Gene-associated microsatellite markers confirm panmixia and indicate a different pattern of spatially varying selection in the endangered Japanese eel *Anguilla japonica**

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Abstract The Japanese eel (Anguilla japonica) is a commercially important fish species in East Asia and its recruitment has been rapidly declining since 1990s. Clarifying the genetic population structure of A. japonica is the basis of multinational cooperation on its management and protection, due to its large distribution range. Gene-associated markers have been proved powerful in delineating fine-scale population genetic structure and spatially varying selection. In the present study, we developed 24 polymorphic gene-associated microsatellite markers including 18 loci associated with the genes under selection in the two North Atlantic eel species (Anguilla anguilla and Anguilla rostrata) and 6 loci based on transcript sequences. A total of 13 geographic populations were sampled across its distribution range, including 11 samples from China (9 from China's mainland and 2 from Taiwan region), and 2 samples from Japan. A total of 416 individuals (mostly glass eels) were collected and genotyped at the 24 microsatellites. All measures of differentiation were accordant with a panmictic scenario (F_{sT} =-0.001) in A. japonica. No footprints of spatially varying selection were found, indicating that the selection pattern in A. japonica might be different from that in the two North Atlantic eel species. We suggest that A. japonica should be managed as a single unit and management and conservation efforts must be coordinated at the international level, as overexploitation in any region will decrease its recruitment across the whole distributional range.

Keyword: Japanese eel; Anguilla japonica; panmixia; gene-associated microsatellite; spatially varying selection

1 INTRODUCTION

The Japanese eel (*Anguilla japonica*) is a temperate catadromous anguillid eel, distributed in the rivers of China (China's mainland and Taiwan region), Japan, and Korea (Tesch, 1977). It spawns in the waters to the west of Mariana Islands near 14°N–16°N, 142°E (Tsukamoto, 1992, 2006), 2 000 to 3 500 km away from the East Asian continent. The leptocephali hatch between April and November (Tsukamoto, 1990; Tzeng, 1990; Tsukamoto et al., 2003), and then drift with the currents, reaching the coasts of East Asia in 4 to 6 months. They metamorphose into glass eels along

the continental shelf and then enter estuaries where they continue to grow as elvers (Tesch, 2003). Elvers grow to yellow eels during their upstream migration, and yellow eels spend the next 5 to 10 years in freshwater rivers. Once eels attain sexual maturity,

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they migrate back to spawning ground for reproduction (Tesch, 1977).

The Japanese eel is a commercially important fish species in East Asia. And the stocks are outside safe biological limits and the fishery has not been sustainable in recent years. The Japanese eel population is currently estimated to be less than 10% of 1970s level (Dekker, 2003), and it is classified as "Endangered" by the International Union for Conservation of Nature (IUCN) Red List (https://www. iucnredlist.org/species/166184/1117791). Although the exact reason for the decline is unclear, there is a direct correlation between the decline in eel catch and population size to the habitat modifications and impediments to upstream migration (Itakura et al., 2015). Other reasons such as pollution, disease, oceanic climate change and habitat losses could also influence the abundance of the Japanese eel population (Arai, 2014). However, the reliance of eel aquaculture on wild-caught juveniles or glass eel seems to be more important reason (Tsukamoto, 2013). Due to the large distribution range of A. japonica, the effective management needs to involve the efforts from all countries within its range. And clarifying the genetic population structure of A. japonica is the basis of multinational cooperation on management and protection. In addition, it is also of vital importance to understand the genetic consequences of external disturbance to this endangered species, such as ongoing climate change, ocean warming and overfishing.

Some work has been conducted to test panmixia of A. *japonica* using different molecular markers; however, the results are not consistent with each other. Panmixia was accepted based on analyses on mtDNA sequences (Sang et al., 1994; Ishikawa et al., 2001). Microsatellite markers, which have more power in studying population structure, were also used to solve this problem. Tseng et al. (2006) divided the populations of A. japonica into two groups based on eight microsatellite markers, and they found no isolation by distance (IBD) pattern in either group. The results challenged the panmixia hypothesis by finding small but significant genetic differences. However, the study used a small number of markers and reported higher temporal than geographic difference. Notwithstanding the results have rejected the null hypothesis of panmixia for the genetic structure of anguillid eels, several studies have challenged this opinion recently. Han et al. (2010) used eight microsatellite markers analyzing the spatial and temporal structure of A. japonica. Although a

