



Minimum Travel Distance Among Publicly Insured Infants with Severe Congenital Heart Disease: Potential Impact of In-state Restrictions

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Abstract

Travel distance to surgical centers may be increased when coverage restrictions prevent children with congenital heart disease (CHD) from receiving care at out-of-state congenital heart surgery centers. We estimated the minimum travel distance to congenital heart surgery centers among publicly insured infants with time-sensitive CHD surgical needs, under two different scenarios: if they were and were not restricted to in-state centers. Using 2012 Medicaid Analytic eXtract data from 40 states, we identified 4598 infants with CHD that require surgery in the first year of life. We calculated the minimum travel distance between patients' homes and the nearest cardiac surgery center, assuming patients were and were not restricted to in-state centers. We used linear regression to identify demographic predictors of distance under both scenarios. When patients were not restricted to in-state centers, mean minimum travel distance was 43.7 miles, compared to 54.1 miles when they were restricted. For 5.9% of patients, the difference in travel distance under the two scenarios exceeded 50 miles. In six states, the difference in mean minimum travel distance exceeded 20 miles. Under both scenarios, distance was positively predicted by rural status, residence in middle-income zip codes, and white/non-Hispanic or American Indian/Alaskan Native race/ethnicity. For some publicly insured infants with severe CHD, facilitating the receipt of out-of-state care could mitigate access barriers. Existing efforts to regionalize care at fewer centers should be designed to avoid exacerbating access barriers among publicly insured CHD patients.

Keywords Congenital heart disease · Health services research · Ethics and policy

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Introduction

Congenital heart disease (CHD) represents the most common and resource-intensive class of birth defects in the US. Approximately 6 in every 1000 live-born children require cardiac heart surgery for CHD [1], and half of these patients are covered by public insurance programs such as Medicaid or the Children's Health Insurance Program, CHIP [2–9]. Given the number and vulnerability of publicly insured CHD patients who require surgery, minimizing barriers to accessing surgical care is an important clinical and policy goal [10, 11].

For many publicly insured CHD patients who require early surgery, such as infants with Tetralogy of Fallot, a potentially important access barrier may be the distances they must travel to congenital heart surgery centers. These patients may have limited financial resources to travel long distance for surgery, or to stay for extended periods in hotels during hospitalizations [12–14]. Additionally, travel distance may be high due to barriers to accessing out-of-state

care. By federal law, state Medicaid programs must pay for out-of-state care if there is an emergency, the service is not available in the state, or the standard of care in the area is to receive care across state lines. When these conditions are not met, some patients may have to travel to in-state centers even when an out-of-state center is much closer [15]. Even when a state Medicaid program is willing to pay for out-of-state care, patients may still be unable to access this care if the out-of-state provider is not enrolled in the Medicaid program of the patient's state [16–18]. For example, some providers do not enroll in out-of-state Medicaid programs when reimbursement rates are lower than that of their home state [19, 20].

A recent study described the distances that children with CHD patients traveled for surgery, and showed that approximately 25% of patients traveled greater than 100 miles [21]. This study was important and highly rigorous, but did not focus on the publicly insured, included patients with both time-sensitive and less urgent surgical needs, and did not assess the degree to which access to out-of-state surgical centers could decrease travel distance. To address these gaps, we used national 2012 Medicaid claims to estimate the magnitude and predictors of travel distance for care among publicly insured infants requiring surgery in the first year of life, assuming these patients were or were not restricted to in-state centers. We focused on this population because their time-sensitive surgical needs may make them particularly vulnerable to access barriers from high travel distance.

Methods

Primary Data Sources

Congenital heart surgical centers were identified using the 2012 directory of 119 centers released by *Congenital Cardiology Today* (Online Appendix 1) [22]. This directory lists providers and includes addresses at all centers in North America that perform open heart surgeries in children. We chose this as our source of provider supply, rather than surgeon-level data (such as the Society of Thoracic Surgeons-Congenital Heart Surgery Registry or the American Medical Association Physician Masterfile) because some congenital heart surgeons operate at multiple centers in different locations but submit data under a single tax identifier associated with one primary location.

