

# Prognostic implications of the rapid recruitment of coronary collaterals during ST elevation myocardial infarction (STEMI): a meta-analysis of over 14,000 patients

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## Abstract

Acute coronary collateralisation of an infarct-related arterial (IRA) territory may be identified during angiography for ST elevation myocardial infarction (STEMI). Whether the presence or absence of these collaterals affects outcomes remains uncertain. A search of EMBASE, MEDLINE and Cochrane Library, using the Preferred Reporting Items for Systematic reviews and Meta-Analysis (PRISMA) guidelines was conducted to identify studies which reported on the association between coronary collaterals and in-hospital and longer term mortality, left ventricular ejection fraction (LVEF), risk of repeat acute myocardial infarction (AMI) and repeat revascularisation. Patients with Rentrop grade 0 or 1 were defined as poor collaterals whilst those with Rentrop grade two or three were defined as those with robust collaterals. Studies were eligible if they included patients  $\geq$  18 years of age who had immediate coronary angiography for STEMI. Included studies were observational which recorded the degree of collateral blood flow to the IRA. Two investigators reviewed all citations using a predefined protocol with final consensus for all studies, the data from which was then independently entered to ensure fidelity of results. Inverse variance random effects model for the meta-analysis along with risk of bias assessment was performed. 20 studies with a total of 14,608 patients were identified and included in the analysis. Patients with robust collaterals had lower mortality (OR 0.55, 95% CI 0.48–0.64), both in-hospital (OR 0.47, 95% CI 0.35–0.63) and longer term (OR 0.58, 95% CI 0.46–0.75). Patients with robust collaterals also had a higher mean LVEF (SMD 0.23, 95% CI 0.10–0.37). There was no difference in the rates of AMI or repeat revascularisation between patients with robust or poor collaterals. The presence of robust collaterals during STEMI is associated with reduced in-hospital and longer term mortality and improved left ventricular function. These findings have implications for prognostication and identifying patients who require close monitoring following STEMI.

Keywords Coronary collaterals · Rentrop · STEMI · Collateral

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# Highlights

- The presence of robust collaterals, visualised during ST elevation myocardial infarction (STEMI), perfusing the territory supplied by the occluded vessel, reduce risk of all-cause mortality.
- The presence of robust collaterals visualised during a STEMI reduce both short and long term mortality.
- The presence of robust collaterals visualised during a STEMI is associated with higher LVEF.
- The degree of collaterals does not predict risk of AMI or revascularisation.
- The molecular and cell signalling pathways by which collaterals are recruited should be further investigated.

## Introduction

During angiography for percutaneous coronary intervention (PCI) in the setting of ST elevation myocardial infarction (STEMI), the finding of coronary collaterals, perfusing the territory distal to the occluded, infarct related artery (IRA), is frequently identified. Whilst the presence of a bystander chronic total occlusion (CTO), with collaterals perfusing the pre-existing CTO is a well described predictor of poorer outcomes [1], the rapid recruitment of collaterals to perfuse the acutely occluded IRA is a separate and distinct entity. Whether the presence of these collaterals is associated with improved prognosis remains uncertain. We performed a systematic review and meta-analysis to examine the impact of rapidly recruited coronary collaterals on risk of mortality, left ventricular function, recurrent acute myocardial infarction (AMI) and repeat revascularisation, in patients presenting with STEMI.

## Methods

## Search strategy

We followed the Preferred Reporting Items for Systematic reviews and Meta-Analysis (PRISMA) guidelines to formulate and conduct the search strategy [2]. We performed a computerised systematic search through MEDLINE, EMBASE, and the Cochrane Library databases. The search was conducted in MEDLINE from 1946–January 10th 2020, EMBASE from 1974–January 10th 2020 and Cochrane Library from 2003–January 10th 2020 and then repeated until March 15th 2020 to identify newer published data. All relevant subject headings as well as free text terms relating to coronary collaterals and STEMI were used. The keywords were searched as text words as well as exploded medical subject headings when feasible. The search strategy for MED-LINE was as follows; (1) Collateral Circulation/(12,137), (2) collateral\*.mp. (45,969), (3) Rentrop.mp. (303), (4) CCS. mp (5746), (5) Collateral Connection Score.mp. (3099), (6) 1 or 2 or 3 or 4 or 5 (51,682), (7) ST Elevation Myocardial Infarction/ (3099), (8) STEMI.mp. (10,410), (9) 7 or 8, (10) 6 and 9 (120). The search strategy for EMBASE was (1) Collateral circulation/ (11,049), (2) collateral artery/(409), (3) coronary artery collateral circulation/(1756), (4) collateral\*. mp. (56,811), (5) Rentrop.mp. (569), (6) CCS.mp. (9630), (7) Collateral Connection Score.mp. (5), (8) 1 or 2 or 3 or 4 or 5 or 6 or 7 (66,357), (9) ST segment elevation myocardial infarction/ (36,557), (10) STEMI.mp. (27,144), (11) 9 or 10 (41,450), (12) 8 and 11 (434). The search strategy is shown in Appendix 1. We further screened references from the retrieved and included studies as well as prior metaanalyses for any relevant studies which were not retrieved through the initial search.

