



# Association between adherence to antihypertensive medications and health outcomes in middle and older aged community dwelling adults; results from the Irish longitudinal study on ageing

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## Abstract

**Purpose** To examine the association between antihypertensive medication (AHTM) implementation adherence and healthcare utilisation in community-dwelling adults aged  $\geq 50$  years in Ireland.

**Methods** This was a prospective cohort study. The Irish Longitudinal Study on Ageing (TILDA) was linked to pharmacy claims data for participants aged  $\geq 50$  years. Participants were included if they had  $\geq 3$  pharmacy claims for one or more AHTM (ATC codes 'C02', 'C03', 'C07', 'C08' or 'C09') within the year preceding the year of self-reported healthcare utilisation outcome occurrence. Outcomes included self-reported general practitioner (GP), emergency department (ED), outpatient department visits and hospital admissions. Implementation adherence was measured using proportion of days covered (PDC), with participants classified as adherent if the average PDC  $\geq 0.8$ . Negative binomial models were used to analyse the association between AHTM adherence and number of GP, ED, outpatient visits and hospitalisations (adjusted IRR and 95% CI are presented).

**Results** One thousand four hundred thirty-one participants were included. The majority of participants (72.6%) were considered adherent. Good implementation adherence to AHTM was associated with a significant decrease in self-reported GP visits (adjusted IRR 0.91, 95% CI 0.83–0.99). Adherence had no significant impact on the number of ED visits, outpatient visits or hospitalisations reported by TILDA participants.

**Conclusions** Good adherence to AHTM was associated with less self-reported GP visits in this population, suggesting improved overall health status. However, the impact of medication non-adherence on the other self-reported healthcare utilisation outcomes (ED, outpatient visits and hospitalisations) was not evident in this study.

**Keywords** Medication adherence · Hospitalisation · Health outcomes · Older

## Introduction

The ABC taxonomy defines adherence to medicines as a process whereby patients take their medication as prescribed and

consists of three core components: initiation, implementation and discontinuation [1]. Initiation refers to the patient taking the first dose as prescribed; implementation describes the fidelity the patient exhibits to the agreed therapeutic regimen, and discontinuation refers to when a patient stops taking their medicine [1].

Medication adherence ranges from 47 to 100% in older populations [2]. Medication non-adherence may represent a greater risk in older people, due to increasing drug burden and comorbidities [3, 4] resulting in poorer health outcomes for this cohort [4, 5]. Medication non-adherence in older populations is multi-faceted, compounded by a combination of drug-related factors such as dosing regimen, side effects and polypharmacy, patient-related factors such as cognitive function, health literacy and multi-morbidity, and the patient-physician relationship [6].

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Non-adherence to antihypertensive medication (AHTM) is common, with some studies suggesting as many as 45% of patients are non-adherent [7], possibly due to the asymptomatic nature of the condition [4] and leading to an increased risk of uncontrolled blood pressure [7–9]. Hypertension is considered the strongest modifier of cardiovascular disease risk, with intensive treatment demonstrating superior benefit to standard antihypertensive therapy in adults aged  $\geq 75$  years [10]. Previous observational database studies have reported the association between AHTM non-adherence and cardiovascular outcomes such as stroke [11–14], congestive heart failure [13, 15], acute myocardial infarction [11, 12], cerebrovascular diseases [16] and all-cause mortality [11]. All studies reported an inverse relationship between good AHTM adherence and the risk of cardiovascular endpoints [11–16]. Studies conducted in the USA have demonstrated a significant relationship between cardiovascular medication adherence and all-cause hospitalisation [17, 18] and emergency department visits [18–20], with variability emerging when analysing therapeutic classes separately [17, 18]. There is limited evidence on the association between AHTM adherence and healthcare utilisation in middle and older aged adults within a European setting. There is uncertainty over the threshold for AHTM adherence required to achieve therapeutic success, with most studies opting for an arbitrary cutoff of 80% [21].

The aims of this study are to establish:

- (i) the level of AHTM implementation adherence measured using pharmacy refill claims data,
- (ii) the association between AHTM implementation adherence and healthcare utilisation outcomes and
- (iii) the impact altering the implementation adherence threshold has on the association between implementation adherence and healthcare utilisation outcomes

in community-dwelling adults aged  $\geq 50$  years in Ireland, who received medications as part of a state-subsidised scheme across a 12-month period.

