#### **ORIGINAL PAPER**



# **Uncertainty analysis of SPI and SRI calculation using bootstrap in the Mediterranean regions of Algeria**

**Mohammad Achite1,2 · Ommolbanin Bazrafshan3  [·](http://orcid.org/0000-0003-2524-3992) Zohreh Pakdaman4 · Andrzej Wałęga<sup>5</sup> · Fateme Pourhaghverdi3 · Tommaso Caloiero6**

Received: 24 September 2023 / Accepted: 20 April 2024 © The Author(s), under exclusive licence to Springer Nature B.V. 2024

## **Abstract**

The standardized precipitation index (SPI) and the standardized runoff index (SRI) are widely used in drought monitoring. In the calculation of these indices, the time scale and distribution functions are especially signifcant. In the current study, the uncertainty in the estimation of the two indices, in terms of time scale and distribution functions, was investigated in the wadi Mina basin (Algeria) using monthly precipitation and runof values based on the gamma-II (GAM-II), extreme value-III (EVD-III), Pierson-III (PEI-III) and Weibull-II (WEI-II) distribution functions. With this aim, precipitation and runoff amounts were calculated considering 12- and a 24-month time scales; then, using the bootstrap method, 1000 random sample were generated for each precipitation and runof event and for each time scale, and the confdence interval of the two indices was calculated at around 95%. The size of the confdence interval was considered as uncertainty and the error rate between the estimated and observed data was calculated. The results showed that all the considered distributions ft the time series acceptably, and that the time scale of the data is not signifcantly correlated with the goodness of ft. Moreover, there is no apparent relationship between the rejection cases and the scale and position of the regional stations or the investigated variables. The lack of signifcant diferences between the observed and estimated time series for a specifc distribution caused the averages estimated in SPI to fall within the same descriptive class. Based on the results, WEI-II and EVD-III showed the lowest estimation error and uncertainty in meteorological and hydrological drought, respectively, at both 12- and 24-month time scales, thus suggesting the use of these two functions for drought monitoring at medium-term and long-term time scales.

**Keywords** Drought monitoring · Uncertainty · Bootstrap · Sampling · Precipitation index · Standardized runof

## **1 Introduction**

Territories overlooking the Mediterranean basin, considered a highly vulnerable area to climate change, are particularly susceptible to the adverse efects of drought due to its irregular precipitation patterns (Tramblay et al. [2020\)](#page-29-0). As an example, in recent years North

Extended author information available on the last page of the article

African regions have witnessed an increase in the frequency and severity of drought events, leading to growing concerns regarding water scarcity, food security, and ecosystem health (Spinoni et al. [2014\)](#page-29-1).

Meteorological and hydrological droughts are complex phenomena driven by a combination of climatic and hydrological factors. Meteorological droughts are primarily characterized by a prolonged defcit in precipitation, while hydrological droughts manifest as decreased streamfow, groundwater levels, and reservoir storage, impacting water availability in various sectors (Achite et al. [2023](#page-27-0)). Generally, drought indices are a widely adopted tool for the evaluation of drought conditions. Among the various indices one of the most applied is the standardized precipitation index (SPI) that was initially formulated by McKee et al. ([1993\)](#page-28-0). The SPI operates on the basic principle of converting the probability of cumulative precipitation on diferent time scales into an index. This fexible approach allows the assessment of diferent types of droughts: shorter time scales are suitable for meteorological and agricultural drought assessments, while longer time scales prove more suitable for characterizing hydrological and water resource-related droughts. This index serves multiple purposes, acting as a valuable tool for assessing drought severity, issuing early warnings about droughts, and discerning the infuence of climate change, as several studies highlighted, e.g. Bordi et al. ([2009\)](#page-28-1), Buttafuoco and Caloiero [\(2014](#page-28-2)), and Buttafuoco et al. [\(2018](#page-28-3)). In addition, the SPI boasts several advantageous features. First, it is based solely on precipitation data, making it possible to assess drought even in the absence of other hydrometeorological measurements. Second, its adaptability to diferent time scales makes it possible to evaluate drought conditions in diferent meteorological, hydrological and agricultural contexts. Finally, the standardization of the SPI ensures that the frequency of extreme drought events remains relatively constant across locations and time scales, as demonstrated by Lloyd-Hughes and Saunders [\(2002](#page-28-4)) and Hayes et al. ([1999\)](#page-28-5). Consequently, it maintains its status as a robust and widely favored index for meteorological drought diagnosis, making it the preferred choice for researchers to unravel and characterize drought events in terms of duration and intensity (Bordi et al. [2004](#page-28-6)). Alongside its advantages, the potential disadvantages associated with the SPI should be considered. These include the difficulty of finding appropriate probability distribution functions to model observed precipitation patterns, as highlighted by Guttmann ([1999\)](#page-28-7), Wu et al. ([2007\)](#page-29-2), and Angelidis et al. ([2012\)](#page-28-8). Another disadvantage in the use of the SPI is that sufficiently lengthy time series data are necessary to ensure dependable estimates, as emphasized by Guttman [\(1994](#page-28-9), [1999](#page-28-7)) and Wu et al. ([2005\)](#page-29-3). In addition, the potential inability to identify the onset of drought (Wu et al. [2007\)](#page-29-2) or its conclusion (Blain [2012](#page-28-10)) should be considered in cases where SPI series deviate from normal distributions, often seen in precipitation data with a high frequency of zero values. Finally, the omission of other climate variables, such as temperature which, as underscored by Vicente-Serrano et al. ([2010\)](#page-29-4), may be crucial in the context of climate warming, because this variable makes it possible to include precipitation loss, like evaporation, in the drought analysis. Nevertheless, given the importance of the SPI, Shukla and Wood ([2008\)](#page-29-5) applied the concept employed by McKee et al. ([1993\)](#page-28-11) for the SPI in defning the standardized runof index (SRI) as the unit standard normal deviate associated with the percentile of accumulated hydrologic runoff over a specifc duration.

As previously evidenced, the identifcation of a probability distribution suitable to describe precipitation or runoff data can be considered a possible source of uncertainty in the SPI or SRI evaluation. In efect as underlined by Guttmann ([1999\)](#page-28-7), if different distributions are used to describe an observed series of precipitation or runof data, as a result, diferent SPI or SRI values may be obtained. Accurate assessment and

quantifcation of the uncertainties associated with drought indices and their impacts are crucial for efective drought management and mitigation strategies. For this reason, some authors have conducted analyses to examine the impact of probability distribution selection and record length on the calculation of drought indices or the evaluation and comparison of all sources of uncertainty inherent in drought index estimates. For example, recent studies (Naumann et al. [2012](#page-29-6); Hu et al. [2015\)](#page-28-12) have introduced bootstrap-based methods to quantify the uncertainty associated with the SPI calculations. Specifcally, in the work of Naumann et al. ([2012](#page-29-6)), the primary focus was to compare the SPI confdence intervals derived from datasets of varying lengths, while Hu et al. ([2015\)](#page-28-12) predominantly discussed the impact of sampling uncertainty on the estimation of SPI value. Vergni et al. ([2018](#page-29-7)) employed a bootstrap-based approach to evaluate how the SPI confdence intervals behave when subjected to variations in multiple sources of uncertainty: the underlying probability distribution, time scale, time series length, and month (for time scales shorter than 12 months). Ghasemnezhad et al. ([2022\)](#page-28-13) analyzed the uncertainty of the SRI in Iran for diferent time scales, from 3 to 48 months, using the Monte Carlo sampling and considering normal, log normal, Weibull and gamma distribution functions. The results showed that the highest uncertainty is related to the normal and log normal distributions and the lowest uncertainty is related to the gamma and Weibull distributions. In addition, increasing the time scale and decreasing the length of the record led to an increase in uncertainty.

In this context, the objective of this scientifc paper is to enhance our understanding of meteorological and hydrological droughts in a Mediterranean area (Algeria) by conducting a rigorous uncertainty analysis using the Bootstrap resampling technique. The findings of this research may offer valuable insights into drought monitoring, early warning systems, and water resource management practices in this climatically vulnerable region, ultimately fostering the development of more resilient and sustainable strategies for addressing drought challenges.

