ORIGINAL ARTICLE

Genetic parameters in subtropical pine F_1 hybrids: heritabilities, between-trait correlations and genotype-by-environment interactions

Pomerayi Mutete¹ · Ruramai Murepa¹ · Washington J Gapare²

Received: 29 April 2015 /Revised: 25 August 2015 /Accepted: 31 August 2015 / Published online: 8 September 2015 \oslash Springer-Verlag Berlin Heidelberg 2015

Abstract Growth and stem straightness traits of 29 Pinus caribaea var. hondurensis×Pinus tecunumanii (PCH×PTEC) and 26 P. caribaea var. hondurensis×Pinus oocarpa (PCH× POOC) hybrid pair-crosses plus a total of 16 intraspecific families were assessed at ages 5, 8 and 15 years from planting at two sites. The PCH×PTEC hybrid was the most productive, yielding 37 % more than a Pinus elliottii local control and was 21 % superior to either parental species in DBH growth. PCH×POOC hybrid was, on average, 16 % superior to either parental species for DBH. Narrow-sense heritability estimates were low to moderate for growth traits (average of 0.27) and stem straightness (0.16). The estimated additive genetic correlations between growth traits and ages within traits were high (>0.8) and positive, providing confidence in early selection based on diameter at breast height. The high proportion of estimated additive genetic variance compared to dominance variance in the F_1 pine hybrids suggests that breeding strategies that maximize the use of additive genetic variance may be effective. The ranking of the 11 PCH parents based on general hybridizing ability predictions (estimated breeding values as hybrids) was somewhat inconsistent between PTEC and POOC hybrid crosses for all traits ($r_{9 \text{ d.f.}}$ =0.38–0.45; p ∼0.15–0.25). There was no evidence of practically important

This article is part of the Topical Collection on Breeding

 \boxtimes Washington J Gapare Washington.Gapare@csiro.au $G \times E$ interaction for the hybrids except for PCH \times PTEC height growth. This study suggests that a single, multihybrid breeding population seems appropriate in Zimbabwe if the trial sites are representative of the planting target zone.

Keywords Tree breeding \cdot Heritability \cdot Pines \cdot F₁ hybrids \cdot Genetic correlation \cdot Genotype by environment interaction \cdot General hybridizing ability (GHA)

Introduction

Zimbabwe is divided into five Provisional Silvicultural Zones (PSZ) (Fig. [1\)](#page-1-0) (Barrett and Mullin [1968](#page-13-0)). Commercial pine and eucalypt forestry are concentrated in PSZ I to III which experience a subtropical to temperate climate due to the modifying effect of altitude with pronounced dry and wet seasons. Predominant pine species include Pinus patula Schiede ex Schlechtendal Chamisso, Pinus taeda L., Pinus elliottii Engelm. and P. kesiya Royle ex Gordon, in total covering 88,000 ha (Timber Producers Federation [2011\)](#page-15-0). Predictions are that, by 2030, the total planted area will reach approximately 120, 000 ha (Arnold and White [1994;](#page-13-0) Timber Producers Federation [2014](#page-15-0)). Of this area, 80 % or 96,000 ha will be planted to pines (Timber Producers Federation [2014\)](#page-15-0). This area is predicted to span across PSZ III to IV. Species trials established on PSZ III and IV indicated that the current commercial pine species were not sufficiently productive (Barnes [1981](#page-13-0); Crockford [1995](#page-14-0)). For example, mean annual increments for P. elliottii in PSZ III and IV are usually around 16 m^3 ha⁻¹ year⁻¹ compared to 26 m^3 ha⁻¹ year⁻¹ in PSZ I and II (Crockford [1995;](#page-14-0) Timber Producers Federation [1999;](#page-15-0) Gotore et al. [2014](#page-14-0)). P. patula and P. taeda are not tolerant of moisture deficits that typify PSZ III and IV while *P. elliottii*, though tolerant, is slow in capturing site and invariably low in volume production

¹ Forest Research Centre, Forestry Commission, P. O. Box HG 595, Highlands, Harare, Zimbabwe

² Commonwealth Scientific and Industrial Research Organisation, GPO Box 1600, Canberra ACT 2601, Australia

Fig 1 Map of Zimbabwe showing Provisional Silvicultural Zones. Modified from Barrett and Mullin [\(1968\)](#page-13-0)

compared to other species (Barnes [1989;](#page-13-0) Mullin [1992\)](#page-14-0). The major limiting factors in PSZ III and IV include low rainfall and high temperature which increases evapotranspiration. It was perceived that some of the limitations of the currently planted pure pine species in Zimbabwe could be overcome by more species introduction and breeding and also through their inter-specific hybridization (Barnes [1989;](#page-13-0) Barnes et al. [1997](#page-13-0); Nyoka [2000\)](#page-14-0).

The breeding programs of P. patula, P. taeda L. and P. elliottii are almost three generations advanced from their wild base populations, and much has been learnt about their genetics (Crockford et al. [1988](#page-14-0); Barnes et al. [1992a,](#page-13-0) [b](#page-13-0); Nyoka and Barnes [1995](#page-14-0); Pswarayi et al. [1996](#page-15-0); Gapare and Musokonyi [2002](#page-14-0); Nyoka et al. [2010\)](#page-14-0). In the early 1990s, new germplasm including several provenances of P. patula that had not been previously included in the breeding program

and also some hitherto untested species (Pinus maximinoi H.E., Pinus tecunumanii Eguiluz & Perry and P. greggii Engelm. ex Parl.) were included in the testing program (Barnes [1981](#page-13-0); [1989;](#page-13-0) [1993\)](#page-13-0). The motivation was to reduce the monoculture of *P. patula* in high-altitude areas as well as reduce the dependence on *P. elliottii* in the low-altitude areas. Introduction of P. maximinoi, P. tecunumanii and Pinus greggii was also perceived to be a long-term solution to finding suitable species for marginal environments (Barnes [1989\)](#page-13-0). For example, P. tecunumanii has demonstrated high growth rates and is tolerant to drought (Nyoka and Barnes [1995;](#page-14-0) Nyoka et al. [2010\)](#page-14-0). The planting of seed from the top 10 families of the low-elevation *P. tecunumanii* gave gains in 8year volume of respectively 13 and 23 % over P. patula at Stapleford and Cashel sites, though some issues of susceptibility to wind damage were also noted (Nyoka et al. [2010\)](#page-14-0).

In order to meet future wood requirements, there has been a steady expansion of commercial pine forestry into marginal areas for most of the tested species. This has led to the development of options for sustainable improvement of forest plantation productivity including research into performance of hybrids. Earlier work on pine hybrids in the 1970s involved P. elliottii and P. taeda (Barnes and Mullin [1978](#page-13-0)). The hybrid was developed because the two species had complementary traits and were likely to cross easily given that they were closely related (Barnes and Mullin [1978](#page-13-0)). Results from the hybrid trials were not encouraging and no investment in hybrid testing was made until 1993 when interspecific hybrids involving Pinus caribaea var. hondurensis, P. elliottii, Pinus oocarpa and P. tecunumanii were tested on two sites. Gwaze ([1999](#page-14-0)) reported that hybrid vigour (heterosis) in all the traits was exhibited in all the hybrids at the two sites, being more clearly expressed at the lowelevation drier site. The volume production of the hybrid between P. caribaea and P. tecunumanii was more than four times that of the commonly grown *P. elliottii* and 52 % more than the best performing pure species (P. tecunumanii) at the wetter site (Gwaze [1999\)](#page-14-0).

By way of precedent, experience with the P . elliottii \times P. caribaea hybrid in Queensland, Australia, where it was planted on almost 70,000 ha, demonstrated that it exceeded the productivity of P. elliottii (Dieters [1999;](#page-14-0) Dieters and Brawner [2007](#page-14-0)). The hybrid appears to inherit the high growth rate from *P. caribaea* and stem straightness, wind firmness, high wood density and adaptability to wet sites from P. elliottii (Nikles [2000](#page-14-0); Dieters and Brawner [2007](#page-14-0)). There is also experience of P. elliottii×P. caribaea hybrid in Argentina where it has shown superior performance compared to the pure spe-cies in field trials (Cappa et al. [2013](#page-14-0)). The *P. elliottii* \times P. caribaea hybrid is planted commercially in Argentina, Australia and South Africa. The use of P. caribaea var. hondurensis in combination with other species such as P. tecunumanii, P. elliottii and P. oocarpa is therefore predicted to give hybrids combining high productivity, adaptation and stem strength (e.g. Dieters et al. [1997\)](#page-14-0).

However, several reviews have noted some of the challenges associated with hybrids. As highlighted in a review by Dungey [\(2001\)](#page-14-0), some challenges include selection of parental species, early crossing and multiplication options. Other notable challenges include hybrid invariability, high development costs when compared with pure species strategies (for fixed resources and hybrids requiring more complex breeding strategies; e.g. Dungey [2001;](#page-14-0) Kerr et al. [2004a](#page-14-0), [b](#page-14-0)). Obtaining hybrid seed in sufficient quantities for commercial forestry can be extremely difficult and expensive (Gwaze [1999\)](#page-14-0).