significant difference was seen among annual cohorts and spatial samples, no specific temporal or spatial scale patterns were observed in the pairwise $F_{\rm ST}$ tests or the phylogenetic tree of all samples and the results of the IBD test and the isolation by time (IBT) test were both insignificant. A stable genetic structure could not be observed. The result provided evidence for panmixia. After that, Minegishi et al. (2012) also found no significant genetic difference among 9 localities using 6 microsatellite markers and reported that Japanese forms a panmictic population.

Traditional approach using neutral markers have reported shallow population structure in many marine fishes, which is assumed to be the result of high levels of gene flow. However, in species with large effective population size, the weak genetic structure could also result from the limited effects of genetic drift (Hauser and Carvalho, 2008). Russello et al. (2012) detected ecotype-level divergence using 8 outlier loci in Okanagan Lake kokanee (Oncorhynchus nerka); however, there was no evidence of divergence at neutral loci. Milano et al. (2014) detected a dramatic divergence between Atlantic and Mediterranean populations and fine-scale significant population structure in European hake (Merluccius merluccius) using outlier loci, which was not revealed using neutral loci. Many marine organisms have relatively large effective population size, and they are under limited effects of genetic drift. In such a scenario, gene-associated markers might have more power in unveiling the cryptic genetic population structure than neutral markers.

Local adaptation is how organism response to selective pressures in their local habitats, acting on genetically controlled fitness differences among individuals (Rellstab et al., 2015). Since eels are panmictic and show no philopatry (Als et al., 2011; Pujolar et al., 2014), long-term local adaptation is not possible in eels but single-generation signatures of local selection still can be detected (Gagnaire et al., 2012; Pujolar et al., 2014), which is a completely different scenario relative to other species. In other fish species, fish can adapt to local conditions and those individuals more resistant to adverse environmental condition in a population will survive, while less adapted individuals will not, and this is passed on to the next generation. However, this is not the case in eels: Individuals may have some good genetic combinations which provide good adaptation for survival in their freshwater habitat, however, habitats for offspring of these individuals would change due to

Table 1 Sampling information of the Japanese eel, including sampling location and date, sampling size (n), and development stage

Category	Sampling locality	Code 1	Development stage	п	Latitude	Longitude	Date
	Dandong Yalu River estuary	DD	Glass eel	30	40.00°	124.36°	May 20, 2014
	Pudong Changjiang (Yangtze) River estuary	PD	Glass eel	36	30.85°	121.85°	March 20, 2014
	Fuyang Qiantang River	FY	Yellow eel	6	30.05°	119.96°	January, 2014
			Yellow eel	12			March 23, 2014
			Glass eel	10			May 4, 2014
			Yellow eel	3			May 4, 2014
	Yuhuan	YH	Glass eel	30	28.12°	121.28°	March 22, 2014
China	Ruian Feiyunjiang estuary	RA	Glass eel	36	27.73°	120.66°	March 22, 2014
	Sansha	SS	Glass eel	36	26.87°	120.04°	March 23, 2014
	Quanzhou Jinjiang estuary	QZ	Glass eel	25	24.86°	118.64°	January 8, 2015
	Shantou Rongjiang estuary	ST	Glass eel	30	23.39°	116.82°	January 8, 2015
	Zhongshan Zhujiang (Pearl) River estuary	ZS	Glass eel	25	22.49°	113.58°	January 6, 2015
	Yilan Lanyang River estuary	MY	Glass eel	39	24.72°	121.84°	December 30, 2014–January 22, 2015
_	Pingtung Gaoping River estuary	MA	Glass eel	28	22.47°	120.44°	December 31, 2014
Japan	Chiba Tone River estuary	MJ	Glass eel	38	35.74°	140.82°	January, 2011
-	Aichi	MC	Glass eel	32	34.73°	137.10°	April, 2016