### Data source and searching

Articles were eligible for inclusion if they included patients  $\geq$  18 years of age who had immediate coronary angiography for STEMI, within 12 h of onset of chest pain or if evidence of ongoing ischaemia. Eligible study designs were ones which recorded the degree of collateral blood flow to the IRA prior to intervention. To standardise the degree of collaterals, we included studies, which quantified the collaterals based upon the Rentrop classification [3], where grade 0 = no filling of any collateral channel; Grade 1 = filling of the side branches of the infarct related artery; Grade 2=Partial filling of the epicardial vessel of the infarct related artery' Grade 3 = complete filling of the epicardial vessel. For the analysis, we grouped patients with Rentrop grade zero or one as poor collateral recruiters, whilst those with Rentrop grade two or three were classified as robust collateral recruiters, a dichotomy which was done in the majority of previously published studies, as this corresponds with the degree of perfusion of the occluded vessel.

Two investigators (U.A., D.N.) reviewed all citations identified through the literature search using a predefined protocol. Articles that clearly did not meet inclusion criteria were excluded at the title and/or abstract level. The remaining articles were selected for full text review. When limited information was available from the abstract, full text was always obtained. If further details were sought, the corresponding authors of the study were contacted to obtain further information. Disagreements regarding the selection and quality assessment of articles were resolved through discussion, and full consensus was achieved at each stage of review. Both investigators (U.A., D.N.) also independently entered results and outcomes into separate predetermined tables to ensure fidelity of results. Ethics approval was not deemed necessary given this was a meta-analysis of previously published data.

The outcomes of interest were all-cause mortality, left ventricular function, risk of recurrent AMI and need for repeat revascularisation. We further classified time course of mortality as in-hospital and short term, defined as  $\leq 6$  months follow up, or longer term mortality, defined as  $\geq 12$  months follow up. The Newcastle-Ottawa Scale (NOS) [4] was used to assess the quality of observational studies and risk of bias. The overall quality of the study was determined semi-quantitatively using the Agency for Healthcare Research and Quality (AHRQ) standards [5], whereby studies were graded of good quality, fair quality or poor quality, based upon scoring of the NOS.

#### Data analysis

For dichotomous outcomes, odds ratio was chosen to present data, whilst for continuous variables, the standardised mean difference (SMD) was used to present data. For continuous variables, the mean and standard deviation was recorded. If studies only presented medians and interquartile ranges, the data was transformed into mean with standard deviations using previously validated methods [6]. The program Review Manager (version 5.3) was used to conduct an inverse variance random effects model for the meta-analysis. The Cochrane Q-statistic  $(I^2)$  was used to assess the consistency among studies, with  $I^2 < 25\%$  considered low,  $I^2 > 50\%$ moderate, and  $I^2 > 75\%$  high heterogeneity [7]. Publication bias was estimated visually by funnel plots, where publication bias was considered unlikely if the plot resembles an inverted funnel, and was further tested by using the weighted regression test of Egger [7].

## Results

After removing duplicate studies, a total of 554 studies were screened of which, a final 20 studies with 14,608 patients were included in the analysis (Fig. 1). Twelve of the studies [8–19] were prospective observational studies, 6 [20–25] were retrospective observational, and 2 [26, 27] were retrospective analysis of randomised control trials of patients presenting with STEMI. One study included only patients  $\geq$  65 years [21], two included only patients with an anterior STEMI, whereby the left anterior descending artery was the IRA [14, 24], one included only patients with inferior STEMI [19] and one included only patients with cardiogenic shock complicating STEMI [18]. Included studies reported on rapidly recruited collaterals, perfusing the territory supplied by the IRA, distinct to preformed collaterals perfusing an existing CTO. All but two studies reported the degree of collateralisation using the Rentrop or equivalent scoring tool. In the study by Alsanjari et al. [8], collaterals were graded as "collaterals" or "no collaterals". Given the fact that patients with visible collaterals would almost certainly be Rentrop grade two or three, we included this group as robust collateral recruiters whilst those without collaterals were defined as poor collateral recruiters. In the study by Wang et al. [24], patients were grouped as those with no collaterals, defined as Rentrop grade zero, and those with collaterals, which were defined as Rentrop grade 1-3, of which 77% were grade one. For the purposes of this analysis, we characterised patients with visible collaterals as having robust collaterals. Two studies [8, 11] reported outcomes at 6 months, and longer term, with time specific data included in the analysis models.

