

The efficiency of mutual funds

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Abstract This paper analyzes the short-term market efficiency of the mutual fund industry around the world. Using a unique database of worldwide domestic equity funds, it employs a parametric (regression model) and non-parametric (data envelopment analysis (DEA) model) approaches to establish a relation between cost (expense ratio, turnover, loads, and risk) and benefit (return) of mutual funds. The empirical results of the parametric approach show a statistically significant negative relationship between expenses and risk-adjusted performance across countries. When we reexamine this relationship using a non-parametric approach, we show, in contrast to our previous result, a positive relationship between expenses and risk-adjusted performance. Thus, using the DEA methodology, we find strong evidence that equity mutual funds around the world are approximately mean–variance efficient.

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1 Introduction

The assessment of the performance of mutual funds has been the normal procedure to choose where to invest and many papers have explained that historical performance record is the main criteria for fund selection (e.g. [Bergstresser and Poterba 2002](#); [Deaves 2004](#); [Busse and Irvine 2006](#)). The question whether mutual funds exhibit performance persistence has received considerable attention in the academic finance literature during the last two decades (e.g., [Hendricks et al. 1993](#); [Carhart 1997](#); [Bollen and Busse 2005](#), and [Vidal-García and Vidal 2016](#)). These studies examine whether it is possible to forecast future returns based on past performance. This is an important question for the mutual fund industry as if past performance has no forecasting power for future performance, investors would not benefit of data collection. However, the mutual fund industry is growing at a fast pace and expanding to more markets. Leading mutual fund companies such as Lipper and Morningstar publish mutual fund rankings that receive coverage from the press around the world.

Many studies have examined mutual funds efficiency. The most common measures of portfolio performance are Jensen's alpha, the Sharpe ratio and the Treynor ratio. However, there are several problems when using these measures that have been thoroughly explained in the literature. [Treynor \(1965\)](#), [Sharpe \(1966\)](#), [Jensen \(1968, 1969\)](#) create portfolio evaluation models that are derived from the capital asset pricing model (CAPM). [Roll \(1978\)](#), [Reilly and Akhtar \(1995\)](#), and [Grinblatt and Titman \(1994\)](#) explain that capital asset pricing models could be sensitive to the selection of the benchmark when examining fund efficiency. They point out that these performance measures are statistically biased against market timing ability. [Lehmann and Modest \(1987\)](#) explain that mutual fund performance evaluation can vary with the selection of benchmarks. They show statistically significant abnormal performance using several benchmarks. [Matallín-Saez \(2007\)](#) examines the difference between factors and benchmarks for portfolio performance evaluation. The author shows that there are similar biases regardless of using either factors or benchmarks and that the selection of an appropriate benchmark is more important than the multifactor model employed.

There are a number of papers that examine mutual fund performance and its relation with fees. Mutual fund fees cover the service offered by funds. They should be influenced by fund performance as the principal service offered by mutual funds is portfolio management. However, the mutual fund literature shows that high fees are the main reason of equity funds underperformance as fund performance improves when before-fee returns are used (see, [Grinblatt and Titman 1989](#); [Malkiel 1995](#); [Gruber 1996](#); [Droms and Walker 1996](#)). [Carhart \(1997\)](#), for instance, shows that net returns are negatively correlated with fees, and that these fees are much larger for actively managed mutual funds than for passive ones. [Gil-Bazo and Ruiz-Verdú \(2009, hereafter GBRV\)](#) find that funds with worse before-fee performance charge higher fees. The authors focus on the link between before-fee performance and fees and assess whether variations in fund fees explain the variations in performance. They find a negative relation between fees and before-fee performance for US equity mutual funds. Their results confirm earlier findings of [Gruber \(1996\)](#), who showed that high fees are related to fund underperformance instead of being related to the ability to outperform the market.

Supporting this evidence, [Berkowitz and Kotowitz \(2002\)](#) find a negative relation between fees and performance for low-quality managers. [Vidal et al. \(2015\)](#) show robust evidence of forecasting power for mutual fund fees. They explain that funds showing either a positive or negative relation with their fees present significant evidence of a negative return predictability using fees as forecasting variable. [Vidal-García and Vidal \(2015\)](#) find that only 4% of US funds reduce their fees in case of poor performance. The authors define underperformance as negative alpha estimates in three consecutive years. Thus, the percentage of fees charged by mutual funds is a relevant aspect in performance valuation ([Ramos 2009](#); [Khorana et al. 2009](#)).

Several additional studies examine the relation between performance-sensitive investors and mutual fund fees. [Christoffersen and Musto \(2002\)](#) explain that fees are determined based on the elasticity of the demand and that funds with less elastic demand incur higher fees. They argue that performance-sensitive investors sell their shares in the fund after a bad performance, thus funds presenting worse past performance attract less investors. Similarly, [Gil-Bazo and Ruiz-Verdú \(2008\)](#) explain that top-performing funds compete for sophisticated investors (performance-sensitive investors), which reduces their fees. Underperforming funds leave, however, that portion of the market, hence attracting unsophisticated investors who are willing to pay higher fees.

An interesting debate in the mutual fund industry is the differential effect of active and passive management on fund performance. Indexed equity funds have lower fees compared with active funds. [Basak and Pavlova \(2013\)](#) explain that active fund managers are concerned about their performance in comparison to benchmark indices, thus increasing the percentage of stocks within their benchmark indices to reduce the risk of underperformance. However, [Cremers et al. \(2016\)](#) explain that investors in active funds benefit from the presence of passive funds. They point out that active funds incur lower fees and include higher active stocks in markets with more explicit indexing. In comparison, active funds increase their fees in markets with more closet indexing. Thus, they conclude that the presence of explicit indexing increases the level of competition in a fund market and closet indexing indicates the opposite.

Management fees are the main component of mutual fund expenses representing around 90% of the total expenses. These fees are charged to investors for portfolio management. Management fees usually represent a fixed percentage of total assets under management. In this sense, managers are compensated for asset growth instead of for performance. Fund management companies do not seem to be willing to establish performance-based fee funds as these fee structures are only suitable for qualified investors. Furthermore, the fund industry has an important lack of competition and is mainly dominated by banks, which allows fund companies to charge fees not linked to performance. Following the economic efficiency theory, fund managers showing performance ability should be remunerated for the costs involved in information acquisition and the costs of trading (see, [Grossman and Stiglitz 1980](#)). Otherwise there would be no incentive for becoming informed. Thus, funds offering better services to investors should charge higher fees to cover their information gathering role, leading to a positive relationship between fund expenses and risk-adjusted fund returns before expenses (see, [Ippolito 1989](#); [Díaz-Mendoza et al. 2014](#)). The results of [GBRV \(2009\)](#) contrast this argument and point out that the negative link between fund expenses and fund returns does not exist for the funds that show better governance, which charge fees in relation to performance. In line with the result of [GBRV \(2009\)](#) is the study of [Berkowitz and Kotowitz \(2002\)](#), who show a positive relation between fees and performance for the best managers.

In this study, we examine the relation between expenses and performance using parametric and non-parametric methodologies for all countries around the world with available data. Our paper aims to provide insightful results on the global efficiency of mutual fund performance, which has not been studied so far. We evaluate the market efficiency of the

mutual fund industry around the world in performance using a sample of domestic mutual funds for 35 countries for the period 1990–2015 that includes daily returns of 16,085 actively managed equity mutual funds. Our objective is to test whether mutual funds are efficient, employing different measurement methodologies. We use the one-factor CAPM, the [Fama and French \(1993\)](#) three factor model and the [Carhart's \(1997\)](#) four-factor model. We also use the data envelopment analysis (DEA) non-parametric methodology to examine the efficiency of mutual funds.

The paper provides new interesting results. Using the four-factor Carhart model, we show that premiums are positive, suggesting that more risky, small, value-focused, and previous-winner stocks achieve higher returns. We also examine fund performance by presenting the percentage of positive and negative values of performance measures and find that less than half of the funds exhibit positive values of performance. The estimation of the performance-expenses relation shows that the coefficients of expenses are always significantly negative for all countries and all risk-adjusted performance measures. Finally, we present the percentage of efficient funds for every variation of the DEA model using gross returns, and alphas from CAPM, Fama–French and Carhart models as output measures. The results show that most of the funds are efficient.

We make several contributions to the literature. First, we find a statistically significant negative relationship between expenses and risk-adjusted performance across countries, indicating that higher expenses are likely to lead to bad performance. Second, the use of DEA shows, in contrast to our previous result, strong evidences that equity mutual funds around the world are almost mean–variance efficient. Thus, DEA confirms the mean–variance efficiency theory. Finally, we depict the areas of operational inefficiency that improve performance of mutual funds. Third, our research would allow us to test the agency theory and the potential agency conflict between managers and investors of mutual funds. As [Jensen and Meckling \(1976\)](#) explain, assuming that both mutual fund managers (the agents) and investors (the principals) focus on maximizing their individual utility, it is possible that the fund managers will not take decisions in the best interest of the fund investors. With our results, we provide evidence whether the manager's and the investor's wealth increase at the same time or if one increases at the expense of the other.

The rest of the paper is organized as follows. Section 2 describes the data, details the construction of the variables used in the analysis and presents descriptive statistics for the sample. Section 3 presents the parametric and non-parametric methodologies used to estimate the results. Section 4 provides the empirical results. Section 5 concludes.

2 Data and summary statistics

2.1 Data

Our sample of mutual funds includes daily returns of 16,085 actively managed equity mutual funds. The funds are domiciled in 35 countries around the world from Asia-Pacific (Australia, China, Hong Kong, India, Indonesia, Japan, Malaysia, New Zealand, Singapore, South Korea, Taiwan, and Thailand), Europe (Austria, Belgium, Denmark, Finland, France, Germany, Italy, Netherlands, Norway, Poland, Portugal, Spain, Sweden, Switzerland, United Kingdom, Ireland, and Luxembourg), North America (Canada and United States), and the rest of the world (Brazil, Chile, Israel, and South Africa). All returns are expressed in local currencies and include dividends. We only include the primary share, as some funds present multiple share classes and might have multiple observations. Fund returns incorporate management

and distribution fees, but not sales loads (fee paid when shares are purchased or sold). Our time period is a 26-year interval from January 1st, 1990 to December 31st, 2015. Our final sample of global funds includes over 90% of the total net assets of equity funds around the world as of December 2015 (see, Investment Company Institute (ICI) 2015 aggregate statistics). We download data from Morningstar Direct database. Morningstar is a leading provider of mutual fund data, which includes a global coverage for a large number of fund variables. We consider time-series observations on total net assets, turnover, funds' fees, loads, and fund age.

