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Analysis of macro nutrient related growth responses using multivariate adaptive regression splines

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Abstract

Strawberry micropropagation is generally based on Murashige and Skoog mineral salts, and many cultivars grow well on this medium. However, the diverse species found in germplasm collections often do not thrive, which indicates a need to optimize the mineral nutrients. In this study, Multivariate Adaptive Regression Splines (MARS), was employed to predict shoot quality, multiplication, and leaf color responses of three strawberry species in response to the major tissue culture nutrients by generating functional associations. MARS is a non-parametric approach that can be used to deal with continuous and categorical data without requiring the strict distributional assumptions of the basic linear models. The MARS algorithm is capable of capturing non-linear patterns between the input and target variables. NH₄NO₃, CaCl₂·2H₂O, MgSO₄·7H₂O, KNO₃ and KH₂PO₄ were tested in a range of $0.5 \times to 3 \times MS$ medium, within a computer-generated optimal design that consisted of 32 treatment combinations. The plant responses were affected by all of the major salts tested and the genotype factor. Multivariate Adaptive Regression Splines captured the significant factors and their interactions to predict optimal major salts suitable for all three strawberry species: 3300 mg L⁻¹ NH₄NO₃, 862.4 mg L⁻¹ CaCl₂, 1110 mg L⁻¹ MgSO₄, 3439 mg L⁻¹ KNO₃, and 329.8 mg L⁻¹ KH₂PO₄. This study identified the major nutrient needs of the three strawberry species and provides an alternative statistical technique for tissue culture data analyses.

Key Message

The MARS statistical approach was used to predict macro nutrient related growth responses of three strawberry species. The objective of the study was to make a gentle introduction to the MARS algorithm and show its potential application to tissue culture research.

Keywords Fragaria · In vitro culture · MARS · Mineral nutrition · Statistical analysis

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Introduction

While strawberries were first grown on a variety of media in the early days of micropropagation, they are currently most commonly grown on MS medium (Murashige and Skoog 1962), or variations such as half-strength MS mineral nutrients. Simpson and Bell (1989) noted that most *Fragaria* × ananassa cultivars grew well on MS medium, but proliferation varied greatly by genotype. They concluded that strawberries do not have a uniform response to any single growth medium recipe due to the differences in response to N⁶ benzyladenine (BA) concentrations. A study of *Fragaria chiloensis* cultivars from Chile found that MS was suitable for the two cultivars studied (Quiroz et al. 2017). Most diverse germplasm or breeding collections can use MS medium for the majority of accessions, but often there are genotypes or species that do not thrive despite changes in plant growth regulators. The cultivated strawberry, *Fragaria*×*ananassa* Duchesne ex Rozier has a very narrow genetic background, but the 20 or more other species are more genetically diverse and include a wide range of ploidy levels (Darrow 1966).

The lack of study on the mineral nutrition of strawberry combined with the genetic diversity present in germplasm collections and breeding programs, provides incentives to improve the mineral nutrition in the medium for micropropagated strawberries. Developing a mineral nutrition model can provide the basis for an improved medium for diverse groups of plants.

The performance of a good statistical model depends on the experimental design, clear understanding of the variable structure, and application of the right statistical method. Understanding the data type is of primary importance to determine the appropriate statistical technique (Mertler and Vannatta 2002; Niedz and Evens 2016). Variables are classified as qualitative (subgrouped as nominal and ordinal), and quantitative (labeled as discrete and continuous). Nominal variables are 'names' with two or more categories without hierarchical order. Ordinal data are similar to nominal except that the categories have distinct order (level A is more severe than level B). Discrete variables are counts that include integers, and continuous variables are measurements along a continuum, which could be divided into smaller fractions (Dettori and Norvell 2018). In vitro data structure could be classified as nominal (genotype), ordinal (shoot quality rated as poor, moderate, and good), discrete (counts of shoot number: 1, 2, 3...), and continuous (measurements of shoot length). Conventional linear models such as ANOVA and regression, should only be used with continuous data that shows a linear relationship between the dependent and explanatory variables (Quinn and Keough 2002). However, plant tissue culture is a complicated biological process that involves multiple design inputs that typically interact in nonlinear ways to affect the tested plant responses. Therefore, the basic linear methods are generally not suitable for analyzing in vitro data.

