



The effects of FDI, technological innovation, and financial development on CO₂ emissions: evidence from the BRICS countries

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Abstract

The scholars of environmental economics have attempted the investigation of the impact of foreign direct investment-growth nexus, but they have missed the essential role played by technological innovation and financial development regarding the environmental costs. The notable economic growth and the consequent speedy process of urbanization in BRICS countries have brought about colossal escalation of energy needs leading to environmental degradation. The present study endeavors to explore the effect of foreign direct investment, technological innovation, and financial development on carbon emissions in BRICS member countries, with data from 1990 to 2017. The results verify a strong cross-sectional dependence within the panel countries. The Augmented Mean Group (AMG) estimator shows that foreign direct investment, technological innovation, and financial development in the BRICS countries possess a negative and statistically significant long-run association with CO₂ emissions, while economic growth, trade openness, urbanization, and energy use are found to contribute statistically significant and positive with carbon emissions. The current study chose to employ the Dumitrescu and Hurlin panel causality test for examining the direction of causality. Findings reveal a bidirectional long-run causality running among financial development, economic growth, trade openness, urbanization, energy use, and CO₂ emissions; on the contrary, unidirectional causality is found between foreign direct investment and carbon emissions. Consequently, for the BRICS member countries, the development of industries, financial institutions, and development of technological innovation are required to attract quality foreign direct investment. Moreover, urbanization contributes enormously to environmental degradation and necessitates urgent policy responses in these countries.

Keywords Foreign direct investment · Technological innovation · Financial development · EKC hypothesis · Environmental degradation

Introduction

Out of the major research domains in the field of environmental economics, climate change has attracted considerable attention from researchers throughout the globe because of its association and likely threats to sustainable development

(Destek and Sarkodie 2019), the consequential ever-increasing industrialization, and the ensuing urbanization, which, in the past few decades, have caused dramatic changes in the world (Dong et al. 2018b). This phenomenon is highly relevant to a group of states called BRICS that includes Brazil, Russia, India, China, and South Africa. The BRICS economies, by virtue of their rapid industrialization, have experienced remarkable economic growth which is evident from their GDP that reached the promising level of 2187 US\$ in 2010 and grew at the rate of 6.5% annually since 1985 (Dong et al. 2017; Azevedo et al. 2018; World Bank 2018; Danish and Wang 2019). The developing economies look upon the BRICS states as role models and a source of guidance and inspiration for economic growth (Danish et al. 2018). Goldman (2003) studies that the BRICS states have the

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capacity to challenge the economic monopoly of the G_6 in the coming few decades. This will be evident in the next 5–6 years when these economies will surpass half of the economic stature of G_6 by 2025. It is highly probable that these states will transform themselves by dint of economic growth, state-of-the-art technological advancement, strict pursuance of environment-friendly guidelines, and economic structural transformation from pollution intensive industrial phase to a highly advanced information exchange service centers (Destek and Sarkodie 2019). However, all that rapid advancement has and will not come without consequences. All these economic achievements have brought with them mammoth issues especially relating to environment, for instance, carbon emissions (Dong et al. 2017).

Foreign direct investment (FDI) denotes the transmission of technology, management skills, and knowledge from developed economies to underdeveloped economies (Doytch and Narayan 2016). FDI has been described as a trusted way for developing domestic production capabilities of an economy, increasing their investments through new finances and accessing innovative technologies (Xu et al. 2018). The FDI can be attracted through financial development and energy use, which, in turn, can stimulate economic development and encourage research-oriented activities for enhancing economic efficiency (Ziaei 2015). Concerning the analysis of possible association between sustainable development and FDI, three major hypotheses are dominant: FDI Halo Hypothesis, Pollution Havens Hypothesis, and Environmental Kuznets Curve (EKC). The FDI Halo Hypothesis suggests that FDI is likely to exercise positive environmental spillover impact as FDI usually brings superior technologies from the developed economies to the underdeveloped countries (Balsalobre-Lorente et al. 2019). Pollution Havens Hypothesis forwards that multinational corporations and trans-boundary plants serve as disseminators of advanced technology from advanced states to the less advanced countries encompassing the environmental sector and the consequent impact and improvement in environmental performance (Bakirtas and Cetin 2017). There has been an extensive advocacy by international monetary organizations advising the developing economies to bring reforms for attracting foreign investment for their economic development. The theoretical evidence of how GDP can boost economic development goes back to neoclassical growth models. The recent models have forwarded that technological advancement, produced through FDI, can yield instant as well as durable economic benefits (Hsiao and Hsiao 2006). It is widely believed that a country with a competent and stable financial sector has the potential to ensure improvement in investment procedure, financial risk mitigation, capital accumulation, and so forth thereby attracting higher degree of FDI, which, in turn, improves technological innovation (TI) and allows the retention of natural environment. The financial development plays a pivotal role

in the allocation of financial resources for the purpose of valuable ventures and mobilization of savings, which improve domestic production and bring about economic development. Many scholars have forwarded the idea that when the financial sector develops in a country, it can invite FDI as well as state-of-the-art technology that is friendly to environment (Birdsall and Wheeler 1993; Frankel and Rose 2002). Hence, it can be safely maintained that the financial sector expansion drastically affects energy consumption (Sadorsky 2011; Islam et al. 2013), which consequently impacts the degree of CO_2 emissions (Tamazian et al. 2009; Alam et al. 2012).

Today's world faces a huge challenge in balancing the environmental sustainability, on one hand, and countries desire for economic prosperity, on the other hand. To address the challenge of keeping the economic growth without risking their environmental pollution, countries in the world have introduced various policies to address climate change and minimize CO_2 emissions (Akadiri Saint et al. 2019a). In this respect, the economic growth model of BRICS countries has extraordinarily inspired the emerging economies. The BRICS countries have received this attention due to the rapid rise of their influence in the global economic development, which is also changing the global environmental governance pattern. The BRICS member countries are the most polluting countries in the world; resultantly, they emit biggest quantity of CO_2 emissions. Two of the BRICS members, China and India, are under immense pressure to control their respective CO_2 emissions because of their status as the main contributing sources for newly added carbon emissions. In the meantime, China and India are also under pressure to boost industrialization and subsequent urbanization in order to address poverty alleviation (Wang et al. 2018). China, being a BRICS member, is globally the biggest consumer of energy and consequently biggest CO_2 emitter (Birol 2016). On account of their limited resources, the developing economies find it difficult to invest in the environmental protection that comes at huge costs (Zhang 2011; Chang 2015; Seetanah et al. 2019). This situation strongly calls for a thorough investigation to assess how financial development and TI operate in developing countries and discover their impact on environment. The study may be helpful in formulating a broad consensus on feasible policies to address the issue of carbon emissions and minimize depletion of ozone layer and global warming (Alam et al. 2012).

