



Chapter 2

Measuring Instruments and Their Properties

2.1 Types of Measuring Instruments

Measuring instruments are the technical objects that are specially developed for the purpose of measuring specific quantities. A general property of measuring instruments is that their accuracy is known. Measuring instruments are divided into material measures, measuring transducers, indicating instruments, recording instruments, and measuring systems.

A *material measure* is a measuring instrument that reproduces one or more known values of a given quantity. Examples of measures are balance weights, measuring resistors, measuring capacitors, and reference materials. Single-valued measures, multiple-valued measures, and collections of measures are distinguished. Examples of multiple-valued measures are graduated rulers, measuring tapes, resistance boxes, and so on. Multiple-valued measures are further divided into those that reproduce discrete values of the corresponding quantities, such as resistance boxes, and those that continuously reproduce quantities in some range, for example, a measuring capacitor with variable capacitance. Continuous measures are usually less accurate than discrete measures.

When measures are used to perform measurements, the measurands are compared with the known quantities reproduced by the measures. The comparison is made by different methods, but so-called *comparators* are a specific means that are used to compare quantities. A comparator is a measuring device that makes it possible to compare similar quantities and has a known sensitivity. The simplest comparator is the standard equal-armed pan balance.

In some cases, quantities are compared without comparators, by experimenters, with the help of their viewing or listening perceptions. For instance, when measuring the length of a body with the help of a ruler, the ruler is placed on the body and

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the observer fixes visually the graduations of the ruler (or fractions of a graduation) at the corresponding points of the body.

A *measuring transducer* is a measuring instrument that converts the measurement signals into a form suitable for transmission, processing, or storage. The measurement information at the output of a measuring transducer typically cannot be directly observed by the experimenter.

One must distinguish measuring transducers and the transforming elements of a complicated instrument. The former are measuring instruments, and as such, they have rated (i.e., listed in documentation) metrological properties (see below). The latter, on the other hand, do not have an independent metrological significance and cannot be used separately from the instrument of which they are a part.

Measuring transducers are diverse. Thermocouples, resistance thermometers, measuring shunts, and the measuring electrodes of pH meters are just a few examples of measuring transducers. Measuring current or voltage transformers and measuring amplifiers are also measuring transducers. This group of transducers is characterized by the fact that the signals at their inputs and outputs are a quantity of the same kind, and only the magnitude of the quantity changes. For this reason, these measuring transducers are called *scaling measuring transducers*.

Measuring transducers that convert an analog signal at the input into a discrete signal at the output are called analog-to-digital converters. Such converters are manufactured either as autonomous, i.e., independent measuring instruments, or as units built into other instruments, in particular, in the form of integrated microcircuits. Analog-to-digital converters are a necessary component of a variety of digital devices, but they are also employed in monitoring, regulating, and control systems.

An *indicating instrument* is a measuring instrument that is used to convert measurement signals into a form that can be directly perceived by the observer. Based on the design of the input circuits, indicating instruments are just as diverse as measuring transducers, and it is difficult to survey all of them. Moreover, such a review and even classification are more important for designing instruments than for describing their general properties.

A common feature of all indicating instruments is that they all have readout devices. If these devices are implemented in the form of a scale and an indicating needle, then the indications of the instrument are a continuous function of the magnitude of the measurable quantity. Such instruments are called analog instruments. If the indications of instruments are in a digital form, then such instruments are called digital instruments.

The above definition of digital instruments formally includes two types of devices. The first type, which includes automatic digital voltmeters, bridges, and similar instruments, performs all measuring transformations in a discrete form; in the second type, exemplified by induction meters for measuring electrical energy, all measuring transformations of signals occur in an analog form and only the output signal assumes a discrete form. The conversions of measurement information into a discrete form have several specific features. Therefore, only instruments in which the measurement conversions occur in a discrete form are usually considered to be digital instruments.

The indications of digital instruments can be easily recorded and are convenient for entering into a computer. In addition, their design usually makes it possible to obtain significantly higher accuracy than the accuracy of analog instruments. Moreover, when digital instruments are employed, no reading errors occur. However, with analog instruments, it is easier to judge trends in the variation of the measurands.

In addition to analog and digital instruments, there also exist analog-discrete measuring instruments. In these instruments, the measuring conversions are performed in an analog form, but the readout means are discrete (but not digital). Analog-discrete instruments combine the advantages of both analog and digital instruments. Mentioned above induction meters for measuring electric energy are examples of such hybrid instruments.

In many cases, measuring instruments are designed to record their indications. Such instruments are called *recording instruments*. Data can be recorded in the form of a continuous record of the variation of the measurand in time, or in the form of a series of discrete points. Instruments of the first type are called automatic-plotting instruments, and instruments of the second type are called printing instruments. Printing instruments can record the values of a measurand in digital form. Printing instruments give a discrete series of values of the measurand with some time interval. The continuous record provided by automatic-plotting instruments can be regarded as an infinite series of values of the measurand.

Sometimes measuring instruments are equipped with induction, photo-optical, or contact devices and relays for purposes of control or regulation. Such instruments are called regulating instruments. Regulating units typically lead to some reduction of the accuracy of the measuring instrument.

Measuring instruments also customarily include null indicators, whose primary purpose is to detect the presence of a nonzero signal. The reason for them to be considered measuring instruments is that a null indicator, such as a galvanometer, can often be used as a highly sensitive indicating instrument.

A *measuring system* is a collection of functionally integrated measuring, computing, and auxiliary devices connected to each other with communication channels.

2.2 Metrological Characteristics of Measuring Instruments

We shall divide all characteristics of measuring instruments into two groups: metrological, which are significant for using a measuring instrument in the manner intended, and secondary. We shall include in the latter such characteristics as mass, dimensions, and degree of protection from moisture and dust. We shall not discuss secondary characteristics because they are not directly related with the measurement accuracy, even though they sometimes influence the selection and application of an instrument.

By metrological characteristics of a measuring instrument, we mean the characteristics that make it possible to judge the suitability of the instrument for performing measurements in a known range with known accuracy. A simple

example of a metrological characteristic common to all measuring instruments except single measures (i.e., measures reproducing a single value of a quantity) is the measurement range of the instrument. We will call metrological characteristics that are established before or during the design and development of the instrument as *nominal metrological characteristics*. Examples of such a characteristic are the nominal value of a measure (10 Ω , 1 kG, etc.), the measurement range of an instrument (0–300 V, 0–1,200 °C, etc.), the conversion range of a transducer, the value of the scale factor of an instrument scale, and so on.

The relation between the input and the output signals of indicating instruments and transducers is determined by the transfer function. For indicating instruments, this relation is determined by the instrument scale, whereas for measuring transducers, it is determined by a graph or an equation. If this graph or equation had been determined and specified before the transducer was developed (or during its development), then the graph or equation represents a nominal metrological characteristic.

The real characteristics of measuring instruments differ from the nominal characteristics because of fabrication inaccuracies and changes occurring in the corresponding properties in time. These differences between nominal and real metrological characteristics lead to the error of the instrument.

Ideally, a measuring instrument would react only to the measured quantity or to the parameter of the input signal of interest, and its indication would not depend on the external conditions, such as the power supply regime, temperature, and so on. In reality, the external conditions do affect the indications of the instrument. The quantities characterizing the external conditions affecting the indications of a measuring instrument are called *influence quantities*.

For some types of measuring instruments, the dependence of the output signal or the indications on a given influence quantity can be represented as a functional dependence, called the *influence function*. The influence function can be expressed in the form of an equation (e.g., the temperature dependence of the EMF of standard cells) or a graph. In the case of a linear dependence, it is sufficient to give the coefficient of proportionality between the output quantity and the influence quantity. We call this coefficient the *influence coefficient*. Influence coefficients and functions make it possible to take into account the conditions under which measuring instruments are used, by introducing the corresponding corrections to the obtained results.

The imperfection of measuring instruments is also manifested because when the same quantity is measured repeatedly under identical conditions, the results can differ somewhat from one another. If these differences are significant, the indications are said to be nonrepeatable.

The inaccuracy of a measuring instrument is usually characterized by its error. Taking an indicating instrument as an example, let the true value of a quantity at the input of the instrument be A_t and the instrument indication be the value A_r . The absolute error of the instrument will be

$$\zeta = A_r - A_t.$$

If the indications of the repeated measurements of A_i are somewhat different, (but not enough to be considered nonrepeatable), one can talk about a random component of instrument error. For analog instruments, the random component of instrument error is normally caused by friction in the supports of a movable part of the instrument and/or by hysteresis phenomena. The limits of this error component can be found directly if the quantity measured by the instrument can be varied continuously, which is the case with, e.g., the electric current or voltage. The common method involves driving the indicator of the instrument continuously up to the same scale marker, once from below and once from above the marker. To compensate for friction (and/or hysteresis), the input signal that drives the indicator to the marker from below needs to be higher than what it would have been without friction; the input signal that drives the indicator to the same marker from above will be smaller. We will call the dead band the absolute value of the difference between the two values of the measurand that are obtained in such a test corresponding to a given scale marker of the instrument. The dead band gives the range of possible values of the random component of instrument error, and one half of this length is the limiting value of the random error.

There are also several instrument types, notably, weighing scales, whose indications cannot vary continuously. The random error of weighing scales is usually characterized by the standard deviation [7]. This characteristic of an instrument is calculated from the changes produced in the indications of the scales by a load with a known mass; the test is performed at several scale markers, including the limits of the measurement range. One method for performing the tests and the computational formula for calculating the standard deviation of weighing scales are presented in [7].

Measuring instruments are created to bring certainty into the phenomena studied and to establish regular relations between the phenomena. Thus, the uncertainty created by the nonrepeatability of instrument indications interferes with using an instrument in the manner intended. For this reason, the first problem that must be solved when developing a new measuring device is to make its random error insignificant, i.e., either negligibly small compared with other errors or falling within permissible limits of error for measuring devices of the given type. We should note here that because uncertainty of instrument indications represents only a random component of its inaccuracy, the term “uncertainty” cannot replace the term “limits of error” as applied to measuring instruments.

If the random error is insignificant and the elements determining instrument accuracy are stable, then by calibration, the measuring device can always be “tied” to a corresponding measurement standard and the potential accuracy of the instrument can be realized.

The value of the measurand corresponding to the interval between two neighboring markers on the instrument scale is called the *value of a scale division*. Similarly, the *value of the least significant digit* is the value of the measurand corresponding to one increment of the least significant digit of a digital readout device.

The *sensitivity* of a measuring instrument is the ratio of the change in the output value of the measuring instrument to the corresponding change in the input value of the quantity that causes the output value to change. The sensitivity can be a nominal metrological characteristic or an actual characteristic of a real instrument.

The *discrimination threshold* is the minimum change in the input signal that causes an appreciable change in the output signal.

The *resolution* is the smallest interval between two distinguishable neighboring discrete values of the output signal.

Instability (of a measuring instrument) is a general term that expresses the change in any property of the measuring instrument in time.

Drift is the change occurring in the output signal (always in the same direction) in the absence of the input signal over a period of time that is significantly longer than the time needed to perform a measurement with a given measuring instrument. The presence of drift entails the need to reset the zero indication of the instrument.

The drift and the instability do not depend on the input signal, but they can depend on the external conditions. The drift is usually determined in the absence of the signal at the input.

