Chapter 14 The Integration of Work and Learning: Tackling the Complexity with Structural Equation Modelling

 Eva Kyndt and Patrick Onghena

 Abstract Researching professional learning within the paradigm of the integration of work and learning is interesting as it captures the complexity of workplace learning. However, it does require advanced statistical techniques that are able to model this complexity. Structural equation modelling (SEM) is one of the techniques that enable the examination of more complex relations. This book chapter aims at providing a basic introduction to SEM without using mathematical formulas and going into all the specific technicalities while at the same time staying true to the complexity of the presented analysis.

 The current book chapter starts with a general and conceptual presentation of SEM, including the advantages and disadvantages of the technique. The goal of this introduction is to address the questions of why and when SEM should be used and which conditions need to be fulfilled for a valid application of this technique. After a discussion of the fit indices that are used to evaluate the models, the analyses of different types of models are presented by means of an authentic dataset. More specifically, data regarding employees' approaches to learning at work were used to illustrate confirmatory factor analysis (including measurement invariance across groups), path analysis, and the analysis of a full SEM model. The chapter concludes by discussing several possible extensions of SEM and their relevance for researching professional learning.

E. Kyndt (\boxtimes)

P. Onghena

Centre for Research on Professional Learning & Development, and Lifelong Learning, KU Leuven - University of Leuven, Leuven, Belgium e-mail: eva.kyndt@ppw.kuleuven.be

Methodology of Educational Sciences Research Group, KU Leuven – University of Leuven, Leuven, Belgium

14.1 Introduction

 Throughout this book, the integration of learning and working takes a central role. Research into the learning potential of the workplace and the characteristics that promote or impede learning has already offered interesting insights (e.g. Hurtz & Williams, 2009; Kyndt & Baert, 2013; Lohman, 2000). Billett (2001) states that the workplace has to be designed in a way that people are invited and stimulated to learn. In addition, Tynjälä (2008) rightly pointed out that learning results from the interaction between the workplace and the individual: 'While the organisation of work sets the context and conditions for learning, it continues to be the reciprocal interaction between the individual and the workplace that determines learning' (Tynjälä, [2008](#page-36-0) , p. 141). When investigating professional learning, it is therefore important to pay attention to possible interactions and reciprocal and mediating relationships. Researching professional learning within the paradigm of the integration of work and learning is interesting as it captures the complexity of workplace learning. However, it does require more advanced statistical techniques that are able to model this complexity. Structural equation modelling (SEM) is one of the techniques that enable the examination of more complex relations. This book chapter aims at providing a basic introduction to SEM without using mathematical formulas and going into all the specific technicalities while at the same time staying true to the complexity of the presented analysis.

 The current book chapter will start with presenting SEM at a conceptual level. We will present why and when SEM could be used, what the advantages and disadvantages are in comparison with regression analysis and which different types of models can be analysed with SEM. Subsequently, the analyses will be illustrated on a dataset that was collected to investigate the approaches to learning of employees in relation with their work motivation, perceived workload and choice independence. Furthermore, several applications of SEM within the research of professional learning will be discussed. Finally, the main conclusions will be summarised.

14.2 Structural Equation Modelling (SEM)

 Structural equation modelling denotes a family of multivariate techniques including and combining factor analysis and path analysis in which the focus lies on theoretical constructs represented by latent factors (Hox & Bechger, [1998](#page-35-0)). Latent factors are unobserved constructs that are reflected by a set of observed variables. Within this section we will first focus on why and when SEM could be used. The different models within SEM will be introduced, and the assumptions of the analysis along with the conditions for conducting SEM will be discussed. Following, the fit indices used to evaluate the models are presented, and the use of modification indices for model improvement is considered.

 The core of the SEM analysis involves specifying a theoretical model and subsequently testing whether this model is plausible given the sample data. This comparison is based on the comparison of the variance-covariance matrix of the theoretical model to the variance-covariance matrix that is observed in the sample data (Crockett, [2012](#page-34-0)). Therefore, SEM is sometimes also known as covariance structure analysis. SEM thus examines a model that represents the linear relationships among variables. Because SEM is based on the analysis of covariances, a SEM model in itself cannot establish causal effects (see Sect. [14.5 \)](#page-31-0).

14.2.1 Why and When Should We Use SEM?

 The main reason for applying SEM instead of traditional regression analysis is the flexibility and ability to model more complex relationships between constructs. With SEM it is possible to specify 'path models with intervening variables between the independent and dependent variables, and latent factor as well' (Hox & Bechger, [1998](#page-35-0) , p. 6). Although the method of Baron and Kenny [\(1986 \)](#page-34-0) offers an alternative for assessing whether or not a variable mediates the relationship between two other variables, from a statistical point of view, analysing the different paths simultaneously will yield better results (Iacobucci, [2009](#page-35-0)). The estimates of the strength of the relationships are more precise, and there is less bias as each effect is estimated together with the other effects (Iacobucci, [2009](#page-35-0)). In other words, the same variance cannot be estimated twice, as it is the case when separate regression analyses are applied.

Chin (1998) states that for any given SEM model, alternative models that fit the data as well as the proposed model can be found that potentially provide substantially different explanations of the data. Therefore, it is important to note that SEM is traditionally not recommended for exploratory purposes. Clear hypotheses about the structure of the data both in terms of factors as in terms of paths between constructs are needed for sound and replicable applications (Hox & Bechger, 1998). The paths included in the model should be theoretically justified (Chin, 1998). Additionally, the more complex the specified model is, the higher the requirements in terms of sample size become. This issue will be discussed within the section focusing on the conditions that need to be fulfilled for the analysis.

 SEM models comprise a measurement model and structural model. The measurement model relates the observed or manifest variables to the latent constructs while the structural or path model denotes the paths/relationships between the constructs. Fitting a measurement model is also known as confirmatory factor analysis (Hox & Bechger, [1998](#page-35-0); Iacobucci, 2009). Full SEM models combine the two into one model. When performing SEM analyses, it is always convenient to start by drawing a path diagram; it can guide the analysis. Within this chapter, the generally accepted notation for representing SEM models will be used (Tacq, 1997). Within this notation, boxes represent the observed variables (e.g. the items in your questionnaire) and circles depict the latent constructs (e.g. the underlying construct

 Fig. 14.2 Conceptual model for a path analysis with single mediation between two observed variables

you are trying to measure using different items). Within the measurement model, the arrows originate from the latent construct and point to the observed variables. The underlying idea is that the latent construct gives rise to or is reflected in the observed variables (Chin, [1998](#page-34-0); Hox & Bechger, 1998; Iacobucci, 2009). Within the path model the single-headed arrows reflect the directional relationship between two constructs; double-headed arrows depict covariances.

 Figure 14.1 represents a measurement model with three latent constructs and eight observed variables. The arrows on the left side of the observed variables indicate the residual error term originating from the fact that the observed variation is not completely explained by the latent construct.

 Figure 14.2 depicts a conceptual model for a path analysis. The illustration shows a simple mediation model in which one variable mediates the relation between two other variables. Within a full SEM model (see below), the paths would connect latent constructs. A path analysis reflects the directional relations between observed variables (Cohen, Manion, & Morrison, 2011).

 Figure [14.3](#page-4-0) shows a full SEM model in which the measurement and the structural model are combined into one analysis.

 Fig. 14.3 Conceptual model for a full SEM model

14.2.2 Which Conditions Should Be Fulfilled?

 When applying SEM and making inferences based on the analysis, it is important to consider the underlying assumptions of the analysis. First, SEM imposes the same statistical assumptions on the data as traditional regression analysis. In addition, *multivariate normality* is assumed. Secondly, SEM assumes that the correct model has been specified, meaning that no relevant variables are missing and that the directional relationships are specified correctly. A correct *model specification* foremost depends on the theoretical grounds of the research at hand. Moreover, the observed variables in SEM are assumed to reflect the latent construct and not to cause it (Kline, 2012). This also explains why the arrows in Figs. [14.1](#page-3-0) and 14.3 point from the latent construct towards the observed variables and not vice versa. Changes in the latent construct should be reflected in all observed variables as they are conceptually related to each other. The observed variables in the SEM model cannot be indicators that compensate each other to form an artificial index or composite score (Kline, [2012](#page-35-0)). In the latter case of *formative indicators*, a change in one observed variable could cause a change in the construct but does not necessarily result in a change in the other observed variables. An example of a formative measure is when the success of an organisation is measured through the combination of an organisation's annual profits, the increase in number of staff members and an indication of the popularity of the company. It could be that the company grows more popular and at the same time a decrease in number of staff members occurs. The profits, number of staff members and popularity of the company might be good indicators, but they do not reflect a latent construct, as they are conceptually unrelated.

After the model has been specified, it is important to check whether the model can be identified, that is, 'whether a unique solutions to the model can be generated' (Crockett, [2012](#page-34-0), p. 36). *Model identification* can be executed by following two guidelines developed by Bollen (1989 in Crockett, [2012](#page-34-0)). First, the structural model should be recursive; this means that all relationships within the structural model are unidirectional and no feedback loops are included. In other words, the dependent variables in the model cannot be a cause and an effect at the same time (Crockett, [2012 ;](#page-34-0) Kline, [2012](#page-35-0)). In addition, the observed variance-covariance matrix must contain more unique elements than the number of parameters that need to be estimated (i.e. factor loadings, latent constructs, paths between latent constructs, etc.). The number of unique elements in the variance-covariance matrix can be calculated using $p(p+1)/2$, where *p* equals the number of observed variables. The latter guideline is also known as the *t* rule (Crockett, 2012).

 The invalidating effect of violating the statistical assumptions that SEM makes can be reduced by fulfilling conditions regarding sample size and missing values.

