ORIGINAL ARTICLE



Correlation of rs9331888 polymorphism with Alzheimer's disease among Caucasian and Chinese populations: a meta-analysis and systematic review

Bin Zhu¹ · Rui Min Wang² · Jian Ting Wang³ · Rui Ling Chen¹ · Yan Fei Zheng⁴ · Lei Zhang⁵ · Zhi Gang Zhao¹

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Abstract Clusterin polymorphism (rs9331888) was reported to be associated with the susceptibility to alzheimer's disease (AD). Nevertheless, the results were inconclusive. To derive a more precise estimation of this association, this meta-analysis was conducted. We've conducted a comprehensive search of PubMed, Embase, CNKI and AlzGene database for casecontrol studies published throughout October, 2016 that evaluated the role of rs9331888 gene variants in AD patients. Odds ratios (ORs) and their 95% confidence intervals (CIs) were calculated to assess the strength of associations between the rs9331888/C > G polymorphism and AD disease. A total of 9 studies were enrolled in the Meta Analysis. The overall analysis revealed a significant association between the rs9331888/C > G polymorphism and AD disease in the recessive model (GG vs. GC + CC: OR = 1.11, 95% CI: 1.05–1.18; P < 0.01). Sub-group analysis revealed that the Caucasian populations which with recessive model (GG vs. GC + CC: OR = 1.12, 95% CI: 1.06–1.2; P < 0.01) were dramatically

Bin Zhu and Rui Min Wang contributed equally to this work.

Zhi Gang Zhao 1022zzg@sina.com

- ¹ Present address: Department of Pharmacy, Beijing Tiantan Hospital Affiliated to Capital Medical University, Beijing 100050, China
- ² Present address: Department of Hospital medical room, Beijing Tiantan Hospital Affiliated to Capital Medical University, Beijing 100050, China
- ³ Department of nephropathy, People's Hospital Affiliated to FuJian University of Traditional Chinese Medicine, Fujian 350004, China
- ⁴ School of Basic Medicine, Beijing University of Chinese Medicine, Beijing 100029, China
- ⁵ Department of Pharmacy, Beijing Shijitan Hospital, Capital Medical University, Beijing 100038, China

related to AD, while no significant association was found in the Chinese populations among the five genetic models. Our meta-analysis demonstrated that the rs9331888/C > G polymorphism in the clusterin gene might contribute to AD susceptibility especially in Caucasian populations. Whereas the relationship of the polymorphism to the disease in Chinese populations was still in controversial. Additional welldesigned studies, with larger sample sizes, are required to further elucidate this association.

Keywords Clusterin · rs933188 · Polymorphism · Alzheimer's disease

Introduction

Alzheimer's disease's (AD) is a chronic, multifactorial and personality changing of neurodegenerative disorder. As one of the most common types of neurodegenerative disease, people suffered from AD manifest as progressive and irreversible memory loss and cognitive decline (Scheltens et al. 2016). Approximately 1% of people who are over 65 years old are under threat of AD, while the number increases to 25%–35% among people over 85 years old (Lu et al. 2014). It is estimated that that there are now 46.8 million people living with dementia worldwide, with numbers projected to nearly double every 20 years, increasing to 74.7 million by 2030 and 131.5 million by 2050(World Alzheimer Report 2015:The Global Impact of Dementia). In American, an estimated 5.3 million Americans have AD, and of them 5.1 million are age 65 years (2015 Alzheimer's disease facts and figures 2015).

Although the potential cause of AD is still unclear, the involvement of heredity genetic risk factors in AD's predisposition and progression is widely acknowledged. Recent advantages in multistage Genome Wide Association Studies (GWAS) have identified several loci conferring risk for AD (Bodily et al. 2016; Cuyvers and Sleegers 2016). Since 2009, a number of susceptibility genes, such as Bridging integrator 1 (BIN1), Clusterin (*CLU*), CD2 associated protein (CD2AP), and ATP binding cassette subfamily A member 7(ABCA7) were gradually reported to be correlated with AD disease (Zhang et al. 2016). Of all the reported loci, rs9331888 in the *CLU* gene was considered to be significantly associated with AD (Xing et al. 2012).