## Methods

### Study design

This was a prospective cohort study using linked data from a national pharmacy claims database, the Health Service Executive-Primary Care Reimbursement Service (HSE-PCRS) and the first wave of The Irish Longitudinal study on Ageing (TILDA) [22, 23].

TILDA is a nationally representative cohort of Irish community-dwelling individuals aged  $\geq 50$  years and records information relating to the health, economic and social

circumstances of this cohort as they progress through the ageing process. The first wave of data collection began in October 2009 through to February 2011, resulting in a sample size of 8504 (participants and proxies). There were 8176 participants aged  $\geq 50$  years. The sampling framework is described in detail elsewhere [22]. Participants complete a computer-aided personal interview (CAPI) as part of the study.

The study population included adults aged  $\geq 50$  years (at time of CAPI), who had participated in wave 1 of TILDA and have a general medical services (GMS) card. Data regarding prescription dispensing claims from GMS patients is collected from the (HSE-PCRS) [24]. The HSE-PCRS GMS scheme provides free health services and prescribed medications to eligible persons in Ireland. At the time of prescription data collection, there was no medication co-payment associated with the scheme. Eligibility of the GMS scheme is based on income-related means-testing, with the exception of those aged  $\geq 70$  years, who were automatically eligible for the scheme until January 2009 [24]. Eligible participants were those who received  $\geq 3$  pharmacy claims for an AHTM within the 12 months preceding the time period referred to in the CAPI interview in wave 1 were included (see Fig. 1) in the study. AHTM were identified from the World Health Organization Anatomical Therapeutic Chemical (ATC) codes, as those with second level codes C02 (anti-hypertensives), C03 (diuretics), C07 (beta blockers), C08 (calcium channel blockers) and C09 (agents acting on the renin-angiotensin system), in line with previous studies [11]. Participants were both incident and prevalent users of AHTM.

The EMERGE (ESPACOMP Medication Adherence Reporting Guideline) guideline was followed in the reporting of this study [25].

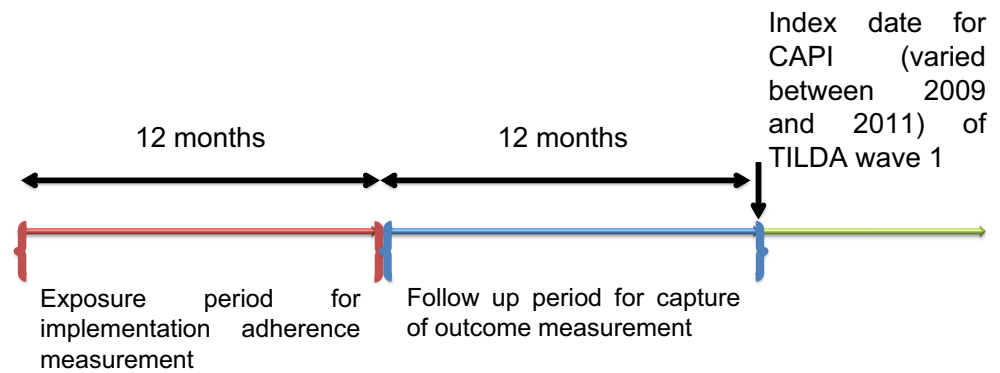
### Exposure variables

Implementation adherence to AHTM was calculated using the proportion of days covered (PDC), which has been validated for providing accurate estimates of adherence using administrative claims databases [26]. The PDC represents the sum of days' supply of each prescription refill, divided by the days in the observation period, taking into consideration overlapping supplies.

Implementation adherence was measured as a dichotomous variable. A patient with a PDC value  $\geq 0.8$  was considered adherent, which is the conventionally used cutoff for adherence [27]. If patients had been taking more than one class of AHTM during the study period, the PDC used was an average of all values across the different classes.

