



Preoperative short-course radiotherapy (5 × 5 Gy) with delayed surgery versus preoperative long-course radiotherapy for locally resectable rectal cancer: a meta-analysis

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Accepted: 17 October 2019 / Published online: 19 November 2019
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Abstract

Purpose Preoperative short-course radiotherapy (PSRT) and preoperative long-course radiotherapy (PLRT) are standard treatment regimens for locally advanced rectal cancer. However, whether the efficacy and safety of PSRT with delayed surgery (more than 4 weeks) are superior to those of PLRT remains unresolved and was explored in this meta-analysis.

Methods Studies published in PubMed, Embase, the Cochrane Library, and ClinicalTrials.gov databases were systematically searched. RevMan 5.3 was used to calculate pooled hazard ratios (HR) and relative risk (RR).

Results Seven studies including 4973 patients were identified in the meta-analysis. Pooled statistics showed that there was no statistically significant difference in overall survival (HR = 1.30, 95% CI 0.58–2.89, $P = 0.52$) or disease-free survival (HR = 1.10, 95% CI 0.73–1.66, $P = 0.64$) between the preoperative short-course and long-course radiotherapy groups. Moreover, pathological complete remission, early postoperative complications, treatment-related grade 3/4 toxicity, local recurrence, and distant metastasis were similar between the two groups. Interestingly, a subgroup analysis revealed that preoperative short-course radiotherapy without adjuvant chemotherapy not only resulted in lower treatment-related grade 3/4 toxicity than the long-course radiotherapy group (RR = 0.19, 95% CI 0.08–0.48, $P < 0.01$) but also resulted in significantly lower overall survival and pathological complete remission ($P = 0.02$, $P < 0.01$, respectively). Disappointingly, pooled statistics observed few advantages over long-course radiotherapy in short-course radiotherapy with the adjuvant chemotherapy subgroup.

Conclusions PSRT with delayed surgery was as effective as PLRT for the management of locally resectable rectal cancer. However, not adding additional chemotherapy to PSRT not only significantly decreased grade 3/4 toxicity but also decreased pathological complete remission and overall survival.

Trial registration The protocol for this meta-analysis was prospectively registered with PROSPERO (CRD42019133641)

Keywords Rectal cancer · Preoperative short-course radiotherapy · Preoperative long-course radiotherapy · Adjuvant chemotherapy · Meta-analysis

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Introduction

According to global cancer statistics in 2018 [1], Asia (followed by Europe and USA) had the highest number of cancer cases and highest cancer-related death rate worldwide (both sexes). Colorectal cancer is the third leading cause of new cases, accounting for 10.2% of the world, and is the second leading cause of cancer-related death following lung cancer. With the rapid development of radiation technology, preoperative radiotherapy, which has been shown to increase local control without improving the survival rate, is currently recognized as the standard treatment modality for locally advanced rectal cancer [2–6].

Although both the latest National Comprehensive Cancer Network guidelines version 3.2018 (NCCN 2018 V3) [7] and the European Society for Medical Oncology (ESMO) guidelines [8] recommend preoperative short-course radiotherapy (PSRT) as one of the standard treatments for locally advanced rectal cancer, neither explicitly provides an optimal interval between the end of short-course radiotherapy and surgery. Additionally comparisons of the efficiency and safety between PSRT with delayed surgery and preoperative long-course radiotherapy (PLRT) have provided controversial results. To be consistent with previous studies [9, 10], in our meta-analysis, we defined the time interval for delayed radiotherapy as at least 4 weeks from the completion of radiation. Randomized controlled trials (RCTs) have provided conflicting results regarding the benefit of PSRT with delayed surgery. Bujko et al. [11] and Markovina et al. [12] reported that patients undergoing PSRT after which surgery was delayed for more than 4 weeks had better survival than those undergoing PLRT concomitant with chemotherapy, while Kairevičė et al. [13] and Latka et al. [14] came to the opposite conclusion. The results regarding treatment-related grade 3/4 toxicity, postoperative complications (PCs), downstaging, and pathological complete remission (pCR) have also been inconsistent [11, 12, 15–18]. Therefore, we performed this meta-analysis to explore the efficiency and safety of PSRT with delayed surgery versus preoperative conventional radiotherapy for locally resectable rectal cancer.

Materials and methods

Search strategy

Eligible studies published in PubMed, Embase, the Cochrane Library, and ClinicalTrials.gov were systematically searched without limits for language. The search date ranged from inception to May 1, 2019. The search strategy included the following terms: “rectal neoplasms/radiotherapy,” “rectal neoplasms/surgery,” “radiotherapy, adjuvant,” “conventional chemoradiotherapy,” “long-course radiotherapy,” “short-

course radiotherapy,” and “preopera*” and combinations of these terms constructed using the Boolean operators “AND” and “OR”. Relevant articles were also extracted manually from the references of retrieved publications. Two reviewers (Qiaoli Wang and Yongping Huang) independently searched the aforementioned databases for eligible articles.

Inclusion criteria

Our inclusion criteria were designed according to Participants, Intervention, Comparison and Outcomes, Study design (PICOS) principles, as follows: (1) P: patients diagnosed with rectal cancer by biopsy; (2) I: experimental group: PSRT (5 × 5Gy) followed by delayed surgery at least 4 weeks later, ignoring whether adjuvant or consolidation chemotherapy was given; control group: PLRT (total dose at least 45 Gy at 1.8–2 Gy/f) followed by surgery at least 4 weeks later regardless of whether the adjuvant or consolidation chemotherapy was given; (3) C: comparison of the efficacy and side effects of the experimental and the control group treatment; (4) O: overall survival (OS), disease-free-survival (DFS), treatment-related grade 3/4 toxicity, early postoperative complications (PCs; defined as reversible complications related to surgery, such as anastomotic leakage, wound infection, intra-abdominal infection, ileus, and other complications requiring re-operation), local recurrence (LR; defined as an intrapelvic recurrence following primary rectal cancer resection, with or without distal metastasis), distant metastasis (DM; documented in any organ) and pathological complete rate (pCR); and studies reporting one of these outcomes were included; and (5) S: randomized controlled trials (RCTs) and non-RCTs.