Clusterin which is also known as apolipoprotein J (Apo J), located at chromosome 8p21-p22 extending over 16 kb, is a multifunctional lipoprotein expressed in almost all mammalian tissues, especially in the brain. The protein binds to $A\beta$ peptides and fibrils to prevent aggregation and also it can involve in AB clearance via binding to megalin receptors and increasing endocytosis by glial cells (Li et al. 2014; Zhou et al. 2014). Large-scale GWAS identified that rs9331888 (G > C) single nucleotide polymorphisms (SNPs) in CLU are significantly associated with AD in populations of Caucasian ancestry. And Merve Alayliog lu et al. also got the same results in Turkey populations (Alaylioglu et al. 2016). However, inconsistent results regarding this variant had been reported in Chinese populations. Lu et al. had reported that there was weak association or no association between the rs9331888 polymorphism and AD in southern Chinese populations (Lu et al. 2014). Meanwhile, Edina et al. found that there was no evidence of plasma clusterin lever correlated with Alzheimer"s disease (Silajdzic et al. 2012). Therefore, to confirm the association between rs933188 polymorphism and AD, we performed a meta-analysis of case-control studies by pooling all eligible studies to evaluate the overall risk and influence of ethnic factors to this disease.

Methods

Literature search and inclusion criteria

To identify all relevant publications focus on the risk for AD and rs9331888 polymorphism, we conducted a comprehensive literature search of electronic databases, including the Pubmed, Embase and China National Knowledge Infrastructure (CNKI). Eligible case-control studies were extracted with the last search update on October 1, 2016. The following terms were used: "Alzheimer's disease", "demential", "Clusterin,", "*CLU*", "APO J", OR "rs933188" and "polymorphism" OR "Variant" without any limitation applied. The reference lists of retrieved studies and recent reviews were also manually searched for further relevant studies. AlzGene database (www.alzgene.org, updated April 18, 2011) was also utilized in our searching process.

Inclusion and exclusion criteria

Studies in this meta-analysis must meet the following inclusion criteria: (1) evaluation of the association between rs933188 polymorphism and the AD; (2) case-control study; (3) studies focusing on human being; (4) providing detail genotype frequencies; Exclusion criteria: (1) duplication of previous publications; (2) comment, review and editorial; (3) study without detailed genotype data.

Data extraction

The data of the eligible studies were extracted in duplicate by two investigators (Zhu and Zhang) independently in duplicate with a standard data-collection form. The following data was collected from each study if available: (1)first author's name; (2) years of publication; (3) country of origin; (4) participants ethnicity; (5) Hardy-Winberg equilibrium; (6) AD diagnosis criteria; (7) genotyping method; (8) numbers of cases and controls; (9) counts of cases and controls for each genotype. When there were multiple publications from the same population, only the one with largest study was included. Any discrepancy was resolved through discussion until a consensus was reached. If the dissent still existed, the third investigators would be involved to resolve the dispute.

Quality assessment

The literature quality was evaluated by using the quality scoring criteria modified from previous study (Zhang et al. 2017) by two authors independently. Quality scores ranged from 0 point (worst) to 10 points (best). Studies scoring higher than 5 points were classified with adequate quality. Disagreement was settled through discussion among four of the investigators.

Statistics analysis

Hardy-Weinberg equilibrium was evaluated using Chi-square test in control groups and a *P* value <0.05 was considered significant disequilibrium. The odds ratio (OR) and 95% confidence interval (95% CI) were calculated to evaluate the strength of AD susceptibility and *CLU* polymorphism. Pooled ORs were obtained from combination of single study by heterozygote comparison (GC vs. CC), homozygote comparison (GG vs. CC), dominant model (GG + GC vs. CC), recessive model (GG vs. GC + CC), and allelic model (G vs. C) respectively. Heterogeneity was evaluated by Q statistic and I² statistic. Once, Q-test >0.10 or I² < 50%, the fixedeffect model (Mantel–Haenszelmethod) was used to calculate the pooled ORs, otherwise, the random-effect model (DerSimonian–Laird method) was used. The significance of the pooled ORs was assessed by Z-test, where P < 0.05 indicated statistically significant. Publication bias was assessed by Begg's funnel plots and Begg's test quantitatively. If publication bias was indicated, we further evaluated the number of missing studies in a meta-analysis by applying the trim and fill method and recalculated the pooled risks estimate with the addition of those missing studies. All statistical analyses were performed using the STATA software version 12.0 (STATA Corporation, College Station, TX, USA).

