

Does public sector efficiency matter? Revisiting the relation between fiscal size and economic growth in a world sample

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Abstract This paper revisits the relationship between fiscal size and economic growth. Our work differs from the empirical growth literature because this relationship depends explicitly on the efficiency of the public sector. We use a sample of 64 countries, both developed and developing, in four five-year time periods between 1980 and 2000. Building on the work of Afonso et al. (Public Choice 123:321–347, 2005), we construct a measure of public sector efficiency in each country and each time period by calculating an output-to-input ratio. In addition, we get an estimate of technical efficiency of public spending for 52 countries from 1995 to 2000 by employing a stochastic frontier analysis. Using these two measures, we find evidence of a non-monotonic relation between fiscal size and economic growth that depends critically on the size-efficiency mix.

Keywords Fiscal policy · Government efficiency · Growth

JEL Classification H1 · E6 · 04

1 Introduction

The relationship between government size and economic growth is not expected to be monotonic. On one hand, governments provide public goods and services and correct market failures. On the other hand, policy intervention generates its own costs, as it requires taxes and distorts incentives. There is thus a tradeoff depending on the size-efficiency mix of the public sector. By efficiency, we mean the ability of the government to transform its revenues

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into public goods and services that benefit the economy and promote growth. After a critically large size, or a critically low efficiency, the costs of a larger public sector outweigh the benefits.¹

This paper revisits the relation between fiscal size and economic growth. Our work differs from the empirical growth literature because this relation depends explicitly on the efficiency of the public sector. We use a sample of 64 countries, both developed and developing, in four five-year periods from 1980 to 2000.

We follow the methodology of Afonso et al. (2005) for the OECD and construct measures of public sector efficiency (PSE). This index measures the cost-effectiveness of the public sector in reaching a range of objectives of government intervention. It is basically the ratio of performance indicators (output) to a measure of public expenditure related to those indicators (input), based on the assumption that the input is used to achieve that output. We construct such indexes of public sector efficiency for four policy areas: administration, stabilization, infrastructure and education. In addition to this measure, focusing on 52 countries for the sub-period 1995–2000 during which more data are available, we also obtain an estimate of the so-called technical efficiency (TE) of the public sector by applying a stochastic production frontier analysis (see e.g., Kumbhakar and Lovell 2000; Greene 2005). The ranking of countries according to the TE measure does not differ substantially from that implied by the PSE measure.

We then incorporate these two measures (PSE or TE) into a simple econometric model in which the size-growth relationship is non-monotonic depending on the size-efficiency mix. This novel feature is included into an otherwise standard growth regression (see e.g., Barro and Sala-i-Martin 2004, Chap. 12).

Our main finding is that, when fiscal size is measured by the government consumption share in GDP, the size-efficiency mix is significant in explaining the size-growth relationship. The latter is indeed non-monotonic as discussed above. This result holds for both efficiency measures constructed and is robust to a number of changes in the econometric specification, as well as to dividing the world sample into two sub-samples consisting of “high-income” and “developing” countries. Among other things, the model provides an endogenously determined efficiency threshold below (resp. above) which the size-growth relationship is negative (resp. positive). In general, this relationship is found to be negative in most countries and time periods. When we use, for instance, the PSE as a measure of efficiency in our world sample for all four five-year periods, our estimates imply that only in 34 out of 159 observations (different countries in different periods) is the size-growth relationship positive.²

Our results imply that what really matters to growth is not government size per se, but the size-efficiency mix. They can also help to explain why the evidence on the growth effects of the overall fiscal size has so far been mixed (see e.g., Levine and Renelt 1992; Tanzi and Zee 1997; Gemmel and Kneller 2001; Mueller 2003). Essentially, our results suggest that

¹ A simple and popular conceptual framework is provided by Barro’s (1990) model, where there is a tradeoff between growth-promoting public goods and the distorting taxes required to finance them. When the government size and its associated tax burden are high (resp. small) relative to the public sector’s efficiency, a larger size is bad (resp. good) for growth. See also Hillman (2003) and Mueller (2003) on the market failures versus policy distortions tradeoff.

² Regarding the causal effect of fiscal size on economic growth, a concern has been the potential endogeneity of fiscal size. The literature so far has not provided a “credible” identification of fiscal size in growth regressions (see e.g., Agell et al. 2006). Although the aim of our paper is not to resolve the causality issue, we also provide some evidence that it can be easier to find a credible identification of the size-efficiency mix, rather than of size alone, in growth regressions.

it is difficult to obtain a “robust” effect of the overall fiscal size on economic growth when important elements that shape the size-growth relationship (in our case, the efficiency of the public sector) are omitted from the analysis.³ In sum, as Levine and Renelt (1992, p. 951) point out, “using simple expenditure data without accounting for government efficiency may yield inaccurate measures of the actual delivery of public services”.

The rest of the paper proceeds as follows. Section 2 develops measures of government efficiency. Section 3 studies the growth effects of the size-efficiency mix. Conclusions are in Sect. 4.

2 Measures of government efficiency

In this section, we present two measures of government efficiency.

2.1 Public sector efficiency

Following Afonso et al. (2005, 2006), we construct sub-indices of relative Public Sector Efficiency (PSE) in certain policy areas in each country and each time period, and then take the average of these sub-indices to obtain an index of aggregate government efficiency in each country and each time period.

Afonso et al. have constructed PSEs for seven policy areas for OECD countries over the eighties and nineties. Here, we focus on four policy areas (education, administration, infrastructure and stabilization) for 64 countries, both industrialized and developing, and four five-year time periods from 1980 to 2000 (obviously, due to data availability, there is a tradeoff between the number of countries and the number of policy areas).⁴ We keep only those observations for which indexes of government efficiency in all four areas are available.

Since the methodology is laid out in Afonso et al. (2005, 2006), here we only discuss the basic insight and point out where we differ. The basic insight of this methodology is to compare the performance of government in certain areas of economic activity (where these areas are influenced directly by government intervention) to the associated expenditure that the government allocates to achieve this particular performance. Thus, to construct a PSE index, we need a measure of Public Sector Performance (PSP) and a measure of the associated Public Sector Expenditure (PEX) for each country in each policy area and each time period. Then, the PSE will be the ratio of PSP to PEX. More details on the construction of PSP and PSE indexes in each policy area are in our [Appendix](#).

To make these PSP and PEX measures (expressed in different units of measurement) comparable across countries, we follow Afonso et al. by expressing each country’s PSP and PEX relative to the average PSP and PEX of all countries in each period, and this is done for all periods and indexes. In other words, each country’s PSP and PEX are expressed as percentages of the respective average (normalized to be 1), and in turn the PSE is obtained

³An additional potential explanation that has received a lot of empirical support is that the overall size of government cannot capture the different implications of different government activities. As has been shown (see e.g., Devarajan et al. 1996; Kneller et al. 1999; Angelopoulos et al. 2007), the growth effects of the different components of government expenditure, as well as of the various types of tax instruments, are not the same. See also Angelopoulos and Philippopoulos (2007) for a single country, time-series study that also supports the result that both the composition and efficiency of the government matter.

⁴Greene (2005) has measured the efficiency of public spending in developing countries focusing on the areas of health and education. Afonso et al. (2006) have also constructed measures of public sector efficiency for a group of 24 upper-middle income countries for the late nineties.

as the ratio of these relative PSPs and PEXs.⁵ Therefore, the resulting PSE is an index that measures the efficiency of a country's government relative to governments in other countries in each period in a particular policy area. The larger the value, the more efficient the country's government is. This is the notion of *relative efficiency* in Afonso et al.

Table A.1 in the Appendix reports the relative PSPs, and the resulting PSEs, in the four policy areas for the countries and time periods for which data are available. The order of countries is alphabetical. The second-from-the-end column in Table A.1 reports the (relative) aggregate efficiency of a country's government obtained as the average of the four (relative) sub-indices. As expected, high-income OECD countries get on average better scores, although the public sectors in nations like South Korea, Thailand or Malaysia appear to be particularly efficient. The most efficient governments during 1995–2000 are those of South Korea (2.221), Canada (2.039), the USA (1.938) and Switzerland (1.813) which are twice as efficient as the average countries, e.g., the United Kingdom or France. At the bottom end, Namibia (0.483), Nicaragua (0.447) and Yemen (0.35) score about half of the average score.⁶

Of course, we have to be cautious with these estimates. For instance, in rich countries, like Finland or Sweden, the cost of resources used for providing public education or capital is higher than in say Uruguay or Lebanon, and this may result in an overestimation of relative efficiency in the latter group of countries. In addition, government performance in a certain policy area may be overestimated when private resources complement government policy; this is the case of education in many countries (e.g., Greece).

In sum, the main advantage of the above output-to-input approach is its simplicity and logical coherence, which allow a meaningful comparison across countries. Its main weakness is that several assumptions have to be made to calculate such a composite index (for a critical assessment of different methodologies and measures of public sector efficiency, see e.g., Afonso et al. 2005, 2006; and the special issue of *European Economy*, no. 3, 2004, on “Public finances in EMU 2004”).

2.2 A stochastic production frontier methodology

As an alternative approach to measuring government efficiency, we estimate a stochastic production frontier for the public sector and then obtain an estimate of the so-called Technical Efficiency (*TE*) of this sector. For a review of this methodology, see e.g., Kumbhakar and Lovell (2000).

Our stochastic frontier model is of the form:

$$\ln y_i = \beta_0 + \beta_1 \ln x_i + v_i - u_i, \quad (1)$$

⁵Since the averages of PSP and PEX are both normalized to be 1, the resulting PSE has an average around 1 (specifically, the PSEs in education and stabilization have an average of about 1.1, whereas the PSEs in infrastructure and administration have an average of about 1.25).

⁶Two countries score suspiciously high in this table. Paraguay, which seems to be the most efficient country in the world, and Argentina, which seems to be the second most efficient country in the last time period. Regarding Paraguay, this result is driven by a very high score in the variable *Electric Power Transmission and Distribution Losses* (see the Appendix), which results in a very high PSP in infrastructure. This score may reflect measurement errors or unusual circumstances, so we drop Paraguay from our regressions in the next section. Regarding Argentina, the high efficiency score for 1995–2000 is probably due to the extended stabilization program implemented by the country in this period. We also choose not to include Argentina in our analysis in the next section. We report, however, that including these two countries does not have a significant effect on the econometric results presented later.

where y_i is a measure of public sector output in country i , x_i is a measure of public sector input, u_i is the nonnegative technical inefficiency component of the error term, and v_i is the noise component assumed to be distributed normally and independently of u_i . Both error components are assumed to be independent of the regressors.

