REVIEW



Meta-Analysis for Cyto-Pathological Outcomes in Endoscopic Ultrasonography-Guided Fine-Needle Aspiration With and Without the Stylet

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Abstract

Background Theoretically, use of the stylet during endoscopic ultrasound-guided fine-needle aspiration (EUS-FNA) will prevent the tip of the needle from being blocked and the sample from being contaminated, improving the quality and diagnostic yield of specimens.

Aim This meta-analysis aimed to evaluate any benefit from the use of a stylet during EUS-FNA.

Methods MEDLINE, EMBASE, WEB of SCIENCE, and the Cochrane Central Register of Controlled Trials were searched for articles published through October 2015. A fixed-effects model was used to calculate the pooled effects when there is no heterogeneity and a random-effects model was used in situations with significant heterogeneity.

Results Five prospective randomized controlled studies and two retrospective studies of a total of 5491 specimens (2504 specimens in the S+ group and 2987 specimens in the S- group) were included in this study. The pooled analysis demonstrated no significant differences in the adequate sample rate between the S+ group (2135/2504, 85.26 %) and S- group (2609/2987, 87.35 %) [odds ratio 0.94 (95 % confidence interval 0.79–1.11), p = 0.45]. Furthermore, the rate of cellularity > 50 %, the contamination rate, and the blood contamination rate were not significantly superior in the S+ group when compared with the S- group.

Se Woo Park britnepak@hallym.or.kr *Conclusions* Whether the use of a stylet improves the proportion of adequate specimens or quality of the specimen is questionable. Our results suggest that the use of a stylet dose not guarantee any advantage during EUS-guided tissue sampling.

Introduction

Endoscopic ultrasound-guided fine-needle aspiration (EUS-FNA) has become widely accepted as an effective standard technique for establishing tissue diagnosis for patients with suspected gastrointestinal tract and adjacent organ (lymph nodes, mediastinal masses, or adrenal tumors), gastrointestinal subepithelial lesions, and pancreatic neoplasms [1-3]. However, especially in regards to pancreatic neoplasm, this technique presents a high diagnostic accuracy but a relatively low negative predictive value (NPV): median figures for sensitivity, specificity, NPV, and diagnostic accuracy are 83 % (range 54-95 %), 100 % (range 71-100 %), 72 % (range 16-92 %) and 88 % (range 65-96 %), respectively [3-7]. Thus far, many technical trials including manipulation of the stylet have been investigated for their potential to improve the diagnostic yield of EUS-FNA, because many confounding factors have been suggested to affect the diagnostic outcomes. These include the degree of technical difficulty, experience of the echoendoscopist and the cytopathologist, type and diameter of the needle, characteristics of the lesion (anatomical location, size, stiffness of the target lesion), number of passes, application of

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suction, method of tissue acquisition, and immediate onsite cytopathology assessment [8–10]. Among these trials, several reports [11-14] have focused on the need for a stylet during EUS-FNA. Theoretically, the puncture for the lesion needs to be performed with a stylet in the needle lumen to prevent blockage or contamination by the intestinal mucosal tissue and to allow more adequate aspiration of the target tissue. Furthermore, a stylet is needed to maintain needle stiffness, allowing a puncture of a fibrotic lesion [12]. Contrary to the above, some intriguing studies recently found that the use of a stylet can lead to increased procedure time and the risk of an unintentional needle stick injury, especially when multiple passes are needed [2, 15]. In particular, controversies exist regarding the sample adequacy and diagnostic yield for EUS-FNA. One study [11] reported that the use of the stylet resulted in lower sample quality with significant difference, although other studies [2, 12–15] reported that the stylet did not affect clinical outcomes, such as sample adequacy or diagnostic yield, and thus its clinical benefit in daily practice may be limited.

Several studies consistently reported that there were no advantages to using a stylet during EUS-FNA, but each of those studies had limitations, such as small sample sizes. Moreover, a meta-analysis on this subject has not been previously performed. Therefore, the aim of our metaanalysis is to discover if there is any benefit from the use of a stylet during EUS-FNA.

Materials and Methods

This systematic review was conducted and reported according to the preferred reporting items for systematic review and meta-analyses (PRISMA) statement [16].

