

Differential effects of paraoxonase 1 (PON1) polymorphisms on cancer risk: evidence from 25 published studies

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Abstract Paraoxonase is an HDL-associated enzyme that plays a preventive role against oxidative stress, which is thought to contribute to cancer development. PON1 activity varies widely among individuals, which is in part related to two common nonsynonymous polymorphisms in the PON1 gene (Q192R and L55M). The polymorphisms in PON1 have been implicated in cancer risk. However, results from the studies to date have been conflicting. To clarify the association, a meta-analysis was performed for 7,073 cases and 9,520 controls from 25 published case-control studies. Odds ratios (ORs) with 95% confidence intervals (CIs) were used to assess the strength of the association. Significant associations between PON1-L55M but not Q192R polymorphism and total cancer were

observed from all the comparisons. In stratified analyses, PON1-55M allele was a risk factor for breast cancer. Similarly, increased risk was observed for prostate cancer (OR = 1.18, 95% CI: 1.01–1.36, $P_{\text{heterogeneity}} = 0.260$) and Caucasian population (OR = 1.18, 95% CI: 1.02–1.38, $P_{\text{heterogeneity}} = 0.1$) of the LM genotype, compared with the LL genotype. For PON1-Q192R polymorphism, PON1-192R allele was a decreased risk factor for cancer in the Asian group (RR vs QQ: OR = 0.61, 95% CI: 0.38–0.98, $P_{\text{heterogeneity}} = 0.268$; QR vs QQ: OR = 0.71, 95% CI: 0.52–0.96, $P_{\text{heterogeneity}} = 0.130$; RR + QR vs QQ: OR = 0.71, 95% CI: 0.53–0.95, $P_{\text{heterogeneity}} = 0.135$). Although some modest bias could not be eliminated, this meta-analysis suggests that the *PON1-55M* allele is a risk factor for the development of cancer, in particular for breast cancer. Future studies with larger sample sizes are warranted to further evaluate these associations.

Dai-Hua Fang and Cong-Hai Fan are contributed equally to this work and should be considered as co-first authors.

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Introduction

Oxygen radicals can oxidize DNA bases, form mutagenic lesions and chromosome aberrations, and activate chemical carcinogens into highly reactive compounds. Elevated oxidative stress has been found to mediate carcinogenesis by causing metabolic malfunction and damage to biological molecules, including DNA [1]. It has been observed that oxidative stress is associated with an increased risk in various types of cancers [2, 3]. The level of oxidative stress is determined by the relative rates at which reactive oxygen species (ROS) are generated and detoxified. It has been reported that ROS may also play a role in human cancer

development [4]. The detoxification of ROS is accomplished by small molecules, such as vitamins C and E, and by enzymes which cleave to the ROS to generate nontoxic products [5]. Human serum paraoxonase1 (PON1) is an esterase enzyme that has lipophilic antioxidant characteristics and that participates in the elimination of ROS. It binds to high-density lipoprotein (HDL) and contributes to the detoxification of organophosphorus compounds, such as paraoxes and carcinogenic lipid-soluble radicals from lipid peroxidation [6–8].