After estimating (1) by maximum likelihood, a measure of technical efficiency for each country i (TE_i) is defined as:

$$TE_i = E[\exp\{-u_i\}/\varepsilon_i], \quad (2)$$

where $\varepsilon_i = v_i - u_i$ (see Kumbhakar and Lovell 2000, Chap. 3). This efficiency score is bounded between zero and one.

To apply the above, we need to measure public sector outputs and inputs (y_i and x_i , respectively). We use the average of the PSP indices as a measure of y_i . As a measure of x_i , we use *Total Government Expenditure* (as a share of GDP) which is available from the World Development Indicators. We estimate (1)–(2) under the assumption that u_i is characterized by a nonnegative half-normal distribution (we have also examined the case where u_i is assumed to follow a truncated normal distribution but, since this gives similar results, we discuss only the nonnegative half-normal case).

Results for each country's technical government efficiency (TE_i) during the 1995–2000 subperiod (where we again look at a five-year period average, as we did with the *PSE* measure above) are reported in Table A.2 in the Appendix.⁷ The ranking results look sensible again. In this cross-section world sample during 1995–2000, Switzerland's government scores the best, followed by Sweden and Finland. Again, as probably expected, governments in OECD countries are more efficient than those in developing countries, although public sectors in fast-growing economies like Thailand, Malaysia, Cyprus and especially South Korea get high scores. Algeria, Nicaragua and Yemen have now the least efficient governments. Therefore, the ranking of countries using the *PSE* measure does not differ substantially from that using the *TE* measure (recall that this refers to the 1995–2000 period during which both measures are available) with the correlation coefficient being 0.75.

In this sample, an LR test of the null that $\sigma_u^2 = 0$ gives a value of 5.64, which rejects it (the respective p -value of the test is 0.009).⁸ This implies that government technical efficiency differs significantly across countries during 1995–2000. We report that we have also estimated government *TE* during the three time periods before 1995 (i.e. the three five-year periods between 1980 and 1995). However, there are significantly less data available for these earlier years (especially in the 1980s when the sample size drops to around 25 to 30, i.e., it mainly consists of the OECD countries). Not surprisingly, we have not been able to reject the null $\sigma_u^2 = 0$ for any of these early periods. Hence, concerning the *TE* measure, we concentrate on the 1995–2000 subperiod.

In sum, the *TE* measure has obvious advantages but, on the other hand, it depends on the assumptions made about the error term. The assumption that government expenditure is uncorrelated with the error term may be strong when governments respond to negative

⁷To examine whether the TE_i estimates in Table A.2 are not biased due to heteroscedasticity in either v_i or u_i (see Kumbhakar and Lovell 2000), we have tested whether the variance functions of v_i or u_i depend (linearly) on *govexp*. Since this is rejected, we can have some faith in the homoscedasticity assumption.

⁸The limiting distribution of the LR test statistic is a mixture of a chi-square with zero degrees of freedom, i.e. a point mass at zero, and a chi-square with 1 degree of freedom (see e.g., Kumbhakar and Lovell 2000). The p -value of the test reported here takes this into account.

shocks by increasing their expenditures. In any case, the *TE* measure of government efficiency provides a useful alternative measure also used below to check the importance of the size-efficiency mix.

3 The size-efficiency nexus matters to growth

This section tests whether there is a non-monotonic relationship between government size and economic growth with this relationship driven by the size-efficiency mix.

3.1 Econometric model

We use the above-constructed measures of government efficiency (*PSE* or *TE*) in a growth regression of the following form (see Dutt and Mitra 2002 for a similar specification in a trade policy context):

$$growth_{it} = \alpha_0 + \alpha_1 size_{it} + \alpha_2 size_{it} * eff_{it} + X_{it}\beta + \varepsilon_{it}, \quad (3)$$

where $growth_{it}$ is the growth rate of country i at time t , $size_{it}$ is a measure of government size, eff_{it} is a measure of government efficiency (*PSE* or *TE*) and X_{it} includes control variables usually included in growth regressions (see below).

The partial derivative with respect to $size_{it}$ is simply:

$$\frac{\partial growth_{it}}{\partial size_{it}} = \alpha_1 + \alpha_2 eff_{it}, \quad (4)$$

where we expect α_2 to be positive in the sense that the more efficient the public sector, the larger the positive effect of government on growth. We also expect α_1 to be negative to catch the adverse effects of government size on growth.

As long as the estimated coefficients α_1 and α_2 in (3) are statistically significant and have the right signs, so that the size-efficiency nexus matters to growth, the above specification can also give an estimate of a (common to all countries) critical level of efficiency, eff^* , where $eff^* \equiv -(\alpha_1/\alpha_2) > 0$ makes the partial in (4) equal to zero. When an individual country's efficiency, eff_{it} , is higher (resp. lower) than eff^* , the positive (resp. negative) effects dominate and the country is placed on the positively (resp. negatively) sloped part of the size-growth curve; this, of course, requires eff^* to lie within the range of values of eff_{it} in the data. Note that (3)–(4) imply that the growth effects of fiscal size can differ among countries and time periods.⁹

3.2 Data and variables used in the regressions

For the eff variable, we use the two measures of government efficiency (*TE* and *PSE*) constructed in Sect. 2 above. The rest of the variables are as in most of the literature. We work with five-year period averages as we did with our eff measures (five-year periods are also used in the growth literature, especially the literature on the growth effects of fiscal policy,

⁹We have also examined a specification like $growth_{it} = \alpha_0 + \alpha_1 size_{it} + \alpha_2 size_{it}^2 + X_{it}\beta + \varepsilon_{it}$, which gives a partial as a function of $size$, so that an “optimal” size can be calculated given the estimated coefficients irrespectively of efficiency. We report that estimation of this equation does not give meaningful results (coefficients are not significant and in some regressions they have wrong signs).

see e.g., Folster and Henrekson 2001; Kneller et al. 1999). The main datasets used are the Penn World Tables (PWT) version 6.1 (see Heston et al. 2002) and the World Development Indicators (WDI) developed by the World Bank.

Our dependent variable, the growth rate of per capita GDP, is from the PWT. In particular, the PWT dataset provides us with the real GDP per capita in constant prices, which is then used to obtain the five-year average of annual growth rates (denoted as *growth* in our regressions). The PWT also provides us with consumption of the general government as a share of GDP in constant prices, which is averaged over five-year periods to give a variable denoted as *govshare* in our regressions. This will be our primary measure of government size.¹⁰ An alternative measure of government size, which is also used below, is total expenditures of the central government as a share of GDP (denoted as *govexp* in our regressions and obtained from WDI). This variable includes transfers and interest payments on public debt, in addition to government consumption (to avoid double counting, we do not include government investment in our *govexp* measure, as government investment is included in the investment share in GDP used as a separate regressor).

Concerning the above two measures of fiscal size, an advantage of *govshare* over *govexp* is that it refers to the general government and can thus capture better the full thrust of fiscal size on economic growth; moreover, it is PPP adjusted and therefore more suitable for international comparisons. The advantage of *govexp*, on the other hand, is that it allows us to examine whether including more types of government expenditure (at the disadvantage of using data at the central level only) gives different results regarding the effect of fiscal size on growth. Ideally, we would like to have a measure of general government spending for all types of government expenditure, but unfortunately, such a measure does not, as far as we know, exist for all the countries and time periods in our world sample. Finally, the fiscal size of government can be also measured by tax revenue or the budget balance, both as shares of GDP (see e.g., Tanzi and Zee 1997; Persson and Tabellini 2003); see below in Sect. 3.4 for details.¹¹

In our choice of the control variables included in X in (3) above, we will follow most of the literature (see e.g., Barro and Sala-i-Martin 2004, Chap. 12, and the review papers mentioned above). Thus, we use the logarithm of the initial level of GDP per capita (denoted as *lgdp*), obtained from PWT, to control for convergence effects; the initial (or the value closest to the beginning of period) secondary school enrollment rate (denoted as *enroll*), obtained from WDI, to proxy for human capital;¹² the investment share of GDP (denoted as *investment*), obtained from PWT and averaged over the five-year period; the logarithm of the fertility ratio (denoted as *fertility*), obtained from WDI; a measure of openness (denoted as

¹⁰This is the general government consumption component of GDP. It does not include public investment, interest payments, subsidies and other transfers. Public investment is included in PWT in the variable “investment share in GDP” (see below). Note however that a large part of government spending on goods and services, included in *govshare*, has investment features (e.g., salaries of teachers, professors and doctors and spending on police or the judiciary system). The variable *govshare* is closer to what Tanzi and Schuknecht (2000) refer to as a measure of “real government expenditure”.

¹¹The tax revenue-to-GDP ratio is generally not preferred to fiscal spending measures, mainly because of tax evasion problems (see e.g., Tanzi and Zee 1997). The same can be said about the budget-to-GDP ratio since it includes tax revenue.

¹²A better proxy for human capital could be a measure of the average years of schooling (see e.g., Barro and Sala-i-Martin 2004). However, such measures are not available for all the countries in our sample and we do not want to restrict our sample for any other reasons than the requirements for the efficiency measure. Hence, we use the *enroll* variable, also used by Levine and Renelt (1992).

openness), obtained from PWT and defined as the sum of exports and imports over GDP.¹³ Finally, we include in our regressions time dummies, as well as regional dummies for countries in Sub-Saharan Africa, East Asia, Latin America and the economies in transition.

3.3 Basic results

Results using the PSE measure of efficiency for the sample of 64 countries from 1980 to 2000 are presented in Table 1. We report standard errors obtained under the assumption of spherical errors and standard errors that are robust to arbitrary heteroscedasticity and arbitrary intra-country serial correlation (see e.g., Wooldridge 2002). The first three columns in Table 1 report estimates when using *govshare* as a measure of fiscal size and the last three when using *govexp*.¹⁴

In column (1) of Table 1, we start with a standard growth regression: the coefficient of *govshare* is significantly negative. In column (2), we add the PSE measure of government efficiency, which is positive but marginally significant, while the coefficient of *govshare* remains significantly negative. To examine whether it is government efficiency that shapes the size-growth relationship, we move to column (3), which presents results for our key equation (3) above.¹⁵ Both estimates of *govshare* and *govshare * eff* are significant with the expected sign (negative and positive respectively), indicating a heterogeneous across countries size-growth relationship depending on government efficiency. Actually, the estimates imply a threshold of $eff^* = 1.358$, which means that only in 34 out of 159 observations (different countries in different time periods), is the size-growth relationship positive.

