# **Assessment of seasonal variations of surface water quality characteristics for Porsuk Stream**

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**Abstract** In this study, the factor analysis technique is applied to surface water quality data sets obtained from Porsuk stream in Turkey, generated during 10 years (1995–2005) monitoring of 29 parameters at one site (Esenkara) for all four seasons. The varifactors obtained from factor analysis indicate that the parameters responsible for water quality variations are mainly related to mineral and inorganic nutrients, organic pollution, microbiological pollution in winter and spring; mineral and nutrients in summer; microbiological and nutrient pollution in fall. This study presents the necessity and usefulness of multivariate statistical assessment of large and complex databases in order to get better information about the quality of surface water.

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#### **Introduction**

A river is a system comprising both the main course and the tributaries, carrying the one-way flow of a significant load of matter in dissolved and particulate phases from both natural and anthropogenic sources. The quality of a river at any point reflects several major influences, including the lithology of the basin, atmospheric inputs, climatic conditions and anthropogenic inputs. On the other hand, rivers plays a major role in assimilation or transporting municipal and industrial wastewater and runoff from agricultural land. Municipal and industrial wastewater discharge constitutes a constant polluting sources, whereas surface runoff is a seasonal phenomenon, largely affected by climate within the basin (Shrestha and Kazama [2007\)](#page-14-0).

The quality of water is identified in terms of its physical, chemical and biological parameters. Polluted surface waters cannot achieve a balanced ecosystem. A balanced ecosystem is one in which living things and the environment interact beneficially with one another. Water quality obviously plays a critical role in this relationship (Ntengw[e](#page-13-0) [2006\)](#page-13-0), as it is key to the maintenance of a wellbalanced environment.

A variety of methods are being used to display the information which is concealed in the quality variables observed in a water quality monitoring network. A large portion of these approaches are statistical methods. When the number of variables is greater than two, employment of multivariate analysis techniques gives simpler and more easily interpretable results for the evaluation of observed quality data (Mazlum et al[.](#page-13-0) [1999\)](#page-13-0). The multivariate statistical techniques are the appropriate tool for a meaningful data reduction and interpretation of multi-constituent chemical and physical measurements (Massart et al[.](#page-13-0) [1988\)](#page-13-0). The multivariate statistical techniques are such as cluster analysis (CA), principal component analysis (PCA), factor analysis (FA), discriminant analysis (DA) and multidimensional scaling (MDS) have been successfully used as unbiased methods in analysis of water quality data for drawing meaningful conclusions (Vega et al[.](#page-14-0) [1998](#page-14-0); Helena et al[.](#page-13-0) [2000](#page-13-0); Wunderlin et al[.](#page-14-0) [2001](#page-14-0); Voncina et al[.](#page-14-0) [2002;](#page-14-0) Simeonov et al[.](#page-14-0) [2003,](#page-14-0) [2004;](#page-14-0) Simeonova et al[.](#page-14-0) [2003;](#page-14-0) Ouyan[g](#page-14-0) [2005;](#page-14-0) Prau[s](#page-13-0) [2005](#page-13-0); Panda et al[.](#page-13-0) [2006;](#page-13-0) Ouyang et al[.](#page-14-0) [2006;](#page-14-0) Kuppusamy and Giridha[r](#page-13-0) [2006;](#page-13-0) Shrestha and Kazam[a](#page-14-0) [2007\)](#page-14-0).

The purpose of this work is to obtain seasonal variations of monitored physico-chemical parameters in the Porsuk stream (Turkey), which was generated under a 10-years (1995–2005) monitoring program, and to obtain the more detailed information about water quality parameters for all four seasons by using the factor analysis method.

#### **Study area**

The Porsuk Stream (length of 460 km) is a tributary of the Sakarya River (length 824 km). It arises in the Murat Mountain to the south of the city of Kütahya, situated in Western Turkey. After taking up various creeks, it flows past Kütahya and enters the reservoir named after itself. Leaving the reservoir, the stream passes through the city of Eskisehir and flowing through a wide valley, again named after itself, it joins the Sakarya River. During its course, the Porsuk Stream travels a distance of 460 km and drains an area of 11 325 km<sup>2</sup> (SH[W](#page-14-0)) [1995;](#page-14-0) Albe[k](#page-13-0) [2003\)](#page-13-0).