Data on patient characteristics and diagnoses were obtained from the 2012 Medicaid Analytic eXtract (MAX) database from 40 “states” (39 states and the District of Columbia). Each of these 40 states contained a congenital heart surgical center. Seven states without a congenital heart surgical center (Alaska, Montana, New Hampshire, North Dakota, South Dakota, Vermont, Wyoming) were excluded

as patients in these states could not be “restricted” to in-state care. We also excluded patients living in Hawaii, given the questionable validity of using straight-line distance across a large body of water to reflect distance to out-of-state care. Finally, we lacked access to MAX data for three states at the time of analysis (Colorado, Idaho, Kansas). Based on 2010 data, these 39 states plus the District of Columbia collectively account for >95% of Medicaid births [23].

MAX data are research-friendly, publicly accessible versions of Medicaid claims, submitted by states to the Centers for Medicare and Medicaid Services (CMS) on annual bases. Data include claims from children enrolled in both fee-for-service and managed Medicaid, as well as those enrolled in the CHIP [24, 25]. MAX data are an important source of data for national studies of pediatric care [26], with more than 40% of children covered by public insurance at some point during each year [27]. Unlike procedure-based registries or state inpatient administrative databases that only include CHD patients who undergo surgery, MAX data can be used to identify all children with selected CHD diagnoses, regardless of whether they underwent surgery. We used data from 2012 as it is the most recent year for which MAX data are available from the majority of states.

Sample

Our sample included infants who were born during 2012 and who had one or more inpatient or outpatient claim during 2012 with one of the 11 ICD-9-CM diagnosis codes (Table 1). These codes corresponded to 10 CHD conditions that always or almost always require surgical repair within the first year of life. Conditions were not mutually exclusive (e.g., patients could have both Tetralogy of Fallot and pulmonary valve atresia).

Study Variables

The main outcome variable was minimum travel distance for care, defined as the straight-line distance (in miles) between zip code centroids of a patient's home and the nearest congenital heart surgical center. Distances between zip code centroids have been found to be a good proxy for driving distance and driving time [28–30]. Although our cohort only included patients from 40 states, we assumed patients could travel to centers in any state. Of note, our outcome does not represent the distance patients actually traveled for care, which cannot be reliably assessed in MAX data and may be affected by considerations such as a center's reputation or volume. Rather, our outcome represented a lower bound of travel distance for care. We calculated our outcome under two scenarios. First, we assumed that patients could travel to the nearest center even if it was out-of-state. Second, we

