

Memantine add-on to antipsychotic treatment for residual negative and cognitive symptoms of schizophrenia: a meta-analysis

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Abstract

Rationale We examined whether memantine add-on to antipsychotic treatment is beneficial in schizophrenia treatment. **Objective** This systematic review and meta-analysis aimed to achieve stronger evidence on the efficacy and safety of memantine add-on for treating schizophrenia. **Methods** We analyzed double-blind, randomized, placebo-controlled trials of memantine add-on treatment in schizophrenia patients receiving antipsychotics. The primary outcomes were amelioration of negative symptoms and all-cause discontinuation. Dichotomous outcomes are presented as risk ratios (RRs), and continuous outcomes are presented as mean differences (MDs) or standardized mean differences (SMDs). **Results** Eight studies ($n = 448$) were included. Although memantine add-on treatment was superior to placebo for ameliorating negative symptoms (SMD = -0.96 , $p = 0.006$, $I^2 = 88\%$; $N = 7$, $n = 367$) in the Positive and Negative Syndrome Scale general subscale (MD = -1.62 , $p = 0.002$, $I^2 = 0\%$; $N = 4$, $n = 151$) and Mini-Mental Status Examination score (MD = -3.07 , $p < 0.0001$, $I^2 = 21\%$; $N = 3$, $n = 83$), there were no statistically significant differences in the amelioration of overall (SMD = -0.75 , $p = 0.06$, $I^2 = 86\%$; $N = 5$, $n = 271$),

positive (SMD = -0.46 , $p = 0.07$, $I^2 = 80\%$; $N = 7$, $n = 367$), and depressive symptoms (SMD = -0.127 , $p = 0.326$, $I^2 = 0\%$; $N = 4$, $n = 201$); all-cause discontinuation (RR = 1.34, $p = 0.31$, $I^2 = 0\%$; $N = 8$, $n = 448$); and individual adverse events (fatigue, dizziness, headache, nausea, constipation) between the groups. For negative symptoms, the significant heterogeneity disappeared when risperidone studies alone were considered ($I^2 = 0\%$). However, memantine add-on treatment remained superior to placebo (SMD = -1.29 , $p = 0.00001$). Meta-regression analysis showed that patient age was associated with memantine-associated amelioration of negative symptoms (slope = 0.171, $p = 0.0206$). **Conclusions** Memantine add-on treatment may be beneficial for treating psychopathological symptoms (especially negative symptoms) in schizophrenia patients. The negative-symptom effect size may be associated with younger adult schizophrenia patients.

Keywords Memantine · Schizophrenia · Negative symptoms · Cognitive function · Systematic review · Meta-analysis

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Introduction

Schizophrenia symptoms are characterized by positive, negative, cognitive, and affective symptoms (van Os and Kapur 2009). Although antipsychotics (most of which are dopamine antagonists) have beneficial effects on positive symptoms in typical schizophrenia patients, antipsychotics are less effective for negative symptoms or cognitive dysfunction (Miyamoto et al. 2012). Thus, research is needed to examine whether therapeutic targets other than dopamine receptors exist, for treating negative symptoms as well as cognitive decline.

Multiple lines of evidence from animal, genetic, and post-mortem studies suggest that glutamate receptor [e.g., the *N*-methyl-D-aspartate (NMDA) receptor] hypofunction in the brain (particularly in the thalamus) may be associated with negative symptoms in schizophrenia patients (Beck et al. 2016; Gray and Roth 2007; Hoflich et al. 2015; Kondziella et al. 2006; Miyamoto et al. 2012; Vukadinovic 2014). NMDA-mediated neuronal cell death may play a role in the pathology of schizophrenia (Lakhan et al. 2013). Some studies reported that NMDA receptor modulators, such as glycine, D-serine, and sarcosine, a nonselective glycine reuptake inhibitor used as an adjunctive therapy in antipsychotic-treated schizophrenia, showed amelioration of negative symptoms (Hashimoto 2014). Memantine is postulated to exert its therapeutic effect through its action as a low-to-moderate affinity noncompetitive (open channel), nonselective, voltage-dependent, NMDA receptor antagonist, which binds preferentially to NMDA receptor-operated calcium channels (Berman et al. 2012; Kishi and Iwata 2013). Memantine has been approved worldwide for treating moderate to severe Alzheimer's disease. Memantine blocks the effects of sustained, pathologically elevated levels of glutamate that may otherwise lead to neuronal dysfunction (Danysz and Parsons 2003; Di Iorio et al. 2017; Sani et al. 2012). Memantine may also upregulate NMDA receptor expression, causing activation in the presence of a strong stimulus (Joshi et al. 2007). From the above studies, memantine add-on to antipsychotic treatment may have a benefit for treating cognitive impairment and negative symptoms in schizophrenia patients (Di Iorio et al. 2017; Sani et al. 2012). Our previous meta-analysis showed that memantine add-on to antipsychotic treatment in schizophrenia showed a trend toward superior efficacy for ameliorating overall symptoms [standardized mean difference (SMD) = -0.99, 95% confidence intervals (95% CIs) = -2.04, 0.06, $p = 0.06$] and negative symptoms (SMD = -1.08, 95% CI = -2.21, 0.04, $p = 0.06$) over placebo (Matsuda et al. 2013). However, because the numbers of patients and studies included were small (four studies including 222 patients) (Matsuda et al. 2013), we considered our inability to accurately estimate the efficacy and safety of memantine add-on treatment because of a low statistical power (i.e., insufficient sample size) as a limitation of our previous meta-analysis (Folstein et al. 1975). Four double-blind, randomized, placebo-controlled trials of memantine add-on treatment in schizophrenia were published recently (Fakhri et al. 2016; Mazinani et al. 2017; Omranifard et al. 2015; Veerman et al. 2016). Because a meta-analysis can increase the statistical power for group comparisons and can overcome the limitation of sample size in underpowered studies (Higgins and Green 2011), we hypothesized that the meta-analysis updated with the four new studies could establish superiority of memantine add-on to antipsychotic treatment over placebo for ameliorating negative symptoms in schizophrenia. Therefore, we conducted an

updated systematic review and meta-analysis to achieve more robust evidence regarding the efficacy for psychopathology (particularly negative symptoms) of memantine add-on to antipsychotic treatment in schizophrenia patients (eight studies, total of 448 patients).