Sample size . The debate about the appropriate *sample sizes* for SEM is ongoing. For every 'rule of thumb' that exists, another occurs. In general, the rule is that the more complex the model, the more parameters that need to be estimated, the larger the sample size needs to be and, of course, larger is better (Iacobucci, [2010](#page-35-0)). The most correct and accurate method to assess the sample size is to assess the power of the analysis, as sample size depends on the specifications of the SEM model at hand (Chin, 1998). More information on power analysis can be found in MacCallum, Browne and Sugawara (1996). Iacobucci (2010) argues that the vague rule of thumb that the sample needs to be larger than 200, which was commonly accepted a while ago, is conservative and oversimplistic. In her article, she therefore argues that small samples of 50–100 could suffice. However, we would like to emphasise that this is only the case when you are testing a simple model with strong effects. One could wonder whether these simple models depict the true effects in an accurate way. The literature does offer some interesting rules of thumb that give an indication of an appropriate *sample size* in which the number of constructs or estimated parameters are taken into account. For the measurement model, the ratio of the sample size to the number of observed variables should at least be 10:1 (Hair, Black, Babin, Anderson, & Tatham, [2006](#page-34-0)). Bentler and Chou [\(1987](#page-34-0)) recommend that the ratio between the sample size and the number of parameters that need to be estimated should also be 10:1 or higher. More information on conducting SEM with small samples can be found in the article of Bentler and Yuan (1999). In addition, alternative estimation methods such as partial least squares (PLS) exist that are appropriate for small samples (see Sect. [14.4 \)](#page-28-0).

Missing values . By default, SEM only uses the data of participants without missing values. This approach assumes that if the dataset contains *missing values* , these values are missing completely at random. If this is not the case, one could adopt more advanced methods for handling missing values. For more information on this topic, the reader is referred to Allison (2003) .

 There is some evidence that SEM is robust to violations of the statistical assumptions if the sample size is large (more than 200 independent observations) and there are no missing values (Hsu, Chen, & Hsieh, [2006](#page-35-0) ; Hu, Bentler, & Kano, [1992 ;](#page-35-0) Yuan & Bentler, [1999](#page-36-0); Yuan & Zhong, 2013). However, with highly discrete and/or skewed data, especially if sample size is small or moderate, it is recommended to apply more robust estimation techniques and alternative statistics (for more information on these alterna-tives, see Bentler & Yuan, [1999](#page-34-0); Jung, [2013](#page-35-0); Kline, 2012; Satorra, [1990](#page-36-0)).

14.2.3 Fit Indices

 A wide variety of *fi t indices* have been proposed to evaluate the proposed model in terms of goodness of fit and simplicity of the model. Some fit indices emphasise the fit of the model to the data, while others take into account whether the model is parsimoni-ous (Hox & Bechger, [1998](#page-35-0); Iacobucci, [2010](#page-35-0)). In general, there is some agreement on which fit indices should be reported. First, there is the chi-square test, which is the only inferential measure. The null hypothesis of the chi-square test is that the model fits the data, meaning that to conclude that your model fits the data the chi-square test should not be significant. However, the chi-square test is very sensitive to sample size (Hox $\&$ Bechger, 1998). As a consequence, when working with large samples, the statistical test will be significant in almost all real data applications (Hox $\&$ Bechger, 1998; Iacobucci, [2010](#page-35-0)). One might think that it would be advisable not to work with large samples; however, this is not a valid advice, as a sufficiently large sample size is necessary to support the precision of the parameter estimation (Iacobucci, 2010). Alternatively, it has been suggested that fit is acceptable when the *ratio of the chi-square test statistic to the degrees of freedom* is not larger than $3: \chi^2/\text{df} \leq 3$ (Iacobucci, [2010](#page-35-0)).

Due to the sensitivity of the chi-square test, the fit of the model is always evaluated based on *multiple alternative indices* . Because all goodness-of-fi t indices are some function of the chi-square test, the majority of these indices are also subjected to the sample size but to a much smaller degree than the chi-square test (Hox & Bechger, [1998 \)](#page-35-0). As the following indices are not inferential, no statistical hypothesis testing is involved, only guidelines or 'rules of thumb' can be offered (Hu & Bentler, 1999; Iacobucci, [2010](#page-35-0)). Table [14.1](#page-7-0) summarises the different guidelines that are offered in the standard methodological literature. Below we will describe the most commonly used cut-off values. In general, authors are advised to report the *comparative fit index* (*CFI*) that captures the relative goodness of fit in comparison to a simpler model. In a sense it indicates whether making your model more complex actually pays off. Preferably, the CFI is close to or higher than .95 (Hu & Bentler, [1999 ;](#page-35-0) Iacobucci, [2010](#page-35-0)), but values starting from .90 are considered acceptable (Iacobucci, 2010). Some authors also report the Tucker-Lewis Index (TLI) also known as the Non-normed Fit Index (NNFI), which follows the same rules of thumb as the CFI. Next to the chi-square test and CFI, the *SRMR* or *Standardised Root Mean square Residual* is usually reported. The SRMR actually indicates to what extent your model does not fit the data. The higher the value, the worse the model fit. This SRMR largely depends on the factor loadings in the measurement model and is less

Type fit index	Fit index	Adequate fit index
Absolute fit indexes	Standardised Root Mean Square Residual (SRMR)	$SRMR \leq .08$ (Hu & Bentler, 1999; MacCallum et al., 1996) $SRMR \leq .05 = good$ (Byrne, 2001; Jaccard & Wan, 1996
	Jöreskog-Sörbom Goodness-of-Fit Index (GFI)	$CFI \ge .95$ (Browne & Cudeck, 1993; Hu & Bentler, 1999)
	Index (AGFI)	Adjusted Goodness-of-Fit AGFI \geq 95 (Hu & Bentler, 1999)
Incremental fit indexes/ comparative fit indexes/ relative fit indices	Bentler Comparative Fit Index (CFI)	$CFI \ge .95$ (Hu & Bentler, 1999) $CFI \geq .90$ (Bentler, 1992; Byrne, 2001)
	Normed Fit Index (NFI) Tucker-Lewis index $(TLI) = NNFI$: Non-normed Fit Index	$NFI \ge .95$ (Hu & Bentler, 1999) TLI \geq .95 (Hu & Bentler, 1999)
Parsimony-adjusted fit indexes	Root Mean Square Error of Approximation (RMSEA)	RMSEA \leq .06 (Hu & Bentler, 1999) RMSEA $.06 \rightarrow .08$ = reasonable error (Browne & Cudeck, 1993) RMSEA $.08 \rightarrow 1$ = reasonable error (MacCallum et al., 1996)

Table 14.1 Cut-off values for fit indices

prone to violations of the distributional assumption. The maximum value of this index equals 1 and (very) low values are preferred. Values below .08 indicate an acceptable model. Finally, most researchers also report the *Root Mean Square Error of Approximation* (*RMSEA*) and its 90 % confidence interval. The RMSEA was developed to provide an indication of the extent to which the model does not match the true model. Small values indicate a good match. An RMSEA smaller than .06 is advised, although a value lower than .08 can also be considered acceptable (Browne & Cudeck, 1993; MacCallum et al., [1996](#page-35-0)).

14.2.4 Modification Indices

When the model fit is not satisfactory, a researcher could choose to modify his model. This modification can be based on the parameter estimates (e.g. removing nonsignificant paths or observed variables with low loadings) or on the mod *ification indices* that indicate which parameter(s) should be added to the model to improve the model fit and how much the chi-square statistic is expected to minimally decrease when that parameter would be added (Hox & Bechger, 1998; Iacobucci, 2009). There is some debate on the use or misuse of *modification indices* (Hox & Bechger, 1998; Iacobucci, [2009](#page-35-0)). At first glance, these modification indices appear to be very helpful. Often these modification indices

are used to improve the model fit sequentially, up till the point that the model meets the requirements. A strong advice regarding the use of the modification indices for improving the model fit is that alterations to the hypothesised model based on the data should only be done when there are theoretical grounds that support these alterations (Chin, 1998 ; Hox & Bechger, 1998). However, some researchers seem to be very creative in evaluating the theoretical justification post hoc, leading to potentially incorrect models or models that are difficult to cross-validate because they rely too much on the data of the sample at hand (Hox & Bechger, 1998; Iacobucci, [2009](#page-35-0)). It is therefore proposed to use modification indices prudently and to compare different a priori constructed models (Hox $& Becher, 1998$ $& Becher, 1998$).

14.3 Analysing and Interpreting Data

Within this section the goal is to offer some guidelines for making decision about the data and the steps that could be followed within the analysis, as well as the interpretation of the results. However, bear in mind that how the model is built and the decisions taken within this process need to be theory driven. Crockett (2012) describes five sequential steps within SEM. The first two steps were discussed above: model specification and model identification. Subsequently, the model is estimated. Different estimation procedures can be used for the estimation of the variance- covariance matrix of the model, within this chapter the maximum likelihood estimation will be used. For an introduction into the different estimation methods, we refer the reader to Crockett (2012) . Below, we will foremost focus on the fourth and fifth step: model testing and model modification. Model testing involves the evaluation of the plausibility of the theoretical model given the sample data (Crockett, 2012). This evaluation is based on multiple fit indices (cf. infra). The final step of model modification was already shortly introduced above and will also be illustrated below.

14.3.1 Illustration: Concept and Sample

 The analyses will be illustrated with a data set that was collected to investigate the relationship between employees' approaches to learning at work and their work motivation, perceived workload and choice independence.

14.3.1.1 Concepts and Measurement Instruments

 Within the literature three approaches to learning at work are distinguished: a deep approach, a surface-rational approach and a surface-disorganised approach. A deep approach to learning refers to the combination of an eagerness to learn and the use

of integrative strategies that contribute to personal understanding. The surfacerational approach reflects a preference for orderly, accurate and detailed work achieved by using surface learning strategies such as memorisation and a step-bystep approach. Finally, the surface-disorganised approach is considered a nonacademic orientation in combination with surface motives. It is associated with feeling overwhelmed and a sense of incompetence when executing task. These approaches to learning at work were measured by means of the *Approaches to learning at Work Questionnaire* (Kirby, Knapper, Evans, Carty, & Gadula, [2003](#page-35-0)).

 Work motivation was conceptualised from the perspective of the self- determination theory. Within this study the focus lied on the reasons why someone does a particular job and a distinction between autonomous and controlled motivation is made. Work motivation was measured with the *Motivation at Work Scale* (Gagné et al., 2010). Perceived workload and choice independence were measured with the *Workplace Climate Questionnaire* (Kirby et al., 2003). The complete theoretical background, the rationale and the results of the actual study can be found in the article of Kyndt, Raes, Dochy and Janssens (2013).