### **1.1 Study area and data collected**

The study area is the Wadi Mina basin, in northwest of Algeria, which covers an area of 4900 km<sup>2</sup> and lies between 00° 22′ 59″ E and 01° 09′ 02″ E and between 34°41′57″ N and 35° 35′ 27″ N (Fig. [1](#page-3-0)). The wadi Mina involves four major tributaries: wadi Mina, wadi Haddad, wadi Abd and wadi Taht. The topography of the basin is complex and rugged, and the altitude varies from 164 to 1327 m. The climate at the study area is continental with cold winters and hot summers with large temperature diferences. Average annual precipitation in the basin ranges from 250 to 500 mm, mostly concentrated between November and March, while the mean annual temperature ranges from about 16–19.5 °C. Almost half of the basin is covered by vegetation of varying densities, such as scrubs (32%), forests (35.8%), and cereal crops (Achite and Ouillon [2007](#page-27-1)). Monthly precipitation and runoff records for a 40-year observation period (1974–2009) were compiled for five precipitation  $(S1, S2, ..., S5)$  and hydrometrics  $(H1, H2, ..., H5)$ stations from the National Agency of the Water Resources (Fig. [1](#page-3-0) and Tables [1](#page-3-1) and [2\)](#page-4-0). Data were examined for homogeneity using the double mass curve, linear regression, and Mann–Whitney test procedures to ensure quality. The technique found a few in homogeneities, and the irregular data were corrected using data from nearby homogeneous stations (Achite et al. [2021](#page-27-0)).



<span id="page-3-0"></span>**Fig. 1** Map of the study area along with located hydro-meteorological stations

<span id="page-3-1"></span>**Table 1** Precipitation stations characteristics

ID Code	Name	Longitude	Latitude	Elevation (m) Mean STDEV		
	S1 013306 Oued Abtal		$0^{\circ}$ 40' 33.97" E 35° 28' 03.59" N 354		21.57 21.73	
	S <sub>2</sub> 013401 Sidi Abdelkader Djillali		$0^{\circ}$ 34' 08.35" E 35° 29' 20.71" N 225		20.11 21.30	
	S3 013302 Ain Hammara		$0^{\circ}$ 39' 16.85" E 35° 23' 15.39" N 288		21.13 21.93	
	S4 013001 Kef Mehboula		$0^{\circ}$ 49' 34.20" E 35° 18' 40.72" N 475		27.29 27.10	
	S5 013304 Takhmaret		$0^{\circ}$ 37' 27.25" E 35 $^{\circ}$ 06' 49.01" N 655		20.63 21.21	

## **2 Methodology**

### **2.1 Standardized precipitation index (SPI) and standardized runof index (SRI)**

The calculation process for the SPI and the SRI is exactly the same but with precipitation data as input for SPI and runoff data as input for SRI (Bazrafshan et al. 2015). The process starts with ftting the appropriate probability distribution function to the data series in any desired time interval. If *X* represents the total amount of the variable during a period of  $\tau$  month in month j, by fitting the desired distribution to the time series of *X*, the density distribution function or marginal CDF  $u = F_{X\tau(x\tau)}$  is obtained. Then the index is calculated for each observation with the inverse normal function or  $\mathcal{Q}^{-1} = (u_\tau)$ . In other words, instead of showing cumulative probability, the index is described by a standard normal variable (with zero mean and 1 standard deviation).

Table [3](#page-5-0) shows the classifcation of drought severity based on SPI or SRI values.



<span id="page-4-0"></span>

<span id="page-5-0"></span>

### **2.2 Drought characteristics**

The drought characteristics investigated this study are duration, severity, magnitude and peak. According to McKee et al. ([1993](#page-28-11)), a drought event is defned as a period in which the SPI or the SRI values are less than zero. Consequently, the drought duration (D) refers to the period during which drought values consistently remain below zero, and the accumulated drought values within each event represent the drought severity (S), which for convenience is considered as an absolute value as follows:

<span id="page-5-1"></span>
$$
S = \left| \sum_{i=1}^{D} SPI_i \right| \tag{1}
$$

The drought magnitude results from dividing the severity by the duration within each event (Azhdari et al. [2021\)](#page-28-14). Finally, the drought peak is the maximum value of S in a period of D months.

#### **2.3 Resampling using bootstrap method**

This method is classifed in the group of non-parametric statistics methods and resampling techniques and is used to estimate the parameter of a statistical population using sampling with placement. Bootstrap is also used to calculate the confdence interval for the estimation (Verdonck et al. [2001](#page-29-8)). In the bootstrap method, subsamples are generated by resampling with replacement from the original sample. Considering that the number of original samples is equal to *n*, it is possible to create infnite sub-samples of size *n* with placement (Dixon [2006](#page-28-15)).

The bootstrap steps in this research are as follows:

*Step 1*: *N* times from the main precipitation or runoff data values  $X = X_1 ... X_n$  should be sampled by the sampling method with placement, then *N* groups resulting from the bootstrap samples will be obtained.

<span id="page-5-2"></span>
$$
X_{(j)}^* = X_{1(j)}^* \dots X_{n(j)}^* \text{for } j = 1 \dots N \tag{2}
$$

*Step 2*: Using the maximum likelihood method, the theta parameter of the distribution function f(x.θ) based on each bootstrap sample  $X^*_{(j)}$ ,  $j = 1...N$  is estimated. Then N sets of  $\theta_{(j)}^*$ : **j** = 1 ... N parameters are obtained.

*Steps 3*: For a given precipitation or runoff event  $X_i$ ,  $i = 1...n$  using eqs. [3](#page-6-0) and [4](#page-6-1) and using N sets of parameters $\theta_{(j)}^*$ , the number N estimate the cumulative probability H(*xi*) as  $H(X_{i,(i)})$ .i = 1 ... n;j = 1 ... N can be calculated.

$$
SPI(or SRI) = \emptyset^{-1}(F(X))
$$
\n(3)

<span id="page-6-1"></span><span id="page-6-0"></span>
$$
H(x) = q + (1 - q)F(x)
$$
 (4)

*Step 4*: Using Eq. [5](#page-6-1), the SPI or the SRI value is estimated for events  $X_i$ ,  $i = 1...n$ according to estimates of the cumulative distribution function that values  $H(X_{i(j)})$ .i = 1 ... n;j = 1 ... N; then to calculate the SPI or SRI for each phenomenon x<sub>1</sub>, N estimates of SPI or SRI are obtained as  $SPI(orSRI)_{i,(j)}$  for  $i = 1...n; j = 1...N$ .

$$
SPI(or SRI) = \emptyset^{-1}(H(X))
$$
\n(5)

*Step 5*: By putting the *SPI*(*orSRI*)<sub>*i*.(*j*)</sub> values for  $i = 1 ... n; j = 1 ... N$  as the sampled SPI or SRI for the data event  $x_i$ , based on it, a point estimate or an interval estimate for SPI or SRI can be obtained (Hu et al. [2015\)](#page-28-12).

#### **2.4 SPI and SRI bootstrap sampling algorithm**

To evaluate SPI or SRI, the frst step was to calculate the amount of cumulative precipitation or runoff in the year *i* and the month *j*, which ranges from 1 to 12. The value of *j* depends on the time scale  $\tau$ , which can be any scale from 1 to 48 months.

<span id="page-6-2"></span>
$$
Y_{i,j}^{\tau} = \sum_{t=0}^{\tau-1} X_{i,j-t}
$$
 (6)

where  $x_{ij}$  is the amount of monthly precipitation or runoff for the year *i* and the month  $j = 1.2 \dots 12$  and;  $Y_{i,j}^t$  is the cumulative precipitation or runoff value for the desired  $\tau$ scale.

As an example,  $Y_{1983,12}^{12}$  means the scale of one month, in the twelfth month (December) 1983 ( $\tau = 12$ ; j = 12 and i = 1983).