Exclusion criteria

The exclusion criteria were as follows: (1) patients with clinical stage IV or synchronous distant metastases; (2) the preoperative radiation dose did not meet the inclusion criteria; (3) the time interval to surgery was less than 4 weeks in either the PSRT or PLRT arm; (4) single arms, reviews, case reports, letters, comments, etc.; and (5) studies lacking relevant statistics.

Data extractions

Data were independently extracted by two reviewers (Wang Qiaoli and Huang Yongping) and recorded in a standard extraction form. The relevant information extracted from the selected studies is listed as follows: (1) baseline characteristics: name of the first author, year of publication, resources and study year, follow-up time, number of patients, study type, clinical stage of the tumor, tumor location, interventions, comparison, and ages; (2) outcomes: OS, DFS, treatment-related grade 3/4 toxicity, PCs, LR, DM, and pCR.

Disagreements between the two authors were resolved by a group discussion.

Data analysis

Where a meta-analysis was judged to be appropriate, pooled statistics analyses were performed using Review Manager software 5.3 (RevMan 5.3) supplied by the Cochrane Collaboration. Dichotomous variables were calculated as the relative risk (RR). Hazard ratios (HR) were selected as the effect indicator to pool survival statistics based on the method published by Tierney et al. [19]. Additionally, 95% confidence intervals (CI) were calculated for count data. If the HR could not be extracted directly from the eligible articles, Engauge Digitizer software was used to extract relevant information. We used the Mantel-Haenszel method to estimate the typical RR and risk difference. χ^2 and I^2 test statistics were used to investigate heterogeneity. If the P value of the χ^2 test was < 0.10 and the I^2 value was $> 50\%$, the random effects model was used (significant heterogeneity). In contrast, the fixed effects model was used in cases with no significant heterogeneity. A sensitivity analysis was performed to confirm the robustness of the results by omitting one study at a time.

The subgroup analysis was based on chemotherapy in the PSRT group to reduce some heterogeneity of the treatment effect. The two subgroups were classified as follows: one, the subgroup was the PSRT plus pre/postoperative adjuvant chemotherapy subgroup, and two, the PSRT without adjuvant chemotherapy subgroup. Thus, the effect of chemotherapy in the PSRT group could be clearly observed.

The protocol for this meta-analysis was prospectively registered with PROSPERO (CRD42019133641).

Quality assessment and risk of bias

The risk of bias was assessed by two reviewers (Wang and Huang), and discrepancies were resolved by a group discussion. The quality of the RCTs was evaluated by the Cochrane risk of bias assessment tool, which includes seven items: random sequence generation, allocation concealment, blinding of participants and personnel, blinding of outcome assessment, incomplete outcome data, selective reporting, and other bias. Each item was assessed as having a high, low, or unclear risk of bias [20]. The quality of the non-RCTs was evaluated by the Newcastle-Ottawa Scale (NOS), which contains three perspectives: selection, comparability, and the outcomes of studies. We identified each “high” quality item with a “star.” “Comparability” had a maximum of two “stars,” and the rest of the items had a maximum of one “star” each [21]. Studies with six or more stars were considered high quality [10].

Results

Study selection

A total of 1278 records were selected by searching PubMed, Embase, the Cochrane Library, and ClinicalTrials.gov databases, and 268 duplicates were eliminated. After the titles and abstracts were screened, 983 irrelevant records were excluded. Further exclusion was achieved by scanning the full texts of the remaining 27 articles. Then, 20 articles were excluded, including 2 meta-analyses, 2 reviews, 4 articles with undesirable fraction doses or tumor clinical stage, 2 articles with incomplete and duplicated data, and 10 articles with unsatisfactory interval times between radiation and surgery. Eventually, we included 7 appropriate studies [11, 13–15, 18, 22, 23] involving a total of 4973 patients. The flow diagram is shown in Fig. 1.

Characteristics of the included studies

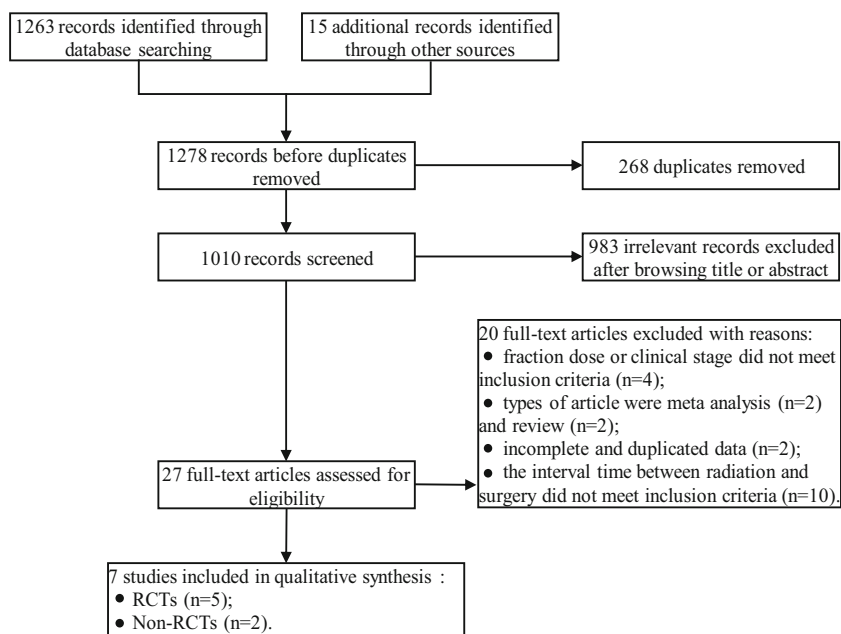
Five of these studies [11, 13, 14, 18, 23] were RCTs, and the other two studies [15, 22] were not RCTs. According to the Cochrane bias assessment (Figs. 2 and 3), four of the five RCTs mentioned randomization and reported the generation of randomized sequences. The performance bias could not be assessed because it was too difficult to keep the radiotherapy regimens a secret between the participants and researchers. Blinding was adopted in only one study. The quality of the two non-RCTs was evaluated by the NOS. As shown in Table 1, the total number of stars of the two non-RCTs was no less than eight each. The basic information on the eligible studies included in our meta-analysis is listed in Table 2.