The high cost of hybrid breeding relative to pure species breeding creates an imperative to identify the most efficient breeding strategy. Proposed and commercially used strategies for breeding pine hybrids have been reviewed in detail (Dungey et al. [1999;](#page-14-0) Shelbourne [2000;](#page-15-0) Kerr et al. [2004a](#page-14-0), [b\)](#page-14-0). For example, Dungey et al. [\(1999\)](#page-14-0) concluded that most strategies are either an adaptation of the original reciprocal recurrent selection (RRS) strategy outlined by Comstock et al. [\(1949\)](#page-14-0) or use recurrent selection for general combining ability (GCA) in the parent species. Kerr et al. ([2004a,](#page-14-0) [b\)](#page-14-0) developed and simulated synthetic species (SYN) and pure species selection (PSS) strategies. Their conclusion was that SYN strategy was the most cost-effective across a wider range of genetic structures, in particular where there is less dominance variance and the pure hybrid correlations in both species are positive (Kerr et al. [2004b\)](#page-14-0).

This paper combines age-5 data reported by Gwaze [\(1999](#page-14-0)) with growth and stem straightness measures at ages 8 and 15 years from planting to compare the productivity trends of hybrids against their pure parental species for completeness. The specific objectives of this study were to (1) estimate genetic parameters, additive and dominance variances and heritability for height, diameter at breast height (DBH) and stem straightness, (2) estimate genetic correlations between studied traits and (3) examine genotype by environment $(G \times E)$ interaction of the same traits. We also used the data to study how well the breeding values of P. caribaea var. hondurensis parents corresponded between the two interspecific hybrid combinations (e.g. Dieters et al. [1997](#page-14-0)). Additionally, we use this information to discuss the implications and selection strategies for genetic improvement of pine hybrids in Zimbabwe in order to increase the profitability of future softwood plantations in Zimbabwe.

Materials and methods

Genetic material and genetic tests

The control-pollinated hybrid families were provided to the Zimbabwe Forestry Commission by the then Queensland Forest Research Institute, Australia, through the then Oxford Forestry Institute, UK in 1992. The families originated from an incomplete factorial design of unrelated first- and secondgeneration parents (Dieters et al. [1997\)](#page-14-0). Eleven unrelated first- and second-generation P. caribaea var. hondurensis (PCH) parents were crossed with six first-generation P. oocarpa (POOC) parents and six first-generation P. tecunumanii (PTEC) parents to form two 11×6 factorial arrays. The PCH parents were used as the female parents in all factorial crosses, and there were three POOC parents from each of Zapotillo and Angeles provenances and three PTEC parents from each of Mountain Pine Ridge and Yucul provenances (Dieters et al. [1997\)](#page-14-0). The parents represented in the genetic crosses were both very few in number and were generally not represented in interspecific crosses precludes any firm conclusions about heterosis.

The two factorials were almost complete, with 61 of the possible 66 F_1 families produced in each factorial. However, not all pair-crosses were made available for the Zimbabwe tests. The actual number of hybrid families included in the tests in Zimbabwe is provided in Table 1. These included several controls (Table 1). Controls of the pure species were included in the tests, but they were generally unrelated to the hybrids. The *P. elliottii* control from Zimbabwe was a full-sib cross between parents of outstanding growth. Five families of P. caribaea var. hondurensis (PCH)×P. elliottii hybrid crosses were also included in the tests as part of the controls. Seedlings were raised at John Meikle Forest Research Station, and the tests were established at two contrasting sites, Cashel and Mukandi in 1993 (Table 2).

Field design at Mukandi was an incomplete block design with 6 replicates and 16 blocks. Each family was planted in five-tree row plots and the spacing between trees was 3×3 m. At Cashel, the design was randomised complete block with six replicates. Spacing and plot size were as at Mukandi.

For this set of trials, survival was assessed at 5, 8 and 15 years and expressed as a percentage of the total number of planted trees for each taxon. Productivity traits were assessed at the same ages: Height was measured on all live trees, denoted as HT5, and HT8 and HT15; tree diameter was measured at breast height (1.3 m above ground level) over bark, denoted as DBH5, DBH8 and DBH15; and stem straightness was assessed using a 7-point absolute visual scale $(1=$ crooked to $7=$ very straight), denoted as STR5, STR8 and STR15. Both trials were thinned to 50 % of initial stocking prior to the age 15 years assessments, leaving too few trees for precise genetic parameter estimation: We therefore only estimated trait means for each taxon at each site at age 15 years. Age-5 data were the subject of an earlier publication (Gwaze [1999\)](#page-14-0), and we include these data in our analysis in order to get a comprehensive overview of the performance from juvenile to later ages.

Table 1 Details of the genetic material used (adapted from Gwaze [1999\)](#page-14-0)

Table 2	Details of hybrid field test sites (adapted from Gwaze 1999)					
	Mukandi	Cashel				
Longitude	$32^{\circ} 51' E$	$32^{\rm o}$ 50' E				
Latitude	18^0 41'S	19^{0} 37'S				
Altitude (m)	1300	1525				
Mean annual rainfall (mm)	1711	745				
Soil parental material	Granite	Shale				

Statistical models and analyses

Comparison of taxa

Taxa differences were tested against family-within taxon variation, the latter being a pooled variance of family means using ANOVA in ASReml R (Butler et al. [2009;](#page-14-0) R Development Core Team [2011\)](#page-15-0). The significance of differences between taxa at $p<0.05$ was tested using the Bonferroni and Newman-Keuls adjusted t tests (Armitage et al. [2001\)](#page-13-0).

Assumptions of analysis based on conventional quantitative genetic model

Some simplifying assumptions are necessary in order to model genetic architecture of hybrid populations. For example, the genetic loci controlling the traits examined may be assumed to be common to PCH, PTEC, POOC and PCH×PTEC and PCH×POOC hybrids populations, and alleles and genotypes segregate freely and randomly within and between populations. We may also assume epistasis to be negligible. Although the assumption of no epistasis is probably unrealistic, this assumption is necessary in most studies using quantitative genetic approaches due to statistical difficulty of measuring interactions among numerous loci (Kerr et al. [2000\)](#page-14-0). However, simulation studies have demonstrated that when analysis of

FRC Forest Research Centre, OFRI Queensland Forest Research Institute, OFI Oxford Forest Institute (seed from Honduras and Nicaragua), OP open pollinated (approximate half-sib)

^a Number of pair-crosses provided for the tests in Zimbabwe

variance is applied, even where epistatic (additive×additive, additive×dominance, dominance×dominance), effects are present, and fitted into the model, much of the variance due to epistatic effects is in fact partitioned into the main effects (e.g. Cheverud and Routman [1995](#page-14-0)). For the purpose of breeding strategy design, the main distinction is that between additive-related and dominance-related gene effects, and so the formal assumption of no epistasis in our models is considered unlikely to result in erroneous conclusions.

Data analyses

The PCH×PTEC (29) and PCH×POOC (27) hybrid crosses at each site were analysed separately. All pure species and PCH×PELL hybrids were excluded from genetic parameter estimation due to the smaller numbers of families (Table [1\)](#page-3-0), except for taxon performance comparisons. A series of genetic analyses were conducted using ASReml R (Butler et al. [2009](#page-14-0); R Development Core Team [2011\)](#page-15-0). Diagnostic plots were used to verify normal distribution of residuals and identify outliers. Univariate models were first fitted to HT, DBH and STR for 5 and 8 years data from each trial. Survival was reported to be above 95 % for all taxa in Gwaze ([1999\)](#page-14-0), and we did not reanalyse any survival data.

For single-site analyses, we followed the parental model [1] similar to one used by Brawner et al. ([2005](#page-13-0)):

$$
Y_{jklm} = \mu + B_j + F_k + M_l + FM_{kl} + FB_{jk} + MB_{jl}
$$

+
$$
P_{jkl} + E_{jklm}
$$
 (1)

where Y_{iklm} is the *m*th tree of the *kl*th family in the *j*th block, μ is the overall mean, B_i is the fixed effect of the *j*th block, F_k is the random effect of the *k*th female parent, ~ $N(0, \sigma^2 j)$, M_l is the random effect of the *l*th male parent, ~ $N(0, \sigma^2_m)$, FM_{kl} is the random effect of the interaction between the kth female parent and the *l*th male parent, ~ N (0, σ_{fin}^2), FB_{jk} is the random effect the interaction between the *jth* block and the *kth* female parent, ~ N (0, σ^2_{b}), MB_{jl} is the random effect of the interaction between the *j*th block and the *lth* male parent, ~ $N(0, \sigma_{bm}^2)$, P_{jkl} is the random effect of variation between plots, ~ $N(0, \sigma_p^2)$ and E_{iklm} is the random error associated with the *m*th observation of the *kl*th family in the *j*th block∼N (0, σ^2_e).

The pooled site model is the same as that given above but with the inclusion of fixed terms for test and block nested within test, as well as random terms for test×female-parent interaction, test×male-parent interaction and female-parent×male-parent× test interaction (Brawner et al. [2005](#page-13-0)). Across-sites analysis used standardized data, which were transformed by dividing each observation by the square root of the within-test error variance (previously estimated from the single-site analyses for each trait) (Brawner et al. [2005\)](#page-13-0). Results from single-site analyses were used to obtain starting values for the pooled-site analyses. Both heterogeneous dominance and error variances were included in the model (e.g. Costa e Silva et al. [2005](#page-14-0)).

