



Does pollution haven hypothesis hold in newly industrialized countries? Evidence from ecological footprint

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Abstract

This study aims to investigate the validity of pollution haven hypothesis for the period from 1982 to 2013 in ten newly industrialized countries. For this purpose, we examine the relationship between real income, foreign direct investment, energy consumption, and ecological footprint using second-generation panel data methodology to take into account the cross-sectional dependence among newly industrialized countries. In doing so, the possible nonlinear relationship between foreign direct investment and environmental degradation is also searched. The results show that increased energy consumption and economic growth lead to increase in ecological footprint. Moreover, the U-shaped relationship between foreign direct investment and ecological footprint is confirmed in newly industrialized countries.

Keywords Pollution haven hypothesis · Ecological footprint · Newly industrialized countries

Introduction

In recent decades, although developing countries have experienced financing constraints in infrastructure investments triggering economic growth, they have overcome this problem through foreign capital investments and achieved high growth performances. Therefore, foreign direct investment (FDI, hereafter) inflows are considered a blessing for developing countries that have a resource shortage in financing high-cost investment projects. Furthermore, the fact that it is less prone to crises due to a long-term investment perspective and that it offers international technology access and know-how management makes foreign direct investment inflows more attractive for developing countries (Nunnenkamp 2001).

Even if developing countries implement many policies to host foreign direct investments, the driving forces and

economic consequences of foreign capital inflows are still a matter of debate. For instance, excessive increases in foreign capital investment inflows of developing countries have been generally explained with the cheap labor and natural resource abundance facility for multinational corporations. However, one of the most recent arguments called as *pollution haven hypothesis* (PHH) associates the FDI attractiveness of developing countries with looser environmental regulations of these countries compared to developed countries. According to this hypothesis, as a result of loose environmental regulations, the developing countries have gained the competitive advantage and dirty industries of developed countries have migrated to developing countries in order to reduce the production costs. On the other hand, the opposite hypothesis which is called as *pollution halo hypothesis* argues that the production structure of these multinational corporations is generally based on clean technology. Therefore, increasing mentioned investments spread out own modern technology to developing countries and reduce the pollution level in the developing countries. However, both hypotheses assume that there is increasing or decreasing linear relationship between FDI and environmental degradation. In fact, a possible nonlinear association between the mentioned variables is generally ignored.

Based on the above reasons, the main motivation of this study is to investigate the possible nonlinear relationship between foreign direct investment and environmental degradation to detect the validity of pollution haven hypothesis for the

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period from 1982 to 2013 in 10 newly industrialized countries. There are some reasons for choosing this country group. In the last three decades, the economic performances of newly industrialized countries (Brazil, China, India, Indonesia, Mexico, Malaysia, Philippines, South Africa, Thailand, and Turkey) have rapidly increased compared with the rest of the world. Namely, the share of the national output of these countries in the global output has risen from 10.7 to 23.1% over the observed period. In addition, the strong economic performance of these countries is mostly associated with foreign investment inflows of the countries because the FDI inflows of the selected countries have risen approximately 20 times in the last 30 years (WDI, 2017).

The contribution of this study is threefold: (i) this is the first study to examine the validity of the pollution haven hypothesis through an ecological footprint indicator instead of carbon emission as an indicator of environmental degradation in newly industrialized countries. The reason why we prefer the ecological footprint is that this indicator represents environmental degradation more accurately than carbon emissions (Ozturk et al. 2016; Destek et al. 2018; Ozcan et al. 2018; Ahmed et al. 2019). (ii) Unlike the linear assumption, this study examines the validity of possible nonlinear relationship between foreign direct investment and ecological footprint. Because, the empirical findings from the nonlinear empirical model allows for more detailed policy implications. (iii) Since the ignorance of the possible cross-sectional dependence may lead to invalid findings, this study employs the second-generation panel data methodologies that take into account the cross-sectional dependence among newly industrialized countries.