In current study, we applied the Gamma (GA), Weibull (WEI), Extreme value (EVD) and Pearson (PE) distributions (Fig. [2](#page-7-0)).

Based on the presented method to calculate the uncertainty, starting from precipitation or runoff data for  $Y_{i,12}^{12}$ , which means 1 month scale in the 12th month of 1974–2009, resampling has been done 1000 times with bootstrap as follows:

$$
Y_{i,12(j)}^{*12} = \left( y_{1974.9(j)}^{*12}, y_{1975,12(j)}^{*12}, y_{1976.12(j)}^{*12}, \dots, y_{2009.12(j)}^{*12} \right) \text{for } j = 1, 2, \dots, 1000 \tag{7}
$$

In the next step, based on each bootstrap sample,  $Y_{i.12(j)}^{*12}$  a set of 1000 parameters of  $\theta_j^*$  of each function has been obtained. Then, using relations Eqs. [1,](#page-5-1) [2](#page-5-2) and [3](#page-6-0), there are 1000 estimates of SPI or SRI according to the amount of precipitation or runoff from  $Y_{i,12}^{12}$ , for which  $SPI($ (or  $SRI$ )<sub>i,j</sub> = 1 ... 1000 can be obtained. Finally, from these thousand estimates, it ispossible to calculate the upper and lower bounds based on the 5% and 95% quantile.



<span id="page-7-0"></span>**Fig. 2** The interval for SPI and SRI, by bootstrap method procedure. For example, for SPI,  $\tau = 12$ ; j=12 and  $i=1983$  precipitation is 9 mm. (**a**) The M ( $M=1000$ ) EVD cumulative distribution functions (fitted to each sample bootstrap from the original precipitation data), and M possible values of cumulative probability P ( $X_{1983,12}^{12}$ ) are estimated, (**b**) the empirical cumulative distribution function of the M values of P( $X_{i,j}^{\tau}$ ), and the 0.5th  $P(X_{i,j}^{\tau}) = 0.0001$  and 0.95th  $P(X_{i,j}^{\tau}) = 0.0018$  percentile discovered. The lower and upper limits of the SPI confdence (*P*=90%) is−4.07 and−2.08 respectively; (**c**, **d**) repeated for SRI

## **2.5 Estimation of error**

Equation  $8$  was used to compare the efficiency of the bootstrap simulation values with the observed values as suggested by Hugh et al. (2015).

$$
ARE = \left| \frac{(SPI_{Obs} - SPI_{Est})}{SPI_{Obs}} \right| \times 100 \tag{8}
$$

in which ARE is the absolute value of the error ratio and  $SPI_{Obs}$  and  $SPI_{Est}$  are the observed value and the average value of the estimates in 1000 times of sampling in each time scale, respectively.

#### **2.6 Kappa Kohen weighted (Kw)**

In order to compare the observed and predicted classes of SPI and SRI, the Cohen's kappa statistic was used. Kappa statistic was frst introduced by Cohen ([1968\)](#page-28-16) as a measure of agreement in psychology. Given that  $P_{ij}$  is the ratio of the total components of the time series that belong to class *i* of one series and to class *j* of the second series, these ratios can be formed inside a matrix whose main diameter contains the unique  $P_{ii}$  ratios of the components that match in both series and  $P_0$  is the sum of  $P_i$  values. Considering  $P_{e,ii}$  as the product of the sum of the respective row and column ratios ( $P_{e,ii} = P \bullet i$ .  $P i \bullet$ ), the expected agreement value is equal to the sum  $P_e$  of the  $P_{e,ii}$  values. The Kappa statistic is then defned as:

$$
K = \frac{P_0 - P_e}{1 - P_e} \tag{9}
$$

Landis and Koch [\(1977\)](#page-28-17) presented Table [4](#page-8-0) to classify agreement into diferent categories. In the comparison of SPI or SRI classes in the observed and predicted series, the disagreement between mild drought and moderate drought is not as big as the disagreement between mild drought and severe drought. Therefore, by considering specifc weights for each of the non-agreement states, a more accurate comparison of the SPI or SRI classes in the observed and predicted series can be made. By considering the weighted  $w_{ii}$  for the disagreement in the  $ij$  domain of the matrix, the weighted disagreement is obtained from the sum of the products of the ratio  $P_{ii} \times w_{ii}$ . Therefore, the weighted kappa statistic is obtained from the following equation:

$$
K = 1 - \frac{\sum w_{ij} P_{ij}}{\sum w_{ij} P_{e,ij}}
$$
\n
$$
\tag{10}
$$

The signifcance test statistic of the weighted kappa statistic with the null hypothesis of disagreement is as follows:

$$
t = \frac{K}{\sqrt{\frac{\sum w_{ij}^2 P_{e,ij} - (\sum w_{ij} P_{e,ij})^2}{n(\sum w_{ij} P_{e,ij})^2}}}
$$
(11)

where *n* is the number of time series observations.



<span id="page-8-0"></span>**Table 4** Bescription of Kw and value of agreement

## **3 Results and discussion**

### **3.1 Goodness of ft**

The evaluation of the efficiency of the four mentioned functions and the goodness-of-fit test on precipitation and runoff data in the five climatological and hydrological stations are shown in Table [5,](#page-9-0) which provides K–S and *P*-value statistics for each distribution.

Rejection percentages are similar among the distributions, with GAM-II and EVD-III performing slightly better (no rejection) than PE-III and WEI-II (5% rejection). Therefore, in general, the performance of all four candidate distribution functions can be considered acceptable. The noteworthy point is that the goodness-of-ft test is not signifcantly related to the time scale, so that 5% rejection occurred in both 12 and 24 month scales. Furthermore, no apparent relationship has emerged between the rejection cases with the scale and position of the regional stations or the investigated variable. The results of Vergni et al. ([2017](#page-29-7)), Stagge et al. ([2015](#page-29-9)), Lloyd-Hughes and Saunders ([2002](#page-28-4)) and Khatun et al. [\(2007\)](#page-28-18), Bazrafshan et al. ([2020\)](#page-28-19) also reported the same results on goodness of ft.

Variable	St.name	Time scale	PE-III		GAM-II		<b>GEV-III</b>		WEI-II	
			$K-S$	$P_{\rm value}$	$K-S$	$P_{\rm value}$	K-S	$\mathbf{P}_{\text{value}}$	$K-S$	$P_{value}$
Precipitation	S1	12	0.02	0.41	0.03	0.14	0.02	0.41	0.05	0.6
		24	0.025	0.51	0.028	0.39	0.02	0.64	0.05	0.3
	S <sub>2</sub>	12	0.02	0.73	0.01	0.91	0.02	0.76	0.02	0.48
		24	0.030	0.30	0.024	0.61	0.025	0.50	0.02	0.81
	S <sub>3</sub>	12	0.03	0.29	0.020	0.79	0.029	0.33	0.021	0.75
		24	0.01	0.95	0.04	0.76	0.02	0.51	0.03	0.23
	S <sub>4</sub>	12	0.024	0.55	0.033	0.22	0.02	0.71	0.021	0.76
		24	0.02	0.77	0.02	0.80	0.03	0.10	0.02	0.40
	S5	12	0.020	0.80	0.01	0.96	0.04	0.022	0.04	0.06
		24	0.01	0.83	0.03	0.24	0.02	0.62	0.02	0.46
Runoff	SH <sub>1</sub>	12	0.023	0.65	0.023	0.63	0.01	0.92	0.026	0.48
		24	0.029	0.34	0.015	0.97	0.33	0.204	0.026	0.50
	SH <sub>2</sub>	12	0.024	0.60	0.024	0.59	0.020	0.78	0.04	0.042
		24	0.018	0.88	0.031	0.26	0.027	0.45	0.021	0.75
	SH <sub>3</sub>	12	0.021	0.761	0.02	0.49	0.024	0.58	0.02	0.56
		24	0.03	0.23	0.034	0.18	0.03	0.32	0.047	0.22
	SH <sub>4</sub>	12	0.032	0.23	0.017	0.91	0.03	0.22	0.03	0.26
		24	0.024	0.359	0.030	0.30	0.02	0.54	0.018	0.89
	SH <sub>5</sub>	12	0.03	0.23	0.022	0.71	0.02	0.76	0.032	0.25
		24	0.04	0.01	0.02	0.34	0.03	0.25	0.017	0.92
Percentage of rejection at 5% 5% significant level		5%		$0\%$		$0\%$		5%		

<span id="page-9-0"></span>**Table 5** Kolmogorov–Smirnov test for precipitation and runof

Bold number: No signifcant at 95%level

Table [6](#page-11-0) shows the descriptive statistics including mean, minimum, maximum and standard deviation of SPI and SRI at the 95% confdence level for the 12- and 24-month time scales and the investigated stations considering the four investigated functions.

