### REVIEW



# Prescription psychostimulants for the treatment of stimulant use disorder: a systematic review and meta-analysis

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### Abstract

**Rationale** Agonist-based pharmacologic intervention is an accepted approach in treatment of opioid and tobacco use disorders. **Objectives** We conducted a systematic review and meta-analysis to evaluate usefulness of an agonist approach as treatment of (psycho)stimulant use disorder (PSUD).

**Methods** We reviewed PubMed/Medline, LILACS, and ClinicalTrials.gov databases searching for randomized, double-blind, placebo-controlled, parallel-design studies evaluating outcomes of individuals treated for cocaine- or amphetamine-type substance use disorder. We combined results of all trials that included the following prescription psychostimulants (PPs): modafinil, methylphenidate, or amphetamines (mixed amphetamine salts, lisdexamphetamine, and dextroamphetamine). The combined sample consisted of 2889 patients. Outcomes of interest included the following: drug abstinence (defined as 2–3 weeks of sustained abstinence and the average maximum days of consecutive abstinence), percentage of drug-negative urine tests across trial, and retention in treatment. We conducted random-effects meta-analyses and assessed quality of evidence using the GRADE system. **Results** Thirty-eight trials were included. Treatment with PPs increases rates of sustained abstinence [risk ratio (RR) = 1.45, 95% confidence interval (CI) = (1.10, 1.92)] and duration of abstinence [mean difference (MD) = 3.34, 95% CI = (1.06, 5.62)] in patients with PSUD, particularly those with cocaine use disorder (very low-quality evidence). Prescription amphetamines were particularly efficacious in promoting sustained abstinence in patients with cocaine use disorder [RR = 2.44, 95% CI = (1.66, 3.58)], and higher doses of PPs were particularly efficacious for treatment of cocaine use disorder [RR = 1.95, 95% CI = (1.38, 2.77)] (moderate-quality evidence). Treatment with prescription amphetamines also yielded more cocaine-negative urines [MD = 8.37%, 95% CI = (3.75, 12.98)]. There was no effect of PPs on the retention in treatment.

**Conclusion** Prescription psychostimulants, particularly prescription amphetamines given in robust doses, have a clinically significant beneficial effect to promote abstinence in the treatment of individuals with PSUD, specifically the population with cocaine use disorder.

Keywords Psychostimulants  $\cdot$  Cocaine  $\cdot$  Methamphetamine  $\cdot$  Prescription psychostimulants  $\cdot$  Substance use disorders  $\cdot$  Modafinil  $\cdot$  Methylphenidate  $\cdot$  Amphetamine  $\cdot$  Agonist  $\cdot$  Replacement

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# Introduction

Global use of cocaine increased two-fold over the 5 years leading to 2017 reaching all-time high of an estimated 18 million that used cocaine in the past year, with an accelerating trend observed in Europe, North America, some South American countries, and Australia. Moreover, cocaine seizures reported in Asia and West Africa suggest that its use might be increasing in these regions as well (United Nations Office on Drugs and Crime 2019). Global market for amphetamine and methamphetamine continues to expand, and in 2017, an estimated 29 million used amphetamines in the past year, with an increasing trend seen in the United States and in East and South-East Asia (United Nations Office on Drugs and Crime 2019). A significant number of individuals who use stimulants will develop a (psycho) stimulant use disorder (PSUD), which includes a cocaine use disorder (CUD) and an amphetamine-type use disorder (AUD), which are among the most prevalent drug-related conditions around the world with highly burdensome health and social consequences (United Nations Office on Drugs and Crime 2019). The prevalence of CUD and AUD has raised globally at 39.7% and 22.5%, respectively, from 1990 to 2016. CUD accounts for 1153.6 thousand disability-adjusted life years (DALYs), a disease burden measure that comprises years of life lost (YLLs) and years of life lived with disability (YLDs), while AUD accounts for 881.4 thousand DALYs (Degenhardt et al. 2016). Only alcohol and opioids pose heavier burdens than CUD and AUD. Stimulant drugs have considerable potential of yielding use disorders among their users, with 16.7% of those with any lifetime cocaine use developing a use disorder. This proportion is of 11.2% for amphetamines, comparable to 15.4% for alcohol and 23.1% for heroin (Anthony et al. 1997). More recent publications also found a lifetime prevalence of PSUD of 1.7% (Grant et al. 2016) and that 15.6% of cocaine abusers transition to dependence at any point of their lives (Florez-Salamanca et al. 2013).

Despite the alarming public health impact of PSUD, the proportion of individuals with PSUD who are receiving treatment is extremely low (United Nations Office on Drugs and Crime 2019), with only 20% of patients with cocaine or amphetamine use disorder receiving specialty treatment in 2018 as patients with CUD and AUD and similarly low rates reported in patients with opioid use disorder (OUD) though rates of OUD treatment have been increasing in the past few years (Substance Abuse Mental Health Services Administration 2019). One of the reasons for such low treatment rates is the absence of the medication-based model of PSUD treatment. This is in contrast to much higher treatment rates among those with opioid use disorder, where medications are an essential treatment component (Schuckit 2016). In the absence of an approved and well-accepted pharmacological intervention to treat PSUD, psychosocial interventions remain the primary model of treatment (De Giorgi et al. 2018; Lee and Rawson 2008; Minozzi et al. 2016).

The most accepted pharmacological intervention for substance use disorders involves treatment with agonist-like medications, the approach known as a replacement or substitution therapy (Darke and Farrell 2016). Agonist-based treatment has been successfully implemented in treatment of opioid use disorder (Mattick et al. 2009) and tobacco use disorder (Hartmann-Boyce et al. 2014). A parallel approach has for long been regarded as a promising approach for treatment of PSUD (Shearer et al. 2002), and several medications were proposed as putative "stimulant agonists" (Amato et al. 2011; Castells et al. 2016; Rush and Stoops 2012; Stoops and Rush 2013).

An "agonist" medication would have a similar pharmacologic and behavioral effect as the drug of abuse, providing relief of craving and other symptoms of acute and protracted withdrawal, which are main factors responsible for the maintenance of drug use and for relapse following periods of abstinence (Darke and Farrell 2016). Another desired feature of an agonist medication is that it may function as a "blocker" in case of the use of the primary drug and therefore diminish euphoric effects and prevent further escalation of use (Shearer 2008). At the same time, for a medication to be acceptable for agonist-based treatment, it will need to have an acceptable safety profile and be feasible for clinical use (Darke and Farrell 2016). Medication taken orally with slow onset of action and a long elimination half-life will have less euphoric and discontinuation subjective effects reducing the risk of abuse, and stable blood concentrations will provide "pharmacological stability" (Darke and Farrell 2016). Moreover, the agonist medication should have few acute side effects and no behavioral organ toxicity in clinically used doses and therefore be acceptable for chronic administration. Preferably, it will have mildly positive subjective effects that will promote medication adherence and have acceptable safety profile when combined with the primary or other drugs of abuse (Darke and Farrell 2016). Effective treatment with agonist medication will facilitate initial abstinence with resulting improved adherence to treatment, as high rates of treatment dropout are one of the main challenges in working with this population. Improved treatment engagement will allow patients to benefit from cognitive and behavioral therapies offered in the treatment program to prevent relapse and extend benefits of drug abstinence, all of which would be difficult in patients who are actively using stimulants. Moreover, individuals receiving agonists are likely to become interested in and access other medical and social interventions, as well as other therapeutic services offered at the program, such as psychoeducation and harm-reduction services (Shearer et al. 2002).

All of the stimulants acutely increase brain levels of dopamine and noradrenaline, producing stimulating and pleasurable euphoric effects, and with chronic use produce long-

lasting changes in brain circuits (Kalivas and O'Brien 2008). However, individuals with PSUD who use stimulants chronically have altered functioning of the dopaminergic system, which may be responsible for the clinical signs and symptoms observed in those individuals, such as low energy, low mood and anhedonia, increased impulsivity, as well as impaired cognition and decision making (London et al. 2015; Sabrini et al. 2019). Medications with diverse pharmacological effects increasing the dopaminergic activity have been proposed as candidates for agonist-type treatment of PSUD (Rush and Stoops 2012) and have been evaluated in controlled clinical trials over the past 30 years. An examination of this evidence appears to show that among the most effective of all tested dopaminergic medications are the prescription psychostimulants (PPs), such as amphetamine salts or methylphenidate. PPs have accepted medical use and favorable safety profile and may therefore be good candidate medications for PSUD agonist treatment (Goldstein and Volkow 2011: Goldstein et al. 2010).

Treatment with PPs should be implemented cautiously due to the potential for euphoric effects and the risk of misuse and diversion (Darke and Farrell 2016). Using extended-release preparations suitable for once-daily administration and administering medication under direct observation, similarly to treatment with opioid agonist methadone, can reduce potential of abuse (Nuijten et al. 2016). Toxicity of PPs when given chronically should also be considered (Darke and Farrell 2016) especially as individuals chronically using high doses of stimulants are at increased risk of cardiovascular disorders (Callaghan et al. 2018) and psychotic symptoms (Fiorentini et al. 2011). However, these medications are widely used for the treatment of attention-deficit hyperactivity disorder (ADHD) and have overall good cardiovascular and psychiatric safety profile when used in healthy children and adults (Cooper et al. 2011; Habel et al. 2011).

Many individuals with PSUD have other co-occurring psychiatric and substance use disorders, including ADHD that occurs in 5%–31% of adults seeking treatment for substance use disorders (van de Glind et al. 2014) and major depression which may affect treatment outcomes (Hellem et al. 2015; Rounsaville 2004; Rounsaville et al. 1991). Moreover, patients seeking treatment for PSUD often present with dependence on other drugs, notably opioids (Marsden et al. 2009). These co-occurring conditions may indicate greater severity of the PSUD, and those individuals may have worse health and treatment outcomes, and such comorbidities may modulate the impact of medications.

Cochrane reviews with meta-analyses assessing the efficacy of PPs on CUDs (Castells et al. 2016) and AUDs (Perez-Mana et al. 2013) were published in 2016 and 2013, respectively. Both publications pooled measures of effect from RCTs comparing PPs to placebo on sustained abstinence, drug use across the study period, and retention to treatment. In those analyses, PPs, especially prescription amphetamines, were found to significantly increase rates of sustained abstinence in patients with CUDs (Castells et al. 2016), but the quality of the evidence was very low due to heterogeneity and inconsistency of the estimates with a large number of medications included. However, no significant effect was found in patients with AUD on measures of drug use, sustained abstinence, or retention in treatment, though attrition for both medication and placebo was high (Perez-Mana et al. 2013). More recently, new high-quality trials with robust sample sizes, sound methodology, and up-to-date designs (e.g., monitored intake, extended-release formulations, and higher dosages) have been published, and therefore, a new meta-analysis is warranted.

The aim of this study was to systematically review trials testing selected PPs as treatment for patients with cocaine or amphetamine-type substance use disorder (PSUD) and pool data with meta-analyses. We included studies that used either prescription amphetamines, methylphenidate, or modafinil. We summarized the evidence from the following outcomes: (a) sustained (2–3 weeks) drug abstinence, (b) percentage of drug-negative urine tests across trial, (c) maximum days of continuous abstinence, and (d) retention in treatment. When possible, we carried out subgroup analyses per drug of abuse, medication used, ADHD status, and comorbid opioid use disorder (OUD). We performed risk of bias and quality assessments and assessed the quality of the evidence using the GRADE system.