Literature review

The pollution haven hypothesis argues that developing countries may attract the multinational corporations from developed countries to lower production costs because of their cheap labor, abundant natural resources, and relax environmental regulations. In addition, the governments of developing countries stimulate these firms to reduce investment-saving deficit, to provide the required external capital, and to reduce the foreign trade deficit. However, it is also assumed that increased foreign direct investment inflows may cause environmental degradation by loose environmental regulations in the host country. This argument is confirmed for some developing or least-developed countries by many empirical studies. For instance, Omri et al. (2014) utilized GMM approach to test the relationship between CO₂ emissions, FDI, capital stock, GDP, trade openness, urbanization, financial development, and real exchange rate of the period spanning from 1990 to 2011 in 54 countries. The results confirmed the existence of PHH in these countries. Shahbaz et al. (2015) probed the linkage between CO₂ emission, foreign direct

investment, GDP, and energy consumption in 99 countries from 1975 to 2012. The results of FMOLS approach confirmed the PHH for 99 countries. In addition, they found an inverted U-shaped relationship between foreign direct investment and CO₂ emissions.

It is seen that many studies examining the PHH have focused on eastern Asian countries; Merican et al. (2007) utilized ARDL approach to probe the impacts of FDI on CO₂ emissions in five ASEAN (Malaysia, Thailand, Indonesia, Singapore, and the Philippines) countries spanning from the period 1970 to 2001. They found that the PHH is valid in Malaysia, Thailand, and Philippines. Lau et al. (2014) attempted to investigate the linkage between CO₂ emission, GDP, the square of GDP, FDI, and trade openness over the period of 1980–2008 in Malaysia using ARDL bounds test. As results of this study, FDI increased environmental pollution in Malaysia. Bakhsh et al. (2017) applied the three-stage least square method to analyze the relationship between environmental degradation, foreign direct investment, and GDP for the period of 1980–2014 in Pakistan. The results showed that FDI increases environmental degradation in Pakistan. Behera and Dash (2017) probed the link between CO₂ emission, primary energy consumption, fossil fuel energy consumption, urbanization, and FDI during the period 1980–2012 in 17 countries from South and Southeast Asian (SSEA) region and for three income-based panels (namely, high-, middle-, and low-income panels) using panel data methods. They found that FDI has significant impacts on CO₂ emissions in total, and high- and middle-income SSEA region. Baek (2016) used panel data method to investigate the validity of PHH for the period of 1981–2010 in five ASEAN countries utilizing FDI, GDP, energy consumption, and CO₂ emission data. According to PMG estimator results, the PHH does exist in these ASEAN countries.

Some studies focused on the developing countries that have cheap labor facility; Kiviyiro and Arminen (2014) used ARDL approach to analyze the linkage between carbon emissions, GDP, GDP square, energy consumption, and FDI over the period of 1971–2007 in six sub-Saharan African countries: the Republic of the Congo, the DRC (Democratic Republic of the Congo), Kenya, South Africa, Zambia, and Zimbabwe. The results of ARDL approach indicated that FDI increases CO₂ emissions in Kenya and Zimbabwe. Solarin et al. (2017) examined the relationship between CO₂ emission, GDP, GDP square, energy consumption, renewable energy consumption, fossil fuel consumption, foreign direct investment, institutional quality, urbanization, and trade openness to test the PHH for Ghana from 1980 to 2012. According to ARDL method, the PHH exists for Ghana. He (2006) aimed to explore the relationship between SO₂ emissions, FDI, and industrial output over the period of 1994–2001 in China's 29 provinces using panel data method. The results of the analysis provided significant evidence for the PHH in China's 29 provinces. Cole

