REVIEW



# Association between dietary intake and risk of ovarian cancer: a systematic review and meta-analysis

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#### Abstract

**Purpose** It is unclear how dietary intake influences the ovarian cancer. The present paper sets out to systematically review and meta-analyze research on dietary intake to identify cases having high- or low-risk ovarian cancer.

**Methods** Scopus, PubMed, and Wiley Online Libraries were searched up to the date November 24, 2019. Two reviewers were requested to independently extract study characteristics and to assess the bias and applicability risks with reference to the study inclusion criteria. Meta-analyses were performed to specify the relationship between dietary intake and the risk of ovarian cancer identifying 97 cohort studies.

**Results** No significant association was found between dietary intake and risk of ovarian cancer. The results of subgroup analyses indicated that green leafy vegetables (RR = 0.91, 95%, 0.85-0.98), allium vegetables (RR = 0.79, 95% CI 0.64–0.96), fiber (RR = 0.89, 95% CI 0.81–0.98), flavonoids (RR = 0.83, 95% CI 0.78–0.89) and green tea (RR = 0.61, 95% CI 0.49–0.76) intake could significantly reduce ovarian cancer risk. Total fat (RR = 1.10, 95% CI 1.02–1.18), saturated fat (RR = 1.11, 95% CI 1.01–1.22), saturated fatty acid (RR = 1.19, 95% CI 1.04–1.36), cholesterol (RR = 1.13, 95% CI 1.04–1.22) and retinol (RR = 1.14, 95% CI 1.00–1.30) intake could significantly increase ovarian cancer risk. In addition, acrylamide, nitrate, water disinfectants and polychlorinated biphenyls were significantly associated with an increased risk of ovarian cancer. **Conclusion** These results could support recommendations to green leafy vegetables, allium vegetables, fiber, flavonoids and

green tea intake for ovarian cancer prevention.

Keywords Dietary intake · Meta-analysis · Ovarian cancer · Systematic review

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#### Introduction

As the seventh most frequently occurring cancer in women, ovarian cancer accounts for 295,414 new cases and 184,799 deaths annually [1] with the scientific history of the disease dating back to more than 150 years ago. Ovarian cancer is subsumed under the category of chronic and degenerative diseases [2]. Epithelial ovarian cancer falls into the category of five histologic subtypes having various molecular alterations, risk factors, cellular origin, pathogenesis, gene expression, and prognosis with the main ovarian cancer characterized by being high-grade serous, clear cell, endometrioid, mucinous and low-grade serous subtypes. Partly due to the heterogeneous nature of ovarian cancer, its etiology is unclear. Numerous carcinogenic mechanisms have been considered for epithelial ovarian cancer. From among these mechanisms, one can refer to incessant ovulation and inflammation, low immunity, hormonal changes, heightened cell senescence and varying glucose homeostasis, de novo fatty acid synthesis and uncontrolled reactive oxygen production. Experimental investigations suggest that surface ovarian epithelial cells contain high 8-oxoguanine levels during ovulation as one of the most salient mutagenic lesions in DNA [3-11]. Major molecular mechanisms and signaling pathways are involved in ovarian cancer cell development. The growth factor pathways, the activator of transcription 3 pathway, the nuclear factor kappalight-chain enhancer of activated B cells (NF-kB) pathway, the proto-oncogene tyrosine protein kinase Src pathway, the mitogen-activated protein kinase pathway, the ErbB activation pathway, the lysophosphatidic acid pathway and the phosphatidylinositol 3-kinases (PI3K) pathway are implicated in ovarian cancer cell growth and differentiation, cell movement and apoptosis, autophagy, metabolic programing, survival, transcription regulation, and angiogenesis, and are directly linked to tumor suppressor and oncogenic genes in ovarian cancer [12-14].

Factors related to reproduction, e.g., high parity, oral contraceptive use, high lactation duration, tubal ligation, hysterectomy, oophorectomy, salpingectomy and bilateral salpingo-oophorectomy, have been shown to be related to a lower ovarian cancer risk. Some researchers recommend transvaginal ultrasound together with the serum-based marker CA-125 as a primary ovarian cancer screening tool [3, 7, 15–19].

Numerous factors have been shown to be related to a higher ovarian cancer risk. These factors include obesity, old age, children later or not having a full-term pregnancy, treatment of fertility, post menopause hormone therapy, family history of ovarian cancer, breast cancer, or colorectal cancer, family cancer syndrome, hereditary breast and ovarian cancer syndrome, Peutz–Jeghers syndrome, MUTYH-associated polyposis, PTEN tumor hamartoma syndrome, hereditary non-polyposis colon cancer, hypo activity, smoking and alcohol use. Moreover, some factors exert unknown effects on ovarian cancer risk, e.g., diet, androgens and talcum powder [19–23].

The effect of dietary intake on ovarian cancer is unclear [23]. Since 2013, two reviews assessing the relationship between dietary intake and ovarian cancer have been published [24, 25]. Nevertheless, the inconclusive results necessitated more detailed and in-depth investigations of relationships of dietary intake and ovarian cancer risk. Thus, the present systematic and meta-analysis review aims to furnish an overview of the relation between dietary intake and ovarian cancer.

#### Methods

The present meta-analysis has been planned, conducted, and reported in accordance with PRISMA (the Preferred Reporting Items for Systematic Reviews and Meta-Analyses) [26].

#### Inclusion criteria for this review

The inclusion criteria for studies included the following: (1) the research design was a cohort study; (2) the relationship between dietary intake and ovarian cancer was investigated; and (3) the relative risk or hazard ratio with corresponding 95% confidence interval (CI) between the dietary intake and the ovarian cancer risk was documented. As exclusion criteria, conference proceedings, dissertations or abstracts only were excluded.

#### **Search strategies**

Related articles from the electronic databases Scopus, Pub-Med and Wiley Online Library were screened up to the date November 24, 2019. The general search terms of exposure for each of the databases are presented in the Supplementary Tables S1-3 in accordance with WCRF (the World Cancer Research Fund) Specification Manual [27]. References for the articles retrieved and previous systematic reviews investigating the relationship between dietary intake and ovarian cancer risk were searched manually. EndNote X9 citation software was used to import citations (Clarivate Analytics; Boston, MA, USA) as a result of which 25,722 articles were imported into a web-based citation screening platform leaving 18,864 articles for review.

#### **Selection of studies**

Subsequent to the removal of duplicates, titles and abstracts of articles were reviewed to specify their relevance with respect to the inclusion and exclusion criteria. In case initial screening revealed suitability of publications, full texts were examined.

#### **Data extraction and management**

Data extracted from each study independently by the two researchers AK and FA included: author, year of publication, region/country, study name, investigation period, follow-up period, number of cases and participants, age of participants, dietary assessment and exposure, risk estimates, study outcomes and confounder adjustments.

# Assessment of risk of bias and applicability in included studies

The Newcastle–Ottawa Scale (NOS; http://www.ohri.ca/ programs/clinical\_epidemiology/oxford.asp) was employed to assess the quality of studies included. A maximum of 9 stars could be awarded to each study and a maximum of 2 stars for comparability. The 2 authors (AK and FA) independently assessed the risk of bias. The senior author (AFAR) was consulted with regard to the discrepancies to resolution purposes.

#### **Statistical analysis**

The association between dietary intake and the risk of ovarian cancer were our main analytical object. Dietary in this meta-analysis was defined as fruits, vegetables and mushrooms, meats and eggs, dairy products, fats and fatty acids, alcohol, micro/macronutrients and plant-based bioactive compounds, coffee and tea, non-food contaminants and sweets and carbohydrate foods. The random-effects model was used when data were significantly heterogeneous and the fixed-effect model was used when data were homogeneous. Heterogeneity was investigated using the Q and the  $I^2$ statistical tests with P value. P < 0.050 was considered statistically significant. Subgroup analyses stratified by types of dietary for each group and geographic region/country were carried out to investigate potential sources of heterogeneity. The STATA/MP 15.1 software (StataCorp LP, College Station, TX, USA) was used for statistical analyses.

#### Results

#### Literature retrieval results

The literature searches initially produced 25,722 records: 11,434 PubMed, 9598 Scopus and 4690 from Wiley Online Library citations. Subsequent to duplicate removal (n = 6858) and application of exclusion criteria (n = 18,864), 114 records were identified as being qualified for full-text review from which 18 were excluded after the application of inclusion and exclusion criteria. The Supplementary Table S4 depicts a list of the excluded full-text articles and justifications for exclusion. This finally led to the inclusion of a total of 97 studies [4, 8, 18, 20, 28–120]. Figure 1 shows detailed information on the selection process.

#### **Characteristics of included studies**

The included studies incorporated data analyses using 97 cohorts covering the European Prospective Investigation into Cancer and Nutrition (EPIC) in the Europe (6 studies) [4, 46, 59, 90, 93, 108], Australian Ovarian Cancer Study (AOCS) in the Australia (1 study) [8], Multi-centers studies in the North America and Europe (6 studies) [18, 48, 55, 71, 78, 85], Japan Collaborative Cohort (JACC) Study in the Japan (1 study) [20], NLCS in the Netherlands (7 studies) [28, 39, 44, 54, 73, 96, 117], UK Women's Cohort Study (UKWCS) in the United Kingdom (1 study) [29], EPIC and

the Netherlands Cohort Study (NLCS) in the Netherlands (1 study) [30], Women's Lifestyle and Health Cohort Study (WLH) in the Sweden (1 study) [31], A Cohort of American Women in the United States (1 study) [32], Oxford Vegetarian Study (OVS) and the EPIC-Oxford Cohort in the United Kingdom (1 study) [33], National Institutes of Health (NIH)-AARP (American Association for Retired Persons) Diet and Health Study in the United States (8 studies) [34, 45, 47, 52, 61, 64, 112, 119], California Teachers Study (CTS) in the United States (3 studies) [35, 36, 69], The Nurses' Health Study (NHS) in the United States (7 studies) [37, 41, 42, 56, 67, 72, 115], Adventist Health Study (AHS) in the United States (1 study) [38], Swedish Mammography Cohort (SMC) in the Sweden (10 studies) [40, 49, 57, 74, 76, 98, 99, 107, 114, 116], Iowa Women's Health Study (IWHS) in the United States (8 studies) [43, 75, 87, 95, 101, 106, 109, 118], Mobile Health Clinic of the Social Insurance Institution in the Finland [50], The NHS and NHS II in the United States (5 studies) [51, 58, 79–81], Breast Cancer Detection Demonstration Project (BCDDP) Follow-up Cohort in the United States (1 study) [53], Norwegian Counties Study in the Norway (1 study) [60], Women's Health Initiative (WHI) in the United States (3 studies) [62, 83, 110], Canadian Study of Diet, Lifestyle, and Health (CSDLH) in the Canada (3 studies) [63, 77, 88], Japan Public Health Centerbased Prospective (JPHC) Study in the Japan (2 studies) [65, 105], Million Women Study (MWS) in the United Kingdom (1 study) [66], Canadian National Breast Screening Study (CNBSS) in the Canada (6 studies) [68, 70, 84, 86, 97, 120], Women's Health Study (WHS) in the United States (1 study) [82], Multiethnic Cohort Study (MEC) in the United States (1 study) [89], Norwegian Women and Cancer (NOWAC) Study in the Norway (1 study) [91], Prostate, Lung, Colorectal, and Ovarian (PLCO) Prospective Study in the United States (1 study) [92], Västerbotten Intervention Project (VIP) in the Sweden (1 study) [94], Prospective Cohort Study in Hangzhou in the China (1 study) [100], Norwegian Prospective Study in the Norway (2 study) [102, 103], Seventh-Day Adventists Cohort Study in the United States (1 study) [104], Diet, Cancer and Health (DCH) Cohort in the Denmark (1 study) [111] and (EPIC)-Norfolk Study in the United Kingdom (1 study) [113]. The studies covered periods from 1984 to 2019, with populations ranging from 254 to 1,280,296. Sixty-five studies incorporated participants whose age was both lower and higher than 50 years, 30 studies included people whose age exceeded 50 years and 2 studies were 30-49 years old. Most studies represented incident ovarian cancer outcome, 4 studies had ovarian cancer death outcome [20, 90, 100, 104] and only one study depicted ovarian cancer incidence and death outcome [70]. Thirty-seven studies presented data on relationships for histologic and ovarian cancer subtypes [4, 18, 30, 31, 36, 37, 40, 41, 46, 48, 49, 51, 54-59, 61, 62, 64, 67, 71, 72, 76, 78-80, 85, 96, 98, 99, 108,

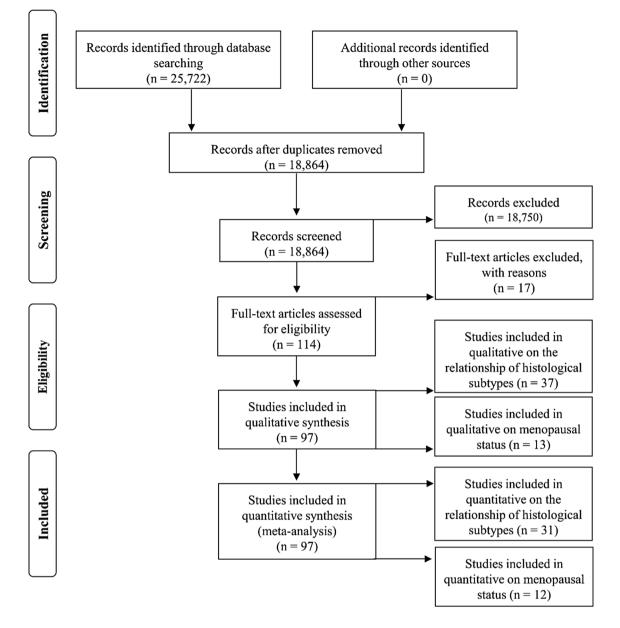


Fig. 1 Study selection process and results in accordance with the PRISMA framework

112, 114–116]. Thirteen studies furnished data on menopausal status [29, 30, 36, 38, 48, 51, 58, 67, 80, 84, 105, 115, 120]. Nineteen studies conducted the relationship between dietary intake and ovarian cancer risk in postmenopausal women participating [28, 39, 43, 44, 54, 62, 63, 73, 75, 83, 87, 95, 96, 101, 106, 109–111, 117].