Based on the results, the averages estimated in SPI are in the same descriptive class, so that the extreme meteorological wet (maximum SPI) is estimated by the PE-III function and the extreme meteorological drought (minimum SPI) is estimated by the EVD-III function, which has the highest standard deviation. In hydrological drought, the most severe hydrological drought is estimated by WEI-II, but the most severe hydrological drought and the highest standard deviation is estimated by GAM-II, which is often seen in a 12-month time scale.

### **3.2 Election of distribution function and time scale on drought monitoring**

SPI and SRI were extracted in the 12- and 24-month time scales with four distribution functions, and drought events were estimated based on the classifcation proposed by McKee et al.  $(1993)$  $(1993)$ . Figure [3](#page-13-0) (a, b) shows the SPI and Fig. 3 (c, d) the calculated SRI at station S1 and SH1 for the 12- and 24-month time scales. The SPI on a 12-month scale has several fuctuations, so that the duration of drought changes with diferent intensity and weakness in diferent years. Several historical wet/dry events have been detected by all functions for the SPI and the SRI at both 12- and 24-month time scales (Table [7](#page-14-0)). For example, on a 12-month scale ( $\tau$ =12), in 1983 (i=1983) in December (j=12), with 9 mm of precipitation, the most severe drought  $(SPI = -3.72)$  has been estimated by the EVD-III, followed by the GAM-II (SPI=  $-3$ ), while the WEI-II and the PE-III estimated  $-2.06$  and  $-2.05$ respectively. Considering the event  $Y^{12}_{19756}$  with  $P=39$  mm, the WEI-III has estimated the most severe wet period (SPI=  $-3.10$ ) followed by the PE-III (SPI=  $-2.81$ ), the GAM-II  $(SPI = -2.67)$  and the EVD-III (SPI =  $-2.23$ ). Considering the 24-month scale, two shortterm drought periods and two long-term drought periods have been observed, in which the historical drought event  $Y^{24}_{2000.5}$  occurred. Similar to the 12-month scale, the EVD-III has estimated the most severe drought (SPI =  $-3.14$ ) followed by the WEI-II (SPI =  $-1.53$ ). In general, for the meteorological drought, EVD-III estimated the most severe drought while the WEI-II has estimated the values of lower intensities.

Regarding hydrological wet and drought events, the SRI showed three long-term drought periods for the 12-month time scale and, in the period 2003–2007, two historical events  $Y_{2005,8}^{12}$  and  $Y_{2007,6}^{12}$  occurred. On a scale of 24 months, the duration of drought is increasing and a historical event  $Y_{2005,12}^{24}$  occurred. The GAM-III estimated the most severe droughts in all the years while the PE-III showed the least severity of drought. On the other hand, considering the historical wet events  $Y_{1981,5}^{12}$ ,  $Y_{1998,5}^{12}$ , and  $Y_{1975,12}^{24}$  were rated as the most severe for the PE-III while the EVD-III evidenced the least severity.

Figure [4](#page-15-0) shows the drought zoning in two historical events of meteorological (Fig. [4](#page-15-0)a) and hydrological (Fig. [4](#page-15-0)b) drought with the help of the EVD-III and the GAM II functions, respectively, in Mina basin. The above occurrences are the most exceptional droughts that have been mentioned by the functions during the study period. In the event  $SPI_{1983,12}^{12}$  the most vulnerable parts of the region in terms of lack of precipitation have been identifed in the southern parts, while in event  $SRI_{2007,6}^{12}$  the most vulnerable points in terms of surface currents have been detected in the northern regions.

Figure [5](#page-15-1) is calculated considering the classifcation (9 categories) proposed by Lloyd-Hughes and Saunders [\(2002](#page-28-4)) for two time scales and two indicators investigated in stations S1 and SH1. Based on the results of the SPI-12 (Fig. [5a](#page-15-1)), most of the frequency of

<span id="page-11-0"></span>



### Natural Hazards



<span id="page-13-0"></span>**Fig. 3** Time series of SPI (**a**,**b**) and SRI (**c**,**d**) in the time scale of 12 at station S1 and SH1



<span id="page-14-0"></span> $\ddot{\phantom{a}}$ 



<span id="page-15-0"></span>**Fig. 4** Spatial drought mapping of  $SPI_{1983,12}^{12}(a)$  *and SRI*<sup>12</sup><sub>2007,6</sub>(*b*) using EVD and Gamma III in the Mina Basin



<span id="page-15-1"></span>**Fig. 5** Drought frequency in each category of wet/drought descriptive class (**a**) SPI and (**b**) SRI at 12 and 24-month time scales (*N*: normal; *EW*: extreme wet; *SW*: severe wet; *MW*: moderate wet; *MIW*: mild wet; *MOD*: moderate drought; *SD*: severe drought; *ED*: extreme drought) (Threshold value of each class presented in Table [3](#page-5-0))

droughts in the study area fall within normal and mild drought classes, with percentages of normal conditions equal to 37% for the PEI-III, 40% for the GAM-II and the EVD-III and 475 for the WEI-II and percentages of mild drought conditions estimated to be 22%, 20%, 17% and 14% for the WEI-III, PEI-III, GAM-II and EVD-III functions, respectively. Similar results have been obtained for the SPI-24. In the two investigated scales, EVD-III and GAM-II functions compared to the other two functions have successfully estimated SD, ED drought events (Fig. [5](#page-15-1)a).

These results confrm past studies, such as Achite et al. ([2023\)](#page-27-0), evidencing that, in the Mediterranean regions of Algeria, due to the distribution of precipitation around the average, the condition of drought generally falls within the normal class.

As regards the SRI-12 (Fig. [5](#page-15-1)b), the PEI-III estimated more than 40% of the conditions to be normal, while the other functions estimated 32% on average. The 24-month time scales presents almost the same results. PEI-III and WEI-II functions have been considerably more successful in estimating historical drought events for both the time scales.

Generally, a higher frequency of hydrological droughts in the severe and extreme classes compared to meteorological droughts have been observed and this can be considered a con-sequence of climate change and anthropogenic activities (Aghakouchak et al. [2021](#page-28-20); Zheng et al. [2023\)](#page-29-10).

