DOI 10.1007/s10958-016-2796-z Journal of Mathematical Sciences, Vol. 214, No. 4, April, 2016 PROBABILITIES OF SMALL DEVIATIONS OF THE WEIGHTED SUM OF INDEPENDENT RANDOM VARIABLES WITH COMMON DISTRIBUTION THAT DECREASES AT ZERO NOT FASTER THAN A POWER

## L. V. Rozovsky<sup>\*</sup>

UDC 519.2

The paper presents estimates of small deviation probabilities of the sum  $\sum_{j\geq 1} \lambda_j X_j$ , where  $\{\lambda_j\}$  are positive numbers and  $\{X_j\}$  are i.i.d. positive random variables satisfying weak restrictions at zero and infinity. Bibliography: 16 titles.

1. Introduction and results. Let  $S = \sum_{j \ge 1} \lambda(j) X_j$ , where  $\{X_i\}$  are independent copies of a positive random variable X with distribution function V(x) and let V(x) be a bounded

a positive random variable X with distribution function V(x), and let  $\lambda(\cdot)$  be a bounded positive *nonincreasing* function defined on the interval  $[1, \infty]$ . We assume that the series S converges almost surely, which is equivalent to the condition

$$\sum_{j\geq 1} \mathbf{E} \min\left(1, \lambda(j) X\right) < \infty.$$
(1.1)

It is known that a convergence rate of the probability  $\mathbf{P}(S < r)$  to zero as  $r \searrow 0$  depends on the behavior of the distribution V at zero and at infinity, and the latter, in a sense, is determined by condition (1.1). So, if  $\lambda(n) \approx n^{-A}$ , A > 1, or  $\lambda(n) \approx q^n$ , 0 < q < 1, then (1.1) is equivalent to the conditions  $\mathbf{E}X^{1/A} < \infty$  and  $\mathbf{E}\log(1+X) < \infty$ , respectively.

Henceforth,  $u(y) \approx v(y) \iff \log \{u(y)/v(y)\} = O(1), y \to \infty.$ 

If the behavior of V at zero is not too restricted, then it is possible to get an optimal *logarithmic* asymptotics for  $\mathbf{P}(S < r)$  (see, for instance, [1–6]).

In order to obtain more exact estimates, it is usually assumed that V decreases at zero as a power, for example, it is a regularly varying function (see [7, 8, 16]), or satisfies essentially milder condition proposed in [9] (see also [10]):

## **L**. There exist constants $b \in (0, 1)$ , $c_1, c_2 > 1$ , and $\varepsilon > 0$ such that for each $r \leq \varepsilon$ ,

$$c_1 V(br) \le V(r) \le c_2 V(br).$$
 (1.2)

Later on, results from [9] were improved in [11] and [12]. In the last paper, the condition  $\mathbf{L}$  was replaced by a more general assumption  $\mathbf{R}$  (see below), which, in particular, allows V(r), in addition, to decrease at zero slower than *any* power of r (for instance, to be a slowly varying function).

For y > 0, set

$$\nu(y) = \frac{1}{y} \int_{0}^{y} u \, dV(u)$$

and introduce the following condition:

**R**. There exist constants  $b \in (0,1)$ ,  $c_1 > b$ ,  $c_2 > 1$ , and  $\varepsilon > 0$  such that for every  $r \leq \varepsilon$ ,

$$c_1 \nu(br) \le \nu(r) \le c_2 \nu(br).$$
 (1.3)

\*St.Petersburg State Chemical Pharmaceutical Academy, St.Petersburg, Russia, e-mail: l\_rozovsky@mail.ru.

540 1072-3374/16/2144-0540 ©2016 Springer Science+Business Media New York

Published in Zapiski Nauchnykh Seminarov POMI, Vol. 431, 2014, pp. 178–185. Original article submitted November 5, 2014.

It was shown in [12] that  $\mathbf{L} \iff \mathbf{R}\Big|_{c_1 > 1}$ . At the same time, if

$$\nu(y) \asymp l(y), \ y \to +0, \tag{1.4}$$

where l(y) is a function slowly varying at zero, then

$$V(y) \asymp \widetilde{l}(y) = \int_{0}^{y} l(u)/u \, du, \ y \to +0 \tag{1.5}$$

and  $\tilde{l}(y)$  is also slowly varying at zero. From (1.4) and (1.5), it follows that (1.3) is satisfied for *b* small enough, whereas the left-hand side of condition (1.2) is violated. Thus, **R** is weaker than **L**.