random dispersal, which could lead to the loss of these adaptations. Trans-generational signatures of selection cannot be detected while within-generation selection acting on genes could be found. Recently, geneassociated markers have shown their power in detecting signature of selection in several organisms. In the case of anguillid eels, signature of selection is much more difficult to find, due to the special migratory features. However, recent studies have found signature of selection in North Atlantic eels. Gagnaire et al. (2012) scanned for footprints of selection in the American eel (Anguilla rostrata) using a panel of 100 single nucleotide polymorphisms (SNPs) and found the allele frequencies of thirteen loci had a correlation with environmental factors. In the study of European eel (Anguilla anguilla) using the same SNP panel, 10 loci were detected showing genetic-by-environment associations (Ulrik et al., 2014). Pujolar et al. (2014) detected signature of selection in the European eel 50 354 restriction-site-associated using DNA sequencing (RAD) generated SNPs, and dozens of loci showed evidence of divergent selection. However, the study on local adaptation is still absent in A. japonica.

The previous studies have led to conflicting results

on panmixia, but all previous studies used a very limited (6-8) number of microsatellites. In the present study, 24 polymorphic gene-associated microsatellite markers were developed for *A. japonica*. These markers were then used to test the hypothesis of panmixia and to detect footprints of selection in *A. japonica*.

2 MATERIAL AND METHOD

2.1 Sampling and DNA extraction

Glass eels or yellow eels were collected from 13 sampling locations spanning the natural species distribution between January 2011 and April 2016 (Table 1), and the whole organism of these eels was preserved in 95% ethanol. Whole genomic DNA was extracted using a standard phenol-chloroform extraction and ethanol precipitation, following the proteinase K digestion.

2.2 Microsatellite markers development and genotyping

Microsatellite loci were obtained in two ways. Firstly, we searched for the transcript sequences of

the genes against the online transcriptome database for A. japonica (http://molas.iis.sinica.edu.tw/jpeel/) built by Hsu et al. (2015) using the names of selected genes in A. rostrata (Gagnaire et al., 2012) and A. anguilla (Pujolar et al., 2014) as keyword. The detail information of the microsatellite loci in each transcript sequence were obtained using MISA (http://pgrc.ipk-gatersleben.de/misa). software Secondly, we developed another panel of microsatellite loci from the transcript sequences randomly chosen from the database of A. japonica. To obtain a longer flanking region for primer design, each transcript sequence with the microsatellite locus was mapped on the draft genome of A. japonica (Henkel et al., 2012). Finally, Premier 5.0 software (Premier Biosoft International, Palo Alto, CA, USA) was used to design primers. 23 primer pairs were designed on the basis of the gene under selection in North Atlantic eels and 12 primer pairs were designed from transcript sequence randomly selected from database. Out of the 36 primer sets, 24 primer pairs consistently amplified fragments of the expected size and were polymorphic. Among the 24 polymorphic loci, 18 loci were developed associated with the gene under selection in North Atlantic eels, and the other 6 loci were developed based on transcript sequence which were randomly selected from the database. PCR was carried out using a three-primer system (Liu and Avise, 2011). All forward primers were 5' tailed with M13 forward sequence а (5'-GGAAACAGCTATGACCAT-3'). An M13 primer of the same sequence but 5'-labelled with three different fluorescent dyes was used in combination with forward primers. All loci were amplified separately on a Mastercycler (Eppendorf). The PCR consisted of approximately 50 ng genomic DNA, $2 \mu L$ of $5 \times$ buffer (Promega, Madison, WI), 0.2 mmol/L of each dNTP, 0.2 µmol/L labeled M13 reverse primer and locus specific primer without tail, 0.02 µmol/L locus specific primer with M13 reverse tail, 0.25 U Tag DNA polymerase (Promega, Madison, WI) and water up to 10 µL. The thermal cycling protocol consisted of 3 min at 95°C followed by 35 cycles of 20 s at 95°C, 52 s at 52°C and 30 s at 72°C, and a final elongation of 10 min at 72°C. PCR products labelled with different fluorescent primers were pooled and electrophoresis was made on an ABI automated sequencer 3730x1 (Applied Biosystems). Size determination of alleles were made by comparison with the GS500-ROX size standard using GeneMarker (SoftGenetics, State College, USA).

