Chapter 10 Quadratic Pricing and Hedging

Abstract We investigate pricing and hedging of the insurance payment process under quadratic objectives. Four types of quadratic loss functions are considered. First, we deal with a minimal hedging error in a mean-square sense. The hedging error is evaluated both under an equivalent martingale measure and the realworld measure. Next, we investigate locally risk minimizing strategies which lead to non-self-financing investment portfolio processes. Finally, we minimize an instantaneous mean-variance risk measure of the insurer's surplus to derive a hedging strategy. The pricing and hedging strategies are characterized by linear and nonlinear BSDEs.

Since a self-financing investment portfolio cannot perfectly replicate the insurance payment process, we aim to find an investment strategy which hedges the payment process with a minimal replication error. It is natural to measure the replication error with a quadratic function. In this chapter we investigate four types of quadratic objectives for pricing and hedging in incomplete markets. Quadratic objectives have gained their importance in portfolio optimization since Markowitz (1952) solved a one-period portfolio selection problem. We point out that the Markowitz mean-variance portfolio selection is used in practice for investment decision making and asset-liability management, see Zenios and Ziemba (2006) and Adam (2007). Let us also remark that replicating portfolios for insurance liabilities are constructed in practice by means of a mean-square criterion, see Boekel et al. (2009).

10.1 Quadratic Pricing and Hedging Under an Equivalent Martingale Measure

Let an equivalent martingale measure $\mathbb{Q} \in \mathcal{Q}^m$ be given. We formulate our pricing and hedging objective under the measure \mathbb{Q} . First, we define a class of admissible strategies under \mathbb{Q} .

Definition 10.1.1 A strategy $\pi := (\pi(t), 0 \le t \le T)$ is called admissible under the measure \mathbb{Q} , written $\pi \in \mathscr{A}^{\mathbb{Q}}$, if it satisfies the conditions:

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- (i) $\pi : [0, T] \times \Omega \to \mathbb{R}$ is a predictable process,
- (ii) $\mathbb{E}^{\mathbb{Q}}\left[\int_{0}^{T} |\pi(t)\sigma(t)|^{2} dt\right] < \infty$,
- (iii) there exists a unique càdlàg, adapted solution X^{π} to (7.11) on [0, T].

If assumptions (C2)–(C4) hold under \mathbb{Q} , then the wealth process X^{π} is square integrable under \mathbb{Q} for any admissible strategy $\pi \in \mathscr{A}^{\mathbb{Q}}$, see (7.12).

The insurer faces the payment process (7.3). We control the wealth of the insurer at time *T* which may denote the planning horizon in an ALM study or the point of time when the insurance portfolio terminates. The goal is to find an initial capital *x* and an admissible investment strategy $\pi \in \mathscr{A}^{\mathbb{Q}}$ which minimize the mean-square hedging error

$$\inf_{x,\pi\in\mathscr{A}^{\mathbb{Q}}} \mathbb{E}^{\mathbb{Q}} \Big[\Big| e^{-\int_0^T r(s)ds} X^{\pi,x}(T) - e^{-\int_0^T r(s)ds} F \Big|^2 \Big],$$
(10.1)

where the investment portfolio process $X^{\pi,x}$ is given by (7.11). Hence, we aim to find an investment portfolio which hedges the insurance payment process with a minimal replication error at time *T*. The use of the mean-square objective for measuring the replication error is very natural. The formulation of the objective under an equivalent martingale measure is mathematically and practically convenient, as we discuss later. The optimal capital, or the initial value of the optimal hedging portfolio, yields the price of the payment process.

Theorem 10.1.1 Let an equivalent martingale measure $\mathbb{Q} \in \mathcal{Q}^m$ be given and assume that (C1)–(C4) from Chap. 7 hold under \mathbb{Q} . We consider the quadratic pricing and hedging problem (10.1).

(a) There exists a unique solution $(Y, Z_1, Z_2, U) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2_N(\mathbb{R})$ to the BSDE

$$Y(t) = \int_{0}^{T} e^{-\int_{0}^{s} r(u)du} dP(s) - \int_{t}^{T} Z_{1}(s)dW^{\mathbb{Q}}(s) - \int_{t}^{T} Z_{2}(s)dB^{\mathbb{Q}}(s) - \int_{t}^{T} \int_{\mathbb{R}} U(s,z)\tilde{N}^{\mathbb{Q}}(ds,dz), \quad 0 \le t \le T,$$
(10.2)

where the square integrability conditions for (Y, Z_1, Z_2, U) hold under \mathbb{Q} .

(b) The optimal initial value of the hedging portfolio x* and the optimal admissible hedging strategy π* ∈ A^Q for the payment process P take the form

$$x^* = Y(0) = \mathbb{E}^{\mathbb{Q}}\left[\int_0^T e^{-\int_0^s r(u)du} dP(s)\right],$$
$$\pi^*(t) = \frac{Z_1(t)}{\sigma(t)} e^{\int_0^t r(s)ds}, \quad 0 \le t \le T.$$

Proof (a) The assertion follows from Theorem 3.1.1.

(b) We recall that the dynamics of $X^{\pi,x}$ is described by (7.11). By the change of measure, the Girsanov's theorem and discounting we obtain

$$e^{-\int_0^T r(s)ds} X^{\pi,x}(T)$$

= $x + \int_0^T e^{-\int_0^s r(u)du} \pi(s)\sigma(s)dW^{\mathbb{Q}}(s)$
 $-\int_0^T e^{-\int_0^s r(u)du} H(s)ds - \int_0^T \int_{\mathbb{R}} e^{-\int_0^s r(u)du} G(s,z)N(ds,dz).$ (10.3)

From (10.3) we conclude that the mean-square error (10.1) is given by

$$\mathbb{E}^{\mathbb{Q}}\left[\left|x+\int_{0}^{T}e^{-\int_{0}^{s}r(u)du}\pi(s)\sigma(s)dW^{\mathbb{Q}}(s)-\int_{0}^{T}e^{-\int_{0}^{s}r(u)du}dP(s)\right|^{2}\right],$$

and by (10.2) we get

$$\mathbb{E}^{\mathbb{Q}}\left[\left|x-Y(0)+\int_{0}^{T}\left(e^{-\int_{0}^{s}r(u)du}\pi(s)\sigma(s)-Z_{1}(s)\right)dW^{\mathbb{Q}}(s)\right.\right.\\\left.\left.-\int_{0}^{T}Z_{2}(s)dB^{\mathbb{Q}}(s)-\int_{0}^{T}\int_{\mathbb{R}}U(s,z)\tilde{N}^{\mathbb{Q}}(ds,dz)\right|^{2}\right].$$

Taking the squares and using moment properties of stochastic integrals, see Theorem 2.3.3, we derive

$$\begin{split} & \left(x - Y(0)\right)^2 + \mathbb{E}^{\mathbb{Q}} \left[\int_0^T \left| e^{-\int_0^s r(u) du} \pi(s) \sigma(s) - Z_1(s) \right|^2 ds \right] \\ & + \mathbb{E}^{\mathbb{Q}} \left[\int_0^T \left| Z_2(s) \right|^2 ds \right] + \mathbb{E}^{\mathbb{Q}} \left[\int_0^T \int_{\mathbb{R}} \left| U(s,z) \right|^2 Q(s,dz) \eta(s) ds \right] \\ & - 2\mathbb{E}^{\mathbb{Q}} \left[\int_0^T \left(e^{-\int_0^s r(u) du} \pi(s) \sigma(s) - Z_1(s) \right) dW^{\mathbb{Q}}(s) \int_0^T Z_2(s) dB^{\mathbb{Q}}(s) \right] \\ & - 2\mathbb{E}^{\mathbb{Q}} \left[\int_0^T \left(e^{-\int_0^s r(u) du} \pi(s) \sigma(s) - Z_1(s) \right) dW^{\mathbb{Q}}(s) \int_0^T \int_{\mathbb{R}} U(s,z) \tilde{N}^{\mathbb{Q}}(ds,dz) \right] \\ & + 2\mathbb{E}^{\mathbb{Q}} \left[\int_0^T Z_2(s) dB^{\mathbb{Q}}(s) \int_0^T \int_{\mathbb{R}} U(s,z) \tilde{N}^{\mathbb{Q}}(ds,dz) \right]. \end{split}$$

Finally, property (5.5) yields the hedging error

$$(x - Y(0))^{2} + \mathbb{E}^{\mathbb{Q}}\left[\int_{0}^{T} \left|e^{-\int_{0}^{s} r(u)du}\pi(s)\sigma(s) - Z_{1}(s)\right|^{2}ds\right]$$
$$+ \mathbb{E}^{\mathbb{Q}}\left[\int_{0}^{T} \left|Z_{2}(s)\right|^{2}ds\right] + \mathbb{E}^{\mathbb{Q}}\left[\int_{0}^{T} \int_{\mathbb{R}} \left|U(s,z)\right|^{2}Q(s,dz)\eta(s)ds\right],$$

and the optimality of (x^*, π^*) follows. The admissability of π^* is clear.

The price and the hedging strategy for the payment process *P* are characterized by the predictable representation of the discounted payments (or by the linear BSDE (10.2)). The price x^* is arbitrage-free. Recalling the interpretation given in Sect. 9.2 and Theorem 4.2.1, we can conclude that the hedging strategy π^* is a delta-hedging strategy. The hedging strategy π^* is updated with the current information on the financial and the insurance risk.

Example 10.1 We consider the classical Black-Scholes model with constant coefficients r, μ , σ and an insurer who issues a portfolio of n unit-linked endowment policies with a capital guarantee. We are interested in pricing and hedging the claim $F = (n - J(T))(K - S(T))^+$ where J is the deaths counting process for the insurance portfolio. We consider a constant mortality intensity λ and we choose the equivalent martingale measure \mathbb{Q}^* defined in Theorem 9.2.1. In order to find the price and the optimal hedging strategy for the claim F, we have to solve the BSDE

$$Y(t) = e^{-rT} (n - J(T)) (K - S(T))^{+} - \int_{t}^{T} Z(s) dW^{\mathbb{Q}^{*}}(s) - \int_{t}^{T} U(s) \tilde{N}^{\mathbb{Q}^{*}}(ds), \quad 0 \le t \le T.$$
(10.4)

From Proposition 8.1.1 and formula (9.34) we conclude that the solution to the BSDE (10.4) is given by the triple

$$\begin{split} Y(t) &= \left(n - J(t)\right) e^{-\lambda(T-t)} e^{-rt} \left(K e^{-r(T-t)} \Phi\left(-d\left(t, S(t)\right) - \sigma \sqrt{T-t}\right) \right. \\ &- S(t) \Phi\left(-d\left(t, S(t)\right)\right)\right), \quad 0 \leq t \leq T, \\ Z(t) &= -\left(n - J(t-)\right) e^{-\lambda(T-t)} \sigma S(t) e^{-rt} \Phi\left(-d\left(t, S(t)\right)\right), \quad 0 \leq t \leq T, \\ U(t) &= -e^{-\lambda(T-t)} e^{-rt} \left(K e^{-r(T-t)} \Phi\left(-d\left(t, S(t)\right) - \sigma \sqrt{T-t}\right) \right. \\ &- S(t) \Phi\left(-d\left(t, S(t)\right)\right)\right), \quad 0 \leq t \leq T. \end{split}$$

By Theorem 10.1.1 the optimal hedging strategy is given by the formula

$$\pi^*(t) = -\left(n - J(t-)\right)e^{-\lambda(T-t)}S(t)\Phi\left(-d\left(t,S(t)\right)\right), \quad 0 \le t \le T,$$

and the price of the claim, and the initial value of the optimal hedging portfolio, is equal to

$$x^* = Y(0) = ne^{-\lambda T} \left(Ke^{-rT} \Phi \left(-d(0, S(0)) - \sigma \sqrt{T} \right) - S(0) \Phi \left(-d(0, S(0)) \right) \right).$$

The value of the optimal hedging portfolio is determined by the process

$$X^{*}(t) = x^{*}e^{rt} + \int_{0}^{t} \pi^{*}(s)e^{r(t-s)} (\theta \sigma ds + \sigma dW(s)), \quad 0 \le t \le T.$$

The solution to the optimization problem (10.1) is easy to derive but the objective (10.1) has two drawback from the financial point of view. Firstly, we have to choose

the equivalent martingale measure \mathbb{Q} . This choice determines the price. When we investigate pricing in incomplete markets we expect to find the equivalent martingale measure. Consequently, the equivalent martingale measure used for pricing should be the output of an optimization problem, rather than the input. Secondly, the use of an equivalent martingale measure for evaluating the hedging error is questionable. Profits and losses or the performance of the hedging portfolio should be evaluated under the real-world measure.