A number of advanced statistical techniques are promoted in tissue culture research as robust alternatives to basic modeling methods. These include Response Surface Methodology (RSM), Decision trees, and Artificial Neural Networks (ANNs) (Akin et al. 2017; Nezami-Alanagh et al. 2018; Poothong and Reed 2014). The RSM is a parametric approach that is able to model non-linear data by showing polynomial relationships between the tested variables. However, RSM can analyze only continuous variables, and therefore it generates genotype (nominal variable) specific models (which means a separate optimal growth area for each genotype) (Akin et al. 2016). Classification and Regression Tree (CART), Chi-square Automatic Interaction Detector

(CHAID), and Exhaustive Chi-square Automatic Interaction Detector (Exhaustive CHAID), are non-parametric data mining methods that are able to evaluate both continuous and categorical data, without requiring the restrictive normality and heteroscedasticity assumptions of the parametric methods (Olden et al. 2008). These techniques generate visual, easy to interpret decision trees and can include nominal variables as factors, and thus provide common optimal factor ranges for the tested genotypes (Akin et al. 2017). Artificial Neural Networks are non-parametric tools that are powerful in detecting non-linear associations, and capable of dealing with continuous and categorical data. However, ANNs do not provide a neat mathematical equation that shows the relative contribution of each independent variable in the neural model. Therefore, ANNs are considered as a "black box" (Olden et al. 2008).

Another promising statistical approach not previously utilized in tissue culture data analyses is Multivariate Adaptive Regression Splines (MARS). The MARS approach is a nonparametric regression technique that reveals complex nonlinear relationships and interactions by a sequence of spline functions of the independent variables, without requiring the restrictive distributional assumptions of the general linear models. The MARS algorithm can handle both categorical and numerical variables. It shows the contribution of each predictor to the response using subsets of piecewise linear regression splines (Everingham and Sexton 2011; Friedman 1991). A spline is a flexible curve (polynomial) that is fixed at various points or knots and shows the association between the target and explanatory variable. The knots are randomly distributed within the range of each explanatory variable and detect the regions of relationship change between the input and output variables. The linear functions that form the MARS model combine additively or interactively, and the relative importance of the inputs on the target variable are illustrated clearly in a mathematical equation (Eyduran et al. 2019; Zhang and Goh 2014).

This study investigated the potential of the MARS algorithm to reveal the intrinsic non-linear and multidimensional relationships of some responses of tissue culture grown strawberries. This process aims to provide a basic introduction to the MARS algorithm and to encourage future application of this promising approach for tissue culture research, and to improve the in vitro culture of strawberries.

Materials and methods

Plant material and culture conditions

Shoots of *Fragaria bucharica* Losinsk. (Local no. CFRA 1906.001; PI 657844), *Fragaria chiloensis* (L.) Mill. subsp. *chiloensis patagonica* Staudt (CFRA 1084.002; PI

616535), and Fragaria x ananassa (Duchesne ex Weston) Duchesne ex Rozier 'Jucunda' (CFRA 256.002; PI 551623) strawberry species from the U.S. Department of Agriculture, National Clonal Germplasm Repository (NCGR), Corvallis, Oregon, were micropropagated on MS medium (Murashige and Skoog 1962), with LS vitamins (Linsmaier and Skoog 1965), and with the following per liter: 80 mg adenine sulfate, 170 mg NaH₂PO₄, 30 g sucrose, 4.44 \propto M BA, $4.92 \propto M$ indole-3-butyric acid (IBA), $0.03 \propto M$ gibberellic acid (GA) and 6 g L^{-1} agar (A1111; PhytoTechnology Laboratories[→], Lenexa, KS) at pH 5.7. The medium pH was adjusted using H₃PO₄ and KOH, and all medium ingredients were added before autoclaving. The plant growth regulators and vitamins were obtained from PhytoTechnology Laboratories®, and all stock solutions were prepared in house. Each MagentaTM box (MagentaTM GA7, Magenta, Chicago, IL) contained 40 mL medium, and was autoclaved for 20 min at 121 °C (118 kPa). Growth room conditions were 80 µmol m²s⁻¹ light intensity with a 16-h photoperiod of half warm-white and half cool-white fluorescent lamps (Ecolux Starcoat, 32 watts; GE, Fairfield, CT) at 25 ± 2 °C.

