Genetic variation and population structuring in two brooding coral species (Siderastrea stellata and Siderastrea radians) from Brazil

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Abstract Siderastrea stellata and S. radians are scleractinian coral species that present a remarkable overlap of diagnostic characteristics and sympatric distribution. Moreover, both are viviparous with similar reproductive strategies and with a gregarious larval behavior. Samples of both species from the Brazilian coast were analyzed using 18 isozymic loci to quantify their genetic variability and populational structure. Results confirmed species identity, high intrapopulational variability and revealed moderate genetic structuring among all samples (S. stellata: F_{ST} = 0.070; S. *radians*: $F_{ST} = 0.092$). Based on genotypic diversity analysis, there was evidence that local recruitment may have a minor role in the populations (mean, $G_o:G_e = 1.00 \pm 0.0003 SD$ for S. stellata and $0.99 \pm 0.0003 SD$ 0.0023 SD for S. radians). Deviations towards heterozygote deficiencies found in both Siderastrea species could be explained by the Wahlund effect, since there was

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V. N. Solferini e-mail: vera.solferini@gmail.com evidence that populations might be composed of colonies of different ages. In S. radians it is also likely that there is some inbreeding occurring in the studied populations. Despite the brooding pattern and the gregarious larval behavior, our data suggest the occurrence of gene flow along the Brazilian coast. This is the first study on population genetics of Brazilian reef corals.

Keywords Brazilian reefs \cdot Gene flow \cdot Genotypic diversity · Population structure · Scleractinian corals · Siderastrea

Introduction

Biological information concerning Brazilian coral fauna is sparse, and somewhat restricted to its reproductive biology (Pires et al. [1999;](#page-11-0) Calderon et al. [2000;](#page-10-0) Neves and Pires [2002](#page-11-0); Neves and da Silveira [2003](#page-11-0)). In Brazil, coral reefs are discontinuously distributed along 3,000 km of the northeastern coast, from the South of Bahia State to Maranhão, and represent the only reefal ecosystem found in the South Atlantic Ocean (Maida and Ferreira [1997](#page-10-0)). So far, approximately 23 species have been recorded along the northeastern and southeastern coasts, six of which are endemic, 16 also found in the West Indies and one in the Indo-Pacific region (Neves et al. [2006\)](#page-11-0). The 'Northeastern Coastal Region' is characterized by the occurrence of reefs and estuaries associated with mangrove vegetation, in which the influence of the warm Brazilian Current and river discharges create optimal temperature and water enrichment for corals (Maida and Ferreira [1997](#page-10-0)). Accord-ing to Leão et al. ([2003\)](#page-10-0), Brazilian coral reefs show peculiar features, being characterized by an exquisite architecture (pinnacles with mushroom-like forms), very

low diversity, remarkable endemism among reef builders, and species that are predominantly massive.

Siderastrea is a small scleractinian coral genus comprising five colonial zooxanthellate species, two of which occur in the Indo-Pacific region, S. savignyana Edwards and Haime 1850 and S. glynni (Budd and Guzmán, [1994](#page-10-0)), while the remainder occur in the Atlantic Ocean, S. radians (Pallas, 1766), S. siderea (Ellis & Solander, 1786), and S. stellata Verrill, 1868 (Laborel [1974](#page-10-0)). Neves ([2004\)](#page-11-0) thoroughly revised the diagnostic characters of the genus, and proposed the sympatric occurrence of Siderastrea radians, and S. stellata on the Brazilian coast. In the literature, S. stellata has been referred to as a synonym of S. siderea and S. radians and it was previously considered to be the only Brazilian siderastreid (Laborel [1974;](#page-10-0) Werner [1996;](#page-11-0) Maida and Ferreira [1997\)](#page-10-0). However, S. radians and S. stellata undergo intratentacular budding, while S. siderea does not. Moreover, S. radians and S. stellata are known to brood their embryos, whereas S. siderea has been reported to shed gametes for external fertilization (Duerden [1904;](#page-10-0) Neves and da Silveira [2003;](#page-11-0) Szmant [1986](#page-11-0)). A recent survey using the nuclear ITS region confirmed the divergence between S. siderea, S. radians and S. stellata (Forsman et al. [2005](#page-10-0)).