In order to determine how well the breeding values of PCH parents corresponded between the two interspecific hybrid combinations (e.g. Dieters et al. [1997](#page-14-0)), best linear unbiased predictions (BLUPs) were obtained for each of the 11 (female) PCH parents, from the pooled-site analyses, separately for PCH×PTEC and PCH×POOC hybrid crosses using the software ASReml R (Butler et al. [2009;](#page-14-0) R Development Core Team [2011\)](#page-15-0). Dieters et al. ([1997](#page-14-0)) noted that such predictions of the average effects of the female parents are estimates of their general hybridizing abilities (GHA), as defined by Nikles and Newton [\(1991\)](#page-14-0), in contrast with the conventional general combining abilities (GCA).

For noninbred parents, these variance components can be interpreted in the following manner (Cockerham [1963](#page-14-0); Becker [1984\)](#page-13-0): σ_{f}^2 and σ_{m}^2 are estimates of one-quarter of the additive genetic variance, σ^2_A ; σ^2_{fm} is an estimate of one-quarter of the dominance variance, σ_{D}^2 , σ_{fs}^2 and σ_{ms}^2 are estimates of onequarter of the additive–site interaction variance, σ_{AE}^2 and σ^2_{fms} is an estimate of one-quarter of the dominance–site interaction variance, σ_{DS}^2 . The corresponding estimates of σ_{f}^2 σ_{m}^{2} , and σ_{fm}^{2} , from the single-site analyses are upwardly biased due to the confounded effects of genotype–environment interactions (Comstock and Moll [1963\)](#page-14-0). Therefore, σ_{f}^2 , (or σ_{m}^{2}) and σ_{fm}^{2} are biased estimates of one-quarter of the additive and dominance variances, respectively.

We repeated the analysis of these data using an individualtree model. Such a model assumes that the additive variances in the population are the same for the female and male parents. Given that the present-day geographic ranges of these species are often found growing alongside PCH, POOC and PTEC on sites from Belize to Nicaragua (Dvorak et al. [2000a,](#page-14-0) [b](#page-14-0)), it may be a reasonable assumption that these species are closely related and therefore their variances in the segregating population are the same for female and male parents. For example, a phylogenetic study by Dvorak et al. [\(2000b\)](#page-14-0) suggested PCH, PTEC and POOC are closely related. We also assumed that it may not be possible to disentangle their likely differences in the additive variance between the male and female parents due to sampling error.

The model assigns a random effect to the breeding value of each tree, both for trees with records and those that are represented as parents or grandparents in the analysis. Each model incorporated the full pedigree including parents and grandparents by inclusion in the additive genetic relationship matrix for the trees (Gilmour et al. [2009](#page-14-0)). However, the female and male parents of the 11 first- and second-generation selections were assumed unknown and coded zero in the pedigree table. The incorporation of this pedigree is critical as it considers dependences (i.e. genetic relationships) that occur in these reduced hybrid populations, hence improving the estimation of genetic parameters.

The statistical model [2] used in the analysis of individualsite data was as follows:

$$
y_{ijklm} = \mu_{+} B_i + tree_j + fam_k + plot_l + e_{ijklm} \tag{2}
$$

where y_{ijklm} is the individual-tree measurement, μ is the overall mean, B_i is the fixed effect of block, tree_i is the random additive genetic effect of individual tree∼ $N(0, \sigma_A^2)$, fam_k is the random effect of full-sib family∼ $N(0, \sigma_{fam}^2)$, plot_l is the random effect of plot∼ $N(0, \sigma_{\text{plot}}^2)$ and e_{ijklm} is the random residual effect∼ $N(0, \sigma_E^2)$.

The pooled-site model is the same as that given in [2] except that trial is an additional fixed effect, and y_i is now defined as the vector of observations for a single trait indexed (i) by trial.

Estimation of genetic parameters

Parental model

The REML variance component estimates from the parental model were used to estimate heritability (biased $h²_b$, and unbiased, h^2 from single- and pooled-site analyses, respectively. Single-site analyses yield an estimate of heritability which is upwardly biased due to the confounded effects of genotypeby-environment interactions (Comstock and Moll [1963](#page-14-0); Hodge and White [1992](#page-14-0)). Two separate estimates of the heritability can be obtained—one from the female parents and one from the male parents (Dieters et al. [1997\)](#page-14-0). Additional formulae can be obtained by substituting the appropriate male and male×site variance components into the equations below. The formulae used to estimate these genetic parameters are listed below:

$$
\hat{h}^2 b = \frac{4\hat{\sigma}_f^2}{\hat{\sigma}_f^2 + \hat{\sigma}_m^2 + \hat{\sigma}_{fm}^2 + \hat{\sigma}_p^2 + \hat{\sigma}_e^2}
$$
(3)

$$
\hat{h}^{2} = \frac{4\hat{\sigma}_{f}^{2}}{\hat{\sigma}_{f}^{2} + \hat{\sigma}_{m}^{2} + \hat{\sigma}_{fs}^{2} + \hat{\sigma}_{ms}^{2} + \hat{\sigma}_{fm}^{2} + \hat{\sigma}_{fms}^{2} + \hat{\sigma}_{f}^{2} + \hat{\sigma}_{e}^{2}} \tag{4}
$$

Biased dominance as a proportion of phenotypic variance $(d²_b)$ was only estimated from single-site estimates as follows:

$$
\hat{d}^2_{b} = \frac{4\hat{\sigma}_{fm}^2}{\hat{\sigma}_f^2 + \hat{\sigma}_m^2 + \hat{\sigma}_{fm}^2 + \hat{\sigma}_p^2 + \hat{\sigma}_e^2}
$$
(5)

Individual-tree model

Observed variance components were used to estimate the causal variance components for each trait and interpreted as follows: $\hat{\sigma}_A^2$ is estimate of additive genetic variance, $\hat{\sigma}_D^2 = 4$

 $\hat{\sigma}_{fam}^2$ is the estimate of dominance genetic variance, $\hat{\sigma}_G^2 = \hat{\sigma}_A^2$ $+\hat{\sigma}_D^2$ is the estimate of total genetic variance assuming no epistasis. Individual narrow-sense heritabilities (denoted $\hat{h}^2 b$ i to indicate it is from individual tree model) from the single-site analyses (upwardly biased due to the confounded effects of genotype×environment interactions (Comstock and Moll [1963\)](#page-14-0)) were estimated as the additive genetic variation divided by the phenotypic variation $(\sigma^2 p)$:

$$
\hat{h}^2{}_{bi} = \hat{\sigma}_A^2 / \hat{\sigma}_P^2 \tag{6}
$$

where phenotypic variance is estimated as:

$$
\hat{\sigma}_P^2 = \hat{\sigma}_A^2 + \hat{\sigma}_{fam}^2 + \hat{\sigma}_{plot}^2 + \hat{\sigma}_E^2 \tag{7}
$$

Pooled-site analysis was conducted for each trait in order to calculate unbiased heritability estimates. In that case, an unbiased estimate of narrow-sense heritability was estimated as:

$$
\hat{h}^2_{\ i} = \frac{\hat{\sigma}_A^2}{\hat{\sigma}_P^2} \tag{8}
$$

where phenotypic variance from pooled-site analyses is estimated as:

$$
\hat{\sigma}_P^2 = \hat{\sigma}_A^2 + \hat{\sigma}_{S_A}^2 + \hat{\sigma}_{Sfam}^2 + \hat{\sigma}_{plot}^2 + \hat{\sigma}_E^2 \tag{9}
$$

 σ_A^2 , σ_{plot}^2 and σ_E^2 are as defined in Eq 2, σ_{SA}^2 and σ_{Sfam}^2 are the interaction variances between site and additive genetic effects and site×full-sib family effects, respectively.

Approximate standard errors of the heritability estimates were derived based on Taylor series approximation using the R pin function (White [2013](#page-15-0)).

The additive genetic correlation estimates between traits 1 and 2 were obtained from the estimated additive covariance and variance components from the individual-tree model as:

$$
r_A = \frac{\hat{\sigma}_{A_1 A_2}}{\sqrt{\hat{\sigma}_{A_1}^2 \hat{\sigma}_{A_2}^2}}
$$
(10)

where:

 $\sigma_{A_1A_2}$ = additive genetic covariance component between trait 1 and trait 2

 $\sigma_{A_1}^2$ = additive genetic variance for trait 1 at each site

 $\sigma_{A_2}^2$ = additive genetic variance for trait 2 at each site

For female, male, dominance (from model 1) and additive genetic effects (model 2), the first indication of their significance was given by the ratio of the variance components to their corresponding standard error. Terms for which this ratio was >2 were regarded as significant. Terms for which the ratio was <1 were regarded as not significant. For ratios between 1

and 2, the likelihood ratio (LR) test was applied (−2×(difference between log likelihoods including and excluding the term)∼ χ^2 ; Gilmour et al. [2009;](#page-14-0) Stram and Lee [1994](#page-15-0)).

