



Gestational age at birth and risk of intellectual disability without a common genetic cause

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Abstract

Preterm birth is linked to intellectual disability and there is evidence to suggest post-term birth may also incur risk. However, these associations have not yet been investigated in the absence of common genetic causes of intellectual disability, where risk associated with late delivery may be preventable. We therefore aimed to examine risk of intellectual disability without a common genetic cause across the entire range of gestation, using a matched-sibling design to account for unmeasured confounding by shared familial factors. We conducted a population-based retrospective study using data from the Stockholm Youth Cohort ($n = 499,621$) and examined associations in a nested cohort of matched outcome-discordant siblings ($n = 8034$). Risk of intellectual disability was greatest among those born extremely early (adjusted $OR_{24 \text{ weeks}} = 14.54$ [95% CI 11.46–18.44]), lessening with advancing gestational age toward term ($aOR_{32 \text{ weeks}} = 3.59$ [3.22–4.01]; $aOR_{37 \text{ weeks}} = 1.50$ [1.38–1.63]); $aOR_{38 \text{ weeks}} = 1.26$ [1.16–1.37]; $aOR_{39 \text{ weeks}} = 1.10$ [1.04–1.17]) and increasing with advancing gestational age post-term ($aOR_{42 \text{ weeks}} = 1.16$ [1.08–1.25]; $aOR_{43 \text{ weeks}} = 1.41$ [1.21–1.64]; $aOR_{44 \text{ weeks}} = 1.71$ [1.34–2.18]; $aOR_{45 \text{ weeks}} = 2.07$ [1.47–2.92]). Associations persisted in a cohort of matched siblings suggesting they were robust against confounding by shared familial traits. Risk of intellectual disability was greatest among children showing evidence of fetal growth restriction, especially when birth occurred before or after term. Birth at non-optimal gestational duration may be linked causally with greater risk of intellectual disability. The mechanisms underlying these associations need to be elucidated as they are relevant to clinical practice concerning elective delivery around term and mitigation of risk in post-term children.

Keywords Intellectual disability · Gestational age · Stockholm Youth Cohort · Regression splines · Siblings · Post-term birth

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Introduction

Intellectual disability is a group of developmental disorders evident early in childhood and characterized by cognitive and functional impairments as a result of delayed or incomplete development of the mind [1]. Individuals with intellectual disability have a reduced ability to understand new or complex information and to learn and apply new skills, resulting in a reduced ability to cope independently [2]. Intellectual disability is thought to affect over 1% of the population [3, 4] although estimates vary with the demographic and socioeconomic composition of study populations [4, 5] and with definitions and study design [5, 6]. The cost of intellectual disability to individuals and society is substantial [7] and people living with these disabilities often face significant stigma [8] while encountering substantial health and social inequalities and early mortality [9].

Although there are many risk factors, a specific cause is identified for less than half of those with mild disabilities (IQ range 50–69) who make up the majority of cases [3, 10]. Mild intellectual disability often clusters within families [10] suggesting that genetic or other shared familial factors may influence risk. When disabilities are more severe, specific causes are identified in over 75% of cases, often involving genetic or chromosomal abnormalities and inborn errors of metabolism [10]. When intellectual disability is present without a specific genetic or chromosomal cause, it is associated with advanced maternal age, maternal risk behaviors or medical problems during pregnancy and fetal growth restriction [11], suggesting that these may be risk factors.

While it is known that children born preterm (< 37 completed weeks) are at greater risk of intellectual disability than those born at term [12], less is known about the development of risk along the gestational course, or about risk among post-term children (> 41 weeks). An early smaller study found no difference in intelligence between children born at term or post-term, although it was limited in terms of statistical power and length of follow-up [13]. More recent, larger, studies have suggested post-term birth may be associated with a range of adverse neurological, developmental, behavioural and emotional outcomes in early childhood [14, 15] and there is increasing evidence to suggest it is associated with cognitive and academic deficits in later childhood and adolescence [16–19], especially when the baby is growth-restricted [18].

The association between the full range of gestational duration, from very early to very late births, and intellectual disability has not yet been examined in population-based studies. Furthermore, the evidence to date is insufficient because of incomplete control of confounding from

shared familial factors and insufficient recognition that genetic causes of intellectual disability may also influence gestational duration [20, 21].

