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How does copayment for health care services affect demand, health and redistribution? A systematic review of the empirical evidence from 1990 to 2011

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Abstract This article reviews the quantitative evidence on the behavioural effects of copayment within the health area across a wide range of countries. The review distinguishes itself from previous similar reviews by having a high degree of transparency for the search strategy used to identify the studies included in the review as well as the criteria for inclusion and by including the most recent literature. Empirical studies were identified by performing searches in EconLit. The literature search identified a total of 47 studies of the behavioural effects of copayment. Considering the demand effects, the majority of the reviewed studies found that copayment reduces the use of prescription medicine, consultations with general practitioners and specialists, and ambulatory care, respectively. The literature found no significant effects of copayment on the prevalence of hospitalisations. The empirical evidence on whether copayment for some services, but not for others, causes substitution from the services that are subject to copayment to the 'free' services rather than lower total use is sparse and mixed. Likewise, the health effects of copayment have only been analysed empirically in a limited number of studies, of which half did not find any significant effects in the short term. Finally, the empirical evidence on the distributional consequences of copayment indicates that individuals with low income and in particular need of care generally reduce their use relatively more than the remaining population in consequence of copayment. Hence, it is clear that copayment involves some important economic and political trade-offs.

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Keywords Copayment · Demand effects · Behavioural effects · Health · Distributional consequences

JEL Classification I11 (Analysis of healthcare markets) · I14 (Health and inequality) · I18 (Government policy; regulation, public health)

Introduction and background

Copayment is frequently brought up as a possible solution to the steadily increasing cost pressure faced by many universal health care systems [1]. Economic theory predicts that copayment may under some circumstances reduce the individual demand for health care services by increasing the price paid by the consumer at the time of consumption, while at the same time contribute to the financing. However, it may also give rise to adverse health effects and distributional consequences.

The use of copayment within the health area varies considerably between countries and largely reflects political choice as well as historical circumstances [2]. The empirical literature on the effects of copayment for health care services has previously been reviewed in a number of studies with different areas of focus and using different methods and delimitations. One group of studies have focussed exclusively on the effects of copayment for prescription medicine [3–6]. Others have reviewed the literature on the effects of copayment for various types of health care services [7–9], one of which restricted the review to consider effects among the elderly [10]. While the reviews focussing on the use of prescription medicine are explicit about their search strategy and criteria for inclusion, this is not the case for the reviews including various types of

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health care services and entire populations. Hence, this review distinguishes itself from these reviews by having a high degree of transparency about the search strategy used to identify the studies included in the review as well as the criteria for inclusion. Moreover, given that new studies of the behavioural effects of copayment appear on a frequent basis, an updated review is justified alone by including the most recent literature.

This article reviews the quantitative evidence on the behavioural effects of copayment within the health area across a wide range of countries. The specific aims are to provide a systematic overview of the extent to which copayment (1) reduces the individual demand for the services on which it is imposed, (2) has adverse health effects or shifts the use to services that are not subject to copayment rather than reduce total use and finally (3) gives rise to distributional consequences.¹ Hereby the review contributes not only to the academic literature, but it also has a policy purpose as national decision makers in many European countries are faced with decisions about whether to introduce copayment for new types of services or expand, restructure or phase out existing copayment schemes.

Within the health area, the copayment debate often takes its point of departure in the RAND health insurance experiment (HIE). This section proceeds by summarising the results of the RAND HIE and briefly outlining the theoretical framework for the analysis of the behavioural effects of copayment. The remainder of the article is organised as follows: the "Methods" section describes the search strategy used to identify the empirical literature included in the review and the criteria for inclusion. The "Results" section reviews the results of the identified studies. Finally, the "Conclusions and perspectives" section concludes and discusses the findings in a broader perspective.

The RAND health insurance experiment

The RAND HIE enjoys special status as one of the most comprehensive social experiments ever with respect to both the number of involved individuals and costs. The experiment was launched by the US government with the purpose of investigating mainly how copayment affects the use of health care services. The experiment ran from November 1974 to February 1977 and involved around 6,000 individuals, who were randomly distributed into five different health insurance plans with varying levels of copayment. The insurance plans varied with respect to both average price elasticity of the individual demand for health care services is estimated to be -0.20 across the different types of health care services included in the experiment [11].

More specific results from the RAND HIE found that:

- Copayment reduces the number of visits but not the intensity of the treatment received. This indicates that the effect of copayment is limited once the individual has contacted the relevant health care provider [12].
- Copayment reduces the effect of effective and ineffective treatments to the same extent, and it does not prevent the occurrence of inappropriate hospitalisations [13, 14].
- The effects of copayment vary for different population groups. Copayment reduces the demand for health care services more for low income groups, and it is particularly effective in reducing the use of health care services among low income children [12].
- For the average participant in the experiment copayment does not affect health as measured by a number of indicators. However, there are negative health effects of copayment for individuals with high blood pressure and low income as well as individuals with poor sight [15].

There are a number of reasons in favour of not basing the current debate about copayment in the European health care systems on the RAND results alone. First, the RAND HIE is conducted within an institutional setting differing markedly from the universal health care systems found in most European countries. Second, the copayment schemes that exist alongside universal health care systems are usually more complex than those implemented in the RAND HIE. Third, the RAND HIE does not include the elderly, who account for a large share of the health expenditures in most countries. Finally, the experiment was conducted more than 3 decades ago. Since then, the incidence of a number of lifestyle illnesses has developed, just as the diagnosis and treatment options have changed for many illnesses.

Brief outline of the theoretical framework

The theoretical framework for the analysis of the behavioural effects of copayment takes it point of departure from the demand side. This requires that the use of the services on which copayment is imposed is demand driven. For several services within the health care area, it may be argued that use is also determined by the supply side. It is thus important to keep in mind that copayment is only one among several instruments with potential to lower the use of services and

¹ Given that demand is approximated by actual use in all of the reviewed studies, the terms demand and use are applied interchangeably throughout the article.

the corresponding expenditure within the health area. Visitation and waiting lists as well as mechanisms seeking to affect the suppliers are also capable of lowering the use of health care services. It may, however, be argued that use in most cases requires the presence of some sort of demand. An important point in this regard is that copayment should not be viewed as a substitute for the alternative rationing mechanisms, but rather as a supplement [16]. Hence, when we refer to the demand for health care services in the following, this is done assuming all else equal, i.e. that potential supply-side initiatives as well as prioritisation through visitation and waiting lists are kept constant.

Demand effects

Economic theory predicts that copayment may under some circumstances reduce the individual demand for health care services by increasing the price paid by the consumer at the time of consumption [17, 18]. The magnitude of this effect depends on the price elasticity of demand. Imposed on health care services for which the demand is largely price inelastic, copayment shifts the burden of financing from the public coffers to the users rather than lower the use.² Imposed on health care services for which the demand is price elastic, copayment may be shown to reduce the demand, thereby potentially also lowering the total health care costs.

