Acceptance rate = 0.373	

* denote probability of mean < 0

Fig. 2 The CUSUM plot of the parameters



economic growth (lnGDP) or human capital (HC) and ecological footprint (EF). At the same time, Fig. 3 also indicated that (there if) uni-directional causality running lnGDP or EF or HC to financial development (lnFD), and running from HC to lnGDP.

Discussion

In frequentist inference, the obtained results demonstrate that the impact of financial development exerts an ambiguous effect in both the short run and long run, which is similar to the finding of Khan et al. (2019b). However, using the Bayesian inference, the obtained outcomes have found

a positive strongly influence of financial development on EF. With a 79% probability, the study may conclude that an expansion financial development policy may be harmful to environmental quality in the context of Singapore. The finding is inconsistent with Saud et al. (2020) and Naqvi et al. (2020), who validate the negative impact of financial development on EF, while it is in line with Destek and Sarkodie (2019), who find that financial development increases EF in Singapore. More importantly, by applying Bayesian analysis, this work overcame the ambiguity found in a previous study (Khan et al. 2019b). Thus, the study confirmed that financial development plays a positive role in the ecological footprint in Singapore. More

Table 8 Robustness check by FMOLS, DOLS, and CCR regression

Method(s)	Variables	Coefficient	Standard error	p value
FMOLS regression	lnFD	0.2331	0.1743	[0.090]
	lnGDP	6.4794	1.4048	[0.000]
	HC	-2.8550	1.0456	[0.010]
	Intercept	-57.062	12.072	[0.000]
DOLS regression	lnFD	1.3655	0.4511	[0.006]
	lnGDP	5.0736	1.7102	[0.007]
	HC	-1.2704	1.4995	[0.006]
	Intercept	-64.094	11.726	[0.000]
CCR regression	lnFD	0.3726	0.2884	[0.006]
	lnGDP	6.4769	1.5500	[0.000]
	HC	-2.8314	1.3089	[0.038]
	Intercept	-59.253	12.914	[0.000]

Table 9 Results of the causality test

Null hypothesis: no causality	Chi_square	p value
lnFD does not Granger cause EF	1.9601	0.5807
EF does not Granger cause lnFD	13.6307	0.0035
lnGDP does not Granger cause EF	11.6575	0.0087
EF does not Granger cause lnGDP	9.0829	0.0282
HC does not Granger cause EF	10.0785	0.0179
EF does not Granger cause HC	7.8212	0.0499
lnGDP does not Granger cause lnFD	17.3014	0.0006
lnFD does not Granger cause lnGDP	5.0863	0.1656
HC does not Granger cause lnFD	14.6027	0.0022
lnFD does not Granger cause HC	1.3994	0.7057
lnGDP does not Granger cause HC	5.0463	0.1684
HC does not Granger cause lnGDP	7.8931	0.0483

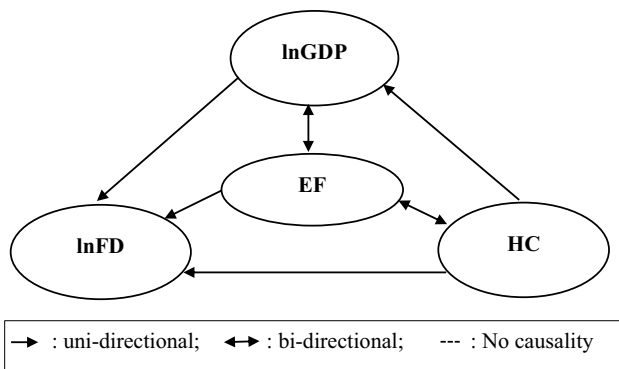


Fig. 3 Plots of the causality test

broadly, the finding is consistent with many studies in various contexts, such as the research of Khan et al. (2019a) in the context of five Belt and Road initiative (BRI) regions or Godil et al. (2020) in Turkey. Nevertheless, it should be noted that some studies show contradicted results, such as Uddin et al. (2017) on 27 leading world EF contributors, Ahmed et al. (2019) in Malaysia, or Omoke et al. (2020) in Nigeria (see also Table 1). In the context of Singapore, the study believes that financial development causes the scale effect by fostering economic liberalization and attracting foreign direct investment (Pazienza 2015; Saud et al. 2020). It means that an increase in financial development in Singapore leads to a rise in economic growth, high manufacturing and energy consumption, especially non-renewable energy consumption due to economic liberalization (Destek and Sarkodie 2019). This argument is consistent with some recent studies that ASEAN countries, including Singapore where non-renewable energy has been largely used in many industries (Kongbuamai et al. 2020; Nathaniel and Khan 2020). Consequently, the financial development is both side-effects on the increase of CO₂ emissions, air and soil pollution, and on ultimate raise level of EF (Pazienza 2015; Saud et al. 2020). As a result, the Singapore Government should considerably improve current policies regarding to financial development.