Our sample includes the most important countries in terms of world market capitalization. It is similar to other papers using global mutual funds data (see, for instance, Ferreira et al. 2012). We eliminate some types of funds from our database, namely, index funds, sector funds, bond and money market funds, funds investing internationally, and funds that invest in financial instruments such as convertible debt. We also apply some filters to the fund return data. First, we limit our sample to open-end domestic equity funds to consider only funds that invest in the same country. Second, we only include equity funds with 24 months of returns, as a sufficient return data period is necessary for multifactor regressions.

To the best of our knowledge, our sample is the largest sample currently available for daily mutual fund returns. We incorporate funds that come into existence at any point during the sample period, making our data not limited and reducing the extent of the selection bias. To address the problem of survivorship bias in our sample, which is the result of including in a sample only surviving funds (see, Elton et al. 1996a, b; Carhart 1997), we incorporate all funds in our sample until they disappear.

2.2 Variable construction

In this section, we describe the methodology employed to design the regression models for each country. We create a national version of the multifactor models. For U.S. funds, we download the factors from Fama and French (1993) website.¹ For the rest of the countries, we create daily factors, implementing the methodology of Fama and French (1993). For this purpose, we use all stocks from the Worldscope database (Thomson Financial Company) available in each country. This database covers above 98% of total market capitalization on a country level. Following Fama and French, we create 6 value-weighted portfolios formed on size and book-to-market. The market excess return is estimated as the value-weighted return of all stocks in the Worldscope database per country minus the Treasury bill rate for that country. We obtain a daily Treasury bill rate dividing the 1-month Treasury bill return by the number of days in the month. The Small minus Big (SMB) factor is the average return on the three small portfolios minus the average return on the three big portfolios. High minus Low (HML) factor is the average return on the two value portfolios minus the average return on the two growth portfolios. We also create a daily version of the momentum factor of Carhart (1997). We use six value-weighted portfolios formed on size and prior returns. The momentum factor is estimated as the average return on the two high prior return portfolios minus the average return on the two low prior return portfolios.

2.3 Descriptive statistics

Table 1 shows the descriptive statistics of the sample on a country-by-country basis. As explained previously, we have selected the countries with the largest market capitalization. The table presents the sample of funds as obtained from the Morningstar database. The second column of Panel A (Table 1) shows the number of mutual funds from each country. The third

¹ The US factors are drawn from French's website: <http://mba.tuck.dartmouth.edu/pages/faculty/ken.french/>.

Table 1 Summary statistics

Country	Number of funds	Raw return (%)	TNA (\$ million)	Fund age (years)
<i>Panel A: Descriptive statistics</i>				
Australia	964	0.803	238	13
Austria	20	−0.039	167	11
Belgium	30	0.032	98	17
Brazil	97	1.007	117	9
Canada	990	0.141	407	10
Chile	57	−0.198	51	9
China	578	−0.572	347	4
Denmark	63	0.378	114	12
Finland	49	0.063	158	13
France	975	0.129	87	11
Germany	80	−0.098	629	16
Hong Kong	14	−0.810	561	13
India	479	0.844	99	8
Indonesia	109	0.281	112	5
Ireland	33	0.278	101	4
Israel	163	0.063	13	12
Italy	66	0.159	145	15
Japan	1000	−1.023	95	12
Korea (South)	1053	−0.127	85	6
Luxembourg	17	−0.333	498	14
Malaysia	203	0.101	85	13
Netherlands	28	−0.355	273	15
New Zealand	20	0.412	154	9
Norway	109	0.021	228	15
Poland	78	−0.810	114	7
Portugal	23	−0.139	18	15
Singapore	19	−0.467	146	14
South Africa	270	−0.483	148	9
Spain	145	−0.446	43	16
Sweden	191	−1.077	455	11
Switzerland	197	−0.774	300	8
Taiwan	247	0.649	57	14
Thailand	237	0.371	66	11
United Kingdom	982	−0.306	797	13
United States	6501	0.146	2327	12
All countries	16,085	−0.062	267	11

Table 1 continued

	μ		σ		S		K	
	Daily	Monthly (%)	Daily (%)	Monthly (%)	Daily (%)	Monthly	Daily	Monthly
<i>Panel B: Daily and monthly statistics</i>								
Australia	0.803	1.198	0.236	1.817	-1.345	-1.094	38.453	7.234
Austria	-0.039	-0.044	0.108	0.831	-1.002	-0.945	34.234	6.765
Belgium	0.032	0.048	0.289	2.225	-1.435	-1.234	38.442	7.432
Brazil	1.007	1.502	0.805	6.198	-2.567	-2.112	45.112	9.342
Canada	0.141	0.210	0.393	3.026	-1.744	-1.234	46.123	8.423
Chile	-0.198	-0.123	0.551	4.242	-1.877	-1.466	48.111	7.232
China	-0.572	-0.334	0.593	4.566	-1.934	-1.645	42.341	6.232
Denmark	0.378	0.564	1.052	8.100	-2.667	-2.277	43.234	6.322
Finland	0.063	0.094	0.175	1.347	-0.674	-0.511	28.332	6.322
France	0.129	0.185	0.359	2.765	-1.667	-1.206	42.671	8.331
Germany	-0.098	-0.017	0.654	3.023	-2.208	-1.576	39.567	7.673
Hong Kong	-0.810	-0.657	1.222	9.409	-3.452	-2.944	49.342	4.622
India	0.844	1.123	1.351	7.402	-2.453	-2.123	52.345	3.954
Indonesia	0.281	0.324	0.782	6.021	-2.456	-2.034	49.343	6.421
Ireland	0.278	0.322	0.774	5.959	-2.156	-1.786	51.343	5.753
Israel	0.063	0.234	0.175	1.347	-0.865	-0.523	38.453	4.975
Italy	0.159	0.328	0.765	3.234	-2.345	-1.667	47.334	5.345
Japan	-1.023	-0.975	1.023	7.877	-3.456	-2.865	49.453	5.322
Korea (South)	-0.127	-0.078	0.827	6.367	-2.345	-1.897	53.232	6.453
Luxembourg	-0.333	-0.231	0.927	7.137	-2.277	-1.867	52.566	5.776
Malaysia	0.101	0.211	0.281	2.163	-1.445	-1.112	46.453	6.422
Netherlands	-0.355	-0.134	0.815	4.345	-2.345	-1.094	48.445	5.678
New Zealand	0.412	0.745	1.147	8.8315	-3.456	-2.987	54.342	5.432
Norway	0.021	0.064	0.358	2.756	-1.556	-1.123	41.233	5.673
Poland	-0.810	-0.654	1.256	9.671	-2.451	-1.956	45.533	6.422
Portugal	-0.139	-0.321	0.387	2.979	-1.776	-1.567	43.233	5.422
Singapore	-0.467	-0.233	1.312	8.102	-2.456	-2.034	49.522	6.411
South Africa	-0.483	-0.211	1.345	8.356	-2.311	-1.945	51.234	5.422
Spain	-0.446	-0.208	1.242	9.563	-2.045	-1.945	47.666	6.542
Sweden	-1.077	-0.544	1.123	8.647	-3.324	-2.456	52.345	3.854
Switzerland	-0.774	-0.345	1.155	8.893	-3.567	-3.123	54.299	3.976
Taiwan	0.649	0.111	0.807	6.213	-2.867	-2.455	45.232	4.123

Table 1 continued

	μ		σ		S		K	
	Daily	Monthly (%)	Daily (%)	Monthly (%)	Daily (%)	Monthly	Daily	Monthly
Thailand	0.371	0.723	1.033	7.954	-3.567	-3.112	46.663	4.563
United Kingdom	-0.306	-0.567	0.896	4.967	-2.765	-1.345	49.566	5.667
United States	0.146	0.222	0.408	3.141	-1.967	-1.765	45.211	4.012

Panel A presents descriptive statistics on our sample of mutual funds. The sample period is from January-1990 to December-2015. *Number of funds* presents the number of actively managed equity mutual funds in each country. *Raw return* presents the average daily fund return over the entire period of study. *TNA* is the total net assets under management and is expressed in millions of dollars. *Fund age* is the average number of years since the fund was created

Panel B shows summary statistics of the fund return distributions. The mean (μ) and standard deviation (σ) are sample estimates. Skewness (S) is estimated as $S = \frac{1}{\sigma^3 T} \sum_{t=1}^T (R_t - \mu)^3$ and excess kurtosis (K) is estimated as $K = \frac{1}{\sigma^4 T} \sum_{t=1}^T (R_t - \mu)^4 - 3$

column presents the mean raw returns, whereas the fourth column reports the mean asset values and the fifth column shows the average number of years since the fund was created The US presents the largest amount of funds with 6501, and Hong Kong shows the smallest number with only 14 funds. Raw returns fluctuate significantly between countries, from -1.077% for Sweden to +1.007% for Brazil. Total net assets under management are valued at 267 million dollars, on average per country, and on average, a fund has lost -0.062% per day. Average asset values per country vary from 18 million dollars in Portugal to 2327 million dollars in the United States. The last column is the fund age variable, which indicates that, on average, a mutual fund has been in existence for a period of 11 years.

Panel B (Table 1) shows summary statistics of the daily and monthly return distributions. Daily returns present larger excess kurtosis and higher negative skewness than monthly observations. The negative skewness might be the result of the stock market collapse in 2000 and other smaller crashes.

