### **ORIGINAL ARTICLE**



# **Lung cancer metabolomic data from tumor core biopsies enables risk‑score calculation for progression‑free and overall survival**

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### **Abstract**

**Introduction** Metabolomics has emerged as a powerful method to provide insight into cancer progression, including separating patients into low- and high-risk groups for overall (OS) and progression-free survival (PFS). However, survival prediction based mainly on metabolites obtained from biofuids remains elusive.

**Objectives** This proof-of-concept study evaluates metabolites as biomarkers obtained directly from tumor core biopsies along with covariates age, sex, pathological stage at diagnosis (I/II vs. III/VI), histological subtype, and treatment vs. no treatment to risk stratify lung cancer patients in terms of OS and PFS.

**Methods** Tumor core biopsy samples obtained during routine lung cancer patient care at the University of Louisville Hospital and Norton Hospital were evaluated with high-resolution 2DLC-MS/MS, and the data were analyzed by Kaplan–Meier survival analysis and Cox proportional hazards regression. A linear equation was developed to stratify patients into low and high risk groups based on log-transformed intensities of key metabolites. Sparse partial least squares discriminant analysis (SPLS-DA) was performed to predict OS and PFS events.

**Results** Univariable Cox proportional hazards regression model coefficients divided by the standard errors were used as weight coefficients multiplied by log-transformed metabolite intensity, then summed to generate a risk score for each patient. Risk scores based on 10 metabolites for OS and 5 metabolites for PFS were signifcant predictors of survival. Risk scores were validated with SPLS-DA classifcation model (AUROC 0.868 for OS and AUROC 0.755 for PFS, when combined with covariates).

**Conclusion** Metabolomic analysis of lung tumor core biopsies has the potential to diferentiate patients into low- and highrisk groups based on OS and PFS events and probability.

**Keywords** Lung cancer · Metabolomics · Overall survival · Progression free survival · Risk score calculator · Tumor core biopsy

#### **Abbreviations**



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- PC1 First principal component PC2 Second principal component
- PC3 Third principal component
- PFS Progression free survival
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### **1 Introduction**

Accounting for an estimated 12–13% of new cancer diagnoses in 2021, lung cancer is one of the most prevalent cancer types in the U.S. Of ~ 6,00,000 deaths attributed to cancer by the end of 2021, 22% are estimated to be lung cancer (Siegel et al., [2021\)](#page-11-0). Survival has increased since the early 1990s, likely due to improvements in early diagnosis and treatment options. Yet, the outlook for any particular patient is dire, with a 5-year survival rate  $< 5\%$  for all stages combined (Howlader et al., [2021](#page-10-0)). Reasons for this poor survival are multifactorial, with two dominant issues being presentation at an advanced stage and poor treatment success rates for patients at an advanced stage (Goldstraw et al., [2007](#page-10-1); Puchades-Carrasco et al., [2016](#page-11-1)). Once patients have been diagnosed, identifying metabolic markers associated with high risk of early death has potential clinical utility for risk stratifcation and management of patient care.

Metabolomics aims to quantify a broad spectrum of small molecule metabolites within a biological sample and has emerged as a powerful tool for phenotyping a variety of diseases, including cancer (Collino et al., [2013\)](#page-10-2). Mass spectrometry (MS) and nuclear magnetic resonance spectroscopy (NMR) are two major techniques commonly employed in metabolomics studies. Specifcally for lung cancer, metabolomics has proven useful in a variety of diagnostic and prognostic roles (Bamji-Stocke et al., [2018\)](#page-10-3), helping to resolve the link between genotype and phenotype. Several studies have investigated the metabolic profles of lung cancer patients compared to healthy controls (Fan et al., [2009;](#page-10-4) Hori et al., [2011;](#page-10-5) Kami et al., [2013\)](#page-10-6), showing profound metabolic dysregulation in lung cancer. Metabolic alterations have also been found in diferent histology types (Hori et al., [2011](#page-10-5); Wikoff et al., [2015\)](#page-11-2) and pathological stages (Hori et al., [2011;](#page-10-5) Puchades-Carrasco et al., [2016](#page-11-1)) indicating that metabolomics is useful for determining disease progression.