2.3 Genetic diversity analysis

Genetic diversity indices including observed heterozygosity (H_0), expected heterozygosity (H_E), the number of alleles (N_a), and inbreeding coefficient (F_{1S}) were calculated using GenAlEX 6.5 (Peakall and Smouse, 2006, 2012). The presence of null alleles was assessed using MICRO-CHECKER software version 2.2.3 (Van Oosterhout et al., 2004). Genotypic linkage disequilibrium (LD) and deviation from Hardy-Weinberg equilibrium (HWE) were estimated using GENEPOP 4.5.1 (Rousset, 2008). The significant thresholds for HWE and LD tests were adjusted using Bonferroni correction.

2.4 Population structure

Global F_{ST} value across all populations was calculated using FSTAT version 2.9.3.2 (Goudet, 2001). Pairwise $F_{\rm ST}$ values between all populations were obtained using ARLEQUIN 3.5 (Excoffier and Lischer, 2010) and significance was assessed by 10 000 bootstrapping permutations. A modified false discovery rate (FDR) correction was applied to adjust the significant threshold of pairwise $F_{\rm ST}$ values (Benjamini and Yekutieli, 2001). The Bayesian clustering program STRUCTURE version 2.3.4 (Pritchard et al., 2000), was used to infer population structure and assign individuals to populations. Admixture model was used, assuming correlated allele frequency, with the option of 'with no prior knowledge of sampling locations'. The program was run with an initial burn-in of 100 000 cycles and 1 000 000 Markov chain Monte Carlo (MCMC) repetitions, and K was modeled from K=1 to K=13. Runs were iterated 10 times for each K. The optimal value of ΔK was determined according to Evanno et al. (2005) method using StructureSelector (Li and Liu, 2018).

2.5 Footprints of selection

The F_{ST} outlier approach (Beaumont and Nichols, 1996) implemented in LOSITAN (Antao et al., 2008) was used to identify outlier loci. The program generated a plot of F_{ST} vs heterozygosity.

One hundred thousand simulations were performed using the stepwise mutation model (SMM) and the neutrality of microsatellite markers was determined. Markers having $F_{\rm ST}$ values higher than the 95% confidence interval were inferred to be subject to divergent selection, and markers having $F_{\rm ST}$ values lower than the 95% confidence interval were inferred to be subject to balancing selection. False discovery rate was set to 0.01.

Spatial analysis method (SAM) was used to assess the association between allelic frequencies and environmental variables, based on multiple univariate logistic regression (Joost et al., 2007). Environmental factors were chosen including degrees north latitude, degrees east longitude, and sea-surface temperature at river mouth averaged across the 10 days, 30 days and 3 months that preceded the sampling date. Temperature data were acquired from a National Oceanic and Atmospheric Administration (NOAA) database containing geo-referenced sea-surface temperatures. The FY population was excluded from this analysis, because the sample included different development stages and the sampling date was also different. Two populations (MJ and MC) from Japan and one population from Taiwan, China (MY) were excluded from genotype-temperature association analysis, because the samples were collected from different date in the same month.

3 RESULT

3.1 Genetic diversity

The average number of alleles per locus across populations ranged from 4.846 (SE=0.274) to 29.231 (SE=1.156) (Table 2). The H_0 and H_E of each sample ranged from 0.675 (SE=0.038) and 0.746 (SE=0.039) to 0.716 (SE=0.033) and 0.769 (SE=0.034) (Table 3). Among a total of 312 sample x locus combinations, 29 were deviated from HWE after Bonferroni corrections (α =0.05, K=13). After Bonferroni correction, no linkage disequilibrium was detected (α =0.05, K=276). Null alleles existed at eight loci in more than three samples (not in all samples).

3.2 Population structure

The global non-significant $F_{\rm ST}$ value of -0.001 (95% CI: -0.002 < $F_{\rm ST}$ < 0.000) indicated that there was no significant genetic differentiation among these samples. None of the pairwise $F_{\rm ST}$ values were significant after FDR correction (α =0.05, K=78) (Table 4). STRUCTURE analysis identified K=6 was the optimal number of clusters (Fig.1). However, all populations showed similar patterns when K=6, demonstrating no genetic structure (Fig.2).

3.3 Test of selection

No outlier was found in outlier analysis (Fig.3). No significant correlations were found in SAM, indicating that the 24 loci were not under spatially varying selection.