Methods

This meta-analysis was performed according to the Preferred Reporting Items for Systematic Reviews and Meta-Analyses guidelines (Moher et al. 2009) (PRISMA 2009 checklist). The review has been registered with PROSPERO (<http://www.crd.york.ac.uk/PROSPERO/>. CRD42017058749).

Search strategy and inclusion criteria

To identify relevant studies, two of the authors (T.K. and Y.M.) independently searched MEDLINE, Cochrane library, and PsycINFO without language restrictions from the inception of their databases to March 14, 2017, using the following search strategy: memantine and schizophrenia. The authors also searched the Cochrane Central Register of Controlled Trials and clinical trial registries (<http://clinicaltrials.gov/>) to include randomized controlled trials as comprehensively as possible and to minimize the possibility of publication bias. Only double-blind, randomized, placebo-controlled trials of memantine add-on to antipsychotic treatment in schizophrenia patients lasting ≥ 2 weeks were included. Studies with a cross-over design were allowed in this systematic review and meta-analysis. The same two authors independently assessed inclusion/exclusion criteria and selected the studies. The references of the included articles and review articles were also searched for citations of additional relevant published and unpublished studies, including conference abstracts.

Data synthesis and outcome measures

Our primary outcome measure for efficacy was the amelioration of negative symptoms, according to the Positive and Negative Syndrome Scale (PANSS) negative subscale (Kay et al. 1989) score and the Brief Psychiatric Rating Scale (BPRS) negative subscale (Overall and Gorham 1962) score. The secondary outcome measures for efficacy included improvements in overall scores (PANSS, BPRS), positive symptoms (PANSS, BPRS), PANSS general subscale score, depressive symptoms [Hamilton Rating Scale for Depression (Hamilton 1960), the Calgary Depression Scale for schizophrenia (Addington et al. 1990)], the Clinical Global Impression Severity Score (Guy and Bonato 1970), and the Mini-Mental Status Examination (MMSE) (Folstein et al. 1983) score as well as discontinuation due to inefficacy. Our primary outcome measure for safety was all-cause

discontinuation. Secondary outcome measures for safety included discontinuation due to adverse events and the incidence of individual adverse events.

Data extraction

Two authors (T.K. and S.M.) independently extracted data from the included studies. Where possible, we used intention-to-treat (ITT) or modified ITT analysis. When such data were unavailable, the results for observed case (OC) analysis were extracted from each study. When the data required for meta-analysis were missing, we contacted the investigators of the relevant study and requested unpublished data.

Meta-analysis methods

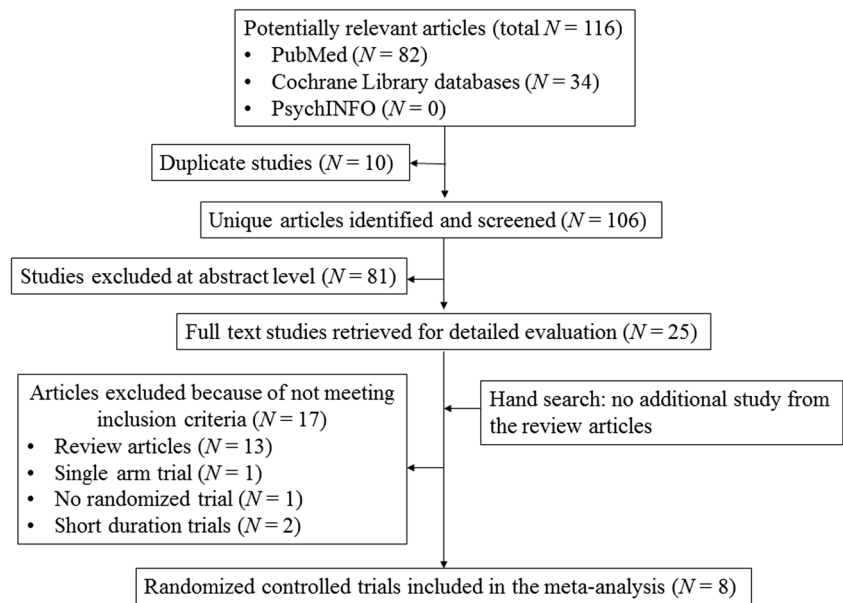
The meta-analysis was conducted using Review Manager software (version 5.3 for Windows, Cochrane Collaboration, <http://tech.cochrane.org/revman>). The random effects model was selected for this meta-analysis because of the potential heterogeneity across studies. Dichotomous outcomes were presented as risk ratios (RRs) with 95% CIs. Continuous outcomes were analyzed using the mean difference (MD) or, when different studies used different scales, the SMD. Because lower numbers are worse for the MMSE scores in the meta-analytic program Review Manager (see below), we reversed the algebraic sign of the outcomes when higher numbers were negative (i.e., for MMSE scores). We assessed the methodological quality of the trials, according to the Cochrane risk-of-bias criteria in the Cochrane Handbook (version 5.1.0. Cochrane Collaboration, http://handbook.cochrane.org/front_page.htm). We also investigated study heterogeneity using the I^2 statistic, considering $I^2 \geq 50\%$ to reflect considerable heterogeneity (Higgins et al. 2003). When considerable heterogeneity was observed in overall, positive, and negative symptoms, we performed sensitivity analyses for the following: scales (BPRS vs. PANSS), geographical region (Asia vs. other regions), second-generation antipsychotic (SGA) studies vs. first-generation antipsychotic (FGA) studies, clozapine studies vs. other antipsychotic studies, risperidone studies vs. other antipsychotic studies, analyzed population (ITT population vs. OC population), and sponsorship (industry vs. non-industry). In addition, we performed a meta-regression analysis to evaluate the association between the result of meta-analysis on the amelioration of the symptoms (overall, positive, and negative symptoms, PANSS general scores and MMSE scores) and certain modulators (patient age, sample size, study duration, publication year, and percent male using Comprehensive Meta-Analysis software version 2 (Biostat Inc., Englewood, NJ, USA)). Finally, we utilized funnel plots to explore potential publication bias. Egger's regression test was used to detect publication bias in meta-analyses using the same software.