In order to determine the extent of genotype×environment interaction for each of the traits, univariate, paired-pooled analysis was conducted. Heterogeneous error terms were fitted for each site for each trait. Type B additive genetic correlation estimates (r_B) were then made following Burdon [\(1977\)](#page-13-0) and higher values (>∼0.8) indicate little genotype× environment interaction, and lower values indicate that practically important genotype-× -environment interactions exist (Robertson [1959](#page-15-0)). One-tailed LR tests were used to test an estimated type B additive genetic correlation against +1. This was done by using a parameterisation of (co)variance matrix based on a correlation form and constraining the correlation parameter to be $+1$ under the null hypothesis to be tested. If $log L_1$ and $log L_2$ are the REML log-likelihoods from the unrestricted and the restricted $(r_B=1)$ models, respectively, the test statistic (D) is given by:

$$
D = 2(\log L_1 - \log L_2) \tag{11}
$$

which is distributed approximately as χ^2 under H_0 , with degrees of freedom given by the difference between the number of parameters estimated under the non-restricted and the restricted models (Costa e Silva et al. [2005;](#page-14-0) Gilmour et al. [2009\)](#page-14-0).

Results

Estimates of trait means by taxon

Phenotypic means and standard errors for all traits observed at each site are presented in Table 3. At both sites, the hybrids outperformed the pure species, with the hybrids between P. caribaea var. hondurensis and P. tecunumanii (PCH×PTEC) being the most productive in terms of height and DBH at all ages. Generally, the hybrids were significantly different from the pure species (p <0.05) for growth traits at all ages (Table 3). Growth (height and DBH) was better at Mukandi than at Cashel for all taxa at age 5 years, but reversed at ages 8 and 15 years. The PCH×PTEC hybrid was the most productive for height and DBH at age 8 years at both sites. The PCH×PTEC hybrid continued to be the most productive taxon, being on average 10 % better than either parental species. The hybrids had better stem straightness than pure species at Mukandi, whereas at Cashel, the pure species had as straight stems as the hybrids. Similar trends were evident at ages 8 and 15 years.

Growth performance of pure species and hybrids

Percent superiority of the hybrids over *P. elliottii* local control (PEE1) at 8 and 15 years at Cashel and Mukandi are shown in

Table 3 Means for growth and form traits for various taxa

Taxon		Cashel Mukandi Cashel Mukandi Cashel Mukandi				
	HT5(m) $DBH5$ (cm)			$STR5(1-7)$		
PEE1	5.2f	6.5e	8.9d	9.7d	3.5ab	3.5cd
PEE ₂	5.5ef	7.2e	9.8cd	10.7d	3.8a	3.8 _{bc}
POOC	6.6cd	8.7d	10.5c	13.7c	3.0c	3.4d
PCH	5.9e	8.8cd	10.0c	13.8c	2.9c	3.6cd
PTEC	6.5d	8.9cd	10.2c	13.7c	3.0c	3.6cd
PCH×PTEC	7.7a	10.0a	12.9a	16.7a	3.4 _b	4.1 _b
PCH×POOC	6.9 _{bc}	9.5 _b	11.3 _b	15.4 _b	3.4 _b	4.1 _b
$PEE \times PCH$	7.2b	9.3 _{bc}	13.3a	15.7 _b	3.7a	4.2a
	HT8(m)		$DBH8$ (cm)		$STR8(1-7)$	
PEE1	13.3d	10.4c	19.2d	13.6d	3.7ab	4.6abc
PEE ₂	14.6cd 11.5c		21.4cd	14.9d	3.9a	4.8ab
POOC	15.2c	13.2 _b	22.6c	19.5c	2.8c	3.4d
PCH	15.7c 13.9b		22.3c	19.4c	3.2bc	4.2d
PTEC	15.2c	13.6b	22.1c	19.7c	2.9c	3.5d
PCH×PTEC	18.0 a	15.6a	27.5a	23.8a	3.4 _b	4.1c
PCH×POOC	17.3 _b	15.4a	25.5 _b	22.5 _b	3.8a	4.5 _{bc}
$PEE \times PCH$	17.0 _b	14.9a	26.8ab	22.8 _b	4.1a	5.2a
	HT15(m)		$DBH15$ (cm)		$STR15(1-7)$	
PEE1		19.7d 17.4e	23.5e	18.8d	4.3 _{bc}	4.6ab
PEE ₂		21.6cd 19.2e	26.8de	22.6d	5.0a	6.2a
POOC	22.8c	21.5d	28.5cde 30.6c		3.5c	4.8d
PCH	22.9c	21.9cd	29.4cd	31.0c	4.0 _{bc}	5.3bc
PTEC		23.7bc 22.0bcd	33.5bc	33.1bc	3.7c	4.8cd
PCH×PTEC	25.8a	23.6a	37.3a	37.1a	4.3 _b	5.7 _b
PCH×POOC	25.4a	23.3ab	35.6b	34.7b	4.8a	6.2a
$PEE \times PCH$		24.8ab 22.8abc	34.8b	34.6b	5.1a	6.6a

The same letter within a subcolumn for each trait indicates significant difference between means at $p < 0.05$ using the Bonferroni and Newman-Keuls adjusted t tests

PEE1 Pinus elliottii control – Zimbabwe select, PEE2 P. elliottii – Australia select, POOC P. oocarpa, PCH P. caribaea var. hondurensis, PTEC P. tecunumanii, $PCH \times PTEC$ P. caribaea var. hondurensis \times P. tecunumanii, PCH×POOC P. caribaea var. hondurensis×P. oocarpa, PEE×PCH P. elliottii×P. caribaea var. hondurensis

Table [4.](#page-7-0) Generally, the hybrids were superior to the local control at age 8 and 15 years from planting, except for PCH× POOC (P. caribaea var. hondurensis \times P. oocarpa) which was inferior for height by almost 14 % at Cashel and Mukandi at age 8 years. Likewise, stem straightness for the hybrids at age 8 years were inferior to the P. elliottii local control. All three hybrids were superior to parental species and mid parent values for all traits at both ages and sites (Tables [5](#page-7-0), [6](#page-7-0) and [7\)](#page-7-0). For example, PCH×PTEC (P. caribaea var. hondurensis× P. tecunumanii) hybrid was on average, 21 % superior to either parental species in DBH growth. PCH×POOC hybrid was on average, 16 % superior to either parental species and mid parent values for DBH. The PEE×PCH (P. elliottii×P. caribaea var. hondurensis) hybrid was on average, 30 % superior to either

Table 4 Superiority $(\%)$ of the hybrids over the local control PEE1 at 8 and 15 years

Taxon				Cashel Mukandi Cashel Mukandi Cashel Mukandi		
	HT ₈		DBH ₈		STR8	
$PCH \times PTEC$ 26.1		33.3	30.2	42.9	-8.8	-12.2
$PCH \times POOC$ -13.9		-13.0	7.8	16.4	-13.6	-2.2
$PEE \times PCH$	21.8	30.2	28.4	40.3	9.8	11.5
	HT15		DBH ₁₅		STR ₁₅	
$PCH \times PTEC$ 23.6		26.3	37.0	49.3	0.0	19.3
$PCH \times POOC$ 22.4		25.3	34.0	45.8	10.4	25.8
$PEE \times PCH$	20.6	23.7	32.5	45.7	15.7	30.3

POOC P. oocarpa, PCH P. caribaea var. hondurensis, PTEC P. tecunumanii; PCH×PTEC P. caribaea var. hondurensis×P. tecunumanii, PCH×POOC P. caribaea var. hondurensis×P. oocarpa, PEE×PCH P. elliottii×P. caribaea var. hondurensis

parental species and mid parent values for DBH at age 8 years and maintained that superiority to age 15 years.

Heritabilities and genetic correlations between traits, ages and sites

Biased narrow-sense heritability estimates for each of the traits at ages 5 and 8 years from both parental and individual tree models are presented in Table [8](#page-8-0). Female heritability estimates for growth traits among hybrids were generally significant $(p<0.05)$, with a few exceptions. For example, significant \hat{h}^2b was observed for HT5 and HT8 for both hybrids at Cashel but PCH×PTEC hybrid was not significant for HT5 and HT8 at Mukandi. Male heritability estimates for all traits were generally insignificant $(p>0.05)$, and in some cases, the estimates were 0 and where heritability estimate was greater than 0, the standard errors were larger than the estimate (Table [8](#page-8-0)). Estimates of female and male heritability for stem straightness (STR5 and STR8) for both hybrids were insignificant $(p>0.05)$. Generally, \hat{h}^2 bi for height and DBH were low to moderate, and significant

Table 5 Superiority (%) of the PCH x PTEC (P. caribaea var. hondurensis \times P. tecunumanii) over pure parent species and mid-parent at 8 and 15 years

Taxon				Cashel Mukandi Cashel Mukandi Cashel Mukandi		
	HT ₈		DBH ₈		STR ₈	
PCH	14.6	12.2	23.3	22.7	6.2	2.4
PTEC	18.4	14.7	24.4	20.8	17.2	17.1
Mid-parent 16.1		13.0	23.9	21.7	9.7	5.1
	HT15		DBH ₁₅		STR ₁₅	
PCH	12.7	7.8	26.9	19.7	7.5	7.5
PTEC	8.9	7.3	11.3	12.1	16.2	18.8
Mid-parent 10.7		7.3	18.4	15.6	10.3	11.8