PON1, a member of a family of proteins including PON2 and PON3, is located on the long arm of chromosome 7q21.3 [6]. PON1 is an esterase that is widely distributed among tissues such as the liver, kidney, and intestine; but it is also present in blood plasma. There is a 10- to 40-fold inter-individual variability in serum PON1 activity, as measured by rates of hydrolysis of the exogenous substrate paraoxon [6]. One source of the variability is the polymorphism of the *PON1* gene. Epidemiologic and molecular studies have identified that there are two important common functional genetic polymorphisms in the coding region of the gene at positions 55 and 192 of the *PON1* gene. Substitution of glutamine (Q genotype) at position 192 in exon 6 of the *PON1* gene by arginine (R genotype) leads to the first polymorphism (rs662A>G, Gln192Arg, Q192R, A192B). Similarly, substitution of leucine (L genotype) at position 55 in exon 3 by methionine (M genotype) leads to the second polymorphism (rs854560T>A, Leu55Met, L55M) [9, 10]. Studies revealed that polymorphisms of the *PON1* gene may affect PON1 activity. The PON1 activity of the PON1 192 Q allele carriers was reported to be lower than that of the R carriers [7, 11, 12]. The PON1-55L allele is correlated with higher PON1 activity and mRNA levels than PON1-55M allele [12, 13]. The study deduced that the lower activity of PON1 was caused by a decreased stability of the PON1-55M protein [14]. Reduced PON1 activities have been reported in different groups of patients, including those with diabetes mellitus, cardiovascular disease, and hypercholesterolemia. Furthermore, a number of epidemiological studies have investigated the associations between these polymorphisms and different cancers, such as lung [15], breast [16], brain [17], and ovarian [18] cancers. However, the results remained controversial [16, 17, 19, 20], partially because of small sample size, the difference in the genotype distribution by ethnicity, study design, assay characteristics, and so on. In light of the extensive role of polymorphisms in PON1, a pooled analysis is needed to accumulate data from different studies and to provide better evidence for or against that association. Therefore, we performed a meta-analysis to evaluate the association of polymorphisms in PON1 with cancer susceptibility in all eligible case–control studies.

Materials and methods

Identification of eligible studies

An electronic literature search was performed with PubMed, Embase, for all relevant reports (the last search update was as of Jun 23, 2011), using the key words “paraoxonase 1” or “PON1,” “polymorphism,” “tumor,” or “cancer.” The search was limited to human-associated studies. We also used the PubMed option “related articles” in each research article to search potentially relevant articles. In addition, studies were identified by a manual search of the reference lists of reviews and retrieved studies. When more than one of the same or overlapping population by different researchers or overlapping data by the same authors were found, only the most recent or complete study was used for this meta-analysis. Studies, regardless of sample size, were enrolled if they met the following inclusion criteria: (i) a study of the PON1 Q192R or L55M polymorphism and cancer risk; (ii) a case–control study; (iii) with available genotype frequency.

Data extraction

Two of the authors (Dai-Hua Fang and Qiang Ji) extracted all data from eligible publications independently that met the inclusion criteria and reached the consensus for any controversy. For each study, the following characteristics were collected: the first author’s last name, year of publication, the country and ethnicity of study population, the number of genotyped subjects in cases and controls, source of control groups (population- or hospital based), the types of cancers, and genotyping methods. Different ethnic descents were categorized as Caucasian, Asian, or mixed ethnicity, including more than one ethnic descent.

Statistical analysis

A meta-analysis was used to examine the overall association of the PON1-L55M and -Q192R polymorphisms with the risk of cancer by odds ratio (ORs) with 95% confidence intervals (CIs). As compared to the wild-type LL or QQ homozygote, the risk of carriage of the M or R allele (i.e., LM and MM or RR and QR genotypes) on cancers was estimated, followed by evaluating the risk of LM +MM versus LL or QR + RR versus QQ on cancer in the dominant model and MM versus LM + LL or RR versus QR + QQ on cancer in recessive effects, respectively. Stratified analyses were also performed by ethnicity, researched methods, and cancer types (if only one cancer type contained less than two individual studies, it was combined into the “Other Cancers” group).

The statistical significance of the pooled OR was determined with the Z test, and a P value of <0.05 was considered significant. Heterogeneity across the studies was evaluated by a χ^2 test based on a Q statistic test [21], and was considered significant if $P < 0.05$. A fixed-effect model using the Mantel–Haenszel method and a random-effects model using the DerSimonian and Laird method were used to pool the results [22]; the fixed-effect model was used as well when there was no heterogeneity across results of the studies, or the random-effect model was used. Moreover, a sensitivity analysis was performed to assess the stability of the results, by which a single study in the meta-analysis was deleted each time to determine the influence of the individual data set to the overall pooled OR [23]. To test the publication bias, Funnel plots and the Egger’s linear regression test were applied [24]. The Hardy–Weinberg equilibrium for controls was also tested by the χ^2 test for accuracy of fit, using a web-based program (<http://ihg.gsf.de/cgi-bin/hw/hwa1.pl>). All statistical tests for this meta-analysis were performed with STATA version 10.0 (Stata Corporation College Station, TX, USA).