Implementation adherence was calculated for the year preceding the time period (12 months) for which the respondent is referring to in the CAPI, as per the methodology used in a previous study (see Fig. 1). [28]

**Fig. 1** Exposure period for adherence measurement and follow-up period for outcome measurement



## Outcome variables

The main outcome(s) was self-reported healthcare utilisation in the 12 months prior to the date of participants CAPI interview including:

1. General practitioner (GP) visits
2. Emergency department (ED) visits
3. Outpatient department visits
4. Number of occasions admitted to hospital overnight (hospitalisations)

The responses were reported on a continuous scale (up to a maximum 200 for GP, ED and outpatient visits and up to a maximum of 50 for hospitalisations). Dichotomous variables were generated for each healthcare utilisation outcome e.g. one or more GP visits, ED visits, outpatient visits or hospitalisations based on the distribution of the data, respectively, in the year prior to the CAPI interview (see Fig. 1).

## Covariates

Sociodemographic and clinical covariates identified as influencing both medication adherence and healthcare utilisation were identified from the literature [4]. Covariates included (i) age group, (ii) gender, (iii) polypharmacy ( $\geq 5$  medications), (iv) level of educational attainment, (v) memory score, (vi) functional disability, (vii) depressive symptoms and (viii) social support. Covariates were based on responses in the CAPI.

## Data analysis

Descriptive statistics including means, medians and variance were calculated for implementation adherence, healthcare utilisation outcomes and covariates. Comparisons between adherent or non-adherent participants for categorical covariates were analysed using chi-squared tests, and Wilcoxon rank sum tests were used for continuous covariates.

Regression models were used to calculate the unadjusted and adjusted association between adherence and healthcare

utilisation, including both dichotomous (logistic regression) and count outcomes (negative binomial). For count outcomes with an excessive number of zeros, standard negative binomial regression models and zero inflation negative binomial models were constructed and model fit was compared using the Vuong test [29] and criteria for assessing goodness to fit (AIC and BIC).

Sensitivity analyses were conducted by altering the PDC cutoff for adherence to  $PDC \geq 0.7$  and  $PDC \geq 0.9$ , respectively, to analyse the impact of different adherence levels on healthcare utilisation outcomes. Analyses were performed using SAS version 9.4 (SAS Institute, Inc., Cary, NC).

## Results

Eight thousand one hundred seventy-six participants aged  $\geq 50$  years were identified from wave 1 of TILDA. Three thousand one hundred seventy-six participants had consented to have their GMS pharmacy claims data linked to their TILDA data, and 1431 (45%) participants had received  $\geq 3$  pharmacy claims for an antihypertensive medication within the previous 12 months. Over half of participants were women (55%). The average age of participants was 72 years, and 45% of participants were aged  $\geq 75$  years. The average PDC was 0.84 (SD 0.21), and 73% had a  $PDC \geq 0.80$  (adherent). There was no significant gender difference in whether participants were considered adherent or non-adherent. However, there was a significant age difference, with a higher proportion of those aged  $\geq 75$  years considered adherent, compared with the younger cohorts (see Table 1).

## Outcome data

Table 2 presents the univariate (unadjusted) and multivariate (adjusted) incident rate ratios (IRR) and 95% confidence intervals for the association between implementation adherence ( $PDC \geq 0.80$ ) and healthcare utilisation. Supplementary Table 1 presents the results of logistic regression models (odds ratios and 95% CI).

