A BK inequality for randomly drawn subsets of fixed size

J. van den Berg · J. Jonasson

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Abstract The BK inequality (van den Berg and Kesten in J Appl Probab 22:556–569, 1985) says that, for product measures on $\{0, 1\}^n$, the probability that two increasing events A and B 'occur disjointly' is at most the product of the two individual probabilities. The conjecture in van den Berg and Kesten (1985) that this holds for all events was proved by Reimer (Combin Probab Comput 9:27–32, 2000). Several other problems in this area remained open. For instance, although it is easy to see that non-product measures cannot satisfy the above inequality for all events, there are several such measures which, intuitively, should satisfy the inequality for all increasing events. One of the most natural candidates is the measure assigning equal probabilities to all configurations with exactly k1's (and probability 0 to all other configurations). The main contribution of this paper is a proof for these measures. We also point out how our result extends to weighted versions of these measures, and to products of such measures.

Keywords BK inequality · Negative dependence

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J. van den Berg (⊠) CWI and VU University Amsterdam, Amsterdam, The Netherlands e-mail: J.van.den.Berg@cwi.nl

J. Jonasson Chalmers University of Technology and Gothenburg University, Göteborg, Sweden e-mail: jonasson@chalmers.se



1 Introduction and statement of results

We start with an illustrative example, where two persons have to divide a random collection of resources:

Example A large box contains thirty items, which we simply name by the numbers $1, \ldots, 30$. Two persons, Alice and Bob, both have a list of those subsets of $\{1, \ldots, 30\}$ that he or she considers 'useful'. For instance, the items could be tools, and Alice may need to do a certain job which can be performed with the combination of tools $\{5, 16, 20\}$, but can also be performed with the combination $\{5, 18, 20\}$, the combination $\{8, 25\}$ etcetera. Similarly for Bob.

Now suppose that a fixed number, say, ten, of items is drawn randomly (uniformly) from the box. These ten items are to be divided between Alice and Bob. For both persons to be satisfied, there must be a pair of disjoint subsets of the ten items, such that one of these subsets is on Alice's list, and the other is on Bob's list. Our main result, Theorem 1.2, says that the probability of this event is at most the product of the probability that the above set of ten items contains a set on Alice's list and the probability of the analogous event concerning Bob's list.

1.1 Definitions, statement of results, and background

Before we state our main result, we recall results from the literature which motivate our current work and are used in our proofs. First some notation and definitions: Throughout this paper, Ω will denote the set $\{0,1\}^n$, and P_p the product distribution on Ω with parameter p. We will often use the notation [n] for $\{1,\ldots,n\}$. For $\omega \in \Omega$ and $S \subset [n]$, we define ω_S as the 'tuple' $(\omega_i, i \in S)$. Further we use the notation $[\omega]_S$ for the set of all elements of Ω that 'agree with ω on S'. More formally,

$$[\omega]_S := {\alpha \in \Omega : \alpha_S = \omega_S}.$$

Now, for $A, B \subset \Omega$, $A \square B$ is defined as the event that A and B 'occur disjointly' in the sense that there are disjoint subsets $K, L \subseteq [n]$ such that, informally speaking, the ω values on K guarantee that A occurs, and the ω values on L guarantee that B occurs. Formally, the definition is:

$$A \square B = \{ \omega \in \Omega : \exists \text{ disjoint } K, L \subset [n] \text{ s.t. } [\omega]_K \subset A \text{ and } [\omega]_L \subset B \}.$$

For ω and $\omega' \in \Omega$ we write $\omega' \geq \omega$ if $\omega'_i \geq \omega_i$ for all $i \in [n]$. An event $A \subset \Omega$ is said to be increasing if $\omega' \in A$ whenever $\omega \in A$ and $\omega' \geq \omega$.

Inequality (1) below was proved for increasing events in [15]. That special case has become a widely used tool in percolation theory and related topics (see e.g. [8] and [9]). The paper [15] also stated the conjecture that (1) holds for all events. Some other special cases were proved in [16] and [14]. There was not much hope for a proof of the general case until finally this was obtained by the then unknown young mathematician Reimer, see [13]:



Theorem 1.1 For all n and all A, $B \subset \{0, 1\}^n$,

$$P_p(A \square B) \le P_p(A)P_p(B). \tag{1}$$

It is easy to see that non-product measures on $\{0, 1\}^n$, cannot satisfy (the analog of) (1) for all events. However, one may intuitively expect that many measures do satisfy the analog of (1) for all *increasing* events. Such measures are said to have the BK property (or, simply, to be BK measures). The most intuitively appealing case where one may expect this property to hold, is the measure corresponding with randomly, uniformly, drawing a subset of fixed size from the set [n]. (See Section 3.1 of [7], and the lines below (4.18) in [9] where this has been conjectured). It seems (oral communication) that several researchers have made efforts to prove this.

