#### **ORIGINAL PAPER**



# **A new diagram for performance evaluation of complex models**

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#### **Abstract**

In the latest years the capacity and complexity of climate and environmental modeling has increased considerably. Therefore, tools and criteria for model performance evaluation are needed to ensure that diferent users can beneft from model selection. Among graphical tools, Taylor's diagram is widely used to provide evaluation and comparison of model performances, with particular emphasis on climate models. Taylor's diagram accounts for diferent statistical features of model outputs and observations, including correlation, variability and centered root mean square error. Not included is model bias, which is an essential feature for climate model evaluations, and it is usually calculated separately to complement the information embedded in Taylor's diagram. In this paper a new diagram is proposed, referred to as Aras' diagram, which allows for visual assessments of the correspondence between model outputs and reference data in terms of total error, correlation, as well as bias and variability ratios through an easy-to-interpret two-dimensional (2D) plot, allowing for proper weighting of diferent model features. The strengths of the new diagram are exemplifed in a case study of performance evaluation of EURO-CORDEX historical experiment over Southern Italy using E-OBS as reference dataset, for three hydrological variables (i.e. daily precipitation, daily surface minimum temperature, and daily maximum surface temperature), and four popular climate indices (i.e. total annual precipitation, annual maxima of daily precipitation, annual minima of daily minimum temperatures, and annual maxima of daily maximum temperatures). The proposed diagram shows interesting properties, in addition to those already included in Taylor's diagram, which may help promoting climate model evaluations based on their accuracy in reproducing the climatological patterns observed in time and space.

Keywords Bias ratio · Variability ratio · Correlation · Climate models · Kling Gupta efficiency · Taylor's diagram · Performance evaluation · EURO-CORDEX experiment · E-OBS · Aras' diagram

# **1 Introduction**

Gaining understanding of natural phenomena has been crucial for advancing knowledge and supporting societal development. Models are vital tools towards this aim, as they allow for simulation of the dynamics of natural processes in both space and time (see e.g. Chaulya and Prasad [2016](#page-18-0)). Under this setting, it is important to evaluate model performances by comparing model outputs with observations (see e.g. Legates and McCabe [1999;](#page-19-0) Flato et al. [2013;](#page-18-1) Paul

et al. [2023\)](#page-19-1). Typically, this step can be carried out using qualitative and quantitative measures supported by the application of criteria and benchmarks for model selection (see e.g. Moriasi et al. [2012](#page-19-2)). The evaluation of models is often an iterative process, in which the model is calibrated and tested against observed data, and then refned and tested again. Sensitivity and uncertainty analyses are also important components of performance evaluation, as they provide insights into the robustness and reliability of model predictions. However, this should be considered only a module of model evaluation, as the latter should be complemented by a detailed investigation of the accuracy of the model structure in reproducing the diferent processes involved (see e.g. Gupta et al. [2008](#page-18-2); Knutti [2010;](#page-19-3) Biondi et al. [2012](#page-18-3); Kaleris & Langousis [2017\)](#page-18-4). Qualitative analysis can involve the use of graphical tools (see e.g. Kundzewicz & Robson [2004\)](#page-19-4), as graphical measures may allow for direct or adjusted comparisons, depending on the transformations imposed to data

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(see e.g. Moriasi et al. [2015](#page-19-5)). Quantitative assessment of model performances, instead, relies on computing statistical measures, aimed at providing an evaluation of the goodnessof-ft of model outputs to observations (see e.g. Ritter & Muñoz-Carpena [2013](#page-19-6)). To this end, a number of criteria has been developed, used, and critically reviewed in the scientifc literature, highlighting the importance of this topic (see e.g. Moriasi et al. [2015](#page-19-5); Krause et al. [2005\)](#page-19-7) as well as the distance from reaching a rigorous procedure for model performance assessment (see e.g. Ritter and Muñoz-Carpena [2013](#page-19-6); Baker and Taylor [2016\)](#page-18-5).

Several quantitative statistical metrics can be used for evaluating the performance of models, including: (1) Bias, which measures the average diference between the model's output and observed data; (2) correlation, (3) root mean square error (RMSE), which is the square root of the average of the squared diferences between the model's output and the observed data; (4) the Nash–Sutcliffe efficiency (NSE, Nash and Sutclife [1970\)](#page-19-8), which is a measure of how well the model predicts the observed data relative to the mean of the observed data;  $(5)$  the Kling–Gupta Efficiency (KGE), proposed by Gupta et al. ([2009](#page-18-6)), a measure of how well a model simulates the observed data, including explicit evaluation of correlation, variability, and bias. The latter study also showed that both NSE and RMSE provide an evaluation of model performance that is conditionally biased by correlation. Evidently, the combined use of diferent metrics should be retained as the most appropriate way to approach model selection (see e.g. Flato et al. [2013;](#page-18-1) Moriasi et al. [2015](#page-19-5)).

On these grounds, a practical and easy way to compare multiple models is by graphically representing diferent and preferentially complementary performance meas-ures in a single diagram (see e.g. Jolliff et al. [2009](#page-18-7)). An example is Taylor's diagram (Taylor [2001](#page-19-9)), which exploits the law of cosines to display the analytical relationship between standard deviations, correlation, and centered root-mean-squared error (CRMSE) of diferent model outputs and reference data. Plotted in a polar coordinate system, Taylor's diagram eases comparison of multiple model outputs (represented by distinct points) by utilizing the conceptual equivalence of their distance from the origin to CRMSE. The radial and azimuthal coordinates of each point, which corresponds to a diferent model, allow for visual assessment of the potential reasons of the discrepancies observed between model outputs and observations. Due to its compactness in yielding an immediate assessment of multiple model performances, Taylor's diagram has been widely used in environmental sciences. Particularly popular is its use in the evaluation and comparison of climate models. The latter are fundamental tools for simulating and predicting the Earth's climate system, including its atmospheric, oceanic, and land surface components.