Eligibility Criteria and Definitions

Both prospective and retrospective cohort studies, casecontrol studies, and randomized clinical trials (RCTs) of adult patients who underwent EUS-FNA, published as full articles, were considered. Studies were eligible for inclusion in the meta-analysis if they met the following criteria: (1) comparative design for evaluating the sample quality and diagnostic ability of the specimen from EUS-FNA with or without a stylet; (2) data not duplicated in another manuscript; and (3) endpoints including at least one appropriate cyto-pathological outcome among the following: adequate specimen, cellularity, contamination, or blood contamination. Inclusion was not restricted by study size or language. Exclusion criteria were: (1) comments, reviews, or guideline articles; (2) studies reporting non-gastrointestinal endoscopy, such as laparoscopic procedures or bronchoscopy; (3) studies providing no data on cyto-pathological outcomes; and (4) studies repeatedly reported (in which case, only the latest reported article was selected).

Adequate samples were defined as passes providing representative tissue from the sampled organ. The widely accepted definition for cellularity was the percentage area of slides that contained cells of the representative lesion: "0," no representative cells present; "1," representative cells present in <25 % area of the slides; "2," representative cells present in 25–50 % area of the slides; "3," representative cells present in >50 % area of the slides. Similarly, contamination was defined as the following: "0," no contamination seen; "1," contamination present in <25 % area of the slides; "2," contamination present in >50 % area of the slides; "3," contamination present in >50 % area of the slides. Finally, blood contamination was defined as the following: "1," minimal; "2," moderate; "3," significant.

Information Sources and Search Strategy

A literature search was conducted to identify all relevant studies that compared the clinical outcomes for EUS-guided tissue sampling with and without a stylet. A systematic literature search of MEDLINE, EMBASE, and WEB of SCIENCE databases and the Cochrane Central Register of Controlled Trials updated was conducted. Furthermore, manual searching through other sources was conducted according to PRIMA flow diagram. The following medical subject heading terms were used: "endoscopic ultrasound," "fine-needle aspiration," and "stylet." Internet search engines, Google Scholar, and Yahoo were also searched with the relevant key words. The latest date for updating the search was 1 October 2015.

Study Selection

After removing duplicate studies, the titles and abstracts of studies found by the searches were examined to exclude articles with irrelevant study designs. If available, the full texts of all selected studies were screened according to the inclusion and exclusion criteria. Selected full text articles were critically appraised for relevance and validity. Two investigators (S.W.P. and J.H.K.) independently evaluated the studies for their eligibility and subsequently resolved any disagreements by discussion, together with clinical expert consultation.

Data Extraction and Study Quality Assessment

The data retrieved from each study included the year of publication, country, research design, number of individuals and specimens in the S+ and S- groups, type of needle used, number of needle passages in each group, and the primary and secondary outcomes. To avoid bias in the data extraction process, two investigators (S.W.P. and J.H.K.) independently evaluated each study's quality and compared their results with one another on blinded state regarding journal titles, author names, and institutional affiliations. In cases of disagreement, the third investigator made a determinative decision. To minimize the risk of bias in the included studies, a formal quality assessment was conducted. The methodological quality of the RCT was assessed by two authors independently (S.W.P. and J.H.K.) using Cochrane Handbook for Systematic Reviews of Interventions Version 5.1.0 (updated March 2011) [17]. Furthermore, the quality of the selected non-randomized retrospective studies was evaluated using the Newcastle-Ottawa Scale (NOS) [18], which evaluated studies with a star-rating system ranging from 0 (the lowest quality) to 9 (the highest quality) stars, based on three study components, including selection, comparability, and outcome assessment. Articles that were awarded five or more stars were considered high-quality studies, whereas the other studies were considered to be of low quality.

Evaluation Criteria for Endpoints

The primary end point was the rate of adequate specimens from EUS-guided tissue sampling with or without a stylet. Secondary end points were: (1) the rate of cellularity >50 %; (2) the contamination rate; (3) the blood contamination rate; (4) diagnostic ability.