Based on the abovementioned brief literature and taking into consideration the existing studies, the current study aims to examine the association between CO_2 emissions and its determinants: namely FDI, financial development, technological innovation, and economic growth in BRICS states. The current study obtains panel data of BRICS states from 1990 to 2017. To our best knowledge, the current study will be pioneering, illuminating, and an addition to the existing literature highlighting the association of TI and FDI with CO_2 emissions. The current study drastically differs from a

recent study conducted by Danish and Wang (2019) on BRICS Countries, where they investigated the association of natural resources and CO₂ emissions. The present study will add to the existing literature in the domain of FDI, technological innovation, financial development, and carbon emissions using the case of BRICS states. The study has highlighted the validity of EKC measuring the association of economic growth and carbon emissions.

The present study is of huge importance because it carries significant guidelines for future following the fourth assessment report of Intergovernmental Panel on Climate Change (IPPC), which maintains that the global atmospheric temperature must not increase beyond 2 °C. The environmental experts have identified a number of affordable alternative energy technologies for keeping the environment at such temperature (IPCC 2007, 2014; Tol 2007). The current study chose the case of BRICS countries owing to the fact that the BRICS member states are considered to be prone to the climatic changes and that in case of rise in global temperature, the coastal areas of these BRICS Countries are likely to face extreme weather conditions for instance floods, shortage of drinkable water, and even droughts. Majority of researchers have underscored the role of economic growth and climate change in emerging economies. The existing literature presents a contradiction, in theoretical and empirical dimensions, on the association between the variables under the study. The current study appears to differ in great deal from the studies carried out previously in the same domain. This difference ranges from time-range and methodology to the selection of variables. Majority of researchers prefer employing diverse range of proxies for financial development (Tamazian et al. 2009; Al-mulali et al. 2015); however, no study has ever used financial development index (FI) for BRICS countries, which comprises wide-ranging factors for measurement. The present study fills the said research gap by employing a single and comprehensive FI for BRICS countries and discover the actual role played by FI and TI across these countries.

This paper is presented in the following way: The second part gives a succinct, systemic review of the existing body of literature; the third section discusses research methodology, data collection, and estimation procedure followed by the section that covers results and findings, while the last part submits conclusions and recommendations.

Review of literature

Out of a host of most significant issues dominating the domain of environmental economics is the investigation of linkage between financial development and environmental degradation. The studies on the said association have come up with mixed findings. Some have portrayed a positive association between these variables, some have brought neutral results,

while the rest have found a negative association. This paper endeavors to explore the work of previous scholars on the main variables for setting the stage for the main theme.

FDI and carbon emissions

The central theme of the current study is focused on exploring the association between FDI and environment. Keeping in view the growing importance of BRICS states and the growth of FDI inflow there, which is likely to rise in the next few years, the study underlines the environmental costs of such gigantic FDI inflow within that region (Goldman 2003; Zeng et al. 2017; Dong et al. 2017; Azevedo et al. 2018; Danish and Wang 2019). The theories on FDI inform about contradictory effects of FDI on environment from being positive, neutral to negative. The FDI impact on environment can be determined through the degree of dominance of the transmission channels. These conflicting views are evident through three types of hypotheses: the pollution haven hypothesis, the pollution halo hypothesis, and the scale effects hypothesis (Pao and Tsai 2010). The findings of various studies carried out in different states and regions inform that FDI can exert varied impacts on environment. For example, the studies conducted by Pao and Tsai (2010) on BRICS states, Al-mulali and Binti Che Sab (2012) on GCC states, Paziienza (2015) on OECD states, and Zhang and Zhou (2016), Liu et al. (2017) and Xing et al. (2017) on China have found that FDI can improve the quality of environment.

In contrast, the studies carried out by He (2006) and Ren et al. (2014) in Chinese, Hitam and Borhan (2012) and Lau et al. (2014) on Malaysian, Solarin et al. (2017) on Ghanaian, Tang (2015) on Vietnamese context, Paramati et al. (2016) on 20 emerging states, Sbia et al. (2014) on Middle East region, Abdouli and Hammami (2017) on MENA, Shahbaz et al. (2015) on a wide variety of countries based on their income status, and a recent study by Shahbaz et al. (2018) using the case of France have concluded that FDI is one of the leading causes triggering environmental degradation. At the same time, some studies have come up with the findings which do not support either the positive or the negative implications of FDI and concluded insignificant results, for example, the study by Kiviyiro and Arminen (2014) on some countries of sub-Saharan region. Unpredictably, the studies that used a number of BRICS states as their cases also submitted contrasting findings (Zeng et al. 2017; Dong et al. 2017; Azevedo et al. 2018; Danish and Wang 2019). However, all these studies excluded the essential variables of financial development and growth, which can offer better findings when employed along with FDI. The present investigators intend to include financial development in a wide-ranging manner. Keeping in view the colossal degree of FDI along with economic and financial growth in BRICS

states, the current study aims to explore the combined effect of all these variables on environment. In order to execute that, we intend to employ several empirical approaches, which are elaborated in the following section.

TI and CO₂ emissions

The concept of technological change is credited to Schumpeter (1942) as quoted by Fields (2004) who presented the theory that a superior technological invention is incorporated into an existing market through three phases: invention, innovation, and diffusion. In his opinion, the process of research and development (R&D) is employed to execute the invention and innovation phases while the diffusion phase is executed when individuals or organizations adopt that TI in order to exploit it. The composite effect of these three phases is called the process of technological change. Owing to the internalization of technology employed as a variable into the model of market-functioning, the latest growth theory is referred to as “endogenous” growth theory. Experts believe that technological change is of vital importance in explaining major issues affecting the environment, which addresses larger picture with respect to time and scale and include climate change (Weitzman 1997; IPCC 2007). A wide range of arguments have been forwarded to recognize the actual magnitude of technological changes, which help in decreasing environmental pollution including changes in the fuel mix, employment of highly energy efficient production technologies, and installation of end-of-pipe technology which is regarded as the most significant among all. Regarding the investigation of climate change, which relates to energy and environment, the most significant theoretical assumptions address the nature and rate of technology change (Yeh and Rubin 2012). Likewise, some researchers contend that reduction in CO₂ emissions may be brought about through investment in R&D and technology change (Jones 2005). Others argue that if the society accepts these increased costs in order to decrease the degree of carbon emissions through technological development, it can surely help resolve the key issue of climate change (Newell and Pizer 2008). A study carried out by Sohag et al. (2015) found economic development and trade openness to increase the degree of energy consumption, while TI was found to augment energy efficiency as well as decrease energy consumption, therefore eventually cause reduction in CO₂ emissions. Conversely, a number of studies have submitted opposing findings about TI and environment. In this respect, the study by Parry (2003) endeavored to investigate the role of TI or the optimal pollution control (Pigouvian) welfare in reducing CO₂ emissions. It concluded that the welfare gains from optimal pollution control were larger in comparison with the welfare gains obtained from TIs. However, Smulders and de Nooij (2003) found that induced innovation could alleviate the per capita income decline but, because of the energy conservation policies, it could not completely counterbalance its impact.