The development and evaluation of these models is crucial for improving our understanding of the climate system, predicting future climate change, and informing policy decisions (see e.g. Deidda et al. [2013;](#page-18-8) Langousis and Kaleris [2014](#page-19-10); Langousis et al. [2016;](#page-19-11) Emmanouil et al. [2022,](#page-18-9) [2023\)](#page-18-10).

However, Taylor's diagram has well known limitations that should be considered when interpreting results. In particular, (1) it does not explicitly consider the overall bias of model outputs (see e.g. Gleckler et al. [2008;](#page-18-11) Hu et al. [2019\)](#page-18-12) and (2) it exploits the root-mean-squared error (RMSE), which is a measure of fit conditional on correlation. More precisely, as shown in [Appendix,](#page-17-0) selection of best performing models via minimization of the RMSE implicitly benefts those that underestimate the variability of observations, unless the correlation coefficient between model results and observations is equal to 1. With regard to point (1), please note that a model characterized by signifcant bias could still be assessed to provide a good fit to observations based on standard deviation and correlation. According to several performance evaluation studies, climate models heavily suffer from large biases and cannot be used directly for impact studies, such as hydrological model applications, unless they are treated using bias correction methods (see e.g. Mamalakis et al. [2017](#page-19-12); Perra et al. [2020;](#page-19-13) Emmanouil et al. [2021,](#page-18-13) [2023](#page-18-10)).To remedy the aforementioned shortcoming, Taylor ([2001](#page-19-9)) proposed a methodology to complement the original diagram with information on bias. In addition, several alternative metrics and tools were progressively developed (see e.g. Xu et al. [2016](#page-19-14); Hu et al. [2019](#page-18-12); Sáenz et al. [2020](#page-19-15); Paul et al. [2023\)](#page-19-1).

Evidently, the development of alternative performance evaluation approaches indicates that the choice of the performance measure to be used should be consistent with the requirements of the application (see e.g. Zhou et al. [2021](#page-19-16)). In this study, we propose Aras' diagram*,* a new tool which exploits in a geometrical setting the structure of the Kling–Gupta efficiency (KGE) index. The choice of KGE is motivated by several advantages over other measures used for model performance assessments. First, KGE accounts for three metrics (i.e. correlation, variability ratio and normalized bias ratio, see Sect. [2\)](#page-2-0). Second, KGE is dimensionless, making it insensitive to the variables considered and their range of values. Third, KGE is a relative measure, allowing for easy comparison of model performances across diferent sites and time periods. Lately, KGE has become a widely used measure for model performance assessments due to its comprehensiveness, simplicity, and ease of interpretation, being used in a variety of studies, including climate change impact assessments (see e.g. Liu [2020](#page-19-17); Agyekum et al. [2022](#page-18-14); Ahmed et al. [2019;](#page-18-15) Castaneda-Gonzalez et al. [2018](#page-18-16); Ta et al. [2018](#page-19-18)), water resources management, and food forecasting (see e.g. Pechlivanidis and Arheimer [2015;](#page-19-19) Mwangi et al. [2016;](#page-19-20) Pool et al. [2018](#page-19-21); Lamontagne et al. [2020](#page-19-22); Brunner et al. [2021](#page-18-17)).

In order to demonstrate the attributes and skills of Aras' diagram, we use it for performance evaluation of EURO-CORDEX historical experiment over Southern Italy using E-OBS as reference dataset, for three hydrological variables (i.e. daily precipitation, daily minimum surface temperature, and daily maximum surface temperature), and 4 climate indices (i.e. total annual precipitation, annual maxima of daily precipitation, annual minima of daily minimum surface temperatures, and annual maxima of daily maximum surface temperatures).

The paper is structured as follows. In Sect. [2,](#page-2-0) the underlying theoretical background of Aras' diagram is discussed. Section [3](#page-4-0) describes the study area and provides a brief summary of the climate model and reference data used. In Sect. [4](#page-5-0) we describe the most important fndings of the pre-sent study, and in Sect. [5](#page-14-0) we discuss the competitive advantages of the newly developed diagram. Concluding remarks are presented in Sect. [6.](#page-17-1)

# <span id="page-2-0"></span>**2 Theoretical foundation and construction of Aras' diagram**

Assume *n* couplets of values  $\{x_{o,i}, x_{m,i}\}_{i=1,...,n}$ , where subscripts *o* and *m* refer to observations and model predictions respectively, and *n* is the length of each time series. Keeping the same subscripts,  $\mu_o$ ,  $\sigma_o$  and  $\mu_m$ ,  $\sigma_m$  stand for the estimated mean and standard deviation of observed and simulated series, respectively, while  $\rho$  denotes their correlation coefficient.