Primary endpoint

Overall survival

Five studies [11, 13–15, 23] reported the overall survival (OS). Of these five, four [11, 13, 14, 23] were RCTs that directly provided HR values for OS. Four articles [11, 13, 14, 23] reported the results of two RCTs at different follow-up times, so we extracted long-term follow-up results [13, 23]. We excluded the study conducted by Chung et al. [15] because we could not extract effective statistics from the step-like survival curve provided by the authors (the 95% CI of the HR was too broad to believe). The quality of pooled statistics was thus higher after this choice. Ultimately, two studies [13, 23], with a total of 655 patients, were analyzed. Significant heterogeneity was observed ($I^2 = 84\%$, $P = 0.01$); therefore, the random effects model was used. There were no statistically significant differences in OS between the PSRT with delayed surgery group and the PLRT group (pooled HR = 1.30, 95%

Fig. 1 Flow diagram of the study selection process



CI 0.58–2.89, $P = 0.52$; Fig. 4). A subgroup analysis was performed to reduce some heterogeneity.

The subgroup analysis found that patients who underwent PLRT had better OS than those who underwent PSRT without pre/postoperative chemotherapy (HR = 2.05, 95% CI 1.13–3.72, $P = 0.02$). Moreover, patients who underwent PSRT with pre/postoperative chemotherapy had a similar OS to that found in patients in the PLRT group (pooled HR = 0.90, 95% CI 0.70–1.15, $P = 0.41$). The details are shown in Fig. 4.

Secondary endpoint

Disease-free survival

Five studies [11, 13–15, 23] reported the outcome of disease-free survival (DFS). We excluded the study conducted by Chung et al. [15] because statistics extracted from the step-like survival curve provided by the authors were unreliable (the 95% CI of HR was too broad to believe). Four of these five articles [11, 13, 14, 23] reported the results of two RCTs at different follow-up times, and we therefore extracted long-term follow-up results [13, 23]. Ultimately, two studies [13, 23], involving a total of 655 patients, were included. One of the articles directly provided the HR of DFS. We extracted the statistics of the other article [13] from the given survival curve based on the method described in Tierney et al. [19] using Engauge Digitizer software. A random effects model was used for high heterogeneity ($I^2 = 52\%$). No significant differences in DFS were found between the PSRT group and the PLRT group (pooled HR = 1.10, 95% CI 0.73–1.66, $P = 0.64$; Fig. 5).

The subgroup analysis showed that there were no significant differences between either the PSRT with or without

adjuvant chemotherapy subgroup and the PLRT group (HR = 0.95, 95% CI 0.75–1.20, $P = 0.67$ vs HR = 1.48, 95% CI 0.85–2.57, $P = 0.16$, respectively).

Pathological complete remission

Five studies [11, 14, 15, 18, 22] were eligible for the analysis of pathological complete remission (pCR). Of these, three [11, 14, 18] were RCTs and the rest [15, 22] were retrospective cohort studies. The random effects model was used to pool the statistics ($I^2 = 80\%$, $P < 0.01$). There was no significant difference in pCR between the two groups (RR = 0.74, 95% CI 0.37–1.48, $P = 0.39$; Fig. 6). The results remained unchanged after sensitivity analysis.

The subgroup analysis showed that pCR was higher in the PLRT group than in the PSRT group without additional chemotherapy group (RR = 0.42, 95% CI 0.30–0.60, $P < 0.01$), whereas pCR showed an increasing trend in the PSRT with chemotherapy group (RR = 1.37, 95% CI 0.90–2.09, $P = 0.14$). Moreover, the subgroup analysis sharply decreased heterogeneity, and the I^2 values in each subgroup were all 0%.

Early postoperative complications

Three studies [11, 14, 18], with a total of 815 patients, were included in the study of postoperative complications (PCs). No heterogeneity was observed ($I^2 = 0\%$, $P = 0.74$), and a fixed effects model was used. The incidence of PCs was similar in the two groups (RR = 1.21, 95% CI 0.93–1.57, $P = 0.16$; Fig. 7).