14.3.1.2 Sample

 The *sample* consisted of 358 employees from diverse companies (59 % female). The majority of the participants were employed in profit organisations (52%) , and 38 % were employed in nonprofit or social profit organisations (e.g. healthcare). The remaining 10 % of the participants were employed within the public sector. Participants were between 20 and 64 years old $(M=37.85, SD=10.64)$; on average they had 11.22 years of seniority (SD = 10.16). Most respondents had a permanent full-time contract (83 $\%$); others worked part time (14 $\%$) or had a temporary contract (3 %). Finally, participants' initial level of education was diverse: '1 % did not obtain a diploma or finished elementary school, 25 % obtained a secondary degree, 40 % obtained a bachelor's degree (professional or academic), and 34 % obtained a master's degree' (Kyndt et al., [2013 ,](#page-35-0) p. 278).

14.3.1.3 Software and Output

 For this illustration the analyses were performed with the *lavaan* package of R (Rosseel, 2012). R is a free software for statistical computing that can be downloaded from [www.R-project.org](http://www.r-project.org) (R Development Core Team, [2012](#page-36-0)). Figure [14.4](#page-10-0) was plotted by means of the *qgraph* package (Epskamp, Cramer, Waldorp, Schmittmann, & Borsboom, [2012](#page-34-0)). The R code of this example can be found in the appendix. SEM analysis can also be performed with the AMOS package (Extension of SPSS; Arbuckle, [2011 \)](#page-34-0), SAS Calis procedure (SAS Institute Inc, [2008](#page-36-0)), Mplus (Muthén & Muthén, [1998–2010](#page-36-0)), EQS (Bentler, [2004](#page-34-0)) or Lisrel (Jöreskog $& Sörbom, 1997$). For this illustration, we chose to present the output as given by R so that the reader would recognise these output when undertaken the analysis themselves. This output presents more information than discussed

Fig. 14.4 Measurement model with standardised coefficients

within this introductory chapter; therefore, we have marked the values on which the interpretations are based. When performing the analysis in R , the first steps that need to be undertaken are setting a working directory, loading the data and installing the necessary packages.

14.3.2 Measurement Model: Confirmatory Factor Analysis

We will illustrate the *confirmatory factor analysis* with the simplest measurement that was included within the study, that is, the measurement of perceived workload and perceived choice independence of the employee (Kirby et al., [2003 \)](#page-35-0). Both constructs were measured by five items (i.e. observed variables) that were scored on a five-point Likert scale (see Table 14.2).

14.3.2.1 Confirming the Model

The results of the CFA show that the hypothesised measurement model fits the data reasonably well (Output 1). Although the chi-square test is statistically significant $(\chi^2 = 98.223, df = 34, p < .001)$, the ratio between the test statistic and degrees of freedom (98.223/34 = 2.89) is below 3. In addition, the *CFI* and *TLI,* respectively, equal .94 and .92, which is above the proposed cut-off of .90. The *SRMR* and *RMSEA* are rather high $(SRMR = .075, RMSEA = .073, CI 90 % [0.056, .09]),$ but both are acceptable, although it is not a good sign that the values within the confidence interval exceed .08.

```
lavaan (0.5-12) converged normally after 31 iterations
Used Total
 Number of observations 354 359
  Estimator<br>Minimum Function Test Statistic 198.223
  Degrees of freedom 34
P-value (Chi-square) 0.000
Model test baseline model:
 Minimum Function Test Statistic 1176.569<br>Degrees of freedom 45
 Degrees of freedom 45<br>P-value 0.000
  P-value 0.000
Full model versus baseline model:
 Comparative Fit Index (CFI) (0.943<br>Tucker-Lewis Index (TLI) (0.925Tucker-Lewis Index (TLI)
Loglikelihood and Information Criteria:
 Loglikelihood user model (H0) -4804.185<br>Loglikelihood unrestricted model (H1) -4755.074
 Loglikelihood unrestricted model (H1)
 Number of free parameters 21<br>Akaike (AIC) 9650.371
 Akaike (AIC) 19650.371<br>Bayesian (BIC) 19731.626
 Bayesian (BIC) 9731.626<br>Sample-size adjusted Bayesian (BIC) 9665.005
 Sample-size adjusted Bayesian (BIC)
Root Mean Square Error of Approximation:
  RMSEA 0.073
90 Percent Confidence Interval 0.056 0.090
 P-value RMSEA \leq 0.05 0.013
Standardized Root Mean Square Residual:
  SRMR \begin{pmatrix} 0.075 \end{pmatrix}Parameter estimates:
 Information Expected
 Standard Errors
               Estimate Std.err Z-value P(>|z|) Std.lv Std.all
Latent variables:
 Workload =~
   WL1<br>
WL2 0.611 0.063 9.686 0.000 0.615 0.512<br>WL3 0.197 0.061 3.218 0.001 0.199 0.181<br>WL4 0.769 0.062 12.432 0.000 0.774 0.633<br>WL5 0.961 0.057 16.854 0.000 0.967 0.859
 Choice.independence =\frac{1000}{C12} 1.000<br>CT2 0.815
   CI1 1.000 0.809 0.743
CI2 0.815 0.079 10.353 0.000 0.659 0.619
   CI3 0.892 0.079 11.263 0.000 0.721 0.680
CI4 0.808 0.072 11.240 0.000 0.653 0.678
CI5 0.836 0.075 11.173 0.000 0.676 0.674
Covariances:
 Workload ~~
   Choic.ndpndnc -0.231 0.054 -4.259 0.000 -0.284 -0.284
Variances:
   WL1 0.310 0.050 0.310 0.234
   WL2 1.061 0.084 1.061 0.737
   WL3 1.166 0.088 1.166 0.967
   WL4 0.896 0.075 0.896 0.600
   WL5 0.332 0.048 0.332 0.262
   CI1 0.529 0.056 0.529 0.447
   CI2 0.698 0.061 0.698 0.616
   CI3 0.605 0.057 0.605 0.538
   CI4 0.500 0.047 0.500 0.540
   CI5 0.549 0.051 0.549 0.546
Workload 1.013 0.106 1.000 1.000
   Choic.ndpndnc 0.654 0.089 1.000 1.000
```
 Output 1 CFA all items

Scale	Item nr.	Ouestion
Workload	WL1	The workload here is too heavy
	WL ₂	It sometimes seems to me that my job requires me to do too many different things
	WL3	In this organisation you're expected to spend a lot of time learning things on your own
	WL4	There seems to be too much work to get through here
	WL5	There's a lot of pressure on you as an employee here
Choice independence	CI1	There is a real opportunity in this organisation for people to choose the particular tasks they work on
	CI2	The organisation really seems to encourage us to develop our own work-related interests as far as possible
	CI3	We seem to be given a lot of choice here in the work we have to do
	CI4	This organisation gives you a chance to go about your work in ways which suit your own way of learning
	CI ₅	Employees here have a great deal of choice over how they learn new tasks

 Table 14.2 Items measurement model

Table 14.3 Model fit without item WL3

Instrument	ν^2/df	CFI	TLI	RMSEA 90 % CI	SRMR
Workload and choice independence	2.40	.97	.95	$.063$ [$.043$; $.083$]	.052

In sum, the measurement model shows an acceptable fit, but there is room for improvement. Because the SRMR is rather high, it could be interesting to examine the factor loadings of the different items and delete item(s) with low factor loadings. Ideally the standardised values of the factor loadings (see std.all in output) are around or above .50 (Hair et al., 2006; Maruyama, 1998). Within this solution, only item WL3 has a factor loading below .50. If we look at the questions in Table 14.2, we can observe that WL3 focuses on a different aspect in comparison with the other four items. WL3 is the only item that considers learning. Therefore, we decided to test a second measurement model in which item WL3 was excluded. The results show an improved fit (Table 14.3 and Output 2).

```
lavaan (0.5-12) converged normally after 31 iterations
                                                 Used Total<br>355 359
 Number of observations 355
  Estimator ML<br>Minimum Function Test Statistic 62.391
 Degrees of freedom P-value (Chi-square) (26P-value (Chi-square)
Model test baseline model:
  Minimum Function Test Statistic 1128.760<br>Degrees of freedom 36
 Degrees of freedom 36<br>
P-value 0.000
  P-value 0.000
Full model versus baseline model:
 Comparative Fit Index (CFI) (0.967)<br>Tucker-Lewis Index (TLI) (0.954)
 Tucker-Lewis Index (TLI)
Loglikelihood and Information Criteria:
 Loglikelihood user model (H0) -4286.378<br>Loglikelihood unrestricted model (H1) -4255.182
 Loglikelihood unrestricted model (H1)
 Number of free parameters 19<br>Akaike (AIC) 8610.755
 Akaike (AIC) 8610.755<br>Bayesian (BIC) 8684.326
 Bayesian (BIC) 8684.326<br>Sample-size adjusted Bayesian (BIC) 8624.049
 Sample-size adjusted Bayesian (BIC)
Root Mean Square Error of Approximation:
  RMSEA 0.063
 90 Percent Confidence Interval 0.043 0.083<br>P-value RMSEA <= 0.05 0.136
 P-value RMSEA \leq 0.05Standardized Root Mean Square Residual:
  SRMR 0.052
Parameter estimates:
 Information Expected<br>Standard Errors Expected Standard
 Standard Errors
                  Estimate Std.err Z-value P(>|z|) Std.lv Std.all
Latent variables:
 Workload =~<br>WL1
    WL1 1.000 1.009 0.878
WL2 0.604 0.063 9.590 0.000 0.609 0.508
WL4 0.765 0.062 12.403 0.000 0.771 0.632
   WL5 0.954 0.057 16.700 0.000 0.962 0.855
 Choice.independence =\sim CI1 1.000
   CI1 1.000 0.806 0.742
   CI2 0.815 0.079 10.315 0.000 0.658 0.617<br>CT3 0.893 0.079 11.242 0.000 0.720 0.679
    CI3 0.893 0.079 11.242 0.000 0.720 0.679
CI4 0.809 0.072 11.230 0.000 0.653 0.678
CI5 0.837 0.075 11.155 0.000 0.675 0.673
```
 Output 2 CFA without WL3