The stream is very important because it provides the two cities mentioned with drinkable water. The agricultural areas in the watershed, especially along the valley after Eskisehir, are irrigated with water supplied from the stream (Albe[k](#page-13-0) [2003\)](#page-13-0).

Porsuk stream has been heavily polluted by domestic and industrial activities such as Nitrogen Fertilizer Factory, Sugar-beet Factory, and Magnesite Factory in Kütahya, Seyitomer Thermic Power Plant, which have no, or insufficient treatment plants. Figure [1](#page-2-0) shows Porsuk stream and the locations of the main pollution sources.

#### **Factor analysis**

Factor Analysis is a multivariate statistical technique, which attempts to extract a lower dimensional linear structure from the data set. These factors can be interpreted in terms of new variables. Factor analysis aims to explain observed relation between numerous variables in terms of simpler relations (Catte[l](#page-13-0) [1965\)](#page-13-0). In FA, the basic concept is expressed in the following formula:

$$
z_j = a_{j1} f_1 + a_{j2} f_2 + \dots + a_{jm} f_m + e_{ij}; \ j = 1, 2, \dots, p
$$
\n(1)

where, *z* is the measured value, *f* the factor score, *a* the factor loading, *e* the residual term accounting for errors or other sources of variation, *i* the sample number, *j* the variable number, and *m* is the total number of factors (Anazawa and Ohmori [2001\)](#page-13-0).

#### **Results and discussion**

In this study, water quality data is obtained from the State Hydraulic Works (SHW) at Esenkara monitoring station. Esenkara monitoring station is one of the monitoring stations along the Porsuk Stream situated after the city of Kutahya. The data was observed at monthly intervals between the years 1995 and 2005. Data collection and analysis based on standard techniques have

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**Fig. 1** Porsuk stream and locations of the main pollution sources (Source: Muhammetoglu et al. [2005\)](#page-13-0)

been carried out by SHW within their long-term monitoring program.

In our study, we selected 29 parameters for the estimation of surface water quality characteristics. The selected parameters are: Flow, water temperature, pH, electrical conductivity, total dissolved solids, suspended solids, total alkalinity, chloride, ammonical nitrogen, nitrite nitrogen, nitrate nitrogen, dissolved oxygen, organic matter, biochemical oxygen demand, chemical oxygen demand, total hardness, orto-phosphorus, sulphate, iron, manganese, sodium, potassium, calcium, magnesium, total coliforms, fecal streptococ, Escherichia coli, boron and fluoride. Because the measurement unit of the selected 29 variables are different, first, they are standardized. Next, Pearson correlation coefficient matrix are used for the determination of the correlation between the variables. Then, these correlation coefficients are examined to decide if factor analysis can be applied to the variables. As a result, factor analysis method was found to be applicable problem. Descriptive statistics results of water quality characteristics according to seasons are given in Table [1.](#page-3-0)

## Seasonal correlation of water quality parameters

Data in Table [2](#page-4-0) provide the seasonal correlation matrix of the water quality parameters obtained from the PCA. In general, the river water temperature had relatively weak to fair correlations, i.e., most of the correlation coefficients are less than 0.75 (absolute value) with other parameters for the entire four seasons. In spring, correlation coefficient between temperature and other parameters were less than or equal to 0.5. Such correlations had not changed in summer. These correlations had changed in fall with a positive increase in correlations with NH<sub>3</sub>-N (0.593), N0<sub>2</sub>-N (0.632) and o- $PO_4^{3-}$  (0.502) as well as a negative increase

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in correlations with  $DO(-0.751)$ . The correlation coefficients between temperature and parameters fell below the absolute value of 0.72 in winter, indicating relatively fair correlations. It should be noted that few efforts have been devoted to investigating the correlations among the variables used in this study. Ouyang et al[.](#page-14-0) [\(2006\)](#page-14-0) investigated the seasonal physical and chemical characteristic changes of 22 monitoring stations located in the main stem of the LSJR during March 1998–March 2001. These authors found that the water temperature had very weak correlations with other water quality parameters such as turbidity, TKN, TOC, DOC, EC, alkalinity, salinity, and  $BOD<sub>5</sub>$ , which were, in general, similar to our findings.