# Methods

### **Medication selection**

In contrast to previous systematic reviews on this topic (Castells et al. 2016; Perez-Mana et al. 2013), we restricted this analysis to trials with medications that are most analogous to cocaine or amphetamine-type substances, with similar behavioral effects and therefore most "potent" pharmacologically to be candidates for PSUD "agonist-based" treatment. This includes amphetamine isomers (dexamphetamine and methamphetamine), which are monoamine release enhancers, and methylphenidate, which is a monoamine transport blocker (Rush and Stoops 2012). Furthermore, we included modafinil, which also blocks dopamine transporter in doses used clinically exerting effect comparable to lower doses of methylphenidate (Kim et al. 2014; Madras et al. 2006; Volkow et al. 2009) and has a behavioral profile typical of psychostimulants, such as cocaine (Andersen et al. 2010). Despite these effects, modafinil generally has low abuse liability (Jasinski 2000; Jerry et al. 2016), and this combination of behavioral and pharmacological effects as a psychostimulant with limited abuse potential makes modafinil

a worthwhile replacement drug candidate. All of the selected medications have similar behavioral effects as abused stimulants (Rush and Baker 2001) and are therefore on the list of controlled substances. This is in contrast to bupropion, another medication that binds to dopamine transporter but has lower receptor occupancy than modafinil or methylphenidate (Griffith et al. 1983; Learned-Coughlin et al. 2003; Meyer et al. 2002), has lower abuse liability (Griffith et al. 1983), and is not a controlled substance, and was therefore not included in this review.

We have also included in the analysis two publications that evaluated treatment with mixed amphetamine salts combined with topiramate (Levin et al. 2020; Mariani et al. 2012). Topiramate was found to have a positive effect on abstinence from cocaine (Singh et al. 2016), and therefore, it is possible that it had an added benefit; thus, we also evaluated the overall effect with these two studies excluded.

### Search strategy and selection criteria

This review and meta-analysis was conducted following a preestablished protocol registered on PROSPERO under the number CRD42019129653 and following the Preferred Reporting Items for Systematic Reviews and Meta-Analysis (PRISMA) guidelines (Moher et al. 2009). We searched all publications in the following databases for bibliography: MEDLINE, PubMed, the Cochrane Library (Cochrane Database of Systematic Reviews, Cochrane Central Register of Controlled Trials and Cochrane Methodology Register), Web of Science, and ClinicalTrials.gov. Also, the trial's references were examined for additional references. No language restriction was used. Publications published up to September 2019 were included.

Our initial search strategy in PubMed/Medline was ("Prescription Psychostimulants" or modafinil or methylphenidate or dextroamphetamine or d-amphetamine or amphetamine or lisdexamphetamine or "mixed amphetamine salts") and (cocaine or methamphetamine or stimulants). This search strategy was adapted for other databases.

We included randomized, parallel-grouped, double-blind, and placebo-controlled clinical treatment trials that used a PP as the pharmacological intervention. All trials included in this review were conducted in outpatient settings and lasted from 8 to 26 weeks. Non-treatment studies, such as human behavioral pharmacology studies, were not included because of differences in studied population and limited generalizability to clinical population.

RevMan 5 (The Nordic Cochrane Centre 2014) software was used to obtain pooled measures of effect as well as graphic displays of the meta-analytic illustrations. We used Cochrane's Risk of Bias Assessment tool, included in RevMan 5, to generate methodological quality graphs and summaries. We used the GRADE (Grading of Recommendations Assessment, Development, and Evaluation) approach (Balshem et al. 2011), with help of the GRADEpro software (McMaster University (developed by Evidence Prime 2015) to assess the quality of evidence. A summary of findings (SoF) table is provided, and results are described according to GRADE judgment.

### Data extraction and quality assessment

The initial plan for the analysis was developed by AB and GG and finalized with input from all authors. Titles or abstracts of publications obtained with the search strategy or from references list were screened independently by VST and FBA. Full texts were further analyzed by VST and FBA, and final decision about study inclusion included AB and TMF. A standardized form was applied to data extracted from the included publications to assess study quality and evidence synthesis. Information on this form included the following: substance of abuse, presence of comorbid substance use disorders, presence of comorbid mental health disorders, medication tested (with maximum dosages), follow-up period, outcome effect measures, and information about risk of bias. VST and FBA assessed the risk of bias of the included studies separately with the Cochrane tool (Higgins et al. 2011). Any discrepancies were resolved via a discussion with other authors until reaching a consensus. Data were extracted directly from each trial's published results and study protocols, when available. If data for the outcomes of interest were not published or inserted in the protocols, authors were contacted or we used data included in the published Cochrane reviews (Castells et al. 2016; Perez-Mana et al. 2012, 2013).

### Outcomes

We selected sustained stimulant abstinence as the primary outcome. Abstinence, particularly at the end of treatment, is an outcome strongly related to cocaine use during follow-up where it outperforms other measured of drug use, such as percent drug-negative urine screens (Carroll et al. 2014; Miguel et al. 2019), though there are also benefits to sustained low-level use during treatment on functioning following treatment (Roos et al. 2019b). Although most of the publications included in the review used a standardized measure of 3 weeks of consecutive abstinence at the end of the study, some publications used 2 weeks (Levin et al. 2007) or reported sustained abstinence at any time point throughout the trial period (Kampman et al. 2015). All studies except for Nuijten et al. (2016) used objective urine toxicology tests to confirm the abstinence, either alone or combined with selfreport.

As for quantitative measures of abstinence, we measured and pooled percentage of drug-negative urine tests per group throughout trials and maximum days of sustained abstinence, using means and standard deviations to obtain a pooled mean difference. We chose mean difference over standardized mean difference because all studies included assessed the referred outcome using the same unit (percentage of drug-negative urine tests and maximum number of days of continuous abstinence). As a secondary outcome, we also assessed retention to treatment and compared the number of patients who finished the study in treatment and control groups. We decided not to include safety outcomes as earlier reviews found no medication and placebo differences in dropouts due to any adverse events, cardiovascular events, and serious adverse events (Castells et al. 2016; Perez-Mana et al. 2013), and the additional studies included in the present review did not report differences in such outcomes. Funnel plots for the three outcomes were not suggestive of publication bias and are available in the Appendix.

### **Statistical analysis**

For the included studies, we calculated summary measures of intervention efficacy providing risk ratios for dichotomous outcomes (sustained stimulants abstinence and retention in treatment) and mean differences for the continuous outcome (percentage of drug-negative urine tests across trial and maximum days of consecutive abstinence).

For dichotomous outcomes, we selected risk ratios (RR) over risk differences (RD) since relative measures are normally more consistent among different studies than absolute measures (Deeks 2002; Engels et al. 2000). In addition, RR satisfy the following three criteria for summary statistics choice in meta-analyses: consistency, mathematical properties required for a proper meta-analysis, and ease of interpretation (Deeks 2002). Finally, RR are easier to interpret than odds ratios because odds ratios are frequently interpreted as if they were risk ratios, which inappropriately overestimates measures of effects (Viera 2008).

When trials used multiple treatment groups, we merged treatment groups into one when these groups used different doses of the same medication, as recommended by the Cochrane manual (Higgins and Green 2011). One study had the following three treatment groups: modafinil, dexamphetamine, and modafinil + dexamphetamine (Schmitz et al. 2012). In this case, we excluded the modafinil + dexamphetamine group from our comparison and included the other two treatment groups as separate analyses (using the same placebo group), as done previously (Castells et al. 2016).

For the percentage of drug-negative urine tests across trial and maximum days of consecutive abstinence outcomes, we used mean differences (MDs) as all trials provided the same unit (% of drug-negative urine tests and days). Trials differed on methods to impute missing data, but most of them imputed missing as positive. RevMan 5 uses the definition of SMD known as Hedges' g, which is adjusted for possible bias caused by small sample sizes. If means and standard deviations were not reported, we contacted the study authors to obtain these and/or used other statistics to calculate the effect sizes according to the procedures implemented in our metaanalysis software. When necessary, we performed transformations on measures of mean spread to harmonize the results between trials.

If we were not able to obtain the desired data directly from the authors, we then used secondary data obtained from authors by the already published Cochrane meta-analyses cited before (Castells et al. 2016; Perez-Mana et al. 2013). In studies that compared individuals under different doses of the same medication as separate study groups, we merged sample sizes and calculated pooled standard deviations. When publications used different PPs as distinct study groups, we included that publication twice, as if each drug versus placebo comparison represented a single study.

We pooled studies comparing the same types of intervention and control and using the same outcome measure using random-effects models for meta-analysis to account for heterogeneity among the treatment effects of different trials (Borenstein and Higgins 2013). Our a priori tolerated alpha level for effect measures was 0.05. Statistical heterogeneity was assessed using Chi square and  $I^2$  statistics. An  $I^2$  value greater than 50% was regarded as indicative of substantial heterogeneity (Higgins and Green 2011). When dichotomous outcome data were missing, we assumed that patients who dropped out after randomization had a negative outcome. Missing continuous outcome data were analyzed on an endpoint basis, including only participants with a final assessment, as reported by the original study authors.

We also calculated the number needed to treat (NNT) from the measures of effect of the outcomes assessed in our metaanalysis. For risk ratios obtained from dichotomous variables, the computation of the NNT proceeds as follows:  $NNT = \left|\frac{1}{ACR \times (1-RR)}\right|$ , where RR is the risk ratio for each outcome and ACR is the assumed control risk (ACR). There are many ways to set this parameter. We used the approach that divides the number of positive events in the control (placebo) group divided by the total number of events (Higgins and Green 2011).

### Quality of evidence

Quality of evidence was determined using the Grading of Recommendations Assessment, Development, and Evaluation (GRADE) approach (Balshem et al. 2011). "Quality" is defined as a measure of confidence in the effect estimate provided. Evidence is rated as of high, moderate, low, or very low quality. A quality rating of "high" means that it is very likely that the true effect lies close to the estimate, while "very low" means that the true effect is likely to be different from the estimated effect. Randomized trials begin with a quality rating of high, and observational studies begin with a quality rating of low. These ratings can be downgraded based on the following five criteria: risk of bias in the included studies; inconsistency in results; indirectness of evidence; imprecision of effect estimates; and risk of publication bias. However, ratings can be upgraded if the effect size is large, there is evidence of a dose–response effect, or all plausible confounding is controlled for.

## Subgroup analyses

We pre-specified subgroup analysis per drug of abuse (cocaine and methamphetamines) providing sensitivity analysis for results overall and for cocaine use disorder when removing trials with add-on topiramate; treatment drug (prescription amphetamines, modafinil, and methylphenidate); and comorbid OUD and attention deficit hyperactivity disorder (ADHD) status, where prescription psychostimulants are notably effective (De Crescenzo et al. 2017). We considered that these factors could partially explain heterogeneity between trials and that the subpopulations specified could respond differently to treatment (Oxman and Guyatt 1992).

In addition, we defined two dose categories for PPs; low and high. We used maximum doses currently approved by the FDA (60 mg/day for prescription amphetamines, 400 mg/day for modafinil, and 60 mg/day for methylphenidate) as the threshold for separating low (below that limit) vs. high doses. We hypothesize that patients with PSUD may be cross-tolerant to PPs and may thus require higher doses and more potent agents to achieve therapeutic response. To explore whether there is a dose–response effect, we conducted a subgroup analysis for the primary outcome of sustained abstinence by the dose. Finally, we performed sensitivity analyses to (a) exclude trials with immediate-release PPs to evaluate its impact on sustained abstinence and (b) elucidate whether the presence of contingency management (CM) (either abstinence or compliancetargeted) changed the effect of PPs on retention to treatment.

### **Risk of bias**

We assessed the validity of the included studies using criteria from the Risk of Bias Assessment tool, developed by the Cochrane Collaboration (Higgins and Green 2011). The tool analyzes risk of bias by classifying it in the following six different domains: generation of allocation sequence, concealment of treatment allocation, blinding of patients and personnel, blinding of outcome assessors, data incompleteness, selective reporting, and other sources of bias.