The metrological characteristics of measuring instruments should also include their dynamic characteristics. These characteristics reflect the inertial properties of measuring instruments. It is necessary to know them to correctly choose and use many types of measuring instruments. The dynamical characteristics are examined below in Sect. 2.5.

The properties of measuring instruments can normally be described based on the characteristics enumerated above. For specific types of measuring instruments, however, additional characteristics are often required. Thus, for the gauge rods, the flatness and degree of polish are important. For voltmeters, the input resistance is important. We shall not study such characteristics, because they refer only to individual types of measuring instruments.

2.3 Rating of the Errors of Measuring Instruments

Measuring instruments can only be used as intended when their metrological properties are known. In principle, the metrological properties can be established by two methods. One method is to find the actual characteristics of a specific instrument. In the second method, the nominal metrological characteristics and the permissible deviations of the real characteristics from the nominal characteristics are given.

The first method is laborious, and for this reason, it is used primarily for the most accurate and stable measuring instruments. Thus, the second method is the main method. The nominal characteristics and the permissible deviations from them are given in the technical documentation when measuring instruments are designed, which predetermines the properties of measuring instruments and ensures that they are interchangeable.

In the process of using measuring instruments, their real properties are checked to determine whether these properties deviate from the established nominal characteristics. If some real property deviates from its nominal value by an amount more than allowed, then the measuring instrument is adjusted, refurbished, or discarded and no longer used.

Thus, the choice of the nominal characteristics of measuring instruments and the designation of permissible deviations of the real characteristics from them – rating of the metrological characteristics of measuring instruments – are of great importance for measurement practice. The practice of rating the metrological characteristics of measuring instruments has evolved over time, and we will examine it next.

Both the production of measuring instruments and the rating of their characteristics initially arose spontaneously in each country. Later, rules that brought order to the rating process were established in all countries with significant instrument production. The recommendations developed at this time by international organizations, primarily Publication 51 of the International Electrotechnical Commission (IEC) and a number of publications by the International Organization of Legal Metrology (OIML), were of great importance for standardizing the expression of rated characteristics [8, 9]. The terminological documents are also extremely valuable for developing rating procedures [1, 10, 12].

We shall now return to the gist of the problem. The values of nominal metrological characteristics, such as the upper limits of measurement ranges, the nominal values of the measures, the scale factors of instruments and so on, are chosen from a standardized series of values of these characteristics. A more difficult task is to rate the accuracy characteristics: errors and instability.

Despite the efforts of designers, the real characteristics of measuring instruments depend to some extent on the external conditions. For this reason, the conditions are determined under which the measuring instruments are to be calibrated and checked, including the nominal values of all influence quantities and the ranges of their permissible deviation from the nominal values. These conditions are called *reference conditions*. The error of measuring instruments under reference conditions is called the *intrinsic error*.

In addition to the reference conditions and intrinsic errors, the rated operating conditions of measuring instruments are also established, i.e., the conditions under which the characteristics of measuring instruments remain within certain limits and the measuring instruments can be employed as intended. Understandably, errors in the rated operating conditions are larger than errors under the reference conditions. This change is characterized by specifying the limits of the additional error (the additional error the instrument can have due to deviation of the corresponding influence quantity from the reference condition), the permissible value of the corresponding influence quantity, or by indicating the limits of the permissible error under the rated operating conditions (the overall possible error of the instrument).

The errors of measuring instruments are expressed not only in the form of absolute and relative errors, adopted for estimating measurement errors, but also in the form of *fiducial errors*. The fiducial error is the ratio of the permissible limits

of the absolute error of the measuring instrument to some standardized value – fiducial value. The latter value is established by standards on separate types of measuring instruments; we discuss these rules later in this section. The fiducial error is somewhat similar to relative error but, since it is normalized to a constant standardized value, the fiducial error is constant across the entire measurement range of the device. The purpose of fiducial errors is that they make it possible to compare the accuracy of measuring instruments that have different measurement ranges. For example, the accuracy of an ammeter with a measurement limit of 1 A and permissible absolute error of 0.01 A has the same fiducial error of 1% (and in this sense, the same accuracy) as an ammeter with a measurement limit of 100 A and permissible absolute error of 1 A.

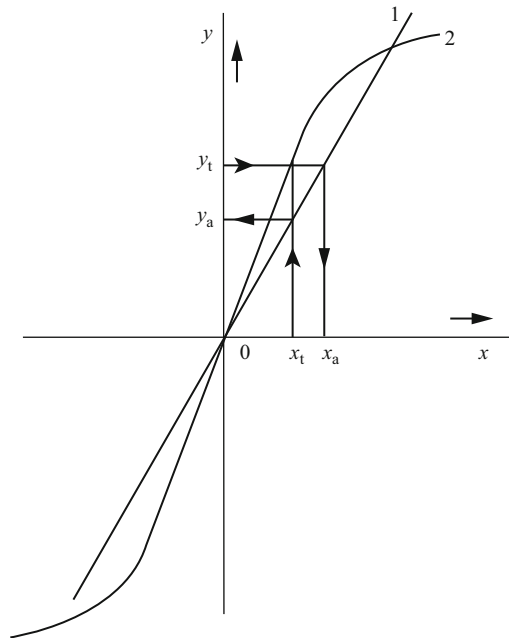
For measuring transducers, the errors can be represented relative to either the input or output signals. Let us consider the relationship between these two error representations.

Figure 2.1 shows the nominal and, let us assume, the real transfer functions of some transducer. The nominal transfer function, as done in practice whenever possible, is assumed to be linear. We denote the input quantity by x and the output quantity by y . They are related by the dependency

$$x = Ky,$$

where K is the nominal transduction constant.

Fig. 2.1 Nominal (curve 1) and real (curve 2) transfer functions of a measuring transducer



At the point with true values of the quantities x_t and y_t , the true value of the transduction constant will be $K_t = x_t/y_t$. Calculations based on the nominal constant K , however, result in an error.

Let $x_a = Ky_t$ and $y_a = x_t/K$ be determined based on y_t and x_t (see Fig. 2.1). Then the absolute transducer error with respect to the input will be

$$\Delta_a = y_a - y_t = \left(\frac{1}{K} - \frac{1}{K_t} \right) x_t.$$

The error with respect to the output is expressed analogously:

$$\Delta y = y_a - y_t = \left(\frac{1}{K} - \frac{1}{K_t} \right) x_t.$$

We note, first, that Δx and Δy always have different signs: If $(K - K_t) > 0$, then $(1/K - 1/K_t) < 0$.

But this is not the only difference. The quantities x and y can also have different dimensions; i.e., they can be physically different quantities, so that the absolute input and output errors are not comparable. For this reason, we shall study the relative errors:

$$\begin{aligned} \delta_x &= \frac{\Delta x}{x_t} = (K - K_t) \frac{y_t}{x_t} = \frac{K - K_t}{K_t}, \\ \delta_y &= \frac{\Delta y}{y_t} = \frac{(K_t - K) x_t}{K K_t y_t} = \frac{K_t - K}{K}. \end{aligned}$$

As $K_t \neq K$, we have $|\delta_x| \neq |\delta_y|$.

We denote the relative error in the transduction constant at the point (x_t, y_t) as δ_k , where $\delta_k = (K - K_t)/K_t$. Then

$$\frac{\delta_x}{\delta_y} = -(1 + \delta_k).$$

However, $\delta_k \ll 1$, and in practice relative errors with respect to the input and output can be regarded as equal in magnitude.

In measures, the rated error is determined as the difference between the nominal value of the measure and the “true value” of the quantity reproduced by the measure; the “true value” is obtained by another, known to be much more precise, measurement. This is analogous to indicating instruments if one considers the nominal value of a measure as the indication of the instrument.

It is interesting to note that single measures that reproduce passive quantities, for example, mass, electric resistance, and so on, have only systematic errors. The error of measures of active quantities (electric voltage, electric current, etc.) can have both systematic and random components. Multiple-valued measures of passive quantities can have random errors introduced by the switching elements.

To summarize, when the errors of measuring instruments are rated, the permissible limits of the intrinsic and all additional errors are specified. At the same time, the reference and rated operating conditions are indicated.

Of all forms enumerated above for expressing the errors of measuring instruments, the best is the relative error, because in this case, the indication of the permissible limit of error gives the clearest idea of the level of measurement accuracy that can be achieved with the given measuring instrument. The relative error, however, usually changes significantly over the measurement range of the instrument, and for this reason, it is difficult to be rated.

The absolute error is frequently more convenient than the relative error. In the case of an instrument with a scale, the limit of the permissible absolute error can be rated with the same numerical value for the entire scale of the instrument. But then it is difficult to compare the accuracies of instruments having different measurement ranges. This difficulty disappears when the fiducial errors are rated.

Let us now consider how the limits of permissible errors are expressed. For our discussion below, we shall follow primarily [9]. The limit of the permissible absolute error can sometimes be expressed by a single value (neglecting the sign):

$$\Delta = \pm a,$$

sometimes in the form of the linear dependence:

$$\Delta = \pm(a + bx), \quad (2.1)$$

where x is the nominal value of the measure, the indication of a measuring instrument, or the signal at the input of a measuring transducer, and a and b are constants, and sometimes by a general equation,

$$\Delta = f(x).$$

When the last dependence is complicated, it is given in the form of a table or graph.

The fiducial error γ (in percent) is defined by the formula

$$\gamma = 100\Delta/x_f,$$

where x_f is the fiducial value.

The fiducial value is assumed to be equal to the following:

1. The value at the end of the instrument scale.
2. The nominal value of the measurand, if it has been established.
3. The length of the scale, if the scale graduations narrow sharply toward the end of the scale. In this case, the error and the length of the scale are expressed in the same units of length (e.g., centimeters).

The rules above are in accordance with Recommendation 34 of OIML [9]. However, Publication 51 of IEC [8] foresees that if the zero marker falls within the scale, the fiducial value is equal to the span of the scale, which is a sum of the end values of the scale (neglecting their signs). This is controversial and we will discuss it in detail below.

A better between these two recommendations is the one by OIML. Indeed, consider, for example, an ammeter with a scale $-100-0-100\text{A}$ and with a permissible absolute error of 1 A. In this case, the fiducial error of the instrument will be 1% according to OIML and 0.5% according to IEC. But when using this instrument, the possibility of performing a measurement with an error of up to 0.5% cannot be guaranteed for any point of the scale, which makes the interpretation of the fiducial error confusing. An error not exceeding 1%, however, can be guaranteed when measuring a current of 100 A under reference conditions.

The tendency to choose a fiducial value such that the fiducial error would be close to the relative error of the instrument was observed in the process of improving IEC Publication 51. Indeed, in the previous edition of this publication, the fiducial value for instruments without a zero marker on the scale was taken to be equal to the difference of the end values of the range of the scale, and now it is taken to be equal to the larger of these values (neglecting the sign). Consider, for example, a frequency meter with a scale $45-50-55\text{ Hz}$ and the limit of permissible absolute error of 0.1 Hz. According to the previous edition of IEC Publication 51, the fiducial error of the frequency meter was assumed to be equal to 1%, and the current edition makes it equal to 0.2%. But when measuring the nominal 50 Hz frequency, the instrument relative error indeed will not exceed 0.2% (under reference conditions), while the 1% error has no relation to the accuracy of this instrument. Thus, the current edition is better. We hope that IEC will take the next step in this direction and take into account the recommendation of OIML for setting the fiducial value of instruments with a zero marker within the scale.