 $\operatorname{Set}$ 

$$f(u) = \mathbf{E}e^{-uX}, \quad L(u) = \sum_{n \ge 1} \log f(u\lambda(n)), \ u \ge 0.$$
 (1.6)

We assume (see [14, (3.2)]) that the distribution V satisfies at infinity the condition

**F**.  $\limsup s^2 \mathbf{P}(X \ge s) / \mathbf{E} X^2 \mathbf{1}[X < s] < \infty.$ 

The condition **F** holds if X belongs to the attraction domain of a stable law and, in particular, has a finite variance. It also implies that  $\mathbf{E}X^{\delta} < \infty$  for some positive  $\delta$  (see [11]).

**Theorem 1.** ([12]). Let conditions (1.1),  $\mathbf{R}$ , and  $\mathbf{F}$  be satisfy and, in addition,

$$\tau^2(u) = u^2 L''(u) \to \infty, \quad u \to \infty.$$
(1.7)

Then

$$\mathbf{P}(S < r) \sim \frac{\exp\left(L(u) + ur\right)}{\tau(u)\sqrt{2\pi}}, \quad r \to 0,$$
(1.8)

where the function u = u(r) is a unique solution of the equation

$$L'(u) + r = 0. (1.9)$$

**Remark 1.** Condition (1.7) follows from **L**. Also, it is valid (without any assumption concerning the behavior of V at zero) if the conditions **F** and  $\lambda(j) \approx e^{-g(j)}$  hold, where the function g is nondecreasing and g(u)/u monotonically tends to zero at infinity (see [12]). On the other hand, if, for example,  $\lambda(j) \approx \exp(-j^2)$ ,  $\mathbf{E}X^2 < \infty$ , and (see (1.4))  $l(y) \approx \log^{-2} y, y \to +0$ , then  $\liminf_{u\to\infty} \tau(u) = 0$ .

For rapidly decreasing weights  $\lambda(j)$ , the necessary condition (1.1) is significantly weaker than the condition **F** playing a key role in the proof of Theorem 1.

In [13], the asymptotics of the probability  $\mathbf{P}(S < r)$  was studied in the case where the condition  $\mathbf{L}$  is satisfied and the condition  $\mathbf{F}$  is not. In particular, the following result was obtained there.

**Theorem 2.** Let conditions (1.1), L, and

$$\sup_{m \ge 1} \lambda(m l) / \lambda(m) \le A \lambda(l), \quad l \ge 1,$$
(1.10)

hold, where A is a positive constant. Then

$$\mathbf{P}(S < r) \approx \frac{\exp\left(L(u) + ur\right)}{\tau(u)}, \quad r \to 0,$$
(1.11)

where the function u = u(r) satisfies equation (1.9).

Note that (1.10) holds if  $\lambda(n) \approx e^{-g(\log n)}$ , where the function g(y)/y does not decrease for all y large enough. For example, as  $\lambda(n)$ , one can take  $n^{-\delta}$  or  $e^{-n^{\delta}}$  with  $\delta > 0$ .

In the present paper, the behavior of the probability  $\mathbf{P}(S < r)$  is studied provided that **L** is replaced by **R** and the condition **F** is relaxed.

Let us formulate the results.

**Theorem 3.** Let conditions (1.1), **R**, and (see (1.6))

$$\limsup_{s \to 0} s |f'''(s)| / f''(s) < \infty$$
(1.12)

hold. Then

$$\mathbf{P}(S < r) \asymp \frac{\exp\left(L(u) + ur\right)}{1 + \tau(u)}, \quad r \to 0,$$
(1.13)

where the function u = u(r) satisfies the equation (1.9).

**Remark 2.** Condition (1.12) holds if the function

$$\beta(x) = \mathbf{E} \, X^2 \, \mathbf{I}[X < x]$$

satisfies the condition

$$\sup_{x \ge u} x^{-k} \beta(x) = O\left(u^{-k} \beta(u)\right), \quad u \to \infty,$$
(1.14)

for some k > 0, or, more generally, if

$$\sup_{t\geq 1}\beta(tu)/\beta(t)=O\left(g(u)\right),\quad u\to\infty,$$

where  $\int_{1}^{\infty} g(u) u e^{-u} du < \infty$ .

Note that (1.14) is weaker than **F**. Indeed, **F** implies

$$x^{-2} \beta(x) \le x^{-2} \mathbf{E} \min(x^2, X^2) \le u^{-2} \mathbf{E} \min(u^2, X^2) = O(u^{-2} \beta(u)), \quad u \to \infty$$

for any  $x \ge u > 0$ .

**Remark 3.** Theorem 3 is still valid if the function u in (1.13) satisfies the conditions

$$\frac{L'(u) + r}{\sqrt{L''(u)}} = O(1), \quad r \to 0,$$
(1.15)

and

$$L'(\varepsilon u) + r \le 0 \le L'(u/\varepsilon) + r, \tag{1.16}$$

where  $\varepsilon \in (0, 1)$  is a constant and r > 0 is small enough.

In other words, the exact solution of equation (1.9) can be replaced by an approximate one.