Therefore, in a large Swedish population-based cohort, we aimed to: (1) examine the associations between gestational age and intellectual disability without a common genetic cause, taking into account a range of potential confounders; (2) examine interactions between gestational duration and fetal growth in relation to risk of intellectual disability; and (3) explore the causal nature of associations between gestational duration and risk of intellectual disability in a nested cohort of matched outcome-discordant siblings.

Methods

Study cohort

The Stockholm Youth Cohort is a register-based cohort of all individuals who lived in Stockholm County for at least 1 year between 2001 and 2011 and were aged between 0 and 17 years during that period ($n = 736,180$) [22]. Using unique personal identification numbers, cohort members and their first-degree relatives were linked with a range of national and regional registers including information on pregnancy- and birth related characteristics, socioeconomic characteristics and medical and psychiatric diagnoses.

We excluded individuals with genetic and inborn metabolic syndromes who had been diagnosed with intellectual disability (13.6% of cases in our study population), children born outside Sweden, multiple births, adoptees, children < 4 years of age by the end of follow up on the 31st of December 2011, with a missing link to biological parents, or with missing data on gestational age or other covariates (Fig. 1). To account for potential recording errors, we excluded individuals with improbable birth weights (< 350 g or > 6000 g) and those with improbable combinations of birth weight and gestational age by deleting observations with values smaller than the 25th percentile minus 3 interquartile ranges, or larger than the 75th percentile plus 3 interquartile ranges, from sex- and week-specific birth weight distributions (Table S8) [18]. This left a cohort of 499,621 individuals to examine population-level associations between gestational age and intellectual disability. To examine associations among matched siblings, we excluded individuals without full siblings in the cohort and families with outcome-concordant offspring ($n = 491,587$) leaving a cohort of 8034 matched outcome-discordant full siblings.