The limits of permissible relative error are rarely listed as rated but can be computed. If the rated error is expressed as the fiducial error γ (in percent), the permissible relative error for each value of the measurand must be calculated according to the formula

$$\delta = \gamma \frac{x_f}{x}.$$

If the rated error is expressed as the limits of absolute error Δ , the limit of permissible relative error δ is usually expressed in percent according to the formula

$$\delta = \frac{100\Delta}{x} = \pm c.$$

For digital instruments, the errors are often rated in the conventional form

$$\pm(b + q), \tag{2.2}$$

where b is the relative error in percent and q is some figure of the least significant digit of the digital readout device. For example, consider a digital millivoltmeter with a measurement range of 0–300 mV and with the indicator that has four digits. The value of one unit in the least significant digit of such an instrument is 0.1 mV. If this instrument is assigned the limits of permissible error (0.5% + 2), then number “2” in the parentheses corresponds to 0.2 mV. Now the limit of the relative error of the instrument when measuring, for example, a voltage of 300 mV can be calculated as follows:

$$\delta = \pm \left(0.5 + \frac{0.2 \times 100}{300} \right) = \pm 0.57\%.$$

Thus, to estimate the limit of permissible error of an instrument from the rated characteristics, some calculations must be performed. For this reason, although the conventional form (2.2) gives a clear representation of the components of instrument error, it is inconvenient to use.

A more convenient form is given in Recommendation 34 of OIML [9]: According to this recommendation, the limit of permissible relative error is expressed by the formula

$$\delta = \pm \left[c + d \left(\frac{x_e}{x} - 1 \right) \right], \quad (2.3)$$

where x_e is the end value of the measurement range of the instrument or the input signal of a transducer and c and d are relative quantities.

In (2.3), the first term on the right-hand side is the relative error of the instrument at $x = x_e$. The second term characterizes the increase of the relative error as the indications of the instrument decrease.

Equation 2.3 can be obtained from (2.2) as follows. To the figure q , there corresponds the measurand qD , where D is the value of one unit in the least significant digit of the instrument’s readout device. In the relative form, it is equal to qD/x . Now, the physical meaning of the sum of the terms b and qD/x is that it is the limit of permissible relative error of the instrument. So,

$$\delta = \left(b + \frac{qD}{x} \right).$$

Using identity transformation, we obtain

$$\delta = b + \frac{qD}{x} + \frac{qD}{x_e} - \frac{qD}{x_e} = \left(b + \frac{qD}{x_e} \right) + \frac{qD}{x_e} \left(\frac{x_e}{x} - 1 \right).$$

If we denote

$$c = b + \frac{qD}{x_e}, \quad d = \frac{qD}{x_e},$$

we obtain (2.3).

In application to the example of a digital millivoltmeter studied above, we have

$$\delta = \pm \left[0.57 + 0.07 \left(\frac{x_e}{x} - 1 \right) \right].$$

It is clear that the last expression is more convenient to use, and in general, it is more informative than the conventional expression (2.2).

Note that for standardization of analog instruments, the error limits are established for the total instrument error and not for the separate components. If, however, the instrument has an appreciable random component, then permissible limits for it are established separately, in addition to the limits of the total error. For example, aside from the limits of the permissible intrinsic error, the limits of the permissible variation are also established.

Additional errors (recall that these are errors due to the deviation of the corresponding influence quantities from their values falling within the reference condition) of measuring instruments are rated by prescribing the limits for each additional error separately. The intervals of variation of the corresponding influence quantities are indicated simultaneously with the limits of the additional errors. The collection of ranges provided for all influence quantities determines the rated operating conditions of the measuring instrument. The limits of permissible additional errors are often represented in proportion to the values of their corresponding influence quantities or the deviation of these quantities from the limits of the intervals determining their reference values. In this case, the corresponding coefficients are rated. We call them the influence coefficients.

In the case of indicating measuring instruments, additional errors are often referred to by the term *variation of indications*. This term is used, in particular, for electric measuring instruments [8].

The additional errors arising when the influence quantities are fixed are systematic errors. For different instruments of the same type, however, systematic errors can have different values and, moreover, different signs. For this reason, the documentation for the overwhelming majority of instrument types sets the limits of additional errors as both positive and negative with equal numerical values. For example, the change in the indications of an electric measuring instrument caused by a change in the temperature of the surrounding medium should not exceed the limits $\pm 0.5\%$ for each 10°C change in temperature under rated operating conditions (the numbers here are arbitrary).

If, however, the properties of different measuring devices of a given type are sufficiently uniform, it is best to standardize the influence function, i.e., to indicate the dependence of the indications of the instruments or output signals of the transducers on the influence quantities and the limits of permissible deviations from each such dependence. If the influence function can be standardized, then it is possible to introduce corrections to the indications of the instruments and thereby to use the capabilities of the instruments more fully.

Figure 2.2 shows how the instrument errors depend on the values of an influence quantity, assuming two basic alternatives for rating the additional errors. The upper

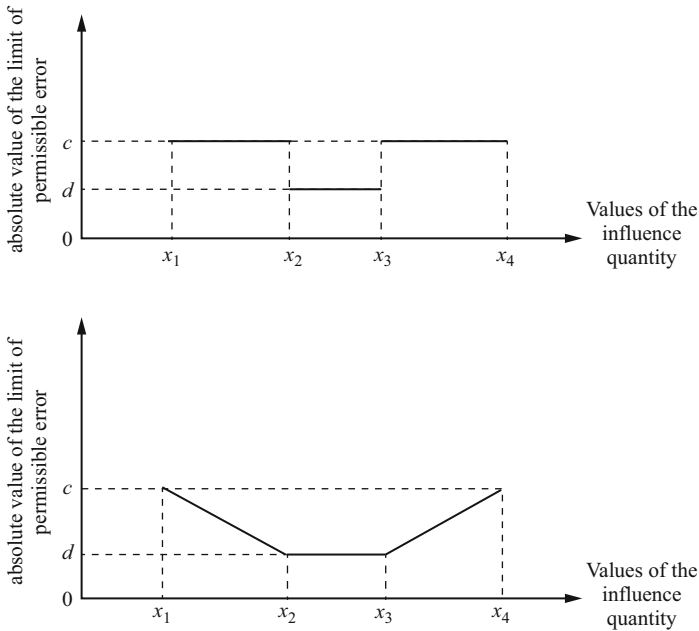


Fig. 2.2 Two variants of rating limits of additional errors of measuring instruments

figure represents the case where the documentation lists the limits of the intrinsic and additional errors. Such rating stipulates that the instrument accuracy is determined by the limits of the intrinsic error as long as the influence quantity is within reference condition and by the sum of the limits of the intrinsic and constant limits of the additional errors if the influence quantity is within rated operating condition. The lower figure depicts the case when the documentation lists the limits of the intrinsic error and the influence coefficients for the additional errors. Here, when the influence quantity is outside the reference condition, the limits of the additional error expand linearly with the deviation of the influence quantity from the reference condition (as long as the influence quantity stays within the rated operating conditions).

The interval (x_2, x_3) corresponds to reference conditions; the interval (x_1, x_4) corresponds to the rated operating conditions; d is the absolute value of the limits of permissible intrinsic error; c is the absolute value of the limits of permissible error in the rated operating conditions; and $(c-d)$ is the absolute value of the limits of permissible additional error.

It should be emphasized that the actual additional errors that can arise in a measurement will depend not only on the properties of the measuring instrument but also on the accuracy of obtaining the values of the corresponding influence quantities.

Often a measuring instrument has an electrical signal on its input. This input signal can be characterized by several parameters. One of them reflects the magnitude of the measurand. This parameter is called the *informative parameter*: By measuring its magnitude, we can find the value of the measurand. All other parameters do not have direct connections with the magnitude of the measurand, and they are called *noninformative parameters*.

Measuring instruments are constructed with the goal to make them insensitive to all noninformative parameters of the input signal. This goal, however, cannot be achieved completely, and in the general case, the effect of the noninformative parameters can only be decreased but not eliminated. But, for all noninformative parameters, it is possible to determine limits such that when the noninformative parameters vary within these limits, the total error of the measuring instrument will change insignificantly, which makes it possible to establish the reference ranges of the noninformative parameters.

If some noninformative parameter falls outside the reference limits, then the error arising is regarded as another additional error. The effect of each noninformative parameter is rated separately, as for influence quantities. Furthermore, rating the additional errors arising from noninformative parameters is done based on the same assumptions as those used for rating the additional errors caused by the influence quantities.

The errors introduced by changes in the noninformative parameters of the input signals are occasionally called dynamic errors. In the presence of multiple parameters, however, this term is not expressive. It is more intuitive to give each error a characteristic name, as is usually done in electric and radio measurements. For example, the change in the indications of an AC voltmeter caused by changes in the frequency of the input signal is called the frequency error. In the case of the measurements of the peak variable voltages, apart from the frequency errors, the errors caused by changes in the widths of the pulse edges, the decay of the flat part of the pulse, and so on are called the shape errors.

Another property of measuring instruments that affects their accuracy and is also rated is stability. Stability, like accuracy, is a positive quality of a measuring instrument. Just as the accuracy is characterized by inaccuracy (error, uncertainty), stability is characterized by instability. An important particular case of instability is drift. Drift is usually not rated. Instead, when it is discovered, the zero indication of the instrument is reset.

The first method of rating the instability involves stipulating the time period after which the instrument must be checked and calibrated if needed. The second method consists of indicating different limits for the error of the instrument for different periods of time after the instrument was calibrated. For example, the following table (taken with modifications from [18]) can be provided in the specifications of a digital instrument:

Time after calibration	24 h	3 months	1 year	2 years
Temperature	$23 \pm 1 \text{ }^\circ\text{C}$	$23 \pm 5 \text{ }^\circ\text{C}$	$23 \pm 5 \text{ }^\circ\text{C}$	$23 \pm 5 \text{ }^\circ\text{C}$
Limits of error	$\pm(0.01\% + 1 \text{ unit})$	$\pm(0.015\% + 1 \text{ unit})$	$\pm(0.02\% + 1 \text{ unit})$	$\pm(0.03\% + 2 \text{ units})$

In the last line entries, the first number in the parentheses specifies the percent of the instrument indication and the second is a figure of the least significant digit (from 0 to 9). The second number lists the absolute error in units of the least significant digit of the instrument. To find the corresponded part of the limits of error of that instrument, one must calculate the value of this number in units of measurement. For example, if the above table is given in the documentation of a millivoltmeter with the range of 300 mV and 4-digit readout device, then the value of the least-significant digit is 0.1 mV. Assume that a user utilizes this instrument 2 years after calibration and the readout is 120.3 mV. Then, the limits of error of this instrument for this measurement are $\pm(120.3 \times 0.0003 + 0.2) = \pm 0.24 \text{ mV}$. The second number is constant for a given instrument range. It was called the *floor error* in [18].

Obviously, specifying how instrument accuracy changes with time since calibration conveys more information about the instrument characteristics than simply rating the interval between calibrations, and this extra information is beneficial to the users.

Below is another example of specification of a digital multirange voltmeter, also from [18] (the specification for only two ranges is shown).