Taylor's diagram (Taylor [2001](#page-19-9)) is based on a graphical representation of the relationship between the centered rootmean-squared error (CRMSE), standard deviation and correlation between observations and predictions:

$$
CRMSE^{2} = \sigma_{m}^{2} + \sigma_{o}^{2} - 2\sigma_{m}\sigma_{o}\rho
$$
 (1)

noting that CRMSE is linked to RMSE and BIAS as:

$$
RMSE^2 = BIAS^2 + CRMSE^2 \tag{2}
$$

with CRMSE, RMSE and BIAS defned as:

$$
CRMSE = \left\{ \frac{1}{n} \sum_{i=1}^{n} \left[ \left( x_{m,i} - \mu_m \right) - \left( x_{o,i} - \mu_o \right) \right]^2 \right\}^{\frac{1}{2}} \tag{3}
$$

$$
RMSE = \left[\frac{1}{n}\sum_{i=1}^{n} (x_{m,i} - x_{o,i})^2\right]^{\frac{1}{2}}
$$
(4)

and.

$$
BIAS = \mu_m - \mu_o \tag{5}
$$

Comparing the structure of Eq. ([1\)](#page-2-1) with the geometrical law of cosines:

$$
d^2 = x^2 + y^2 - 2xy\cos\phi \tag{6}
$$

a graphical representation of CRMSE,  $\sigma_m$ ,  $\sigma_o$  and  $\rho$  can be obtained by setting:

$$
d = \text{CRMSE}, \, x = \sigma_m, \, y = \sigma_o \text{ and } \phi = \arccos(\rho) \tag{7}
$$

The Kling–Gupta efficiency (KGE, Gupta et al. [2009\)](#page-18-6) can be regarded as a statistical measure of the accuracy of diferent model simulations. It is a dimensionless index that ranges between  $-\infty$  and 1, where the value of 1 indicates perfect agreement between model simulations and observed data (see e.g. Lamontagne et al. [2020](#page-19-22)).

Applying the same notation as in Gupta et al. [\(2009](#page-18-6)), we defne the variability ratio:

$$
\alpha = \frac{\sigma_m}{\sigma_o} \tag{8}
$$

and the normalized bias ratio:

$$
\beta = \frac{\mu_m}{\mu_o} \tag{9}
$$

<span id="page-2-2"></span>The Kling–Gupta efficiency is defined as:

$$
KGE = 1 - \sqrt{[(\alpha - 1)]^2 + [(\beta - 1)]^2 + [(\rho - 1)]^2}
$$
 (10)

In a geometrical interpretation, Eq. ([10](#page-2-2)) involves the concept of Euclidian distance from two points in a threedimensional space. Equation  $(10)$  $(10)$  can also be formulated by introducing the total error, *E*:

$$
E = (1 - KGE)^{2} = [(\alpha - 1)]^{2} + [(\beta - 1)]^{2} + [(\rho - 1)]^{2}
$$
\n(11)

<span id="page-2-3"></span><span id="page-2-1"></span>Equation  $(11)$  $(11)$  can also be interpreted as the three-dimensional formulation of the Pythagorean theorem,

$$
d^2 = y^2 + x^2 + z^2 \tag{12}
$$

where  $d=1 - KGE$ ,  $y = \alpha - 1$ ,  $x = \beta - 1$  and  $z = -1$ .

In such a 3D representation, *d* is the Euclidean distance between the point representing model performance and the origin of the diagram (the ideal point) and is equal to the error measure  $E = 1 - KGE$ .

Gupta et al. [\(2009](#page-18-6)) also introduced a general "scaled version" of KGE, by incorporating into Eq.  $(11)$  $(11)$  $(11)$  the scaling factors  $s_{\alpha}$ ,  $s_{\beta}$  and  $s_{r}$ , aimed to rescale the criteria space before computing the Euclidian distance from the ideal point, as a means of weighting diferently the components of the total error:



<span id="page-3-0"></span>**Fig. 1** Model error representation used in Aras' diagram

$$
KGE_s = 1 - \sqrt{[s_\alpha(\alpha - 1)]^2 + [s_\beta(\beta - 1)]^2 + [s_r(\rho - 1)]^2}
$$
\n(13)

Practical implications of the aforementioned concepts can be represented in a Cartesian coordinate system using a twodimensional diagram, where  $\beta$  -1 and  $\alpha$  -1 take values on the *x* and *y* axes, respectively; see Fig. [1](#page-3-0) and discussion below.

To do so, one first exploits the scaled  $KGE<sub>s</sub>$  formulation by assuming  $s_\alpha = s_\beta = 1$  and  $s_r = 0$ . Under this setting, Eq.  $(11)$  $(11)$  reduces to:

$$
E_{\alpha\beta} = (1 - KGE_{\alpha\beta})^2 = (\alpha - 1)^2 + (\beta - 1)^2
$$
 (14)

and by setting  $y = \alpha - 1$ ,  $x = \beta - 1$  and using the Pythagorean theorem, one obtains a graphical representation on a 2D Cartesian coordinate plane of the error component  $E_{\alpha\beta}$  induced by the bias and variability ratios, as:

$$
E_{\alpha\beta} = d^2 = y^2 + x^2 \tag{15}
$$

where *d* denotes the Euclidean distance of point  $E_{\alpha\beta}$  in Fig. [1](#page-3-0) from the origin.

At a second step, one calculates the total percentage error  $E$  given by Eq.  $(11)$  $(11)$  $(11)$ , and draws a line segment that starts from point  $E_{\alpha\beta}$ , has length equal to the difference  $E-E_{\alpha\beta}$ (computed from Eqs.  $(11)$  $(11)$  and  $(14)$ ), and slope:

$$
\tan \phi = \frac{a-1}{\beta - 1} \tag{16}
$$

Figure [1,](#page-3-0) which forms the foundation of Aras' diagram (named by the frst name of the frst Author), illustrates the construction of the aforementioned error representation, where the length of the segment that links point *E* (i.e. total percentage error) to point  $E_{\alpha\beta}$  (i.e. percentage error induced by the bias and variability ratios) provides an indication of the error component induced by discrepancies in correlation (i.e. the longer the segment the larger the disagreement). The slope of the segment indicates the relative contribution of variance and bias ratios (i.e. larger slopes correspond to smaller bias-induced errors). The endpoint of the segment, which represents the error  $E_{\alpha\beta}$ , is marked using a circle which is flled (empty) in case of positive (negative) correlation.