10.2 Quadratic Pricing and Hedging Under the Real-World Measure

We aim to find an investment portfolio which hedges the insurance payment process (7.3) with a minimal replication error at time *T* and we evaluate the performance of the hedging portfolio under the real-world measure. The goal is to find an initial capital *x* and an admissible investment strategy $\pi \in \mathscr{A}$ which minimize the mean-square hedging error

$$\inf_{x,\pi\in\mathscr{A}} \mathbb{E}\left[\left|X^{\pi,x}(T)-F\right|^{2}\right],\tag{10.5}$$

where $X^{\pi,x}$ is given by (7.11), and we use the class of admissible strategies \mathscr{A} from Definition 7.3.1. If one prefers to minimize the hedging error (10.5) for the discounted quantities as in (10.1), then obvious modifications should be introduced.

We deal with two backward stochastic differential equations

$$Y(t) = 1 + \int_{t}^{T} \left(2Y(s)r(s) - \frac{|Z(s)|^{2}}{Y(s)} - |\theta(s)|^{2}Y(s) - 2\theta(s)Z(s) \right) ds$$
$$- \int_{t}^{T} Z(s)dW(s), \quad 0 \le t \le T,$$
(10.6)

and

$$\mathscr{Y}(t) = F + \int_{t}^{T} \left(-\mathscr{Y}(s) + H(s) + \int_{\mathbb{R}} G(s, z) Q(s, dz) \eta(s) - \mathscr{Z}_{1}(s) \theta(s) \right) ds - \int_{t}^{T} \mathscr{Z}_{1}(s) dW(s) - \int_{t}^{T} \mathscr{Z}_{2}(s) dB(s) - \int_{t}^{T} \int_{\mathbb{R}} \mathscr{U}(s, z) \tilde{N}(ds, dz), \quad 0 \le t \le T.$$
(10.7)

Equation (10.6) is called a stochastic Riccati equation. First, we establish existence and uniqueness results for these two BSDEs.

Proposition 10.2.1 Assume that (C1)–(C2) from Chap. 7 hold. There exists a unique solution $(Y, Z) \in \mathbb{S}^{\infty}(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R})$ to the BSDE (10.6) such that $Y(t) \geq k > 0$, $0 \leq t \leq T$. Moreover, the process $(\int_0^t Z(s)dW(s), 0 \leq t \leq T)$ is a BMO-martingale.

Proof Since we look for a process Y which is uniformly bounded away from zero, we can introduce new variables

$$\hat{Y}(t) = \frac{1}{Y(t)}, \quad \hat{Z}(t) = \frac{-Z(t)}{|Y(t)|^2}, \quad 0 \le t \le T.$$

The Itô's formula yields

$$\hat{Y}(t) = 1 + \int_{t}^{T} \left(\left(\left| \theta(s) \right|^{2} - 2r(s) \right) \hat{Y}(s) - 2\theta(s) \hat{Z}(s) \right) ds - \int_{t}^{T} \hat{Z}(s) dW(s), \quad 0 \le t \le T.$$
(10.8)

We end up with a linear BSDE. By Proposition 3.3.1 there exists a unique solution $(\hat{Y}, \hat{Z}) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R})$ to the linear BSDE (10.8). From the representation of the solution \hat{Y} and boundedness of θ and r we deduce that $0 < k \leq \hat{Y}(t) \leq K$, $0 \leq t \leq T$. Hence, we conclude that there exists a unique solution (Y, Z) to the nonlinear BSDE (10.6) and $0 < k \leq Y(t) \leq K$, $0 \leq t \leq T$. We now prove the *BMO* property. By the Itô's formula we obtain

$$d(|Y(t)|^2) = 2Y(t)dY(t) + |Z(t)|^2dt,$$

and

$$\int_{t}^{T} |Z(s)|^{2} ds$$

$$= 1 - |Y(t)|^{2}$$

$$+ \int_{t}^{T} 2Y(s) \left(2Y(s)r(s) - \frac{|Z(s)|^{2}}{Y(s)} - |\theta(s)|^{2}Y(s) - 2\theta(s)Z(s) \right) ds$$

$$- \int_{t}^{T} 2Y(s)Z(s) dW(s)$$

$$\leq K + K \int_{t}^{T} |Z(s)| ds - \int_{t}^{T} 2Y(s)Z(s) dW(s)$$

$$\leq K + K \frac{1}{\alpha} \int_{t}^{T} |Z(s)|^{2} ds + K\alpha - \int_{t}^{T} 2Y(s)Z(s) dW(s), \quad 0 \leq t \leq T,$$
(10.9)

where we use boundedness of r, θ and Y, strict positivity of Y and inequality (3.10). Choosing α sufficiently large and taking the expectation, we derive

$$\mathbb{E}\left[\int_{t}^{T} \left|Z(s)\right|^{2} ds |\mathscr{F}_{t}\right] \le K, \quad 0 \le t \le T,$$
(10.10)

and the BMO property is proved.

We will need higher moment estimates for the control process Z. From estimate (10.9) we deduce that for any $p \ge 2$ we have

$$\left(\int_0^T |Z(s)|^2 ds\right)^p \le K + K \frac{1}{\alpha^p} \left(\int_0^T |Z(s)|^2 ds\right)^p + K \alpha^p + \left|\int_0^T 2Y(s)Z(s)dW(s)\right|^p.$$

We choose α sufficiently large, take the expected value, apply the Burkholder-Davis-Gundy inequality and we obtain the inequality

$$\mathbb{E}\left[\left(\int_0^T \left|Z(s)\right|^2 ds\right)^p\right] \le K_p + K_p \mathbb{E}\left[\left(\int_0^T \left|Z(s)\right|^2 ds\right)^{p/2}\right].$$
 (10.11)

Starting with (10.10), by iteration we can derive the moment estimate

$$\mathbb{E}\left[\left(\int_0^T \left|Z(s)\right|^2 ds\right)^p\right] \le K_p, \quad p \ge 1.$$
(10.12)

We use estimate (10.12) in the sequel.

We now investigate the second BSDE.

Proposition 10.2.2 Assume that (C1)–(C4) from Chap. 7 hold.

(a) There exists a unique solution $(\mathscr{Y}, \mathscr{Z}_1, \mathscr{Z}_2, \mathscr{U}) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R})$ to the BSDE (10.7). The process \mathscr{Y} has the representation

$$\mathscr{Y}(t) = \mathbb{E}^{\mathbb{Q}^*} \bigg[e^{-\int_t^T r(s)ds} F \\ + \int_t^T e^{-\int_t^s r(u)du} \bigg(H(s) + \int_{\mathbb{R}} G(s,z)Q(s,dz)\eta(s) \bigg) ds |\mathscr{F}_t \bigg] \\ = \mathbb{E}^{\mathbb{Q}^*} \bigg[\int_t^T e^{-\int_t^s r(u)du} dP(s) |\mathscr{F}_t \bigg], \quad 0 \le t \le T,$$
(10.13)

where the equivalent martingale measure \mathbb{Q}^* is given by

$$\frac{d\mathbb{Q}^*}{d\mathbb{P}}\Big|\mathscr{F}_t = e^{-\int_0^t \theta(s)dW(s) - \frac{1}{2}\int_0^t |\theta(s)|^2 ds}, \quad 0 \le t \le T.$$

(b) *If*

$$\mathbb{E}\left[\left|F\right|^{p} + \left(\int_{0}^{T} \left|H(s)\right|^{2} ds\right)^{p/2} + \left(\int_{0}^{T} \int_{\mathbb{R}} \left|G(s,z)\eta(s)\right|^{2} Q(s,dz) ds\right)^{p/2}\right] < \infty,$$

for some p > 2, then

$$\mathbb{E}\left[\sup_{0\leq t\leq T}\left|\mathscr{Y}(t)\right|^{p}\right]<\infty.$$
(10.14)

Proof (a) The existence and the uniqueness follow from Theorem 3.1.1. Representation (10.13) can be derived by following the arguments from Proposition 3.3.1 and the fact that the compensator of N remains unchanged under \mathbb{Q}^* , see also the proof of Theorem 9.4.1.

(b) By the Burkholder-Davis-Gundy inequality, the dynamics of the BSDE (10.7), the Cauchy-Schwarz inequality and boundedness of r, θ we obtain

$$\begin{split} & \mathbb{E}\bigg[\bigg(\int_{0}^{T} |\mathscr{Z}_{1}(s)|^{2}\bigg)^{p/2}\bigg] \\ &\leq \mathbb{E}\bigg[\bigg(\int_{0}^{T} |\mathscr{Z}_{1}(s)|^{2}ds + \int_{0}^{T} |\mathscr{Z}_{2}(s)|^{2}ds + \int_{0}^{T} \int_{\mathbb{R}} |\mathscr{U}(s,z)|^{2}N(ds,dz)\bigg)^{p/2}\bigg] \\ &\leq K\mathbb{E}\bigg[\sup_{0 \leq t \leq T} \bigg|\int_{0}^{t} \mathscr{Z}_{1}(s)dW(s) + \int_{0}^{t} \mathscr{Z}_{2}(s)dB(s) + \int_{0}^{t} \int_{\mathbb{R}} \mathscr{U}(s,z)\tilde{N}(ds,dz)\bigg|^{p}\bigg] \\ &= K\mathbb{E}\bigg[\sup_{0 \leq t \leq T} \bigg|\mathscr{U}(t) - \mathscr{U}(0) + \int_{0}^{t} \bigg(-\mathscr{U}(s)r(s) + H(s) \\ &+ \int_{\mathbb{R}} G(s,z)Q(s,dz)\eta(s) - \mathscr{Z}_{1}(s)\theta(s)\bigg)ds\bigg|^{p}\bigg] \\ &\leq K\mathbb{E}\bigg[\big(1+T^{p}\big)\sup_{0 \leq t \leq T} |\mathscr{U}(t)|^{p} + T^{p/2}\bigg(\int_{0}^{T} |H(s)|^{2}ds\bigg)^{p/2} \\ &+ T^{p/2}\bigg(\int_{0}^{T} \int_{\mathbb{R}} |G(s,z)\eta(s)|^{2}Q(s,dz)ds\bigg)^{p/2} + T^{p/2}\bigg(\int_{0}^{T} |\mathscr{Z}_{1}(s)|^{2}ds\bigg)^{p/2}\bigg]. \end{split}$$

Assume now that the time horizon T is sufficiently small. We get the inequality

$$\mathbb{E}\left[\left(\int_{0}^{T} |\mathscr{Z}_{1}(s)|^{2}\right)^{p/2}\right] \leq \frac{KT^{p/2}}{1 - KT^{p/2}} + \frac{K(1 + T^{p})}{1 - KT^{p/2}} \mathbb{E}\left[\sup_{0 \leq t \leq T} |\mathscr{Y}(t)|^{p}\right].$$
(10.15)

From representation (10.7), the Cauchy-Schwarz inequality and boundedness of r and θ we deduce