Very strong correlations between EC and other parameters (TDS, M-Al, PV, TH and Mg) were also found in winter and spring, but such correlations were not found in summer and fall. That is, the correlation coefficients between EC and TDS, EC and M-Al, EC and PV, EC and TH and EC and Mg 2 <sup>+</sup> were respectively 0.904, 0.908, 0.827, 0.951 and 0.924 in winter, 0.960, 0.946, 0.657, 0.971 and 0.909 in spring, 0.566, 0.668, −0.017, 0.542 and 0.259 in summer, and 0.491, 0.457, −0.268, 0.435 and 0.232 in fall.

Very strong correlations between DO and other parameters were not found in four seasons.

Results from the PCA showed that large seasonal variations in correlation coefficients between BOD<sub>5</sub> and other parameters were not found. The correlation coefficients between BOD 5 and other parameters were less than or equal to 0.587.

Very strong correlations between TDS and other parameters (i.e., M-Al, TH and Mg) were found in spring, but the correlations were moderately reduced in summer and profoundly in fall, and finally recovered in winter. That is, the correlation coefficients between TDS and M-Al, TDS and TH and TDS and Mg were respectively 0.891, 0.949 and 0.908 in spring, 0.647, 0.673 and 0.299 in summer, 0.373, 0.431 and 0.017 in fall and 0.841, 0.917 and 0.903 in winter. These data imply that TDS was not always highly correlated with M-Al, TH and Mg in Porsuk.

The M-Al had strong correlations with TH (0.972) and Mg (0.874) in winter, TH (0.978) and Mg  $(0.857)$  in spring and TH  $(0.831)$  and Mg (0.324) in summer. Such correlations became very poor in fall (TH (−0.106) and Mg (0.010)).

Very strong correlations between SS and other parameters (i.e., Fe, *T-coli* and *E-coli*) were found in fall, but the correlations were moderately reduced spring and profoundly in summer, and finally recovered in winter. That is, the correlation coefficients between SS and Fe, SS and *T-coli* and SS, and SS and E-coli were respectively 0.820, 0.878 and 0.876 in fall, 0.676, 0.419 and 0.489 in spring, 0.578, 0.028 and 0.102 in summer and 0.878, 0.115 and 0.163 in winter. These data imply that SS was not always highly correlated with Fe, *T-coli* and *E-coli* in Porsuk.

The Cl<sup>−</sup> had strong correlations with Mg (0.827) in winter and Mg (0.754) in spring, such correlations became very poor in summer and fall. The NH<sub>3</sub>-N had strong correlations with  $NO<sub>2</sub>-N$  $(0.755)$  and o-PO $^{3-}_{4}$   $(0.882)$  in spring, such correlations became very poor in winter, summer and fall. The  $NO<sub>2</sub>-N$  had strong correlations with  $o\text{-}PO_4^3(0.851)$  in spring and  $o\text{-}PO_4^3(0.708)$  in fall, such correlation very poor in winter and summer. The  $NO<sub>3</sub>$ -N and COD had not correlations with other parameters in four seasons. The PV had negative correlations with TH  $(-0.828)$ , Mg (−0.805), EC (−0.827) and TDS (−0.819) in winter, TH (−0.725), Mg (−0.614), EC (−0.657) and TDS (−0.585) in spring but such correlations were not found in summer and fall. The TH had strong correlations with Mg (0.944) in winter Mg (0.904) in spring, such correlations were not found in summer and fall. The  $SO_4^{3-}$  had good correlations with Na (0.719) in winter but such correlations were not in spring, summer and fall. The F.strp had correlation E-coli (0.778) in winter, E-coli (0.945) in spring and E-coli (0.944) in fall, such correlation was not found in summer. The *T-coli* had good correlations F.strp (0.900) and E-coli (0.952) in spring and F.strp (0.940) and *E-coli* (0.998) in fall and such correlations were not found in winter and summer.