Figure [2](#page-3-2) presents a guide to Aras' diagram. The origin of the diagram (i.e. point (0, 0)) indicates perfect model performance relative to observations. The closer the model mark (i.e. point  $E$  in Figs. [1,](#page-3-0) [2\)](#page-3-2) is to the origin the better is model performance. In the diagram, circles representing  $%$  errors are drawn (e.g. 10%, 25% and 50%). Any model mark located inside the inner circle has error below 10%, any model mark located inside the second circle has error below 25% and any model mark located inside the third circle has error below 50%.

The two axes split the entire diagram into four quadrants. With respect to the *y*-axis, positive and negative *y*-values correspond to variability ratios  $\alpha > 1$  and  $\alpha < 1$ , respectively. Therefore, the quadrants above and below the *x*-axis are indicative of over- and under-estimation of the variability of observed data, respectively. Similarly, the quadrants on the right and left of the *y*-axis correspond to bias ratios  $\beta > 1$ , (i.e. positive bias) and  $\beta$  < 1 (i.e. negative bias), respectively, indicating over- and under-estimation of the mean value of the observed data.

<span id="page-3-1"></span>

<span id="page-3-2"></span>**Fig. 2** Aras' diagram guide

# <span id="page-4-0"></span>**3 Case study and data**

In what follows, we use Aras' diagram to assess the accuracy of climate model outputs from EURO-CORDEX (Jacob et al. [2014](#page-18-18)) historical experiments over southern Italy, using E-OBS (Haylock et al. [2008a](#page-18-19), [b\)](#page-18-20) gridded data as reference dataset.

# **3.1 Description of the study area**

The study area is the entire southern part of the Italian peninsula (latitude 37.5–42.5 N; longitude 11–19E) (see Fig. [3\)](#page-4-1) with approximate area of  $90,000 \text{ km}^2$  crossed by the mountainous ridge of Southern Italy, and a long coastal-line on the order of 2500 km. It is characterized by a wide range of elevations, with peaks reaching up to 2912 m and it includes the "Tavoliere delle Puglie", the second larger plain in Italy. Precipitation in southern Italy varies depending on elevation and location. The regional average of mean annual rainfall ranges from 1000 to 1500 mm. Precipitation generally increases with elevation, with some of the highest elevations

receiving more than 2000 mm of rain annually. Seasonality also afects precipitation which is mainly distributed in the autumn and winter months, with dry summers. The average temperature in the entire area varies between 7–9 °C in winter and 21–23 °C during summer. Temperature is infuenced by elevation and proximity to the sea. The coastal areas of the region have a Mediterranean climate, with mild winters and hot summers. At higher elevations, temperatures generally decrease, and the climate becomes more continental, with colder winters and milder summers. The region is an important area for agriculture, with crops such as olives, grapes, and citrus.

### **3.2 Climate data**

The only way to assess the magnitude and probable causes of climate change is by understanding the natural climate variability (Tsonis et al. [2017\)](#page-19-23). Climate models are computer-based algorithms used to simulate and predict climate conditions under diferent deterministic scenarios. They are fundamental tools for simulating and predicting the Earth's climate system, including its atmospheric, oceanic, and land



<span id="page-4-1"></span>**Fig. 3** Digital elevation model (DEM) of the study area

surface components. The development and evaluation of these models is crucial for improving our understanding of the climate evolution, analyze and predict climate change, and inform policy decisions (Wetterhall et al. [2009](#page-19-24); Palatella et al. [2010](#page-19-25)).

The evaluation of climate models performance with respect to ground observations is a complex and challenging task, due to the highly nonlinear and dynamic nature of the climate system, the large spatial and temporal scales involved, and the complexity of the physical processes that govern the system. The evaluation process can be used also to diagnose and identify model strengths and weaknesses. Moreover, by evaluating the historical period and choosing the best performing climate model, one may assume that such a model should be the best option for future projections. The increasing complexity of climate models, incorporating more physical processes and components, also requires larger and powerful computational resources (Tsonis and Kirwan [2023\)](#page-19-26).

A global climate model (GCM) is designed to represent the Earth's climate system on a global scale. It contains complex mathematical equations that simulate the interactions between the atmosphere, oceans, land surface, ice and other components of the climate system. GCMs are used to understand the long-term climate patterns, project future climate scenarios and assess the impact of greenhouse gas emissions on global climate (see e.g. Kirchmeier-Young and Zhang [2020;](#page-18-21) Moustakis et al. [2021\)](#page-19-27). A regional climate model (RCM), on the other hand, is a more specialized model that focuses on a smaller geographical area, typically a region within a global model grid. RCMs use higher resolution and a more detailed representation of physical processes compared to GCMs, which allows for a more accurate representation of local and regional climate characteristics including orographic efects. RCMs are often driven by the output of a GCM that provides larger scale boundary conditions for the regional model. By downscaling the GCM output, RCMs can provide more localized information for assessing climate impacts, such as assessing changes in precipitation patterns, temperature extremes or regional climate variability (see e.g. Vrac et al. [2007](#page-19-28); Fowler et al. [2007;](#page-18-22) Johnson and Sharma [2009](#page-18-23); Mujumdar et al. [2009\)](#page-19-29).

In summary, GCMs provide a global overview of the Earth's climate, while RCMs focus on smaller regions in a global context and provide more detailed and localized climate information. RCMs are often used in conjunction with GCMs to bridge the gap between global climate projections and regional climate assessments.

As mentioned in the Introduction, for the purpose of this paper a performance evaluation of EURO-CORDEX historical experiment over Southern Italy has been carried out considering historical reference periods of all available combinations of GCMs and RCMs. The investigation focuses on the hydrological variables of precipitation and temperature, by considering four climate indices: annual total precipitation (PRCP), annual maxima of daily precipitation (RX1day), annual minima of daily minimum surface temperatures (TNn) and annual maxima of daily maximum surface temperatures (TXx).