Output 2 (continued)

 Table 14.4 Measurement invariance for males and females

Model		Model comparison						
	χ^2 (df)	CFI	RMSEA	BIC		$\Delta \gamma^2$ (Δdf)	<i>p</i> -value \triangle CFI	
Model 1	$88.810**$ (52)	.966	.063	8833.54				
Model 2 (equal loadings)	$100.039**$ (59)	.962	.063	8803.69	Model 1 vs. 11.228 (7)		.129	.004
Model 3 (+equal intercepts)	$103.398**$ (66)	.965	.057	8765.96	Model 2 vs.	3.359(7)	.850	$-.003$
Model 4 (+equal errors)	$113.432***$ (68)	.958	.061	8764.25	Model $3 \text{ vs. } 10.034(2)$ 4		< 01	.007

* *p* < .05; ** *p* < .01; *** *p* < .001

Because both models show an acceptable fit, both could be accepted. The decision for the most appropriate measurement model should foremost be guided by theory. In this case, one could consider whether or not investing time in learning is an important aspect of the construct of perceived workload. In addition, you could also test which model is superior in a statistical sense. Because the first model can be considered an extension of the second model (i.e. the models are nested), the *chi- square test for model comparison* can be applied. By using a simple formula in Excel (=CHIDIST(Δ*chi-statistics*¹ ; Δ*degrees of freedom*)), you can calculate whether the difference between the chi-square statistics of both models is statistically significant. In this example, the difference between the chi-square statistics equals 98.223−62.391 = 35.832, and the difference between the degrees of freedom equals 34–36 = 8. The significance test reveals that this difference is statistically significant $(p<.001)$ indicating the second model is statistically superior to the first (a smaller chi-square value indicates a better fit). From a theoretical point of view, the conclusion is also supported because the emphasis within perceived workload lies on general workplace conditions, regardless of the fact that the effort had to be undertaken for learning specifically. Figure [14.4](#page-10-0) depicts the measurement model.

 Δ = difference.

14.3.2.2 The Stability of the Model: Measurement Invariance

 Recently, more and more attention has been given to the stability of the measurement model (Boeve-de Pauw, Jacobs, & Van Petegem, [2012](#page-34-0) ; Coertjens, Donche, De Maeyer, Vanthournout, $\&$ Van Petegem, [2013](#page-34-0)). If the goal is to compare the constructs across groups and/or over time – which is often the case – it is important to determine whether a questionnaire measures the same constructs with the same structure across groups and/or over time. When *measurement invariance* is established, it can be accepted that different groups of participants (e.g. males and females) or the same participants across different measurement moments (longitudinal measurement invariance) interpret the individual questions and underlying constructs in a similar way.

 Different levels (less to more demanding) of measurement invariance are described: configural, metric, scalar, and strict invariance. Configural invariance (Model 1) refers to the fact that the basic model structure is invariant across groups (Boeve-de Pauw et al., [2012 \)](#page-34-0). It shows that the participants conceptualise the constructs similarly. It basically shows that the same structure – the same items belong to the same construct – holds for both groups or both measurement moments. Configural invariance, however, does not ensure that the separate items are interpreted similarly because the factor loadings of the items can be different across groups or measurement moments.

To test whether metric invariance is achieved, one can compare the configural model to a model in which the factor loadings are constrained (Model 2) to be equal for each group or at each measurement moment. If constraining the factor loadings does not result in a significantly less good fit of the model, metric invariance can be claimed. A model is considered as good as the previous model if the difference between the *CFI* s of both models is smaller than .01. Ideally the difference in the chi-square test is also not significant; however, it is known that the chi-square statistic is sensitive to the size of the sample (Iacobucci, 2010).

 Metric invariance indicates that the items are interpreted in a similar way across groups or measurement moments. When the aim is to compare means of the latent constructs across groups, it is necessary to achieve scalar invariance. Scalar invariance indicates that differences in means of the observed items are a consequence of the differences in means of the latent constructs. To identify scalar invariance, the model in which both the loadings and the intercepts of the items are constrained to be equal across groups or measurement moments (model 3) is compared to the model in which only the loadings (model 2) are constrained. When the model with constrained loadings and intercepts results into a too large decrease of the CFI, one could explore the option of freeing some of the intercepts and achieving partial intercept invariance. Advanced statistical models such as multiple-indicator growth analysis in case of longitudinal analysis can take these intercept variances into account (e.g. Coertjens et al., [2013](#page-34-0)). Finally, one can also check for invariance in error variances (strict invariance). However, in practice full measurement invariance is rarely achieved nor necessary (Boeve-de Pauw et al., [2012](#page-34-0)).

 Measurement invariance *across groups* can be tested relatively easy in R with the *semTools* package (Pornprasertmanit, Miller, Schoemann, & Rosseel, [2013 \)](#page-36-0). The specific commands can be found in the Appendix.

```
Measurement invariance tests:
Model 1: configural invariance:
   chisq df pvalue cfi rmsea bic
88.810 52.000 0.001 0.966 0.063 8833.541
Model 2: weak invariance (equal loadings):
 chisq df pvalue cfi rmsea bic
100.039 59.000 0.001 0.962 0.063 8803.685
[Model 1 versus model 2]
delta.chisq delta.df delta.p.value delta.cfi (11.228) (0.000) (0.129) (0.004)1 (1.00) (Model 3: strong invariance (equal loadings + intercepts):<br>
chisq df pvalue cfi rmsea bic
 chisq df pvalue cfi rmsea bic
103.398 66.000 0.002 0.965 0.057 8765.959
[Model 1 versus model 3]
   delta.chisq delta.df delta.p.value delta.cfi
14.588 14.000 0.407 0.001
[Model 2 versus model 3]
   delta.chisq delta.df delta.p.value delta.cfi<br>
3.359 7.000 0.850 -0.003
Model 4: equal loadings + intercepts + means:
 chisq df pvalue cfi rmsea bic
113.432 68.000 0.000 0.958 0.061 8764.254
[Model 1 versus model 4]
   delta.chisq delta.df delta.p.value delta.cfi
24.622 16.000 0.077 0.008
[Model 3 versus model 4]
   delta.chisq delta.df delta.p.value delta.cfi<br>
(10.034) (2.000 0.007 0.007
```
 Output 3 Measurement invariance

 Within our illustration, we checked whether males and females interpreted the items and constructs in a similar way. In other words the measurement invariance across males and females was tested. The results can be found in Output 3 and are preferably reported by means of a table (see Table [14.4](#page-14-0)). The results in Table [14.4](#page-14-0) show that the measurement of perceived workload and choice independence reaches scalar invariance. The differences between the *CFI* s are below .01 and the chi- square test for model comparison show that the different models do not differ significantly from each other.

 Establishing *longitudinal measurement invariance* follows the same procedure as establishing measurement invariance over groups:

- Testing configural invariance
- Constraining factor loadings to be equal and comparing this to the configural invariance model (metric invariance)
- Constraining factor loadings and intercepts to be equal and comparing this to the metric invariance model (scalar invariance)

 However, the difference with establishing measurement invariance across groups is that longitudinal measurement invariance is assessed for each scale separately and not the instrument as a whole (Coertjens, Donche, De Maeyer, Vanthournout & Van Petegem, [2012 \)](#page-34-0). Currently, longitudinal measurement invariance cannot be assessed by means of an R-package. This type of analysis is foremost executed with the Mplus

Fig. 14.5 Path diagram illustration path analysis (*Note*: *WL* workload, *CI* choice independence, *Auto.* autonomous motivation, *Contr.* controlled motivation, *DA* deep approach)

Fig. 14.6 Path diagram illustration path analysis without nonsignificant paths

software (Muthén & Muthén, [1998–2010](#page-36-0)). In our article on the development of a self-assessment instrument for the generic competences of vocational education students, an example of longitudinal measurement invariance testing relevant for the field of professional learning can be found (Kyndt et al., [accepted](#page-35-0)).

14.3.3 Structural Model: Path Analysis

 An illustration of a model including the two motivational scales as mediating variables is provided below. Figure 14.5 represent the path diagram of the model under examination.

The *path analysis* including the two mediating variables shows a fit that is not optimal. Moreover, because the model includes almost every possible relationship, the model cannot be considered parsimonious. The poor model fit is evident from a ratio between the chi-square and degrees of freedom that is too large $(\chi^2/df = 5.92)$ and an *RSMEA* of .12 which is also too large. Simplifying the model by excluding nonsignificant paths could be a solution to this problem. Based on the output (Output 4), we decided to exclude the path between controlled motivation and a deep approach to learning, and the path between workload and autonomous motivation. Because the interest lies in predicting employees' deep approach to learning and controlled motivation was inserted as a mediator between the perception of workload and choice independence, the paths from workload and choice independence to controlled motivation were also removed (Output 5). The new model that will be tested is represented by a path diagram in Fig. 14.6 .

```
lavaan (0.5-12) converged normally after 19 iterations
  Number of observations 359
  Estimator ML<br>Minimum Function Test Statistic 3.923
 Degrees of freedom \begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix}P-value (Chi-square)
Model test baseline model:
 Minimum Function Test Statistic 249.403
 Degrees of freedom 9<br>
P-value 0.000
  P-value 0.000
Full model versus baseline model:
  Comparative Fit Index (CFI)<br>Tucker-Lewis Index (TLI) 0.816
 Tucker-Lewis Index (TLI)
Loglikelihood and Information Criteria:
  Loglikelihood user model (H0) -5099.399<br>Loglikelihood unrestricted model (H1) -5096.438
 Loglikelihood unrestricted model (H1)
 Number of free parameters 11<br>Akaike (AIC) 10220.798
 Akaike (AIC) 10220.798<br>Bayesian (BIC) 10263.515
 Bayesian (BIC) 10263.515<br>Sample-size adjusted Bayesian (BIC) 10228.617
 Sample-size adjusted Bayesian (BIC)
Root Mean Square Error of Approximation:
  RMSEA 0.117
90 Percent Confidence Interval 0.041 0.215
 P-value RMSEA \le 0.05 0.069
Standardized Root Mean Square Residual:
  SRMR 0.029
Parameter estimates:
 Information Expected
 Standard Errors
                 Estimate Std.err Z-value P(>|z|) Std.lv Std.all
Regressions:
 DA ~<br>Autonomous
    Autonomous 0.266 0.062 4.313 0.000 0.266 0.219
Controlled 0.085 0.056 1.535 0.125 0.085 0.071
 Autonomous ~
    WL 0.067 0.049 1.349 0.177 0.067 0.063
CI 0.529 0.051 10.400 0.000 0.529 0.485
 \begin{array}{c} \texttt{Controlled}~\sim \\ \texttt{WL} \end{array}WL 0.242 0.055 4.410 0.000 0.242 0.227
CI 0.181 0.056 3.208 0.001 0.181 0.165
 DA \sim \frac{1}{NT}WL 0.199 0.059 3.347 0.001 0.199 0.155
CI 0.488 0.069 7.115 0.000 0.488 0.369
Variances:
    DA 20.246 1.511 20.246 0.711
   Autonomous 14.805 1.105 14.805 0.768<br>Controlled 18.208 1.359 18.208 0.931
                   18.208 1.359
```