The estimated coefficients α_1 and α_2 also allow us to calculate the growth effect of fiscal size in each country and each time period, as implied by (4). Results are reported in the last column of Table A.1 in the Appendix. As can be seen, the estimated effect differs substantially across countries. There is a small group of countries where public sectors are efficient meaning a positive growth effect from fiscal size. This group includes Canada, Japan, South Korea and Switzerland in all time periods for which we have data; and Australia, Finland and the USA in most time periods (here we report those countries with more than one observation/time period; see Table A.1 for all countries). However, for most countries and time periods, this effect is negative. Therefore, the general picture that emerges is that fiscal sizes have grown too much—relative to public sector efficiency—in the last decades. This finding is similar to the arguments made in e.g., Gwartney et al. (1998) and Tanzi and Schuknecht (2000), although these papers do not take account of efficiency explicitly.

Regarding the control variables that enter significantly, *lgdp* is negative, implying (conditional) convergence, while *investment* and *openness* are positive. The effect of *fertility* is negative (see also Barro and Sala-i-Martin 2004, Chap. 12) but not robustly significant. The effect of *enroll* is positive but not significant. Regarding the regional dummies, those for the economies in transition are significantly negative, while those for Latin American countries

¹³We have also used the average annual growth rate of the labour force, obtained from the WDI, in the growth regressions, but it is always insignificant.

¹⁴We do not include a dummy for each country (and thus we do not estimate fixed effects regressions) as this would result in losing all cross-country variation. This is important because the measure of efficiency developed here is a *relative* one across countries. It would make little sense to use this variation to explain differences *within* countries only.

¹⁵We do not include *eff* together with *size * eff* in the same regression, as they are highly correlated and as a result both *eff* and *size * eff* become insignificant. In this specification, the growth effect of government efficiency takes place only via government size, assuming that efficiency is independent of size.

Table 1 Growth regressions using PSE: 62 countries, 1980–2000

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
Growth rate						
<i>Govshare</i>	-0.052 [0.023]** (0.027)*	-0.054 [0.022]** (0.026)**	-0.106 [0.031]** (0.037)**	-	-	-
<i>Govshare * pse</i>	-	-	0.078 [0.030]** (0.039)**	-	-	-
<i>Govexp</i>	-	-	-	-0.049 [0.023]** (0.025)*	-0.037 [0.028] (0.028)	-0.053 [0.025]** (0.028)*
<i>Govexp * pse</i>	-	-	-	-	-	0.009 [0.028] (0.027)
<i>Pse</i>	-	0.919 [0.477]* (0.567)	-	-	0.471 [0.576] (0.596)	-
<i>Lgdp</i>	-2.108 [0.503]** (0.720)**	-2.392 [0.519]** (0.671)**	-2.325 [0.501]** (0.636)**	-1.736 [0.479]** (0.686)**	-1.879 [0.576]** (0.664)**	-1.804 [0.521]** (0.670)**
<i>Investment</i>	0.119 [0.039]** (0.045)**	0.107 [0.039]** (0.047)**	0.118 [0.038]** (0.047)**	0.109 [0.039]** (0.049)**	0.106 [0.040]** (0.049)**	0.108 [0.040]** (0.050)**
<i>Enroll</i>	0.015 [0.014] (0.020)	0.016 [0.014] (0.019)	0.019 [0.014] (0.017)	0.025 [0.015] (0.021)	0.024 [0.015] (0.021)	0.025 [0.015] (0.021)

Table 1 (Continued)

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
<i>Fertility</i>	-1.677 [0.813]** (1.012)	-1.522 [0.810]** (1.003)	-1.275 [0.814] (0.985)	-2.136 [0.810]** (1.041)**	-2.017 [0.824]** (1.064)*	-2.078 [0.830]** (1.067)*
<i>Openness</i>	0.011 [0.004]** (0.005)**	0.013 [0.004]** (0.005)**	0.012 [0.004]** (0.005)**	0.011 [0.004]** (0.005)**	0.011 [0.004]** (0.005)**	0.011 [0.004]** (0.005)**
<i>Sub-Saharan Africa</i>	-0.227 [0.996] (0.706)	-0.204 [0.987] (0.690)	-0.051 [0.980] (0.683)	-0.372 [0.992] (0.708)	-0.414 [0.995] (0.720)	-0.365 [0.995] (0.704)
<i>East Asia</i>	-0.838 [0.720] (0.916)	-1.794 [0.869]** (1.101)	-1.629 [0.773]** (0.883)*	-1.156 [0.784] (0.944)	-1.449 [0.863] (1.085)	-1.207 [0.801] (0.978)
<i>Latin America</i>	-0.873 [0.523]* (0.751)	-0.994 [0.522]* (0.739)	-0.956 [0.515]* (0.678)	-0.801 [0.519] (0.737)	-0.797 [0.520] (0.738)	-0.729 [0.522] (0.729)
<i>Transition economies</i>	-3.601 [0.742]** (0.974)**	-3.478 [0.738]** (1.011)**	-3.329 [0.736]** (0.984)**	-3.280 [0.756]** (1.191)**	-3.289 [0.757]** (1.170)**	-3.240 [0.768]** (1.211)**
<i>Constant</i>	19.327 [4.878]** (6.248)**	20.898 [4.902]** (5.908)**	20.269 [6.609]** (5.685)**	16.342 [4.592]** (5.898)**	16.826 [4.635]** (5.763)**	16.801 [4.807]** (5.840)**
R^2	0.378	0.394	0.405	0.374	0.377	0.375

Notes: 1. The estimation method is Least Squares. The sample consists of 62 countries, in five-year periods over 1980–2000. There is a total of 159 observations. All regressions include time dummies. 2. Standard errors obtained under the assumption of spherical errors are shown in brackets below the estimated coefficients. Standard errors that are robust to arbitrary heteroscedasticity and arbitrary intra-country serial correlation are shown in parentheses. 3. An asterisk denotes significance at the 10% level and two asterisks at the 5% level

are negative but not significant when we use robust standard errors. An interesting result is the negative dummy for East Asian countries, as this variable usually has a positive effect in similar regressions (see e.g., Barro and Sala-i-Martin 2004, Chap. 12). However, East Asian countries, in general, are ranked highly in our efficiency measures (see Table A.1 above), so that a large part of the positive regional effect has been already controlled for by our fiscal measure.

The results are less clear when we enter the other widely used measure of fiscal size, *govexp* (see the last three columns in Table 1). The coefficient of *govexp* is negative but not robustly so (see column (5) that includes *pse*). More importantly, in column (6), there is no significant evidence of a nonlinear relationship like the one found in column (3); namely, the coefficient of *govexp* * *pse* is not significant (although it has the right sign). Recall that the key difference between *govshare* and *govexp* is that the latter includes redistributive transfers and interest payments on public debt. Both items (i.e., transfers and interest payments) do not involve a direct use of real resources by the state sector (recall the economy's resource constraint). We thus do not find it surprising that *govexp* does not give as clear results as *govshare*. In a sense, these new results indicate that both the size-efficiency mix and the composition of government expenditure matter to growth.¹⁶

3.4 Robustness of basic results

We now examine the robustness of the basic results presented above by extending the empirical specification in two dimensions. First, we test whether our results—regarding the importance of the size-efficiency mix on growth—are sensitive to the financing assumption of government spending (see e.g., Miller and Russek 1997; Kneller et al. 1999). Given that we do not have detailed tax and spending data for all countries and time periods in our sample, we use a general form of government budget that equates aggregate spending to tax revenue and deficit (see e.g., Miller and Russek 1997). In principle, in the absence of Ricardian equivalence, the effect of spending on growth can be different depending on whether higher spending is financed by more tax revenues or by a larger budget deficit (higher debt). If, for instance, we include a measure of taxation, together with spending, in a growth regression, we would expect the effect of the tax measure to be negative capturing the adverse implications of a larger fiscal size, whereas the effect of the spending measure to be positive capturing the positive effects of e.g. more public good provision. It is therefore interesting to see whether our results are robust to the inclusion of a finance instrument (because of multicollinearity problems, we cannot include both tax revenues and public deficits in the regressions).

For our sample, we obtain data for tax revenues, as a share of GDP, from the WDI database (we denote the respective measure, which is again expressed in five-year period averages, as *tax*). We then rerun the basic regressions of Table 1 by including *tax* as an additional

¹⁶We have also used another potential measure of the extent of government involvement in the economy, the so-called Economic Freedom index as developed by the Fraser Institute (see e.g., Gwartney et al. 2006). The Economic Freedom (EF) index is a rather general measure of government involvement than includes the size of government; the degree of regulation of credit, labor and business by the government; the legal structure; the security of property rights; the freedom to trade, etc. We report that, when we use the EF index as a measure of fiscal size in our regressions for the world sample (i.e., instead of *govshare* and *govexp*), then (a) it has a negative growth effect (see also De Haan et al. 2006) although this effect is not always significant (b) the estimated α_2 is not significant in (3)–(4) above. We believe this is not surprising given that this index contains more variables than the size of the government, while equations like (3) test whether the growth effect of size depends on the size-efficiency mix. Besides, the EF index may be correlated with government efficiency.

explanatory variable. Results for the main variables are shown in Table 2 (since the estimates for the control variables are not generally affected, we do not include them in Table 2 to save on space—these results are available upon request). As can be seen, the results of Table 1 remain essentially unchanged when we include *tax*, which, itself, is not significant. We report that these results again do not change if we use deficits instead of taxes.¹⁷

Second, we also test whether the inclusion of lagged growth rates changes our results. Although our basic specification (see Table 1) is common in the empirical growth-policy literature working with five-year averages (see e.g., Kneller et al. 1999; Folster and Henrekson 2001), dynamic effects from past growth may persist even after five years. Therefore, we now examine whether the size-efficiency mix retains its significance in explaining economic growth, even after controlling for lagged growth rates (see also Miller and Russek 1997). Results obtained from including the lagged-once growth rate (denoted as *grolag* in our regressions) as an explanatory variable in the regressions of Table 1 are reported in Table 3 (again, we present results for the main variables only to save on space). Note that the sample size drops from 159 to 98 observations (there are now 46 instead of 62 countries). The lagged-once growth rate is generally significant, but the results for the main variables of interest are not qualitatively affected. Actually, in column (3), where we present our key results using *govshare*, *grolag* is not found to be significant.

Finally, in Table 4, we present results for the main variables by including both *grolag* and *tax* in our regressions. As can be seen, the previous results and analysis remain robust to this specification as well.