**Table 1** Characteristics of participants based on implementation adherence

	Total, <i>n</i> (%)	Non-adherent (PDC ≤ 0.80), <i>n</i> (%)	Adherent (PDC ≥ 0.8) <i>n</i> , (%)	<i>P</i> value
All participants	1431	392 (27%)	1039 (73%)	
Gender				
Female	786 (55%)	210(27%)	576 (73%)	0.53
Male	645(46%)	182 (28%)	463 (72%)	
Age				
Participants age (median years, IQR)	74 (67–79)	72(64–78)	74(68–79)	0.08
50–64 years	287 (20%)	103 (36%)	184 (64%)	0.0004*
65–74 years	496 (35%)	137 (28%)	359 (72%)	
≥ 75 years	648 (45%)	152 (23%)	496 (77%)	
Functional disability				
Functional disability <sup>a</sup>	339 (24%)	93 (27%)	246 (73%)	0.98
No functional disability <sup>a</sup>	1092 (76%)	299 (27%)	793 (73%)	
Memory score				
Immediate recall score (median score, IQR) <sup>b</sup>	12 (9–14)	12(10–14)	12 (9–14)	0.70
Delayed recall score (median score, IQR) <sup>c</sup>	5 (3–6)	5 (3–7)	5(3–6)	0.11
Highest education level attained				
Primary level of education/no formal education	745 (52%)	192 (26%)	553 (74%)	0.13
Secondary level of education	470 (33%)	129 (27%)	341 (73%)	
Third or higher level of education	214 (15%)	70 (33%)	144 (67%)	
Depressive symptoms				
No depressive symptoms <sup>d</sup>	975 (69%)	272(28%)	703 (72%)	0.77
Subclinical depression <sup>d</sup>	284 (20%)	73 (26%)	211 (74%)	
Clinical depression <sup>d</sup>	149(11%)	41 (28%)	108 (72%)	
Social support				
Lives with others	950(66%)	271 (29%)	679 (71%)	0.18
Lives alone	481(34%)	121(25%)	360 (75%)	
Polypharmacy				
Taking ≥ 5 medications	750(52%)	197 (26%)	553 (74%)	0.27
Taking < 5 medications	668(47%)	193 (29%)	475(71%)	

\*  $p < 0.05$ <sup>a</sup> Any IADL (instrumental activities of daily living) or ADL (activities of daily living) disability<sup>b</sup> Participants complete two learning trials comprising each of a 10-word list learning task. The results of each of these tasks are added together to give an overall sum of immediate recall. Higher scores indicate better immediate memory, up to a maximum score of 20. Missing values ( $n = 4$ )<sup>c</sup> The delayed recall task is the same as the 10-word list learning task used to measure immediate recall but is carried out at a later point in the CAPI. Higher scores indicate better-delayed memory, up to a maximum score of 10. Missing values ( $n = 48$ )<sup>d</sup> Depression severity classified using the Centre for Epidemiologic Depression Scale [30]. Reference category was no depressive symptoms (score 0–7). Sub-clinical depression had a score 8–15, and clinical depression was scored > 15. Missing values ( $n = 23$ )

SD Standard deviation

### General practitioner visits

Ninety-eight percent of participants reported visiting the GP at least once in the 12 months subsequent to adherence measurement ( $n = 1404$ ). The mean number of GP visits reported for the previous 12 months was 4 (SD 6.58).

A negative binomial regression analysis demonstrated that adherence was associated with a significant decrease in the number of GP visits in the following year (9% versus non-adherence). This relationship remained significant after adjustment for covariates. Participants who reported being on ≥ 5 medications

had significantly more GP visits in comparison to those who were on fewer medications, as did those who reported having a functional disability or depression. Being older, female and living with others was associated with fewer GP visits.

### Emergency department visits

One thousand four hundred twenty-eight participants reported the number of ED visits they had in the year preceding the CAPI. Twenty percent of participants reported having ≥ 1 ED visit.

**Table 2** Univariate and multivariate analysis of the association between implementation adherence and covariates and the rate of healthcare utilisation outcomes