Health effects and substitution

Even when copayment reduces use of the relevant services, it does not necessarily lower the total health care costs. First, it is possible that copayment for some services within a given area, but not for others, causes substitution from services that are subject to copayment to services that are free or subject to less copayment rather than lower the total use. Second, the introduction of copayment for services, for which the individual demand is price elastic, may lead to a deterioration in public health, if the copayment causes citizens to drop effective and necessary care and treatment. Besides this being inconvenient for the affected individuals, it implies that savings brought along by copayment in one area may be fully or partly cancelled out by increases in the use of other types of services. However, it is also possible that copayment for one type of health care service reduces the use of complementary services.³

Distributional effects

Finally, economic theory suggests that copayment may have distributional consequences [16]. First, it is expected that low income individuals reduce their use of health care services more than the remaining population in consequence of copayment for two reasons. One reason, which is especially expected to affect the use of services accounting for a large share of the budget, is that low income individuals have smaller consumption possibility sets. The other reason is that subject to the standard assumption of decreasing marginal utility of income, the marginal utility of income is higher for low income individuals. This implies that the utility loss associated with copayment is higher for low income individuals, all else equal, and they are therefore expected to reduce their use of the services that are subject to more copayment. Second, individuals with a high demand for services on which copayment is imposed are hit particularly hard when part of the burden of payment is distributed according to use rather than income. Even when the demand is inelastic for low income groups and use is not affected by copayment, there are still distributional consequences because copayment lowers the income left for remaining consumption disproportionally more for low income individuals.

Considering the distributional consequences of copayment, an important point is that the group of individuals with a high demand in the reference point with no copayment is likely to be heterogeneous with regards to the need for health care services as well as the corresponding price elasticity of demand. Some individuals have a limited need for the service in question and in consequence of this their demand is very price sensitive. This group is expected to reduce their use notably in consequence of copayment, which was the exact point of introducing copayment. Another group of individuals have a regular need for the service in question and are thus not expected to reduce their use notably in consequence of copayment. This group, typically made up of chronic patients, is hit particularly hard by the copayment, which may be considered undesirable from an equity point of view. In either case, the distributional consequences of copayment may be mitigated by making the copayment income dependent, exempting groups in particular need of the service in question, or put a ceiling on the annual copayment, after which additional use is paid in full by the public coffers.

Methods

Inclusion criteria

The literature review includes quantitative studies of the behavioural effects of user charges for health care services.

 $^{^2}$ Examples of health care services for which the price elasticity of demand is most likely limited are chemotherapy, bypass operations and other major interventions, which may be lifesaving but pose considerable health risks in themselves.

³ For example, it is possible that increased copayment for prescription medicine reduces the demand for general practice consultations, which often result in a prescription.

Studies published as articles in peer-reviewed journals and as working papers, scientific dissertations,⁴ books and book chapters in English, Danish or Swedish are included. In order to ensure some degree of overall homogeneity across the health care systems and institutional settings considered, we include only studies based on data from high income countries cf. the World Bank [19]. Only original studies, i.e. studies that perform an independent analysis of existing data or data collected specifically for the purpose, are considered. Timewise, the review is restricted to include studies published from 1990 and up until December 2011. Moreover, the focus of the review is on direct copayment, which may take the form of a fixed rate per unit of service, a share of the costs of providing the service or a deductible (i.e. the user pays the full costs of providing the service up until a given threshold) [5], as opposed to indirect copayment in the form of reference or tiered pricing, which is mainly used within the pharmaceutical area.⁵

Finally, studies that have identified the effect of copayment by using variation in the copayment rate between individuals who have purchased private health insurance on a voluntary basis are not included in the present review.⁶ The main reason for this limitation is that the voluntary element in these schemes implies a selection, which is tricky to handle empirically. The level of copayment an individual is facing in such a setting may to a large extent be expected to depend on unobservable individual characteristics, such as risk preferences and expected use of the health care services that are subject to copayment. The methodological challenges associated with handling such selection are judged to necessitate a thorough quality assessment of the identified studies, which is outside the scope of this review.

Search strategy and identification of relevant studies

The reviewed studies were identified by performing electronic searches in the bibliographic database EconLit. The search was restricted to words included in the title, abstract and keywords. The search terms used were [payment* OR copayment* OR co-payment* OR user fee* OR user charge* OR co-finance OR user pays OR cost sharing OR out-of-pocket OR private payment* OR finance OR financing OR funding] AND [health* OR medical OR somatic OR doctor* OR dental OR medicine OR drug* OR primary care OR general practice OR specialist OR physio* OR rehabilitation OR hospital OR ambulatory].⁷ The identification of the relevant studies among the studies identified by the search was based on title and abstract or summary. In case of doubt, the decision was based on a reading of the entire study. The relevance of the studies was assessed based on the inclusion criteria outlined in the previous section and whether they answered at least one of the research questions outlined in the introduction. Subsequently, the reference lists of the studies identified as being relevant were searched for additional literature.⁸

Results

The initial search in EconLit resulted in 3,289 hits. A screening of the many hits resulted in 37 studies being identified as relevant.⁹ Next, the reference lists of the selected studies were searched for additional literature, resulting in 10 additional studies. Hence, the search strategy identified 47 studies of the behavioural effects of copayment in total. It is ensured that the same study is not included more than once, e.g. both in the form of working paper and journal article.

Table 1 provides key information about the reviewed studies, the majority of which are published within the recent decade. The reviewed studies are described with respect to the data, methods, focus of analysis and type of publication.

It is seen from Table 1 that the vast majority of the reviewed studies identify the effects of copayment in a national context. The health care system in Canada provides the background for most of the studies, although these are often restricted to consider individual states, but

⁴ Master Theses are not systematically available from all institutes and faculties and are therefore not included.

⁵ A review of the empirical literature on the consequences of different types of copayment within the pharmaceutical area found that indirect copayment has the ability to shift use from expensive to cheaper types of medicine, although often with modest savings in consequence [5].

⁶ See Atherly [20] for a review of effect studies for the elderly in the US, for whom variation in the level of copayment is mainly cause by voluntary purchase of supplementary private health insurance and Kiil [21] for an overview of effect studies of voluntary private health insurance in universal health care systems.

 $^{7 \}times 10^{-7}$ * Indicates a truncated search, i.e. a search on the stem of the word as well as its different endings.

⁸ Relevant employees at a number of Danish research institutions, ministries and interest groups were also contacted in order to ensure that all relevant literature was included. However, this did not result in the identification of additional studies.