Experimental design

An I-optimal experimental design (DOE) was built by Design-Expert software (Design-Expert 2010) using NH₄NO₃, CaCl₂·2H₂O, MgSO₄·7H₂O, KNO₃, and KH₂PO₄ as inputs within 32 treatment combinations. The factors were varied from 0.5×to 3×MS medium salts. Treatment 32 was the control NCGR *Fragaria* medium consisting of MS salts (1×) with adenine sulfate and NaH₂PO₄ (Table 1). All treatment

Treatments	NH ₄ NO ₃	CaCl ₂ ·2H ₂ O	MgSO ₄ ·7H ₂ O	KNO ₃	KH ₂ PO ₄
1	1.75	1.66	0.50	1.75	0.50
2	3.00	3.00	0.50	0.50	3.00
3	2.23	2.36	1.17	0.50	1.50
4	3.00	0.50	0.50	0.50	0.50
5	0.50	3.00	3.00	0.50	3.00
6	0.50	0.50	3.00	3.00	3.00
7	0.50	3.00	0.50	1.93	1.94
8	0.50	0.50	0.50	0.50	3.00
9	0.50	0.50	3.00	0.50	0.50
10	0.50	3.00	0.50	0.50	0.50
11	1.27	1.43	2.25	2.50	1.42
12	1.27	1.43	2.25	2.50	1.42
13	1.58	3.00	1.59	3.00	3.00
14	0.50	3.00	3.00	3.00	0.50
15	3.00	3.00	3.00	3.00	3.00
16	2.00	1.45	0.84	1.50	3.00
17	1.58	3.00	1.59	3.00	3.00
18	3.00	0.50	1.69	1.81	1.65
19	3.00	3.00	0.50	3.00	0.50
20	0.85	0.50	3.00	0.85	3.00
21	3.00	0.50	3.00	3.00	0.50
22	0.50	1.96	0.50	3.00	3.00
23	2.00	1.45	0.84	1.50	3.00
24	0.50	3.00	0.50	1.93	1.94
25	1.27	1.43	2.25	2.50	1.42
26	3.00	0.50	0.50	3.00	3.00
27	3.00	0.50	3.00	0.50	3.00
28	2.13	3.00	3.00	0.50	1.38
29	0.50	0.50	0.50	3.00	0.50
30	3.00	3.00	2.14	1.38	0.50
31	0.50	1.65	1.75	0.50	1.75
32	1.00	1.00	1.00	1.00	1.00

Table 1I-optimal designconsisting of MS (Murashigeand Skoog 1962) major salts

The concentration ranges of the treatment combinations are expressed as ×MS levels

combinations contained the same type and concentration of plant growth regulators, vitamins and sucrose as noted above. For each treatment, two boxes with five plantlets for each geno-type were used (n=10). Boxes were randomized on the growth room shelf. Cultures were grown on each treatment for 9 week, with transfers to fresh medium every 3 week.

Data structure

Three plantlets selected on a diagonal, from each MagentaTM box (n=6) were evaluated, and the remaining four shoots were photographed for additional documentation. Shoot quality, which is a subjective visual assessment of shoot growth was evaluated as 1=poor (no growth or bad appearance), 2=moderate (some growth and moderate appearance), and 3=good (well-developed shoot culture with expanded leaves and good color). Multiplication was also quantified with the same scale of quality depending on the proliferation rate of the plant, with 1=no shoot production, 2=2 to 4 shoots, and 3=>4 shoots. Leaf color was estimated as >50% of leaves on the shoot as follows: 1=yellow or red, 2=light green, 3=green (Niedz et al. 2007; Reed et al. 2013).

Statistical analyses

Although shoot quality, multiplication, and leaf color responses are ordinal-type data, these inputs were treated as continuous by taking the mean response of six shoots per treatment for each strawberry species. The outputs were converted to continuous variables to balance the within group variation. The genotype as a nominal variable was also included within the analyses. A multiple-response model was built by the MARS approach (Table 2), using STATISTICA 7.1 software (StatSoft 2005).

The MARS algorithm is constructed by piecewise linear basis functions (*BFs*) of the following form:

$$BF1 = \max(0, x - t) \begin{cases} x - t, x > t \\ 0, x \le t \end{cases}$$
(1a)

$$BF2 = \max(0, t - x) \begin{cases} t - x, t > x \\ 0, t \le x \end{cases}$$
(1b)

where *x* is the variable range and *t* is the knot.

The MARS model is a linear combination of basis functions:

$$Yi = a_0 + a_1 BF1_i + a_2 BF2_i + \dots$$
(2)

where Y_i is the dependent variable, a_0 is the intercept, and a_1 and a_2 are the coefficients of the related basis functions (Emangolizadeh et al. 2015; Everingham and Sexton 2011). For better understanding of the model, a simple MARS

example consisting of three piecewise linear functions or splines (BF1, BF2, and BF3), connected at two points or knots generated by RStudio software (R Core Team 2017) is illustrated in Fig. 1. The MARS mathematical equation is expressed as:

$$Yi = a_0 + a_1 BF1_i + a_2 BF2_i + a_3 BF3_i$$
(3)

in which $BF1 = \max(0, x - 3)$, and $BF2 = \max(0, 3-x)$, and $BF3 = \max(0, x - 6)$. In this case the knot points are t = 3 and 6. These two knots delimit the *x* input range in three regions, in which different linear relationships between the response and the *x* explanatory variable are detected (Fig. 1).