The subgroup analysis revealed no significant differences between the two PSRT subgroups (RR = 0.89, 95% CI 0.36–

Fig. 2 Risk of bias summary of RCTs: review authors' judgments about each risk of bias item for each included study

	Random sequence generation (selection bias)	Allocation concealment (selection bias)	Blinding of participants and personnel (performance bias)	Blinding of outcome assessment (detection bias)	Incomplete outcome data (attrition bias)	Selective reporting (reporting bias)	Other bias
Bujko 2016	+	+	?	+	+	+	+
Ciseł 2019	+	+	?	+	+	+	+
Kairevičė 2017	+	+	?	+	+	-	+
Latkauskas 2016	?	?	?	+	+	+	+
Xiao 2018	+	+	-	?	?	+	+

Fig. 3 Risk of bias graph of RCTs: review authors' judgments about each risk of bias item presented as percentages across all included studies

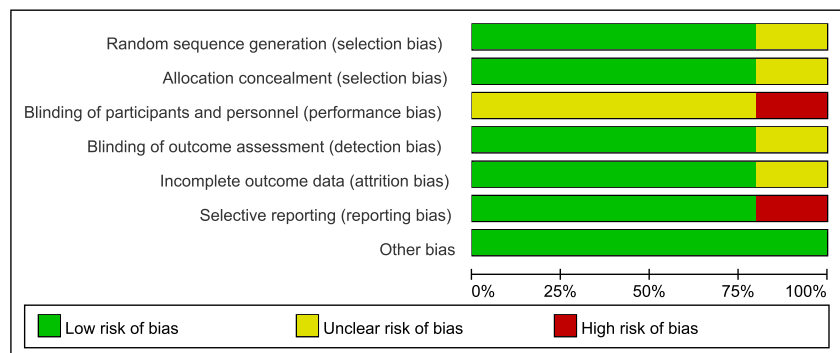


Table 1 Scores of 2 cohort studies according to the Newcastle-Ottawa Scale

Study	Selection		Comparability		Outcomes		Total number of stars		
	Representativeness of the exposed cohort	Selection of the nonexposed cohort	Ascertainment of exposure	Demonstration that the outcome of interest was not present at the start of the study	Comparability of cohorts on the basis of the design or analysis	Assessment of outcome		Was follow-up long enough for outcomes to occur	Adequacy of the follow-up of cohorts
Chung 2016 [15]	★	★	★	★	★★	★	★	0	8
Hoendervangers 2018 [22]	★	★	★	★	★★	★	★	★	9

2.21, $P = 0.80$; RR = 1.25, 95% CI 0.95–1.65; $P = 0.11$, respectively).

Grade 3/4 toxicity

Three trials [11, 15, 18], with a total of 783 patients, reported treatment-related grade 3/4 toxicity. As shown in Fig. 8, treatment-related grade 3/4 toxicity was similar between the PSRT group and the PLRT group (RR = 0.78, 95% CI 0.24–2.49, $P = 0.68$). The random effects model was employed due to high heterogeneity ($I^2 = 87\%$, $P < 0.01$). The results changed after the sensitivity analysis; therefore, a further subgroup analysis was needed.

Intragroup heterogeneity decreased through the subgroup analysis ($I^2 = 39\%$, $P = 0.20$). As shown in Fig. 8, the PSRT without pre/postoperative chemotherapy group had a lower treatment-related grade 3/4 toxicity events than the PLRT group (RR = 0.19, 95% CI 0.08–0.48, $P < 0.01$). Patients who were treated with PSRT plus chemotherapy had similar treatment-related grade 3/4 toxicity to those treated with PLRT (RR = 1.31 95% CI 0.75–2.27, $P = 0.34$).

Local recurrence

Two articles [13, 15], with a total of 212 participants, were included in our meta-analysis. Heterogeneity was not present ($I^2 = 0$, $P = 0.53$); therefore, a fixed effects model was employed. Pooled statistics showed a lack of a significant difference between the PSRT group and the PLRT group (RR = 1.27, 95% CI 0.39–4.15, $P = 0.70$; Fig. 9). Sensitivity analysis proved the synthesis results were robust.

The subgroup analysis showed no significant differences in the two subgroups (RR = 1.06, 95% CI 0.28–4.07, $P = 0.93$ vs RR = 2.79, 95% CI 0.18–42.42, $P = 0.46$, respectively).

Distant metastasis

Four articles [11, 13, 15, 23], with a total of 727 patients, were eligible for the distant metastasis (DM) analysis. Two [11, 23] of the four articles reported the results of an RCTs at different follow-up times, so we extracted long-term follow-up results [23]. Ultimately, three articles were analyzed. Considering the low heterogeneity, a fixed effects model was employed, as shown in Fig. 10 ($I^2 = 37\%$, $P = 0.21$). No significant differences were observed between the two groups (RR = 1.06, 95% CI 0.85–1.33, $P = 0.58$). The results remained unchanged after sensitivity analysis.

The subgroup analysis showed no significant differences in the two subgroups (RR = 1.39, 95% CI 0.79–2.43, $P = 0.25$ vs RR = 1.01, 95% CI 0.79–1.28, $P = 0.94$, respectively).

