



Pre-stimulus theta power is correlated with variation of motor evoked potential latency: a single-pulse TMS study

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Abstract

There has been a growing interest in the role of pre-stimulus oscillations on cortical excitability in visual and motor systems. Prior studies focused on the relationship between pre-stimulus neuronal activity and TMS-evoked motor evoked potentials (MEPs) have reported heterogeneous results. We aimed to assess the role of pre-stimulus neural activity on the latency of MEPs, which might enhance our understanding of the variability of MEP signals, and potentially provide information on the role played by cortical activity fluctuations in the excitability of corticospinal pathways. Near-threshold single-pulse TMS (spTMS) was applied at random intervals over the primary motor cortex of 14 healthy participants while they sat passively, to trigger hand muscle contractions. Multichannel EEG was recorded during spTMS blocks. Spearman correlations between both the variation in MEP onset latencies and peak-to-peak MEP amplitudes, and the pre-stimulus power of EEG oscillations were calculated across participants. We found that the variation in MEP latency was positively correlated with pre-stimulus power in the theta range (4–7 Hz) in a broad time window (–3.1 to –1.9 s) preceding the spTMS generating the MEP. No correlation between pre-stimulus power in any frequency band and MEP amplitude was found. Our results show that pre-stimulus theta oscillations are correlated with the variation in MEP latency, an outcome measure determined by fiber conduction velocity and synaptic delays along the corticospinal tract. This finding could prove useful for clinicians using MEP latency-based information in pre- or intra-operative diagnostics of corticospinal impairment.

Keywords Brain stimulation · Pre-target onset · Ongoing activity · Latency variability · Single-pulse TMS · Motor evoked potentials · Electroencephalography

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Introduction

Fluctuations in the power of brain oscillations at different frequency bands have been shown to reflect cyclic modulations of cortical excitability, affecting the outcomes of visual detection (Thut et al. 2006; Busch et al. 2009; Romei et al. 2010; Busch and VanRullen 2010), discrimination (van Dijk et al. 2008), and perception in general (Romei et al. 2008a, b; Dugué et al. 2011; Jaegle and Ro 2014; Tomassini et al. 2017). In motor systems, particularly during movement preparation, event related synchronization/desynchronization in the alpha and beta bands has also been reported consistently in the literature (Serrien et al. 2003; Deiber et al. 2012; van Wijk et al. 2012; Zaepffel et al. 2013; Kajihara et al. 2015). In this context, methods that probe the state of cortical excitability and its temporal dynamics preceding the onset of a stimulus (either a visual target or a descending corticospinal volley) could contribute to our understanding of these phenomena.

In motor systems, the delivery of single-pulse transcranial magnetic stimulation (spTMS) to the primary motor cortex induces motor evoked potentials (MEPs) in peripheral muscles and is widely used to probe the integrity of descending corticospinal tract fibers. spTMS can also be used to estimate the level of corticospinal excitability (Kujirai et al. 1993; Nakamura et al. 1997; Chen et al. 1998; Sanger et al. 2001; Chen 2004) while quantifying the impact of motor training or the level of impairment due to brain lesions or disease. More sophisticated versions of these same approaches have also served to probe the causal implications of local frequency-specific brain rhythms in motor control processes. For example, spTMS to the primary motor cortex preceded by pulses at different intervals to prefrontal regions during a Go/No-go task has served to reveal a causal role for beta (20 Hz) rhythms in the ‘top-down’ control of inhibitory signals modulating MEP amplitudes (Picazio et al. 2018). Similarly, in the visual/attentional domain, the ability of occipital spTMS to elicit reportable visual phosphenes (Elkin-Frankston et al. 2011) has been applied to reveal alpha excitability fluctuations in primary visual areas and demonstrate periodic attentional sampling of visual stimuli according to ongoing theta (5 Hz) cycles (Dugue et al. 2015).

One disadvantage of MEPs to evaluate the physiological mechanisms subtending primary motor and corticospinal function is the large trial-to-trial variability in the amplitude of these motor evoked responses (Kiers et al. 1993; Wassermann 2002; Saisanen et al. 2008). Previous studies that tried to associate this variability with ongoing oscillatory brain activity have reported heterogeneous results (Zarkowski et al. 2006; Sauseng et al. 2009; Mäki and Ilmoniemi 2010; Berger et al. 2014; Keil et al. 2014; Schulz et al. 2014; Iscan et al. 2016). Zarkowski et al. (Zarkowski et al. 2006) showed that MEP amplitude correlated negatively with pre-TMS stimulus power in the high alpha (10–13 Hz) frequency band and positively with power in the low gamma (30–60 Hz) band. The ratio of power in the low gamma to high alpha band during the pre-pulse period was the best predictor of MEP amplitude. Sauseng et al. (2009) confirmed this relationship between power and MEP amplitude for the alpha band only. Maki and Ilmoniemi (2010) found that power in the low beta (15–18 Hz) range measured over the stimulated primary motor cortex (M1) (but not over other non-stimulated areas) was weaker just before large versus small amplitude MEPs. Schulz et al. (2014) found a negative correlation between pre-TMS (–550 to 0 ms) power in beta (15–35 Hz) band and MEP amplitude. They also found a positive correlation between pre-TMS alpha (8–14 Hz) band corticomuscular coherence and MEP amplitude in centro-parietal regions. Following their study, Keil et al. (2014) also found a negative correlation between EEG power in beta frequency (18 Hz) and pre-TMS power in a parietal electrode

cluster. However, they did not find a relationship between corticomuscular coherence and MEP amplitude in the alpha band. Instead, they found a linear trend in the high beta (34 Hz) band in similar areas (centro-parietal regions). However, other studies failed to replicate these findings (Iscan et al. 2016). Mitchell et al. (2007) and Berger et al. (2014) found no correlation between pre-stimulus EEG power in various frequency bands and MEP amplitude across trials. Similarly, van Elswijk et al. (2010) did not find a significant relationship between pre-TMS EEG power or pre-TMS EEG phase and MEP amplitude. However, they reported that the pre-TMS phase of the EMG determines MEP amplitude. Mitchell et al. (2007) also showed that the background EMG correlated with the MEP amplitude.