	GP visits			ED visits <sup>i</sup>			Outpatient visits <sup>j</sup>			Hospitalisations						
	Univariate model		Multivariate model	Univariate model		Multivariate model	Univariate model		Multivariate model	Univariate model		Multivariate model				
	IRR	95% CI	IRR	95% CI	IRR	95% CI	IRR	95% CI	IRR	95% CI	IRR	95% CI				
Adherent vs non-adherent	0.91	0.83–1.00*	0.91	0.83–0.99*	0.87	0.59–1.29	0.88	0.59–1.32	0.96	0.77–1.18	1.09	0.88–1.35	0.92	0.67–1.26	0.83	0.60–1.14
Female vs male	0.91	0.84–0.99*	0.89	0.82–0.96*	1.04	0.72–1.49	0.95	0.65–1.39	0.79	0.65–0.96*	0.85	0.71–1.03	0.86	0.65–1.14	0.83	0.62–1.11
Age: 65–74 years <sup>a</sup>	0.74	0.67–0.83*	0.73	0.65–0.81*	0.71	0.44–1.15	0.61	0.37–1.00	0.74	0.57–0.96*	0.79	0.60–1.03	1.15	0.77–1.72	1.16	0.77–1.76
Age: ≥75 years <sup>a</sup>	0.81	0.73–0.90*	0.76	0.68–0.84*	0.50	0.32–0.79*	0.49	0.30–0.81*	0.68	0.53–0.87*	0.62	0.47–0.80	1.14	0.78–1.67	1.12	0.74–1.68
Polyparmacy <sup>b</sup>	1.38	1.27–1.49*	1.24	1.15–1.35*	0.95	0.65–1.39	0.89	0.57–1.40	1.81	1.46–2.24*	1.84	1.45–2.34*	2.15	1.61–2.85*	1.76	1.30–2.38*
Secondary level education <sup>c</sup>	0.92	0.84–1.00	0.96	0.88–1.05	0.82	0.54–1.24	0.84	0.55–1.28	1.02	0.83–1.26	1.11	0.89–1.38	1.09	0.80–1.50	1.17	0.83–1.64
Third level or higher education <sup>c</sup>	0.87	0.77–0.97*	0.89	0.79–1.00	1.22	0.76–1.95	1.04	0.63–1.71	1.36	1.04–1.78*	1.54	1.16–2.03*	1.48	1.00–2.20	1.82	1.21–2.76*
Functional disability <sup>d</sup>	1.46	1.34–1.60*	1.32	1.20–1.46*	1.29	0.86–1.90	1.49	0.98–2.25	1.60	1.30–1.97*	1.48	1.19–1.84*	2.06	1.52–2.80*	1.58	1.12–2.23*
Subclinical depression <sup>e</sup>	1.13	1.03–1.25*	1.11	1.00–1.23*	0.57	0.35–0.92*	0.64	0.37–1.13	1.09	0.86–1.37	1.07	0.85–1.35	2.25	1.62–3.13*	2.11	1.49–2.97*
Clinical depression <sup>e</sup>	1.43	1.26–1.62	1.17	1.03–1.37*	0.90	0.56–1.44	0.80	0.47–1.34	1.65	1.23–2.19	1.49	1.00–2.00	1.99	1.29–3.06*	1.54	0.95–2.49
Immediate recall score (memory) <sup>f</sup>	0.98	0.97–1.00*	0.99	0.97–1.00	1.01	0.96–1.06	1.00	0.93–1.08	0.98	0.95–1.00	0.98	0.95–1.00	0.99	0.95–1.03	1.02	0.96–1.08
Delayed recall score <sup>g</sup>	0.98	0.97–1.00*	1.00	0.97–1.02	1.00	0.92–1.09	0.98	0.87–1.11	1.02	0.98–1.06	NA <sup>j</sup>	NA <sup>j</sup>	0.96	0.90–1.03	0.95	0.87–1.04
Social support <sup>h</sup>	0.94	0.87–1.03	0.91	0.83–0.99*	0.96	0.66–1.39	0.75	0.50–1.12	1.16	0.95–1.42	0.93	0.76–1.15	0.80	0.60–1.07	0.80	0.59–1.09

\*  $p < 0.05$

<sup>a</sup> Reference age category 50–64 years

<sup>b</sup> Participants who self-reported being on  $\geq 5$  medications

<sup>c</sup> Reference primary level education or no formal education

<sup>d</sup> Any IADL (instrumental activities of daily living) or ADL (activities of daily living) disability

<sup>e</sup> Depression severity classified using the Centre for Epidemiologic Depression Scale. [30] Reference category was no depressive symptoms (score 0–7). Sub-clinical depression had a score 8–15 and clinical depression was scored  $> 15$

<sup>f</sup> Participants complete two learning trials comprising each of a 10-word list learning task. The results of each of these tasks are added together to give an overall sum of immediate recall. Higher scores indicate better immediate memory, up to a maximum score of 20

<sup>g</sup> The delayed recall task is the same as the 10 word list learning task used to measure immediate recall but is carried out at a later point in the CAPI. Higher scores indicate better delayed memory, up to a maximum score of 10