Life history and reproductive behavior may be highly variable among sessile marine invertebrates (Levin and Bridges [1995\)](#page-10-0). Dispersal differences are expected to produce distinct patterns of genetic exchange among populations. In contrast to externally developed planulae, brooded planulae are generally released at an advanced stage and are expected to settle immediately (Harrison and Wallace [1990](#page-10-0)), supporting local recruitment. Siderastrea radians and S. stellata undertake internal fertilization and brooding (Duerden [1904](#page-10-0); Szmant [1986](#page-11-0); Soong [1991\)](#page-11-0). In the laboratory, planulae of *S. radians* remained mostly competent for 24–48 h, while S. stellata displayed an extended competency period, being unable to metamorphose before 72 h (Duerden [1904;](#page-10-0) Neves and da Silveira [2003](#page-11-0)). Besides larval development (phylopatric versus teleplanic larvae), physical-environmental factors (e.g. ocean currents, wind regime) have been also considered important in the dispersal of coral species (Jokiel [1990](#page-10-0); Veron [1995](#page-11-0)).

The reproductive mode, e.g. gamete spawners or brooded planulae, also may have an important effect on gene flow and population structure (e.g. Ayre et al. [1997;](#page-10-0) Ayre and Hughes [2000](#page-10-0); Nishikawa and Sakai [2005\)](#page-11-0). Limited dispersal is generally associated with inbreeding, loss of heterozygosity and population subdivision due to genetic drift. We investigated if Brazilian Siderastrea populations fit these predictions, also using hierarchical surveys of genetic variation to quantify local and regional patterns of genetic variability along the northeastern and southeastern coast of Brazil. We also tested whether the genetic

structure of reef communities in Brazil conforms to the "stepping stone" model, in which only adjacent populations exchange genes (Kimura and Weiss [1964\)](#page-10-0). This relationship between gene flow and geographic distance is usually recognized as the mechanism that could explain differentiation of coral species (Hellberg [1994](#page-10-0); Palumbi [2003](#page-11-0)). Samples of S. radians and S. stellata from Brazilian reef systems and coral communities were subject to populational analysis using isozymes, which seem to be a suitable molecular marker for characterizing and understanding population structure and migration in coral reefs (Ridgway [2005\)](#page-11-0).

Material and methods

Study sites and collection

Nine locations were sampled, six distributed among the reefs of the northeastern coast (NE), and three in the coral communities of the southeastern coast (SE), covering about 2,000 km of the Brazilian littoral (Fig. 1). Respectively, eight and seven populations of S. stellata and S. radians were sampled; detailed information concerning locations and dates is given in Table 1. In this study, we use the term 'collection' to refer to each sample or population to avoid confusion, as the term 'sample' could also suggest a single coral individual. Due to the low abundance on the reefs, some collections have a very small number of colonies sampled, especially from Siderastrea stellata. Colonies with diameter of about 10 cm were collected randomly at each sampling site during low tide, at depths of 2–5 m with a hammer and chisel. In order to reduce the mucus production due to handling during field activities, colonies

Fig. 1 Sampling sites in the reefal systems and coral communities from the northeastern (NE) and southeastern (SE) regions of the Brazilian coast. See Table 1 for abbreviations of the collections

were kept alive in foam boxes and left to rest for approximately 24 h. Fragments of skeleton with tissue were scraped off and frozen in liquid nitrogen until electrophoresis. Pieces from the same colonies were bleached, and species identification was performed using the corallites structures. Voucher specimens were deposited in the Museu de Zoologia da Universidade de São Paulo, São Paulo (Cabo Branco/Ponta do Seixas: MZUSP476, MZUSP477; São José da Coroa Grande: MZUSP478, MZUSP479; Porto de Galinhas: MZUSP480, MZUSP481; Maragogi: MZUSP482, MZUSP483; Picãozinho: MZUSP484, MZUSP485; Nova AlmeidaS: MZUSP488, MZUSP489; Praia do Forno: MZUSP486, MZUSP487; Praia da Tartaruga: MZUSP490, MZUSP491).