To be precise, let $k \le n$ and let $\Omega_{k,n}$ be the set of all $\omega \in \Omega$ with exactly k1's. Let $P_{k,n}$ (which we call the k-out-of-n measure) be the distribution on Ω that assigns equal probability to all $\omega \in \Omega_{k,n}$ and probability 0 to all other elements of Ω . Our main result, Theorem 1.2 below, is that such measures indeed have the BK property. As far as we know, this is the first substantial example of a non-product BK measure.

Theorem 1.2 For all n, all $k \le n$, and all increasing $A, B \subset \{0, 1\}^n$,

$$P_{k,n}(A \square B) \le P_{k,n}(A) P_{k,n}(B). \tag{2}$$

Remark In Sect. 3 we explain that this result and its proof extend to certain weighted versions of $P_{k,n}$ (also called conditional Poisson measures) and to products of such measures.

The rest of the paper is organized as follows: In Sect. 1.2 we state Proposition 1.3, an intermediate result by Reimer which was of crucial importance in his proof of Theorem 1.1 (and which is also very interesting in itself). In Sect. 2 we first state and prove Proposition 2.1. This is an analog of (and its proof uses) the above mentioned Proposition 1.3. Then we derive Theorem 1.2 from Proposition 2.1 in a way similar to that in which Reimer derived Theorem 1.1 from Proposition 1.3. We end the current section with some remarks which are of general interest but are not necessary for understanding the proof of Theorem 1.2.

Remarks and discussion

(a) The example in the beginning of this section corresponds with the case n=30 and k=10 in Theorem 1.2: Take

 $A = \{ \omega \in \Omega : \operatorname{supp}(\omega) \text{ contains a set on Alice's list} \},$

where supp $(\omega) = \{i \in [n] : \omega_i = 1\}$. Take B similarly, but now with Bob's list.

(b) One of the most widely used notions of negative dependence is NA (Negative Association), which is defined as follows. First, two events $A, B \subset \Omega$ are said to be orthogonal if there are two disjoint subsets $K, L \subset [n]$ such that A is defined in terms



of the indices in K and B in terms of the indices in L. Now, a measure μ on Ω is said to be NA if for all increasing, orthogonal events $A, B \subset \{0, 1\}^n$, $\mu(A \cap B) \leq \mu(A)\mu(B)$.

Note that if two events A and B are orthogonal, then clearly $A \square B = A \cap B$. Hence BK implies NA. The reverse is not true (see [11]).

In the last 12 years there has been a lot of research activity aiming at a general theory of negative dependence. This started with the papers [12] and [6]. The understanding of NA has enormously increased, in particular by an algebraic/(complex-)analytic approach involving the zeroes of the generating polynomials (see [1,3,4]). Other techniques to study NA-related problems can be found in [10] and [5]. However, so far these approaches do not work for the BK property and it is unclear how this property would fit in a general framework.

(c) For some non-product measures, in particular Ising models, the following question makes sense: can the \Box -operation be *modified* in a natural way such that (1.1) holds for *all* events? This is investigated in J. van den Berg and A. Gandolfi (in preparation).

1.2 Reimer's intermediate result

The following result, Proposition 1.3 below, is essentially, but in different terminology, Theorem 1.2 (or the equivalent Theorem 1.3) in [13].

As before, Ω denotes $\{0,1\}^n$. Some more notation is needed: For $\omega = (\omega_1, \ldots, \omega_n) \in \Omega$, we denote by $\bar{\omega}$ the configuration obtained from ω by replacing 1's by 0's and vice versa:

$$\bar{\omega} = (1 - \omega_1, \dots, 1 - \omega_n).$$

Further, for $A \subset \Omega$, we define $\bar{A} = \{\bar{\omega} : \omega \in A\}$. Finally, if V is a finite set, |V| denotes the number of elements of V. Now we state Reimer's 'intermediate' result to which we referred before:

Proposition 1.3 (Reimer [13]) For all n and all $A, B \subset \{0, 1\}^n$,

$$|A\Box B| \le |A \cap \bar{B}|. \tag{3}$$

Remarks

- (a) The very ingenious, linear-algebraic, proof of this proposition was the crucial part of Reimer's paper [13]. The fact that a result of the form of this proposition would imply Theorem 1.1 had already been discovered independently by other researchers (see [14]).
- (b) The language/terminology in [13] is somewhat unusual (for probabilists). This makes it, at first sight, difficult to see that Proposition 1.3 above is indeed equivalent to Theorem 1.2 in [13]. Several authors have reviewed Reimer's paper with additional explanation (see [2]).