Recent studies have investigated the use of metabolic profles from serum, plasma, urine, and tissue samples for predicting survival in a variety of cancers (Berker et al., [2019](#page-10-7); Huang et al., [2019](#page-10-8); Jin et al., [2014;](#page-10-9) Xie et al., [2017](#page-11-3)). In particular, a study using high-resolution magic angle spinning nuclear magnetic resonance (HR-MAS NMR) metabolomics data derived from paired serum-tissue samples discovered 9 spectral regions with signifcant diferentiation between short and prolonged survival groups of early stage lung cancer patients in a Kaplan–Meier survival analysis ( $p = 0.0100$ ) (Berker et al., [2019](#page-10-7)). Recently, four metabolites in were included in a weighted risk score calculation and signifcant diferences in survival probability were found between low- and high-risk advanced non-small cell lung cancer (NSCLC) patients receiving first-line chemotherapy (Shen et al., [2017](#page-11-4)). Based on the literature, there is a promising outlook on the use of metabolomics for survival prediction in lung cancer patients, and the area is relatively unexplored.

In contrast to recent cancer patient serum metabolomic (Peng et al., [2018](#page-11-5); Tian et al., [2018](#page-11-6)) or radiomic (Dercle et al., [2020](#page-10-10)) profling, we report patient survival based on metabolomic analysis of tumor core biopsies routinely obtained during lung cancer patient care. In (Berker et al., [2019](#page-10-7)), tissue samples were found to result in better prediction accuracy for diferentiating lung cancer subtypes ( $AUROC<sub>Test</sub> 0.82$ ) compared to serum samples (AUROC<sub>Test</sub> 0.73). In our previous study (Miller et al., [2021\)](#page-11-7), metabolomic profles of patient tissue samples successfully predicted lung cancer disease control and progressive disease groups as response to frst-line therapy. In the current study, tissue-derived metabolites along with the covariates pathological staging at diagnosis, age, sex, histology subtype, and treatment vs. non-treatment were evaluated as markers for determining hazard ratios and survival probability in lung cancer patients. Our hypothesis is that metabolomic analysis of the tumor tissue itself has the potential to reveal biomarkers for distinguishing patients based on survival. These metabolic biomarkers would have clinical utility for calculating a weighted risk score to separate patients into low- and high-risk groups for OS and PFS.

### **2 Methods**

The data for this study were obtained from an on-going study (Miller et al., [2021](#page-11-7)). Experimental workflow and data wrangling pipeline are summarized below and in Fig. [1](#page-2-0).

### **2.1 Tissue collection**

Patients were recruited from an ongoing study, as described in (Miller et al., [2021](#page-11-7)). Informed consent was obtained to participate. All specimens were collected following approved Internal Review Board protocols at University of Louisville Hospital (IRB 05.0523) and Norton Hospital (IRB 18.0264) from patients with known or suspected NSCLC. Demographic information, including age, sex, race, smoking history, personal history of malignancy, and relevant family history were recorded. Samples were collected by the clinical team, blinded to the research analysis.



<span id="page-2-0"></span>**Fig. 1** Diagram of study workflow

#### **2.2 Tissue processing and metabolite extraction**

Tumor core biopsies were immediately placed into 1 mL  $13C$  labeled glucose medium in 24-well cell culture plates and incubated at 5%  $CO<sub>2</sub>$  and 37 °C on a shaker for 24 h. Following incubation, a liquid–liquid metabolite extraction was performed. Briefy, 500 mL acetonitrile was added, and tissue was homogenized with a pellet mixer. After 2–3 min homogenization, 376 mL of DNase/RNase free water and 250 mL chloroform were added. Contents were vortexed and centrifuged 180×*g* for 20 min. The top (polar) layer was aspirated and frozen at − 80 °C. A quality control sample was prepared from an NSCLC tissue biopsy incubated in unlabeled glucose media for 24 h and processed likewise. Sample polar layers were flash frozen in liquid  $N<sub>2</sub>$ , then lyophilized for 24–48 h until dried and transported on ice to CREAM core facility for 2DLC-MS/MS analysis.