#### Assessment of risk of bias

Supplementary Tables S5–13 depict the assessment of risk of bias for the 97 cohort studies undergoing additional metaanalyses. Most studies did not include any missing outcome data, or presented missing data as being insignificant for the study to be assessed at high-risk of attrition bias. The entirety of the included studies was deemed to be at low risk of selective reporting bias considering the fact that all the outcomes mentioned in method sections were incorporated in the results.

# Intake of fruits, vegetables and mushrooms and ovarian cancer risk

Supplementary Table S14 shows that 20 cohort studies [4, 8, 18, 20, 28–43] were considered in meta-analysis of the relationship between intake of fruits, vegetables and mushrooms and ovarian cancer risk. Comparison of the highest and

lowest intake categories yielded a summary RR of 0.98 (95% CI 0.91–1.05) without heterogeneity ( $P_{\text{heterogeneity}} = 0.866$ ,  $I^2 = 0.0\%$ ) (Fig. 2a). These results suggested no association between intake of fruits, vegetables and mushrooms and ovarian cancer risk (P=0.513). Results of the subgroup analyses were consistent with the findings of the main analysis; there was no association between intake of fruits and risk of ovarian cancer (Table 1). From among the 20 studies on fruit consumption, 5 revealed a significantly lower risk of ovarian cancer [28, 30, 32, 33, 38] while 15 [4, 8, 18, 20, 29, 31, 34–37, 39–43] reported non-significant relationships. The results of subgroup analyses indicated that green leafy vegetables (RR = 0.91, 95% CI 0.85-0.98,  $P_{\text{heterogeneity}} = 0.115, I^2 = 35.5\%, P = 0.009$ ) and allium vegetables (RR = 0.79, 95% CI 0.64–0.96,  $P_{\text{heterogeneity}} = 0.996$ ,  $I^2 = 0.0\%$ , P = 0.021) intake could significantly decrease ovarian cancer risk (Table 1). However, there was no significant association between intake of fruits, vegetables and mushrooms and ovarian cancer when subgroup analysis was performed according to geographic location/country (Table 2). From among 20 studies on vegetable consumption, 7 reported a significantly lower risk of ovarian cancer [8, 28, 32, 33, 40, 41, 43], while 11 [4, 18, 29–31, 34–39] reported no significant relationships. Fairfield et al. [42] reported a moderate risk reduction with high intake of vegetable but this was not found to be statistically significant. In one study [20], a lower ovarian cancer risk was reported with soy bean curd (tofu) consumption while observing a significant higher risk with Chinese cabbage. A lower risk of ovarian cancer was reported with green leafy vegetables [8, 38-40, 43], bananas [30], yellow and cruciferous vegetables [32], e.g., broccoli [37] and cabbage and carrots, beets [40] and tomato [38, 40]. Considering the 2 studies [4, 29] on mushrooms consumption, 1 [29] was found to reveal a significantly higher risk, while the other [4] reported no significant relationships. The data are indicative of the fact that no significant associations with histologic subtypes of ovarian cancer are reported with fruit, vegetable and mushroom consumption (Table 1) [4, 18, 30]. No significant relationships were observed with menopausal status and intake of fruits, P = 0.229 [29, 30, 38]. For tomatoes and other fruits, the statistically significant relatively strong effects were larger among postmenopausal women [38].

#### Intake of meats and eggs and ovarian cancer risk

Supplementary Table S15 outlines the 17 cohort studies [8, 20, 29, 30, 32, 33, 36, 38, 41, 43–50] considered in the meta-analysis of the relationship between intake of meats and eggs and the risk of ovarian cancer. Figure 2b, reported an association between intake of meats and eggs and the risk of ovarian cancer, and there was no significant increase association (RR = 1.06, 95% CI 0.97–1.16,  $P_{\text{heterogeneity}} = 0.277$ ,  $I^2 = 15.1\%$ , P = 0.196). A moderate risk increasing was found with intake of red meat, pork, eggs and dried and salted fish and canned tuna but this was not statistically significant. The results indicated that a moderate risk reduction was obtained by intake of fish and fish dishes (Table 1). Results of the subgroup analyses indicated that intake of meats and eggs could increase ovarian cancer risk in United States populations (RR = 1.15, 95% CI 1.10–1.30,  $P_{\text{heterogeneity}} = 0.426$ ,  $I^2 = 0.5\%$ , P = 0.028) (Table 2). From among the 14 studies [8, 20, 29, 30, 32, 38, 41, 43–47, 49, 50] on total meat consumption, 2 studies revealed a significant higher risk, especially the red and cured or processed meat subgroups [32] and total meat, total beef, poultry [38], and fried meat [50], while 11 [8, 20, 29, 30, 41, 43-47, 49] reported no significant relationships. Two [8, 33] of the 11 studies [8, 20, 29, 30, 32, 33, 38, 44–46, 49] on fish consumption and the risk of ovarian cancer found significantly lower risks, while 6 [29, 30, 32, 44, 46, 49] reported no significant relationships. Three studies also found a significant higher risk with dried or salted fish consumption [20], fish [38] and canned tuna [45]. From among the 9 studies [20, 29, 36, 38, 41, 43, 46, 48, 49] on egg consumption, 1 study [43] revealed a statistically significant higher risk, while 7 [20, 29, 36, 38, 46, 48, 49] reported no significant relationships. Bertone et al. [41] reported higher risks of ovarian cancer with frequent egg intake. No significant relationships were reported between histologic subtypes of ovarian cancer and meat and egg consumption [36, 48, 49]. A statistically significant relationship was obtained between serous tumors and total meat and poultry. The RR for total meat and poultry intake was 1.29 (95% CI 1.11–1.50, P = 0.001), suggesting a significant increase association between total meat and poultry consumption and serous ovarian cancer risk [46]. No significant relationships were reported with menopausal status and meat and egg consumption (RR = 1.64, 95% CI 0.83–3.24, P=0.155) [29, 38]. The statistically significant strong effect reported for meat was larger for postmenopausal women [38].

#### Intake of dairy products and ovarian cancer risk

The meta-analysis included 17 cohort studies [8, 20, 29, 30, 32, 36, 38, 41, 43, 46, 51–57] on the relationship between intake of dairy products and the risk of ovarian cancer as detailed in Supplementary Table S16. Dietary intake of dairy products was not significantly associated with the risk of ovarian cancer (RR = 1.01, 95% CI 0.94–1.08,  $P_{\text{heterogeneity}} = 0.881$ ,  $I^2 = 0.0\%$ , P = 0.762) (Fig. 3a). Subgroup analyses revealed a moderate risk increasing of ovarian cancer with the intake of high-fat dairy products and total cheese; while, a moderate risk reduction was achieved with the intake of low-fat dairy

Study	Risk (95% CI)
Nieuwenhuis and van den Brandt 2019 [28]	0.84 (0.53, 1.34)
Dunneram et al. 2018 [29]	0.98 (0.71, 1.42)
Playdon et al. 2017 [8]	0.92 (0.71, 1.19)
Merritt et al. 2016 [30]	0.95 (0.78, 1.16)
Merritt et al. 2016 [30]	1.14 (0.86, 1.51)
Hedelin et al. 2011 [31]	1.29 (0.82, 2.00)
Dolecek et al. 2010 [32]	0.82 (0.50, 1.33)
Key et al. 2009 [33]	0.69 (0.45, 1.07)
George et al. 2009 [34]	1.11 (0.82, 1.50)
Chang et al. 2008 [35]	• 1.25 (0.82, 1.89)
Chang et al. 2007 [36]	0.91 (0.61, 1.34)
Sakauchi et al. 2007 [20]	♦ 1.98 (0.55, 9.86)
Gates et al. 2007 [37]	0.95 (0.59, 1.54)
Kiani et al. 2006 [38]	0.50 (0.23, 1.10)
Koushik et al. 2005 [18]	1.00 (0.77, 1.35)
Mommers et al. 2005 [39]	• 1.08 (0.67, 1.73)
Schulz et al. 2005 [4]	• 0.98 (0.87, 1.11)
Larsson et al. 2004 [40]	0.96 (0.65, 1.43)
Bertone et al. 2002 [41]	0.92 (0.62, 1.36)
Fairfield et al. 2001 [42]	1.06 (0.69, 1.63)
Kushi et al. 1999 [43]	0.78 (0.44, 1.36)
Overall (I-squared = $0.0\%$ , p = $0.866$ )	0.98 (0.91, 1.05)
I I .2 .5 T Reduces risk of OC	
	Increases risk of OC
	Relative %
Study	
	Relative %
Study Dunneram et al. 2018 [29]	Relative % Risk (95% CI) Wei
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8]	Relative         %           Risk (95% CI)         Wei           0.83 (0.34, 4.25)         0.50
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30]	Relative         %           Risk (95% CI)         Wei           0.83 (0.34, 4.25)         0.50           0.97 (0.74, 1.27)         10.9           0.98 (0.80, 1.21)         18.5
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44]	Relative         %           Risk (95% CI)         Wei           0.83 (0.34, 4.25)         0.50           0.97 (0.74, 1.27)         10.5           0.98 (0.80, 1.21)         18.3           -         1.02 (0.68, 1.56)         4.66
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45]	Relative         %           Risk (95% CI)         Wei           0.83 (0.34, 4.25)         0.50           0.97 (0.74, 1.27)         10.5           0.98 (0.80, 1.21)         18.3           -         1.02 (0.68, 1.56)         4.67           1.14 (0.91, 1.43)         15.5
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.5         -       1.02 (0.68, 1.56)       4.65         1.14 (0.91, 1.43)       15.5         •       1.65 (0.91, 2.98)       2.26
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32]	Relative         %           Risk (95% CI)         Wei           0.83 (0.34, 4.25)         0.50           0.97 (0.74, 1.27)         10.5           0.98 (0.80, 1.21)         18.3           -         1.02 (0.68, 1.56)         4.67           1.14 (0.91, 1.43)         15.5
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.5         -       1.02 (0.68, 1.56)       4.65         1.14 (0.91, 1.43)       15.5         •       1.65 (0.91, 2.98)       2.26
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33] Chang et al. 2007 [36]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.5         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.65         1.14 (0.91, 1.43)       15.3         1.65 (0.91, 2.98)       2.20         0.37 (0.18, 0.77)       1.50
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2007 [36] Chang et al. 2007 [36]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.66         1.14 (0.91, 1.43)       15.3         •       1.65 (0.91, 2.98)       2.20         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33] Chang et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [20]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.63         1.14 (0.91, 1.43)       15.3         1.65 (0.91, 2.98)       2.20         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         1.38 (0.55, 3.48)       0.93
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33] Chang et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [20] Cross et al. 2007 [47]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.67         1.14 (0.91, 1.43)       15.3         -       1.65 (0.91, 2.98)       2.20         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         -       1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.92         -       1.21 (0.91, 1.61)       9.77
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33] Chang et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [47] Genkinger et al. 2006 [48]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.66         1.14 (0.91, 1.43)       15.3         -       1.65 (0.91, 2.98)       2.26         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         -       1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.93         1.21 (0.91, 1.61)       9.75         -       1.18 (0.89, 1.57)       9.85
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33] Chang et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [46] Cross et al. 2007 [47] Genkinger et al. 2006 [48] Kiani et al. 2006 [38]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.67         1.14 (0.91, 1.43)       15.3         -       1.65 (0.91, 2.98)       2.20         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         -       1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.92         -       1.21 (0.91, 1.61)       9.77
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33] Chang et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [46] Cross et al. 2007 [47] Genkinger et al. 2006 [48] Kiani et al. 2006 [38]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.66         1.14 (0.91, 1.43)       15.3         -       1.65 (0.91, 2.98)       2.26         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         -       1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.93         1.21 (0.91, 1.61)       9.75         -       1.18 (0.89, 1.57)       9.85
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2007 [36] Schulz et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [46] Cross et al. 2007 [47] Genkinger et al. 2006 [48] Kiani et al. 2006 [48] Larsson and Wolk 2005 [49]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.5         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.66         1.14 (0.91, 1.43)       15.5         4       1.65 (0.91, 2.98)       2.20         0.37 (0.18, 0.77)       1.56         0.78 (0.53, 1.15)       5.30         -       1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.92         -       1.18 (0.89, 1.57)       9.87         1.28 (0.65, 2.53)       1.72
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33] Chang et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [47] Genkinger et al. 2006 [48] Kiani et al. 2006 [48] Larsson and Wolk 2005 [49] Bertone et al. 2002 [41]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.3         1.02 (0.68, 1.56)       4.63         1.14 (0.91, 1.43)       15.3         1.65 (0.91, 2.98)       2.20         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.92         1.21 (0.91, 1.61)       9.77         1.18 (0.89, 1.57)       9.83         1.28 (0.65, 2.53)       1.72         1.01 (0.67, 1.53)       4.60
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33] Chang et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [47] Genkinger et al. 2006 [48] Kiani et al. 2006 [48] Larsson and Wolk 2005 [49] Bertone et al. 2002 [41] Kushi et al. 1999 [43]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.5         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.6         1.14 (0.91, 1.43)       15.3         -       1.65 (0.91, 2.98)       2.20         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         -       1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.92         -       1.21 (0.91, 1.61)       9.77         -       1.18 (0.89, 1.57)       9.83         1.28 (0.65, 2.53)       1.77         -       1.01 (0.67, 1.53)       4.66         -       1.08 (0.76, 1.54)       6.37         -       1.71 (0.89, 3.28)       1.83
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [46] Sakauchi et al. 2007 [47] Genkinger et al. 2006 [48] Kiani et al. 2006 [48] Larsson and Wolk 2005 [49] Bertone et al. 2002 [41] Kushi et al. 1999 [43] Knekt et al. 1994 [50]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.3         1.02 (0.68, 1.56)       4.66         1.14 (0.91, 1.43)       15.3         1.65 (0.91, 2.98)       2.26         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.92         1.18 (0.89, 1.57)       9.83         1.21 (0.91, 1.61)       9.77         1.18 (0.89, 1.57)       9.83         1.28 (0.65, 2.53)       1.77         1.01 (0.67, 1.53)       4.66         1.08 (0.76, 1.54)       6.33         1.71 (0.89, 3.28)       1.83         1.73 (0.48, 6.17)       0.49
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2007 [36] Chang et al. 2007 [36]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.5         0.98 (0.80, 1.21)       18.3         -       1.02 (0.68, 1.56)       4.6         1.14 (0.91, 1.43)       15.3         -       1.65 (0.91, 2.98)       2.20         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         -       1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.92         -       1.21 (0.91, 1.61)       9.77         -       1.18 (0.89, 1.57)       9.83         1.28 (0.65, 2.53)       1.77         -       1.01 (0.67, 1.53)       4.66         -       1.08 (0.76, 1.54)       6.37         -       1.71 (0.89, 3.28)       1.83
Study Dunneram et al. 2018 [29] Playdon et al. 2017 [8] Merritt et al. 2016 [30] Gilsing et al. 2011 [44] Daniel et al. 2011 [45] Dolecek et al. 2010 [32] Key et al. 2009 [33] Chang et al. 2007 [36] Schulz et al. 2007 [46] Sakauchi et al. 2007 [46] Sakauchi et al. 2007 [47] Genkinger et al. 2006 [48] Kiani et al. 2006 [48] Larsson and Wolk 2005 [49] Bertone et al. 2002 [41] Kushi et al. 1999 [43] Knekt et al. 1999 [43]	Relative       %         Risk (95% CI)       Wei         0.83 (0.34, 4.25)       0.50         0.97 (0.74, 1.27)       10.9         0.98 (0.80, 1.21)       18.3         1.02 (0.68, 1.56)       4.66         1.14 (0.91, 1.43)       15.3         1.65 (0.91, 2.98)       2.26         0.37 (0.18, 0.77)       1.50         0.78 (0.53, 1.15)       5.30         1.04 (0.70, 1.54)       5.11         1.38 (0.55, 3.48)       0.92         1.18 (0.89, 1.57)       9.83         1.21 (0.91, 1.61)       9.77         1.18 (0.89, 1.57)       9.83         1.28 (0.65, 2.53)       1.77         1.01 (0.67, 1.53)       4.66         1.08 (0.76, 1.54)       6.33         1.71 (0.89, 3.28)       1.83         1.73 (0.48, 6.17)       0.49