4 DISCUSSION

4.1 Confirmed panmixia in A. japonica

Japanese eel has long been considered having a panmictic population structure because sexually matured eels from their distribution range migrate and spawn in a single spawning site and the larvae disperse from the spawning site via the North Equatorial Current (NEC) and the Kuroshio Current (KC) and are transported back to the freshwater habitats randomly. However, there is still debate on the hypothesis of panmixia for A. japonica. Chan et al. (1997) detected spatial genetic heterogeneity in A. japonica using allozyme markers, the significant clinal shift in allelic frequency was putatively attributed to single-generation selection along an environment gradient (Avise, 2003). The geographic range of the sampling sites in our study was similar to that in Chan et al. (1997), however, we found no spatial heterogeneity in the gene frequency of the 24 microsatellites. Tseng et al. (2006) divided the populations of A. japonica into two groups including a low-latitude group (South China and Taiwan, China) and a high-latitude group (Japan, Korea, and North China) using eight microsatellite markers. However, their study used a small number of markers and reported higher temporal difference than geographic difference. Han et al. (2010) suggested that the genetic partitioning detected in Tseng et al. (2006) could be the results of 'chaotic genetic patchiness' and random variations in parental contributions to reproductive activity, incomplete mixing of larvae and kin aggregation might contribute to the subdivision of the Japanese eel population. Han et al. (2010) conducted the most comprehensive study on A. japonica using microsatellite markers and confirmed panmixia of this species.

In the present study, all measures of differentiation were accordant with a panmixia scenario. Our results supported the hypothesis of panmixia in *A. japonica*. Great efforts have been put into the study of genetic population structure for *A. japonica*, however, no one has ever tried to use gene-associated markers. Geneassociated makers can be a good option to detect subtle genetic population differentiation in marine organisms. On one hand, traditional neutral markers may be ineffective when populations are recently isolated, and divergence is not yet reflected at neutral loci (Russello et al., 2012). On the other hand, genetic drift has little effects in species with large effective population size (such as marine fishes), which may

	GenBank		e 2 Characterization of 24 Interosatence foet for <i>Anguna Jupon</i>	<i>nca</i> aci				H		H	
Locus	accession No.	Motif	Primer sequence $(5^{-}\rightarrow 3^{\prime})$	$T_{\rm a}$	R	Mean	SE	Mean	SE	Mean	SE
AJ_SSR_102673	AVPY01102674.1	(TG)15	GGAAACAGCTATGACCATGGGAAATGTGACAAGAGGG ATCGGCGTTAAAAAAGA	50	210-432	19.385	0.636	0.763	0.028	0.926	0.002*
AJ_SSR_124106	AVPY01124107.1	(CT)15	GGAAACAGCTATGACCATGACTCAAATAACACAATGAAGCA TCACATAGACAGTGGAAGACAG	54	155–195	5.692	0.414	0.340	0.024	0.345	0.022
AJ_SSR_12536	AVPY01012537.1	(TG)15	GGAAACAGCTATGACCATGTATCTGAAATAGACTATGCATCC ACATCTGTTAATTGTTGTACCTT	57	382-434	8.769	0.361	0.786	0.014	0.802	0.006
AJ_SSR_180064	AVPY01180065.1	(CT)18	GGAAACAGCTATGACCATGGAACTCCACATCCTTTTT CACACTTTACCCATCAAC	55	378-442	15.846	0.576	0.821	0.016	0.895	0.005*
AJ_SSR_200878	AVPY01200879.1	(GT)9	GGAAACAGCTATGACCATGAGTATTGTACTGGTTCCCT TAGATTCTCTCAGATGTGG	55	282-406	11.308	0.263	0.737	0.023	0.850	0.006*
AJ_SSR_205239	AVPY01205240.1	(CT)21	GGAAACAGCTATGACCATGCAGTGACGCAAAACTCCAGAG AAACGTGCATGTTCAACCAGG	55	345-409	10.538	0.501	0.811	0.014	0.841	0.005
AJ_SSR_209945	AVPY01209946.1	(GT)14	GGAAACAGCTAFGACCAFGCTGGAGTTAGGGGGACAAG GCAGACAGTGAAAAGGAA	58	250–298	15.077	0.571	0.874	0.011	0.901	0.003
AJ_SSR_221451	AVPY01221452.1	(CA)8	GGAAACAGCTATGACCATGGTTCCCTGACTCTCCTTACG GCATTTTATCTGGTTGCTGT	50	158–196	8.462	0.447	0.678	0.029	0.794	0.007*
AJ_SSR_38976	AVPY01038977.1	(TC)8	GGAAACAGCTATGACCATGCTACGGCCTACTCCGCTT GGGCATGTCCAACAACGA	55	220–238	5.462	0.268	0.434	0.026	0.479	0.012*
AJ_SSR_38976_1	AVPY01038977.1	(CA)7	GGAAACAGCTATGACCATGGAGGACACAAGGCATACA ATTCACAATTGCTCAGGG	55	212–246	6.769	0.231	0.564	0.017	0.604	0.009*
AJ_SSR_404232	AVPY01404233.1	(CA)18	GGAAACAGCTATGACCATGCCTTTGCCTCTTG GGGATCATCGGGATTTGTG	55	263–393	14.154	0.465	0.741	0.023	0.874	0.004*
<i>T</i> _a : optimized anneal was developed based <i>japonica</i> ; the numbe	ling temperature (°C); <i>R</i> 1 on the genes under sel r at the end of the locus	: allele size (b ection in <i>Angu</i> name stands f	p); N_a : number of alleles; H_0 : observed heterozygosity; H_E : expected heterozyguilla rostrata or Anguilla anguilla; "G" means the locus was developed from the or the contig number in draft genome of Anguilla japonica. * Significant deviati	gosity; SH e transcr tion from	3: standard pt sequence HWE after	error. The error. es randoml Bonferror	composition y chosen fre i correction	t of the locu om the trans in at least o	is name: "A scriptome d one populat	J" means at a base of $(\alpha=0.0)$	the locus Anguilla 5, K=13)