#### **2.3 2DLC‑MS/MS analysis and data pre‑processing**

All samples were analyzed in random order on a Thermo Q Exactive HF Hybrid Quadrupole-Orbitrap Mass Spectrometer coupled with a Thermo DIONEX UltiMate 3000 HPLC system (Thermo Fisher Scientifc, Waltham, MA, USA) equipped with a reversed-phase chromatography (RPC) column and a hydrophilic interaction chromatography (HILIC) column confgured to form a parallel 2DLC-MS system(Klavins et al., [2014](#page-10-11)). To obtain full MS data,

every sample was analyzed by parallel 2DLC-MS in positive (+) and negative (−) modes. One unlabeled sample in each group was analyzed by 2DLC-MS/MS in positive and negative modes to acquire MS/MS spectra at 20, 40, and 60 eV collision energies as a quality control for metabolite identifcation. Data of unlabeled samples were frst matched to an in-house database that contains parent ion m/z, MS/ MS spectra, and retention time of authentic standards (MSI Level 1 identifcation). Data without a match (MSI Level 2 identifcation) were analyzed using Compound Discoverer software v2.0 (Thermo Fisher Scientifc, Germany).

#### **2.4 Organization of MS peak intensity data**

2DLC-MS data was presented as an alignment table for each batch with retention time, m/z, signal intensity, stable isotope labeling, name of identifed metabolite, and database used for metabolite identifcation. Once metabolites from positive and negative modes were combined, a preliminary step to handle missing values was performed by removing features that contained more than 50% missing values and resulted in a dataset of 52 metabolites with approximately 25.7% missing values. Based on previous analyses, some data appeared to be missing not at random (MNAR) while some was missing at random (MAR). It is clear that certain metabolites were not detected in the quality control samples when particular batches of samples were analyzed, but there were other instances of seemingly random missing values.

Therefore, we chose to handle missing values by imputation rather than limit of detection (LOD) substitution. Data were log-transformed (van den Berg et al., [2006\)](#page-11-8), which is a common method to normalize biological data by centering it and correcting for heteroscedasticity (van den Berg et al., [2006](#page-11-8)), and imputed by probabilistic principal component analysis (PPCA) (Tipping & Bishop, [1999\)](#page-11-9). A complete subset of the original data was taken and simulated missing values were introduced. A variety of substitution and imputation methods were tested, where imputation by PPCA resulted in the lowest RMSE (Miller et al., [2021\)](#page-11-7). The R package *pcaMethods* was used to perform PPCA. The function was seeded with an integer of '1234' to allow for reproducibility and "maxIterations" was set to 1000.

### **2.5 Patient clinical data**

Out of 54 eligible patients, 46 had follow-up and survival time information, 44 of which had staging information. Data collected included subject (age, sex, primary ethnicity, primary race, status, age at death, overall survival, cancer description, histology subtypes, cancer stage and substage, and progression-free survival in days), treatment (therapy type, therapy details, days of therapy, chemotherapy agent type, chemotherapy agent repeat units, number of chemotherapy cycles completed, surgery results, overall response to surgery/chemotherapy, days since diagnosis and response assessment types), and other information regarding the specimens.

#### **2.6 Data analysis**

All statistical analyses were conducted in the R programming language version 4.1.0. Hazard ratios were calculated by Cox proportional hazards regression model (*coxph* function in *survival* package). Survival curves were generated by Kaplan–Meier analysis with p-values found by log-rank test using high and low log transformed metabolite intensity values to stratify patient groups, defned as metabolite levels above or below the median (*survft* function in *survival* package; *ggsurvplot* function in *survminer* package). Transformed metabolite intensity values were visualized by *heatmap.2* function in *gplots* package. Metabolite correlations were calculated by the *cor* function in the base R *stats* package. All possible combinations of three signifcant metabolites from the univariable Cox regression models were investigated in addition to covariates in the multivariable models. The optimal model was chosen based on the lowest average p-value across all three statistical tests [likelihood ratio test, Wald test, score (logrank) test]. Dimensionality reduction was achieved with partial least squares discriminant analysis (PLS-DA) to ensure the discrete dependent responses were taken into account. For PLS-DA

component analysis, the dependent response variable was found by stratifying patients according to event (death or progression) vs. non-event. PLS-DA was implemented with *plsda* function in *mdatools* package. Pathological stage at diagnosis, sex, histological subtype, and treatment vs. nontreatment included as covariates, were one-hot encoded and classifed as stage I/II (1) vs. stage III/IV (2), female (1) vs. male (2), adenocarcinoma (1) vs. squamous cell carcinoma (2), and non-treatment (1) vs. treatment (2).