The MARS model building process consists of forward and backward stepwise selections to define the knots and splines. First, the model is overfitted by forward selection of more basis functions than required to express the response variable, and subsequently backward pruning is performed by deleting the least significant splines one at a time until an optimal model is obtained. The model is refitted after each basis function removal, and each reduced sub-optimal model is tested with the Generalized Cross-Validation (GCV) method to prevent overfitting. (Everingham and Sexton 2011). The model with the lowest GCV score is considered to be the best (Zakeri et al. 2010). An optimal MARS model is also expected to have a Pearson correlation coefficient value close to one, as the coefficient approaches zero the predictive ability of the model decreases. The MARS algorithm is sensitive to multicollinearity, which refers to correlation between inputs. When polynomial terms are included to model curvature multicollinearity problem could occur in regression type methods (Friedman 1991). Pearson correlation coefficients between pairs of the salt inputs were estimated to test for multicollinearity.

MARS is an adaptive procedure that allows users to specify many options within the model, including a maximum number of basis functions, the level of interactions between variables, and a minimum number of terms to leave between each knot depending on the data. A four-way interaction was specified, and the algorithm was assigned a maximum of 74 basis functions and a minimum of 32 terms. These values were selected to generate the best model with the highest Pearson correlation and lowest GCV score.

Results

The design of the experiment provided a diversity of responses for the tested strawberries which varied with the mineral nutrients (Fig. 2). The MARS model was able to reveal the complex relationships between the tested salt inputs and the responses by showing the relevance of each predictor on the target variable in a mathematical equation

Table 2 Basis functions and corresponding coefficients of the Multivariate Adaptive Regression Splines (MARS) model for plant quality, multiplication, and leaf color prediction in Fragaria bucharica, Fragaria chiloensis and *Fragaria x ananassa* 'Jucunda'