Results

Characteristics of studies

A total of 25 eligible studies were enrolled for this study according to the preset inclusion criteria (Fig. 1), in which

7,073 cases and 9,520 controls were included for the pooled analysis. For PON1-Q192R polymorphism, all 25 studies reported the available data, including five brain tumors studies, five breast cancer studies, three prostate cancer studies, and the others, which were categorized into the “Other Cancers” group. There were 15 studies of Caucasian descendents, two of Asian descendents, and eight with mixed ethnicity. Cancers were diagnosed histologically or pathologically in most studies. The polymerase chain reaction–restriction fragment length polymorphism (PCR–RFLP) and TaqMan assay were performed in 12 and 11 studies, respectively, and sequencing was applied by two studies. In addition, most of the controls were sex- and age-matched for the case groups, of which 20 were population based and five were hospital based. For PON1-L55M polymorphism, 16 studies with available data from ten Caucasian descendents and one of Asian descendents, along with five with mixed ethnicity, were collected for the pooled analysis, including four breast cancer studies, three prostate cancer studies, and nine studies that were categorized into the other cancers groups (shown in Supplemental table). The genotype distributions among the controls of all studies were not deviated from the Hardy–Weinberg equilibrium, with the exception of two studies [25, 26] on PON1-L55M polymorphism and seven studies [5, 19, 25, 27–30] on PON1-Q192R polymorphism; however Hussein et al described that the distribution of Q192R genotypes of their study was also consistent with the Hardy–Weinberg equilibrium in controls ($P = 0.710$) (Table 1).

