ON THE STRENGTH OF CONNECTEDNESS OF A RANDOM GRAPH

By

P. ERDŐS (Budapest), corresponding member of the Academy, and A. RÉNYI (Budapest), member of the Academy

Let G be a non-oriented graph without parallel edges and without slings, with vertices $V_1, V_2, \ldots, V_n$. Let us denote by $d(V_k)$ the *valency* (or degree) of a point V_k in G, i. e. the number of edges starting from V_k. Let us put

$$c(G) = \min_{1 \leqq k \leqq n} d(V_k). \tag{1}$$

If G is an arbitrary non-complete graph, let $c_p(G)$ denote the least number k such that by deleting k appropriately chosen vertices from G (i. e. deleting the k points in question and all edges starting from these points) the resulting graph is not connected. If G is a complete graph of order n, we put $c_p(G) = n-1$. Let $c_e(G)$ denote the least number l such that by deleting l appropriately chosen edges from G the resulting graph is not connected. We may measure the strength of connectedness of G by any of the numbers $c_p(G)$, $c_e(G)$ and in a certain sense (if G is known to be connected) also by $c(G)$. Evidently one has

$$c(G) \geqq c_e(G) \geqq c_p(G). \tag{2}$$

It is known further that any two points of G are connected by at least $c_p(G)$ paths having no point in common, except the two endpoints (theorem of MENGER—WHITNEY, see [1] and [2]) and by at least $c_e(G)$ paths having no edge in common (theorem of FORD and FULKERSON, see [3]).

We shall denote by $\nu_r(G)$ the number of vertices of G which have the valency r $(r = 0, 1, 2, \ldots)$.

As in two previous papers ([4], [5]) we consider the random graph $\Gamma_{n,N}$ defined as follows: Let there be given n labelled points $V_1, V_2, \ldots, V_n$. Let us choose at random N edges among the $\binom{n}{2}$ possible edges connecting these n points, so that each of the $\binom{\binom{n}{2}}{N}$ possible choices of these edges should be equiprobable. We denote by $\Gamma_{n,N}$ the random graph thus obtained. We shall denote by $\mathbf{P}(\cdot)$ the probability of the event in the brackets.

The aim of this note is to investigate the strength of connectedness of the random graph $\Gamma_{n,N}$ when n and N both tend to $+\infty$, $N=N(n)$ being a function of n. As it has been shown in [4], the following theorem holds:

THEOREM 1. *If we have* $N(n)=\frac{1}{2}n\log n+\alpha n+o(n)$ *where* α *is a real constant, then the probability of* $\Gamma_{n,N(n)}$ *being connected tends to* $\exp(-e^{-2\alpha})$ *for* $n\to+\infty$.

In this paper we shall prove the following theorem:

THEOREM 2. *If we have* $N(n)=\frac{1}{2}n\log n+\frac{r}{2}n\log\log n+\alpha n+o(n)$ *where* α *is a real constant and* r *a non-negative integer, then*

$$\lim_{n\to+\infty}\mathbf{P}(c_p(\Gamma_{n,N(n)})=r)=1-\exp\left(-\frac{e^{-2\alpha}}{r!}\right), \tag{3}$$

further

$$\lim_{n\to+\infty}\mathbf{P}(c_e(\Gamma_{n,N(n)})=r)=1-\exp\left(-\frac{e^{-2\alpha}}{r!}\right) \tag{4}$$

and

$$\lim_{n\to+\infty}\mathbf{P}(c(\Gamma_{n,N(n)})=r)=1-\exp\left(-\frac{e^{-2\alpha}}{r!}\right). \tag{5}$$

REMARK. Clearly Theorem 2 can be considered as a generalization of Theorem 1. As a matter of fact, any of the statements $c_p(G)=0$ or $c_e(G)=0$ is equivalent to G not being connected and thus for $r=0$ (3) and (4) reduce to the statement of Theorem 1. It has been shown further in [4] that if $N(n)=\frac{n}{2}\log n+\alpha n+o(n)$ and $\Gamma_{n,N(n)}$ is not connected, then it consists almost surely of a connected component and of a few isolated points. Therefore (5) is for $r=0$ also equivalent to the statement of Theorem 1. Thus in proving Theorem 2 we may restrict ourselves to the case $r\geqq 1$.