Terms	Coefficients (a_i)			Basis functions		
_	Quality	Multiplication	Leaf color			
1	2.68	3.06	2.74	Intercept		
2	- 0.75	- 1.13	- 0.73	$\max(0, \mathrm{NH}_4\mathrm{NO}_3-2)$		
3	- 0.89	- 1.37	- 0.87	max(0, 2-NH ₄ NO ₃)		
4	0.28	0.46	0.26	max(0, KNO ₃ -1.81)		
5	- 0.35	- 0.4	- 0.34	$max(0, 2-NH_4NO_3) \times max(0, Jucunda)$		
6	0.52	0.74	0.55	$max(0, 2-NH_4NO_3) \times max(0, MgSO_4-1.75)$		
7	0.42	0.73	0.49	$max(0, 2-NH_4NO_3) \times max(0, 1.75-MgSO_4)$		
8	0.6	0.78	0.54	max(0, NH ₄ NO ₃ -2)×max(0, 1.81-KNO ₃)		
9	0.41	0.45	0.33	$\max(0, 2-\text{NH}_4\text{NO}_3) \times \max(0, 1.81-\text{KNO}_3)$		
10	- 0.61	- 0.87	- 0.56	$\max(0, 2-\mathrm{NH}_4\mathrm{NO}_3) \times \max(0, \mathrm{MgSO}_4-1.75) \times \max(0, F. bucharica)$		
11	- 0.2	- 0.25	- 0.27	max(0, CaCl ₂ -1.96)		
12	0.15	0.22	0.2	$\max(0, 1.81\text{-}\text{KNO}_3) \times \max(0, \text{KH}_2\text{PO}_4\text{-}0.5) \times \max(0, \text{Jucunda})$		
13	- 1.11	- 1.35	- 1.1	$\max(0, \mathrm{NH}_4\mathrm{NO}_3\text{-}2) \times \max(0, 1.81\text{-}\mathrm{KNO}_3) \times \max(0, F. \ bucharica)$		
14	- 0.26	- 0.18	- 0.35	max(0, F. bucharica)		
15	- 0.46	- 0.71	- 0.46	$\max(0, \text{KNO}_3-1.81) \times \max(0, F. bucharica)$		
16	- 0.47	- 0.64	- 0.44	$\max(0, 1.81\text{-KNO}_3) \times \max(0, F.chiloensis)$		
17	0.21	0.25	0.19	$\max(0, \text{MgSO}_4-0.5) \times \max(0, \text{KH}_2\text{PO}_4-0.5) \times \max(0, F.chiloensis)$		
18	- 0.23	- 0.31	- 0.27	$max(0, 1.96-CaCl_2) \times max(0, KNO_3-1.81)$		
19	0.24	0.41	0.26	$\max(0, \text{KNO}_3-1.81) \times \max(0, \text{KH}_2\text{PO}_4-1.94)$		
20	- 0.98	- 1.39	- 1	max(0, KH ₂ PO ₄ -1.94)		
21	- 1.34	- 1.67	- 1.28	$\max(0, \operatorname{CaCl}_2-1.96) \times \max(0, F. bucharica)$		
22	- 0.39	- 0.4	- 0.36	$max(0, 1.94-KH_2PO_4) \times max(0, Jucunda)$		
23	1.08	1.49	1.04	max(0, CaCl ₂ -1.96)×max(0, KNO ₃ -0.5)×max(0, <i>F. bucharica</i>)		
24	- 0.61	- 0.78	- 0.56	$\max(0, \text{NH}_4\text{NO}_3-0.5) \times \max(0, \text{CaCl}_2-1.96) \times \max(0, \text{KNO}_3-0.5) \times \max(0, F. bucharica)$		
25	0.87	1.01	0.83	$\max(0, \mathrm{NH}_4\mathrm{NO}_3-0.5) \times \max(0, \mathrm{CaCl}_2-1.96) \times \max(0, F. bucharica)$		
26	0.53	0.76	0.51	$\max(0, 2-\text{NH}_4\text{NO}_3) \times \max(0, \text{MgSO}_4-1.75) \times \max(0, \text{KH}_2\text{PO}_4-0.5) \times \max(0, F. bucharica)$		
27	0.42	0.56	0.47	$\max(0, \text{NH}_4\text{NO}_3\text{-}0.5) \times \max(0, \text{KH}_2\text{PO}_4\text{-}1.94) \times \max(0, F. \ bucharica)$		
28	0.15	0.28	0.2	max(0, 1.81-KNO ₃)×max(0, KH ₂ PO ₄ -0.5)×max(0, <i>F. chiloensis</i>)		
29	- 0.03	- 0.05	- 0.03	$\max(0, \operatorname{CaCl}_2-0.5) \times \max(0, \operatorname{MgSO}_4-0.5) \times \max(0, \operatorname{KH}_2\operatorname{PO}_4-0.5) \times \max(0, F. \ chiloensis)$		
30	- 0.27	- 0.28	- 0.3	$\max(0, 2-\text{NH}_4\text{NO}_3) \times \max(0, \text{MgSO}_4-1.75) \times \max(0, F. chiloensis)$		
31	- 0.09	- 0.09	-0.07	$\max(0, \text{NH}_4\text{NO}_3-0.5) \times \max(0, \text{MgSO}_4-0.5) \times \max(0, \text{KH}_2\text{PO}_4-0.5) \times \max(0, F. chiloensis)$		
32	0.12	0.11	0.1	$\max(0, \text{NH}_4\text{NO}_3\text{-}0.5) \times \max(0, \text{KH}_2\text{PO}_4\text{-}0.5) \times \max(0, F. \text{ chiloensis})$		

The model consists of 32 terms leaved between each knot and 74 basis functions with four-way interactions. Four cut-offs, 2.0, 1.96, 1.81, and 1.94 are the most frequent in the multi-response MARS model equation, which constituted the basis functions with the corresponding MS salts. These cut-offs were selected to make the multi-response MARS model simpler by deleting the related basis functions. Consequently, the simplified MARS model consisted of the marked coefficients and basis functions

(Table 2). The algorithm successfully incorporated the genotype-nominal variable within the analysis, and therefore was able to detect the common nutrient requirements for all three strawberry species. The predictive ability of the model for the tested outcomes was also high, with a Pearson correlation coefficient about 0.90 (p < 0.01). The GCV value was approximately 0.49. The estimated Pearson correlation coefficients between the pairs of nutrient salt factors were found to be insignificant, which indicates that no multicollinearity problem was present. Plant quality, multiplication, and leaf color were affected by all of the tested salt inputs and the genotype factor (Table 2).