The risk of bias for each specific domain is assessed either as "low risk," "unclear risk," or "high risk." We assumed that all of the included trials had unclear risk of bias from blinding of patients and personnel since study medications had behavioral effects that could be noticed by both patients and clinicians. Risk of bias graph and summary with author's judgments for each trial are available in the Appendix.

# Results

### **Study selection**

Of the 164 abstracts retrieved from the initial search, 21 studies were included in the full-text review stage. After screening of reference lists from included original publications and reviews, 31 additional studies were added to the full-text reading stage (Fig. 1). At this initial stage, we excluded studies that did not include outcomes of interest or did not meet methodological requirements. A total of 16 studies were excluded after full-text review. Initial discrepancies at the abstract and full-text review stages were later resolved by consensus of study authors.

Thirty-eight RCTs were included in the final analysis with 26 trials conducted in patients with cocaine use disorder (CUD) and 12 in patients with amphetamine-type use disorder (AUD). Eighteen trials used as an outcome 3-week sustained abstinence and one trial assessed 2-week sustained abstinence. Out of those, seven trials evaluated prescription amphetamines, eight evaluated modafinil, and four evaluated meth-ylphenidate. Five trials included patients with comorbid substance use disorders, and four trials included patients with ADHD.

Characteristics of all trials included are described in Table 1.

# Effect of prescription psychostimulants on abstinence: overall and in patients with CUD vs AUD

We found an overall significant benefit of PPs when compared to placebo on promoting 2–3 weeks of sustained abstinence (risk ratio [RR] = 1.45, 95% confidence interval [CI] (1.10, 1.92),  $l^2 = 37\%$ ); NNT = 16, 95% CI (8, 70) (Fig. 2). When analyzing subgroups per drug of abuse (CUD vs. AUD), this benefit is dragged away from the null by CUD studies. The effect in CUD studies is not only statistically significant but also clinically meaningful considering both confidence interval bounds (RR = 1.70, 95% CI (1.26, 2.31),  $l^2 = 24\%$ ); NNT = 12, 95% CI (7, 32). No benefit is shown for AUD studies (RR = 0.89, 95% CI (0.62, 1.27),  $l^2 = 0\%$ ).

We performed sensitivity analysis to evaluate whether the benefit of PPs for PSUDs remained when removing two trials that administered topiramate as an add-on medication, both for the treatment of CUDs and conducted by the same research team (Levin et al. 2020; Mariani et al. 2012). The overall effect of PPs for PSUDs remained statistically significant (RR = 1.34, 95% CI (1.01, 1.79),  $I^2 = 34\%$ ); NNT = 20, 95% CI (9, 660). The effect of PPs for CUDs also remained

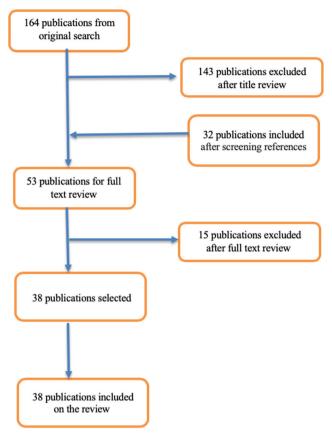


Fig. 1. Studies included in the review with number of studies included after each stage of the screening process

statistically significant (RR = 1.59, 95% CI (1.14, 2.23),  $l^2$  = 26%); NNT = 14, 95% CI (7, 58), compared to RR = 1.70, 95% CI (1.26, 2.31),  $l^2$  = 24%); NNT = 12, 95% CI (7, 32) when the studies were included. See the Appendix for details.

We conducted an additional analysis of studies that reported abstinence in the final 3 weeks, which is an outcome measure adopted in the recent trials, and four trials were included (Dackis et al. 2012; Kampman et al. 2015; Levin et al. 2015b; Levin et al. 2020). We found a significant effect of PP abstinence with a very high RR (RR = 3.01, 95% CI (1.58, 5.75),  $l^2 = 0\%$ ). See the Appendix for details.

We performed sensitivity analysis excluding a trial with immediate-release methylphenidate for CUD (Dursteler-MacFarland et al. 2013) from the overall and drug-specific analyses. The overall effect of PPs for PSUD remained statistically significant (RR = 1.47, 95% CI (1.10, 1.96),  $l^2 = 41\%$ ); NNT = 18, 95% CI (9, 81). The effect of PPs for CUDs also remained statistically significant (RR = 1.74, 95% CI (1.27, 2.39),  $l^2 = 28\%$ ); NNT = 11, 95% CI (6, 30).

### Effect by the medication type

A subgroup analysis by medication shows a clinically and statistically significant effect for prescription amphetamines

(see Fig. 3), when compared to placebo (RR = 2.44, 95% CI (1.66, 3.58),  $I^2 = 0\%$ ); NNT = 7, 95% CI (4, 14). There was no effect for modafinil (RR = 1.22, 95% CI (0.83, 1.77),  $I^2 = 29\%$ ) or methylphenidate (RR = 0.90, 95% CI (0.60, 1.37),  $I^2 = 0\%$ ).

### The effect of co-occurring disorders

To assess whether patients with co-occurring ADHD or OUD had a different response to PPs, we conducted separate analyses for trials where patients had a co-occurring disorder vs. those that did not (Fig. 4). We found a significant benefit of PPs in trials that did not report an ADHD diagnosis (RR = 1.55, 95% CI (1.14, 2.11),  $I^2 = 33\%$ ); NNT = 14, 95% CI (7, 53), while no benefit was observed in trials that included patients with comorbid ADHD (RR = 1.17, 95% CI (0.61, 2.25),  $I^2 = 48\%$ ). When restricting analyses to prescription amphetamines, there was a significant benefit in the non-ADHD group (RR = 2.33, 95% CI (1.55, 3.51),  $I^2 = 0\%$ ); NNT = 7, 95% CI (4, 15). We did not conduct this analysis for the ADHD group as there was only one trial (Levin et al. 2015b) with these characteristics.

To assess the impact of the co-occurring OUD, we separately evaluated three trials where patients also had an OUD and were treated with an opioid agonist (methadone or diacetylmorphine) and those that did not include co-occurring OUD (Fig. 5). The effect of psychostimulants in trials with comorbid OUD was robust (RR = 2.03, 95% CI (1.24, 3.33),  $l^2 = 0\%$ ); NNT = 8, 95% CI (4, 32) while there was no benefit of PPs in trials without co-occurring OUD (RR = 1.34, 95% CI (0.98, 1.83),  $l^2 = 39\%$ ). However, when restricting analyses to prescription amphetamines, the statistically significant results were detected in both the OUD+ (RR = 2.41, 95% CI (1.39, 4.17),  $l^2 = 0\%$ ); NNT = 6, 95% CI (3, 21) and the OUD– (RR = 2.46, 95% CI (1.43, 4.24),  $l^2 = 0\%$ ); NNT = 6, 95% CI (3, 19) groups.

# The effect of the dose

To assess the impact of the PP dose, we separately evaluated trials that used low doses and those that used high doses. Four of the 17 trials included used dose of PPs that are lower than FDA's maximum recommended doses (for approved conditions) while 15 used the maximum doses or higher. One trial (Dackis et al. 2012) tested two different doses of modafinil, a low 200 mg/d dose and a high 400 mg/d dose, and we included effect estimates of each of those groups compared to placebo separately. One trial (Levin et al. 2015b) tested two doses of dexamphetamine, 60 and 80 mg/d, but since both were in the high dose range, we merged both treatment groups into a same comparison as recommended by the Cochrane handbook (Higgins and Green 2011).

Table 1 Characte	Characteristics of included studies	ncluded s	studies											
Author, year	Primary disorder	Other SUDs <sup>a</sup>	Other psychiatric conditions	Primary medication	Other medications	Slow release	Psychosocial treatment	Maximum daily dose	Sample size	Follow-up period (weeks)	Outcome: continuous abstinence	Outcome: maximum days of continuous abstinence	Outcome: retention in treatment	Outcome: drug- negative urine tests throughout trial
Grabowski et al. (1994)	Cocaine	No	None	MPH <sup>b</sup>	None	No	None	45	L	10	No	No	Yes	No
Grabowski et al. (1997)	Cocaine	No	None	MPH	None	No	CBT—all groups	45	49	11	No	No	Yes	Yes
Grabowski et al. (2001)	Cocaine	No	None	D-AMP <sup>c</sup>	None	Yes	CBT—all groups	60	128	12	No	No	Yes	No
Schubiner et al. (2002)	Cocaine	No	ADHD <sup>d</sup>	MPH	None	No	CBTall groups	06	48	12	No	No	Yes	Yes
Shearer et al. (2003)	Cocaine	Opioid	None	D-AMP	None	Yes	None	60	30	14	Yes	No	Yes	Yes
Grabowski et al. (2004)	Cocaine	Opioid	None	D-AMP	Methadone	Yes	CBT—all groups	60	120	26	Yes	No	Yes	Yes
Dackis et al. (2005)	Cocaine	No	None	MOD <sup>e</sup>	None	Yes	CBT-all groups	400	62	8	Yes	No	Yes	No
Levin et al. (2006)	Cocaine	Opioid	ADHD	MPH	Methadone	Yes	CBT—all groups	80	98	12	Yes	No	No	No
Tiihonen et al. (2007)	$AMP^{f}$	No	None	HdM	None	Yes	Unstructured—all	54	34	20*	No	No	Yes	No
Levin et al. (2007)	Cocaine	No	ADHD	HdM	None	Yes	groups CBT—all groups	09	106	14	Yes	No	Yes	Yes
Shearer et al. (2009)	Meth <sup>g</sup>	No	None	MOD	None	Yes	CBT—all groups	200	80	10	No	No	Yes	No
Anderson et al. (2009)		No	None	MOD	None	Yes	CBT-all groups	400	210	12	Yes	Yes	Yes	No
Mooney et al. (2009)	Cocaine	No	None	Meth	None	Used	CBTall groups	30	82	8	No	No	Yes	No
Longo et al. (2010)	Meth	No	None	D-AMP	None	poun Yes	CBT-all groups	110	49	12	No	No	Yes	No
Galloway et al. (2011)		No	None		None	Yes	Motivational	09	60	8	No	No	Yes	Yes
							therapy—all groups							
Konstenius et al. (2010)	AMP	No	ADHD	HdM	None	Yes	Individual skills training	72	24	12	Yes	No	Yes	Yes
							program—all							
Heinzerling et al. (2010)	Meth	No	None	MOD	None	Yes	CBT + abstinence CM—all groups	400	71	12	Yes	No	Yes	Yes
Anderson et al. (2012)	Meth	No	None	MOD	None	Yes	CBT-all groups	400	210	12	Yes	Yes	Yes	Yes
Dackis et al. (2012)	Cocaine	No	None	MOD	None	Yes	CBT—all groups	200/400	210	8	Yes	No	Yes	No
Mariani et al. (2012)	Cocaine	No	None	ERMS-AMP <sup>h</sup>	Topiramate	Yes	Compliance and abstinence CM—	60	81	12	Yes	No	Yes	Yes
								001001	č	2	;		;	
Melan at al (2012)	<b>Module</b>	N0	None	U-AMP	None	Y es	CB1—all groups	400/60	6/ 0F	10	Yes	No	Yes	No
Milles et al. (2013)	Mem	N	INONE	HIM	lvone	I CS	supervised make— all groups	4C	0/	77	INO	INO	I CS	0N
Dursteler-MacFarland et al. (2013)	Cocaine	Opioid	None	HdM	Diacetylmorphine	No	Supervised intake— all groups; group	60	62	12	Yes	Yes	Yes	No
Konstenius et al.	Meth	None	ADHD	HdM	None	Yes	CBT—two attits CBT—all groups	180	54	26	No	No	Yes	No
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Outcome: drug- negative urine tests throughout trial	No	No	No	No	Yes	No	Yes	Yes		Yes	Yes	No	No	No	No		
Outcome: retention in treatment	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes		Yes	Yes	Yes	Yes	Yes	No		
Outcome: maximum days of continuous abstinence	No	Yes	No	No	No	No	Yes	Yes		No	No	No	No	No	Yes		
Outcome: continuous abstinence	Yes	No	No	Yes	Yes	No	No	Yes		Yes	No	No	No	No	Yes		
Follow-up period (weeks)	12	10	10	8	13	14	9	12		14	13	10	24	8	12		
Sample size	40	110	56	94	123	43	57	73		127	164	6	51	123	91		
Maximum daily dose	400	54	54	300	60/80	70	400	09		60	400	400	200/400	400	400		
Psychosocial treatment	CBT + compliance CM—all groups	CBT + abstinence CM—all groups	Supervised intake— all groups	CBT + compliance CM—all groups	CBT-all groups	CBT-all groups	CBT + attendance	CM—all groups Supervised intake—	all groups	CBT + compliance CM—all groups	CBT-all groups	Psychotherapy	Not mentioned	CBT—all groups	CBT—all groups + abstinence CM—	two arms	
Slow release	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes		Yes	Yes	Yes	Yes	Yes	Yes		
Other medications	Levodopa/ carbidopa, 5naltrexone	None	None	None	None	None	None	Methadone/	Diacetylmorphine	Topiramate	Naltrexone	None	Citalopram	None	Methadone		
Primary medication	MOD	HdM	HdM	MOD	D-AMP	LD-AMP <sup>i</sup>	MOD	D-AMP		D-AMP	MOD	MOD	MOD	MOD	MOD		
Other psychiatric conditions	None	None	None	None	ADHD	None	None	None		None	None	None	None	None	None		
Other SUDS <sup>a</sup>	No	No	No	No	No	No	No	No		No	Alcohol	No	Opioid	No	Opioid		
Primary disorder	Cocaine	Meth	Meth	Cocaine	Cocaine	Cocaine	Cocaine	Crack	cocaine	Cocaine	Cocaine	Meth	Cocaine	Cocaine	Cocaine		se disorde:
Author, year	Schmitz et al. (2014) Cocaine	Ling et al. (2014)	Rezaei et al. (2015)	Kampman et al. (2015)	Levin et al. (2015a)	Mooney et al. (2015)	Morgan et al. (2016)	Nuijten et al. (2016)		Levin et al. (2020)	NCT00142818	NCT00859573	NCT00218036	NCT00218387	NCT00838981		<sup>a</sup> SUD, substance use disorder