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Table 2 Continue	pa										
Loone	GenBank	Matif	Duitmon continued (5/ \21)	Ŀ	0	V	a	Н	_0	H	
rocas	accession No.	MOUL	$(c \leftarrow c)$ seduence	I_{a}	¥	Mean	SE	Mean	SE	Mean	SE
AJ_SSR_45576	AVPY01045577.1	(TG)8	GGAAACAGCTATGACCATGGGAGAAGTGAAGACCCGC GCCAACGACACAAAGCAA	57	187–311	10.769	0.303	0.667	0.028	0.754	0.011
AJ_SSR_45576_1	AVPY01045577.1	(AC)7	GGAAACAGCTATGACCATGGCCTCCGTTTCCTTTATA TGATGTGGGTCAGTTCTG	55	221–267	6.308	0.308	0.630	0.024	0.632	0.011
AJ_SSR_57531	AVPY01057532.1	(TG)6	GGAAACAGCTATGACCATGCCTTTCCCATCAAACCAG TTCCCACATAGGCAGCAC	60	394-420	6.923	0.265	0.447	0.024	0.561	0.018
AJ_SSR_57531_1	AVPY01057532.1	(AC)8	GGAAACAGCTATGACCATGGACCAGGGAAGATACTCAGCAC TGGAAGAAGCACACACAGTCA	63	327–373	4.846	0.274	0.550	0.033	0.594	0.012
AJ_SSR_63975	AVPY01063976.1	(AC)6	GGAAACAGCTATGACCATGGACAAAGACAGTCAGAAATACACGC CACTACCCAACACACACTTACAAAT CACTACCCAACACACAC	55	177–235	14.154	0.529	0.870	0.016	0.880	0.004*
AJ_SSR_82658	AVPY01082659.1	(CA)14	GGAAACAGCTATGACCATGGGAAAAACCTTCAGTGTAA CTGGCAAAAATCCTCTAT	55	216-422	29.231	1.156	0.888	0.023	0.945	0.003*
AJ_SSR_83825	AVPY01083826.1	(AC)11	GGAAACAGCTATGACCATGTGTGTGTGTGAGGACAGGGAGA TTTAGAGGGTTGAATAGGGT	50	357–389	11.462	0.418	0.854	0.018	0.863	0.003*
G_289450	AVPY01289451.1	(AT)10	GGAAACAGCTATGACCATGTACAAGATGCAGGAGATG CAGAAGCACTAACCAAAT	54	318–336	5.308	0.328	0.411	0.027	0.403	0.022
G_5145_1	AVPY01005146.1	(GT)11	GGAAACAGCTATGACCATGATGTTCTGTGGTAGCTGGCTC ACTCCAAGTTCATCCTTTCAG	58	197–245	12.154	0.541	0.727	0.023	0.837	0.004^{*}
G_5145	AVPY01005146.1	(AC)12	GGAAACAGCTATGACCAFGCTTCCCTACACTTTCACCC CATAATCCCTACACTGGCC	58	376-436	15.385	0.500	0.809	0.019	0.893	0.004
G_57494	AVPY01057495.1	(CT)15	GGAAACAGCTATGACCATGTTTCTCTTTTCACCCTCACTC CTTTTTCTTTCTGGCTTTTCAC	55	129–161	11.231	0.378	0.811	0.024	0.855	0.005*
G_445365	AVPY01445366.1	(GT)9	GGAAACAGCTATGACCATGTGTCCCTGACCAGCCAAC ACACCGACCTCAACGC	58	272–312	8.462	0.386	0.660	0.019	0.768	0.005*
G_168306	AVPY01168307.1	(TG)8	GGAAACAGCTATGACCATGGAACTGCTCAACGTGAAAAC CATTGTGTGTCATTGGAAGG	51	255-379	16.923	0.645	0.849	0.020	0.904	0.005*