10.2 Quadratic Pricing and Hedging Under the Real-World Measure

$$\begin{split} \left|\mathscr{Y}(t)\right|^{p} &= \left|\mathbb{E}\bigg[F + \int_{t}^{T} \left(-\mathscr{Y}(s)r(s) + H(s)\right) \\ &+ \int_{\mathbb{R}} G(s,z)Q(s,dz)\eta(s) - \mathscr{Z}_{1}(s)\theta(s) \right) ds \left|\mathscr{F}_{t}\right]\bigg|^{p} \\ &\leq \hat{K}\bigg(\mathbb{E}\bigg[\left|F\right|^{2} + T^{2}\sup_{0 \leq s \leq T}\left|\mathscr{Y}(s)\right|^{2} + T\int_{0}^{T}\left|H(s)\right|^{2} ds \\ &+ T\int_{0}^{T}\int_{\mathbb{R}}\left|G(s,z)\eta(s)\right|^{2}Q(s,dz) ds + T\int_{0}^{T}\left|\mathscr{Z}_{1}(s)\right|^{2} ds \left|\mathscr{F}_{t}\right]\bigg)^{p/2}, \\ &\quad 0 \leq t \leq T, \end{split}$$

which leads to the estimate

$$\begin{split} & \mathbb{E}\Big[\sup_{0 \le t \le T} \left|\mathscr{Y}(t)\right|^{p}\Big] \\ & \le \hat{K}\mathbb{E}\Big[\left|F\right|^{p} + T^{p/2}\left(\int_{0}^{T} \left|H(s)\right|^{2}ds\right)^{p/2} \\ & + T^{p/2}\left(\int_{0}^{T} \int_{\mathbb{R}} \left|G(s,z)\eta(s)\right|^{2}\mathcal{Q}(s,dz)ds\right)^{p/2} \\ & + T^{p}\sup_{0 \le t \le T} \left|\mathscr{Y}(t)\right|^{p} + T^{p/2}\left(\int_{0}^{T} \left|\mathscr{Z}_{1}(s)\right|^{2}\right)^{p/2}\Big] \\ & \le \hat{K}\Big(1 + T^{p/2} + T^{p}\mathbb{E}\Big[\sup_{0 \le t \le T} \left|\mathscr{Y}(t)\right|^{p}\Big] + T^{p/2}\mathbb{E}\Big[\left(\int_{0}^{T} \left|\mathscr{Z}_{1}(s)\right|^{2}\right)^{p/2}\Big]\Big) \\ & \le \hat{K}\Big(1 + T^{p/2} + \frac{T^{p}}{1 - KT^{p/2}} + \left(T^{p} + T^{p/2}\frac{1 + T^{p}}{1 - KT^{p/2}}\right)\mathbb{E}\Big[\sup_{0 \le t \le T} \left|\mathscr{Y}(t)\right|^{p}\Big]\Big), \end{split}$$

where we use the Doob's martingale inequality, the Jensen's inequality and inequality (10.15). If *T* is sufficiently small, then we conclude that $\mathbb{E}[\sup_{0 \le t \le T} |\mathscr{Y}(t)|^p] < K_{T,p}$. To prove estimate (10.14) for an arbitrary *T*, we divide the interval [0, *T*] into sufficiently small subintervals $[T_i, T_{i+1}]$ and we consider the BSDEs

$$\begin{aligned} \mathscr{Y}(t) &= \mathscr{Y}(T_{i+1}) + \int_{t}^{T_{i+1}} \left(-\mathscr{Y}(s-)r(s) + H(s) + \int_{\mathbb{R}} G(s,z)Q(s,dz)\eta(s) \right. \\ &\left. - \mathscr{Z}_{1}(s)\theta(s) \right) ds - \int_{t}^{T_{i+1}} \mathscr{Z}_{1}(s)dW(s) \\ &\left. - \int_{t}^{T_{i+1}} \mathscr{Z}_{2}(s)dB(s) - \int_{t}^{T_{i+1}} \int_{\mathbb{R}} \mathscr{U}(s,z)\tilde{N}(ds,dz), \quad T_{i} \leq t \leq T_{i+1}. \end{aligned}$$

We can establish the assertion $\mathbb{E}[\sup_{T_i \leq t \leq T_{i+1}} |\mathscr{Y}(t)|^p] < K_{i,p}$ on each subinterval. The global estimate for \mathscr{Y} on [0, T] follows by combining a finite number of local estimates.

We now derive the solution the quadratic optimization problem (10.5).

Theorem 10.2.1 Assume that (C1)–(C4) from Chap. 7 hold, and let

$$\mathbb{E}\bigg[|F|^{p} + \left(\int_{0}^{T} |H(s)|^{2} ds\right)^{p/2} + \left(\int_{0}^{T} \int_{\mathbb{R}} |G(s,z)\eta(s)|^{2} Q(s,dz) ds\right)^{p/2}\bigg] < \infty,$$

for some p > 2. We consider the quadratic hedging problem (10.5). The strategy of the form

$$\pi^{*}(t) = \frac{\mathscr{Z}_{1}(t)}{\sigma(t)} - \left(\frac{\mu(t) - r(r)}{\sigma^{2}(t)} + \frac{Z(t)}{Y(t)\sigma(t)}\right) (X^{*}(t-) - \mathscr{Y}(t-)), \quad 0 \le t \le T, (10.16)$$

where (Y, Z) and $(\mathscr{Y}, \mathscr{Z})$ solve the BSDEs (10.6) and (10.7), and the process X^* is given by

$$dX^{*}(t) = \pi^{*}(t) (\mu(t)dt + \sigma(t)dW(t)) + (X^{*}(t-) - \pi^{*}(t))r(t)dt,$$

- $H(t)dt - \int_{\mathbb{R}} G(t,z)N(dt,dz), \quad X(0) = x > 0,$

are the optimal admissible hedging strategy $\pi^* \in \mathcal{A}$ and the optimal hedging portfolio for the payment process P.

Proof 1. The optimality. The proof is based on the method of completing the squares. Consider the BSDEs (10.6) and (10.7) written in the shorthand notation

$$\begin{split} dY(t) &= -f(t)dt + Z(t)dW(t), \quad Y(T) = 1, \\ d\mathscr{Y}(t) &= -f'(t)dt + \mathscr{Z}_1(t)dW(t) + \mathscr{Z}_2(t)dB(t) \\ &+ \int_{\mathbb{R}} \mathscr{U}(t,z)\tilde{N}(dt,dz), \quad \mathscr{Y}(T) = F, \end{split}$$

where the generators f and f' are appropriately defined. Key properties of the solutions to the BSDEs (10.6) and (10.7) are established in Propositions 10.2.1 and 10.2.2. We use these properties in the proof. We introduce the process

$$\mathscr{Y}(t) = -2Y(t)\mathscr{Y}(t), \quad 0 \le t \le T.$$

The Itô's formula yields the dynamics

$$d\hat{\mathscr{Y}}(t) = -\hat{f}(t)dt + \hat{\mathscr{Z}}_{1}(t)dW(t) + \hat{\mathscr{Z}}_{2}(t)dB(t) + \int_{\mathbb{R}}\hat{\mathscr{U}}(t,z)\tilde{N}(dt,dz),$$
$$\hat{\mathscr{Y}}(T) = -2F,$$

where

$$\hat{f}(t) = -2(Y(t)f'(t) + \mathscr{Y}(t)f(t) - \mathscr{Z}_{1}(t)Z(t)), \quad 0 \le t \le T,$$

$$\hat{\mathscr{Z}}_{1}(t) = -2(Y(t)\mathscr{Z}_{1}(t) + \mathscr{Y}(t)Z(t)), \quad 0 \le t \le T,$$

$$\hat{\mathscr{Z}}_{2}(t) = -2Y(t)\mathscr{Z}_{2}(t), \quad 0 \le t \le T,$$

$$\hat{\mathscr{Y}}(t, z) = -2Y(t)\mathscr{U}(t, z), \quad 0 \le t \le T,$$
(10.17)

and

$$\hat{f}(t) = \hat{\mathscr{Y}}(t)r(t) - 2Y(t)H(t) - 2Y(t) \int_{\mathbb{R}} G(t,z)Q(t,dz)\eta(t) - \theta(t)\frac{\hat{\mathscr{Y}}(t)(\mu(t) - r(t)) + \hat{\mathscr{Z}}_{1}(t)\sigma(t)}{\sigma(t)} - Z(t)\frac{\hat{\mathscr{Y}}(t)(\mu(t) - r(t)) + \hat{\mathscr{Z}}_{1}(t)\sigma(t)}{Y(t)\sigma(t)}, \quad 0 \le t \le T.$$
(10.18)

We choose an admissible strategy $\pi \in \mathscr{A}$ and we consider the investment portfolio process X^{π} given by (7.11). By the Itô's formula we derive the dynamics

$$\begin{aligned} d\big(Y(t)\big(X^{\pi}(t)\big)^{2}\big) \\ &= Y(t)\bigg(2X^{\pi}(t-)\pi(t)\big(\mu(t)dt + \sigma(t)dW(t)\big) \\ &+ 2X^{\pi}(t-)\big(X^{\pi}(t-) - \pi(t)\big)r(t)dt - 2X^{\pi}(t-)H(t)dt \\ &- \int_{\mathbb{R}} 2X^{\pi}(t-)G(t,z)N(dt,dz) + \big|\pi(t)\sigma(t)\big|^{2}dt + \int_{\mathbb{R}} \big|G(t,z)\big|^{2}N(dt,dz)\bigg) \\ &+ \big|X^{\pi}(t-)\big|^{2}\big(-f(t)dt + Z(t)dW(t)\big) + 2Z(t)X^{\pi}(t)\pi(t)\sigma(t)dt, \end{aligned}$$

and

$$\begin{split} d\big(\hat{\mathscr{Y}}(t)X^{\pi}(t)\big) \\ &= \hat{\mathscr{Y}}(t-) \Big(\pi(t)\big(\mu(t)dt + \sigma(t)dW(t)\big) \\ &+ \big(X^{\pi}(t-) - \pi(t)\big)r(t)dt - H(t)dt - \int_{\mathbb{R}} G(t,z)N(dt,dz)\Big) \\ &+ X^{\pi}(t-) \Big(-\hat{f}(t)dt + \hat{\mathscr{X}}_{1}(t)dW(t) + \hat{\mathscr{X}}_{2}(t)dB(t) + \int_{\mathbb{R}} \hat{\mathscr{U}}(t,z)\tilde{N}(dt,dz)\Big) \\ &+ \hat{\mathscr{X}}_{1}(t)\pi(t)\sigma(t)dt - \int_{\mathbb{R}} \hat{\mathscr{U}}(t,z)G(t,z)N(dt,dz). \end{split}$$

We can notice that the stochastic integrals with respect to W, B and \tilde{N} are locally square integrable local martingales since X^{π} and $\hat{\mathscr{Y}}$ are càdlàg, Y is bounded, π , $Z\hat{\mathscr{Z}_1}, \hat{\mathscr{Z}_2}$ and $\hat{\mathscr{U}}$ are square integrable, see Theorem 2.3.1. Furthermore, the stochastic integrals with respect to N have locally integrable compensators since X^{π} and $\hat{\mathscr{Y}}$ are càdlàg, Y is bounded, G and $\hat{\mathscr{U}}$ are square integrable. Consequently, the compensated integrals are local martingales, see Theorem 2.3.2. Let $(\tau_n)_{n\geq 1}$ denote a localizing sequence for the local martingales such that $\tau_n \to T$, $n \to \infty$. We obtain the expectations

$$\mathbb{E}[Y(\tau_{n})|X^{\pi}(\tau_{n})|^{2}]$$

$$= Y(0)x^{2} + \mathbb{E}\left[\int_{0}^{\tau_{n}} \left\{Y(t)\left(2X^{\pi}(t-)\pi(t)(\mu(t)-r(t))\right) + 2|X^{\pi}(t-)|^{2}r(t) - 2X^{\pi}(t-)H(t) - \int_{\mathbb{R}} 2X^{\pi}(t-)G(t,z)Q(t,dz)\eta(t) + |\pi(t)\sigma(t)|^{2} + \int_{\mathbb{R}} |G(t,z)|^{2}Q(t,dz)\eta(t)\right) - |X^{\pi}(t-)|^{2}f(t) + 2Z(t)X^{\pi}(t-)\pi(t)\sigma(t)\right\}dt\right],$$
(10.19)

and

$$\mathbb{E}[\hat{\mathscr{Y}}(\tau_{n})X^{\pi}(\tau_{n})] = \hat{\mathscr{Y}}(0)x + \mathbb{E}\left[\int_{0}^{\tau_{n}} \left\{\hat{\mathscr{Y}}(t-)\left(\pi(t)(\mu(t)-r(t)\right) + X^{\pi}(t-)r(t) - H(t) - \int_{\mathbb{R}} G(t,z)Q(t,dz)\eta(t)\right) - X^{\pi}(t-)\hat{f}(t) + \hat{\mathscr{Z}}_{1}(t)\pi(t)\sigma(t) - \int_{\mathbb{R}} G(t,z)\hat{\mathscr{U}}(t,z)Q(t,dz)\eta(t)\right\}dt\right].$$
(10.20)