3.5 High-income and developing countries

So far—although we allowed for the effect of fiscal size to differ across countries depending on the efficiency of the public sector in each country—we have studied rich and developing countries jointly in a single sample. We now divide countries into two subgroups to study whether the size-efficiency mix matters differently in high income and developing countries (where we classify countries as high income following the classification in the WDI dataset). For each group, we first calculate the measure of public sector efficiency (PSE) separately, repeating the steps described in Sect. 2.1 above (since the efficiency measure is re-constructed for more homogeneous groups of countries, this can provide an additional robustness test).

Using these new *PSE* measures, Tables 5 and 6 rerun the basic regressions of Table 1 for high income and developing countries respectively (again, we present results for the main variables only to save on space). As can be seen, the results remain practically unchanged for the subgroup of high-income countries in Table 5. For the subgroup of developing countries in Table 6, the main story, regarding the importance of the size-efficiency mix, is again supported when we use *govshare* as a measure of fiscal size (see column (3) in Table 6),

¹⁷Notice, when we compare Tables 1 and 2, that the inclusion of *tax* does not alter the negative effects of *govshare* and *govexp* in columns (1) and (4) respectively. Thus, the effect of government size itself, as measured by *govshare* or *govexp*, remains negative even if we add a measure of the tax burden, *tax*. This is probably because tax revenues, as an ex post measure, is not an ideal proxy for the distortions imposed by the tax system; higher tax revenue may e.g. reflect less tax evasion and better institutions (see Tanzi and Zee 1997; Angelopoulos et al. 2007). Thus, the basic size-efficiency specification in Table 1 appears to be good enough to capture the tradeoffs in fiscal policy at least in our sample. In other words, to the extent that we allow the effect of the fiscal size to depend on the size-efficiency mix, we view our basic specification as an alternative to including both spending (see positive effects) and taxation (see negative effects) to capture the tradeoff in fiscal policy.

Table 2 Growth regressions using PSE (controlling for *tax*): 62 countries, 1980–2000

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
Growth rate						
<i>Govshare</i>	-0.057 [0.023]**	-0.056 [0.023]**	-0.103 [0.032]**	-	-	-
	(0.026)*	(0.026)**	(0.041)**			
<i>Govshare * pse</i>	-	-	0.072 [0.035]**	-	-	-
			(0.048)			
<i>Govexp</i>	-	-	-	-0.067 [0.037]*	-0.060 [0.037]	-0.078 [0.042]*
				(0.038)*	(0.037)	(0.049)
<i>Govexp * pse</i>	-	-	-	-	-	0.016 [0.029]
						(0.032)
<i>Pse</i>	-	0.748 [0.594]	-	-	0.656 [0.609]	-
		(0.699)			(0.693)	
<i>Tax</i>	-0.039 [0.026]	-0.015 [0.032]	-0.011 [0.029]	0.024 [0.040]	0.039 [0.042]	0.032 [0.042]
	(0.030)	(0.038)	(0.036)	(0.050)	(0.055)	(0.056)
R^2	0.388	0.395	0.405	0.376	0.381	0.377

Notes: 1. The estimation method is Least Squares. The sample consists of 62 countries, in five-year periods over 1980–2000. There is a total of 159 observations. All regressions include time dummies, regional dummies and the control variables of the regressions in Table 1. 2. Standard errors obtained under the assumption of spherical errors are shown in brackets below the estimated coefficients. Standard errors that are robust to arbitrary heteroscedasticity and arbitrary intra-country serial correlation are shown in parentheses. 3. An asterisk denotes significance at the 10% level and two asterisks at the 5% level

Table 3 Growth regressions using PSE (controlling for *grolag*): 46 countries, 1985–2000

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
Growth rate						
<i>Govshare</i>	-0.071 [0.033]**	-0.082 [0.033]**	-0.165 [0.053]**	-	-	-
	(0.038)*	(0.036)**	(0.049)**			
<i>Govshare * pse</i>	-	-	0.103 [0.0474]**	-	-	-
			(0.057)			
<i>Govexp</i>	-	-	-	-0.054 [0.031]*	-0.047 [0.036]	-0.055 [0.035]
				(0.024)**	(0.031)	(0.030)*
<i>Govexp * pse</i>	-	-	-	-	-	0.001 [0.034]
						(0.039)
<i>Pse</i>	-	1.120 [0.679]	-	-	0.282 [0.796]	-
		(0.674)			(0.837)	
<i>Grolag</i>	0.198 [0.104]*	0.182 [0.103]*	0.131 [0.106]	0.225 [0.104]**	0.223 [0.105]**	0.225 [0.105]**
	(0.139)	(0.135)	(0.124)	(0.132)*	(0.131)*	(0.137)
<i>R</i> ²	0.458	0.475	0.487	0.448	0.448	0.448

Notes: 1. The estimation method is Least Squares. The sample consists of 46 countries, in five-year periods over 1985–2000. There is a total of 98 observations. All regressions time dummies, regional dummies and the control variables of the regressions in Table 1. 2. Standard errors obtained under the assumption of spherical errors are shown in brackets below the estimated coefficients. Standard errors that are robust to arbitrary heteroscedasticity and arbitrary intra-country serial correlation are shown in parentheses. 3. An asterisk denotes significance at the 10% level and two asterisks at the 5% level

Table 4 Growth regressions using PSE (controlling for *tax* and *grolag*): 46 countries, 1985–2000

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
Growth rate						
<i>Govshare</i>	-0.074 [0.034]** (0.036)**	-0.079 [0.033]** (0.036)**	-0.190 [0.058]** (0.061)**	-	-	-
<i>Govshare * pse</i>	-	-	0.140 [0.057]** (0.077)*	-	-	-
<i>Govexp</i>	-	-	-	-0.151 [0.049]** (0.056)**	-0.150 [0.049]** (0.056)**	-0.207 [0.063]** (0.072)**
<i>Govexp * pse</i>	-	-	-	-	-	0.054 [0.034] (0.049)
<i>Pse</i>	-	1.585 [0.868] (1.026)	-	-	1.373 [0.847] (0.940)	-
<i>Tax</i>	-0.011 [0.033] (0.030)	-0.035 [0.041] (0.046)	0.043 [0.039] (0.043)	0.127 [0.051]** (0.063)**	0.168 [0.057]** (0.069)**	0.168 [0.059]** (0.077)**
<i>Grolag</i>	0.198 [0.104]* (0.139)	0.176 [0.104]* (0.138)	0.107 [0.108] (0.125)	0.208 [0.101]** (0.128)	0.191 [0.101]* (0.129)	0.213 [0.101]** (0.135)
<i>R</i> ²	0.458	0.479	0.495	0.485	0.501	0.497

Notes: 1. The estimation method is Least Squares. The sample consists of 46 countries, in five-year periods over 1985–2000. There is a total of 98 observations. All regressions include time dummies, regional dummies and the control variables of the regressions in Table 1. 2. Standard errors obtained under the assumption of spherical errors are shown in brackets below the estimated coefficients. Standard errors that are robust to arbitrary heteroscedasticity and arbitrary intra-country serial correlation are shown in parentheses. 3. An asterisk denotes significance at the 10% level and two asterisks at the 5% level

Table 5 Growth regressions using PSE (high income countries): 26 countries, 1980–2000

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
Growth rate						
<i>Govshare</i>	-0.065 [0.028]** (0.038)*	-0.057 [0.027]** (0.032)*	-0.173 [0.042]** (0.050)**	-	-	-
<i>Govshare * pse</i>	-	-	0.124 [0.038]** (0.040)**	-	-	-
<i>Govexp</i>	-	-	-	-0.079 [0.016]** (0.020)**	-0.072 [0.022]** (0.024)**	-0.080 [0.016]** (0.020)**
<i>Govexp * pse</i>	-	-	-	-	-	0.004 [0.023] (0.019)
<i>Pse</i>	-	1.161 [0.358]** (0.345)**	-	-	0.211 [0.465] (0.314)	-
<i>R</i> ²	0.599	0.649	0.649	0.673	0.674	0.674

Notes: 1. The estimation method is Least Squares. The sample consists of 26 countries, in five-year periods over 1980–2000. There is a total of 85 observations. All regressions include time dummies and the control variables of the regressions in Table 1. 2. Standard errors obtained under the assumption of spherical errors are shown in brackets below the estimated coefficients. Standard errors that are robust to arbitrary heteroscedasticity and arbitrary intra-country serial correlation are shown in parentheses. 3. An asterisk denotes significance at the 10% level and two asterisks at the 5% level

Table 6 Growth regressions using PSE (developing countries): 36 countries, 1980–2000

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
Growth rate						
<i>Govshare</i>	-0.022 [0.035]	-0.010 [0.036]	-0.069 [0.041]*	-	-	-
	(0.037)	(0.038)	(0.049)			
<i>Govshare * pse</i>	-	-	0.078 [0.037]**	-	-	-
			(0.044)*			
<i>Govexp</i>	-	-	-	-0.026 [0.036]	0.012 [0.047]	-0.056 [0.039]
				(0.040)	(0.045)	(0.046)
<i>Govexp * pse</i>	-	-	-	-	-	0.090 [0.050]*
						(0.057)
<i>Pse</i>	-	1.127 [0.848]	-	-	1.362 [1.072]	-
		(1.131)			(1.327)	
<i>R</i> ²	0.191	0.213	0.243	0.192	0.212	0.238

Notes: 1. The estimation method is Least Squares. The sample consists of 36 countries, in five-year periods over 1980–2000. There is a total of 74 observations. All regressions include time dummies and the control variables of the regressions in Table 1. 2. Standard errors obtained under the assumption of spherical errors are shown in brackets below the estimated coefficients. Standard errors that are robust to arbitrary heteroscedasticity and arbitrary intra-country serial correlation are shown in parentheses. 3. An asterisk denotes significance at the 10% level and two asterisks at the 5% level

which is calculated as in the world sample above. It is interesting to note that, in developing countries, public expenditure is not significantly related to economic growth in the first two columns, but significance is restored in column (3) which explicitly allows for the size-efficiency mix. All this suggests that in both subgroups, our story—that the size-efficiency mix matters—is confirmed by the data.

We finally report that these results are robust to the inclusion of *tax* as an explanatory variable (see Sect. 3.4 above). On the other hand, including *grolag* reduces the sample size in both subgroups too much to give any reliable results.

3.6 Can the size-efficiency mix help with endogeneity?

When looking for a causal effect from fiscal policy in a growth regression, a usual concern is that there might be a reverse causality when e.g. governments respond to negative shocks by increasing their expenditures (see e.g., Tanzi and Zee 1997; Agell et al. 2006). Although this problem is to some extent mitigated here since we work with five-year averages, such reverse causality cannot be ruled out. In addition, our fiscal size variables, and especially the measure of government efficiency, may be correlated with the error term due to omitted variables or measurement error.