Table 1 Demographic characteristics and travel distance by demographic category

Independent variable	<i>n</i> (% of total)	Mean distance in miles (SD)	Median distance in miles (25–75th percentile)
Sex			
Female	2048 (44.5)	43.8 (49.3)	23.8 (7.5, 67.4)
Male	2537 (55.2)	48.0 (55.8)	26.7 (8.3, 75.9)
Missing	13 (0.3)	54.7 (52.7)	28.1 (14.8, 98.3)
Race/ethnicity			
White, non-Hispanic	1495 (45.9)	58.6 (56.4)	44.3 (16.4, 84.9)
Black, non-Hispanic	762 (23.3)	32.2 (39.1)	11.9 (4.7, 52.5)
Hispanic	800 (24.5)	37.4 (46.6)	15.8 (5.8, 52.6)
Asian or Pacific Islander	99 (3.0)	20.3 (27.5)	9.1 (6.8, 17.5)
American Indian or Alaskan Native	55 (1.7)	230.0 (374.7)	119.1 (46.6, 191.0)
Native Hawaiian or other Pacific Islander	15 (0.5)	12.5 (13.4)	8.1 (3.3, 16.6)
More than one race	35 (1.1)	76.1 (72.7)	56.1 (13.7, 107.1)
Missing	1337 (29.1)	41.8 (47.1)	20.6 (7.8, 65.2)
Age (months)			
0–2.9	959 (20.9)	47.4 (54.9)	25.7 (7.7, 72.0)
3–5.9	1137 (24.7)	47.5 (55.5)	27.6 (7.8, 74.0)
6–8.9	1170 (25.5)	45.5 (50.4)	25.3 (7.9, 69.6)
9–12	1332 (29.0)	44.8 (51.8)	24.6 (8.2, 68.9)
Mean-adjusted gross income (by zip code of residence)			
1 (\$19,212–35,202)	1126 (25.0)	38.7 (77.7)	8.9 (4.8, 61.9)
2 (\$35,203–43,602)	1125 (25.0)	62.4 (77.1)	39.8 (9.6, 91.0)
3 (\$43,603–54,069)	1125 (25.0)	64.2 (99.0)	41.2 (15.8, 89.2)
4 (\$54,070–593,608)	1125 (25.0)	47.9 (83.5)	22.0 (10.6, 60.5)
Missing	97 (2.1)	87.3 (114.1)	30.6 (63.2, 119.0)
Urban/rural status			
Urban	1634 (35.5)	14.3 (24.6)	6.9 (3.5, 12.8)
Rural–urban	2602 (56.6)	58.9 (52.6)	43.9 (18.9, 85.9)
Rural	363 (7.9)	94.3 (69.6)	75.2 (49.3, 119.3)
Diagnosis code (ICD-9-CM)^a			
Truncus arteriosus (common truncus, 745.0)	471 (10.1)	47.8 (58.4)	26.0 (7.5, 68.7)
Complete transposition of the great vessels (745.1, 745.10)	806 (17.3)	47.3 (50.3)	28.2 (9.4, 74.5)
Double outlet right ventricle (745.11)	778 (16.7)	42.9 (51.3)	21.7 (7.5, 63.9)
Tetralogy of Fallot (745.2)	1453 (31.1)	42.5 (51.5)	22.3 (7.5, 63.0)
Common ventricle (745.3)	707 (15.2)	48.8 (48.2)	30.3 (9.3, 79.2)
Atresia of pulmonary valve, congenital (746.01)	686 (14.7)	45.3 (56.7)	22.9 (7.2, 63.0)
Tricuspid atresia and stenosis, congenital (746.1)	413 (8.9)	44.8 (50.9)	20.9 (7.6, 66.6)
Hypoplastic left heart syndrome (746.7)	1077 (23.1)	43.9 (52.9)	21.2 (7.4, 63.7)
Interruption of aortic arch (747.11)	310 (6.7)	45.0 (55.6)	22.7 (6.8, 71.0)
Total anomalous pulmonary venous connection (747.41)	360 (7.7)	41.7 (44.3)	23.8 (8.4, 58.4)

^aRefers to the number of patients with ≥ 1 claim with the diagnosis code in 2012. The percentages in parentheses do not sum to 100% because some patients have more than one of these 11 diagnosis codes on claims in 2012

assumed that patients could not receive out-of-state care [13–18].

From the MAX data, we obtained information on patient age in months at the time of their first claim with one of the 11 CHD diagnosis codes, sex, zip code, and race/ethnicity. The MAX race/ethnicity variable indicates whether patients

are Hispanic or belong to the following non-Hispanic categories: white, black, Asian/Pacific Islander, American Indian/Alaskan Native, Native Hawaiian/Pacific Islander, more than one race, or unknown race.

From 2012 Internal Revenue Service data, we obtained mean-adjusted gross income by patient zip code of

residence [31]. From 2010 US Census Bureau data, we obtained the urban/rural designation of patient zip code of residence. The Census Bureau defines urban areas as zip codes with $\geq 50,000$ people, urban–rural areas as zip codes with 2500–49,999 people, and rural areas as zip codes with < 2500 people [32].

Statistical Analysis

For both scenarios of travel distance (restricting and not restricting patients to in-state centers), we used standard descriptive statistics to calculate the national mean and median of minimum travel distance. We also calculated travel distance by each demographic characteristic and diagnosis. For both scenarios, we also determined the proportion of patients with travel distance of 0–24.9, 25–49.9, 50–99.9, and ≥ 100 miles.