Results

Study characteristics

Of the 116 results obtained in our literature search, we excluded the following: 10 because they were duplicates, 81 after a review of the abstract or title review, and 17 articles after a review of the full text [13 review articles (de Bartolomeis et al. 2012; Di Iorio et al. 2017; Francis 2009; Kavirajan 2009; Kishi and Iwata 2013; Koch et al. 2005; Koola et al. 2014; Matsuda et al. 2013; Sani et al. 2012; Shim and Nadeem 2014; Stys and Lipton 2007; Veerman et al. 2014; Zdanys and Tampi 2008), one single-arm study (Veerman et al. 2017), one nonrandomized trial (Cerullo et al. 2007), and two short-duration studies (Bhakta et al. 2016; Swerdlow et al. 2016)]. We did not retrieve any studies by searching through the review articles (Fig. 1). In total, eight trials including 448 patients (study details in Table 1) (de Lucena et al. 2009; Fakhri et al. 2016; Lee et al. 2012; Lieberman et al. 2009; Mazinani et al. 2017; Omranifard et al. 2015; Rezaei et al. 2013; Veerman et al. 2016) were included in the meta-analysis. The mean duration of the studies was 10.25 weeks (range = 6–12 weeks), and the mean patient age was 38.6 years. All studies were double-blind, randomized, placebo-controlled trials, and all were published in English. There were no studies with a cross-over design. All included patients in this study received memantine add-on vs. placebo to ongoing antipsychotic treatment. Two of the eight studies were clozapine studies (de Lucena et al. 2009; Veerman et al. 2016), two were risperidone studies (Mazinani et al. 2017; Rezaei et al. 2013), two were studies on various SGAs (Lieberman et al. 2009; Omranifard et al. 2015), one was an olanzapine study (Fakhri et al. 2016), and one study involved various FGAs (Lee et al. 2012). The dose of memantine was 20 mg/day in all studies. Four of the eight studies were conducted in Iran (Fakhri et al. 2016; Mazinani et al. 2017; Omranifard et al. 2015; Rezaei et al. 2013). Although two studies used OC populations in their analyses (Mazinani et al. 2017; Omranifard et al. 2015), we included these data in our meta-analysis to increase the sample size as much as possible. There were two studies that did not report the primary outcome of our study (Fakhri et al. 2016; Omranifard et al. 2015). Data from Omranifard et al. 2015 was used only for the following outcomes: discontinuation rate and individual adverse events (Omranifard et al. 2015). One of the eight studies was sponsored by a pharmaceutical company (Lieberman et al. 2009). Evaluations regarding the methodological quality of the included studies were performed according to the Cochrane risk-of-bias criteria and are shown in Supplementary Fig. 1.

Fig. 1 Flowchart of literature review process



Results of the meta-analysis

Efficacy outcomes

Memantine add-on treatment was superior to placebo for ameliorating negative symptoms (SMD = -0.96 , 95% CIs = -1.64 to -0.27 , $p = 0.006$, $I^2 = 88\%$; $N = 7$, $n = 367$; Fig. 2a), PANSS general subscale score (MD = -1.62 , 95% CIs = -2.65 to -0.59 , $p = 0.002$, $I^2 = 0\%$; $N = 4$, $n = 151$; Fig. 2b), and MMSE scores (MD = -3.07 , 95% CIs = -4.46 to -1.69 , $p < 0.0001$, $I^2 = 21\%$; $N = 3$, $n = 83$; Fig. 2c; Table 2). Although memantine add-on treatment showed a trend toward superiority over placebo for ameliorating overall (SMD = -0.75 , 95% CIs = -1.52 to 0.03 , $p = 0.06$, $I^2 = 86\%$; $N = 5$, $n = 271$) and positive symptoms (SMD = -0.46 , 95% CIs = -0.96 to 0.05 , $p = 0.07$, $I^2 = 80\%$; $N = 7$, $n = 367$), there were no statistically significant differences in depressive symptoms or CGI-S scores between the treatment groups (Table 2). The data in each treatment group were simulated with no publication bias (Egger's test p values; overall symptoms = 0.163 , positive symptoms = 0.298 , negative symptoms = 0.0713).

Sensitivity analysis on efficacy outcomes

Results of sensitivity analyses are shown in Table 3. Significant heterogeneities remained in all sensitivity analyses of overall symptoms. For positive symptoms, the significant heterogeneity disappeared when performing a sensitivity analysis using data from the risperidone studies alone ($I^2 = 0\%$). Memantine add-on treatment was no more efficacious than placebo in these studies. For negative symptoms, the significant heterogeneity disappeared only when performing a

sensitivity analysis using data from risperidone studies alone ($I^2 = 0\%$). However, the superiority of memantine add-on treatment over placebo remained (SMD = -1.29 , 95% CIs = -1.79 to -0.79 , $p = 0.00001$; $N = 2$, $n = 76$) for negative symptoms.

Meta-regression analysis on efficacy outcomes

A meta-regression analysis showed that the effect size of memantine with respect to negative (slope = 0.171 , 95% CIs = 0.0262 – 0.315 , $p = 0.0206$; $N = 7$, $n = 367$; Fig. 3a) and overall symptoms (slope = 0.171 , 95% CIs = 0.09835 – 0.324 , $p = 0.0206$; $N = 5$, $n = 271$; Fig. 3b) increased in schizophrenia patients who were younger adults (Table 4).

Safety outcomes

There were no statistically significant differences in all-cause discontinuation (RR = 1.34 , 95% CIs = 0.76 – 2.37 , $p = 0.31$, $I^2 = 0\%$; $N = 8$, $n = 448$; Fig. 2d), discontinuation due to adverse events, or the incidence of individual adverse events (fatigue, dizziness, headache, nausea, constipation) between the groups (Table 2). The data for all-cause discontinuation in each treatment group were simulated with no publication bias (Egger's test p value = 0.707).