In order to test the historical experiment of EURO-COR-DEX over Southern Italy, we used the E-OBS reference dataset (Haylock et al. [2008a,](#page-18-19) [b\)](#page-18-20): a high-resolution gridded dataset of daily climate over Europe, widely used for climate model evaluations. EURO-CORDEX [\(http://www.](http://www.EURO-CORDEX.net/) [EURO-CORDEX.net/\)](http://www.EURO-CORDEX.net/) climate model outputs at the highest available resolution (i.e. 0.11°) and for a total number of 55 GCM-RCM combinations (see Fig. [4\)](#page-6-0) were compared with the reference dataset of E-OBS at the same resolution  $(0.11^{\circ})$ , using two types of analysis. The first, which we refer to as "temporal analysis", was conducted by averaging the time series of model outputs and reference data over the entire study area, while preserving the temporal resolution of the original series. The second type of analysis, which we refer to as "spatial analysis", was conducted by averaging the model outputs and reference data over the entire reference period, while preserving the spatial resolution of the original felds.

# <span id="page-5-0"></span>**4 Results**

The frst variable considered for performance evaluation in this study is surface temperature. The climate indices chosen are the annual minima of daily minimum temperatures, TNn, and the annual maxima of daily maximum temperatures, TXx (i.e. the lowest and highest temperature values from each year, respectively).

Figure [5](#page-7-0) shows the climatological spatial pattern of the annual minima of daily minimum surface temperatures TNn (after averaging the corresponding time series over the entire reference period), where diferent colours (marks) denote diferent RCM models (GCM drivers). According to Taylor's diagram (Fig. [5](#page-7-0)a), all GCM–RCM model combinations display high correlations in space, on the order of 0.9. A visual ranking of model performances is feasible, looking at the model distances from the origin. The entirety of models shows a quite homogeneous trend in performance indicators, from better to lower performances, with all of them being very close to the 0.9 radius line. Aras' diagram (Fig. [5b](#page-7-0)) also indicates high correlations, as all models display short segments, but additional patterns of model performance become also visible. More precisely, no model exhibits total percentage error below 10%, while only 3 out of all GCM-RCM model combinations show total percentage error less than 30%. These 3 combinations are all driven by the same GCM (NCC-NorESM1-M), with best performing RCMs (in



<span id="page-6-0"></span>**Fig. 4** EURO-CORDEX GCM-RCM combinations used in this study. TN corresponds to daily minimum surface temperature, TX to daily surface maximum temperature and Pr to daily total precipitation

descending performance order): IPSL-WRF381P, CLMcom-ETH-COSMO-crCLIM-v1-1 and ICTP-RegCM4-6. Overall, 15 model combinations display performances with total error lower than 50%. The generally good agreement of model simulations and reference data in terms of correlation in space for all GCM-RCM combinations, leads to a bias-variability percentage error  $(E_{\alpha\beta}$  in Figs. [1,](#page-3-0) [2](#page-3-2)) that is very close to the total percentage error (i.e. *E* in Figs. [1](#page-3-0), [2](#page-3-2)). Moreover, almost all models overestimate both the mean and variability (i.e. they are displayed in the frst quadrant of the diagram, where  $x = \beta - 1$  and  $y = \alpha - 1$  coordinates are positive). Exceptions include RCM model CNRM-ALADIN53 driven by GCM model CNRM-CERFACS-CNRM-CM5 (the only one underestimating variability), and three model combinations with negative bias. Two of the latter have the same RCM: GERICS-REMO2015 and are driven by GCMs NCC-NorESM1-M and MPI-M-MPI-ESM-LR.

In general, patterns or clusters of model combinations sharing the same RCMs (same colour but diferent marks) are also very clear in Aras' diagram. For instance, similar performances are shown by RCMs: MOHC-HadREM3- GA7-05, KNMI-RACMO22E, CNRM-ALADIN63 and SMHI-RCA4, independent of the driving GCM. Worst performances in terms of bias are provided by all models sharing RCM KNMI-RACMO22E regardless of the GCM drivers.

Figure [6](#page-8-0) shows diagrams for the spatial climatological pattern of the annual maxima of daily maximum surface temperatures TXx. In this case, the correlation values that range from somewhat below 0.8 to more than 0.9 still display an overall good agreement of model data with observations. This diference in correlation performance is strongly highlighted in Aras' diagram, where the few models characterized by correlation less (or close to) 0.8 are associated with much longer segments. Most of them are driven by the same RCM (ICTP-RegCM4-6) nested in 5 diferent GCMs. The same 5 models are also among the 7 models showing the lowest bias-variability percentage error (less than 10%).

Both Taylor's and Aras' diagrams coincide in indicating GERICS-REMO2015 driven by CNRM-CERFACS-CNRM-CM5 as the best performing model, having a total percentage error approximately equal to 10%. This is the best score we obtained in this study, over all indices and models evaluated. Aras' diagram also shows that: a) all models display performances in simulating the spatial climatological pattern of TXx with total percentage error lower than 50%, b) 9 models



<span id="page-7-0"></span>**Fig. 5** Spatial climatological pattern analysis for the annual minima of daily minimum surface temperatures using: **a** Taylor's diagam, and **b** Aras' diagram



<span id="page-8-0"></span>**Fig. 6** Spatial climatological pattern analysis for the annual maxima of daily maximum surface temperatures using: **a** Taylor's diagram, and **b** Aras' diagram

exhibit negative bias, and c) all models except for 5 (all of which being among the 7 best performing models regarding bias-variability percentage error) overestimate the spatial variability of TXx feld.