Electrophoresis

A total of 143 colonies of S. stellata and 177 of S. radians were analyzed. Small tissue pieces $(3-5 \text{ mm}^2)$ were squashed in extraction buffer (0.1 M Tris pH 8.0, 0.5% β mercaptoethanol). The extracts were blotted onto Whatman #3 filter paper wicks and loaded on to horizontal 8.5% starch gels (hydrolyzed potato starch, Sigma, St. Louis, MO, USA). Four buffer systems were used: (I) electrode: 0.25 M Tris and 0.057 M citric acid, pH 8.0, and gel: electrode solution diluted 1:25, 50 mA/4 h (Ward and Warwick [1980](#page-11-0)) (II) electrode: 0.04 M citric acid and N(3 aminopropyl) morpholine, pH 6.1, and gel: electrode solution diluted 1:20, 50 mA/5 h (Clayton and Tetriak [1972\)](#page-10-0); (III) electrode: 0.01 lithium hydroxide, 0.095 boric acid and 0.003 EDTA, pH 8.0, and gel: electrode solution diluted 1:40 50 mA/4 h; (IV) electrode: 0.3 M boric acid, 60 mM NaOH, pH 8.0, and gel: 10 mM Tris, pH 8.5, 180 V/4 h (Shaw and Prasad [1970](#page-11-0), with modifications).

Eleven enzyme systems were surveyed: leucine aminopeptidase (EC 3.4.11.1, LAP), L-leucyl-L-glycylglycine peptidase (EC 3.4.11-, PLGG), L-leucyl-L-alanine peptidase (EC 3.4.13.18, PLA), malic enzyme (EC 1.1.1.40, ME), phosphoglucomutase (EC 5.4.2.2, PGM), malate dehydrogenase (EC 1.1.1.37, MDH), isocitric dehydrogenase (EC 1.1.1.42, IDH), esterase (EC 3.1.1.1, EST), fumarase (EC 4.2.1.2, FUM), glucose-6-phosphate 1-dehydrogenase (EC 1.1.1.49, 6PGD) and hexokinase (EC 2.7.1.1, HK). Staining procedures were adapted from Shaw and Prasad [\(1970](#page-11-0)), and Alfenas et al. ([1991](#page-10-0)).

Data analysis

Alleles were identified by their mobility relative to the most common allele of the Maragogi collection. The BIOSYS-1 program (Swofford and Selander [1981](#page-11-0)) was used to estimate the genetic variability, namely, the percentage of polymorphic loci (P, 95% criterion), the mean number of alleles per locus (A_p) , the gene diversity (H_e) and the observed heterozygosity (H_o) per collection (unbiased estimate, Nei [1978\)](#page-10-0). BIOSYS-1 was also used to calculate identity matrices (Nei [1978\)](#page-10-0).

Departures from Hardy–Weinberg expectations were calculated using the TFPGA program (Miller [1997\)](#page-10-0). An exact test for each locus was performed for all collections using the conventional Monte Carlo method (adapted from Guo and Thompson [1992\)](#page-10-0), performing 10 batches of 1,000 permutations (10,000 permutations). The sequential Bonferroni procedure was applied to correct type 1 errors introduced by multiple tests (Rice [1989\)](#page-11-0).

The F_{IS} coefficient was calculated using Weir and Cockerham's [\(1984](#page-11-0)) f estimate. About 95% confidence intervals were determined for the F_{IS} value by

bootstrapping 10,000 times across loci using the GDA program (Lewis and Zaykin [1999](#page-10-0)). GENETIX v. 4.02 (Belkhir [2001](#page-10-0)) was used to calculate the F_{IS} coefficient per collection and to make random permutations in the matrix individuals versus genotypes in order to obtain the expected distribution of F_{IS} under the null hypothesis of random mating ($F_{IS} = 0$). The probability of a value equal or larger than the estimated value was calculated by $P = (n + 1)$ $(N + 1)$, where n is the number of pseudo-values larger than or equal to the estimate, and N is the number of random permutations (Sokal and Rohlf [1995](#page-11-0)). Sequential Bonferroni correction was also used here.