<sup>h</sup> Either lives with spouse or other people (vs living alone)

<sup>i</sup> Indicates that zero-inflation negative binomial regression models were used

<sup>j</sup> In the multivariate zero inflation negative binomial model, delayed recall was removed as a covariate, as the algorithm did not converge when it was included

Multivariate GP model ( $n = 1346$ ), multivariate ED model ( $n = 1347$ ), multivariate outpatient visits model ( $n = 1388$ ), multivariate hospitalisation model ( $n = 1348$ );  $n$  is smaller due to missing values across the covariates

A zero inflation negative binomial regression model showed that adherence was not significantly associated with the number of ED visits, either in the unadjusted or adjusted analysis (Table 2). Participants aged  $\geq 75$  years had significantly fewer ED visits than participants aged 50–64 years in adjusted analysis. Adherence was not significantly associated with the likelihood of having an ED visit or not (Supplementary Table 1).

### Outpatient visits

One thousand four hundred twenty-eight participants reported the number of outpatient visits they had in the year preceding CAPI measurement. Fifty-four percent of participants reported having one or more outpatient visits. Adherence to AHTM was not a significant predictor of the number of outpatient visits, (Table 2) but was significantly associated with a reduced likelihood of experiencing an outpatient visit or not, in comparison to non-adherence (Supplementary Table 1).

Polypharmacy, higher level of education and functional disability were significant independent predictors of a higher rate of outpatient visits.

### Hospitalisations

One thousand four hundred thirty participants responded to the CAPI question regarding the number of hospital overnight admissions they had experienced in the past year. Twenty percent of participants reported having at least one hospitalisation in the past year.

Implementation adherence to AHTM was not significantly associated with the number of hospital admissions (Table 2). Being on more than 5 medications, having a higher level of education, having a functional disability and subclinical depression were all significantly associated with a greater number of hospitalisations. Implementation adherence was not significantly associated with the risk of having a hospital admission, either in the univariate or adjusted logistic model (Supplementary Table 1).

## Sensitivity analyses

### Adherence (PDC $\geq 0.7$ )

Similar to the traditional cutoff point, having a PDC  $\geq 0.7$  was not significantly associated with either the number of ED visits or the likelihood of having an ED visit (Supplementary Table 2). Implementation adherence (PDC  $\geq 0.7$ ) was associated with a similar rate of GP visits as the standardised PDC  $\geq 0.80$  cutoff point (Supplementary Table 3). Adherence was associated with a significantly lower

likelihood of experiencing  $\geq 1$  outpatient visit (adjusted OR 0.63, 0.47–0.85), but not with the number of outpatient visits reported. Having a PDC  $\geq 0.70$  was also significantly associated with a lower risk of hospitalisation (adjusted OR 0.65, 95% CI 0.47–0.91) and less overnight admissions (adjusted IRR 0.64, 95% CI 0.45–0.91).

### Adherence (PDC $\geq 0.9$ )

Adherence was not significantly associated with the number of ED visits experienced or the likelihood of experiencing  $\geq 1$  ED admission when set at PDC  $\geq 0.90$  (Supplementary Table 4). Adherence at the higher cutoff of PDC  $\geq 0.90$  was associated with a slightly lower rate of GP visits than the standardised cutoff (Supplementary Table 5). No significant relationship was observed between either the likelihood of experiencing an outpatient visit or the number of outpatient visits reported. Adherence at PDC  $\geq 0.9$  was not significantly associated with a lower risk of hospitalisation or number of admissions.

## Discussion

Implementation adherence to AHTM was high at 73% in this community-based population aged  $\geq 50$  years and was found to be associated with fewer GP visits and less likelihood of having an outpatient visit. Implementation AHTM adherence (PDC  $\geq 0.8$ ) was not significantly associated with a lower number of ED visits or hospitalisations in this cohort.

This study estimates that over a quarter of middle and older aged adults who received AHTM as part of a state-subsidised scheme are non-adherent to their AHTM. This is similar to previous analyses using pharmacy refill claims data [28, 31] and self-report questionnaires [32]. The older sub-population in our study showed higher levels of adherence in comparison to participants'  $\leq 65$  years, in concordance with previous literature [28].