Table 2 Characteristics of the studies included in the meta-analysis

Study	Resources and study year	Clinical tumor stage	Type of study	Outcomes	Follow-up time (months)	Group Sample size	Age (years)	Distance from the anal verge to the lower border (cm) (M-H)/L	Intervention
Xiao 2018 [18]	2014–2017 (China)	cT3–4N0–1, M0	RCTs	PCs, grade 3/4 toxicity, pCR, hospital stay, cost	NA	1	60 (19–75) ^c	42/56	25 Gy/5F, TME surgery in 6 to 8 weeks
						2	58 (23–75) ^c	41/57	50 Gy/25F, concurrent CT, TME surgery in 6 to 8 weeks, CT
Laitkuskas 2016 [14]	2007–2013 (Europe)	cII/III	RCTs	OS, DFS, pCR, downstaging, LR, DM, PCs, LR	39.7 (4.9–79.7) ^c	1	65.6 ± 9.5 ^a	34/34	25 Gy/5F, TME surgery in 6 to 8 weeks, CT
						2	63.14 ± 10.1 ^a	42/30	50 Gy/25F, concurrent CT, TME surgery in 6 to 8 weeks, CT
Kairevičė 2017 [13]	2007–2013 (Europe)	cII/III	RCTs	OS, DFS, LR, DM	60.5 (5–108) ^c	1	65.6 ± 9.51 ^a	34/34	25 Gy/5F, TME surgery in 6 to 8 weeks, CT
						2	63.1 ± 10.13 ^a	42/30	50 Gy/25F, concurrent CT, TME surgery in 6 to 8 weeks, CT
Bujko 2016 [11]	2008–2014 (Poland)	cT3–4M0	RCTs	OS, DFS, pCR, RO, PCs, DM, grade 3/4 toxicity	35 (21–50) ^b	1	60 (54–60) ^b	113/148	25 Gy/5F, CT in 7 weeks, TME surgery at least 4 weeks later
						2	60 (56–65) ^b	115/138	50.4 Gy/28F, concurrent CT, TME surgery at least 4 weeks later, CT
Cisel 2019 [23]	2008–2014 (Poland)	cT3–4M0	RCTs	OS, DFS, DM	35 (21–50) ^b	1	60 (54–60) ^b	113/148	25 Gy/5F, CT in 7 weeks, TME surgery at least 4 weeks later
						2	60 (56–65) ^b	115/138	50.4 Gy/28F, concurrent CT, TME surgery at least 4 weeks later, CT
Chung 2016 [15]	2010–2015 (Korea)	cT3–4N0–2, M0	Retrospective study	OS, DFS, LR, DR, DM, grade 3/4 toxicity, downstaging, pCR	25 (3–58) ^c	1	72 ^c	7/12	25 Gy/5F, CT, surgery in 6 to 8 weeks, CT
						2	72 ^c	65/18	50.4 Gy/28F, concurrent CT, surgery in 6 to 8 weeks
Hoendervangers 2018 [22]	2008–2014 (Netherlands)	cT2N2M0, T3N1–2M0, T4N–anyM0	Retrospective study	pCR, near pCR, T/N downstaging	28.8 (18–42) ^b	1	76 (67–85) ^b	189/176	25 Gy/5F, CT, surgery at least 4 weeks later
						2	63 (53–73) ^a	1761/1680	45–50 Gy/25F, concurrent CT, surgery at least 4 weeks later

^a Values are presented as the mean ± standard deviation. ^b Values are presented as the interquartile range. ^c Values are presented as the median (range).

Group 1: patients treated with PSRT with delayed surgery.

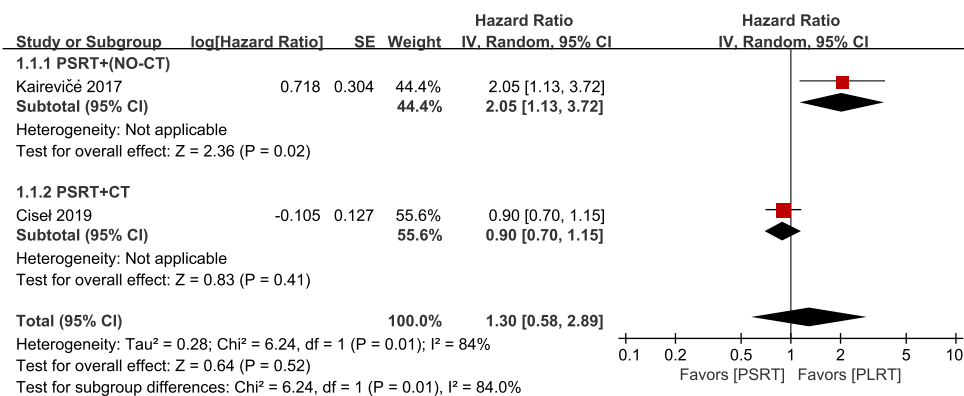
Group 2: patients treated with PLRT.

TME total mesorectal excision, CT chemotherapy, RCT randomized controlled trial, PCs postoperative complications, pCR pathological complete remission, OS overall survival, DFS disease-free survival, LR local recurrence, DM distant metastasis

(M-H): The number of patients whose distance from the anal verge to the lower border of tumor was more than 5 cm

L: The number of patients whose distance from the anal verge to the lower border of tumor was less than 5 cm

Fig. 4 Forest plot showing the comparison of PSRT with delayed surgery and PLRT regarding OS



Discussion

Main results: Our meta-analysis revealed no statistically significant differences in terms of OS, DFS, pCR, treatment-related grade 3/4 toxicity, PCs, local recurrence (LR), or DM. The two subgroups were then divided based on whether additional chemotherapy was available in the PSRT arm. The subgroup analysis revealed no significant differences with regard to DFS, DM, PCs, or LR. The subgroup analysis revealed that compared with the long-course radiotherapy group, the preoperative short-course radiotherapy without adjuvant chemotherapy group had not only a lower treatment-related grade 3/4 toxicity (RR = 0.19, 95% CI 0.08–0.48, *P* < 0.01) but also a significantly lower OS and pCR (HR = 2.05, 95% CI 1.13–3.72, *P* = 0.02; RR = 0.42, 95% CI 0.30–0.60, *P* < 0.01, respectively). Interestingly, pCR showed an increasing trend in the PSRT with adjuvant chemotherapy group, but the difference was not significant (RR = 1.37, 95% CI 0.90–2.09, *P* = 0.14). Furthermore, patients treated with PLRT suffered from more treatment-related grade 3/4 toxicity than those treated with PSRT without chemotherapy (RR = 0.19, 95% CI 0.08–0.48, *P* < 0.01). In summary, our meta-analysis found that PSRT with delayed surgery had similar efficacy and side effects compared with PLRT and that adding additional chemotherapy to the PSRT group resulted in few advantages over the long-course radiotherapy group.