3 Methodology

3.1 Models of mutual fund performance

To estimate fund performance, we use the one-factor CAPM, the Fama and French (1993) three factor model and the Carhart’s (1997) four-factor model:

$$R_{pt} = \alpha_{pt} + \beta_{1,pt}(RM_t - RF_t) + \varepsilon_{pt} \tag{1}$$

$$R_{pt} = \alpha_{pt} + \beta_{1,pt}(RM_t - RF_t) + \beta_{2,pt}SMB_t + \beta_{3,pt}HML_t + \varepsilon_{pt} \tag{2}$$

$$R_{pt} = \alpha_{pt} + \beta_{1,pt}(RM_t - RF_t) + \beta_{2,pt}SMB_t + \beta_{3,pt}HML_t + \beta_{4,pt}MOM_t + \varepsilon_{pt} \tag{3}$$

where R_{pt} is the return on fund p for month t , RF_t is the risk-free rate and RM_t is the market return, SMB_t and HML_t are the Fama and French (1993) size and book-to-market factors and MOM_t is the period t value of the Carhart (1997) momentum return, e_{it} is the residual from the regression, and α_i is the average return above the benchmark. Regression (1) is the CAPM model, regression (2) is the Fama–French three-factor model, and regression (3), including the MOM_t factor, is the Carhart’s four-factor model. Following the recommendation

of [Dimson \(1979\)](#), we add lagged values of the four-factor model to control for the influence of infrequent trading of stocks on daily fund returns. To estimate each fund t -statistic, we apply the [Newey and West \(1987\)](#) heteroskedasticity and autocorrelation consistent estimator of the standard deviation.

[Carhart \(1997\)](#) defends that the four-factor model is more suitable to explain the differences in performance of past winners and past losers. The author points out that this momentum factor accounts for most of this difference. We use a set of control variables that were frequently employed in the relevant finance literature: (i) AGE: The number of years that the fund is in existence. (ii) VOLAT: Volatility is estimated as the standard deviation of the previous 12 month returns. (iii) ASSETS: Fund size measured as total net assets. (iv) Expenses: The total fund expenses. FLOW: Is the net inflows into the funds as defined by [Sirri and Tufano \(1998\)](#). It is estimated as the net growth in assets in excess of returns: $FLOW_{p,t} = [ASSETS_{p,t} - ASSETS_{p,t-1}(1 + NRET_{p,t})]/ASSETS_{p,t-1}$.

3.2 Performance-expenses relation

From an efficiency perspective, higher expenses should be associated with better fund performance, which is in contrast with earlier studies such as [GBRV \(2009\)](#), [Elton et al. \(1993\)](#) and [Carhart \(1997\)](#). [Otten and Bams \(2002\)](#) and [Vidal-Garcia \(2013\)](#) find consistent results with these studies for the European countries whereas [Ippolito \(1989\)](#) shows that returns are not linked to expenses for US funds. The aim of this paper is to provide evidence regarding the relationship between expenses and performance of mutual funds using different measurement methodologies.

We use the following estimation model to empirically examine the expense-performance relation:

$$\begin{aligned} \text{PERF}_{pt} = & \lambda_0 + \lambda_1 \text{EXP}_{p,t} + \lambda_2 \text{LOWPERF}_{p,t-1} \text{EXP}_{p,t} + \lambda_3 \text{MIDPERF}_{p,t-1} \text{EXP}_{p,t} \\ & + \lambda_4 \text{HIGHPERF}_{p,t-1} \text{EXP}_{p,t} + \delta_2 \text{LOWPERF}_{p,t-1} D_{p,t-1} \text{EXP}_{p,t} \\ & + \delta_3 \text{MIDPERF}_{p,t-1} D_{p,t-1} \text{EXP}_{p,t} + \delta_4 \text{HIGHPERF}_{p,t-1} D_{p,t-1} \text{EXP}_{p,t} \\ & + \text{CV}_{pt} + \upsilon_{pt} \end{aligned} \quad (4)$$

where PERF_{pt} are the different performance measures: gross returns (including the expenses), net returns (excluding the expenses), and the values of the risk-adjusted returns (Jensen's alpha) from the CAPM, [Fama and French \(1993\)](#) model, and [Carhart \(1997\)](#) multifactor models; EXP_{pt} is total expenses; and CV_{pt} is a group of control variables including age, volatility, the natural logarithm of assets, and net inflows into the fund. $\text{LOWPERF}_{p,t-1}$, $\text{MIDPERF}_{p,t-1}$, and $\text{HIGHPERF}_{p,t-1}$ are the performance fractional ranks of fund p in period $t-1$, and $D_{p,t-1}$ is a dummy variable to account for the effect of small funds; υ_{pt} is the error term.

We divide the excess returns into partial ranks to control for the asymmetric flow-to-performance relation reported in the finance literature. Following [Sirri and Tufano \(1998\)](#), we sort funds from 0 (bottom) to 1 (top) as a result of their performance in the previous year. We set the following fractional classification: $\text{LOWPERF}_{i,t-1}$ is the worst return quintile, determined as $\text{Min}(\text{RANK}_{t-1}, 0.2)$; $\text{MIDPERF}_{i,t-1}$ represent the middle three return quintiles, determined as $\text{Min}(0.6, \text{RANK} - \text{LOWPERF})$; and $\text{HIGHPERF}_{i,t-1}$ is the best return quintile. By separating fund returns into different sorts, we divide the sensitivity of the mutual fund fees to the performance ranks.

3.3 The data envelopment analysis model

We measure the efficiency of domestic equity funds using the Data Envelopment Analysis (DEA) non-parametric methodology employed in the resolution of production functions. This technique was initially developed by [Charnes et al. \(1978\)](#) to evaluate the performance of educational institutions and since then it has been widely used to evaluate the performance of decision-making units (DMUs) determined by several inputs-outputs structures. It is a useful methodology for examining performance as it is possible to include multiple inputs and outputs that can be measured in different units. The DEA evaluates the highest potential output for a certain number of inputs. It sets an efficiency measure for each decision-making unit related to the best operating unit within a given sample. The procedure examines how efficiently a decision-making unit uses the available resources to create the outputs. The performance of these decision-making units is examined in DEA applying the concept of efficiency described as a ratio of total weighted outputs to total weighted inputs. [Castelli et al. \(2010\)](#) explain that the targets of a DMU are the levels of outputs (inputs) that a given DMU should reach by increasing (decreasing) its yield (consumption) to become efficient.

Efficiencies estimates using DEA are relative to the top performing DMUs. The most efficient DMUs are assigned an efficiency score of unity (100%) and the performance of the rest of DMUs ranges from 0 to 100% compared to the top performers. In this sense, [Angulo-Meza and Lins \(2002\)](#) note that the advantage of this model is that it gives a set of efficient solutions compared to the linear programming model that gives only one optimal solution.

The DEA methodology has been implemented to examine mutual fund performance in the U.S. by several authors, including [Murthi et al. \(1997\)](#), [Morey and Morey \(1999\)](#) and [Basso and Funari \(2001\)](#). To the best of our knowledge, our study is the first one to examine mutual funds’ performance around the world using DEA. The DEA procedure might be used to describe mutual fund indexes with various inputs as risk measures and fees.

If the efficiency is unity, then the DEA technique represents a Pareto efficiency measure² and the efficient units are located on the efficiency frontier. As [Chen and Zhu \(2003\)](#) point out DEA models have been proven as an effective methodology for estimating efficient frontiers.

As described by [Charnes et al. \(1994\)](#), to estimate the DEA efficiency model for a DMU, we should find the optimal solution to the following fractional linear programming problem:

$$\max\{v_i, u_r\} \quad h_{jo} = \frac{\sum_{r=1}^t \mu_r \gamma_{rjo}}{\sum_{i=1}^m v_i x_{ijo}} \tag{5}$$

$$\begin{aligned} \text{subject to} \quad & \frac{\sum_{r=1}^t \mu_r \gamma_{rjo}}{\sum_{i=1}^m v_i x_{ijo}} \leq 1, \quad j = 1, \dots, n \tag{6} \\ & \mu_r \geq \varepsilon \quad r = 1, \dots, t \\ & v_i \geq \varepsilon \quad i = 1, \dots, m \end{aligned}$$

where ε is a small positive number to make sure that the weights are not negative. From equation (6), we obtain the value of the optimal objective function, which is the efficiency measure for unit j_o . We can obtain an equivalent linear programming problem by converting the fractional problem explained above ([Charnes and Cooper 1962](#)). We set $\sum_{i=1}^m v_i x_{io} = 1$, obtaining the Charnes, Cooper and Rhodes (CCR) model:

² An efficient economic outcome is a situation where one party’s position cannot be improved without making another party’s position worse.

$$\max \sum_{r=1}^t u_r \gamma_{rjo} \quad (7)$$

$$\text{subject to } \sum_{i=1}^m v_i x_{ij0} = 1 \quad (8)$$

$$\sum_{r=1}^t u_r \gamma_{rj} - \sum_{i=1}^m v_i x_{ij} \leq 0, \quad j = 1, \dots, n \quad (9)$$

$$\mu_r \leq -\varepsilon \quad r = 1, \dots, t$$

$$v_i \leq -\varepsilon \quad i = 1, \dots, m$$

The optimization problem can be solved by estimating the values of $t+m$ variables, which means the weights u_r and v_i , conditional to $n+t+m+1$ restrictions.

3.4 Data envelopment analysis model for mutual fund performance

The DEA methodology has been applied to examine mutual fund performance in several studies. The DEA technique allows determining mutual fund performance while considering various inputs (variables) like risk measures and expenses. The costs related to mutual fund investment are an important element when examining the fund performance. However, the popular ratios and multifactor models used do not always take into account fund expenses.

[Murthi et al. \(1997\)](#) were the first to apply the DEA approach in a mutual fund efficiency index called DPEI (DEA portfolio efficiency index). They used fund returns as output and expense ratio, turnover, load and standard deviation of returns as inputs. [Basso and Funari \(2001\)](#) develop the $I_{\text{DEA}-1}$ index, which is similar to the DPEI with some difference in the investment costs taken into account in the model. They only consider subscription and redemption costs without other expenses as they have already been subtracted from fund returns.

Using the DEA methodology allows the inefficient funds to know which other fund they could imitate to improve their efficiency. The efficient fund could be a target benchmark and the inefficient fund could obtain better performance imitating the behavior of the efficient one as both have the same input-output characteristics.