⁹ In consequence of the rather general search terms used, the main reason that many studies were classified as being irrelevant and discarded was that they dealt with something else than copayment. More specific groups of discarded studies include studies that investigate effects of supply-side payment systems, studies based on data collected prior to 1990, studies of indirect copayment, and studies that identified the effect of copayment by using variation in copayment rates between individuals who have voluntarily selected themselves into different schemes or purchased complementary private health insurance.

Table 1 Key information about the reviewed studies of the behavioural effects of user charges

References	Country	Analysis			Data	Pub. type		
		Focus	Design	Method	Туре	Years	Level	
Almarsdóttir et al. [22]	ICE	DE	NE	REG	R-TI	1993–1998	AGG	PA
Atella et al. [23]	ITA	DE/HS	NE	DID	R-PA	1997-2002	IND	PA
Augurzky et al. [24]	GER	DE/EQ	NE	DID	S-PA	2003/2005	IND	WP
Blais et al. [25]	CAN	DE	NE	REG	R-TI	1992-1996	IND	PA
Bolin et al. [26]	EUR	DE	OB	REG	S-CR	2004	IND	PA
Cherkin et al. [27]	USA	DE/EQ	NE	REG	S-RC	1984/1985	IND	PA
Christensen [28]	DEN	DE	EX	DC	S-CR	1990	IND	PA
Christensen [29]	DEN	DE	EX	DC	S-CR	1990	IND	PA
Cockx and Brasseur [30]	BEL	DE	NE	DID	R-RC	1993–1994	AGG	PA
Contoyannis et al. [31]	CAN	DE/EQ	NE	IV	R-PA	1995-1997	IND	PA
Costa-Font et al. [32]	SPA	DE	OB	REG	S-CR	2004	IND	PA
Elofsson et al. [33]	SWE	DE/EQ	OB	REG	S-CR	1995	IND	PA
Fiorio and Siciliani [34]	ITA	DE	NE	DID	R-PA	2001/2003/2006	AGG	PA
Grootendorst and Levine [35]	CAN	DE/EQ	NE	REG	S-RC	1994–1997	IND	WP
Gundgaard [36]	DEN	EQ	OB	IN	S; R-CR	2000-2001	IND	PA
Jordin and Engholm [37]	SWE	DE	NE	DC	R-TI	1996-1999	AGG	PA
Kim et al. [38]	KOR	DE/EQ	OB	REG	S-CR	1998	IND	PA
Kupor et al. [39]	JAP	DE/EQ	OB	REG	S-RC	1984/1989	AGG	PA
Krutilová [40]	CZE	DE	NE	DC	R-TI	2007-2009	AGG	PA
Layte et al. [41]	IRL	DE	NE	DID	S-RC	2000/2004	IND	PA
Li et al. [42]	CAN	DE/EQ/HS	NE	IV	R-PA	1996-2002	IND	PA
Lostao et al. [43]	FRA/GER/SPA	DE/EQ	CC	REG	S-RC	1990/2000	IND	PA
Lostao et al. [44]	FRA/GER/SPA	DE/EQ	CC	REG	S-RC	1990/2000	IND	PA
Lundberg et al. [45]	SWE	DE/EQ	OB	REG	S-CR	1995	IND	PA
McManus et al. [46]	AUS	DE	NE	REG	S-TI	1987–1994	AGG	PA
Nexøe et al. [47]	DAN	DE	EX	DC	R-CR	1995	IND	PA
Nolan [48]	IRL	DE	NE	DID	S-PA	1995–2001	IND	PA
O'Reilly et al. [49]	NIRL/IRL	DE/EQ	CC	REG	S-CR	2003	IND	PA
Or et al. [50]	EUR	EQ	CC	REG	S-CR	1998–2004	IND	WP
Pilote et al. [51]	CAN	DE/EQ/HS	NE	REG	R-PA	1994–1998	IND	PA
Poirier et al. [52]	CAN	DE/EQ	NE	DC	R-PA	1991–1993	IND	PA
Puig-Junoy et al. [53]	SPA	DE/HS	NE	DID	R-PA	2004–2006	IND	WP
Rosen et al. [54]	ISR	DE	NE	DID	R-PA	1997–2001	IND	PA
Rückert et al. [55]	GER	DE/EQ	OB	REG	S-RC	2004–2006	IND	PA
Schreyögg and Grabka [56]	GER	DE/EQ DE/EQ	NE	DID	S-PA	2000-2003	IND	PA
Senicyogg and Grabka [50]	OLK	DL/LQ	ILL.	DID	5-1 A	2005-2005	IND	17
Scott et al. [57]	NZE	EQ	OB	REG	S-CR	1996–1997	IND	PA
Simonsen et al. [58]	DEN	DE/EQ	OB	RD	R-PA	2000–2003	IND	WP
Skipper [59]	DEN	DE/EQ DE/EQ	NE	RD	R-PA	1999–2001	IND	WP
Street et al. [60]	RUS	DE/EQ DE	OB	REG	S-CR	1999–2001	HOU	PA
	CAN	DE/HS		REG		1990	IND	
Tamblyn et al. [61]	BEL	DE/HS DE	NE NE	REG	R-PA R PA	1993–1997 1986–1995		PA Pa
van de Voorde et al. [62]					R-PA		AGG	PA DA
van Doorslaer et al. [63]	EUR/USA	EQ	OB OR	IN IN	S-CR	1980-1993	HOU	PA DA
Wagstaff et al. [64]	EUR/USA	EQ	OB NE	IN DEC	S-CR	1980–1993	HOU	PA
Wang et al. [65]	CAN	HS	NE	REG	R-PA	1997-2004	IND	PA
Winkelmann [66]	GER	DE	NE	DID	S-PA	1995–1996	IND	PA
						1998–1999		

Table 1 continued

References	Country	Analysis	Analysis			Data		
		Focus	Design	Method	Туре	Years	Level	
Zápal [67]	CZE	DE	NE	DID	R-PA	2009	IND	PA
Zhong [68]	CAN	EQ	OB	IN	S-RC	1990 1996–1997 2000–2001	IND	РА

Country: AUS = Australia; BEL = Belgium; CAN = Canada; CZE = Czech Republic; DEN = Denmark; EUR = Europe; FRA = France; GER = Germany; ICE = Iceland; IRL = Ireland; ISR = Israel; ITA = Italy; JAP = Japan; KOR = South Korea; NIRL = Northern Ireland; NZE = New Zealand; RUS = Russia; SPA = Spain; SWE = Sweden; USA = United States of America

Focus: DE = demand effects; EQ = distributional effects (equity); HS = health effects and substitution

Analysis design: EX = experiment; NE = natural experiment; OB = observational study; CC = cross-country comparison

Analysis method: DC = descriptive comparison; DID = difference-in-difference; IN = various concentration and inequity indices; IV = instrument variable estimation; REG = regression; RD = regressions-discontinuity estimation

Data type: S- = survey; R- = register; RC = repeated cross-sections; PA = panel data; TI = times series data; CR = cross-sectional data

Level: AGG = aggregate; HOU = household; IND = individual

Publication type: PA = peer-reviewed article; WP = working paper

also Belgium, Denmark, Ireland, Italy, Spain, Sweden, the USA, Czech Republic and Germany are represented in several studies. Hence, a wide selection of countries is represented. While most of the studies are restricted to consider a single country, the distributional consequences of copayment are analysed across several countries in four studies [26, 50, 63, 64]. Moreover, a few studies use variation in the general copayment levels between countries to identify the effect of copayment through cross-country comparisons [43, 44, 49].