2 Proof of Theorem 1.2

2.1 An analog of Proposition 1.3 for k-out-of-n measures

Let $k \le n$, and recall the notation in Sect. 1.1. The key to Theorem 1.2 is the (proof of the) following 'analog' of Proposition 1.3:

Proposition 2.1 For all even m and all increasing $A, B \subset \{0, 1\}^m$,

$$P_{\frac{m}{2},m}(A\square B) \le P_{\frac{m}{2},m}(A\cap \bar{B}). \tag{4}$$

We will show in Sect. 2.2 that Theorem 1.2 follows from Proposition 2.1. Finally, in Sect. 2.3, we will prove Proposition 2.1 by writing $P_{\frac{m}{2},m}$ as a suitable convex combination of measures for which the analog of (4) can be derived from Proposition 1.3 via a suitable 'encoding'.

2.2 Proof that Proposition 2.1 implies Theorem 1.2

Proof The proof below is quite similar to the proof (in [13]) that Proposition 1.3 implies Theorem 1.2.

Let A and $B \subset \Omega$ be increasing, and let $k \leq n$. We first rewrite the desired inequality, (2), in an obvious way:

$$(P_{k,n} \times P_{k,n})((A \square B) \times \Omega) \le (P_{k,n} \times P_{k,n})(A \times B). \tag{5}$$

For each $K \subset [n]$ and $\alpha \in \{0, 1\}^K$, define the 'cell'

$$W_{\alpha} = \left\{ (\omega, \omega') \in \Omega_{k,n} \times \Omega_{k,n} \, : \, \omega_K = \omega'_K = \alpha, \; \omega_{K^c} = \overline{\omega'_{K^c}} \right\},$$

where K^c denotes $[n] \setminus K$.

It is easy to see that these cells form a partition of $\Omega_{k,n} \times \Omega_{k,n}$. Hence it is sufficient to prove that, for each α of the form mentioned above

$$|((A \square B) \times \Omega) \cap W_{\alpha}| < |(A \times B) \cap W_{\alpha}|. \tag{6}$$

Using the notation $|\alpha| = \sum_{i=1}^{n} \alpha_i$, notice that if $W_{\alpha} \neq \emptyset$ then (since for every $(\omega, \omega') \in W_{\alpha}$ one has $2k = |\omega| + |\omega'| = 2|\alpha| + n - |K|$),

$$|\alpha| = k - (n - |K|)/2,$$
 (7)

and hence

$$|\omega_{K^c}| = |\omega'_{K^c}| = (n - |K|)/2 \quad \text{for all } (\omega, \omega') \in W_{\alpha}. \tag{8}$$



Before going on, we introduce more notation. Let $\Omega_{(K^c)}$ be the set of all $\omega \in \{0, 1\}^{K^c}$ for which the number of 1's and the number of 0's are equal (and hence equal to (n - |K|)/2). Define, for $\gamma \in \{0, 1\}^{K^c}$, $\gamma \circ \alpha$ as the element of Ω for which

$$(\gamma \circ \alpha)_{K^c} = \gamma$$
, and $(\gamma \circ \alpha)_K = \alpha$.

Further, define for every event $H \subset \Omega$,

$$H(\alpha) = \{ \gamma \in \{0, 1\}^{K^c} : \gamma \circ \alpha \in H \}.$$

From now on we assume, without loss of generality, that α satisfies (7). Now suppose that (ω, ω') belongs to the set in the r.h.s. of (6). This holds if and only if $\omega_K = \omega_K' = \alpha$, $\omega_{K^c} \in \Omega_{(K^c)}$, $\omega \in A$, $\omega' \in B$ and $\omega_{K^c} = \overline{\omega_{K^c}'}$.