#### **2.7 Metabolite‑based risk score analysis**

A linear equation was developed to stratify patients into low and high risk groups based on signifcant metabolites. Univariable Cox proportional hazards regression model coeffcients divided by the standard errors were used as weight coefficients multiplied by log-transformed metabolite intensity and summed to generate a risk score for each patient (Eq. [1](#page-3-0)). Patients categorized as high risk are expected to experience increased risk of event (death or progression), as defined by the median risk score as the cut-off:

<span id="page-3-0"></span>
$$
Score_i = \sum_{k=1}^{num} w_k M_{i,k}
$$
 (1)

where  $Score_i$  is the weighted risk score of each patient  $i$ ,  $w_k$  is the weight coefficient assigned to metabolite  $k$ ,  $M_{ik}$  is the log-transformed metabolite intensity for sample *i* and metabolite  $k$ , and *num* is the total number of significant metabolites identifed by univariable Cox regression.

### **2.8 Validation of metabolite‑based risk scores**

A sparse partial least squares discriminant analysis (SPLS-DA) classifcation model was created using the covariates only and covariates combined with the metabolite-based risk scores for predicting event vs. non-event OS and PFS with fvefold cross validation and 100 resampling iterations. Classifcation model was trained using *train* function in *caret* package with method "spls". P-values between covariates only and covariates+risk scores AUROC values were calculated using unpaired T-test (assuming equal variances).

### **3 Results**

### **3.1 Patient characteristics**

Patient clinical characteristics are summarized in Table [1.](#page-4-0) Out of 46 patients with follow-up information, 14 expired and 32 were alive at the end of the respective follow-up period. The median overall survival was 20.50 months, and median progression-free survival was 17.41 months. Patients

<span id="page-4-0"></span>



were evenly distributed in terms of pathological stage, while 44 of the 46 patients had pathological staging information, 30 of 46 patients received systemic therapy as some form of chemotherapy or immunotherapy, 17 had a surgical resection and 23 received radiation treatment. Primary histology of all 46 patients was non-small cell. Histology subtypes included 38 with adenocarcinoma and 8 with squamous cell carcinoma. All patients with pathological staging information  $(n=44)$  were considered for analysis.

#### **3.2 Metabolomic data and correlation analysis**

A heatmap of the dataset with missing values present is in Supplementary Fig. 1. The maximum, minimum, and number of missing values in each patient category for the log-transformed metabolite intensity values are shown in Supplementary Table 1. Correlations were found between all metabolites in the dataset, displayed as a heat map (Supplementary Fig. 2A). Several pairs of strongly correlated ( $\geq$  0.75) metabolites were identifed (Supplementary Fig. 2B) and these pairs were fltered from univariable and multivariable Cox regression models. Correlations between metabolites identifed as signifcant in the univariable Cox proportional hazards model analysis are reported for OS and PFS (Supplementary Table 2).

#### **3.3 Univariable and multivariable analysis of metabolomic profles**

Metabolites were considered as separate variables during Cox proportional hazards model analysis, independent of patient characteristics. Covariates signifcant in the univariable models were adjusted for in the multivariable models. Cox regression coefficients and hazard ratios for metabolites related to OS and PFS are in Table [2](#page-5-0).

In univariable analysis, 6 of 52 metabolites were identified as significant (0.01  $\leq p \leq 0.05$ ) and 4 were identified as highly significant ( $p \le 0.01$ ) for OS, while 4 metabolites were identifed as signifcant and 1 was identifed as highly signifcant for PFS. The frst 3 components of PLS-DA analysis (PC1, PC2, PC3) were significant covariates for OS, while only PC2 was signifcant for PFS. The covariates pathological stage and sex were signifcant in the OS model, while none of the covariates were signifcant in the PFS model. Male patients and those having pathological stage III/IV had an increased risk of death (HR 4.75; HR 6.65). A similar trend (not signifcant) was observed for PFS. Signifcant metabolites from the univariable models were included in multivariable analysis along with the signifcant covariates (Table [2](#page-5-0)). All possible combinations of three signifcant metabolites from the univariable Cox regression models were investigated in addition to covariates in the multivariable models. The optimal model was chosen based on the lowest average p-value across all three statistical tests (likelihood ratio test, Wald test, score (logrank) test). Multivariable models achieved signifcance for OS (likelihood ratio test *p*=7E−05, Wald test *p*=0.002, log-rank test  $p=4E-05$ ) and for PFS (likelihood ratio test  $p=0.008$ , Wald test  $p=0.002$ , log-rank test  $p=4E-04$ ).