Statistical Analysis

The meta-analysis was carried out with Review Manager 5.3 software (provided by the Cochrane Collaboration). Statistical analysis for dichotomous variables including the rate of adequate specimens, the rate of cellularity > 50 %, the contamination rate, and the blood contamination rate was performed using the odds ratio (OR) as the summary statistic [with a corresponding 95 % confidence interval (CI)]. For outcome data regarding diagnostic ability, we calculated mean differences or the standardized mean difference. Heterogeneity among the studies was assessed using the Chi squared test. A p value < 0.05 was considered to suggest significant heterogeneity. At the same time, I^2 was also used to assess heterogeneity. An $I^2 > 50$ % was considered to indicate statistical significance. Meta-analysis was to be calculated by using a fixed-effects model when no heterogeneity was detected. If any heterogeneity existed, heterogeneity was to be explored by subgroup analysis or a random-effects model.

Results

Study Selection

Our search strategy yielded a total of 37 potential studies for the meta-analysis (Fig. 1). After searching the titles and abstracts, we excluded 11 repeated articles. Then, based on the inclusion and exclusion criteria, nine articles were excluded: five studies were unrelated to the present metaanalysis and four studies dealt with only non-gastrointestinal endoscopy, such as percutaneous methods or bronchoscopy. Among the remaining 17 papers, an additional ten articles were excluded due to inappropriate study design (review articles). Finally, seven studies were found to be appropriate for meta-analysis [2, 11-14, 19, 20]. These studies described a total of 5491 specimens: 2504 specimens in the S+ group and 2987 specimens in the Sgroup. All studies were published in the past 6 years, from 2010 to 2015. Five studies [2, 11, 13, 15, 20] were prospective randomized controlled studies (RCTs) (one [15] of the five RCTs was a prospective single-blind randomized crossover study) and two studies [12, 14] were retrospective studies. Of the seven studies, three [2, 12, 13] were conducted in the USA, two [11, 14] in Canada, and one each in India [20] and Japan [15]. All studies were reported as full-text articles in English. The baseline characteristics of the studies included in the meta-analysis are summarized in Table 1.

Risk of Bias in Included Studies

A summary of the risk of bias across the five RCTs [2, 11, 13, 15, 20] is given in Figs. 2 and 3. Randomization and allocation concealment were adequately described in four [2, 11, 13, 15] of the included studies and were judged as unclear in the remaining one study [20]. The studies were not double blind because the endoscopist must know the technique to be applied. However, the review authors judged that the outcome and the outcome measurement were not likely to be influenced by lack of blinding. Only one study [20] was judged as unclear for incomplete outcome data; all remaining studies were judged as having low risk of bias. A total of four studies [2, 11, 13, 15] were rated as having low risk of bias for selected outcome reporting. The remaining one study was rated as having unclear risk of bias. However, in the absence of initial study protocols, it was not clear whether all studies reported outcomes as pre-specified in the trial protocol. Other potential sources of bias were noted in the design of studies in terms of interventions used and outcomes assessed. Although some studies used different definitions of outcome measurement, almost all studies included in the





meta-analysis that reported adequacy of specimens, cellularity, and contamination criteria were very similar, suggesting that the definition of these outcomes was not a source of bias. Furthermore, the result of the NOS assessment is shown in Table 2. According to the protocol, two articles [12, 14] were regarded as high quality reports with seven stars.

Adequate Specimens

All of the included studies [2, 11–15, 20] reported the rate of adequacy of specimens, and the pooled analysis using the fixed-effects model showed a higher adequacy rate in the S- group (2609/2987, 87.35 %) than in the S+ group (2135/2504, 85.26 %), but the difference was not statistically significant [pooled OR 0.94 (95 % CI 0.79–1.11), p = 0.45] (Fig. 4a). The χ^2 and I^2 were 9.02 (p = 0.17) and 33 %, respectively, indicating no significant heterogeneity among the studies with respect to the rate of adequacy of specimens.