This more parsimonious model shows a good fit $(\chi^2/\text{df} = 1.82, \text{CFI} = .99, \text{TLI} = .98,$ RMSEA = 0.048 , CI 90 % [0; .058], SRMR = .02). All included paths are significant, indicating that perceived workload and choice independence predict employee's deep approaches to learning significantly. In addition, autonomous motivation was found to mediate this relationship.

Output 5 Path analysis mediating variables without nonsignificant paths

 Fig. 14.7 Path diagram full SEM model

Full SEM Model

 Finally, a *full SEM model* combines a measurement model and structural model that were presented above. As an illustration, the full SEM model of final path analysis will be examined (see Fig. 14.7).

The results show that this model does not adequately fit the data $(\chi^2/df = 2.47)$, CFI = .85, TLI = .84, RMSEA = 0.065, CI 90 % [.059; .071], SRMR = .075). The modification indices were checked to examine whether the model could be improved (Output 6).

```
lavaan (0.5-12) converged normally after 41 iterations
                                              Used Total<br>346 359
 Number of observations
  Estimator ML
Minimum Function Test Statistic 665.646
 Degrees of freedom 270<br>P-value (Chi-square) 270
 P-value (Chi-square)
Model test baseline model:
 Minimum Function Test Statistic 3011.588<br>Degraes of freedom 300
 Degrees of freedom 300<br>P-value 0.000
  P-value 0.000
Full model versus baseline model:
 Comparative Fit Index (CFI) (0.854<br>Tucker-Lewis Index (TLI) (0.838Tucker-Lewis Index (TLI)
Loglikelihood and Information Criteria:
 Loglikelihood user model (H0) -11147.345
 Loglikelihood unrestricted model (H1)
 Number of free parameters 55<br>Akaike (AIC) 22404.690
 Akaike (AIC) 22404.690<br>Bayesian (BIC) 22616.244
 Bayesian (BIC) 22616.244
 Sample-size adjusted Bayesian (BIC)
Root Mean Square Error of Approximation:
  RMSEA 0.065
 90 Percent Confidence Interval P-value RMSEA \leq 0.05P-value RMSEA \leq 0.05Standardized Root Mean Square Residual:
  SRMR \bigcirc 0.075
Parameter estimates:
 Information Expected
 Standard Errors
                 Estimate Std.err Z-value P(>|z|) Std.lv Std.all
Latent variables:
 Workload =~<br>WL1
    WL1 1.000 1.004 0.873
WL2 0.613 0.064 9.593 0.000 0.615 0.515
WL4 0.762 0.063 12.171 0.000 0.765 0.631
   WL5 0.962 0.059 16.390 0.000 0.966 0.856
 Choice.independence =\sim CI1 1.000
    CI1 1.000 0.796 0.735
   CI2 0.839 0.080 10.543 0.000 0.667 0.628
    CI3 0.897 0.079 11.363 0.000 0.714 0.680
CI4 0.804 0.072 11.200 0.000 0.640 0.669
   CI5 0.815 0.075 10.908 0.000 0.649 0.651
 Deep.approach =~
    ALWD1 1.000 0.521 0.446
ALWD2 0.559 0.132 4.243 0.000 0.292 0.292
ALWD3 0.463 0.104 4.461 0.000 0.241 0.312
ALWD4 0.350 0.107 3.260 0.001 0.182 0.213
   ALWD5 0.778 0.132 5.873 0.000 0.405 0.471
    ALWD6 1.065 0.162 6.588 0.000 0.555 0.590
ALWD7 0.773 0.134 5.785 0.000 0.403 0.459
```
 Output 6 Full SEM model

	1.160	0.181	6.417	0.000	0.605	0.557	
ALWD8							
ALWD9	0.691 1.046	0.126 0.170	5.498	0.000	0.360	0.422	
ALWD10			6.153	0.000	0.545	0.513	
Autonomous.motivation =~							
Intrins1	1,000				0.883	0.903	
Intrins2	0.970	0.042	23.369	0.000	0.857	0.916	
Intrins3	0.631	0.051	12.330	0.000	0.557	0.597	
Ident1	0.426	0.063	6.771	0.000 0.000	0.376	0.361	
Ident2	0.342	0.070	4.900		0.301	0.267	
Ident3	0.716	0.052	13.800	0.000	0.632	0.648	
Regressions:							
Deep.approach ~							
Workload	0.093	0.032	2.901	0.004	0.180	0.180	
Choic.ndpndnc	0.387	0.069	5.573	0.000	0.591	0.591	
Autonms.mtvtn	0.163	0.044	3.706	0.000	0.276	0.276	
Autonomous.motivation ~							
Choic.ndpndnc		0.579 0.071 8.208		0.000	0.522	0.522	
Covariances:							
Workload ~~							
$Choice. ndpndnc -0.211 0.053$			-3.948	0.000	-0.264	-0.264	
Variances:							
WL1	0.316	0.052			0.316	0.239	
WL2	1.048	0.084			1.048	0.735	
WL4	0.886	0.075			0.886	0.602	
WL5	0.339	0.050			0.339	0.266	
CI1	0.540	0.054			0.540	0.460	
CI ₂	0.685	0.060			0.685	0.606	
CI3	0.593	0.055			0.593	0.538	
CI4	0.505	0.046			0.505	0.552	
CI5						0.577	
	0.573	0.051			0.573		
ALWD1	1.093	0.089			1.093	0.801	
ALWD2	0.909	0.071			0.909	0.914	
ALWD3	0.540	0.042			0.540	0.903	
ALWD4	0.703	0.054			0.703	0.955	
ALWD5	0.576	0.048			0.576	0.778	
ALWD6	0.577	0.052			0.577	0.652	
ALWD7	0.608	0.050			0.608	0.789	
ALWD8	0.811	0.071			0.811	0.689	
ALWD9	0.599	0.048			0.599	0.822	
ALWD10	0.833	0.071			0.833	0.737	
Intrins1	0.176	0.025			0.176	0.184	
Intrins2	0.141	0.023			0.141	0.161	
Intrins3	0.560	0.045			0.560	0.643	
Ident1	0.944	0.073			0.944	0.869	
Ident2	1.180	0.090			1.180	0.928	
Ident3	0.552	0.045			0.552	0.580	
Workload	1.007	0.108			1,000	1.000	
Choic.ndpndnc	0.633	0.086			1,000	1.000	
Deep.approach	0.120	0.034			0.442	0.442	
Autonms.mtvtn	0.567	0.059			0.728	0.728	

Output 6 (continued)

The modification indices (Output 7) showed that the model could be improved by adding covariances between several observed variables. Only combinations of observed variables that reflected the same latent constructs were included because they can be considered in accordance with our theoretical model. For presentation purposes, only an excerpt of the output was included.