Compared to MEP amplitude, which has been shown to have high intra-individual and inter-individual variability, MEP latency is a much more stable measure within individuals (Kiers et al. 1993; Cacchio et al. 2009, 2011), with the potential to provide information on signal conduction processes operating at the cortical and subcortical level (Kallioniemi et al. 2015). In spite of its stability, there are considerably fewer papers on MEP latency compared to MEP amplitude. Latency is highly influenced by conduction properties and synaptic delays at central (cerebral and spinal) and peripheral divisions of the nervous system (Fernández et al. 2013). Bergmann et al. (2012) triggered spTMS over M1 during non-REM sleep according to the phase of spontaneous (> 1 Hz) neocortical slow oscillations (SO) to show that MEPs were larger and their latencies were shorter during SO up-states than during down-states, which revealed that MEP latency can be modulated by brain states. In another spTMS study, Moscatelli et al. (2016) showed that latency and MEP amplitude were sensitive to training level such that on average karate athletes had shorter latency and larger MEPs compared to controls.

Hamada et al. (2013) showed that inter-individual variability in the offline effects (also referred to as after-effects following the end of a repetitive TMS pattern) generated by continuous theta burst stimulation (cTBS) over M1 could be predicted by the latency of MEPs recorded before stimulation. This result was also replicated by Huang and Mouraux (2015). Nonetheless, clinically, no interhemispheric differences in MEP latency have been found in healthy participants (Saisanen et al. 2008) or in brain tumor patients (Picht et al. 2012). Indeed, no effect of clinical variables on MEP latency has been found when such are corrected for the subjects’ height (determining the distance that corticospinal volleys need to complete from M1 to reach the monitored peripheral muscles). Observations in healthy participants (Saisanen et al. 2008) or healthy hemisphere of brain tumor patients (Picht et al. 2012) found large inter-subject differences in MEP latency. Nonetheless, it should be noted that Huang and Mouraux (2015) did not find a significant

correlation between latency and height. On the contrary, Saisanen et al. (2008) reported a significant effect of height on latency. Saisanen et al. (2008) also found an effect of age on latency. Recently, latency variability between subjects was analyzed during TMS motor mapping in a large cohort of neurosurgical patients. The authors identified, for the three muscles they analyzed, gender and antiepileptic drug intake as muscle-specific factors influencing MEP latency together with resting motor threshold, tumor side and tumor location (Sollmann et al. 2017).

In the current study, we used spTMS to test the role of pre-stimulus cortical oscillations on MEP latency. Our hypothesis is that MEP latency as an outcome measure could give more consistent results regarding the role of pre-stimulus oscillations since it is a more stable measure than amplitude. Therefore, we focused our analysis on MEP latency and checked pre-stimulus activity in different frequency bands to see which bands covary with this outcome measure. We also measured MEP amplitudes as in other studies for completeness.

Contributing a novel finding to prior research, we found that pre-stimulus power in the theta range (4–7 Hz) was positively correlated with the variation in MEP latencies in a pre-stimulus time window (−3.1 to −1.9 s). Our results show that the power of pre-stimulus theta oscillations is correlated with the variation in MEP latency.

Methods

Participants

Fourteen right-handed young adults (3 males and 11 females) between 18 and 39 years of age (mean: 25 ± 6 , SD) participated in the study. They had no history of psychiatric or neurological conditions, they were not taking any medication known to affect motor cortical excitability at the time of the study and did not have any contraindications for TMS. All tolerated the procedure without any side effects or complications. They gave written informed consent to participate in the study, which followed international guidelines and recommendations for the safe use of TMS (Rossi et al. 2009). The study had been approved by the local ethics committee (CPP Ile de France 1) and was conducted in accordance with the Declaration of Helsinki.

Material

The experiment was controlled using Matlab (MathWorks, Natick MA, USA) and the Psychophysics Toolbox (Brainard 1997) running on a PC (operating system: Windows). Synchronization between the computer, TMS stimulator, EEG, and EMG systems was provided by a trigger generator

(Master-8, AMPI, Jerusalem, Israel), USB, and coaxial connections.

TMS pulses were delivered using a biphasic repetitive magnetic stimulator (Super Rapid², Magstim, Whitland, UK) and a 70-mm diameter figure-of-eight coil. Individual anatomical MRI (T1-weighted TR = 2300 ms, TE = 4.18 ms, matrix size = 256×256 , and 176 sagittal slices with 1 mm thickness, acquired on a 3T Siemens scanner), when available, or template anatomical MRI, were fed to a neuronavigation system (Brainsight, Rogue Inc., Montreal, Canada). Neuronavigation was used at all times to ensure that once localized, the same M1 spot was consistently stimulated throughout the experiment.

EEG was continuously acquired from 13 scalp contacts (10–20 EEG system locations F3, Fz, F4, FC1, FC2, C3, Cz, C4, CP1, CP2, P3, Pz, P4), using a TMS-compatible EEG system (BrainAmp DC and BrainVision Recording Software, Brain Products GmbH, Munich, Germany) and TMS-compatible electrodes. The reference and ground were placed on the nose and right earlobe contralateral to the stimulation site near electrode C3. Skin-electrode impedance was maintained at all times below 5 k Ω . The signal was digitized at a sampling rate of 5 kHz.

EMG was acquired using surface electrodes attached to the skin in a belly tendon montage over the first dorsal interosseus (FDI) and the abductor pollicis brevis (APB) muscles in the right hand, contralateral to the cortical stimulation site in participants' dominant left primary motor cortex using a data acquisition system [PowerLab 26T (LTS) and Chart Software, ADInstruments, New Zealand]. The signal was digitized at a sampling rate of 40 kHz, and then down-sampled offline to 5 kHz prior to data analysis.