The statement (5) of Theorem 2 gives information about the *minimal* valency of points of $\Gamma_{n,N}$. In a forthcoming note we shall deal with the same question for larger ranges of N (when $c(\Gamma_{n,N})$ tends to infinity with n), further with the related question about the *maximal* valency of points of $\Gamma_{n,N}$.

We shall prove further the following

THEOREM 3. *If we have* $N(n)=\frac{1}{2}n\log n+\frac{r}{2}n\log\log n+\alpha n+o(n)$ *where* α *is a real constant and* r *a non-negative integer, then we have*

$$\lim_{n\to+\infty}\mathbf{P}(\nu_r(\Gamma_{n,N(n)})=k)=\frac{\lambda^k e^{-\lambda}}{k!}\quad \textit{for}\quad k=0,1,\ldots \tag{6}$$

where $\lambda = \frac{e^{-2\alpha}}{r!}$; *in other words, the distribution of* $\nu_r(\Gamma_{n, N(n)})$ *tends to a Poisson distribution.*

PROOF OF THEOREMS 2 AND 3. Let $r \geqq 1$ be an integer and $-\infty < \alpha < +\infty$. Let us suppose that

$$N(n) = \frac{1}{2} n \log n + \frac{r}{2} n \log \log n + \alpha n + o(n). \tag{7}$$

Let $\Gamma_{n,N}$ be a random graph with the n vertices $V_1, V_2, \ldots, V_n$ and having N edges. Let $P_k(n, N, r)$ denote the probability that by removing r suitably chosen points from $\Gamma_{n,N}$ there remain two disjoint graphs, consisting of k and $n-k-r$ points, respectively. We may suppose $k < \left[\frac{n-r}{2}\right]$. First we have clearly

$$P_k(n, N, r) \leqq \binom{n}{r}\binom{n-r}{k} \frac{\binom{\binom{n}{2} - k(n-k-r)}{N}}{\binom{\binom{n}{2}}{N}}.$$

It follows by some obvious estimations that

$$\sum_{(r+3)\frac{\log n}{\log \log n} < k \leqq \left[\frac{n-r}{2}\right]} P_k(n, N(n), r) = O\left(\frac{1}{n}\right). \tag{8}$$

Now we consider the case $k \leqq (r+3)\frac{\log n}{\log \log n}$. Let $P_k^*(n, N, r)$ denote the probability that by removing r suitably chosen points (the set of which will be denoted by Δ) $\Gamma_{n,N}$ can be split into two disjoint subgraphs Γ' and Γ'' consisting of k and $n-k-r$ points, respectively, but that $\Gamma_{n,N}$ can not be made disconnected by removing only $r-1$ points. If $\Gamma_{n,N}$ has these properties and if s denotes the number of edges of $\Gamma_{n,N}$ connecting a point of Δ with a point of Γ', then we have clearly $s \geqq r$. Otherwise, by definition, $s \leqq rk$. Thus we have

$$P_k^*(n, N, r) \leqq \sum_{s=r}^{rk} \binom{n}{r}\binom{n-r}{k}\binom{rk}{s} \frac{\binom{\binom{n}{2} - k(n-k)}{N-s}}{\binom{\binom{n}{2}}{N}}. \tag{9}$$

It follows that

$$(10)\qquad \sum_{k=2}^{\left[(r+3)\frac{\log n}{\log\log n}\right]} P_k^*(n, N(n), r) = O\left(\frac{1}{\log n}\right).$$