Considering the data and methods of the remaining studies, it is seen from Table 1 that the vast majority are based on individual or household level data, while a few studies use aggregate data in which individual and household-specific characteristics cannot be identified. The studies are distributed evenly across survey and register data. The distributional consequences of copayment are analysed in a handful of studies by estimating various concentration and inequity indices [36, 63, 64, 68]. A large group of studies use panel data combined with natural experiments to identify the effect of copayment through difference-in-difference estimation [23, 24, 34, 48, 53, 54, 56, 66, 67], while two studies implement a difference-indifference estimator using data from natural experiments and repeated cross-sections [30, 41]. Other studies analyse data from natural experiments using instrument variables [31, 42] or discontinuity [59] or kink [58] regression designs to overcome selection problems. Another large group of studies identifies the effects of copayment by estimating various other types of regressions [22, 25, 27, 46, 51, 61, 62, 65] or performing descriptive comparisons [37, 40, 52] based on panel or time series data combined with natural experiments. Three studies analyse data from actual experiments by descriptive comparisons [28, 29, 47], while the remainder identify effects using observational data and various types of regressions [26, 32, 33, 38, 39, 45, 55, 57, 58, 60].

As regards publication type, about one-fourth of the studies are published in the form of working papers at universities or other research institutions, while the rest are published as articles in peer-reviewed journals.¹⁰

The results of the empirical studies are accounted for by topic as follows: The first section reviews the evidence on the presence of demand effects. The next section accounts for the evidence regarding the prevalence of health and substitution effects. The evidence on the distributional consequences of copayment is reviewed in the final section. Significance refers to a 5 % level if nothing else is mentioned. The order of magnitude of the effects are not summarised systematically across the reviewed studies for two reasons. First, the effects are not measured uniformly across the studies. Second, a number of studies do not identify the effect of a specific copayment but consider differences in copayment levels described in general terms.¹¹

Demand effects

Table 2 summarises the results of the identified studies of how copayment affects the use of health care services by type of service. For all types of health care services except hospitalisations, the majority of the reviewed studies find a

¹⁰ Some of the more recent working papers may reasonably be expected to be published in article format at a later point in time.

¹¹ This group of studies includes [22, 26, 32, 33, 37, 43, 44, 49].

Table 2 The effect of copayment on the use of various types of health care services

Type of health care service	Effect	Studies
General practice consultations –		 Bolin et al. [23] (EUR, OB, REG, S-CR, 2004, IND); Cherkin et al. [27] (USA, NE, REG, S-RC, 1984/1985, IND); Cockx and Brasseur [26] (BEL, NE, DID, R-RC, 1993–1994, AGG); Elofsson et al. [29] (SWE, OB, REG, S-CR, 1995, IND); Jordin and Engholm [33] (SWE, NE, DC, R-TI, 1996–1999, AGG); Krutilová [35] (CZE, NE, DC, R-TI, 2007–2009, AGG); Layte et al. [36] (IRL, NE, DID, S-RC, 2000/2004, IND); Lostao et al. [38] (FRA/GER/SPA, CC, REG, S-RC, 1990/2000, IND); Nolan [43] (IRL, NE, DID, S-PA, 1995–2001, IND); O'Reilly et al. [44] (NIRL/IRL, CC, REG, S-CR, 2003, IND); Rückert et al. [50] (GER, OB, REG, S-RC, 2004–2006, IND); van de Voorde et al. [57] (BEL, NE, REG, R-PA, 1986–1995, AGG); Winkelmann [61] (GER, NE, DID, S-PA, 1995–1996/1998–1999, IND)
	0	Augurzky et al. [21] (GER, NE, DID, S-PA, 2003/2005, IND); Schreyögg and Grabka [51] (GER, NE, DID, S-PA, 2000–2003/2005–2006, IND); Zápal [62] (CZE, NE, DID, R-PA, 2009, IND)
	+	Rosen et al. [49] (ISR, NE, DID, R-PA, 1997-2001, IND)
Ambulatory care	_	Kim et al. [34] (KOR, OB, REG, S-CR, 1998, IND); Krutilová [35] (CZE, NE, DC, R-TI, 2007–2009, AGG); Kupor et al. [39] (JAP, OB, REG, S-RC. 1984/1989, AGG)
Hospitalisations	_	Kupor et al. [39] (JAP, OB, REG, S-RC. 1984/1989, AGG)
	0	Krutilová [35] (CZE, NE, DC, R-TI, 2007–2009, AGG); Lostao et al. [38] (FRA/GER/ SPA, CC, REG, S-RC, 1990/2000, IND)
Prescription medicine	_	 Almarsdóttir et al. [19] (ISL, NE, REG, R-TI, 1993–1998, AGG); Atella et al. [20] (ITA, NE, DID, R-PA, 1997–2002, IND); Blais et al. [22] (CAN, NE, REG, R-TI, 1992–1996, IND); Contoyannis et al. [27] (CAN, NE, IV, R-PA, 1995–1997, IND); Costa-Font et al. [28] (SPA, OB, REG, S-CR, 2004, IND); Fiori and Siciliani [30] (ITA, NE, DID, R-PA, 2001/2003/2006, AGG); Grootendorst and Levine [31] (CAN, NE, REG, S-RC, 1994–1997, IND); Krutilová [35] (CZE, NE, DC, R-TI, 2007–2009, AGG); Li et al. [37] (CAN, NE, IV, R-PA, 1996–2002, IND); Lundberg et al. [40] (SWE, OB, REG, S-CR, 1995, IND); McManus et al. [41] (AUS, NE, REG, S-TI, 1987–1994, AGG); Puig-Junoy et al. [48] (SPA, NE, DID, R-PA, 2004–2006, IND); Simonsen et al. [53] (DEN, OB, RD, R-PA, 2000–2003, IND); Skipper [54] (DEN, NE, RD, R-PA, 1999–2001, IND); Street et al. [55] (RUS, OB, REG, S-CR, 1996, HOU); Tamblyn et al. [56] (CAN, NE, REG, R-PA, 1993–1997, IND)
	0	Pilote et al. [46] (CAN, NE, REG, R-PA, 1994–1998, IND); Poirier et al. [47] (CAN, NE, DC, R-PA, 1991–1993, IND)
Prevention	-	Christensen [24] (DEN, EX, DC, S-CR, 1990, IND); Christensen [25] (DEN, EX, DC, S-CR, 1990, IND); Nexøe et al. [42] (DEN, EX, DC, R-CR, 1995, IND)