The last two rows in the above table give the limits of error of the instrument depending on the time from the calibration. The numbers in parentheses specify limits of two additive parts of the error in ppm. A confusing aspect here is that the first part is expressed as a relative error since the first number gives the limits of error relative to the indication of the instrument for a given measurement, while the second number specifies the error relative to the instrument range, the same as the floor error in the previous example.

Time after calibration	24 h	90 days	12 months	Temperature coefficient
Temperature	$23 \pm 1 \text{ }^\circ\text{C}$	$23 \pm 5 \text{ }^\circ\text{C}$	$23 \pm 5 \text{ }^\circ\text{C}$	$0\text{--}18^\circ$ and $28\text{--}55 \text{ }^\circ\text{C}$ Per $1 \text{ }^\circ\text{C}$
10 V	–	–	$\pm(35 + 5 \text{ ppm})$	$\pm(5 \text{ ppm} + 1 \text{ ppm})$
1,000 V	$\pm(20 + 6 \text{ ppm})$	$\pm(35 + 10 \text{ ppm})$	$\pm(45 + 10 \text{ ppm})$	$\pm(5 \text{ ppm} + 1 \text{ ppm})$

The last column specifies the limits of the additional error due to temperature deviation from reference conditions. These limits are rated in the form shown in the lower graph of Fig. 2.2: the limits of the additional error grow by the specified amount for each $1 \text{ }^\circ\text{C}$ of temperature deviation.

We provide examples of using this table in Sect. 4.6 for a measurement under reference temperature conditions and in Sect. 4.7 for a measurement under rated conditions.

The above excerpts of instrument specifications show the importance of understanding conventions used by the manufacturer of the instrument in specifying the instrument accuracy in its certificate. This is especially true if the manufacturer does not follow recommendations for rating the accuracy of instruments that have been issued by organizations such as OIML.

Rating of errors predetermines the properties of measuring instruments and is closely related with the concept of *accuracy classes* of measuring instruments. The purpose of this concept is the unification of the accuracy requirements of measuring instruments, the methods for determining them, and the accuracy-related notation in general, which is certainly useful to both the manufacturers of measuring instruments and to users. Indeed, such unification makes it possible to limit, without harming the manufacturers or the users, the list of instruments, and it makes it easier to use and check the instruments. We shall now discuss this concept in greater detail.

Accuracy classes were initially introduced for indicating electric measuring instruments [8]. Later this concept was also extended to all other types of measuring instruments [9]. In [1], the following definition is given for the term accuracy class: The accuracy class is a class of measuring instruments or measuring systems that meet certain stated metrological requirements intended to keep instrumental errors or uncertainties within specified limits under specified operating conditions.

Unfortunately, this definition does not entirely reflect the meaning of this term. Including measurement systems into the definition is incorrect because systems are usually unique and thus are not divided into classes. Further, instrumental errors and uncertainties are properties of measurements – not instruments – and hence should not be used to define instrument classes. A better definition is given in the previous edition of VIM: The accuracy class is a class of measuring instruments that meets certain metrological requirements that are intended to keep errors within specified limits.



Every accuracy class has conventional notation, established by agreement – the class index – that is presented in [8, 9]. On the whole, the accuracy class is a generalized characteristic that determines the limits for all errors and all other characteristics of measuring instruments that affect the accuracy of measurements performed with their help.

For measuring instruments whose permissible limits of intrinsic error are expressed in the form of relative or fiducial errors, the following series of numbers, which determine the limits of permissible intrinsic errors and are used for denoting the accuracy classes, was established in [9]:

$$(1, 1.5, 1.6, 2, 2.5, 3, 4, 5, \text{ and } 6) \times 10^n,$$

where $n = +1, 0, -1, -2, \dots$; the numbers 1.6 and 3 can be used, but are not recommended. For any one value of n , not more than five numbers of this series

Table 2.1 Designations of accuracy classes

Form of the expression for the error	Limit of permissible error (examples)	Designation of the accuracy class (for the given example)
Fiducial error, if the fiducial value is expressed in units of the measurand	$\gamma = \pm 1.5\%$	1.5
Fiducial error, if the fiducial value set equal to the scale length	$\gamma = \pm 0.5\%$	
Relative error, constant	$\delta = \pm 0.5\%$	
Relative error, increasing as the measurand decreases	$\delta = \pm [0.02 + 0.01(\frac{x_c}{x} - 1)]\%$	0.02/0.01

(i.e., no more than five accuracy classes) are allowed. The limit of permissible intrinsic error for each type of measuring instrument is set equal to one number in the indicated series.

Table 2.1 gives examples of the adopted designations of accuracy classes of these measuring instruments.

In those cases when the limits of permissible errors are expressed in the form of absolute errors, the accuracy classes are designated by Latin capital letters or roman numerals. For example, [41] gives the accuracy classes of block gauges as Class X, Y, and Z. Gauges of Class X are the most accurate; gauges of Class Y are less accurate than Class X, and gauges of Class Z are the least accurate.

If (2.3) is used to determine the limit of permissible error, then both numbers c and d are introduced into the designation of the accuracy class. These numbers are selected from the series presented above, and in calculating the limits of permissible error for a specific value of x , the result is rounded so that it would be expressed by not more than two significant digits.

In conclusion, we shall formulate the basic rules for rating errors of measuring instruments:

1. All properties of a measuring instrument that affect the accuracy of the results of measurements must be rated.
2. Every property that is to be rated should be rated separately.
3. Rating methods must make it possible to check experimentally, and as simply as possible, how well each individual measuring instrument corresponds to the established requirements.

Sometimes, exceptions must be made to these rules. In particular, an exception is necessary for strip strain gauges that can be glued on an object only once. Since these strain gauges can be applied only once, the gauges that are checked can no longer be used for measurements, whereas those that are used for measurements cannot be checked or calibrated.

In this case, it is necessary to resort to regulation of the properties of a *collection* of strain gauges, such as, for example, the standard deviation of the sensitivity and mathematical expectation of the sensitivity. The sensitivity of a particular strain gauge, which is essentially not a random quantity in the separate device, is a random quantity in a collection of strain gauges. Since we cannot check all the gauges, a random sample, representing a prescribed p percent of the entire collection (which could be, e.g., all gauges produced in a given year), is checked. Once the sensitivity of every selected gauge has been determined, it is possible to construct a *statistical tolerance interval*, i.e., the interval into which the sensitivity of any random sample of p percent of the entire collection of strain gauges will fall with a chosen probability α . As $\alpha \neq 1$ and $p \neq 1$, there is a probability that the sensitivity of any given strain gauge falls outside these tolerance limits. For this reason, the user must take special measures that address such a case. In particular, several strain gauges, rather than one, should be used.

2.4 Dynamic Characteristics of Measuring Instruments

The dynamic characteristics of measuring instruments reflect the relation between the change in the output signal and an action that produces this change. The most important such action is a change in the input signal. In this case, the dynamic characteristic is called the dynamic characteristic for the input signal. Dynamic characteristics for various influence quantities and for a load (for measuring instruments whose output signal is an electric current or voltage) are also studied.

Complete and partial dynamic characteristics are distinguished [28].

The complete dynamic characteristics determine uniquely the change in time of the output signal caused by a change in the input signal or by other action. Examples of such characteristics include a differential equation, transfer function, amplitude-and phase-frequency response, and the transient response. These characteristics are essentially equivalent, but the differential equation is the basic characteristic from which the other characteristics are derived.

A partial dynamic characteristic is a parameter of the full dynamic characteristic (introduced shortly) or the response time of the instrument. Examples are the response time of the indications of an instrument and the transmission band of a measuring amplifier.

Measuring instruments¹ can most often be regarded as inertial systems of first or second order. If $x(t)$ is the signal at the input of a measuring instrument and $y(t)$ is the corresponding signal at the output, then the relation between them can be expressed with the help of first-order (2.4) or second-order (2.5) differential

¹The rest of this section requires familiarity with control theory. The reader can skip this portion without affecting the understanding of the rest of the book.

equations , respectively, which reflect the dynamic properties of the measuring instrument:

$$Ty'(t) + y(t) = Kx(t), \quad (2.4)$$

$$\frac{1}{\omega_0^2}y''(t) + \frac{2\beta}{\omega_0}y'(t) + y(t) = Kx(t). \quad (2.5)$$

The parameters of these equations have specific names: T is the time constant of a first-order device, K is the transduction coefficient in the static state, ω_o is the angular frequency of free oscillations, and β is the damping ratio. An example of a real instrument whose properties are specified by the second-order differential equation is a moving-coil galvanometer. In this instrument type, $\omega_o = 2\pi/T_o$, where T_o is the period of free oscillations (the reverse of the natural frequency) and β is the damping ratio, which determines how rapidly the oscillations of the moving part of the galvanometer will subside.

Equations 2.4 and 2.5 reflect the properties of real devices, and for this reason, they have zero initial conditions: for $t \leq 0$, $x(t) = 0$ and $y(t) = 0$, $y'(t) = 0$ and $y''(t) = 0$.

To obtain transfer functions from differential equations, it is first necessary to move from signals in the time domain to their Laplace transforms, and then to obtain the ratio of the transforms. Thus,

$$\begin{aligned} \mathcal{L}[x(t)] &= x(s) & \mathcal{L}[y(t)] &= y(s), \\ \mathcal{L}[y'(t)] &= sy(s) & \mathcal{L}[y''(t)] &= s^2y(s), \end{aligned}$$

where s is the Laplace operator.

For the first-order system, in accordance to (2.4), we obtain

$$W(s) = \frac{y(s)}{x(s)} = \frac{K}{1 + sT},$$

and for the second-order system, from (2.5), we obtain

$$W(s) = \frac{y(s)}{x(s)} = \frac{K}{(1/\omega_0^2)s^2 + (2\beta/\omega_0)s + 1}. \quad (2.6)$$

Let us consider the second-order equation in more detail. If in the transfer function the operator s is replaced by the complex frequency $j\omega$ ($s = j\omega$), then we obtain the complex frequency response. We shall now study the relation between the named characteristics for the example of a second-order system. From (2.5) and (2.6), we obtain

$$W(j\omega) = \frac{K}{(1 - \omega^2/\omega_0^2) + j2\beta\omega/\omega_0}, \quad (2.7)$$

where $\omega = 2\pi f$ is the running angular frequency.