Temporal variations of water quality parameters

Eigenvalues are normally used to determine the number of principal components that can be retained for further study. First four principal components have eigenvalues greater than or close to unity and explain 80.859%, 81.543%, 81.839% and 82.433% of the total variances of information contained in the original dataset for winter, spring, summer and fall, respectively (Table [3\)](#page-9-0).

The component loadings are the linear combinations for each principal component, and express the correlation between the original variables and the newly formed components. The component loadings can be used to determine the relative importance of a variable as compared to other variables in a principal component and do not reflect the importance of the component itself.

Component loadings of principal components for each season are presented in Fig. [2.](#page-10-0)

In winter, Factor 1(F1) explained 35.619% of the total variance and was strong positive loading EC, TDS, M-Al, Cl<sup>−</sup>, TH, Mg; moderate positive loading  $SO_4^{3-}$  and strong negative loading on PV; and moderate negative loading SS, Fe and Mn. This factor seems to measure the preponderance of mineral and inorganic nutrient-related water quality parameters over the organic-related water quality parameters (Table [4\)](#page-11-0).

Factor-2 (F2) explained 16.751% of the total variance and was positively and largely contributed to by T, COD,  $NO_2-N$ , and  $o\text{-}PO_4^{3-}$  and was negatively and largely due to pH and DO. This organic factor can be interpreted as representing influences from point sources, such as of discharges from wastewater treatment plants and industrial effluents.

Factor-3 (F3) explained 7.643% of the total variance and is related to the three bacteriological contamination parameters (loading 0.859, 0.853 and 0.794 for E.coli, F.strp *T. coli* respectively). This factor assigned as the "microbiological factor". F3 may involve an urban origin, where waste disposal from populated areas increases fecal contents in the affected waters.

Factor-4 (F4) is named organic factors, account for 6.932% of the total variance and was strong positive loading for BOD5. This factor represents pollution from domestic waste. Factor-5 (F5) explained 5.672% of the total variance and was positively and largely contributed to by B, and F<sup>−</sup>. Factor-6 (F6) explained 4.615% of the total variance and was positively due to Q and was negatively due to  $NO<sub>3</sub>-N$ . Factor-7 (F7) explained



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3.628% of the total variance and was positively contributed to by  $NH_3-N$  and  $K^+$  and was negatively due to  $Ca^{++}$ .

In spring F1 explained 34.087% of the total variance and was strong positively loaded by mineral-related parameters (i.e TH, EC, M-Al,  $Mg^{2+}$ , Na<sup>+</sup>, TDS and Cl<sup>-</sup>); moderate positive loadings, B,  $Ca^{2+}$  and strongly negative loadings on PV; moderately negative loadings on Fe. This component seems to measure the preponderance of mineral and inorganic nutrient-related water quality parametres.

The second factor F2 (which explained 14.557% of the total variance) was positively and largely due to parameters (i.e,  $o\text{-}PO_4^{3-}$ , NH<sub>3</sub>-N and  $NO<sub>2</sub>-N$ ). This "nutrient" factor represents influences from nonpoint sources such as agricultural runoff and atmospheric deposition.

The third factor, F3 explained 11.471% of the total variance and is related to the three bacteriological contamination parameters (loading 0.975, 0.939 and 0.935 for E.coli, T. coli, F.strp., respectively). This factor assigned as the "microbiological factor". Coliform bacteria are excreted in large numbers in the faeces of humans and other warm-blooded animals. Consequently, water contaminated by faecal matter is identified as being potentially dangerous due to the indicator organisms co-existing with E. coli, which causes cholera (Hammer [1986](#page-13-0)). Some examples of diseases caused by drinking or swimming in faecal contaminated water are diarrhea, cholera, dysentery and skin, eye, ear, nose and throat infections (WH[O](#page-14-0) [1993\)](#page-14-0). The fourth factor F4 explained 7.763% of the total variance and was positively due to Q and SS. This factor explains the erosion from upland areas during rainfall events. The fifth factor F5 explained 5.248% of the total variance and was positively contributed to by COD and Mn and was negatively due to F. COD is a parameter in the amount of organic material observable in the water. F6 explained 4.813% of the total variance and was positively due to T and was negatively contributed to by DO and BOD <sup>5</sup>. The inverse relationship between temperature and dissolved oxygen is a natural process because warmer water becomes saturated more easily with oxygen and it can hold less dissolved oxygen. F7 explained 3.604% of the total variance and was positively <span id="page-10-0"></span>contributed to by  $pH$  and NO<sub>3</sub>-N and was negatively due to  $K^+$ .