We calculated the difference in minimum travel distance for each patient assuming they were or were not restricted to in-state centers, then calculated the proportion of patients for whom this difference exceeded 50 miles (corresponding to roughly 1 h of additional travel time). We averaged this difference over all patients residing in each state to calculate the mean difference at the state level, and used a non-parametric test (Wilcoxon signed-rank test) to assess the significance of this difference for each state. We calculated the number of states for which the difference exceeded 20 miles (corresponding to roughly 1 standard deviation above the mean of the 39 state differences). To assess predictors of minimum travel distance under both scenarios, we used linear regression to model log-transformed travel distance as a function of patient age in months, sex, race/ethnicity, mean-adjusted gross income by zip code (by quartiles), and urban/rural status of zip code of residence. For ease of interpretation, we exponentiated coefficients from regressions to obtain the multiplicative changes in distances associated with changes in predictors.

Data on race/ethnicity, sex, and mean-adjusted gross income by zip code were missing for 29.1%, 0.3%, and 2.1% of the cohort, respectively. Missingness of race/ethnicity was strongly predicted by state, supporting the assumption that these data were missing at random. We imputed missing data using multiple imputation by chained equations with 20 imputations. Race/ethnicity, sex, mean-adjusted gross income by zip code, patient age, urban/rural status, and state were used in imputation models.

Analyses were performed using Stata 15 (College Station, TX). The Institutional Review Boards of the University of Chicago and Columbia University Irving Medical Center exempted this study from review.

Sensitivity Analyses

To assess the robustness of findings to changes in patient cohort selection, we repeated our analyses with the following modifications: (1) We only included infants with two or more instances of the 11 ICD-9-CM diagnosis codes, separated by a day or more ($n = 3266$) (2) we additionally included 1337 infants with CHD conditions that often, but do not always, require surgical repair in the first year of life, such as endocardial cushion defect (745.6, $n = 147$), atresia and stenosis of the aorta (747.22, $n = 172$), and pulmonary artery coarctation and atresia (747.31, $n = 1018$), and (3) we excluded infants with missing sex, race/ethnicity, or zip code income data ($n = 1420$).

We also calculated travel distance (assuming that patients could travel outside their state) when additionally including data from seven states that lacked a congenital heart surgical center (Online Appendix 2).

Results

The sample included 4598 infants. Of all infants, 21% were less than 3 months old at the time of identification, 45% less than 6 months, and 71% less than 9 months. The majority of patients lived in urban–rural zip codes (57%) or urban zip codes (35%). The three most common diagnoses were Tetralogy of Fallot (31%), hypoplastic left heart syndrome (23%), and transposition of the great vessels (17%, Table 1).

The 119 congenital heart surgical centers are listed in Online Appendix 1. Figure 1a and Table 2 display the number of patients in each state. Patient density generally correlated with changes in the population of each state. Figure 1b shows the mean distances by state, assuming that patients could cross state lines.

Distance

Assuming that patients could cross state lines, mean minimum travel distance was 43.7 (SD 46.3) miles, and median minimum travel distance was 24.7 (25–75th percentile 7.8–68.7) miles. The proportion of patients traveling 0–24.9, 25–49.9, 50–99.9, and 100+ miles was 50.2%, 15.9%, 20.8%, and 13.1%, respectively.

Assuming that patients could not cross state lines, mean minimum travel distance was 54.1 (SD 86.2) miles, and median minimum travel distance was 27.4 (25–75th percentile 8.2–79.3) miles. The proportion of patients traveling 0–24.9, 25–49.9, 50–99.9, and 100+ miles was 48.2%, 14.6%, 19.5%, and 17.7%, respectively.