Fig. 1 Studies identified with criteria for inclusion and exclusion

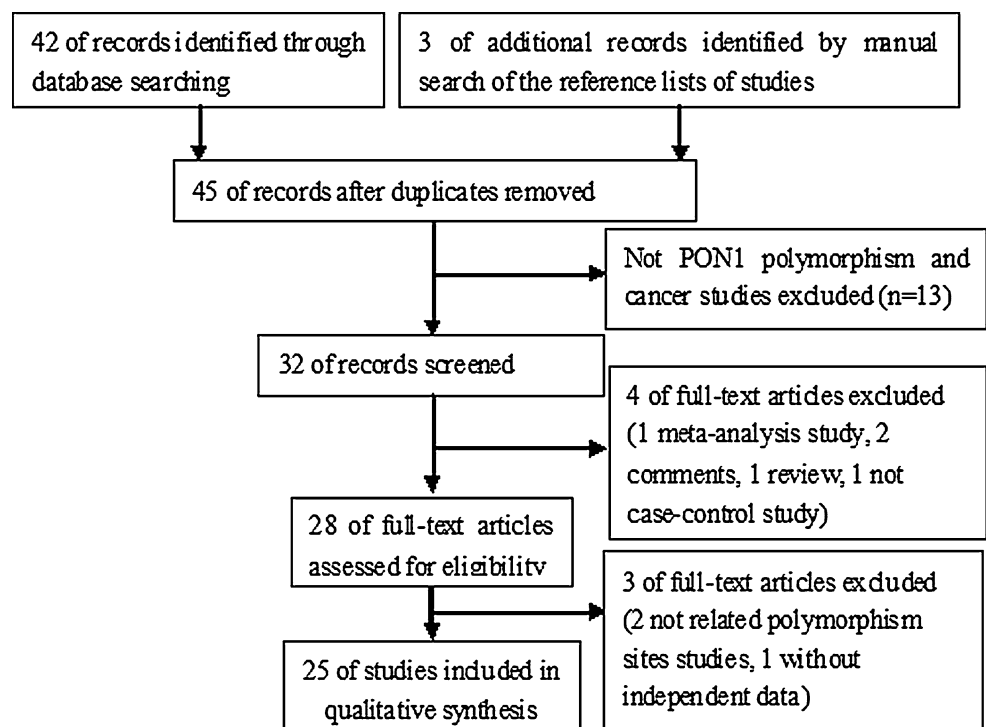


Table 1 Stratified analyses of the PON1-R192Q polymorphism on cancer risk

Variables	Case/control	RR versus QQ			QR versus QQ		
		OR (95% CI)	<i>P</i> ^a	<i>I</i> ²	OR (95% CI)	<i>P</i> ^a	<i>I</i> ²
Total	7073/9520	0.89 (0.69, 1.15)	0.000	74.5	0.89 (0.73, 1.09)	0.000	84.5
Cancer types							
Prostate cancer	1661/2460	0.52 (0.24, 1.13)	0.013	76.9	1.38 (0.83, 2.28)	0.003	82.7
Brain tumors	936/1244	1.01 (0.76, 1.35)	0.417	0.0	1.15 (0.88, 1.51)	0.706	0.0
Breast cancer	1575/2283	0.54 (0.29, 1.02)	0.002	77.2	0.59 (0.34, 1.02)	0.000	89.9
Other	2901/3533	1.23 (0.86, 1.76)	0.000	70.7	0.90 (0.68, 1.20)	0.000	78.8
Ethnicities							
Caucasian	2332/4238	0.85 (0.50, 1.45)	0.000	81.5	0.82 (0.54, 1.26)	0.000	89.8
Asian	564/429	0.61 (0.38, 0.98)*	0.268	18.4	0.71 (0.52, 0.96)*	0.130	56.4
Mix	4177/4853	1.00 (0.81, 1.23)	0.055	49.3	1.03 (0.94, 1.13)	0.193	29.5
Source of controls							
Population based	5942/7567	0.92 (0.70, 1.21)	0.000	72.6	0.82 (0.66, 1.02)	0.000	84.0
Hospital based	1131/1953	0.85 (0.40, 1.79)	0.000	80.8	1.25 (0.76, 2.06)	0.000	82.8
Genotype method							
PCR–RFLP	2492/2396	0.75 (0.42, 1.28)	0.000	82.2	0.74 (0.47, 1.18)	0.000	91.2
TaqMan	4485/6154	0.96 (0.78, 1.18)	0.043	46.9	0.98 (0.85, 1.13)	0.026	51.0
Sequencing	86/970	0.55 (0.11, 21.3)	0.023	80.6	1.59 (0.47, 5.39)	0.048	74.4
Variables	Case/control	RR + QR versus QQ			RR versus QR + QQ		
		OR (95% CI)	<i>P</i> ^a	<i>I</i> ²	OR (95% CI)	<i>P</i> ^a	<i>I</i> ²
Total	7073/9520	0.89 (0.73, 1.08)	0.000	85.5	0.93 (0.74, 1.15)	0.000	70.8
Cancer types							
Prostate cancer	1661/2460	1.14 (0.99, 1.30)	0.273	22.9	0.45 (0.17, 1.15)	0.001	85.3
Brain tumors	936/1244	1.09 (0.92, 1.30)	0.576	0.0	0.97 (0.74, 1.28)	0.513	0.0
Breast cancer	1575/2283	0.58 (0.32, 1.03)	0.000	92.1	0.63 (0.36, 1.10)	0.007	71.8
Other	2901/3533	0.95 (0.72, 1.26)	0.000	81.2	1.26 (1.00, 1.60)	0.028	48.9
Ethnicities							
Caucasian	2332/4238	0.82 (0.55, 1.24)	0.000	90.3	0.87 (0.54, 1.40)	0.000	79.2
Asian	564/429	0.71 (0.53, 0.95)*	0.135	55.3	0.89 (0.63, 1.25)	0.840	0.0
Mix	4177/4853	1.01 (0.93, 1.11)	0.116	39.4	0.99 (0.82, 1.21)	0.056	49.1
Source of controls							
Population based	5942/7567	0.83 (0.67, 1.05)	0.000	87.0	1.00 (0.80, 1.23)	0.000	63.4
Hospital based	1131/1953	1.11 (0.76, 1.61)	0.005	73.3	0.73 (0.36, 1.50)	0.000	81.9
Genotype method							
PCR–RFLP	2492/2396	0.75 (0.48, 1.14)	0.000	91.5	0.79 (0.47, 1.34)	0.000	81.6
TaqMan	4485/6154	0.99 (0.86, 1.13)	0.016	54.2	0.97 (0.86, 1.10)	0.162	29.8
Sequencing	86/970	1.56 (0.37, 6.58)	0.017	82.4	1.52 (0.76, 3.01)	0.119	58.8