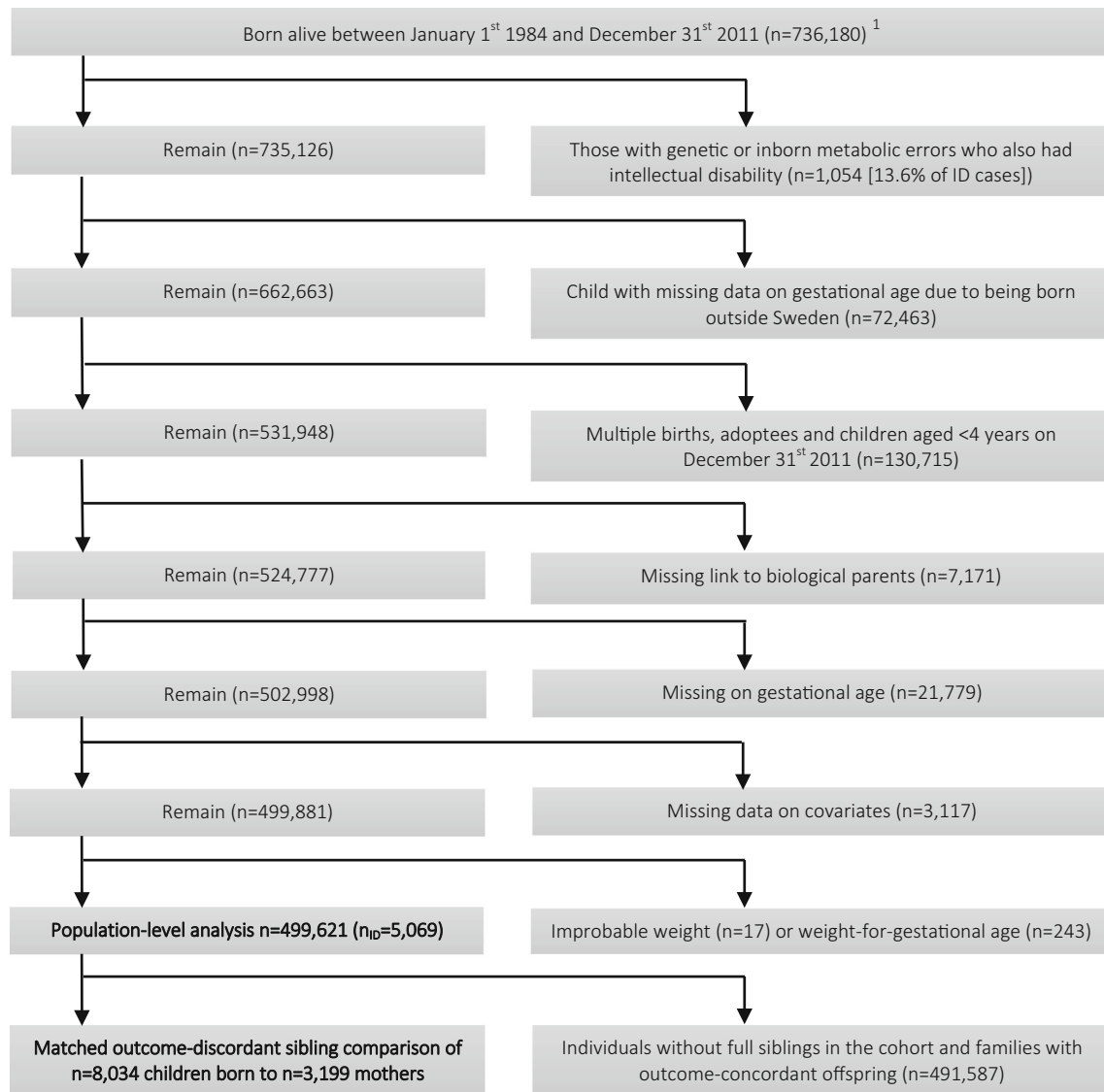


Fig. 1 Selection of the study cohort

Exposure

We obtained information on gestational age at birth from the Medical Birth Register (MBR), constructing a categorical variable to define extremely to very preterm births (21–31 completed weeks), moderately to late preterm births (32–36 weeks), term births (37–41 weeks), post-term births (42 weeks) and very post-term births (43–45 weeks) for use in descriptive statistics and as an exposure variable in regression analyses. We also used a continuous definition of gestational age (in days) for regression analyses.

Outcome

We used a multisource ascertainment approach to identify cohort members with intellectual disability, similar to the case identification for autism described elsewhere [22]: We used the national patient register, the Stockholm county child and adolescent mental health register, the Stockholm County healthcare database (VAL) and the Stockholm adult psychiatric register to identify all inpatient or outpatient diagnoses of intellectual disability recorded using ICD-10 (F70-79) and DSM-IV (317-318) codes and supplemented these diagnoses with a record of care at specialist habilitation services for individuals with intellectual disability in Stockholm County. We identified individuals with genetic defects and inborn errors of metabolism commonly associated with intellectual disability to identify

cases where a known genetic or metabolic cause was present (Table S1).

Covariates

To control for secular change in obstetric and diagnostic practice, we obtained year of birth from the Medical Birth Register (MBR). We then identified additional covariates which in the literature have been associated with pregnancy duration and risk of intellectual disability in offspring. From the MBR, we extracted data for offspring sex [5, 23], parity (1/2/3/4 +) [11, 24], maternal age (< 20/20–24/25–29/30–34/35–39/40–44/45 +) [11, 24], gestational diabetes [11, 25] and gestational hypertension or preeclampsia [11, 26]. We obtained birth weights [11] from the MBR to construct a measure of weight-for-gestational age by examining week- and sex-specific birth weight distributions, and identifying those in the lower and upper deciles of these distributions as born small or large for gestational age respectively. Weight-for-gestational age is therefore conceptualized as the distance between the birth weight of an individual and the average birth weight of all who were born in the same gestational week as that individual. As weight-for-gestational age is orthogonal to gestational age itself, this avoids the issue of collinearity when the measure is included as a covariate in the analytical model. To examine potential interactions between gestational age and weight-for-gestational age, we constructed a categorical measure to identify those born pre-term (< 37 weeks) and small for gestational age, appropriate for gestational age (11th centile to 90th centile) or large for gestational age; those born at term (37–41 weeks) and small, appropriate or large for gestational age; and those born post-term (≥ 42) and small, appropriate or large for gestational age. We also extracted information for maternal and paternal country of birth (Sweden/other Nordic/other European/Russia or Baltic States/Africa/Middle East/Asia or Oceania/North America/South America) [5, 27], maternal and paternal history of psychiatric treatment [28, 29], quintiles of disposable family income adjusted for inflation and family size [30, 31], and parental educational attainment (≤ 9 years/10–12 years/ ≥ 13 years) [31, 32] at (or as close as possible to) birth.