Fig. 2 Forest plot for cohort studies evaluating pooled relative risk (RR) of intake of **a** fruits, vegetables and mushrooms; **b** meats and eggs and ovarian cancer risk. *OC* ovary cancer

 Table 1
 Summary of the pooled risk estimates of the association between dietary intake and the risk of ovarian cancer according to subgroups analysis by types of dietary

Subgroup analysis by types of dietary	No. of studies	RR (95% CI)	Р	Pheterogeneity	$I^{2}(\%)$
Fruits, vegetables and mushrooms	·	·			
Total [4, 8, 18, 20, 28–43]	20	0.92 (0.99-1.06)	0.707	0.798	0.0
Fruits and vegetables [4, 18, 32, 39, 40, 42]	6	1.00 (0.93-1.07)	0.952	0.268	22.0
Fruits [4, 8, 18, 20, 29, 30, 32, 34, 38-40, 42, 43]	13	1.01 (0.96–1.07)	0.604	0.142	29.5
Apples, pears, pomes, applesauce [18, 29, 30, 37, 40]	5	1.01 (0.94–1.09)	0.789	0.625	0.0
Berries [29, 31]	2	1.09 (0.60–1.97)	0.769	0.039	76.5
Bananas [18, 29, 30, 40]	4	1.02 (0.88–1.19)	0.762	0.088	50.5
Citrus family fruits [18, 20, 29, 30, 32, 40, 42]	7	0.98 (0.91-1.06)	0.673	0.974	0.0
Juices [18, 20, 29, 30]	4	1.00 (0.94–1.06)	0.902	0.899	0.0
Vegetables [4, 8, 18, 29, 32, 34, 39, 40, 42, 43]	10	0.96 (0.89–1.03)	0.215	0.207	21.2
Leafy vegetables [4, 38, 39]	3	0.98 (0.90-1.07)	0.681	0.496	0.0
Green leafy vegetables [8, 18, 20, 29, 32, 37, 38, 40, 42, 43]	10	0.91 (0.85-0.98)	0.009	0.115	35.5
Red/yellow vegetables [8, 32]	2	0.77 (0.53-1.14)	0.192	0.131	56.1
Root vegetables [4, 18, 29, 30, 37, 40]	6	0.99 (0.90-1.07)	0.737	0.284	19.8
Cruciferous vegetables [8, 18, 29, 32, 36, 39, 42]	7	1.01 (0.90-1.12)	0.916	0.739	0.0
Broccoli [18, 36, 37]	3	0.84 (0.69–1.02)	0.075	0.466	0.0
Cabbages [4, 18, 36, 40]	4	1.00 (0.91-1.10)	0.969	0.836	0.0
Chinese cabbage [20]	1	10.28 (1.38-76.84)	N/A	N/A	N/A
Fruiting vegetables [4, 18, 20, 29, 37, 38, 40]	7	0.94 (0.85-1.05)	0.275	0.195	30.6
Allium vegetables [4, 29, 37]	3	0.79 (0.64–0.96)	0.021	0.996	0.0
Potatoes [18, 20, 29, 30]	4	0.96 (0.85–1.08)	0.474	0.849	0.0
Legumes [18, 29, 32, 37, 39, 42]	6	1.01 (0.88–1.15)	0.932	0.251	21.6
Soybean [20, 29, 36]	3	0.86 (0.72–1.04)	0.131	0.736	0.0
Nuts [28–31, 41]	5	1.00 (0.88–1.12)	0.949	0.708	0.0
Grains [8, 29–32]	5	1.01 (0.95–1.08)	0.661	0.331	12.2
Mushrooms [4, 29]	2	1.22 (0.80–1.84)	0.355	0.022	81.1
Others [18, 20, 29, 33, 35, 42]	6	1.03 (0.94–1.13)	0.513	0.390	5.6
Meats and eggs	0		01010	0.070	010
Total [8, 20, 29, 30, 32, 33, 36, 38, 41, 43–50]	17	1.14 (0.96–1.35)	0.124	0.906	0.0
Total meat [8, 29, 32, 38, 41, 43, 44, 46]	8	1.07 (0.94–1.22)	0.316	0.041	47.2
Red meat [8, 29, 30, 32, 44, 46, 47, 49]	8	1.08 (0.96–1.22)	0.211	0.162	33.3
Pork [20, 30, 44]	3	1.11 (0.93–1.32)	0.245	0.746	0.0
Beef [20, 38, 44]	3	1.15 (0.85–1.56)	0.364	0.978	0.0
Poultry [8, 20, 29, 30, 38, 41, 44–46]	9	0.98 (0.89–1.07)	0.653	0.873	0.0
Fish and fish dishes [8, 20, 29, 30, 33, 38, 44–46, 49]	10	0.92 (0.83–1.02)	0.108	0.287	15.6
Processed meat [8, 20, 29, 30, 32, 41, 44, 46, 47]	9	1.05 (0.94–1.17)	0.361	0.278	18.5
Eggs/eggs dishes [20, 29, 36, 38, 41, 43, 46, 48, 49]	9	1.12 (0.96–1.30)	0.140	0.278	23.8
Offal [29, 44]	2	1.24 (0.64–2.42)	0.523	0.332	0.0
Dried/salted fish/canned tuna [20, 45]	2	1.24(0.04-2.42) 1.67(0.71-3.45)	0.167	0.098	63.4
	1	1.73 (0.48–6.17)		0.098 N/A	03.4 N/A
Fried meat [50] Others [32, 44]	2	0.90 (0.66–1.22)	N/A 0.491	0.714	0.0
	2	0.90 (0.00–1.22)	0.491	0.714	0.0
Dairy products	17	1.04 (0.05, 1.12)	0.452	0.094	0.0
Total [8, 20, 29, 30, 32, 36, 38, 41, 43, 46, 51–57]	17	1.04 (0.95–1.13)	0.452	0.984	0.0
Dairy foods [32, 36, 43, 46, 52, 53, 57]	7	1.13 (0.87–1.46)	0.369	0.013	58.8
Low fat dairy products [8, 29]	2	0.95 (0.89–1.02)	0.142	0.843	0.0
High fat dairy products [8, 29]	2	1.06 (1.00–1.13)	0.059	0.722	0.0
Total milk [20, 30, 32, 36, 46, 51, 53, 55, 57]	9	1.03 (0.93–1.15)	0.546	0.193	27.3
Whole milk [32, 38, 51, 53, 55, 56]	6	1.04 (0.85–1.27)	0.683	0.749	0.0
Skim/low/reduced-fat milk [32, 38, 43, 51, 53, 55, 56]	7	1.03 (0.86–1.25)	0.739	0.014	56.6