From (8) and (10) it follows that for $n \to +\infty$

$$(11)\qquad \mathbf{P}(c_p(\Gamma_{n, N(n)}) = r) \sim \mathbf{P}(c(\Gamma_{n, N(n)}) = r).$$

As a matter of fact, (8) and (10) imply that if by removing r suitably chosen points (but not by removing less than r points) $\Gamma_{n, N(n)}$ can be split into two disjoint subgraphs Γ' and Γ'' consisting of k and $n-k-r$ points, respectively, where $k \leqq \left[\frac{n-r}{2}\right]$, then only the case $k=1$ has to be considered, the probability of $k>1$ being negligibly small. It remains to prove (5). This can be done as follows. First we prove that

$$(12)\qquad \lim_{n\to+\infty} \mathbf{P}(c(\Gamma_{n, N(n)}) \leqq r-1) = 0.$$

For $r=1$ this follows already from Theorem 1. Thus we may suppose here $r \geqq 2$. We have

$$\mathbf{P}(c(\Gamma_{n,N}) \leqq r-1) \leqq \sum_{h=1}^{r-1} n\binom{n-1}{h} \frac{\binom{\binom{n}{2}-(n-1)}{N-h}}{\binom{\binom{n}{2}}{N}},$$

and thus

$$(13)\qquad \mathbf{P}(c(\Gamma_{n, N(n)}) \leqq r-1) = O\left(\frac{1}{\log n}\right)$$

which proves (12).

Now let $\nu_r(\Gamma_{n,N})$ denote the number of vertices of $\Gamma_{n,N}$ which have the valency r. Then we have clearly by (12)

$$(14)\qquad \mathbf{P}(c(\Gamma_{n, N(n)}) = r) \sim \mathbf{P}(\nu_r(\Gamma_{n, N(n)}) \neq 0).$$

Now evidently

$$(15)\qquad \mathbf{P}(\nu_r(\Gamma_{n, N(n)}) \neq 0) = \sum_{j=1}^{n} (-1)^{j-1} S_j$$

where

$$(16)\qquad S_j = \sum\sum_{1\leqq k_1<k_2<\cdots<k_j\leqq n} \cdots \sum \mathbf{P}(d(V_{k_1}) = r, d(V_{k_2}) = r, \ldots, d(V_{k_j}) = r).$$

Evidently, if we stop after taking an even or odd number of terms of the sum on the right-hand side of (15), we obtain a quantity which is greater or smaller, respectively, than the left-hand side of (15). Now clearly

$$\mathbf{P}(d(V_k)=r)=\binom{n-1}{r}\frac{\binom{\binom{n}{2}-(n-1)}{N(n)-r}}{\binom{\binom{n}{2}}{N(n)}}\sim\frac{e^{-2\alpha}}{nr!},$$

and thus

$$\lim_{n\to+\infty} S_1=\frac{e^{-2\alpha}}{r!}. \tag{17}$$

Now let us consider $\mathbf{P}(d(V_{k_1})=r,\ d(V_{k_2})=r)$ where $k_1\neq k_2$. If both V_{k_1} and V_{k_2} have valency r, three cases have to be considered: a) either V_{k_1} and V_{k_2} are not connected, and there is no point which is connected with both V_{k_1} and V_{k_2}; b) or V_{k_1} and V_{k_2} are not connected, but there is a point connected with both; c) V_{k_1} and V_{k_2} are connected. We denote the probabilities of the corresponding subcases by $\mathbf{P}_a(d(V_{k_1})=r,\ d(V_{k_2})=r)$, $\mathbf{P}_b(d(V_{k_1})=r,\ d(V_{k_2})=r)$ and $\mathbf{P}_c(d(V_{k_1})=r,\ d(V_{k_2})=r)$, respectively. We evidently have

$$\mathbf{P}_a(d(V_{k_1})=r,\ d(V_{k_2})=r)=\frac{(n-2)!}{r!^2(n-2r-2)!}\frac{\binom{\binom{n}{2}-(2n-3)}{N(n)-2r}}{\binom{\binom{n}{2}}{N(n)}}\sim\left(\frac{e^{-2\alpha}}{n\cdot r!}\right)^2,$$

and thus

$$\sum\sum_{1\leq k_1<k_2\leq n}\mathbf{P}_a(d(V_{k_1})=r,\ d(V_{k_2})=r)\sim\frac{1}{2}\left(\frac{e^{-2\alpha}}{r!}\right)^2. \tag{18}$$