An important advantage of the DEA is that it does not need any theoretical model as a measurement benchmark as it is based on a non-parametric analysis. DEA is also useful to account for the issue of endogeneity of transaction costs as it includes the expense ratio, loads, turnover and returns simultaneously in the analysis. The model can examine many outputs and inputs at the same time. For example, the usual case when managers are in charge of the returns and size of the fund.

The DEA model has some advantages over traditional performance measures as explained by [Basso and Funari \(2001\)](#). The DEA model results are not sensitive to the selected investment period using logarithmic returns and assuming stationarity and independence of returns over time. In contrast, the traditional performance models are affected by the investment period and it is possible to obtain different results depending on the frequency of the observations used (daily, monthly, etc). Thus, there is a systematic bias using a time horizon which is not the same as the one considered by the investors.

Another useful characteristic of the DEA model is the possibility to improve the inefficient units with the evidence from their peers. The inefficient units could imitate a unit on the efficient frontier in order to be more efficient as the ones on the efficient frontier have the same input variables as the inefficient units. This efficient unit could be a benchmark for the inefficient funds. In this sense, [Morey and Morey \(1999\)](#) point out that this benchmark is a fund of funds that investors could buy.

4 Results

4.1 Performance evaluation

Table 2 shows the daily statistics for the four-factor Carhart (1997) models for the period from January 1990 to December 2015. Most premiums (alphas) are positive, suggesting that more risky, small, value-focused, and previous-winner stocks achieved higher returns. This fact indicates that the Carhart factors could explain most of the cross-sectional variation in average daily returns around the world over the period under consideration. The size (SMB) factor is always positive, the book-to-market (HML) factor presents mixed results, being significantly positive for most countries, while the momentum factor (MOM) is also significant for most countries, showing mixed evidence. The size factor (SMB) suggests that small stocks have had a satisfactory performance and fund returns are driven by smaller stocks during the sample period. The book-to-market factor (HML) suggests that funds follow a value-oriented style. Momentum strategies only add value in 14 out of 35 countries, while 9 countries present contrarian strategies. The average alpha across the different countries is positive in 21 out of 35 countries. This is in contrast to most mutual fund literature presenting persistent underperformance (e.g. Carhart 1997).

Table 3 reports performance results from the estimation of regressions (1) to (3). Most countries present a positive performance (alpha) irrespective of the multifactor model used in the analysis. The before-expenses measures of performance do not change the sign of the coefficients. As expected, the gross risk-adjusted returns are positive and larger for all models than the net risk-adjusted returns (see, Table 3). The best result is obtained when we use the Carhart (1997) model to measure fund risk-adjusted performance. The daily mean gross risk-adjusted return (for the Carhart model) ranges from 0.178% for Brazil to -0.041% for Sweden, while the daily mean net risk-adjusted return varies from 0.089% for Brazil to -0.082% for Sweden. In Table 3 we do not find significant differences when considering fund expenses across countries. Our results suggest similar behavior for mutual funds in terms of performance assessment.

Next, we examine funds per country using risk-adjusted returns (CAPM, Fama–French and Carhart) and non-risk adjusted performance evaluation (net and gross returns). In Table 4, we examine fund performance by presenting the percentage of positive and negative values of performance measures. Panel A reports the proportion of positive values, while Panels B and C show the proportion of statistically significant (at the 5%) positive and negative values, respectively. Less than half of the funds exhibit positive values of performance as shown in Panel A. We use a paired t-test to reject the null hypothesis that the true mean difference is zero. Considering the gross risk-adjusted performance (Carhart (G)), the estimations range from 40.4% for Denmark (Carhart model) to 61.6% for Indonesia. There is no significant difference when we compare the risk-adjusted values after expenses (Carhart (N)) with the results before expenses (Carhart (G)). Panel B of Table 4 confirms the previous results. Before-expenses measures of performance (gross returns (G)) obtain a larger percentage of positive alphas compared to results after expenses (net returns (N)). Panel B of Table 4 presents the percentage of funds with significant positive values of performance (gross risk-adjusted) vary from 4.9% for Indonesia to 15.4% for Austria (Panel B, last column). Panel C presents the percentage of significantly negative risk-adjusted estimations. The proportion of significantly positive alphas (Panel B) is larger than the proportion of significantly negative ones (Panel C) only before expenses (gross returns (G)). For instance, 10.2% performance values in the Carhart model (Carhart (G)) for the US are significantly positive, while only

Table 2 Summary statistics for the Carhart model

	4-Factor model					Adj R ²
	Alpha	RMRF	SMB	HML	MOM	
Australia	0.814***	0.436***	0.028***	0.107***	0.167***	0.745
Austria	0.013**	0.749***	−0.009	0.094	0.109***	0.734
Belgium	0.198***	0.712***	0.056***	0.098***	0.121	0.749
Brazil	0.892*	0.543***	0.034***	0.221***	0.161	0.797
Canada	0.538**	0.456***	0.034**	0.078	−0.089**	0.793
Chile	−0.160***	0.234***	0.089**	0.167***	−1.407	0.834
China	−0.566**	0.534***	0.086***	−0.033	0.030	0.824
Denmark	0.776***	0.365***	0.035***	−0.078***	0.109***	0.745
Finland	0.046**	0.474***	0.059***	−0.044**	0.176***	0.834
France	0.487***	0.336***	0.084***	−0.079***	0.142***	0.875
Germany	−0.087**	0.469***	0.085***	−0.077**	0.177***	0.783
Hong Kong	−0.576***	0.364***	0.128***	−0.034***	−0.156***	0.791
India	0.874***	0.459***	0.058**	−0.078	0.029	0.723
Indonesia	0.624***	0.543**	0.069***	−0.067*	−0.207**	0.922
Ireland	0.587**	0.623***	0.289***	0.189***	0.067***	0.834
Israel	0.046***	0.534***	0.117***	0.134***	−0.022**	0.818
Italy	0.507**	0.555***	0.255***	0.078***	0.034	0.741
Japan	−0.712**	0.522***	0.087***	0.089**	0.019**	0.856
Korea (South)	0.117***	0.548***	0.083***	0.085***	0.080	0.638
Luxembourg	0.189**	0.632***	0.087***	0.078***	0.065***	0.878
Malaysia	0.316***	0.234***	0.127***	0.089***	−0.089***	0.834
Netherlands	−0.675*	0.481***	0.074***	0.098***	0.007**	0.791
New Zealand	0.738**	0.087**	0.066***	0.019	0.094*	0.855
Norway	0.238***	0.477***	0.128***	0.078***	0.019	0.588
Poland	−0.750**	0.589***	0.089***	0.074***	−0.022	0.841
Portugal	−0.180***	0.241***	0.096***	0.011	−0.021	0.866
Singapore	−0.398**	0.354***	0.195***	−0.039***	−0.038	0.981
South Africa	−0.423***	0.567***	0.166***	0.076***	0.055**	0.854
Spain	−0.305***	0.658***	0.154***	0.176***	−0.067**	0.845
Sweden	−0.822***	0.523***	0.077**	0.156***	0.043***	0.876
Switzerland	−0.719***	−0.007*	0.052*	0.134***	0.028	0.784
Taiwan	0.648**	0.415***	0.156***	0.133***	−0.019***	0.767
Thailand	0.698**	0.419***	0.167***	0.078	−0.017**	0.845
United Kingdom	−0.274**	0.152***	0.109***	0.078***	0.018***	0.862
United States	0.448***	0.176***	0.78***	0.208***	−0.014***	0.867

This table presents the results for the estimation of regression (3) along with their adjusted R-square for the period January-1990 to December-2015. *RMRF* is the market return minus the risk-free rate, *SMB* and *HML* are, respectively, the Fama–French (1993) size and book-to-market factors and *MOM* is the Carhart (1997) momentum factor

***, **, and * indicate significance at the 1, 5, and 10% level, respectively

Table 3 Risk exposure estimates

	CAPM(N)	CAPM(G)	F–F(N)	F–F(G)	Carhart (N)	Carhart (G)
Australia	0.062	0.135	0.054	0.060	0.081	0.162
Austria	0.003	0.007	0.002	0.003	0.004	0.008
Belgium	0.015	0.033	0.013	0.014	0.019	0.039
Brazil	0.068	0.148	0.059	0.066	0.089	0.178
Canada	0.041	0.089	0.035	0.039	0.053	0.107
Chile	–0.020	–0.009	–0.025	–0.022	–0.016	–0.008
China	–0.073	–0.033	–0.090	–0.079	–0.056	–0.028
Denmark	0.059	0.129	0.051	0.057	0.077	0.155
Finland	0.012	0.026	0.010	0.011	0.016	0.032
France	0.037	0.081	0.032	0.036	0.048	0.097
Germany	–0.011	–0.005	–0.013	–0.012	–0.008	–0.004
Hong Kong	–0.074	–0.034	–0.092	–0.080	–0.057	–0.028
India	0.067	0.145	0.058	0.064	0.087	0.174
Indonesia	0.048	0.104	0.041	0.046	0.062	0.124
Ireland	0.045	0.097	0.039	0.043	0.058	0.117
Israel	0.007	0.014	0.005	0.006	0.008	0.017
Italy	0.039	0.084	0.033	0.037	0.050	0.101
Japan	–0.093	–0.042	–0.113	–0.099	–0.071	–0.035
Korea (South)	0.009	0.019	0.007	0.008	0.011	0.023
Luxembourg	0.015	0.031	0.012	0.014	0.018	0.037
Malaysia	0.024	0.052	0.021	0.023	0.031	0.063
Netherlands	–0.087	–0.040	–0.108	–0.094	–0.067	–0.033
New Zealand	0.056	0.123	0.049	0.054	0.073	0.147
Norway	0.018	0.039	0.015	0.017	0.023	0.047
Poland	–0.097	–0.045	–0.120	–0.105	–0.075	–0.037
Portugal	–0.023	–0.010	–0.028	–0.025	–0.018	–0.009
Singapore	–0.051	–0.023	–0.063	–0.055	–0.039	–0.019
South Africa	–0.055	–0.025	–0.067	–0.059	–0.042	–0.021
Spain	–0.039	–0.018	–0.048	–0.042	–0.030	–0.015
Sweden	–0.106	–0.049	–0.131	–0.115	–0.082	–0.041
Switzerland	–0.093	–0.043	–0.115	–0.101	–0.071	–0.035
Taiwan	0.049	0.108	0.043	0.048	0.064	0.129
Thailand	0.053	0.116	0.046	0.051	0.069	0.139
United Kingdom	–0.035	–0.016	–0.043	–0.038	–0.027	–0.013
United States	0.034	0.074	0.029	0.033	0.044	0.089