Although the most recent NCCN and ESMO guidelines both recommend PSRT and PLRT as standard treatment regimens for locally advanced rectal cancer, the oncology radiologists’ preferences for radiation vary significantly among different countries. In the USA, PSRT is rarely performed, and no more than 5% of locally advanced rectal cancer patients accepted PSRT compared with PLRT [24, 25] due to the lack of economic incentives, unfamiliarity with short discharge techniques, and uncertainty about the efficacy and side effects of such treatment [26]. Additionally, PSRT accounts for only 15% of radiation therapy methods in England [27]. For a large number of patients, PLRT is the first option. PSRT still has a low status in Europe, and patients receiving PSRT account for less than 20% of patients receiving PLRT [28, 29]. Therefore, the role of PSRT in the application of cancer treatment must be further elucidated.

Moreover, neither the NCCN guidelines nor the ESMO guidelines clearly provide an optimal time interval for the gap between the end of PSRT and surgery. There are two common time intervals at which short-course radiotherapy and surgery are performed: short-course radiotherapy followed immediately by surgery within 10 days (this method is more commonly practiced in Europe) and short-course radiotherapy followed by delayed surgery (at least 4 weeks after the last radiotherapy is completed).

One study reported that there was no survival advantage in patients treated with surgery within 2 weeks in the PSRT arm

Fig. 5 Forest plot showing the comparison of PSRT with delayed surgery and PLRT regarding DFS

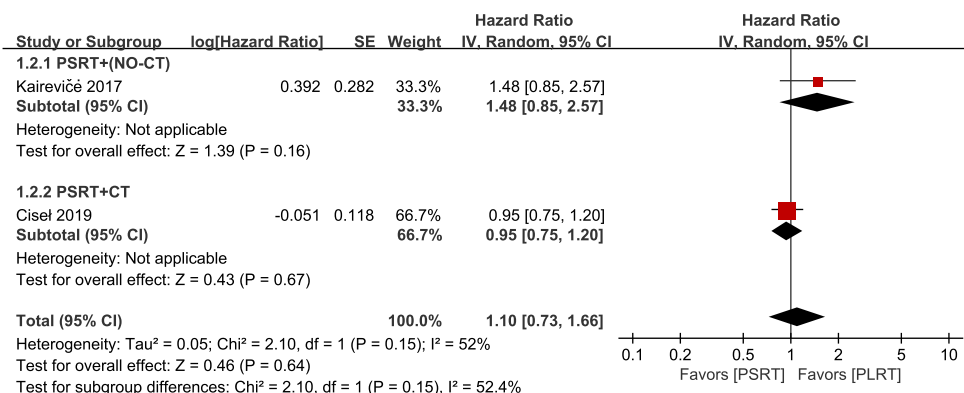
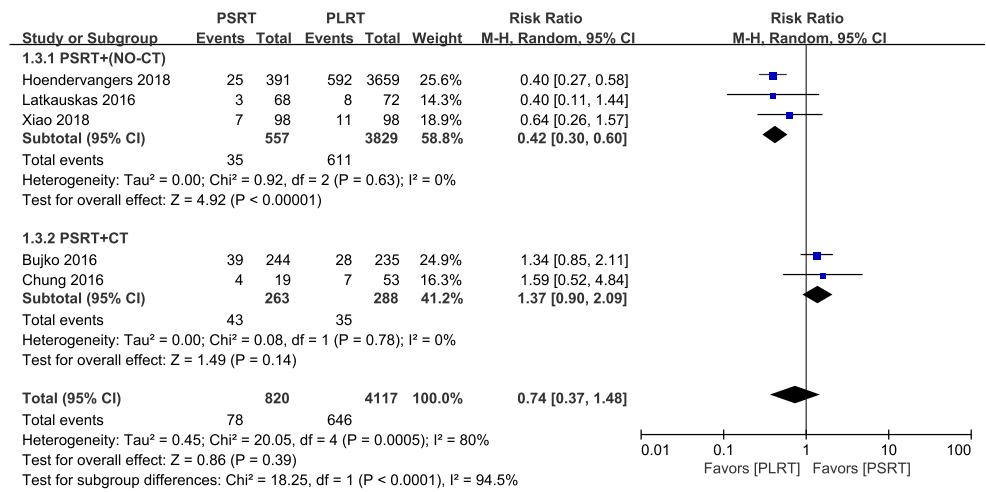


Fig. 6 Forest plot showing the comparison of PSRT with delayed surgery and PLRT regarding pCR



and that a 2-week time interval did not achieve a desirable R0 resection rate [30]. Some RCTs have shown that compared with PSRT followed by immediate surgery (within 10 days) PSRT followed by delayed surgery (at a less than a 4-week time interval) significantly improved tumor pCR, downstaging, and the R0 resection rate, and resulted in fewer PCs [16, 31]. The Stockholm III Trial [32] found that PSRT with delayed surgery (4 to 6 weeks after PSRT) resulted in greater tumor downstaging than PSRT with immediate surgery. A meta-analysis conducted by Wu et al. [10] showed conclusions consistent with these previous findings: PSRT followed by surgery (at least 4 weeks later) was superior to PSRT followed by immediate surgery in terms of pCR, downstaging, the R0 resection rate, and PCs. Disappointingly, however, no conclusive evidence has been presented to demonstrate that delaying surgery can improve survival.