Among the included studies, one study [15] dealt with only histopathological adequacy of specimens instead of cytological adequacy of specimens. Therefore, cytological adequacy of specimens was examined in six studies [2, 11– 14, 20]; five studies [2, 11, 12, 14, 20] used a standard endhole 22 gauge needle and the other one study [13] used a standard end-hole 22 or 25 gauge needle. The subgroup analysis of the six studies that dealt with only cytological adequacy of specimens showed similar results with no statistically significant difference in the rate of adequacy of specimens [pooled OR 0.92 (95 % CI 0.76–1.12), p = 0.42]. The χ^2 and I^2 were 8.93 (p = 0.11) and 44 %, respectively, indicating no significant heterogeneity among the studies with respect to cytological adequacy of specimens (Fig. 4b).

Cellularity

Four studies [2, 12, 13, 17] reported cellularity from EUS-FNA with or without the stylet. The pooled analysis using the fixed-effects model demonstrated no significant differences between the S+ (305/837, 36.44 %) and Sgroups (311/853, 36.46 %) regarding the rate of cellularity > 50 % [pooled OR 1.01 (95 % CI 0.80–1.27), p = 0.92) (Fig. 5a). No heterogeneity was identified among the studies (p = 0.89, $I^2 = 0$ %).

Among four included studies [2, 12, 13, 15], one study [15] used a different definition for cellularity. Thus, cellularity using the same definition was examined in three studies [2, 12, 13]. The subgroup analysis of three studies for the rate of cellularity > 50 % using the same definition showed similar results, with no statistically significant difference between the S+ (143/617, 23.18 %) and S- groups (151/633, 23.85 %) [pooled OR 1.00 (95 % CI

Table 1 C	haracterist	tics of include	sd studies									
Study	Country	Recruitment	t time	Study design	No. of 1 patients/ lesions	Used endoscope	No. of patients 1 of each group 6 (S+ vs. S-) (Vo. of specimens of each group S+ vs. S-)	No. No. of end center	of osonographer	No. of cytologists	ROSE
Abe et al. [15, 19]	Japan	2012.8-2013.		Prospective single-blind crossover RCT	107/110	GF-UC240P-AL5, GF- UCT240-AL5, or GF- UCT260; Olympus	107 versus 107 2	20 versus 220	5 18		2V	Yes
Sandeep (2014)	India	2013.10-201	4. 7	Prospective, RCT	115/128	GF-UCT180; Olympus	NA (6 versus 62	1 1		1	No
Gimeno- Garcia et al. [14]	Canada	2000.5-2010	.10	Retrospective	3078/3364	GF-UC30P; Olympus EG-3870UTK; Pentax	NA	483 versus 1881	-		_	Yes
Wani et al. [12]	USA	2010.12–201	1.9	Prospective, RCT	100/100	GFUC140P or GF- UCT140; Olympus	NA	175 versus 275	4		ε	Yes
Rastogi et al. [2]	USA	2009.8–2010	.5	Prospective, RCT	101/118 d	GF-UC140P or GF- UCT140; Olympus 3G-3630U; Pentax	NA	36 versus 236	2 2		4	Yes
Wani et al. [13]	USA	S+ group 2006.1–200 S- group 2007.10–20)7.9 and)09.4	Retrospective	162/228	GF-UC30P or GF- UCT30; Olympus	72 versus 90	06 versus 122	1 2		7	Yes
Sahai et al. [11]	Canada	2004.1–2004.	ė	Prospective, RCT	111/135	AN	NA	18 versus 191	1 1		1	Yes
Used needl	e Su apl (S-	ction plication + vs. S-)	Age (yea or mean (S-)	urs) [mean ± SD (range)] (S+ vs.	Sex $(n, \%)$ c male $(S + vs$ S -)	of Site of lesions . (pancreas/LN/hollow viscus/other)	No. of passes (mean ± SD) (S vs. S-)	Size of lesion (ran (ran S-)	$r [mean \pm SD]$ ge)] (S+ vs.	Major outcol	mes	
22-gauge Fr (Boston Scientific)	NA Ye	ss versus yes	67.6 (26-7) (26-87)	-87) versus 67.6	52 (47.2) ver 52 (47.2)	sus Overall, 66/23/12/9	2 versus 2	20.8 (8.0–53.((8.0–53.0))) versus 20.8	Acquisition ra and sufficie	ate of an appr nt specimen	opriate
22 gauge Fr (Cook Medical)	NA Ye	s versus yes	51.9 ± 12 $56.8 \pm$	2.5 versus 16.8	40 (60.6) ver 41 (66)	sus 23/33/2/8 versus 20/3 2/6	4/ 2.16 \pm 0.94 vers 2.34 \pm 1.06	us 24.6 ± 14.2 v 22.8 ± 15.1	/ersus	Adequacy of blood numb diagnostic y	specimen Am ber of passes yield	ount of
22 gauge Fr (Cook Medical)	NA Ye	s versus yes	61.4 (33– (32–83)	85) versus 60.7	833 (56.2) versus 1018 (54.1)	604/688/NA/191 vers 684/890/NA/307	us 2.21 ± 1.22 vers 2.06 ± 1.22	us Overall, 30 (5	(08-2	Adequacy of passes diagr	specimen Nur nostic yield	nber of
22 or 25 gal FNA (Coc Medical)	uge Ye Jk	s versus yes	Overall, 6	55.5 ± 10.9	Overall, 54 (54 %)	Overall, 58/25/9/8	NA	Overall, 29.5	± 21.8	Diagnostic yi Degree of cel Degree of add Degree of con	eld Ilularity equacy ntamination	
										Degree of am	nount of bloou	