Financial development and carbon emissions

The current tendency of research in the domain of environmental economics has focused largely on the investigation of linkage between financial development and environment. The fundamental factor for the achievement of economic growth is the role played by developed financial sectors. Many scholars believe that the developed financial markets can motivate the pace of the economic development if these markets can invite FDI and seek higher investment in R&D (Frankel and Romer 1999) thereby influencing the dynamics of environment. In the same way, scholars also believe that financial growth can produce eco-friendly technological advancement, which aims at keeping the environment least polluted, manufacturing eco-friendly products resulting in the rise of developmental sustainability at national, regional, and global levels (Birdsall and Wheeler 1993; Frankel and Rose 2002). The financial development may be blamed for causing rise in CO₂ emissions on account of its encouragement for production activities. In this regard, the growth in financial sector is considered to carry direct effect on the consumption of energy (Sadorsky 2011; Islam et al. 2013), and thus, it can affect the pattern of CO₂ emissions (Tamazian et al. 2009). On the same footing, a number of scholars have claimed that the growth in financial sector is capable to reduce the cost of borrowing, provide greater investment opportunities (Shahbaz et al. 2012), and boost the energy sector effectiveness, which, in turn, is highly likely to decrease CO₂ emissions (Tamazian et al. 2009; Tamazian and Bhaskara Rao 2010). Scholars also believe that all government sectors can get benefit from financial growth, which can finance development projects leading to TI (King and Levine 1993), which, in turn, result in significant reduction of CO₂ emissions by employing efficient energy utilization (Kumbaroughlu et al. 2008).

Financial development, according to Tamazian and Bhaskara Rao (2010), markedly influences in changing CO₂ emissions. However, a Turkish study by Ozturk and Acaravci (2013) found CO₂ emissions not to produce any significant impact from financial development. A recent investigation by Uddin et al. (2017) using the Kuwaiti case found financial development to impact negatively on environment. Surprisingly, there is an opposite research direction that has come up with findings suggesting insignificant association between financial development and degradation of environment. Such a direction includes Turkish study by Ozturk and Acaravci (2013), Middle Eastern and North African (MENA) study by Omri et al. (2015), study on GCC states by Bekhet et al. (2017), Emeriti study by Charfeddine and Ben Khediri (2016), and a European study by Coban and Topcu (2013) reporting either insignificant or mixed findings.

The above-discussed brief review of literature clearly signifies that no investigation has ever been carried out to associate the variables that the present study proposes employing the case of BRICS countries. Resultantly, the

findings of the current study are likely to add insights to the existing literature addressing the association of TI and CO₂ emissions for BRICS countries, which may be replicated using similar cases in future. The present study is also likely to offer addition to literature about the role of financial development and significance of economic development in affecting environment. It can offer assistance in identifying the role played by efficient energy consumption in order to check CO₂ emissions. The above discussion on the theoretical perspective of how FDI, TI, and financial development affect CO₂ emission guides us on the path to develop following hypotheses:

- Hypothesis a: There exists of a negative relationship between FDI and CO₂ emissions in BRICS countries.
- Hypothesis b: There exists of a negative relationship between TI and CO₂ emissions in BRICS countries.
- Hypothesis c: There exists of a negative relationship between FI and CO₂ emissions in BRICS countries.
- Hypothesis d: The existence of the environmental Kuznets curve (EKC) hypothesis in BRICS countries.

Data, methodology, and model specification

In order to analyze empirically, we employ panel data set for BRICS countries 1990–2017 acquired from the World Development Indicators (WDI) (World Bank 2018). Our dependent variable is CO₂ emissions per capita (metric tons), and independent variables are foreign direct investment (net inflows of GDP %), financial development (index), and patent applications as a proxy for technological innovation (TI). Our control variables include GDP, trade openness, urban population, and energy consumption. Table 1 shows the description of variables. The variable of TI is measured as the number of patent applications in accordance with the

guidelines supplied by Alam and Murad (2020) and Madsen et al. (2010). According to Levine (2005), two major elements make up the financial sector of an economy: Financial institutions comprising banking, mortgage, and insurance firms, which serve as financial intermediaries, and financial market consisting of capital market and other derivative markets. As there are four broad measures each for financial institution and financial market, for instance, financial depth, financial efficiency, financial stability, and financial access. This study uses the FI, which measures financial development of each category (Shoab et al. 2020). In financial terms, private credit to GDP serves as the proxy of financial depth. Net interest margin acts as a proxy of financial efficiency of economies. Z-score is employed to measure the stability of the financial system. Stock market capitalization to GDP variable measures the total value of shares of all the listed stock market companies to the total GDP. It acts as the proxy of financial depth in financial markets. Stock market turnover ratio implies the ratio between total shares traded during a financial year and the average market capitalization in the economy. It is the proxy of financial efficiency in financial markets. Most of the recent studies employ various proxies for financial development. For instance, stock market value addition, bank asset, and capital account liberalization were employed by Tamazian and Rao (2009) and Tamazian et al. (2009). While, percentage of domestic credit to the private sector was employed by Shahbaz et al. (2013a, b, c) and Al-mulali et al. (2015). The study by Abbasi and Riaz (2016) employed a combination of proxies: total credit as a percentage of GDP and stock market capitalization and the market traded turn-over measured as a percentage of GDP (Khan et al. 2017). We believe that this is the first work that employs FI index for BRICS countries, which consists wide-ranging measurement factors. The variable of urbanization has been incorporated in the proposed model because at the early stage of urbanization, the goods that consume more electricity are known to increase energy demand (Danish and Wang 2018). The process of urbanization can be stimulated by speedy development in economy, which can bring multidimensional structural changes throughout the economy, eventually resulting in affecting energy consumption (Danish et al. 2018). The studies by Islam et al. (2013) and Danish and Wang (2018) enlighten that urbanization, because of the movement and settlement of large numbers, supports economic activities; therefore, it intensifies the use of energy.