From (10.19)–(10.20), after some tedious calculations, we deduce the formula

$$\begin{split} & \mathbb{E} \Big[Y(\tau_n) \Big| X^{\pi}(\tau_n) - \mathscr{Y}(\tau_n) \Big|^2 \Big] \\ &= \mathbb{E} \Big[Y(\tau_n) \Big| X^{\pi}(\tau_n) \Big|^2 + \hat{\mathscr{Y}}(\tau_n) X^{\pi}(\tau_n) + Y(\tau_n) \Big| \mathscr{Y}(\tau_n) \Big|^2 \Big] \\ &= Y(0) x^2 + \hat{\mathscr{Y}}(0) x \\ &\quad + \mathbb{E} \Big[\int_0^{\tau_n} Y(t) \Big| \sigma(t) \Big|^2 \Big\{ \pi(t) + \left(\frac{\mu(t) - r(t)}{\sigma^2(t)} + \frac{Z(t)}{Y(t)\sigma(t)} \right) X^{\pi}(t-) \\ &\quad + \frac{\hat{\mathscr{Y}}(t)}{2Y(t)} \frac{\mu(t) - r(t)}{\sigma^2(t)} + \frac{\hat{\mathscr{X}}_1(t)}{2Y(t)\sigma(t)} \Big\}^2 dt \end{split}$$

$$\begin{split} &+ \int_{0}^{\tau_{n}} \left| X^{\pi}(t) \right|^{2} \bigg(-f(t) + 2Y(t)r(t) - \frac{|Z(t)|^{2}}{Y(t)} - |\theta(t)|^{2}Y(t) - 2\theta(t)Z(t) \bigg) dt \\ &+ \int_{0}^{\tau_{n}} X^{\pi}(t) \bigg(-\hat{f}(t) + \hat{\mathscr{Y}}(t)r(t) - 2Y(t)H(t) - 2Y(t) \int_{\mathbb{R}} G(t,z)Q(t,dz)\eta(t) \\ &- \theta(t) \frac{\hat{\mathscr{Y}}(t)(\mu(t) - r(t)) + \hat{\mathscr{Z}}_{1}(t)\sigma(t)}{\sigma(t)} \\ &- Z(t) \frac{\hat{\mathscr{Y}}(t)(\mu(t) - r(t)) + \hat{\mathscr{Z}}_{1}(t)\sigma(t)}{Y(t)\sigma(t)} \bigg) dt \bigg] \\ &+ \mathbb{E} \bigg[Y(\tau_{n}) |\mathscr{Y}(\tau_{n})|^{2} - \int_{0}^{\tau_{n}} \bigg| \frac{\hat{\mathscr{Y}}(t)}{2Y(t)} \frac{\mu(t) - r(t)}{\sigma^{2}(t)} + \frac{\hat{\mathscr{Z}}_{1}(t)}{2Y(t)\sigma(t)} \bigg|^{2} Y(t) |\sigma(t)|^{2} dt \\ &+ \int_{0}^{\tau_{n}} \bigg(Y(t) \int_{\mathbb{R}} |G(t,z)|^{2} Q(t,dz)\eta(t) - \hat{\mathscr{Y}}(t-)H(t) \\ &- \hat{\mathscr{Y}}(t-) \int_{\mathbb{R}} G(t,z)Q(t,dz)\eta(t) - \int_{\mathbb{R}} G(t,z)\hat{\mathscr{U}}(t,z)Q(t,dz)\eta(t) \bigg) dt \bigg], \end{split}$$

and by the definition of the generators f and \hat{f} from (10.6) and (10.18) we obtain

$$\mathbb{E}[Y(\tau_{n})|X^{\pi}(\tau_{n}) - \mathscr{Y}(\tau_{n})|^{2}]$$

$$= Y(0)x^{2} + \hat{\mathscr{Y}}(0)x$$

$$+ \mathbb{E}\left[\int_{0}^{\tau_{n}} Y(t)|\sigma(t)|^{2}\left\{\pi(t) + \left(\frac{\mu(t) - r(t)}{\sigma^{2}(t)} + \frac{Z(t)}{Y(t)\sigma(t)}\right)X^{\pi}(t-)\right.$$

$$\left. + \frac{\hat{\mathscr{Y}}(t)}{2Y(t)}\frac{\mu(t) - r(t)}{\sigma^{2}(t)} + \frac{\hat{\mathscr{Y}}_{1}(t)}{2Y(t)\sigma(t)}\right\}^{2}dt\right]$$

$$\left. + \mathbb{E}\left[Y(\tau_{n})|\mathscr{Y}(\tau_{n})|^{2} + \int_{0}^{\tau_{n}}\varphi(t)dt\right], \qquad (10.21)$$

where the process φ collects all terms independent of π . We let $n \to \infty$. We have to justify the interchange of the limit $n \to \infty$ and the expectation. First, we recall that *Y* is uniformly bounded, $\mathbb{E}[\sup_{t\in[0,T]} |\mathscr{Y}(t)|^2] < \infty$ since \mathscr{Y} solves (10.7) and $\mathbb{E}[\sup_{t\in[0,T]} |X^{\pi}(t)|^2] < \infty$ by (7.12). From these properties we can conclude that $Y(\tau_n)|X^{\pi}(\tau_n)|^2$, $\hat{\mathscr{Y}}(\tau_n)X^{\pi}(\tau_n)$ and $Y(\tau_n)|\mathscr{Y}(\tau_n)|^2$ are bounded uniformly in *n* by an integrable random variable. Next, we deduce from the moment estimates (10.12) and (10.14) that the process $\hat{\mathscr{X}}_1$ is square integrable. Finally, under our assumptions the process φ is integrable i.e. $\mathbb{E}[\int_0^T |\varphi(t)| dt] < \infty$. Hence, we apply the monotone convergence theorem to the first expectation on the right hand side of (10.21), the dominated convergence theorem to the second expectation on the right hand side of (10.21) and the dominated convergence theorem on the left hand side of (10.21). We end up with

$$\mathbb{E}[|X^{\pi}(T) - F|^{2}] = \lim_{n \to \infty} \mathbb{E}[Y(\tau_{n})|X^{\pi}(\tau_{n}) - \mathscr{Y}(\tau_{n})|^{2}] = Y(0)x^{2} + \widehat{\mathscr{Y}}(0)x + \mathbb{E}\left[\int_{0}^{T} Y(t)|\sigma(t)|^{2}\left\{\pi(t) + \left(\frac{\mu(t) - r(t)}{\sigma^{2}(t)} + \frac{Z(t)}{Y(t)\sigma(t)}\right)X^{\pi}(t-) + \frac{\widehat{\mathscr{Y}}(t)}{2Y(t)}\frac{\mu(t) - r(t)}{\sigma^{2}(t)} + \frac{\widehat{\mathscr{X}}_{1}(t)}{2Y(t)\sigma(t)}\right\}^{2}dt\right] + \mathbb{E}\left[F^{2} - \int_{0}^{T} \varphi(t)dt\right].$$
 (10.22)

Since the last term in (10.22) does not depend on π , the optimal strategy can be immediately derived from (10.22) and (10.17).

2. The admissability. Let $\pi^*(t) = A(t)X^{\pi^*}(t-) + B(t)$. The dynamics of the investment portfolio (7.11) under the candidate strategy (10.16) is given by

$$dX^{\pi^{*}}(t) = X^{\pi^{*}}(t-)(r(t) + (\mu(t) - r(t))A(t))dt + X^{\pi^{*}}(t-)A(t)\sigma(t)dW(t) + B(t)(\mu(t) - r(t))dt + B(t)\sigma(t)dW(t) - H(t)dt - \int_{\mathbb{R}} G(t,z)N(dt,dz).$$
(10.23)

Since (10.23) is a linear forward SDE, there exists a unique càdlàg, adapted solution X^{π^*} to (10.23), see Theorem V.7 in Protter (2004). We can now conclude that π^* is a predictable process.

We are left with proving the square integrability of the strategy π^* . From (10.21) and integrability of the processes *Y*, \mathscr{Y} and φ we can deduce the uniform estimate

$$\mathbb{E}\left[Y(\tau_n)\Big|X^{\pi^*}(\tau_n) - \mathscr{Y}(\tau_n)\Big|^2\right]$$

= $Y(0)x^2 + \hat{\mathscr{Y}}(0)x + \mathbb{E}\left[Y(\tau_n)\Big|\mathscr{Y}(\tau_n)\Big|^2 + \int_0^{\tau_n} \varphi(t)dt\right] \leq K,$

and by Fatou's lemma we get

$$K \geq \lim_{n \to \infty} \mathbb{E} \Big[Y(\tau_n \wedge t) \Big| X^{\pi^*}(\tau_n \wedge t) - \mathscr{Y}(\tau_n \wedge t) \Big|^2 \Big]$$

$$\geq \mathbb{E} \Big[Y(t) \Big| X^{\pi^*}(t) - \mathscr{Y}(t) \Big|^2 \Big], \quad 0 \leq t \leq T.$$

Consequently, we can prove square integrability of the investment portfolio by noticing that

$$\mathbb{E}\left[\left|X^{\pi^{*}}(t)\right|^{2}\right] \leq 2\mathbb{E}\left[\left|X^{\pi^{*}}(t)-\mathscr{Y}(t)\right|^{2}+\left|\mathscr{Y}(t)\right|^{2}\right]$$
$$\leq 2\mathbb{E}\left[\frac{1}{k}Y(t)\left|X^{\pi^{*}}(t)-\mathscr{Y}(t)\right|^{2}+\left|\mathscr{Y}(t)\right|^{2}\right] \leq K, \quad 0 \leq t \leq T,$$

where we use the lower bound $k \leq Y(t)$. Consider now the BSDE

$$\bar{Y}(t) = X^{\pi^*}(T) + \int_t^T \left(-\bar{Y}(s-)r(s) - \bar{Z}(s)\theta(s)\right) ds$$

+
$$\int_t^T H(s)ds + \int_t^T \int_{\mathbb{R}} G(s,z)N(ds,dz)$$

-
$$\int_t^T \bar{Z}(s)dW(s) - \int_t^T \int_{\mathbb{R}} \bar{U}(s,z)\tilde{N}(ds,dz), \quad 0 \le t \le T. \quad (10.24)$$

Since $\mathbb{E}[|X^{\pi^*}(T)|^2] < \infty$, there exists a unique solution $(\bar{Y}, \bar{Z}, \bar{U}) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2_N(\mathbb{R})$ to the BSDE (10.24). By (10.23) we get $\bar{Y}(t) = X^{\pi^*}(t), \bar{Z}(t) = \pi^*(t)\sigma(t), \bar{U}(t, z) = 0$. The admissibility of π^* has been proved. \Box

We remark that the arguments from the proof of Theorem 10.2.1 can be applied to solve a general linear quadratic control problem, see Lim (2004), Lim (2005) and Øksendal and Hu (2008).

Equation (10.22) gives the minimal hedging error under the optimal strategy π^* . We can easily find the optimal initial capital.