Included studies were published between 1998 to 2020 with follow up periods between in-hospital alone to 5 years. The number of patients included in each study was between 96 to 3,340 with a mean age of 62.3. The percentage of patients with robust collaterals ranged between 10% and 41.3% with an average of 25.0% across all the studies. The thrombolysis in myocardial infarction (TIMI) flow of 0 in the infarct related artery at the time of angiography, whereby there is complete, persisting occlusion of the IRA, and is a strong predictor of collateral recruitment [20], ranged between 67.3% and 100%. The results of each study included in the analysis are summarised in Table 1. The NOS score for each study is presented in Supplementary Table 1, with 12 studies of a good quality, two fair quality and 6 of a poor quality.

#### Mortality

Patients with robust collateral recruitment had a lower risk of mortality compared to those with poor collateral recruitment (OR 0.55, 95% CI 0.48–0.64) throughout all included studies, with a very low degree of heterogeneity between studies ( $I^2 = 17\%$ , p = 0.24) (Fig. 2). Asymmetrical appearance of the funnel plot suggested that smaller studies, which do not show a significant mortality difference, were unpublished. Publication bias was supported by Eggers regression analysis (p < 0.01).

In the 11 studies assessing in-hospital and short term mortality, there was a lower rate of mortality in those patients with robust collaterals compared to those with poor collaterals, (OR 0.47, 95%CI 0.35–0.63) (Fig. 3). There was a low degree of heterogeneity between these studies ( $I^2=23\%$ , p=0.22). Publication bias was suggested by both an asymmetrical appearance of the funnel plot, whereby smaller studies which do not show a significant mortality difference were unpublished, and Eggers regression analysis (p<0.01).

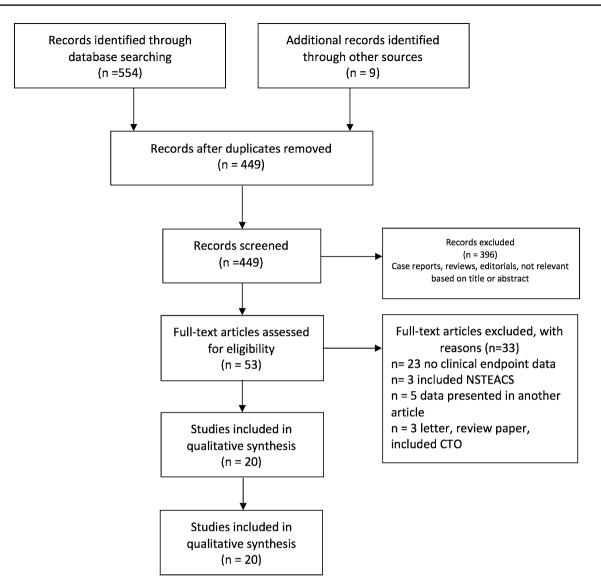


Fig. 1 Flow diagram of search strategy

In the 11 studies assessing longer term mortality, patients with robust collateral recruitment had a lower risk of mortality compared to patients with poor collateral recruitment (OR 0.58, 95% CI 0.46–0.75) (Fig. 4). In these studies, there was a very low degree of heterogeneity ( $I^2 = 22\%$ , p = 0.23), whilst there was no evidence of publication bias, with a symmetrical appearance of the funnel plot and Eggers regression analysis (p = 0.13).

As the study of Alsanjari et al. [8] and Hara et al. [11] accounted for 65.3% of the weighted analysis for mortality, sensitivity analysis was performed excluding these 2 heavily weighted studies. After their exclusion, patients with robust collateral recruitment had a lower risk of mortality compared to those with poor collateral recruitment, (OR 0.40, 95% CI 0.31–0.52, p < 0.0001) with no evidence of heterogeneity ( $I^2 = 0\%$ , p = 0.58) and no evidence of publication bias, with a symmetrical appearance of the funnel plot and by Eggers regression analysis (p = 0.35). Similarly, patients with robust collaterals had a lower risk of in-hospital and short term mortality compared to those with poor collaterals (OR 0.36, 95% CI 0.26–051,p < 0.0001) with no evidence of heterogeneity ( $I^2 = 0\%$ , p = 0.94) or evidence of publication bias with a symmetrical appearance of the funnel plot and by Eggers regression analysis (p = 0.47). Patients with robust collaterals had a lower risk of longer term mortality compared to those with poor collaterals (OR 0.47, 95% CI 0.28–0.78, p < 0.01) with low degree of heterogeneity ( $I^2 = 27\%$ , p = 0.20) and no evidence of publication bias, with a symmetrical appearance of the funnel plot and by Eggers regression (p = 0.34) (Supplementary Table 2).