Table 1 Description of variables

Variable	Symbol	Description and measurement	Source
Carbon dioxide emissions	lnCO ₂	Metric tons per capita	WDI
Foreign direct investment	FDI	Net inflows (% of GDP)	WDI
Technological innovation	lnTI	Total patent applications	WDI
Gross domestic product	lnGDP	(Constant 2010 US\$)	WDI
Financial development index	FI	Index	IMF
Urban population	lnURPOP	Total urban population size (number)	WDI
Energy use	lnENR	kg of oil equivalent per capita	WDI

Econometric methods

Specification of model

The formal version of the proposed model may be composed in the following way:

$$\begin{aligned}
CO_{2it} = & \alpha_0 + \alpha_1 FDI_{it} + \alpha_2 FDI_{it}^2 + \alpha_3 TI_{it} + \alpha_4 GDP_{it} \\
& + \alpha_5 GDP_{it}^2 + \alpha_6 FI_{it} + \alpha_7 TOP_{it} + \alpha_8 URPOP_{it} \\
& + \alpha_9 ENR_{it} + \epsilon_{it}
\end{aligned} \quad (1)$$

Shahbaz et al. (2012) recommend that if the series of variables are transformed into natural logarithm, they can provide reliable and consistent results. Following Shahbaz et al. (2013a), taking log for variables, the estimation model for current study may now be composed in the following way:

$$\begin{aligned}
\ln CO_{2it} = & \alpha_0 + \alpha_1 FDI_{it} + \alpha_2 FDI_{it}^2 + \alpha_3 \ln TI_{it} \\
& + \alpha_4 \ln GDP_{it} + \alpha_5 \ln GDP_{it}^2 + \alpha_6 FI_{it} \\
& + \alpha_7 \ln TOP_{it} + \alpha_8 \ln URPOP_{it} + \alpha_9 \ln ENR_{it} \\
& + \epsilon_{it}
\end{aligned} \quad (2)$$

In the above Eq. (1), CO₂ shows carbon dioxide emissions per capita, FDI indicates foreign direct investment, and FDI² is the square of FDI implying that FDI > 0 and FDI² < 0 directed U-shaped between FDI. Similarly, GDP is the proxy for economic growth; GDP² is square of GDP shows a non-linear association between CO₂ emissions and income. FI is a financial development index, TI is proxy for technological innovation, TOP is trade openness, ENR is for energy use, and URPOP depicts urbanization. *i* and *t* illustrate the number of states and time span chosen for study, respectively. Since the studies that employ panel data use both dimensions: time series and cross-sectional, and the fact that the estimation methodologies based on panel data are efficient in controlling endogeneity, heteroscedasticity, serial correlation, and multicollinearity (Baltagi 2013), we use panel data analysis techniques for improved results.

Econometric procedures

CD tests

The detection of cross-sectional dependence (CD) serves as the primary step of panel data empirical analysis, which should be determined before the panel unit root tests are carried out (Rauf et al. 2018; Dong et al. 2019; Shuai et al. 2019). CD aims to remove the means in the correlation computation. The null hypothesis presumes the presence of cross-sectional independence in the panel, while the presence of CD illustrates the rejection of null hypothesis (Rauf et al. 2018). Breusch-Pagan Lagrange multiplier (LM) test might be inconsistent, so the bias-adjusted LM test (Pesaran et al. 2008) is used to explore the existence of CD in the panel series, which can be shown in the following way:

$$LM^* = \sqrt{\frac{2T}{N(N-1)}} \left(\sum_{i=1}^{N-1} \sum_{j=i+1}^N \hat{\rho}_{ij} \right) \frac{(T-k)\hat{\rho}_{ij}^2 - E(T-k)\hat{\rho}_{ij}^2}{\text{Var}(T-k)\hat{\rho}_{ij}^2} \quad (3)$$

where $\hat{\rho}_{ij}^2$ shows the residual pairwise correlation sample estimate, estimated by using a simple linear regression equation. The above-described models are asymptotically distributed as standard normal if the null hypothesis considers $T_{ij} \rightarrow \infty$ and $N \rightarrow \infty$.

Panel unit root tests

It is of vital importance to identify the order of the integration for each variable because the cointegration tests necessitate that all the variables must be integrated into order one; therefore, the panel unit root test should be carried out to accomplish this objective (Al-mulali et al. 2015). The recent literature suggests numerous panel root tests, which have been broadly bifurcated: The first group includes first generation tests for example LLC (Levin Lin Chu), Breitung, and Hadri panel unit root tests. They are all derived from diverse cross-sectional properties and depend on a common unit root process. The second group includes second-generation tests: IPS (IMPesaran Shin), Fisher ADF, and Fisher PP unit root tests. They control the homogeneity problem. Since the BRICS member countries possess wide-ranging economic structures and diverse CO₂ emissions levels and there exists CD across the panel countries, the current study chooses the second generation of unit root test and opts to apply a few, such as the Pesaran cross-sectionally Augmented Dickey-Fuller (CADF) and the Pesaran cross-sectionally Augmented Im, Pesaran, and Shin (CIPS) (Pesaran 2007).

Westerlund panel cointegration test

The study uses Westerlund (2005) and Kao (1999) owing to the presence of CD in panel. The test establishes a model with panel specific-AR test statistic and the same-AR test statistic, which can be estimated with the following equation, respectively (Wang and Dong 2019):

$$VR = \sum_{i=1}^N \sum_{t=1}^T \hat{E}_{it}^2 \hat{R}_i^{-1} \quad (4)$$

$$VR = \sum_{i=1}^N \sum_{t=1}^T \hat{E}_{it}^2 \left(\sum_{i=1}^N \hat{R}_i \right)^{-1} \quad (5)$$

where $\hat{E}_{it}^2 = \sum_{j=1}^t \hat{e}_{ij}$, $\hat{R}_i = \sum_{t=1}^T \hat{e}_{it}^2$, and \hat{e}_{it}^2 are the residuals from the panel regression model, while VR shows the group means variance-ratio statistic.

Panel long-run parameters estimation methods

A number of studies have suggested that if there is presence of long-run cointegrating associations of the time series, the estimation of long-run parameters should be carried out in the second phase. The key approach that we have employed in this study is fully modified OLS (FMOLS) for the reason that it uses the Newey-West to correct the autocorrelation of the error term U_{it} . Nonetheless, if the lagged and lead variables are chosen in the proposed models to control the errors of autocorrelation on the error term U_{it} , we can choose Dynamic OLS (DOLS). The mathematical expression for the panel FMOLS estimator given by Pedroni (1996, 2001) can be explained as given below:

$$\hat{\beta}_{\text{GMOLS}} = N^{-1} \sum_{n=1}^N \hat{\beta}_{\text{FMOLS},n} \tag{6}$$

where $\hat{\beta}_{\text{FMOLS},n}$ is the FMOLS estimator applied to applied to country n and the associated t-statistic can be written in the following manner:

$$t_{\text{GMOLS}} = N^{-1/2} \sum_{n=1}^N t_{\text{FMOLS},n} \tag{7}$$

However, since the FMOLS and DOLS estimators ignore the CD in the panel, they are likely to supply contradictory estimation. Therefore, the current study also employs Augmented Mean Group (AMG) estimator developed by Eberhardt and Bond (2009), for the estimation of the long-run parameter. The AMG estimator considers CD by including the common dynamic effect parameter, which can be estimated by a two-stage method, which may be written in the following way:

AMG – Stage 1

$$\Delta y_{it} = \alpha_i + \beta_i \Delta x_{it} + \gamma_i f_i + \sum_{l=2}^T \delta_l \Delta D_l + \varepsilon_{it} \tag{8}$$

AMG – Stage 2

$$\hat{\beta} = N^{-1} \sum_{i=1}^N \hat{\beta}_i \tag{9}$$

where Δ illustrates the first difference operator; x_{it} and y_{it} indicates observables; β_i indicates the country-specific estimators of coefficients; f_i indicates the unobserved common factor with the heterogeneous factor; δ_l indicates the coefficient of the time dummies and referred to as the common dynamic process; $\hat{\beta}_{\text{AMG}}$ indicates the mean group estimator for AMG; α_i indicates the intercept, while ε_{it} illustrates the error term.

Panel causality test

If there exists cointegration in panel data, the direction of causality should be evaluated. In view of the presence of CD among panels, the Dumitrescu and Hurlin (D-H) panel causality test was

chosen to identify the direction of causality among variables of interests in the developing economies. This test is derived from individual Wald statistics of assuming non-causality averaged across the cross-sectional units and may be described mathematically through the following equation:

$$y_{it} = \alpha_i + \sum_{j=1}^J \lambda_i^j y_{i(t-j)} + \sum_{j=1}^J \beta_i^j x_{i(t-j)} + \varepsilon_{it} \tag{10}$$

where y and x stand for observables; λ_i^j symbolizes the autoregressive parameters, while β_i^j symbolizes regression coefficient estimates, and they are supposed to differ across cross-sections. The null hypothesis is: there exists no causal association for any subgroup; while the alternative hypothesis is: there exists a causal connection for at least one subgroup of the panel. The above-stated hypothesis can be tested through an average Wald statistic in the following way:

$$W_{N,T}^{\text{HNC}} = N^{-1} \sum_{i=1}^N W_{i,T} \tag{11}$$

where $W_{i,T}$ indicates the individual Wald statistic for each cross-section unit.

Results and discussion

Descriptive statistics of variables and correlation results

Table 2 displays the summary statistics of the selected variables, whereas Table 3 describes the correlations among variables. The results reveal that the largest and smallest mean value of CO₂ emissions is (2.45) for Russia and (0.64) for Brazil. On the basis of mean GD p values, Russia is the richest (27.84) and South Africa is the poorest country (26.44). Besides, out of the BRICS countries, China has the highest mean value for FDI (3.48), while India possesses minimum mean value. Regarding the TI mean value, South Africa possesses highest mean value (5.23), while Russia possesses the lowest (4.50). CO₂ emissions are highly associated with energy consumption (56%), urban population (50%), TOP (45%), FI (33%), TI (18%), GDP (15%), and FDI (4%). We checked the multicollinearity through variance inflation factor (VIF) technique shown in Table 4. As a general rule, below 10 VIF value of a variable suggests absence of multicollinearity problem. The results exhibit that the VIF value is less than 10, so the issue of multicollinearity is no more applicable.

The analysis of panel data estimation suggests that currently CD appears to be the center of scholarly attention in the domain of environmental economics. The numerous tests we conducted reveal that ignoring the CD would render the results unreliable (Ahmad and

Table 2 Descriptive analysis of selected variables

Country	Variables	lnCO ₂	FDI	lnTI	lnGDP	FI	TOP	lnURPOP	lnENR
Brazil	Mean	0.64	2.564	4.772	28.177	0.369	22.76	18.804	7.064
	Std.Dev.	0.178	1.462	0.198	0.24	0.084	4.355	0.147	0.161
	Min	0.338	0.183	4.29	27.805	0.2	15.162	18.517	6.841
	Max	0.96	5.034	5.011	28.516	0.47	29.678	19.005	7.31
China	Mean	1.407	3.48	4.694	28.773	0.364	42.806	20.053	7.147
	Std.Dev.	0.474	1.317	0.703	0.773	0.139	11.127	0.303	0.423
	Min	0.766	0.966	2.565	27.442	0.001	24.273	19.52	6.602
	Max	2.063	6.187	5.533	29.947	0.556	64.479	20.505	7.713
India	Mean	0.109	1.183	4.754	27.727	0.463	34.872	19.585	6.139
	Std.Dev.	0.302	0.871	0.427	0.515	0.101	13.337	0.215	0.205
	Min	0.344	0.027	4.043	26.953	0.34	15.506	19.223	5.858
	Max	0.652	3.621	5.342	28.609	0.652	55.794	19.924	6.456
Russia	Mean	2.454	1.716	4.502	27.846	0.483	54.265	18.488	8.456
	Std.Dev.	0.09	1.275	1.129	0.257	0.102	14.308	0.014	0.102
	Min	2.315	0.175	0.693	27.424	0.323	26.257	18.47	8.289
	Max	2.638	4.503	5.318	28.156	0.627	110.577	18.508	8.688
South Africa	Mean	2.16	1.226	5.238	26.441	0.471	53.206	17.126	7.857
	Std.Dev.	0.076	1.317	0.326	0.241	0.116	9.261	0.197	0.061
	Min	1.942	0.066	4.369	26.099	0.184	37.487	16.768	7.737
	Max	2.301	5.983	5.505	26.778	0.615	72.865	17.441	7.99
Panel	Mean	1.354	2.034	4.792	27.793	0.43	41.582	18.811	7.332
	Std.Dev.	0.928	1.526	0.683	0.892	0.12	16.095	1.031	0.818
	Min	0.344	0.066	0.693	26.099	0.001	15.162	16.768	5.858
	Max	2.638	6.187	5.533	29.947	0.652	110.577	20.505	8.688

Zhao 2018; Dong et al. 2018a, b). Table 5 displays the results of the Breusch-Pagan LM and the Bias adjusted LM tests. The two statistics show the rejection of cross-sectional independence and confirmation of the presence of CD. This clearly suggests that if there arises a shock in one of the sample countries, it can spread out to other countries.

Table 6 exhibits the results of CIPS and CADF unit root tests, applied for two cases (a) at the level form (b) at first difference form (Δ). The results of the unit root

tests confirm that all the selected variables are stationary at first difference form.