To verify possible associations among the studied loci, genotypic data were tested for linkage disequilibrium within collections using the GENEPOP program, without collapsing less frequent alleles (Raymond and Rousset [1995\)](#page-11-0). One hundred batches of 5,000 iterations per batch with 10,000 dememorization steps were made. Again, the sequential Bonferroni procedure was applied. Significant associations among loci could denote episodes of inbreeding, asexual reproduction, selection, recent colonization and the Wahlund effect (in which case interloci associations will vary among collections, Crow and Kimura [1970](#page-10-0)).

The relative frequencies of sexual or asexually derived colonies were inferred using the approach described by Stoddart and Taylor [\(1988](#page-11-0)). First, multi-locus genotypes of each colony (N) were obtained and the number of the unique multi-locus genotypes (N_g) was counted. The ratio N_g : N provides the simplest index of the effects of asexual reproduction on genotypic diversity. The ratio of observed multilocus genotypic diversity (G_o) was then obtained and compared to that expected (G_e) under conditions of sexual reproduction with free recombination. Departures of $G_0:G_e$ from unity provide an index of the combined effects of departure from Hardy–Weinberg equilibrium and multilocus linkage equilibrium. We assessed the statistical significance of differences among G_0 and G_e using an unpaired t-test (Stoddart and Taylor [1988](#page-11-0)).

A hierarchical analysis was performed to partition genetic variability within and among major sampling regions (NE and SE) for S. stellata: F_{SR} , variation among sites within regions; F_{RT} , total variation among regions; F_{ST} , total variation among all sampling sites. The coefficients were calculated using the Weir and Cockerham ([1984\)](#page-11-0) θ parameter. Ninety-five percent confidence intervals were determined for the total F_{ST} value by bootstrapping 10,000 times across the loci using the GDA program (Lewis and Zaykin [1999\)](#page-10-0). Gene flow was inferred for each hierarchical level using Wright's [\(1969\)](#page-11-0) Island model $(N_e m = [(1/\sqrt{N_e m})^2]$ θ) – 1]/4), where N_e is the effective population size and *m* is the proportion of migrants per generation.

To test if there is a correlation between geographic distance and F_{ST} , a Mantel test was performed comparing each pair of S. stellata and S. radians collections. The matrices were compared by the Mantel test using the TFPGA program (Miller [1997](#page-10-0)). The GENETIX v.4.02 program (Belkhir [2001\)](#page-10-0) uses random permutations on the matrix of individuals versus genotypes to find the expected distribution of F_{ST} under the null hypothesis $(F_{ST} = 0)$. The probabilities of each F_{ST} estimated were calculated as presented above for F_{IS} .

Results

Genetic variability within populations

For the 11 enzymes surveyed, 18 loci were scored. None of them were diagnostic, but most loci presented exclusive alleles for each species (Appendix).

A total of 14 and 13 loci were polymorphic (95% criterion) for S. stellata and S. radians, respectively. The most polymorphic loci were Idh for S. stellata, with seven alleles, and Fum and Idh for S. radians, with eight alleles. Percentage of polymorphic loci and gene diversity indicate some genetic variability in the collections (S. stellata $P = 38\%, H_e = 0.120; S.$ radians $P = 45\%, H_e = 0.168$, as seen in Table 2. There is no relationship between the collection size and the percentage of polymorphic loci (S. stellata $R^2 = 0.09$, $P = 0.47$; S. radians $R^2 = 0.04$, $P = 0.65$). The H_o values were very low and F_{IS} per location were very high (Table 2), the averages for S. stellata and S. *radians* were, respectively, $F_{IS} = 0.330$ (95% CI = 0.170–0.592) and 0.580 (95% CI = 0.414–0.720).