<sup>b</sup> MPH, methylphenidate

<sup>c</sup> D-AMP, dexamphetamine

<sup>d</sup> ADHD, attention-deficit-hyperactivity disorder

<sup>e</sup> MOD, modafinil

<sup>f</sup> AMP, amphetamines

<sup>g</sup> Meth, methamphetamines

<sup>h</sup> ERMS-AMP, extended-release mixed amphetamine salts

<sup>i</sup> LD-AMP, lisdexamphetamine \* Early terminated

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	Psychostim	ılants	Place	bo		Risk Ratio		Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	Year	M-H, Random, 95% Cl
1.1.1 Cocaine								
Shearer 2003	7	16	4	14	5.2%	1.53 [0.56, 4.15]	2003	3
Grabowski 2004	24	54	7	40	7.4%	2.54 [1.22, 5.30]	2004	۱
Dackis 2005	10	30	4	32	4.9%	2.67 [0.94, 7.60]	2005	5 +
Levin 2006	3	21	2	13	2.4%	0.93 [0.18, 4.84]	2006	i <u> </u>
Levin 2007	8	53	9	53	6.1%	0.89 [0.37, 2.13]	2007	·
Anderson 2009	22	138	7	72	6.8%	1.64 [0.74, 3.65]	2009	) +•
Schmitz 2012	2	22	1	8	1.4%	0.73 [0.08, 6.97]	2012	2
Schmitz 2012	1	20	1	8	1.0%	0.40 [0.03, 5.65]	2012	2
Dackis 2012	11	135	4	75	4.5%	1.53 [0.50, 4.63]	2012	2
Mariani 2012	13	39	7	42	6.7%	2.00 [0.89, 4.49]	2012	2 +
Dürsteler-MacFarland 2013	3	30	3	32	2.8%	1.07 [0.23, 4.88]	2013	3
Schmitz 2014	9	22	10	18	8.4%	0.74 [0.38, 1.41]	2014	↓ <u> </u>
Kampman 2015	11	47	4	47	4.7%	2.75 [0.94, 8.02]	2015	5 +
Levin 2015	20	83	3	43	4.2%	3.45 [1.09, 10.98]	2015	5
Nuijten 2016	11	38	2	35	3.0%	5.07 [1.21, 21.27]	2016	i
Levin 2019	14	64	4	63	4.8%	3.45 [1.20, 9.90]	2018	
Subtotal (95% CI)		812		595	74.4%	1.70 [1.26, 2.31]		◆
Total events	169		72					
Heterogeneity: Tau <sup>2</sup> = 0.09; Cl	hi² = 19.85, df =	= 15 (P =	0.18); l² =	= 24%				
Test for overall effect: Z = 3.44	(P = 0.0006)							
1.1.2 Meth								
Heinzerling 2010	9	34	10	37	7.1%	0.98 [0.45, 2.12]	2010	) —
Konstenius 2010	8	12	9	12	10.1%	0.89 [0.53, 1.49]	2010	)
Anderson 2012	21	142	12	68	8.4%	0.84 [0.44, 1.60]	2012	2
Subtotal (95% CI)		188		117	25.6%	0.89 [0.62, 1.27]		<b>•</b>
Total events	38		31					
Heterogeneity: Tau <sup>2</sup> = 0.00; Cl	hi² = 0.09, df = 1	2 (P = 0.9	95); I <sup>2</sup> = 0	%				
Test for overall effect: Z = 0.63	(P = 0.53)							
Total (95% CI)		1000		712	100.0%	1.45 [1.10, 1.92]		◆
Total events	207		103					
Heterogeneity: Tau <sup>2</sup> = 0.13; Cl		= 18 (P =		= 37%				
Test for overall effect: Z = 2.61		,						0.01 0.1 1 10 100 Favours Placebo Favours Psychostimulants
Test for subgroup differences	: Chi <sup>2</sup> = 7.32, d	f=1 (P=	0.007), 1	<sup>2</sup> = 86.3	3%			ravours riacebo Favours Esychosumulants

Fig. 2. Overall and by dependence drug effect of prescription psychostimulants compared to placebo for outcome sustained abstinence

Trials that use low doses did not show benefit on promoting sustained abstinence (RR = 1.25; 95% CI [0.71, 2.21]), with low heterogeneity ( $I^2 = 0\%$ ). Trials with that used high doses showed a statistically significant benefit (RR = 1.50; 95% CI [1.10, 2.06]); NNT = 14, 95% CI (7, 67); however, the clinical benefit was marginal as the confidence interval's lower bound was not considered clinically meaningful and heterogeneity was intermediate ( $I^2 = 44\%$ ) (see Fig. 6). On the other hand, there is moderate quality evidence for the benefit of PPs in maximum higher dosages when restricting analyses to CUDs (RR = 1.95; 95% CI [1.38, 2.77]); NNT = 9, 95% CI (5, 22), with a clinically meaningful CI lower bound and low heterogeneity ( $I^2 = 30\%$ ) (see Fig. 7).

### Percentage of drug-negative urine tests across trials

We analyzed data from 15 trials to compare the percentage of drug-free urine tests across trials in PPs vs. placebo groups (see Fig. 8). We found a significant difference of 2.40% (95% CI (0.07, 4.73),  $I^2 = 29\%$ ) favoring PPs when compared to placebo.

As the main effects found in the dichotomous abstinence analyses were from prescription amphetamines for CUD, we conducted a separate analysis for this subgroup (see Fig. 9). We found a more robust difference of 8.37% (95% CI (3.75, 12.98),  $l^2 = 0\%$ ) favoring PPs.

### Maximum days of consecutive abstinence

We gathered data from seven trials to compare average maximum days of consecutive abstinence between PPs and placebo groups (see Fig. 10). Due to the scarcity of studies available, subgroup analyses were not possible. We found a significant difference of 3.34 (95% CI (1.06, 5.62),  $l^2 = 41\%$ ) days favoring psychostimulants when compared to placebo.

### **Retention in treatment**

All but one of the 38 trials included in this review included the outcome of retention in treatment (see Fig. 11). We did not find a significant benefit of psychostimulants when compared to placebo on promoting retention to treatment in overall stimulants use disorders (RR = 1.04, 95% CI (0.97, 1.11),  $I^2$  = 10%), CUD (RR = 1.03, 95% CI (0.96, 1.11),  $I^2$  = 7%) and AUD (RR = 1.08, 95% CI (0.93, 1.27),  $I^2$  = 22%). No differences were found when analyzing subgroups by medication drug, ADHD status, or comorbid dependences. We also compared retention in treatment between treatment groups using number of days in treatment as outcome. Again, we did not find a statistically significant difference between PPs and placebo (SMD = 0.11, 95% CI (-0.27, 0.50),  $I^2$  = 65%). See the Appendix for more details.

	Psychostimu	lants	Place	bo		Risk Ratio	Risk Ratio
Study or Subgroup	Events				Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
1.3.1 Prescription Amphetam	nines						
Grabowski 2004	24	54	7	40	7.4%	2.54 [1.22, 5.30]	
Levin 2015	20	83	3	43	4.2%	3.45 [1.09, 10.98]	
Levin 2019	14	64	4	63	4.8%	3.45 [1.20, 9.90]	
Mariani 2012	13	39	7	42	6.7%	2.00 [0.89, 4.49]	
Nuijten 2016	11	38	2	35	3.0%	5.07 [1.21, 21.27]	
Schmitz 2012	2	22	1	8	1.4%	0.73 [0.08, 6.97]	
Shearer 2003	7	16	4	14	5.2%	1.53 [0.56, 4.15]	
Subtotal (95% CI)		316		245	32.8%	2.44 [1.66, 3.58]	•
Total events	91		28				
Heterogeneity: Tau <sup>2</sup> = 0.00; Cł	hi² = 4.05, df = 8	6 (P = 0.	67); I <sup>2</sup> = 0	%			
Test for overall effect: Z = 4.52	(P < 0.00001)						
1000 1 5 3							
1.3.2 Modafinil							
Anderson 2009	22	138	7	72	6.8%	1.64 [0.74, 3.65]	
Anderson 2012	21	142	12	68	8.4%	0.84 [0.44, 1.60]	
Dackis 2005	10	30	4	32	4.9%	2.67 [0.94, 7.60]	•
Dackis 2012	11	135	4	75	4.5%	1.53 [0.50, 4.63]	
Heinzerling 2010	9	34	10	37	7.1%	0.98 [0.45, 2.12]	
Kampman 2015	11	47	4	47	4.7%	2.75 [0.94, 8.02]	
Schmitz 2012	1	20	1	8	1.0%	0.40 [0.03, 5.65]	
Schmitz 2014	9	22	10	18	8.4%	0.74 [0.38, 1.41]	
Subtotal (95% CI)		568		357	45.8%	1.22 [0.83, 1.77]	<b>—</b>
Total events	94		52				
Heterogeneity: Tau <sup>2</sup> = 0.08; Cl		7 (P = 0)	20); l <sup>2</sup> = 2	9%			
Test for overall effect: Z = 1.02	(P = 0.31)						
1.3.3 Methylphenidate							
Dürsteler-MacFarland 2013	3	30	3	32	2.8%	1.07 [0.23, 4.88]	
Konstenius 2010	8	12	9	12	10.1%	0.89 [0.53, 1.49]	
Levin 2006	3	21	2	13	2.4%	0.93 [0.18, 4.84]	
Levin 2007	8	53	9	53	6.1%	0.89 [0.37, 2.13]	
Subtotal (95% CI)	0	116	5	110	21.4%	0.90 [0.60, 1.37]	•
Total events	22		23				
Heterogeneity: Tau <sup>2</sup> = 0.00; Cl		3(P = 1)		%			
Test for overall effect: Z = 0.48		s (i − i.	00),1 = 0	~			
Total (95% CI)		1000		712	100.0%	1.45 [1.10, 1.92]	•
Total events	207		103				-
Heterogeneity: Tau <sup>2</sup> = 0.13; Cl		18 (P =		= 37%			
Test for overall effect: Z = 2.61							0.01 0.1 1 10 100
Test for subgroup differences		df = 2 P	= 0.002)	$l^2 = 84$	3%		Favours Placebo Favours Psychostimulants