The difference in minimum travel distance when patients were and were not restricted to in-state centers ranged from 0 to 621.5 miles (mean and median 9.2 and 0.0 miles). The

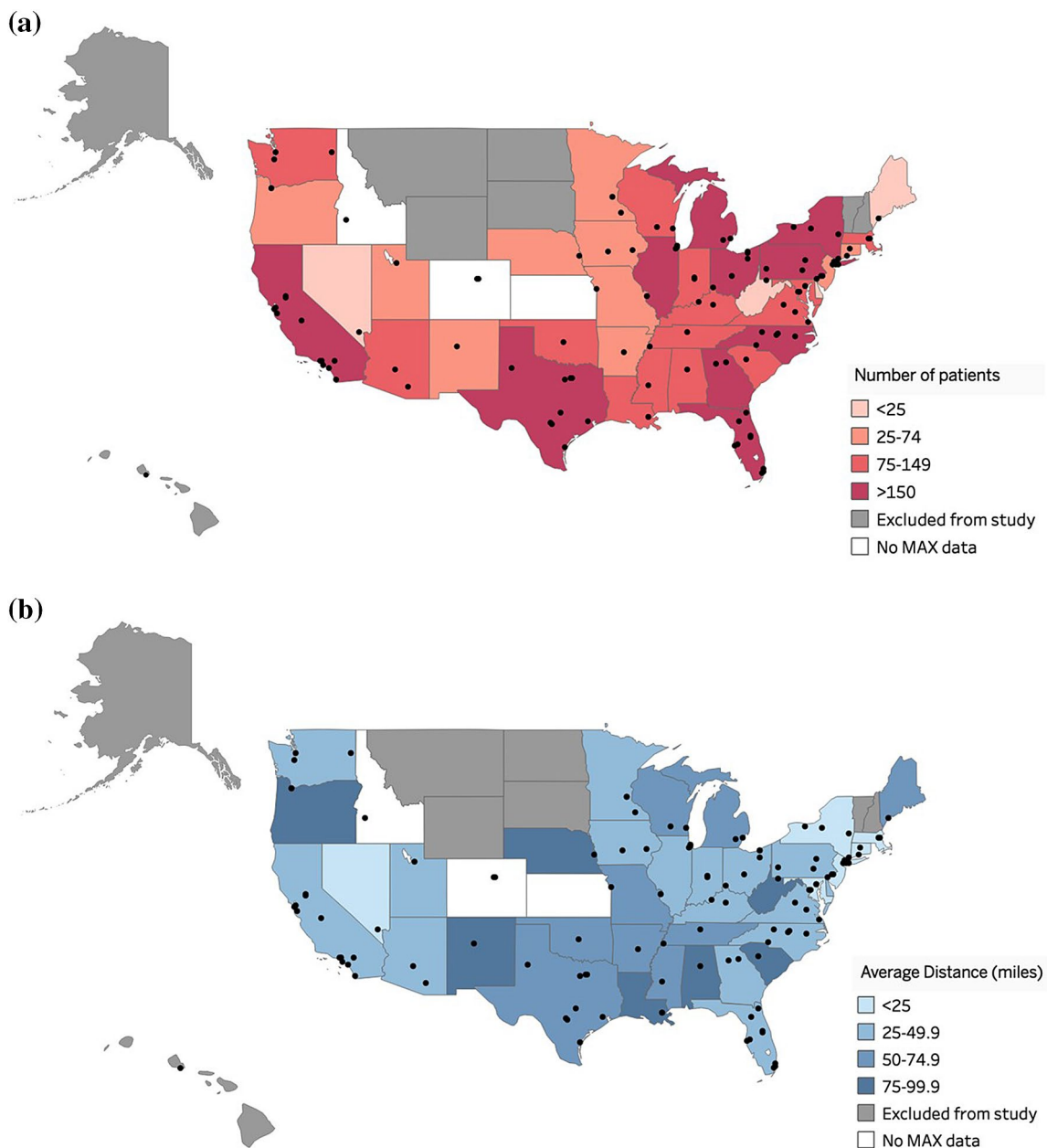


Fig. 1 **a** Number of patients by state. **b** Mean minimum travel distance by state when allowing patients to cross state lines

median was 0.0 because for 87.0% of patients, the in-state option was the closest option. For 5.9% of patients in the sample, the difference in minimum travel distance exceeded 50 miles. This proportion varied widely between states, ranging from 0.6% in New York to 25.3% in Illinois.