This table presents the results for the estimation of regressions (1) to (3) along with their adjusted R-square for the period January-1990 to December-2015. The table shows returns for the CAPM, Fama–French and Carhart models using Net (N) and Gross (G) returns. *CAPM(N)* presents the CAPM alpha estimates using net returns. *CAPM(G)* presents the CAPM alpha estimates using gross returns. *F–F(N)* shows the Fama–French alpha estimates using net returns. *F–F(G)* shows the Fama–French alpha estimates using gross returns. *Carhart (N)* presents the Carhart alpha estimates using net returns. *Carhart (G)* presents the Carhart alpha estimates using gross returns

***, **, and * indicate significance at the 1, 5, and 10% level, respectively

Table 4 Performance measure distribution

	NRET	GRET	CAPM (N)	CAPM (G)	F-F (N)	F-F (G)	Carhart (N)	Carhart (G)
<i>Panel A: Proportion of funds with positive values of performance measures</i>								
Australia	47.3	51.4	39.5	55.3	28.2	43.8	35.9	56.4
Austria	46.7	52.3	34.2	47.8	24.4	38.3	31.1	48.8
Belgium	52.3	55.6	37.7	52.7	26.9	41.8	34.3	53.8
Brazil	54.1	56.3	33.1	46.3	23.6	36.7	30.2	47.2
Canada	42.3	45.7	40.5	56.7	28.9	45.4	36.8	57.8
Chile	49.6	53.3	31.1	43.5	22.2	34.5	28.2	44.4
China	52.2	55.7	29.8	41.7	21.3	33.1	27.2	42.5
Denmark	54.3	57.2	28.3	39.6	20.2	31.4	25.7	40.4
Finland	51.7	54.8	35.1	45.6	25.1	39.6	31.9	50.1
France	50.3	54.6	33.4	43.4	25.7	37.1	30.3	55.6
Germany	42.3	46.1	32.2	41.8	24.8	35.7	29.2	53.6
Hong Kong	54.3	56.3	30.4	39.5	23.4	33.7	27.6	50.6
India	39.8	44.3	26.4	34.3	20.3	29.3	24.4	44
Indonesia	52.3	56.2	41.2	53.5	31.7	45.7	37.4	61.6
Ireland	50.1	53.7	32.2	41.8	24.7	35.7	29.2	53.6
Israel	47.5	51.2	34.5	44.8	26.5	38.3	31.4	57.5
Italy	44.5	47.5	32.0	41.6	24.6	35.5	29.2	53.3
Japan	49.6	53.3	28.9	37.5	22.2	32.1	26.2	48.2
S. Korea	54.4	55.4	34.5	44.8	26.5	38.3	28.7	57.5
Luxembourg	50.3	52.6	34.5	44.8	26.5	38.3	28.7	57.5
Malaysia	52.7	53.5	33.6	43.6	24.7	37.3	28.6	48.1
Netherlands	55.1	57.3	34.5	44.8	24.6	38.3	28.7	49.2
N. Zealand	53.7	55.4	31.2	40.5	22.4	34.6	26.9	44.5
Norway	54.1	55.8	31.8	44.5	22.7	35.3	26.5	45.4
Poland	52.7	54.5	30.4	42.5	21.7	33.7	25.3	43.4
Portugal	55.2	57.1	31.7	44.3	22.6	35.2	26.4	45.3
Singapore	58.1	59.6	35.6	49.8	25.4	39.5	29.6	50.8
South Africa	52.7	54.3	31.2	43.6	22.3	34.6	26.2	44.5
Spain	53.2	55.8	33.6	47.0	23.8	37.3	28.8	48.1
Sweden	54.3	56.7	32.1	44.9	22.9	35.6	26.7	45.8
Switzerland	55.8	56.4	35.6	49.8	25.4	39.5	29.6	50.8
Taiwan	57.2	58.4	34.2	47.8	24.4	38.3	28.5	48.6
Thailand	47.4	49.3	38.4	53.7	27.4	42.6	32.8	54.8
United Kingdom	48.4	50.4	29.6	41.4	21.1	32.8	24.6	42.2
United States	54.2	56.7	32.1	44.9	22.9	35.6	26.7	45.8

Table 4 continued

	CAPM(N)	CAPM(G)	F–F(N)	F–F(G)	Carhart (N)	Carhart (G)
<i>Panel B: Proportion of funds with significant positive values of performance measures</i>						
Australia	5.6	10.1	2.5	10.1	4	8.4
Austria	10.3	18.5	4.6	18.5	7.3	15.4
Belgium	4.5	8.1	2.2	8.1	3.2	6.7
Brazil	6.5	11.7	2.9	11.7	4.6	9.7
Canada	3.4	6.1	1.5	6.1	2.4	5.1
Chile	5.3	9.5	2.4	9.5	3.7	7.9
China	4.8	8.6	2.1	8.6	3.4	7.2
Denmark	5.6	10.2	2.5	10.1	4.1	8.4
Finland	7.3	13.1	3.3	13.1	5.2	10.9
France	5.2	9.3	2.3	9.3	3.7	7.8
Germany	6.2	11.1	2.8	11.1	4.4	9.3
Hong Kong	5.4	9.7	2.4	9.7	3.8	7.5
India	4.8	8.6	2.1	8.6	3.4	6.3
Indonesia	3.5	6.3	1.5	6.3	2.5	4.9
Ireland	3.7	6.6	1.6	6.6	2.8	5.1
Israel	5.9	10.6	2.6	10.6	4.5	8.2
Italy	6.1	9.7	2.6	6.5	4.6	8.5
Japan	6.5	10.4	2.8	6.9	5	9.1
Korea (South)	4.5	7.2	1.9	4.8	3.4	6.3
Luxembourg	3.8	6.2	1.6	4.2	2.9	5.3
Malaysia	4.7	7.5	2.1	5.2	3.6	6.5
Netherlands	7.4	11.8	3.2	7.9	5.6	10.3
New Zealand	4.7	7.5	2.2	5.2	3.6	6.5
Norway	5.6	8.9	2.4	5.9	4.3	7.8
Poland	5.8	9.2	2.5	6.2	4.4	8.7
Portugal	4.6	7.3	2.1	4.9	3.5	6.9
Singapore	5.7	9.1	2.4	6.2	4.3	8.5
South Africa	5.4	8.6	2.3	5.7	4.2	8.1
Spain	4.5	7.2	1.9	4.8	3.4	6.7
Sweden	4.3	6.8	1.8	4.6	3.3	6.4
Switzerland	5.4	8.6	2.3	5.7	4.2	8.1
Taiwan	4.7	7.5	2.1	5.3	3.6	7.2
Thailand	5.4	8.6	2.3	5.7	4.2	8.1
United Kingdom	4.3	6.8	1.8	4.6	3.3	6.4
United States	6.8	10.8	2.9	7.2	5.2	10.2

Table 4 continued

	CAPM(N)	CAPM(G)	F–F(N)	F–F(G)	Carhart(N)	Carhart (G)
<i>Panel C: Proportion of funds with significant negative values of performance measures</i>						
Australia	7.2	5.8	8.9	4	5.2	1.8
Austria	13.3	10.8	16.4	7.3	9.1	3.3
Belgium	5.8	4.7	7.2	3.2	4.1	1.4
Brazil	8.4	7	10.4	4.6	5.8	2.1
Canada	4.4	3.5	5.4	2.4	3.3	1.1
Chile	6.8	5.5	8.4	3.7	4.7	1.7
China	6.2	5.2	7.6	3.4	4.2	1.5
Denmark	7.2	5.8	8.9	4.3	5.3	1.8
Finland	9.4	7.6	11.6	5.2	6.5	2.3
France	6.7	5.4	8.3	3.7	4.6	1.6
Germany	8.1	6.5	9.9	4.4	5.5	2
Hong Kong	7.1	5.6	8.6	3.8	4.8	1.7
India	6.7	5.0	7.6	3.4	4.2	1.5
Indonesia	4.9	3.6	5.9	2.5	3.1	1.1
Ireland	5.1	3.8	6.2	2.6	3.3	1.2
Israel	8.2	6.2	10.1	4.2	5.2	1.9
Italy	8.5	6.4	10.3	4.3	5.4	1.9
Japan	9.1	6.8	11.1	4.6	5.8	2.2
Korea (South)	6.3	4.6	7.3	3.2	4.2	1.4
Luxembourg	5.3	3.9	6.4	2.7	3.3	1.2
Malaysia	6.5	4.8	7.6	3.3	4.1	1.5
Netherlands	10.3	7.6	12.4	5.2	6.6	2.3
New Zealand	6.5	4.8	7.9	3.3	4.1	1.5
Norway	7.8	5.8	9.5	4.3	5	1.8
Poland	8.1	6.1	9.8	4.1	5.1	1.9
Portugal	6.4	4.7	7.8	3.2	4.1	1.4
Singapore	7.9	5.9	9.6	4.1	5.2	1.8
South Africa	7.5	5.6	9.1	3.8	4.8	1.7
Spain	6.3	4.6	7.6	3.2	4.1	1.4
Sweden	6.2	4.4	7.3	3.1	3.8	1.3
Switzerland	7.5	5.6	9.1	3.8	4.8	1.7
Taiwan	6.5	4.8	7.9	3.3	4.1	1.5
Thailand	7.5	5.6	9.1	3.8	4.8	1.7
United Kingdom	6.2	4.4	7.3	3.1	3.8	1.3
United States	9.5	7.1	11.5	4.8	6.1	2.2

This table shows the distribution of fund-month performance measure observations in our sample according to quantity. Panel A presents the percentage of positive values. Panels B and C report the percentage of statistically significant positive and negative estimations, respectively. *NRET* present results using net returns. *GRET* present results using gross returns. *CAPM(N)* presents results using net returns based on CAPM. *CAPM(G)* presents results using gross returns based on CAPM. *F–F(N)* shows the results using net returns based on Fama–French model. *F–F(G)* shows the results using gross returns based on Fama–French model. *Carhart (N)* presents estimation results using net returns based on Carhart model. *Carhart (G)* presents estimation results using gross returns based on Carhart model

2.2% are negative. From Panel C we can appreciate that opposite estimates are shown if net risk-adjusted values are estimated (Carhart (N)).