Three S. stellata collections and six of S. radians showed departures from Hardy–Weinberg expectations in several loci; all deviations were due to deficit of heterozygotes. Using the Monte Carlo method, 4 out of 65 tests were significant for S. stellata and 18 out of 70 were significant for S. radians (Table 3).

Siderastrea stellata presented only four significant loci associations ($P < 0.05$) in 169 tests. When the sequential Bonferroni correction was applied, none of these were significant. The test of linkage disequilibrium revealed that only 15 in 234 tests in S. *radians* are significant ($P < 0.05$). After applying the Bonferroni procedure, one association (Pgm-1/ Pgm-2) was significant in two collections, SJ and NA.

In both Siderastrea species, each collection contained high genotypic diversity: the majority of colonies displayed unique multi-locus genotypes (N_g :N ranged from 0.87 to 1; Table 4). The collections showed >93% of the multi-locus genotypic diversity (G_o) expected for random mating and free recombination (G_e) . For both species, G_o was never significantly different from G_e (mean $G_o:G_e = 1.00 \pm$ 0.0003 SD for S. stellata and 0.99 ± 0.0023 SD for S. *radians*; unpaired Student's t-tests, $P > 0.1$.

Table 2 Average genetic variability for Siderastrea species

See Table 1 for abbreviation of the collections

N, mean collection size per locus; P, percentage of polymorphic loci (95% criterion); A_p, mean number of alleles per locus. H_e, gene diversity (unbiased estimate, Nei [1978](#page-10-0)); H_0 , observed heterozygosity; F_{1S} , fixation index (significant values in bold)

Table 3 F_{IS} values of loci with deviations from the expected values in HW equilibrium

Species	Collection	Mdh	$Pgm-1$	$Pgm-2$	$Pgm-3$	Fum	$Hk-1$	Idh	Me
S. stellata	Ma		-0.142	0.726	-	-0.050	-	-0.066	
	PT	1.00	1.00	0.00		-0.055	1.00	0.745	-0.100
	PF	$\qquad \qquad$	-	0.00	$\overline{}$	0.662	0.732	0.383	-0.154
S. radians	PN	0.663	0.488	0.868	0.869	-0.021	0.650	0.130	0.492
	$P_{\rm C}$	0.650	0.622	$\overline{}$	$\qquad \qquad -$	0.814	-	0.00	0.036
	CB	1.00	0.217	1.00		0.555	0.661	-0.055	0.375
	SJ	1.00	0.814	1.00	$\overline{}$	$\qquad \qquad$	0.784	0.00	0.257
	Ma	$\overline{}$	0.907	0.919	1.00	0.231	1.00	0.00	0.747
	NA	0.900	0.882	0.767	0.728	0.365	-0.052	1.00	0.621

In bold, loci with significant deviation after sequential Bonferroni correction ($P < 0.001$); – monomorphic loci. See Table 1 for abbreviation of the collections

Population structure and genetic identity

Siderastrea stellata

There was significant genetic differentiation ($F_{ST} = 0.070$, 95% $CI = 0.021 - 0.100$ among all collections and the number of migrants per generation ($N_{\rm e}$ m) was 3.3. The $F_{\rm SR}$ values between NE and SE collections were moderate: $F_{SR} = 0.048$ and 0.091 (95% CI = 0.002–0.092; 0.009– 0.158), respectively. Estimated $N_{\rm e}m$ was 4.9 for NE collections, and 2.5 for SE collections. There was no inter-

region differentiation ($F_{RT} = -0.006$, 95% CI = -0.040-0.045; $N_{\rm e}$ m = 41.9), suggesting no hierarchical structuring.

Geographic distances ranged from 7.56 km to 1748 km among the collections. The F_{ST} parameter ranged from -0.006 (Pc – PF) to 0.404 (NA – PT) in the pairwise estimates. The correlation between geographic distance and genetic structuring was not significant ($r = 0.10$, $P = 0.28$). The F_{ST} estimates calculated for each pair of populations are presented in Table 5, and eight out 28 tests were significant when sequential Bonferroni correction was applied.