et al. (2011) used data for 112 Chinese cities between 2001 and 2004 to examine the relationship between FDI, economic growth, four industrial water pollution emissions, and four industrial air pollution emissions as environmental pollution indicators. The obtained results validate the PHH because FDI increases environmental pollution in China. Wang and Chen (2014) analyzed the impacts of FDI and institutional development on industrial SO₂ emissions for the period of 2002–2009 in 287 Chinese cities. The study showed that FDI generally induces environmental pollution. Ren et al. (2014) explored the effects of trade openness, exports, imports, GDP, and FDI on CO₂ emissions over the period of 2000–2010 in China's industrial sectors using the two-step GMM estimation. The results suggested that FDI increases China's CO₂ emission. Sun et al. (2017) utilized the ARDL approach to investigate the PHH for the period 1980–2012 in China using CO₂ emission, GDP, GDP square, energy consumption, foreign direct investment, economic freedom, urbanization, financial development, and trade openness. The results of the ARDL method revealed that the PHH is valid in China.

In addition, some researchers focused on high-income emerging countries; Pao and Tsai (2011) investigated the relationship between CO₂ emission, GDP, GDP square, energy consumption, and FDI in BRIC (Brazil, Russia, India, and China) countries for the period between 1980 and 2007, except for Russia (1992–2007). The result of the empirical analysis is consistent with the PHH. Zakarya et al. (2015) examined the relationship between CO₂ emissions, energy consumption, GDP, and FDI over the period of 1990 to 2012 in BRICS (Brazil, Russian Federation, India, China, and South Africa) countries. According to the analysis, FDI increases CO₂ emissions in BRICS countries; thus, this result is consistent with the PHH. Moreover, Al-Mulali (2012) examined the link between CO₂ emissions, GDP, energy consumption, total trade, and FDI for the period of 1990–2009 in 12 Middle Eastern countries which have abundant natural resource using with panel data methods. The results of the analysis validate the PHH in these countries.

On the other hand, some studies found that the pollution halo hypothesis is valid for developing countries. The pollution halo hypothesis argues that environmental degradation will be reduced along with eco-friendly technology transfer from developed countries to developing countries by foreign direct investment inflow. This hypothesis is also confirmed by some studies; Tamazian and Rao (2010) employed GMM approach to determine the impacts of GDP, inflation, FDI, financial liberalization, trade openness, price liberalization, forex and trade liberalization, institutional quality, energy consumption, and energy imports on CO₂ emissions in 24 transitional countries for the period of 1993–2004. The results indicated that higher levels of FDI help to achieve lower CO₂ per capita emissions in these countries so the pollution halo hypothesis is valid. Kirkulak et al. (2011) purposed to test the impacts of

FDI on SO₂ emissions in 286 Chinese cities from 2001 to 2007 utilizing annual data of FDI, GDP, population, proportion of people working for science in the total population, proportion of FDI's output value in gross industrial output value, and SO₂ emissions. The regression results indicated that FDI has negative effect on SO₂ emission suggesting that the pollution halo hypothesis is valid. Al-Mulali and Tang (2013) employed the cointegration test of Pedroni (1999) and FMOLS method to estimate the effects of FDI, energy consumption, and GDP on CO₂ emissions over the period 1980–2009 in the Gulf Cooperation Council (GCC) countries. The researchers found that FDI has negative impact on CO₂ emissions; thus, the pollution halo hypothesis is confirmed in the GCC countries. Tang and Tan (2015) purposed to determine the effects of energy consumption, GDP, and FDI on CO₂ emission spanning from the period 1976 to 2009 in Vietnam. The results of the analysis revealed that FDI reduces carbon emission in Vietnam; thus, the pollution halo hypothesis is supported in Vietnam. Zhu et al. (2016) looked the effects of FDI, economic growth, energy consumption, on carbon emissions for the period of 1980–2010 in five ASEAN countries employing the panel quantile regression method. According to results of the analysis, the pollution halo hypothesis is supported in the five ASEAN countries. Zhang and Zhou (2016) aimed to determine the effect of FDI on CO₂ emission from 1995 to 2010 in 29 Chinese provinces using STIRPAT model. The results of the panel data confirmed that FDI decreases CO₂ emissions; thus, the pollution halo hypothesis is confirmed.