Kaplan–Meier survival analysis (Fig. [2](#page-6-0)) revealed that several metabolites contributed to OS and PFS probability over time, where the maximum OS was 101.6 months. Metabolites contributing to lower OS probability included hypoxanthine, L-pyroglutamic acid, N6,N6,N6-trimethyl-L-lysine, proline, adipic acid, benzoic acid, and inosine (Fig. [2](#page-6-0)A). For PFS, low levels of adenine and histamine and high levels of inosine contributed to shorter survival (Fig. [2](#page-6-0)B).

#### **3.4 Risk score analysis**

Metabolite based risk scores were calculated using Eq. [1](#page-3-0) for OS and PFS (Figs. [3](#page-7-0) and [4,](#page-8-0) respectively). For both OS

### <span id="page-5-0"></span>**Table 2** Cox proportional hazards regression models of metabolites, PLS-DA components (PC's) and covariates



#### **Table 2** (continued)



Pathological stage at diagnosis, sex, histological subtype, and treatment vs. non-treatment included as covariates, were one-hot encoded and classified as stage I/II (1) vs. stage III/IV (2), female (1) vs. male (2), adenocarcinoma (1) vs. squamous cell carcinoma (2), and non-treatment (1) vs. treatment (2). Univariable analysis revealed 10 signifcant metabolites, 3 signifcant PC's and 2 signifcant covariates for Overall Survival (OS) while 5 significant metabolites and one significant PC were found for Progression Free Survival (PFS). Multivariable analysis resulted in a signifcant model after adjusting for sex and pathological stage, where guanosine remained as the only signifcant metabolite for OS, while glutamic acid remained as the only signifcant metabolite for PFS



<span id="page-6-0"></span>**Fig. 2** Kaplan–Meier survival analysis. Survival curves of lung cancer patients with high metabolite levels (red; *>* median) and low metabolite levels (blue;  $\leq$  median) for overall survival (**A**) and progression-free survival (**B**). P-values from log-rank test

and PFS, key metabolites were selected by signifcance in the univariable Cox proportional hazards regression model and correspond to the transformed intensity for each sample. In each case, the cut-off was defined as the median risk score to stratify high and low risk patients (Figs. [3A](#page-7-0), [4](#page-8-0)A). Probability of survival for low and high risk patients was determined by Kaplan–Meier analysis, resulting in a marginally signifcant survival diference for OS  $(p=0.0764)$  and highly significant difference for PFS  $(p = 0.0012)$  (Figs. [3](#page-7-0)B, [4B](#page-8-0)). Transformed intensity values of key metabolites among low and high risk patients are visualized as a heatmap (Figs. [3](#page-7-0)C, [4C](#page-8-0)). A Cox proportional hazards regression model was generated to determine hazard ratios of metabolite-based risk score and pathological stage at diagnosis (Figs. [3](#page-7-0)D, [4D](#page-8-0)). In univariable models, metabolite-based risk scores for OS [HR 1.15 (1.08–1.23), *p* = 6.36E−06] and for PFS [HR 1.29  $(1.14-1.46)$ ,  $p = 7.97E-05$ ] were significant predictors of survival. In OS multivariable model, metabolite-based risk score remained signifcant [HR 1.12 (1.06–1.19), *p* = 2.46E−04] after adjusting for sex and pathological stage. The multivariable model overall was signifcant for OS (likelihood ratio test *p*=2E−06; Wald test *p*=2E−05; log-rank test  $p = 8E-08$ ).