Paper	Year Patie	ints Robust colla	Year Patients Robust collaterals Rentrop grade	Study type	Follow up	% Females	Age LAD IKA		TIMI 0 flow in IKA NOS qualit
Allahwala	2020 1625	345 (21.2%)	Yes	Retrospective observational	In-Hospital	371 (22.8%)	64.7	7 (47.2%)	767 (47.2%) 1093 (67.3%)*
Alsanjari	2019 1944	. 322 (17%)	No	Prospective observational	In-hospital & 5yrs	469 (24.1%) 65.1 748 (38.5%) 1783 (91.7%)	65.1 74	8 (38.5%)	1783 (91.7%)
Antoniucci	2002 1164	. 264 (23%)	Yes	Prospective observational	6 months	158 (22.1%)	63.8	8 (50.5%)	588 (50.5%) 950 (81.6%)*
Chu	2019 346	108 (31.2%	Yes	Retrospective observational	12 months	72 (20.8%)	67.7 15	67.7 156 (45.1%) N/A	N/A
Desch	2010 235	69 (29.4%)	Yes	Prospective observational	2.2 years	58 (24.7%)	65 11	115 (48.9%)	208 (88.5%)
Elsman	2004 1059	106 (10%)	Yes	Retrospective observational	12 months	205 (19.4%)	59.4 50	9 (48.1%)	$509~(48.1\%)$ 1059 $(100\%)^{*}$
Freund	2020 95	33 (34.7%)	Yes	Retrospective analysis of ran- domised control trial	4 years	26 (27.4%)	66 38	38 (40%)	90 (94.7%)
Hara	2016 3340	770 (23%)	Yes	Prospective observational	In-hospital & 5yrs	772 (23.1%)	65 14	1470 (44%)	3340~(100%)
Hernandez	2017 947	212 (22%)	Yes	Retrospective observational	2.4 years	186 (19.6%)	60.8 36	363 (38.4%)	880 (92.9%)
Kajiya	2014 96	33 (34.4%)	Yes	Prospective observational	4 years	31 (32.3%)	68.5 19	19 (29.7%)	N/A
Kim	2006 247	54 (21.9%)	Yes	Prospective observational	12 months	50 (20.2%)	59 11	118 (47.8%)	228 (92.3%)
Perez-castellano 1998	1998 180	32 (17.8%)	Equivalent	Prospective observational	In-Hospital	33 (18.3%)	64 18	180 (100%)	205 (86%)
Rechciński	2013 330	78 (23.6%)	Yes	Prospective observational	2 years	97 (29.4%)	56.2 12	122(37.0%)	217 (67.8%)
Sen	2017 1375	278 (20.2%)	Yes	Prospective observational	30 days	327 (23.8%)	57.4 57	573 (41.7%)	1375~(100%)*
Shen	2014 389	60 (15.4%)	Yes	Prospective observational	6 months	74 (19%)	63.6 19	63.6 197 (50.6%)	389 (100%)*
Sorajja	2007 318	119 (37.4%)	Yes	Retrospective analysis of ran- domised control trial	30 day & 6 months 65 (20.4%)	65 (20.4%)	60 12	0 (37.7%)	120 (37.7%) 272 (85.6%)
Valim	2011 105	22 (20.9%)	Yes	Prospective observational	In-hospital	36 (34.3%)	64.2 N/A	A'	N/A
Wang	2011 189	78 (41.3%)	No	Retrospective observational	12 months	32 (16.9%)	56.2 18	189 (100%)	N/A
Yaylak	2015 235	88 (37.4%)	Yes	Prospective observational	In-hospital	45 (19.1%)	55 0(	0 (0%)	210~(89.4%)