Table 1 continu	ed						
Used needle	Suction application (S+ vs. S-)	Age (years) [mean \pm SD or mean (range)] (S+ vs. S-)	Sex (<i>n</i> , %) of male (S+ vs. S-)	Site of lesions (pancreas/LN/hollow viscus/other)	No. of passes (mean ± SD) (S+ vs. S−)	Size of lesion [mean \pm SD or mean (range)] (S+ vs. S-)	Major outcomes
22 gauge FNA (Cook Medical)	Yes versus yes	Overall, 63.1 ± 13.8	Overall, 70 (69 %)	Overall, 61/31/8/18	2 versus 2	Overall, 27.5 ± 18.0	Diagnostic yield Degree of cellularity Degree of adequacy Degree of contamination Degree of amount of blood
22 gauge FNA (Cook Medical)	Yes versus yes	65.1 (43–89) versus 64.4 (33–88)	71 (98.6 %) versus 87 (96.6 %)	14/62/NA/30 versus 27/63/NA/32	Overall, 3.7 ± 1.3	25.2 (2-70) versus 23.2 (5-120)	Diagnostic yield Degree of cellularity Degree of adequacy Degree of contamination Degree of amount of blood
22 gauge FNA (Cook Medical)	Yes versus yes	Overall, 62.9 (30–86)	NA	NA	Overall, 2.3 (1-6)	NA	Diagnostic yield
No. number, RO.	SE rapid on-site e	valuation, RCT randomized o	controlled trial, N	A not available, SD stan	dard deviation, LN ly	mph node, FNA fine needle a	spiration

0.76–1.31), p = 0.98]. No heterogeneity was identified among the studies (p = 0.75, $l^2 = 0$ %) (Fig. 5b).

Contamination

The contamination rate was reported in four studies [2, 12, 13, 15]. Pooled analysis using a fixed-effects model demonstrated no significant differences between the groups regarding the contamination rate [S+ group (20/837, 2.39 %) vs. S- group (26/853, 3.04 %): OR 0.79 (95 % CI 0.44–1.44), p = 0.45] (Fig. 6). There was no heterogeneity among the studies (p = 0.63, $l^2 = 0$ %).

Blood Contamination

Five studies [2, 11–13, 20] reported the blood contamination rate ("3," significant), including 801 specimens in the S+ group and 886 specimens in the S- group. In a pooled analysis with a random-effects model, the blood contamination risk was slightly higher in the S+ group (249/801, 31.09 %) than in the S- group (270/886, 30.47 %), but this difference was not statistically significant [OR 1.19 (95 % CI 0.75–1.90), p = 0.46] (Fig. 7). The χ^2 and I^2 were 15.35 (p = 0.004) and 74 %, respectively, which indicated significant heterogeneity among the studies with respect to the blood contamination rate.