Shifting to the larger picture, a general observation one can make is that the overall EURO-CORDEX experiment provides better results for the maximum than for the mini-mum daily surface temperature in space (compare Figs. [6,](#page-8-0) [5](#page-7-0)), with the majority of model combinations overestimating both the mean and variability of the reference data.

The annual minima of daily minimum temperatures TNn, and the annual maxima of daily maximum temperatures TXx averaged over the entire study area were also subjected to temporal analysis with results shown in Figs. [7,](#page-10-0) [8,](#page-11-0) respectively.

The frst impression one gets from both diagrams is that performances in terms of correlation in time are well below those in space for all models and indices analyzed (compare Figs. [5](#page-7-0), [6](#page-8-0) to Figs. [7,](#page-10-0) [8](#page-11-0)). The two Taylor diagrams in Figs. [7](#page-10-0)a, [8](#page-11-0)a show that correlation is always below 0.3–0.4 and a signifcant number of models exhibit negative correlation with the reference data in time. The insufficient overall performance in terms of correlation error is also shown in Aras' diagram, where all models display long segments, with those exhibiting negative correlations being marked with an empty circle at their endpoints  $E_{\alpha\beta}$ .

No model in both Aras' diagrams shown in Fig. [7b](#page-10-0) and Fig. [8b](#page-11-0) exhibits total percentage error below 50%, with most of the insufficient performance being due to errors in correlation. In fact, when considering the bias-variability percentage error  $E_{\alpha\beta}$ , a good number of models (i.e. about half for the minimum and almost all for the annual maximum temperatures) are encompassed within the 50% error circle.

More precisely, with specifc regard to annual minima of daily minimum surface temperatures TNn (see Fig. [7](#page-10-0)b), 22 models exhibit bias-variability percentage errors (i.e.  $E_{\alpha\beta}$ ) less than 50%, 7 models less than 25% and 1 model less than 10%. The latter, corresponds to RCM CLMcom-ETH-COSMO-crCLIM-v1-1 driven by GCM ICHEC-EC-EARTH. All except for 2 models overestimate the interannual variability, and one model (CNRM-ALADIN63 driven by MOHC-HadGEM2-ES) reproduces the interannual variability almost perfectly. All except 6 models exhibit positive bias, and a general tendency is observed towards overestimation of both the mean and variability of annual minimum daily temperatures.

The best model according to Aras' diagram is diferent from the best one indicated by Taylor's diagram; i.e. according to Taylor's diagram, RCM CNRM-ALADIN53 driven by GCM CNRM-CERFACS-CNRM-CM5 performs best (see Fig. [7a](#page-10-0)), whereas Aras' diagram indicates RCM CNRM-ALADIN63 nested in GCM MOHC-HadGEM2-ES as the best performing model combination (see Fig. [7](#page-10-0)b). This

diference is attributed to the fact that, contrary to Taylor's diagram, Aras' diagram accounts for model biases, which in the case of annual minima of daily surface minimum temperatures play an important role; i.e. 117% for the best model according to Taylor's diagram, and 87% for the best model according to Aras' diagram (see Fig. [7](#page-10-0)b).

With specifc regard to annual maxima of daily maximum temperatures (see Fig. [8](#page-11-0)), TXx, Taylor's and Aras' diagrams indicate the same best performing model: RCM GERICS-REMO2015 driven by GCM IPSL-IPSL-CM5A-MR, with total percentage error of 62% mostly due to low correlation.

When looking at the bias-variability percentage error, 11 models exhibit errors below 10%, 4 of them sharing the same RCM (ICTP-RegCM4-6) and other 3 sharing RCM IPSL-WRF381P. Also, 35 models exhibit bias-variability percentage errors below 25%.

A general tendency to overestimate both the mean and variability of maximum daily temperatures is also evident: 82% of the models have positive bias, and 78% of the models overestimate interannual variability with respect to observations.

With regard to precipitation over southern Italy, the EURO-CORDEX models show a wide range of performances in terms of both the spatial and temporal patterns of annual precipitation totals (PRCP) and annual maxima of daily rainfall (Rx1day). Results from the spatial analysis are shown in Figs. [9](#page-12-0), [10](#page-13-0).

The Taylor diagrams in Fig. [9a](#page-12-0) and Fig. [10](#page-13-0)a show that all models exhibit positive correlation with the reference data, indicating better performance for PRCP relative to Rx1day. This is also visible from the length of the segments and their flled endpoints shown in Aras' diagrams in Figs. [9b](#page-12-0), [10b](#page-13-0).

With regard to PRCP, the Aras diagram in Fig. [9](#page-12-0)b indicates that only one model exhibits total error less than 25% (i.e. RCM CLMcom-CCLM4-8–17 driven by GCM ICHEC-EC-EARTH). The same model is the second best in terms of bias-variability error (i.e.  $E_{\alpha\beta}$  less than 10%), outperformed only by the model sharing the same RCM and driven by GCM MOHC-HadGEM2-ES. Also, all combinations of CLMcom-ETH-COSMO-crCLIM-v1-1 RCM perform well (i.e. bias-variability percentage error  $E_{\alpha\beta}$  less than 25%), independent of the driving GCM. Worst performances in terms of overestimation of variability are associated with RCM DMI-HIRHAM5. Overall, 48 out of 55 models overestimate the variability of the reference data and 45 out of 55 models overestimate the observed mean.