Among the 39 states, the mean difference in minimum travel distance when patients could and could not cross state lines ranged from 0.0 to 33.4 miles (mean and median 10.1 and 8.5 miles, Table 2). For six states, this difference exceeded 20 miles—Illinois, West Virginia, Indiana, Nevada, Tennessee, and Utah. The difference was

statistically significant for 26 states. Northeastern states demonstrated the smallest increase in travel distance. In a few states (e.g., Minnesota), the difference was zero, implying that for patients in these states, the nearest center would be the same regardless of whether there was an out-of-state restriction.

Patient Demographic Predictors of Distance

Assuming that patients could cross state lines, non-Hispanic black and Hispanic race/ethnicity was associated

Table 2 Minimum travel distance with and without an in-state travel restriction

State	Number of patients	Mean distance in miles (SD) allowing travel across state lines	Mean distance in miles (SD) with in-state travel restriction	Difference in mean distance (miles)	<i>p</i> value
Illinois	198	31.5 (34.5)	65.0 (84.6)	33.4	<0.01 [†]
Indiana	115	44.4 (32.4)	70.3 (47.8)	25.9	<0.01 [†]
Nevada	23	23.0 (48.1)	48.4 (110.6)	25.4	0.08
Tennessee	130	65.9 (48.4)	89.6 (109.1)	23.7	<0.01 [†]
Utah	45	46.7 (48.8)	67.5 (111.9)	20.8	0.08
Mississippi	83	66.6 (33.2)	86.4 (51.4)	19.8	<0.01 [†]
Louisiana	75	85.4 (56.0)	104.0 (82.6)	18.6	<0.01 [†]
Nebraska	31	80.9 (67.9)	99.4 (108.1)	18.5	0.05
New Jersey	74	16.9 (15.3)	33.5 (35.7)	16.6	<0.01 [†]
Iowa	35	38.9 (28.1)	53.1 (37.6)	14.1	0.07
Wisconsin	93	50.0 (52.7)	62.4 (73.0)	12.4	<0.01 [†]
Arkansas	66	67.6 (53.8)	79.5 (54.5)	11.8	<0.01 [†]
Florida	302	34.9 (42.1)	45.9 (109.9)	11.0	<0.01 [†]
Connecticut	27	13.9 (14.1)	24.8 (43.6)	10.9	<0.01 [†]
Kentucky	105	42.9 (39.1)	53.6 (50.4)	10.8	<0.01 [†]
Oklahoma	80	67.9 (53.2)	78.1 (88.9)	10.2	<0.01 [†]
Massachusetts	100	19.5 (16.4)	29.2 (29.2)	9.7	<0.01 [†]
Maryland	103	19.2 (22.8)	28.1 (29.9)	8.9	<0.01 [†]
Alabama	84	75.3 (46.3)	83.8 (60.3)	8.6	<0.01 [†]
Georgia	153	45.3 (38.5)	53.7 (55.0)	8.4	<0.01 [†]
Ohio	161	30.1 (24.9)	38.2 (51.2)	8.1	<0.01 [†]
Virginia	109	30.5 (27.6)	36.8 (43.8)	6.3	<0.01 [†]
California	304	34.3 (35.8)	40.6 (87.6)	6.3	0.08
South Carolina	99	75.8 (41.7)	81.3 (41.9)	5.5	<0.01 [†]
Michigan	290	60.6 (52.0)	65.4 (62.4)	4.8	<0.01 [†]
New Mexico	31	80.4 (70.7)	85.1 (75.9)	4.6	0.16
Texas	546	51.7 (56.0)	55.9 (75.6)	4.2	<0.01 [†]
New York	333	16.2 (23.5)	20.2 (61.9)	4.0	<0.01 [†]
Oregon	55	80.4 (71.1)	83.4 (75.1)	3.1	0.16
Arizona	125	35.9 (49.5)	38.8 (56.9)	3.0	<0.01 [†]
Washington	81	38.9 (43.7)	41.9 (45.0)	2.9	0.08
Pennsylvania	152	29.6 (25.1)	31.4 (28.2)	1.8	<0.01 [†]
North Carolina	197	37.6 (29.0)	39.4 (32.3)	1.7	<0.01 [†]
Missouri	74	63.2 (57.1)	64.0 (58.2)	0.8	0.05
Delaware	17	34.3 (25.0)	34.7 (25.8)	0.4	0.96
Maine	18	58.0 (59.2)	58.0 (59.2)	0.0	1.0
Minnesota	54	39.3 (55.0)	39.3 (55.0)	0.0	1.0