*I*²: 0–25, no heterogeneity; 25–50, modest heterogeneity; >50 high heterogeneity

PCR polymerase chain reaction, RFLP restriction fragment length polymorphism

^a *P* value of *Q* test for heterogeneity test, * Statistically significant, with *P* < 0.05

Meta-analysis

Overall, there was no association between PON1-R192Q polymorphism and the risk of cancer for any overall comparison. However, subgroup analysis revealed that there were significantly decreased risks of cancer in Asian

population for comparison of RR versus QQ (OR = 0.61, 95% CI: 0.38–0.98, *P*_{heterogeneity} = 0.268), RQ versus QQ (OR = 0.71, 95% CI: 0.52–0.96, *P*_{heterogeneity} = 0.130) and dominant-model comparison (RR + RQ vs. QQ) (OR = 0.71, 95% CI: 0.53–0.95, *P*_{heterogeneity} = 0.135). For the polymorphism of PON1-L55M, increased cancer risks were

observed for the overall comparisons (Table 2; Fig. 2). Furthermore, in the cancer type subgroup analysis, we observed increased risks of breast cancer for all comparisons, prostate cancer for the comparison of LM versus LL (OR = 1.18, 95% CI: 1.01–1.36, $P_{\text{heterogeneity}} = 0.260$), and other cancer types for the comparison of MM versus LL + LM (OR = 1.20, 95% CI: 1.02–1.42, $P_{\text{heterogeneity}} = 0.260$). Similarly, in the source-of-control subgroup analysis, increased cancer risk was observed in the hospital based control group for all comparisons, and in the population based control group for the comparison of LM versus LL (OR = 1.09, 95% CI: 1.00–1.20, $P_{\text{heterogeneity}} = 0.201$). In

addition, increased cancer risk was also observed in the Caucasian population for the comparison of LM versus LL (OR = 1.18, 95% CI: 1.02–1.38, $P_{\text{heterogeneity}} = 0.157$) and in the research method based on the polymerase chain reaction and restriction fragment length polymorphism (PCR–RFLP), as summarized in Table 2.

There was significant heterogeneity across the studies in overall comparisons for two polymorphism sites. For this, the source of heterogeneity was explored for the heterozygote comparison (RQ vs. QQ for PON1-R192Q and LM vs. LL for PON1-L55M) by between subgroups (Cancer type, ethnicity, and source of controls). For PON1-R192Q