Species	Collection	N_g	N	N_{σ} : N	$G_{\rm o}$: $G_{\rm e}$	Species	Collection	N_g	N	N_{σ} :N	$G_{o}:G_{e}$
S. stellata	P_{c}	12	12	1.00	1.000	S. radians	PN	34	34	1.00	1.000
	CB	3	3	1.00	1.000		Pc	16	16	1.00	1.000
NE	PG	13	13	1.00	1.000	NE	CB	24	26	0.92	1.000
	SJ	12	12	1.00	1.000		PG	16	16	1.00	0.937
	Ma	18	18	1.00	1.000		SJ	18	18	1.00	1.000
SE.	NA	6	6	1.00	1.000		Ma	19	19	1.00	1.000
	PT	32	35	0.91	1.000	SE	NA	42	48	0.87	1.000
	PF	39	44	0.88	1.001						

Table 4 Estimates of contribution of asexual reproduction in each reefal system of Siderastrea species

See Table 1 for abbreviation of the collections

 N_g , number of unique multi-locus genotypes; N, number of colonies sampled; G_o , observed multi-locus genotypic diversity; G_e , expected multilocus genotypic diversity

The genetic identities (I, Nei [1978\)](#page-10-0) ranged from 0.912 $(Ma - NA)$ to 1.00 $(CB - PT)$ in *S. stellata*, with an average of 0.976.

Siderastrea radians

Collections presented significant structuring ($\theta = 0.092$, 95% CI = $0.024 - 0.15$), with 2.5 migrants per generation (N_em) . The F_{SR} obtained only among NE collections was moderate, $\theta = 0.078$ (95% CI = 0.017–0.143) and the estimated $N_{e}m$ were 3. Similarly to S. stellata, S. radians did not present a correlation between geographic distance and genetic structuring $(r = 0.22, P = 0.17)$ in a geographic distance range of 186.10 km to 1,672 km. The F_{ST} pairwise

estimates calculated varied from -0.042 (NA $-$ PT) to 0.197 (CB $-$ Ma), and only four out of 21 were significant (Table 6).

The mean genetic identity of S. radians collections was also high $(I = 0.974)$, and ranged from 0.935 (Ma – Pc) to 1.00 (PN $-$ PG). The mean genetic identity among Siderastrea species was also very high $(I = 0.92)$.

Discussion

Intrapopulational genetic variability

In our study, the percentage of polymorphic loci in both species varied significantly among collections. Some

collections, as the one from Porto de Galinhas of S. stellata, presented a high value (66%), as well as the S. radians samples from Cabo Branco/Ponta do Seixas and Nova Almeida (61.1% and 55.5%, respectively). However, most of the collections presented polymorphism values below 40%, which is unlikely to be an effect of the sampling size. Nevertheless, both species also presented a strikingly high diversity of genotypes (Table 4). Some of them, as Porto de Galinhas of S. stellata, presented a high value (66%), as well as S. radians collection from Cabo Branco/Ponta do Seixas and Nova Almeida (61.1% and 55.5%, respectively). Most of the polymorphism values were below 40%, which might not be an effect from the sampling size. Nevertheless, both species also presented strikingly high diversity of genotypes (Table 4). A model proposed by Bengtsson ([2003\)](#page-10-0) showed that even a small frequency of sexual reproduction per generation would be sufficient to make a population highly genotypically variable. In this model, a population started by a number of sexuallyderived propagules may retain its initial genotypic variation for a very long period of time. The population might keep a record of its earlier genetic history, a phenomenon that the author designated as 'memory-effect'. Thus, besides the population size, the sexual recruitment and the possible effect of environmental heterogeneity, the longevity of the colonies may play an important role in maintenance of genetic variability in Siderastrea collections.