Discussion

We investigated seven studies to determine whether the use of the stylet during EUS-FNA was associated with clinical outcomes, such as sample adequacy or diagnostic yield. Although the use of a stylet has been found to improve the quality of specimens obtained by preventing blockage by the intestinal mucosa, thereby reducing contamination and allowing for an improved overall diagnostic yield of EUS-FNA [4, 11, 21], there is no consensus regarding the role of the stylet, and few studies with large numbers of patients have compared the endoscopic and cytopathological outcomes with or without a stylet. To date, most endosonographers tend to use a stylet during EUS-FNA based on the unsubstantiated premise that the use of a stylet improves the quality of the specimen [22]. Contrary to this premise, recent data suggests that the use of a stylet does not guarantee any advantage with regard to the diagnostic yield of malignancy or to the quality of specimens obtained [2, 11, 12]. Thus, this meta-analysis is meaningful because we were the first to compare important cytopathological parameters such as adequacy, amount of blood, and cellularity among specimens obtained with and without a stylet during EUS-FNA.

In terms of the adequacy of specimens, our pooled data indicates that there were no significant differences between





Fig. 3 Risk of bias summary: review authors' judgements about each risk of bias item for each included study

the S+ and S- groups, although only one study [11] among the seven reported that the rate of adequate specimens without the stylet was higher than with the stylet during EUS–FNA. In this prospective analysis, contrary to

our results, Sahai et al. found that sample adequacy was lower and bloodiness higher in the S+ group, suggesting that the quality of samples obtained with a stylet may actually be inferior. Our lack of significant results between the S+ and S- groups can be attributed to several factors: not all studies provided an explicit definition of adequacy of specimen; therefore, the results may be influenced to a greater or lesser extent. Also, there may be a discrepancy between the method of tissue sampling (the aim of one study [15] was histological adequacy and that of the other six studies [2, 11–14, 20] was cytological adequacy). Hence, it can be difficult to provide a precise assessment of the adequacy of specimens, although we conducted subgroup analysis for studies with only cytological outcomes after excluding one study [15] regarding histological adequacy and the results of the subgroup analysis was similar to that of all seven studies.

Another concern when using a stylet to prevent blockage or contamination by the intestinal mucosa and allow more adequate aspiration is the qualitative improvement of cellularity. Among the RCTs analyzed in our review, cellularity was defined differently (one study [15] used the definition of cellularity as a scoring system from 0 to 5, while the other three studies used percentage area of slides that contained cells of the representative lesion from 0 to 3); we did not find any differences in the cellularity in either the S+ or S- groups.

Although there is limited data comparing the use of EUS-FNA with or without a stylet, none of the included

Table 2Newcastle–OttawaScale scores for non-
randomized studies

Author group	Sele	ction			Compa	arability	Outcome			Total
	1	2	3	4	5A	5B	6	7	8	
Gimeno-Garcia et al. [14]	☆	ራ			☆	\$	Δ	አ	አ	7
Wani et al. [12]	አ	አ			Δ	Δ	\$	Δ	Δ	7

1 =case definition adequate?, 2 = representativeness of the cases, 3 = selection of controls, 4 = definition of controls, 5A = comparability of cases and controls, main factor: age/sex, 5B = comparability of cases and controls, secondary factor, 6 = ascertainment of exposure, 7 = same method of ascertainment for cases and controls, 8 = non-response rate