Discussion

This is an updated systematic review and meta-analysis of memantine add-on treatment to antipsychotic-treated schizophrenia patients. Memantine add-on to antipsychotic treatment was well tolerated, with no significant differences in

Table 1 Study, patient and treatment characteristics of included double-blind, placebo-controlled trial

(1) Study name (2) total number, (3) country, (4) sponsorship	Patients (1) inclusion criteria. (2) Analyzed population	Duration	Drug (mg/day)	Number	Mean age \pm SD (years)	% Male	%Race	Antipsychotic (mg/day)	Outcomes ^b
(1) de Lucena 2009, (2) 22, (3) Brazil, (4) non--industry	(1) SZ (100%, DSM-IV). Outpatients (100%), refractory SZ: partial remission of negative symptoms (mean BPSD total at baseline = 40.28). CLO treatment >10 years. (2) ITT/mITT	12 weeks	MEM 20	11	34.60 \pm 9.99	80%	NR	CLO (100%): 540.00 \pm 211.87	MEM > PLA: BPRS total, BPRS positive, BPRS negative, CGI-S, and MMSE
			PLA	11	34.73 \pm 8.57	100%	CLO (100%): 659.09 \pm 185.55		
(1) Fakhri 2016, (2) 60, (3) Iran, (4) non--industry	(1) SZ (100%, DSM-IV-TR). Inpatients (100%), PANSS total >50. (2) ITT/mITT	6 weeks	MEM 20	30	36.46 \pm 2.5	50%	Iranian 100%	OLA (100%): 15–20	MEM > PLA: PANSS positive, PANSS negative
			PLA	30	37.6 \pm 2.8	50%	OLA (100%): 15–20		
(1) Lee 2012, (2) 26, (3) Korea, (4) non--industry	(1) SZ (100%, DSM-IV using SCID-I). Inpatients (100%), stabilized FGA before the trial \geq 3 months (mean PANSS total at baseline = 74.58), MMSE = 18–24. (2) ITT/mITT	12 weeks	MEM 20	15	44.3 \pm 4.3	73.3%	Korean 100%	FGAs (100%) (CHL eq. dose): 1261.7 \pm 1078.-6	MEM = PLA: PANSS total, PANSS positive, PANSS negative, PANSS general, HAMD, CGI-S, and MMSE
			PLA	11	43.4 \pm 3.9	45.5%	FGAs (100%) (CHL eq. dose): 986.4 \pm 831.6		
(1) Lieberman 2009, (2) 138, (3) USA, (4) industry	SZ (99.3%, DSM-IV using SCID) and SA (0.7%). Outpatients (100%), BPRS total score >26 and at least one of the BPRS psychosis factor items >4. Residual positive symptoms \geq 3 months with no exacerbation in the last 4 weeks. SGA monotherapy (OLA, RIS, ARI, ZIP, or QUE) >3 months (with a stable dose >4 weeks) (mean PANSS total at baseline = 74.00). (2) ITT/mITT	8 weeks	MEM 20	70	40.9 \pm 9.8	59.4%	Caucasian 63.8%	OLA (34.8%), RIS (33.3%), ARI (15.9%), ZIP (8.7%), QUE (7.2%)	MEM = PLA: PANSS total, PANSS positive, PANSS negative, PANSS general, CDSS, and CGI-S
			PLA	68	40.1 \pm 11.3	79.1%	Caucasian 56.7%	OLA (37.3%), RIS (31.3%), ARI (10.4%), ZIP (11.9%), QUE (9%)	
(1) Mazinani 2017, (2) 46, (3) Iran, (4) non--industry	(1) SZ (100%, DSM-IV). Inpatients (100%). (2) OC	12 weeks	MEM 20	11	44.8 \pm 6.6	100%	Iranian 100%	RIS (100%): 4–6	MEM > PLA: PANSS negative, MMSE MEM = PLA: PANSS positive, PANSS general.
			PLA	11	45.3 \pm 6.2	100%	RIS (100%): 4–6		
(1) Omranifard 2015, (2) 64, (3) Iran, (4)	(1) SZ (100%, DSM-IV-TR). Inpatients (100%), BPRS total score >26 and SGA therapy in the past 3 months. (2) OC	12 weeks	MEM 20	32	32.3 \pm 9.9	60%	Iranian 100%	OLA (33.3%), RIS (50%), ARI (3.3%), CLO (13.3%)	MEM > PLA: GAF and QLS

Table 1 (continued)

(1) Study name (2) total number, (3) country, (4) sponsorship	Patients (1) inclusion criteria. (2) Analyzed population	Duration	Drug (mg/day)	Number	Mean age \pm SD (years)	% Male	%Race	Antipsychotic (mg/day)	Outcomes ^b
	non--industry		PLA	32	34.2 \pm 10.6	46.7%		OLA (23.3%), RIS (53.3%), ARI (6.7%), CLO (16.6%)	
(1) Rezaei 2013, (2) 40, (3) Iran, (4) non--industry	(1) SZ (100%, DSM-IV-TR). Outpatients (100%), RIS treatment for \geq 8 weeks and clinical stable (<20% PANSS total change) \geq 4 weeks (mean PANSS total at baseline = 45.25). (2) ITT/mITT	8 weeks	MEM 20 PLA 20	20 20	33.5 \pm 6.9 33.0 \pm 6.9	60% 55%	Iranian 100%	RIS (100%): 6 (fixed) RIS (100%): 6 (fixed)	MEM > PLA: PANSS total, <i>PANSS negative</i> , PANSS general MEM = PLA: PANSS positive, HAMD
(1) Veerman 2016 ^a , (2) 52, (3) Netherlands, (4) non--industry	SZ (100%, DSM-IV using MINI-Plus). Outpatients (100%). Failed to achieve remission criteria proposed by Andreasen et al. ^a (mean PANSS total at baseline = 81.21). CLO treatment \geq 6 months with a CLO plasma level >350 ng/ml >12 weeks or intolerability to achieve this threshold. (2) ITT/mITT	26 weeks Phase 1:12 weeks	MEM 20 PLA 26	26 26	42.35 \pm 9.55	75%	NR	CLO (100%): 350.00 \pm 182.84	MEM > PLA: PANSS negative, MEM = PLA: PANSS total, PANSS positive, PANSS general, CGI-S