The number of pairs (ω, ω') that satisfy this is clearly $|A(\alpha) \cap \overline{B(\alpha)} \cap \Omega_{(K^c)}|$. Similarly, it is easy to see that the l.h.s. of (6) is equal to $|(A \square B)(\alpha) \cap \Omega_{(K^c)}|$. Further, it is easy to check from the \square -definition that

$$(A \square B)(\alpha) \subset A(\alpha) \square B(\alpha) \tag{9}$$

So the l.h.s. of (6) is at most $|(A(\alpha) \square B(\alpha)) \cap \Omega_{(K^c)}|$. Hence, sufficient for (6) to hold is

$$|(A(\alpha) \square B(\alpha)) \cap \Omega_{(K^c)}| \le |A(\alpha) \cap \overline{B(\alpha)} \cap \Omega_{(K^c)}|. \tag{10}$$

Finally, note that this last inequality follows immediately from Proposition 2.1. (Replace the m in (4) by n - |K|, and replace A and B by $A(\alpha)$ and $B(\alpha)$ respectively; note that $A(\alpha)$ and $B(\alpha)$ are increasing because A and B are increasing). This completes the proof that Proposition 2.1 implies Theorem 1.2.

2.3 Proof of Proposition 2.1

We first state and prove Proposition 2.2 below. Let m be even. Let $\hat{\Omega}_m$ be the set of all $\omega \in \{0, 1\}^m$ with the property that, for all $1 \le i \le m/2$, $(\omega_{2i-1}, \omega_{2i})$ is equal to (1, 0) or (0, 1). Let \hat{P}_m be the probability distribution on $\{0, 1\}^m$ which assigns equal probabilities to all $\omega \in \hat{\Omega}_m$, and probability 0 to all other ω .

The following is, as we will see, an 'encoded form' of Proposition 1.3.

Proposition 2.2 For all even m and all increasing $A, B \subset \{0, 1\}^m$,

$$\hat{P}_m(A \square B) \le \hat{P}_m(A \cap \bar{B}). \tag{11}$$

Proof Let A, B $\subset \{0, 1\}^m$ be increasing. Note that (11) is equivalent to

$$|(A\Box B)\cap \hat{\Omega}_m| \le |A\cap \bar{B}\cap \hat{\Omega}_m|. \tag{12}$$



Consider the following bijection $T: \hat{\Omega}_m \to \{0, 1\}^{\frac{m}{2}}$:

$$T(\omega_1,\ldots,\omega_m)=(f((\omega_1,\omega_2),f(\omega_3,\omega_4),\ldots,f(\omega_{m-1},\omega_m)),$$

where f(1, 0) = 1 and f(0, 1) = 0.

We claim that

$$T((A \square B) \cap \hat{\Omega}_m) \subset T(A \cap \hat{\Omega}_m) \square T(B \cap \hat{\Omega}_m), \tag{13}$$

and that

$$T(A \cap \bar{B} \cap \hat{\Omega}_m) = T(A \cap \hat{\Omega}_m) \cap \overline{T(B \cap \hat{\Omega}_m)}.$$
 (14)

The first part of this claim, the inclusion (13), can be seen as follows: Let $\omega = (\omega_1, \dots, \omega_n) \in (A \square B) \cap \hat{\Omega}_m$. By the definition of the \square -operation, there are $K, L \subset [n]$ such that $K \cap L = \emptyset$, $[\omega]_K \subset A$, and $[\omega]_L \subset B$. It is easy to see that this implies

$$[T(\omega)]_{T(K)} \subset T(A \cap \hat{\Omega}_m), \text{ and } [T(\omega)]_{T(L)} \subset T(B \cap \hat{\Omega}_m),$$
 (15)

where $T(K) = \{\lceil i/2 \rceil : i \in K\}$ and T(L) is defined analogously. So far, the argument holds for all events. However, since A and B are increasing, we can even find K and L such that, on top of the above properties, $\omega \equiv 1$ on K and $\omega \equiv 1$ on L, and hence, since $\omega \in \hat{\Omega}_m$,

For all
$$1 \le i \le \frac{m}{2}$$
, $K \cap \{2i - 1, i\} = \emptyset$ or $L \cap \{2i - 1, i\} = \emptyset$. (16)

From (16) it follows immediately that $T(K) \cap T(L) = \emptyset$, which, together with (15), gives

$$T(\omega) \in T(A \cap \hat{\Omega}_m) \square T(B \cap \hat{\Omega}_m),$$

completing the proof of (13). We omit the proof of (14) (which is straightforward). Now, using, in this order, (13), Proposition 1.3 and (14), immediately gives (12). This completes the proof of Proposition 2.2.