Results

Characteristics of studies

The initial search identified a total of 179 citations, 44 of which were duplicated. 19 citations were included for further review by title and abstract screening of the remaining 135 citations. After examination of full text articles, 9 case-control studies were identified as being eligible for analysis. The PRISMA flow diagram is illustrated in Fig. 1. Among the studies, 8 were Caucasians and 5 were Chinese population (including five sub studies). Multiple genotyping methods were performed in the studies, including PCR-RFLP, TaqMan and DNA sequencing. All studies were complied with Hardy–Weinberg equilibrium (HWE) in the controls and the genotyping distribution was in agreement with HWE. The characteristics of involved articles were summarized in Tables 1 and 2.

Meta-analysis results

Heterogeneity was firstly identified by Q-test and I-squared statistic in five genetic models. As is showed in the Fig. 2, serious heterogeneity were found in Allele model ($I^2 = 70.8\%$), Homozygote model ($I^2 = 70.4\%$) and Dominant model ($I^2 = 65.1\%$), thus random effect model was applied. The results revealed that no significant associations between the rs9331888/C > G polymorphism and AD disease in above three genotype distributions (G vs. C: OR = 0.968,95% CI: 0.887–1.056, P = 0.464; GG vs. CC: OR = 0.959, 95% CI: 0.790–1.163, P = 0.669; GG + GC vs.



Fig. 1 Flow sheet summarizing study identification and selection

CC:0.987, 95% CI: 0.839–1.160, P = 0.872). Fixed effect model were used in the Heterozygote model ($I^2 = 46.7\%$) and Recessive model ($I^2 = 47.7\%$), and results showed that the recessive model (GG vs. GC + CC: OR = 1.11, 95% CI: 1.05–1.18; P < 0.01) was significantly associated with AD whereas no effect was found in Heterozygote model (CG vs. CC: OR = 0.962, 95% CI: 0.876–1.056, P = 0.416) with AD development.

Subgroups based on ethnicity were utilized to further analyze the relationship of polymorphism with AD. In Caucasian populations, AD was proved to be correlated with rs9331888 polymorphism under the recessive model (GG vs. GC + CC: OR = 1.12, 95% CI: 1.06–1.2; P < 0.01). Nevertheless, no significant differences were observed in any other genetic models (G vs. C: OR = 0.939, 95% CI: 0.863-1.023, P = 0.149; CG vs. CC: OR = 0.942, 95% CI: 0.797–1.112, *P* = 0.479; GG vs. CC: OR = 0.905, 95% CI: 0.732–1.118, *P* = 0.353; GG + GC vs. CC:OR = 0.923, 95% CI: 0.764– 1.114, P = 0.404); In Chinese populations, we did not observed any correlations of the clusterin polymorphism under five genetic models (G vs. C: OR = 0.991, 95% CI: 0.805-1.220, P = 0.931; GG vs. CC: OR = 1.008, 95% CI: 0.683-1.490,P = 0.966; CG vs. CC: OR = 1.098, 95% CI: 0.888-1.358, P = 0.390; GG + GC vs. CC:OR = 1.066, 95% CI: 0.812–1.399, P = 0.644; GG vs. GC + CC: OR = 0.931, 95% CI: 0.850 - 1.019, P = 0.711) (Fig. 3).

Publication bias

Publication bias was evaluated by using the Begg and Egger tests. Allele model and Homozygote model were found with significant evidence of publication bias (Allele model: Begg P = 0.044,Egger P = 0.028; Homozygote model: Begg P = 0.059, Egger P = 0.018) (Fig. 4). Nevertheless, the application of the trim and fill method did not change the risk estimate (Allele model: P = 0.859, 95% confidence interval:-0.168 to 0.201;Homozygote model:P = 0.669, 95% confidence interval:-0.235 to 0.151). No missing studies were imputed in the contour enhanced funnel plots.

Sensitivity analysis

We performed sensitivity analysis to assess the influence of each individual study on the pooled OR by sequentially removing each eligible study. The results indicated that the removal of the study by Lambert et al. led to the loss of a significant association with the risk of AD in the overall pooled population under a recessive model. However, this effect was attributed to loss of power of the meta-analysis due to the overall high weight of this study (30.25%). We did not find that any single study influence the quality of the pooled ORs in the sensitivity analyses(Fig. 5).