In summer, F1, which accounted for 16.412% of the total variance, was positively influenced by M-Al, TH, TDS, EC parameters; moderate positive loadings  $NO<sub>2</sub>$ -N. This factor assigned as the "mineral and inorganic nutrient factor". In summer, while F2, which accounted for 10.477% of the total variance, was positively influenced by  $Mg^{2+}$  parameter and was negatively due to  $Ca^{2+}$ .

F3 explained 9.727% of the total variance and was positively contributed to by  $NO<sub>3</sub>$ -N and was negatively due to T and  $NO<sub>2</sub>-N$ . This factor assigned as the "nutrient" factor. As long as nitrification is occurring in the water, an increase in  $NO<sub>3</sub>-N$  and a reduction in  $NO<sub>2</sub>-N$  is a normal event. Accordingly in the nitrification process,  $NO<sub>2</sub>-N$  is an unstable intermediary which is inclined to transform to  $NO<sub>3</sub>-N$  (Filik Iscen et al[.](#page-13-0) [2008\)](#page-13-0).

F4 explained 8.335% of the total variance and was positively contributed to by  $SO_4^{3-}$  and was negatively due to DO. Sulphate is considered to be a permanent solute of water that reacts with organic matter, producing sulphur  $(S)$ , H<sub>2</sub>O and  $CO<sub>2</sub>$ . The majority of sulphates are soluble in water. The quality of lake water is affected by sulphates. In most natural surface waters, sulphate levels in Turkey range from 200 to over 400 mg/l (Turkish Standard[s](#page-14-0) [2004](#page-14-0)).

F5 explained 7.115% of the total variance and was positively due to B and F<sup>−</sup>. F6 explained



**Fig. 2** Screen plot of the eigenvalues of principal components for all four seasons

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6.464% of the total variance and was positively influenced by Na and was negatively influenced by COD and *T-coli*. F7 explained 5.610% of the total variance and was positively due to  $K^+$  and was negatively due to E-coli. F8 explained 5.301% of the total variance and was positively and largely contributed to F. strp. F9 explained 4.502% of the total variance and was positively influenced by Mn and was negatively influenced by Cl<sup>−</sup>. F10 explained 4.304% of the total variance and was positively and largely contributed to by BOD<sub>5</sub>. This factor represents pollution from domestic waste, so F10 represents the "organic pollution" factor. F11 explained 3.591% of the total variance was positively due to Q and PV.

The F1 in fall, which explained 18.464% of the total variance, was positively and largely influenced by parameters (i.e, *T-coli*, E-coli, F.strp, SS and Fe). The highest burden in this factor was revealed by T-coli, E-coli and F. strp. It revealed that intensive discharge had occurred into the river from sewage.

F2, which explained 17.658% of the total variance, was positively participated by parameters  $(i.e, T, NO<sub>2</sub>-N NH<sub>3</sub>-N, and o-PO<sub>4</sub><sup>3−</sup>)$  and was negatively due to DO. This factor can be attributed to a seasonal changes.