<span id="page-7-0"></span>**Fig. 3** Metabolite-based risk score analysis of lung cancer patient overall survival (OS). Patients were stratifed into low and high risk based on the median risk score as the cut-of (**A**). Kaplan–Meier survival analysis revealed that low and high risk patients had a signifcant diference in survival probability over time (**B**). Heatmap of log transformed intensity values of key metabolites for low risk (left) and high risk (right) patients (**C**). Cox proportional hazards regression results with metabolite-based risk score, sex and pathological

stage at diagnosis as predictors (**D**). Multivariable analysis indicates that the metabolite-based risk score was a more signifcant predictor of overall survival compared to sex or pathological stage at diagnosis. Pathological stage at diagnosis, sex, histological subtype, and treatment vs. non-treatment included as covariates, were one-hot encoded and classifed as stage I/II (1) vs. stage III/IV (2), female (1) vs. male (2), adenocarcinoma (1) vs. squamous cell carcinoma (2), and nontreatment (1) vs. treatment (2)

#### **3.5 Risk score validation**

SPLS-DA classifcation model was employed for validation of metabolite-based risk scores in predicting event vs. nonevent. With covariates only (sex, age, pathological stage, histological subtype, treatment vs. non-treatment), OS and PFS predictions achieved AUROC 0.819 and AUROC 0.652, respectively (Supplementary Fig. 3A). Classifcation model performance when covariates were combined with metabolite-based risk scores, resulted in statistically signifcantly (p<1E−10) higher AUROC 0.868 and AUROC 0.755 for OS and PFS, respectively (Supplementary Fig. 3B).

## **4 Discussion**

This proof-of-concept study tests the hypothesis that metabolomic analysis of tumor core biopsy samples routinely obtained during lung cancer patient care can reveal biomarkers for distinguishing patients based on OS and PFS. The study adhered to REMARK guidelines (McShane et al., [2005](#page-11-10)). The univariable analysis revealed metabolite-based risk scores derived from 10 key metabolites for OS and 5 key metabolites for PFS as signifcant predictors of survival. For OS, the multivariable model was signifcant after adjusting for sex and pathological stage at diagnosis, and the metabolite-based risk score was a more signifcant predictor than either of these covariates. In the multivariable model, guanosine remained as a signifcant metabolite for OS and glutamic acid remained signifcant for PFS. Five metabolites were identifed in common for OS by both Kaplan–Meier survival analysis and Cox proportional hazards regression models (l-pyroglutamic acid, proline, benzoic acid, inosine, and N6, N6, N6-Trimethyl-L-lysine) and one metabolite was identifed in common by both methods for PFS (inosine). The metabolite-based risk scores proved useful for predicting event vs. non-event in OS and PFS; when the risk scores were combined with covariates age, sex, pathological stage, histological subtype, and treatment vs. nontreatment, the AUROC improved compared to the AUROC obtained through the covariates alone. To our knowledge, this is the frst study to show that tumor core biopsy-derived



<span id="page-8-0"></span>**Fig. 4** Metabolite-based risk score analysis of lung cancer patient progression-free survival (PFS). Patients were stratifed into low and high risk based on the median risk score as the cut-off (A). Kaplan– Meier survival analysis revealed that low and high risk patients had a signifcant diference in survival probability over time (**B**). Heatmap of log transformed intensity values of key metabolites for low risk (left) and high risk (right) patients (**C**). Cox proportional hazards regression results with metabolite-based risk score as predictor (**D**).

metabolomic profles analyzed by 2DLC-MS/MS can risk stratify lung cancer patients based on survival probability.

Several metabolites related to nucleic acids were identifed as signifcant by the models in this study, including the nucleobase adenine, nucleosides guanosine and inosine, and the purine derivative hypoxanthine. A recent study which investigated serum/plasma metabolomics for survival prediction in lung cancer patients receiving frst-line chemotherapy found 4 metabolites which difered signifcantly  $(p < 0.05)$  between a discovery and validation cohort, including cafeine, paraxanthine, stachydrine, and methyl glucopyranoside (Shen et al., [2017\)](#page-11-4). Metabolites related to cafeine metabolism generally had higher levels in patients with poor survival. In the current study, hypoxanthine was signifcantly associated with OS (Fig. [2A](#page-6-0)). Interestingly, hypoxanthine is also related to cafeine metabolism (Suzuki & Takahashi, [1975\)](#page-11-11). However, the link to cafeine metabolism may be coincidental. Hypoxanthine guanine phosphoribosyltransferase (HPRT) was found to be expressed in cell membranes of NSCLC cell lines H460 and A549 (Townsend et al., [2017](#page-11-12)). H460 cells have nearly double HPRT expression than A549 and have roughly double the growth rate, indicating that HPRT expression may be directly related to cell proliferation (Townsend et al., [2017](#page-11-12)). HPRT is an