Table 2 Stratified analyses of the PON1-L55M polymorphism on cancer risk

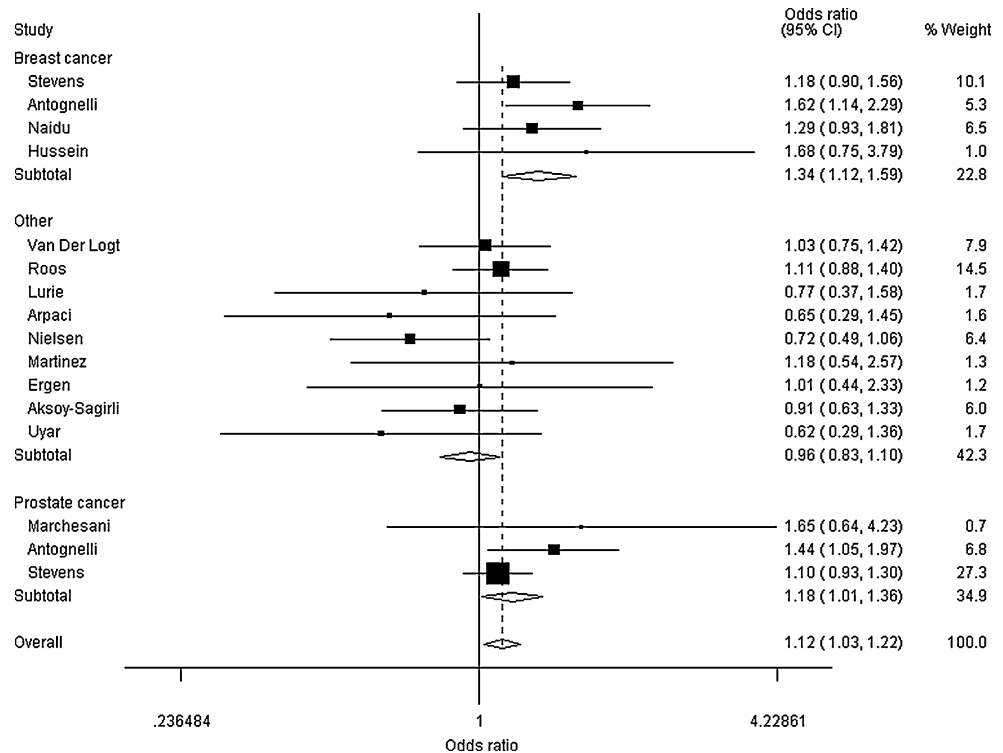
Variables	Cases/controls	MM versus LL			LM versus LL		
		OR (95% CI)	P^a	I^2	OR (95% CI)	P^a	I^2
Total	5175/6147	1.33 (1.03, 1.71)*	0.000	69.6	1.12 (1.03, 1.22)*	0.149	27.3
Cancer types							
Prostate cancer	1659/2457	1.20 (0.62, 2.33)	0.014	76.8	1.18 (1.01, 1.36)*	0.260	25.7
Breast cancer	1517/1379	2.63 (1.33, 5.21)*	0.242	28.4	1.34 (1.12, 1.60)*	0.522	0.0
Other	1999/2311	1.20 (0.91, 1.38)	0.319	13.8	0.96 (0.83, 1.10)	0.572	0.0
Ethnicities							
Caucasian	1872/2809	1.31 (0.87, 1.95)	0.001	67.7	1.18 (1.02, 1.38)*	0.157	31.5
Mix	2916/3086	1.12 (0.95, 1.32)	0.151	40.6	1.07 (0.95, 1.20)	0.237	27.7
Source of controls							
Population based	4691/5687	1.22 (0.92, 1.61)	0.000	69.9	1.09 (1.00, 1.20)*	0.201	23.4
Hospital based	484/460	2.12 (1.45, 3.08)*	0.452	0.00	1.47 (1.09, 1.97)*	0.725	0.00
Genotype method							
PCR–RFLP	2166/2066	1.47 (0.99, 2.18)	0.001	69.9	1.19 (1.04, 1.37)*	0.115	38.1
TaqMan	2989/3306	1.12 (0.95, 1.31)	0.240	25.9	1.07 (0.95, 1.20)	0.348	10.6
Variables	Cases/controls	MM +LM versus LL			MM versus LL + LM		
		OR (95% CI)	P^a	I^2	OR (95% CI)	P^a	I^2
Total	5175/6147	1.18 (1.00, 1.39)*	0.000	65.6	1.28 (1.04, 1.57)*	0.000	65.2
Cancer types							
Prostate cancer	1659/2457	1.25 (0.92, 1.71)	0.080	60.3	0.97 (0.80, 1.19)	0.034	70.4
Breast cancer	1517/1379	1.66 (1.22, 2.25)*	0.021	69.3	1.84 (1.53, 2.20)*	0.403	0.0
Other	1999/2311	0.99 (0.87, 1.14)	0.472	0.0	1.20 (1.02, 1.42)	0.230	24.0
Ethnicities							
Caucasian	1872/2809	1.21 (0.91, 1.61)	0.000	71.2	1.24 (0.90, 1.71)	0.006	61.0
Mix	2916/3086	1.08 (0.97, 1.21)	0.275	0.0	1.20 (0.94, 1.53)	0.035	61.3
Source of controls							
Population based	4691/5687	1.12 (0.95, 1.32)	0.001	64.3	1.21 (0.97, 1.52)	0.000	67.1
Hospital based	484/460	1.65 (1.26, 2.17)*	0.270	18.0	1.72 (1.23, 2.40)*	0.425	0.0
Genotype method							
PCR–RFLP	2166/2066	1.23 (0.93, 1.63)	0.000	74.4	1.40 (1.03, 1.91)*	0.010	60.4
TaqMan	2989/3306	1.08 (0.97, 1.21)	0.395	3.4	1.17 (0.95, 1.46)	0.064	52.0