Proposition 10.2.3 Under the assumptions of Propositions 10.2.1, 10.2.2 and Theorem 10.2.1 the initial value of the investment portfolio which minimizes the quadratic hedging error (10.5) takes the form

$$x^* = \mathscr{Y}(0) = \mathbb{E}^{\mathbb{Q}^*} \left[\int_0^T e^{-\int_0^s r(u) du} dP(s) \right],$$
 (10.25)

where the equivalent martingale measure \mathbb{Q}^* is given by

$$\frac{d\mathbb{Q}^*}{d\mathbb{P}}\Big|\mathscr{F}_t = e^{-\int_0^t \theta(s)dW(s) - \frac{1}{2}\int_0^t |\theta(s)|^2 ds}, \quad 0 \le t \le T.$$

The price and the hedging strategy for the payment process *P* can be obtained by solving the nonlinear BSDE (10.6) and the linear BSDE (10.7). The price (10.25), the initial value of the hedging portfolio, is arbitrage-free. We notice that the risk premiums for the systematic insurance risk and the unsystematic insurance risk are equal to zero and the insurance risk is not priced under the measure \mathbb{Q}^* . In contrast to Sect. 10.1 where the pricing measure were assumed to be given, here the pricing measure \mathbb{Q}^* is derived by solving the optimization problem (10.5). We remark that the pricing measure which arises from solving the quadratic pricing and hedging problem (10.5) is called a minimal variance measure, see Lim (2004) and Schweizer (2010). The optimal hedging strategy (10.16) consists of two terms: the first term is a delta hedging strategy and the second term is a correction factor taking into account discrepancies between the optimal hedging portfolio *X*^{*} and the market-consistent value \mathscr{Y} of the insurance payment process (the reserve required for *P*). The second term adjusts the delta hedging strategy according to the shortfall or the surplus of the assets over the liabilities.

Example 10.2 We consider the classical Black-Scholes model with constant coefficients r, μ , σ and an insurer who issues a portfolio of n unit-linked endowment policies with a capital guarantee. We are again interested in pricing and hedging the claim $F = (n - J(T))(K - S(T))^+$ where J is the deaths counting process for the insurance portfolio. We consider a constant mortality intensity λ . In order to find the price and the optimal hedging strategy for the claim F, we have to solve the BSDEs

$$Y(t) = 1 + \int_{t}^{T} \left(2Y(s)r - \frac{|Z(s)|^{2}}{Y(s)} - \theta^{2}Y(s) - 2\theta Z(s) \right) ds$$

- $\int_{t}^{T} Z(s) dW(s), \quad 0 \le t \le T,$ (10.26)

$$\mathscr{Y}(t) = (n - J(T))(K - S(T))^{+} + \int_{t}^{T} (-\mathscr{Y}(s)r - \mathscr{Z}(s)\theta)ds$$
$$-\int_{t}^{T} \mathscr{Z}(s)dW(s) - \int_{t}^{T} \mathscr{U}(s)\tilde{N}(ds), \quad 0 \le t \le T.$$
(10.27)

It is straightforward to notice that the solution to the BSDE (10.26) is of the form

$$Y(t) = e^{(2r - \theta^2)(T - t)}, \quad 0 \le t \le T,$$

 $Z(t) = 0, \quad 0 \le t \le T.$

Recalling the results from Example 9.5, we conclude that the solution to the BSDE (10.27) is given by the triple

$$\begin{aligned} \mathscr{Y}(t) &= \left(n - J(t)\right)e^{-\lambda(T-t)} \\ &\cdot \left(Ke^{-r(T-t)}\Phi\left(-d\left(t,S(t)\right) - \sigma\sqrt{T-t}\right) - S(t)\Phi\left(-d\left(t,S(t)\right)\right)\right), \\ &0 \le t \le T, \\ \mathscr{Z}(t) &= -\left(n - J(t-)\right)e^{-\lambda(T-t)}\sigma S(t)\Phi\left(-d\left(t,S(t)\right)\right), \quad 0 \le t \le T, \\ \mathscr{U}(t) &= -e^{-\lambda(T-t)} \\ &\cdot \left(Ke^{-r(T-t)}\Phi\left(-d\left(t,S(t)\right) - \sigma\sqrt{T-t}\right) - S(t)\Phi\left(-d\left(t,S(t)\right)\right)\right), \\ &0 \le t \le T. \end{aligned}$$

By Theorem 10.2.1 and Proposition 10.2.3 the optimal hedging strategy is given by the feedback formula

$$\pi^*(t) = -(n - J(t-))e^{-\lambda(T-t)}S(t)\Phi\left(-d(t, S(t))\right)$$
$$-\frac{\mu - r}{\sigma^2}\left(X^*(t-) - \mathscr{Y}(t-)\right), \quad 0 \le t \le T,$$

where the optimal hedging portfolio is determined by the process

$$X^{*}(t) = x^{*}e^{rt} + \int_{0}^{t} \pi^{*}(s) \big(\theta \sigma ds + \sigma dW(s)\big), \quad X^{*}(0) = x^{*},$$

and the price of the claim, and the initial value of the optimal hedging portfolio, is equal to

$$x^* = \mathscr{Y}(0) = ne^{-\lambda T} \left(Ke^{-rT} \Phi\left(-d\left(0, S(0)\right) - \sigma\sqrt{T} \right) - S(0)\Phi\left(-d\left(0, S(0)\right) \right) \right).$$

The quadratic objective (10.5) can be generalized by keeping its mathematical tractability. We can solve the quadratic optimization problem

$$\min_{\pi} \mathbb{E} \left[\int_0^T \left(X^{\pi}(s) - \beta(s) \right)^2 ds + \left(X^{\pi}(T) - \xi \right)^2 \right],$$
(10.28)

where a running cost is added. Under the objective (10.28) an investment strategy is chosen in such a way that the investment portfolio is as close as possible, in the mean-square sense, to the targets β and ξ . The targets β and ξ may represent solvency constraints or profit expectations, see Detemple and Rindisbacher (2008) and Delong (2010). We can also solve the Markowitz portfolio selection problem

$$\min_{\pi} \operatorname{Var} \left[X^{\pi}(T) \right]$$

$$\mathbb{E} \left[X^{\pi}(T) \right] = L,$$
(10.29)

where the variance of the terminal wealth is minimized given the constraint on the expected return. We remark that the Markowitz portfolio selection is commonly used for investment decision making and asset-liability management, see Chap. 4 in Zenios and Ziemba (2006) and Sect. 23.2 in Adam (2007).

Notice that in order to apply the investment strategy (10.16) we have to estimate the drift μ . The advantage of the quadratic optimization under an equivalent martingale measure (10.1) over the quadratic optimization under the real-world measure (10.5) is that the drift of the stock does not enter the optimal strategy in the former case. It is known that the estimation of the drift is challenging. Consequently, the quadratic pricing and hedging under the real-world measure may be difficult to implement in practice, see Sect. 10.4.3 in Cont and Tankov (2004) for a discussion.

10.3 Quadratic Pricing and Hedging Under Local Risk-Minimization

We now investigate the objective of local risk-minimization which was developed in Schweizer (1991) and Schweizer (2008). Local risk-minimization is an important alternative to global quadratic hedging. We will see that the locally risk-minimizing strategy is derived by optimizing an objective formulated under the real-world measure but at the same time the optimal strategy does not depend on the drift of the stock.

Let $\hat{S}(t) = e^{-\int_0^t r(s)ds} S(t)$. The discounted stock price satisfies the dynamics

$$\hat{S}(t) = \hat{S}(0) + \int_0^t \hat{S}(u) \big(\mu(u) - r(u) \big) du + \int_0^t \hat{S}(u) \sigma(u) dW(u), \quad 0 \le t \le T.$$

We introduce a class of admissible investment strategies.

Definition 10.3.1 A strategy $\Pi := (\Pi(t), 0 \le t \le T)$ is called admissible, written $\Pi \in \mathscr{A}_s^{local}$, if it satisfies the conditions:

- (i) $\Pi: \Omega \times [0, T] \to \mathbb{R}$ is a predictable process,
- (ii) $\mathbb{E}[\int_0^T |\hat{S}(t)\Pi(t)\sigma(t)|^2 dt] < \infty.$

A strategy $\Gamma := (\Gamma(t), 0 \le t \le T)$ is called admissible, written $\Gamma \in \mathscr{A}_b^{local}$, if it satisfies the conditions:

- (i) $\Gamma: \Omega \times [0, T] \to \mathbb{R}$ is an adapted process,
- (ii) the process $\mathscr{X}^{\Pi,\Gamma}(t) = \Pi(t)\hat{S}(t) + \Gamma(t)$ is right-continuous and square integrable.

A strategy (Π, λ) is called admissible for the local risk-minimization problem, written $(\Pi, \lambda) \in \mathscr{A}^{local}$, if $\Pi \in \mathscr{A}^{local}_{s}$ and $\Gamma \in \mathscr{A}^{local}_{b}$

We remark that the strategy Π denotes the number of stock which are held in the investment portfolio, Γ denotes the position in the bank account and \mathscr{X} is the discounted value of the investment portfolio. We point out that the investment portfolio $\mathscr{X}^{\Pi,\Gamma}$ is not self-financing under $(\Pi, \Gamma) \in \mathscr{A}^{local}$.

We define the cost process and the risk process of a hedging strategy.

Definition 10.3.2 Assume that (C1)–(C4) from Chap. 7 hold. The cost process of an admissible strategy $(\Pi, \Gamma) \in \mathscr{A}^{local}$ related to hedging the payment process *P* is given by

$$C^{\Pi,\Gamma}(t) = \int_0^t e^{-\int_0^s r(u)du} dP(s) + \mathcal{X}^{\Pi,\Gamma}(t) - \int_0^t \Pi(s)d\hat{S}(s), \quad 0 \le t \le T.$$

The risk process of an admissible strategy $(\Pi, \Gamma) \in \mathscr{A}^{local}$ related to hedging the payment process *P* is given by

$$R^{\Pi,\Gamma}(t) = \mathbb{E}\left[\left|C^{\Pi,\Gamma}(T) - C^{\Pi,\Gamma}(t)\right|^2 |\mathscr{F}_t\right], \quad 0 \le t \le T.$$

Since $(\Pi, \Gamma) \in \mathscr{A}^{local}$, the integral $\int_0^t \Pi(s) d\hat{S}(s)$ is well-defined and $C^{\Pi, \Gamma}$ is square integrable.

The cost process $C^{\Pi,\Gamma}$ describes accumulated discounted costs or profits (cash inflows or outflows) for the insurer who applies an investment strategy (Π, Γ) , pays

the benefits *P* and holds the investment portfolio $\mathscr{X}^{\Pi,\Gamma}$. If the payment process can be perfectly replicated by (Π, Γ) , then

$$\mathcal{X}^{\Pi,\Gamma}(t) = \mathcal{X}(0) + \int_0^t \Pi(s) d\hat{S}(s) - \int_0^t e^{-\int_0^s r(u) du} dP(s), \quad 0 \le t \le T,$$

$$\mathcal{X}^{\Pi,\Gamma}(T) = 0,$$

and the cost process related to hedging *P* is equal to the initial premium $\mathscr{X}(0)$ - the cost of setting the replicating portfolio which matches the liability. If the payment process cannot be perfectly replicated, then the insurer has to inject capital or withdraw capital during the lifetime of the policy in order to match the assets with the liabilities. These inflows and outflows of capital from the investment portfolio are modelled by the cost process.

The idea of local risk-minimization is to find an admissible hedging strategy $(\Pi, \Gamma) \in \mathscr{A}^{local}$ which minimizes the risk process $R^{\Pi,\Gamma}$ for all $t \in [0, T]$. The precise definition of the hedging objective is very technical and it involves limit considerations and local perturbations of investment strategies, see Schweizer (1991) and Schweizer (2008). We can say that under the local risk-minimization we aim to find an asset portfolio which perfectly matches the liability with a minimal mean-square cost of matching. We point out that the hedging objective, the risk process, is evaluated under the real-world measure, as it should be.

We give the key result which characterizes a locally risk-minimizing strategy.