The complex frequency response is often represented with its real and imaginary parts,

$$W(j\omega) = P(\omega) + jQ(\omega).$$

In our case,

$$P(\omega) = \frac{K(1 - (\omega^2/\omega_0^2))}{(1 - (\omega^2/\omega_0^2))^2 + 4\beta^2(\omega^2/\omega_0^2)},$$

$$Q(\omega) = \frac{2\beta(\omega/\omega_0)K}{(1 - (\omega^2/\omega_0^2))^2 + 4\beta^2(\omega^2/\omega_0^2)}.$$

The complex frequency response can also be represented in the form

$$W(j\omega) = A(\omega)e^{j\varphi(\omega)},$$

where $A(\omega)$ is the amplitude-frequency response and $\varphi(\omega)$ is the frequency response of phase. In the case at hand,

$$A(\omega) = \sqrt{P^2(\omega) + Q^2(\omega)} = \frac{K}{\sqrt{(1 - (\omega^2/\omega_0^2))^2 + 4\beta^2(\omega^2/\omega_0^2)}} \quad (2.8)$$

$$\varphi(\omega) = \arctan \frac{Q(\omega)}{P(\omega)} = -\arctan \frac{2\beta(\omega/\omega_0)}{1 - (\omega^2/\omega_0^2)}.$$

Equation (2.8) has a well-known graphical interpretation using the notion of transient response. The transient response is the function $h(t)$ representing the output signal produced by a unit step function $1(t)$ at the input. (The unit step function, which we denote $1(t)$, is a function whose value is 0 for $t < 0$ and 1 for $t \geq 0$.) As the input is not periodic, $h(t)$ is calculated with (2.4) or (2.5). Omitting the technical but, unfortunately, complicated calculations, we arrive at the final form of the transient response of the instrument under study:

$$h(t) = \begin{cases} 1 - e^{-\beta\tau} \frac{1}{\sqrt{1 - \beta^2}} \sin \left(\tau \sqrt{1 - \beta^2} + \arctan \frac{\sqrt{1 - \beta^2}}{\beta} \right) & \text{if } \beta < 1, \\ 1 - e^{-\tau}(\tau + 1) & \text{if } \beta = 1, \\ 1 - e^{-\beta\tau} \frac{1}{\sqrt{\beta^2 - 1}} \sinh \left(\tau \sqrt{\beta^2 - 1} + \operatorname{arctanh} \frac{\sqrt{\beta^2 - 1}}{\beta} \right) & \text{if } \beta > 1. \end{cases}$$

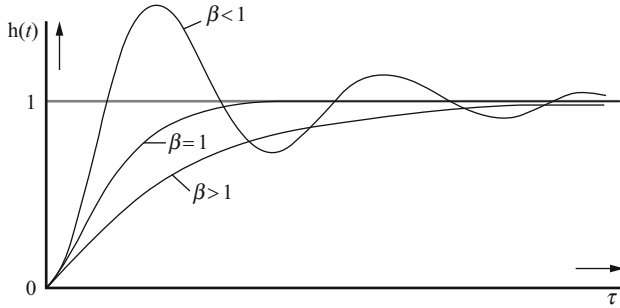


Fig. 2.3 The transient response of an instrument described by a second-order differential equation; β is the damping ratio

(Note that the last case utilizes hyperbolic trigonometric functions.) In this expression, $\tau = \omega_o t$ is normalized time, and the output signal is normalized to make its steady-state value equal to unity, i.e., $h(t) = y(t)/K$. Thus, the formulas above and the corresponding graphs presented in Fig. 2.3 are universal in the sense that they do not depend on the specific values of ω_o and K .

It should be noted that some types of measuring instruments do not have dynamic characteristics at all; these include measures of length, weights, vernier calipers, and so on. Some measuring instruments, such as measuring capacitors (measures of capacitance), do not have an independent dynamic characteristic by themselves. But when they are connected into an electric circuit, which always has some resistance and sometimes an inductance, the circuit always acquires, together with a capacitance, definite dynamic properties.

Measuring instruments are diverse. Occasionally, to describe adequately their dynamic properties, it is necessary to resort to nonlinear equations or equations with distributed parameters. However, complicated equations are used rarely, and it is not an accident. After all, measuring instruments are created specially to perform measurements, and their dynamic properties are made to guarantee convenience of use. For example, in designing a recording instrument, the transient response is made to be short, approaching the steady state level monotonically or oscillating insignificantly. In addition, the scale of the recording instrument is typically made to be linear. But when these requirements are met, the dynamic properties of the instrument can be described by one characteristic corresponding to a linear differential equation of order no higher than second.

Rating of the dynamic characteristics of measuring instruments is performed in two stages. First, an appropriate dynamic characteristic to be rated must be chosen, and second, the nominal dynamic characteristic and the permissible deviations from it must be established.

For recording instruments and universal measuring transducers, a complete dynamic characteristic, such as transient response, must be rated: Without having the complete dynamic characteristic, a user cannot effectively use these instruments.

For indicating instruments, it is sufficient to rate the response time. In contrast to the complete characteristics, this characteristic is a partial dynamic characteristic. The dynamic error is another form of a partial dynamic characteristic. Rating the limits of a permissible dynamic error is convenient for the measuring instruments employed, but it is justified only when the shape of the input signals does not change much.

For measuring instruments described by linear first- and second-order differential equations, the coefficients of all terms in the equations can be rated. In the simplest cases, the time constant is rated in the case of a first-order differential equation, and the natural frequency and the damping ratio of the oscillations are standardized in the case of a second-order differential equation.

When imposing requirements on the properties of measuring instruments, it is always necessary to keep in mind how compliance will be checked. For dynamic characteristics, the basic difficulties have to do with creating test signals of predetermined form (with sufficient accuracy), or with recording the input signal with a dynamically more accurate measuring instrument than the measuring instrument whose dynamic properties are being checked.

If adequately accurate test signals can be created and used to obtain the dynamic characteristic, i.e., a transient response as a response of a unit step function signal and frequency response as a response of a sinusoidal test signal, then in principle the instrument can be checked without any difficulties.

But sometimes the problem must be solved with a test signal that does not correspond to the signal intended for determining the complete dynamic characteristic. For example, one would think that tracing of signals at the input and output of a measuring instrument could solve the problem. In this case, however, special difficulties arise: small errors in recording the test signal and reading the values of the input and output signals often render the dynamic characteristic obtained from them physically meaningless and not corresponding to the dynamic properties of the measuring instrument. Such an unexpected effect occurs because the problem at hand is a so-called improperly posed problem. A great deal of attention is currently being devoted to such problems in mathematics, automatics, geophysics, and other disciplines. Improperly posed problems are solved by the methods of regularization, which essentially consist of the fact that the necessary degree of filtering (smoothing) of the obtained solution is determined based on a priori information about the true solution. Improperly posed problems in dynamics in application to measurement engineering are reviewed in [28, 51].

A separate problem, which is important for some fields of measurement, is the determination of the dynamic properties of measuring instruments directly when the instruments are being used. An especially important question here is the question of the effect of random noise on the accuracy with which the dynamic characteristics are determined.

This section, then, has been a brief review of the basic aspects of the problem of rating and determining the dynamic properties of measuring instruments.

2.5 Calibration and Verification of Measuring Instruments

Every country wishes to have trustworthy measurements. One of the most important arrangements to achieve this goal is to have a system for keeping errors of all measuring instruments within permissible limits. Therefore, all measuring instruments in use are periodically checked. In the process, working standards are used either to verify that the errors of the measuring instruments being checked do not exceed their limits or to recalibrate the measuring instruments.

The general term for the above procedures is *calibration*. But one should distinguish between a real calibration and a simplified calibration.

Real calibration results in the determination of a relation between the indications of a measuring instrument and the corresponding values of a working measurement standard. This relation can be expressed in the form of a table, a graph, or a function. It can also be expressed in the form of the table of corrections to the indications of the measuring instrument. In any case, as the result of real calibration, the indications of the working standard are mapped to the instrument being calibrated. Consequently, the accuracy of the instrument becomes close to the accuracy of the working standard.

Real calibration can be expensive, complex, and time-consuming.

Therefore, calibration is mostly used for precise and complex instruments. For other instruments, the simplified calibration suffices.

The simplified calibration (also called *verification*) simply reveals whether the errors of a measuring instrument exceed their specified limits. Essentially, verification is a specific case of quality control, much like quality control in manufacturing. And because it is quality control, verification results do have some rejects.

Further, verification can take the form of a complete or element-wise check. A complete check determines the error of the measuring instrument as a whole. In the case of an element-wise check, the errors of the individual elements comprising the measuring instrument are determined. The overall error of the measuring instrument is then calculated using methods that were examined in [44].

A complete check is always preferable as it gives the most reliable result. In some cases, however, a complete check is impossible to perform and one must resort to an element-wise check. For example, element-wise calibration is often employed to check measuring systems when the entire system cannot be delivered to a calibration laboratory and the laboratory does not have necessary working standards that could be transported to the system's site.

The units of a system are verified by standard methods. When the system is verified, however, in addition to checking the units, it is also necessary to check the serviceability of the system as a whole. The methods for solving this problem depend on the arrangement of the system, and it is hardly possible to make general recommendations here. As an example, the following procedure can be used for a system with a temperature-measuring channel comprising a platinum–rhodium–platinum thermocouple as the primary measuring transducer and a voltmeter.

After all units of the system have been checked, we note the indication of the instrument at the output of the system. Assume that the indication is $+470\text{ }^{\circ}\text{C}$. For the most common types of thermocouples, there exists known standardized transfer function, while specific brands of thermocouple products have rated limits of deviation from the standardized function.

From the standardized transfer function of the primary measuring transducer, we obtain the output signal that should be observed for the given value of the measured quantity. For our thermocouple, when the temperature of $470\text{ }^{\circ}\text{C}$ is measured, the EMF at the output of the thermocouple must be equal to 3.916 mV . Next, disconnecting the wires from the thermocouple and connecting them to the voltage exactly equal to the nominal output signal of the thermocouple, we once again note the indication of the voltmeter. If it remains the same or has changed within the limits of permissible error of the thermocouple and voltmeter, then the system is serviceable.

Of course, this method of checking will miss the case in which the errors of both the thermocouple and voltmeter are greater than their respective permissible errors but these errors mutually cancel. However, this result can happen only rarely. Moreover, such a combination of errors is in reality permissible for the system.

Let us now consider complete check verification in more detail. Here, the values represented by working standards are taken as true values, and the instrument indication is compared to these values. In fact, a working standard has errors. Therefore, some fraction of serviceable instruments, i.e., instruments whose errors do not exceed the limits established for them, is rejected in a verification – false rejection – and some fraction of instruments that are in reality unserviceable are accepted – false retention. This situation is typical for monitoring production quality, and just as with quality control, a probabilistic analysis of the procedure is useful to understand the extent of a potential issue.

Without loss of generality, suppose for simplicity that the complete check verification is performed by measuring the same quantity simultaneously using a working standard (which in this case is an accurate measuring instrument) and the instrument being checked. Accordingly, we have

$$A = x - \zeta = y - \gamma,$$

where A is the true value of the quantity, x and y are the indications of the instrument and working standard, and ζ and γ are the errors of the instrument and working standard. It follows from the above equation that the difference z between the indications of the instrument and the standard is equal to the difference between their errors,

$$z = x - y = \zeta - \gamma. \quad (2.9)$$

We are required to show that $|\zeta| \leq \Delta$, where Δ is the limit of permissible error of the instrument. From the experimental data (i.e., from the indications), we can find z ; because γ is supposed to be much smaller than ζ , we shall assume that if $|z| \leq \Delta$, then the checked instrument is serviceable, and if $|z| > \Delta$, then it is not serviceable.