Similar to our study, some studies used the ecological footprint as an indicator environmental degradation to observe the validity of pollution haven hypothesis. For instance, Solarin and Al-Mulali (2018) employed the Augmented Mean Group (AMG) estimation procedure to observe the nexus between FDI and environmental degradation for the period from 1982 to 2013 in 20 countries and found that there is not significant relationship between foreign investments and environment. However, Baloch et al. (2019) investigated the nexus between FDI inflows and ecological footprint in 50 Belt and Road countries spanning from the period 1990 to 2016 and utilized the Driscoll-Kraay panel regression model. The results of this study reveal that pollution haven hypothesis is valid for the observed countries. Similarly, Fakher (2019) used the Bayesian model to observe the validity of PHH in developing countries for the period 1996–2016 and confirmed the validity of PHH.

Although different results are obtained according to the used methodology and observed countries, it is seen that most of previous studies focused on carbon dioxide emission as an indicator of environmental degradation. It is known that multinational corporations producing in developing countries do not only increase carbon dioxide emissions but also exploit the various environmental wealth

of these countries. Based on this reason, this study employed ecological footprint developed by Wackernagel and Rees (1998) which includes cropland, grazing land, fishing grounds, forest land, built-up land, and carbon footprint as indicators of environmental degradation instead of carbon dioxide emission. In addition, the limited number of studies that used the ecological footprint as an indicator of environmental degradation ignored the possible quadratic relationship between FDI and ecological footprint.

Data and methodology

This study aims to examine the effect of foreign direct investment on environmental degradation in ten newly industrialized countries for the period 1982–2013. For this purpose, we investigate the relationship between ecological footprint, economic growth, foreign direct investment, and energy consumption. Following Shahbaz et al. (2015), we also utilized the square of foreign direct investment to take into account the nonlinear term of foreign direct investment. The panel version of the empirical model is as follows:

$$\ln EF_{it} = a_0 + a_1 \ln Y_{it} + a_2 \ln FDI_{it} + a_3 \ln FDI_{it}^2 + a_4 \ln EC_{it} + e_{it} \quad (1)$$

where $\ln EF$, $\ln Y$, $\ln FDI$, $\ln FDI^2$, and $\ln EC$ are the natural log of the ecological footprint per capita, real GDP per capita, foreign direct investment per capita, the square of per capita foreign direct investment inflow, and energy consumption per capita. In addition, i , t , and e_{it} indicate cross-section, time period, and residual term, respectively. The ecological footprint per capita is measured in the sum of cropland, grazing land, fishing grounds, forestland, carbon, and built-up land footprints; the real GDP and the real foreign direct investment are measured in 2010 constant US dollars; and energy consumption per capita is measured in kilogram of oil equivalent. The data of real GDP per capita, FDI per capita, and energy consumption per capita are obtained from World Development Indicators, and ecological footprint data is retrieved from Global Footprint Network. In regard to the examined hypothesis, the pollution haven hypothesis is confirmed if FDI has monotonically increasing effect on ecological footprint ($a_2 > 0$ and $a_3 > 0$); the pollution halo hypothesis is supported in case if the FDI has monotonically decreasing effect on environmental degradation ($a_2 < 0$ and $a_3 < 0$); the U-shaped curve is confirmed in case of $a_2 < 0$ and $a_3 > 0$ and the inverted U-shaped curve is valid if $a_2 > 0$ and $a_3 < 0$.