Due to restrictions in the Centers for Medicare and Medicaid Services data use agreement, data for two states with <10 patients are not displayed. One of these states had a difference in mean distance exceeding 20 miles; thus, there were a total of six such states

[†]Statistically significant

with 21% (CI 13–28%) and 16% (CI 7–22%) lower distance compared to white/non-Hispanic race/ethnicity, respectively. Residence in a rural area was associated with a 689% greater distance compared to residence in an urban area (CI 593–799%). Residence in zip codes in the second and third highest quartile of mean-adjusted gross income was associated with a 30% (CI 19–43%) and 31%

(19–44%) increase in distance compared to the poorest quartile, respectively. Sex and age were not significant predictors of distance (Table 3). When assuming that patients were restricted to in-state centers, demographic predictors of distance were similar.

Table 3 Demographic predictors of travel distance, when restricting to in-state centers

Independent variable	Adjusted ratio ^a	95% CI
Sex		
Female	Reference	Reference
Male	1.04	0.97–1.10
Race/ethnicity		
White, non-Hispanic	Reference	Reference
Black, non-Hispanic	0.79 [†]	0.72–0.87
Hispanic	0.86 [†]	0.78–0.93
Asian or Pacific Islander	0.77 [†]	0.60–0.98
American Indian or Alaskan Native	1.19	0.89–1.59
Native Hawaiian or other Pacific Islander	0.53	0.26–1.07
More than one race	1.01	0.72–1.43
Age (months)	1.00	0.99–1.01
Mean-adjusted gross income quartile, by zip code of residence		
1 (\$19,212–35,202)	Reference	Reference
2 (\$35,203–43,602)	1.30 [†]	1.19–1.43
3 (\$43,603–54,069)	1.31 [†]	1.19–1.44
4 (> \$54,069)	1.06	0.97–1.17
Urban/rural status		
Urban	Reference	Reference
Urban–rural	4.54 [†]	4.23–4.88
Rural	7.89 [†]	6.93–8.99

[†]Statistically significant

^aCoefficients from regressions were exponentiated and are interpreted as ratios of distance. For example, black patients have a mean travel distance that is 0.79 times as high as white patients, holding all other factors constant

Sensitivity Analyses

Conclusions were not substantively different when excluding patients with only one instance of an ICD-9-CM code, when including infants with CHD that do not always require surgical repair in the first year of life, and when excluding infants with missing data (Online Appendices 3, 4, and 5).

Discussion

Using national data from 39 states and the District of Columbia, we assessed the minimum travel distance to CHD centers among publicly insured infants with CHD conditions requiring early surgery, assuming they were or were not restricted to in-state centers. The difference in minimum travel distance under these scenarios exceeded 50 miles for 5.9% of patients. This proportion varied widely between states, ranging from 0.6% in New York to 25.3% in Illinois. White/non-Hispanic or American Indian/Native race/ethnicity, with residence in rural areas, and with residence in

middle-income zip codes were each associated with higher minimum travel distance. These findings suggest that many publicly insured infants with time-sensitive surgical needs must travel long distances for care, particularly infants in certain demographic and geographic groups.