To perform probabilistic analysis of when the above assumption is wrong, it is necessary to know the probability distribution for the errors of the checked and standard instruments. Let us suppose we know these distributions. The probability of a false rejection is

$$p_1 = P\{|\zeta - \gamma| > \Delta_{|\zeta| \leq \Delta}\},$$

and the probability of a false retention is

$$p_2 = P\{|\zeta - \gamma| \leq \Delta_{|\zeta| > \Delta}\}.$$

A false rejection is obtained for $|\zeta| \leq \Delta$ when $|\zeta - \gamma| > \Delta$, i.e.,

$$\zeta - \gamma > \Delta, \quad \zeta - \gamma < -\Delta,$$

or

$$\gamma < \zeta - \Delta, \quad \gamma > \zeta + \Delta.$$

If the probability density functions of the errors of the instrument and working standard are $f(\zeta)$ and $\varphi(\gamma)$, respectively, then

$$p_1 = \int_{-\Delta}^{\Delta} f(\zeta) \left(\int_{-\infty}^{\zeta - \Delta} \varphi(\gamma) d\gamma + \int_{\zeta + \Delta}^{+\infty} \varphi(\gamma) d\gamma \right) d\zeta.$$

A false retention is possible when $|\zeta| > \Delta$, i.e., when $\zeta > +\Delta$ and $\zeta < -\Delta$. In this case, $|\zeta - \gamma| \leq \Delta$, i.e.,

$$\zeta - \gamma \leq \Delta, \quad \zeta - \gamma \geq -\Delta,$$

or

$$\zeta - \Delta \leq \gamma \leq \zeta + \Delta.$$

Therefore,

$$p_2 = \int_{-\infty}^{-\Delta} f(\zeta) \left(\int_{\zeta - \Delta}^{\zeta + \Delta} \varphi(\gamma) d\gamma \right) d\zeta + \int_{\Delta}^{+\infty} f(\zeta) \left(\int_{\zeta - \Delta}^{\zeta + \Delta} \varphi(\gamma) d\gamma \right) d\zeta.$$

Thus, if the probability densities are known, then the corresponding values of p_1 and p_2 can be calculated; one can furthermore understand how these probabilities

depend on the difference between the limits of permissible errors of the instrument being checked and the working standard.

If, in addition, cost considerations are added, then one would think about the problem of choosing the accuracy of the working standard that would be suitable for checking a given instrument. In reality, when the accuracy of working standards is increased, the cost of verification increases also. A rejection also has a certain cost. Therefore, by varying the limits of error of working standards, it is possible to find the minimum losses, and this accuracy is regarded as optimal.

Mathematical derivations aside, it is unfortunately difficult to estimate the losses from the use of instruments whose errors exceed the established limits, when these instruments pass the verification. In general, it is hard to express in monetary terms the often-significant economic effect of increasing measurement accuracy. For this reason, it is only in exceptional cases that economic criteria can be used to justify the choice of the relation between the limits of permissible error of the working standard and the checked instruments.

In addition, as has already been pointed out above, the fundamental problem is to determine the probability distribution of the errors of the instruments and standards. The results, presented in Sect. 2.7 below, of the statistical analysis of data from the verification of a series of instruments showed that the sampling data of the instrument errors are statistically unstable. Therefore, the distribution function of the instrument errors cannot be found from these data. However, there are no other data; it simply cannot be obtained anywhere.

Thus, it is impossible to find a sufficiently convincing method for *choosing* the relation between the permissible errors of the working standard and the instrument to be checked. For this reason, in practice, this problem is solved by a volitional method, by *standardizing* the relation between the limits of permissible errors. In practice, the calibration laboratories accept that the accuracy of a working standard must be four times higher than the accuracy of the checked instrument [18]. This means that some instruments that pass the verification may have errors exceeding by 25% the permissible level. Yet more aggressive ratios between the limits of permissible errors of the standard and the instrument, such as 1:10, are usually technically difficult to achieve.

It turns out, however, that a change in the verification process can eliminate this problem. Let us consider this method.

By definition, a serviceable instrument is an instrument for which $|x - A| \leq \Delta$ and an instrument is unserviceable if $|x - A| > \Delta$. Analogous inequalities are also valid for a working standard: $|y - A| \leq \Delta_s$, if the instrument is serviceable and $|y - A| > \Delta_s$, if it is not serviceable.

For $x > A$, for a serviceable instrument, $x - A \leq \Delta$. But $y - \Delta_s \leq A \leq y + \Delta_s$. For this reason, replacing A by $y - \Delta_s$, we obtain for a serviceable instrument,

$$x - y \leq \Delta - \Delta_s. \quad (2.10)$$

Analogously, for $x < A$, for a serviceable instrument,

$$x - y \geq -(\Delta - \Delta_s). \quad (2.11)$$

Repeating the calculations for an unserviceable instrument, it is not difficult to derive the corresponding inequalities:

$$x - y > \Delta + \Delta_s. \tag{2.12}$$

$$x - y < -(\Delta + \Delta_s). \tag{2.13}$$

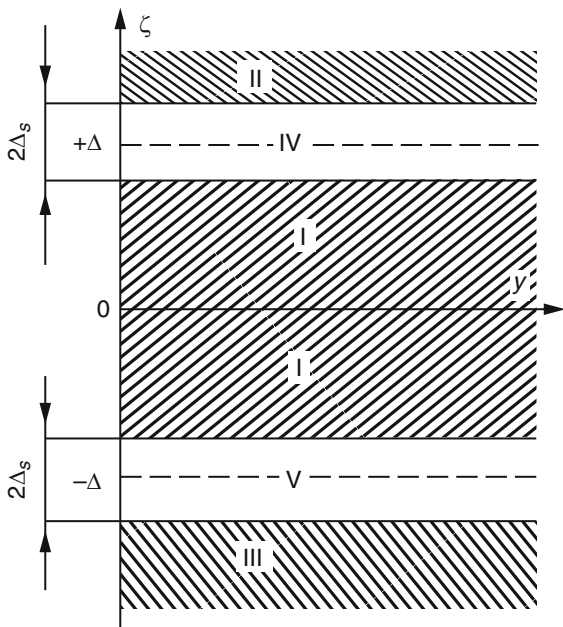
Figure 2.4 graphically depicts the foregoing relations. Let the scale of the checked instrument be the abscissa axis. On the ordinate axis, we mark the points $+\Delta$ and $-\Delta$, and around each of these points, we mark the points displaced from them by $+\Delta_s$ and $-\Delta_s$. If Δ and Δ_s remain the same for the entire scale of the instrument, then we draw from the marked points on the ordinate axis straight lines parallel to the abscissa axis.

Region I corresponds to inequalities (2.10) and (2.11). The instruments for which the differences $x - y$ fall within this region are definitely serviceable irrespective of the ratio of the errors of the standard and checked instruments. Inequalities (2.12) and (2.13) correspond to regions II and III. The instruments for which the differences $x - y$ fall within the regions II or III are definitely unserviceable.

Some instruments can have errors such that

$$\Delta - \Delta_s < |x - y| < \Delta + \Delta_s.$$

Fig. 2.4 Zones of definite serviceability (I), definite rejection (II and III), and uncertainty (IV and V) during verification of measuring instruments with the limit of permissible error Δ based on a working standard whose limit of permissible error is Δ_s ,



These errors correspond to regions IV and V in Fig. 2.4. Such instruments essentially cannot be either rejected or judged to be serviceable, because in reality, they include both serviceable and unserviceable instruments. If they are assumed to pass verification, then the user will get some unserviceable instruments. This can harm the user. If, however, all such doubtful instruments are rejected, then in reality, some serviceable instruments will be rejected.

For instruments that are doubtful when they are manufactured or when they are checked after servicing, it is best that they be judged unserviceable. This tactic is helpful for the users and forces the manufacturers to employ more accurate standards to minimize the rejects. But this approach is not always practical. When the percentage of doubtful instruments is significant and the instruments are expensive and difficult to fix, it is best to check them again. Here, several variants are possible. One variant is to recheck the doubtful instruments with the help of more accurate working standards. When this is impossible, the verification can also be performed using other samples of working standards that are rated at the same accuracy as those used in the initial check. As different working standards have somewhat different errors, the results of comparing the instruments with them will be somewhat different. Thus, some doubtful instruments will move to the regions in Fig. 2.4 that allow definitive verification outcomes.

Ideally, the best way to deal with the doubtful instruments is to increase the accuracy of the working standard. However, the question then arises as to how much the accuracy of the standard instruments should be increased. If there are no technical limitations, then the accuracy of the working standard can be increased until the instrument can be judged as being either serviceable or unserviceable. However, the limits of permissible error of the standard instrument rarely need to be decreased beyond about ten times less than the limit of permissible error of the instrument: The errors of instruments are usually not stable enough to be estimated with such high accuracy.

Rejection of instruments under verification is eliminated completely if instead of verification the instruments are recalibrated. The accuracy of the newly calibrated instrument can be almost equal to the accuracy of the working standard, which makes this method extremely attractive. The drawback of this method is that the result of a calibration is most often presented in the form of a table of corrections to the indications of the instrument, which is inconvenient for using the instrument.

2.6 Designing a Calibration Scheme

Calibration is a metrological operation whose goal is to transfer decreed units of quantities from a primary measurement standard to a measuring instrument. To protect the primary standards and to support calibration of large numbers of instruments, this transfer is performed indirectly, with the help of intermediate standards. In fact, intermediate standards may themselves be calibrated against primary standards not directly but through other intermediary standards. Thus, the

sizes of units reproduced by primary standards are transferred to intermediary standards and through them to measuring instruments.

The hierarchical relations of standards with each other and with measuring instruments that are formed to support calibration can be represented as a *calibration scheme*. Note that the discussion in this section also fully applies to verification and *verification schemes*, which are the analog of calibration schemes in the context of verification. The standards at the bottom of the calibration schemes, which are used to calibrate measuring instruments, are called working standards; the intermediate standards, situated between the primary and working standards in the scheme, are called secondary standards. For the purpose of the discussion in this section, we will refer to secondary standards, working standards, and measuring instruments together as *devices*.

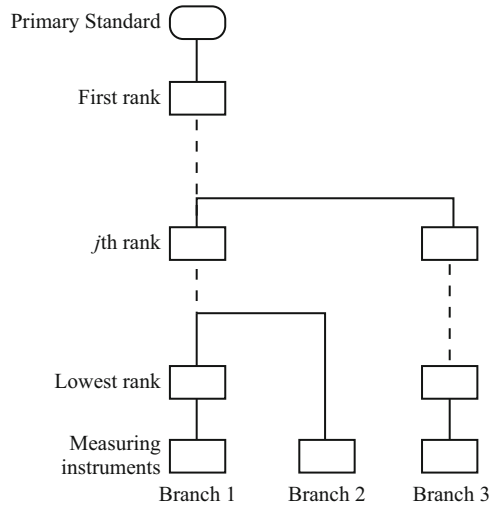
Measurement standards belonging to a calibration scheme are divided into ranks. The rank of a standard indicates the number of steps included in transferring the size of a unit from the primary measurement standard to a given standard, i.e., the number of standards on the path from this standard to the primary standard in the calibration scheme.

One of the most difficult questions arising in the construction of calibration schemes is the question of how many ranks of standards should be provided. Three main factors play a role in deciding this question: accuracy, cost, and capacity. As the number of ranks increases, the error with which the size of a unit is transferred to the measuring instrument increases, because some accuracy is lost at every calibration step. For this reason, to obtain high accuracy, the number of ranks of standards should be reduced to a minimum. On the other hand, the more the number of ranks the greater the overall capacity of the scheme in terms of the number of measuring instruments it can calibrate. In addition, the higher the accuracy of standards, the more expensive they are, and the more expensive they are to use. Thus, from the cost perspective, it is desirable to reduce the number of high-accuracy standards by increasing the number of ranks in the scheme.