Metab Brain Dis (2017) 32:981-989

	Year	Country	Age		Methods	HWE	Quality
			AD	Control			
Kamboh te al. (2012)	2010	USA	72.6 + 6.4	74.7 + 6.5	TaqMan	Y	5
Lambert et al. (2009)	2009	Belgium	78.6 + 8.1	67.0 + 12.9	TaqMan/Sequenom	Y	7
Lambert et al. (2009)	2009	Finland	71.4 + 7.5	69.2 + 6.0	TaqMan/Sequenom	Y	7
Lambert et al. (2009)	2009	Italy	76.6 + 8.7	72.3 + 8.9	TaqMan/Sequenom	Y	7
Lambert et al. (2009)	2009	Spain	75.3 + 9.3	76.9 + 10.9	TaqMan/Sequenom	Y	7
Lambert et al. (2009)	2009	France	73.7 + 8.9	73.8 + 5.4	TaqMan/Sequenom	Y	7
Alaylioglu et al. (2016)	2015	Turkey	76.5 + 6.13	75.4 + 7.03	TaqMan	Y	4
Gu et al. (2011)	2011	USA&German	76.7 + 7.0	76.1 + 7.1	PCR	Y	5
Lu et al. (2014)	2014	China	69.99 + 9.9	68.93 + 9.3	Sequenom	Y	4
Liu (2014)	2014	China	74.57 + 5.9	72.02 + 5.5	Sequenom	Y	6
Wang (2014)	2011	China	78.17 + 5.4	74.56 + 6.3	TaqMan	Y	4
Yu et al. (2010)	2010	China	76.87 + 5.6	75.93 + 4.7	PCR	Y	6

Table 1 Characteristics of the studies included in this meta-analysis

Discussion

Alzheimer's disease is the most common form of dementia in elder, accounting for 50% of all dementia. It had been proved to be one of common diseases with strong genetic component. Study on SNP provided a compelling evidence for a role of genetic variation in AD risk assessment, diagnosis and development of new therapies (Jiao et al. 2015). Clusterin gene located in chromosome 8p21-p12 which is a chromosomal region of interest in AD and it may explain around 9% of the late-onset AD attributable risk (Tan et al. 2016). The study by Lacour et al. found the rs11136000and rs9331888 polymorphism of *CLU* gene showed strong association with AD (Lacour et al. 2017). Haplotype analyses found that rs9331888 in combination with rs11136000 is likely to play a key role in the progress of AD. Mengel et al. had found that patients who carriers of the common rs11136000 and rs9331888 TTC haplotype in the *CLU* gene performed cognitively better than non-carriers and carriers of a rare TCC haplotype performed worse on the cognitive composite score(Mengel-From et al. 2013). Considering of the importance of rs9331888 in AD, we believed that it is critically important to explore the relationship of haplotype and the genotype in the morbidity of AD.

The rs9331888 polymorphism is located in the first exon of transcript NM_203339 and is one of the functional DNA variants underlying this association between CLU and AD. It regulates A β fibril formation and toxicity and facilitates

P for HWE Ethnicity AD Control MAF G CC CG GG CC CG GG AD Control Kamboh Caucasian 125 564 633 125 563 635 0.31 0.31 0.99 Lambert Caucasian 93 429 550 39 197 265 0.71 0.73 0.78 Lambert Caucasian 83 284 219 77 314 247 0.62 0.63 0.13 Lambert Caucasian 128 619 727 86 474 681 0.7 0.74 0.78 Lambert 57 408 Caucasian 315 358 62 330 0.71 0.72 0.67 Lambert 416 2773 0.90 Caucasian 211 854 960 2139 0.68 0.72 Merve Alay Caucasian 9 84 90 21 64 69 0.721 0.656 0.32 10 49 47 Gu Hui Ying Caucasian 13 49 36 0.675 0.617 0.56 Lu Shen Ji Chinese 116 265 114 137 293 150 0.498 0.501 0.79 121 Liu Xiao Yan Chinese 66 52 42 103 67 0.472 0.575 0.83 Wang Bei 53 80 82 172 104 0.549 0.531 0.50 Chinese 145 Jin Tai Yu Chinese 63 158 103 110 184 94 0.562 0.479 0.33 0.5055 Lu Hua Chen Chinese 109 235 114 79 177 86 0.5102 0.51