On the other hand (denoting by l the number of points which are connected with both V_{k_1} and V_{k_2}), we have

$$\mathbf{P}_b(d(V_{k_1})=r,\ d(V_{k_2})=r)=$$

$$=\sum_{l=1}^{r}\frac{(n-2)!}{l!(r-l)!(n-2r+l-2)!}\frac{\binom{\binom{n}{2}-(2n-3)}{N(n)-2r}}{\binom{\binom{n}{2}}{N(n)}}=O\left(\frac{1}{n^3}\right). \tag{19}$$

Similarly one has

$$\mathbf{P}_c(d(V_{k_1})=r,\ d(V_{k_2})=r)=$$

(20)
$$=\sum_{l=0}^{r-1}\frac{(n-2)!}{l!(r-l-1)!^2(n-2r+l)!}\frac{\binom{\binom{n}{2}-(2n-3)}{N(n)-2r}}{\binom{\binom{n}{2}}{N(n)}}=O\left(\frac{1}{n^4}\right).$$

Thus we obtain

$$\lim_{n\to+\infty}S_2=\frac{1}{2}\left(\frac{e^{-2\alpha}}{r!}\right)^2.$$

The cases $j>2$ can be dealt with similarly. Thus we obtain

$$\lim_{n\to+\infty}S_j=\frac{1}{j!}\left(\frac{e^{-2\alpha}}{r!}\right)^j\qquad(j=1,2,3,4,\ldots). \tag{21}$$

It follows from (16) and (21) that

$$\lim_{n\to+\infty}\mathbf{P}(\nu_r(\Gamma_{n,N(n)})\neq 0)=1-\exp\left(-\frac{e^{-2\alpha}}{r!}\right). \tag{22}$$

In view of (2), (11) and (14) Theorem 2 follows.

To prove Theorem 3 it is sufficient to remark that by the well-known formula of CH. JORDAN

$$\mathbf{P}(\nu_r(\Gamma_{n,N(n)})=k)=\sum_{j=0}^{n-k}(-1)^j\binom{j+k}{j}S_{j+k}, \tag{23}$$

and thus by (21), putting $\lambda=\dfrac{e^{-2\alpha}}{r!}$, we obtain for $k=0,1,\ldots$

$$\lim_{n\to+\infty}\mathbf{P}(\nu_r(\Gamma_{n,N(n)})=k)=\frac{\lambda^k}{k!}\sum_{j=0}^{\infty}\frac{(-1)^j\lambda^j}{j!}=\frac{\lambda^k e^{-\lambda}}{k!}. \tag{24}$$

Thus Theorem 3 is proved.

Our thanks are due to T. GALLAI for his valuable remarks.

(Received 12 October 1960)

References

[1] D. König, *Theorie der endlichen und unendlichen Graphen* (Leipzig, 1936).
[2] C. Berge, *Théorie des graphes et ses applications* (Paris, 1958).
[3] L. R. Ford and D. R. Fulkerson, Maximal flow through a network, *Canadian Journal of Math.*, **8** (1956), p. 399.
[4] P. Erdős and A. Rényi, On random graphs. I, *Publ. Math. Debrecen*, **6** (1959), pp. 290—297.
[5] P. Erdős and A. Rényi, On the evolution of random graphs, *Publ. Math. Inst. Hung. Acad. Sci.*, **5** (1960), pp. 17—61.