One would think that it should be possible to find an economically optimal number of ranks of the calibration scheme. Such optimization, however, would require information about the dependence between the cost of the equipment and labor and the accuracy of calibration. This information is usually not available. For this reason, in practice, the optimal calibration schemes cannot be determined, and calibration schemes are commonly constructed in an ad hoc manner. However, a method proposed below allows designing a calibration scheme in a methodical way at least to satisfy its capacity requirements with the minimum number of ranks, and hence with the highest possible calibration accuracy. Accuracy constrains permitting; one can always then increase the number of ranks in the resulting scheme to reflect specific economic considerations.

Figure 2.5 shows a typical structure of a calibration scheme. In the simplest case, when all measuring instruments in the calibration scheme have similar accuracy, a calibration scheme can be represented as a chain; for example, the entire calibration scheme on Fig. 2.5 would consist of just branch 1. The chain has the primary standard at the root, then certain number of secondary standards of the rank 1 below

Fig. 2.5 A typical calibration scheme structure



that are periodically calibrated against the primary standard, followed by a larger number of secondary standards of rank 2, each periodically calibrated against one of the standards of rank 1, and so on until the measuring instruments at the leaves of the hierarchy.

However, some measuring instruments may be more accurate than others and cannot be calibrated by working standards at the bottom of the chain. These instruments must be “grafted” to the middle of the first branch, at the point where they can be calibrated by a standard of sufficient accuracy. These instruments form branch 2 on Fig. 2.5. The standard at the branching point in the calibration scheme acts as a secondary standard for one branch and a working standard for another.

Finally, there may be instruments of significantly different type than those in other branches, whose calibration requires some auxiliary devices between them and their working standards (such as scaling transducers in front of high-accuracy voltmeter for high voltage). The auxiliary devices introduce accuracy loss in calibration, and therefore they require the working standard to have a higher accuracy to account for this loss. In other words, if normally the accuracy ratio of the measuring instrument to working standard must be at most 1:4, (see Sect. 2.5 for the discussion on this accuracy relationship), this ratio must be lower (e.g., 1:10) for these instruments. To avoid the confusion, we place these instruments, along with the auxiliary devices, into distinct branches in the calibration scheme (such as branch 3 in Fig. 2.5). Such a branch can be grafted to another branch at an intermediary standard such that the ratio of its accuracy to the accuracy of the instruments corresponds to the requirement specific to the instruments’ branch.

Secondary standards are usually calibrated with the highest possible accuracy, so that they can be also used as working standards for more accurate types of measuring instruments if needed. However, there is inevitable loss of accuracy with each calibration step. Consequently, different types of secondary standards are

typically used for different ranks, and calibration at different ranks has different performance characteristics, such as time required to calibrate one device or time to prepare a standard for calibration (see below). At the same time, the types of devices that can be used at a given rank are usually known in advance, and it is only necessary to decide how many of them to procure and how to arrange them in an appropriate calibration scheme. Therefore, one can assume that the calibration frequency of secondary and working standards of a given rank, and how long each calibration takes, is known. Furthermore, we assume that the calibration frequency and time required to calibrate are known for all measuring instruments. Finally, the keepers of primary standards typically impose their own usage limits (e.g., they limit the number of calibrations that can be performed against the primary standard in 1 year). We assume that these limits are known as well.

We begin by considering the branch leading to the least accurate instruments as if it were the only branch in the scheme (e.g., branch 1 in Fig. 2.5). We call this branch a *stem*.

In such a single-branch calibration scheme, if the j th rank has N_j standards, then the maximum number of devices in the rank $(j + 1)$ that can be supported will be

$$N_{j+1} = N_j \frac{\eta_j T_{j+1}}{t_{j+1}} \quad (2.14)$$

where T_{j+1} is the time interval between calibrations of a device of rank $j + 1$, t_{j+1} is the time necessary to calibrate one device in the rank $(j + 1)$, and η_j is the utilization factor of the standards of rank j , considered below. Note that at the first calibration step, the number of secondary standards of rank 1 is determined as the minimum between the number given by (2.14) and the restrictions imposed by the keepers of the primary standards as mentioned earlier.

The utilization factor η_j reflects the fraction of time a corresponding standard can be used for calibration. In particular, η_j reflects the fact that the standard may only be used during the work hours; any losses of work time must also be taken into account. For example, if some apparatus is used 8 h per day and 1 h is required for preparation and termination, and preventative maintenance, servicing, etc. reduce the effective working time by 10%, then

$$\eta = \frac{8 - 1}{24} 0.9 = 0.3375.$$

Applying (2.14) to every step of the chain, we determine the capacity of the stem, which is the maximum number of standards of each rank and ultimately the number of measuring instruments $N_m^{(\max)}$ that can be supported by this calibration chain:

$$N_m^{(\max)} = N_0^{(\max)} N_1^{(\max)} \dots N_{m-1}^{(\max)} = \prod_{j=0}^{m-1} \eta_j \frac{T_{j+1}}{t_{j+1}}, \quad (2.15)$$

where m is the total number of steps in transferring the size of a unit from the primary standard to the measuring instrument, inclusively and $N_j^{(\max)}$ is the maximum number of devices at each rank that a “full” calibration scheme can have.

On the other hand, to design a calibration chain, that is, to decide on the number of ranks in the calibration chain that can support a given number N_{instr} of instruments, one can use the following procedure.

To protect the primary standards, they are never used to calibrate the working standards directly. Thus, at least one rank of secondary standards is always needed. We compute the maximum number of the secondary standards of rank 1 N_1 , which could be calibrated against the primary standard in our calibration chain, using (2.14). Next, we check using (2.14) again, if N_1 secondary standards can support calibration of N_{instr} instruments. If not, we know that we need more ranks in the calibration scheme.

In the latter case, we first check if the accuracy of the secondary standards of the new rank will still be sufficient to calibrate the instruments, given the instruments’ accuracy. If not, we have to assume that the calibration of the given number of instruments is impossible with the required calibration frequency (this outcome is extremely rare in practice). Otherwise, we apply (2.14) again to compute the maximum number of secondary standards of rank 2, N_2 , which can be supported by N_1 standards of rank 1. [Note that we apply (2.14) twice because the calibration time of a measuring instrument and secondary standard can be – and typically is – different]. We continue in this manner until we find the smallest number of ranks of secondary standards that can support N_{instr} measuring instruments.

We should mention that, after each iteration of the above algorithm, if the resulting capacity of the calibration scheme is close to required, an alternative to increasing the number of ranks is to raise the efficiency of calibration. This could be achieved by either increasing standard utilization η_j or by reducing the calibration time t_j . If the desired number of supported instruments cannot be achieved by increasing calibration efficiency, we proceed to increment the number of ranks.

Once we have determined the required number of ranks in the scheme, we compute the actual necessary number of standards at each rank in the bottom-up manner, starting from N_{instr} and computing the number of the next rank up by a resolving (2.14) relative to N_j :

$$N_j = N_{j+1} \frac{t_{j+1}}{\eta_j T_{j+1}}. \quad (2.16)$$

Once we are done with the stem of the calibration scheme, we can add remaining branches one at a time as follows. Let j_{attach} be the rank of the lowest-accuracy secondary standards on the stem suitable to calibrate the instruments of the new branch, and $N_{j_{attach}+1}^{(\max)}$ be the maximum number of devices that could be serviced by standards at this rank according to (2.15). Then, $N^{(slack)} = N_{j_{attach}+1}^{(\max)} - N_{j_{attach}+1}$ gives the number of devices that could be added.

If the number of instruments at the new branch according to (2.16) does not exceed $N^{(\text{slack})}$, we attach the new branch at rank j_{attach} , add the necessary number of standards at rank j_{attach} , and, moving from this rank up one step at a time, add the necessary number of standards at each rank (we are guaranteed that there will be enough capacity at each higher rank because the total number of devices at rank $j_{\text{attach}}+1$ does not exceed $N_{j_{\text{attach}}+1}^{(\text{max})}$).

Otherwise, that is, if the existing slack is insufficient, we must increase the capacity of the stem by adding an extra rank to add capacity. Accordingly, we recompute the number of devices at each rank of the stem in the bottom-up manner using (2.16), for the new number of ranks. After that, we repeat an attempt to attach the new branch from scratch.

If at some point we are unable to increment the number of ranks of the stem because the standard at the newly added rank would have insufficient accuracy, we would have to conclude that the given set of instruments is impossible to calibrate with the required accuracy using the available types of standards and the limitations on the use of the primary standard. However, given that the capacity of calibration schemes grows exponentially with the number of ranks, this outcome is practically impossible.

As the number of ranks increases, the capacity of the calibration network, represented by the checking scheme, increases rapidly. The calibration schemes in practice have at most five of ranks of standards, even for fields of measurement with large numbers of measuring instruments.

The relations presented above pertained to the simplest case, when at each step of transfer of the size of the unit, the period of time between calibrations and the calibration time were the same for all devices. In reality, these time intervals can be different for different types of devices. Taking this into account makes the calculations more complicated, but it does not change their essence. We consider these calculations next.

First, it is necessary to move from different time intervals between calibrations of different types of devices to one *virtual constant* time interval T_{vc} and to find the number of measuring instruments of each type N_k^{vc} that must be checked within this period. This is done using the obvious formula:

$$N_k^{vc} = N_k \frac{T_{vc}}{T_k}.$$

Next, it is necessary to find the average time t_j^{av} required to check one device for each step of the checking scheme:

$$t_j^{av} = \frac{\sum_{k=1}^n t_k N_k^{vc}}{\sum_{k=1}^n N_k^{vc}} \quad (2.17)$$

Here n is the number of different types of devices at the j -th step of the checking scheme.

We shall give a numerical example. Suppose it is required to organize a calibration of instruments of types A and B and the following data are given:

1. *Instruments of type A*: $N_A = 3 \times 10^4$; the time interval between calibrations $T_{A1} = 1$ year for $N_{A1} = 2.5 \times 10^4$ and $T_{A2} = 0.5$ year for $N_{A2} = 5 \times 10^3$; the calibration time $t_A = 5$ h.
2. *Instruments of type B*: $N_B = 10^5$; $T_B = 1$ year; the calibration time $t_B = 2$ h.
3. *Primary measurement standard*: Four comparisons per year are permitted, and the utilization factor of the primary standard is $\eta_0 = 0.20$.
4. *Secondary standards*: the frequency of the calibration of secondary standards of rank 1 is 2 years; i.e., $T_1 = 2$ years; the time to perform one calibration is 60 h, and utilization factor $\eta_1 = 0.25$. For the devices of rank 2, $T_2 = 2$ years, $t_2 = 40$ h, and $\eta_2 = 0.25$. The calibration parameters of higher-rank standards are the same as those of the rank-2 standards.

The possible number of first-rank standards in this case is limited by the restrictions on the primary standards use and can be found as

$$N_1^{(\max)} = N_0 f T_1 = 8$$

because $N_0 = 1$; $f = 4$ is the maximum number of comparisons with a reference standard per year, and $T_1 = 2$. Obviously, eight standards are not enough to check 130,000 measuring instruments. We shall now see how many ranks of standards will be sufficient.

As the time between calibrations is different for different instruments, we pick the illusory constant time interval $T_{vc} = 1$ year and find the number of instruments that must be checked within this time period. Conversion is necessary only for instruments of type A with $T_{A2} = 0.5$ years, since the calibration interval of the rest of the instruments matches T_{vc} :

$$N_{A2}^{vc} = N_{A2} \frac{T_{vc}}{T_2} = 5 \times 10^3 \times \frac{1}{0.5} = 10 \times 10^3$$

Therefore,

$$\sum_{k=A,B} N_k^{vc} = N_{AB} = N_{A1} + N_{A2}^{vc} + N_B = 135 \times 10^3$$

instruments must be calibrated within the time T_{vc} .