ARI aripiprazole, *BPRS* Brief Psychiatric Rating Scale, *CDSS* Calgary Depression Scale for Schizophrenia, *CGI-S* Clinical Global Impression Severity, *CHL eq.* chlorpromazine equivalent, *CLO* clozapine, *DSM-IV (TR)* Diagnostic and Statistical Manual of Mental Disorders, Fourth Edition (Text Revision), *FGA* first generation antipsychotic, *GAF* Global Assessment of Functioning, *HAMD* Hamilton Rating Scale for Depression, (*m*)*ITT* (modified) intention-to-treat, *MEM* memantine, *MINI-Plus* Mini International Neuropsychiatric Interview Plus, *MMSE* Mini-Mental Status Examination, *Number* number of patients, *NR* not report, *OC* observed case, *OLA* olanzapine, *PANSS* Positive and Negative Syndrome Scale, *PLA* placebo, *QLS* quality of life scale, *QUE* quetiapine, *RIS* risperidone, *SD* standard deviation, *SGA* second generation antipsychotic, *SCID* Structured Clinical Interview, *SZ* schizophrenia, *ZIP* ziprasidone,

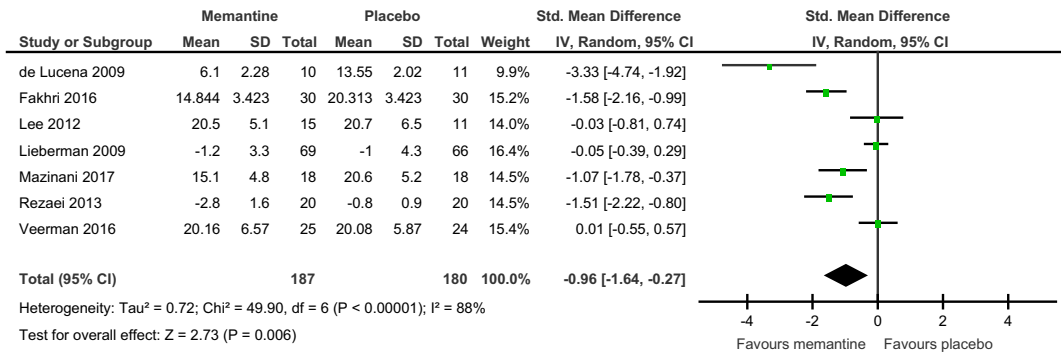
^a Crossover study

^b Italic font indicates the primary outcomes

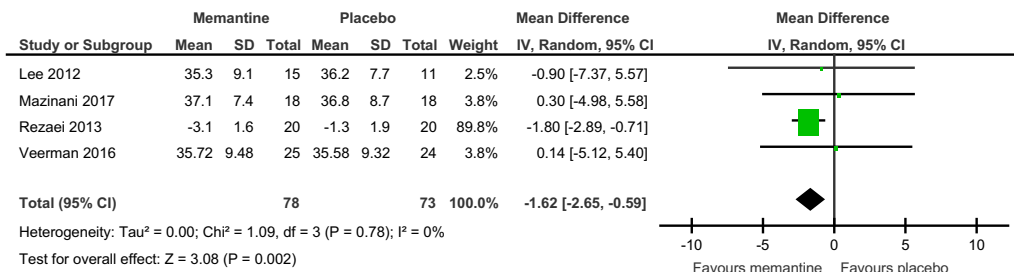
any safety outcomes between the memantine add-on and placebo groups. Moreover, memantine add-on was more efficacious than placebo for ameliorating negative symptoms (effect size was reasonably large), PANSS general subscale score (effect size was small), and MMSE scores (effect size was reasonably large). Memantine add-on treatment also tended to be superior to placebo for ameliorating overall and positive symptoms. However, we detected significant heterogeneities in overall, positive, and negative symptoms. Sensitivity analyses (e.g., methodological quality: analyzed population and sponsorship) to detect confounding factors

for the meta-analyses of these outcomes did not reveal any apparent explanation for the significant heterogeneity for ameliorating overall symptoms. However, the significant heterogeneity for the amelioration of positive and negative symptoms disappeared when sensitivity analyses were performed on risperidone studies alone. Because some medications (such as antipsychotics) may interfere with the actions of memantine on NMDA receptors (Di Iorio et al. 2017), pooling various antipsychotic studies may increase the significant heterogeneities in this outcome. Both risperidone studies were also similar to the current study characteristics in that

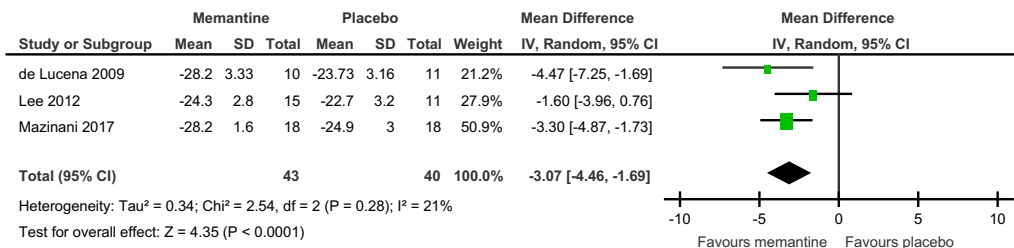
a. Negative symptoms



b. PANSS general subscale score



c. MMSE score



d. All-cause discontinuation

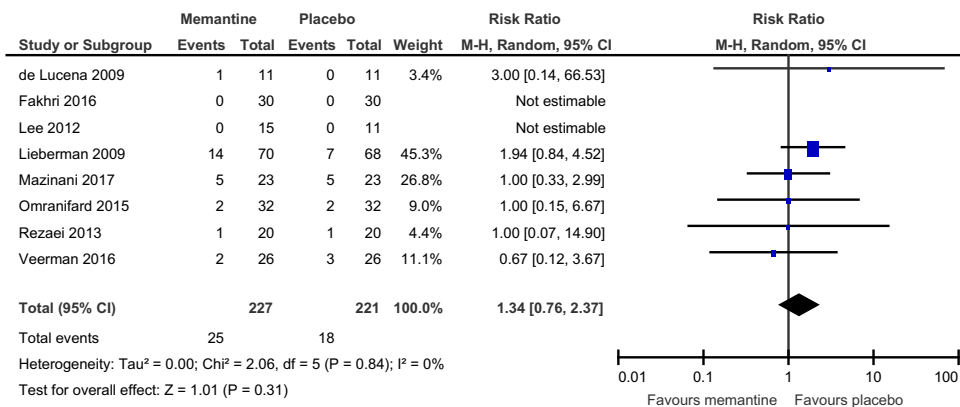


Fig. 2 The forest plots. MMSE Mini-Mental Status Examination, PANSS Positive and Negative Syndrome Scale