Statistical analyses

Analyses were conducted in Stata/MP version 14.2. We examined the characteristics of the study cohort by gestational duration at birth. To examine population-level associations between gestational duration and risk of intellectual disability, we used generalized estimating equation (GEE) multivariable regression models with a

logit link function, exchangeable correlation structure and robust variance estimators to ensure that the standard errors of our estimates were robust against clustering of intellectual disability within families [33]. We calculated restricted cubic regression splines based on five knot locations (5th, 27th, 50th, 73rd and 95th percentiles of the gestational age distribution) to allow for non-linear associations between continuously varying gestational duration and later risk of intellectual disability [34]. We statistically adjusted our estimates for covariates and calculated odds ratios by continuously varying gestational age at birth to estimate risk of intellectual disability associated with birth at specific moments along the gestational course. We investigated potential interactions between gestational age and fetal growth using GEE models with a categorical exposure variable to assess risk of intellectual disability among those born at varying gestational duration (preterm/post-term) and weight for gestational age (small/appropriate/large) with statistical adjustment for confounders.

In a nested cohort of matched outcome-discordant siblings we examined associations between continuously varying gestational age and risk of intellectual disability with conditional likelihood logistic regression models. This allowed us to explore the potential influence of unobserved familial traits, e.g. residual genetic risk/unmeasured socioeconomic factors/parental health behaviors, which may have confounded associations between gestational length and risk of intellectual disability. If we were to observe associations at the population level, non-association within families would suggest confounding by these shared familial traits. Conversely, replication of population-level associations within families would suggest they were robust against shared familial confounding, thereby allowing stronger causal inference from our result [35]. We statistically adjusted within-family associations for non-shared confounding characteristics including sex, parity, gestational diabetes, hypertension or preeclampsia, weight for gestational age, maternal and paternal age, disposable family income quintile, and parental educational attainment.

Sensitivity analyses

We compared characteristics for those with missing and complete data to assess whether our estimates may have been affected by selection bias (Table S2). To ensure that the association between gestational age and intellectual disability was not driven by presence of co-occurring autism spectrum disorder or attention deficit hyperactivity disorder (which are associated with intellectual disability [36–38] and for which risk may also vary by gestational age [39, 40]) we examined associations in a subset of the cohort without a record of these conditions (Figure S1 and

Table S3). We examined whether the risks of intellectual disability associated with preterm or post-term birth varied with mode of delivery (Tables S4 and S5) using categorical measures to identify those born vaginally or by Caesarean section and in unassisted or forceps-/ventouse-assisted deliveries at varying gestational duration. Finally, we conducted post hoc analyses to assess whether risk varied among children born in spontaneous or induced deliveries at varying gestational duration (Table S6).

Results

Prevalence of intellectual disability without a common genetic cause was estimated at 1% in our study population (Fig. 1). Characteristics of the study cohort are described in Table 1. Prevalence among those born at term gestation was 0.9%. By contrast, 5.6% of children born extremely to

very preterm and 1.6% of those born very post-term had intellectual disability.

Examining associations between gestational duration and risk of intellectual disability in a model using a continuous exposure variable with statistical adjustment for potential confounders (Fig. 2), the adjusted odds ratio (aOR) for risk at extremely preterm birth (at 24 weeks) was estimated at 14.54 [95% CI 11.46–18.44]. This risk decreased with gestational age towards term (aOR_{32 weeks} = 3.59 [3.22–4.01]; aOR_{37 weeks} = 1.50 [1.38–1.63]; aOR_{38 weeks} = 1.26 [1.16–1.37]; aOR_{39 weeks} = 1.10 [1.04–1.17]) after which it increased with gestational age post-term (aOR_{42 weeks} = 1.16 [1.08–1.25]; aOR_{43 weeks} = 1.41 [1.21–1.64]; aOR_{44 weeks} = 1.71 [1.34–2.18]; aOR_{45 weeks} = 2.07 [1.47–2.92]).

We report associations using a categorical exposure variable in an online supplement (Table S7). Irrespective of gestational length, risk of intellectual disability was