Regarding the annual maximum daily precipitation, Rx1day, the Aras diagram in Fig. [10](#page-13-0)b shows that all models exhibit total errors above 50%, and positive bias with respect to the spatial climatological pattern averaged over time. When considering only the bias-variability percentage error  $E_{\alpha\beta}$ , best models (with error under 25%) are those resulting from RCM ICTP-RegCM4-6 driven by



<span id="page-10-0"></span>**Fig. 7** Temporal analysis for the annual minima of daily minimum temperatures using: **a** Taylor's diagram, and **b** Aras' diagram



<span id="page-11-0"></span>**Fig. 8** Temporal analysis for the annual maxima of daily maximum surface temperatures using: **a** Taylor's diagram, and **b** Aras' diagram





<span id="page-12-0"></span>**Fig. 9** Spatial analysis of annual total precipitation using: **a** Taylor's diagram, and **b** Aras' diagram



# (a)RX1day spatial - Taylor's diagram

<span id="page-13-0"></span>**Fig. 10** Spatial analysis of annual maxima of precipitation using: **a** Taylor's diagram, **b** Aras' diagram

5 diferent GCMs. This can be seen as an evidence that regional climate model selection is of major importance when interest is in the spatial distribution of annual maxima of daily rainfall.

The results of the temporal analysis of the two indices related to precipitation (i.e. PRCP and Rx1day) are shown in Figs. [11](#page-15-0), [12.](#page-16-0) The Taylor diagrams in Figs. [11](#page-15-0)a, [12](#page-16-0)a are somewhat similar to those of temperature related indices (Figs. [7](#page-10-0)a, [8](#page-11-0)a): all models show low correlation (always below 0.5) with an important number exhibiting negative correlations. The Aras diagrams in Figs. [11](#page-15-0)b, [12b](#page-16-0) confrm this and, in addition, indicate that none of the available GCM-RCM combinations exhibits total percentage error lower than 50%, for both the annual rainfall totals (PRCP) and the annual maxima of daily rainfall (Rx1day).

With specific regard to PCRP (Fig. [11\)](#page-15-0), 7 models exhibit bias-variability percentage error  $E_{\alpha\beta}$  below 10%, and 15 models below 25%. The best performing model in term of bias-variability percentage error corresponds to RCM CLMcom-ETH-COSMO-crCLIM-v1-1 driven by GCM MOHC-HadGEM2-ES, the second and third best models in terms of  $E_{\alpha\beta}$  are RCMs CLMcom-ETH-COSMO-crCLIM-v1-1 and GERICS-REMO2015, both nested in GCM NCC-NorESM1-M. The majority of models (i.e. 48 out of 55) exhibit positive bias, and 41 out of 55 overestimate variability. The worst performing model according to Taylor's diagram is RCM CNRM-ALADIN63 driven by GCM NCC-NorESM1-M, while according to Aras' diagram the worst model combination results from the same RCM (i.e. CNRM-ALADIN63) nested in CNRM-CERFACS-CNRM-CM5. This diference is attributed to the fact that, diferently from Taylor's diagram, Aras' diagram accounts for model biases, which in the latter case exceed 150%.

Results relative to temporal analysis of Rx1day averaged over Southern Italy (Fig. [12](#page-16-0)b), show that all models are associated with positive biases and, in addition, overestimate the interannual variability. No model exhibits total error below 50%, with the best model in terms of bias-variability error (below 25%) resulting from the combination of RCM ICTP-RegCM4-6 nested in GCM NCC-NorESM1-M. Models with  $E_{\alpha\beta}$  below 50% include: RCM ICTP-RegCM4-6 driven by GCMs ICHEC-EC-EARTH, CNRM-CERFACS-CNRM-CM5 and MPI-M-MPI-ESM-LR, and RCM IPSL-WRF381P driven GCM NCC-NorESM1-M, and RCM CNRM-ALA-DIN53 driven by GCM CNRM-CERFACS-CNRM-CM5. The aforementioned same-RCM-clusters, regardless of the driving GCM, indicate, once more, the critical role of RCMs when it comes to modeling rainfall maxima. The worst model (in both diagrams) in terms of interannual variability of daily rainfall maxima is RCM MOHC-HadREM3- GA7-05 driven by GCM CNRM-CERFACS-CNRM-CM5, with associated error that exceeds 200%.

#### <span id="page-14-0"></span>**5 Discussion**

Application of Aras' diagram to four indices related to extreme temperature and precipitation highlight the added value of including in a single 2-dimensional graphical tool three important and independent metrics for performance evaluation of climate models: correlation, bias and variability ratios. By displaying the total percentage error and the percentage error due to bias and variability, it is possible to easily identify models with similar performances, as well as the sources of observed discrepancies. The three homocentric circles corresponding to diferent values of relative error (i.e. 10%, 25% and 50%) ease selection of models with better overall performance, whereas decomposition of the observed error to bias-, variability- and correlation-related components becomes visually possible.

By visually detecting clusters of model combinations with similar performances (see Sect. [4\)](#page-5-0), we found a general dominance of RCM selection in model performance; i.e. models using the same RCM driven by diferent GCMs tend to perform similarly. Also, similarly to other studies (see e.g., Deidda et al. [2013](#page-18-8); Sillmann et al. [2013;](#page-19-30) Kotlarski et al. [2014](#page-19-31); Mascaro et al. [2018](#page-19-32); Vautard et al. [2021](#page-19-33)), we found that climate models performance for temperature is better than that for precipitation, in particular when looking at spatial patterns, see discussion on Figs. [5](#page-7-0)[–12.](#page-16-0) The component of total error due to lack of correlation proved to be one of the most important sources of poor model performance for temporal analysis of all three variables (precipitation, minimum temperature, and maximum temperature), see Figs. [7](#page-10-0), [8](#page-11-0), [11](#page-15-0), [12](#page-16-0). Very good results in terms of correlation were obtained by all models of the EURO-CORDEX experiment in the analysis of the spatial distribution of maximum and minimum temperatures. Not so high but still positive correlations were obtained by all EURO-CORDEX models in the spatial analysis of annual rainfall totals (PRCP) and annual maxima of daily rainfall (Rx1day); see discussion on Figs. [9](#page-12-0), [10](#page-13-0). Lack of correlation becomes the most important error source in the temporal analysis, with about half of the models exhibiting negative correlations for all studied variables. Another general result, is that we could not identify a single best performing model for all studied variables.