weaken the power of neutral markers (Milano et al., 2014). At gene-associated markers putatively under selection, genetic differentiation could be larger than that at loci in which among-population differences are caused by genetic drift (Holsinger and Weir, 2009).

There are several studies showing the power of outlier loci in detecting population structure. Russello et al. (2012) successfully detected ecotype-level divergence using eight outlier microsatellite markers, whereas there was no evidence of divergence at neutral loci. By using outlier loci, Milano et al. (2014)



Fig.1 The scatter plot of possible number of clusters against ad hoc statistic ΔK

revealed a dramatic divergence between Atlantic and Mediterranean populations of European hake and fine-scale significant population structure among regional populations, which was not detected by using neutral markers. Another advantage of the markers in this study is their potential transferability across taxonomic boundaries. Compared with traditional microsatellite markers, microsatellite markers based on transcript sequences are potentially more transferable across taxonomic boundaries (Chagné et al., 2004; Liewlaksaneeyanawin et al., 2004; Gutierrez et al., 2005; Pashley et al., 2006).

For the first time, our study confirmed the panmixia of *A. japonica* using gene-associated microsatellite markers. We suggested that *A. japonica* should be managed as a single unit. Management and conservation efforts must be coordinated at the transnational level, as over-exploitation in any region will decrease recruitment across its whole distributional range.

4.2 Different pattern of spatially varying selection from North Atlantic eels

Both LOSITAN and SAM analysis detected no signals of spatially varying selection. These results indicated that the 24 loci were not targets of spatially varying selection and were nearly neutral genetic markers for Japanese eel. Since most of the



Fig.2 STRUCTURE result based on 24 gene-associated microsatellite markers for 13 geographic populations (K=6)



Fig.3 Outlier test result from LOSITAN software

microsatellite markers used are associated with the genes under selection in the North Atlantic eels, our results suggested that different selection pattern existed between Japanese eel and the North Atlantic eels. Furthermore, two recent studies found no overlap of loci under selection in the two North Atlantic eel species either. Gagnaire et al. (2012) scanned for footprints of selection in the American eel using a panel of 100 SNPs and 13 loci showed significant correlations between allele frequencies and environmental variables. Ulrik et al. (2014) found evidence for spatially varying selection at 10 loci in the European eel using the same SNP panel; however, none of these were the same loci that showed significant associations in the American eel. The contrasting results in the two species suggested that there were two different patterns of selection in the North Atlantic eels, at least at the level of the individual genes assessed (Hemmer-Hansen et al., 2014). Our results suggested that A. japonica might occupy a third selection pattern, which was different from both the North Atlantic eel species. Due to the latitudinal (rather than longitudinal) spread of freshwater habitats and the longer migration distance to the spawning site of Japanese eel (Chan et al., 1997), the selection pattern between Japanese eel and the North Atlantic eels could be different. Different migration patterns causes distinct selective pressures. With the advances of next-generation sequencing technology, genome wide SNPs have been adopted in population genomic studies of marine fishes. Pujolar et al. (2014) detected signatures of selection using

50 354 RAD-generated SNPs and found dozens of loci associated with the highly variable environmental conditions experienced by European eel along its geographic range. However, study of local adaptation in *A. japonica* by using genome wide genetic variations is still absent. A better understanding for genetic mechanism of local adaptation in *A. japonica* could be achieved when genome wide markers were applied in the near future.