Output 7 Modification indices full SEM model

357	ALWD5	\sim \sim	Ident1	0.382	0.025	0.025	0.028	0.028
358	ALWD5 \sim		Ident2	0.805	0.041	0.041	0.042	0.042
359	$ALWD5$ ~~		Ident3		$0.013 - 0.004$	-0.004	-0.004	-0.004
360	ALWD6 ~~		ALWD6	0.000	0.000	0.000	0.000	0.000
361	$ALWD6$ ~~			ALWD7 10.839 0.121		0.121	0.147	0.147
362	$ALWD6$ ~~		ALWD8		0.05000.010	0.010	0.010	0.010
363	$ALWD6$ ~~		ALWD9		$0.667 - 0.029$	-0.029	-0.037	-0.037
364	ALWD6 ~~		ALWD10		$0.750 - 0.038$	-0.038	-0.038	-0.038
365	ALWD6 ~~		Intrins1		$0.196 - 0.010$	-0.010	-0.011	-0.011
400	$ALWD10$ ~~		Intrins2	0.357	0.015	0.015	0.015	0.015
401	$ALWD10$ ~~		Intrins3	2.201	0.058	0.058	0.059	0.059
402	$ALWD10$ ~~		Ident1	3.023	0.087	0.087	0.078	0.078
403	$ALWD10 \sim$		Ident2	2.416	0.087	0.087	0.072	0.072
404	$ALWD10$ ~~		Ident3		$1.036 - 0.040$	-0.040	-0.038	-0.038
405	Intrinsl \sim		Intrins1	0.000	0.000	0.000	0.000	0.000
406	Intrins1	\sim \sim	Intrins 22.639		0.195	0.195	0.213	0.213
407	Intrins1 \sim		Intrins3		$3.710 - 0.049$	-0.049	-0.053	-0.053
408	Intrinsl \sim		Ident1		$2.662 - 0.047$	-0.047	-0.046	-0.046
409	Intrins1 ~~		Ident2	0.006	0.003	0.003	0.002	0.002
410	Intrins1 ~~		Ident3	1.611	0.034	0.034	0.036	0.036
411	Intrins2 \sim		Intrins2	0.000	0.000	0.000	0.000	0.000
412	Intrins2 \sim		Intrins3	2.425	0.038	0.038	0.043	0.043
413	Intrins2 \sim			Ident(12.902)-0.099		-0.099	-0.102	-0.102
414	Intrins2 \sim		Idente		$10.067 - 0.096$	-0.096	-0.091	-0.091
415	Intrins2 \sim		Ident3		$6.506 - 0.066$	-0.066	-0.072	-0.072
416	Intrins3 \sim		Intrins3	0.000	0.000	0.000	0.000	0.000
417	Intrins3 ~~		Ident1	4.778	0.088	0.088	0.091	0.091
418	Tntrins3 \sim		Ident2		$1.755 - 0.060$	-0.060	-0.057	-0.057
419	Intrins3 \sim		Ident3		$0.078 - 0.009$	-0.009	-0.010	-0.010
420	Ident1 ~~		Ident1	0.000	0.000	0.000	0.000	0.000
421	Ident1 \sim			Ident ₂ 46.922	0.392	0.392	0.334	0.334
422	Ident1 $~\sim~$		Ident3	7.649	0.112	0.112	0.110	0.110
423	Ident2 \sim		Ident2	0.000	0.000	0.000	0.000	0.000
424	Ident2 \sim		Ident3		$0.000 - 0.001$	-0.001	-0.001	-0.001
425	Ident3 \sim		Ident3	0.000	0.000	0.000	0.000	0.000
426	Workload ~~		Workload	0.000	0.000	0.000	0.000	0.000
427	Workload ~~		Choice.independence	0.000	0.000	0.000	0.000	0.000
428	Workload ~~		Deep.approach	NA	NA	NA	NA	NA
429			Workload ~~ Autonomous.motivation	0.285	0.025	0.028	0.028	0.028
430	Choice.independence ~~		Choice.independence	0.000	0.000	0.000	0.000	0.000
431	Choice.independence ~~		Deep.approach	NA	NA	NA	NA	NA
432			Choice.independence ~~ Autonomous.motivation	0.285	0.075	0.107	0.107	0.107
433	Deep.approach ~~		Deep.approach	0.000	0.000	0.000	0.000	0.000
434			Deep.approach ~~ Autonomous.motivation	NA	NA	NA	NA	ΝA
	435 Autonomous.motivation ~~ Autonomous.motivation			0.000	0.000	0.000	0.000	0.000
436	Deep.approach		~ Autonomous.motivation	0.000	0.000	0.000	0.000	0.000
437	Deep.approach	\sim	Workload	0.000	0.000	0.000	0.000	0.000
438	Deep.approach	\sim	Choice.independence	0.000	0.000	0.000	0.000	0.000
439	Autonomous.motivation	\sim	Deep.approach	0.285	0.285	0.168	0.168	0.168
	440 Autonomous.motivation	\sim	Workload	0.284	0.027	0.030	0.030	0.030
441	Autonomous.motivation	\sim	Choice.independence	0.000	0.000	0.000	0.000	0.000
442	Workload	\sim	Deep.approach	0.285	0.270	0.140	0.140	0.140
443	Workload		~ Autonomous.motivation	0.285	0.044	0.039	0.039	0.039
444	Workload		Choice.independence	NA	NA	NA	NA	NA
445	Choice.independence	\sim	Deep.approach	0.284	0.811	0.531	0.531	0.531
446	Choice.independence		~ Autonomous.motivation	0.285	0.132	0.147	0.147	0.147
447	Choice.independence	\sim	Workload	NA	NA	NA	NA	NA

Output 7 (continued)

These modifications (Output 8) result in an acceptable model fit $(\chi^2/df = 1.94,$ CFI=.91, TLI=.90, RMSEA=0.052, CI 90 % [.045; .059], SRMR=.064). When reporting on SEM analysis, it is not necessary to provide all coefficients of every model that was tested. It is however important that the reader gets an overview (with fit indices) of the different models that were tested. Only the coefficients of the final model should be reported. A table containing the coefficients, standardised coefficients, criti-cal ratio and significance (level) is usually included. Table [14.5](#page-27-0) provides an example of how the results of the final full SEM model could be presented.

```
Minimum Function Test Statistic 3011.588<br>Degrees of freedom 300
 Degrees of freedom 300<br>
P-value 0.000
 P-value
Full model versus baseline model:
 Comparative Fit Index (CFI) \begin{pmatrix} 0.910 \\ 0.896 \end{pmatrix}Tucker-Lewis Index (TLI)
Loglikelihood and Information Criteria:
 Loglikelihood user model (H0) -11067.206
 Loglikelihood unrestricted model (H1) -10814.522
 Number of free parameters 65<br>Akaike (AIC) 22264.411
 Akaike (AIC) 22264.411<br>Bayesian (BIC) 2254.430
 Bayesian (BIC) 22514.430<br>Sample-size adjusted Bayesian (BIC) 22308.231
 Sample-size adjusted Bayesian (BIC)
Root Mean Square Error of Approximation:
  RMSEA 0.052
 90 Percent Confidence Interval 0.045 0.059<br>P-value RMSEA <= 0.05P-value RMSEA \leq 0.05Standardized Root Mean Square Residual:
  SRMR 0.064
Parameter estimates:
 Information Expected
 Standard Errors
               Estimate Std.err Z-value P(>|z|) Std.lv Std.all
Latent variables:
 Workload =~
   WL1 1.000 1.011 0.879
   WL2 0.583 0.064 9.040 0.000 0.589 0.493
   WL4 0.738 0.063 11.694 0.000 0.746 0.615
   WL5 0.957 0.061 15.706 0.000 0.968 0.858
 Choice.independence = \sim CI1 1.000
   CI1 1.000 0.790 0.729
   CI2 0.854 0.080 10.623 0.000 0.674 0.634
   CI3 0.904 0.080 11.333 0.000 0.714 0.680<br>CI4 0.810 0.073 11.171 0.000 0.640 0.669
   CI4 0.810 0.073 11.171 0.000 0.640 0.669<br>CI5 0.821 0.075 10.883 0.000 0.649 0.651
                               10.883
lavaan (0.5-12) converged normally after 46 iterations
                                          Used Total<br>346 359
 Number of observations 346
  Estimator ML
  Minimum Function Test Statistic \left(505.367\right)Degrees of freedom 260<br>P-value (Chi-square) 260<br>0.000
 P-value (Chi-square)
Model test baseline model:
```
Output 8 Full SEM model after modification

Output 8 (continued)

Output 8 (continued)

	Regression		Standardised	Critical
	weight	Standard error	regression weight	ratio ^a
Measurement model				
Workload = \sim				
WL1	$\mathbf{1}$	$_{\rm b}$.88	$_{\rm b}$
WL ₂	.58	.06	.49	9.04
WL4	.74	.06	.62	11.69
WL5	.96	.06	.86	15.71
Choice independence $=$ \sim				
CI1	1	$\rm b$.73	$_{\rm b}$
CI2	.85	.08	.63	10.62
CI3	.90	.08	.68	11.33
CI ₄	.81	.07	.67	11.17
CI ₅	.82	.08	.65	10.88
Deep approach = \sim				
ALWD1	$\mathbf{1}$	$_{\rm b}$.46	$_{\rm b}$
ALWD ₂	.57	.13	.30	4.39
ALWD3	.38	.10	.26	3.90
ALWD4	.27	.10	.17	2.67
ALWD5	.79	.13	.49	6.09
ALWD6	.98	.15	.55	6.50
ALWD7	.63	.12	.38	5.18
ALWD8	1.16	.18	.57	6.58
ALWD9	.61	.12	.38	5.20
ALWD10	1.02	.16	.51	6.23
Autonomous = \sim				
Intrins1	1	$_{\rm b}$.79	b
Intrins2	1.08	.06	.89	19.70

Table 14.5 Coefficients final full SEM model

(continued)

Note: Estimation Method: Maximum Likelihood

^aAll critical ratios except b: $p < .001$

^bValue fixed at 1.00 for model identification purpose; hence, no standard error was computed Critical ratio: *p* < .01

14.4 Extensions of SEM and Their Application in Research on Professional Learning

Confirmatory factor analysis, path analysis and structural equation modelling have been commonly applied by researchers in various fields, including the field on professional learning. 'Basic' SEM models already offer a variety of possibilities because different types of relationships can be modelled based on the theoretical foundations of the study. In addition various extensions or specific forms of SEM offer a wide range of possibilities. Within this section, these extensions will be presented accompanied by examples of how these analyses can be applied within the field of professional learning.

Within the section on confirmatory factor analysis, we discussed the measurement invariance across groups. If this measurement invariance across groups is established, you could also investigate whether the SEM model is equal or different for different groups of participants. Establishing measurement invariance is important in this regard, as you would want to make sure that the differences that you find in the model reflect differences in the true constructs and relationships and not mere differences in measurement. With *multiple-group SEM* , it is possible to establish whether models are different for different groups of participants. The study of Hurtz and Williams (2009) , for example, applied a multiple-group path analysis. This study examined attitudinal and motivational antecedents of employees' participation in development activities based on data collected within four organisations. A multiple-group path analysis was used to investigate whether the model differed for the four different organisations that were included in their study. They concluded that although the strengths of the different paths differed across organisations, the patterns were the same as the full SEM model that was based on the data from all four organisations (Hurtz & Williams, 2009). Within their study, they also examined whether their measurement was invariant across the groups and concluded that this was not the case. However, because the focus lied on identifying a general model looking at the relationship of the antecedents with participation and not on explaining organisational differences, Hurtz and Williams [\(2009](#page-35-0)) applied group-mean centring to remove organisational mean differences: they computed the difference between the employee score and the organisational mean of the scores and took this difference as their primary outcome variable. After applying this group-mean centring, measurement invariance was reached. However, it must be noted that no conclusions can be drawn regarding possible organisational differences. Alternatively to this approach, on the condition that metric invariance is achieved and the sample size is large enough, one could also adopt a full SEM model, sometimes also called multiple-indicator SEM model, because this allows the modelling of intercept variances. Multiple-group SEM analysis can also be applied to compare the models of males and females, high- and low-qualified employees, etc. In contrast, to multilevel SEM analysis (see below), multiple-group analysis does not require that the different groups are sampled at random. Multiple- group SEM can be conducted with the lavaan package (Rosseel, 2012).