Table 1 Characteristics of the sample by exposure status

	Extremely to very preterm 21–31	Moderate to late preterm 32–36	Term 37–41	Post-term 42	Very Post-term 43–45
Gestational weeks	21–31	32–36	37–41	42	43–45
Number of observations	2601	20,271	438,215	34,828	3706
Percentage of the cohort	0.5 %	4.1 %	87.7 %	7.0 %	0.7 %
Female child	45.2	46.3	49.3	44.1	43.9
Mother's number of prior pregnancies					
0	53.4	53.4	44.2	53.8	61.5
1	27.1	28.4	37.0	30.1	23.5
2	12.5	11.8	13.6	11.6	10.0
3 +	7.0	6.4	5.3	4.5	5.0
Birth weight in grams					
< 2500	99.7	37.6	1.1	0.1	0.1
2500–4500	0.4	62.5	98.6	98.5	98.0
> 4500	0.0	0.0	0.3	1.4	1.9
Gestational hypertension or preeclampsia	23.6	12.9	3.2	1.9	1.5
Gestational diabetes	1.5	1.7	0.8	0.4	0.4
Delivery by Caesarean Section	58.9	32.4	13.3	16.5	22.5
Delivery assisted with ventouse or forceps	1.4	4.7	8.0	14.1	14.9
Maternal psychiatric history	38.6	37.0	32.7	31.5	33.0
Paternal psychiatric history	23.7	22.6	20.9	20.4	21.2
Family disposable income quintile at birth					
Lowest	14.3	15.1	14.7	13.3	15.4
Second	21.7	20.4	20.8	19.4	19.1
Third	19.9	20.4	21.6	21.1	19.3
Fourth	22.3	22.3	21.6	22.6	22.1
Highest	21.8	21.8	21.4	23.6	24.2
Parental educational attainment at birth					
≤ 9 years	9.0	8.0	6.5	5.8	7.0
10–11 years	42.0	43.3	40.5	39.9	40.7
≥ 13 years	49.0	48.7	53.0	54.2	52.3

Table 1 continued

	Extremely to very preterm 21–31	Moderate to late preterm 32–36	Term 37–41	Post-term 42	Very Post-term 43–45
Gestational weeks	21–31	32–36	37–41	42	43–45
Number of observations	2601	20,271	438,215	34,828	3706
Percentage of the cohort	0.5 %	4.1 %	87.7 %	7.0 %	0.7 %
Maternal age					
≤ 20	1.8	2.6	1.8	1.6	1.9
20–24	13.0	15.7	14.7	13.4	16.3
25–29	26.6	30.3	31.0	31.3	32.0
30–34	32.1	31.1	33.8	34.9	31.8
35–39	21.4	16.4	15.7	16.0	15.1
40 +	5.2	4.0	3.2	2.8	2.9
Paternal age					
≤ 20	0.7	0.8	0.5	0.5	0.4
20–24	7.2	8.9	7.2	7.0	8.2
25–29	20.8	24.2	23.5	23.5	25.0
30–34	30.8	31.5	33.5	33.7	32.5
35–39	22.8	20.4	22.0	22.0	21.4
40 +	17.7	14.2	13.3	13.3	12.6
Maternal country of birth					
Sweden	72.2	75.2	76.0	78.5	78.5
Other Nordic	5.1	4.6	4.2	4.2	4.6
Other European	4.3	3.6	3.6	3.5	3.3
Baltic States /Russia	0.6	0.5	0.5	0.7	0.5
Africa	5.3	3.1	3.4	4.9	5.5
Middle East	7.3	6.7	7.0	4.8	4.5
Asia /Oceania	3.2	3.7	2.8	1.7	1.7
North America	0.5	0.6	0.6	0.6	0.4
South America	1.5	2.0	1.9	1.3	1.1
Paternal country of birth					
Sweden	70.7	74.7	74.7	77.3	77.6
Other Nordic	4.3	3.5	3.2	3.0	3.4
Other European	4.6	4.5	4.7	4.6	4.5
Baltic States /Russia	0.2	0.2	0.2	0.2	0.2
Africa	6.2	4.1	4.1	5.7	5.8
Middle East	8.6	7.6	8.3	5.8	5.5
Asia/Oceania	2.4	2.3	1.9	1.2	1.1
North America	0.8	0.7	0.8	0.7	0.70
South America	2.3	2.3	2.2	1.5	1.21
Intellectual disability	5.6	1.8	0.9	1.0	1.6

greatest among those showing evidence of fetal growth restriction (Table 2).

This difference was most pronounced in the preterm group, but our results suggest risk of intellectual disability was also increased among children born post-term and growth-restricted. Associations between gestational length and risk of intellectual disability persisted when we

repeated our analysis in a nested cohort of outcome-discordant siblings (Fig. 3, Table S7).