Table 2 Summary of Genotypefrequencies of rs9331888 amongAD cases and controls











Fig. 2 Forest plots of the rs9331888/C > G polymorphism under five genetic models. **a** is the Allele model (G vs. C), **b** is the Homozygote model (GG vs. CC), **c** is the Heterozygote model (GG vs. GC), **d** is the Dominant model (GG + GC vs. CC) and **e** is the Recessive model (GG vs. GC + CC)

amyloid- β (A β) transport across the blood-brain barrier (Bettens et al. 2015; Lidstrom et al. 1998). However,

inconsistent results reported in Chinese and Caucasian populations. Thambisetty et al. showed that *CLU* influenced

Δ		
Study		%
ID	OR (95% CI)	Weight
Caucasus		
Kamboh (2012)	1.00 (0.89, 1.12)	10.02
Lambert (2009)	0.94 (0.80, 1.11)	8.41
Lambert (2009)	0.93 (0.79, 1.09)	8.52
Lambert (2009)	0.83 (0.74, 0.94)	9.93
Lambert (2009)	0.95 (0.81, 1.11)	8.75
Lambert (2009)	0.84 (0.78, 0.91)	11.11
Gu Hui Ying (2011)	1.28 (0.86, 1.93)	3.32
Merve Alay (2015)	- 1.36 (0.98, 1.89)	4.46
Subtotal (I-squared = 59.6%, p = 0.016)	0.94 (0.86, 1.02)	64.53
China		
Lu Hua Chen (2011)	0.98 (0.80, 1.20)	7.47
Lu Shen Ji (2014)	0.95 (0.80, 1.12)	8.34
Liu Xiao Yan (2014)	0.66 (0.51, 0.86)	5.72
Wang Bei (2011)	1.07 (0.86, 1.34)	6.79
Jin Tai Yu (2010)	1.39 (1.13, 1.72)	7.14
Subtotal (I-squared = 79.8%, p = 0.001)	0.99 (0.80, 1.22)	35.47
Overall (I-squared = 70.8%, p = 0.000)	0.97 (0.89, 1.06)	100.00
NOTE: Weights are from random effects analysis		
507 1	197	



Study		%
D	OR (95% CI)	Weigh
Caucasus		
Kamboh (2012)	1.00 (0.77, 1.30)	9.79
Lambert (2009)	0.89 (0.60, 1.31)	7.46
Lambert (2009)	0.83 (0.60, 1.16)	8.45
Lambert (2009)	0.78 (0.59, 1.04)	9.34
Lambert (2009)	0.99 (0.68, 1.44)	7.71
Lambert (2009)	0.73 (0.61, 0.87)	11.38
Gu Hui Ying (2011)	1.47 (0.61, 3.52)	2.72
Merve Alay (2015)	* 3.05 (1.35, 6.88)	3.05
Subtotal (I-squared = 57.8%, p = 0.020)	0.92 (0.76, 1.11)	59.91
China		
Lu Hua Chen (2011)	0.96 (0.69, 1.34)	8.48
Lu Shen Ji (2014)	1.01 (0.76, 1.34)	9.37
Liu Xiao Yan (2014)	0.65 (0.42, 1.01)	6.66
Wang Bei (2011)	- 1.26 (0.86, 1.86)	7.49
Jin Tai Yu (2010)	1.64 (1.15, 2.33)	8.09
Subtotal (I-squared = 66.2%, p = 0.019)	1.07 (0.81, 1.40)	40.09
Overall (I-squared = 65.1%, p = 0.001)	0.99 (0.84, 1.16)	100.00
NOTE: Weights are from random effects analysis		
145 1	6.88	

C		
Study		%
ID	OR (95% CI)	Weight
Caucasus		
Kamboh (2012)	1.00 (0.76, 1.32)	11.36
Lambert (2009)	• 0.91 (0.61, 1.38)	5.37
Lambert (2009)		7.64
Lambert (2009)	0.88 (0.65, 1.18)	10.32
Lambert (2009)	1.04 (0.70, 1.54)	5.47
Lambert (2009)	0.79 (0.66, 0.95)	27.71
Gu Hui Ying (2011) -	* 1.30 (0.52, 3.24)	0.90
Merve Alay (2015)	* 3.06 (1.31, 7.14)	0.72
Subtotal (I-squared = 41.7%, p = 0.100)	0.90 (0.80, 1.01)	69.50
China		
Lu Hua Chen (2011)	0.96 (0.68, 1.36)	7.15
Lu Shen Ji (2014)		9.32
Liu Xiao Yan (2014)	• 0.75 (0.47, 1.19)	4.55
Wang Bei (2011)	1.30 (0.87, 1.97)	4.48
Jin Tai Yu (2010)	• 1.50 (1.03, 2.18)	5.00
Subtotal (I-squared = 38.4%, p = 0.165)	1.10 (0.93, 1.30)	30.50
	· · · · ·	
Overall (I-squared = 46.7%, p = 0.032)	0.96 (0.88, 1.06)	100.00
.14	1 7.14	