Paper	Year Patien	Year Patients Robust collaterals Rentrop gra	Rentrop grade	grade Study type	Follow up	% Females	Age LAD	% Females Age LAD IRA TIMI 0 flow in IRA NOS	IRA NOS
									quanty
Ying	2014 389	2014 389 60 (15.4%)	Yes	Retrospective observational	6 months	74 (19.0%) 63.6 197 (50.6%) N/A	63.6 197 (;	0.6%) N/A	8 Good
*TIMI 0 or 1 flow	×								

RA infarct related artery, LAD left anterior descending artery, N/A not applicable, NOS newcastle ottawa scale, TIMI thrombolysis in myocardial infarction. See supplementary Table 1 for complete newcastle ottawa scale for each study

# Left ventricular function

Four studies reported on rates of left ventricular impairment, which was defined as < 50% in three studies [8, 13, 20] and not defined in another [18]. There was no difference in rates of left ventricular impairment between those with robust collateral recruitment compared to those with poor collateral recruitment (OR 0.69, 95% CI 0.29-1.66) There was, however, a high degree of heterogeneity between studies ( $I^2 = 93\%$ , p < 0.0001) with no evidence of publication bias, with a symmetrical appearance of the funnel plot and by Eggers regression (p=0.12) (Fig. 5). A total of 13 studies reported mean left ventricular ejection fractions (LVEF), which was performed in-hospital in 11 studies [10, 15, 16, 19, 20, 22, 23, 25–28], and 6 months in two studies [9, 17]. Patients with robust collaterals had a significantly higher standard mean LVEF compared to those with poor collaterals (standard mean difference (SMD) 0.23, 95% CI 0.10–0.37). There was however a high degree of heterogeneity between studies ( $I^2 = 81\%$ , p < 0.0001) with no evidence of publication bias, with a symmetrical appearance of the funnel plot and by Eggers regression (p=0.41) (Fig. 6).

# AMI and repeat revascularisation

Ten studies analysed the effect of collateral recruitment on risk of recurrent AMI. There was no difference between patients with robust collaterals or poor collaterals on risk of AMI (OR 0.82, 95% CI 0.50-1.34). There was very low heterogeneity between studies ( $I^2 = 0\%$ , p = 0.94) with no evidence of publication bias, with a symmetrical appearance of the funnel plot and by Eggers regression (p=0.06)(Fig. 7).

Seven studies analysed the effect of collateral recruitment on risk of repeat revascularisation. There was no difference between patients with robust collaterals or poor collaterals on risk of repeat revascularisation (OR 0.99, 95% CI 0.73-1.34). There was very low heterogeneity between studies ( $I^2 = 0\%$ , p = 0.57) with no evidence of publication bias, with a symmetrical appearance of the funnel plot and by Eggers regression (p=0.15) (Fig. 8).

# Discussion

In this meta-analysis of 20 studies with more than 14,500 patients, we found that in those presenting with STEMI, the angiographic appearance of robust collaterals is associated with lower mortality, both in the short term and longer term, with an associated higher LVEF. More robust coronary collaterals, with partial, or complete retrograde perfusion of the occluded epicardial artery, allow oxygenated blood to perfuse the myocardial territory subtended by the IRA.