Another study demonstrated that a delaying surgery by at least 6 weeks achieved maximal tumor regression and that the tumors in the PSRT group continued to shrink over the

following 1 to 2 weeks [33]. Furthermore, the tumor regression grade was more significant when the time interval to surgery was prolonged to more than 8 weeks in the PSRT group [34]. Delayed surgery not only potentially provides sufficient time for tumor shrinkage but also allows opportunities for systemic therapy. A research conducted by Pettersson et al. [35] showed that performing surgery at an interval of 4–8 weeks in the PSRT group resulted in similar PCs compared with PLRT. Our study did not find advantages in the PSRT group with regard to pCR, and the results were not reversed even in the PSRT with chemotherapy subgroups. One possible reason for this finding was that the time interval between the procedures was not long enough, and the tumor response thus did not reach its peak. If a sufficient time interval is provided, the tumors in the PSRT group may be reduced to a minimum, which may increase pCR and the possibility of R0 resection and potentially improved local control.

Previous studies have supported the notion that PSRT with immediate surgery is as effective as PLRT in terms of efficacy and safety [36–38], whereas the results of

Fig. 7 Forest plot showing the comparison of PSRT with delayed surgery and PLRT regarding PCs

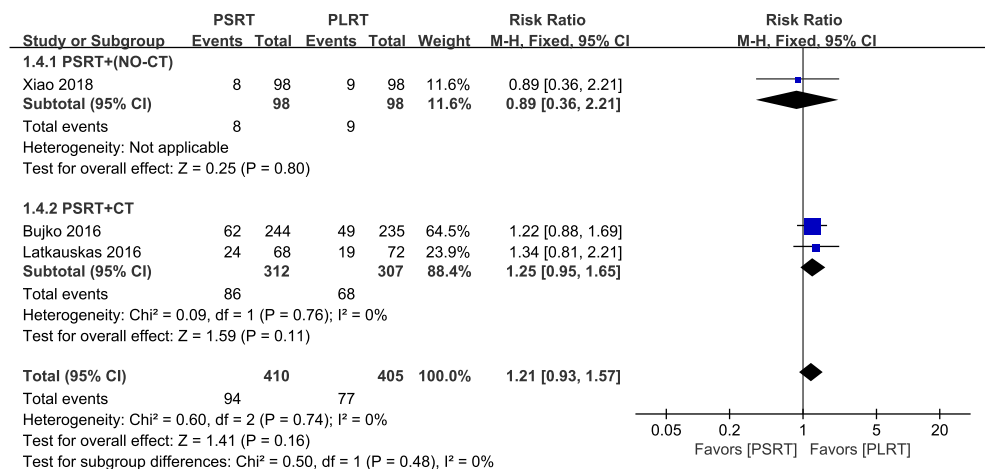
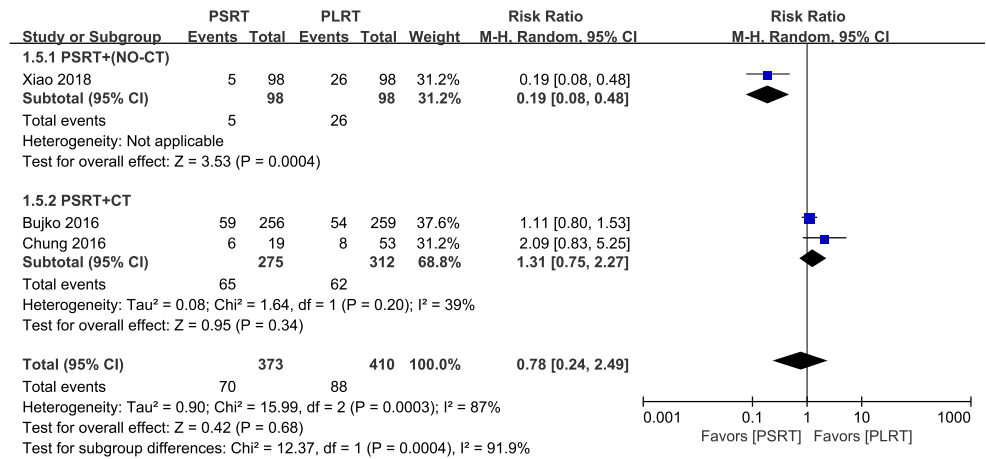


Fig. 8 Forest plot showing the comparison of PSRT with delayed surgery and PLRT regarding grade 3/4 toxicity



comparison of PSRT with delayed surgery and traditional PLRT are still controversial. Two articles [13, 14] reported the results of the same randomized study with different follow-up times; both observed that PLRT resulted in longer DFS than the PSRT with delayed surgery group. Nevertheless, a retrospective cohort study showed the opposite result [12]. According to another study, 3-year OS

was higher in patients who received PSRT with delayed surgery than in those who received conventional PLRT (73% vs 65%, respectively, $P < 0.05$) [11]. However, most of the studies have found that PSRT with delayed surgery is as effective as PLRT in terms of OS [12, 14–16, 39]. Our meta-analysis focused on the PSRT in which surgery was completed at least 4 weeks later versus conventional long-

Fig. 9 Forest plot showing the comparison of PSRT with delayed surgery and PLRT regarding LR