Multilevel SEM requires a random sampling of groups because it assumes that the differences between organisations in terms of the intercept and slope are normally distributed around the average intercept or slope that holds for the population. For a basic introduction into multilevel analysis within professional learning, the reader is referred to Kyndt and Onghena (2014). In short, multilevel SEM combines multilevel analysis and SEM analysis. SEM analysis is not able to take the nested structure of the data into account (i.e. employees nested within organisations), while multilevel analysis is not able to examine more complex models. Multilevel analysis is comparable to regression analysis with regard to the type of relationships they investigate. In addition, multilevel analysis does not provide goodness-of-fi t indices such as CFI, SRMR or RMSEA. One can only conclude that one model fits the data better in comparison with another model (Kyndt & Onghena, [2014](#page-35-0)). Within a

 multilevel SEM, the model is estimated while the organisational clustering is taken into account. Taking the organisational clustering into account is important because 'if the nested structure of the data is ignored, it is more likely that statistical relations are observed in the sample that are in fact not true (type-1 error), in addition it might be that it is concluded that a relationships holds for individuals when they are actually true for groups (ecological fallacy)' (Kyndt & Onghena, 2014, p. 339). Moreover, within multilevel SEM, predictors on the level of the organisation can be combined with predictors at the individual level. A final difference between multilevel SEM and multilevel analysis is that multilevel SEM can also predict outcomes at the organisational level, whereas in traditional multilevel analysis, the outcomes or dependent variables need to be situated at the lowest level, that is, the individual level. Sometimes this latter issue is resolved by aggregating the individual scores to the organisational level (after the within-group agreement has been checked); however, by doing this a lot of statistical power is lost as well as potentially interesting individual differences within organisations. To our knowledge it is not possible to conduct multilevel SEM with the R software, it is possible with Mplus (Muthén $\&$ Muthén, [1998–2010](#page-36-0)). More information on multilevel SEM can be found in the article of Kaplan and Elliott ([1997 \)](#page-35-0). Within educational sciences and labour psychology, different examples of empirical studies using multilevel SEM can be found (e.g. Johnsrud, Heck, & Rosser, 2000; Mauno, Kiuru, & Kinnunen, 2011; Sebastian & Allensworth, 2012). However, we were not able to detect a specific example within the field of professional learning.

 The two above-presented extensions of SEM have something in common that they add to the complexity of the model; as a consequence, both techniques usually require very large samples (many organisations and many employees within the organisations). However, *partial least squares* (*PLS*) might be an interesting alternative to SEM if the sample size requirements cannot be met. Where SEM combines factor analysis and path models, PLS combines principal component analysis and path models (Garthwaite, [1994](#page-34-0); Goutis, [1996](#page-34-0); Hoyle, 1999; Iacobucci, 2010). Principal component analysis does not aim at reflecting latent constructs; rather, it tries to predict the component as good as possible. Factor analysis most commonly uses the maximum likelihood estimation, which considers the variance the different observed variables have in common, while principal component analysis takes all variance into account when estimating the component (Garthwaite, 1994; Hoyle, 1999). Because the focus lies on maximising the prediction and capturing as much variance as possible from the dependent variable, PLS is better suitable for exploratory rather than confirmatory purposes (Hoyle, [1999](#page-35-0); Iacobucci, [2010](#page-35-0)). The loadings tend to be overestimated and path coefficients underestimated (Goutis, 1996; Iacobucci, [2010](#page-35-0)). In addition, no goodness-of-fi t indices are provided. Similarly to multilevel analysis, it can only be judged which model is more suitable in comparison with another model. However, it is an interesting approach when you want to explore complex models with a limited number of observations. Gegenfurtner (2013) illustrates the use of PLS in his study on the relationship between motivation to transfer, retention, transfer and attitudes.

14.5 Discussion

 Throughout this chapter we have tried to introduce the reader to SEM by means of an illustration within the field of professional learning. The models that are tested within this chapter are solely for the purpose of illustrating the method and should not be used for interpreting the relationships between the constructs. These results and interpretations can be found in the publication of the empirical study in which the relationships including control variables were examined (Kyndt et al., 2013). SEM offers researchers a lot of possibilities to investigate complex models. Within the paradigm of the integration of work and learning, advanced techniques such as multiple-group and multilevel SEM might be especially relevant, because these techniques allow the simultaneous examination of individual and organisational differences. These techniques are interesting if the goal is to investigate professional learning conceptualised as a reciprocal interaction between individual and organisation (Tynjälä, [2008 \)](#page-36-0).

 However, despite all the possibilities SEM offers, it also has its limitations and possible pitfalls. The two most evident limitations are the necessity for a strong theoretical basis and the large samples that are needed especially when comprehensive models with many variables need to be estimated. Although simple models can be estimated with moderate to small samples, the added value of SEM foremost lies in estimating complex models (Hox $&$ Bechger, 1998).

 One of the most common pitfalls of SEM is that many researchers are tempted to interpret SEM models as causal models due to the impression the direction of the paths give. However, merely applying SEM to your data does not provide proof for the causality of the relationship (Bollen & Pearl, [2013](#page-34-0)). A SEM model can however raise doubts about a causal theory; when the SEM model is correctly specified and the covariance structure does not support the theoretical causal structure, it seems less plausible that the causal relationships exists. One can say that establishing covariance or correlations between the variables is a necessary but not sufficient condition for establishing causality. To be able to prove the causality of the relationships empirically, data that allow this type of conclusions need to be collected (e.g. longitudinal or experimental data). If SEM is applied to correlational data, the SEM model cannot be interpreted as a causal model (Iacobucci, [2009](#page-35-0)).

 The aim of the current book chapter was to introduce the reader with the possibilities that SEM can offer within the field of research on professional learning. Within the paradigm of the integration of work and learning, the possibility to analyse more complex models can contribute to the theory development and further understanding of how learning and working are intertwined.

Appendix: R Code Illustration

```
## Setting up working directory and loading data
setwd("/Users/evakyndt/Book chapter SEM")
data<-read.table("chapter SEM.csv", header=TRUE, sep=";")
```

```
## Loading packages 
install.packages("lavaan")
library("lavaan")
install.packages("qgraph")
library("qgraph")
install.packages("semTools")
library("semTools")
## Measurement model: Confirmatory factor analysis (Output 1) 
CFAModel1 <- 'Workload=~WL1+WL2+WL3+WL4+WL5
Choice.independence=~CI1+CI2+CI3+CI4+CI5'
Fit1 <- cfa(CFAModel1, data=data)
summary(Fit1, fit.measures=TRUE, standardized=TRUE)
modindices(Fit1)
## CFA without WL3 (Output 2) 
CFAModel2 <- 'Workload=~WL1+WL2+WL4+WL5
Choice.independence=~CI1+CI2+CI3+CI4+CI5'
Fit2 <- cfa(CFAModel2, data=data)
summary(Fit2, fit.measures=TRUE, standardized=TRUE)
## Plot CFA model (Figure 4) 
qgraph.lavaan(Fit1,layout="tree", vsize.man=5, vsize.
  lat=12, include=4, curve=-0.4, edge.label.cex=0.6, 
  titles=F)
## Testing measurement invariance across groups (Output 3)
measurementInvariance(CFAModel1, data=data, group="Sex")
## Path analysis mediation (Output 4) 
SEMModel2 <- 'DA ~ Autonomous + Controlled
Autonomous \sim WL + CI
Controlled ~ WL + CI
DA \sim WL + CI'FitSEM2 <- sem(SEMModel2, data=data)
summary(FitSEM2, fit.measures=TRUE, standardized=TRUE)
## Path analysis mediation without non-significant paths 
(Output 5) 
SEMModel3 <- 'DA~Autonomous
Autonomous \sim CI
DA \sim WL + CI'FitSEM3 <- sem(SEMModel3, data=data)
summary(FitSEM3, fit.measures=TRUE, standardized=TRUE)
## Full SEM model (Output 6) 
FullModel <- 'Workload=~WL1+WL2+WL4+WL5
```

```
Choice.independence=~CI1+CI2+CI3+CI4+CI5
Deep.approach=~ALWD1+ALWD2+ALWD3+ALWD4+ALWD5+ALWD6+ALWD
  7+ALWD8+ALWD9+ALWD10
Autonomous.motivation=~ Intrins1+Intrins2+Intrins3+Iden
  t1+Ident2+Ident3
Deep.approach~Workload+Choice.independence
Deep.approach~Autonomous.motivation
Autonomous.motivation~Choice.independence'
FitSEMFull <- sem(FullModel, data=data)
summary(FitSEMFull, fit.measures=TRUE, standardized=TRUE)
## Modification indices (Output 7) 
modindices(FitSEMFull) 
## Full SEM model after modification (Output 8) 
FullModel2 <- 'Workload=~WL1+WL2+WL4+WL5
Choice.independence=~CI1+CI2+CI3+CI4+CI5
Deep.approach=~ALWD1+ALWD2+ALWD3+ALWD4+ALWD5+ALWD6+ALWD
7+ALWD8+ALWD9+ALWD10
Autonomous.motivation=~ Intrins1+Intrins2+Intrins3+Iden
  t1+Ident2+Ident3
Deep.approach~Workload+Choice.independence
Deep.approach~Autonomous.motivation
Autonomous.motivation~Choice.independence
ALWD3~~ALWD9
ALWD6~~ALWD7
ALWD4~~ALWD7
WL2 \sim WL4Intrins1~~Intrins2
Ident1~~Ident2
Intrins2~~Ident1
Intrins2~~Ident2
Intrins1~~Ident1
Intrins1~~Ident3'
FitSEMFull2 <- sem(FullModel2, data=data)
summary(FitSEMFull2, fit.measures=TRUE, standardized=TRUE)
## references
citation("lavaan")
citation("qgraph")
citation("semTools")
```
 References