Table 1 (	continued)
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Subgroup analysis by types of dietary	No. of studies	RR (95% CI)	Р	Pheterogeneity	$I^{2}(\%)$	
Butter [20, 29, 30, 41, 53]	5	1.00 (0.89–1.14)	0.962	0.160	39.2	
Total cheese [20, 30, 38, 43, 46, 53, 54, 57]	8	1.11 (0.97–1.28)	0.132	0.653	0.0	
Hard cheese [41, 55, 56]	3	0.88 (0.56-1.39)	0.580	0.016	75.7	
Cottage cheese [53, 55, 56]	3	0.90 (0.69–1.18)	0.464	0.788	0.0	
Ice cream [30, 53, 55, 56]	4	0.93 (0.79-1.09)	0.353	0.322	14.0	
Lactose [38, 43, 51, 53–56]	7	1.02 (0.84–1.24)	0.855	0.010	64.2	
Yogurt [20, 30, 46, 54–57]	7	1.03 (0.93–1.14)	0.620	0.761	0.0	
Others [32, 54]	2	0.80 (0.61-1.05)	0.109	0.384	0.0	
Fats and fatty acids						
Total [8, 28–30, 32, 35, 36, 38, 41, 43, 44, 48, 51, 54, 58–61]	18	1.06 (0.98–1.14)	0.139	0.972	0.0	
Total fat [8, 30, 36, 41, 43, 44, 48, 58, 59, 61]	10	1.10 (1.02–1.18)	0.009	0.435	0.6	
Animal fat [30, 41, 43, 44, 48, 58, 59, 61]	8	1.08 (1.00-1.16)	0.059	0.050	46.7	
Dairy fat [38, 41, 44, 51, 54]	5	1.09 (0.95-1.25)	0.235	0.327	13.7	
Plant fat [30, 41, 43, 44, 48, 58, 59, 61]	8	1.04 (0.95–1.13)	0.411	0.118	37.6	
Peanut butter [28, 41]	2	0.88 (0.65-1.19)	0.418	0.320	0.0	
Margarine [29, 30, 41, 44]	4	1.03 (0.92–1.15)	0.636	0.673	0.0	
Saturated fat [8, 36, 41, 43, 48, 58, 59, 61]	8	1.11 (1.01–1.22)	0.023	0.380	6.6	
Saturated fatty acid [30, 41, 44]	3	1.19 (1.04–1.36)	0.010	0.520	0.0	
Monounsaturated fat [8, 41, 43, 48, 58–61]	7	1.01 (0.92–1.11)	0.834	0.474	0.0	
Monounsaturated fatty acid [30, 41, 44]	3	1.09 (0.92–1.29)	0.332	0.625	0.0	
Polyunsaturated fat [8, 41, 43, 48, 58–61]	7	1.04 (0.90–1.20)	0.588	0.038	52.8	
Polyunsaturated fatty acid [30, 41, 44]	3	1.01 (0.91–1.13)	0.798	0.228	24.2	
<i>Trans</i> fat [58, 61]	2	0.99 (0.85–1.16)	0.931	0.142	45.0	
Trans fatty acid [60]	1	1.15 (0.74–1.78)	0.991 N/A	N/A	N/A	
Trans unsaturated fat [48]	1	1.04 (0.84–1.28)	N/A	N/A	N/A	
Trans unsaturated fatty acid [44]	1	1.51 (1.04–2.20)	N/A	N/A	N/A	
Cholesterol [30, 41, 43, 48, 58, 59]	6	1.13 (1.04–1.22)	0.002	0.263	21.0	
Others [8, 29, 32, 35, 41, 44, 59]	7	0.96 (0.87–1.06)	0.002	0.203	41.8	
Alcohol	7	0.90 (0.87–1.00)	0.410	0.079	41.0	
Total [20, 29, 30, 32, 43, 62–76]	20	1.06 (0.91–1.24)	0.474	0.903	0.0	
	16	1.00 (0.91–1.24)	0.990	0.505	0.0	
Alcohol [20, 30, 32, 43, 62–69, 71, 73–75]	7	1.07 (0.96–1.18)	0.990	0.310	4.5	
Wines [29, 30, 67, 69, 71, 73, 74]	7	1.05 (0.93–1.19)	0.214	0.397	4.5 0.0	
Beer and cider [29, 30, 67, 69, 71, 73, 74]	3	0.90 (0.71–1.13)	0.437	0.478	0.0	
Liquor [67, 69, 73]		· · · · · ·				
Spirits [29, 71]	2	1.21 (0.99–1.49)	0.062	0.693	0.0	
Alcohol and folate [70, 72, 75, 76]	4	1.00 (0.87–1.15)	0.994	0.059	42.5	
Others [29, 72]	2	1.06 (0.88–1.28)	0.562	0.273	22.1	
Micro/macronutrients and plant-based bioactive compounds	20	0.07 (0.00, 1.07)	0.504	0.071	0.0	
Total [8, 29–31, 36–38, 42, 43, 51–53, 55, 68, 70, 72, 75–87]	29	0.97 (0.89–1.07)	0.584	0.971	0.0	
Vitamin A [42, 43, 78, 83, 86]	5	0.99 (0.91–1.08)	0.822	0.869	0.0	
Retinol [8, 30, 36, 42, 43]	5	1.14 (1.00–1.30)	0.048	0.839	0.0	
Vitamin B [8, 30, 68, 72, 77]	5	0.98 (0.89–1.08)	0.697	0.264	20.1	
Thiamin [30, 68, 77]	3	0.84 (0.70–1.02)	0.082	0.623	0.0	
Riboflavin [30, 68, 77]	3	0.77 (0.49–1.19)	0.238	0.026	72.5	
Folate [36, 43, 68, 70, 72, 75–78]	9	0.99 (0.85–1.15)	0.856	0.013	54.0	
Vitamin C [8, 30, 36, 42, 43, 78, 83, 86]	8	1.01 (0.94–1.08)	0.787	0.579	0.0	
Vitamin D [30, 36, 43, 53, 55, 80]	6	1.02 (0.92–1.13)	0.731	0.793	0.0	
Vitamin E [8, 30, 36, 42, 43, 78, 83, 86]	8	1.05 (0.98–1.13)	0.150	0.498	0.0	
Carotenoids [8, 30, 36, 42, 43, 83, 85, 86]	8	0.98 (0.93-1.02)	0.296	0.997	0.0	
Fiber [8, 29–31, 36, 43, 84]	7	0.89 (0.81-0.98)	0.014	0.423	2.5	

 Table 1 (continued)

Subgroup analysis by types of dietary	No. of studies	RR (95% CI)	Р	P <sub>heterogeneity</sub>	$I^{2}(\%)$	
Flavonoids [31, 36, 37, 79, 82, 87]	6	0.83 (0.78–0.89)	0.0001	0.178	22.0	
Protein [30, 36, 38, 43]	4	0.96 (0.88-1.06)	0.430	0.741	0.0	
Calcium [30, 36, 43, 51–53, 55]	7	1.00 (0.92-1.08)	0.925	0.142	31.1	
Others [8, 30, 31, 36, 68, 70, 72, 80, 81, 83]	10	1.00 (0.94–1.06)	0.994	0.996	0.0	
Coffee and tea						
Total [8, 29, 30, 36, 37, 67, 88-104]	23	0.92 (0.80-1.07)	0.283	0.689	0.0	
Tea [8, 29, 36, 67, 88, 89, 92, 93, 96, 97, 99, 101]	12	0.96 (0.89-1.03)	0.252	0.286	16.1	
Green tea [8, 100]	2	0.61 (0.49-0.76)	0.0001	0.200	31.5	
Coffee [8, 29, 30, 36, 88–98, 102–104]	18	1.00 (0.85-1.17)	0.970	0.0001	69.9	
Caffeinated coffee [67, 88, 89, 93, 95]	5	0.98 (0.65-1.49)	0.929	0.002	75.9	
Decaffeinated coffee [67, 88, 89, 93, 95]	5	0.87 (0.67-1.14)	0.311	0.941	0.0	
Caffeine [67, 88, 89, 92, 95]	5	0.91 (0.69–1.19)	0.487	0.059	56.0	
Soft drinks [29, 30, 67]	3	1.04 (0.93–1.18)	0.475	0.309	16.4	
Others [29, 37]	2	0.92 (0.79-1.08)	0.324	0.171	43.4	
Non-food contaminants						
Total [105–118]	14	1.06 (0.91-1.23)	0.474	0.719	0.0	
Acrylamide [105, 108, 115–117]	5	1.18 (1.02–1.37)	0.027	0.003	54.4	
Atrazine [106]	1	0.95 (0.58-1.55)	N/A	N/A	N/A	
Cadmium [110, 111, 114]	3	0.93 (0.79-1.09)	0.389	0.219	32.2	
Nitrate [109, 112, 118]	3	1.36 (1.02–1.80)	0.034	0.0001	69.1	
Nitrite [109, 112, 113]	3	1.07 (0.96–1.16)	0.227	0.126	41.8	
<i>N</i> -Nitroso compound [113]	1	0.90 (0.63-1.31)	N/A	N/A	N/A	
Others [107, 109, 112]	3	1.17 (1.03–1.33)	0.019	0.789	0.0	
Sweets and carbohydrate foods						
Total [29, 30, 32, 35, 36, 43, 119, 120]	8	0.99 (0.87–1.13)	0.939	0.635	0.0	
Biscuits [29, 30]	2	0.96 (0.85-1.09)	0.531	0.873	0.0	
Bread [30, 43]	2	0.99 (0.86–1.14)	0.924	0.904	0.0	
Cakes [29, 30]	2	1.03 (0.86–1.24)	0.726	0.808	0.0	
Carbohydrate [30, 35, 36, 43, 119, 120]	6	0.94 (0.81-1.08)	0.376	0.003	61.2	
Sweets [32, 43]	2	1.19 (0.82–1.72)	0.358	0.302	6.2	
Others [29]	1	0.84 (0.59-1.25)	N/A	N/A	N/A	

Data marked in bold are statistically significant

products (Table 1). However, there was no significant association between intake of dairy products and the risk of ovarian cancer when subgroup analysis was performed according to geographic location/country (Table 2). From among the 17 studies on consumption of dairy products, 9 [8, 20, 29, 30, 36, 41, 46, 52, 54] reported no significant relationships. Two studies reported a statistically significant higher risk of dairy products especially all types of milk [32], skim milk and cheese [43] consumption and ovarian cancer. A higher ovarian cancer risk was reported with larger whole fat milk and cheese intake. A low protective relationship has also been reported with low fat milk [38]. Two studies revealed that the intake of lactose was positively but not significantly related with an higher risk [43, 55] with the other 2 studies showing that the intake of lactose was inversely related to ovarian cancer risk [39, 51]. Larger intake of total dairy food was found to be related to a statistically significant lower ovarian cancer risk. A decreased ovarian cancer risk was reported with frequent butter, ice cream, and 2% milk intake [53]. Table 3 presents the results of histologic subtypes and ovarian cancer risk. No significant relationships were found between histologic subtypes of ovarian cancer and dairy product consumption (Table 3) [41, 51, 54–57]. One study suggested that high lactose or total milk intake was associated with higher endometrioid ovarian cancer risk [51]. Two studies demonstrated that high lactose and dairy product intake, especially milk, is related to the high serous ovarian cancer risk [56, 57]. The RR for intake of dairy products was 0.94 (95% CI 0.80-1.12, P=0.507), suggesting a no significant association between dairy product consumptions and menopausal status [29, 38, 51]. In

 Table 2
 Summary of the pooled risk estimates of the association between dietary intake and the risk of ovarian cancer according to subgroups analysis by geographic location/country

Subgroup analysis by geographic location/country	No. of studies	RR (95% CI)	Р	P heterogeneity	$I^{2}(\%)$	
All studies						
Total [4, 8, 18, 20, 28–120]	97	1.03 (0.96–1.11)	0.451	1.00	0.0	
Europe [4, 28–31, 33, 39, 40, 44, 46, 49, 50, 54, 57, 59, 60, 66, 73, 74, 76, 90, 91, 93, 94, 96, 98, 99, 102, 103, 107, 108, 111, 113, 114, 116, 117]	36	1.02 (0.90–1.15)	0.771	0.998	0.0	
United States [18, 32, 34–38, 41–43, 45, 47, 48, 51–53, 55, 56, 58, 61, 62, 64, 67, 69, 71, 72, 75, 78–83, 85, 87, 89, 92, 95, 101, 104, 106, 109, 110, 112, 115, 118, 119]	47	1.07 (0.94–1.21)	0.310	0.980	0.0	
Other [8, 20, 63, 65, 68, 70, 77, 84, 86, 88, 97, 100, 105, 120]	14	1.00 (0.97–1.15)	0.967	0.882	0.0	
Fruits, vegetables and mushrooms						
Total [4, 8, 18, 20, 28–43]	20	1.00 (0.79–1.26)	0.971	0.779	0.0	
Europe [4, 28-31, 33, 39, 40]	8	0.97 (0.89-1.06)	0.519	0.715	0.0	
United States [18, 32, 34–38, 41–43]	10	0.98 (0.86–1.11)	0.735	0.706	0.0	
Other [8, 20]	2	0.94 (0.73-1.21)	0.646	0.306	4.8	
Meats and eggs						
Total [8, 20, 29, 30, 32, 33, 36, 38, 41, 43–50]	17	1.21 (0.89–1.66)	0.231	0.963	0.0	
Europe [29, 30, 33, 44, 46, 49, 50]	7	0.96 (0.83-1.12)	0.613	0.251	23.4	
United States [32, 36, 38, 41, 43, 45, 47, 48]	8	1.15 (1.10–1.30)	0.028	0.426	0.5	
Other [8, 20]	2	1.00 (0.77-1.29)	0.984	0.472	0.0	
Dairy products						
Total [8, 20, 29, 30, 32, 36, 38, 41, 43, 46, 51–57]	17	1.08 (0.86-1.36)	0.500	0.791	0.0	
Europe [29, 30, 46, 54, 57]	5	0.99 (0.90-1.08)	0.756	0.575	0.0	
United States [32, 36, 38, 41, 43, 51–53, 55, 56]	10	1.05 (0.92-1.19)	0.466	0.817	0.0	
Other [8, 20]	2	1.09 (0.84–1.41)	0.529	0.386	0.0	
Fats and fatty acids						
Total [8, 28–30, 32, 35, 36, 38, 41, 43, 44, 48, 51, 54, 58–61]	18	1.04 (0.89–1.23)	0.612	0.961	0.0	
Europe [28–30, 44, 54, 59, 60]	7	1.13 (1.00–1.27)	0.045	0.821	0.0	
United States [32, 35, 36, 38, 41, 43, 48, 51, 58, 61]	10	1.01 (0.92–1.11)	0.799	0.804	0.0	
Other [8]	1	1.03 (0.80–1.33)	N/A	N/A	N/A	
Alcohol						
Total [20, 29, 30, 32, 43, 62–76]	20	1.00 (0.79–1.27)	0.991	0.753	0.0	
Europe [29, 30, 66, 73, 74, 76]	6	0.98 (0.89-1.09)	0.759	0.397	3.0	
United States [32, 43, 62, 64, 67, 69, 71, 72, 75]	9	1.01 (0.88–1.16)	0.842	0.405	3.6	
Other [20, 63, 65, 68, 70]	5	1.00 (0.99-1.01)	0.992	0.517	0.0	
Micro/macronutrients and plant-based bioactive compounds						
Total [8, 29–31, 36–38, 42, 43, 51–53, 55, 68, 70, 72, 75–87]	29	0.98 (0.83-1.16)	0.801	0.910	0.0	
Europe [29–31, 76]	4	0.95 (0.82-1.11)	0.533	0.413	0.0	
United States [36–38, 42, 43, 51–53, 55, 72, 75, 78–83, 85, 87]	19	0.96 (0.90-1.02)	0.212	0.805	0.0	
Other [8, 68, 70, 77, 84, 86]	6	0.91 (0.77-1.06)	0.224	0.678	0.0	
Coffee and tea						
Total [8, 29, 30, 36, 37, 67, 88–104]	23	0.91 (0.67-1.24)	0.554	0.915	0.0	
Europe [29, 30, 90, 91, 93, 94, 96, 98, 99, 102, 103]	11	0.84 (0.64–1.11)	0.215	0.0001	82.9	
United States [36, 37, 67, 89, 92, 95, 101, 104]	8	0.92 (0.77-1.10)	0.370	0.275	19.5	
Other [8, 88, 97, 100]	4	0.89 (0.59–1.35)	0.591	0.063	58.9	
Non-food contaminants		. ,				
Total [105–118]	14	1.12 (0.84–1.49)	0.430	0.466	0.0	
Europe [107, 108, 111, 113, 114, 116, 117]	7	0.96 (0.85–1.09)	0.542	0.186	31.7	
United States [106, 109, 110, 112, 115, 118]	6	1.03 (0.89–1.20)	0.697	0.204	30.9	
Other [105]	1	0.81 (0.45–1.51)	N/A	N/A	N/A	