Table 2 The results of meta-analysis

Efficacy						
	<i>N</i>	<i>n</i>	SMD	95% CI	<i>p</i>	<i>I</i> ²
Overall symptoms	5	271	−0.75	−1.52 to 0.03	0.06	86%
Positive symptoms	7	367	−0.46	−0.96 to 0.05	0.07	80%
Negative symptoms	7	367	−0.96	−1.64 to −0.27	<i>0.006</i>	88%
Depressive symptoms	4	201	−0.127	−0.380 to 0.127	0.326	0%
	<i>N</i>	<i>n</i>	MD	95% CI	<i>p</i>	<i>I</i> ²
PANSS general subscale score	4	151	−1.62	−2.65 to −0.59	<i>0.002</i>	0%
CGI-S score	4	226	−0.19	−0.83 to 0.45	0.56	81%
MMSE score	3	83	−3.07	−4.46 to −1.69	<i><0.0001</i>	21%
	<i>N</i>	<i>n</i>	RR	95% CI	<i>p</i>	<i>I</i> ²
Discontinuation due to inefficacy	7	408	No patients during the study			
Safety						
	<i>N</i>	<i>n</i>	RR	95% CI	<i>p</i>	<i>I</i> ²
All-cause discontinuation	8	448	1.34	0.76–2.37	0.31	0%
Discontinuation due to adverse events	7	408	2.53	0.72–8.86	0.15	0%
Serious adverse events	4	271	1.00	0.34–2.93	0.99	0%
Fatigue	3	236	1.86	0.83–4.15	0.13	0%
Dizziness	6	332	1.64	0.79–3.42	0.19	0%
Headache	5	322	1.48	0.82–2.68	0.19	0%
Nausea	5	317	1.35	0.60–3.05	0.47	0%
Constipation	5	311	1.78	0.86–3.71	0.12	0%

Italic font indicates the significant results

95% CI 95% confidence interval, CGI-S Clinical Global Impression Severity, MD mean difference, MMSE Mini-Mental Status Examination, *N* number of comparisons, *n* number of patients, *na* not applicable, PANSS Positive and Negative Syndrome Scale, RR risk ratio, SMD standardized mean difference

the studies were conducted in Iran, duration of the studies was 4 and 6 weeks, sample size of the studies was 40 and 46 patients, and the studies were not industry sponsored. This risperidone subgroup also showed that the advantages of memantine add-on treatment over placebo remained for negative symptoms. Therefore, we conclude that memantine may have a benefit for negative symptoms in schizophrenia patients treated with risperidone, although the sample size in the risperidone subgroup was small ($n = 76$). Further studies are required to determine which antipsychotic should be used in association with memantine, because there are differences in the affinities for a variety of neurotransmitter receptor subtypes among the antipsychotics (Kishi et al. 2013). Further, although memantine is approved for use in those with moderate to severe Alzheimer's disease, it was surprising that the effect size of memantine with respect to overall and negative symptoms increased in schizophrenia patients who were younger adults. However, we did not address multiple comparisons; this significant effect in younger adults may be a false-positive error arising from the number of meta-regression analyses performed (Higgins and Green 2011). However, there was no evidence for an association between the effect size of memantine and age in Alzheimer's disease patients (Matsunaga et al. 2014, 2015).

There are strong lines of evidence indicating that dysfunction of glutamate receptors, such as NMDA receptors, may explain the pathophysiology of positive, negative, and cognitive symptoms of schizophrenia (Beck et al. 2016; Gray and Roth 2007; Hoflich et al. 2015; Kondziella et al. 2006; Miyamoto et al. 2012; Vukadinovic 2014). Our meta-analysis showed that memantine add-on was superior to placebo for the amelioration of negative symptoms, PANSS general subscale score, and MMSE scores. However, some phase III trials of other drugs (e.g., bitopertin and LY2140023) which are associated with the glutamate receptor hypofunction hypothesis for schizophrenia failed to significantly decrease schizophrenia symptoms compared to placebo (Beck et al. 2016). The discrepancy in the results may be explained by the following discussion: (1) the superiority of memantine add-on treatment over placebo in our meta-analysis may be influenced by a small study effect (Moreno et al. 2009). There was only one double-blind, placebo-controlled trial of memantine add-on treatment which was similar to the phase III trials of bitopertin and LY2140023; i.e., the memantine study included 138 schizophrenia patients and was conducted with industry sponsorship (Lieberman et al. 2009). This study showed that memantine add-on was not superior to placebo for ameliorating schizophrenia