F3 (which explained 13.507% of the total variance) was positively affected due to  $NO<sub>3</sub>-N$ , TDS, EC, M-Al, and TH. This factor assigned as the "mineral and inorganic nutrient factor". F4 (which explained 7.984% of the total variance) was positively affected due to Cl<sup>−</sup>, PV and Na<sup>+</sup> and was negatively affected due to F. F5 explained 7.050% of the total variance and was positively contributed to by  $SO_4^{3-}$  and Mg and was negatively due to Ca. F5 is strongly correlated with Mg which is mainly originated from agricultural uses. F6, which explained 5.331% of the total variance, was positively participated by pH and was negatively participated by COD. F7 explained 5.069% of the total variance and was negatively contributed by K. F8, which explained 3.812% of the total variance, was positively due to Mn and B. F9 explained 3.559% of the total variance and was negatively influenced by BOD<sub>5</sub>.

Vega et al[.](#page-14-0) [\(1998](#page-14-0)) investigated the seasonal and polluting effects on water quality of the Pisuerga River (Spain) using exploratory data analysis. The first factor in this study was mostly contributed by mineral and inorganic nutrient-related water quality parameters. Also, Ouyang et al[.](#page-14-0) [\(2006](#page-14-0)) investigated the seasonal variations on water quality of the St. Johns River, Florida using principal factor analysis technique. However, in their study, the first factor was mostly contributed by organicrelated parameters and physical parameters. We attribute the discrepancies to the different river environments and different water quality parameters as well as to the different time periods used in each study. In our work, we found similar results with Vega et al[.](#page-14-0) [\(1998\)](#page-14-0) for all seasons except the fall.

### **Conclusion**

- 1. In this study, surface water quality data for 29 parameters collected from Esenkara monitoring station along the Porsuk Stream in Eskisehir, Turkey from 1995–2005 were analyzed, using the factor analysis. Results from the factor analysis show that river water temperature had relatively weak to fair correlations with other parameters for the entire for seasons except for DO in fall, which had a correlation coefficient of −0.751.
- 2. Very strong correlations between EC and other parameters (TDS, M-Al, PV, TH and Mg) were also found in winter and spring, but such correlations were not found in summer and fall.
- 3. Ouyang et al[.](#page-14-0) [\(2006](#page-14-0)), found strong correlation between the DO and the organic-related parameters (TKN, TOC and DOC) in spring, but the correlations were reduced moderately in summer and sharply in fall and finally recovered in winter. In this study, strong correlations between DO and other parameters were not found in four seasons. In addition, large seasonal variations in correlation coefficients between  $BOD<sub>5</sub>$  and other parameters were not found in this study. But, Ouyang et al[.](#page-14-0) [\(2006](#page-14-0)) found large seasonal variation in correlations between  $BOD<sub>5</sub>$  and organicrelation parameters in LSJR. These discrepancies might be due to climate differences

<span id="page-13-0"></span>because Turkey is cold area as compared to Florida.

- 4. Very strong correlations between TDS and other parameters (M-Al, TH and Mg) were found in spring, but the correlations were moderately reduced in summer and profoundly in fall, and finally recovered in winter. That is, the correlation coefficient between TDS and M-Al, TDS and TH and TDS and Mg were respectively 0.891, 0.949 and 0.908 in spring, 0.647, 0.673 and 0.299 in summer, 0.373, 0.431 and 0.017 in fall and 0.841, 0.917 and 0.903 in winter. These data imply that TDS was not always highly correlated with M-Al, TH and Mg in the Porsuk stream.
- 5. Large seasonal variation in correlations between SS and other parameters (Fe, *T-coli* and *E-coli*) were found in fall, but the correlations were moderately reduced spring and profoundly in summer, and finally recovered in winter. That is, the correlation coefficients between SS and Fe, SS and *T-coli* and SS and *E-coli* were respectively 0.820, 0.878 and 0.876 in fall, 0.676, 0.419 and 0.489 in spring, 0.578, 0.028 and 0.102 in summer and 0.878, 0.115 and 0.163 in winter. These data imply that SS was not always highly correlated with Fe, T-coli and E-coli in Porsuk stream.
- 6. The results of present study point to the existence of three main types of contamination of Porsuk Stream: mineral and inorganic nutrients, organic pollution, microbiological pollution each respectively of agricultural, domestic-industrial effluents and urban origin. These forms of contamination are in turn related to different terrestrial features including rivers and urban feature and so on.
- 7. Selection of correct parameters in water quality monitoring studies is an important point for technique and financial feasibility of monitoring programmes (Champman 1992). That's why, to continue the water monitoring studies of Porsuk stream at Esenkara station by using the parameters reduced from the factor analysis may result in technically and economically feasible outcome. In addition, these parameters can also be used in environmental studies to prevent the pollution.