Univariable analysis indicates that the metabolite-based risk score was a signifcant predictor of progression-free survival. Pathological stage at diagnosis, sex, histological subtype, and treatment vs. nontreatment included as covariates, were one-hot encoded and classifed as stage I/II (1) vs. stage III/IV (2), female (1) vs. male (2), adenocarcinoma (1) vs. squamous cell carcinoma (2), and non-treatment (1) vs. treatment (2)

enzyme that functions by transferring phosphoribose from phosphoribosyl pyrophosphate (PRPP) to hypoxanthine or guanine bases to form inosine monophosphate (IMP) and guanine monophosphate (GMP) (Stout & Caskey, [1985](#page-11-13); Wilson et al., [1983](#page-11-14)). Therefore, it is unsurprising that high levels of hypoxanthine and inosine were strongly associated with shorter survival (Fig. [2A](#page-6-0)) as well as progression and death (Table [2](#page-5-0)). Interestingly, the opposite trend was observed with guanosine and adenine. High levels of adenine resulted in greater PFS probability over time (Fig. [2B](#page-6-0)), while increased guanosine was associated with OS (HR  $0.41, p = 0.0202$  (Table [2\)](#page-5-0).

Several amino acids and amino acid derivatives were identifed as signifcant, including glutamic acid, proline, L-pyroglutamic acid, N6,N6,N6-trimethyl-L-lysine, and pipecolinic acid. Hori et al. discovered that glutamic acid levels were lower in serum and higher in lung tissue samples of lung cancer patients compared to healthy controls (Hori et al., [2011](#page-10-5)). However, another study found that glutamic acid was signifcantly increased in serum of lung and breast cancer patients compared to healthy controls (Cascino et al., [1995](#page-10-12)). In the current study, increased glutamic acid in lung tumor tissue was associated with decreased risk of progression (HR  $0.42$ ,  $p = 0.0082$ ) (Table [2](#page-5-0)). This suggests a complex relationship between lung cancer survival and glutamic acid levels in tumor tissue compared to serum, and a more targeted experimental approach is required to elucidate the mechanisms involved.

The results further show that elevated proline was associated with death (HR 3.63,  $p = 0.0124$ ) (Table [2](#page-5-0)), consistent with its role supporting cancer cell survival and proliferation (Tanner et al., [2018](#page-11-15)). High levels of pyroglutamic acid (PGA) were also associated with death (HR 3.90,  $p = 0.0265$ ) (Table [2\)](#page-5-0). PGA is known to accumulate during oxidative stress (Gueta et al., [2020](#page-10-13)) due to aberrant cancer cell redox homeostasis. Elevated N6, N6, N6-Trimethyl-L-lysine strongly associated with death (HR 18.21, *p*=2.59E−04), highlighting its known role in cancer progression (Maas et al., [2020\)](#page-11-16). Pyruvate was associated with decreased PFS (HR 3.47,  $p = 0.0468$ ), likely due to its promotion of angiogenesis (Jung et al., [2011\)](#page-10-14), tumor proliferation (Sellers et al., [2015](#page-11-17)), and immune downstaging (Abusalamah et al., [2020](#page-10-15)). Itaconic acid was also associated with decreased PFS (HR  $2.95, p=0.0257$ .