PCH P. caribaea var. hondurensis, PTEC P. tecunumanii

Table 6 Superiority (%) of the PCH x POOC (P. caribaea var. hondurensis \times P. oocarpa) over pure parent species and mid-parent at 8 and 15 years

POOC P. oocarpa, PCH P. caribaea var. hondurensis

 \hat{h}^2 bi was observed for these traits at both sites, except for stem straightness (STR5 and STR8) and HT5 for PCH×PTEC hybrid. For example, $\hat{h}^2 bi$ for PCH×PTEC hybrid for STR5 were not significant. Significant values of \hat{h}^2bi for height (HT5 and HT8) ranged from 0.14 to 0.51. $\hat{h}^2 bi$ for diameter at breast height (DBH5, DBH8) ranged from 0.15 to 0.48, with higher estimates observed at Cashel than Mukandi. For example, significant \hat{h}^2 bi for DBH5 at Cashel was almost double that at Mukandi for PCH×PTEC hybrid. In cases where female heritability was significant for a trait, so was the estimate from the individual tree model. For example, $(h²bi)$ from the individualtree model for growth traits were significant $(p<0.05)$ in most cases where $\hat{h}^2 b$ were also significant.

Pooled-site estimates of narrow-sense heritabilities for each of the traits at ages 5 and 8 years from the two models are shown in Table [9.](#page-9-0) Pooled-site heritability estimates were generally intermediate between the two heritability estimates from the respective single-site analyses. For example, significant pooled-site \hat{h}^2 was observed for HT5 and HT8 for PCH× POOC (Table [9](#page-9-0)). Pooled-site analyses showed narrow-sense heritability estimates that ranged from 0.02 to 0.40, with the

Table 7 Superiority (%) of the PEE x PCH (P. elliottii × P. caribaea var. hondurensis) over pure parent species and mid-parent at 8 and 15 years

Baseline				Cashel Mukandi Cashel Mukandi Cashel Mukandi		
	HT ₈		DBH ₈		STR ₈	
PEE	16.4	29.6	25.2	53.0	5.1	8.3
PTEC	8.3	7.2	20.2	17.5	28.1	23.8
Mid-parent	11.8	17.3	22.9	33.3	17.1	15.6
	HT15		DBH15		STR ₁₅	
PEE	14.8	18.8	29.9	53.1	2.0	6.5
PTEC	8.3	4.1	18.4	11.6	27.5	24.5
Mid-parent	11.7	10.7	24.3	29.1	13.3	15.8

PCH P. caribaea var. hondurensis, PTEC P. tecunumanii

Table 8 Estimates of narrow-sense heritabilities \pm approximate standard errors for single-site $(h²b)$ analyses for female and male and from individual tree model (h^2bi) , and dominance as a proportion of

phenotypic variance $(\hat{d}^2 b)$ of *P. caribaea* var. *hondurensis* (PCH) by P. oocarpa (POOC) and P. tecunumanii (PTEC) F_1 hybrids

Significance levels based on one-tailed LR tests that were used to test the departure of female, male and dominance variance from zero. Statistical significance for estimated heritability is same as for additive genetic variance

NE not estimable and assumed to be zero

*Significant (p <0.05), otherwise not significant (p >0.05)

^a Estimated from Eq [3](#page-5-0)

^b Estimated from Eq [6](#page-5-0)

^c Estimated from Eq [5](#page-5-0)

lowest value for HT5 for PCH×PTEC hybrid and the largest

for HT5 and DBH5 for PCH×POOC hybrid (Table [9](#page-9-0)). \hat{d}^2 b was non-significant $(p>0.05)$, and in some cases, the estimates were 0 and where heritability estimate was greater than 0, the standard errors were larger than the estimate (Table 8).

Age-age and trait-trait genetic correlation estimates (r_A) from single-site analyses are presented in Table [10.](#page-9-0) Generally, r_A were statistically significant and followed expectations for all hybrids and traits at both sites. For example, r_A between HT5 and HT8 were significant and averaged 0.96, and r_A between HT5 and DBH5 averaged 0.87. For DBH and STR, r_A was generally significant and ranged from 0.18 to 0.74. This correlation is favourable, indicating that selection for larger diameter would improve stem straightness.

The importance of $G \times E$ interaction was assessed for all traits for the two hybrids through the magnitude of a common estimate of the genetic correlation between the performances of the same trait measured in different trials. The results showing estimated type B additive genetic correlation that are significantly different from +1 based on LR tests are presented in Table [9](#page-9-0). We also used Robertson's [\(1959\)](#page-15-0) threshold of 0.8 to indicate practical significance of G×E interaction, i.e. values below 0.8 are deemed to indicate presence of practically important G×E interaction. For PCH×PTEC hybrid, r_B between sites for height averaged 0.62, suggesting $G \times E$ interaction. For height, r_B between sites averaged 0.90 for PTEC×POOC hybrid. For DBH, r_B between sites for all hybrids were >0.83. Similar trends were also observed for stem straightness, $r_B > 0.83$.

Table 9 Estimates of narrowsense heritabilities±approximate standard errors for pooled-site (h^2) analyses for female and male, and from individual-tree model (h^2) and type B genetic correlations (r_B) of P. caribaea var. hondurensis (PCH) by P. oocarpa (POOC) and P. tecunumanii (PTEC) F₁ hybrids

 \overline{a}

 \overline{a}

Significance levels based on one-tailed LR tests that were used to test the departure of female, male and dominance variance from zero and r_B from +1. Statistical significance for estimated heritability is same as for additive genetic variance

NE not estimable and assumed to be zero

* significant (p <0.05), otherwise not significant (p >0.05)

^a Estimated from Eq [4](#page-5-0)

^b Estimated from Eq [8](#page-5-0)

Table 10 Trait-trait genetic correlations (r_A) for PCH×PTEC and PCH×POOC hybrids

Significance levels were based on two-tailed LR tests to test the departure of r_A from zero ns not significant

 $*_{p<0.05;}$ $*_{p<0.01;}$ $*_{p<0.001}$

Relationships between the general hybridizing abilities (GHA) of 11 PCH parents used in the two sets of hybrids at ages 5 and 8 years from planting were of similar magnitude. Figure 2a–c shows the relationships for traits at age 8 years. There was no consistent pattern in the rankings of the parents used in the two sets of hybrids for all traits. The correlation coefficients between the two sets of breeding value predictions (breeding value= $2 \times$ GHA) were 0.45, 0.38 and 0.40 for HT8, DBH8 and STR8, respectively (Fig. 2a–c). Although the correlations between the GHAs for all traits were positive, they were not significant at the $p=0.05$ level (Steel and Torrie [1980,](#page-15-0) p. 597).

Fig. 2 a–c Relationships between the general hybridizing abilities (GHA) of 11 P. caribaea var. hondurensis (PCH) parents that were predicted from age-8 data of F_1 hybrid crosses with six P. oocarpa (POOC) and six P. tecunumanii (PTEC) across two sites for height (HT8), diameter at breast height (DBH8), and straightness (STR8), respectively. The significance of the r value (correlation between the two sets of GHA predictions) was not significant at $p=0.05$ (Steel and Torrie [1980,](#page-15-0) page 597) However, there were four parents that consistently ranked above average for HT8 and DBH8 when crossed with P. oocarpa. Only three of those parents ranked above average for STR8. The four PCH parents were among the second-generation parents (Dominic Kain – personal communication, 2014).

Discussion

The results from this study explore several aspects of the genetic architecture of the tested pine hybrids that are relevant to

define breeding strategies. P. elliottii and P. tecunumanii are some of the currently recommended commercial species for areas represented by Cashel site (Silvicultural Zones III and IV) while *P. elliottii* is also one of the recommended species for areas represented by Mukandi site. This study has shown that the hybrids had better growth performance than the pure species at both sites, with pure species being more productive at Mukandi than Cashel due to better site conditions (higher rainfall, more fertile soils that are deep and acidic) at Mukandi. Soils at Cashel are derived from sedimentary rocks with varying proportions of other minerals, mainly shale and have moderately shallow soils compared to Mukandi. PCH×PTEC (P. caribaea var. hondurensis×P. tecunumanii) hybrid was the most productive hybrid. For example, PCH×PTEC hybrid was 37 % better than *P. elliottii* local control, the currently recommended species at Cashel and Mukandi sites. Likewise, PEE×PCH (P. elliottii×P. caribaea var. hondurensis) hybrid was almost 35 % superior to P. elliottii control. Barnes and Mullin ([1978](#page-13-0)) found that the hybrid between P. elliottii and P. taeda outperformed the pure species on sites marginal for the pure species, but not on sites optimal for the pure species. P. elliottii is known to struggle at low elevation, drier sites like Cashel (Table [2\)](#page-3-0), and therefore, PCH×PTEC and PEE×PCH hybrids are obvious candidates for such areas covering Silvicultural Zones III and IV. These hybrids could also replace P. patula and P. taeda in these zones because they are not tolerant of moisture deficits that typify these two zones.