Theorem 10.3.1 Assume that (C1)–(C4) from Chap. 7 hold. The payment process P admits an admissible locally risk-minimizing strategy $(\Pi^*, \Gamma^*) \in \mathscr{A}^{local}$ if and only if the discounted payment process has the representation

$$\int_0^T e^{-\int_0^s r(u)du} dP(s) = \xi_0 + \int_0^T \zeta(s)d\hat{S}(s) + \mathscr{H}(T),$$
(10.30)

where $\zeta \in \mathscr{A}_s^{local}$, and \mathscr{H} is a right-continuous, square integrable martingale which is strongly orthogonal to the martingale component of \hat{S} and verifies $\mathscr{H}(0) = 0$. In this case, we define

$$\Pi^*(t) = \zeta(t), \quad 0 \le t \le T,$$

$$\mathscr{X}^*(t) = \xi_0 + \int_0^t \zeta(s) d\hat{S}(s) + \mathscr{H}(t) - \int_0^t e^{-\int_0^s r(u) du} dP(s), \quad 0 \le t \le T,$$

$$\Gamma^*(t) = \mathscr{X}^*(t) - \Pi^*(t)\hat{S}(t), \quad 0 \le t \le T.$$

Proof The result follows from Proposition 5.2 from Schweizer (2008).

We remark that two square integrable martingales are strongly orthogonal if $\mathbb{E}[M^1(t)M^2(t)] = 0.$

Decomposition (10.30) is called the Föllmer-Schweizer decomposition. The locally risk-minimizing strategy (Π^* , Γ^*) is called mean-self-financing as the optimal cost process $C^*(t) = \xi_0 + \mathcal{H}(t)$ is a martingale, i.e. the average future cost is equal to zero. The strategy (Π^* , Γ^*) is also called 0-achieving as $\mathcal{X}^*(T) = 0$ and all claims are covered.

The next result shows that the Föllmer-Schweizer decomposition can be derived from a BSDE.

Proposition 10.3.1 Assume that (C1)–(C4) from Chap. 7 hold.

(a) The Föllmer-Schweizer decomposition of the discounted payment process is given by

$$\int_{0}^{T} e^{-\int_{0}^{s} r(u)du} dP(s)$$

= $Y(0) + \int_{0}^{T} \frac{Z_{1}(s)}{\sigma(s)\hat{S}(s)} d\hat{S}(s)$
+ $\int_{0}^{T} Z_{2}(s) dB(s) + \int_{0}^{T} \int_{\mathbb{R}} U(s, z) \tilde{N}(ds, dz)$

where $(Y, Z_1, Z_2, U) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2_N(\mathbb{R})$ is the unique solution to the BSDE

$$Y(t) = \int_{0}^{T} e^{-\int_{0}^{s} r(u)du} dP(s) - \int_{t}^{T} Z_{1}(s)\theta(s)ds$$

- $\int_{t}^{T} Z_{1}(s)dW(s) - \int_{t}^{T} Z_{2}(s)dB(s)$
- $\int_{t}^{T} \int_{\mathbb{R}} U(s, z)\tilde{N}(ds, dz), \quad 0 \le t \le T.$ (10.31)

(b) The admissible locally risk-minimizing strategy $\Pi^* \in \mathscr{A}^{local}$ and the optimal discounted hedging portfolio for the payment process P are given by

$$\Pi^*(t) = \frac{Z_1(t)}{\sigma(t)\hat{S}(t)}, \quad 0 \le t \le T,$$

$$\mathscr{X}^*(t) = Y(t) - \int_0^t e^{-\int_0^s r(u)du} dP(s), \quad 0 \le t \le T.$$

Proof (a) Consider the BSDE

$$Y(t) = \int_0^T e^{-\int_0^s r(u)du} dP(s) + \int_t^T f(s)ds - \int_t^T Z_1(s)dW(s) - \int_t^T Z_2(s)dB(s) - \int_t^T \int_{\mathbb{R}} U(s,z)\tilde{N}(ds,dz), \quad 0 \le t \le T, \quad (10.32)$$

where the generator f will be specified in the sequel. Assume there exists a solution $(Y, Z_1, Z_2, U) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2_N(\mathbb{R})$ to (10.32). From the BSDE (10.32) we get the representation of the discounted payments

$$\int_{0}^{T} e^{-\int_{0}^{s} r(u)du} dP(s)$$

= $Y(0) - \int_{0}^{T} f(s)ds + \int_{0}^{T} Z_{1}(s)dW(s)$
+ $\int_{0}^{T} Z_{2}(s)dB(s) + \int_{0}^{T} \int_{\mathbb{R}} U(s,z)\tilde{N}(ds,dz), \quad 0 \le t \le T, \quad (10.33)$

From property (5.5) we deduce that the stochastic integrals driven by *B* and \tilde{N} are strongly orthogonal to the martingale component of \hat{S} , i.e. the stochastic integral driven by *W*. The discounted payment process (10.33) would now satisfy the Föllmer-Schweizer decomposition (10.30) if for some $\zeta \in \mathscr{A}_{s}^{local}$ we had

$$\int_0^T \zeta(s) d\hat{S}(s) = \int_0^T Z_1(s) dW(s) - \int_0^T f(s) ds.$$
(10.34)

Since

$$\int_0^T \zeta(s) d\hat{S}(s) = \int_0^T \zeta(s) \hat{S}(s) \big(\mu(s) - r(s)\big) ds + \int_0^T \zeta(s) \hat{S}(s) \sigma(s) dW(s),$$

we should choose

$$f^*(s) = -Z_1(s)\theta(s), \quad 0 \le s \le T$$

$$\zeta^*(s) = \frac{Z_1(s)}{\sigma(s)\hat{S}(s)}, \quad 0 \le s \le T.$$

With this choice, there exists a unique solution $(Y, Z_1, Z_2, U) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2_N(\mathbb{R})$ to (10.32) and $\zeta^* \in \mathscr{A}_s^{local}$. From (10.33) and (10.34) we deduce the Föllmer-Schweizer decomposition.

(b) The formulas for Π^* and \mathscr{X}^* follow from item (a), (10.32)–(10.34) and Theorem 10.3.1.

We state an important corollary.

Corollary 10.3.1 Under the assumptions of Theorem 10.3.1 and Proposition 10.3.1 the optimal hedging portfolio process has the representation

$$X^*(t) = \mathbb{E}^{\mathbb{Q}^*}\left[\int_t^T e^{-\int_t^s r(u)du} dP(s)|\mathscr{F}_t\right], \quad 0 \le t \le T,$$
(10.35)

where

$$\frac{d\mathbb{Q}^*}{d\mathbb{P}}\Big|\mathscr{F}_t = e^{-\int_0^t \theta(s)dW(s) - \frac{1}{2}\int_0^t |\theta(s)|^2 ds}, \quad 0 \le t \le T,$$

and the optimal amount invested in the stock is given by

$$\pi^*(t) = \frac{Z_1(t)}{\sigma(t)} e^{\int_0^t r(s)ds}, \quad 0 \le t \le T.$$
(10.36)

Proof Since $X^*(t) = \mathscr{X}^*(t)e^{\int_0^t r(s)ds}$, the representation of X^* follows from Proposition 10.3.1 and the representation of the solution *Y* to the linear BSDE (10.31). Since $\pi^*(t) = \Pi^*(t)S(t)$, the formula for π^* follows immediately from Proposition 10.3.1.

The hedging portfolio (10.35) and the optimal amount invested in the stock (10.36) can be obtained by solving the linear BSDE (10.31). The price of the insurance payment process, which is the initial value of the hedging portfolio, is arbitrage-free. The insurance risk is not priced under the measure \mathbb{Q}^* . The hedging strategy π^* is a delta-hedging strategy which is updated with the current information on the financial and the insurance risk. We notice that the optimal hedging portfolio and the optimal hedging strategy are characterized under the equivalent martingale measure \mathbb{Q}^* , hence the drift of the stock does not enter the solution. We point out that under the locally risk-minimizing strategy (Π^* , Γ^*) there is no mismatch between the assets and the liabilities. The hedging portfolio is forced to match the market-consistent value of the liabilities, see Corollary 10.3.1. This equivalence holds since the investment portfolio process \mathscr{X}^* is not self-financing. By Proposition 10.3.1 we obtain

$$\begin{split} dX^*(t) &= r(t)e^{\int_0^t r(s)ds} \mathscr{X}^*(t)dt + e^{\int_0^t r(s)ds}d\mathscr{X}^*(t) \\ &= X^*(t)r(t)dt + e^{\int_0^t r(s)ds} \big(dY(t) - e^{-\int_0^t r(s)ds}dP(t)\big) \\ &= X^*(t)r(t)dt \\ &+ e^{\int_0^t r(s)ds} \big(\Pi^*(t)\hat{S}(t)\big(\mu(t) - r(t)\big)dt + \Pi^*(t)\hat{S}(t)\sigma(t)dW(t)\big) \\ &- dP(t) + e^{\int_0^t r(s)ds}Z_2(t)dB(t) + \int_{\mathbb{R}} e^{\int_0^t r(s)ds}U(t,z)\tilde{N}(dt,dz). \end{split}$$

Using π^* , we derive the dynamics of the hedging portfolio

$$dX^{*}(t) = \pi^{*}(t) (\mu(t)dt + \sigma(t)dW(t)) + (X^{*}(t) - \pi^{*}(t))r(t)dt$$

- $dP(t) + e^{\int_{0}^{t} r(s)ds} Z_{2}(t)dB(t) + \int_{\mathbb{R}} e^{\int_{0}^{t} r(s)ds} U(t,z)\tilde{N}(dt,dz).$
(10.37)

The stochastic integrals driven by *B* and \tilde{N} are interpreted as cash inflows/outflows which guarantee that $X^*(t) = \mathbb{E}^{\mathbb{Q}^*}[\int_t^T e^{-\int_t^s r(u)du} dP(s)|\mathscr{F}_t]$ for all $0 \le t \le T$. As already noticed, the expected value of the cash inflows/outflows is zero

$$\mathbb{E}\left[\int_0^T e^{\int_0^t r(s)ds} Z_2(t)dB(t) + \int_0^T \int_{\mathbb{R}} e^{\int_0^t r(s)ds} U(t,z)\tilde{N}(dt,dz)\right] = 0$$

and the hedging portfolio (10.37) is mean-self-financing. We remark that the locally risk-minimizing strategy (10.36) for the stock would coincide with the minimum mean-square error strategy for the stock derived in Theorem 10.1.1 if the measure \mathbb{Q}^* from Corollary 10.3.1 were chosen in the global quadratic hedging problem. However, the investment strategy from Theorem 10.1.1 is obtained in the framework of self-financing portfolios and $X^*(t) \neq \mathbb{E}^{\mathbb{Q}^*}[\int_t^T e^{-\int_t^s r(u)du} dP(s)|\mathscr{F}_t]$ except at the inception of the contract.

10.4 Quadratic Pricing and Hedging Under an Instantaneous Mean-Variance Risk Measure

As discussed at the end of Sect. 10.2, the Markowitz mean-variance objective is often used by financial institutions. In this chapter we investigate a local version of the mean-variance objective.

In Sects. 10.1–10.2 where we were only interested in finding the price of the payment process P at the initial time t = 0. We are now interested in dynamic pricing of the payment process P over the period [0, T]. We assume that the price Y of the payment process P solves the BSDE

$$Y(t) = F + \int_{t}^{T} H(s)ds + \int_{t}^{T} \int_{\mathbb{R}} G(s, z)N(ds, dz) + \int_{t}^{T} f(s)ds$$
$$- \int_{t}^{T} Z_{1}(s)dW(s) - \int_{t}^{T} Z_{2}(s)dB(s)$$
$$- \int_{t}^{T} \int_{\mathbb{R}} U(s, z)\tilde{N}(ds, dz), \quad 0 \le t \le T,$$
(10.38)

with a generator *f*. By Propositions 3.3.1 and 3.4.1 an arbitrage-free price of *P*, i.e. $Y(t) = \mathbb{E}^{\mathbb{Q}}[\int_{t}^{T} e^{-\int_{t}^{s} r(u)du} dP(s)|\mathscr{F}_{t}], \mathbb{Q} \in \mathscr{Q}^{m}$, has to solve a BSDE of the form (10.38). Hence, it is reasonable to assume a priori that the price process satisfies (10.38). In the sequel we use a local mean-variance objective to derive the generator *f* of the BSDE (10.38), the price process and the hedging strategy.