Procedure

Participants were seated in a comfortable chair with their elbows flexed at approximately 90° and both hands resting on a table. M1 of the dominant (left) hemisphere was stimulated. The coil was consistently held tangentially to the scalp, with the handle pointing posterior with a 45° angle from the sagittal plane of the head for optimal orientation (Brasil-Neto et al. 1992). In this orientation, induced electrical current travels in posterior to anterior direction (Janssen et al. 2015). Once the motor hot spot was identified, it was stimulated with consistent coil location, orientation and tilt throughout the rest of the experiment. The resting motor threshold (RMT) was determined as the lowest intensity capable of inducing five out of ten MEPs of at least 50 μ V peak-to-peak amplitude (Rossini et al. 1994). For our group of 14 participants the average RMT was $70 \pm 5\%$ of the maximum stimulator output (MSO).

Following the RMT determination procedure, 4 blocks of 30 single-pulses at 110% of RMT were delivered at random

intervals of approximately 10 s drawn from a Poisson distribution (98% of intervals in the range of 9–11 s), while EEG and EMG were recorded. A mean interval of 10 s was chosen to strike a balance between the following three considerations: (1) Allowing enough time for effects of the prior pulse to decay completely; (2) Leaving a “clean” ~5 s window of time before each pulse; and (3) Maximizing the number of trials in the session. In three participants stimulation intensity was increased to 120%, 125% and 135% of RMT, respectively, to ensure the presence of consistent MEPs.

The instruction given to participants was to relax, fixate a small cross on a computer display in front of them and avoid eye movements and eye-blinks several seconds before and immediately after each TMS pulse. Blocks were separated by a 2–5 min break.

Data analysis and statistics

Data analysis and statistics were performed using Matlab (MathWorks, Inc) and the Statistics Toolbox for Matlab. Preprocessing and epoching of the EEG data were carried out using the FieldTrip toolbox for Matlab (Oostenveld et al. 2011). Epochs were extracted from 5 s before each trigger (the TMS pulse onset). Since analyses focused on the EEG time course before the TMS pulse was delivered, the artifact created by the TMS pulse on subsequent EEG data (i.e., affecting recordings in the time period immediately following the TMS pulse) were not a concern. No filtering was applied to the data during preprocessing. Trials with ocular artifacts during the final 3 s prior to the TMS pulse were rejected by visual inspection. EEG recordings and MEPs from trials on which the TMS coil was not well positioned (i.e., stimulation was delivered more than 5 mm away from our FDI pre-defined hotspot) were identified in

the neuronavigation log file and excluded from subsequent data analyses. Trials with excessive noise due to TMS coil contact or friction on EEG contacts, muscle contraction or eye-blinks were removed from the analysis. Trials that did not elicit MEPs were also excluded. Out of 120 trials, the average number of trials used in the analyses was 61 ± 22 (see Table 1 for subject-wise number of valid trials).

EMG data were first down-sampled by a factor of 8, from 40 to 5 kHz, to match the sampling rate of the EEG data. EMG data epochs extended from -0.5 to +0.5 s with respect to the TMS trigger pulse. The MEP is a very brief event occurring within the first 100 ms after the TMS pulse (at ~ 20–25 ms), but well after the very brief artifact in the EMG time series introduced by the TMS pulse (see Fig. 1).

For each epoch of EMG data, we subtracted the mean over the baseline interval from 8 to 18 ms after the TMS pulse, defined as the “quiet” interval in between the TMS artifact and the onset of the motor evoked potential. This results in a very precise alignment of the amplitude of the signal at the time of MEP onset. We then computed the

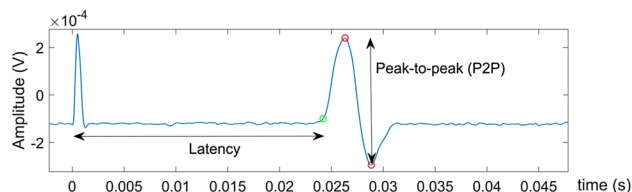


Fig. 1 An example of EMG record including a single TMS pulse (at time=0) and the corresponding Motor Evoked Potential (MEP). The two main parameters that are extracted to characterize the MEP, are the peak-to-peak (P2P) amplitude and the latency

Table 1 Summary statistics for Latency, MEP amplitude, and pre-stimulus power across subjects

Subject	Latency (ms)			Amplitude (mV)			Pre-stimulus power ($\mu\text{V}^2/\text{Hz}$)					Valid trials
	Mean	SD	CQV	Mean	SD	CQV	Delta	Theta	Alpha	Beta	Gamma	
S1	25.13	1.13	0.02	0.30	0.23	0.54	7.53	0.70	1.09	0.29	5.99	65
S2	21.99	1.60	0.05	0.34	0.24	0.43	25.72	2.58	1.06	0.17	2.28	79
S3	21.49	1.45	0.02	0.11	0.08	0.30	6.07	1.05	0.33	0.14	0.13	65
S4	23.73	1.21	0.02	0.64	0.87	0.69	24.03	1.34	1.11	0.24	0.16	36
S5	22.36	1.09	0.02	0.24	0.22	0.53	7.09	1.09	0.57	0.16	0.19	92
S6	24.26	3.27	0.08	0.39	0.31	0.49	46.75	1.57	0.35	0.14	0.24	35
S7	22.92	1.02	0.02	0.13	0.10	0.38	4.25	0.87	4.24	0.30	0.13	36
S8	23.08	1.89	0.05	0.26	0.22	0.47	4.28	2.99	2.79	0.22	0.07	103
S9	21.76	1.01	0.04	0.39	0.56	0.57	21.27	2.72	1.47	0.19	0.16	57
S10	23.17	1.44	0.04	0.50	0.59	0.65	7.58	2.26	0.55	0.12	0.08	67
S11	24.37	1.34	0.05	0.33	0.29	0.49	4.44	1.17	0.53	0.19	0.43	81
S12	23.57	0.79	0.02	0.31	0.33	0.63	3.75	0.87	0.27	0.09	0.13	51
S13	24.73	0.87	0.02	0.22	0.12	0.36	8.55	1.61	3.86	0.29	0.09	34
S14	25.07	1.67	0.04	0.31	0.30	0.60	5.99	1.19	0.73	0.15	0.08	54

peak-to-peak MEP amplitude (P2P) and MEP onset latency, as follows:

P2P was computed as the difference between the maximum and the minimum EMG amplitude from 20 to 35 ms after the TMS pulse. Latency was defined as the first time-point after the TMS pulse for which the amplitude exceeded 5% of the baseline-to-first peak amplitude (in absolute value) (Huang and Mouraux 2015). In Fig. 2, latency and MEP amplitude (P2P) distributions of subjects are presented using histogram plots. Individual means and standard deviations (SD) for latency and MEP amplitudes are presented in Table 1.

In Fig. 3, single trial MEP responses of a representative participant (Subject #5) are presented.

Variability of MEP onset latencies and MEP amplitude

Coefficient of quartile variation (CQV) (Bonett 2006) was used to measure the variability of MEP amplitudes (Iskan et al. 2016), and MEP onset latencies within participants.

This is a more robust method than the coefficient of variation, as it is less sensitive to the deviations from normality (Bonett 2006). Quartiles are the points (i.e., Q_1, Q_2, Q_3) that divide any ranked data set into four groups that have the same number of data values. Q_1 is the middle number between the smallest number and the median (Q_2) of the data set. The third quartile Q_3 is the middle value between Q_2 and the highest value of the data set. Briefly, CQV is the ratio of interquartile range ($Q_3 - Q_1$) to the sum of first (Q_1) and third (Q_3) quartiles in data:

$$CQV = \frac{Q_3 - Q_1}{Q_3 + Q_1} \tag{1}$$

For each participant, CQV generated a single value (for either MEP amplitudes or latencies) using the single trial outputs (see Table 1).

Fig. 2 Latency (bottom panel) and MEP amplitude (P2P, top panel) distributions of the 14 subjects (S1-S14) participating in the study. Colors represent the number of occurrences (number of trials)

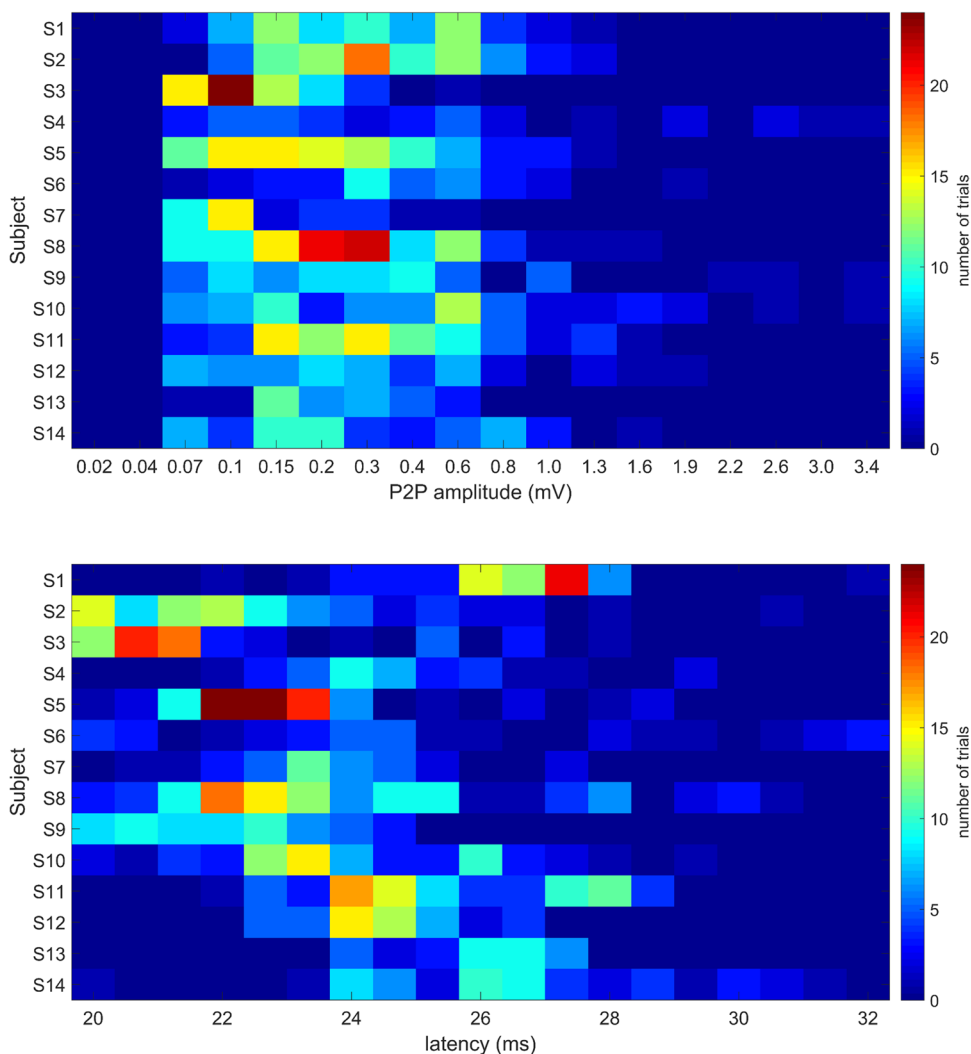
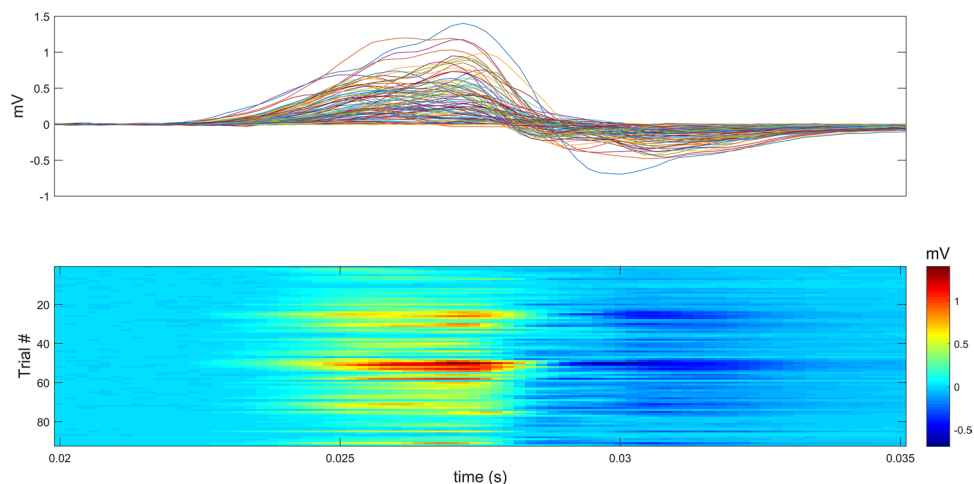