### **3.3 Uncertainty analysis**

Based on the introduced methodology, for both the time scales, year and month, the values of SPI and SRI were calculated with 1000 production sampling and error values, bands and average. Table [8](#page-16-0) shows some historical drought events in the investigated scales. The results of the 12-month historical drought with 9 mm rainfall, SPI estimated by the PEI-III is−2.06, SPI estimated by bootstrap is−2.19, estimation error (ARE) is 0.06, upper and lower band are−1.70 and−2.87 respectively, and the degree of uncertainty or band diference is 1.16. Considering the GAM-II, SPI is−3.00, bootstrap SPI is−2.65, estimation error is 0.11, upper and lower bands are−2.07 and−3.35

Value SPI OBS		<b>EST</b>	ARE	UB	LB.	UN	<b>OBS</b>	<b>EST</b>	RE	UB	L <sub>B</sub>	<b>UN</b>
$Y_{i,j}^{\tau}$	$Y_{1983,12}^{12}$						$Y_{2005,5}^{24}$					
PEI-III		$-2.06$ $-2.19$ 0.06									$-1.70$ $-2.87$ 1.16 $-0.84$ $-0.97$ 0.15 $-0.50$ $-1.55$ 1.04	
GAM-II		$-3.00$ $-2.65$ 0.11									$-2.07$ $-3.35$ 1.27 $-0.94$ $-0.99$ 0.05 $-0.44$ $-1.65$ 1.21	
<b>EVD-III</b>		$-3.72$ $-2.86$ 0.23									$-2.08$ $-4.07$ 1.99 $-0.98$ $-0.95$ 0.03 $-0.35$ $-1.68$ 1.33	
WEI-II				$-2.05$ $-2.03$ $0.008$ $-1.67$ $-2.44$ $0.76$ $-0.74$ $-0.93$ $0.24$ $-0.51$ $-1.40$ $0.88$								
<b>SRI</b>	$Y_{2007,6}^{12}$						$Y_{2005,12}^{24}$					
PEI-III		$-2.22 - 2.39 0.07$									$-1.70$ $-3.70$ $2.04$ $-2.51$ $-3.11$ $0.23$ $-1.94$ $-4.52$ $2.58$	
GAM-II		$-3.05$ $-4.79$ 0.56									$-3.79$ $-6.01$ $2.21$ $-2.49$ $-4.30$ $0.72$ $-3.08$ $-5.92$ $2.83$	
<b>EVD-III</b>		$-2.28$ $-3.90$ 0.70									$-2.99$ $-4.64$ $1.64$ $-2.30$ $-4.19$ $0.81$ $-3.25$ $-4.67$ $1.42$	
WEI-II		$-2.86$ $-3.41$ 0.19									$-2.93$ $-3.99$ $1.05$ $-2.22$ $-2.91$ $0.30$ $-2.33$ $-3.55$ 1.21	

<span id="page-16-0"></span>**Table 8** The observed, estimated by bootstrap, simulation error, upper and lower bands of SPI and SRI in several historical drought events at station S1 and SH1

*OBS* SPI/SRI observed, *EST* SPI/SRI estimated, *AER* Absolute error ratio, *UB* Upper band, *LB* Lower band, *UN* Uncertainty

respectively, and the degree of uncertainty is 1.27. For the EVD-III, SPI is−3.72, bootstrap SPI is -2 0.86, estimation error is 0.23, upper and lower bands are−2.08 and−4.07, respectively, and the uncertainty level is 1.99. Finally, as regards the WEI-II, SPI is−2.05, bootstrap SPI is−2.03, estimation error is 0.008, upper and lower band are−1.67 and−2.44, respectively, and uncertainty level is 0.76. The changes in band values are shown in Fig.  $6$  (a1–a4). Based on the results, all the estimates for SPI are



<span id="page-17-0"></span>**Fig. 6** Comparison of observed, estimated, upper and lower bond  $SPI_{1983.12}^{12}$  (a1-a4) and  $SRI_{2007.6}^{12}$  (b1-b4) in station S1 and SH1 Zoning of upper band (**a**), lower band (**b**) and uncertainty (**c**) of  $\mathbf{SPI}^{12}_{1983.12}$  event in EVD-III function and zoning of  $SRI^{12}_{2007.6}$  upper band (**d**), lower band (**e**) and uncertainty (**f**) in GAM-II are shown in Fig. [7](#page-19-0) The interpolation is based on the kriging method

within the confdence range and the bootstrap estimated values are in good agreement with the observed values in all functions and are in the defned confdence band.

The same values were calculated for SRI historical events. For example, in the historical event of  $Y_{2007,6}^{12}$  with a runoff of 0.1 m<sup>3</sup>/s, the estimated SRI with the PEI-III is −2.22, the SRI estimated by bootstrap is−2.39, the estimated error is 0.07, the upper and lower bands are−1.70 and−3.70, and the amount of uncertainty or diference between the bands is 2.04. For the GAM-II, with SRI is−3.05, SPI estimated with bootstrap is−4.79, estimated error is 0.56, upper and lower band are−3.79 and−6.01, respectively, and uncertainty level is 2.21. In the EVD-III results, observed SRI is−2.28, bootstrap SRI is−3.90, estimated error is 0.70, upper and lower bands are−2.99 and−4.64, and the uncertainty is 1.64. Considering the WEI-II, observed SRI is−2.86, bootstrap SRI is−3.41, estimated error is 0.19, upper and lower band are−2.93 and−3.99, respectively, and the uncertainty is 1.05. The changes in upper and lower bands are shown in Fig. [6](#page-17-0) (b1–b4). Based on the results, only for the PEI-III estimated and observed value fall within the uncertainty band, while for the other functions, although the SRI estimated is within the confidence band, the observed values are outside the band, and there is a great diference in the estimated and the observed event. In Fig. [6](#page-17-0) (b1-b4), this discrepancy in historical event  $Y_{2007,6}^{12}$  can be clearly understood. Most of the estimated values have maintained the observed trend, but in the minimum value of SRI that occurred (historical drought) on the mentioned date, the signifcance diference between observed and estimated is well evident.

As a result, considering the upper band (Fig. [7a](#page-19-0)) the central areas of the Mina Basin are the most vulnerable areas suffering from severe droughts  $(SPI < -1.9)$ . These areas constitute 13% of the region surface, while more than 60% of the region in the northern and southern parts is under normal conditions ( $SPI < -0.5$ ). In the lower band of the mentioned historical event, the central regions with an area of more than 15% are under severe and extreme droughts, while more than 50% of the region is under normal conditions. Therefore, from the comparison of the upper and lower bands, it emerges that the spatial pattern of drought does not change signifcantly among the maps, but in this example, the only factor afecting the size of the confdence interval is shown by the variability of the rainfall data of each station. Considering that the central parts have the highest intensity of drought, therefore the greatest distance between the upper and lower band is also in these places with an uncertainty of 2.2, while the northern and southern parts have the lowest uncertainty.

Regarding the historical  $SRI^{12}_{2007,6}$  event, the most vulnerable points in this event are the northern parts of the Mina basin, where the spatial pattern of the upper and lower band is similar, which indicates the highest variability of the runof. The uncertainty of the estimates is also high in the northern parts.

Figure [8](#page-20-0) shows the uncertainty and error rate for the diferent functions and time scales in the entire time range under investigation. The average uncertainties of SPI-12 (Fig. [8](#page-20-0)a) in the whole basin using PEI-III, GAM-II, EVD-III, WEI-II function are 0.8, 0.742, 1.14 and 0.742, respectively, while the errors (Fig. [8b](#page-20-0)) are equal to 12.26, 20.8, 62.08 and 1.91 respectively. Therefore, on a 12-month scale in the case of SPI, WEI-II in the frst degree and PEI-III in the second degree have the least uncertainty and error in meteorological stations. For the 24-month time scale, the average uncertainties of the SPI (Fig. [8c](#page-20-0)) in the whole basin using PEI-III, GAM-II, EVD-III, WEI-II function are 1.67, 1.102, 1.066 and 1.09, respectively, with errors (Fig. [8](#page-20-0)d) equal to 11, 38.2, 55.04 and 26.2 respectively. Therefore, for this time scale, WEI-II and PEI-III functions have the least uncertainty and error. The comparison of uncertainty and error between the time scales showed that with the increase of the time scale and the decrease of the sample size, the amount of



<span id="page-19-0"></span>**Fig. 7** Map of upper band (**a**), lower band (**b**) and uncertainty (**c**)  $SPI_{1983.12}^{12}$  in in EVD-III and  $SRI_{2007.6}^{12}$  zoning of upper band (**d**), lower band (**e**) and uncertainty (**f**) in the GAM -II

uncertainty and estimation error increases. In fact, on average, the amount of uncertainty and error in the SPI-12 are 0.85 and 24.24, respectively, while in the SPI-24 they are 1.106 and 32.61 respectively.