The natural approach to dealing with such an endogeneity is to use instruments for the endogenous variables in IV methods. A fundamental concern with IV regression methods, however, is whether the instruments are valid and relevant. As far as we know, the relevant literature has not yet provided a credible identification of fiscal policy so that the instruments used are both exogenous and strongly correlated with the endogenous variables (see e.g., Agell et al. 2006). We now investigate whether accounting for the size-efficiency mix can help in this direction. We will build upon the basic specification of Sect. 3.3.

We need instruments for *size* and *size * eff* in 2SLS regressions. As such instruments, we use variables usually considered as potential determinants of fiscal policy (see e.g., Persson and Tabellini 2003, Chap. 3). In particular, we use the age dependency ratio (*agedep*) and two measures of country size (population and surface area, denoted respectively as *pop* and *surface*). All these three variables are obtained from WDI and, except for *surface*, are averaged over the five-year periods. In Table 7, we present results for the core variables when we re-estimate the basic regressions of Table 1 by using these instruments in 2SLS methods (the results for the control variables do not change significantly, so we do not present them to save on space).

We start again with the *govshare* variable. When we do not account for efficiency (column (1) in Table 7), the Sargan over-identifying restrictions test rejects the null that the instruments are uncorrelated with the error term. However, when efficiency is included as an endogenous variable, either on its own (column (2)) or multiplicatively with *govshare* (column (3)), the null clearly cannot be rejected (the *p-value* is very low in both cases). Therefore, in this sample, the instruments affect growth only indirectly through the size-efficiency mix. Note also that the Anderson (1984) canonical correlations, and the Cragg and Donald (1993) tests of whether the equation is under-identified, reject the null, thus lending some support to the relevance of the instruments.¹⁸ More importantly, the first-stage *F*-statistic is very high for the *govshare * eff* variable, which indicates that the instruments are strongly correlated with this variable. Although the first-stage *F*-statistic for *govshare* is not as large, it is clear that the diagnostics favor the key regression in column (3) that controls for the size-efficiency mix. In this regression (in column (3)), the critical $eff^* = 1.238$

¹⁸These tests have been implemented using the routines written by Baum et al. (2006).

Table 7 Growth regressions using PSE: 2SLS for 62 countries, 1980–2000

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
Growth rate						
<i>Govshare</i>	-0.126 [0.067]* (0.088)	-0.160 [0.082]* (0.096)*	-0.208 [0.071]** (0.081)**	-	-	-
<i>Govshare * pse</i>	-	-	0.168 [0.053]** (0.054)**	-	-	-
<i>Govexp</i>	-	-	-	-0.096 [0.050]* (0.059)	0.028 [0.094] (0.123)	-0.225 [0.106]** (0.107)**
<i>Govexp * pse</i>	-	-	-	-	-	0.314 [0.198] (0.205)
<i>Pse</i>	-	4.826 [1.902]** (1.775)**	-	-	4.880 [2.848] (2.718)	-
Sargan over-identification test	$\chi^2_{(2)} = 9.513$ (0.008)	$\chi^2_{(1)} = 0.110$ (0.739)	$\chi^2_{(1)} = 0.029$ (0.865)	$\chi^2_{(2)} = 9.894$ (0.007)	$\chi^2_{(1)} = 4.144$ (0.041)	$\chi^2_{(1)} = 2.902$ (0.088)
Cragg-Donald under-identification	$\chi^2_{(3)} = 20.78$ (0.000)	$\chi^2_{(2)} = 15.12$ (0.000)	$\chi^2_{(2)} = 19.88$ (0.000)	$\chi^2_{(3)} = 42.17$ (0.000)	$\chi^2_{(2)} = 8.96$ (0.011)	$\chi^2_{(2)} = 5.56$ (0.061)

Table 7 (Continued)

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
Anderson canonical correlations	$\chi^2_{(3)} = 19.53$ (0.000)	$\chi^2_{(2)} = 14.44$ (0.000)	$\chi^2_{(2)} = 18.73$ (0.000)	$\chi^2_{(3)} = 37.40$ (0.000)	$\chi^2_{(2)} = 8.72$ (0.012)	$\chi^2_{(2)} = 5.469$ (0.064)
First-stage F (gov)	$F(3, 143) = 6.23$	$F(3, 143) = 6.23$	$F(3, 143) = 6.23$	$F(3, 143) = 12.64$	$F(3, 143) = 12.64$	$F(3, 143) = 12.64$
First-stage F (pse)	-	$F(3, 143) = 4.91$	-	-	$F(3, 143) = 4.91$	-
First-stage F (gov * pse)	-	-	$F(3, 143) = 11.95$	-	-	$F(3, 143) = 3.59$

Notes: 1. The estimation method is 2SLS. The sample consists of 62 countries, in five-year periods over 1980–2000. There is a total of 159 observations. All regressions include time dummies, regional dummies and the control variables of the regressions in Table 1. 2. Standard errors obtained under the assumption of spherical errors are shown in brackets below the estimated coefficients. Standard errors that are robust to arbitrary heteroscedasticity and arbitrary intra-country serial correlation are shown in parentheses. 3. An asterisk denotes significance at the 10% level and two asterisks at the 5% level. 4. The instruments used are: agedep, pop, surface. 5. The Sargan test is a test of over-identifying restrictions. Under the null, the test statistic is distributed as chi-squared in the number of over-identifying restrictions (the p -value is reported in parentheses). 6. The Anderson (1984) canonical correlation is a likelihood-ratio test of whether the equation is identified. The Cragg and Donald (1993) test statistic is also a chi-squared test of whether the equation is identified. Under the null of underidentification, the statistics are distributed as chi-squared with degrees of freedom = $(L - K + 1)$ where L = number of instruments (included + excluded) and K is the number of regressors (the p -values are reported in parentheses). 7. The 1st stage F -statistic tests the hypothesis that the coefficients on all the excluded instruments are zero in the 1st stage regression of the endogenous regressor on all instruments (the p -value is reported in parenthesis)

implies that in 46 countries/periods there is a positive effect on growth from *govshare*. The fact that the critical efficiency level is lower in the 2SLS regressions indicates that the estimate of fiscal size is biased downwards when endogeneity is not accounted for, so that the “true” effect of fiscal size may in fact be less negative (or more positive) than implied in Table 1 for many countries.

As in Table 1 above, the results are not so promising when we use the *govexp* variable as a measure of government size. Although the Sargan test does not reject the validity of the instruments, the Anderson (1984) canonical correlations and the Cragg and Donald (1993) tests cannot reject the null that the equation in column (6) of Table 7 is under-identified.

Therefore, although further research is clearly required concerning the issue of causality in the fiscal policy-growth relation in cross-country growth regressions, our results suggest that taking account of the size-efficiency mix can help in identifying the growth effects of fiscal policy.

3.7 An alternative measure of government efficiency

To further examine the robustness of our results, we replace *PSE* with the *TE* measure of efficiency. Again, we build upon the basic specification of Sect. 3.3.

As explained in Sect. 2, we have been able to construct the *TE* measure for the 1995–2000 period only. In Table 8, we present results focusing on this period. Actually, in this table, we report results for both the *PSE* and *TE* indices of government efficiency, and both the *govshare* and *govexp* measures of fiscal size. This has the additional advantage of checking whether there has been a structural break in the size-efficiency-growth relationship of (3). The regressions in Table 8 are the same as those in Table 1, except that now we do not include time dummies.

We start again with *govshare* (columns (1)–(3)). The average effect of *govshare* is negative (column (1)), while the size-efficiency mix (when we use the *PSE* measure for efficiency) is important (column (2)). In other words, the nonmonotonic relationship holds for both the whole period and the 1995–2000 subperiod. The critical level of efficiency is now $eff^* = 1.216$, which implies that for 24 out of 51 countries in this sample the size-growth relationship is positive. Note also that the regression with the size-efficiency mix is much better than the regression without it, as can be seen by both the increase in R^2 and the fact that the coefficients of *lgdp*, *openness* and *East Asia* become significant. Regarding *lgdp*, in particular, this implies that the size-efficiency mix is an important long-run determinant of economic growth that has to be conditioned on so that convergence can be captured in the data (see e.g., Barro and Sala-i-Martin 2004, Chap. 12).

Then, we estimate (3) for the 1995–2000 sample by using *TE* as the efficiency measure. The results are in column (3). The coefficients are again significant with the right signs. The critical efficiency is now $eff^* = 0.889$, which implies that only in 8 out of 51 countries in this period is the size-growth relationship positive (see the last column in Table A.2 in the Appendix for the estimated growth effect in each country in this case). These are Finland, South Korea, Sweden and Switzerland, as well as (but only marginally) Canada, Germany, Iceland and Uruguay. Note, however, that the regression with the *PSE* measure in the size-efficiency mix explains about 10% more of the variation in the growth rate than the regression with the *TE* measure.

In columns (4)–(6) of Table 8, we rerun the same regressions by using *govexp* as a measure of government size. As before, *govexp* is negative and significant, while the *size * eff* variables have a positive sign but are not significant.