Higher abundance of cystine, composed of two cysteine molecules, was strongly associated with OS (HR 0.21, *p*=1.89E−04) (Table [2\)](#page-5-0). Cysteine is critical for cancer metabolic remodeling (Serpa, [2020\)](#page-11-18), and abundance potentially due to reduced uptake may indicate decreased tumor metabolic dysregulation—consistent with previous fndings that enhanced cystine processing is associated with more aggressive cancers (Ji et al., [2018\)](#page-10-16). Similarly, elevated histamine (Medina & Rivera, [2010\)](#page-11-19) was also associated with improved survival (PFS) probability (Fig. [2](#page-6-0)B), On the other hand, elevated acetyl-l-carnitine was associated with death (HR 4.73,  $p = 0.0175$ ) (Table [2](#page-5-0)), consistent with its role in promoting angiogenesis (Cooke & Ghebremariam, [2008](#page-10-17)) and cancer cell proliferation (Yu et al., [2017\)](#page-11-20). Increased creatine was further associated with death (HR  $44.57$ ,  $p = 0.0020$ ) (Table [2\)](#page-5-0). Creatine can be converted into phosphocreatine, which acts as an energy reservoir and is involved in the production of ATP from ADP (Loo et al., [2015\)](#page-10-18). Increased creatine in tumor tissue may lead to an increase in ATP production, which is associated with high cancer metabolic activity (Puchades-Carrasco et al., [2016\)](#page-11-1). In another study, creatine was signifcantly increased in lung cancer tissue compared to normal tissue, measured by NMR (Wald et al., [1997\)](#page-11-21). In contrast, although several benzoic acid derivatives have shown anti-tumor activity (Anantharaju et al., [2017](#page-10-19); Girouard et al., [2020](#page-10-20); Yue et al., [2016](#page-11-22); Zhang et al., [2021](#page-11-23)), little is known about how benzoic acid is involved in lung cancer metabolism. Previously, increased benzoic acid was found in late stage lung cancer patients relative to early stage and healthy controls (Hori et al., [2011](#page-10-5)), consistent with our fndings that increased benzoic acid levels were associated with progression (HR  $6.17$ ,  $p = 0.0190$ ) and death (HR 5.55, *p*=0.0257) (Table [2\)](#page-5-0).

PLS-DA component analysis was performed as an alternative to the metabolite based risk score calculated in Eq. [1,](#page-3-0) where the PC1 and PC2 scores were used for OS and PFS, respectively. The PLS-DA component was inefective for separating low risk and high risk patients based on PFS in Kaplan–Meier analysis  $(p=0.165)$  (data not shown). To increase separation, follow-up time was restricted to 0–48 months ( $p = 0.0615$ , Supplementary Figs. 4 and 5). Although PLS-DA component analysis resulted in signifcant Cox proportional hazards regression models (Supplementary Figs. 4D, 5D), they were not as signifcant as the metabolite-based risk scores (Figs. [3](#page-7-0)D, [4D](#page-8-0)). Further, to ensure that missing value imputation was not contributing to the survival analysis, the metabolomic data was onehot-encoded as detected (0) vs. not-detected (1) and a Cox proportional hazards model was performed (Supplementary Table 3). None of the metabolites were signifcant, while only one and three diferent metabolites were respectively marginally significant  $(0.05 \le p \le 0.1)$  for OS and PFS, indicating that the presence or absence of metabolites in patient tissue samples was not afected by the events of disease progression or death.

This study includes a small and heterogeneous patient population, which limits the interpretation of results. Potential confounding factors include type and duration of chemotherapy/ immunotherapy cycles, smoking status, and diferences in biopsy collection date relative to date of diagnosis. However, it was previously reported that adjusting for confounding factors such as age and cancer stage at diagnosis, smoking status, hypertension, diabetes mellitus, BMI, HDL/LDL levels, and time between sample collection and diagnosis had minimal signifcant efect on metabolite hazard ratios associated with prostate cancer-specifc mortality (Huang et al., [2019\)](#page-10-8). Our study along with (Huang et al., [2019](#page-10-8)) found that metabolomic profles were able to uncover biomarkers relevant to cancer survival prediction without adjusting for these potential confounding factors, emphasizing the predictive capability of the metabolomics data. Additionally, it was previously found that number of chemotherapy cycles was a weak predictor of survival in epithelial ovarian cancer, being overpowered by pathological stage and a metabolite-based risk score (Xie et al., [2017](#page-11-3)). Further limitations of this study are that the data on association of metabolic score and OS or PFS would have to be validated in a larger and independent cohort. Considerations for future studies include integrating machine learning algorithms into a more comprehensive multivariable predictive modeling framework for high-accuracy predictions of patients in survival groups and simultaneous determination of a reliable set of key metabolites as biomarkers.

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**Data availability** Datasets used are in [Metabolomics Workbench (MetWB), RRID:SCR\_013794, [www.metabolomicsworkbench.org](http://www.metabolomicsworkbench.org)] (Study ST001527).

#### **Declarations**

**Conflict of interest** The authors declare that they have no competing interests.

**Informed consent** Informed consent was obtained from all individual participants included in the study.

**Research involving human and animals participants** This study was performed in line with the principles of the Declaration of Helsinki. Approval was granted by Internal Review Board protocols at University of Louisville Hospital (IRB 05.0523) and Norton Hospital (IRB 18.0264).

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