Shoot quality model

The MARS model prediction equation for shoot quality is presented in Supplement 1. To further optimize the target responses and provide common nutrient concentrations for the tested strawberry species, ranges for the multipleresponse MARS model were selected as $2 \times NH_4NO_3$,

Fig. 1 Graphical representation of a simple Multivariate Adaptive Regression Splines (MARS) model represented by three piecewise linear basis functions (BF1, BF2, and BF3) connected to each other with knots (indicated in blue). The knots show the regions of association change between the explanatory and target variable. These two knots delimit the x input range in three regions, in which different linear relationships between the response and the x explanatory variable are detected





Fig. 2 Growth response of shoots of *Fragaria bucharica*, *Fragaria chiloensis* and *Fragaria x ananassa* 'Jucunda' grown on treatments with variations in MS major salt concentrations (Table 1). Treatments

were chosen to show the range of plant quality, multiplication, and leaf color produced by changes in mineral nutrition.

 $1.96 \times CaCl_2$, $3 \times MgSO_4$, $1.81 \times KNO_3$, and $1.94 \times KH_2PO_4$. These values were defined by looking at the MARS equation (Table 2). Four cut-offs, 2.0, 1.96, 1.81, and 1.94, were the most frequent in the equation, which constituted the basis functions with the corresponding salts [max (0, NH₄NO₃-2), max (0, CaCl₂-1.96), max (0, KNO₃-1.81), and max (0, KH₂PO₄-1.94)]. These cut-offs were selected to make the model simpler by deleting the related basis functions. Finally, MgSO₄ was assigned a value of 3 to maximize the responses. After assigning these values, most of the basis functions were deleted according to the rules in Eq. (1a) and (1b) (max (0, NH₄NO₃-2) was equal to 0 when NH₄NO₃<2), and the model for plant quality was simplified to:

Quality = $2.68 - 0.26 \times \max(0, F. bucharica) + 0.21 \times \max(0, MgSO_4 - 0.5) \times \max(0, KH_2PO_4 - 0.5) \times \max(0, F. chiloensis) - 0.03 \times \max(0, CaCl_2 - 0.5) \times \max(0, CaC$

 $\begin{array}{l} MgSO_{4}\text{-}0.5) \times \max (0, \text{KH}_{2}\text{PO}_{4}\text{-}0.5) \times \max (0, F. chiloensis) \\ -0.09 \times \max (0, \text{NH}_{4}\text{NO}_{3}\text{-}0.5) \times \max (0, \text{MgSO}_{4}\text{-}0.5) \times \max \\ (0, \text{KH}_{2}\text{PO}_{4}\text{-}0.5) \times \max (0, F. chiloensis) + 0.12 \times \max \\ (0, \text{NH}_{4}\text{NO}_{3}\text{-}0.5) \times \max (0, \text{KH}_{2}\text{PO}_{4}\text{-}0.5) \times \max (0, F. chiloensis) \\ \end{array}$

The model was further reduced based on the rules of Eq. (1a) and (1b), by subtracting the assigned salt values and the knot points (*e.g.*, max (0, MgSO₄-0.5) was 3-0.5=2.5 when MgSO₄>0.5).

The final model for shoot quality was:

Quality = $2.68 - 0.26 \times \max(0, F. bucharica) + 0.76 \times \max(0, F. chiloensis) - 0.16 \times \max(0, F. chiloensis) - 0.49 \times \max(0, F. chiloensis)$

 $(0, F. chiloensis) + 0.26 \times max (0, F. chiloensis)$

The mean quality of *F. bucharica* was predicted to be:

 $2.68 - 0.26 \times 1 + 0.76 \times 0 - 0.16 \times 0 - 0.49 \times 0 + 0.26 \times 0 =$ 2.68 - 0.26 = 2.42

Genotype was treated as a nominal categorical variable, thus 1 was used for max (0, *F. bucharica*) and 0 otherwise [in this case for max (0, *F. chiloensis*)].

Using the same approach, the average shoot quality of *F*. *chiloensis* was:

 $2.68 - 0.26 \times 0 + 0.76 \times 1 - 0.16$ $\times 1 - 0.49 \times 1 + 0.26 \times 1 = 3.05$

The predicted mean quality of 'Jucunda' was

 $2.68 - 0.26 \times 0 + 0.76 \times 0 - 0.16 \times 0 - 0.49 \times 0 + 0.26 \times 0 = 2.68$

Shoot multiplication model

The MARS model for predicting multiplication (Table 2), after inserting $2 = NH_4NO_3$, $1.96 = CaCl_2$, $3 = MgSO_4$, $1.81 = KNO_3$ and $1.94 = KH_2PO_4$ values, was reduced to:

Multiplication = $3.06-0.18 \times \max(0, F. bucha$ $rica) + 0.9 \times \max(0, F. chiloensis) - 0.26 \times \max(0, F. chiloensis) - 0.49 \times \max(0, F. chiloensis) + 0.24 \times \max(0, F. chiloensis)$