Table 2 (continued)						
Subgroup analysis by geographic location/country	No. of studies	RR (95% CI)	Р	P heterogeneity	$I^{2}(\%)$	
Sweets and carbohydrate foods						
Total [29, 30, 32, 35, 36, 43, 119, 120]	8	1.03 (0.82–1.29)	0.827	0.578	0.0	
Europe [29, 30]	2	0.95 (0.80-1.14)	0.607	0.711	0.0	
United States [32, 35, 36, 43, 119]	5	1.10 (0.80–1.52)	0.555	0.029	63.0	
Other [120]	1	1.25 (0.75–2.07)	N/A	N/A	N/A	

Data marked in bold are statistically significant

the case of cheese, the statistically significant relatively strong effect was larger among postmenopausal women [38].

# Intake of dietary fats and fatty acids and ovarian cancer risk

Supplementary Table S17 outlines the 18 cohort studies [8, 28-30, 32, 35, 36, 38, 41, 43, 44, 48, 51, 54, 58-61] incorporated in the meta-analysis of the relationship between dietary fats and fatty acids consumption and the risk of ovarian cancer. Intake of dietary fats and fatty acids was not statistically significantly associated with an increased risk of ovarian cancer (RR = 1.06, 95% CI 0.99-1.13,  $P_{\text{heterogeneity}} = 0.849, I^2 = 0.0\%, P = 0.084$ ) (Fig. 3b). The subgroup analyses also showed total fat (RR = 1.10, 95% CI 1.02–1.18,  $P_{\text{heterogeneity}} = 0.435$ ,  $I^2 = 0.6\%$ , P = 0.009), saturated fat (RR = 1.11, 95% CI 1.01–1.22,  $P_{\text{heterogeneity}} = 0.380$ ,  $I^2 = 6.6\%$ , P = 0.023), saturated fatty acid (RR = 1.19, 95% CI 1.04–1.36,  $P_{\text{heterogeneity}} = 0.520$ ,  $I^2 = 0.0\%$ , P = 0.010) and cholesterol (RR = 1.13, 95% CI 1.04–1.22,  $P_{\text{heterogeneity}} = 0.263, I^2 = 21.0\%, P = 0.002)$  intake could significantly increase ovarian cancer risk (Table 1). In the subgroup analysis according to geographic location/country, a statistically increased risk for ovarian cancer was observed in the European populations (RR = 1.13, 95% CI 1.00–1.27,  $P_{\text{heterogeneity}} = 0.821, I^2 = 0.0\%, P = 0.045$ ) (Table 2). From among the 18 studies on consumption of dietary fats and fatty acids, 12 [28, 29, 32, 35, 36, 38, 41, 43, 48, 51, 59, 60] revealed non-significant relationships with ovarian cancer. Four studies reporting a significant high risk of cholesterol [30, 58], polyunsaturated and saturated fat [30], trans unsaturated fatty acids [44], total dietary fat intake especially from animal sources [58, 61] and ovarian cancer. A weak positive relationship was observed for the intake of polyunsaturated fat [61]. The highest dairy fat intake quintile was related to higher risks as compared with the lowest intake quintile while a positive relationship was not observed for serous epithelial ovarian cancer [54]. A lower risk of ovarian cancer was reported with poly- to mono-unsaturated fat ratio [8] and unsaturated fat [30]; while, a positive relationship was reported between intake of animal and saturated fat [48] and polyunsaturated fat [59] and ovarian cancer risk. Overall, no significant associations were obtained between histologic subtypes of ovarian cancer and consumption of dietary fats and fatty acids (Table 3) [30, 41, 48, 54, 58, 59, 61]. The high dairy fat intake was related to a very low significance increase with the risk of serous ovarian cancer [41]. The positive association between animal fat and cholesterol intakes and risk of ovarian cancer were similar for both serous tumors and non-serous tumors. In addition, the positive association between saturated fat intake and risk of ovarian cancer was observed for non-serous tumors [58]. Total fat and fat originating from animal sources were positively related with the risk of serous ovarian cancer [61]. No significant relationships with menopausal status were reported of intake of dietary fats and fatty acids (RR = 1.03, 95% CI 0.93-1.15, P=0.532) [29, 30, 38, 48, 51, 58]. A positive association between saturated fat intake and risk of ovarian cancer was reported among postmenopausal women [30].

#### Alcohol intake and ovarian cancer risk

As Supplementary Table S18 depicts, the meta-analysis of the relation between the intake of alcohol and the risk of ovarian cancer included 20 cohort studies [20, 29, 30, 32, 43, 62–76]. Figure 4a, reported no significant association between alcohol intake and the risk of ovarian cancer (RR = 1.00, 95% CI 0.90–1.00,  $P_{\text{heterogeneity}} = 0.601$ ,  $I^2 = 0.0\%$ , P = 0.985). Subgroup analyses indicated that the intake of wines and spirits was at a moderately increased risk of ovarian cancer, but not statistically significant (Table 1). Subgroup analysis by geographic location/country showed that there was no significant association between the intake of alcohol and the risk of ovarian cancer (Table 2). From among the 20 studies on the intake of alcohol, 11 studies [20, 29, 30, 32, 63, 65–68, 71, 73] revealed non-significant relationships with 1 study [43] finding decreased risk of ovarian cancer with the intake of alcohol. Chang et al. [69] also reported that the intake of alcohol is unlikely to influence the risk of ovarian cancer, while drinking wine was related to the higher ovarian cancer risk in the year prior to the baseline. Larsson and Wolk [74] asserted that light-tomoderate consumption of wine lowered the ovarian cancer

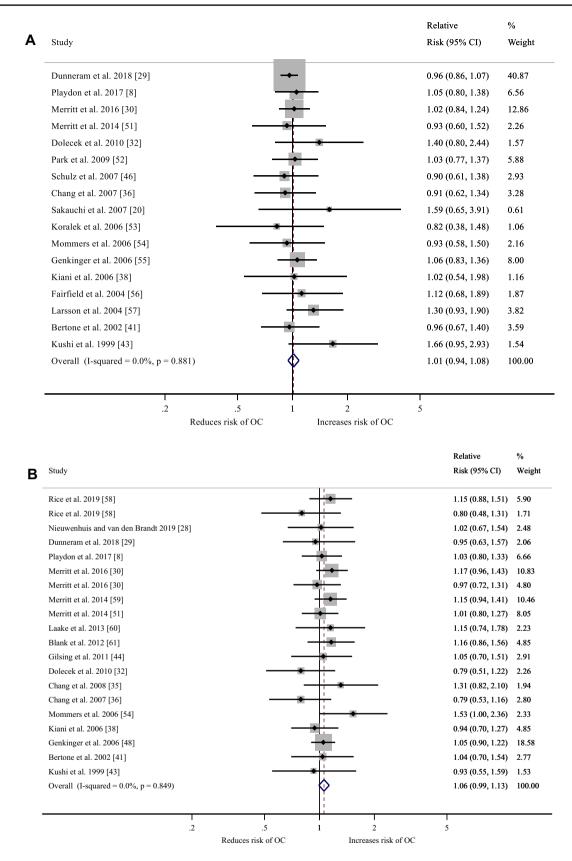


Fig. 3 Forest plot for cohort studies evaluating pooled relative risk (RR) of intake of **a** dairy products; **b** dietary fats and fatty acids and ovarian cancer risk. *OC* ovary cancer

Table 3 Relative risk (and 95% confidence interval) of ovarian cancer according to histologic subtypes and dietary intake

Types of dietary	No. of stud- ies								
		Serous tumor End		Endometroid tum	Endometroid tumor		Mucinous tumor		
		RR (95%CI)	Р	RR (95%CI)	Р	RR (95%CI)	Р	RR (95%CI)	Р
Fruits, vegetables and mush- rooms [4, 18, 30]	3	1.00 (0.96–1.05)	0.937	1.03 (0.95–1.11)	0.494	1.04 (0.94–1.16)	0.416	N/A	N/A
Meats and eggs [46]	1	1.29 (1.05–1.61)	N/A	N/A	N/A	N/A	N/A	N/A	N/A
Dairy products [41, 51, 54–57]	6	1.09 (0.95–1.25)	0.209	0.95 (0.71-1.26)	0.707	0.99 (0.65-1.50)	0.962	1.15 (0.70–1.88)	0.579
Fats and fatty acids [30, 41, 48, 54, 58, 59, 61]	7	1.06 (0.99–1.14)	0.114	1.02 (0.85–1.23)	0.826	1.00 (0.80–1.26)	0.991	1.13 (0.71–1.78)	0.611
Alcohol [62, 64, 71, 76]	4	1.06 (0.97–1.15)	0.208	0.98 (0.87-1.11)	0.773	1.07 (0.85–1.35)	0.560	0.36 (0.11-1.22)	0.102
Micro/macronutrients and plant-based bioactive com- pounds [31, 37, 55, 72, 76, 78, 79, 85]	8	1.00 (0.96–1.04)	0.834	1.08 (1.02–1.15)	0.015	1.01 (0.91–1.12)	0.923	0.89 (0.64–1.24)	0.488
Coffee and tea [98, 99]	2	0.85 (0.64–1.12)	0.255	N/A	N/A	N/A	N/A	0.79 (0.56–1.11)	0.174
Non-food contaminants [108, 112, 114–116]	5	0.97 (0.93–1.02)	0.241	1.00 (0.92–1.08)	0.960	0.95 (0.83–1.08)	0.438	1.01 (0.67–1.52)	0.971
Sweets and carbohydrate foods	0	N/A	N/A	N/A	N/A	N/A	N/A	N/A	N/A

Data marked in bold are statistically significant

risk. The relationship between alcohol and ovarian cancer varies with the total folate [70, 76] or not by the folate, methionine and pyridoxine intake [72]. Non-significant relationships were found between the histologic subtypes of ovarian cancer and alcohol consumption (Table 3) [62, 64, 71, 76]. Lower consistencies in subtype-specific relationships were reported in terms of the quantity of alcohol intake [64]. The intake of alcohol had a weak positive association with risk of serous ovarian cancer [62]. Non-significant associations were reported between menopausal status and the intake of alcohol [29]. The RR for alcohol intake was 1.12 (95% CI 0.60–2.10, P=0.723). Results from Kelemen et al. [75] demonstrate that the consumption of alcohol is inversely associated with postmenopausal ovarian cancer, and that the relationship between folate with ovarian cancer varies by the quantity of alcohol consumption.