Remark At first sight one may think that in (13) even equality holds, but this is false: Take n = 4, $A = \{\omega_1 = 1\} \cap \{\omega_3 = 1 \text{ or } \omega_4 = 1\}$, and $B = \{\omega_3 = 1\} \cap \{\omega_1 = 1 \text{ or } \omega_2 = 1\}$. Then $A \square B$ (and hence the l.h.s. of (13)) is \emptyset , while the r.h.s. of (13) is the subset of $\{0, 1\}^2$ which contains only the element $\{1, 1\}$.

Now we are ready to prove Proposition 2.1:

Proof Let m be even, and recall the definition of $P_{\frac{m}{2},m}$. Let $A, B \subset \{0, 1\}^m$ be increasing. Let, for π a permutation of [m], $\hat{\Omega}_{m,\pi}$ be the set of all $\omega \in \{0, 1\}^m$ with the property that, for each $1 \le i \le m/2$, $(\omega_{\pi(2i-1)}, \omega_{\pi(2i)})$ is equal to (0, 1) or (1, 0).



Let $\hat{P}_{m,\pi}$ be the probability distribution on $\{0, 1\}^m$ which assigns equal probabilities to all $\omega \in \hat{\Omega}_{m,\pi}$ and probability 0 to all other $\omega \in \{0, 1\}^m$. It follows immediately from Proposition 2.2 (by relabelling the indices) that for each π

$$\hat{P}_{m,\pi}(A \square B) \le \hat{P}_{m,\pi}(A \cap \bar{B}). \tag{17}$$

Now observe that if we first randomly (and uniformly) draw a permutation π and then randomly draw a configuration ω according to the distribution $\hat{P}_{m,\pi}$, then ω is a 'typical' random configuration drawn according to $P_{\frac{m}{2},m}$. So $P_{\frac{m}{2},m}$ is a convex combination of the $\hat{P}_{m,\pi}$'s. Finally, since each $\hat{P}_{m,\pi}$ satisfies (17), every convex combination of the $\hat{P}_{m,\pi}$'s also satisfies (17). Hence (4) holds. This completes the proof of Proposition 2.1, and hence the proof of Theorem 1.2.

3 Some extensions of Theorem 1.2

3.1 Weighted *k*-out-of-*n* measures

Let, as before, $1 \le k \le n$, and recall the definition of $\Omega_{k,n}$ and the k-out-of-n measure $P_{k,n}$. Let w_1, \ldots, w_n be non-negative numbers, let $w = (w_1, \ldots, w_n)$, and define the probability measure $P_{k,n}^w$ (a weighted version of $P_{k,n}$ which is sometimes called Conditional Poisson measure)) as follows:

$$P_{k,n}^{w}(\omega) = C \operatorname{I}(\omega \in \Omega_{k,n}) \prod_{i=1}^{n} w_i^{\omega_i}, \ \omega \in \{0, 1\}^n,$$

where C is a normalizing constant and I denotes indicator function. It is not difficult to see that the analog of Theorem 1.2 holds for the weighted measures defined above. In fact, the proof remains practically the same by the following observation: Let W_{α} be a cell (as in Sect. 2.2) with α satisfying (7), and let $(\omega, \omega') \in W_{\alpha}$. Since, for each index $i \in K^c$, exactly one of ω_i and ω_i' equals 1 and the other equals 0, each index $i \in K^c$ contributes exactly a factor w_i to the $(P_{k,n}^w \times P_{k,n}^w)$ measure. Moreover, each $i \in K$ contributes (by the definition of W_{α}) exactly a factor $(w_i)^{2\alpha_i}$. Hence, the proof again reduces to showing (10).

3.2 Products of (weighted) k-out-of-n measures

The proof of the BK inequality for increasing events under k-out-of-n measures extends straightforwardly to that for products of such measures: By the arguments of Sect. 2.2 the proof reduces to showing that Proposition 2.1 holds for products of measures of the form $P_{\frac{m}{2},m}$. Now recall from Sect. 2.3 that the reason that Proposition 2.1 holds is, essentially, that $P_{\frac{m}{2},m}$ is a convex combination of measures of the form $\hat{P}_{m,\pi}$. Now, of course, products of measures $P_{\frac{m_1}{2},m_1}, \ldots P_{\frac{m_l}{2},m_l}$, are convex combinations of measures of the form $\hat{P}_{M,\pi}$, where $M=m_1+\cdots+m_l$, and π is a permutation of [M].