In a subset of the cohort without a diagnosis of ASD or ADHD, pre- and post-term birth remained associated with increased risk of intellectual disability (Figure S1, Table S3). Among those born at 21–31 completed weeks of gestation, risk of intellectual disability was lesser when the baby was delivered by Caesarean section, while Caesarean

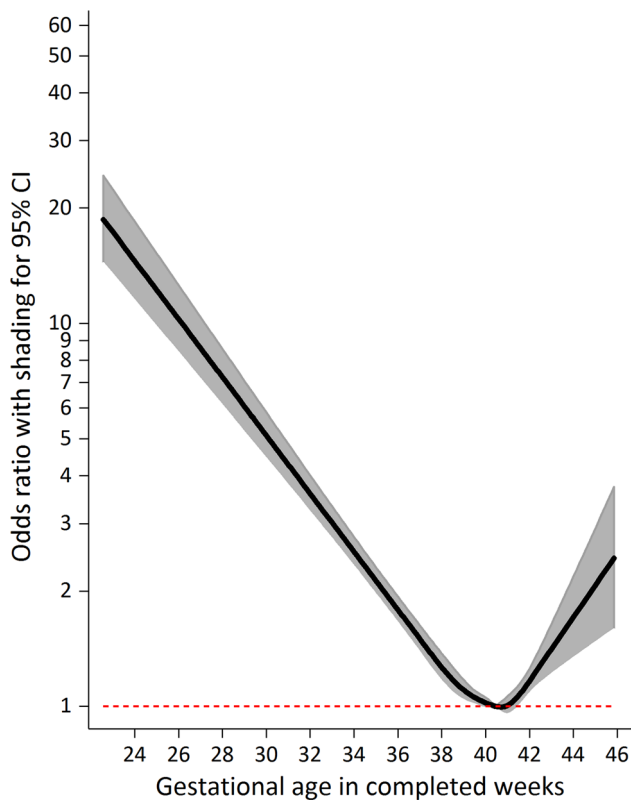


Fig. 2 Population-level association between gestational duration and risk of intellectual disability. Notes: The population-level association ($N = 499,621$) was estimated using a generalized estimating equations model with a logit link, and adjusted statistically for year of birth, child sex, parity, gestational hypertension or preeclampsia, gestational diabetes, birth weight for gestational age, maternal and paternal age, maternal and paternal psychiatric history, maternal and paternal country of birth, family disposable income quintile at birth, and parental educational attainment at birth. Those born at 40 weeks and 3 gestational days are the referent

birth was associated with greater risk than vaginal birth between 37 and 41 weeks gestation (Table S4). There was no consistent variation in risk due to unassisted versus assisted delivery within gestational age categories (Table S5). Importantly, risk of intellectual disability associated with early or late birth remained when considering those born in vaginal or unassisted deliveries (Tables S4 and S5). Among those born between 37 and 41 weeks, risk of intellectual disability was greater when birth was induced (Table S6). This effect existed independently of the influence of fetal growth restriction or other potential confounders. Finally, children born in induced post-term deliveries were at greater risk of intellectual disability than children born spontaneously at term, while the increase in risk associated with spontaneous post-term birth was lesser (Table S6).

Discussion

In this large population-based study, we found a greater risk of intellectual disability without a common genetic cause among preterm and post-term births compared with term births. Risk also varied within the term period and was lowest when the child was born at 40–41 completed weeks' gestation. These associations were evident in analyses using the full sample, as well as in a nested cohort of matched outcome-discordant siblings. Risk of intellectual disability was greatest among those showing evidence of fetal growth restriction, especially when born before or after term. To our knowledge, this is the first total-population study to estimate risk of intellectual disability without a common genetic cause over the entire range of gestation using high-quality prospectively measured data. In addition to a range of measured confounders, this study explored the influence of unmeasured familial effects using a matched sibling design. This allowed us to take into account unmeasured familial confounding of the association between intellectual disability and gestational length, as these traits are heritable within families [10, 41, 42].

There were several limitations. First, 5% of the study cohort had missing data on gestational age at birth or other covariates. Although we cannot know with certainty how these exclusions may have affected our result, sensitivity analyses suggest that our estimates may have been conservative as we may have excluded preterm children with higher prevalence of intellectual disability (Table S2). Second, we did not have information on whether gestational length was calculated by the mother's report of her last menstrual period or based on ultrasound measurement in specific pregnancies. As our sample includes births from 1984 onwards, it is likely that there is greater measurement error in earlier cohort years, where gestational length would have been estimated by last menstrual period for a larger proportion of pregnancies. This may have resulted in overestimation of rates of post-term birth [43–45] and underestimation of population-level [46] and within-family associations [47] between gestational length and later risk of intellectual disability. Third, while the matched-sibling design provides a powerful method to examine the influence of shared confounding, it is more sensitive than traditional methods to confounders not perfectly shared by the siblings. Selection based on exposure-discordance could also prompt discordance in terms of non-shared confounding characteristics, which may bias the within-family effect [47]. The size and direction of such bias depends on the similarity or dissimilarity of matched siblings in terms of exposure and confounding characteristics [47]. Given that measurement error in the gestational age variable would have downwardly biased our estimate of the within-