4.2 Performance-expenses relation

This section analyzes the economic efficiency of the mutual fund industry per country. We empirically test the relationship between performance and the fees paid by investors. Grossman and Stiglitz (1980) explain that there should be a positive link between fees and before-expenses performance. Thus, we expect the performance-expenses relation to be positive. From an efficiency perspective, higher expenses paid by investors should be related to greater performance. Efficiency means that fund services to investors would cover their costs, and thus net performance should be similar between funds after expenses. The mutual fund industry should show a robust link between expenses and gross performance.

Table 5 shows the results for the performance-expenses relation for the Carhart (1997) model using net returns. The coefficients of expenses are always significantly negative (at better than the 5% level) for all countries and risk-adjusted performance measures. Very similar results are found using other multifactor models³. We also find a negative relationship between returns and expenses for fractional ranks of past performance ($LOWP_{(t-1)}$). This negative relation between returns and expenses from previous performance is linked to the bottom performing funds ($LOWP_{(t-1)}$) while it is not significant for the medium and high fractional ranks. Moreover, to examine whether our findings are influenced by the performance of small funds, we add a dummy variable to each fractional rank that is equal to one if the fund size is in the lowest 10% of the fund size distribution, and zero otherwise ($L \times Small$, $M \times Small$, $H \times Small$). We do not obtain a relevant significance for our sample of countries.

As previously documented in the literature, we find that funds with poor risk-adjusted performance incur higher expenses (see Table 5, column $LOWP_{(t-1)}$) which means that funds charging high expenses tend to underperform, in contrast to the indication of the efficiency theory.

Table 5 allows us to examine the effect of mutual fund characteristics on risk-adjusted performance. In line with prior studies of fund performance, younger funds obtain higher performance than older ones. We show a negative link between fund volatility and fund performance, indicating that more volatile funds present worse performance. We also find a significant positive relation between fund assets and fund risk-adjusted performance, which indicate possible economies of scale for mutual fund markets around the world. Finally, we document a negative relation between net inflows and fund performance, which shows an improvement in performance when net inflows are negative.

4.3 Data envelopment analysis (DEA)

We use DEA methodology, a technique used in operations research to evaluate relative measures of efficiency, to examine mutual fund performance. Using this methodology, we address some of the main problems in portfolio evaluation since it does not require a benchmark and allow considering mutual fund expenses. DEA is a flexible methodology for performance evaluation as it permits a model with many inputs and outputs. This technique allows us to examine whether the fund manager uses the available resources to achieve the maximum level of output (scale efficiency). It is a suitable methodology to use for portfolio evaluation, as investors are willing to invest in a fund that maximizes returns and minimizes expenses.

³ We omit the tables for the sake of brevity. They are, however, available from the authors upon request.

Table 5 Fund performance-expenses relation

	λ_0	EXP	VOLAT	AGE	ASSET	FLOW
Australia	1.043***	-0.083***	-0.036***	-0.659***	0.244**	-0.114***
Austria	0.022***	-1.011***	-0.198***	-0.089***	0.065**	-0.067***
Belgium	0.336***	-1.082***	-0.566**	-0.457***	0.852***	-0.078**
Brazil	1.346***	-0.089**	-0.096***	-0.239***	0.095***	-0.089**
Canada	0.914***	-2.134***	-1.423*	-0.563***	1.339**	-0.221**
Chile	-0.094***	-2.052***	-0.267***	-0.229***	0.014**	-0.078***
China	-0.334***	-0.793***	-0.078**	-0.510***	0.776**	-0.216***
Denmark	1.319***	-0.665***	-0.084***	-0.159***	-0.004	-0.124***
Finland	0.078***	-0.028**	-0.012***	-0.089***	0.293**	-0.220***
France	0.827***	-1.083***	-0.329***	-0.439***	0.482***	-0.114***
Germany	-0.051***	-1.017***	-0.385**	-0.310***	1.024**	-0.056***
Hong Kong	-0.339***	-0.982**	-0.011***	-0.609***	0.589***	-0.078***
India	1.485***	-0.214***	-0.048**	-0.511***	0.112***	-0.088***
Indonesia	1.060***	-1.033**	-0.087***	-0.449***	0.595***	-0.233***
Ireland	0.997***	-0.062***	-0.087***	-0.349***	0.928**	-0.313**
Israel	0.078***	-1.872***	-0.378***	-0.279***	0.079***	-0.214***
Italy	0.861***	-0.771**	-0.629***	-0.612***	1.392***	-0.218***
Japan	-0.418***	-1.241**	-0.089**	-0.339***	0.431***	-0.119***
Korea (South)	0.198***	-1.049**	-0.355**	-0.459***	0.471***	-0.214***
Luxembourg	0.321***	-1.710***	-0.321***	-0.099***	1.879***	-0.134***
Malaysia	0.537***	-1.945***	-1.256**	-0.110***	1.391**	-0.116**
Netherlands	-0.397***	-0.776***	-0.135**	-0.709***	1.781***	-0.214***
New Zealand	1.254***	-0.016**	-0.078***	-0.208***	1.011**	-0.212***
Norway	0.404***	-2.634**	-0.278***	-0.089***	0.317**	-0.114***
Poland	-0.441***	-1.048***	-0.328***	-0.716***	0.727***	-0.226***
Portugal	-0.106***	-0.349***	-0.209**	-0.089***	0.158***	-0.078***
Singapore	-0.234***	-3.011***	-0.346***	-0.412***	1.769**	-0.056***
South Africa	-0.248***	-1.043**	-0.222***	-0.209***	0.745***	-0.074***
Spain	-0.179***	-2.081***	-0.298**	-0.315***	0.681**	-0.073***
Sweden	-0.483***	-0.073***	0.155	-0.279***	0.526***	-0.178***
Switzerland	-0.422***	-1.563***	-0.067***	-0.112***	0.758**	-0.088***
Taiwan	1.101***	-0.032***	0.012	-0.229***	0.028	-0.073***
Thailand	1.186***	-0.531***	-0.207***	-0.310***	0.673***	-0.178***
United Kingdom	-0.162***	-0.295***	-0.288***	-0.409***	0.012***	-0.266***
United States	0.761***	-0.629***	-1.230***	-0.308***	1.182**	-0.117**

Table 5 continued

	LOWP. (t-1)	MIDP. (t-1)	HIGHP. (t-1)	L × Small	M × Small	H × Small	Adjusted R ²
Australia	-0.178***	0.149	0.078	0.034*	0.057*	0.078	0.556
Austria	-0.089***	0.112*	0.126	0.067	0.078	0.165	0.387
Belgium	-0.123***	0.088	0.089	0.167	0.087	0.098	0.488
Brazil	-0.077***	0.045	0.875	0.089**	0.178	0.023	0.353
Canada	-0.178***	0.456	0.126	0.456	0.213	0.234	0.345
Chile	-0.089***	0.056*	0.556	0.178	0.112	0.167	0.556
China	-0.218***	0.843	0.078	0.346	0.189	0.133	0.423
Denmark	-0.099***	0.678	0.045	0.178	0.288	0.267	0.593
Finland	-0.067***	-0.028	0.123	0.211	0.188	0.188	0.266
France	-0.188***	0.067	0.078	0.178	0.078*	0.145	0.367
Germany	-0.213***	0.187	0.234	0.133	0.133	0.378	0.614
Hong Kong	-0.066***	0.457	0.223	0.038	0.047	0.211	0.245
India	-0.234***	0.089*	0.121	0.089	0.178	0.198	0.213
Indonesia	-0.156***	0.177	0.122	0.189	0.084	0.234	0.523
Ireland	-0.217***	0.217	0.078	0.188	0.178	0.178	0.567
Israel	-0.055***	0.489	0.233	0.194	0.277	0.109	0.214
Italy	-0.067***	0.067	0.167	0.056	0.067*	0.122	0.312
Japan	-0.078***	0.056	0.667	0.188	0.056*	0.245	0.542
Korea (South)	-0.312***	0.055	0.217	0.277	0.134	0.187	0.298
Luxembourg	-0.213***	0.078*	0.178	0.167	0.045*	0.256	0.543
Malaysia	-0.217***	0.433	0.216	0.189	0.089	0.189	0.598
Netherlands	-0.067***	0.312	0.178	0.154	0.024	0.132	0.478
New Zealand	-0.213***	0.255	0.189	0.178	0.011**	0.213	0.334
Norway	-0.068***	0.122	0.234	0.090*	0.036	0.187	0.388
Poland	-0.213***	0.178*	0.233	0.178	0.067	0.278	0.394
Portugal	-0.188***	0.324	0.178	0.295	0.023*	0.122	0.543
Singapore	-0.198***	0.045	0.376	0.156**	0.278	0.034	0.433
South Africa	-0.087***	0.156	0.178	0.167	0.178*	0.028	0.453
Spain	-0.244***	0.067	0.211	0.156	0.278	0.067	0.598
Sweden	-0.099***	0.078	0.189	0.067	0.167	0.034	0.289
Switzerland	-0.056***	0.134	0.278	0.166	0.078	0.011	0.312
Taiwan	-0.287***	0.167	0.189	0.155	0.144	0.078	0.345
Thailand	-0.199***	0.167	0.354	0.067	0.213	0.166	0.367
United Kingdom	-0.078***	0.167	0.289	0.278	0.154	0.034	0.456
United States	-0.358***	0.245	0.188	0.177	0.078	0.056	0.245