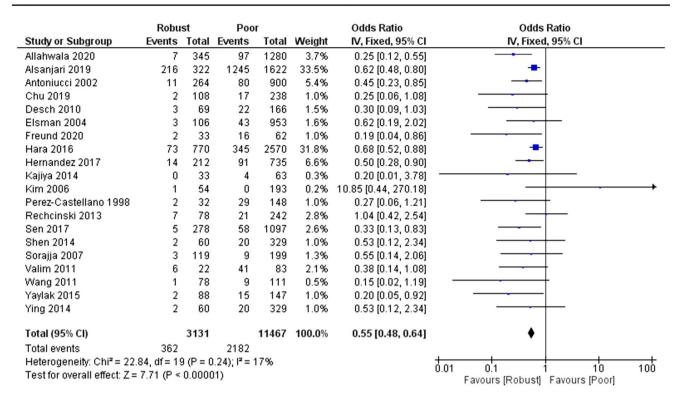


Fig. 2 Risk of mortality in patients with robust vs poor collaterals

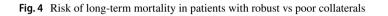
	Robu	st	Poo	r		Odds Ratio	Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl
Allahwala 2020	7	345	97	1280	10.6%	0.25 [0.12, 0.55]	<b>_</b>
Alsanjari 2019	12	322	114	1622	15.0%	0.51 [0.28, 0.94]	
Antoniucci 2002	11	264	80	900	13.8%	0.45 [0.23, 0.85]	
Hara 2016	44	770	182	2570	27.2%	0.80 [0.57, 1.12]	
Perez-Castellano 1998	2	32	29	148	3.5%	0.27 [0.06, 1.21]	
Sen 2017	5	278	58	1097	8.1%	0.33 [0.13, 0.83]	
Shen 2014	2	60	20	329	3.6%	0.53 [0.12, 2.34]	
Sorajja 2007	3	119	9	199	4.3%	0.55 [0.14, 2.06]	
Valim 2011	6	22	41	83	6.7%	0.38 [0.14, 1.08]	
Yaylak 2015	2	88	15	147	3.5%	0.20 [0.05, 0.92]	
Ying 2014	2	60	20	329	3.6%	0.53 [0.12, 2.34]	
Total (95% CI)		2360		8704	100.0%	0.47 [0.35, 0.63]	◆
Total events	96		665				
Heterogeneity: Tau <sup>2</sup> = 0.0	5; Chi² =	12.98, 0	df = 10 (P	= 0.22	); <b>I<sup>2</sup> =</b> 239	6	0.05 0.2 1 5 20
Test for overall effect: Z =	•	•					0.05 0.2 1 5 20 Favours (Robust) Favours (Poor)

Fig. 3 Risk of in-hospital and short term mortality in patients with robust vs poor collaterals

It is perhaps unsurprising that increased perfusion would result in improved LVEF. Whilst the rates of left ventricular impairment did not meet statistical significance, three of the four studies [8, 13, 18] did not report LVEF. In the studies which did report mean ejection fractions, this was significantly higher in patients with robust collaterals compared to those with poor collaterals. It is possible that a greater proportion of those patients with poor collaterals had more severe left ventricular impairment, which may explain these apparent conflicting findings. Whilst the absolute difference in LVEF was small, there was a sustained benefit derived from all studies suggesting the protective effects of collaterals is maintained.

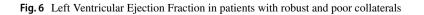
Perhaps driven by the protective effect on left ventricular function, mortality, both in-hospital and short term, as well as longer term was lower in those with robust

	Robu	st	Poo	r		Odds Ratio	Odds Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl	
Alsanjari 2019	216	322	1245	1622	32.7%	0.62 [0.48, 0.80]	-	
Chu 2019	2	108	17	238	2.6%	0.25 [0.06, 1.08]		
Desch 2010	3	69	22	166	3.7%	0.30 [0.09, 1.03]		
Elsman 2004	3	106	43	953	4.0%	0.62 [0.19, 2.02]		
Freund 2020	2	33	16	62	2.4%	0.19 [0.04, 0.86]		
Hara 2016	73	770	345	2570	32.1%	0.68 [0.52, 0.88]	-	
Hernandez 2017	14	212	91	735	13.2%	0.50 [0.28, 0.90]		
Kajiya 2014	0	33	4	63	0.7%	0.20 [0.01, 3.78]		
Kim 2006	1	54	0	193	0.6%	10.85 [0.44, 270.18]		<b></b> →
Rechcinski 2013	7	78	21	242	6.6%	1.04 [0.42, 2.54]		
Wang 2011	1	78	9	111	1.4%	0.15 (0.02, 1.19)		
Total (95% CI)		1863		6955	100.0%	0.58 [0.46, 0.75]	•	
Total events	322		1813					
Heterogeneity: Tau <sup>2</sup> =	0.03; Ch	i <sup>2</sup> = 12.	82, df = 1	0 (P = 0	0.23); I <sup>2</sup> =	22%	0.01 0.1 1 10	100
Test for overall effect:	Z = 4.28	(P < 0.0	001)				Favours [Robust] Favours [Cor	



	Robu	st	Poo	r		Odds Ratio		Odds Ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	IV, Random, 95% Cl		IV, Random, 95% Cl	
Allahwala 2020	127	338	843	1218	27.8%	0.27 [0.21, 0.34]		-	
Alsanjari 2019	62	116	278	530	26.8%	1.04 [0.70, 1.56]		-+-	
Kim 2006	21	54	89	193	24.9%	0.74 [0.40, 1.38]			
Valim 2011	7	22	21	83	20.4%	1.38 [0.49, 3.84]			
Total (95% CI)		530		2024	100.0%	0.69 [0.29, 1.66]		-	
Total events	217		1231						
Heterogeneity: Tau <sup>2</sup> =	0.69; Ch	i² = 40.1	10, df = 3	(P < 0.	00001); P	² = 93%	0.01	0.1 1 10 10	7
Test for overall effect:	Z = 0.82	(P = 0.4	11)				0.01	Favours [Robust] Favours [Poor]	U