As we did in Table 7, we have also run 2SLS regressions for the equations in Table 8 by using the same set of instruments for the size-efficiency mix. The estimated coefficients

Table 8 Growth regressions using PSE and TE: OLS for 51 countries, 1995–2000

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
Growth rate						
<i>Govshare</i>	-0.088 [0.049]* (0.048)*	-0.202 [0.055]** (0.055)**	-0.225 [0.088]** (0.083)**	-	-	-
<i>Govshare * pse</i>	-	0.166 [0.050]** (0.046)**	-	-	-	-
<i>Govshare * te</i>	-	-	0.253 [0.137]* (0.130)*	-	-	-
<i>Govexp</i>	-	-	-	-0.138 [0.055]** (0.049)**	-0.148 [0.047]** (0.056)**	-0.187 [0.089]** (0.088)**
<i>Govexp * pse</i>	-	-	-	-	0.047 [0.060] (0.045)	-
<i>Govexp * te</i>	-	-	-	-	-	0.081 [0.113] (0.098)
<i>Lgdp</i>	-1.143 [1.027] (1.344)	-1.185 [0.943]* (1.094)*	-1.899 [1.079]* (1.275)	-0.676 [0.908] (0.992)	-0.940 [0.972] (1.078)	-0.994 [1.015] (1.091)
<i>Investment</i>	0.040 [0.076] (0.093)	0.040 [0.068] (0.089)	0.024 [0.074] (0.091)	0.035 [0.073] (0.095)	0.031 [0.074] (0.093)	0.026 [0.075] (0.094)

Table 8 (Continued)

Dep. variable	(1)	(2)	(3)	(4)	(5)	(6)
<i>Enroll</i>	0.015 [0.027]	0.029 [0.025]	0.014 [0.026]	0.046 [0.030]	0.039 [0.031]	0.043 [0.304]
<i>Fertility</i>	(0.034) 0.025	(0.028) 0.502	(0.033) -0.069	(0.030) -1.388	(0.032) -1.087	(0.031) -1.409
<i>Openness</i>	[1.679] (1.359)	[1.508] (1.260)	[1.632] (1.384)	[1.528] (1.105)	[1.582] (1.158)	[1.538] (1.122)
	0.012 [0.008]**	0.019 [0.008]**	0.012 [0.008]**	0.018 [0.009]**	0.017 [0.009]**	0.018 [0.009]**
	(0.010) -0.929	(0.009)** -0.535	(0.011) -1.262	(0.010)* -0.348	(0.010)* -0.347	(0.011)* -0.247
<i>Sub-Saharan Africa</i>	[1.803] (0.960)	[1.617] (1.000)	[1.761] (0.974)	[1.776] (0.877)	[1.785] (0.905)	[1.793] (0.906)
<i>East Asia</i>	-1.903 [1.709]	-3.923 [1.645]**	-2.774 [1.726]	-4.179 [2.020]**	-3.895 [2.061]**	-4.122 [2.034]**
	(1.814) -1.191	(1.893)** -1.470	(1.934) -1.341	(1.768)** -1.316	(1.875)** -1.181	(1.966)** -1.199
<i>Latin America</i>	[0.980] (1.039)	[0.881] (0.903)	[0.956] (1.056)	[0.942] (0.958)	[0.962] (0.922)	[0.976] (0.950)
<i>Transition economies</i>	-3.617 [1.608]**	-3.854 [1.440]**	-3.738 [1.564]**	-4.104 [1.551]**	-3.871 [1.586]**	-3.768 [1.629]**
	(1.705)** 12.004	(1.590)** 16.093	(1.697)** 19.240	(2.063)** 8.906	(2.025)* 10.497	(2.029)* 12.026
<i>Constant</i>	[9.003] (10.020)	[8.146]* (7.823)	[9.588] (9.638)	[7.980] (7.423)	[8.257] (7.951)*	[8.986] (8.143)*
<i>R</i> ²	0.291	0.446	0.347	0.335	0.346	0.344

Notes: 1. The estimation method is OLS. The sample consists of 51 countries, 1985–2000. There are 51 observations. 2.-3. As in Table 1

are again supportive of the importance of the size-efficiency nexus, at least for the *govshare* measure, but the first-stage regression diagnostics reveal that the instruments are not strongly correlated with the endogenous variables. Since the small sample size does not help us to draw any safe conclusions, we find the results of Table 7 to be more reliable. In any case, as discussed above, the identification of fiscal policy remains a challenge in this literature. Finally, we report that with the *TE* measure of efficiency, we cannot divide countries into rich and developing, as we did in Sect. 3.5 (the subsamples are now too small). Concerning the addition of *tax* in the regressions (as we did in Sect. 3.4 above), we report that once more the main results are not affected.

Therefore, the main result from this subsection is that the relationship between the size-efficiency mix and economic growth is robust to the time period and the measure of government efficiency used.

4 Concluding remarks

We revisited the relationship between fiscal size and economic growth and provided evidence that this relationship depends on the size-efficiency mix of the public sector. The policy implication is that what matters to growth is not the government's size per se, but its size-efficiency mix. Of course, improving the efficiency of the public sector is not an easy task. It requires, among other things, the reallocation of government resources, as well as the effective and efficient use of those resources towards identified and transparent strategic priorities.

The measurement of government efficiency is still an open issue. The measures developed here, although plausible, cannot be treated as definitive. Future research may provide alternative measures to test the robustness of our results. Further research is also needed to investigate the causal effects of fiscal policy on growth in cross-country regressions. We nevertheless believe that we have contributed to these important policy issues.

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Appendix: Construction of PSE

We construct measures of public sector efficiency (PSE) for 64 countries, in four five-year periods between 1980 and 2000, as output-to-input ratios by working as in Afonso et al. (2005). Afonso et al. have focused on OECD countries, where the available data cover both government performance and the associated public expenditure. Although we have tried to follow Afonso et al. in the choice of policy areas and variables used, the construction of such a rich PSE for a broader group of countries runs into data limitations, especially when looking for decomposed public expenditure data. Thus, some deviations from the variables used by Afonso et al. are inevitable. Nevertheless, the variables used here are the same in spirit.

In the policy area of education, the PSP can be measured by the variable *Secondary School Enrollment*, while the associated PEX is the average of the variable *Public Spending*

Table A.1 Public Sector Efficiency (PSE) in 64 countries, 1980–2000

Country	Period	PSP-Admin	PSP-Educ	PSP-Infrastr	PSP-Stabil	PSE-Admin	PSE-Educ	PSE-Infrastr	PSE-Stabil	PSE-average	Growth effect
Algeria	1990–1995	0.557	0.793	0.797	0.258	0.466	0.648	0.749	0.242	0.526	-0.065
Algeria	1995–2000	0.586	0.803	0.452	0.233	0.522	0.73	0.459	0.237	0.487	-0.068
Argentina	1995–2000	0.592	1.013	0.528	2.092	1.763	1.366	1.043	4.136	2.077	-
Australia	1980–1985	1.266	1.108	0.972	0.78	2.374	1.048	1.357	1.089	1.467	0.008
Australia	1985–1990	1.309	1.169	1.015	0.552	2.053	1.079	1.296	0.705	1.283	-0.006
Australia	1990–1995	1.224	1.226	0.997	1.316	1.754	1.122	1.243	1.64	1.44	0.006
Australia	1995–2000	1.306	1.222	1.26	1.149	1.954	1.218	1.603	1.462	1.559	0.016
Austria	1985–1990	1.309	1.354	1.214	1.409	1.257	1.074	0.981	1.139	1.113	-0.019
Austria	1990–1995	1.224	1.255	1.043	1.701	1.196	1.083	0.831	1.355	1.116	-0.019
Austria	1995–2000	1.257	1.212	1.247	1.709	1.197	1.042	0.956	1.31	1.127	-0.018
Belgium	1980–1985	1.384	1.254	1.332	1.22	1.181	1.069	0.814	0.745	0.952	-0.032
Belgium	1985–1990	1.285	1.303	1.279	1.108	1.243	1.12	0.791	0.686	0.96	-0.031
Belgium	1990–1995	1.156	1.24	1.155	1.154	1.242	1.114	0.735	0.734	0.956	-0.031
Bolivia	1985–1990	0.359	0.437	0.691	0.23	0.404	0.935	1.584	0.527	0.863	-0.039
Brazil	1990–1995	0.778	0.275	0.42	0.553	1.599	0.78	0.41	0.539	0.832	-0.041
Bulgaria	1990–1995	0.78	1.056	0.724	0.24	0.565	0.934	0.499	0.166	0.541	-0.064
Bulgaria	1995–2000	0.836	1.111	0.618	0.239	0.792	1.606	0.502	0.194	0.773	-0.046
Canada	1980–1985	1.384	1.247	1.107	0.728	2.715	0.921	1.559	1.026	1.555	0.015
Canada	1985–1990	1.429	1.318	1.105	0.686	2.712	0.924	1.432	0.889	1.489	0.010
Canada	1990–1995	1.336	1.284	1.037	1.379	2.457	0.873	1.197	1.592	1.53	0.013
Canada	1995–2000	1.428	1.288	1.187	1.266	3.654	1.13	1.631	1.741	2.039	0.053
Chile	1985–1990	0.715	0.811	0.638	0.374	0.936	1.08	0.783	0.459	0.815	-0.042
Chile	1990–1995	0.668	0.773	0.575	0.756	1.08	1.363	0.856	1.126	1.106	-0.020
Chile	1995–2000	0.952	0.966	0.962	0.748	1.5	1.411	1.377	1.071	1.34	-0.002

Table A.1 (Continued)

Country	Period	PSP-Admin	PSP-Educ	PSP-Infrastr	PSP-Stabil	PSE-Admin	PSE-Educ	PSE-Infrastr	PSE-Stabil	PSE-average	Growth effect
Colombia	1990–1995	0.778	0.643	0.302	0.451	2.23	1.09	0.719	1.075	1.278	-0.006
Costa Rica	1985–1990	0.956	0.531	0.917	0.524	0.757	0.568	1.145	0.655	0.781	-0.045
Costa Rica	1990–1995	0.892	0.575	0.698	0.832	0.751	0.743	1.001	1.193	0.922	-0.034
Costa Rica	1995–2000	0.959	0.563	0.868	0.672	0.892	0.613	1.216	0.941	0.915	-0.035
Cyprus	1980–1985	0.805	1.116	1.248	1.251	0.565	1.481	1.326	1.328	1.175	-0.014
Cyprus	1985–1990	0.832	1.182	1.247	1.306	0.663	1.472	1.289	1.35	1.193	-0.013
Cyprus	1990–1995	1.113	1.204	1.139	1.845	0.897	1.438	1.084	1.757	1.294	-0.005
Cyprus	1995–2000	1.19	1.002	1.491	1.619	0.941	0.977	1.314	1.427	1.165	-0.015
Czech Rep.	1990–1995	0.89	1.228	0.995	1.244	1.043	1.148	0.827	1.033	1.013	-0.027
Czech Rep.	1995–2000	0.952	1.089	1.086	0.777	1.806	1.119	0.954	0.683	1.14	-0.017
Denmark	1980–1985	1.384	1.179	0.952	0.705	1.579	0.829	0.758	0.562	0.932	-0.033
Denmark	1985–1990	1.429	1.291	1.073	0.756	1.755	0.839	0.88	0.62	1.023	-0.026
Denmark	1990–1995	1.336	1.236	1.093	1.791	1.58	0.77	0.837	1.371	1.139	-0.017
Denmark	1995–2000	1.428	1.23	1.62	1.234	1.801	0.758	1.31	0.998	1.217	-0.011
Dominican Rep.	1990–1995	0.668	0.405	0.239	0.29	1.302	1.169	0.528	0.639	0.909	-0.035
Dominican Rep.	1995–2000	0.836	0.725	0.32	0.397	1.172	1.783	0.616	0.764	1.084	-0.021
Egypt	1990–1995	0.757	0.919	0.8	0.521	0.61	1.011	0.723	0.471	0.704	-0.051
Egypt	1995–2000	0.592	0.927	0.701	0.563	0.465	0.978	0.677	0.543	0.666	-0.054
El Salvador	1995–2000	0.647	0.515	0.648	0.694	0.539	1.089	1.287	1.378	1.073	-0.022
Finland	1985–1990	1.429	1.383	1.257	0.868	2.288	1.186	1.318	0.911	1.426	0.005
Finland	1990–1995	1.336	1.305	1.248	1.22	1.786	0.909	0.99	0.967	1.163	-0.015
Finland	1995–2000	1.312	1.304	1.773	1.598	1.911	0.905	1.485	1.338	1.41	0.004
France	1980–1985	1.384	1.156	1.163	0.667	1.092	1.067	0.91	0.521	0.898	-0.036
France	1985–1990	1.309	1.275	1.168	0.742	1.108	1.056	0.859	0.545	0.892	-0.036