We exploited Aras' diagram also to gain information about underestimation/overestimation of the mean and variability of the reference data. A systematic overestimation of the mean and variability can be observed in all analyses we performed; see Figs. [5](#page-7-0)[–12](#page-16-0) and their discussion. In particular, in the spatial analysis of minimum and maximum temperatures and annual rainfall totals, almost all models overestimate both the mean and variability of observed felds. In the spatial analysis of annual maxima



<span id="page-15-0"></span>**Fig. 11** Temporal analysis of annual total precipitation using: **a** Taylor's diagram, and **b** Aras' diagram



<span id="page-16-0"></span>**Fig. 12** Temporal analysis of annual maxima of daily precipitation using: **a** Taylor's diagram, and **b** Aras' diagram

Taylor's and Aras' diagrams allow for diferent views of model performances. For instance, with regard to temporal analysis of annual rainfall totals (Fig. [11](#page-15-0)), best performing models according to Taylor's diagram are: RCM ICTP-RegCM4-6 driven by CNRM-CERFACS-CNRM-CM5, and RCMs MOHC-HadREM3-GA7-05, CLMcom-ETH-COSMO-crCLIM-v1-1 nested in CNRM-CER-FACS-CNRM-CM5. Among those, Aras' diagram allows to clearly detect RCM CLMcom-ETH-COSMO-crCLIM-v1-1 driven by CNRM-CERFACS-CNRM-CM5 as the best performing one. This is due to the value added by accounting for biases in models' total error.

# <span id="page-17-1"></span>**6 Conclusions**

Diferent types of diagrams may be used to evaluate climate models. Among those, the most popular one is Taylor's diagram. While very useful, this diagram does not include information on bias-induced errors. In this context, we propose a two-dimensional diagram, referred to as Aras' diagram, which allows for visual evaluation of the relative performances of complex models. This is done by exploiting the Kling–Gupta efficiency (KGE) through a decomposition to its three main components: correlation, bias and variability ratios, allowing for quick assessments of possible overestimation or underestimation of both the mean and variability of observed felds.

Aras' diagram can be used in many aspects of model evaluation and diagnosis, for any type of variable, index, temporal or spatial climatological pattern, seasonality, etc. In this paper Aras' diagram has been used successfully for performance evaluation of climate models over Southern Italy for the hydrological variables of precipitation and temperature, by considering four climate indices: annual total precipitation, annual maxima of daily precipitation, annual minima of daily minimum temperatures, and annual maxima of daily maximum temperatures, in both space and time. The proposed diagram, based on the Kling–Gupta efficiency equation and its decomposition, has shown interesting properties which may assist climate model evaluations based on their accuracy in reproducing the climatological patterns observed in time and space.

An important note to be made here is that, as Aras' diagram is based on the standardization used in deriving the KGE index (see Sect. [2](#page-2-0)), it may lead to inconclusive results in cases when the observations exhibit low mean value and/ or standard deviation (see e.g. Clark et al. [2021\)](#page-18-24), resulting in high values of the bias and variability ratios, respectively. In such cases, Aras' and Taylor's should be used in conjunction, as their combined use should allow for a better overall assessment of the accuracy of model results relative to observations.

# <span id="page-17-0"></span>**Appendix**

Denote by  $(\mu_m, \sigma_m)$  and  $(\mu_o, \sigma_o)$  the first two statistical moments (i.e. means and standard deviations) of model results and observations, respectively. As shown by Gupta et al. ([2009\)](#page-18-6), the mean-squared-error (MSE) and the Nash–Sutcliffe efficiency (NSE) are linked through the following strictly monotonically decreasing relationship:

<span id="page-17-2"></span>
$$
NSE = 1 - \frac{MSE}{\sigma_o^2}
$$
 (17)

In addition, Gupta et al. ([2009\)](#page-18-6) showed that NSE can be decomposed as:

$$
NSE = 2a\rho - a^2 - b^2 \tag{18}
$$

where  $a = \sigma_m / \sigma_o$ ,  $b = (\mu_m - \mu_o) / \sigma_o$ , and  $\rho$  stands for the correlation coefficient between observations and model results.

<span id="page-17-4"></span><span id="page-17-3"></span>By combining equations  $(17)$  $(17)$  and  $(18)$  $(18)$ , one obtains:

$$
MSE = \sigma_o^2 (1 - 2a\rho + a^2 + b^2)
$$
 (19)

By taking the 1st and 2nd derivatives of equation ([19\)](#page-17-4) with respect to *a*, and considering that  $\sigma_o^2$  is a positive constant, one obtains:

<span id="page-17-5"></span>
$$
\frac{d(MSE)}{da} = 2\sigma_o^2(a - \rho), \frac{d^2(MSE)}{da^2} = 2\sigma_o^2 > 0
$$
 (20)

According to equation  $(20)$  $(20)$ , the minimum MSE (corresponding to  $d(MSE)/da = 0$ ) is attained when  $a = \sigma_m / \sigma_0 = \rho \leq 1$ . Therefore, unless  $\rho = 1$ , model selection through minimization of MSE and/or  $RMSE = MSE^{1/2}$ , would favor models that tend to underestimate the variability of observations.

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#### **Declarations**

**Conflict of interest** The authors declare no competing interests.

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