Fig. 3 Single trial MEP responses of a representative participant (Subject #5) zoomed into 20–35 ms after TMS pulse



Power of pre-stimulus EEG oscillations

Pre-stimulus power between -4.5 and -0.5 s was estimated from the spectrum calculated with a Fast Fourier Transform (FFT), with a Hanning window to remove discontinuities in the edges of the signal. We used the Fieldtrip method: `mtmfft` (multi-taper method fast Fourier transform). We were interested in low frequencies. Therefore, we did not use a high-pass filter. In the single trial analysis the power was calculated separately for each pre-stimulus interval in each frequency band, electrode and participant. For the group-level analysis, the power was averaged across all trials in each frequency band, electrode and participant. These frequency bands were delta (1–3 Hz), theta (4–7 Hz), alpha (8–13 Hz), beta (14–30 Hz) and gamma (31–100) bands. Mean pre-stimulus power values of subjects averaged across all electrodes for each band are given in Table 1.

Outlier detection

First, we inspected data for the presence of multivariate outliers. In our study we employed the commonly used Mahalanobis distance (MD) method to detect outliers (Lepage et al. 2012; Ballard et al. 2014). Data points whose MD value were bigger than the critical χ^2 value were considered to be outliers.

Correlations between power of pre-stimulus oscillations and mean or variability of MEP amplitude

We computed a Spearman correlation between pre-stimulus power in delta (1–3 Hz), theta (4–7 Hz), alpha (8–13 Hz), beta (14–30 Hz) and gamma (31–100) bands as the independent variable and the mean or CQV of MEP amplitude as the dependent variable across participants using the Fieldtrip function: `ft_statfun_correlationT`. The time window for the calculation was between 4.5 and 0.5 s before the stimulus.

Correlations between power of pre-stimulus oscillations and mean or variability of MEP onset latencies

We computed a Spearman correlation between pre-stimulus power in the delta (1–3 Hz), theta (4–7 Hz), alpha (8–13 Hz), beta (14–30 Hz) and gamma (31–100) bands as the independent variable and the mean or CQV of MEP onset latencies as the dependent variable across participants using the Fieldtrip function: `ft_statfun_correlationT`. The time window for the calculation was between 4.5 and 0.5 s before the stimulus.

Time–frequency analysis for correlations

We performed a post hoc cluster-based permutation test for correlations to isolate more precisely the effect in time using time–frequency analysis. A step size of 0.1 s and a sliding window of 1 s were chosen. To evaluate the changes in frequency content of the signal over time we used the Fieldtrip method: `mtmconvol` (multi-taper-method convolution). This method convolves the signal with a complex wavelet. Again we used a Hanning window to remove discontinuities in the edges of the signal.

Multiple comparisons

The significance probability was estimated using the Monte Carlo Method. The multiple comparisons problem was addressed using cluster-based permutation statistics (Maris and Oostenveld 2007). The neighborhood for the cluster was defined with the triangulation method based on a two-dimensional projection of the electrode position. The number of permutations was set to 10,000.

Results

Correlation between pre-stimulus band power and the mean or variability of MEP amplitude across participants

No correlation between power and MEP amplitudes across participants was found in any frequency sub-bands, in the delta (1–3 Hz), theta (4–7 Hz), alpha (8–13 Hz), beta (14–30 Hz) and gamma (31–100) range (see "Methods"). Moreover, no correlation of pre-stimulus power was found with the MEP amplitude variability (i.e., CQV of MEP amplitude).

Correlation between pre-stimulus theta power and CQV of MEP onset latencies

Figure 4 shows the topography of the Spearman correlation between pre-stimulus theta power and CQV of MEP onset latencies using cluster-based permutation statistics for all

participants in our cohort ($N=14$) and in a subcohort of 13 subjects [one subject who had an MD value ($MD=37.75$) bigger than the critical χ^2 value of 36.12 ($df=14$; $\alpha=0.001$), was excluded from the analysis]. It can be observed that for both analyses over the precentral and parietal regions, theta power correlated positively with the variability (i.e., CQV) of latencies. A scatter plot from channel FC2 is given in the bottom part of the figure.

Time–frequency analysis for correlations between pre-stimulus theta power and latency of MEPs across subjects

In Fig. 5 correlations between power of pre-stimulus theta oscillations and latency of MEPs are provided with significant cluster channels in the 3.1–1.9 s range before the TMS pulse. When the time–frequency analysis was performed for all (i.e., 14) subjects, the significant cluster ($p=0.021$) spanned the time range 3.2–1.6 s before the TMS pulse (see Fig. 6).