For providing the cointegration analysis, we use Kao (1999) test of no cointegration between a group of variables (see Table 7). It may be noticeably observed that all the tests unanimously reject the null hypothesis of no cointegration following the Kao (1999) approach. Thus, it can be concluded that there exists cointegration association in the study variables. This particular finding yields huge significance and supplies a strong support in favor of the variables that possess

Table 3 Correlation for variables

Variables	lnCO ₂	FDI	lnTI	lnGDP	FI	TOP	lnURPOP	lnENR
lnCO ₂	1.00							
FDI	0.04**	1.00						
lnTI	0.18**	0.01*	1.00					
lnGDP	0.15*	0.48**	-0.24***	1.00				
FI	-0.33***	0.04**	0.18**	0.08**	1.00			
TOP	0.45***	0.02**	0.28***	-0.12*	0.49***	1.00		
lnURPOP	-0.50***	0.36**	-0.21***	0.48***	-0.01*	-0.25***	1.00	
lnENR	0.56***	-0.08*	-0.02**	-0.10*	0.31***	0.58***	-0.52***	1.00

*** significant at 1%, ** significant at 5%, * significant at 10%

Table 4 Test of multicollinearity

Model	VIF	Tolerance
FDI	1.34	0.746
lnTI	1.09	0.591
lnGDP	1.42	0.572
FI	3.09	0.819
TOP	1.64	0.608
lnURPOP	1.41	0.710
lnENR	2.05	0.486

DV is carbon dioxide emissions (metric tons per capita). The tolerance values are not less than 0.2 and VIF values are all less than 5, implying there is absence of multicollinearity

a long run association. It is of vital importance to supply additional cointegration testing, which can prove the robustness for CD.

In this respect, Westerlund (2005) is believed to be one of the highly reliable tests with the null hypothesis being no cointegration for all panels. Displayed through Table 8, the results of Westerlund test recommend that on the basis of the critical values generating from bootstrapped robust, there are three out of four tests, which reject the null hypothesis. These findings further confirm a long-run association among variables on the basis of cointegration (see Tables 9 and 10).

Results of panel long-run parameters estimation methods

Table 9 shows the long-run parameters of the estimation for FMOLS and AMG estimators. The two estimation methodologies offer comparable results. Nevertheless, the absolute value of the coefficients of *FDI* estimated by FMOLS are far smaller in comparison to those offered by AMG estimation, while that of *lnGDP* in FMOLS is significantly larger in comparison with that offered by AMG estimation. This suggests that there may be overestimation or underestimation of parameters because of the existence of CD. As a result, the AMG estimator is employed as the benchmark estimation method.

The AMG estimator suggests that the coefficient of *FDI* on *CO₂* emissions is negatively significant at 1% level. The

Table 5 Results of cross-section independence tests

Test	Statistic	<i>p</i> value
Breusch-Pagan LM	23.530***	0.000
Pesaran scaled LM	6.058***	0.000

*** statistical significance at 1% level

Table 6 Results of CIPS and CADF panel unit root test

Variables	CIPS		CADF	
	At level	Δ	At level	Δ
lnCO ₂	-1.06	-3.38***	-1.91	-2.07**
FDI	-2.14	-5.27***	-1.42	-4.12***
lnTI	-2.96	-5.31***	-4.93	-7.42***
lnGDP	-0.73	-1.35**	-4.89	-1.14**
FI	-2.08	-5.19***	-0.69	-4.33***
TOP	-2.41	-4.78***	-3.03	-5.79***
lnURPOP	-1.59	-1.34**	-1.44	-0.86**
lnENR	-1.37	-3.03***	-3.53	-2.21**

Asterisks (***, **) indicate the statistical significance at 1% and 5% level. Δ indicates at first difference form

decrease in *CO₂* emissions is 7.3% due to 1% increase in *FDI*. The negative coefficient of *FDI* supports the pollution halo hypothesis. These finding aligns with the studies by Pao and Tsai (2010) on BRICS states, Al-Mulali and Tang (2013) on GCC states, Pazienza (2015) on OECD states, and Zhang and Zhou (2016), Xing et al. (2017), and Liu et al. (2017) on China that have concluded that *FDI* develops the quality of environment. *FDI* inflowing is a major factor in assisting to expand environmental and advanced technological skills and *TI* in production. Conversely, the developing economies can serve as “pollution havens” for global polluting industries. Developed countries have strict pollution regulations, so it is likely that industrial pollution may be transferred to developing states. The long-run findings reveal that technological development possesses a negative association with pollution. Technological innovation can reduce *CO₂* emissions by 2% at 5% level of significant. The production process is the key component of economic development and causes more environmental corrosion. The development in technological stature has a positive and healthy effect on environment through its role in the reduction of pollution; however, this beneficial effect is at its initial stages, where it demands some time to furnish satisfactory outcomes. The effect of *FI*, in due course, is also friendly to environment and stands significant at 5%

Table 7 Panel cointegration tests by Kao (1999)

Test statistics	Statistic	<i>p</i> value
Modified Dickey-Fuller	-1.704**	0.044
Dickey-Fuller	-1.785**	0.037
Augmented Dickey-Fuller	0.726**	0.033
Unadjusted modified Dickey-Fuller	-5.084***	0.000
Unadjusted Dickey-Fuller	-3.098***	0.001

Asterisks (***, **) indicate the statistical significance at 1% and 5% level

Table 8 Results of Westerlund cointegration test

Statistic	Gt	Ga	Pt	Pa
Value	− 3.127***	− 4.849***	− 6.418***	− 5.433***
Z-value	− 2.848	3.487	− 2.801	0.445
p value	0.000	0.000	0.000	0.000

*** level of rejection of no cointegration at 1% level of significance

level since a rise in a one percentage of FI will decrease 6.7% in CO₂ emissions. The current study results align with previous studies conducted by Tamazian et al. (2009) and Tamazian and Bhaskara Rao (2010) who incorporated the variable of financial development in their models and concluded that financial development has the potential to reduce environmental pollution. This may be observed that the impact of economic growth on CO₂ emissions is positive. The coefficient of GDP is highly significant and carries a positive sign, which raises in conjunction with the rise in carbon emissions. Therefore, the economic growth causes increase in CO₂ emissions in the developing countries. The study outcomes also suggest the conclusion that the developing states must focus on two core interlinked issues: economic growth and environment. The 1% rise in GDP (lnGDP) can contribute to 8.1% rise in carbon emissions. These study findings align with previous studies by Shahbaz et al. (2013c) and Shahbaz et al.