Table A.1 (Continued)

Country	Period	PSP-Admin	PSP-Educ	PSP-Infrastr	PSP-Stabil	PSE-Admin	PSE-Educ	PSE-Infrastr	PSE-Stabil	PSE-average	Growth effect
France	1990–1995	1.112	1.331	1.113	1.477	0.953	1.117	0.782	1.038	0.972	-0.030
France	1995–2000	1.068	1.295	1.289	1.438	0.945	1.075	0.869	0.969	0.965	-0.031
Germany	1990–1995	1.336	1.255	1.592	1.33	1.356	1.272	1.588	1.326	1.386	0.002
Germany	1995–2000	1.306	1.207	1.547	1.467	1.234	1.272	1.445	1.371	1.331	-0.002
Greece	1980–1985	0.739	1.142	0.86	0.559	0.477	2.826	0.77	0.5	1.143	-0.017
Greece	1985–1990	0.98	1.229	0.875	0.41	0.555	2.431	0.64	0.3	0.981	-0.029
Greece	1990–1995	1.002	1.215	0.729	0.555	0.839	2.303	0.626	0.477	1.061	-0.023
Greece	1995–2000	1.074	1.188	1.172	0.577	1.058	1.971	1.144	0.563	1.184	-0.014
Hungary	1980–1985	0.922	0.987	0.79	1.289	0.861	0.98	0.474	0.773	0.772	-0.046
Hungary	1985–1990	1.049	1.111	0.777	0.342	0.964	0.93	0.438	0.193	0.631	-0.057
Hungary	1990–1995	1.102	1.212	0.766	0.472	0.963	0.905	0.439	0.27	0.644	-0.056
Hungary	1995–2000	1.19	1.166	0.72	0.456	1.545	1.224	0.498	0.315	0.895	-0.036
Iceland	1985–1990	1.429	1.268	0.83	2.745	0.842	1.201	0.862	2.854	1.44	0.006
Iceland	1990–1995	1.336	1.228	0.904	1.258	0.787	1.117	0.858	1.193	0.989	-0.029
Iceland	1995–2000	1.428	1.174	1.228	1.638	0.881	1.003	1.238	1.652	1.193	-0.013
India	1995–2000	0.83	0.534	0.779	0.324	2.335	0.87	1.602	0.667	1.369	0.001
Indonesia	1985–1990	0.154	0.563	0.679	0.536	0.315	3.07	1.051	0.829	1.316	-0.003
Indonesia	1990–1995	0.668	0.597	0.721	0.703	1.487	2.196	1.3	1.268	1.563	0.016
Iran	1990–1995	0.89	1.003	0.621	0.293	0.744	1.038	0.898	0.424	0.776	-0.045
Ireland	1980–1985	1.153	1.151	0.794	0.459	1.281	0.998	0.552	0.319	0.788	-0.045
Ireland	1985–1990	1.191	1.188	0.934	0.632	1.429	0.986	0.656	0.443	0.878	-0.037
Ireland	1990–1995	1.224	1.219	0.803	1.335	1.597	1.11	0.639	1.062	1.102	-0.020
Ireland	1995–2000	1.183	1.058	1.013	1.083	1.878	1.079	0.905	0.968	1.207	-0.012
Israel	1995–2000	1.153	1.163	1.616	0.604	0.702	0.748	1.058	0.395	0.726	-0.049

Table A.1 (Continued)

Country	Period	PSP-Admin	PSP-Educ	PSP-Infrastr	PSP-Stabil	PSE-Admin	PSE-Educ	PSE-Infrastr	PSE-Stabil	PSE-average	Growth effect
Italy	1995–2000	1.068	1.214	1.089	0.771	1.269	1.272	0.738	0.522	0.95	-0.032
Jamaica	1980–1985	0.574	0.814	0.551	0.342	0.381	0.646	0.463	0.287	0.444	-0.071
Jamaica	1985–1990	0.594	0.945	0.408	0.24	0.282	0.851	0.334	0.196	0.416	-0.074
Jamaica	1990–1995	0.778	0.9	0.4	0.265	0.593	0.974	0.587	0.389	0.636	-0.056
Jamaica	1995–2000	0.83	1.085	0.819	0.309	0.451	0.905	0.721	0.272	0.587	-0.060
Japan	1980–1985	1.266	1.349	1.479	1.993	5.247	1.232	2.676	3.606	3.19	0.143
Japan	1985–1990	1.309	1.439	1.475	2.489	5.594	1.352	2.816	4.753	3.629	0.177
Japan	1990–1995	1.224	1.39	1.371	2.784	5.47	1.816	2.128	4.322	3.434	0.162
Jordan	1985–1990	0.715	0.488	0.627	0.523	0.346	0.46	0.56	0.467	0.458	-0.070
Jordan	1990–1995	0.89	0.584	0.85	0.859	0.455	0.368	0.807	0.816	0.611	-0.058
Jordan	1995–2000	0.952	0.818	0.992	0.682	0.465	0.551	0.956	0.657	0.658	-0.055
Korea, Rep.	1980–1985	0.687	1.191	1.193	0.811	0.89	1.54	2.266	1.539	1.559	0.016
Korea, Rep.	1985–1990	0.711	1.275	1.263	1.152	1.131	1.542	2.542	2.319	1.883	0.041
Korea, Rep.	1990–1995	1.113	1.353	1.176	1.645	1.96	1.653	2.225	3.112	2.237	0.069
Korea, Rep.	1995–2000	1.068	1.333	1.512	1.165	2.254	1.757	2.753	2.121	2.221	0.067
Lebanon	1995–2000	0.354	1.039	0.548	0.771	0.311	2.071	0.46	0.648	0.872	-0.038
Luxembourg	1980–1985	1.372	0.936	1.136	1.704	1.418	0.8	0.867	1.3	1.096	-0.020
Luxembourg	1985–1990	1.429	0.953	1.044	2.003	1.585	0.958	0.871	1.671	1.271	-0.007
Luxembourg	1990–1995	1.336	0.944	0.418	2.34	1.426	1.396	0.331	1.85	1.251	-0.008
Luxembourg	1995–2000	1.342	0.929	0.299	2.274	1.441	1.12	0.238	1.81	1.152	-0.016
Malaysia	1995–2000	0.952	1.277	0.933	1.559	1.012	1.357	1.378	2.301	1.512	0.012
Mexico	1980–1985	0.624	0.649	0.755	0.165	1.003	0.741	1.165	0.254	0.791	-0.044
Mexico	1985–1990	0.715	0.666	0.731	0.94	1.325	0.888	0.91	1.17	1.074	-0.022
Mexico	1990–1995	0.668	0.722	0.725	1.077	1.319	0.811	1.467	2.18	1.444	0.007

Table A.1 (Continued)

Country	Period	PSP-Admin	PSP-Educ	PSP-Infrastr	PSP-Stabil	PSE-Admin	PSE-Educ	PSE-Infrastr	PSE-Stabil	PSE-average	Growth effect
Mexico	1995–2000	0.592	0.771	0.726	0.964	1.51	0.764	1.455	1.931	1.415	0.004
Namibia	1990–1995	1.113	0.504	1.128	0.403	0.417	0.269	0.991	0.354	0.508	-0.066
Namibia	1995–2000	1.068	0.432	1.115	0.343	0.425	0.254	0.959	0.295	0.483	-0.068
Netherlands	1980–1985	1.384	1.254	1.501	1.114	1.65	0.907	0.912	0.677	1.037	-0.025
Netherlands	1985–1990	1.429	1.243	1.47	2.888	1.762	0.901	0.873	1.714	1.313	-0.004
Netherlands	1990–1995	1.336	1.279	1.291	1.528	1.744	1.115	0.799	0.945	1.15	-0.016
Netherlands	1995–2000	1.428	1.273	1.561	1.312	1.947	1.271	1.037	0.871	1.281	-0.006
New Zealand	1980–1985	1.384	1.185	0.86	0.744	1.235	1.165	0.7	0.605	0.926	-0.034
New Zealand	1985–1990	1.429	1.264	0.924	0.599	1.269	1.042	0.684	0.443	0.86	-0.039
New Zealand	1990–1995	1.336	1.256	0.875	1.583	0.893	0.885	0.704	1.273	0.939	-0.033
New Zealand	1995–2000	1.306	1.241	0.71	1.355	0.784	0.872	0.677	1.293	0.907	-0.035
Nicaragua	1995–2000	0.721	0.449	0.315	0.355	0.588	0.626	0.271	0.305	0.447	-0.071
Norway	1980–1985	1.293	1.217	0.978	1.207	1.716	0.993	0.92	1.135	1.191	-0.013
Norway	1985–1990	1.312	1.304	1.017	1.06	1.763	0.916	0.853	0.889	1.105	-0.020
Norway	1990–1995	1.336	1.346	0.998	1.649	1.337	0.833	0.738	1.221	1.082	-0.022
Norway	1995–2000	1.306	1.325	1.129	1.488	1.698	0.849	0.949	1.251	1.187	-0.013
Panama	1980–1985	0.348	0.677	0.525	1.003	0.205	0.787	0.541	1.034	0.642	-0.056
Panama	1985–1990	0.359	0.754	0.363	3.949	0.209	0.712	0.419	4.555	1.474	0.009
Paraguay	1985–1990	0.117	0.384	3.504	0.489	0.228	1.539	12.52	1.747	4.008	-
Paraguay	1990–1995	0.556	0.503	10.84	0.698	0.822	1.043	28.17	1.812	7.96	-
Peru	1980–1985	0.579	0.693	0.717	0.11	0.633	1.202	1.252	0.192	0.82	-0.042
Peru	1990–1995	0.557	0.749	0.356	0.356	0.97	1.137	0.603	0.603	0.828	-0.041
Peru	1995–2000	0.598	0.845	0.587	0.606	0.84	1.317	0.993	1.025	1.044	-0.025
Philippines	1980–1985	0.192	0.701	1.417	0.604	0.259	1.952	3.714	1.583	1.877	0.040