 Table 3 Genetic diversity parameters for each sampled population across 24 microsatellite loci

Pop	Ν	r a	H	I _o	H	IE	F	, IS
	Mean	SE	Mean	SE	Mean	SE	Mean	SE
DD	11.125	1.146	0.710	0.033	0.769	0.034	0.070	0.023
FY	10.667	1.143	0.699	0.040	0.756	0.037	0.078	0.023
MA	11.042	1.084	0.675	0.038	0.746	0.039	0.089	0.027
MJ	12.167	1.278	0.679	0.037	0.758	0.036	0.104	0.021
MY	12.542	1.323	0.701	0.036	0.760	0.038	0.074	0.016
PD	11.958	1.300	0.688	0.039	0.752	0.038	0.081	0.026
QZ	10.833	1.259	0.715	0.038	0.762	0.034	0.066	0.025
RA	12.000	1.283	0.716	0.033	0.766	0.033	0.064	0.022
SS	12.000	1.317	0.685	0.034	0.753	0.039	0.077	0.025
ST	10.958	1.021	0.703	0.039	0.751	0.035	0.069	0.023
YH	11.000	0.917	0.688	0.034	0.761	0.034	0.090	0.028
ZS	10.417	1.034	0.690	0.039	0.753	0.039	0.077	0.031
MC	12.042	1.218	0.710	0.036	0.769	0.034	0.081	0.018

 $N_{\rm a}$: number of alleles; $H_{\rm O}$: observed heterozygosity; $H_{\rm E}$: expected heterozygosity; $F_{\rm IS}$: inbreeding coefficient; SE: standard error.

	DD	FY	MA	MJ	MY	PD	QZ	RA	SS	ST	YH	ZS	MC
DD													
FY	-0.001												
MA	0.001	0.002											
MJ	0.002	0.004	0.003										
MY	-0.001	0.000	-0.004	0.003									
PD	0.002	0.005	0.008*	0.005	0.004								
QZ	0.000	0.002	0.006	0.003	0.001	0.005							
RA	-0.004	0.001	0.002	-0.001	0.000	0.001	0.000						
SS	0.000	-0.001	0.002	0.000	-0.001	0.004	-0.001	0.000					
ST	0.000	0.004	0.001	-0.001	0.002	0.006	0.003	0.000	0.002				
YH	-0.004	0.003	-0.002	0.002	0.000	0.004	0.002	-0.003	-0.002	-0.002			
ZS	-0.003	0.000	0.003	0.000	0.000	0.005	-0.002	-0.001	-0.001	-0.002	0.000		
MC	-0.002	-0.001	0.003	0.003	-0.001	0.004	-0.001	0.000	-0.001	0.000	0.002	0.005	

Table 4 Pairwise F_{ST} values among 13 geographic samples

*P<0.05, ** significant after FDR correction (K=78, P<0.009 87).

5 CONCLUSION

For the first time, our study developed twenty-four gene-associated microsatellite markers and confirmed panmixia of A. japonica by using these microsatellites. The results suggested that A. japonica should be managed as a single unit. Management and conservation efforts must be coordinated at the transnational level, as over-exploitation in any region will decrease the recruitment of this species. Efforts were also made to detect signature of selection and no footprints of spatially varying selection were found, indicating that A. japonica might occupy a different pattern of selection from North Atlantic eels. The study on population genetic structure and local adaptation in A. japonica is far from complete. A whole-genome level study using population genomic approaches is needed in the future, which could help to elucidate the population structure and genetic mechanism underlying the spatially varying selection for geographic populations of A. japonica.

6 DATA AVAILABILITY STATEMENT

The datasets used and/or analyzed during the current study are available from the corresponding author on reasonable request.

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