## **References**

- Albek, E. (2003). Estimation of point and diffuse contaminant loads to streams by non-parametric regression analysis of monitoring data. *Water, Air, and Soil Pollution, 147*, 229–243. doi[:10.1023/A:1024592815576.](http://dx.doi.org/10.1023/A:1024592815576)
- Anazawa, K., & Ohmori, H. (2001). Chemistry of surface water at a volcanic summit area, Norikura, Central Japan: Multivariate statistical approach. *Chemosphere, 45*, 807–816. doi[:10.1016/S0045-6535](http://dx.doi.org/10.1016/S0045-6535(01)00104-7) [\(01\)00104-7.](http://dx.doi.org/10.1016/S0045-6535(01)00104-7)
- Cattel, R. B. (1965). Factor analysis: Introduction to essentials. *Biometrics, 21*, 190–215. doi[:10.2307/2528364.](http://dx.doi.org/10.2307/2528364)
- Champman, J. S. (1992). *Water quality assessments* (pp. 38– 39). London: Champman and Hall.
- Filik Iscen, C., Emiroglu, O., Ilhan, S., Arslan, N., Yilmaz, V., & Ahiska, S. (2008). Application of multivariate statistical techniques in the assessment of surface water quality in Uluabat Lake, Turkey. *Environmental Monitoring and Assessment, 144*, 269–276. doi[:10.1007/s10661-007-9989-3.](http://dx.doi.org/10.1007/s10661-007-9989-3)
- Hammer, J. M. (1986). *Water and Wastewater Technology* 3rd ed., (pp. 473–477). New Delhi, India: Prentice Hall.
- Helena, B., Pardo, R., Vega, M., Barrado, E., Fernandez, J. M., & Fernandez, L. (2000). Temporal evolution of groundwater composition in an alluvial (Pisuerga river, Spain) by principal component analysis. *Water Research, 34*, 807–816. doi[:10.1016/](http://dx.doi.org/10.1016/S0043-1354(99)00225-0) [S0043-1354\(99\)00225-0.](http://dx.doi.org/10.1016/S0043-1354(99)00225-0)
- Kuppusamy, M. R., & Giridhar, V. V. (2006). Factor analysis of water quality characteristics including trace metal speciation in the coastal environmental system of Chennai Ennore. *Environment International, 32*, 174–179. doi[:10.1016/j.envint.2005.08.008.](http://dx.doi.org/10.1016/j.envint.2005.08.008)
- Massart, D. L., Vandeginste, B. G. M., Deming, S. N., Michotte, Y., & Kaufman, L. (1988). *Chemometrices: A textbook*. Amsterdam: Elsevier.
- Mazlum, N., Ozer, A., & Mazlum, S. (1999). Interpretation of water quality data by principal components analysis. *Turkish Journal Of Engineering and Environmental Science, 23*, 19–26.
- Muhammetoglu, A., Muhammetoglu, H., Oktas, S., Ozgokcen, L., & Soyupak, S. (2005). Impact assessment of different management scenarios on water quality of porsuk river and dam system-Turkey. *Water Resources Management, 19*, 199–210, doi[:10.](http://dx.doi.org/10.1007/s11269-005-3473-z) [1007/s11269-005-3473-z.](http://dx.doi.org/10.1007/s11269-005-3473-z)
- Ntengwe, F. W. (2006). Pollutant loads and water quality in streams of heavily populated and industrialised towns. *Physics and Chemistry of the Earth, 31*, 832–839.
- Panda, U. C., Sundaray, S. K., Rath, P., Nayak, B. B., & Bhatta, D. (2006). Application of factor and cluster analysis for characterization of river and estuarine water systems—A case study: Mahanadi River (India). *Journal of Hydrology* (*Amsterdam*), *331*(3–4), 434– 445. doi[:10.1016/j.jhydrol.2006.05.029.](http://dx.doi.org/10.1016/j.jhydrol.2006.05.029)