Table 3 The results of sensitivity analysis

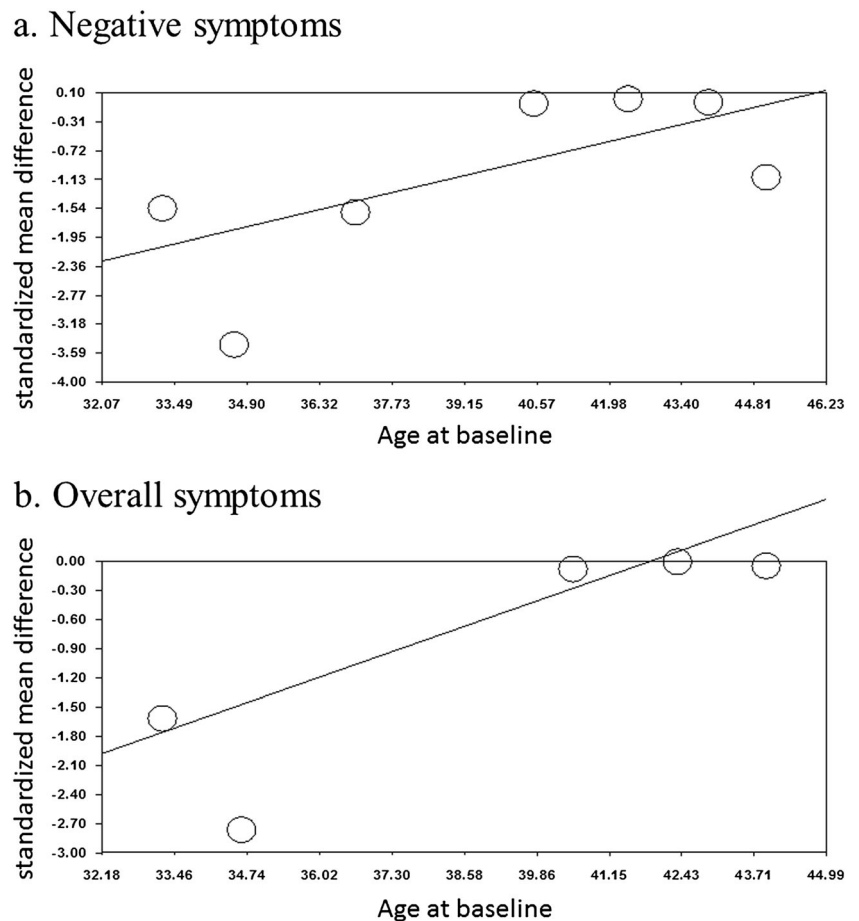
Overall symptoms									
Variable	Subgroup	<i>N</i>	<i>n</i>	<i>I</i> ²	SMD	95% CI	<i>p</i>	Test for subgroup differences	
								<i>p</i>	<i>I</i> ² (%)
Scale	BPRS	1	21	na	-2.65	-3.89 to -1.42	<0.0001	0.001	90.1%
	PANSS	4	250	80%	-0.39	-1.04 to 0.25	0.23		
Geographical region	Asia	2	66	88%	-0.82	-2.33 to 0.69	0.29	0.89	0%
	Other region	3	205	88%	-0.70	-1.70 to 0.31	0.17		
SGA vs. FGA	SGA	5	245	90%	-0.94	-1.90 to 0.01	0.05	0.15	51.5%
	FGA	1	26	na	-0.04	-0.82 to 0.74	0.92		
Clozapine	Clozapine	2	70	93%	-1.27	-3.87 to 1.33	0.34	0.61	0%
	Other antipsychotics	3	201	86%	-0.54	-1.48 to 0.39	0.25		
Sponsorship	Industry	1	135	na	-0.08	-0.41 to 0.26	0.66	0.12	58.9%
	Non-industry	4	136	86%	-0.99	-2.09 to 0.11	0.08		
Positive symptoms									
Variable	Subgroup	<i>N</i>	<i>n</i>	<i>I</i> ²	SMD	95% CI	<i>p</i>	Test for subgroup differences	
								<i>p</i>	<i>I</i> ² (%)
Scale	BPRS	1	21	na	-1.27	-2.22 to -0.31	0.009	0.61	63.1%
	PANSS	6	346	80%	-0.35	-0.87 to 0.17	0.18		
Geographical region	Asia	4	162	82%	-0.55	-1.34 to 0.23	0.17	0.58	0%
	Other region	3	205	69%	-0.28	-0.87 to 0.32	0.36		
Analyzed population	ITT/mITT	6	331	83%	-0.46	-1.04 to 0.13	0.13	0.94	0%
	OC	1	36	na	-0.49	-1.15 to 0.18	0.15		
SGA vs. FGA	SGA	6	341	82%	-0.54	-1.10 to 0.02	0.06	0.21	36.7%
	FGA	1	26	na	0.08	-0.70 to 0.86	0.85		
Clozapine	Clozapine	2	70	79%	-0.59	-1.79 to 0.61	0.33	0.81	0%
	Other antipsychotics	5	297	84%	-0.42	-1.05 to 0.21	0.19		
Risperidone	Risperidone	2	76	0%	-0.28	-0.73 to 0.17	0.23	0.55	0%
	Other antipsychotics	5	291	86%	-0.53	-1.24 to 0.17	0.14		
Sponsorship	Industry	1	135	na	0.03	-0.31 to 0.37	0.85	0.09	65.9%
	Non-industry	6	232	78%	0.06	-1.15 to 0.03	0.06		
Negative symptoms									
Variable	Subgroup	<i>N</i>	<i>n</i>	<i>I</i> ²	SMD	95% CI	<i>p</i>	Test for subgroup differences	
								<i>p</i>	<i>I</i> ² (%)
Scale	BPRS	1	21	na	-3.33	-4.74 to -1.92	<0.00001	0.0008	91.2%
	PANSS	6	346	86%	-0.69	-1.31 to -0.07	0.03		
Geographical region	Asia	4	162	73%	-1.08	-1.74 to -0.42	0.001	0.74	0%
	Other region	3	205	90%	-0.85	-2.00 to 0.31	0.15		
Analyzed population	ITT/mITT	6	331	90%	-0.95	-1.73 to -0.16	0.02	0.82	0%
	OC	1	36	na	-1.07	-1.78 to -0.37	0.003		
SGA vs. FGA	SGA	6	341	90%	-1.12	-1.89 to -0.34	0.005	0.05	73.3%
	FGA	1	26	na	-0.03	-0.81 to 0.74	0.93		
Clozapine	Clozapine	2	70	95%	-1.59	-4.87 to 1.68	0.34	0.66	0%
	Other antipsychotics	5	297	87%	-0.84	-1.57 to -0.11	0.02		
Risperidone	Risperidone	2	76	0%	-1.29	-1.79 to -0.79	0.00001	0.37	0%
	Other antipsychotics	5	291	90%	-0.84	-1.70 to 0.03	0.06		
Sponsorship	Industry	1	135	na	-0.05	-0.39 to 0.29	0.76	0.01	84.2%
	Non-industry	6	232	86%	-1.14	-1.92 to -0.36	0.004		

95% CI 95% confidence interval, BPRS Brief Psychiatric Rating Scale, FGA first generation antipsychotic, (m)ITT (modified) intention-to-treat, *n* number of patients, *N* number of studies, *na* not applicable, OC observed case, PANSS Positive and Negative Syndrome Scale, SGA second generation antipsychotic, SMD standardized mean difference

symptoms (Lieberman et al. 2009). (2) There were differences in those drugs' profiles which may have influenced the results

of our meta-analysis, because those drugs generally have affinities for a variety of neurotransmitter receptor subtypes,