Country: AUS = Australia; BEL = Belgium; CAN = Canada; CZE = Czech Republic; DEN = Denmark; EUR = Europe; FRA = France; GER = Germany; ICE = Iceland; IRL = Ireland; ISR = Israel; ITA = Italy; JAP = Japan; KOR = South Korea; NIRL = Northern Ireland; RUS = Russia; SPA = Spain; SWE = Sweden

Analysis design: EX = experiment; NE = natural experiment; OB = observational study; CC = cross-country comparison Analysis method: DC = descriptive comparison; DID = difference-in-difference; IN = various concentration- and inequity indices; IV = instrument variable estimation; REG = regression; RD = regressions-discontinuity estimation

Data type: $S_{-} =$ survey; $R_{-} =$ register; RC = repeated cross-sections; PA = panel data; TI = times series data; CR = cross-sectional data Level: AGG = aggregate; HOU = household; IND = individual

negative effect of copayment. This indicates that a higher level of copayment generally reduces the use of health care services, while a lower level increases use.

Table 3 summarises the results of the studies that have estimated price elasticities of the demand for different types of health care services. The estimated price elasticities are all negative and numerically less than one, which means that the individual demand drops by less than 1 % following a 1 % increase in the price paid by the consumer at the time of consumption. Other aspects of the studies included in Table 3 are discussed in the relevant subsections, which are organised by type of health care service.

General practice

The three studies that use variation in the general copayment levels between countries to identify the effect of copayment through cross-country comparisons all find a negative

Table 3 Estimates of the price elasticity of demand for various types of health care services

Study	Population group	Type of health care service	Price elasticity of the demand
Cockx and Brasseur [26] (BEL, NE, DID, R-RC, 1993–1994,	Low income	GP	-0.06 for men -0.01 for women
AGG)		HOME	-0.18 for men
			-0.08 for women
		SPEC	-0.14 for men
			-0.02 for women
Contoyannis et al. [27] (CAN, NE, IV, R-PA, 1995–1997, IND)	Elderly chronic patients	MED	From -0.12 to -0.16
Kim et al. [34] (KOR, OB, REG,	Users of ambulatory care	AMB (hospital)	-0.09
S-CR, 1998, IND)		AMB (clinic)	-0.18
Li et al. [37] (CAN, NE, IV, R-PA, 1996–2002, IND)	Total population	MED	-0.11 (-0.20 for low income)
Simonsen et al. [53] (DEN, OB, RD, R-PA, 2000–2003, IND)	Users of prescription medicine	MED	From -0.08 to -0.25 (dep. on model spec.)
Skipper [54] (DEN, NE, RD, R-PA, 1999–2001, IND)	Users of penicillin	MED (penicillin)	From -0.18 to -0.35 (dep. on model spec.)
van de Voorde et al. [57] (BEL,	Total population	GP	From -0.12 to -0.16
NE, REG, R-PA, 1986–1995,		HOME	From -0.20 to -0.39
AGG)		SPEC	-0.10

Type of health care service: AMB = ambulatory care; GP = general practice consultations; HOME = medical home visits; MED = prescription medicine; SPEC = specialist consultation

Country: BEL = Belgium; CAN = Canada; DEN = Denmark; KOR = South Korea

Analysis design: NE = natural experiment; OB = observational study

Analysis method: DID = difference-in-difference; IV = instrument variable estimation; REG = regression; RD = regressions-discontinuity estimation

Data type: S- = survey; R- = register; RC = repeated cross-sections; PA = panel data; CR = cross-sectional data

Level: AGG = aggregate; IND = individual

association between the copayment level and the use of general practice [43, 44, 49], as do the studies using observational data and regression [26, 33, 55]. However, one of these studies finds that the effect tails off about 2.5 years after the introduction of copayment [55]. Moreover, it is found that the tendency to refrain from seeking medical help due to copayment is more pronounced among younger people [33, 55]. One study was based on a cross-country comparison [49] and two of the studies used observational data and regression to measure changes in the use of general practice as perceived by the individual (as opposed to actual use) [33, 55].

The majority of the reviewed studies use natural experiments combined with panel [24, 48, 54, 56, 62, 66, 67], repeated cross-section [27, 30, 41] or time series [37, 40] data to identify the effect of copayment for general practice on the use. The results of these studies are mixed: the majority find evidence of a negative demand effect of copayment [27, 30, 37, 40, 41, 48, 62, 66], some find no effect [24, 41, 56, 67], while one study finds that copayment leads to an increase in use [54]. The one study that finds an increase in the number of doctor visits following the introduction of copayment notes that this is most likely due to some structural circumstances related to the payment of suppliers [54].¹² The two studies using aggregate data and descriptive comparisons to identify the effect of an increase in copayment caused by a reform both find evidence of a negative association between the level of copayment and the use of general practice [37, 40]. Moreover, the effect is found to be relatively larger for home visits [40].

Likewise, the two studies that identify the effect of copayment using a sizeable increase in the copayment for office visits to general practitioners and specialists as well as medical home visits in Belgium both find a negative effect of copayment on the average number of all types of visits per year, but disagree on the order of magnitude cf. Table 3 [30, 62].¹³ While one of these studies argues that

¹² Specifically, while there is an upper limit for the total earnings of doctors, they are allowed to exceed the threshold with income stemming from copayment. This structure may unintentionally have created an incentive for the doctors to seek to increase the visit rate among the patients who are required to copay in order to increase total earnings.

¹³ Cockx and Brasseur [30] adjust the estimates reported in van de Voorde et al. [62] by taking into account that the increases in the relative copayments for the three types of doctor visits considered were not proportional and decomposing the demand effect into an income effect and a substitution effect.

the price elasticity of demand cannot be estimated for the upper income groups based on the observed change in copayment and limits itself to report elasticity estimates for the lower income groups [30], the other reports elasticity estimates for the total population [62]. Along a similar line, two studies analyse the effect of copayment for general practice based on natural experiments within the setting of the Irish health care system [41, 48]. One study exploits variation in the level of copayment that occurs because individuals with income below a certain threshold are exempted from copayment and find that this shift in the level of copayment affects the use of general practice significantly [48]. Moreover, this study finds that the negative effect of being charged copayment due to an increase in income is larger than the positive effect of being exempted due to low income. The other study uses exogenous variation stemming from the fact that individuals aged 70 and above are exempted from copayment regardless of their income and finds that while this increases the probability of having at least one annual GP visit significantly, it does not affect the frequency of visits [41].¹⁴

On the contrary, two studies based on a natural experiment within the setting of the German health care system find no significant effect of introducing a small copayment for visits to general practitioners on neither the probability of having at least one visit nor the number of visits within a 3-month period [24, 56]. Both of these studies use a control group of privately insured, for whom no copayment is charged.