This study used national data to estimate minimum travel distance in a cohort of CHD patients who may be particularly vulnerable to access barriers due to high travel distances. Previous studies examining travel distance to congenital heart surgical centers have focused on single institutions or have included adult CHD patients, patients without CHD, patients with CHD that does not require surgery, or privately insured patients [19, 33–35]. We also report for the first time that residence in a rural area was positively associated with travel distance, and that racial/ethnic minority status other than Native American was associated with lower travel distance compared to white race/ethnicity, likely due to the concentration of racial/ethnic minorities in urban areas.

Our findings have implications for the current debate about regionalization of CHD surgical care [36–42]. Studies have demonstrated up to sevenfold variation in risk-adjusted outcomes between congenital heart surgery centers [43–48]. Consequently, many advocate establishing “centers of excellence” for CHD care, which could divert patients to high-volume centers, minimize practice variation, and potentially improve outcomes and decrease costs [49–52]. Such a model assumes that patients can travel long distances and bypass their nearest surgical center. While studies demonstrate that some patients do bypass the nearest center [19, 33, 35], these studies included privately and publicly insured patients with both elective and emergent surgical needs. Our results demonstrate that certain demographic groups of publicly insured infants with time-sensitive surgical needs already live far from care. Closing centers under a regionalization model might exacerbate access barriers for these patients, depending on whether these centers disproportionately serve these patients and the availability of alternative nearby centers. Future research is needed to assess whether the benefits of regionalization may outweigh any increased access barriers from longer travel distances, using approaches such as simulation exercises [40], or cost-effectiveness analysis.

Our study is the first to quantify the degree to which minimum travel distance would differ if publicly insured infants with CHD were or were not always restricted to in-state centers. The magnitude of the difference in distance under these two scenarios varied widely. Our findings suggest that facilitating the receipt of out-of-state care may be an important step toward alleviating current access barriers, particularly for patients in certain states. Of note, facilitating such care would also be crucial if CHD surgical care continues to be regionalized. Because regionalization could decrease the number of available surgical centers, fewer states might have in-state options. Thus, policies that guarantee out-of-state

Medicaid coverage, such as automatically enrolling providers in the Medicaid programs of nearby states, may be essential to minimize access barriers under a regionalization model [17–19].

Our study has some limitations. First, we calculated the minimum distances patients must travel for care, which may be lower than the distance they actually travel and which therefore may underestimate access barriers. There might be practical reasons why some patients elect to bypass their nearest center or be referred elsewhere. Some patients may select centers based on reputation or outcomes [53], prior relationships with providers, or established referral practices between their pediatric cardiologist (or pediatrician or insurance provider) and specific surgeons. Likewise, some patients might have family or other supports closer to other centers.

Second, we lacked information on the center at which patients ultimately received care. Therefore, we could not assess whether in-state restrictions might have impeded access to higher-quality out-of-state centers. Third, we did not assess whether in-state restrictions might worsen outcomes by increasing travel distance. MAX data lack detailed outcomes data (e.g., morbidity, quality of life), and the date-of-death fields are not always complete. Furthermore, there are many potential confounders of the relationship between outcomes and travel distance that cannot be controlled for using these data.

Finally, we used 2012 data. Currently, 2012 is the latest year in which MAX data from almost every state are available. Future studies should repeat our analyses using more recent data as they are released. Finally, MAX data submitted by comprehensive managed care organizations may not perfectly capture diagnoses, though the quality of these data is improving [54]. In sensitivity analyses, findings were not affected substantially by changes to our cohort. Finally, our study focused on inpatient surgical care and did not examine post-operative outpatient care that may be required throughout a patient's lifetime.

Conclusion

Among publicly insured infants with CHD who require early surgery, many live far away from surgical centers that can provide definitive care, with some demographic and geographic groups at a particular disadvantage. For many of these infants, facilitating the receipt of out-of-state care could alleviate access barriers. Efforts to regionalize CHD surgical care at fewer centers should be designed to avoid exacerbating access barriers for patients, particularly those who already live far from care.

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Compliance with Ethical Standards

Conflict of interest All authors declare that they have no conflict of interest.

Ethical Approval This article does not contain any studies with human participants or animals performed by any of the authors.

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