- Allison, P. D. (2003). Missing data techniques for structural equation modeling. *Journal of Abnormal Psychology, 112* , 545–557.
- Arbuckle, J. L. (2011). *IBM SPSS AMOS 20 user's guide* . Armonk, NY: IBM Corporation.
- Baron, R. M., & Kenny, P. A. (1986). The moderator-mediator variable distinction in social psychology research: Conceptual, strategic and statistical considerations. *Journal of Personality and Social Psychology, 51* , 1173–1182.
- Bentler, P. M. (1992). On the fi t of models to covariances. *Psychological Bulletin, 88* , 588–606.
- Bentler, P. M. (2004). *Eqs 6 structural equations program book* . Encino, CA: Multivariate Software, Inc.
- Bentler, P. M., & Chou, C. (1987). Practical issues in structural modeling. *Sociological Methods & Research, 16* , 78–117.
- Bentler, P. M., & Yuan, K.-H. (1999). Structural equation modelling with small samples: Test statistics. *Multivariate Behavioral Research, 34* , 181–197.
- Billett, S. (2001). Learning through work: Workplace affordances and individual engagement. *Journal of Workplace Learning, 13* , 209–214.
- Boeve-de Pauw, J., Jacobs, K., & Van Petegem, P. (2012). Gender differences in environmental values: An issue of measurement? *Environment and Behaviour* . doi[:10.1177/0013916512460761](http://dx.doi.org/10.1177/0013916512460761).
- Bollen, K. A., & Pearl, J. (2013). Eight myths about causality and structural equation models. In S. L. Morgan (Ed.), *Handbook of causal analysis for social research* (pp. 301–328). Dordrecht, The Netherlands: Springer.
- Browne, M. W., & Cudeck, R. (1993). Alternative ways of assessing model fit. In K. A. Bollen $\&$ J. S. Long (Eds.), *Testing structural equation models* (pp. 136–162). Beverly Hills, CA: Sage.
- Byrne, B. M. (2001). *Structural equation modeling with Amos: Basic concepts, applications, and programming* . Mahwah, NJ: Lawrence Erlbaum.
- Chin, W. W. (1998). Issues and opinion on structural equation modeling. *MIS Quarterly, 22* , vii–xvi.
- Coertjens, L., Donche, V., De Maeyer, S., Vanthournout, G., & Van Petegem, P. (2012). Longitudinal measurement invariance of Likert-type learning strategy scales: Are we using the same ruler at each wave? Journal of Psychoeducational Assessment, 30, 577-587.
- Coertjens, L., Donche, V., De Maeyer, S., Vanthournout, G., & Van Petegem, P. (2013). Modeling change in learning strategies throughout higher education: A multi-indicator latent growth perspective. *PLoS One*, 8(7), e67854.
- Cohen, L., Manion, L., & Morrison, K. (2011). *Research methods in education* . Abingdon, UK: Routledge.
- Crockett, S. A. (2012). A five-step guide to conducting SEM analysis in counseling research. *Counseling Outcome Research and Evaluation, 3* , 30–47.
- Epskamp, S., Cramer, A. O. J., Waldorp, L. J., Schmittmann, V. D., & Borsboom, D. (2012). qgraph: Network visualizations of relationships in psychometric data. *Journal of Statistical Software, 48* , 1–18.
- Gagné, M., Forest, J., Gilbert, M. H., Aubé, C., Morin, E., & Malorni, A. (2010). The motivation at work scale: Validation evidence in two languages. *Educational and Psychological Measurement, 70* , 628–646.
- Garthwaite, P. H. (1994). An interpretation of partial least squares. *Journal of the American Statistical Association, 89* , 122–127.
- Gegenfurtner, A. (2013). Dimensions of motivation to transfer: A longitudinal analysis of their influences on retention, transfer, and attitude change. *Vocations and Learning*, 6, 187–205.
- Goutis, C. (1996). Partial least squares algorithm yields shrinkage estimators. *The Annals of Statistics, 24* , 816–824.
- Hair, J., Black, W., Babin, B., Anderson, R., & Tatham, R. (2006). *Multivariate data analysis* (6th ed.). Upper Saddle River, NJ: Pearson Prentice Hall.
- Hox, J. J., & Bechger, T. M. (1998). An introduction to structural equation modeling. *Family Science Review, 11* , 354–373.
- Hoyle, R. H. (1999). *Statistical strategies for small sample research* . Thousand Oaks, CA: Sage.
- Hsu, S. H., Chen, W. H., & Hsieh, M. J. (2006). Robustness testing of PLS, LISREL, EQS and ANN-based SEM for measuring customer satisfaction. *Total Quality Management & Business Excellence, 17* , 355–371.
- Hu, L., & Bentler, P. M. (1999). Cutoff criteria for fit indexes in covariance structure analysis: Conventional criteria versus new alternatives. *Structural Equation Modeling, 6* , 1–55.
- Hu, L., Bentler, P. M., & Kano, Y. (1992). Can test statistics in covariance structure analysis be trusted? *Psychological Bulletin, 112, 351-362.*
- Hurtz, G. M., & Williams, K. J. (2009). Attitudinal and motivational antecedents of participation in voluntary employee development activities. *Journal of Applied Psychology, 94* , 635–653.
- Iacobucci, D. (2009). Everything you always wanted to know about SEM (structural equations modeling) but were afraid to ask. *Journal of Consumer Psychology*, 19, 673–680.
- Iacobucci, D. (2010). Structural equations modeling: Fit indices, sample size, and advanced topics. *Journal of Consumer Psychology, 20, 90-98.*
- Jaccard, J., & Wan, C. K. (1996). *LISREL approaches to interaction effects in multiple regression* . Thousand Oaks, CA: Sage.
- Johnsrud, L. K., Heck, R. A., & Rosser, V. J. (2000). Morale matters: Midlevel administrators and their intention to leave. *The Journal of Higher Education, 71* , 34–59.
- Jöreskog, K. G., & Sörbom, D. (1997). *LISREL 8: User's reference guide*. Chicago: Scientific Software International.
- Jung, S. (2013). Structural equation modeling with small sample sizes using two-stage ridge leastsquares estimation. *Behavior Research Methods*, 45, 75–81.
- Kaplan, D., & Elliott, P. R. (1997). A didactic example of multilevel structural equation modeling applicable to the study of organizations. *Structural Equation Modeling: A Multidisciplinary Journal, 4* , 1–24.
- Kirby, J. R., Knapper, C. K., Evans, C. J., Carty, A. E., & Gadula, C. (2003). Approaches to learning at work and workplace climate. *International Journal of Training and Development, 7* , 31–52.
- Kline, R. B. (2012). Assumptions in structural equation modeling. In R. Hoyle (Ed.), *Handbook of structural equation modeling* (pp. 111–125). New York: Guilford Press.
- Kyndt, E., & Baert, H. (2013). Antecedents of employees' involvement in work-related learning: A systematic review. *Review of Educational Research*, 83(2), 273-313. doi[:10.3102/0034654313478021](http://dx.doi.org/10.3102/0034654313478021).
- Kyndt, E., & Onghena, P. (2014). Hierarchical linear models for researching professional learning: Relevance and implications. In S. Billett, C. Harteis, & H. Gruber (Eds.), *International handbook of research in professional practice-based learning* (pp. 337–370). Dordrecht, The Netherlands: Springer.
- Kyndt, E., Raes, E., Dochy, F., & Janssens, E. (2013). Approaches to learning at work: Investigating work motivation, workload and choice independence. *Journal of Career Development, 40* , 271–291.
- Kyndt, E., Janssens, I., Coertjens, L., Gijbels, D., Donche, V., & Van Petegem, P. Vocational education students' working life competencies: A self-assessment instrument. *Vocations and Learning* (accepted).
- Lohman, M. C. (2000). Environmental inhibitors to informal learning in the workplace: A case study of public school teachers. Adult Education Quarterly, 50, 83-101.
- MacCallum, R. C., Browne, M. W., & Sugawara, H. M. (1996). Power analysis and determination of sample size for covariance structure modeling. *Psychological Methods, 1* , 130–149.
- Maruyama, G. (1998). *Basics of structural equation modeling* . Thousand Oaks, CA: Sage.
- Mauno, S., Kiuru, N., & Kinnunen, U. (2011). Relationships between work-family culture and work attitudes at both the individual level and the department level. *Work & Stress: An International Journal of Work, Health, and Organisations, 25* , 147–166.
- Muthén, L. K., & Muthén, B. O. (1998–2010). *Mplus user's guide* (6th ed.). Los Angeles: Muthén & Muthén.
- Pornprasertmanit, S., Miller, P., Schoemann, A., & Rosseel, Y. (2013). *semTools: Useful tools for structural equation modeling. R package version 0.3-0.* Retrieved July 17, 2013, from [http://](http://cran.r-project.org/package=semTools) [CRAN.R-project.org/package=semTools](http://cran.r-project.org/package=semTools)
- R Development Core Team. (2012). *R: A language and environment for statistical computing.* Vienna, Austria: R Foundation for Statistical Computing. Retrieved July 19, 2012, from [http://](http://www.r-project.org/) [www.R-project.org/](http://www.r-project.org/)
- Rosseel, Y. (2012). Lavaan: An R package for structural equation modeling. *Journal of Statistical Software, 48* , 1–36.
- SAS Institute Inc. (2008). *SAS/STAT 9.2 user's guide* . Cary, NC: Author.
- Satorra, A. (1990). Robustness issues in structural equation modelling: A review of recent developments. *Quality & Quantity, 24* , 367–386.
- Sebastian, J., & Allensworth, E. (2012). The influence of principal leadership on classroom instruction and student learning: A study of mediated pathways to learning. *Educational Administration Quarterly, 48* , 626–663.
- Tacq, J. (1997). *Multivariate analysis techniques in social science research: From problem to analysis* . London: Sage.
- Tynjälä, P. (2008). Perspectives into learning at the workplace. *Educational Research Review, 3* , 130–154.
- Yuan, K.-H., & Bentler, P. M. (1999). On asymptotic distributions of normal theory MLE in covariance structure analysis under some nonnormal distributions. *Statistics & Probability Letters, 42* , 107–113.
- Yuan, K.-H., & Zhong, X. (2013). Robustness of fit indices to outliers and leverage observations in structural equation modeling. *Psychological Methods, 18* , 121–136.