Table A.1 (Continued)

Country	Period	PSP-Admin	PSP-Educ	PSP-Infrastr	PSP-Stabil	PSE-Admin	PSE-Educ	PSE-Infrastr	PSE-Stabil	PSE-average	Growth effect
Philippines	1985–1990	0.359	0.842	0.449	0.485	0.467	1.768	0.946	1.021	1.05	-0.024
Philippines	1990–1995	0.557	0.829	0.449	0.622	0.649	1.459	0.733	1.015	0.964	-0.031
Poland	1990–1995	1.058	1.191	0.731	0.271	0.93	1.127	0.547	0.203	0.702	-0.051
Portugal	1985–1990	0.98	1.038	0.84	0.472	0.825	1.246	0.709	0.398	0.794	-0.044
Portugal	1990–1995	0.946	1.095	0.849	0.926	0.577	1.058	0.643	0.701	0.744	-0.048
Portugal	1995–2000	1.074	1.205	1.015	1.033	0.661	1.094	0.797	0.811	0.841	-0.040
Romania	1990–1995	0.669	1.029	0.736	0.419	0.58	1.495	0.663	0.377	0.779	-0.045
Romania	1995–2000	0.598	1.038	0.845	0.512	0.546	1.282	0.805	0.487	0.78	-0.045
South Africa	1985–1990	1.309	0.751	1.302	0.234	0.805	0.626	1.361	0.244	0.759	-0.047
South Africa	1990–1995	1.069	0.812	1.02	0.403	0.673	0.615	1.015	0.401	0.676	-0.053
South Africa	1995–2000	1.074	0.768	1.17	0.352	1.381	0.605	1.186	0.357	0.882	-0.037
Spain	1995–2000	1.19	1.259	1.02	0.667	2.132	1.337	0.929	0.607	1.251	-0.008
Sweden	1985–1990	1.429	1.267	1.105	1.299	2.461	0.824	0.863	1.015	1.291	-0.005
Sweden	1990–1995	1.336	1.388	0.99	1.031	1.979	0.869	0.694	0.723	1.066	-0.023
Sweden	1995–2000	1.428	1.368	1.234	2.248	2.129	0.877	0.894	1.628	1.382	0.002
Switzerland	1980–1985	1.384	1.119	1.239	4.532	2.379	1.128	2.089	7.642	3.31	0.152
Switzerland	1990–1995	1.336	1.178	1.099	1.95	1.672	1.009	1.346	2.389	1.604	0.019
Switzerland	1995–2000	1.306	1.142	1.41	2.735	1.575	1.006	1.59	3.083	1.813	0.035
Syria	1980–1985	0.461	0.72	0.767	0.8	0.224	0.642	0.577	0.602	0.511	-0.066
Thailand	1995–2000	0.83	0.759	0.999	2.025	0.849	0.837	1.551	3.145	1.596	0.018
Trinidad & Tobago	1990–1995	0.668	0.917	0.631	0.525	0.431	1.138	0.7	0.583	0.713	-0.050
Trinidad & Tobago	1995–2000	0.714	0.996	0.984	0.557	0.488	1.51	1.089	0.616	0.926	-0.034
Tunisia	1985–1990	0.715	0.632	0.647	0.502	0.567	0.498	0.567	0.441	0.518	-0.066
Tunisia	1995–2000	0.714	0.754	0.863	0.739	0.551	0.536	0.836	0.716	0.66	-0.055

Table A.1 (Continued)

Country	Period	PSP-Admin	PSP-Educ	PSP-Infrast	PSP-Stabil	PSE-Admin	PSE-Educ	PSE-Infrast	PSE-Stabil	PSE-average	Growth effect
Turkey	1980–1985	0.692	0.51	0.728	0.258	0.798	1.076	1.169	0.413	0.864	-0.039
Turkey	1985–1990	0.594	0.615	0.731	0.28	0.877	1.895	1.327	0.509	1.152	-0.016
Turkey	1990–1995	0.846	0.723	0.728	0.419	0.772	1.201	1.058	0.609	0.91	-0.035
Turkey	1995–2000	0.707	0.705	0.705	0.518	0.702	1.544	0.736	0.541	0.881	-0.037
United Kingdom	1980–1985	1.384	1.132	0.963	0.673	1.074	1.06	0.782	0.547	0.866	-0.038
United Kingdom	1985–1990	1.309	1.175	0.936	0.582	1.117	1.093	0.798	0.497	0.876	-0.038
United Kingdom	1990–1995	1.224	1.29	0.894	0.991	0.995	1.17	0.677	0.751	0.898	-0.036
United Kingdom	1995–2000	1.232	1.288	1.048	0.975	1.09	1.278	0.844	0.785	0.999	-0.028
Uruguay	1995–2000	0.83	0.901	0.472	0.382	0.89	1.662	0.484	0.392	0.857	0.001
USA	1980–1985	1.266	1.291	1.044	0.851	1.778	1.001	1.494	1.218	1.373	0.006
USA	1985–1990	1.309	1.275	1.161	0.878	1.821	1.093	1.612	1.219	1.436	0.014
USA	1990–1995	1.224	1.268	0.962	1.287	1.955	1.154	1.307	1.748	1.541	0.045
USA	1995–2000	1.183	1.24	1.288	1.282	2.637	1.218	1.954	1.944	1.938	-0.039
Venezuela	1980–1985	0.692	0.224	0.617	0.585	0.684	0.213	0.92	0.872	0.672	-0.054
Venezuela	1985–1990	0.715	0.276	0.462	0.283	0.996	0.265	0.72	0.44	0.606	-0.059
Venezuela	1990–1995	0.668	0.271	0.341	0.424	1.042	0.287	0.539	0.669	0.634	-0.057
Venezuela	1995–2000	0.714	0.307	0.603	0.305	1.534	0.302	0.978	0.494	0.827	-0.041
Yemen	1995–2000	0.714	0.484	0.332	0.115	0.479	0.446	0.351	0.122	0.35	-0.079

Key: PSP: Public Sector Performance; PSE: Public Sector Efficiency; Admin: Administration; Educ: Education; Infrast: Infrastructure; Stabil: Stabilization; * See footnote 6

Table A.2 Technical Efficiency (*TE*) of public spending in 52 countries, 1995–2000

Country	<i>TE</i>	Growth effect	Country	<i>TE</i>	Growth effect
Algeria	0.363355	-0.133	New Zealand	0.7823	-0.027
Argentina (see footnote 6)	0.830471	-	Nicaragua	0.312448	-0.146
Australia	0.875214	-0.004	Norway	0.858392	-0.008
Austria	0.867984	-0.005	Peru	0.509566	-0.096
Bulgaria	0.466058	-0.107	Portugal	0.706914	-0.046
Canada	0.910333	0.005	Romania	0.513156	-0.095
Chile	0.672651	-0.055	South Africa	0.582557	-0.078
Costa Rica	0.56748	-0.081	Spain	0.697047	-0.049
Cyprus	0.872052	-0.004	Sweden	0.934942	0.012
Czech Rep.	0.653997	-0.060	Switzerland	0.965281	0.019
Denmark	0.885186	-0.001	Thailand	0.857647	-0.008
Dominican Rep.	0.453656	-0.110	Trinidad & Tobago	0.573398	-0.080
Egypt	0.47918	-0.104	Tunisia	0.527729	-0.091
El Salvador	0.500757	-0.098	Turkey	0.461675	-0.108
Finland	0.928959	0.010	United Kingdom	0.745181	-0.036
France	0.802157	-0.022	Uruguay	0.451386	0.004
Germany	0.903573	0.004	USA	0.903279	-0.111
Greece	0.686213	-0.051	Venezuela	0.372336	-0.131
Hungary	0.565051	-0.082	Yemen	0.292314	-0.151
Iceland	0.906244	0.004			
India	0.496827	-0.099			
Ireland	0.727425	-0.041			
Israel	0.713245	-0.045			
Italy	0.657281	-0.059			
Jamaica	0.513118	-0.095			
Jordan	0.589832	-0.076			
Korea, Rep.	0.927815	0.010			
Lebanon	0.454145	-0.110			
Luxembourg	0.791004	-0.025			
Malaysia	0.866256	-0.006			
Mexico	0.608441	-0.071			
Namibia	0.496673	-0.099			
Netherlands	0.866055	-0.006			

in Education as a percentage of GDP (both variables are available from the World Development Indicators, WDI), where we use the end of period values (or the closest to the end available) of Secondary School Enrollment.¹⁹ The resulting PSE is then a measure of government efficiency in the policy area of education.

¹⁹Afonso et al. (2005) use the same PEX, but they also include a measure of the quality of education when they construct the PSP.

In the policy area of administration, the PSP is measured by the end of period values of the variables *Corruption in Government* and *Bureaucratic Quality* (both obtained from the IRIS-3 dataset)²⁰ with higher scores denoting better outcomes, while the PSE is obtained as in Afonso et al. (2005) by dividing this variable by the average public spending on goods and services (available from WDI).

In the policy area of infrastructure, the PSP is measured by the average of *Diesel Locomotives in Use* as a percentage of total locomotives, and the average of the inverse of *Electric Power Transmission and Distribution Losses* (both variables are available from WDI). These measures have also been used by Tanzi and Davoodi (1998) as indicators of the quality of infrastructure (see also Angelopoulos and Philippopoulos 2007). A problem here is that the relevant PEX for infrastructure quality, which has been used by Afonso et al. for the OECD countries, is not available for the larger group of countries we work with. We therefore choose to use *Total Government Expenditure* (as a percentage of GDP) for PEX (this is also available from WDI), again averaged over the five-year period.

Finally, in the policy area of stabilization, the PSP is measured by the average of the inverse of the variables *Inflation Rate* and *Unemployment Rate* (obtained from WDI), while the relevant PSE is calculated by dividing this PSP by *Total Government Expenditure* (as a percentage of GDP), averaged over the five-year period. Afonso et al. also use total government spending as a measure of public sector expenditures that are associated with stabilization and economic performance indicators, such as inflation and unemployment.

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²⁰Afonso et al. (2005) have used very similar variables (measures of corruption, red tape, quality of judiciary and shadow economy). We prefer the IRIS-3 indexes because they are available for the counties and time periods we work with.

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