Adherence to AHTM was associated with less GP visits, which may be indicative of better health status [33]. In this population, all participants had access to free GP care; therefore cost barriers do not explain a reduced number of visits reported. Many participants were prevalent, chronic users, and their adherence estimate are likely to be an indicator of long-term adherence, contributing to this protective effect. The results are similar to a recent study of AHTM adherence in Irish adults aged  $\geq 65$  years recruited from community pharmacies, although only 70% of the cohort were GMS eligible [34]. In Ireland, GPs are reimbursed at a capitation rate per GMS patient treated and not paid per patient consultation [35]. Thus, it is not possible to estimate the direct cost of an increase in GP visits to the Irish healthcare payer. However, the burden on GP

resource use (i.e increased workload, waiting times) and indirect costs to the healthcare payer should be considered.

Implementation adherence did not significantly influence the number of ED visits reported or the likelihood of experiencing an ED visit in this cohort. This is in contrast to smaller studies in socio-economically disadvantaged cohorts in the USA that documented a significant inverse association between chronic heart failure (CHF) medication adherence and ED visit rate [18–20]. Similar findings were reported in a large scale administrative database study in the USA; non-adherent hypertensive adults had more ED visits, and the rate of ED visits increased with the extent of non-adherence [36]. In general, the literature detailing the relationship between medication non-adherence and ED visits is mixed. A study of high-risk COPD patients found that those who were adherent were more likely to have an ED visit, but non-adherent were more likely to have higher ICU use and healthcare costs [37]. The authors postulated that adherent patients may exhibit more health-seeking behaviours and thus attending the ED even when symptoms are mild or manageable.

There is a deficit in the literature in relation to the association between AHTM adherence and outpatient visits or use of ancillary care services. In contrast to the results obtained in this analysis, a previous study showed that non-adherence (MPR < 0.80) was associated with a significantly reduced risk of both generic and osteoporosis-related outpatient service utilisation in women aged  $\geq 55$  years [38]. This may signal the presence of the “healthy adherer” bias referred to previously [39]. However, a study of anti-epileptic medication adherence in adults aged  $\geq 65$  years did not find a statistically significant association between non-adherence (MPR < 0.80) and both the risk and number of ancillary care visits [40].

With regard to hospitalisation, previous evidence has suggested that sub-optimal adherence to  $\beta$  blockers and calcium channel blockers in adults aged  $\geq 66$  years predicts an increase in the likelihood of all-cause readmissions, respectively [17]. Participants in this study were using these medications for secondary prevention of cardiovascular events, whereas in the current study, there was no distinction between primary or secondary prevention. Adherence in secondary prevention may be more influential on health outcomes and healthcare utilisation. A statistically significant increase in hospitalisation rate was associated with a 10% increase in AHTM adherence measured using PDC in a study of older Irish community pharmacy patients [34]. However, covariates such as functional disability and depression, found to be significantly associated with hospitalisation rate in the current analysis, were not adjusted for.

Cohort studies in the USA have found that poor adherence to AHTM was significantly associated with an increased risk of cardiovascular-related healthcare utilisation and costs in patients [13, 28]. A Korean study found that poor, but not intermediate adherence to AHTM, measured using

cumulative medication adherence was associated with an increased risk of the first hospitalisation for CVD in new users of AHTM [31].

There are several strengths to this study. Firstly, there is little evidence in the literature on the impact of suboptimal implementation of AHTM on healthcare utilisation. AHTM adherence was calculated using a large nationally representative sample of middle and older aged Irish people, and its association with self-reported healthcare utilisation was reported. Our findings provide useful information to researchers conducting economic evaluations of adherence interventions in community settings.

In addition, the sensitivity of the arbitrary adherence threshold of 80% was tested in this study. We found that lowering this adherence threshold to 70% resulted in a significant reduction in hospitalisation risk and the rate of hospitalisations reported for those who were considered adherent.