Ignoring the cross-sectional dependence for panel data may lead to wrong estimations due to the high integration all over the world. Therefore, we first test the validity of cross-sectional dependence among newly industrialized

countries using with Pesaran's (2004) cross-sectional dependence (CD hereafter) test. The CD test is constructed as follows:

$$CD = \sqrt{\left(\frac{2T}{N(N-1)}\right) \sum_{i=1}^{N-1} \sum_{j=i+1}^N (\hat{\rho}_{ij})} \square N(0, 1) \quad (2)$$

where N and T state respectively the cross-sectional dimension and the time period. In addition, $\hat{\rho}_{ij}$ is the sample estimate of the pairwise correlation of the residuals.

In order to consider the cross-sectional dependence, we used well-known and frequently used unit root test developed by Pesaran (2007). The computation of the cross-sectional ADF (CADF) regression is as follows:

$$\Delta y_{it} = a_i + \rho_i y_{it-1} + \beta_i \bar{y}_{t-1} + \sum_{j=0}^k \gamma_{ij} \Delta \bar{y}_{it-1} + \sum_{j=0}^k \delta_{ij} y_{it-1} + \varepsilon_{it} \quad (3)$$

where a_i is the deterministic term, k is the lag order, and \bar{y}_t is the cross-sectional mean of time t . Following the above equation, t statistics are obtained with the computation of individual ADF statistics. Furthermore, CIPS is retrieved from the average of CADF statistic for each i as follows:

$$CIPS = \left(\frac{1}{N}\right) \sum_{i=1}^N t_i(N, T) \quad (4)$$

The critical values of CIPS for different deterministic terms are given by Pesaran (2007).

To test the validity of the long-run relationship between variables, we used error correction-based cointegration method developed by Westerlund (2007). In the testing procedure, there are four statistics (G_b , G_α , P_b , P_α) to test the null hypothesis if there is no cointegration. The test can be performed by testing the significance of the error correction term in the constrained panel error correction model. The main error correction model of the test can be written as follows:

$$\Delta Y_{it} = \delta'_i d_t + a_i Y_{i,t-1} + \lambda'_i X_{i,t-1} + \sum_{j=1}^{p_i} a_{ij} \Delta Y_{i,t-j} + \sum_{j=-q_i}^{p_i} \gamma_{ij} X_{i,t-1} + \mu_{it} \quad (5)$$

where d_t refers to the deterministic terms; and $d_t = 0$ (no deterministic term), $d_t = 1$ (with constant term), and $d_t = (1, t)'$ (with constant term and trend). Moreover, a_i determines the speed at which the system returns to the equilibrium, after an unpredictable shock.

Pesaran (2006) developed a new estimator that takes into account the cross-sectional dependence. Based on Eq. 1, the residual term (e_{it}) is a multifactor residual term and is constructed as follows:

$$e_{it} = \lambda'_i UF_t + u_{it} \quad (6)$$

where UF_t is the $m \times 1$ vector of unobserved common factors. In addition, Pesaran (2006) utilizes cross-sectional averages to deal with cross-sectional dependence of residuals as observable proxies for common factors.

Empirical results

In the first step, we examine the possible cross-sectional dependence among newly industrialized countries by using the CD test of Pesaran (2004). As a seen in Table 1, the null hypothesis if there is no cross-sectional dependence is rejected; thus, we confirm the cross-sectional dependency among the selected countries. This means a shock in one of the newly industrialized countries may easily be transmitted to the other countries. In this direction, we should use the second-generation panel data methodologies which take into account the cross-sectional dependence among countries.

In the second step, we investigate the stationary properties of the variables using augmented IPS (CIPS) unit root test of Pesaran (2007). As a shown in Table 2, the null hypothesis of unit root is not rejected for the level form of variables. However, the null hypothesis is rejected and all variables have become stationary for first differenced forms of variables. This finding means all variables are integrated of order one.