Test for subgroup differences: Chi<sup>2</sup> = 12.76, df = 2 (P = 0.002), l<sup>2</sup> = 84.3%

Fig. 3. Overall and by treatment drug effect of prescription psychostimulants compared to placebo for outcome sustained abstinence

We conducted a separate analysis to assess whether providing CM to the patients undergoing treatment with PPs would modify their effect on retention to treatment. The retention in treatment was comparable in the studies with CM vs. those that did not use CM (61.4% vs. 52.0%, t = 1.09, p =0.28). In the trials where CM was provided for all study participants, PPs were not efficacious on promoting retention to treatment vs. placebo (RR = 0.95, 95% CI (0.84, 1.07),  $I^2 =$ 0%). This analysis was included in the Appendix.

#### Medication adherence

We conducted an analysis to address the possible impact of medication adherence on abstinence. Of the 13 trials that included sustained abstinence outcome, all reported overall medication adherence mostly using direct pill count. Medication adherence ranged from 51% to 99% but was generally high with a median of 91%, and we did a median split to compare abstinence in trials with adherence below vs. above the median. Trials above the median had significantly higher

rates of abstinence (RR = 1.74, CI = (1.01, 3.00)) as compared to those below the median (RR = 1.26, CI = (0.83, 1.91)). This analysis is included in the Appendix as Fig. 10.

### Grade

We conducted GRADE evaluation of the quality of evidence. Evidence was of moderate quality for the following: (1) benefit of prescription amphetamines for treatment of patients with CUD; (2) benefit of PPs for treatment of patients with CUD and comorbid heroin dependence; (3) benefit of higher doses of PPs for treatment of patients with CUD; (4) benefit of prescription amphetamines for CUD patients with and without comorbid opioid use disorder; and (5) benefit of prescription amphetamines for patients with PSUD and no ADHD.

We found low-quality evidence of the benefit of prescription amphetamines on increasing cocaine-negative urines across trials.

Evidence was of very low-quality for the following: (1) overall benefit of PPs on promoting sustained

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	Psychostim	lants	Place	bo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
1.21.1 No ADHD							
Anderson 2009	22	138	7	72	6.8%	1.64 [0.74, 3.65]	
Anderson 2012	21	142	12	68	8.4%	0.84 [0.44, 1.60]	<b>_</b>
Dackis 2005	10	30	4	32	4.9%	2.67 [0.94, 7.60]	
Dackis 2012	11	135	4	75	4.5%	1.53 [0.50, 4.63]	<del>,</del>
Dürsteler-MacFarland 2013	3	30	3	32	2.8%	1.07 [0.23, 4.88]	
Grabowski 2004	24	54	7	40	7.4%	2.54 [1.22, 5.30]	
Heinzerling 2010	9	34	10	37	7.1%	0.98 [0.45, 2.12]	
Kampman 2015	11	47	4	47	4.7%	2.75 [0.94, 8.02]	
Levin 2019	14	64	4	63	4.8%	3.45 [1.20, 9.90]	
Mariani 2012	13	39	7	42	6.7%	2.00 [0.89, 4.49]	<b></b>
Nuijten 2016	11	38	2	35	3.0%	5.07 [1.21, 21.27]	
Schmitz 2012	2	22	1	8	1.4%	0.73 [0.08, 6.97]	
Schmitz 2012	1	20	1	8	1.0%	0.40 [0.03, 5.65]	
Schmitz 2014	9	22	10	18	8.4%	0.74 [0.38, 1.41]	
Shearer 2003	7	16	4	14	5.2%	1.53 [0.56, 4.15]	
Subtotal (95% CI)		831		591	77.1%	1.55 [1.14, 2.11]	◆
Total events	168		80				
Heterogeneity: Tau <sup>2</sup> = 0.12; C		= 14 (P =	0.10); l²:	= 33%			
Test for overall effect: Z = 2.77	' (P = 0.006)						
1.21.2 ADHD							
Konstenius 2010	8	12		12	10.1%	0.89 [0.53, 1.49]	
Levin 2006	8	21	9 2	13	2.4%	• • •	
Levin 2008 Levin 2007	3	53				0.93 [0.18, 4.84]	
Levin 2007 Levin 2015	20	53 83	9 3	53 43	6.1% 4.2%	0.89 [0.37, 2.13] 3.45 [1.09, 10.98]	
Subtotal (95% CI)	20	169	3	121	4.2 % 22.9%	1.17 [0.61, 2.25]	
Total events	39	105	23	121	22.370	1.17 [0.01, 2.25]	
Heterogeneity: Tau <sup>2</sup> = 0.21; C		2/0 - 0		00%			
Test for overall effect: Z = 0.48		3 (F = 0.	12),1 - 4	0 70			
Testion overall ellect. Z = 0.40	(r = 0.03)						
Total (95% CI)		1000		712	100.0%	1.45 [1.10, 1.92]	◆
Total events	207		103				
Heterogeneity: Tau <sup>2</sup> = 0.13; C	hi² = 28.77, df =	= 18 (P =	0.05); l <sup>2</sup> :	= 37%			0.01 0.1 1 10 100
Test for overall effect: Z = 2.61	(P = 0.009)						0.01 0.1 1 10 100 Favours Placebo Favours Psychostimulants
Test for subgroup differences	: Chi² = 0.57, d	f=1 (P=	: 0.45), l²	= 0%			

Fig. 4. Overall and by ADHD status effect of prescription psychostimulants compared to placebo for outcome sustained abstinence

abstinence from stimulants; (2) benefit of PPs on increasing drug-negative urine tests across study; (3) and benefit of PPs for patients with PSUD on promoting maximum days of continuous abstinence. It is unlikely that PPs will have beneficial effect on retention in treatment in patients with PSUD.

The level of evidence for RCTs was downgraded twice for all outcomes due to risk of bias caused by high attrition rates in most of the studies and by potential detection bias due to the behavioral effects of the medication that could hinder blinding. Approximately half of the trials included had significantly more behavioral side effects (most frequently sleep disturbances, anxiety, headache, and dizziness) in the treatment group compared to placebo, and the few trials that evaluated blinding directly had mixed findings. No studies included an active control which might have minimized the risk of detection bias. Downgrades were made also due to inconsistency and imprecision. Upgrades in the level of evidence judgment were made solely based on large effects (RR > 2.0). GRADEPro objective criteria were used to assess quality of evidence of each of the pre-specified outcomes and subgroup analyses. See the summary of findings (SoF) for further information. A summary of the main findings can be seen in Table 2.

### Discussion

This meta-analysis found that prescription psychostimulants likely promote sustained drug abstinence and may reduce stimulant use throughout trial and extend duration of abstinence when used in treatment of individuals with PSUDs. The overall effect is primarily influenced by studies that used prescription amphetamines, mostly dextroamphetamine, for treatment of individuals with cocaine use disorder. The present analysis offers preliminary evidence that medications with a more "potent" agonist effect (i.e., dextroamphetamine) are more effective than medications that are less "potent" (i.e., modafinil) and that patients treated with higher doses of agonist medications benefit more than patients treated with lower doses, further supporting the hypothesis that the "agonist effect" is mainly responsible for the clinical benefit.

The quality of evidence that supports results of this metaanalysis varies depending on the medication, condition, and outcome. A moderate quality evidence supports the large benefits of using prescription amphetamines and the benefit of higher doses of prescription psychostimulants when used in the treatment of individuals with cocaine use disorder where the achievement of sustained abstinence is the desired treatment outcome.

The results of this study are consistent with and further extend results of prior meta-analyses published by the Study or Subgroup

Anderson 2009

Anderson 2012

Heinzerling 2010

Kampman 2015

Konstenius 2010

Dackis 2005

Dackis 2012

Levin 2007

Levin 2015 Levin 2019

Mariani 2012

Schmitz 2012

Schmitz 2012

Schmitz 2014

Subtotal (95% CI)

1.23.1 No comorbid Dependence

Psychostim	ulants	Place	bo		Risk Ratio	Risk Ratio
Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% CI
ence						
22	138	7	72	6.8%	1.64 [0.74, 3.65]	
21	142	12	68	8.4%	0.84 [0.44, 1.60]	
10	30	4	32	4.9%	2.67 [0.94, 7.60]	
11	135	4	75	4.5%	1.53 [0.50, 4.63]	
9	34	10	37	7.1%	0.98 [0.45, 2.12]	
11	47	4	47	4.7%	2.75 [0.94, 8.02]	
8	12	9	12	10.1%	0.89 [0.53, 1.49]	<b>_</b>
8	53	9	53	6.1%	0.89 [0.37, 2.13]	
20	83	3	43	4.2%	3.45 [1.09, 10.98]	
14	64	4	63	4.8%	3.45 [1.20, 9.90]	
13	39	7	42	6.7%	2.00 [0.89, 4.49]	
1	20	1	8	1.0%	0.40 [0.03, 5.65]	
2	22	1	8	1.4%	0.73 [0.08, 6.97]	
9	22	10	18	8.4%	0.74 [0.38, 1.41]	
	841		578	79.2%	1.34 [0.98, 1.83]	•

Total events 159 85 Heterogeneity: Tau<sup>2</sup> = 0.13; Chi<sup>2</sup> = 21.44, df = 13 (P = 0.06); I<sup>2</sup> = 39% Test for overall effect: Z = 1.83 (P = 0.07)

1.23.2 Comorbid Heroin Depend	dence										
Dürsteler-MacFarland 2013	3	30	3	32	2.8%	1.07 [0.23, 4.88]			+		
Grabowski 2004	24	54	7	40	7.4%	2.54 [1.22, 5.30]					
Levin 2006	3	21	2	13	2.4%	0.93 [0.18, 4.84]					
Nuijten 2016	11	38	2	35	3.0%	5.07 [1.21, 21.27]					
Shearer 2003	7	16	4	14	5.2%	1.53 [0.56, 4.15]					
Subtotal (95% CI)		159		134	20.8%	2.03 [1.24, 3.33]			•		
Total events	48		18								
Heterogeneity: Tau <sup>2</sup> = 0.00; Chi <sup>2</sup>	= 3.84, df = 4	(P = 0.43)	$  ^2 = 09$	Ж							
Test for overall effect: Z = 2.83 (F	° = 0.005)										
Total (95% CI)		1000		712	100.0%	1.45 [1.10, 1.92]			•		
Total events	207		103								
Heterogeneity: Tau <sup>2</sup> = 0.13; Chi <sup>2</sup>	= 28.77, df =	18 (P = 0.0	05); I² =	37%			0.01	01	1	10	100
Test for overall effect: Z = 2.61 (F	° = 0.009)						0.01	Favours Placebo	Favours P		
Test for subgroup differences: C	;hi² = 1.99, df	'= 1 (P = 0.	16), I <sup>2</sup> =	= 49.79	Х6			1 4104/01 140000	1 4154151	0,01000000	

Fig. 5. Overall and by comorbid dependence status effect of prescription psychostimulants compared to placebo for outcome sustained abstinence

Cochrane editorial group of PP efficacy in the treatment of cocaine (Castells et al. 2016) and amphetamine use disorders (Perez-Mana et al. 2013). Similar to Castells et al. (2016), we found that treatment with prescription amphetamines improved sustained abstinence in patients with cocaine use disorder, and we found similar results analyzing subgroups per medication, ADHD status, and concurrent opioid use disorders. We further show that PPs increase maximum days of sustained abstinence from stimulants, though with a small number of studies. This is particularly relevant in light of the association between sustained abstinence and maximum days of abstinence during treatment for cocaine use disorder and the decreased cocaine use and better functioning in the long term (Carroll et al. 2014). However, contrary to the previous reviews, this new metaanalysis shows a significant benefit of PPs on reducing drug use across trial period. This efficacy is more pronounced in trials with prescription amphetamines for patients with CUD.