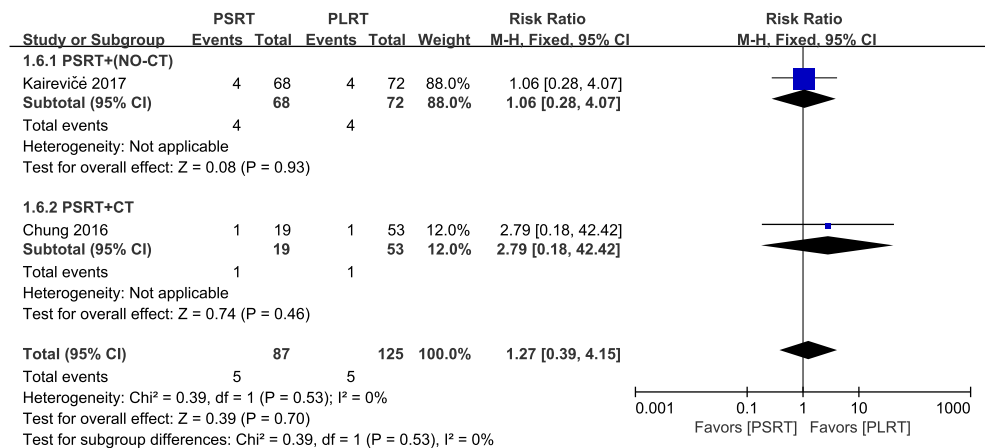
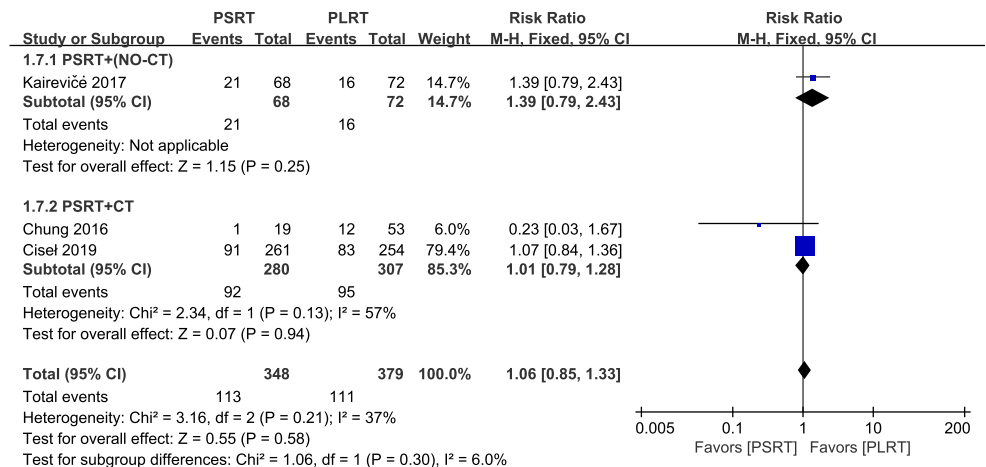


Fig. 10 Forest plot showing the comparison of PSRT with delayed surgery and PLRT regarding DM



course chemoradiotherapy. Pooled statistics showed no significant differences between the two arms regarding survival.

Interestingly, our subgroup analysis found that additional chemotherapy cannot significantly improve OS in the PSRT with delayed surgery group (Fig. 4), and long-course radiotherapy concurrent with chemotherapy based on 5-Fu leads to higher OS and pCR compared with no additional chemotherapy in the PSRT with delayed surgery group. Unfortunately, the improvement of the OS reported in the PLRT group came at the cost of increased treatment-related grade 3/4 toxicity. In contrast, a single-arm retrospective study performed in the UK showed that PSRT plus adjuvant chemotherapy improved OS only in patients with two or more risk factors [40]. This observation may imply that chemotherapy provides advantages only in selected patients; otherwise, patients may suffer more from overtreatment.

A Cochrane systematic review demonstrated that preoperative chemoradiotherapy only improved local control, without providing a survival benefit [41]. In our subgroup analysis, we did not find that additional chemotherapy in the PSRT arm provides advantages with regard to LR but we did observe that OS was lower in the PSRT without chemotherapy subgroup. We believe that additional adjuvant chemotherapy is, theoretically, more conducive to the elimination of micrometastatic disease [25], but our meta-analysis does not indicate that the DM rate is lower in clinical practice.

Two studies showed that adjuvant chemotherapy following preoperative neoadjuvant therapy provided no survival benefit over preoperative neoadjuvant therapy alone [42, 43], while another study showed that adjuvant chemotherapy increased OS [44]. According to another study, adding consolidation chemotherapy added to postoperative therapy provided no benefit to OS in patients treated with PSRT who underwent an operation after 3 to 7 days [45]. This finding may also imply that adding additional chemotherapy before surgery may improve the survival rate more than adding it after surgery. Therefore, the optimal time to add chemotherapy to patients' treatment plan remains unclear, and the time interval between the completion of short-course radiation and surgery may play an important role in survival. More RCTs are needed to illustrate whether the addition of chemotherapy to the PSRT with delayed surgery regimen is favorable. We expect the ongoing RAPIDO trial (NCT: 01558921) [46] in Sweden and a phase III ongoing trial in China (NCT: 02533271) [47] to clarify this issue.

The limitations of the current study should not be neglected, since five RCTs and two non-RCTs were included in our meta-analysis. First, the Cochrane risk of bias assessment showed that all RCTs exhibited some potential risk of bias (details shown in Fig. 2). Second, clinical heterogeneity was present in the included studies. Patients in most of the eligible studies were in clinical stage II–III, and the location

of the tumor varied among patients. Third, the types of specific chemotherapy drugs were different across studies. Additionally, follow-up times and methods varied among different studies. For example, the results were tracked by telephone follow-up in Xiao's [18] study, and two [13, 14] articles came from different stages of the same RCT (NCT: 00597311). Finally, the operative methods and quality were different in each study.

Acknowledgments The authors would like to thank Yunnan Cancer Hospital.

Funding information This study received financial support from Kunming Medical University Cancer Radiation Therapy Technology Innovation Team Construction Project Funding and Funding for the "Ten Thousand Plan" Youth Talent Project in Yunnan Province.

Compliance with ethical standards

All the authors have complied with all applicable ethical standards.

Conflict of interest The authors declare that they have no conflict of interest.

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