Both the ARDL and Bayesian analyses confirm that economic growth has a positive impact on EF (probability of 100%). The finding is consistent with the trade-off theory and the empirical evidence between economic growth and EF since economic growth has increased energy (Nathaniel and Khan 2020; Ozcan et al. 2018; Udemba 2020; Zafar et al. 2019), facilitation of urbanization (Ahmed et al. 2019; Nathaniel et al. 2020; Wu et al. 2019), and pressure on infrastructure and ecological assets (Sharma et al. 2020). Consequently, economic growth may break the balance in the biodiversity and increase EF, supported by the hypothesis of scale effect. As mentioned above, the increase in economic growth has generated a high level of non-renewable energy

consumption in ASEAN countries, including Singapore (Kongbuamai et al. 2020; Nathaniel and Khan 2020). To maintain environmental quality, therefore, more effective economic growth strategies should be implemented.

Finally, human capital mitigates EF in both the short and long run. The results are firmly validated by both the ARDL and Bayesian analyses with the probability of 100%. The finding is justifiable since some scholars (for instance, Nathaniel 2020; Nathaniel et al. 2021; Pata and Caglar 2021), who have discussed and empirically confirmed that human capital decreases EF. The results show that an increase in human capital leads to a corresponding rise in technological change adoption (Ackah and Kizys 2015; Consoli et al. 2016) as well as the awareness towards environmental quality problems (Adil 2018; Asongu 2018; Reynolds et al. 2010; Ulucak and Li 2020); and thus, activities of environmental destruction should be reduced and sustainable development could be more promoted in Singapore. The findings also demonstrate the consistency with other contexts that emphasize the importance of human capital and have appropriate policies to foster human capital, such as the USA (Zafar et al. 2019), G7 countries (Ahmed et al. 2020a; S. P. Nathaniel 2020), and Central and Eastern European Countries (Chen et al. 2019), among the others. The results likely prove how human capital is important and beneficial in the context of Singapore.

Conclusion and policy implications

Conclusions

We aim to uncover the association between financial development and EF by incorporating economic growth and human capital as control variables. To achieve this purpose, the study used the ARDL approach to provide preliminary results, and then a Bayesian analysis was employed to provide more insightful outcomes. As expected, the Bayesian analysis enables us to firmly conclude that both financial development and economic growth have a positive impact on EF, while human capital decreases EF. These findings indicated that the trade-off hypothesis between financial development and ecological assets is valid. Hence, monetary expansion policies should be associated with improving human capital to achieve sustainable development goals in the context of Singapore.

The findings confirm that the current financial development policies of Singapore may threaten environmental quality. They should be better assimilated and used more effectively in policy instruments regarding financial development. As noted by the previous studies, financial development may improve environmental quality by boosting research and development and technological advancement

(Ahmed et al. 2019). Thus, the Singapore Government should emphasize the vital role of the financial sector in providing funds for an adaptation to use eco-friendly technologies in new ventures and also to in the existing businesses as a replacement of outdated technologies (Ahmed et al. 2019; Usman et al. 2020a). Furthermore, the study also recommends that Singapore Government should facilitate the optimal utilization of energy by boosting financial support for eco-friendly projects at a minimum interest rate (Usman et al. 2020a). Finally, additional funds should be allocated to support environmental policies as well as to raise the environmental awareness of residents (Pata and Yilanci 2020).

Based on the positive impact of economic growth on EF in Singapore, the study suggests that the economic development policies should be focused to achieve less dependence on non-renewable energy (Nathaniel and Khan 2020; Udemba 2020; Ulucak and Apergis 2018; Zafar et al. 2019). Furthermore, the policies of financial development and economic growth in the context of Singapore should be integrated seamlessly to effectively foster environmental quality. More specifically, financial development should allocate more resources to developing and applying cleaner and more eco-friendly technologies. Novel technology is likely to remove dirty and obsolete technologies while fostering advanced and smart technologies in the manufacturing and residential sectors.

Furthermore, policies to improve environmental problems awareness are also significant in improving environmental quality (Adil 2018; Asongu 2018; Reynolds et al. 2010; Ulucak and Li 2020). Similarly, the study emphasizes an integrated approach to formulate policies about financial development, economic growth and human capital (Ahmed et al. 2021; Croes et al. 2021). Therefore, the work suggests that the Singapore Government should have a comprehensive vision about why financial development, economic growth, and human capital influence ecological footprint and provide an integrated yet effective way to improve environmental quality.

Availability of data and materials The datasets used and/or analyzed during the current study are available from the corresponding author on reasonable request.

Author contribution BHN conducted conceptualization, data curation, formal analysis, and writing — original draft. AA conducted the writing — review and editing — supervision, and project administration.

Declarations

Ethics approval Not applicable

Consent to participate Not applicable

Consent for publication Not applicable

Competing interests The authors declare no competing interests.

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