Table 9 Results of panel AMG, FMOLS, DOLS, and FE estimators

Dependent variable: lnCO ₂				
Variables	AMG	FMOLS	DOLS	FE
FDI	− 0.073*** (0.023)	− 0.072*** (0.021)	− 0.064*** (0.020)	− 0.042*** (0.016)
FDI ²	0.042*** (0.002)	0.066*** (0.003)	0.092*** (0.004)	0.051*** (0.002)
lnTI	− 0.020** (0.010)	− 0.034*** (0.012)	− 0.105*** (0.016)	− 0.014** (0.007)
lnGDP	0.081*** (0.029)	0.092*** (0.030)	0.088*** (0.026)	0.048*** (0.017)
lnGDP ²	− 0.019** (0.009)	− 0.015*** (0.005)	− 0.015*** (0.004)	− 0.014*** (0.005)
FI	− 0.067** (0.030)	− 0.069* (0.041)	− 0.036* (0.021)	− 0.028* (0.016)
TOP	0.035*** (0.012)	0.066*** (0.008)	0.013* (0.007)	0.012* (0.007)
lnURPOP	0.292*** (0.058)	0.362*** (0.028)	0.315*** (0.022)	0.141* (0.084)
lnENR	0.110** (0.047)	0.115*** (0.023)	0.109*** (0.017)	0.119*** (0.044)

Standard errors in parentheses *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$

(2018). The study also submits that the coefficient of GDP² (lnGDP²) is negative and highly significant. The findings of the BRICS states case study reveal that the greater the economic growth is, the higher the carbon emissions will be, and that the EKC hypothesis problem can be supported. The likely explanation to this dilemma can be that the developing countries may have obtained their income levels later than their development stage. Generally, the level of income and social development in a country is reflected through GDP.

The rise in CO₂ emissions was 11% due to 1% rise in the use of energy (lnENR) (Akadiri Saint et al. 2019c). It implies that the use of environment-friendly energy is highly desirable since it contributes more towards carbon emissions. The effect of TOP on carbon emissions is positive and significant. Concerning the effect of urbanization (lnURPOP) on CO₂ emissions, the results show a positively significant association between urbanization and CO₂ emissions. The effect of urbanization can be explained through the notion that during the early phase of urbanization, electronic commodities are purchased in large numbers, the city transportation system is expanded, and increased number of financial institutions are established. All these pursuits increase energy consumption, which, in turn, increase the CO₂ emissions levels (Akadiri Saint et al. 2019d; Akadiri Saint et al. 2019b; Saint Akadiri et al. 2020).

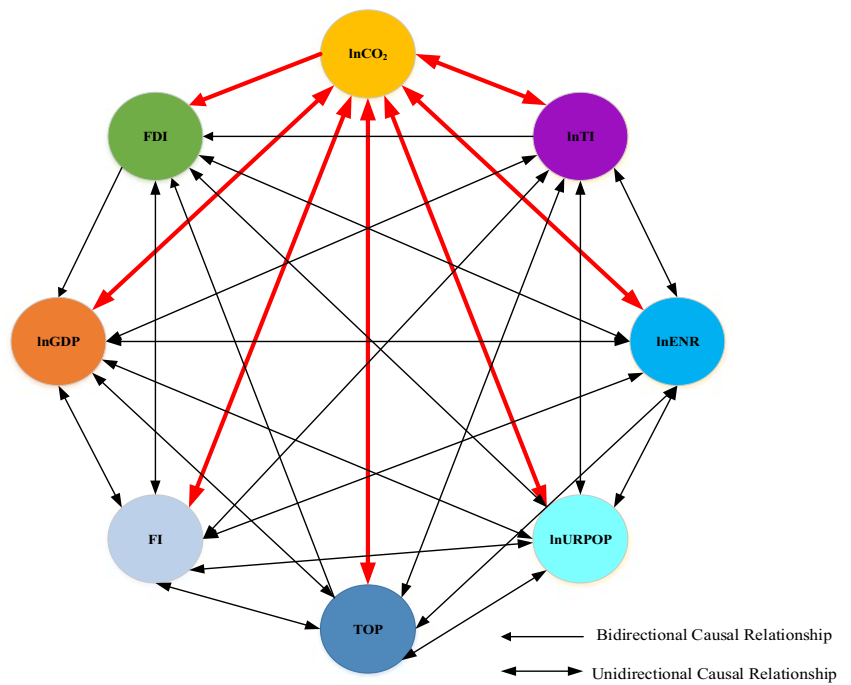
Fixed effects model and DOLS estimations offered the same results, which were almost identical to those obtained by AMG and FMOLS though with different coefficients values in these methods. Therefore, it can be maintained that the findings of this study are robust. Table 10 reports that the results of the estimation using the interaction term (FDI*lnTI) have positive and significant impact in all estimations included. This aligns with the results from estimations shown in Table 9. The finding reveals that elevated degree of technological development in the host country can cause higher capability of the host country to absorb any potential spillovers from FDI.

We investigate the causal analysis among study variables by using the D-H causality estimation method. Table 11 and Fig. 1 display the findings showing a unidirectional causality between FDI and carbon emissions. Moreover, bidirectional causal association exists among lnCO₂, lnTI, lnGDP, FI, TOP, lnURPOP, and lnENR. The results point out that the majority of the variables have two-way causality.

Conclusions and policy recommendations

The key current study objective is to evaluate the impact of FDI, technological innovation and financial development on CO₂ emissions using the case of BRICS states from 1990 to 2017. Considering the CD that may exist within cross countries, the current study employs a series of econometric

Fig. 1 Causality relationship flows



techniques, which results confirm the CD in panel countries. According to AMG and FMOLS results, the influence of FDI, TI, and financial development are negative and significant on carbon emissions. However, the impact of GDP, TOP, urbanization, and energy use are positively significant on carbon emissions. The estimated results reveal an inverted U-shaped relationship between GDP and CO_2 emissions. The application of DOLS and fixed effects model yield results, which are almost same than those obtained by FMOLS and AMG though with different coefficients values in these methods. Furthermore, a unidirectional causality between FDI and carbon emissions is found. Furthermore, bidirectional causal association exists between GDP, financial development, TI, TOP, urbanization, energy use, and carbon emissions.

The study results indicate that economic growth increases environmental degradation. This study finding verifies empirically the EKC existence for the BRICS countries (Bakirtas and Cetin 2017; Balsalobre-Lorente et al. 2019). We find that FDI helps to attain lower CO_2 emissions, thus validating the pollution halo hypothesis (Shahbaz et al. 2019; Balsalobre-Lorente et al. 2019). This implies that FDI inflows in conjunction with green technological transfer and improved labor and environmental management methods are highly likely to assist the BRICS states in achieving the sustainable development goals. Financial development is likely to perform a decisive role for environmental disclosure in BRICS countries. Our study also reflects the impact of urbanization and energy use on the CO_2 emissions. The empirical results of the current study offer suggestions for promoting more efficient system of urban transportation, environment-friendly architecture, and more sustainable energy use in urbanized localities

(Akadiri Saint et al. 2019d; Sarkodie and Strezov 2019; Balsalobre-Lorente et al. 2019).