Different amounts of time are required to calibrate instruments of types A and B. The average calibration time t_{instr}^{av} of these working instruments, in accordance with (2.17), is

$$t_{instr}^{av} = \frac{(N_{A1} + N_{A2}^{vc})t_A + N_B t_B}{N_{AB}} = \frac{35 \times 10^3 \times 5 + 100 \times 10^3 \times 2}{135 \times 10^3} = 2.78\text{h.}$$

Now, using (2.14), we shall find the maximum number of second-rank standards:

$$N_2^{(\max)} = N_1 \frac{\eta_1 T_2}{t_2} = 8 \times \frac{0.25 \times 2 \times 6 \times 10^3}{40} = 600.$$

The maximum number of instruments that can be calibrated with the above number of rank-2 secondary standards is

$$N_{instr}^{(\max)} = N_2^{(\max)} \frac{\eta_2 T_{vc}}{t_{instr}^{av}} = 600 \times \frac{0.25 \times 365 \times 24}{2.78} = 472661.$$

Here, $T_{vc} = 365 \times 24 = 8.76 \times 10^3$ because 1 year = 365 days and η_2 was calculated for 24 h. The above number exceeds the total number of instruments N_{AB} to be calibrated; we thus conclude that two ranks are sufficient.

Next, we perform bottom-up calculations to find the necessary number of standards at each rank. The number of rank-2 standards is

$$N_2 = N_{AB} \frac{t_{instr}^{av}}{\eta_2 T_{vc}} = 135 \times 10^3 \times \frac{2.78}{0.25 \times 365 \times 24} = 171.$$

Similarly, one can check that all eight rank-1 secondary standards are needed, thus concluding the design of this calibration scheme.

Calculations similar to those in the above example allow one to choose in a well-grounded manner the structure of a calibration scheme and to estimate the required number of secondary standards of each rank. Calibration schemes in practice usually have extra capacity, which makes it possible to distribute secondary and working standards to limit their transport, to maximize the efficiency of calibration.

2.7 Statistical Analysis of Measuring Instrument Errors

A general characteristic of the errors of the entire population of measuring instruments of a specific type could be their distribution function. An important question then is if it is possible to find this function from experimental data. The studies in [47, 55] have addressed this question using the data provided by calibration laboratories on instrument errors they observed during calibration. These data thus reflected the sample of instruments that were calibrated; because it is impossible to obtain the errors of all instruments of a given type that are in use, the use of a sampling method is unavoidable.

To establish a property of an entire group (general population) based on a sample, the sample must be representative. Sample homogeneity is a necessary indicator of representativeness. In the case of two samples, to be sure that the samples are homogeneous, it is necessary to check the hypothesis $H_0: F_1 = F_2$, where F_1 and F_2 are distribution functions corresponding, respectively, to the first and second sample.

The results of a calibration, as is well known, depend not only on the error of the measuring instrument being calibrated but also on the error of the standard. For this reason, measuring instruments calibrated with not less than a fivefold margin of accuracy (i.e., using a standard at least five times more accurate than the instrument) were selected for analysis.

In addition, to ensure that the samples are independent, they were formed either based on data provided by calibration laboratories in different regions of the former USSR or, in the case of a single laboratory, on the data separated by a significant time interval. The sample sizes were maintained approximately constant. Errors exceeding twice the limit of permissible error were deemed outliers and eliminated from the analysis.

The test of hypothesis H_0 was performed using the Wilcoxon and Siegel-Tukey criteria with a significance level $q = 0.05$. The technique of applying these criteria is described in Chap. 3. Table 2.2 shows the result of these tests obtained in the study of [47]. The table includes two samples, obtained at different times, for each

Table 2.2 The homogeneity hypothesis testing for samples of six types of measuring instruments

Instrument type	Samples			Result of hypothesis testing	
	Year collected	Size	Tested point on scale	Wilcoxon	Siegel–Tukey
Э 59 Ammeter	1974	160	30-graduation mark	+	–
			60-graduation mark	0	–
	1976	160	80-graduation mark	0	–
			100-graduation mark	+	+
Э 59 Voltmeter	1974	120	70-graduation mark	–	0
	1976	108	150-graduation mark	+	+
Д 566 Wattmeter	1974	86	70-graduation mark	+	+
	1976	83	150-graduation mark	+	+
TH-7 Thermometer	1975		100 °C	0	–
			150 °C	–	+
	1976		200 °C	+	+
Standard spring manometer	1973	250	9.81 kPa		
	1976	250	9.81 kPa	+	+
P331 resistance measure	1970	400	10 kΩ	0	–
	1975	400	100 kΩ	0	–
		400	10 kΩ	0	–

instrument type. Rejection of the hypothesis is indicated by a minus sign, and acceptance is indicated by a plus sign. The symbol 0 means that a test based on the given criterion was not performed.

The Wilcoxon and Siegel–Tukey criteria are substantially different: The former is based on comparing averages, and the latter is based on comparing variances. For this reason, it is not surprising that there are cases when the hypothesis H_0 is rejected according to one criterion but accepted according to the other. The hypothesis of sample homogeneity must be rejected if even one of the criteria rejects it.

Both samples of instruments of a given type were found to be homogeneous only for the Д566 wattmeters and standard manometers. For other measuring instruments, the compared samples were often found to be nonhomogeneous. It is interesting that the samples can be homogeneous on one scale marker, and inhomogeneous on another (see Э59 voltmeters and ammeters). TH-7 thermometers had homogeneous samples in one range of measurement and inhomogeneous samples in a different range. The calculations were repeated for significance levels of 0.01 and 0.1, but the results were generally the same in both cases.

The above experiment was formulated to check the stability of the distribution functions of the errors, but because the instruments in the compared samples were not always the same, the result obtained has a different but no less important meaning: It indicates that the samples are inhomogeneous. It means that the parameters of one sample are statistically not the same as these parameters of another sample of the same type of measuring instruments.

Thus, the results obtained show that samples of measuring instruments are frequently nonhomogeneous with respect to errors. For this reason, they cannot be used to determine the distribution function of the errors of the corresponding instruments.

This result is also confirmed by the study of [55], which compared samples obtained from the data provided for Э59 ammeters by four calibration laboratories in different regions of the former USSR. The number of all samples was equal to 150–160 instruments. The errors were recorded at the markers 30, 60, 80, and 100 of the scale. The samples were assigned the numbers 1, 2, 3, and 4, and the hypotheses $H_0: F_1 = F_2, F_2 = F_3, F_3 = F_4,$ and $F_4 = F_2$ were checked (the pairs of samples to compare were selected arbitrarily). The hypothesis testing was based on the Wilcoxon criterion with $q = 0.05$. The analysis showed that we can accept the hypothesis $H_0: F_1 = F_2$ only, and only at the marker 100. In all other cases, the hypothesis had to be rejected.

Thus, sampling does not permit finding the distribution function of the errors of measuring instruments. Moreover, the fact that the sampling data are unstable could mean that the distribution functions of the errors of the instruments change in time. There are definite reasons for this supposition.

Suppose that the errors of a set of measuring instruments of some type, at the moment they are manufactured, have a truncated normal distribution with zero mean. For measures (measuring resistors, shunts, weights, etc.), a measure with a too large positive error makes this measure impossible to repair (one could fix a

weight whose mass exceeds the target by removing some material but one cannot repair a weight whose mass is too low). Furthermore, as measures age, their errors trend toward positive errors (e.g., weights lose some material due to polishing off with use). This is taken into account when manufacturing measures. For example, if in the process of manufacturing of a weight its mass is found to be even slightly less than the nominal mass then the weight is discarded. As a result, the distribution of the intrinsic errors of measures as they leave the factory is usually asymmetric.

Instrument errors change in the course of use. Usually the errors only increase. In those cases in which, as in the case of weights, the direction of the change of the errors is known beforehand and is taken into account during manufacturing, the errors can at first decrease, but then they will still increase. Correspondingly, changes in the instrument errors deform the distribution functions of the errors. This process, however, does not occur only spontaneously. At the time of routine checks, measuring instruments whose errors exceed the established limits are discarded, which again affects the distribution function of the errors of the remaining instruments.

The right-hand side of Fig. 2.6 shows the approximate qualitative picture of the changes occurring in the probability distribution of errors of a batch of weights in time. It shows the initial distribution of errors with all the errors being negative. With time, as the measures wear off, their errors decrease, with some positive errors starting to appear. As this trend continues, at some point some instruments start being discarded (which is shown in the figure by a vertical cut-off line at $+\Delta$ error limit). The process ultimately terminates when the measuring instruments under study no longer exist: either their errors exceed the established limits or they are no longer serviceable for other reasons.

The left-hand side of this figure shows an example of changes in error distribution in a batch of measuring instruments. In this example, the errors generally increase in time but the change is biased toward positive errors. Again, at some

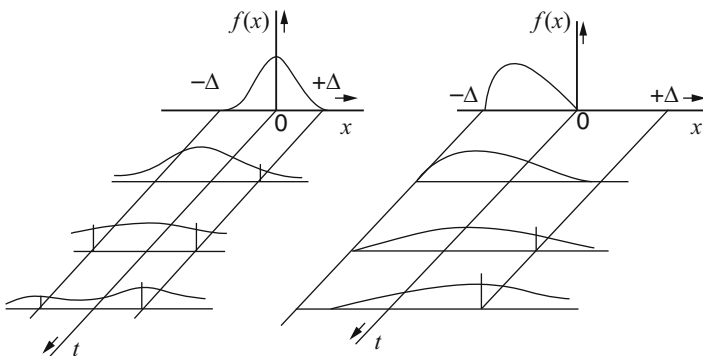


Fig. 2.6 Examples of possible changes in the probability densities of the errors of measuring devices in time. The figure on the *left* shows an example of changes in error distribution of a batch of measurement instruments; the figure on the *right* shows a possible change in error distribution of a batch of weights

point instruments start to be discarded, but most of the discarded instruments are those with positive errors.

There are other evident reasons for this result. One reason is that the stock of instruments of each type is not constant. On the one hand, new instruments that have just been manufactured are added to the stock. On the other hand, in the verification, some instruments are rejected, and some instruments are replaced. The ratio of the numbers of old and new instruments is constantly changing. Another reason is that groups of instruments are often used under different conditions, and the conditions of use affect differently the rate at which the instrumental errors change.

The temporal instability of measuring instruments raises the question of whether the errors of measuring instruments are in general sufficiently stable so that a collection of measuring instruments can be described by some distribution function. At a fixed moment in time, each type of instruments without doubt can be described by distribution function of errors. But the problem is how to find this distribution function. The simple sampling method, as we saw above, is not suitable. Moreover, even if the distribution function could be found by some complicated method, after some time, it would have to be redetermined, because the errors, and the composition of the stock of measuring instruments, change. Therefore, we have to conclude that the distribution of errors of measuring instruments cannot be found based on the experimental data.

The results presented above were obtained in the former USSR, and instruments manufactured in the former USSR were studied. However, there is no reason to expect that instruments manufactured in other countries will have different statistical properties.