Α	\$ +		S			Odds Ratio		Odds Ratio
Study or Subaroup	Events	Total	5 - Events	Total	Weight	M-H Fixed 95% CL	Үеаг	M-H Fixed 95% Cl
Sahai 2010	89	118	166	191	11 7%	0.46 (0.26, 0.84)	2010	
Wani 2011	100	106	111	122	2.2%	1.65 [0.59, 4.63]	2011	
Rastogi 2011	133	236	146	236	24.0%	0.80 (0.55, 1.15)	2011	
Wani 2012	197	275	188	275	20.1%	1.17 [0.81, 1.68]	2012	
Gimeno-Garcia 2013	1436	1483	1821	1881	19.1%	1.01 [0.68, 1.48]	2013	-+-
Sandeep 2014	59	66	55	62	2.3%	1.07 [0.35, 3.26]	2014	
Abe 2015	121	220	122	220	20.7%	0.98 [0.67, 1.43]	2015	-+-
Total (95% CI)		2504		2987	100.0%	0.94 [0.79, 1.11]		•
Total events	2135		2609					
Heterogeneity: Chi ² = 9. Test for overall effect: Z	02, df = 6 = 0.76 (P	(P = 0. = 0.45	.17); I² = :)	33%				0.01 0.1 1 10 100 Favours [S -] Favours [S +]
В	c .		c			Odda Datia		Odda Datio
Study or Subaroup	5 T Evonte	Total	S - Evonte	Total	Mojaht		Voar	
Sabai 2010	00	110	166	101	1/ 0%		2010	
3anai 2010 Wani 2011	100	106	111	127	2 9 96	1 65 (0.20, 0.04)	2010	
Rastogi 2011	133	236	146	236	30.2%	0.80 (0.55, 1.15)	2011	
Wani 2012	197	275	188	275	25.3%	1.17 [0.81, 1.68]	2012	 _
Gimeno-Garcia 2013	1436	1483	1821	1881	24.1%	1.01 [0.68, 1.48]	2013	_ _
Sandeep 2014	59	66	55	62	2.9%	1.07 [0.35, 3.26]	2014	
Total (95% CI)		2284		2767	100.0%	0.92 [0.76, 1.12]		•
Total events	2014		2487					
Heterogeneity: Chi ² = 8.	93, df = 5	(P = 0.	11); I ² = 4	44%				
Test for overall effect: Z	= 0.80 (P	= 0.42)					
								Favours (S-) Favours (S+)

Fig. 4 a Forest plots for adequate specimens, b forest plots for cytological adequacy of specimen (subgroups analysis). S+: with stylet, S-: without stylet, CI: confidence interval

Δ								
	S +		S -			Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	Year	M-H, Fixed, 95% Cl
Rastogi 2011	36	236	37	236	21.4%	0.97 [0.59, 1.59]	2011	-+-
Wani 2011	53	106	56	122	17.8%	1.18 [0.70, 1.98]	2011	
Wani 2012	54	275	58	275	31.9%	0.91 [0.60, 1.38]	2012	
Abe 2015	162	220	160	220	28.9%	1.05 [0.69, 1.60]	2015	
Total (95% CI)		837		853	100.0%	1.01 [0.80, 1.27]		•
Total events	305		311					
Heterogeneity: Chi ² =	0.61, df=	3 (P =	0.89); l ² =	= 0%				
Test for overall effect:	Z = 0.10	(P = 0.9	32)					0.01 0.1 1 10 100 Eavoure [S_1] Eavoure [S+1]
B								
D	S +		S -			Odds Ratio		Odds Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Fixed, 95% Cl	Year	M-H, Fixed, 95% Cl
Rastogi 2011	36	236	37	236	30.1%	0.97 [0.59, 1.59]	2011	
Wani 2011	53	106	56	122	25.0%	1.18 [0.70, 1.98]	2011	
Wani 2012	54	275	58	275	44.8%	0.91 [0.60, 1.38]	2012	
Total (95% CI)		617		633	100.0%	1.00 [0.76, 1.31]		•
Total events	143		151					
Heterogeneity: Chi ² =	0.58, df=	2 (P =	0.75); l ² =	= 0%				
Test for overall effect:	Z = 0.02	(P = 0.9	98)					

Fig. 5 a Forest plots for cellularity, **b** Forest plots for cellularity using the same definition (subgroups analysis). S+: with stylet, S-: without stylet, CI: confidence interval

studies in our review has shown any differences between the groups. In one retrospective analysis of four studies that reported cellularity, Wani et al. [12] compared EUS-FNA with and without a stylet for 228 lesions. The authors did not find a significant difference in cellularity [15]. Similarly, the same group also performed a similar prospective RCT [13] with 550 lesions to compare the performance of EUS-FNA with or without a stylet and concluded that there was no significant difference between the S+ and S- groups for cellularity. This lack of a significant difference may be attributed to a certain amount of subjectivity in their assessment by cytopathologists, although predefined criteria were used to compare the cytologic characteristics of the specimens and the cytopathologists were blinded to



Fig. 6 Forest plots for contamination. S+: with stylet, S-: without stylet, CI: confidence interval



Fig. 7 Forest plots for blood contamination. S+: with stylet, S-: without stylet, CI: confidence interval

the technique used to obtain the specimens in order to obviate any bias. Also, most studies conducted by two or more cytopathologists did not assess the interobserver agreement among cytopathologists in the assessment of EUS-FNA specimens. Variability in the diagnosis of malignancy among pathologists and its impact on the results cannot be excluded.