In the next step, we examine the existence of the long-run relationship between variables. In doing so, we utilized the cointegration test of Westerlund (2007). The results are illustrated in Table 3. When we evaluated the results of error correction–based cointegration test of Westerlund (2007), we found that G_t and G_a statistics do not reject the null hypothesis. On the other hand, P_t and P_a statistics confirm the long-run relationship between variables. These conflicting results may be sourced from that while G_t and G_a statistics are computed with the assumption of unit-specific error correction parameters and accepted as mean-group tests. However, P_t and P_a tests are calculated under the assumption of common error–correction parameter across cross-sectional units. Based on these reasons, our results can be interpreted as the existence of the long-run relationship between ecological footprint, real income, foreign direct investment, and energy consumption is supported with the assumption of common error–correction parameters. Fortunately, the finding of weak cointegration does not pose a

Table 1 Cross-sectional dependence test results

	lnEF	lnY	lnFD	lnFD ²	lnEC
CD test	18.580***	32.880***	18.110***	18.100***	21.990***
<i>p</i> value	0.000	0.000	0.000	0.000	0.000

***Statistical significance at 1% level

Table 2 Panel unit root test results

	CIPS-stat (level)	CIPS-stat (first differences)
lnEF	− 2.371	− 5.637***
lnY	− 2.169	− 3.945***
lnFDI	− 2.004	− 5.029***
lnFDI ²	− 2.133	− 5.469***
lnEC	− 1.798	− 5.244***

***Statistical significance at 1% level

problem in that the CCE estimator can be used without the need for pre-testing procedure such as cointegration.

Next, we used common correlated effect (CCE) estimator of Pesaran (2006) to obtain the coefficients of variables. In case of the panel results, as a seen in Table 4, CCE-mean group estimation results show that increasing real income per capita increases environmental degradation. This finding shows that the increase in economic activity still accelerates environmental degradation and is consistent with the studies of Ahmed et al. (2019) and Baloch et al. (2019) which found the ecological footprint increasing impact of economic growth. Moreover, the environmental degradation increasing impact of energy consumption is also validated. These findings mean that the energy portfolio of newly industrialized countries is still predominantly composed of fossil energy sources, and this finding is consistent with Solarin and Al-Mulali (2018). In addition, the significant signs of the coefficient of foreign direct investment and the coefficient of the square of foreign direct investment are found opposite. Therefore, the validity of pollution haven hypothesis and pollution halo hypothesis is rejected and the existence of nonlinear form is confirmed. The negative sign of the coefficient of the foreign direct investment and the positive sign of the square of foreign direct investment imply the U-shaped relationship between foreign direct investment and ecological footprint. This finding is inconsistent with the work of Shahbaz et al. (2015) which found the inverted U-shaped relationship between foreign direct investment and environmental degradation. This inconsistency may be sourced from used environmental degradation indicator. Namely, Shahbaz et al. (2015) utilized with CO₂ emission as an indicator of environmental degradation, but this study used

Table 3 Panel cointegration test results

	Statistic	<i>p</i> value
G_t	− 2.387	0.120
G_a	− 6.128	0.220
P_t	− 7.854*	0.050
P_a	− 6.592*	0.050

*Statistical significance at 10% level

Table 4 CCE estimation results

	lnY	lnFDI	lnFDI ²	lnEC
Brazil	0.541 [0.582]	− 0.149* [0.077]	0.019** [0.009]	0.303 [0.549]
China	0.216*** [0.064]	− 0.045*** [0.011]	0.007*** [0.002]	0.489*** [0.055]
India	− 0.079 [0.169]	− 0.007 [0.005]	0.002 [0.011]	0.805*** [0.158]
Indonesia	0.339*** [0.125]	− 0.003 [0.002]	0.001 [0.002]	0.122 [0.114]
Malaysia	0.272 [0.373]	− 0.185*** [0.071]	0.025** [0.011]	0.217 [0.244]
Mexico	0.603 [0.715]	0.004 [0.563]	0.002 [0.058]	0.418 [0.661]
Philippines	0.210*** [0.069]	− 0.005 [0.017]	− 0.001 [0.004]	0.096 [0.199]
S. Africa	0.817*** [0.160]	− 0.006 [0.003]	0.001 [0.001]	0.225*** [0.052]
Thailand	0.122*** [0.031]	− 0.045*** [0.006]	0.009** [0.005]	0.098*** [0.027]
Turkey	0.688*** [0.201]	− 0.086** [0.034]	0.010** [0.004]	0.530** [0.205]
CCE-MG	0.442*** [0.188]	− 0.052** [0.021]	0.007** [0.002]	0.166* [0.089]