- Praus, P. (2005). Water quality assessment using SDVbased principal component analysis of hydrological data. *Water South Africa 31*(4), 417–422.
- <span id="page-14-0"></span>Ouyang, Y. (2005). Evaluation of river water quality monitoring stations by principal component analysis. *Water Research, 39*, 2621–2635. doi[:10.1016/j.](http://dx.doi.org/10.1016/j.watres.2005.04.024) [watres.2005.04.024.](http://dx.doi.org/10.1016/j.watres.2005.04.024)
- Ouyang, Y., Nkedi-Kizza, P., Wu, Q. T., Shinde, D., & Huang, C. H. (2006). Assessment of seasonal variations in surface water quality. *Water Research, 40*, 3800–3810. doi[:10.1016/j.watres.2006.08.030.](http://dx.doi.org/10.1016/j.watres.2006.08.030)
- Shrestha, S., & Kazama, F. (2007). Assessment of surface water quality using multivariate statistical techniques: A case study of the Fuji river basin; Japan. *Environmental Modelling & Software, 22*, 464–475. doi[:10.1016/j.envsoft.2006.02.001.](http://dx.doi.org/10.1016/j.envsoft.2006.02.001)
- SHW (1995). *Water quality assessment in the Porsuk watershed for the year 1994* (p. 75). Ankara: Department of Technical Research and Quality Control (in Turkish).
- Simeonova, P., Simeonov, V., & Andreev, G. (2003). Analysis of the struma river water quality. *Central European Journal of Chemistry, 2*, 121–126. doi[:10.2478/](http://dx.doi.org/10.2478/BF02479264) [BF02479264.](http://dx.doi.org/10.2478/BF02479264)
- Simeonov, V., Stratis, J. A., Samara, C., Zachariadis, G., Voutsa, D., Anthemidis, A., et al. (2003). Assessment of the surface water quality in Northern Greece. *Water Research, 37*, 4119–4124. doi[:10.1016/](http://dx.doi.org/10.1016/S0043-1354(03)00398-1) [S0043-1354\(03\)00398-1.](http://dx.doi.org/10.1016/S0043-1354(03)00398-1)
- Simeonov, V., Simeonova, P., & Tsitouridou, R. (2004). Chemometric quality assessment of surface waters two case studies. *Chemical and Engineering Ecology, 11*(6), 449–469.
- Turkish Standards (2004). *Su kirliligi yönetmeligi, 25687 sayılı resmi gazete*.
- Vega, M., Pardo, R., Barrado, E., & Deban, L. (1998). Assessment of seasonal and polluting effects on the quality of river water by exploratory data analysis. *Water Research, 32*, 3581–3592. doi[:10.1016/S0043-](http://dx.doi.org/10.1016/S0043-1354(98)00138-9) [1354\(98\)00138-9.](http://dx.doi.org/10.1016/S0043-1354(98)00138-9)
- Voncina, D. B., Dobcnik, D., Novic, M., & Zupan, J. (2002). Chemometric characterization of the quality of river water. *Analytica Chimica Acta, 462*, 87–100. doi[:10.1016/S0003-2670\(02\)00298-2.](http://dx.doi.org/10.1016/S0003-2670(02)00298-2)
- World Health Organization (WHO) (1993). *Guidelines for drinking water quality* (2nd ed., Vol. 1, pp. 13). Geneva: Recommendations.
- Wunderlin, D. A., Diaz, M. P., Ame, M. V., Pesce, S. F., Hued, A. C., & Bistoni, M. A. (2001). Pattern recognition techniques for the evaluation of spatial and temporal variations in water quality. A case study: Suquia river basin (Cordoba-Argentina). *Water Research, 35*, 2881–2894. doi[:10.1016/S0043-](http://dx.doi.org/10.1016/S0043-1354(00)00592-3) [1354\(00\)00592-3.](http://dx.doi.org/10.1016/S0043-1354(00)00592-3)