Researchers such as Guenang and Kamga [\(2014](#page-28-21)), Pieper et al. [\(2020](#page-29-11)) and Sherif et al. ([2014\)](#page-29-12) introduced WEI-II as the best function in SPI modeling while Zhang and Li [\(2020](#page-29-13)) believe that WEI-II has not been successful in estimating SPI in long and medium term periods and it is better to use it only for a short-term scale (1–3 months).

In the case of hydrological drought, the average uncertainties of SRI-12 (Fig. [8](#page-20-0)e) in the whole basin using PEI-III, GAM-II, EVD-III, WEI-II function are 1.064, 0.846, 0.828 and 0.872, respectively, with errors equal to 1.704, 1.542, 0.574 and 1.07 respectively (Fig. [8f](#page-20-0)). Therefore, on a 12-month scale in the case of SRI, EVD-III distribution has the lowest uncertainty and error in hydrometric stations. For the 24-month time scale, the average uncertainties of SRI-24 (Fig. [8](#page-20-0)g) in the whole basin using PEI-III, GAM-II, EVD-III, WEI-II function are 1.604, 1.34, 1.238 and 1.378 respectively, with errors equal to 1.73, 1.742, 1.57 and 1.66 respectively (Fig. [8h](#page-20-0)). Therefore, for this time scale, EVD-III has the least uncertainty. Regarding hydrological drought, increasing the time scale or reducing the sample size causes the uncertainty to increase from 0.90 to 1.36 and the error from 1.22 to



<span id="page-20-0"></span>**Fig. 8** Average of estimated error and uncertainty for the diferent functions and time scales. **a** and **c**: uncertainty of SPI 12 and 24; **b** and **d**: ARE of SPI 12 and 24; **e** and **g**: uncertainty of SRI 12 and 24; **f** and **h**: ARE of SRI 12 and 24

1.68. It is noteworthy that the uncertainty in SPI is lower and its error is higher compared to SRI. The greater the band diference, the lower the certainty and the greater the uncertainty (Vergni et al. [2015\)](#page-29-7). The higher SPI error compared to SRI is due to the fact that this index refects only precipitation data, since it does not include other factors, such as evaporation in the water balance. By contrast, the SRI refects all hydrological processes infu-encing runoff; therefore, the error can be lower (Cerpa Reyes et al. [2022\)](#page-28-22). Consequently, a longer period of time (more number of sampling) has less uncertainty, since the number of samples decreases as of the time scale increases, so the uncertainty bandwidth increases as well. Perhaps one of the most important factors of higher ARE in SPI than in SRI is the greater precipitation variability compared to runof.

Researchers such as Zhu et al. [\(2019](#page-29-14)), Vicente-Serrano et al. [\(2012](#page-29-15)), Shukla and Wood ([2008\)](#page-29-5) and Engeland et al. [\(2004](#page-28-23)) introduced the EDV-III function as the best distribution function in SRI modeling.

## **3.4 Comparison of the agreement between observed and estimated values in the severity of drought classes**

To compare the agreement of the observed values with the bootstrap estimate in the assessment of the drought severity classes presented in Table [9,](#page-21-0) the weighted kappa statistic was used. The highest agreement between observed and estimated drought values in meteorological drought (almost perfect agreement) has been calculated in the WEI-II for both the SPI-12 and SPI-24. Then, the EVD-III and the GAM-II are placed in the class of substantial and almost perfect agreement, and the PE-II is in the class of substantial and moderate agreement.

The lowest agreement (slight and poor agreement) was observed for SRI-12 and SRI-24 by the EVD-III. Other functions have substantial and moderate agreement. In terms of the time scale, the lowest agreement in both indicators was seen in the 24-month time scale. Therefore, in a summary, it can be said that the WEI-II has the highest agreement in meteorological and hydrological drought in a time scale of 12 months (longer sample size).

#### **3.5 Drought characteristics in observed and estimated values**

Based on the presented methodology, drought characteristics including severity, duration, magnitude and peak were estimated for sample stations S1 and SH1. Figure [9,](#page-22-0) [10](#page-22-1), [11](#page-23-0) and [12](#page-23-1) show the Max, Min and average characteristics for the observed and estimated values in diferent functions and time scales of SPI, SRI. It can be said that the characteristics of drought in the estimated values by bootstrap are far more than the observed values, especially in the 24-month scale in two indices.

Paired t-test was used to compare drought characteristics in observed and estimated characteristics. The statistical comparison of the mentioned characteristics is presented in Table [10.](#page-24-0) Based on the results, there is a signifcant diference between the severity and peak of SPI-12 estimated by the WEI-II and the PEI-III, while this signifcant diference does not emerge among any of the characteristics in diferent functions in SPI-24. An interesting result has been obtained considering the characteristics of drought in SRI-24 caused by the WEI-II. In fact, the results of the WEI-II have a signifcance diference with other distributions with the estimates in hydrological drought that are sometimes overestimated

<span id="page-21-0"></span>



<span id="page-22-0"></span>**Fig. 9** Drought severity of SPI 12 (**a**), SPI 24 (**b**), SRI12 (**c**), SRI 24 (**d**) for diferent distribution functions



<span id="page-22-1"></span>**Fig. 10** Drought duration of SPI 12 (**a**), SPI 24 (**b),** SRI12 (**c**), SRI 24 (**d**); for diferent distribution functions



<span id="page-23-0"></span>**Fig. 11** Drought magnitude of SPI 12 (**a**), SPI 24 (**b**), SRI12 (**c**), SRI 24 (**d**); for diferent distribution functions



<span id="page-23-1"></span>**Fig. 12** Drought peak of SPI 12 (**a**), SPI 24 (**b**), SRI12 (**c**), SRI 24 (**d**); for diferent distribution functions

or underestimated compared to the other distribution functions. This situation is not evident in meteorological drought.

According to Vergni et al.  $(2017)$  $(2017)$ , as expected, the uncertainty increases when each record length decreases. Therefore, the main source of uncertainty is the length of the record, and the efects due to the time scale are insignifcant. Because the uncertainties in time scale, record length and drought characteristics did not show signifcant diference.

<span id="page-24-0"></span>

Table 10 (continued)



Natural Hazards

The results of uncertainty of hydrological drought showed that two-parameter distributions (WEI-II, GAM-II) produce shorter confdence intervals than three-parameter distributions (EVD-III, PE\_III). However, due to the poorer ft of two-parameter distributions to runoff values, estimates from these two-parameter distributions may provide less reliable runoff probability. Therefore, drought characteristics will be significantly different from the values of two-parameter distribution functions.

Regardless of the type of distribution function and the number of parameters, the estimated SRI related to extreme values of runoff is always associated with an uncertainty. This uncertainty is due to the variability of the hypothetical distribution function that is used and this situation occurs when the data (runoff) approaches the tails of this hypothetical distribution function.

As regards the meteorological drought uncertainty, the WEI-II distribution function had the lowest uncertainty, and its drought characteristics showed a signifcant diference with the PEI-III. According to Zamani and Bazarfashan [\(2020](#page-29-16)), in the regions with high fuctuations of precipitation regimes, the standard deviation is high; therefore, the coefficient of variation of precipitation (the ratio of standard deviation to the average precipitation) is high, causing an increase in kurtosis and light tail. This situation indicates the high frequency of irregular rainfall events throughout the year in Mediterranean regions. In this case, the WEI-II distribution function will be the best estimate of drought (Table [10\)](#page-24-0).

## **4 Conclusion**

Drought investigation is now one of the most important challenges, especially in arid and semi-arid regions. SPI and SRI are simple and commonly used indicators for meteorology and hydrology drought. Despite their simplicity, both indicators are based on the statistical distribution of precipitation and runoff time series. The properties of these time series can depend on the estimates of parameters in statistical distributions, ultimately afecting the uncertainty of drought analysis. This is particularly important in the context of spatial and temporal drought analysis and projections.