**Remark 4.** Set 
$$I_0(u) = \int_{1}^{\infty} \log f(u\lambda(t)) dt$$
. If the function  $\lambda(\cdot)$  satisfies the condition

$$\int_{1}^{\infty} |\log'' \lambda(t)| \, dt < \infty \tag{1.17}$$

and, moreover,

$$\int_{1}^{\infty} |(s \log' f(s))'| \, ds < \infty, \tag{1.18}$$

542

then (see (1.13))

$$L(u) = I_0(u) + 0.5 \log f(u) + O(1), \quad u \to \infty.$$
(1.19)

We remark that (1.18) holds if the function  $s \log' f(s)$  is monotone at infinity. Some other conditions can be found in [10] and [15].

The following consequence of Theorem 3 is similar to [13, Theorem 4].

**Theorem 4.** Let conditions (1.1), **R**, (1.12), (1.17), (1.18), and  $s^2 I_0''(s) \to \infty$ ,  $s \to \infty$ , hold. Then

$$\mathbf{P}(S < r) \asymp \sqrt{\frac{f(u)}{u^2 I_0''(u)}} e^{I_0(u) + ur}, \quad r \to 0,$$

where the function u = u(r) satisfies the conditions

$$I_0'(\varepsilon u) + r \le 0 \le I_0'(u/\varepsilon) + r \quad and \quad \left|\frac{I_0'(u) + r}{\sqrt{I_0''(u)}}\right| < 1/\varepsilon$$

with constant  $\varepsilon \in (0,1)$  and all positive r small enough.

2. Proofs. Below, we essentially use results from [12] and [13].

Let a random variable X(u),  $u \ge 0$ , have distribution  $e^{-ur} V(dr)/f(u)$ . From [13, (3.4), (3.5)] it follows that

From 
$$[15, (5.4), (5.5)]$$
, it follows that

$$\mathbf{P}(S < r) \le \frac{e^{L(u) + ur}}{\tau(u)} \left(\frac{1}{\sqrt{2\pi}} + 6\,\tau(u)\,\mu(u)\right),\tag{2.1}$$

where

$$\mu(u) = \frac{1}{\sigma^3(u)} \sum_{j \ge 1} \lambda_j^3 \mathbf{E} |X(u\,\lambda(j)) - \mathbf{E}X(u\,\lambda(j))|^3$$
(2.2)

and the function u satisfies condition (1.9). Moreover,

$$\mathbf{P}(S < r) \le e^{L(u) + ur}.\tag{2.3}$$

We have,

$$\mu(u) \le 8 \sum_{j\ge 1} \lambda^3(j) \operatorname{\mathbf{E}} X^3(u\lambda(j)) \le 8A/\tau(u),$$
(2.4)

where owing to [12, (2.5) for  $h \ge 1$ ] and (1.12) (see also [14, Lemma 2.2]),

$$A = \sup_{\gamma > 0} \gamma \mathbf{E} X^3(\gamma) / \mathbf{Var} X(\gamma) < \infty.$$
(2.5)

The upper bound in (1.13) follows from (2.1) and (2.3)–(2.5).

Next, from [13, (3.6), (3.7), and (3.10)] with regard to (2.2), (2.4), and (2.5), it follows that for any K > 0,

$$\mathbf{P}(S < r) \ge e^{L(u) + ur - 2K} \max\left(1 - \frac{\tau^2(\bar{u})}{K^2}, \frac{1}{\tau(\bar{u})} \left(K \bar{\Phi}(K/\tau(\bar{u})) - 48A\right)\right),$$
(2.6)

where u and  $\bar{u}$  satisfy conditions (1.9) and  $\bar{u}(r + L'(\bar{u})) = K$ , respectively,  $\bar{\Phi}(t) = \frac{1}{t} \int_{0}^{t} d\Phi(X)$ , and  $\Phi(\cdot)$  is the standard normal distribution. Moreover,

$$|\tau^{2}(\bar{u}) - \tau^{2}(u)| \le (2 + 8A) K.$$
(2.7)

The lower bound in (1.13) follows from (2.6) and (2.7), provided that K is large enough. Theorem 3 is proved. Let us consider Remark 1. The first statement follows from [9]. To check the second one, we observe that if  $u \lambda(n+1) < 1 \leq u \lambda(n)$ , then

$$\tau^2(u) \le \left(\sum_{1 \le j \le n} + \sum_{j > n}\right) (u\,\lambda(j))^2 \operatorname{\mathbf{E}} X^2(u\,\lambda(j)) = J_1 + J_2.$$

In our example,

$$J_1 = O\left(u^2 \sum_{j>n} \lambda^2(j)\right), \quad J_2 = O\left(u^2 \sum_{1 \le j \le n} \frac{1}{|\log u\lambda(j)|}\right), \quad u \to \infty$$

Let  $u = u_N = e^{(N+1/2)^2}$ , where N = 1, 2, ... Then n = N and, accordingly to the previous estimates, for  $z_N = (N + 1/2)^2$  and  $N \to \infty$ , we have

$$J_1 = O\left(\sum_{j \ge N} e^{-2(j^2 - z_N)}\right) = O(e^{-cN}), \quad J_2 = O\left(\sum_{1 \le j \le N} \frac{1}{z_N - j^2}\right) = O(\log N/N).$$

Thus,  $\lim_{N\to\infty} \tau(u) = 0$ . Remark 1 is completely checked.