Hardy–Weinberg deviations are commonly found in corals (e.g. Brazeau and Harvell [1994;](#page-10-0) Márquez et al. [2002;](#page-10-0) Ng and Morton [2003](#page-11-0); Miller and Ayre [2004\)](#page-10-0). In the present study, heterozygote deficiency was observed in both species, with different intensities in many loci and collections. This could reflect complex interactions among a range of factors, including the occurrence of null alleles (Gardner [1992](#page-10-0)), selective mortality (Zouros and Foltz [1984\)](#page-11-0), the Wahlund effect (Ayre and Dufty [1994](#page-10-0)), differences in the time of spawning ('genotype-dependent spawning', Zouros and Foltz [1984](#page-11-0)) and inbreeding (Smith and Potts [1987\)](#page-11-0). There is the possibility of the occurrence of more than one species or sub-species in our collections. If so, preferential intraspecific crosses would generate linkage disequilibrium, but significant or consistent associations among genotypes were very scarce. The occurrence of inbreeding seems to be very likely in our collections for S. radians. Indeed, the collections of S. radians presented the higher number of loci with deviation and the highest F_{IS} values. The shorter planktonic phase of this species associated with the phylopatric behavior of the larvae can explain the high Hardy–Weinberg deviation observed. However, if inbreeding had a major role in our collections, it would affect all loci simultaneously to a similar degree (Lewontin and Krakauer [1973\)](#page-10-0). Thus,

associated with inbreeding events, the Wahlund effect caused by temporal variation of reproductive cohorts would be also a possible explanation for the observed heterozygote deficiency. In this case, the sampled populations of both species would be composed of colonies from different ages (which could also explain the observed genetic variability), generated by the settlement of different cohorts whose parents did not interbreed before. One possible reason for the formation of these populations would be the temporally different sources of larvae, which seems to be a common feature in scleractinians (Jokiel [1990](#page-10-0); Veron [1995](#page-11-0)), due to the variation of the hydrodynamic conditions.

Local recruitment and structuring

Among scleractinians, recent electrophoretic investigations have revealed small values of F_{ST} and genetic distance, indicating that the genetic structure may be independent of geographic distance, breeding system and mode of larval development (Márquez et al. [2002](#page-10-0); Ng and Morton [2003](#page-11-0)). A moderate to low structure has also been reported among broadcast-spawning species (e.g. Pocillopora verrucosa, Ridgway, Hoegh-Guldberg & Ayre, 2001), suggesting that local recruitment may have a restricted role in the maintenance of genetic structure of coral populations. Genetic evidence from nine coral species showed that population variability of brooding and broadcasting species could be maintained by a balance between localized settlement of larvae and high gene flow along the Great Barrier Reef caused by long-dispersal larvae (Ayre and Hughes [2000](#page-10-0)). In that study, the authors found significant genetic variation among local populations within geographic regions for all species, but not among the geographic groups on a larger scale (1,800 km) for two broadcasting and three brooding species. Nishikawa and Sakai ([2005\)](#page-11-0) also found evidence of high gene flow along a macrogeographic scale despite of a substantial proportion of local recruitment in the brooding species Goniastrea aspera.

Reproductive behavior, phylopatric behavior and early development might explain the moderate structure documented for S. stellata $(F_{ST} = 0.07)$ and S. radians $(F_{ST} = 0.09)$. However, our data do not support the prediction that these brooding species have a very restricted dispersal, regardless of the analyzed scale. Indeed, in all evaluated hierarchical levels, the inferred number of migrants per generation is substantially high. Siderastrea stellata is likely to be lecitotrophic, depending on yolk reserves and photosynthetic products during the early larval stage (Neves and da Silveira [2003](#page-11-0)). The long competency period of S. stellata (up to 15 days) would be expected to enhance homogenization, reducing the population differentiation. The competency period of S. radians larvae is shorter (up to 48 h, Duerden [1904](#page-10-0)), which could explain the higher structuring; however the values of F_{ST} for both species are not statistically different. Virtually nothing is known about the dispersal capabilities of Siderastrea radians and S. stellata under field conditions. Thus, the competency time observed in laboratory conditions may be an underestimation and in their natural environmental, with suitable physical and biotic conditions, larvae could potentially live on plankton for a longer time.