The table presents the results for the estimation of equation (4) for the period January-1990 to December-2015. The table uses estimations (Jensen's alpha) from the Carhart model. λ_0 is the alpha of the model. *EXP* represents total expenses, *VOLAT* is fund volatility, *AGE* is the average number of years the fund exists in the sample, *ASSET* is the natural logarithm of total assets, *FLOW* are net inflows into the fund. *LOWP(t-1)*, *MIDP(t-1)*, and *HIGHP(t-1)* are the performance fractional ranks in period *t-1*. *L × Small*, *M × Small*, and *H × Small* are dummy variables to account for the effect of small funds. *Adj R-sq* is the adjusted R-square from the regression

***, **, and * indicate significance at the 1, 5, and 10% level, respectively

The DEA compares each fund to the best available funds in the same country. With known expenses and risk taken, we examine for each country the most profitable fund. Investors are interested to find the fund return net of expenses at a given level of risk. Thus, we consider return as the only output, and four inputs, namely expense ratio, turnover, load and standard deviation of returns. DEA examines the efficiency of a fund compared to the group of funds that consider the same inputs to achieve the same outputs. DEA separates the efficient funds from the inefficient ones depending on whether they are on the Pareto-efficient frontier or not. The separation from the efficient frontier provides an estimation of its relative inefficiency. This supports the mean–variance theory of [Markowitz \(1952\)](#), who explains that market portfolios are efficient when they have the highest expected return for a given variance. Thus, the DEA shows whether a fund can improve its performance compared to a group of similar funds. When the efficiency is 100% and the slack variables (input factors which represent performance inefficiencies) are zero, the output of a fund cannot be expanded without increasing its inputs. In this case, the inputs cannot be lowered with the current output level. Then, the fund would be Pareto-efficient with output efficiency of 1, while a DEA measure below 1 show that the fund is inefficient. A fund's inefficiency is estimated as the difference between the efficiency value and 1.

We examine the variation in average DEA across countries. For this purpose, we estimate the DEA score for each fund within each country. [Table 6](#) presents the percentage of efficient funds across countries for our sample period 1990–2015 for every variation of the DEA model using gross returns, and alphas from CAPM, Fama–French and Carhart models as output measures. It is clear that most funds are efficient using any return measure as output and across countries. It is worth noting that our sample period is long enough to include extreme market events like recession periods and financial crisis across countries, which increases volatility in stock markets. We address this issue in a subsequent analysis. Gross returns have higher efficiency rates, while the proportion of efficient funds tends to decrease when using multifactor models, although it increases with larger factor models. Using gross returns, the percentage of efficient funds ranges from 64% for Australia to 85.2% for Finland, while using the four-factor model of Carhart the proportion of efficient funds ranges from 58.8% from Australia to 79.1% for Spain. We do not find evidence of any relation between number of funds per country and efficiency, as the country with the largest number of funds, the U.S., shows an efficiency percentage using gross returns of 70.5%, while the country with the lowest number of funds, Hong Kong, presents an efficiency percentage of 78.4%. Our results support the evidence shown by [Ippolito \(1989\)](#) that in an efficient market fund returns allow to cover loads and expenses.

[Table 7](#) shows the average DEA efficiency scores by country. We present the score varying between 0 and perfectly efficient funds scoring 1. We find that funds from all countries present a degree of efficiency above 0.60 using resources, although they are not Pareto-efficient funds on the efficiency frontier as their value is below 1. The degree of inefficiency is measured as the difference between one and the score. To be on the efficiency frontier funds would need to reduce some of its inputs. However, if the efficiency is 1 and the slack variables are zero, the input level would need to increase to expand the output. Although most funds are efficient (score of 1) as we have seen in [Table 6](#), the average is below this value due to the inefficient ones in each country. The efficiency scores differ depending on the output measure. Gross return and CAPM give higher efficiency scores. The efficiency scores of funds tend to decrease when using multifactor models and when the performance measure is more sophisticated. The efficiency scores, using gross returns as output measure, range from 0.66 for France to 0.88 for Denmark, Canada and Poland. When considering the Carhart model, the efficiency scores range from 0.58 for France and Indonesia to 0.76 for Canada. As with previous results,

Table 6 Percentage of efficient funds

	Gross returns (%)	CAPM (%)	Fama–French (%)	Carhart
Australia	64.0	53.3	55.6	58.8
Austria	74.3	61.9	64.6	68.3
Belgium	84.4	70.3	73.4	77.6
Brazil	72.1	60.2	62.7	66.3
Canada	68.3	56.9	59.4	62.8
Chile	75.4	62.8	65.5	69.3
China	71.7	55.2	62.3	65.9
Denmark	74.6	57.3	64.9	68.6
Finland	85.2	65.5	74.1	78.3
France	69.4	53.3	60.3	65.2
Germany	63.3	48.6	55.3	59.5
Hong Kong	78.4	60.3	68.2	73.6
India	73.1	56.2	63.5	68.7
Indonesia	72.2	60.2	62.8	67.8
Ireland	69.3	57.7	60.2	65.1
Israel	64.3	53.5	56.5	60.4
Italy	75.0	62.5	66.3	71.2
Japan	82.2	68.5	72.3	78.4
Korea (South)	74.3	61.9	65.3	70.5
Luxembourg	77.5	64.5	68.2	73.6
Malaysia	72.2	60.3	63.5	68.5
Netherlands	79.3	66.2	69.7	75.3
New Zealand	67.5	56.2	59.4	64.5
Norway	64.5	53.7	56.7	61.2
Poland	69.9	58.2	62.2	66.4
Portugal	70.1	58.4	62.3	66.5
Singapore	73.4	58.7	65.3	69.7
South Africa	79.4	63.5	70.6	75.4
Spain	83.3	66.6	74.1	79.1
Sweden	81.2	64.9	72.2	77.1
Switzerland	69.7	55.7	62.3	66.2
Taiwan	73.3	58.6	65.2	69.6
Thailand	78.5	62.8	69.8	74.5
United Kingdom	72.3	57.8	64.3	68.6
United States	70.5	56.4	62.7	66.9

This table presents the results of efficient funds, according to the DEA output formulation. *Gross returns* present results using gross returns. *CAPM*, *Fama–French*, and *Carhart* show results using the alpha estimates from their model, respectively. The sample period is from January-1990 to December-2015

we do not find evidence of any relation between the number of funds per country and the efficiency score. We interpret our results of high efficiency scores as the mutual fund industry is a competitive market with investors having access to any fund around the world. Inefficient mutual funds can follow their efficient peer group as they have achieved an efficiency level in

Table 7 Efficiency scores

	Gross returns	CAPM	Fama–French	Carhart
Australia	0.86	0.82	0.78	0.74
Austria	0.82	0.78	0.74	0.71
Belgium	0.78	0.74	0.70	0.67
Brazil	0.68	0.65	0.61	0.60
Canada	0.88	0.83	0.80	0.76
Chile	0.69	0.65	0.62	0.60
China	0.78	0.74	0.70	0.67
Denmark	0.88	0.83	0.80	0.75
Finland	0.82	0.78	0.75	0.70
France	0.66	0.63	0.60	0.58
Germany	0.86	0.82	0.79	0.73
Hong Kong	0.81	0.77	0.74	0.69
India	0.74	0.71	0.68	0.63
Indonesia	0.68	0.64	0.62	0.58
Ireland	0.71	0.67	0.64	0.61
Israel	0.86	0.81	0.77	0.73
Italy	0.69	0.65	0.62	0.60
Japan	0.84	0.79	0.75	0.73
Korea (South)	0.69	0.65	0.62	0.60
Luxembourg	0.74	0.70	0.66	0.64
Malaysia	0.84	0.79	0.75	0.73
Netherlands	0.86	0.82	0.77	0.74
New Zealand	0.68	0.65	0.61	0.61
Norway	0.77	0.73	0.69	0.66
Poland	0.88	0.84	0.79	0.75
Portugal	0.85	0.81	0.76	0.73
Singapore	0.78	0.74	0.70	0.67
South Africa	0.76	0.72	0.68	0.65
Spain	0.77	0.73	0.69	0.66
Sweden	0.79	0.75	0.71	0.67
Switzerland	0.82	0.77	0.75	0.70
Taiwan	0.83	0.78	0.76	0.71
Thailand	0.86	0.81	0.79	0.73
United Kingdom	0.87	0.82	0.80	0.74
United States	0.84	0.79	0.77	0.72

This table presents the mean efficiency scores of the sample of equity funds. *Gross returns* present results using gross returns. *CAPM*, *Fama–French*, and *Carhart* show results using the alpha estimates from their model, respectively. The sample period is from January-1990 to December-2015

the same conditions as the inefficient ones. A challenge for fund managers in increasing the efficiency is to deal with the dilution effect (Greene and Hodges 2002), which is a result of money inflows that funds receive due to their good performance. Large new inflows dilute the

overall performance of the fund until fund managers can efficiently invest this new available cash to match the fund's performance record. Another potential issue to consider in increasing the efficiency scores is related to regulatory obligations and the open-end structure of most funds to have an important part of their assets in cash to timely meet investor redemptions.

We can determine the sources of inefficiency by analyzing the slacks (performance inefficiencies) of the cost variables. The slacks evaluate where funds use resources inefficiently and indicate the degree that each input can be reduced to obtain an efficiency score of one. Panel A of Table 8 shows the average of the absolute slacks, while Panel B presents the relative average slacks (absolute average slack in input divided by the average value of the inputs). Relative slacks are useful to compare the marginal impact of inputs on fund returns across countries. We report the results only for the Carhart output measure. The other output measures present qualitatively similar results and have been omitted for brevity. An interesting result is that the fund risk (standard deviation of returns) presents only small slacks across countries, which supports the idea that most funds are mean–variance efficient. The risk variable shows that fund portfolios are properly diversified and equity funds have eliminated the non-systemic part of their portfolio risk. Turnover and loads present larger slacks, suggesting that more funds are inefficient on these aspects. This is consistent with the fund literature using the DEA model. For instance, [Murthi and Choi \(2001\)](#) show that loads and turnover are the main sources of inefficiencies across all fund categories. The slacks from fund loads suggest that investors should not consider funds that charge any load as this reduces profitability. These slacks vary significantly across countries from 0.075 for Australia to 1.556 for Indonesia. Taiwanese funds are much more inefficient in turnover since their fund managers spend more than double in turnover activity than any other manager. Expense ratios do not show significant slacks, which indicates that funds that charge higher fees appear to earn enough returns to compensate the expenses. This confirms earlier empirical findings of [Ippolito \(1989\)](#), who employs a CAPM model to explain that, in an efficient market, mutual funds are expected to obtain enough risk adjusted return to cover expenses for information gathering.