We recall that the investment portfolio X^{π} under an admissible investment strategy $\pi \in \mathscr{A}$ is given by (7.11). We assume that there exists a solution $(Y, Z_1, Z_2, U) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2_N(\mathbb{R})$ to the BSDE (10.38). We introduce a surplus process $\mathscr{S} := (\mathscr{S}(t), 0 \le t \le T)$ which models the excess of the wealth of the insurer over the price of the payment process. We set

$$\mathscr{S}(t) = X^{\pi}(t) - Y(t), \quad 0 \le t \le T.$$

We can derive the dynamics of the surplus process. We get the equation

$$d\mathscr{S}(t) = X^{\pi}(t-)r(t)dt + \pi(t)(\mu(t) - r(t))dt + f(t)dt + (\sigma(t)\pi(t) - Z_1(t))dW(t) - Z_2(t)dB(t) - \int_{\mathbb{R}} U(t,z)\tilde{N}(dt,dz),$$
(10.39)

with the initial condition $\mathscr{S}(0) = 0$, since we should have X(0) = Y(0). We remark that \mathscr{S} is square integrable for any $\pi \in \mathscr{A}$. Since *Y* can be interpreted as the marketconsistent reserve for the liabilities, the surplus process \mathscr{S} models the excess of the assets over the liabilities—the net asset wealth. The surplus process \mathscr{S} models the profit earned by the insurer. We recall that the net asset wealth is the key object investigated in Solvency II Directive, see European Commission QIS5 (2010). In Leitner (2007) the process \mathscr{S} is called a tracking error.

We define the mean-variance Markowitz risk measure

$$\rho(\xi) = L\sqrt{\operatorname{Var}[\xi]} - \mathbb{E}[\xi], \qquad (10.40)$$

where the parameter L is a risk aversion coefficient which sets the trade-off between variance minimization and expected return maximization. Following Leitner (2007), we apply the risk measure (10.40), with a time-varying risk aversion coefficient L, to the infinitesimal change in the surplus process \mathscr{S} . We investigate the instantaneous mean-variance risk measures

$$\rho(d\mathscr{S}(t))/dt$$

$$= L(t)\sqrt{\mathbb{E}\left[d[\mathscr{S},\mathscr{S}](t)|\mathscr{F}_{t-}\right]/dt} - \mathbb{E}\left[d\mathscr{S}(t) - \mathscr{S}(t-)r(t)dt|\mathscr{F}_{t-}\right]/dt}$$

$$= L(t)\sqrt{\left|\pi(t)\sigma(t) - Z_1(t)\right|^2 + \left|Z_2(t)\right|^2 + \int_{\mathbb{R}}\left|U(t,z)\right|^2 Q(t,z)\eta(t)}$$

$$- \left(Y(t-)r(t) + \pi(t)\left(\mu(t) - r(t)\right) + f(t)\right), \quad 0 \le t \le T.$$
(10.41)

The quadratic variation is now used for modelling the instantaneous variance. The moments in (10.41) are derived by Theorems 2.3.2–2.3.3. The goal is to find an admissible hedging strategy $\pi \in \mathscr{A}$ which minimizes the instantaneous risk measures $\rho(d\mathscr{S}(t))$ for all $t \in [0, T]$ and choose a generator f of the price dynamics Y which makes the instantaneous risk measures vanish $\rho(d\mathscr{S}(t)) = 0$ for all $t \in [0, T]$. This is a reasonable pricing and hedging objective. The insurer should be interesting in applying an investment strategy under which the expected excess return on the surplus over the risk-free return on the surplus is maximized. At the same time, the insurer should choose an investment strategy under which the return on the surplus (the tracking error) is not volatile. Hence, the mean-variance objective is used for choosing the hedging strategy. The insurance payment process is next priced by

requiring that the instantaneous Sharpe ratio of the surplus process equals a prespecified target L, i.e. the insurance payment process is priced by requiring that

Sharpe Ratio
$$(\mathscr{S}(t))$$

= $\frac{\mathbb{E}[d\mathscr{S}(t) - \mathscr{S}(t-)r(t)dt|\mathscr{F}_{t-}]/dt}{\sqrt{\mathbb{E}[d[\mathscr{S},\mathscr{S}](t)/dt|\mathscr{F}_{t-}]}} = L(t), \quad 0 \le t \le T.$ (10.42)

We can now interpret *L* as a process which controls the ratio of the expected earned surplus (the net asset wealth) to its deviation over time or as a profit, specified by a Sharpe ratio, which is demanded by the insurer who sells the insurance contract. We assume that $L(t) \ge \theta(t) + \varepsilon$, $\varepsilon > 0$. The former inequality is obvious since the investment in the payment process *P* carries an additional risk compared to the investment in the risky stock *S* and the insurer is interested in earning a risk premium strictly above θ , which is the risk premium earned by investing in the stock. For the instantaneous Sharpe ratio pricing we refer to Bayraktar and Young (2007), Young (2008), Bayraktar and Young (2008) and Bayraktar et al. (2009), for the connection with the cost of capital pricing we refer to Pelsser (2011).

Theorem 10.4.1 Assume that (C1)–(C4) from Chap. 7 hold and let L be a predictable process such that $L(t) \ge \theta(t) + \varepsilon$, $\varepsilon > 0$, and $L(t) \le K$, $0 \le t \le T$. The admissible investment strategy $\pi^* \in \mathscr{A}$ which minimizes the risk measures (10.41) for all $t \in [0, T]$ and the generator f^* which makes the risk measures (10.41) vanish for all $t \in [0, T]$ take the form

$$\pi^{*}(t) = \frac{1}{\sigma(t)} \left(Z_{1}(t) + \sqrt{\frac{|\theta(t)|^{2}}{|L(t)|^{2} - |\theta(t)|^{2}}} \sqrt{|Z_{2}(t)|^{2} + \int_{\mathbb{R}} |U(t,z)|^{2} Q(t,dz)\eta(t)} \right),$$

$$0 \le t \le T,$$

$$f^{*}(t) = -Y(t-)r(t) - Z_{1}(t)\theta(t) + \sqrt{|L(t)|^{2} - |\theta(t)|^{2}} \sqrt{|Z_{2}(t)|^{2} + \int_{\mathbb{R}} |U(t,z)|^{2} Q(t,dz)\eta(t)}, \quad 0 \le t \le T,$$

$$(10.43)$$

where $(Y, Z_1, Z_2, U) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2_N(\mathbb{R})$ is the unique solution to the BSDE

$$Y(t) = F + \int_t^T H(s)ds + \int_t^T \int_{\mathbb{R}} G(s, z)N(ds, dz)$$
$$+ \int_t^T \left(-Y(s-)r(s) - Z_1(s)\theta(s)\right)$$

$$+\sqrt{|L(s)|^{2} - |\theta(s)|^{2}}\sqrt{|Z_{2}(s)|^{2} + \int_{\mathbb{R}}|U_{2}(s,z)|^{2}Q(s,dz)\eta(s)}ds$$

$$-\int_{t}^{T}Z_{1}(s)dW(s) - \int_{t}^{T}Z_{2}(s)dB(s)$$

$$-\int_{t}^{T}\int_{\mathbb{R}}U(s,z)\tilde{N}(ds,dz), \quad 0 \le t \le T.$$
 (10.44)

Proof The form of the generator and the strategy are deduced from properties of the function

$$w(\pi) = L\sqrt{(\pi\sigma - z)^2 + u^2 + v^2} - ry - \pi(\mu - r) - f.$$

It is straightforward to find a unique minimizer π^* of the function w and f^* such that $w(\pi^*) = 0$. We now study the BSDE (10.44). By the Schwarz inequality

$$\left| zz' + \int_{\mathbb{R}} u(z)u'(z)Q(t,dz)\eta \right|$$

$$\leq \sqrt{|z|^{2} + \int_{\mathbb{R}} |u(z)|Q(t,dz)\eta} \sqrt{|z'|^{2} + \int_{\mathbb{R}} |u'(z)|^{2}Q(t,dz)\eta}, \qquad (10.45)$$

we can prove the following inequality

$$\left| \sqrt{|z|^{2} + \int_{\mathbb{R}} |u(z)|^{2} Q(t, dz) \eta} - \sqrt{|z'|^{2} + \int_{\mathbb{R}} |u'(z)|^{2} Q(t, dz) \eta} \right|^{2} \\ \leq |z - z'|^{2} + \int_{\mathbb{R}} |u(z) - u'(z)|^{2} Q(t, dz) \eta.$$
(10.46)

Hence, the generator (10.44) is Lipschitz continuous in the sense of Theorem 3.1.1. From (3.22) and Theorem 3.1.1 we conclude that there exists a unique solution to the BSDE (10.44). It is clear that the investment strategy (10.43) is admissible.

The price process of the payment process P solves the nonlinear BSDE (10.44). It is important to point out that the price process (10.44) may not be represented as the expected value of the future discounted claims under an equivalent martingale measure. We also remark that the price process (10.44) may not satisfy the property of monotonicity with respect to the claim in the sense that a more severe claim may have a lower price. Hence, the price derived in Theorem 10.4.1 may give rise to arbitrage opportunities. In the language of BSDEs the solution to the BSDE (10.44) does satisfy the comparison principle with respect to the terminal condition and the BSDE (10.44) does not have a measure solution. This problem, which arises for BSDEs with jumps, was pointed out in Sect. 3.2.

Example 10.3 We consider the classical Black-Scholes model with constant coefficients r, μ , σ and an insurer who issues an endowment policy. We are interested in pricing and hedging the claim F = 1 - J(T) where J is the death counting process for the policy. We consider a constant mortality intensity λ and a constant Sharpe ratio L. In order to find the optimal price process and the optimal investment strategy for the claim F, we have to solve the BSDE

$$Y(t) = 1 - J(T) + \int_{t}^{T} \left(-Y(s-)r - Z(s)\theta + \sqrt{L^{2} - \theta^{2}} |U(s)| \sqrt{(1 - J(s-))\lambda} \right) ds - \int_{t}^{T} Z(s) dW(s) - \int_{t}^{T} U(s)\tilde{N}(ds), \quad 0 \le t \le T.$$
(10.47)

Clearly, we can set Z(t) = 0, $0 \le t \le T$. Let us guess that U(t) = -Y(t-), $Y(t) \ge 0$, $0 \le t \le T$. We end up with the equation

$$dY(t) = Y(t-)rdt - \sqrt{L^2 - \theta^2}Y(t-)\sqrt{(1 - J(t-))\lambda}dt - Y(t-)\tilde{N}(dt),$$

$$Y(T) = 1 - J(T).$$

First, it is straightforward to derive the dynamics

$$d(Y(t)e^{-\int_{0}^{t}(r-\sqrt{L^{2}-\theta^{2}}\sqrt{(1-J(s-))\lambda})ds})$$

= $-e^{-\int_{0}^{t}(r-\sqrt{L^{2}-\theta^{2}}\sqrt{(1-J(s-))\lambda})ds}Y(t-)\tilde{N}(dt),$ (10.48)
 $Y(T) = 1 - J(T).$

Integrating and taking the expected value, we can obtain the candidate solution

$$Y(t) = \mathbb{E}\Big[\Big(1 - J(T)\Big)e^{-\int_{t}^{T}(r - \sqrt{L^{2} - \theta^{2}}\sqrt{(1 - J(s -))\lambda})ds}|\mathscr{F}_{t}\Big]$$