The predicted mean multiplication of *F. bucharica* was: $3.06-0.18 \times 1 + 0.9 \times 0 - 0.26 \times 0 - 0.49 \times 0 + 0.24 \times 0 = 3$. 06-0.18 = 2.88

The predicted average multiplication of *F. chiloensis* was: $3.06-0.18 \times 0 + 0.9 \times 1 - 0.26 \times 1 - 0.49 \times 1 + 0.24 \times 1 = 3.0$ 6+0.9-0.26-0.49+0.24=3.45

The predicted mean multiplication of 'Jucunda' was estimated as 3.06.

Leaf color model

The MARS model to predict leaf color (Table 2), after replacing $2 = NH_4NO_3$, $1.96 = CaCl_2$, $3 = MgSO_4$, $1.81 = KNO_3$ and $1.94 = KH_2PO_4$ values, was shortened to:

Leaf color = $2.74 - 0.35 \times \max(0, F. buchar$ $ica) + 0.68 \times \max(0, F. chiloensis) - 0.16 \times \max(0, F.$ chiloensis) $-0.38 \times max (0, F. chiloensis) + 0.22 \times max (0, F. chiloensis)$

The mean leaf color of *F. bucharica* was: 2.74 - 0.35 = 2.39The average leaf color of *F. chiloensis* was:

2.74 + 0.68 - 0.16 - 0.38 + 0.22 = 3.1

The predicted mean leaf color of "Jucunda" was estimated as 2.74. The mean quality, multiplication, and leaf color of 'Jucunda' depended only on the intercepts of the model under $2 \times \text{NH}_4\text{NO}_3$, $1.96 \times \text{CaCl}_2$, $1.81 \times \text{KNO}_3$ and $1.94 \times \text{KH}_2\text{PO}_4$. Therefore, it can be concluded that the responses of "Jucunda" won't change under any MgSO₄ concentration. However, the aim was to define a common medium for the tested strawberry species. Therefore, $3 \times \text{MgSO}_4$ was specified for the final medium.

The software user doesn't need to make all of those tedious calculations, because only stating the factor values will be enough to obtain the predicted responses. Those equations have been demonstrated to show how transparent and computationally practical the MARS algorithm is for analyzing in vitro culture experiments.

Shoot response to treatment combinations

As expected, none of the treatment combinations (Table 1) exactly matched the suggested optimal macronutrient ranges predicted by the multiple response MARS algorithm. However, growth of shoots of *F. bucharica*, *F. chiloensis* and 'Jucunda' all showed improvement on some treatments, with some of the nutrients in similar ranges compared to the control NCGR *Fragaria* medium treatment (Fig. 3).

Discussion

Developing a growth medium for diverse species of *Fra-garia* provides a useful tool for germplasm collections and breeding stocks. Many of the commonly grown cultivars can be cultured on standard MS medium, however many of the wild species are more difficult. This is the first time that modeling has been used to determine improved medium components for strawberry germplasm.

Modeling using non-parametric machine learning techniques has been utilized in only a few tissue culture studies (Akin et al. 2017; Kovalchuk et al. 2018; Nezami-Alanagh et al. 2018). This is mainly because the plant science researchers are not familiar with this type of statistical modelling. Therefore, this study aimed to introduce the MARS algorithm, which is a promising new tool to analyze in vitro data. The MARS algorithm is a powerful data mining technique that is able to reveal high dimensional, non-linear, and interactional effects among multiple inputs and output



Fig.3 A comm of shoots of *Fragaria bucharica*, *Fragaria chiloensis* and *Fragaria x ananassa* 'Jucunda' grown on the modified MS (NCGR medium) control (treatment 32) and two improved treatments

(Table 1). These improved treatments included some, but not all, of the suggested optimal macronutrient ranges predicted by the Multivariate Adaptive Regression Splines (MARS) algorithm

without requiring the restrictive distributional assumptions of the traditional linear models (Friedman 1991).