Finally, two studies use the fact that general practice visits often result in prescriptions [66, 67]. One of these studies finds that exempting children for copayment for general practice does not affect use approximated by number of prescriptions [67]. The other study finds that a sharp increase in copayment for prescription medicine curbs the demand for general practice and argues that this may be explained by the fact that the increased copayment for prescription medicine increases the average total costs associated with visits to the doctor [66].

Ambulatory care and hospitalisations

The four studies that analyse the effect of copayment on the use of ambulatory care and hospitalisations, respectively, find that while copayment reduces the use of ambulatory care [38–40], hospitalisations are less affected [39, 40, 43]. One study uses an increase in the general copayment level caused by a reform to identify the effect of copayment on

the use of a broad selection of health care services [40]. This study finds that hospitalisations are the health care service that is the least affected by copayment, while the use of ambulatory care in the form of visits to accident and emergency departments is reduced. Another study uses aggregate data from Japan's 47 prefectures, which differ in terms of copayment levels [39]. This study finds that while copayment reduces the use of both hospital (inpatient) care and ambulatory (outpatient) care, the use of ambulatory care is most sensitive to the copayment rate [39]. The study conducted within the setting of the South Korean health care system finds that while the individual demand for ambulatory care drops in consequence of increasing copayment, the demand for ambulatory care provided in a hospital setting is less price sensitive than the corresponding demand for ambulatory care at clinics [38]. Moreover, this study also finds that low income patients reduce their use of ambulatory care provided in hospitals as well as clinics relatively more than high income patients in consequence of copayment, cf. Table 3.

Prescription medicine

About one-third of the studies considering the effect of copayment on the use of prescription medicine estimate effects for the entire adult population [22, 34, 40, 45, 46, 58, 60]. With the exception of [22], these studies all find that copayment has a negative effect on use [34, 40, 45, 46, 58, 60]. One study considers hypothetical changes in the use of prescription medicine as perceived by the individual [45], while the remaining studies are based on data on actual use.

Another third of the studies exploit a shift in the level of copayment for prescription medicine at the transition from adult to elderly to identify demand effects of copayment [32, 35, 53, 61]. These studies unambiguously find that the drop in copayment levels around the age of retirement brings along an increase in the use of prescription medicine. One study also estimates demand effects for welfare recipients and finds that they reduce their use of prescription medicine relatively more than the elderly in consequence of copayment [35, 61].

The final third of the reviewed studies have estimated effects for various groups with a documented need for medicine due to illness. The results of this group of studies are mixed: some find no effect [51, 52], other find clear evidence of a negative demand effect of copayment [31, 42, 59], and yet other find that effects vary across different groups of patients and medicine types [23, 25]. Considering first the two studies that find no effect of copayment for prescriptions on the use of prescription medicine used

 $^{^{14}}$ The authors note that the fact that general practitioners receive a fixed fee for patients aged 70 and above creates an incentive for the general practitioners to seek to limit the number of visits for this patient group.

to treat hypertension and anxiety, respectively, among the elderly users of these two types of medicine [52]. The other study focusses on the probability of complying with medical advice on medication for elderly who had a coronary [51]. Among the group of studies that find clear evidence of a negative demand effect of copayment for prescription medicine, one analyses the use of prescription medicine for elderly patients with rheumatoid arthritis [42], another considers a broader group of elderly chronic patients [31], while the most recent study analyses the use of insulin used to treat diabetes and penicillin mainly used to treat non-chronic illness [59]. In addition to a negative demand effect of copayment, the study of insulin users also finds that this patient group reacts to the price change by stockpiling medicine immediately before the rise in the copayment level [59]. The remaining two studies find varying effects across different groups of patients and medicine types. One study of hypertension patients treated with ACE-inhibitors finds that an abolition of copayment for prescriptions brings along an increase in compliance with medical advice on medication for the group of patients who do not take enough medicine to begin with, while the behaviour of the remaining patients is unaffected [23]. Another study of adult welfare recipients finds that an increase in the copayment level causes a significant drop in the use of inhaled corticosteroids used to treat asthma, while the use of neuroleptics used to treat schizophrenia and anticonvulsants used to treat epilepsy is not significantly affected [25]. Along a similar line, some studies find that copayment reduces the use of nonessential medicine relatively more than the use of essential medicine prescribed to specific diagnosis groups with the purpose of preventing deterioration in health and prolonging life [35, 46, 58, 61].

Overall, the empirical evidence thus indicates that the extent to which copayment affects the use of prescription medicine depends on the type of medicine as well as the part of the population in focus.

Prevention

The three studies based on controlled randomised experiments all find that copayment decreases the use of preventive health care services. One of these studies analyses the probability of getting a flu vaccination [47], while the other two consider prevention in the form of screening for ischaemic heart disease [28, 29].

Health effects and substitution

Table 4 summarises the results of the sparse literature that investigates whether copayment leads to deterioration in the public health and the extent to which copayment for some services within a given area, but not for others, causes substitution from the services that are subject to copayment to the 'free' services.

Prescription medicine

All studies included in Table 4 focus on the consequences of copayment for prescription medicine. Overall, the studies that tend to find negative effects of copayment on

Table 4 Health effects of copayment and substitution to other health care services

Type of health care service subject to copayment	Study	Demand effect (from Table 2)	Effect on mortality	Substitution to other types of health carservices	
Prescription medicine	Atella et al. [23] (ITA, NE, DID, R-PA, 1997-2002, IND)	_	+	HOS	+
	Li et al. [42] (CAN, NE, IV, R-PA, 1996-2002, IND)	_		GP	+
	Pilote et al. [51] (CAN, NE, REG, R-PA, 1994–1998, IND)	0	0	GP; HOS; AE	0
	Puig-Junoy et al. [53] (SPA, NE, DID, R-PA, 2004–2006, IND)	_		HOS	0
	Tamblyn et al. [61] (CAN, NE, REG, R-PA, 1993–1997, IND)				
	Essential	_	+	HOS; AE; LTC	+
	Non-essential	_	0	HOS; AE; LTC	0
	Wang et al. [65] (CAN, NE, REG, R-PA, 1997-2004, IND)			GP; HOS; LTC	0

Type of health care service: AE = accident and emergency department; GP = general practice consultation; HOS = hospitalisation; LTC = long-term care; MED = prescription medicine

Country: CAN = Canada; ITA = Italy; SPA = Spain

Analysis design: NE = natural experiment

Analysis method: DID = difference-in-difference; IV = instrument variable estimation; REG = regression

Data type: R- = register; PA = panel data

Level: IND = individual

the use of prescription medicine find a positive effect of copayment on mortality and substitution to other types of health care services.¹⁵

The majority of the reviewed studies consider the consequences of copayment for prescription medicine specific groups of elderly Canadians. Two of these studies find no evidence of either health effects or substitution for elderly who had a coronary [51] and users of antidepressants [65], respectively. The other two studies find evidence of substitution from medicine to other types of health care services for elderly patients with rheumatoid arthritis [42] and to some extent also for elderly welfare recipients [61]. More specifically, for elderly welfare recipients it is found that a relatively modest drop in the use of essential medicine following an increase in copayment leads to an increase in mortality as well as the use of accident and emergency departments, hospitalisations and long-term care, while a larger drop in the use of non-essential medicine does not affect any of these outcomes [61].