I^2 : 0–25, no heterogeneity; 25–50, modest heterogeneity; >50 high heterogeneity

PCR polymerase chain reaction, RFLP restriction fragment length polymorphism

^a P value of Q-test for heterogeneity test, * Statistically significant, with $P < 0.05$

Fig. 2 ORs (log scale) of cancer associated with L55M for the LM genotype compared with the LL genotype. For each study, the estimate of OR and its 95% CI is plotted with a box and a horizontal line. Filled diamond pooled OR and its 95% CI



polymorphism, cancer type ($\chi^2 = 49.66$, $df = 3$, $P = 0.000$) and source of controls ($\chi^2 = 12.89$, $df = 1$, $P = 0.000$) but not ethnicity ($\chi^2 = 5.77$, $df = 2$, $P = 0.056$) contributed substantially to that heterogeneity. For PON1-L55M polymorphism, cancer type ($\chi^2 = 9.01$, $df = 2$, $P = 0.011$) but not ethnicity ($\chi^2 = 1.92$, $df = 2$, $P = 0.373$) and source of controls ($\chi^2 = 3.54$, $df = 1$, $P = 0.060$) contributed substantially to that heterogeneity.

Sensitivity analyses

For PON1-R192Q polymorphism, sensitivity analysis revealed that nine independent studies were the main source of heterogeneity; two were prostate cancer [19] and breast cancer [25] related studies, respectively. In the other cancer type subgroup, two studies were related to lung cancer [15, 31] and five studies were related to ovarian cancer [18], bladder cancer [28], osteosarcoma [29], non-Hodgkin's lymphoma [32] and multiple myeloma [33]. The heterogeneity was decreased when these seven studies were removed (RR vs. QQ: $P_{\text{heterogeneity}} = 0.124$, QR vs. QQ: $P_{\text{heterogeneity}} = 0.149$, RR + QR vs. QQ: $P_{\text{heterogeneity}} = 0.060$, RR vs. QQ + QR: $P_{\text{heterogeneity}} = 0.201$). Significant association were observed in subgroup of breast cancer (QR vs. QQ: OR = 0.79, 95% CI: 0.66–0.95, $P_{\text{heterogeneity}} = 0.555$; RR + QR vs. QQ: OR = 0.79, 95% CI: 0.66–0.94, $P_{\text{heterogeneity}} = 0.659$; RR vs. QQ + QR: OR = 0.76, 95% CI: 0.58–1.00, $P_{\text{heterogeneity}} = 0.512$) after the study carried by

Antognelli et al [25] removed. For race subgroup analysis, seven studies [15, 18, 19, 25, 28, 29, 33] were the main source of heterogeneity of Caucasian population, and found to impact the pooled OR (data not shown). In addition, no other single study was found to impact the pooled OR as indicated by sensitivity analyses.

For PON1-L55M polymorphism, sensitivity analysis revealed that three independent studies were the main source of heterogeneity; two were prostate cancer [5, 19] and one was breast cancer [25] related studies, respectively. The heterogeneity was decreased when these three studies were removed (MM vs. LL: $P_{\text{heterogeneity}} = 0.067$, MM + LM vs. LL: $P_{\text{heterogeneity}} = 0.096$, MM vs. LL + LM: $P_{\text{heterogeneity}} = 0.131$). In addition, no other single study was found to impact the pooled OR as indicated by sensitivity analyses.

Publication bias

Begg's funnel plot and the Egger's test were performed to assess the publication bias of the currently available literature. The shape of the funnel plots did not reveal any evidence for obvious asymmetry in all comparison models. Then, the Egger's test was used to provide statistical evidence for funnel plot symmetry. The results still did not show any evidence of publication bias ($t = -1.15$, $P = 0.263$ for QR vs. QQ; $t = -0.620$, $P = 0.545$ for LM vs. LL).