There was no hierarchical structuring in S. stellata, but we found significant structuring within regions, especially in region SE. The estimated values of F_{SR} and F_{ST} in S. stellata are also not significantly different, suggesting that there is homogeneous gene flow among collections. Additional evidence of extensive gene flow in both species is provided by the shared rare alleles among geographically distant collections in most of the loci (Appendix). The lack of correlation between geographical distance and genetic differentiation (F_{ST}) suggested that the observed genetic differentiation among the populations distributed along the coast cannot be explained solely by geographical isolation.

Considering our results, there are no apparent physical or ecological barriers between the north and southeastern coastal regions. We also did not find any evidence of a ''stepping stone'' pattern in the collections of either species. Siderastrea stellata populations are most likely distributed discontinuously along the north/southeastern coast (Laborel [1974;](#page-10-0) Neves pers. obs.). There is no such record for S. radians, but it is expected that this species generally has a similar distribution, since it usually occurs in the same natural conditions as those in which S. stellata populations are found. Thus, despite the discontinuous distribution structure of reef communities in Brazil (Maida and Ferreira [1997\)](#page-10-0), there is strong evidence that Siderastrea populations exchange genes along the coast. Some long distance gene flow might also occur via a series of short distance recruitment along generations. If there is local recruitment in the Brazilian siderastreids, the gene flow seems to somewhat override its effect, preventing the development of high geographic structuring at the evaluated scale.

Interspecific genetic identities in Siderastrea

Recognition of interspecific limits among scleractinians can be a difficult task (Laborel [1974;](#page-10-0) Lang [1984](#page-10-0); Veron

[1995](#page-11-0); Forsman et al. [2005](#page-10-0)). In this context, the presence of exclusive alleles may provide a relevant support to taxonomical analyses, particularly when the species are closely related and with similar morphological traits. Although diagnostic loci were not found, some alleles were restricted or particularly predominant in one of the studied species (see loci Pgm and Mdh in Appendix). This is evidence of genetic divergence between the two species and confirms the maintenance and acceptance of their specific status.

The mean interspecific genetic identity between the Brazilian Siderastrea species was high, and not very different from intraspecific values. Our data showed that the species are distinct although closely related. Indeed, a similar pattern has been reported in other coral species (Weil and Knowlton [1994;](#page-11-0) McFadden [1999\)](#page-10-0). Contrasting with the southeastern coral communities, where S. stellata was the only species identified, S. radians and S. stellata have a sympatric distribution along the northern shallowwater reefal system and are also very similar morpho-logically (Neves [2004](#page-11-0)). Thus, these species could be considered 'pseudo-sibling species', as they are genetically distinct but only recognizable when the appropriate morphological attributes are considered (Knowlton [1994](#page-10-0)).

This study is the first evaluation of the genetic variability of two reef-building species of scleractinian corals from Brazil. Moreover, the results contribute to the maintenance of the specific status of the two congeners, supporting Neves ([2004\)](#page-11-0) morphological analysis and refuting Werner [\(1996](#page-11-0)), who suggested S. stellata as synonym of S. radians. On the other hand, finer-scale investigations would be necessary to confirm whether the spatial and temporal patterns found are consistent. Further studies might determine if local recruitment is the main influence on the structure of the Sideratrea populations located at the extremes of species distribution.

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Appendix

Allelic frequencies from loci of Siderastrea stellata and S. radians $(N = total$ collection size)

Appendix continued

Appendix continued

Siderastrea stellata									Siderastrea radians						
Collection N	P_{C} 12	CB 3	PG 13	SJ 12	Ma 18	NA 6	PT 35	PF 44	PN 34	Pc 16	CВ 26	PG 16	SJ 18	Ma 19	NA 48
Me															
								0.074	0.037	0.031	0.048	0.067			0.125
2	0.182		0.273	0.286		1.00	0.109	0.278	0.444	0.250	0.357	0.567	0.281	0.533	0.475
3	0.591	1.00	0.455	0.714	1.00		0.891	0.519	0.482	0.656	0.595	0.367	0.719	0.467	0.350
$\overline{4}$	0.227		0.273					0.130	0.037	0.063					0.050

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