Fig. 3 Forest plots of the rs9331888/C > G polymorphism in subgroup analysis under five genetic models. **a** is the Allele model (G vs. C), **b** is the Homozygote model (GG vs. CC), **c** is the Heterozygote model (CG vs. CC), **d** is the Dominant model and **e** was the Recessive model (GG + GC vs. CC)

longitudinal changes in brain function and faster cognitive decline and GWAS studies have also identified that

rs9331888 in *CLU* was substantially associated with AD risk in individuals of Caucasian ancestry and other populations





(Thambisetty et al. 2013). Tan et al. showed that AD risk rs9331888 allele was associated with a decrease in *CLU* plasma levels in Chinese populations (Tan et al. 2016). Nevertheless, other researches such as Lu et al. showed that the rs9331888 variants may not be an AD susceptibility factor in southern Chinese Han population. A meta-analysis by Zhang et al. also reported that there was no significant association with East Asian population (Zhang et al. 2016). To solve the problem, we conduct a meta-analysis to investigate the pooled effect size of association between *CLU* SNP rs9331888 and AD in different ethnic backgrounds.

Our results implied a conspicuously significant relationship between the *CLU* polymorphism and AD risk under the recessive genetic model but not in other model. To further analysis the influence of ethnic factor to AD risk, a subgroup was employed in our research. The results revealed that in recessive model Caucasian populations are even more easily to progress to AD (P < 0.05), which is in agreement with previous observations (Jiao et al. 2015). Nevertheless, we did not found any association of polymorphism of rs933188 with AD risk in Chinese population, which was in accordance with the results drawn by Shuai et al., and these suggesting that the SNP may be an ethnicity-dependent factor in AD progress. Additional, subgroup analysis results of Caucasian populations were with difference, they had just showed that there were significant association of SNP rs9331888 with AD among allelic model, homozygote model, recessive model, and dominant model, whereas in our research we had showed that only the recessive model were significant associated with AD risk.

There were fewer researches to study the relationship of rs9331888 polymorphism with AD. Prior to this study, two

Fig. 5 Sensitivity analysis examining the association between the rs9331888/C > G polymorphism and risk of AD under the recessive model



groups performed meta-analyses to detect the correlation of rs9331888 polymorphism in patients with AD, and our results are in accordance with the two studies which suggest that rs9331888 polymorphism contributes to Alzheimer's disease susceptibility in Caucasian but not in Chinese populations (Shuai et al. 2015). However, compared with the prior study, our study was an update of the former researches. In this metaanalysis, we had enrolled more researches not only including in the Alzgene database but also recently published studies especially in Chinese populations. Totally, 9292 AD patients and 11,958 controls were included in the study, which can provide enough statistical power and strengthened the reliability of our results. Upon including eligible studies, a methodological quality assessment was conducted and all studies had acceptable quality.

Due to several limitations of the present meta-analysis, the results of our study should be interpreted with caution. To be specific, publication biases exist in the 9 studies in the five genetic models, which indicated that there might be problems due to insufficient articles. In addition, small sample size in each study might be the cause of the failure to achieve statistical significance. Additional, as there were not unified detection methods, serious heterogeneity were observed in our study even in subgroup analysis which may be one of the main risk factor explain for the negative results.

In summary, our data from an independent and large casecontrol sample revealed that SNP rs9331888 displayed significant association with AD in Caucasian populations under the Recessive model. However, limitations still exist, such as that we did not observed the influence of rs11136000 and rs9331888 TTC haplotype in the progress of AD development. Additional, as with limited sample size, the relationship of rs9331888 polymorphism with AD susceptibility in Chinese populations was worth further exploration. In future, more studies with genotype and haplotype data are required to further verify the results.

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Compliance with ethical standards

Conflict of interest The authors declare that they have no conflicts of interest.

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