Hence the proof goes through as before. The above argument also goes through for products of *weighted k*-out-of-*n* measures.

3.3 Some ideas for further generalizations

With $w=(w_1,\ldots,w_n)$ as in Sect. 3.1, and X a random variable taking values in $\{0,\ldots,n\}$, define $P_{X,n}^w$ as the measure of the configuration ω resulting from the following procedure. First draw a number k from the same distribution as X. Then draw an $\omega \in \Omega_{k,n}$ according to the distribution $P_{k,n}^w$. Motivated by the search for other examples of BK measures it is natural to ask: for which X is $P_{X,n}^w$ BK? Of course, by Sect. 3.1, this is the case if X is with probability 1 equal to some constant k. Other examples of such X can be easily obtained from the result in Sect. 3.1 by adding 'dummy' indices and then projecting: Let $m \geq 0$, and introduce auxiliary weights $w_{n+1},\ldots w_{n+m} \geq 0$. Let $0 \leq k \leq n+m$. From Sect. 3.1 we have that $P_{k,n+m}^{(w_1,\ldots,w_{n+m})}$ (a measure on $\{0,1\}^{n+m}$) is BK. From the definition of BK it follows immediately that the BK property is preserved under projections. Hence, the projection of $P_{k,n+m}^{(w_1,\ldots,w_{n+m})}$ on $\{0,1\}^n$ is also BK. In other words, if we let, for a random configuration $(\omega_1,\ldots,\omega_{n+m})$ drawn under $P_{k,n+m}^{(w_1,\ldots,w_{n+m})}$, the random variable X denote $\sum_{i=1}^n \omega_i$, then $P_{X,n}^w$ is BK. It is not hard (but also, at this stage, not very helpful) to write a general form for the distribution of an X of this type. It would be interesting to find 'natural' random variables X which are not of this type but yet have the property that $P_{X,n}^w$ is BK.

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References

- Borcea, J., Brändén, P., Liggett, T.M.: Negative dependence and the geometry of polynomials. J. Am. Math. Soc. 22, 521–567 (2009)
- Borgs, C., Chayes, J.T., Randall, D.: The van den Berg-Kesten-Reimer inequality: a review. In: Bramson, M., Durrett, R. (eds.) Perplexing Problems in Probability (Festschrift in honor of Harry Kesten), pp. 159–175 (1999)
- Brändén, P.: Polynomials with the half-plane property and matroid theory. Adv. Math. 216, 302–320 (2007)
- Brändén, P., Jonasson, J.: Negative dependence in sampling. Preprint. www.math.chalmers.se/ jonasson/recent.html (2011)
- Dubhashi, D., Jonasson, J., Ranjan, D.: Positive influence and negative dependence. Combin. Probab. Comput. 16, 29–41 (2007)
- Dubhashi, D., Ranjan, D.: Balls and bins: a study in negative dependence. Random Struct. Alg. 13, 99– 124 (1998)
- Grimmett, G.R.: Percolative problems. In: Grimmett, G.R. (ed.) Probability and Phase Transition, pp. 69–86. Kluwer, Dordrecht (1994)
- 8. Grimmett, G.R.: Percolation, 2nd edn. Springer, Berlin (1999)
- 9. Grimmett, G.R.: Probability on graphs. In: IMS Textbooks. Cambridge University Press (2010)
- 10. Kahn, J., Neiman, M.: Negative correlation and log-concavity. Random Struct. Alg. 37, 367–388 (2010)
- Markström, K.: Closure properties and negatively associated measures violating the van den Berg-Kesten inequality. Elect. Comm. Probab. 15, 449–456 (2009)



- 12. Pemantle, R.: Towards a theory of negative dependence. J. Math. Phys. 41, 1371-1390 (2000)
- 13. Reimer, D.: Proof of the Van den Berg-Kesten conjecture. Combin. Probab. Comput. 9, 27-32 (2000)
- Talagrand, M.: Some remarks on the Berg-Kesten inequality. In: Probability in Banach Spaces, vol. 9, pp. 293–297. Birkhäuser, Boston (1994)
- van den Berg, J., Kesten, H.: Inequalities with applications to percolation and reliability. J. Appl. Probab. 22, 556–569 (1985)
- van den Berg, J., Fiebig, U.: On a combinatorial conjecture concerning disjoint occurrences of events. Ann. Probab. 15, 354–374 (1987)