DEA confirms the mean–variance efficiency theory, which is defined as the ability of a fund to obtain the maximum return for a given level of risk. [Markowitz \(1952\)](#) and [Tobin \(1958\)](#) developed the concept of optional portfolio selection, which states that investors aim to maximize their utility by selecting among the possible mean–variance efficient portfolios according to their risk tolerance. We use a DEA model based on a mean–variance framework, in which the return is the output of the model and the variance of the fund is used as input. Our results confirm previous evidence (see, [Murthi et al. 1997](#)) of the mean–variance efficiency theory for mutual funds around the world, as most funds are on the efficient frontier and the standard deviation shows only small slacks.

Return distribution for some assets might show excess kurtosis (fat tails). The commonly used measure of risk (standard deviation) relies on a normal distribution of returns and this distribution does not account for the fat tails of possible stock market returns. To solve this problem, the DEA model projects the efficient frontier from the sample through non-linear forms of data. This methodology allows examining the mutual fund performance by measuring the distance of the optimal projection from the efficient frontier.

Some authors explain that investors might be more interested in extreme values (skewness and kurtosis) than in central tendencies (standard deviation) as they argue that an investor's expected utility depends positively on skewness and negatively on kurtosis (See, [Scott and Horvath 1980](#)). However, we do not think that there is a significant recent empirical evidence to consider that most investors have these expected utility preferences across countries.

Table 8 Mean slacks in inputs

	Expense ratio	Turnover	Load	Standard deviation
<i>Panel A: Mean slacks in inputs</i>				
Australia	0.045	22.31	0.075	0.005
Austria	0.034	8.37	1.234	0.002
Belgium	0.010	11.23	0.673	0.011
Brazil	0.006	14.56	0.563	0.008
Canada	0.054	18.34	0.842	0.005
Chile	0.109	15.11	1.321	0.013
China	0.034	12.39	1.432	0.000
Denmark	0.076	16.23	1.322	0.006
Finland	0.034	9.56	1.234	0.009
France	0.067	10.75	1.443	0.006
Germany	0.089	14.67	0.456	0.014
Hong Kong	0.031	17.45	0.674	0.017
India	0.042	24.56	1.234	0.005
Indonesia	0.033	31.22	1.556	0.008
Ireland	0.101	19.11	1.234	0.012
Israel	0.045	8.56	1.034	0.000
Italy	0.022	19.21	0.856	0.001
Japan	0.033	11.25	0.367	0.010
Korea (South)	0.065	18.45	1.123	0.007
Luxembourg	0.044	14.37	0.668	0.005
Malaysia	0.054	19.23	0.345	0.009
Netherlands	0.012	45.38	0.677	0.015
New Zealand	0.034	14.23	0.781	0.008
Norway	0.056	14.55	0.892	0.004
Poland	0.045	18.33	1.345	0.018
Portugal	0.071	12.38	1.422	0.000
Singapore	0.012	15.44	0.845	0.003
South Africa	0.034	29.56	0.784	0.009
Spain	0.045	36.56	1.345	0.011
Sweden	0.078	37.65	1.478	0.008
Switzerland	0.028	10.45	1.356	0.017
Taiwan	0.034	47.56	1.341	0.000
Thailand	0.045	24.03	0.856	0.008
United Kingdom	0.040	27.01	0.923	0.005
United States	0.039	12.22	0.665	0.004

Table 8 continued

	Expense ratio	Turnover	Load	Standard deviation
<i>Panel B: Relative slacks (absolute slack/mean value of inputs)</i>				
Australia	0.014	0.202	0.037	0.002
Austria	0.023	0.160	0.398	0.001
Belgium	0.006	0.215	0.336	0.005
Brazil	0.004	0.282	0.281	0.004
Canada	0.037	0.166	0.421	0.002
Chile	0.035	0.137	0.660	0.006
China	0.011	0.112	0.716	0.000
Denmark	0.024	0.147	0.661	0.003
Finland	0.011	0.183	0.617	0.004
France	0.021	0.206	0.721	0.003
Germany	0.028	0.282	0.228	0.007
Hong Kong	0.010	0.337	0.337	0.008
India	0.013	0.331	0.617	0.002
Indonesia	0.010	0.421	0.607	0.004
Ireland	0.032	0.173	0.482	0.005
Israel	0.014	0.077	0.403	0.000
Italy	0.015	0.174	0.334	0.000
Japan	0.022	0.102	0.143	0.004
Korea (South)	0.045	0.247	0.438	0.003
Luxembourg	0.030	0.194	0.260	0.002
Malaysia	0.037	0.369	0.134	0.004
Netherlands	0.008	0.872	0.264	0.006
New Zealand	0.023	0.273	0.305	0.003
Norway	0.038	0.279	0.287	0.001
Poland	0.014	0.352	0.433	0.008
Portugal	0.022	0.238	0.458	0.000
Singapore	0.003	0.208	0.272	0.001
South Africa	0.011	0.399	0.252	0.004
Spain	0.031	0.494	0.433	0.005
Sweden	0.054	0.508	0.476	0.003
Switzerland	0.019	0.141	0.437	0.007
Taiwan	0.024	0.642	0.432	0.000
Thailand	0.031	0.324	0.276	0.003
United Kingdom	0.027	0.363	0.297	0.002
United States	0.027	0.165	0.214	0.001

The table presents the mean of the absolute slacks in Panel A and the relative mean slacks (absolute mean slack in input/mean value of the inputs) in Panel B. The results presented in this table refer to the Carhart output DEA model. *Expense ratio* is the annual fee that funds charge their shareholders. *Turnover* is a measure of the fund trading activity. *Load* is the fund sales charge. *Standard deviation* is a measure of the volatility of a fund. The sample period is from January-1990 to December-2015

4.4 Robustness Analysis

We examine the robustness of our results from the regression and DEA analysis using the methodology of [Fama and MacBeth \(1973\)](#). [Coval and Stafford \(2007\)](#) point out that the estimates from Fama-MacBeth regression are more accurate than OLS results. We check whether our OLS estimates are similar to those from the Fama-MacBeth regression for both the performance-expense regression and the DEA approach.

We estimate the performance-expenses relation in equation (4) and the DEA model using the coefficients of the performance measures obtained from the Fama-MacBeth two-step approach to avoid cross-fund correlation in the residuals as a result of systematic misspecification that affect performance estimates. First, we regress risk-adjusted returns against the factors from the different multifactor models, then we take a time series average of the estimates. We find statistically and qualitatively similar results for the performance-expense regression (listed in [Table 5](#)) and the DEA model ([Tables 6, 7](#)).⁴

We also test whether recessions influence our results about the efficiency of mutual funds. For this purpose, we use a recession variable that is equal to one if the economy at a given day is in a recession as defined by the OECD, and zero otherwise. We include this recession indicator in equation (4) and as an input in the DEA model (equation (11)). We find very similar results to our previous ones across countries. The recession indicator is significantly positive for all countries using equation (4) but does not significantly change the coefficient estimates of the expense variable ([Table 5](#)). For the DEA model, the inclusion of the recession variable only shows slight variation in the efficiency scores ([Table 6](#)) and the percentage of efficient funds ([Table 7](#)). Thus, considering recessions do not present significant differences in mutual fund efficiency across countries.

4.5 Limitations

The results of this research, as in all studies that use the DEA methodology, should be interpreted with precaution. In fact, DEA presents some limitations that might create biased results. The DEA model has high sensitivity to data errors and outliers and takes for granted that the inputs and outputs are free of errors. An issue to consider is that the sample should include a minimum number of funds to obtain reliable estimations. Although the main limitation of DEA is that it can show how a unit is doing in comparison to its peers but it cannot be compared to a theoretical maximum. In this sense, [Miller and Noulas \(1996\)](#) point out that examining whether hedge funds are inefficient compared to others does not result in a maximum output. Another problem of DEA is the ranking of units based on efficiency scores from different DEA models. Ranking of units is a difficult task when several markets are considered as the inputs from the models are affected by different macroeconomic situations. Another nontrivial limitation of using DEA models to measure efficiency is that efficient units do not show any difference using DEA. Additionally, outside influence is not considered to estimate DEA models as efficiency scores are estimated only in comparison to their peers in the sample. [Taylor and Harris \(2004\)](#) state that a problem with DEA is that measures relative efficiency instead of absolute efficiency, as the model assumes that each input and output variable is considered identical.

⁴ We do not report the coefficient estimates here for brevity but they are available from the authors upon request.

5 Conclusion

This paper examines the market efficiency of the mutual fund industry around the world in short-term mutual fund performance. Employing a unique database of worldwide domestic equity funds, the paper uses parametric and non-parametric evaluation approaches where a relation between cost (input variables) and benefit (output measure) is established.

The results show a negative and statistically significant relationship between expenses and risk-adjusted performance across countries after controlling for the effect of volatility, age, net inflow, and fund size. This result indicates that higher expenses are associated with poorer performance. The study reexamines this negative relation, which is usually widely documented in prior literature, with a non-parametric approach (DEA). The advantages of the DEA methodology are that it solves the benchmark problems in multifactor models and helps determine the causes of inefficiency. The DEA technique can consider different factors in the examination of fund performance. It can use different inputs and outputs and does not assume a functional form in the relationship between them. In contrast to our previous result, the use of DEA shows strong evidence that equity mutual funds around the world are approximately mean–variance efficient, which means that are close to the mean–variance efficient frontier. Thus, DEA confirms the mean–variance efficiency theory. We also show the areas of operational inefficiency for funds to improve their performance. Turnover and loads present larger slacks, suggesting that funds could improve their performance on these aspects.

Our results have important implications. First, there is a strong incentive to increase fees among funds in order to improve performance. Thus, our results support the agency theory explanations. Second, our results have practical relevance for potential investors, as they might consider some of the fund's characteristics in their investment decisions.

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