	R	obust		1	Poor			Std. Mean Difference	Std. Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl	IV, Random, 95% Cl
Allahwala 2020	51.5	9	337	46.1	9.9	1217	9.5%	0.56 [0.43, 0.68]	
Antoniucci 2002	48.1	11.9	264	48.3	12.2	900	9.3%	-0.02 [-0.15, 0.12]	
Chu 2019	45.7	10.5	108	43.9	11.3	238	7.9%	0.16 [-0.07, 0.39]	+
Desch 2010	47.6	9.5	39	45.4	11.9	104	5.7%	0.19 [-0.18, 0.56]	
Elsman 2004	47	10	106	46	11	391	8.1%	0.09 [-0.12, 0.31]	- <del>-</del>
Freund 2020	48.7	11.8	26	41.4	8.2	47	4.2%	0.75 [0.26, 1.25]	
Hernandez 2017	51.7	11.2	212	48.3	11.1	735	9.1%	0.31 [0.15, 0.46]	
Rechcinski 2013	45.4	10.9	78	46.4	9.4	242	7.5%	-0.10 [-0.36, 0.15]	
Sen 2017	51.4	8.5	278	49.7	9	1097	9.4%	0.19 [0.06, 0.32]	
Shen 2014	49.4	3.4	60	48.8	6	329	7.1%	0.11 [-0.17, 0.38]	- <b>-</b>
Sorajja 2007	48.3	11.2	119	45	14.9	199	7.9%	0.24 [0.01, 0.47]	<b>⊢</b> ⊷−
Yaylak 2015	43.4	2.9	88	41.3	3.4	147	7.2%	0.65 [0.38, 0.92]	
Ying 2014	49.4	3.4	60	48.8	6	329	7.1%	0.11 [-0.17, 0.38]	
Total (95% CI)			1775			5975	100.0%	0.23 [0.10, 0.37]	•
Heterogeneity: Tau <sup>2</sup> =	= 0.04: CI	hi² = 6	4.20. dt	f = 12 (F	< 0.0	0001);	<sup>2</sup> = 81%		
Test for overall effect									-1 -0.5 0 0.5 1 Favours (Poor) Favours (Robust)



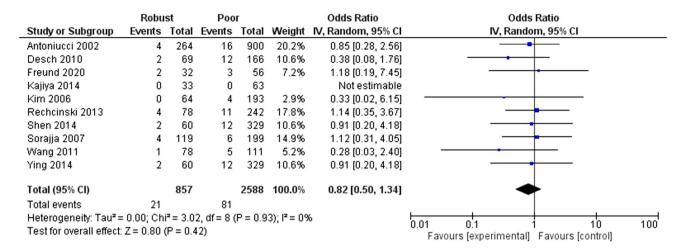


Fig. 7 Risk of recurrent acute myocardial infarction in patients with robust vs poor collaterals

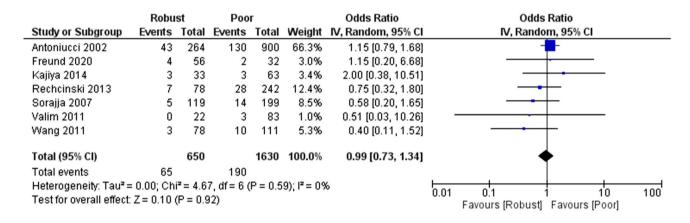


Fig. 8 Risk of repeat revascularisation in patients with robust vs poor collaterals

collaterals. Despite contemporary advancements in management of STEMI, as well as improved post infarct management, mortality rates at 1 year remain at 4.3–4.5% [29]. It is possible that the degree of intrinsic collateral recruitment may explain this remaining mortality penalty, despite optimal management and timely revascularisation. With respect to short-term mortality outcomes, there was high agreement amongst the studies. Furthermore, two studies [8, 11] reported both short and longer term outcomes. The study by Hara et al. [11] showed a trend toward reduced in-hospital mortality in those with robust collaterals (OR 0.80 95% CI 0.57–1.12), whilst at long term follow up of 5 years, there was a clear survival advantage (OR 0.68 95% CI 0.52–0.88). This delay may be explained by the protective effects of collaterals on left ventricular function, with subsequent longer term survival advantage [30]. Alsanjari et al. [8] showed a significant in-hospital mortality benefit which was sustained out to 5 years, suggesting the acute benefit of myocardial salvage persists over time. Even after excluding the studies by Alsanjari et al. [8] and Hara et al. [11] which accounted for a large percentage of weight for the analysis, the presence of robust collaterals remained predictive of total, short-term and long term mortality without evidence of heterogeneity or publication bias, suggesting a sustained and ubiquitous protective effect of collaterals on mortality.