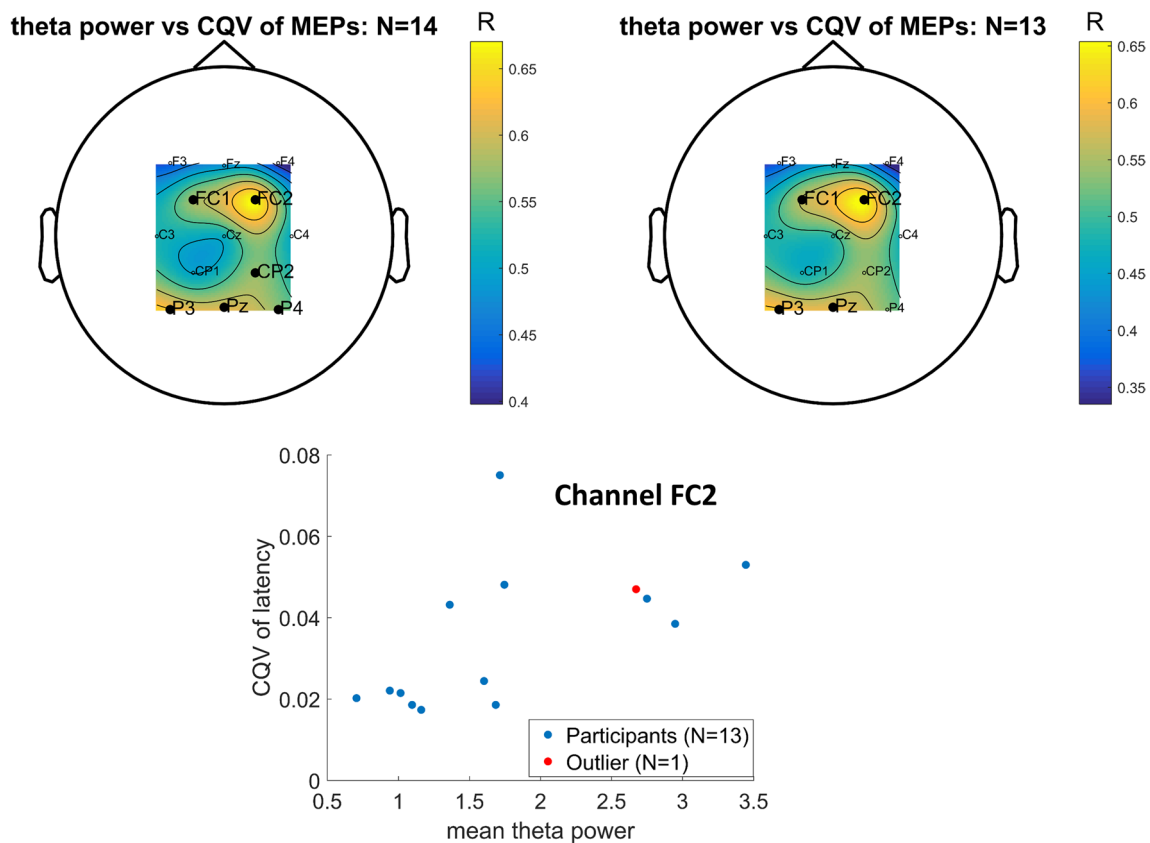


Fig. 4 Topography of the correlation between pre-stimulus (–4.5 to –0.5 s) theta oscillations and CQV of MEP onset latencies for the complete cohort of 14 subjects (left panel) and for a subset of 13 participants (right panel, dataset excluding the participant identified as an outlier). Significantly clustered channels are marked with big-

ger labels ($N=14$), one cluster (FC1, FC2, CP2, P3, Pz, P4) with $p=0.019$; $N=13$, two clusters (FC1 & FC2 with $p=0.048$; P3 & Pz with $p=0.049$). A scatter plot from channel FC2 is given in the bottom part of the figure