Additionally, our review revealed that contamination, including blood contamination, in the S+ group tended to be higher than in the S- group, though not at a level reaching significance, similar to the differences between both groups for other outcomes. Theoretically, the use of a stylet will prevent the tip of the needle from being blocked by a plug of digestive wall tissue and/or contamination of the sample before entering the FNA target lesion, thereby producing improved overall outcomes. In clinical practice, Sahai et al. [11] initially reported the effect of the stylet on EUS-FNA in a prospective study in which a single endosonographer performed EUS-FNA of 135 lesions in 111 patients with a 22-gauge needle. However, in that study, the proportion of bloody samples was higher in the S+ group than in the S- group (75 vs. 52 %, p < 0.0001), suggesting that the use of the stylet during EUS-FNA could hinder cytological diagnosis. Similarly, our results suggest that this theory is invalid. Use of the stylet did not change the contamination rate or the frequency of bloody and/or inadequate samples; however, this observation may be attributed to type II error, as none of these studies are powered sufficiently to detect a true benefit or the interobserver disagreement among cytopathologists in the assessment of EUS-FNA specimens.

Furthermore, the lack of differences between all the endpoints assessed (quality of specimen and cytological characteristics) may be related to different lesion sites because the hardness of each lesion was different. In other words, the included studies did not use the same lesion to compare the two groups. For example, the tissue of the lymph node is softer and has less fibrotic tissue than pancreatic malignancies (especially pancreatic ductal adenocarcinoma). Finally, although suction for negative pressure, which may have increased the amount of blood, was applied equally in all patients of both groups, suction effects can in no way be negligible.

Although we included relevant studies in our review, several limitations of our study need to be considered, as in any meta-analysis. Firstly, we did not conduct publication bias, which may distort the outcome of the overall analysis. Obviously, it did not guarantee that there was no publication bias in our meta-analysis, although Cochrane library may not recommend publication bias when the number of included studies is < 10. The use of different needle types (in three studies [2, 12, 20], a standard 22-gauge needle by Cook Medical was used; in one study [13], a standard 22- or 25-gauge needle by Cook Medical was used; in one study [15], a standard 22-gauge needle by Boston Scientific was used; in the other two studies [11, 14], a standard 22-gauge needle from an unknown company was used) may contribute to heterogeneity that could affect the results. Therefore, these results should be interpreted with caution, as the aim and

performance of each needle were not uniform. In addition, not all of the studies provided an explicit definition or criteria for outcomes, especially for the definition of specimen adequacy, cellularity, contamination, or bloodiness. This might have influenced our results to a certain extent. Finally, one of the main goals of our study was to improve the quality of specimens, but the available studies had limited information regarding histological adequacy (in only one study) beyond cytological adequacy.

In conclusion, this meta-analysis exploring the role of the stylet during EUS-FNA procedures shows that the practice of using a stylet during EUS-FNA is questionable as to whether it improves the quality of the specimen obtained. Also, the results of our review challenge the current belief that the use of a stylet during EUS-FNA helps prevent clogging of the lumen of the needle by gut wall tissue. In other words, the use of a stylet during EUS-FNA does not appear to confer any advantage with regards to the adequacy of specimen or diagnostic yield. Echoendoscopists should be aware that it would be reasonable to not use a stylet during EUS-FNA, potentially making the procedure easier and less labor intensive, as well as more time- and cost-efficient.

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Compliance with ethical standards

Conflict of interest All authors have no conflicts of interest or financial ties to disclose. Furthermore, this report is a meta-analysis. The author states that the report includes every item in the PRISMA checklist for meta-analysis clinical studies. Also, this manuscript was screened for plagiarism using Turnitin (www.turnitin.com).

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