The study by Kim et al. [13], was the only study which suggested a trend toward higher mortality with those with robust collaterals, although this did not reach significance. This may be an issue with respect to appropriate power, as the study was designed as an MRI based study, and found beneficial effects of robust collaterals with respect to smaller infarct size and smaller area at risk. Alternatively, the phenomenon of "collateral decay [31]" which has been described in the setting of a CTO, whereby robust collaterals seen at the time of an acute infarct are no longer evident on repeat angiography may conceivably apply in the STEMI setting to explain these apparent incongruous findings. There were no differences in risk of recurrent acute myocardial infarction or repeat revascularisation in patients with robust or poor coronary collaterals. Whilst the ability to recruit collaterals is stable in subsequent STEMI presentations [32], it appears that spontaneous collateral recruitment may be protective following onset of acute myocardial infarction rather than influencing risk of subsequent vascular events. Therefore, the innate ability to recruit collaterals appears to be protective once an infarct has commenced rather than preventative, as it does not preclude acute plaque rupture and thromboembolic occlusion.

The incidence of angiographically evident robust coronary collaterals was on average 25% of patients presenting with STEMI. Whilst the mechanisms by which some patients are able to recruit collaterals is uncertain, a combination of patient specific factors (younger age), anatomical considerations (right coronary artery as the culprit vessel) as well as history of angina and presence of more severe bystander coronary disease, mediated through multiple growth factors and cytokines have been previously postulated [20]. Given the protective effects of collateral recruitment, further research is required to ascertain whether adjuvant pharmacological or mechanical treatment may facilitate in their recruitment, to improve outcomes.

The findings of the study are similar to a previous metaanalysis on the impact of collaterals in STEMI [28], although our meta-analysis includes six more studies with 4000 more patients Furthermore, the previous analysis included one study which appraised the effect of collaterals to a concomitant CTO during STEMI, a quite distinct situation to rapidly recruited collaterals to the IRA during STEMI, which may introduce some degree of bias in the results. This confirmation of results, further emphasizes the need to identify biological and mechanistic basis for collateral recruitment.

## Limitations

Despite the relative consistent findings between studies of our analysis, there are important limitations which need to be considered. Firstly, all included studies were observational studies, including two which were based on the initial angiographic findings in randomised control trials, prior to intervention, and presented unadjusted outcomes, which may introduce bias into the selection of data although the very nature of spontaneously recruited collaterals suggests that only observational studies are possible. The baseline demographics were relatively similar; however the IRA was variable and in some studies, excluded specific IRA (51,923). Furthermore patient's background, in particular prior history of angina and or coronary artery disease, particularly disease in the contralateral donor vessel and prior history of angina were not considered and as such may have influenced the degree of collateral recruitment and hence outcomes. Although these collaterals were supplying the territory subtended by an acutely occluded vessel in the setting of a STEMI, it is possible that they were preformed rather than acutely recruited. However, given that the majority of the culprit lesions in patients with STEMI have a mild to moderate degree of stenosis prior to occlusion [33, 34], this is unlikely to be the case. Given included studies ranged from 1998 to 2020, it is conceivable that patient characteristics and management strategy changes may have influenced outcomes, which is an inherent limitation of pooled results. However, whilst studies have shown that patients with STEMI are younger and more likely to have traditional cardiovascular risk factors as compared to 20 years ago, there has been no change in mortality in patients who underwent reperfusion therapy [35]. Given the protective effects of collaterals is likely in the short term prior to revascularisation, it is unlikely however that there is significant bias from including these studies. Another anatomical consideration was that included studies looked at the presence of retrograde filling via contralateral collaterals, and the impact of the presence of anterograde, bridging collaterals was not taken into consideration. Whilst the majority of observed collaterals are from the contralateral vessel, nevertheless this may impact on coronary perfusion and may independently affect prognosis. Finally, it is possible, that there may have been insufficient statistical power to detect associations for some outcomes.

## Conclusions

The ability to recruit robust coronary collaterals during a STEMI is associated with a lower in-hospital, short term and longer term mortality as compared to those patients who cannot recruit sufficient collaterals. Similarly, robust collateral recruitment is associated with improved left ventricular function following STEMI, which may be the mechanism by which this survival advantage is achieved. These findings have implications in identifying patients who may benefit from closer monitoring and prognostication in the post infarct setting.

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#### **Compliance with ethical standards**

**Conflicts of interest** None of the authors have relevant conflicts of interest.

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