P. tecunumanii has demonstrated high growth rates and tolerance to drought and could be developed into a major commercial species in Zimbabwe (Nyoka et al. [1996](#page-14-0); Tembani et al. [2014](#page-15-0)). Its drawback is susceptibility to stem breakage, even in light winds (Dvorak et al. [1993;](#page-14-0) Nyoka and Barnes [1995](#page-14-0); Nyoka et al. [2010\)](#page-14-0). Growth rates for PCH in Zimbabwe are unimpressive when compared to other unimproved pine species like P oocarpa or P. tecunumanii (Gapare and Musokonyi [2002](#page-14-0)). P. tecunumanii used in combination with P. elliottii, P. oocarpa and P. caribaea var. hondurensis has shown that it could provide hybrids combining high productivity, adaptation and strong stems (Tables [5,](#page-7-0) [6](#page-7-0) and [7\)](#page-7-0). For example, PCH×PTEC hybrid was 21 % superior to either parental species and mid parent values for DBH and showed same superiority at age 15 years. P. caribaea var. hondurensis provides resistance to stem breakage (Gwaze [1999\)](#page-14-0). One obstacle to commercial development of this hybrid is that P. caribaea var. hondurensis does not flower at high altitudes in Zimbabwe (Gapare and Musokonyi [2002\)](#page-14-0). Gwaze [\(1999\)](#page-14-0) suggested that importing seed from Australia or improving flowering of the species by planting at low altitudes are potential solutions.

The results of the comparison of pure species and hybrids should be treated as indicative only. In this study, the presumed heterosis was present in all hybrids at both sites and was expressed more at Cashel (marginal site) than the wetter site (Mukandi), where strong dominance effects were exhibited at both ages (Gwaze [1999\)](#page-14-0). The growth results at ages 8 and 15 years confirm to some extent the results of Gwaze [\(1999\)](#page-14-0) that there is potential gain for hybrid species. Heterosis reported here may be inflated, particularly that of the hybrid between PCH×PTEC, because one of the pure parental species, P. tecunumanii, was from natural stands. Brawner et al. [\(2005\)](#page-13-0) reported that PCH×PTEC hybrids showed evidence of hybrid superiority for growth at two locations in Queensland, Australia. For example, PCH×PTEC grew well at the both locations with an average increase in diameter at age 10 years of 14 and 11.5 % over PCH and PTEC, respectively.

Reports on genetic parameter estimates for pine interspecific hybrid populations remain scarce (Powell and Nikles [1996;](#page-14-0) Dieters et al. [1997](#page-14-0); Gwaze et al. [2000\)](#page-14-0), and in cases where they are available, the sample size is small and there is a general lack of pure species controls. This means that there is little information on whether hybrid populations behave similarly to pure species populations and conform to current quantitative genetic models (e.g. Kain [2003](#page-14-0)). Generally, the estimates of female and male heritability were non-significant and had large standard errors, perhaps due to the limited sample size. $\hat{h}^2 b$ for the female parents (PCH) might be expected to be significant compared to male heritability due to the higher number of female (11) than male (6) parents. However, no consistent pattern emerged in the results. The pooled-site heritability estimates for PCH×POOC hybrids for DBH were generally higher than the companion estimates for PCH× PTEC hybrids. Brawner et al. ([2005](#page-13-0)) observed similar patterns for these hybrids grown at two sites in Queensland, Australia. The lack of significant female $\hat{h}^2 b$ for stem straightness for the hybrids is in contrast with results reported by Dieters et al. [\(1997\)](#page-14-0) for the same material but grown in Queensland, Australia. They reported female $\hat{h}^2 b$ of 0.41 ± 0.20 and 0.48 ± 0.23 for PCH×POOC and PCH×PTEC, respectively. However, we note that heritability estimates are specific to sites and also that the hybrids were developed in Queensland, Australia.

The estimates of female and male heritability need to be interpreted with caution, given the small number of female and male parents and families in the tests. The parents of each species also represented a mix of provenances—three POOC parents from each of Zapotillo and Angeles provenances and three PTEC parents from each of Mountain Pine Ridge and Yucul provenances (Dieters et al. [1997\)](#page-14-0). Such a mix and the small number of female and male parents would make it difficult to meaningfully detangle their likely differences in the additive variance between the male and female parents. The observed differences in female and male heritability estimates may have been inflated by difference provenance origin. The limitation in these data also makes it impossible to definitively recommend using the parental or individual-tree model. Generally, use of a parental model which provides heritability estimates for both female and male parents would provide

the breeder with details in terms of breeding strategy to adopt—either focus on female or make parent selections depending on heritability.

Estimates of genetic parameters of the hybrids show individual-tree, narrow-sense heritabilities for height, DBH and stem straightness to be low to moderate (Powell and Nikles [1996;](#page-14-0) Dieters et al. [1997;](#page-14-0) Gwaze et al. [2000](#page-14-0)). For example, Dieters et al. ([1995](#page-14-0); [1997\)](#page-14-0) reported narrow-sense heritability estimates for PEE×PCH and PCH×POOC hybrids for DBH, height and straightness in the same range as observed in our study. Dominance variance was negligible and less precisely estimated, an expected result given the very low numbers of parents and small sample sizes (White and Hodge [1989](#page-15-0)). Trends, however, indicated that dominance was greater for stem straightness at age 5 years (STR5) than for the other growth traits and greatest at Mukandi (Table [8\)](#page-8-0). In the pooled-site analyses, dominance was relatively unimportant compared with additive variance (details not shown). We would expect dominance to be relatively low due to increased heterozygosity and absence from inbreeding in the population (e.g. Wu [1997](#page-15-0)). The predominance of additive genetic variance in hybrids is consistent with reports by Dieters et al. [\(1997](#page-14-0)) and Powell and Nikles ([1996\)](#page-14-0) in pine hybrids and Madhibha et al. ([2013](#page-14-0)) in eucalypt hybrids. These results suggest that breeding strategies which maximize the use of additive genetic variance may be effective.

Genetic correlation estimates between height and DBH at both ages were large and positive, above 0.80 (Table [9](#page-9-0)). These correlations indicate both traits at the two ages are likely controlled by the same set of genes: The result gives confidence in early selection. Genetic correlations between DBH and STR at both sites were low but significant, indicating larger diameter associated with straight stems. This is a favourable correlation, suggesting that selection for DBH may also give candidates with straighter stems. Similar patterns have been observed in parental species and also the hybrids (Dieters et al. [1997](#page-14-0); Gwaze et al. [2000;](#page-14-0) Gapare and Musokonyi [2002](#page-14-0)).

Heritability estimates from the data pooled across the two sites were generally intermediate between those from individual sites. There was no evidence of practically important $G \times E$ interaction for the hybrids except for height growth for PCH× PTEC hybrid. This suggests that PCH×PTEC and PCH× POOC hybrids appear to be stable across the two sites. It is not clear what is driving $G \times E$ for height growth in PCH \times PTEC hybrid. Our results suggest that the importance of $G \times$ E interaction was trait and hybrid taxon dependent, for example, height for PCH×PTEC hybrid. While G×E interaction in hybrid populations is not well understood, the developmental stability of hybrids has been defined from two different models, epistasis and pleiotropy (Wu [1997\)](#page-15-0). Wu ([1997](#page-15-0)) postulated that low developmental stability may result from reactions to the new environment and from a breakdown of coadapted gene complexes. However, we note that P. tecunumanii parents for the PCH×PTEC hybrid originated from low-elevation provenances (Yucul and Mountain Pine Ridge (MPR)). Nyoka et al. ([2010](#page-14-0)) reported the MPR provenance to be interactive for growth at Cashel. For example, the provenance showed exceptional height growth at two years, where it was ranked among the best but its growth rate subsequently declined with increasing age, to be ranked lowest for both growth and stem straightness at age five and eight years (Nyoka et al. [2010](#page-14-0)). P. caribaea var. hondurensis, in particular the MPR provenance is also known to exhibit $G \times E$ interaction (e.g., Woolaston et al. [1991](#page-15-0)) and also between Mukandi and Cashel (Gapare and Musokonyi, [2002](#page-14-0)).

Implications for hybrid breeding strategy

The hybrids studied here appear to be robust and well adapted to the target areas, expressing favorable genes from both parents. In this study, we recommend that synthetic (SYN) hybrid strategy would be the most cost-effective strategy, given that there is less dominance variance and the pure–hybrid correlations in both species are greater than zero (e.g. Kerr et al. [2004b\)](#page-14-0). The creation of a synthetic breed by intermating advanced generation hybrids was found to provide the most genetic gain per breeding cycle when there is less dominance variance than additive variance (Brawner et al. [2005](#page-13-0)). Using outstanding material in F_1 and subsequent hybrid generations to advance a breeding program is expected to stabilise a synthetic population after two or three generations of mating due to the exponential reduction in linkage disequilibrium between unlinked genes (Falconer and Mackay [1996\)](#page-14-0). The ultimate effect of continued selection within a synthetic would be to increase the most favourable double homozygote and decrease the frequency of all others.