We separated the adherence exposure period from the outcome measurement period, therefore minimising the risk of contamination bias. Previous studies have assessed the impact of medication adherence on healthcare utilisations using cross-sectional designs. Of course, by using observational methods, one cannot conclude that non-adherence was the causal factor in healthcare utilisation if a significant association is observed. However, by having the correct temporality one can be more confident in the direction of the relationship.

This study has several potential limitations. Firstly, the study population only includes participants who are eligible to receive medications and healthcare as part of a state-subsided scheme, and so may include a more socially disadvantaged population in those aged 50–69 years. However, 69% of the study cohort was aged  $\geq 70$  years. The HSE-PCRS scheme is representative of 90% of this population [24]. Education level was used as a proxy variable for socioeconomic status and was controlled for in the multivariate analysis.

Secondly, using pharmacy claims data is an indirect method of measuring adherence as dispensation data reflects medication availability as opposed to true exposure. However, it is considered objective, inexpensive, and not subject to the same degree of recall bias as self-reported methods and cost as Medication Event Monitoring Systems (MEMs) [21]. Studies have indicated that information regarding medication use from the pharmacy claims database is more accurate in comparison to other methods [41].

Thirdly, information on participants' healthcare utilisation in the year preceding medication adherence measurement was not available and could not be accounted for in multivariate analyses. Self-reported outcome measurement was used in this study, which may be subject to recall bias. However, a recent Irish study using TILDA data has shown good reliability between self-reported and electronically measured GP and outpatient visits, but not for ED visits [42]. Future research will

document the change in adherence from wave 1 of TILDA to subsequent waves, accounting for baseline self-reported healthcare utilisation.

While we adjusted for a number of covariates in the analysis, there may be unmeasured confounding due to lack of control for illness severity, psychosocial variables such as self-efficacy and health-seeking behaviours (healthy adherer bias). However, including engagement with preventative healthcare screening services in post hoc sensitivity analyses did not significantly alter the results.

Finally, the reasons for the healthcare utilisation episode remain unknown and whether it was cardiovascular-related. Availability of clinical notes would facilitate judgement of this, but the lack of an integrated electronic healthcare record in Ireland prevents the availability of this information.

## Implications for future research

Further population studies, using a combination of adherence measures and disease-specific outcome measures should be conducted to establish the association between non-adherence and healthcare utilisation. Studies should test the validity of the arbitrary 80% adherence threshold across different therapeutic areas, as we have shown that different thresholds may significantly impact association estimates. Consideration should be given to the temporality of medication adherence measurement with respect to health outcome measurement. Finally, over half of the population in this sample of middle and older aged adults were using  $\geq 5$  medications, indicating the presence of multimorbidity. Future research should endeavour to estimate adherence to all chronic medications in people with multimorbidity and establish the association with healthcare utilisation outcomes.

## Conclusion

Irish adults in receipt of state-subsidised healthcare, who are adherent to their antihypertensive medication, report visiting their GP less frequently and are less likely to visit outpatient departments than those who are non-adherent. Having good adherence may represent overall improved health status for the patient and decreased resource burden on the physician. However, no significant relationship was evident between implementation adherence and tertiary healthcare utilisation in this population, which may be due to the lack of specificity for disease-specific healthcare events.

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**Authors' individual contributions** CW, CC and KB were involved in the concept and design of the study. Data was provided by TILDA and the HSE-PCRS. CW carried out the statistical analysis. All authors were involved in interpretation of the data. CW wrote the first draft of the manuscript and all authors contributed to subsequent drafts.

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**Data availability** The data that support the findings of this study are available from TILDA and the HSE-PCRS but restrictions apply to the availability of these datasets, which were used under license for the current study, and so are not publicly available. Researchers interested in using TILDA data may access the anonymised dataset for free from the following sites: Irish Social Science Data Archive (ISSDA) at University College Dublin <http://www.ucd.ie/issda/data/tilda/>; Interuniversity Consortium for Political and Social Research (ICPSR) at the University of Michigan <http://www.icpsr.umich.edu/icpsrweb/ICPSR/studies/34315>.

## Compliance with ethical standards

**Ethical approval** Ethical approval for each wave of TILDA was obtained from the Trinity College Research Ethics Committee. Provision was made within this application to allow for the linking of participants GMS dispensing data, subject to participants' consent.

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