Response Surface Methodology is a parametric method, which is able to describe non-linear associations and is used in tissue culture research. However, this method cannot include categorical variables such as genotype within the model. Therefore, RSM projects genotype-specific input requirements for the corresponding dependent variable (Akin et al. 2016). In a study of raspberry mineral nutrition, RSM generated individual models for each raspberry genotype. This resulted in five graphs for plant quality, one for each genotype, and each with a different result rather than a common model. The best treatments were with high NH₄NO₃ and KNO₃, as well as low minors and iron (compared to MS medium) for 'Canby'; high mesos, low minors and iron for 'Indian Summer'; high mesos and low iron for 'Nootka' and 'Trailblazer'; and high mesos and minors for 'Willamette'. Low KNO₃ and high mesos were required for higher multiplication, and depending on genotype NH₄NO₃ and iron were also significant (Poothong and Reed 2014). In a study of hazelnut shoot culture (Akin et al. 2016), analysis using the CHAID decision tree predicted higher plant quality for three genotypes at $NH_4NO_3 < 1.701 \times and KH_2PO_4 < 2.012 \times DKW$ medium (Driver and Kuniyuki 1984). Multiplication was predicted to be optimal at $NH_4NO_3 < 0.5 \times and CaNO_3 < 1.725 \times (Akin$ et al. 2016). The CHAID algorithm can include nominal variables within the analyses and provides common nutrient requirements for the tested genotypes but is not capable of performing multiple responses simultaneously such as the MARS algorithm. The CHAID algorithm constructs separate visual trees for each target variable, and the user must determine the common ranges for the tested outcomes (Akin et al. 2016). Medium improvements for wild apricot cultures from Kazakhstan (Kovalchuk et al. 2017) using the Classification and Regression Tree (CART) algorithm, mapped $KH_2PO_4 < 2.41 \times and MgSO_4 < 0.75 \times for$ plant quality, $0.77 \times < KH_2PO_4 < 0.91 \times for$ multiplication, $KH_2PO_4 < 2.41 \times and K_2SO_4 < 1.22 \times WPM$ (Lloyd and McCown 1980) for leaf color. In addition, the CART regression tree is not able to generate a multi-response tree but forms separate models for each target variable. The neurofuzzy logic technique described the $SO_4^{2-} \times Cl^{-}$ interaction and the Fe^{2+} , K^+ , NH_4^+ linear independent effects, as significant for quality of two pistachio rootstocks. $K^+ \times EDTA^- \times$ SO_4^{2-} and $Fe^{2+} \times BO_3^{-}$ interactions played crucial roles for multiplication (Nezami-Alanagh et al. 2018). Artificial neural networks can utilize both qualitative and quantitative variables and reveal complex non-linear effects. However, these methods do not provide a transparent mathematical equation that shows the relative importance of the inputs on the output (Olden et al. 2008). None of these techniques have been used to improve strawberry growth medium.

Murashige and Skoog (1962) medium with some modifications was used as the standard for multiplication of a wide range of in vitro cultures at the NCGR, however a large variation in the response indicated that it is not ideal for many species and cultivars (Reed and Hummer 1995). As seen in the current study, a medium with increased amounts of five major MS nutrients $(2 \times NH_4NO_3, 1.96 \times CaCl_2, 3 \times MgSO_4, 1.81 \times KNO_3$ and $1.94 \times KH_2PO_4$), is needed to provide improved growth for the three species tested without changing growth regulator concentrations. The next step in this study would be to test this medium on additional accessions to determine the applicability of this medium to large germplasm collections.

Conclusions

The MARS algorithm is a promising tool to advance understanding and prediction of biological events. This method can reveal non-linear and multidimensional relationships, and typically outcompete basic linear models, which makes it ideal to use for in vitro data. MARS is a non-parametric approach that is able to handle continuous and categorical data, without the need to satisfy the restrictive distributional assumptions required by the traditional parametric techniques. The MARS model clearly showed the direct effect and the complex interactions of the parameters. The target variables were affected by all of the tested mineral nutrients and the genotype input. This study was able to define the common major nutrient requirements of the tested strawberry species, due to the inherent characteristic of the MARS algorithm to incorporate nominal variables as predictors within the model. Thanks to the ability of MARS to analyze multiple responses simultaneously, the optimal major salts for good shoot quality, multiplication, and greener leaves were predicted to be: 3300 mg L⁻¹ NH₄NO₃, 862.4 mg L^{-1} CaCl₂, 1110 mg L^{-1} MgSO₄, 3439 mg L^{-1} NO₃ and $329.8 \text{ mg L}^{-1} \text{ KH}_2 \text{PO}_4.$

Author contributions MA planned and executed the experiment, collected the data, drafted the manuscript and assisted with analysis of the data. SPE helped with data analysis interpretation and writing the manuscript. EE performed the statistical analysis. BMR assisted with planning and analysis, supervised the study, and edited the manuscript.

Compliance with ethical standards

Conflict of interest The authors declare that there are no potential conflicts of interest regarding the research, authorship and publication of this manuscript.

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