Discussion

In this meta-analysis, we investigated the association between PON1-L55M and Q192R polymorphisms and risk of cancer. The results revealed that PON1-L55M but not Q192R polymorphism was associated with increased risk for developing cancers. The LM/MM genotype and M allele increased the risk of cancer occurrence, especially for breast cancer. Given the important roles of PON1 in the regulation of the oxidative stress and inflammation, which were believed to be important in carcinogenesis, we deduced that PON1-L55M polymorphism may modulate the risk of the development of cancer, especially in a type of breast cancer.

Oxidative stress and free radicals have been associated with an increased risk in various types of cancers [2, 34]. The human body has a number of endogenous free-radical scavenging systems. PON1, an antioxidant enzyme, may cause defects in the antioxidant/oxidant balance [35]. This can trigger oxidative stress and the formation of ROS. Studies have revealed that PON1 expression is depressed in human lung cancer [36], pancreatic [37], and gastric cancer [38]. In the case of PON1, PON1-55M allele has been found to have significantly lower PON1 expression than PON1-55L [13, 39], resulting in an increased risk of developing cancers [5, 18, 19, 25, 40] and suggesting PON1-55M allele was the risk of cancer. Furthermore our results were confirmed by the genetic model-free approach, which reported by Minelli et al. [41]. In the present study, significant associations were observed between the PON1-L55M polymorphism and prostate cancer for the comparison of LM versus LL and other cancer types for the comparison of MM versus LL + LM; however, a significant association was observed in breast cancer for the all comparisons, suggesting that different effects PON1-L55M polymorphism in different cancer types. The risk factors for breast cancer, including BRCA1 mutations and increased estrogen metabolism, are known to influence oxidative stress [42]. Moreover, since oxidative stress may be involved in cell proliferation and malignant conversion during the development of breast cancer [43], it is reasonable to expect that PON1, as a part of the lipid peroxidation scavenging systems, may influence breast cancer development. In the subgroup analysis, a significant association was observed between PON1-L55M polymorphism and cancer risk in that two studies based on hospital controls. In contrast, there was no association was observed in studies based on population controls. This association may be the reason that the two studies based on hospital control were the main source of heterogeneity for PON1-L55M; furthermore, the significant associations were pooled from two studies based on hospital control, and they may be affected by limited studies.

In contrast to L55M polymorphism, our data failed to observe statistical significance in the distribution of overall allele and genotype frequencies of Q192R between cancers, and the results were confirmed by the genetic model-free approach [40]. However, in the subgroup analysis, we observed that PON1-192 RR/QR or R allele acts as a decreased risk for cancer from two studies based on the Asian population. Studies revealed that PON1 192 Q allele carriers were reported to be lower than that of the R carriers [7, 11, 12], and a lower PON1 level was regarded as a risk for cancer; furthermore, allele distributions in control subjects varied significantly by ethnic group and were similar to those reported by the National Center of Biotechnology Information for Caucasian (Q: 0.668) and Asian population (Q: 0.430). Further cumulated studies would be needed to reveal the association in the Asian population, since the result presented in that study was from the data of two Asian population studies.

However, results derived from this meta-analysis should be interpreted with caution. First, when a stratified analysis of tumor type, ethnicity, or control recourse was performed, the number of each subgroup seemed to be smaller. Second, there was no further evaluation of potential gene–gene interactions and gene–environment interactions (i.e., the interaction between insecticide exposure and PON1 was deduced with higher risk for brain cancers), due to the lack of original data from the studies. Finally, the number of published studies was not enough for a comprehensive analysis, particularly for any kind of the cancers and ethnicities. In spite of these limitations, this meta-analysis had several strengths. First, no publication biases were detected, indicating that summary results may be unbiased. Second, in the sensitivity analysis, no individual study affected the pooled OR which indicated that our results were statistically trustworthy.

In conclusion, this meta-analysis suggests that the PON1-L55M polymorphism may contribute to the genetic susceptibility of cancer, particularly for breast cancer. However, further studies would be needed to clarify the possible roles of the PON1 polymorphisms in the etiology of cancer.

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