Table 2 Interaction between gestational duration and fetal growth in relation to risk of intellectual disability

Gestational duration ^a /weight-for-gestational age category ^b	Odds ratio ^c	95% CI		<i>p</i>	n (%) ^d	N ^e	Percentage of the cohort (%)
		Lower	Upper				
Preterm/small	3.77	2.97	4.84	< 0.001	73 (3.8)	1935	0.4
Preterm/appropriate	2.24	2.01	2.51	< 0.001	380 (2.1)	18,256	3.7
Preterm/large	2.36	1.81	3.07	< 0.001	61 (2.3)	2681	0.5
Term/small	1.88	1.73	2.05	< 0.001	731 (1.8)	41,579	8.3
Term/appropriate	1.00				3008 (0.9)	352,016	70.5
Term/large	1.06	0.95	1.18	0.27	402 (0.9)	44,620	8.9
Post-term/small	2.29	1.83	2.85	< 0.001	85 (2.2)	3822	0.8
Post-term/appropriate	1.11	0.98	1.25	0.10	292 (1.0)	30,840	6.2
Post-term/large	1.22	0.87	1.69	0.24	37 (1.0)	3872	0.8

^aPreterm was defined as birth < 37 completed weeks of gestation. Term birth was defined as birth between 37 and 41 completed weeks of gestation. Post-term birth was defined as birth at ≥ 42 completed weeks of gestation

^bFetal growth categories were defined as small-for-gestational age [in the lowest decile of the gestational age-specific birthweight distribution], appropriate-for-gestational age [in the 11th to 90th decile of the gestational age-specific birthweight distribution] and large-for-gestational age [in the upper decile of the gestational age-specific birthweight distribution]

^cPopulation-level associations were estimated using a generalized estimating equations model with a logit link, and adjusted statistically for year of birth, child sex, parity, gestational hypertension or preeclampsia, gestational diabetes, maternal and paternal age, maternal and paternal psychiatric history, maternal and paternal country of birth, family disposable income quintile at birth, and parental educational attainment at birth

^dNumber and percentage of ID cases within gestational duration/fetal growth category

^eNumber of observations within gestational duration/fetal growth category

^fN = 499,621

family effect, additional bias due to sibling non-shared confounding would have either offset this downward bias or further enhanced it. Fourth, there may be bias due to omitted non-shared confounding characteristics in our matched sibling analyses. For example, it is possible that prenatal infection [48], maternal obesity [49, 50], or use of drugs or alcohol [51] may have influenced gestational length and resulted in greater risk of offspring intellectual disability in as far as these factors were present in one pregnancy but not the other.

The mechanisms underlying our findings are likely to differ depending on whether birth occurred before or after the due date. With regards to preterm birth, perturbations in development of the fetal brain because of shortened gestation can increase risk for longer-term neurodevelopmental problems [52, 53]. Our findings for preterm small-for-gestational age children would suggest that these effects might become particularly apparent if the fetus is already growth-restricted. After birth, further injury to the brain could result from respiratory support for preterm infants with immature pulmonary function [54]. Mechanisms linking post-term birth with later risk of intellectual disability might involve placental deterioration or insufficiency causing fetal hypoxia or nutritional deficiencies [55], which in turn could result in injury to the fetal brain. Meconium aspiration, which is more common in post-term

birth [55], may result in neonatal asphyxia thereby incurring risk for brain injury and later neurodevelopmental problems [56].

Our finding of associations among those born in unassisted or vaginal deliveries suggested that adverse obstetric circumstances did not explain the higher risk of intellectual disability associated with birth at < 37 or > 42 weeks. Furthermore, our findings suggest that risk of intellectual disability increases with induction of labor at further post-term gestation, although these estimates are likely to be biased by the higher risk nature of induced pregnancies as a whole (Table S6). Risk of intellectual disability may have also increased with advancing post-term gestational age when delivery started spontaneously, although our data may have been underpowered to detect these subtler effects (Table S6). Importantly, due to the observational nature of our study, and given that the decision to induce labor will be informed by other factors than gestational length alone, we cannot infer from our data whether the risks associated with post-term delivery could be curtailed by induction of labor around term. This question will therefore be better answered by randomized studies designed specifically for the purpose of comparing outcomes of children induced at late term with those born post-term by expectant management [57]. Finally, the generalisability of our findings may vary with regional differences in practice regarding the