#### Intake of micro/macronutrients and plant-based bioactive compounds and ovarian cancer risk

As Supplementary Table S19 depicts, the meta-analysis of the relationship between micro/macronutrients and plantbased bioactive compounds intake and the risk of ovarian cancer incorporated 29 cohort studies [8, 29–31, 36–38, 42, 43, 51–53, 55, 68, 70, 72, 75–87]. Micro/macronutrients and plant-based bioactive compounds intake were not significantly associated with a reduced risk of ovarian cancer (RR = 0.96, 95% CI 0.91–1.01,  $P_{\text{heterogeneity}}$  = 0.906,  $I^2$  = 0.0%, P = 0.102) (Fig. 4b). The subgroup analyses represent significant decreases in ovarian cancer risk with intake of fiber (RR = 0.89, 95% CI 0.81–0.98,  $P_{\text{heterogeneity}}$  = 0.423,  $I^2 = 2.5\%$ , P = 0.014) and flavonoids (RR = 0.83, 95% CI 0.78–0.89,  $P_{\text{heterogeneity}} = 0.178$ ,  $I^2 = 22.0\%$ , P = 0.0001) (Table 1). In contrast, intake of retinol (RR = 1.14, 95% CI 1.00–1.30,  $P_{\text{heterogeneity}} = 0.839$ ,  $I^2 = 0.0\%$ , P = 0.048) significantly increased the risk of ovarian cancer. The intake of micro/macronutrients and plant-based bioactive compounds was at a moderately decreased risk of ovarian cancer in the United States, Australia and Canada populations, but not statistically significant (Table 2). From among the 29 studies on micro/macronutrients and plant-based bioactive compound intake, 16 [29, 30, 37, 42, 51, 55, 72, 78, 80-87] non-significant relationships and a positive relationship were reported for supplemental folate and calcium (diet) intake [43]. A lower risk of ovarian cancer was reported with fiber [8], isoflavones [36], kaempferol and flavonoids [37], dairy protein [38], calcium [52, 53], riboflavin [68], folate [70, 76, 77], pyridoxine [77], flavonols and flavanones [79] consumption. Isoflavonoids, coumestrol, total fiber, cereal fiber, or vegetable fiber intake exerted a protective influence for borderline subtype of ovarian cancer, but not for invasive ovarian cancer [31]. A high intake of dietary folate can be influential in lowering the ovarian cancer risk, specifically in alcohol-consuming women [70, 75, 76]. As shown in Table 3, significant positive association was found between intake of micro/macronutrients and plant-based bioactive compounds and endometrioid ovarian tumor risk (RR = 1.08, 95% CI 1.02–1.15; P = 0.015). Non-significant associations between other histologic subtypes of ovarian cancer and consumption of micro/macronutrients and plant-based bioactive compounds were reported [31, 37, 55, 72, 76, 78, 79, 85]. The highest to the lowest quartile of isoflavonoids,

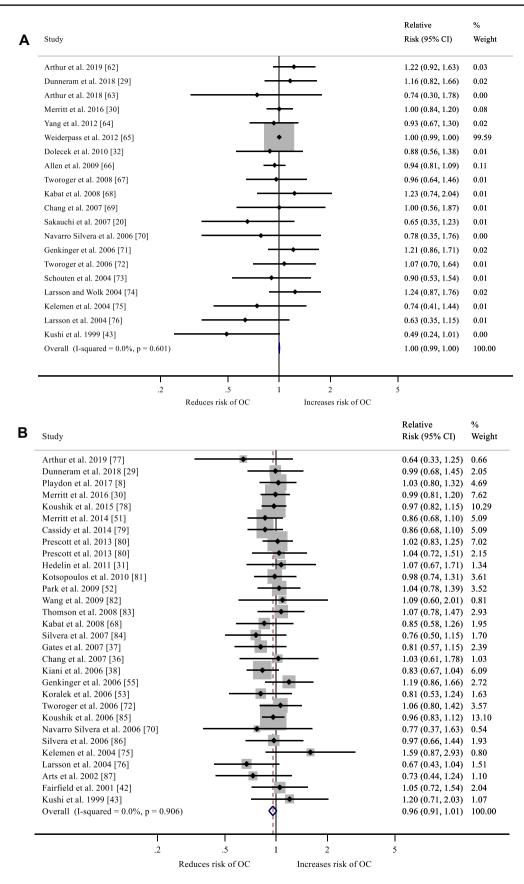


Fig. 4 Forest plot for cohort studies evaluating pooled relative risk (RR) of intake of a alcohol; b micro/macronutrients and plant-based bioactive compounds and ovarian cancer risk. OC ovary cancer

coumestrol, total fiber, cereal fiber or vegetable fiber intake differed between the two subtypes of ovarian cancer leading to a protective effect for borderline ovarian cancer but not for invasive cancer [31]. The larger dietary folate was related to a marginally reduced risk and for the serous subtype [72]. Larger intakes of vitamins were related to marginally larger risks of endometrioid tumors, but not other histological types [78]. The relationship for the intake of flavanone was stronger for serous invasive and mildly differentiated tumors as compared to non-serous and less-aggressive tumors [79]. Relationships for histologic subtypes of ovarian cancers were not significantly different for the entirety of carotenoids with the exception of lutein/zeaxanthin that was related to a statistically significant high mucinous ovarian cancer risk [85]. Non-significant associations were obtained with menopausal status and with consumption of micro/macronutrients and plant-based bioactive compounds (RR=0.88, 95% CI 0.76-1.02; P=0.093) [29, 38, 51, 80, 84].

#### Coffee and tea intake and ovarian cancer risk

Supplementary Table S20 suggests that the meta-analysis of the relationship between intake of coffee and tea and the

risk of ovarian cancer incorporates 23 cohort studies [8, 29, 30, 36, 37, 67, 88–104]. Figure 5 reported no significant association between coffee and tea intake and the reduced risk of ovarian cancer (RR = 0.89, 95% CI 0.75-1.05,  $P_{\text{heterogeneity}} = 0.0001, I^2 = 71.3\%, P = 0.167$ ). The subgroup analyses also showed that green tea consumption significantly decreased the risk of ovarian cancer (RR = 0.61, 95%CI 0.49–0.76,  $P_{\text{heterogeneity}} = 0.200$ ,  $I^2 = 31.5\%$ , P = 0.0001) (Table 1). As shown in Table 2, a moderately decrease risk was found in the European and United States populations. From among the 23 studies on the consumption of coffee and tea, 13 [8, 29, 30, 36, 88, 91, 93, 94, 98, 101–104] revealed non-significant relationships. A lower risk of ovarian cancer was reported with tea (non-herbal) [37], caffeinated coffee [67], coffee [89] and tea [92, 96, 99] consumption. Also, Gunter et al. [90], revealed a statistically significant positive relationship between coffee and ovarian cancer mortality. In 1 research investigation [97], a positive relationship was found to exist between coffee consumption and the risk of ovarian cancer. Elevated levels of green tea post-diagnosis consumption can increase ovarian cancer survival rates [100]. No significant associations between histologic subtypes of ovarian cancer and coffee and tea consumption were

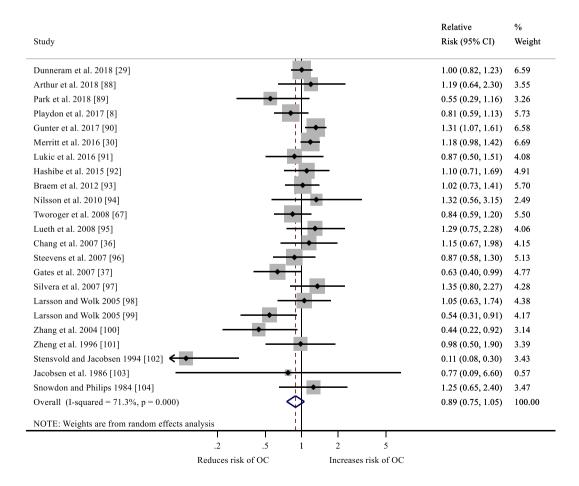


Fig. 5 Forest plot for cohort studies evaluating pooled relative risk (RR) of coffee and tea intake and ovarian cancer risk. OC ovary cancer

found (Table 3) [98, 99]. Steevens et al. [96] demonstrated that tea was inversely related with serous tumors, while it was associated with high risk of mucinous and endometrioid tumors. Menopausal status and coffee and tea consumption were not significantly related (RR = 0.99, 95% CI 0.72–1.35; P = 0.939) [29, 67]. Caffeine was found to be inversely related to postmenopausal ovarian cancer; while, it was positively related to premenopausal ovarian cancer [67]. One study [95] indicated that a coffee component other than caffeine, or in conjunction with caffeine, is related with higher risk of ovarian cancer among postmenopausal women drinking five or more than five cups of coffee a day.

### Intake of non-food contaminants and ovarian cancer risk

Supplementary Table S21 suggests that the meta-analysis of the relationship between intake of non-food contaminants and the risk of ovarian cancer incorporated 14 cohort studies [105–118]. Intake of non-food contaminants was not significantly associated with the risk of ovarian cancer (RR = 0.98, 95% CI 0.90–1.08,  $P_{\text{heterogeneity}} = 0.204$ ,  $I^2 = 23.1\%$ , P = 0.740) (Fig. 6a). The results of subgroup analysis demonstrated that intake of acrylamide, nitrate and other non-food contaminants including water disinfectants (dichloroacetic acid, trichloroacetic acid, haloacetic acid 5, haloacetic acid 6, bromochloroacetic acid, bromodicholoromethane, trihalomethanes and chloroform) and polychlorinated biphenyls were significantly associated with an increased risk of ovarian cancer (Table 1). Subgroup analysis by geographic location/country showed that no statistically significant association was found between intake of non-food contaminants and ovarian cancer risk (Table 2). Nine of the 14 studies on non-food contaminant intake, acrylamide [105, 108, 116], atrazine [106], polychlorinated biphenyls [107], cadmium [110, 111, 114], N-nitroso compounds [113] and ovarian cancer risk were not significantly related. Dietary nitrate was inversely related to the risk of ovarian cancer; while dietary nitrite from processed meats was positively related to the risk [109]. In 2 studies, a positive relationship was reported between the risk of ovarian cancer and acrylamide [117] and nitrate [118] intake. The highest intake category of animal sources of nitrite experienced a 34% increase in ovarian cancer risk [112]. Wilson et al. [115] reported a non-statistically significant indication of a higher risk of ovarian cancer among women experiencing the highest acrylamide intakes subsequent to confounder adjustment especially the intake of caffeine. As shown in Table 3, no significant associations between histologic subtypes of ovarian cancer and non-food contaminant intake were obtained [108, 112, 114-116]. Non-significant associations were obtained between menopausal status and non-food contaminant intake (RR = 1.10, 95% CI 0.69–1.78; P = 0.686) [105, 115]. The high nitrate quantities in public drinking water and private well use may raise the risk of ovarian cancer among postmenopausal women [109].

#### Intake of sweets and carbohydrate foods and ovarian cancer risk

Supplementary Table S22 shows that the meta-analysis of the relationship between intake of sweets and carbohydrate foods and risk of ovarian cancer incorporated 8 cohort studies [29, 30, 32, 35, 36, 43, 119, 120]. Intake of sweets and carbohydrate foods was not significantly associated with the risk of ovarian cancer (RR = 1.00, 95% CI 0.88-1.13,  $P_{\text{heterogeneity}} = 0.101, I^2 = 41.6\%, P = 0.966$ ) (Fig. 6b). Subgroup analyses of type of sweets and carbohydrate foods reveled no significantly associated with the risk of ovarian cancer (Table 1). In the subgroup analysis based on geographic location/country, no significant association was observed in the subgroups (Table 2). In 5 [29, 30, 32, 35, 36] out of 8 studies on sweet and carbohydrate food intake, non-significant relationships were obtained. In 2 studies, a positive and higher risk of ovarian cancer with carbohydrate intake [43] and diets with high glycemic load values [120] were reported; while in 1 study [119], sugars were found to be inversely related to ovarian cancer risk. No significant association between menopausal status and sweet and carbohydrate food consumption was reported (RR = 1.20, 95%CI 0.65-2.21; P=0.562) [29, 120].

#### Discussion

The present meta-analysis is supportive of a significant heterogeneity between dietary intake and ovarian cancer. In general, the authors observed that dietary intake exerts low effects on the risk of ovarian cancer. Some studies evaluated risk by histologic subtypes of ovarian cancer and menopausal status, in this case, the authors observed a less effect of dietary intake. The World Cancer Research Fund Continuous Update Project Report on diet and ovarian cancer risk considers the assertion that diet has a role as only limited and inconclusive as the findings to date are scare and inconsistent [121]. Nevertheless, there is evidence to suggest relationships between diet and ovarian cancer survival [57, 122–130].

#### Intake of fruits, vegetables and mushrooms and ovarian cancer risk

The results emanating from the present meta-analysis demonstrate that vegetable, fruit and mushroom intake exert low effects on reduce risk of ovarian cancer. Previous metaanalyses have documented significant relationships between

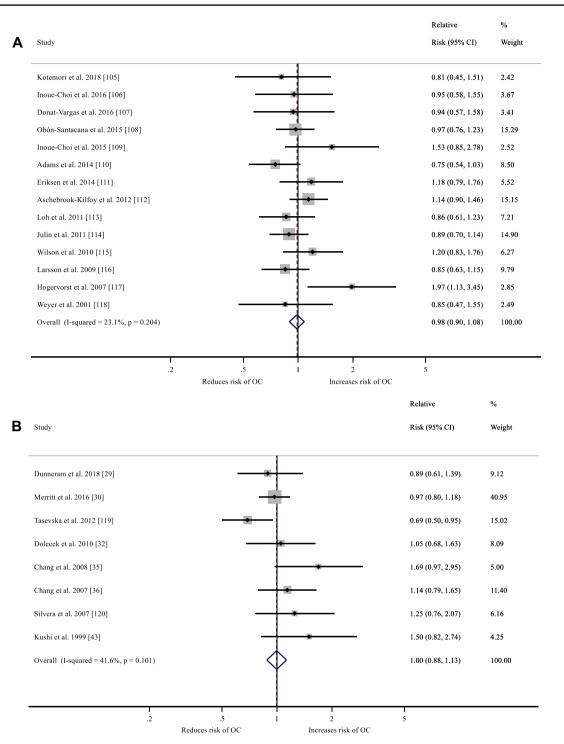


Fig. 6 Forest plot for cohort studies evaluating pooled relative risk (RR) of intake of **a** non-food contaminants; **b** sweets and carbohydrate foods and ovarian cancer risk. *OC* ovary cancer

fruit, vegetable and mushroom consumption and the lower risk of cancer in various anatomical locations [131–136].