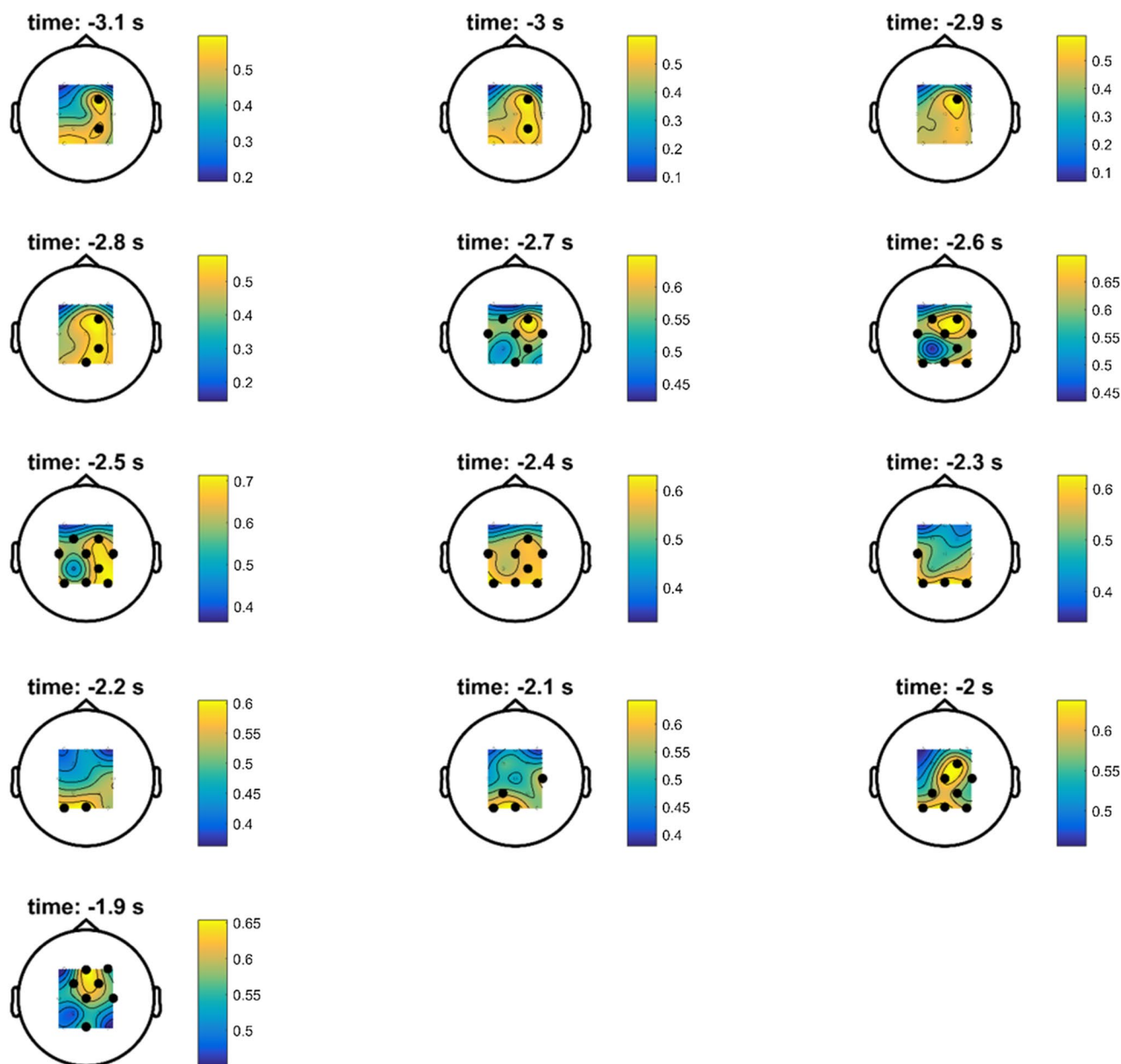


Fig. 5 Correlations between pre-stimulus (–3.1 to –1.9 s) theta oscillations and CQV of MEP onset latencies. Significant cluster channels ($p=0.038$) are labeled with black dots. Analyses were performed on a cohort of $N=13$ (i.e., the participant identified as an outlier was excluded)

Discussion

Driven by a growing interest in the role of pre-stimulus oscillations on cortical excitability, the current study uses EEG recordings to explore potential associations between pre-stimulus rhythmic activity in different frequency bands and parameters of the MEPs generated by stimulating the primary motor cortex. We focused on testing the role of cortical pre-stimulus rhythmic neural activity on MEP amplitude and its latency, a less studied outcome measure which might help further elucidate the source of variability for

motor evoked signals, and inform on how this influences excitability across corticospinal pathways.

We found that pre-TMS power in the theta range (4–7 Hz) was positively correlated with variation in latency of MEPs across a broad time window (–3.1 to –1.9 s) preceding the stimulation pulse. We did not find evidence for an association between power in any frequency band and MEP amplitude.

Figure 2 shows the variation of peak-to-peak (P2P) amplitudes and latencies between subjects (see also Fig. 3 for MEP signals on a representative subject, #5). We tried