Unlike studies done by the Cochrane investigators, we evaluated a more restricted range of psychostimulant medications. We limited them to prescription amphetamines, methylphenidate, and modafinil, all controlled substances, to rule out medications with insufficient dopaminergic potency, thereby ensuring a clinically meaningful "agonist effect" (Herin et al. 2010). Among those, the results of trials with modafinil were consistently disappointing, and this may not be a PP with promise on promoting abstinence of cocaine and amphetamines. Our publication included studies published after the most recent review (Castells et al. 2016), most of which used higher doses and extended-release preparations of PPs (Levin et al. 2020; Nuijten et al. 2016). Both of these features had been recommended for clinical use by recent trials (Mariani et al. 2012) and reviews (Mariani and Levin 2012; Rush and Stoops 2012) on this topic.

Similar to Pérez-Mañá and colleagues (2013), we found no effect of PPs on promoting sustained amphetamine abstinence, which is expected since none of the trials published after that review used sustained amphetamine abstinence as an outcome, though we evaluated fewer medications. It should be noted however that earlier studies with methylphenidate used formulations with poor bioavailability which may account for negative findings (Levin et al. 2007; Levin et al. 2006) or used lower, less-effective doses (Konstenius et al. 2010). In addition, there are studies which showed beneficial effects of methylphenidate (Konstenius et al. 2014; Ling et al. 2014; Tiihonen et al. 2007) but did not report sustained abstinence as an outcome and were therefore not included in this analysis. Results of future studies that use improved methodology (Ezard et al. 2018) may change the overall assessments of PP effectiveness in patients with AUD. We found low-

	Psychostim	lants	Place	bo		<b>Risk Ratio</b>	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
1.26.1 Submaximum							
Dackis 2012	3	65	4	75	2.9%	0.87 [0.20, 3.72]	
Dürsteler-MacFarland 2013	3	30	3	32	2.7%	1.07 [0.23, 4.88]	
Kampman 2015	11	47	4	47	4.6%	2.75 [0.94, 8.02]	
Levin 2007	8	53	9	53	6.0%	0.89 [0.37, 2.13]	
Subtotal (95% CI)		195		207	16.1%	1.25 [0.71, 2.21]	<b>•</b>
Total events	25		20				
Heterogeneity: Tau <sup>2</sup> = 0.00; Cl	hi² = 2.97, df =	3 (P = 0.	40); I <sup>z</sup> = 0	1%			
Test for overall effect: Z = 0.77	' (P = 0.44)						
1.26.2 Maximum or above							
Anderson 2009	22	138	7	72	6.6%	1.64 [0.74, 3.65]	
Anderson 2012	21	142	12	68	8.2%	0.84 [0.44, 1.60]	
Dackis 2005	10	30	4	32	4.7%	2.67 [0.94, 7.60]	
Dackis 2003	8	70	4	75	4.1%	2.14 [0.67, 6.80]	
Grabowski 2004	24	54	7	40	7.3%	2.54 [1.22, 5.30]	
Heinzerling 2010	9	34	10	37	6.9%	0.98 [0.45, 2.12]	
Konstenius 2010	8	12	9	12	9.9%	0.89 [0.53, 1.49]	
Levin 2006	3	21	2	13	2.3%	0.93 [0.18, 4.84]	
Levin 2005	20	83	3	43	4.1%	3.45 [1.09, 10.98]	
Levin 2019	14	64	4	63	4.7%	3.45 [1.20, 9.90]	
Mariani 2012	13	39	7	42	6.5%	2.00 [0.89, 4.49]	
Nuiiten 2016	11	38	2	35	2.9%	5.07 [1.21, 21.27]	
Schmitz 2012	2	22	1	8	1.3%	0.73 [0.08, 6.97]	
Schmitz 2012	1	20	1	8	1.0%	0.40 [0.03, 5.65]	
Schmitz 2014	9	22	10	18	8.2%	0.74 [0.38, 1.41]	
Shearer 2003	7	16	4	14	5.0%	1.53 [0.56, 4.15]	
Subtotal (95% CI)		805		580	83.9%	1.50 [1.10, 2.06]	◆
Total events	182		87				
Heterogeneity: Tau <sup>2</sup> = 0.17; Cl	hi² = 26.90, df =	= 15 (P =	0.03); I <sup>2</sup> :	= 44%			
Test for overall effect: Z = 2.52	? (P = 0.01)	,					
Total (95% CI)		1000		787	100.0%	1.45 [1.10, 1.90]	•
Total events	207		107				•
Heterogeneity: Tau <sup>2</sup> = 0.13; Cl		= 19 (P =		= 36%			
Test for overall effect: Z = 2.65			0.00/11	00.0			0.01 0.1 1 10 100
Test for subgroup differences		f = 1 (P =	: 0.58) IZ	= 0%			Favours [experimental] Favours [control]
restion subgroup unclences	. om – 0.01, u		0.007.1	- 0 /0			

Fig. 6. Overall and by dose effect of prescription psychostimulants compared to placebo on outcome sustained abstinence—overall PSUD

quality evidence that PPs increase sustained abstinence for patients with PSUDs in general, but that was attributable to results of studies in cocaine rather than amphetamine use disorder.

As with the earlier finding (Castells et al. 2016), we did not find evidence that treatment with PPs increases the retention in treatment. Retention was measured as a binary variable for treatment completion, and we also carried an additional analysis of retention as a continuous variable of duration in treatment; none of them revealed differences between the treatment and control groups. Subgroup analyses per drug and per medication also did not find significance for treatment retention for any subgroup. This further strengthens the positive effect found on abstinence, especially at the end of treatment, as differential retention might bias toward the treatment arm with the better retention.

There is very limited evidence that any medication improves retention to treatment (Chan et al. 2019; Indave et al. 2016) for PSUDs, unlike the effect of opioid agonists that does increase retention in treatment for opioid use disorder (Mattick et al. 2009). On the other hand, CM is a psychosocial intervention that has been shown to increase treatment retention rates in many different contexts (Garcia-Rodriguez et al. 2009; Miguel et al. 2016; Petry et al. 2005). Combining pharmacotherapy with CM for PSUDs may be an effective strategy to decrease attrition and promote further treatment benefits (Tardelli et al. 2018) as the usually high attrition rates undermine the effect of any therapeutic intervention for PSUDs, though it can be challenging to implement. Nonetheless, PPs had a significant effect on promoting cocaine abstinence, and trials with lower attrition rates found greater effect on sustained abstinence (Nuijten et al. 2016).

Maximum days of continuous abstinence during treatment is an outcome measure predicting long-term endpoints on cocaine use and global functioning (Carroll et al. 2014), but it is not often used in clinical trials. A pooled estimate of this outcome is included in our meta-analysis, but we were limited by the small number of trials and high heterogeneity, partially explained by the fact that we merged CUD and AUD trials. Also, pooled continuous variables usually carry more heterogeneity than binary outcomes in meta-analyses (Alba et al. 2016). But, even with these limitations, we found a significant effect of PPs on the continuous abstinence, and we recommend it for inclusion in the future trials of PSUD pharmacotherapy.

Our review brings an innovative dose range subgroup comparison. The necessity to use higher doses of PPs for PSUDs than those established for other conditions has been a matter of debate

	Psychostimu	lants	Place	bo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
1.27.1 Submaximum							
Dackis 2012	3	65	4	75	3.6%	0.87 [0.20, 3.72]	
Dürsteler-MacFarland 2013	3	30	3	32	3.4%	1.07 [0.23, 4.88]	
Kampman 2015	11	47	4	47	5.9%	2.75 [0.94, 8.02]	
Levin 2007	8	53	9	53	7.9%	0.89 [0.37, 2.13]	
Subtotal (95% CI)		195		207	20.9%	1.25 [0.71, 2.21]	<b></b>
Total events	25		20				
Heterogeneity: Tau <sup>2</sup> = 0.00; C	hi <sup>2</sup> = 2.97, df = 3	3 (P = 0.4	40); I <sup>2</sup> = 0	%			
Test for overall effect: Z = 0.77	7 (P = 0.44)						
1.27.2 Maximum or higher							
Anderson 2009	22	138	7	72	8.9%	1.64 [0.74, 3.65]	
Dackis 2005	10	30	4	32	6.1%	2.67 [0.94, 7.60]	
Dackis 2012	8	70	4	75	5.3%	2.14 [0.67, 6.80]	
Grabowski 2004	24	54	7	40	9.9%	2.54 [1.22, 5.30]	
Levin 2015	7	40	3	43	4.5%	2.51 [0.70, 9.04]	
Levin 2015	13	43	3	43	5.1%	4.33 [1.33, 14.13]	
Levin 2019	14	64	4	63	6.1%	3.45 [1.20, 9.90]	
Mariani 2012	13	39	7	42	8.8%	2.00 [0.89, 4.49]	
Nuijten 2016	11	38	2	35	3.7%	5.07 [1.21, 21.27]	
Schmitz 2012	1	20	1	8	1.2%	0.40 [0.03, 5.65]	
Schmitz 2012	2	22	1	8	1.7%	0.73 [0.08, 6.97]	
Schmitz 2014	9	22	10	18	11.3%	0.74 [0.38, 1.41]	<b>_</b>
Shearer 2003	7	16	4	14	6.6%	1.53 [0.56, 4.15]	
Subtotal (95% CI)		596		493	79.1%	1.95 [1.38, 2.77]	◆
Total events	141		57				
Heterogeneity: Tau <sup>2</sup> = 0.12; C	hi <sup>2</sup> = 17.23, df =	12 (P =	0.14); I <sup>2</sup> =	= 30%			
Test for overall effect: Z = 3.77		•					
Total (95% CI)		791		700	100.0%	1.77 [1.31, 2.40]	◆
Total events	166		77				
Heterogeneity: Tau <sup>2</sup> = 0.10; C		16 (P =		= 26%			ter al de la de
Test for overall effect: Z = 3.74							0.01 0.1 1 10 100
Test for subgroup differences		f = 1 (P =	: 0.19), I <sup>2</sup>	= 41.89	%		Favours Placebo Favours Psychostimulants

Fig. 7. Overall and by dose effect of prescription psychostimulants compared to placebo on outcome sustained abstinence-CUD only

(Levin et al. 2015a), as people with PSUDs may be cross-tolerant to the psychostimulant effect of these medications. Our metaanalysis has confirmed that higher but not lower doses of PPs promoted sustained abstinence. Results of trials that used low doses may have underestimated the true potential of PPs, especially for treatment of CUD. Dose of the "agonist" medication plays a major role in its effectiveness as it has been shown with opioid agonists (Faggiano et al. 2003) and also shown in preclinical and human laboratory studies which further underscores the benefit of agonist approach (Czoty et al. 2016). According to the GRADE guidelines, the presence of a dose-response gradient increases the quality of the evidence of a treatment (Higgins and Green 2011). Therefore, since this dose-range analysis is novel, it lowers the risk of bias of results that are already more statistically consistent than those shown in prior analyses which already regarded PPs for CUD as a "promising treatment for cocaine dependence" (Castells et al. 2016). Our findings elevate the quality of the evidence supporting the use of PPs, particularly prescription amphetamines for promoting abstinence in patients with CUD. The NNT of 7 for prescription amphetamines on the treatment of CUDs and 8 for PPs in the treatment of patients with co-occurring OUD is comparable to that found for acamprosate (NNT = 8) and naltrexone (NNT = 9), medications approved for the treatment of alcohol use disorder (Maisel et al. 2013).