Finally, the present study submits some policy implications for BRICS countries. Above all, it offers an understanding of the estimated dynamic associations among the variables that can offer assistance to the BRICS states in their endeavors to advance FDI, financial development, economic growth, urbanization, and technological progress for achieving the goal of reduction in CO_2 emissions. For instance, technological progress and efficient energy use can improve the overall environmental health of the BRICS member states. Financial growth can improve the quality of environment through larger investments in state-of-the-art and environment-friendly technologies; therefore, the governments should pay attention towards investments in environment-friendly technologies. We firmly believe that policies highlighting financial openness and liberalization for attracting elevated degrees of R&D associated FDI can reduce the degradation of environment. The governments must introduce regulations aimed at seeking guarantees from the foreign investment companies regarding implementation of green and environment-friendly technologies. Concerning energy consumption, the energy development programs must be transformed from non-renewable to renewable energy. Furthermore, the energy policies that aim at boosting the production and use of renewable energy will leave a positive and sustainable effect on economic growth. Thus, the actions aimed at controlling CO_2 emissions and policy recommendations ought to be shaped differently keeping in

view the degree of CO₂ emissions in a particular country.

While the current study has used FDI inflows, the future researchers may study the role of international trade and technological innovation in assessing pollution levels employing both EKC and the pollution haven hypothesis. This would help in understanding the determinants shaping the EKC.

Appendix

Table 10 Results of panel AMG, FMOLS, DOLS, and FE estimators with interaction term

Dependent variable: lnCO2				
Variables	AMG	FMOLS	DOLS	FE
FDI	-0.108** (0.055)	-0.151*** (0.034)	-0.294*** (0.037)	-0.105*** (0.023)
FDI*lnTI	0.014* (0.008)	0.024*** (0.007)	0.051*** (0.008)	0.018*** (0.004)
lnTI	-0.029* (0.017)	-0.026* (0.015)	-0.021* (0.011)	-0.048*** (0.012)
lnGDP	0.063*** (0.022)	0.091*** (0.051)	0.089*** (0.033)	0.083*** (0.031)
lnGDP ²	-0.013*** (0.004)	-0.057*** (0.021)	-0.054*** (0.003)	-0.031*** (0.008)
FI	-0.019*** (0.007)	-0.014* (0.008)	-0.018 (0.015)	-0.014* (0.008)
TOP	0.076*** (0.001)	0.067*** (0.008)	0.010** (0.005)	0.036*** (0.007)
lnURPOP	0.355*** (0.096)	0.358*** (0.029)	0.316*** (0.016)	0.095 (0.080)
lnENR	0.111*** (0.035)	0.115*** (0.024)	0.109*** (0.013)	0.131*** (0.042)

Standard errors in parentheses ****p* < 0.01, ***p* < 0.05, **p* < 0.1

Table 11 Results of D-H panel causality test

Null hypothesis	W-Stat.	Zbar-Stat.	Prob.
FDI ≠ lnCO ₂	5.832	-0.697	0.485
lnCO ₂ ≠ FDI	9.808	4.592	0.000
lnTI ≠ lnCO ₂	10.548	2.935	0.003
lnCO ₂ ≠ lnTI	12.083	3.926	0.000
lnGDP ≠ lnCO ₂	9.961	14.169	0.000
lnCO ₂ ≠ lnGDP	3.576	4.074	0.000
FI ≠ lnCO ₂	2.960	3.100	0.001
lnCO ₂ ≠ FI	2.545	2.443	0.014

Table 11 (continued)

Null hypothesis	W-Stat.	Zbar-Stat.	Prob.
TOP ≠ lnCO ₂	8.685	5.189	0.000
lnCO ₂ ≠ TOP	5.010	6.340	0.000
lnURPOP ≠ lnCO ₂	1.973	1.192	0.023
lnCO ₂ ≠ lnURPOP	2.205	1.506	0.013
lnENR ≠ lnCO ₂	3.312	3.655	0.004
lnCO ₂ ≠ lnENR	5.869	7.698	0.000
lnTI ≠ FDI	3.815	2.030	0.042
FDI ≠ lnTI	0.654	-0.546	0.584
lnGDP ≠ FDI	1.948	1.500	0.133
FDI ≠ lnGDP	2.178	1.862	0.062
FI ≠ FDI	2.425	2.253	0.024
FDI ≠ FI	9.352	2.164	0.030
TOP ≠ FDI	2.474	1.870	0.052
FDI ≠ TOP	1.412	0.434	0.665
lnURPOP ≠ FDI	3.180	2.823	0.005
FDI ≠ lnURPOP	4.544	4.667	0.000
lnENR ≠ FDI	4.566	5.639	0.000
FDI ≠ lnENR	4.695	3.013	0.002
lnGDP ≠ lnTI	3.588	4.092	0.000
lnTI ≠ lnGDP	7.580	6.239	0.000
FI ≠ lnTI	7.091	2.444	0.014
lnTI ≠ FI	6.850	5.422	0.000
TOP ≠ lnTI	2.368	2.163	0.030
lnTI ≠ TOP	10.249	2.742	0.006
lnURPOP ≠ lnTI	3.365	3.740	0.000
lnTI ≠ lnURPOP	8.365	6.074	0.000
lnENR ≠ lnTI	3.183	4.125	0.000
lnTI ≠ lnENR	4.154	2.408	0.016
FI ≠ lnGDP	3.830	4.812	0.000
lnGDP ≠ FI	8.910	10.568	0.000
TOP ≠ lnGDP	5.707	4.887	0.000
lnGDP ≠ TOP	3.157	4.089	0.000
lnURPOP ≠ lnGDP	12.391	15.272	0.000
lnGDP ≠ lnURPOP	11.671	16.979	0.000
lnENR ≠ lnGDP	6.379	9.962	0.000
lnGDP ≠ lnENR	9.611	2.054	0.000
TOP ≠ FI	2.567	1.346	0.001
FI ≠ TOP	7.749	2.242	0.045
lnURPOP ≠ FI	4.782	5.096	0.000
FI ≠ lnURPOP	3.029	3.971	0.000
lnENR ≠ FI	3.498	4.254	0.000
FI ≠ lnENR	5.213	7.165	0.000
lnURPOP ≠ TOP	7.586	10.669	0.000
TOP ≠ lnURPOP	9.253	8.031	0.000
lnENR ≠ TOP	10.503	15.795	0.000
TOP ≠ lnENR	10.528	7.761	0.000
lnENR ≠ lnURPOP	5.289	7.619	0.000
lnURPOP ≠ lnENR	12.236	18.547	0.000

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