As rich sources of nutrients and pharmaceutically bioactive compounds, vegetables, fruits, and mushrooms possess anti-carcinogenic properties [32, 34, 39, 134, 135] having antioxidant and anti-proliferative activities, are able to modulate steroid hormone concentrations and metabolism, enzymatic carcinogen detoxification, and to maintain intracellular matrix integrity, to stimulate the immune system and to inhibit metastasis and synthesis and methylation of DNA [32, 34, 39, 135]. The result, which is in agreement with the present systematic review, suggests that vegetable consumption rather than fruit consumption may be more beneficial [24]. Considering the intake of vegetables, previous systemic reviews and meta-analyses of cohort studies documented significant inverse relationships with ovarian cancer [24, 25, 124]. These findings were consistent with the results in specific epidemical studies [4, 8, 24, 25, 38-40, 43], which reported inverse associations between intake of green leafy green leafy and allium vegetables and ovarian cancer risk. Consistent with the specific epidemical studies [20, 24, 25, 32, 37–40], and an earlier meta-analysis [124], we observed a moderately reduce associations of consumption of red/yellow vegetables, broccoli, fruiting vegetables and soybean and risk of ovarian cancer. In 1 study, negative relationships were reported between the ovarian cancer risk and high vegetable intake (Chinese cabbage). The reason may be that the Japanese often consume pickled Chinese cabbage. It is worth mentioning that processing vegetables lowers the levels of micronutrients, e.g., antioxidants and vitamins [20]. Pickled vegetables make it possible for fermentation and growth of fungi and yeasts to take place and do possess the potential to produce carcinogenic N-nitroso compounds and mycotoxins [137, 138].

The anti-carcinogenic bioactivity of constitutive compounds within vegetables and fruits such as glucosinolates, polyphenols, flavonoids (diosmin, hesperidin, flavonols, flavones, anthocyanidins, catechins, flavanones, isoflavones and isoflavonoids), carotenoids and other bioactive compounds, cause apoptosis and hinder cell growth through pleiotropic mechanisms containing the capability to support adjust metabolic reprogramming and anti-inflammatory attributes [82, 139–143]. Phytoestrogens has both estrogenic and antiestrogenic effects and multiple actions within cancer cells [122, 144, 145]. Phytoestrogens inhibit tumor angiogenesis, tyrosine kinase and topoisomerase II. Other signal pathways are mediated by receptors such as GnRH receptor, FSH or LH receptors and GFR which adjust hormone concentrations and the related gene and protein expressions, e.g., protein kinase B (Akt), Raf, caspase3, NF-KB, Bcl-2, inhibit apoptosis, metastasis and cell proliferation of ovarian cancer cells [122, 123]. Organosulfur compounds, substances in allium vegetables, are considered to be associated with cancer protection through modulation of metabolizing enzymes and anti-proliferative activity thus inhibiting cell proliferation and tumor growth [4]. Isothiocyanates which found among the cruciferous vegetables including broccoli are capable of inhibiting metabolic activation and enforcing detoxification of carcinogens, altering apoptosis, protecting against oxidative damage, and producing antiestrogenic effects, probably in synergy with isoflavones [36, 140]. Antagonistic activity of isothiocyanates against cancer is associated with direct or indirect interaction with Nrf2 and NF-kB protein, phosphorylation of EGFR at Tyr1068 and lower phosphorylation of Akt [140, 146, 147].

Previous meta-analyses have reported on the protective role played by dietary carotenoids and ovarian cancer risk [128, 148]. Antioxidants such as carotenoids, vitamin E and vitamin C are considered to be influential in the control of cellular differentiation or proliferation and to affect reactive oxygen species, to initiate lipid peroxidation, inhibitory effect on insulin-like growth factor-1 (IGF-1)-stimulated cell multiplication, cell damage, to disrupt cell signaling and to express all salient events in carcinogenesis [5, 8, 29, 32, 78, 149].

Intake of mushroom was associated with a higher risk of ovarian cancer [29] whose result was contrary to other studies [4, 150]. Medicinal mushrooms can be considered as a source of ergothioneine, selenium, fiber, and several other vitamins and minerals, having anti-tumor and anticarcinogenic effects by inhibiting NF- $\kappa$ B signaling pathway components [151–153]. The salient safety concerns for mushrooms include the toxicity and carcinogenicity of agaritine and its derivatives and their possible contamination with toxic metals [154].

#### Intake of meats and eggs and ovarian cancer risk

The results emanating from the present meta-analysis demonstrate that meat and egg consumption is less influential on increase risk ovarian cancer. The results from systematic review and meta-analyses suggest that red and processed meat consumption is not related to the ovarian cancer risk [25, 155]. There is other meta-analytical research to indicate that low intake of processed meat and high consumption of poultry and fish can lower the ovarian cancer risk [24, 156]. The results from the present study could provide moderate non-significant support to the idea that a low ovarian cancer risk was associated with high fish consumption. The metaanalysis demonstrated that total fish consumption was not significantly related to ovarian cancer risk [157]. Mechanisms recommended for a relationship between red and processed meat and ovarian cancer include high fat intake, as shown to be associated with high levels of circulating estrogens or mediated through obesity, leading to high biologically active estrogen levels [155]. Preservation, cooking or processing methods may lead to the introduction of mutagens and carcinogens to meat, e.g., N-nitroso compounds, heterocyclic amines and polycyclic aromatic hydrocarbons. Heme iron, more commonly existing in red meat than white meat, is also capable of stimulating endogenous N-nitroso production [24, 155]. Heat-resistant infectious agents can get involved in synergic interactions with the chemical carcinogens [158]. A mechanism for the protective effect of fish intake is the rich source of omega-3 fatty acids and their anti-apoptotic and anti-inflammatory properties [8, 157]. The results from Tavani et al.'s [159] meta-analysis demonstrated that there was a significantly inverse relationship

between intake of omega-3 polyunsaturated fatty acid and risk of ovarian cancer. One cohort study [41] and meta-analyses suggest that dietary intake of total omega-3 fatty acids is not significantly related with the ovarian cancer risk [160, 161]. Omega-3 fatty acid anti-proliferative and anti-carcinogenic exerts an effect on epithelial ovarian cancer cell lines [161]. Carcinogenic and mutagenic *N*-nitroso compounds and heterocyclic amines in processed fish can raise the risk of cancer [157, 161].

Results from epidemiologic cohort studies investigating the relationship between egg intake and ovarian cancer risk are inconclusive. Meta-analyses of observational investigations suggested that dietary egg intake is capable of raising the risk of ovarian cancer [162, 163]. The biological mechanism through which eggs exert a harmful effect on cancer risk probably involves the large cholesterol content and large quantities of protein per energy content in egg [163, 164]. Chlorine, used to wash eggs, interacts with the organic substances within the eggs turning into potentially carcinogenic organochlorines which interrupt estrogen-related pathways [24, 163]. Furthermore, eggs can be considered as a source of heterocyclic amines formed during high-temperature cooking [165].

#### Intake of dairy products and ovarian cancer risk

In Jeyaraman et al.'s [166] overview of reviews, twenty-six meta-analyses showed no statistically significant associations between all-dairy products, whole milk, milk, low-fat/ skim milk, yogurt, cheese, hard cheese, cottage cheese, butter, ice cream or lactose, intake and risk of ovarian cancer. Three meta-analyses showed an increased risk of ovarian cancer with higher consumption of whole milk or lactose exposure. Lu et al.'s [167] meta-analysis demonstrates that total dairy products intake exerts no significant effects on raising the mortality rates of all cancers, while low total dairy intake even lowered the relative risk. Another metaanalysis suggested that high dairy product/constituent intakes, e.g., lactose, can raise the ovarian cancer risk [168]. The authors' results indicate that consumption of dairy products exerts no statistically significant associations with the ovarian cancer. Inconsistent results from this meta-analysis showed that lactose intake was related to both increased and decreased risk of ovarian cancer [38, 43, 51, 55]. Galactose produced by lactose has been shown to raise the ovarian cancer risk by direct toxicity to the ovarian germ cells. The high dairy product intake is also likely to increase the ovarian cancer risk within the population through genetic or biochemical properties of galactose. The intense impairment of the galactose-1-phosphate uridyltransferase (GALT) gene leads to accumulation of galactose and other metabolites within the body, e.g., the ovary [54–57, 162, 169–171]. These findings indicate that the consumption of high-fat dairy products and total cheese exert moderate but not statistically significantly increased association with the ovarian cancer. The estrogens of cow milk, dairy contaminants, processing or storage of dairy products and dairy additives can raise the risk of cancer [170, 171].

# Intake of dietary fats and fatty acids and ovarian cancer risk

Some meta-analytical researches suggest that high dietary fat intake represents a significant risk factor for the development of ovarian cancer [172, 173]. Hou et al.'s [174] metaanalysis reveals non-significant associations between dietary fat and fatty acid intakes and ovarian cancer. In the present review, intake of dietary fats and fatty acids was found to be not statistically significantly associated with an increased risk of ovarian cancer. The subgroup findings indicated that the intake of total fat, saturated fat and fatty acid and cholesterol exerts significantly increased association with the ovarian cancer. Dietary fats were hypothesized to affect ovarian carcinogenesis mainly through hormone-related mechanisms via mitogenic effects on ER $\alpha$ -positive or -negative tumor cells. Obese females suffer from insulin resistance, and concurrent hyperinsulinemia with excess IGF-1 receptor may lead to androgen steroidogenesis and development of tumors [8, 30, 160]. Fatty acid metabolism plays a significant role in ovarian cancer tumorigenesis [11]. Cholesterol can exert an effect on the ovarian cancer risk through high circulating estrogen or progesterone [175].

#### Alcohol intake and ovarian cancer risk

The results indicate that the intake of alcohol exerts no significant association with the increased risk of ovarian cancer. Previous meta-analytical research demonstrates that the intake of alcohol is not related to a high ovarian cancer risk [175–181]. Our findings demonstrated that the intake of wines and spirits was at a moderately increased risk of ovarian cancer. Mechanisms of alcohol-related ovarian carcinogenesis include high cumulative estrogen exposure, alteration of gonadotropin levels, promotion of DNA damage, impaired folate metabolism, oxidative stress, acetaldehyde, DNA hypomethylation, inhibition of carcinogen detoxification or clearance, and increased metastatic potential of tumor cells. Antioxidants, e.g., polyphenols and resveratrol contained in red wine, were suggested to justify the inverse relationship between red wine and the ovarian cancer risk [177, 178, 182]. Conversely, alcohol is documented to exert a protective potential against ovarian carcinogenesis through lowering follicle stimulating hormone, luteinizing hormone and gonadotropin levels [182]. Phytochemicals in red wine have varied effects, e.g., cancer preventive and pro-estrogenic activity and genotoxicity [69].

#### Intake of micro/macronutrients and plant-based bioactive compounds and ovarian cancer risk

The results emanating from the present meta-analysis indicate that micro/macronutrients and plant-based bioactive compounds exert lower effects on reduce risk of ovarian cancer. Consistent with the specific epidemical studies [8, 36, 37, 79], and prior meta-analyses [24, 126, 127, 179], we observed a significantly reduced association of consumption of fiber and flavonoids, and risk of ovarian cancer. The increasing dietary fiber intake can lower the bioavailability of steroid hormones through influencing bacterial macroflora, increasing sex hormone-binding globulin causing smaller circulating levels of unbound or biologically available estrogen, lessening circulating estrogens through inhibition of bile reabsorption, increasing faecal excretion, and increasing the protection of phytoestrogens, which can preclude the development of ovarian cancer [8, 24, 84, 126, 127, 183]. Dietary fiber can exert an effect on inflammation lowering glycemic load and improving insulin sensitivity, favorably regulating IGF-1 [8, 127, 183]. Numerous mechanisms can inhibit ovarian carcinogenesis through flavonoids, e.g., free radical scavenging or enhancing the body's antioxidant systems, such as through upregulating activity of glutathione S-transferase and other detoxifying enzymes, thus enhancing carcinogen clearance. Flavonoids have also been mentioned to prevent the incorporation of some metabolic precursors, which shows that they could be preventative in DNA, RNA, and/or protein synthesis, and they have been mentioned to incorporate into DNA, according to their similarities with nucleosides structures [184]. Certain flavonoids, e.g., quercetin, luteolin and apigenin, are likely to decrease inflammation through their inhibitory effect on the enzymes cyclooxygenase-2 and inducible nitric oxide synthase which are both salient mediators of inflammatory reactions. Quercetin has the capability to prevent the aromatase activity and topoisomerase I-catalyzed DNA religation [185]. Flavonoids can modulate sex steroid hormone levels through both estrogenic and anti-estrogenic effects. Flavonoids may inhibit proliferation in human cancer cell lines and angiogenesis through suppression of expression of VEGF and triggering the non-apoptotic cell death, interrupting cell signaling and cell cycle and inducing apoptosis [31, 36, 37, 79, 82, 186, 187]. Kaempferol can lower the level of proliferation and significantly decrease the expression of VEGF through suppression of extracellular signal-regulated kinase (ERK)-NFkB-cMyc-p21-VEGF pathway in cancer cells [141, 188]. Taking a molecular viewpoint, flavones and isoflavone can be considered as competitive inhibitors of cytochrome *P*<sub>450</sub> [189].