We envisage that the creation of a synthetic breed would be facilitated if parents for advanced-generation crosses could be selected without testing the candidate parents in a specific hybrid combination. For example, if a parent consistently ranks above average regardless of the species with which it is combined, it would be considered stable against different genetic backgrounds and the correlation between pure and hybrid species performance would be high. In this study, there were non-significant correlations between the GHAs of the two interspecific combinations for all traits at ages 5 and 8 years. This may be due to imprecise estimates of the GHA of the respective hybrids because of a smaller sample size. For example, the correlation between the true and predicted breeding values were very low ranging from 0.11 to 0.33 for HT8 and DBH8. Our results differ from those of Dieters et al. [\(1997\)](#page-14-0) who reported identical rankings of the parents for stem straightness, but not so for DBH at age 5 years. Dieters et al. [\(1997\)](#page-14-0) also attributed their correlations to the level of dominance variance which was small relative to additive variance. Brawner et al. ([2005](#page-13-0)) reported positive and high correlations

(>0.75) between PCH parents used in combination with either PTEC or POOC for growth traits. They attributed the increased correlations compared to those reported by Dieters et al. ([1997](#page-14-0)) to more precise breeding value predictions from a larger number of tests and parents. However, we identified four second-generation PCH parents that consistently ranked above average for HT8 and DBH8 when crossed with P. oocarpa. The statistical significance of the correlations between GHAs of the two interspecific combinations could be expected to be increased by increasing the sample size. However, the 'true' genetic correlations between parental performances in hybrid combinations between the two species would be higher than our estimates because of independence of errors of estimating breeding values in the respective species combinations.

As a follow-up to the proposed strategy, several species and interspecific hybrids using locally bred parents are being tested by the Research and Development Division of the Zimbabwe Forestry Commission in Zimbabwe. The focus is on interspecific hybrids of the central American and Mexican closed cone pines as well as hybrids based on P. caribaea. The central American and Mexican closed-cone pines included P. patula, P. tecunumanii, P. oocarpa, P. greggii and P. pringlei. For P. patula, P. tecunumanii and P. oocarpa, three outstanding parents in progeny and also in provenance/progeny tests were selected, while those of P. greggii and P. pringlei were the available parents that have not been field-tested as there is no proper breeding programme for these two species in Zimbabwe (Barnes et al. 1997; Nyoka [2000\)](#page-14-0).

However, the adoption of the hybrids as commercial tree species has its limitations and careful consideration is needed before it is implemented. For example, breeding of PCH and subsequent hybrids may be a challenge in Zimbabwe because PCH does not flower at high altitudes. A notable option suggested by Barnes (1993) was for Zimbabwe to run a joint program with Mozambique to establish breeding seedling orchards of PCH in low-elevation areas of Mozambique to produce seeds for both Zimbabwe and Mozambique. The vegetative propagation facility at Mukandi Nursery may be used for vegetative multiplication of very juvenile material in order to extend the small amount of control-crossed seed (e.g. Madhibha et al. [2013\)](#page-14-0). Another option could be by importing seed from other countries as mentioned by Gwaze [\(1999\)](#page-14-0), but this option is likely to be even more expensive and may not be sustainable for the local industry. While growth traits will remain important, other hybrid programs elsewhere have started focusing on wood properties and are generally inherited in an additive manner (Kain [2003](#page-14-0)). This is particularly important given the trend towards shorter rotations in pines which often result in larger amounts of corewood (Gapare et al. [2006\)](#page-14-0). This will need to be taken into account in future breeding strategies.

Acknowledgments The authors acknowledge many members (past and present) of the Zimbabwe Forestry Commission who have been involved in the establishment, maintenance, measurement and assessments of the trials. In particular we wish to thank Messrs Titus Mangezi, Caleb Mhongweh and Golden Gadaga for good management of the hybrid trials. We also thank Dr. Garth Nikles of the then Queensland Forest Research Institute, Australia for providing the control-pollinated hybrid seed to the Zimbabwe Forestry Commission through the then Oxford Forestry Institute UK. Special thanks to the late Dr. Richard Barnes; without the foresight and insistence in the late 1990s, the idea of hybrid trials in Zimbabwe would never have been implemented. We also thank Drs. Milos Ivković and David Bush and an anonymous reviewer for their constructive suggestions on an earlier version of this paper. Special thanks and patience to the Associate Editor, Dr. Rowland Burdon for his useful suggestions on several versions of this paper.

Conflict of interest The authors declare that they have no conflict of interest.

Data archiving statement Height, diameter and stem straightness data used in this manuscript will be made available as an electronic supplement to this publication upon acceptance of the manuscript.