Remark 2 is a consequence of the obvious estimates  $f''(s) \ge \beta(1/s)/e$  and  $s |f''(s)| \le \beta(1/s) + \int_{1}^{\infty} \beta(x/s) e^{-x}(x-1) dx$ .

Remark 3 is a consequence of the following relations (see [13, (3.10) and below]), in which the function h = h(r) satisfies the equation L'(h) + r = 0:

$$\left|\log\frac{\tau^{2}(u)}{\tau^{2}(h)}\right| = \left|\int_{h}^{u} \frac{d\tau^{2}(u)}{\tau^{2}(u)}\right| \le (2+8A) \left|\int_{h}^{u} \frac{du}{u}\right| = (2+8A) \left|\log\frac{u}{h}\right|,$$
$$\le L(u) + ur - (L(h) + hr) = \int_{h}^{u} (L'(t) + r) dt \le |(u-h)| |L'(u) + r| = (u-h)^{2} L''(\widetilde{u}), \ \widetilde{u} \in (u,h).$$

Remark 4 and Theorem 4 with regard to (2.5) are checked in the same way as similar statements from [10].

This research was supported by the Russian Foundation for Basic Research (project No. 13-01-00256a) and by the program "Leading Scientific Schools" (project 2504.2014.1).

Translated by the author.

## REFERENCES

- 1. F. Aurzada, "On the lower tail probabilities of some random sequences in  $l_p$ ," J. Theoret. Probab., **20**, 843–858 (2007).
- 2. F. Aurzada, "A short note on small deviations of sequences of i.i.d. random variables with exponentially decreasing weights," *Statist. Probab. Letters*, **78**, No. 15, 2300–2307 (2008).
- A. A. Borovkov and P. S. Ruzankin, "On small deviations of series of weighted random variables," J. Theoret. Probab. 21, 628–649 (2008).
- 4. L. V. Rozovsky, "Small deviations of series of weighted i.i.d. non-negative random variables with a positive mass at the origin," *Statist. Probab. Letters*, **79**, 1495–1500 (2009).
- L. V. Rozovsky, "On small deviations of series of weighted positive random variables," J. Math. Sci., 176, No. 2, 224–231 (2011).
- L. V. Rozovsky, "Small deviations of series of independent nonnegative random variables with smooth weights," *Theory Probab. Appl.*, 58, No. 1, 121–137 (2014).

0

- R. Davis and S. Resnick, "Extremes of moving averages of random variables with finite endpoint," Ann. Probab., 19, 312–328 (1991).
- A. A. Borovkov and P. S. Ruzankin, "Small deviations of series of independent positive random variables with weights close to exponential," *Siber. Adv. Math.*, 18, No. 3, 163–175 (2008).
- M. A. Lifshits, "On the lower tail probabilities of some random series," Ann. Probab., 25, 424–442 (1997).
- T. Dunker, M. A. Lifshits, and W. Linde, "Small deviations of sums of independent variables," in: *High Dimensional Probability, Progress Probab.*, 43, Birkhauser, Basel (1998), pp. 59–74.
- L. V. Rozovsky, "On small deviation probabilities for sums of independent positive random variables," J. Mat Sci., 147, No. 4, 6935–6945 (2007).
- 12. L. V. Rozovsky, "Small deviation probabilities of weighted sums of independent positive random variables with a common distribution that deacreases at zero not faster then a power," *Teor. Veroyatn. Primen.*, **60**, No. 1, 178–186 (2015).
- L. V. Rozovsky, "Small deviation probabilities of weighted sums under minimal moment assumptions," *Statist. Probab. Letters*, 86, No. 1, 1–6 (2014).
- 14. N. C. Jain and W. E. Pruitt, "Lower tail probability estimates for subordinators and nondecreasing random walks," Ann. Probab., 15, No. 1, 76–101 (1987).
- L. V. Rozovsky, "Comparison theorems for small deviations of weighted series," Probab. Math. Stat., 32, No. 1, 117–130 (2012).
- L. V. Rozovsky, "Small deviation probabilities of weighted sums with fast decreasing weights," *Probab. Math. Stat.*, 35, No. 1, 161–178 (2015).