It is noteworthy that we have found PPs to be ineffective for AUDs, with no heterogeneity between studies. However, very few trials with AUDs have used sustained abstinence as an outcome, so that only three trials, one with methylphenidate (unpublished data) (Konstenius et al. 2010) and two with modafinil (Anderson et al. 2012; Heinzerling et al. 2010), were pooled, while studies with positive findings were not included (Konstenius et al. 2014; Ling et al. 2014; Longo et al. 2010).

In the analysis per medication, neither modafinil nor methvlphenidate was effective on promoting sustained abstinence for PSUDs. Prescription amphetamines were responsible for the overall effect of PPs; however, only two trials up to this point used prescription amphetamines to treat AUDs (Galloway et al. 2011; Longo et al. 2010), and while findings were partially positive, none assessed sustained abstinence. On the other hand, a study conducted in patients with amphetamine use disorder and ADHD found that high dose of extendedrelease methylphenidate reduced use of amphetamine as compared to placebo, though this study was not included in the efficacy outcomes of this meta-analysis as it did not include the outcome of abstinence (Konstenius et al. 2014). It is possible that, as for CUD, trials with high doses and extended release formulation of prescription amphetamines could promote sustained abstinence from methamphetamine, and at least one of such studies is ongoing (Ezard et al. 2018).

PPs were particularly efficacious on promoting sustained abstinence for the subgroup of individuals with comorbid cocaine and opioid use disorders. Studies in this patient population

		PPs		PI	acebo			Mean Difference	Mean Difference
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% CI
1.5.1 Cocaine									
Grabowski 1997	30.7	40	25	43.8	41	24	1.0%	-13.10 [-35.79, 9.59]	
Grabowski 2004	21	16.8	41	16.1	16.9	19	5.4%	4.90 [-4.28, 14.08]	
Levin 2007	27	29	53	30	29	53	3.9%	-3.00 [-14.04, 8.04]	
Levin 2015	45.4	36.4	83	33.2	33.7	43	3.0%	12.20 [-0.56, 24.96]	———
Levin 2019	24.8	29.6	64	14.8	21.8	63	5.5%	10.00 [0.97, 19.03]	
Mariani 2012	43.9	31.3	39	33.7	35.8	42	2.4%	10.20 [-4.42, 24.82]	+
Morgan 2016	52	49.3	30	26	36.4	27	1.1%	26.00 [3.65, 48.35]	
NCT00142818 (Kampman)	9.7	8.9	82	8.9	9.1	82	21.8%	0.80 [-1.96, 3.56]	+
Nuijten 2016	10.6	25.1	38	3.9	17.9	35	4.7%	6.70 [-3.24, 16.64]	+
Schubiner 2002	50	50	24	42	32	24	0.9%	8.00 [-15.75, 31.75]	
Shearer 2003	38.6	34.3	16	27.1	30	14	1.0%	11.50 [-11.51, 34.51]	
Subtotal (95% CI)			495			426	<b>50.6</b> %	5.24 [0.97, 9.51]	◆
Heterogeneity: Tau <sup>2</sup> = 16.00;	Chi <sup>2</sup> = 18	6.13, di	f = 10 (i	P = 0.10	); <b>I</b> ² = 3	38%			
Test for overall effect: Z = 2.4	1 (P = 0.0	)2)							
1.5.2 Amphetamine									
Anderson 2012	36.9	38.7	142	33.1	37	68	4.0%	3.80 [-7.06, 14.66]	_ <del>_</del>
Galloway 2011	2.9	4.3	30	3.2	5	30	23.8%	-0.30 [-2.66, 2.06]	+
Heinzerling 2010	13.1	11.5	34	12.7	13.2	37	10.8%	0.40 [-5.35, 6.15]	+
Konstenius 2010	15.2	7.7	12	13.7	6.6	12	10.8%	1.50 [-4.24, 7.24]	+
Subtotal (95% CI)			218			147	49.4%	0.14 [-1.86, 2.15]	<b>•</b>
Heterogeneity: Tau <sup>2</sup> = 0.00; C	chi² = 0.7	9, df =	3 (P = 1	0.85); I <sup>z</sup>	= 0%				
Test for overall effect: $Z = 0.1$	4 (P = 0.8	9)							
Total (95% CI)			713			573	100.0%	2.40 [0.07, 4.73]	•
Heterogeneity: Tau <sup>2</sup> = 4.50; C	Chi² = 19.	59.df:	= 14 (P	= 0.14	$ ^{2} = 29$	3%			
Test for overall effect: Z = 2.0			110	0.14/	2.				-100 -50 0 50 100
Test for subgroup differences		·	lf = 1 (P	= 0.03)	$ ^{2} = 7$	7.7%			Favours [experimental] Favours [control]

Test for subgroup differences: Chi<sup>2</sup> = 4.49, df = 1 (P = 0.03), l<sup>2</sup> = 77.7%

Fig. 8. Overall mean difference in percentage of drug-negative urine tests throughout trial comparing prescription psychostimulants to placebo

included potent medications (dextroamphetamine) given at higher doses which is likely the main contributor to overall findings. Moreover, patients already taking methadone could have better adherence to a new medication and a synergic effect between opioid agonist and PPs could promote higher abstinence rates. Higher efficacy of PPs in a population with OUDs could also be explained by a different model of care. The opioid treatment program (OTP) model requires daily attendance to clinic, where study staff could motivate patients at a regular basis and supervise methadone intake, features that ensure patients are highly adherent to the experimental treatment. The OTP model has likely enhanced the adherence to medication in these trials, which could have partially explained the positive results among this specific subgroup (especially when considering that the overall quality of evidence for PPs for PSUD was severely undermined by elevated attrition rates).

Nuijten and colleagues conducted a trial using high doses of extended-release dexamphetamine for individuals dependent on crack-cocaine using a structure similar to those of the OTP and found significant results for sustained abstinence and maximum days of abstinence, with low attrition rates for both groups (Nuijten et al. 2016). Beyond the effect of medication itself, this well-designed trial successfully addressed common problems in trials with PSUDs, such as elevated dropout rates. This suggests that a structured model of care using PPs similar to methadone clinics could be an alternative for outpatient medication-based intervention for patients with PSUDs, yet other studies have shown that prescription amphetamines can be effective when given in a traditional outpatient treatment setting.

Patients with ADHD are more prone than the general population to develop substance use disorders, and ADHD is a common comorbidity among patients seeking treatment for

		PPs		PI	acebo			Mean Difference		Mean Diff	erence		
Study or Subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% Cl		IV, Randon	n, 95% CI		
Shearer 2003	38.6	34.3	16	27.1	30	14	4.0%	11.50 [-11.51, 34.51]			•		
Nuijten 2016	10.6	25.1	38	3.9	17.9	35	21.5%	6.70 [-3.24, 16.64]		+			
Mariani 2012	43.9	31.3	39	33.7	35.8	42	10.0%	10.20 [-4.42, 24.82]		+	-		
Levin 2019	24.8	29.6	64	14.8	21.8	63	26.1%	10.00 [0.97, 19.03]		F	-		
Levin 2015	45.4	36.4	83	33.2	33.7	43	13.1%	12.20 [-0.56, 24.96]		+			
Grabowski 2004	21	16.8	41	16.1	16.9	19	25.3%	4.90 [-4.28, 14.08]		-	-		
Total (95% CI)			281			216	100.0%	8.37 [3.75, 12.98]			•		
Heterogeneity: Tau <sup>2</sup> : Test for overall effect					0.94);	l² = 0%				50 0		50	100
									Favo	urs Placebo I	- avours Pl	~S	

Fig. 9. Overall mean difference in percentage of drug-negative urine tests throughout trial comparing prescription psychostimulants to placebo restricted to prescription amphetamines for CUDs

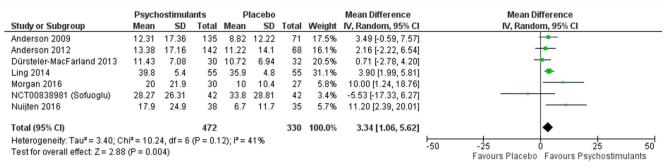


Fig. 10. Overall mean difference in maximum days of sustained abstinence comparing prescription psychostimulants to placebo

	Psychostim	ulants	Place	bo		Risk Ratio	Risk Ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% Cl	M-H, Random, 95% Cl
1.6.1 Cocaine							
Anderson 2009	83	136	42	71	6.2%	1.03 [0.82, 1.31]	+
Dackis 2005	19	30	21	32	2.9%	0.97 [0.67, 1.40]	
Dackis 2012	83	135	37	75	5.1%	1.25 [0.96, 1.63]	+
Dürsteler-MacFarland 2013	18	30	26	32	3.4%	0.74 [0.53, 1.03]	
Grabowski 1994	3	4	3	3	0.9%	0.80 [0.40, 1.58]	
Grabowski 1997	12	25	12	24	1.3%	0.96 [0.54, 1.70]	
Grabowski 2001	23	93	8	35	0.9%	1.08 [0.53, 2.19]	
Grabowski 2004	24	54	10	40	1.1%	1.78 [0.96, 3.29]	
Kampman 2015	34	47	37	47	6.4%	0.92 [0.73, 1.16]	-+
Levin 2007	23	53	24	53	2.2%	0.96 [0.63, 1.47]	
Levin 2015	64	83	29	43	6.0%	1.14 [0.90, 1.45]	+
Levin 2019	42	64	37	63	4.9%	1.12 [0.85, 1.47]	+-
Mariani 2012	29	39	35	42	6.5%	0.89 [0.71, 1.12]	
Mooney 2009	17	55	8	27	0.9%	1.04 [0.52, 2.11]	
Mooney 2015	12	22	15	21	1.9%	0.76 [0.48, 1.22]	+
NCT00142818 (Kampman)	52	82	39	82	4.6%	1.33 [1.01, 1.77]	
NCT00218036 (Schmitz)	9	17	3	9	0.4%	1.59 [0.57, 4.43]	
NCT00218387 (Malcolm)	50	83	17	40	2.5%	1.42 [0.95, 2.12]	
NCT00838981 (Sofuoglu)	30	45	28	46	3.9%	1.10 (0.80, 1.49)	- <b>-</b> -
Nuijten 2016	34	38	31	35	10.5%	1.01 [0.86, 1.19]	+
Schmitz 2012	7	22	3	16	0.3%	1.70 [0.52, 5.57]	
Schmitz 2012	4	20	3	16	0.2%	1.07 [0.28, 4.09]	
Schmitz 2014	9	22	11	18	1.1%	0.67 [0.36, 1.25]	
Schubiner 2002	11	24	14	24	1.4%	0.79 (0.45, 1.36)	
Shearer 2003	6	16	5	14	0.5%	1.05 [0.41, 2.70]	
Subtotal (95% CI)		1239		908	75.8%	1.03 [0.96, 1.11]	•
Total events	698		498				
Heterogeneity: Tau <sup>2</sup> = 0.00; Cl		= 24 (P =	0.37); l² =	= 7%			
Test for overall effect: Z = 0.86	i (P = 0.39)						
1.6.2 Amphetamine							
Anderson 2012	76	142	36	68	4.9%	1.01 [0.77, 1.33]	+
Galloway 2011	26	30	25	30	7.2%	1.04 [0.84, 1.29]	+
Heinzerling 2010	14	34	13	37	1.2%	1.17 [0.65, 2.12]	_ <del></del>
Konstenius 2010	7	12	10	12	1.4%	0.70 [0.41, 1.20]	
Konstenius 2014	10	27	4	27	0.4%	2.50 [0.89, 7.00]	
Ling 2014	29	55	31	55	3.3%	0.94 [0.66, 1.32]	
Longo 2010	15	23	8	26	1.0%	2.12 [1.11, 4.06]	
Miles 2013	17	39	10	39	1.0%	1.70 [0.89, 3.23]	<u> </u>
NCT00859573 (Mancino)	1	6	1	3	0.1%	0.50 [0.05, 5.51]	
Rezaei 2015	18	28	16	28	2.3%	1.13 [0.74, 1.72]	
Shearer 2009	11	38	15	42	1.0%	0.81 [0.43, 1.54]	
Tiihonen 2007	6	17	4	17	0.4%	1.50 [0.51, 4.38]	
Subtotal (95% CI)		451		384	24.2%	1.08 [0.93, 1.27]	•
Total events	230		173				
Heterogeneity: Tau <sup>2</sup> = 0.02; Cl		= 11 (P =	0.22); l² :	= 22%			
Test for overall effect: Z = 0.99	(P = 0.32)						
Total (95% CI)		1690		1292	100.0%	1.04 [0.97, 1.11]	•
Total events	928		671				
Test for suprell offset 7 = 1.20 (B = 0.22) 0.01 0.01 0.01 10 100							
Test for subgroup differences: Chi <sup>2</sup> = 0.29, df = 1 (P = 0.59), i <sup>2</sup> = 0%							