*, **, *** indicate statistical significance at 10, 5, and 1% level, respectively. Numbers in brackets are standard errors

ecological footprint which covers carbon footprint as well as different degradation indicators.

In case of country-specific results, it can be seen that the U-shaped relationship between foreign direct investment and ecological footprint is found in Brazil, China, Malaysia, Thailand, and Turkey. In addition, increased real income per capita increases the ecological footprint per capita in China, Indonesia, Philippines, South Africa, Thailand, and Turkey. Similarly, increasing energy consumption increases the ecological footprint in China, India, South Africa, Thailand, and Turkey. However, based on the finding of cross-sectional dependence among countries, the group-mean results are more reliable for policy implications.

Conclusions and policy implications

This paper examines the validity of the pollution haven hypothesis by investigating the relationship between real income, foreign direct investment, energy consumption, and ecological footprint for the period from 1982 to 2013 in ten newly industrialized countries: Brazil, China, India, Indonesia, Malaysia, Mexico, Philippines, South Africa, Thailand, and Turkey. In addition, the square of foreign direct investment is used as explanatory variable to take into account the possible nonlinear relationship between foreign direct investment and ecological footprint. In doing so, the second-generation panel data methodologies are employed to consider the cross-sectional dependence among the selected countries.

In case of group mean results, we found that increased energy consumption and economic growth lead to increase in ecological footprint. Moreover, we found that the significant signs of the coefficients of foreign direct investment and the square of the foreign direct investment are opposite; thus, the validity of both pollution haven hypothesis and pollution halo hypothesis is rejected and the U-shaped relationship

between foreign direct investment and ecological footprint is confirmed. This finding means that environmental degradation decreases to a certain level with increasing foreign direct investment, and after this level, increasing foreign direct investment increases environmental degradation. In case of individual country results, the results show that the sign of the coefficient of foreign direct investment is negative and the sign of the coefficient of the square of foreign direct investment is positive in Brazil, China, Malaysia, Thailand, and Turkey. Therefore, the U-shaped relationship between foreign direct investment and ecological footprint is found for these countries. In addition, we found that increased energy consumption increases ecological footprint in China, India, South Africa, Thailand, and Turkey and economic growth leads to increase in ecological footprint in China, Indonesia, Philippines, South Africa, Thailand, and Turkey.

The existence of the U-shaped relationship between foreign direct investment and environmental degradation may be sourced from the production structure of foreign firms. It is a well-known fact that the production structure of the mentioned countries is mainly based on fossil energy consumption and the energy efficiency level of domestic companies is lower than that of foreign companies. Therefore, the clean technology-based production activities of foreign companies initially reduce the environmental degradation in newly industrialized countries because of the relatively modern production structure. However, in the later stages, the loose environmental regulations of the host country led foreign companies to aim low-cost production instead of considering environmental quality. This situation creates a difficult dilemma for governments of these countries because the foreign investments have a key role on economic development processes of these countries.

In regard to policy implications, this study suggests that the laws which mandate the use of clean technology should be applied to the domestic investors who are producing with fossil energy consumption in order not to reduce the domestic

competitive power of foreign investors. In addition, the governments should provide tax incentives and subsidies for foreign investors to use eco-friendly technology instead of taking prohibitive measures. Moreover, the implementation of these laws should not only focus on the targets of reducing carbon emission but also prevent the activities causing damage to water resources, forest areas, and agricultural lands.

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