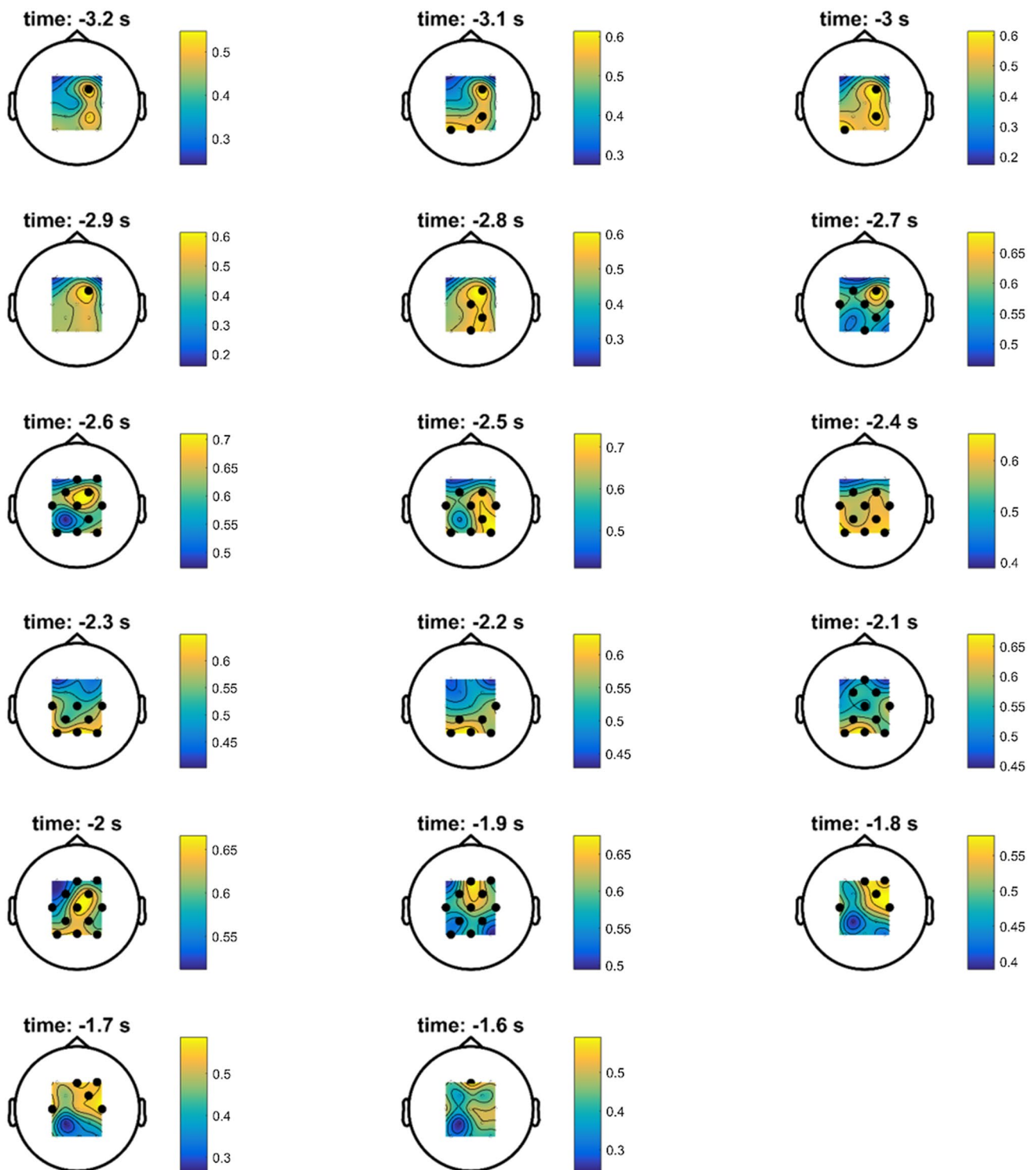


Fig. 6 Correlations between pre-stimulus (–3.2 to –1.6 s) theta oscillations and CQV of MEP onset latencies. Significant cluster channels ($p=0.021$) are marked with black dots. $N=14$

to explain this variability by testing for a potential role of pre-stimulus power in different EEG frequency bands. Nonetheless, no significant correlation was found between pre-stimulus power and MEP amplitude. This result is in

line with those reported by some previous studies (Mitchell et al. 2007; van Elswijk et al. 2010; Berger et al. 2014; Iscan et al. 2016) using similar approaches. In contrast, two previous studies (Zarkowski et al. 2006; Sauseng et al. 2009) did

report a significant correlation, a discrepancy which could be explained by the low reliability of such studies based on a very small sample of participants ($N=4$ and $N=6$, respectively).

The main novelty of the present study is the potential relationship between the power of pre-stimulus oscillations and MEP onset latencies during prolonged spTMS sessions. Our results (Figs. 4, 5) reveal a significant correlation between pre-stimulus theta power and the CQV of latencies across subjects. The time window (i.e., -3.1 to -1.9 s) in which this correlation was found shows that the variation in latency depends on the pre-stimulus EEG up to several seconds before the TMS pulse. It should be noted that the duration of the pre-stimulus window showing significant correlations remained similar even when an outlier participant was included in the analysis (Fig. 6). Regarding the topography of the relationship, the correlations were most pronounced over right precentral and left parietal regions. Previously these two regions were shown to be important for sensorimotor integration, intentions, planning, selection and programming of movements, and were identified in pre-TMS beta power and MEP amplitude correlations (Schulz et al. 2014). Here we show that same regions are involved in the pre-TMS theta power and latency variability.

Two main questions are worth-discussing. First, our data reveal that power in the theta, but not beta (as these reflect motor-related oscillatory activity), alpha, or delta bands are correlated with variability in MEP latency. In direct support of this finding, a recent paper, showed that cortical excitability during long periods of wakefulness correlated significantly with oscillations in theta power, whereas the power of delta, alpha, or beta rhythms failed to predict the variability of MEP amplitude (Ly et al. 2016). Second, in our study, MEP latency but not MEP amplitude correlated significantly with theta rhythms. The selective impact of theta power on MEP latency variability, a parameter depending on conduction and spinal (and also neuromuscular) synaptic delays for corticospinal descending volleys could be explained by the higher stability of MEP latency measures compared to amplitude, which was the case in terms of CQV (See Table 1) in our study. Indeed, a recent report (Tomassini et al. 2017), showed that the phase of theta oscillations predicted perceptual performance more than one second before a movement. Although that study was based on theta phase rather than theta power and the correlation was found for perceptual performance instead of MEP latency, it is relevant in terms of showing correlations in a pre-stimulus time range very far in advance of stimulus onset.

Our interpretations might be tempered by two potential methodological shortcomings. First, for practical reasons, our study recorded EEG from 13 scalp contacts. Second, in some participants the intensity of TMS output had to be adjusted to ensure that MEPs were evoked by stimulation,

a fact that could have increased the variability and added noise to our data. The number of subjects that participated in our study ($N=13$) was relatively modest, but within the range of several other studies looking for a relationship between oscillations and MEP measures [(Mäki and Ilmoniemi 2010), $N=14$; (Sauseng et al. 2009), $N=6$; (Iskan et al. 2016), $N=17$]. A substantial proportion of the TMS trials collected with EEG had to be excluded from the analysis due to several types of artifacts. However, on average we still had more than twice the optimal number of trials (Goldsworthy et al. 2016).

To the best of our knowledge, this is the first study to find a correlation between pre-stimulus theta power and the latency of MEPs. The exploitation of this link could prove useful to clinicians using latency information for pre-operative diagnostics of patients with tumors (Sollmann et al. 2017) or as an indicator of lower extremity and gait impairment (Peters et al. 2017). It could also serve to better assess latency information by including the variation stemming from changes in theta power. Moreover, our results could contribute to the development of brain-state triggered stimulation (Walter et al. 2012; Gharabaghi et al. 2014), a promising therapeutic methodology (Gharabaghi et al. 2014; Zrenner et al. 2016; Karabanov et al. 2016) based on tailoring the delivery of TMS to motor cortical regions (in terms of intensity, timing/phase and/or frequency) based on ongoing patterns of cortical activity to maximize its modulatory impact on corticospinal systems (Romei et al. 2016).

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Compliance with ethical standards

Conflict of interest The authors declare no conflict of interest.

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