Fig. 11. Overall and by drug of abuse effect of prescription psychostimulants compared to placebo on outcome retention to treatment

**Table 2.** Box with a summary ofthe main findings

Overall findings:	Subgroup analyses:		
• PPs are efficacious on promoting sustained abstinence in patients with PSUD.	• PPs do not seem to promote sustained abstinence in patients with AUDs. However, trials with prescription amphetamines showed promise for outcomes that are not abstinence-based.		
• PPs seem to have a small effect on extending maximum days of sustained abstinence among			
patients with PSUD.	Prescription amphetamines are highly efficacious or		
• PPs do not seem efficacious on promoting retention to treatment between patients with PSUD.	promoting sustained abstinence in patients with CUD.		
• This did not differ by abuse drug and was not modified by the presence of Contingency Management.	<ul> <li>PPs are especially efficacious on promoting sustained abstinence in patients with concurrent opioid use disorder. This may be due to adherence features present in the OTP treatment.</li> <li>PPs are efficacious on promoting sustained abstinence in patients without ADHD. However, studies with ADHD populations have methodological limitations.</li> <li>Maximum current dosages of PPs or higher are mor efficacious than lower dosages on promoting sustained abstinence of stimulants, especially cocaine.</li> </ul>		
• Abstinence and Compliance-targeted Contingency Management may improve overall retention and should be incorporated by future trials.			
• Quality of evidence was impaired by high attrition rates in most of the trials.			
<ul> <li>The combination of PPs and Topiramate has shown promising efficacy on promoting sustained cocaine abstinence and should be further tested.</li> </ul>			

substance use disorders (Lee et al. 2011; van Emmerik-van Oortmerssen et al. 2012). PPs are the gold standard treatment for ADHD (Faraone and Buitelaar 2010) and were found to improve ADHD symptoms in trials with comorbid ADHD and SUDs (Konstenius et al. 2010; Levin et al. 2007). However, the impact on PSUD is mixed. Factors that might have led to early negative results include use of medications with poorer bioavailability (Levin et al. 2006, 2007) or inadequate dosing (Levin et al. 2006). Levin and colleagues hypothesized that patients with ADHD and heavy cocaine use could require even higher dosages of dopaminergic medications than those with high cocaine use but no ADHD and conducted a trial comparing a regimen of 80 mg/day of dextroamphetamine to the usual dose of 60 mg/day and placebo (Levin et al. 2015b). Both medication groups achieved significantly higher rates of sustained abstinence compared to placebo; the 80 mg groups had larger odds of sustained abstinence than the 60 mg group, although that difference was not significant. Consistent with this approach, Konstenius et al. (2014) found that high doses of extended-release methylphenidate (up to 180 mg/day) resulted in clinically significant improvement in ADHD and reduction in amphetamine use among those with amphetamine use disorder (Konstenius et al. 2014) whereas the lower dose of 72 mg/ day did not result in a reduction in ADHD symptoms or amphetamine use (Konstenius et al. 2010). Because the 2014 study with higher doses of methylphenidate did not provide data on the sustained abstinence outcome, their efficacy measures were not included in this meta-analysis. Low dosages in the other two trials pooled might explain PPs' low efficacy in the ADHD subgroup. Importantly, patients in the PPs group showed improvements in the ADHD outcomes in all of three trials mentioned above which refutes the argument that drug-related clinical improvements in patients with ADHD are mostly mediated

by treatment of the ADHD symptoms. Another caveat of this subgroup analysis is that most studies reported did not assess ADHD status and did not exclude these individuals. This way, it is possible that many patients seeking treatment for PSUDs also have ADHD but are not diagnosed. Given that there are limited though encouraging data, further studies with higher PPs dosages are necessary to evaluate the effect of PPs in patients with co-occurring PSUD and ADHD (Woon et al. 2018).

Psychosocial treatments, particularly CM targeting abstinence and treatment adherence, have shown benefits in patients with PSUD (Kampman 2019; Lussier et al. 2006). Using CM in combination with a pharmacologic treatment has been shown to have synergic effect in decreasing drug use and enhancing treatment (Penberthy et al. 2010; Tardelli et al. 2018). However, in the present meta-analysis, we did not find that CM (targeting either abstinence or adherence) improved treatment retention. Incorporating CM elements to decrease treatment attrition may be considered in future trials as high rates of treatment attrition remain a prominent issue in trials conducted in patients with PSUD although it is possible that the effect of well-designed and implemented CM may create the ceiling effect, with increase in placebo response, which lowers the possibility of detecting the medication effect.

The current meta-analysis has several limitations. Many of the trials included were likely underpowered, which might have hindered positive findings about the efficacy of PPs when looking at the trials individually but the meta-analysis that also includes adequately powered trials usually corrects this limitation (Nuesch et al. 2010). Since almost all of the trials evaluated excluded individuals with severe psychiatric comorbidities from their sample or did not assess for other common conditions, such as ADHD, it was not possible to conduct sub-group analyses on depression or other highly prevalent conditions in patients with PSUD. Due to the heterogeneity of outcome measurement among the included studies, it was not possible to pool relevant outcomes, such as other continuous drug use variables, e.g., reduction in drug use, reduction in drug use category (Roos et al. 2019a), or other non-abstinence endpoints. Also, due to the scarcity of data, it was not possible to compare trials using higher dosages of PPs to those using the current maximum dosages approved by the FDA for other conditions. The duration of sustained abstinence as an outcome differed across trials. Moreover, elevated attrition rates might have underestimated treatment effects and decisively influenced on the GRADE judgment of the quality of evidence. Few trials with features that are known to improve adherence in trials of pharmacologic treatment of substance use disorders, such as CM targeting drug use or attendance (Tardelli et al. 2018) and daily supervised intake and motivational enhancement (Weiss 2004), were available. Also, quality assessment was downgraded for all outcomes due to possible detection bias, since the intervention has potential behavioral effects that may hinder blinding of clinicians, patients, and outcome assessors. Therefore, it is vital that the interpretations of our findings for clinical practice consider that the nature of the intervention makes high-quality evidence methodologically impossible at this point. This warrants future trials with comparable methods to those that showed best evidence (prescription amphetamines in higher dosages for CUD, for example), since downgrades in the quality of evidence are not due to lack of efficacy but to properties of the intervention and to methodological issues of the trials included which may be difficult to overcome.

The present meta-analysis has several strengths. We included the GRADE approach for each of the subgroup analyses which might help in dealing with group specificities when elaborating new public policies and in designing future clinical trials on this field. The new subgroup analysis by dose range suggests a dose–response relationship and strengthens the quality of the evidence supporting PPs for PSUDs. Lastly, calculating NNTs provides a clinically translatable measure of effect that might facilitate the dissemination findings in clinician guiding activities, though the NNT obtained from metaanalyses should be interpreted with caution due to differences in treatment effects between studies (Marx and Bucher 2003).

The results of the present study have implications for further trials that aim to test the efficacy of PPs for PSUDs. An optimal trial should test at least one high dose of PPs with extended-release formulations to ensure the maximum potential benefit and to test efficacy of the agonist-based therapy. Moreover, features that minimize attrition and maximize adherence with the medication, a common problem of trials with PSUDs, are warranted and may include CM, daily attendance to clinic, and supervised medication intake. Use of extendedrelease preparations and once-daily dosing, preferably under direct observation, should be considered to maximize safety and minimize diversion potential. Unstable cardiovascular disorder and a history of a psychotic disorder should be considered as exclusionary. Abstinence-based outcomes (sustained abstinence and maximum days of abstinence) should be combined with continuous drug use outcomes.

Trials with higher dosages of PPs, especially prescription amphetamines, are particularly needed for the treatment of AUD. The combination of PPs with topiramate has shown promise on the treatment of PSUD, and therefore, combination treatments should be considered for future trials. We believe that feasibility studies of the "agonist-type" pharmacological intervention, namely outpatient-based supervised treatment with high-dose, extended-release preparation of amphetamines, should be considered particularly in countries with high rates of PSUDs that currently do not provide any medical treatment for those individuals. In real-world scenarios, one possible strategy is to offer treatment with prescription amphetamines in the setting of an OTP. Many patients with PSUD are already enrolled in OTPs (those with co-occurring OUD and PSUD), and others can be referred to OTPs for the medical management of PSUD as the traditional PSUD outpatient programs do not offer consistent medical and multi-professional oversight. Alternatively, mobile technology solutions for monitoring and increasing adherence to the medication may be considered.

# Conclusion

Recent trials with extended-release formulations and higher dosages of PPs, particularly prescription amphetamines, have shown promising results promoting abstinence from cocaine and reducing drug use. PPs' potential as an "agonist-type" treatment seems to be better explored with higher dosage regimens and at clinical settings that have direct observed dosing available. The results from patients with comorbid opioid use disorders are particularly encouraging, and this may be due to the fact that high dosages of potent PPs were used, and this population is already enrolled to a healthcare facility that offers daily attendance, supervised medication intake, evidencebased psychosocial interventions, and a wide-range of ancillary services. A widely used and successful model of treating opioid use disorder or incorporating mobile technology solutions to monitor and enhance medication adherence may now be assessed for treatment of individuals with psychostimulant use disorder and incorporate prescription amphetamines as an agonist intervention. Considering the major public health impact of untreated PSUD, and the absence of the widely accepted pharmacological intervention, there is an urgent need to conduct implementation studies of this treatment approach.

# Compliance with ethical standards

Conflict of interest The authors have no conflict of interest to disclose.

**Ethical approval** The present article does not necessarily express the view of the United Nations.

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