The objectives of this work were to analyze the uncertainty of meteorological and hydrological droughts based on SPI and SRI in the wadi Mina basin in northwest Algeria. The following distributions of precipitation and runoff time series were considered: Gamma (GA), Weibull (WEI), Extreme value (EVD), and Pearson (PE). Bootstrap methods were used to generate samples of data, and subsequently, the uncertainty of both drought indicators was assessed.

The results show that all considered distributions were acceptable fts for the time series, and the time scale of the data was not signifcantly related to the goodness of ft. Moreover, there was no apparent relationship between the rejection cases and the scale and position of the regional stations or the investigated variable. The lack of signifcant diferences between time series for a particular distribution caused the averages estimated in SPI to fall within the same descriptive class. The most severe meteorological droughts were observed for the PE-III function and the EVD-III function. In the case of hydrological drought, the most severe drought was estimated by the WEI-II and the GAM-II. The type of periods (dry or wet) depends on the type of distribution of the data. In general, for meteorological drought, the EVD-III represents the most severe drought, and the WEI-II has estimated values of lower intensities. Regarding hydrological drought, all events are rated as the most severe for the EVD-III.

In the uncertainty analysis, most of the observed SPI values fell within the confdence range for all distributions, both for the SPI and the SRI. Only in the case of the SRI, results from the PEI-III were above the true value in the bootstrap band. Based on the weighted Kappa statistics, the highest agreement between observed and estimated values for both kinds of drought occurred with the WEI-II at SPI 12 and SPI 24. Following this function, the EVD-III and the GAM-II showed substantial and almost perfect agreement, while the PEI-III fell into the category of substantial and moderate agreement.

The results of the results, WEI-II has the lowest estimation error and uncertainty in meteorological drought in the two investigated time scales, but in hydrological drought it is related to EVD-III. Investigating the use of the two mentioned functions for drought monitoring in two medium-term and long-term time scales is recommended.

These results are important from a methodological perspective because they highlight the necessity of including uncertainty in drought analysis, especially now when future climate projections are subject to strong uncertainty. It is also crucial to select the statistical distribution for calculating SPI and SRI appropriately, especially for longer time-periods.

Considering the temporal distribution of precipitation and its efect on runof in Mediterranean regions, the results of this study can help to develop a meteorological/hydrological drought warning and monitoring system in similar regions.

**Acknowledgements** We thank the National Agency of the Water Resources (**ANRH**) for the collected data and the General Directorate of Scientifc Research and Technological Development of Algeria (**DGRSDT**).

**Author contributions** All authors contributed to the study's conception and design. MA, OB: Collecting data, preparing and calculation, editing and reviewing manuscripts; OB, ZP and FP: Data analysis, plotting, and manuscript preparation; TC: Writing introduction; AW: Writing conclusion. All authors read and approved the fnal manuscript.

**Funding** The authors declare that no funds, grants, or other support, were received during the preparation of this manuscript.

**Data availability** Not applicable.

## **Declarations**

**Confict of interest** The authors declare that they have no known competing fnancial interests or personal relationships that could have appeared to infuence the work reported in this paper.

**Ethical approval** We confrm that this article is original research and has not been published previously in any journal in any language.

**Consent to participate** All the authors read and approved the fnal version of the manuscript.

**Consent to publish** The authors confrm that the work described has not been published before, and it is not under consideration for publication elsewhere.

## **References**

<span id="page-27-1"></span>Achite M, Ouillon S (2007) Suspended sediment transport in a semiarid watershed, Wadi Abd, Algeria (1973–1995). J Hydrol 343(3–4):187–202

<span id="page-27-0"></span>Achite M, Simsek O, Adarsh S, Hartani T, Caloiero T (2023) Assessment and monitoring of meteorological and hydrological drought in semiarid regions: the Wadi Ouahrane basin case study (Algeria). Phys Chem Earth 130:103386