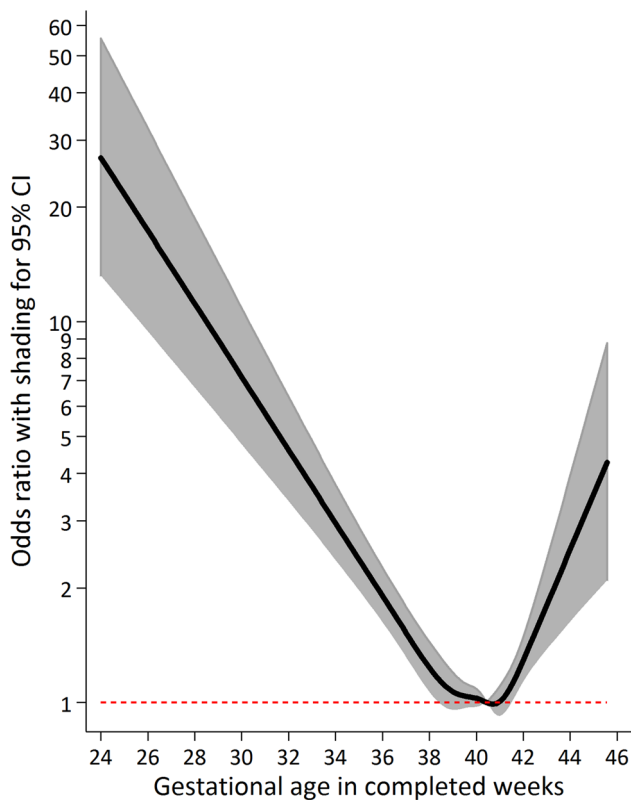


Fig. 3 Within-family association between gestational duration and risk of intellectual disability. Notes: The within-family association ($N = 8034$) was estimated using a conditional likelihood logistic regression model, and adjusted statistically for year of birth, child sex, parity, gestational hypertension or preeclampsia, gestational diabetes, birth weight for gestational age, maternal and paternal age, family disposable income quintile at birth, and parental educational attainment at birth. Those born at 40 weeks and 3 gestational days are the referent.

management of pre- or post-term pregnancy and in the quality of obstetric and neonatal care.

Our findings are consistent with other studies examining risk of cognitive deficit in relation to birth before or after term gestational duration [12, 16–19, 58–64]. These studies suggest there may be increased risk of intellectual disability [12, 59], special educational needs [17, 62, 63], poorer performance in school [12, 18, 58, 60, 63] and lower IQ in childhood [16, 64] or adulthood [19, 61] when children are born before or after term. Furthermore, our findings are consistent with those of other studies investigating variation within the term period and reporting that birth at 40–41 weeks' gestation was optimal in relation to IQ scores at age six [16], risk of special educational needs in primary or secondary school [17], and end-of-secondary school performance outcomes [18]. The independent risk of intellectual disability associated with being born small-for-gestational age is consistent with earlier studies examining other outcomes for fetal growth-restriction in

infants born at preterm or post-term gestational duration [16, 58, 65, 66].

In conclusion, our findings suggest that delivery at non-optimal gestational age is associated with greater risk of intellectual disability in offspring in the absence of common genetic causes. This association existed independently of a range of measured potential confounders as well as unmeasured confounding from shared familial factors. While this study cannot provide conclusive evidence for causality, our use of a matched sibling design helped to address confounding due to unmeasured shared familial factors, thereby providing a better estimate of the causal effect than in studies using traditional methods for dealing with confounding. As birth at non-optimal gestational duration may be linked causally with greater risk of intellectual disability, it is important that the mechanisms underlying these associations are elucidated because of their relevance to clinical practice concerning elective delivery within the term period and the mitigation of risk in children who are born post-term. Our work highlights the need for randomized controlled studies to establish whether offspring neurodevelopmental outcomes would improve if women in post-term pregnancies were routinely induced.

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Compliance with ethical standards

Conflict of interest statement The authors declare having no conflicting interests.

Ethical approval Ethical approval for this study was granted by the research ethics committee at Karolinska Institute [2010/1185-31/5 and 2013/1118-32], allowing record linkage without personal consent when the confidentiality of the individuals is maintained. The personal identity of participants was replaced with a serial number before the research group were given access to these data.

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