The two remaining studies are based on data from Italy and Spain, respectively. One study of Italian patients with hypertension finds that the probability of complying with medical advice on medication increases in the absence of copayment for the group of patients who do not take enough medicine to begin with and that this brings along a drop in mortality and hospitalisations [23]. The other study finds that exempting the elderly in Spain for copayment for prescription medicine does not reduce the number of hospitalisations significantly [53].

Distributional effects

The empirical literature takes two different approaches in investigating the distributional consequences of copayment.

The first approach involves estimating various concentration and inequity indices of the financing and use of health care services and decomposing the total degree of inequity into various factors. The two studies considering the financing side both find that individuals in the lower end of the income distribution pay a relatively large share of their income in the form of copayment in all of the countries considered, i.e. the financing of health care services through copayment is regressive [63, 64]. Likewise, the two studies of use both find evidence that copayment brings about horizontal inequity in the use of health care services. One study finds that while the use of health care services is generally not associated with income, there is horizontal inequity in favour of high income individuals in the use of dental services and prescription medicine in a setting where both of these services are characterised by a high level of copayment [36]. The other study investigates the distributional consequences of copayment by comparing the degree of inequity in the use of prescription medicine over a number of years with varying copayment levels. Horizontal inequity indices indicate that while the degree of inequity is generally lower among the elderly (who are facing a lower copayment level than the remaining population), it increases when the elderly are charged higher copayments [68]. Finally, Or et al. [50] analyse the effect of various characteristics of the health care system on social inequities in the use of health care services across a number of European countries and find that copayment as a share of the total health expenditures has a small but statistically significant effect on the degree of social inequity in the use of general practice and specialist consultations.

The second approach is to compare the effect of copayment across different population groups in order to reveal whether vulnerable population groups reduce their use of health care services more than the general population in consequence of copayment. Approximately one-third of the demand studies summarised in Table 2 have taken this approach and extended their analyses to also consider the distributional consequences of copayment. Table 5 summarises the results of these studies with regard to the association between the price sensitivity of demand and various measures of vulnerability. A positive association between the price sensitivity of demand and low income, education level, social status or bad health, respectively, means that individuals with these characteristics reduce their use of health care services relatively more in consequence of copayment. Social status is defined by occupation and health is measured by self-perceived health and the prevalence of chronic conditions in the reviewed studies.

General practice

The majority of the reviewed studies find that vulnerable population groups reduce their use of general practice more than the general population in consequence of copayment [33, 43, 44, 49, 55, 57]. Two of the studies that do not find any differences in use between different population groups [24, 56] are characterised by not finding any demand effects of copayment for either the general population or vulnerable subgroups. The third study [27] considered a population of state government and higher education employees, a group that enjoys moderate incomes, and therefore does not provide any information on individuals with very low incomes.

Prescription medicine

A similar pattern applies to the studies of prescription medicine, where the majority of the reviewed studies also

¹⁵ With the exception of Puig-Junoy et al. [53].

Type of health care service	Study		Association between the price sensitivity of demand and:			
		Low income	Low education level	Low social status	Bad health	
General practice	Augurzky et al. [24] (GER, NE, DID, S-PA, 2003–2005, IND)	0	0			
consultations	Cherkin et al. [27] (USA, NE, REG, S-RC, 1984/1985, IND)	0				
	Elofson et al. [33] (SWE, OB, REG, S-CR, 1995, IND)	+			+	
	Lostao et al. [43] (FRA/GER/SPA, CC, REG, S-RC, 1990/2000, IND)			+		
	Lostao et al. [44] (FRA/GER/SPA, CC, REG, S-RC, 1990/2000, IND)		+			
	O'Reilly et al. [49] (NIRL/IRL, CC, REG, S-CR, 2003, IND)	+			+	
	Rückert et al. [55] (GER, OB, REG, S-RC, 2004–2006, IND)	+				
	Schreyögg and Grabka [56] (GER, NE, DID, S-PA, 2000-2003/2005-2006, IND)	0			0	
	Scott et al. [57] (NZE, OB, REG, S-CR, 1996-1997, IND)	+				
Ambulatory care	Kim et al. [38] (KOR, OB, REG, S-CR, 1998, IND)	+				
	Kupor et al. [39] (JAP, OB, REG, S-RC, 1984/1989, AGG)	0				
Hospitalisations	Kupor et al. [39] (JAP, OB, REG, S-RC, 1984/1989, AGG)	+				
Prescription	Contoyannis et al. [31] (CAN, NE, IV, R-PA, 1995-1997, IND)	_			-	
medicine	Li et al. [42] (CAN, NE, IV, R-PA, 1996-2002, IND)	+				
	Lundberg et al. [45] (SWE, OB, REG, S-CR, 1995, IND)	+	+		+	
	Pilote et al. [51] (CAN, NE, REG, R-PA, 1994-1998, IND)			0		
	Poirier et al. [52] (CAN, NE, DC, R-PA, 1991-1993, IND)	0				
	Simonsen et al. [58] (DEN, OB, RD, R-PA, 2000-2003, IND)	+	+			
	Skipper [59] (DEN, NE, RD, R-PA, 1999–2001, IND)	+				

Table 5 The effect of copayment on the use of health care services for vulnerable population groups

Country: CAN = Canada; DEN = Denmark; FRA = France; GER = Germany; IRL = Ireland; JAP = Japan; KOR = South Korea; NIRL = Northern Ireland; NZE = New Zealand; SPA = Spain; SWE = Sweden

Analysis design: NE = natural experiment; OB = observational study; CC = cross-country comparison

Analysis method: DC = descriptive comparison; DID = difference-in-difference; IV = instrument variable estimation; REG = regression; RD = regressions-discontinuity estimation

Data type: S_{-} = survey; R_{-} = register; RC = repeated cross-sections; PA = panel data; CR = cross-sectional data Level: IND = individual

find that copayment leads to a larger reduction in the use of prescription medicine for vulnerable population groups than for the general population [42, 45, 58, 59]. Two studies do not find any demand effects of copayment neither for a population of elderly users of various types of prescription medicine nor for vulnerable subgroups [51, 52]. Finally, a single study finds the use of prescription medicine to be less price sensitive for low income individuals and individuals in bad health than for the remaining population [31]. This result runs counter to the results of the remaining empirical literature.