= $\Big(1 - J(t)\Big)e^{-(r + \lambda - \sqrt{L^{2} - \theta^{2}}\sqrt{\lambda})(T - t)}, \quad 0 \le t \le T,$ (10.49)

where we use the fact that J(T) = 0 implies $J(t) = 0, 0 \le t \le T$. Next, from (10.49) we derive the dynamics

$$dY(t) = -e^{-(r+\lambda-\sqrt{L^2-\theta^2}\sqrt{\lambda})(T-t)}dJ(t) + \left(r+\lambda-\sqrt{L^2-\theta^2}\sqrt{\lambda}\right)Y(t-)dt,$$

and

$$d(Y(t)e^{-\int_0^t (r-\sqrt{L^2-\theta^2}\sqrt{(1-J(s-))\lambda})ds})$$

= $e^{-\int_0^t (r-\sqrt{L^2-\theta^2}\sqrt{(1-J(s-))\lambda})ds}(-e^{-(r+\lambda-\sqrt{L^2-\theta^2}\sqrt{\lambda})(T-t)}dJ(t)$
+ $(r+\lambda-\sqrt{L^2-\theta^2}\sqrt{\lambda})Y(t-)dt)$

$$- Y(t-)e^{-\int_0^t (r-\sqrt{L^2-\theta^2}\sqrt{(1-J(s-))\lambda})ds} \left(r-\sqrt{L^2-\theta^2}\sqrt{(1-J(t-))\lambda}\right)dt$$

$$= e^{-\int_0^t (r-\sqrt{L^2-\theta^2}\sqrt{(1-J(s-))\lambda})ds} \left(-e^{-(r+\lambda-\sqrt{L^2-\theta^2}\sqrt{\lambda})(T-t)}(1-J(t-))dJ(t) + \lambda Y(t-)(1-J(t-))\right)dt$$

$$= -e^{-\int_0^t (r-\sqrt{L^2-\theta^2}\sqrt{(1-J(s-))\lambda})ds}Y(t-)\tilde{N}(dt),$$

which agrees with (10.48). Our candidate solution (Y, Z, U), where Y is given by (10.49), Z(t) = 0, U(t) = -Y(t-), $0 \le t \le T$, is square integrable. Hence, the unique solution to the BSDE (10.47) is found. By Theorem 10.4.1 the optimal price process is defined by

$$Y(t) = (1 - J(t))e^{-(r + \lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T - t)}, \quad 0 \le t \le T,$$
(10.50)

the optimal investment strategy is given by

$$\pi^*(t) = \left(1 - J(t-)\right) \frac{\theta\sqrt{\lambda}}{\sigma\sqrt{L^2 - \theta^2}} e^{-(r+\lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T-t)}, \quad 0 \le t \le T$$

We have obtained the unique solution to the pricing and hedging problem under the instantaneous mean-variance risk measure. If $\lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda} \le 0$, then the price of the endowment policy is larger than $e^{-r(T-t)}$ which is the price of the bond paying 1 at the maturity. The property of monotonicity of the pricing operator with respect to the claim is not satisfied. The insurer can buy the bond which hedges the payment from the endowment policy and earns a risk-free profit. We can conclude that the price (10.50) may give rise to arbitrage opportunities. To get the arbitragefree price, we have to introduce the constraint $\lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda} > 0$, or equivalently $L^2 < \lambda + \theta^2$.

We are now interested in pricing and hedging the claim $F = (1 - J(T))(K - S(T))^+$. In order to find the optimal price process and the optimal hedging strategy for the claim *F*, we have to solve the BSDE

$$Y(t) = (1 - J(T))(K - S(T))^{+} + \int_{t}^{T} (-Y(s) - r - Z(s)\theta + \sqrt{L^{2} - \theta^{2}} |U(s)| \sqrt{(1 - J(s))\lambda}) ds$$
$$- \int_{t}^{T} Z(s) dW(s) - \int_{t}^{T} U(s) \tilde{N}(ds), \quad 0 \le t \le T.$$
(10.51)

We can conclude that the unique solution to the BSDE (10.51) is given by the triple

$$Y(t) = (1 - J(t))e^{-(\lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T - t)}$$

$$\cdot \left(Ke^{-r(T - t)}\Phi\left(-d(t, S(t)) - \sigma\sqrt{T - t}\right) - S(t)\Phi\left(-d(t, S(t))\right)\right),$$

$$0 \le t \le T,$$

$$Z(t) = -(1 - J(t -))e^{-(\lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T - t)}\sigma S(t)\Phi(-d(t, S(t))), \quad 0 \le t \le T,$$
$$U(t) = -(1 - J(t -))e^{-(\lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T - t)} \cdot (Ke^{-r(T - t)}\Phi(-d(t, S(t)) - \sigma\sqrt{T - t}) - S(t)\Phi(-d(t, S(t)))),$$
$$0 \le t \le T,$$

where we use (9.6) and (9.7). The optimal price process is defined by

$$Y(t) = (1 - J(t))e^{-(\lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T - t)}$$

$$\cdot \left(Ke^{-r(T - t)}\Phi\left(-d(t, S(t)) - \sigma\sqrt{T - t}\right)\right)$$

$$- S(t)\Phi\left(-d(t, S(t))\right), \quad 0 \le t \le T, \quad (10.52)$$

and the optimal hedging strategy is given by

$$\pi^*(t) = -(1 - J(t -))e^{-(\lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T - t)}S(t)\Phi\left(-d(t, S(t))\right)$$

+ $(1 - J(t -))e^{-(\lambda - \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T - t)}\frac{\theta\sqrt{\lambda}}{\sigma\sqrt{L^2 - \theta^2}}$
 $\cdot \left(Ke^{-r(T - t)}\Phi\left(-d(t, S(t)) - \sigma\sqrt{T - t}\right) - S(t)\Phi\left(-d(t, S(t))\right)\right),$
 $0 \le t \le T.$

The price (10.52) is arbitrage-free if $L^2 < \lambda + \theta^2$.

We also give an example which shows that the price from Theorem 10.4.1 may be arbitrage-free without any additional assumptions on the parameters. It turns out that arbitrage-free prices arise for specific types of claims.

Example 10.4 Let the assumptions from Example 10.3 hold. We consider an insurer who issues a life insurance policy paying 1 at maturity of the contract provided that the policyholder dies within the duration of the contract. We are interested in pricing and hedging the claim F = J(T) where J is the death counting process for the policy. In order to find the optimal price process and the optimal investment strategy for the claim F, we have to solve the BSDE

$$Y(t) = J(T) + \int_{t}^{T} \left(-Y(s-)r - Z(s)\theta + \sqrt{L^{2} - \theta^{2}} |U(s)| \sqrt{(1 - J(s-))\lambda} \right) ds - \int_{t}^{T} Z(s) dW(s) - \int_{t}^{T} U(s)\tilde{N}(ds), \quad 0 \le t \le T.$$
(10.53)

Clearly, we can set Z(t) = 0, $0 \le t \le T$. Let us guess that $U(t) \ge 0$, $0 \le t \le T$. We end up with the BSDE

$$Y(t) = J(T) + \int_{t}^{T} \left(-Y(s-)r + \sqrt{L^{2} - \theta^{2}}U(s)\sqrt{\left(1 - J(s-)\right)\lambda}\right) ds - \int_{t}^{T} U(s)\tilde{N}(ds), \quad 0 \le t \le T.$$

$$(10.54)$$

By Proposition 3.4.1 and a similar reasoning as in Example 10.3, the solution to the BSDE (10.54) is of the form

$$Y(t) = e^{-r(T-t)} - (1 - J(t))e^{-(r+\lambda + \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T-t)}, \quad 0 \le t \le T,$$

$$U(t) = e^{-(r+\lambda + \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T-t)}, \quad 0 \le t \le T,$$
(10.55)

and we see that $U(t) \ge 0, 0 \le t \le T$. Our candidate solution (Y, Z, U), where (Y, Z) is given by (10.55), $Z(t) = 0, 0 \le t \le T$, is square integrable. Hence, the unique solution to the BSDE (10.53) is found. By Theorem 10.4.1 the arbitrage-free optimal price process is defined by

$$Y(t) = e^{-r(T-t)} - (1 - J(t))e^{-(r+\lambda + \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T-t)}, \quad 0 \le t \le T,$$

and the optimal investment strategy is given by the formula

$$\pi^*(t) = \left(1 - J(t-)\right) \frac{\theta \sqrt{\lambda}}{\sigma \sqrt{L^2 - \theta^2}} e^{-(r+\lambda + \sqrt{L^2 - \theta^2}\sqrt{\lambda})(T-t)}, \quad 0 \le t \le T.$$

By Theorem 10.4.1 the hedging strategy for the payment process P is characterized by the control processes of the nonlinear BSDE (10.44). The investment strategy (10.43) is studied in Chap. 12 where the results of this chapter are derived again under different pricing and hedging objectives. In Chap. 12 we also give conditions which guarantee that the price process (10.44) derived under the instantaneous mean-variance risk measure is arbitrage-free.

Instead of the mean-variance minimization, we can use the approach proposed by Bayraktar and Young (2007), Young (2008), Bayraktar and Young (2008), Bayraktar et al. (2009), and we can find a hedging strategy π^* which minimizes the instantaneous variation $\mathbb{E}[d[\mathscr{S}, \mathscr{S}](t)|\mathscr{F}_{t-}]$ for all $t \in [0, T]$ together with a generator f^* which makes the instantaneous mean-variance risk measures vanish $\rho(d\mathscr{S}(t)) = 0$ for all $t \in [0, T]$. We obtain

$$\pi^{*}(t) = \frac{1}{\sigma(t)} Z_{1}(t), \quad 0 \le t \le T,$$

$$f^{*}(t) = -Y(t-)r(t) - Z_{1}(t)\theta(t) \qquad (10.56)$$

$$+ L(t)\sqrt{\left|Z_{2}(t)\right|^{2} + \int_{\mathbb{R}} \left|U(t,z)\right|^{2} Q(t,dz)\eta(t)}, \quad 0 \le t \le T,$$

where $(Y, Z_1, Z_2, U) \in \mathbb{S}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2(\mathbb{R}) \times \mathbb{H}^2_N(\mathbb{R})$ is the unique solution to the BSDE

$$Y(t) = F + \int_{t}^{T} H(s)ds + \int_{t}^{T} \int_{\mathbb{R}} G(s, z)N(ds, dz) + \int_{t}^{T} \left(-Y(s-)r(s) - Z_{1}(s)\theta(s) + L(s)\sqrt{|Z_{2}(s)|^{2} + \int_{\mathbb{R}} |U_{2}(s, z)|^{2}Q(s, dz)\eta(s)}\right)ds - \int_{t}^{T} Z_{1}(s)dW(s) - \int_{t}^{T} Z_{2}(s)dB(s) - \int_{t}^{T} \int_{\mathbb{R}} U(s, z)\tilde{N}(ds, dz), \quad 0 \le t \le T.$$
(10.57)

Again, the price process (10.57) may lead to arbitrage opportunities.

Bibliographical Notes For a review on quadratic pricing and hedging we refer to Schweizer (2010). The proof of Theorem 10.1.1 is classical, see Sect. 10.4 in Cont and Tankov (2004) and Ankirchner and Imkeller (2008). Riccati backward stochastic differential equation is studied in Lim (2004), Kohlmann and Tang (2002), Kohlmann and Tang (2003). Higher moment estimates for a solution to a BSDE driven by a Brownian motion are investigated in El Karoui et al. (1997b). The proof of the optimality in Theorem 10.2.1 is taken from Delong (2010), see also Lim (2004) and Øksendal and Hu (2008) for the method of completing the squares. In the proof of the admissability in Theorem 10.2.1 we closely follow the arguments from Lim (2004). For Markowitz mean-variance optimization problems in finance and insurance we refer to Lim (2004), Øksendal and Hu (2008), Xie (2009), Xie et al. (2008), Kharroubi et al. (2012). It is worth mentioning that quadratic hedging problems in a financial model with assets driven by quasi-left continuous semimartingales with bounded jumps and in a financial model with assets driven by general semimartingales were solved in Kohlmann et al. (2010) and Jeanblanc et al. (2012) where the authors applied semimartingale BSDEs. The proof of Proposition 10.3.1 is inspired by Ankirchner and Heine (2012). For locally risk-minimizing strategies we refer to Ankirchner and Heine (2012), Dahl and Møller (2006), Dahl et al. (2008), Møller (2001) and Vandaele and Vanmaele (2008). Section 10.4 is taken from Delong (2012a).