Fig. 3 The meta-regression analysis



including glutamate, serotonin, dopamine, nicotinic, and acetylcholine receptors. Memantine is an antagonist for various receptors (NMDA receptors, serotonin-3 receptors, and nicotinic acetylcholine receptors, including alpha-7 receptor). Our previous meta-analysis showed that serotonin-3 receptor antagonist add-on to antipsychotics was superior to placebo for ameliorating negative symptoms (Kishi et al. 2014). The hypodopaminergic state in the prefrontal cortex and mesocortical pathways is related to the negative symptoms and cognitive dysfunction (Juckel 2016; Kambeitz et al. 2014). Because memantine is a dopamine D₂ receptor agonist (Seeman et al. 2008) as well, memantine may restore the antipsychotic-induced hypodopaminergic state in the prefrontal cortex and mesocortical pathway. (3) There was no evidence from a meta-analysis of other glutamate-related drugs, such as bitopertin. The placebo response of recent clinical trials was large (Beck et al. 2016). Thus, because the effect size between drug and placebo responses was small, it may be difficult for a single trial to estimate the superiority of drugs over placebo on efficacy due to low statistical power (i.e., insufficient sample size) (Beck et al. 2016; Folstein et al. 1975).

Although we analyzed the combined data from studies of memantine and amantadine to obtain greater statistical power

in our previous meta-analytic study (Matsuda et al. 2013), we found several differences in the pharmacological profiles of memantine and of amantadine (Shim and Nadeem 2014). Memantine is an antagonist for NMDA receptors, serotonin-3 receptors, and nicotinic acetylcholine receptor, including alpha-7 receptors; in addition, it is a dopamine D₂ receptor agonist. On the other hand, amantadine is a weak NMDA receptor antagonist and an alpha-7 nicotinic acetylcholine receptor antagonist with its effects, including the release of dopamine and norepinephrine. This was a serious limitation of our previous study. However, because four additional memantine studies were published since then, we conducted this updated meta-analysis using only data from memantine studies to accurately estimate the efficacy and safety of memantine add-on to antipsychotic treatment.

Our previous meta-analysis showed a significant heterogeneity in overall, positive, and negative symptoms (Matsuda et al. 2013). However, when excluding the de Lucena et al. (2009) study, all significances of heterogeneity disappeared. Only the de Lucena et al. (2009) study used BPRS in the evaluation of psychopathology from all the studies examined. Therefore, the subgroup, excluding the de Lucena et al. (2009) study, was same as the subgroups using only PANSS data (Table 3). When excluding the de Lucena et al. (2009) study

Table 4 The results of meta-regression analysis

Overall symptoms			
	Slope	95% CI	<i>p</i>
Age	0.211	0.09835 to 0.324	0.0002
Sample size	0.0241	−0.0196 to 0.0677	0.281
Study duration	−0.00427	−0.521 to 0.512	0.987
Publication year	0.149	−0.255 to 0.554	0.469
%Male	0.0530	−0.0209 to 0.127	0.160
Positive symptoms			
	Slope	95% CI	<i>p</i>
Age	0.0643	−0.0581 to 0.187	0.303
Sample size	0.00919	−0.0203 to 0.0387	0.541
Study duration	0.0845	−0.142 to 0.311	0.465
Publication year	−0.0375	−0.212 to 0.137	0.673
%Male	0.0105	−0.0185 to 0.0394	0.478
Negative symptoms			
	Slope	95% CI	<i>p</i>
Age	0.171	0.0262 to 0.315	0.0206
Sample size	0.0263	−0.0124 to 0.0651	0.183
Study duration	0.0383	−0.284 to 0.361	0.816
Publication year	0.0526	−0.207 to 0.312	0.692
%Male	0.0212	−0.0197 to 0.0621	0.310
PANSS general subscale scores			
	Slope	95% CI	<i>p</i>
Age	0.165	−0.159 to 0.489	0.319
Sample size	0.00844	−0.347 to 0.364	0.963
Study duration	0.438	−0.419 to 1.30	0.317
Publication year	0.521	−0.554 to 1.60	0.342
%Male	0.0580	−0.0607 to 0.177	0.338
MMSE scores ^a			
	Slope	95% CI	<i>p</i>
Age	0.163	−0.210 to 0.537	0.392
Sample size	0.0259	−0.318 to 0.370	0.882
Publication year	0.0636	−0.583 to 0.710	0.847
%Male	0.00941	−0.0698 to 0.0886	0.816

95% CI 95% confidence interval, MMSE Mini-Mental Status Examination, PANSS Positive and Negative Syndrome Scale

^a Because duration of all studies included in this outcome was 12 weeks, we did not perform a meta-regression analysis with respect to study duration

(i.e., subgroup using only PANSS data), all significances of heterogeneity remained in this study (Table 3). For MMSE score, when excluding the de Lucena et al. (2009) study, the superiority of memantine add-on over placebo remained (MD = −2.69, 95% CIs = −4.29 to −1.09, $p = 0.001$, $I^2 = 27%$; $N = 2$, $n = 62$).

This study had several limitations. First, although a greater number of patients were included in the meta-analysis than in the previous meta-analyses, the number of studies and patients remained insufficient (Trikalinos et al. 2004). Therefore, we cannot rule out a “small study effect,” in which smaller studies tend to show larger treatment effects than larger studies (Moreno et al. 2009). Second, patient characteristics differed between the studies examined, including symptom severity, inclusion criteria, race and ethnicity, and study duration. These differences could generate heterogeneity, when combining data for systematic review and meta-analysis. Third, because all studies evaluated had a short trial duration (mean = 10.25 weeks), we could not determine whether memantine add-on to antipsychotic treatment had any long-term effect on schizophrenia symptoms. Fourth, although we utilized a funnel plot to explore potential publication bias, this technique is generally used only when ≥ 10 studies are included in a meta-analysis and the current study examined only eight studies. Nevertheless, we did not detect any publication bias using Egger’s test (Table 4).

Conclusions

Our results suggest that memantine add-on to antipsychotic treatment demonstrated treatment efficacy for negative symptoms in schizophrenia patients and was well tolerated. The effect size of negative symptoms may be associated with schizophrenia patients who are younger adults. However, due to study limitations, a long-term study of memantine on a larger sample of schizophrenia patients is required.

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Compliance with ethical standards

Conflicts of interest Drs. Kishi, Matsuda, and Iwata declare that they have no direct conflicts of interest relevant to this study. No grant support or other sources of funding were used to conduct this study or prepare this manuscript.

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