References

- Armitage P, Berry G, Matthews JNS (2001) Statistical methods in medical research, 4th edn. Blackwell Science, Oxford, p 817
- Arnold M, White T (1994) Review of the tree breeding programme of the Zimbabwe Forestry Commission., p 47
- Barnes RD (1981) A review of the tree breeding programme in Zimbabwe and recommendations for its development. Technical Cooperation Scheme, Overseas Development Administration, London, United Kingdom, p 33
- Barnes RD (1989) The forest genetics programme in Zimbabwe. Plan of work for 1989/90: Part II. ODA Technical Cooperation Scheme Report. OFI, Oxford, p 229
- Barnes RD (1993) The forest genetics programme in Zimbabwe. Plan of work for 1993: Part II. ODATechnical Cooperation Scheme Report. OFI, Oxford, p 308
- Barnes RD, Mullin LJ (1978) Three-year height performance of Pinus elliottii Engelm. var. elliottii \times P. taeda L. hybrid families on three sites in Rhodesia. Silvae Genet 27:217–223
- Barnes RD, Mullin LJ, Battle G (1992a) Genetic control of fifth year traits in Pinus patula Schiede and Deppe. Silvae Genet 41:242–248
- Barnes RD, Mullin LJ, Battle G (1992b) Genetic control of eight year traits in Pinus patula Schiede and Deppe. Silvae Genet 41:318–326
- Barnes RD, Nyoka BI, Gapare WJ (1997) The forest genetic programme in Zimbabwe: Part I: Principles, Structure, Procedure, Techniques, Practice and Literature., pp 162–163
- Barrett RL, Mullin LJ (1968) A review of introductions of forest trees in Rhodesia. The Rhod Bull For Res 1:63
- Becker WA (1984) Manual of quantitative genetics, 4th edn. Academic Enterprises, Pullman, Washington D.C, p 170
- Brawner JT, Dieters MJ, Nikles DG (2005) Mid-rotation performance of Pinus caribaea var. hondurensis hybrids with both P. oocarpa and P. tecunumanii: hybrids superiority, stability of parental performance and potential for a multi-species synthetic breed. For Genet 12:1–13
- Burdon RD (1977) Genetic correlation as a concept for studying genotype–environment interaction in forest tree breeding. Silvae Genet 26:168–175
- Butler DG, Cullis BR, Gilmour AR, Gogel BJ (2009) ASReml-R reference manual, release 3. Technical report. NSW Department of Primary Industries. 160
- Cappa EP, Marco M, Nikles DG, Last IS (2013) Performance of Pinus elliottii, Pinus caribaea, their F-1, F-2 and backcross hybrids and Pinus taeda to 10 years in the Mesopotamia region, Argentina. New For 44:197–218
- Cheverud JM, Routman EJ (1995) Epistasis and its contribution to genetic variance components. Genet 63:511–520
- Cockerham CC (1963) Estimation of genetic variances. In: Hanson WD, Robinson HF (eds) Statistical genetics and plant breeding. NAS/ NRC Publ. 982, Washington, D.C, pp 53–93
- Comstock RE, Moll RH (1963) Genotype–environment interactions. In: Hanson WD, Robinson HF (eds) Statistical genetics and plant breeding. NAS/NRC Publ. 982, Washington, D.C, pp 164–194
- Comstock RE, Robinson HF, Harvey PH (1949) A breeding procedure designed to make maximum use of both general and specific combining ability. Agronomy J 41:360–367
- Costa e Silva J, Dutkowski GW, Borralho NMG (2005) Across-site heterogeneity of genetic and environmental variances in the genetic evaluation of Eucalyptus globulus trials for height growth. Ann For Sci 62:183–191
- Crockford KL (1995) Yield models and tables for Pinus patula, Pinus elliottii and Pinus taeda in Zimbabwe. Zimbabwe Bulletin of Forestry Research. No. 12. FRC, Highlands, Harare, p 51
- Crockford KS, Birks JS, Barnes RD (1988) Family within provenance trials of Pinus caribaea, P. oocarpa and P. patula ssp. tecunumanii: Early results and implication for breeding strategy. In: Gibson GL, Griffin AR, Matheson AC (eds) Populations structure and genetic improvement strategies in clonal and seedling forestry. Proceedings of IUFRO conference, Pattaya, Thailand, pp 176–183
- Dieters MJ (1999) Later-age performance of P. elliottii var. elliottii and P. caribaea varieties and some of their hybrids. Unpubl. report, Queensland Forestry Research Institute, Gympie, p 28
- Dieters MJ, Brawner J (2007) Productivity of Pinus elliottii, P. caribaea and their F_1 and F_2 hybrids to 15 years in Queensland. Ann For Sci 64:691–698
- Dieters MJ, White TL, Hodge GR (1995) Genetic parameter estimates for volume from full-sib tests of slash pine (Pinus elliottii). Can J For Res 25:1397–1408
- Dieters MJ, Nikles DG, Toon PG, Pomroy P (1997) Genetic parameters for F_1 hybrids of *Pinus caribaea* var. *hondurensis* with both *Pinus* oocarpa and Pinus tecunumanii. Can J For Res 27:2024–2031
- Dungey HS (2001) Pine hybrids a review of their use performance and genetics. For Ecol Manag 148:243–258
- Dungey HS, Dieters MJ, Nikles DG (1999) Hybrid breeding strategies a review. In: Proceedings of the 11th Australian Plant Breeding Conference, Adelaide, 19–23 April 1999. pp 36–37
- Dvorak WS, Lambeth CC, Li B (1993) Genetic and site effects on stem breakage in Pinus tecunumanii. New For 7:237–253
- Dvorak WS, Cutierreze A, Osorio F, Hodge GR, Brawner JT (2000a) Conservation and Testing of Tropical and Subtropical Forest Tree Species by the CAMCORE Cooperative. pp. 128–147. College of Natural Resources, NCSU, Raleigh NC. USA
- Dvorak WS, Jordan AP, Hodge GR, Romero JL (2000b) Assessing evolutionary relationships of pines in the Oocarpae and Australes subsections using RAPD markers. New For 20: 163–192
- Falconer DS, Mackay TFC (1996) Introduction to Quantitative Genetics, 4th edn. Longman, London and New York, p 464
- Gapare WJ, Musokonyi C (2002) Provenance performance and genetic parameters for Pinus caribaea var. hondurensis planted at three sites in Zimbabwe. For Genet 9:183–189
- Gapare WJ, Wu HX, Abaquez A (2006) Genetic control of the time of transition from juvenile to mature wood in Pinus radiata D. Don. Ann For Sci 63:871–878
- Gilmour AR, Gogel BJ, Cullis BR, Thompson R (2009) ASReml User Guide Release 3.0. VSN International Ltd, Hemel Hempstead, p 372
- Gotore T, Murepa R, Gapare WJ (2014) Effects of nitrogen, phosphorus and potassium on the early growth of Pinus patula and Eucalyptus grandis. J Trop For Sci 26:22–31
- Gwaze DP (1999) Performance of some interspecific F1 pine hybrids in Zimbabwe. For Genet 6(4):61–67
- Gwaze DP, Dungey HS, Dieters MJ, Toon PG, Nikles DG (2000) Interspecific pine hybrids. 1. Genetic parameter estimates in Australia. For Genet 7:11–20
- Hodge GR, White TL (1992) Genetic parameter estimates for growth traits at different ages in slash pine and some implications for breeding. Silvae Genet 41:252–262
- Kain DP (2003) Genetic parameters and improvement strategies for the Pinus elliottii var. elliottii×Pinus caribaea var. hondurensis hybrid in Queensland, Australia. Ph.D. thesis, Australian National University, ACT, Australia, p 361
- Kerr RJ, Dungey HS, Dieters MJ (2000) A simulation of the comparative gains from three different hybrid breeding schemes. (1) Theory. In: Dungey HS, Dieters MJ, Nikles DG (eds) Hybrid breeding and genetics of forest trees: proceedings of the QFRI/CRC-SPF symposium, 9–14 April 2000, Noosa, Queensland, Australia. Department of Primary Industries, Brisbane, pp 343–350
- Kerr RJ, Dieters MJ, Tier B, Dungey HS (2004a) Simulation of hybrid forest tree breeding strategies. Can J For Res 34:195–208
- Kerr RJ, Dieters MJ, Tier B, Dungey HS (2004b) Simulation of comparative gains from four different hybrid tree breeding strategies. Can J For Res 34:209–220
- Madhibha T, Murepa R, Musokonyi C, Gapare W (2013) Genetic parameter estimates for interspecific Eucalyptus hybrids and the implication for hybrid breeding strategy. New For 44:63–84
- Mullin LJ (1992) Introduction to tree breeding. In: Piearce GD, Shaw P (eds) Proceedings of the anniversary seminar on forestry research in Zimbabwe, Mutare, August 27–31, 1990. The Forestry Commission, Harare, Zimbabwe, pp 47–51
- Nikles DG (2000) Experience with some Pinus hybrids in Queensland, Australia. In: Dungey HS, Dieters MJ, Nikles DG (eds) Hybrid breeding and genetics of forest trees: proceedings of the QFRI/ CRC-SPF symposium, 9–14 April 2000, Noosa, Queensland, Australia. Department of Primary Industries, Brisbane, pp 27–43
- Nikles DG, Newton RS (1991) Correlations of breeding values in pure and hybrid populations of hoop pine and some southern pines in Queensland and relevance to breeding strategies, In Proceedings of the 11th Meeting of Research Working Group No. 1 (Forest Genetics), 11–15 March 1991. Coonawarra, South Australia, pp 192–196
- Nyoka B (2000) The status of pine interspecific hybrid research in Zimbabwe. In: Dungey HS, Dieters MJ, Nikles DG (eds) Hybrid breeding and genetics of forest trees: proceedings of the QFRI/CRC-SPF symposium, 9–14 April 2000, Noosa, Queensland, Australia. Department of Primary Industries, Brisbane, pp 120–124
- Nyoka BI, Barnes RD (1995) Genetic parameters and provenance productivity of Pinus oocarpa and Pinus patula ssp. tecunumanii. S Afri For J 173:1–7
- Nyoka BI, Gapare WJ, Mangezi T, Mhongwe J, Maruzane D, Barnes RD (1996) The forest genetics programme in Zimbabwe Part 1. Plan of work for 1996/97. Zimbabwe Forest Research Centre, Harare, p 198
- Nyoka BI, Tongoona P, Gumbie CM (2010) Provenance productivity of high and low elevation Pinus tecunumanii in Zimbabwe. Silvae Genet 59:189–199
- Powell MB, Nikles DG (1996) Performance of Pinus elliottii var. elliottii and P. caribaea var, hondurensis, and their F_1 , F_2 and backcross hybrids across a range of sites in Queensland. Dieters, M. J., Matheson, A. C., Nikles, D. G., Harwood, C. E. and Walker, S. M. (Eds). In proc. "Tree improvement for sustainable tropical forestry", Queensland, Australia, 27–1 November 1996. pp 382–383
- Pswarayi IZ, Barnes RD, Birks JS, Kanowski P (1996) Genetic parameter estimates for production and quality traits of Pinus elliottii Engelm, var. elliottii in Zimbabwe. Silvae Genet 45:216–222
- R Development Core Team (2011) R: A language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria [\(http://www.gbif.org/resource/81287\)](http://www.gbif.org/resource/81287)
- Robertson A (1959) The sampling variance of genetic correlation 1179 coefficient. Biometrics 15:469–485
- Shelbourne CJA (2000) Some insights on hybrids in forest tree improvement. In: Dungey HS, Dieters MJ, Nikles DG (eds) Hybrid Breeding and Genetics of Forest Trees, Proceedings of a QFRI/CRC-SPF Symposium, 9–14 April 2000, Noosa, Queensland, Australia. Department of Primary Industries, Brisbane, pp 53–62
- Steel GD, Torrie JH (1980) Principles and procedures in statistics: a biometrical approach, 2nd edn. McGraw-Hill Book Company, Singapore, 633 p
- Stram DO, Lee JW (1994) Variance components testing in the longitudinal mixed effects model. Biogeosciences 50(4):1171–1177
- Tembani M, Madhibha T, Marunda CT, Gapare WJ (2014) Sustaining and improving forest genetic resources for Zimbabwe: Lessons from 100 years. Intl For Rev 16:15–632
- Timber Producers Federation (1999) Zimbabwe timber industry statistics for the year ended 31 March 1999. Timber Producers' Federation, Mutare, Zimbabwe, p 28
- Timber Producers Federation (2011) Industry performance. Timber producers federation, Mutare, Zimbabwe, p 33
- Timber Producers Federation (2014) Industry performance. Timber producers federation, Mutare, Zimbabwe, p 38
- White I (2013) The R pin function. Available from: [http://www.](http://www.homepages.ed.ac.uk/iwhite//asreml/useofpin.pdf) [homepages.ed.ac.uk/iwhite//asreml/useofpin.pdf](http://www.homepages.ed.ac.uk/iwhite//asreml/useofpin.pdf). Accessed: 12/2014
- White TL, Hodge GR (1989) Predicting breeding values with applications in forest tree improvement. Forestry Sciences Series, Vol. 33. Kluwer Academic Publishers, Dordrecht, Netherlands, p 367
- Woolaston RR, Kanowski PJ, Nikles DG (1991) Genotype–environment interaction in Pinus caribaea var. hondurensis in Queensland, Australia. II. Family-site interaction. Silvae Genet 40:228–232
- Wu RL (1997) Genetic control of macro- and micro environment sensitivities in Populus. Theor Appl Genet 94:104–114