Ambulatory care and hospitalisations

The study of ambulatory care conducted within the setting of the South Korean health care system finds that low income patients reduce their use of ambulatory care provided in hospitals as well as clinics relatively more than high income patients in consequence of copayment [38]. Another study based on aggregate data from Japan finds that low income patients reduce their use of hospital (inpatient) care more than the remaining population in consequence of copayment, while there are no systematic differences for ambulatory (outpatient) care [39].

Conclusions and perspectives

This article reviews the quantitative evidence on the behavioural effects of copayment for different types of health care services across a wide range of countries. The present review distinguishes itself from previous reviews, also including studies of various types of health care services and entire populations, by having a high degree of transparency about the search strategy used to identify the studies included in the review as well as the criteria for inclusion, and by including the most recent literature.

The main results of the review can be summarised and discussed as follows: Considering the demand effects, the majority of the reviewed studies find that copayment reduces the use of prescription medicine, consultations with general practitioners and specialists, and ambulatory care, respectively, thereby potentially lowering the total health care costs. In general the literature finds no significant effects of copayment on the prevalence of hospitalisations, which implies that copayment for this type of treatment mainly shifts the burden of financing from the public coffers to the users rather than reduces demand. One of several reasons why the demand for consultations with general practitioners reacts more to copayment than hospitalisations could be due to differences in accessibility. For example, the finding that the demand for hospital care is largely price inelastic probably has something to do with the fact that access is usually restricted by some sort of visitation, either through general practice or on site. This implies that hospital patients have a regular need for care and therefore do not reduce their use notably in consequence of copayment. Along a similar line, another explanatory factor is that contacts to general practitioners are usually initiated by the patient, whereas hospitalisations to some extent may be argued to be a providerinitiated service, over which the patient has far less control [69].

The empirical evidence on whether copayment for some services, but not for others, causes substitution from the services that are subject to copayment to the 'free' services rather than lower total use is sparse and mixed. Likewise, the health effects of copayment have only been analysed empirically in a limited number of studies, of which half did not find any significant effects in the short term. This result is, however, restricted by the fact that the reviewed studies primarily consider effects on mortality, which is a rather coarse measure of health. It can thus not be ruled out that copayment affects health as measured by more nuanced indicators. As regards copayment for prescription medicine, one study finds that copayment reduces the use of non-essential medicine relatively more than the use of essential medicine [61]. However, most studies are unable to say anything about whether the eliminated use is necessary from a medical point of view. Moreover, none of the data sets used in the empirical literature spans a sufficient sequence of years to be able to make inference about the long-term consequences.

Finally, the empirical literature confirms almost unambiguously that distributional consequences are a real matter of concern and not just a theoretical construct. On the financing side, the reviewed studies find that copayment represents an unequal type of financing vertically as well as horizontally. The empirical evidence likewise indicates that vulnerable groups, including individuals with low income and in particular need of care, reduce their use relatively more than the remaining population in consequence of copayment. This highlights the need for proparticularly vulnerable groups tecting from the consequences of copayment, e.g. by making the copayment income dependent or exempting groups in particular need of the services in question. The strategy of exempting certain groups from copayment suffers from the weakness that it may result in rather arbitrary dividing lines [69]. Moreover, exempting vulnerable groups from copayment implies that the burden of financing primarily falls on the working population, who already contribute considerably to the health system through their tax payment. If the tax on labour is held constant and only the working population is required to copay, this group will experience an increase in their total payment towards the health system, which may over time reduce their support for the public system. Another possibility is to encourage people to purchase supplementary private health insurance policies that cover the copayment. This, however, only benefits those who can afford to pay the insurance premium. Moreover, economic theory predicts that supplementary private health insurance counteracts any demand effects of copayment that may occur by lowering the price paid by the consumer.

Hence, the results of the empirical literature indicate that introducing copayment for new types of health care services or extending existing schemes involves some important economic and political trade-offs. It is thus important for the responsible policymakers to be clear about what they wish to obtain by introducing copayment and how to counter the possible side effects.

Considering the welfare consequences of copayment, economic theory predicts that under perfect competition individuals who are charged the full cost consume until the point where the marginal cost of producing the service equals their willingness to pay given preferences and income. Based on this line of reasoning, it may be argued that copayment brings the level of use closer to the socially optimal level than tax funding. This argument, however, presupposes that individuals are able to value the benefits and disadvantages of use, which is often not the case for health care services. Moreover, market solutions are not always optimal within the health area, where the possible presence of information problems, externalities and myopic preferences may well mean that the optimal use for the individual does not equal the optimal use for society as a whole. Adding further complications is the fact that the need for and thus also the individual demand for health care services is characterised by uncertainty, which is an important reason why these services were financed through taxes and social insurance systems in the first place [17].

Hence, the case of full copayment should not necessarily be used as benchmark for efficiency within the health area. Another issue that should also be taken into account in an overall assessment of the distributional consequences of copayment is that the alternative to copayment, i.e. funding through taxes or social insurance systems, has been shown theoretically as well as empirically to redistribute income from high to low income groups [70, 71].

Other considerations than equity and efficiency also apply to general health policy [16]. For one thing, the cost and feasibility of administering a given copayment scheme is an important consideration. In this regard it is worth noting that the complexity and costs of administration are likely to increase if efforts are made to preserve equity by exempting vulnerable populations from payment. Another issue that may constitute an important constraint on implementing and changing copayment arrangements is the question of public acceptability. A survey of public attitudes in several countries indicates that copayment is a hotly debated political issue and illustrates a range of contrasting public attitudes [16].

Finally, it is important to keep in mind that copayment is only one of several instruments with potential to lower the use of services and the corresponding expenditure within the health area. Visitation and waiting lists as well as mechanisms seeking to affect the suppliers, such as reimbursement schemes, are also capable of lowering the use of health care services, which often involves a complex interaction of factors [16, 69]. This means that politicians contemplating copayment as a means to curb the increasing health expenditures also need to consider the incentives facing, e.g. the pharmaceutical industry, hospitals, general practitioners, specialists, nursing homes and so on, given that these actors also have a share in the responsibility for decision making on accessibility, use and expenditures related to the health system. In an overall reform perspective, it is thus of crucial importance that copayment is seen in connection with other possible steering instruments.

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