- <span id="page-28-20"></span>AghaKouchak A, Mirchi A, Madani K, Di Baldassarre G, Nazemi A, Alborzi A, Wanders N (2021) Anthropogenic drought: defnition, challenges, and opportunities. Rev Geophys 59(2):1–23
- <span id="page-28-8"></span>Angelidis P, Maris F, Kotsovinos N, Hrissanthou V (2012) Computation of drought index SPI with alternative distribution functions. Water Resour Manag 26:2453–2473
- <span id="page-28-14"></span>Azhdari Z, Bazrafshan O, Zamani H, Shekari M, Singh VP (2021) Hydro-meteorological drought risk assessment using linear and nonlinear multivariate methods. Phys Chem Earth Parts a/b/c 123:103046
- Bazarafshan O, Mahmoudzadeh F, Bazarafshan J (2015) Evaluation of drought changes based on standardized precipitation index and standardized evapotranspiration index in southern coasts of Iran. Sci Res J Desert Manag. <https://doi.org/10.22034/JDMAL.2017.24662>
- <span id="page-28-19"></span>Bazrafshan O, Zamani H, Shekari M (2020) A copula-based index for drought analysis in arid and semiarid regions of Iran. Nat Res Model 33(1):e12237
- <span id="page-28-10"></span>Blain GC (2012) Revisiting the probabilistic defnition of drought: strengths, limitations and an agrometeorological adaptation. Bragantia Campinas 71(1):132–141
- <span id="page-28-6"></span>Bordi I, Fraedrich K, Jiang J-M, Sutera A (2004) Spatio-temporal variability of dry and wet periods in eastern China. Theor Appl Climatol 79:81–91
- <span id="page-28-1"></span>Bordi I, Fraedrich K, Sutera A (2009) Observed drought and wetness trends in Europe: an update. Hydrol Earth Syst Sci 13:1519–1530
- <span id="page-28-2"></span>Buttafuoco G, Caloiero T (2014) Drought events at diferent timescales in southern Italy (Calabria). J Maps 10:529–537
- <span id="page-28-3"></span>Buttafuoco G, Caloiero T, Ricca N, Guagliardi I (2018) Assessment of drought and its uncertainty in a southern Italy area (Calabria region). Measurement 113:205–210
- <span id="page-28-22"></span>Cerpa Reyes LJ, Ávila Rangel H, Herazo LC (2022) Adjustment of the standardized precipitation index (SPI) for the evaluation of drought in the Arroyo Pechelín basin, Colombia, under zero monthly precipitation conditions. Atmosphere 13(2):236. <https://doi.org/10.3390/atmos13020236>
- <span id="page-28-16"></span>Cohen J (1968) Weighted kappa: Nominal scale agreement provision for scaled disagreement or partial credit. Psychol Bulletin 70(4):213
- <span id="page-28-15"></span>Dixon PM (2006) Bootstrap resampling. Encyclopedia Environmetrics
- <span id="page-28-23"></span>Engeland K, Hisdal H, Frigessi A (2004) Practical extreme value modelling of hydrological foods and droughts: a case study. Extremes 7:5–30
- <span id="page-28-13"></span>Ghasemnejad F, Fazeli M, Bazarafshan O, Parveen Nia M (2018) Uncertainty analysis of hydrological drought characteristics using Latin hypercube (case study: Minab Dam watershed). Pasture and Watershed Sci-Res J 72(2):527–542. <https://doi.org/10.22059/jrwm.2019.276424.1356>
- Ghasemnezhad F, Fazeli M, Bazrafshan O, Parvinnia M, Singh VP (2022) Uncertainty analysis of hydrological drought due to record length, time scale, and probability distribution functions using Monte Carlo simulation method. Atmosphere 13(9):1390
- <span id="page-28-21"></span>Guenang GM, Kamga FM (2014) Computation of the standardized precipitation index (SPI) and its use to assess drought occurrences in Cameroon over recent decades. J Appl Meteorol Climat 53(10):2310–2324
- <span id="page-28-9"></span>Guttman NB (1994) On the sensitivity of sample L moments to sample size. J Clim 7(6):1026–1029
- Guttman NB (1999) Accepting the standardized precipitation index: a calculation algorithm. J Am Water Res as 35:311–322
- <span id="page-28-7"></span>Guttman NB (1999) Accepting the standardized precipitation index: a calculation algorithm. J Am Water Res 35:311–322
- <span id="page-28-5"></span>Hayes MJ, Svoboda MD, Wilhite DA, Vanyarkho OV (1999) Monitoring the 1996 drought using the standardized precipitation index. Bull Am Meteorol Soc 80:429–438
- <span id="page-28-12"></span>Hu YM, Liang ZM, Liu YW, Wang J, Yao L, Ning Y (2015) Uncertainty analysis of SPI calculation and drought assessment based on the application of bootstrap. Int J Climatol 35:1847–1857
- <span id="page-28-18"></span>Khatun J, Ramkissoon K, Giddings MC (2007) Fragmentation characteristics of collision-induced dissociation in MALDI TOF/TOF mass spectrometry. Anal Chem 79(8):3032–3040
- <span id="page-28-17"></span>Landis JR, Koch GG (1977) The measurement of observer agreement for categorical data. Biometrics 159–174
- <span id="page-28-4"></span>Lloyd-Hughes B, Saunders MA (2002) A drought climatology for Europe. Int J Climatol 22:1571–1592
- <span id="page-28-0"></span>Mckee TB, Doseken NJ, Kleist J, (1993) The Relationship of Drought Frequency and Duration to Time Scales In: proc. 8th Conf. On Applied Climatology, American Meteorological Society, Massachusetts pp 179–184.
- <span id="page-28-11"></span>McKee TB, Doesken NJ and Kleist J (1993). The relationship of drought frequency and duration to time scales. Pp 179–184. In Proceedings of the 8th Conference of Applied Climatology. 17–22, Anaheim, California.
- Mishra AK, Desai VR (2005) Drought forecasting using stochastic models. Stoch Environ Res Risk Assess 19:326–339
- <span id="page-29-6"></span>Naumann G, Barbosa P, Carrao H, Singleton A, Vogt J (2012) Monitoring drought conditions and their uncertainties in Africa using TRMM data. J Appl Meteorol Clim 51:1867–1874
- <span id="page-29-11"></span>Pieper P, Düsterhus A, Baehr J (2020) A universal Standardized Precipitation Index candidate distribution function for observations and simulations. Hydrol Earth Syst Sci 24(9):4541–4565
- <span id="page-29-12"></span>Sherif M, Almulla M, Shetty A, Chowdhury RK (2014) Analysis of rainfall, PMP and drought in the United Arab Emirates. Int J Climatol 34(4):1318–1328
- <span id="page-29-5"></span>Shukla S and Wood AW (2008) Use of a standardized runoff index for characterizing hydrologic drought. Geophysical research letters. 35(2)
- Shukla S, Wood AW (2008) Use of a standardized runoff index for characterizing hydrologic drought. Geophys Res Lett 35:L02405
- <span id="page-29-1"></span>Spinoni J, Naumann G, Carrao H, Barbosa P, Vogt J (2014) World drought frequency, duration, and severity for 1951–2010. Int J Climatol 34:2792–2804
- <span id="page-29-9"></span>Stagge JH, Tallaksen LM, Gudmundsson L, Van Loon AF, Stahl K (2015) Candidate distributions for climatological drought indices (SPI and SPEI). Int J Climatol 35(13):4027–4040
- <span id="page-29-0"></span>Tramblay Y, Koutroulis A, Samaniego L, Vicente-Serrano SM, Volaire F, Boone A, Le Page M, Llasat MC, Albergel C, Burak S et al (2020) Challenges for drought assessment in the Mediterranean region under future climate scenarios. Earth-Sci Rev 210:103348
- <span id="page-29-8"></span>Verdonck FA, Jaworska J, Thas O, Vanrolleghem PA (2001) Determining environmental standards using bootstrapping, Bayesian and maximum likelihood techniques: a comparative study. Anal Chim Acta 446(1–2):427–436
- <span id="page-29-7"></span>Vergni L, Di Lena B, Todisco F, Mannocchi F (2017) Uncertainty in drought monitoring by the standardized precipitation index: the case study of the Abruzzo region (central Italy). Theor Appl Climatol 128(1–2):13–26
- <span id="page-29-4"></span>Vicente-Serrano SM, Beguería S, López-Moreno JI (2010) A multiscalar drought index sensitive to global warming: the standardized precipitation evapotranspiration index. J Clim 23:1696–1718
- <span id="page-29-15"></span>Vicente-Serrano SM, López-Moreno JI, Beguería S, Lorenzo-Lacruz J, Azorin-Molina C, Morán-Tejeda E (2012) Accurate computation of a streamfow drought index. J Hydrol Eng 17(2):318–332
- <span id="page-29-3"></span>Wu H, Hayes MJ, Wilhite DA, Svoboda MD (2005) The efect of the length of record on the standardized precipitation index calculation. Int J Climatol: A J Royal Meteorol Soc 25(4):505–520
- <span id="page-29-2"></span>Wu H, Svoboda MD, Hayes MJ, Wilhite DA, Wen F (2007) Appropriate application of the standardised precipitation index in arid locations and dry seasons. Int J Climatol 27:65–79
- <span id="page-29-16"></span>Zamani H, Bazrafshan O (2020) Modeling monthly rainfall data using zero-adjusted models in the semiarid, arid and extra-arid regions. Meteorol Atmos Phys 132(2):239–253
- <span id="page-29-13"></span>Zhang Y, Li Z (2020) Uncertainty analysis of standardized precipitation index due to the efects of probability distributions and parameter errors. Front Earth Sci 8:76
- <span id="page-29-10"></span>Zheng J, Zhou Z, Liu J, Yan Z, Xu CY, Jiang Y, Wang H (2023) A novel framework for investigating the mechanisms of climate change and anthropogenic activities on the evolution of hydrological drought. Sci Total Environ 900:165685
- <span id="page-29-14"></span>Zhu S, Xu Z, Luo X, Wang C, Zhang H (2019) Quantifying the contributions of climate change and human activities to drought extremes, using an improved evaluation framework. Water Res Manag 33:5051–5065

**Publisher's Note** Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Springer Nature or its licensor (e.g. a society or other partner) holds exclusive rights to this article under a publishing agreement with the author(s) or other rightsholder(s); author self-archiving of the accepted manuscript version of this article is solely governed by the terms of such publishing agreement and applicable law.

## **Authors and Afliations**

**Mohammad Achite1,2 · Ommolbanin Bazrafshan3  [·](http://orcid.org/0000-0003-2524-3992) Zohreh Pakdaman4 · Andrzej Wałęga<sup>5</sup> · Fateme Pourhaghverdi3 · Tommaso Caloiero6**

 $\boxtimes$  Ommolbanin Bazrafshan o.bazrafshan@hormozgan.ac.ir

> Mohammad Achite m.achite@univ-chlef.dz

Zohreh Pakdaman zpakdaman@hormozgan.ac.ir

Andrzej Wałęga andrzej.walega@urk.edu.pl

Fateme Pourhaghverdi fpourhaghverdi@hormozgan.ac.ir

Tommaso Caloiero tommaso.caloiero@cnr.it

- <sup>1</sup> Laboratory of Water and Environment, Faculty of Nature and Life Sciences, Hassiba Benbouali University of Chlef, 02180 Chlef, Algeria
- <sup>2</sup> Georessources, Environment and Natural Risks Laboratory, University of Oran, 2 Mohamed Ben Ahmed, P.Box 1015 El M'naouer, 31000 Oran, Algeria
- <sup>3</sup> Department of Natural Resources Engineering, Faculty of Agriculture and Natural Resources Engineering, University of Hormozgan, Bandar-Abbas 7916193145, Iran
- <sup>4</sup> Department of Statistics, Faculty of Science, University of Hormozgan, Bandar-Abbas 7916193145, Iran
- <sup>5</sup> Department of Sanitary Engineering and Water Management, University of Agriculture in Krakow, Mickiewicza 24/28 Street, 30-059, Krakow, Poland
- <sup>6</sup> National Research Council of Italy, Research Institute for Geo-Hydrological Protection (CNR˗IRPI), Rende, Italy