Although epidemiological studies have suggested an inverse correlation between cancer development and dietary consumption of vitamin A, our results showed that retinol consumption was positively related to the ovarian cancer risk. Retinol and vitamin A derivatives influence cell differentiation, proliferation, and apoptosis and play an important physiologic role in a wide range of biological processes [190].

These findings were consistent with the results in specific epidemical studies [30, 68, 77], which reported significantly/ modestly decreased risk associations between intake of thiamin and riboflavin and ovarian cancer. Thiamin may also play a role in carcinogenesis, by their independent roles in DNA synthesis. Additionally, riboflavin can be influential in carcinogenesis as cofactors in folate metabolism [68].

The results from a meta-analysis [191] suggest that large levels of dietary protein consumption are not related to the ovarian cancer risk. Although our results were not statistically significant, we found the lower decrease of diatery proteins with the ovarian cancer risk. A number of dairy product constituents, e.g., proteins and calcium are likely to be responsible for the protective relationships with cancers [38, 52, 53]. For instance, lactoferrin is known for its inhibitory action on cell proliferation and its anti-inflammatory, antioxidant abilities and protection against cancer development and metastasis [192]. Casein, e.g., whey proteins, is capable of enhancing cellular levels of glutathione, an antioxidant. Furthermore, whey proteins can lower cancer risk through elevating hormonal and cell-mediated immune responses [193]. Our findings demonstrated that the intake of calcium and total vitamin D was not significantly related with the ovarian cancer risk. The inadequate calcium intake and low vitamin D status can be considered as salient risk factors in numerous types of cancer [194, 195]. The meta-analysis suggested that higher calcium intake may be inversely related to the ovarian cancer risk [125]. The high calcium intake is likely to raise or lower the ovarian cancer risk. High calcium intakes can depress 1,25-(OH)<sub>2</sub>D, leading to an enhancement of cellular proliferation and, hence, tumorigenesis [55]. Tissue-specific expression of the CYP27B1-encoded 25-hydroxyvitamin D-1α-hydroxylase and of the extracellular calcium-sensing receptor (CaR) locally produced  $1,25(OH)_2D_3$  and extracellular Ca<sup>2+</sup> act jointly as main regulators of cellular proliferation, differentiation and function. Therefore, impairment of antimitogenic, proapoptotic and pro-differentiating signaling from the 1,25(OH)<sub>2</sub>D<sub>3</sub>activated vitamin D receptor (VDR) and from the CaR in vitamin D and calcium insufficiency has been suggested in the pathogenesis of the above-mentioned types of cancer [194, 195]. Furthermore, vitamin D has a non-genomic impact, i.e., regulation of calcium and phosphate homeostasis pathways, activating protein kinase C, protein kinase A, PI3K and phospholipase C. The VDR variant is also likely to be involved in ovarian cancer carcinogenesis which expands VDR results in lower NF-kB transcriptional activation, causing lower IL-12 expression and a weaker immune response [195]. Higher calcium levels are characterized by down-regulation of circulating parathyroid hormone, which can lower hepatic and osteoblastic synthesis of IGF-1 [125, 194, 196]. Vitamin D plays a potential role in cell cycle and apoptosis, angiogenesis, invasion and metastasis, inflammatory response and tumor metabolism [190]. No conclusive or strong evidence exists to support the claim made in a large number of review articles that vitamin D exposures lower the ovarian cancer risk occurrence or mortality [7, 197, 198].

#### Coffee and tea intake and ovarian cancer risk

The results emanating from the present meta-analysis demonstrate that coffee and tea exert less effect on reduced risk of ovarian cancer. Our results demonstrated a significant relationship between green tea consumption and a reduced risk of ovarian cancer. On the other hand, modestly lower risk of ovarian cancer was found regarding total tea intake. Prior meta-analytical research revealed an inverse relationship between tea and green tea consumption and the risk of ovarian cancer [96, 129]. Furthermore, non-significant relationships were obtained between tea and green tea consumption and the risk of ovarian cancer in other meta-analyses [93, 130, 199-201]. Tea polyphenols, e.g., catechins, flavanols, theaflavins and thearubigins, are beneficial components of tea which may down-regulate the expression of a variety of tumor genes, causing tumor cell apoptosis, blocking tumor cell cycle, up-regulating the body metabolism, eradicating excess free radicals, influencing the prevention and inhibition of tumors. Some polyphenols may cause tumor cell apoptosis inhibiting tumor angiogenesis [93, 129]. Catechins and gallocatechins, having low quantities of methylxanthines detected mainly in green tea, and theaflavin digallate, a major component of black tea, are the two most effective anti-cancer factors in tea. Signaling proteins affected by epigallocatechin gallate in ovarian cancer, i.e., JUN, FADD, NFKB1, Bcl-2, HIF1a, and MMP, influence the cell cycle, cellular assembly and organization, and DNA replication. The effect of tea varies for the difference in genetic heterogeneity and lifestyle among populations. Moreover, the numerous ingredients in tea may have varying anti-cancer activities and effects on various types and subtypes of ovarian cancer [93, 129, 202, 203].

Consistent with the results emanating from other metaanalytical research studying [93, 96, 130, 204–207], we suggest no relationship between coffee intake and ovarian cancer. Furthermore, our findings indicated a lower risk of ovarian cancer with intake of decaffeinated coffee and lesser with caffeine. In contrast, a lower positive risk of ovarian cancer was observed with intake of soft drinks. An inverse association between coffee consumption and cancer risks can be mediated by various mechanisms, e.g., hormonal homeostasis, reduction of oxidative stress and DNA damages, detoxification of carcinogens, inhibition of carcinogenesis, and induction of apoptosis. Coffee holds many bioactive components such as polyphenols, caffeine, diterpenes, melanoidins, which lower oxidative stress, exert anti-cancerogenic properties, through triggering defense mechanisms, carcinogenic detoxification, and activation or suppression of onco-suppressors and proto-oncogenes, respectively [204, 205]. A number of studies in the authors' results assert that coffee consumption can be related to higher ovarian cancer risks [90, 95, 97]. Conversely, coffee also holds acrylamide and caffeine, which exert potential hormonal and carcinogenic effects [93].

# Intake of non-food contaminants and ovarian cancer risk

Non-food contaminants exert less effect on increase risk ovarian cancer. These findings were consistent with the results in specific epidemical studies [109, 112, 117, 118], which reported significantly/modestly increased risk associations between intake of acrylamide and nitrate, nitrite and ovarian cancer. The systematic review and meta-analyses of epidemiological investigations suggest that high acrylamide intake levels are strongly related to the ovarian cancer risk [208]. Other extensive systematic review and meta-analysis revealed no increases in the risk of most types of cancer as related to acrylamide exposure [209]. Foods, as sources of acrylamide, hold various nutrients and most foods are highenergy sources, causing obesity and increasing the risk of numerous cancers, and exposure to acrylamide [210]. The major pathway to acrylamide carcinogenesis is through its oxidization to glycidamide, by the Cyp2e1 enzyme system. Acrylamide can play a carcinogenic role in the selected body sites through affecting hormonal balances in addition to their detrimental and mutagenic effects on DNA [105, 108, 117, 208, 210]. The literature on dietary nitrate or nitrite and cancer risk has been expanding but the results have been inconclusive. The meta-analysis of epidemiological investigations revealed that non-significant relationships exist between dietary nitrate or nitrite and ovarian cancer [211]. Nitrate and nitrite can be considered as precursors in the endogenous formation of N-nitroso compounds, potential human carcinogens. Nitrosamides directly alkylate DNA causing tumors in numerous organs, while nitrosamines need to be activated by specific cytochrome  $P_{450}$  enzymes to be carcinogenic. Certain nutrients are likely to influence endogenous N-nitroso compounds formation in the stomach. Conversely, heme iron has been demonstrated to promote total N-nitroso compounds formation. Nevertheless, epidemiologic evidence for these interactions on cancer risk is yet to be evolved. Antioxidants, e.g., vitamins C and E, reveal *N*-nitroso compounds formation through converting nitrite to nitric oxides lowering the level of *N*-nitroso compoundsinduced DNA adducts [109].

These findings of subgroup analysis showed that intake of dichloroacetic acid, trichloroacetic acid, haloacetic acid 5. haloacetic acid 6. bromochloroacetic acid. bromodicholoromethane, trihalomethanes, chloroform and polychlorinated biphenyls was significantly associated with an increased risk of ovarian cancer. Some studies revealed non-significant relationships between atrazine, polychlorinated biphenyls, cadmium and N-nitroso compound intake and the risk of ovarian cancer [106, 107, 110, 111, 113, 114]. The mechanism of polychlorinated biphenyls works through disruption of hormone-dependent pathways [107]. Generally, the available epidemiological research does not furnish conclusive and scientifically compelling evidence of causality between exposure to atrazine or triazine herbicides and cancer in humans [212]. Atrazine exposure mainly affected antioxidant defenses and, to a lesser degree, lipid, protein, carbohydrates and nucleic acids oxidation processes by cytochrome  $P_{450}$  enzymes, generation of redox-active metabolites and impairment of the electron transport cascades in mitochondria [213, 214]. Numerous mechanisms potentially associate cadmium with cancer, e.g., aberrant gene expression, oxidative stress and inflammation interference with DNA, alterations of DNA methylation and enhanced proliferation, depressed apoptosis or epigenetic alterations [110, 111, 114]. Mechanistic and epidemiologic evidence indicates that estrogen-mimicking contaminants, e.g., the environmental and dietary pollutant cadmium, are likely to contribute to the development of ovarian cancer [111, 114].

#### Intake of sweets and carbohydrate foods and ovarian cancer risk

The authors observed no significant effect of sweet and carbohydrate food consumption on the ovarian cancer risk. We observed a moderately increased association of consumption of sweets and risk of ovarian cancer. Interestingly, our subgroup analysis consistent with the results in specific epidemical study [119] showed a lower decrease association of consumption of carbohydrates and risk of ovarian cancer. While our results contrast with some studies suggesting that a diet high in carbohydrates increases ovarian cancer risk [43] and ovarian cancer risk null relationships [120], showing that most individual foods and nutrients are not associated with ovarian cancer risk. The positive relationship with high levels of carbohydrate, e.g., sugar intake, may be related to their elevated glycemic index, which has been associated with hyperinsulinemia and lower IGF binding protein concentrations, thus enhancing IGF-1 levels. Increased IGF-1 bioactivity hinders apoptosis, stimulates cell proliferation and sex steroid synthesis and hampers sex-hormone binding globulin synthesis, whose phenomenon can be suggestive of the development of ovarian cancer. Furthermore, the results suggest that the acute glucose fluctuations evoke oxidative stress, accompanied with subsequent oxidative DNA damage, suggestive of cancer development [119, 120, 215–217].

We analyzed the association between dietary intake and the risk of ovarian cancer subtypes and menopausal status. We found that endometrioid ovarian cancer incidence was more susceptible to intake of micro/macronutrients and plant-based bioactive compounds. Furthermore, we observed no statistically significant association between menopausal status and dietary intake and the risk of ovarian cancer menopausal status.

The main limitation of our systematic review was the lack of consistency in risk associations between dietary intake and ovarian cancer. Additionally, data evolving for histologic subtypes of ovarian cancer and menopausal status were limited.

#### Conclusions

The conclusions emanating from the present review reveal the potential effects of dietary intake on ovarian cancer risk. It is likely that the few observed relationships have to do with the multifarious effects of components in dietary intake. Some findings are suggesting the potential for a lower ovarian cancer risk among women consuming green leafy green leafy, allium vegetables, fiber, flavonoids, green tea and possibly vegetables, red/yellow vegetables, broccoli, fruiting vegetables, soybean, fish, low fat dairy products, thiamin, riboflavin, carotenoids and tea. Conversely, for women consuming total fat, saturated fat and fatty acid, cholesterol, retinol, acrylamide and nitrate, dichloroacetic acid, trichloroacetic acid, haloacetic acid 5, haloacetic acid 6, bromochloroacetic acid, bromodicholoromethane, trihalomethanes, chloroform, polychlorinated biphenyls and probably red meat, pork, eggs, dried and salted fish, canned tuna, high-fat dairy products, total cheese, animal fat, dairy fat, wines, spirits, vitamin E and nitrite the risk may increase. In this regard, a global project using both quantitative and qualitative dietary analysis and varied information is needed. This global project must, of necessity, include the specific dietary, composition of food ingredients and human genetic population to assess the role of dietary intake in the risk of ovarian cancer.

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Author contributions All authors have made substantial contributions to the conception and design of the study. AFAR, AK and FA conducted the search and data extraction; AK and FA collected the data; AK and FA analyzed the data and wrote the manuscript. All authors read and approved the final content.

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#### **Compliance with ethical standards**

**Conflict of interest** The authors declare that they have no conflict of interest.

Ethical approval Ethical approval was not obtained as the data emanated from previously published studies.

**Informed consent** Informed consent of patients was not obtained as the data emanated from previously published studies.

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